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Abstract

Notwithstanding the fact that advertising is one of the most used marketing tools, little is known about what is driving (i) the *timing* and (ii) the *magnitude* of advertising actions. Building on normative theory, the authors develop a parsimonious model that captures this dual investment process. They explain advertising spending patterns as observed in the market, and investigate the impact of company, competitive, and category-related factors on these decisions, thereby introducing the novel concept of Ad-sensor. Analyses are based on a unique combination of (i) weekly advertising data on 748 CPG brands in 129 product categories in the UK, (ii) household panel purchase data, and (iii) data on new product introductions. The analyzed brands include both large and small brands, both frequent and infrequent advertisers, thus providing a more complete and correct overview of the market. The results show that advertising spending patterns can be explained as real-life applications of the normative literature, in which advertising and advertising goodwill management are embedded in dynamic (s,S) inventory systems. Adstock and Ad-sensor show a positive effect on both timing and magnitude decision. Competitive reasoning is found to have little to no effect on advertising decisions, whereas category-related factors do show an impact. The extent to which campaigning strategies are more or less the outcome of advertising goodwill management systems, however, varies across brands as a function of their relative size and advertising frequency.

Key words: Advertising, timing, competition, Tobit-II, Bayesian inference

1. Introduction

Advertising is one of the most important marketing instruments. For example, in 2006, US adspend totaled \$285.1 billion, representing 2.2% of the country's GDP. Companies as Procter & Gamble and AT&T spend billions of dollars per year on advertising (AdAge, 2007). Given its prominent position, it should come as no surprise that advertising, and the way it affects people's decisions, has been the subject of an extensive body of prior research (see e.g. Tellis and Ambler 2007 for a recent review). The main focus of these studies was on the quantification of advertising's *effectiveness*. Studies explaining observed advertising *spending patterns*, in contrast, have received much less attention. Still, insights into why brands start/stop advertising are very relevant to advertising media and advertisers alike. The former will benefit from a profound understanding of the purchase behavior of their customers, i.e. the spending patterns of advertising brands. Advertisers, in turn, are interested in accurate predictions on the expected timing and size of their competitors' spending in order to gauge the extent of competitive interference they may expect (cfr. Danaher et al., 2008).





As shown in Figure 1, considerable variation exists along both the timing and size dimension. The first three panels exhibit the weekly advertising expenditures for three soft-drink brands in the UK. Brand A is a frequent and heavy advertiser (100% of the time, average spending £347,348 per week), while brand C is situated at the other end of the spectrum. It engages only occasionally in advertising actions (42% of the time), and when doing so, spends only small amounts (£44,784 on average). Brand B takes an intermediate position: it advertises less often than brand A (59% of the time), but spends a larger amount on these sparse actions than C (£134,481 on average). The bottom panels of Figure 1 depict three brands in the UK cleanser market. Also in that market, considerable variability is observed in both the timing and the size dimension. Moreover, the absolute spending level appears to be considerably lower than in the soft-drink market. What is driving this over-time variation within a given brand? Why do we find such substantial differences across brands? Or across industries?

Some features of these observed patterns may have emerged as the result of applying the guidelines of a series of *normative* studies which have shown that, in most instances, pulsed advertising is an optimal strategy (e.g. Mahajan and Muller, 1986; Mesak, 1992; Park and Hahn, 1992; Villas-Boas, 1993; Naik et al.,1998). These studies, however, although insightful on the optimality of pulsed over even spending, remain vague on some crucial implementation issues, including (i) how often to advertise, (ii) how many weeks an advertising pulse or campaign should last, and (iii) how much should be spent. Moreover, (iv) they do not provide insights on the observed behavioral differences between brands and categories. As such, and differing from our study, their objective is not to explain the variation found in observed behavior. Such real world behavior, in contrast, was the focus of a body of *empirical* studies (e.g. Metwally, 1978; Jones, 1990; Hanssens, 1980a+b; Chandy et al., 2001; Steenkamp et al., 2005). These studies manage to capture and explain very well the behavior under examination, but are weaker in the theoretical foundations of their explanations, thus almost being the opposite of the normative studies.

We build on the normative literature, and develop a framework which allows us to describe and understand the advertising behavior as observed in the market, and this along two dimensions: (i) the *timing* of advertising actions (i.e. whether or not to advertise), and (ii) the

magnitude of these actions. We subsequently relate observed differential behavior across brands to the size of the brands, and the experience they have in advertising.

We begin with an overview of the relevant literature. We subsequently present our conceptual framework, and introduce the core concepts of this study. We describe our econometric model, and give some initial insights in our data. We then present our empirical results, and conclude with a discussion of the key managerial insights and some areas for future research.

2. Relevant Literature

The current paper can be positioned against two research streams: (i) normative studies on optimal advertising behavior, (ii) empirical studies on advertising and its effectiveness.

Normative literature

Over the past decades, the preponderance of the prescriptions from normative studies on the optimal timing of advertising has shifted from constant advertising schedules (Zielske, 1959; Sasieni, 1971; 1989) to pulsing advertising schedules (e.g. Mahajan and Muller, 1986) as more and more real-world effects were included in the analyses. For example, Katz (1980) introduced learning and forgetting effects, while Mesak (1992) and Naik et al. (1998) added, respectively, wear-out effects and quality restoration. Park and Hahn (1992), Villas-Boas (1993) and Dubé et al. (2005), in turn, extended the analyses to competitive settings. In most instances, pulsed advertising is now considered to be the optimal strategy for firms. Whereas pulsing is used as a generic term describing advertising schedules alternating high and low levels of advertising, flighting (e.g. Katz, 1980) is more strict in its definition as it refers to alternation between high and zero levels of advertising. As such, it is an extreme case of pulsing. Finally, the concept of campaigning (Doganoglu and Klapper, 2006) was introduced to describe the fact that advertising pulses are often not one-time spikes, but regularly last several weeks.

Pulsing strategies appear to be frequently applied by managers (Feinberg, 1992). Recently, Dubé et al. (2005) found evidence that the observed behavior in the US Frozen Entrée category could be explained as a pulsing strategy based on a dynamic competitive game. Doganoglu and Klapper (2006), covering the German Detergent Market, found similar support for the application of the normative guidelines in the real world decisions they studied. Finally, such patterns are also widely present in our own dataset (cfr figure 1), in which only 6 of the 748 included brands advertise permanently, alternating between high and low levels of advertising.

However, in contrast with their general agreement on the optimality of pulsed advertising strategies, these normative studies provide less clarity on a number of issues related to the actual implementation of the advocated strategies. Analysis of real world advertising spending patterns indeed showed that, although pulsing strategies are often encountered, large differences are observed both within and between brands in (i) the actual timing of advertising campaigns, (ii) the number of weeks such campaigns last, and (iii) the monetary value of campaigns. Overall, very few normative studies go into that level of detail, and three limitations of these studies thus appear. First, these studies mainly focus on the timing of advertising actions within campaigns, thereby leaving especially decisions on the magnitude of these actions uncovered. Second, they provide guidelines for a single brand, thereby ignoring differences in individual brands' advertising preferences as well as factors that may systematically affect advertising decisions across different brands and categories. Finally, as a corollary of this single-brand focus, only very few studies allow advertising decisions to be correlated with the decisions of other brands (Park and Hahn, 1992; Villas-Boas, 1993; Dubé et al., 2005).

Empirical literature

In a wide series of econometric studies on advertising, measuring the *effectiveness of advertising* takes a central position. Performance was treated as a function of advertising expenditures in so-called single equation models (e.g. Lambin and Palda, 1969; Lambin, Naert and Bultez, 1975). These models treated advertising as exogenous, without investigating how spending patterns were formed. This exogeneity assumption was relaxed in subsequent simultaneous equation models, starting with Bass (1969) and including work by Bass and Parsons (1969) and Hanssens (1980a), as in more recent VAR models (Dekimpe and Hanssens, 1995a; 1999). The latter not only allow for feedback effects (when past own performance helps explain current spending), but also *competitive interactions*. A recent study in this field is Steenkamp et al. (2005), who used vector-autoregressive models to study advertising reaction strategies in 442 packaged goods categories.

A major strength of these studies is that they, in contrast to the normative studies, look further than the explanation of the behavior of just one brand, and try to explain patterns across brands and categories. However, this body of research shares three important limitations. First of all, the theoretical background in these studies is often rather limited. Often, observed patterns are explained without a concise and consistent theoretical framework grounded in the normative literature. Second, although advertising expenditures are no longer treated as exogenous, no distinction is made between the decisions to advertise or not (timing), and how much to spend when advertising (magnitude), even though the factors that drive both decisions could (partly) be different or have different weights. Finally, most (if not all) of these studies show a bias towards large and frequently acting brands. This is due to the fact that most time-series techniques have problems with large numbers of zeros and irregular patterns (e.g. in advertising spending), as is often the case with smaller brands. Steenkamp et al. (2005), for example, limited themselves to those top-3 brands in each category that also had an average share larger than 5%, and that advertised at least more than 12% of the time (25 out of 208 weeks). Zanutto and Bradlow (2006) showed that such data pruning may bias the overall inferences, as the included brands are only representative for a small subset of all brands in the market. Hence, the empirical generalizations derived in these studies may only be valid for that specific subset.

Our study

We build upon the insights of the existing normative literature on optimal advertising scheduling by including campaigning in our framework. Two crucial elements in our work are the concepts of Adstock and Ad-sensor, capturing the campaigning state dependence of brands and their dynamic pulsing behavior, respectively.

The definition of campaigning implies two basic conditions. First of all, campaigns are defined as prolonged periods of advertising, alternated by periods without advertising. Once advertising has begun, it will be continued for some time. This state dependence will be captured by the Adstock concept, a concept which is widely known and used in advertising research (e.g. Broadbent, 1984; Hanssens et al., 2001). The probability of a new advertising action will be higher if a brand was also advertising in the previous weeks, which has resulted in higher Adstock. Conversely, the probability of no new advertising action will be higher if the brand was

not advertising in the previous weeks, during which Adstock was depleting, resulting in a lower Adstock level. The second condition holds that, at a certain moment in time, one has to start a new campaign, i.e. start advertising. Conversely, at a certain moment in time, a campaign will end, and one has to stop advertising. Managers want to keep their Adstock above a certain level. As soon as the Adstock, built by previous campaigns, has depreciated to a that level, one will start advertising. Similarly, one is likely to stop advertising once a desired (higher) level has been reached again. Such advertising reasoning shows close resemblance to so-called (s,S) inventory management systems. Such systems keep the stock of a certain item between a minimum level s and a maximum level S, by repurchasing if the stock becomes too small, up to the desired maximum level. Although very popular and widely used in logistics (e.g. Silver et al., 1998), applications in advertising research are rather scarce (Zufryden, 1979; Doganoglu and Klapper, 2006).

As such, we build a parsimonious and flexible model which captures in a straightforward way how observed advertising spending patterns could result from dynamic advertising adjustment strategies. We address weaknesses of previous research by incorporating four main challenges in our model specification: (i) we allow for differential processes driving the timing and magnitude decisions, (ii) we accommodate heterogeneous preferences and behavior across brands, (iii) we examine the effect of moderating variables across brands and categories, and (iv) we accommodate correlations between the brands decisions within the same category.

Our study, in addition, is unique in its empirical scope, covering advertising decisions on a weekly basis for 748 brands in 129 CPG categories. We include all brands in these categories that advertise in our study, regardless of their advertising intensity (provided that they advertise at least once, otherwise the issue becomes moot) and their Brand market share. An illustration of the relative importance of this issue can be found in figure 2. This figure depicts the consequences of the application of the decisions rules as used in Steenkamp et al. (2005) to our dataset. We categorize all brands according to their compliance (Y/N) with the size (top 3, minimum 5% market share) and frequency limitation (minimum advertising frequency of 12%).

Figure 2. Application of the Steenkamp et al. (2005) data pruning rules to our dataset

	Advertising Frequency > 12%	Advertising Frequency ≤ 12%
Top 3	N = 151 Combined share of advertising = 63.9%	N = 71 Combined share of advertising = 0.8%
AND Market share > 5%	Mean market share = 18.4% Mean combined market share = 32.7% Mean advertising share in category = 41.2% Mean number of advertising weeks = 82	Mean market share = 17.7% Mean combined market share = 20.3% Mean advertising share in category = 40.6% Mean number of advertising weeks = 7
Not Top 3	N = 229 Combined share of advertising = 33.7%	N = 297 Combined share of advertising = 1.6%
OR Market share ≤ 5%	Mean market share = 2.1% Mean combined market share = 7.7% Mean advertising share in category = 11.5% Mean number of advertising weeks = 57	Mean market share = 1.0% Mean combined market share = 3.2% Mean advertising share in category = 3.9% Mean number of advertising weeks = 6

For each block we report the Number of brands (N); their Combined share of advertising in our dataset; Mean market share; Mean combined market share; Mean advertising share in their category; and the Mean number of advertising weeks. If we only include the brands in our empirical dataset that fulfill both requirements, we would cover only 63.9% of all advertising expenditures. Brands would have an average market share of 18.4%, and an advertising share of 41.2%, advertising on average 82 out of 156 weeks. However, these brands, on average, account for only 32.7% of the total market, covering, on average, between 0.6% and 99% of category advertising expenditures as included in our dataset. Limiting the number of included brands would clearly lead to the omission of a major part of the observed advertising actions and expenditures from our analyses. Relaxing the aforementioned requirements hence appears recommended. Relaxing the size limitation, would enable us to cover over 98% of all advertising expenditures. In addition, although infrequent advertisers represent only a very small percentage of advertising expenditures, they still account for almost 50% of the advertising brands, and nearly 10% of all advertising *actions*. Understanding how their behavior may differ from more frequently acting brands is consequently warranted if we want to understand advertising as we observe it in the market. The unique dataset we thus obtain, allows us to formulate a set of insightful empirical generalizations on the timing and magnitude of observed advertising patterns.

3. Drivers of Advertising Investment Decisions

Advertising decisions can be seen as a multiple decision process. Two key decisions which have to be taken, are *when* to advertise, and *how much* to spend (Tellis, 2004 p 72; Danaher, 2007 p 645; Danaher et al., 2008). This dual advertising decision is treated as an investment decision process, which is in line with the growing stream of literature claiming that marketing expenditures are more and more considered to be investments (Srivastava et al., 1999). At each point in time, the brand therefore chooses (i) whether or not to advertise (*timing*), and, (ii) conditional upon this decision, how much to spend (*magnitude*) (e.g. Bar-Ilan and Strange, 1999).

Adstock

Central to our model are the concepts of Adstock and Ad-sensor, which in itself is also derived from Adstock. These concepts capture two crucial aspects of advertising investments. The tendency to concentrate advertising investments in longer pulses or campaigns, as we will show, is captured by our Adstock variable (Broadbent (1984). Ad-sensor is subsequently introduced as a feedback variable mimicking the brand's decision rule to start and stop advertising campaigns. This basic advertising investment decision process is depicted in figure 3.





The Adstock concept was originally developed to assess the dynamic effects of advertising. It rests on the assumption that advertising helps to build a stock of advertising goodwill (Broadbent, 1984). In the absence of further advertising spending, however, this Adstock decays at a constant rate (see e.g. Dekimpe and Hanssens, 2007). In the past, it has been

used in studies on e.g. advertising awareness (Brown, 1986), television advertising effectiveness (Tellis and Weiss, 1995), television scheduling (Broadbent et al., 1997; Ephron and McDonald, 2002), trial of new products (Steenkamp and Gielens, 2003), product-harm crises (Cleeren et al., 2007) and competitive advertising interference (Danaher et al., 2008). In line with these studies, we follow the definition by Broadbent (1984) and operationalize Adstock of brand b as:

(1)
$$Adstock_{b,t} = (1 - \lambda_b)Advertising_{b,t} + \lambda_bAdstock_{b,t-1}$$

Advertising is often scheduled in campaigns of several consecutive weeks, followed by zero advertising during a number of weeks. The likelihood of a brand advertising in week t will consequently be higher when it was also advertising in the weeks before. During this period, Adstock will be built by means of advertising actions. So, a brand which is in a campaign keeps advertising when its Adstock is at relatively high levels. Conversely, the likelihood of a brand not advertising in week t will be lower when it was not advertising in the previous weeks. During such periods, Adstock will be considerably lower as it decays when no new advertising investments are made. We therefore expect a positive effect of Adstock on the likelihood of a new advertising action in week t.

Conditional upon their decision to advertise, brands still have to determine the amount they will spend on their action. As decision making processes are often characterized by a strong preference for the status quo (e.g. Samuelson and Zeckhauser, 1988), previous behavior is a particularly good predictor of new actions (e.g. Srinivasan et al., 2007). Higher advertising levels during the previous weeks of a campaign are therefore likely to be followed by higher levels in a subsequent action, provided that the brand chooses to advertise. Higher previous advertising expenditures during a campaign, in turn, are reflected in relatively higher Adstock levels, whereas lower advertising levels will have resulted in relatively lower Adstock levels. We therefore can expect that brands, conditional on the decision to advertise, will spend more on such new actions when their Adstock level is higher.

Together, these two effects, both in timing and magnitude of advertising actions, show how Adstock captures the state dependence of brands in their advertising decisions.

Ad-sensor

Although Adstock provides a better understanding of the campaigning state dependencies, it is less insightful on why brands would start or stop an advertising campaign. What triggers the launch of new campaigns? Why and when do they end? Answers to these questions can be found by analyzing the goals of advertising investments. By means of advertising, brands build advertising goodwill among consumers. This goodwill is expected to translate into sales, and should consequently not fall below a certain level. If this, however, would be the case, new advertising investments become necessary to preserve and strengthen sales. At that moment, a new campaign will be launched. The ultimate goal of any brand would be to achieve unlimited goodwill. However, advertising budgets are not unlimited. We therefore assume that a specific target level of goodwill will be determined for each campaign, the level of which is unknowng to us. In the beginning of a campaign, when goodwill build-up has just started, incentives to stop will be rather small. The closer to that target level, however, the lower will be the need to continue investing. In addition, once it has been reached, there is a clear incentive to stop investing. The pressure a brand feels to start a campaign as soon as its goodwill among consumers becomes too low, and to stop that same campaign when the desired (higher) goodwill level is reached, is the essence of our Ad-sensor concept.

We define our Ad-sensor by integrating the Adstock concept in an (s,S) stock replenishment system. As such, s represents the minimum Adstock level a brand wants to maintain, whereas S is the (higher) target level. The implicit goal of such (s,S) systems is hence to maintain the stock between the two levels. These levels are known to the brand manager, but unknown to the researcher. In addition, managers can apply dynamic strategies in their choice of these minimum and maximum levels, thus allowing for different levels in different campaigns. We therefore consider the observed minimum and maximum levels as the actual outcomes of managers' utility maximization calculi.

To model this (s,S) behavior, we first consider what happens when the Adstock of brand b falls below the minimum desired level. At that moment, the brand should start advertising again. The brand will launch a new campaign, and will continue to invest in advertising until the target is reached. At that maximum, the first derivative of Adstock to time is zero, at least in continuous time. Since we are using discrete time observations, however, at time t the researcher can only observe up to time t-1, so the first order condition is equivalent to:

(3)
$$\frac{\Delta A dstock_{b,t-1}}{\Delta t} = 0$$

this becomes:

(4)
$$\frac{Adstock_{b,t-1} - Adstock_{b,t-2}}{\Delta t} = 0$$

(5)
$$\Leftrightarrow Adstock_{b,t-1} - Adstock_{b,t-2} = 0.$$

Given our additive Adstock function, this is satisfied if:

(6)
$$(1 - \lambda_b) A dvert_{b,t-1} + \lambda_b A dstock_{b,t-2} - A dstock_{b,t-2} = 0$$

(7)
$$\Leftrightarrow (1 - \lambda_b) A dvert_{b,t-1} = (1 - \lambda_b) A dstock_{b,t-2}$$

(8)
$$\Leftrightarrow Advert_{b,t-1} = Adstock_{b,t-2}$$

The second-order condition for a maximum requires that in the period before the maximum, t-1, Adstock is still increasing, which requires that $Advert_{b,t-2} > Adstock_{b,t-3}$. Based on these observations, we define our Ad-sensor variable as the difference between Adstock in time *t*-1 and Adstock in *t*-2:

(9)
$$Adsensor_{b,t} = Adstock_{b,t-1} - Adstock_{b,t-2}$$

The full mathematical derivation of Ad-sensor is given in appendix A. As such, the Ad-sensor allows us to capture the evolution of a brand's Adstock, and the associated pressure to advertise. During the build-up of Adstock, Ad-sensor will have positive values. Over time, as Adstock increases, these values start to decrease. Once the target maximum Adstock level S was reached, Adstock becomes negative, a clear incentive to stop advertising. Over time, Adstock depletes, and Ad-sensor increases again, implying an increasing pressure for the brand to advertise again.

To provide better insights in the functioning of this system, we simulate a series of advertising actions. In our simulation, we impose values on the carry-over parameter λ , which is used to calculate Adstock and consequently Ad-sensor, and the minimum and maximum Adstock levels s and S. In practice, however, we estimate all parameters based on the observed advertising patterns. The numerical build-up of this example is included in appendix B. The evolution of the associated Ad-sensor is represented in figure 4. The solid grey line represents the advertising expenditures, the dotted black line the created Adstock and the solid black line the Ad-sensor. As indicated by equation 9, the latter represents the recent change in Adstock due to advertising investments.

Figure 4. Advertising, Adstock and Ad-sensor



In week 4, the brand launches a new campaign, as Adstock has fallen below the allowed minimum level s. As a corollary of our Adstock definition, Adstock will increase as long as the advertising investments are larger than the previous Adstock level (see appendix A). The first investments start building Adstock, but at the same time as well increase the pressure not to stop the campaign prematurely (captured by the Ad-sensor), as the desired level S is still far away. By period 6, the Adstock level is increasingly approaching the advertising level, and increases in Adstock start becoming smaller. Gradually, the brand is approaching the target Adstock level S. This is also reflected in the Ad-sensor, which slowly starts to decay from period 7 on. Stopping becomes less harmful, as the target level is getting closer. In week 10, the maximum (desired) Adstock level has been reached. By week 11, Ad-stock starts to decay. Ad-sensor, in turn, becomes strongly negative in the next period: the target of the campaign was attained, continued investments make little sense. However, still some smaller amounts of advertising are typically found at the end of advertising campaigns. These are often due to so-called make-goods (see e.g. Doganoglu and Klapper, 2006), smaller actions often added at the end of campaigns in order to compensate for lost opportunities during the campaign itself. Over time, Adstock depletes at a constant rate λ (see equation 1), but not in constant absolute terms. When Adstock levels are still high, depletion will be large in absolute terms, causing strong negative Ad-sensor values. Over time, the Adstock level becomes smaller, and depletion will be smaller in absolute terms. This results in a gradual increase of the Ad-sensor.

In sum: the Ad-sensor variable is essentially a feedback variable mimicking the brand's decision rule to start and stop advertising. We thus expect a positive effect of the Ad-sensor on the timing decision (yes or no). In addition, a similar effect is hypothesized for the magnitude decision (given timing). A wider gap between the actual and the target Adstock level requires

stronger efforts to rapidly bridge this gap. As the brand gets closer to the target level, however, this pressure becomes smaller, as the gap has become much smaller. Beyond the point in time during which the target level was reached, the relevance of continued spending can be questioned. The Ad-sensor tells brand that, if they still would be spending on advertising actions, it should at least be small amounts.

Moderating factors of Adstock and Ad-sensor

The combination of Adstock and Ad-sensor creates a new model for the analysis of advertising decisions. A subsequent investigation of the general applicability of this model across a large set of brands is hence asked for. However, the extent to which Adstock is managed between s and S may differ across brands. In this paper, we focus on two important brand characteristics: Brand market share and Advertising frequency. The motivation for this choice is threefold. First of all, both factors have commonly been used in the past as a basis for data selection rules (e.g. Steenkamp et al., 2005). Brands typically included in previous studies on the basis of such selection rules, however, are not representative for the market as a whole (Slotegraaf and Pauwels, 2008), and inferences based on these subsets could be biased (Zanutto and Bradlow, 2006). As we include all types of brands, understanding to what extent behavior may differ for types of brands which were previously excluded from analyses is essential, and can in addition provide insights on the extent to which data pruning rules could have altered our findings. In addition, two other motivations explain our choice. Brand market share has emerged from previous research as a key characteristic in advertising decisions (e.g. Patti and Blasko, 1981; Lynch and Hooley, 1990). Advertising decision making and its outcome will depend on the market share a brand has and wants to maintain. Advertising frequency, in turn, may create learning effects. An examination of the effect of advertising frequency consequently helps us to understand if, and to what extent, more experienced brands have gained a relative advantage over time in managing their Adstock.

Larger brands have more means at their disposal, and marketing budgets, moreover, are often determined on a percentage of sales basis (Allenby and Hanssens, 2005). This will affect the advertising decision process in two ways. These brands can reserve larger budgets for marketing research compared to their smaller counterparts. They thus can be expected to be better informed on how their advertising goodwill level is evolving, and can consequently better react

to it. We therefore expect a positive effect of *Brand market share* on the effect of Ad-sensor in both decisions. In addition, their larger budgets allow them to pursue longer and more intense advertising campaigns. The state dependence effect as implied by the Adstock will consequently be higher for such brands.

Experience enables brands to adapt their organizations and processes in order to perform optimally. The advertising decision process is no exception to this. More experienced brands have become more efficient through learning effects and have established effective advertising decision processes. On the organizational level, these decision processes tend to stay very similar over time (Frederickson and Iaquinto, 1989), as organizations have to be reliable, accountable and reproducible (Boeker, 1988). What has proven to be effective, will be continued. A consistent closer monitoring of the advertising goodwill evolution as well as reactance to it can consequently be expected. We therefore expect the effect of Ad-sensor to be positively affected by *Advertising frequency*. This experience, built by more intense advertising strategies in the past, moreover, will as well enable brands to better pursue longer and more intense advertising campaigns. The effect of Adstock is consequently expected to be stronger for more experienced brands.

Covariates

However, it is unlikely that advertising decisions are only influenced by brands' own internal advertising reasoning. Three main sets of factors are commonly considered: (i) Company factors, (ii) Competitive factors, and (iii) Category or marketplace factors (Montgomery et al., 2005).

First of all, brands will look at themselves. As advertising theory tells that *New product introductions* should be combined with more intense advertising campaigns (e.g. Rossiter and Percy, 1997; Kotler and Armstrong, 2004), especially because advertising has shown to be more effective for new products (Lodish et al., 1995), we can expect a positive effect of such introductions on the advertising decisions, resulting in a higher likelihood of advertising and higher actual expenditures. Overall, however, advertising budgets are often set on a yearly basis (e.g. Farris and West, 2007). In the course of the year, these budgets get depleted, most often faster than expected, sometimes slower, creating relative shortage or slack resources by the end of the year respectively. Given the common knowledge that having money leads to spending it,

we expect brands to spend their budgets relatively faster in the beginning of the year. This leaves them with relatively fewer means at the end, and thus a negative impact of an *End of year* dummy on the advertising decisions, resulting in fewer and smaller advertising actions, is hypothesized.

Next to themselves, brands will monitor their competitors and their own performance relative to the latter. *Competitive adstock* captures the likelihood of competitive advertising campaigns. The effect of this factor on the decision outcome is not clear a priori, as arguments in both directions can be found, with brands clearly reacting on each other (e.g. Metwally, 1978; Chen and MacMillan, 1992), or trying to avoid competitive clutter (Danaher et al., 2008). In addition, brands frequently make decisions in order to perform well relative to their competitors, on e.g. market share (e.g. Metwally, 1978; Armstrong and Collopy, 1996). However, as budgets are often set as a percentage of past sales (see e.g. Allenby and Hanssens, 2005), a negative *Relative performance evolution* versus competitors as expressed by a decrease in market share may at the same time create a stronger urge to react and lower the ability to react. Here as well, the effect on the advertising decisions is not clear a priori.





Finally, brands are looking at the markets they are operating in. High *Category growth* not only engenders larger current profits and hence marketing budgets, it can also be regarded as an indicator of potential future profits, leading companies to defend their positions in such categories even more fiercely (Gatignon et al., 1990). At the same time, however, if category growth is near zero, competitive actions can become a zero-sum game. Hence, such categories

can be characterized by intense competition, also with advertising, to protect sales volumes (Aaker and Day, 1986). Given the well-known detrimental impact of such strategies on the profits of the brands, we hypothesize that the indicator-of-future-profit effect will be stronger than the zero-sum effect, leading to a positive effect of category growth on the advertising decisions. Categories with higher *Category concentration* are more open to collusive behavior, leading to lower competitive interactions (e.g. Steenkamp et al., 2005), and thus likely as well to lower advertising spending. We therefore expect to find a negative effect of concentration on the advertising decisions.

The conceptual framework presented here is summarized in figure 5. We investigate the ability of Adstock and Ad-sensor to explain advertising decisions, test the influence of moderating factors Brand market share and Advertising frequency, while controlling for an extensive set of other variables. We relate our results further to the literature when reporting our empirical findings.

4. Model Development

Our conceptual framework implies four modeling requirements. First, we need to model both the timing (yes/no) and spending decisions (monetary value), while allowing for different response parameters for both decisions¹. Second, these response parameters are allowed to vary across brands. Third, we need to accommodate the effects of the moderating variables, preferably in a simultaneous estimation step for maximal statistical efficiency. Fourth, the decisions of when and how much to spend may be interrelated between brands within a category, and hence we need to specify a full error covariance structure.

To meet these requirements we link the drivers to the two decision variables (i.e. timing and magnitude) through a new multivariate Hierarchical Tobit-II model, which extends the models of Fox et al. (2004) and Van Heerde et al. (2008) as these models do not comply with all

¹ A similar framework, investigating timing and magnitude of capital stock investments, was introduced by Bar-Ilan and Strange (1999). The authors allowed the company to decide on both when and how much to invest, thereby going beyond existing literature that focused on either timing or intensity of investments. Bowman, Farley and Schmittlein (2000) extended this reasoning to the choice and level of use of international service providers. Dekimpe, Parker and Sarvary (2000), in turn, introduced a similar dual but also more sequential reasoning to international new product adoption processes. Finally, Gielens and Dekimpe (2007) applied a resembling framework to the entry strategy of retail firms into transition economies.

four requirements². In the subsequent exposition, *c* refers to category (c=1,...,C) and *b* to brand $(b=1,...,B_c)$.

Timing

An advertising decision in category c by brand b in week t (z_{cbt}) is described by a multivariate probit model:

(10)
$$z_{cbt} = \begin{cases} 1 & if \ z_{cbt}^* > 0 \\ 0 & otherwise \end{cases}$$

Previous work suggests analyzing the decisions at a weekly level. In general, less data aggregation provides more accurate results (Tellis and Franses, 2006). In addition, the managerial survey reported by Steenkamp et al. (2005) indicated that brands can react to events as fast as within one week, but generally not faster.

The latent variable z_{cbt}^* , describing the timing decision process of the brand, is modeled through a linear model:

$$z_{cbt}^{*} = \zeta_{1,0}^{cb} + \zeta_{1,1}^{cb} Adstock_{cb,t-1} + \zeta_{1,2}^{cb} Ad - sensor_{cb,t} + \zeta_{1,3}^{cb} CompAdstock_{cb,t-1} + \zeta_{1,4}^{cb} PerfEvol_{cb,t-1} + \zeta_{1,5}^{cb} NPI_{cb,t} + \zeta_{1,6}^{cb} EOY_{cb,t} + \zeta_{1,7}^{cb} CatGrowth_{c,t-1} + \zeta_{1,8}^{cb} CatConc_{c,t-1} + \zeta_{2,1}^{cb} Holiday_{t} + \zeta_{2,2}^{cb} QrtrI_{t} + \zeta_{2,3}^{cb} Qrtr2_{t} + \zeta_{2,4}^{cb} Qrtr3_{t} + \zeta_{2,5}^{cb} Trend_{t} + \mu_{cbt}$$

In equation (11) we first include a set of *time-varying* variables, i.e. Adstock and Ad-sensor and the covariates Competitive adstock, Relative performance evolution, New product introduction, End of year budget depletion, Category growth and Category concentration. In addition, we *control* for holidays, seasonality, and possible trending behavior. As advertising decisions for time t will be based on information available up to time t-1, we include 1 period lagged versions of most time-varying explanatory variables.

Amount spent

Conditional on the decision to advertise ($z_{cbt} = 1$), we model y_{cbt} , the natural logarithm of the amount spent on advertising by brand *b* in category *c* during week *t* as:

 $^{^{2}}$ These models do not allow for effects of moderating variables on the response parameters. In addition, in the work of Fox et al. (2004) no full error covariance structure is specified across the two equations; error covariance structures are separated for incidence and timing.

(12)

$$y_{cbt} = \omega_{1,0}^{cb} + \omega_{1,1}^{cb} Adstock_{cb,t-1} + \omega_{1,2}^{cb} Ad - sensor_{cb,t} + \omega_{1,3}^{cb} CompAdstock_{cb,t-1} + \omega_{1,4}^{cb} PerfEvol_{cb,t-1} + \omega_{1,5}^{cb} NPI_{cb,t} + \omega_{1,6}^{cb} EOY_{cb,t} + \omega_{1,7}^{cb} CatGrowth_{c,t-1} + \omega_{1,8}^{cb} CatConc_{c,t-1} + \omega_{2,1}^{cb} Holiday_{t} + \omega_{2,2}^{cb} Qrtr1_{t} + \omega_{2,3}^{cb} Qrtr2_{t} + \omega_{2,4}^{cb} Qrtr3_{t} + \omega_{2,5}^{cb} Trend_{t} + \varepsilon_{cbt}$$

We include the same explanatory variables in the magnitude equation as in the timing equation. Although there is no specific requirement to have the same set in both equations, we include them in an exploratory way to investigate whether all factors have an effect in both decisions.

Moderating factors

In their specific baseline advertising preferences (the intercepts) and their reactions to their Adstock and Ad-sensor, brands may be guided by a number of own-company, and category factors, as these may shape both the ability and the necessity to react. We therefore relate a subset of the response parameters ζ_1^{cb} and ω_1^{cb} to a set of moderator variables:

(13)
$$\zeta_{1,0}^{cb} = \overline{\zeta}_{1,0,0} + \overline{\zeta}_{1,0,1} BrandMarketShare_{cb} + \overline{\zeta}_{1,0,2} AdvFreq_{cb} + \overline{\zeta}_{1,0,3} Food_c + \overline{\zeta}_{1,0,4} Drinks_c + \overline{\zeta}_{1,0,5} Cosmetics_c + u_{1,0}^{cb}$$

(14)
$$\zeta_{1,i}^{cb} = \overline{\zeta}_{1,i,0} + \overline{\zeta}_{1,i,1} BrandMarketShare_{cb} + \overline{\zeta}_{1,i,2} AdvFreq_{cb} + u_{1,i}^{cb}, \quad for \ i = 1, 2$$

and

(15)
$$\omega_{1,0}^{cb} = \overline{\omega}_{1,0,0} + \overline{\omega}_{1,0,1} BrandMarketShare_{cb} + \overline{\omega}_{1,0,2} AdvFreq_{cb} + \overline{\omega}_{1,0,3} Food_c + \overline{\omega}_{1,0,4} Drinks_c + \overline{\omega}_{1,0,5} Cosmetics_c + e_{1,0}^{cb}$$

(16)
$$\omega_{l,i}^{cb} = \overline{\omega}_{l,i,0} + \overline{\omega}_{l,i,1} BrandMarketShare_{cb} + \overline{\omega}_{l,i,2} AdvFreq_{cb} + e_{l,i}^{cb}, \quad for \ i = 1, 2.$$

The included moderator variables are *time-invariant* and allow us to measure the cross-sectional variance in baseline advertising preferences as expressed by the intercept included in (11)-(12). The categories included in our sample, can be categorized under four main product classes, i.e. Household Products, Food, Drinks and Cosmetics. We control for the preferences of these four product classes, using Household Products as reference category. In addition, we investigate the possible moderating effect Brand market share and Advertising frequency can have on both the baseline advertising preferences and the impact of Adstock and Ad-sensor on the advertising decisions. Mean-centering of both variables allows us to examine the effects of deviations relative to an 'average' brand. The effects of the covariates, captured by $\zeta_{1,i}^{cb}$ and $\omega_{1,i}^{cb}$ (for i =

3...8), are related to the hyperparameters $\overline{\zeta}_{1,i,0}$ and $\overline{\omega}_{1,i,0}$, and the brand-specific error terms $u_{1,i}^{cb}$ and $e_{1,i}^{cb}$.

Decisions by one brand on when and how much to advertise are likely to impact those of other brands. We therefore assume that the error vectors $\boldsymbol{\mu}_{ct} = (\mu_{c1t}, ..., \mu_{cB_{ct}})'$ and $\boldsymbol{\varepsilon}_{ct} = (\varepsilon_{c1t}, ..., \varepsilon_{cB_{ct}})'$ follow a joint multivariate normal distribution, with a full variance-covariance matrix: $(\boldsymbol{\varepsilon}_{ct}, \boldsymbol{\mu}_{ct}')' \sim MVN(0, \boldsymbol{\Sigma}_{c})$. Finally, unobserved drivers of model parameters may cause the error terms in (13) up to (16) to be correlated as well: $(\boldsymbol{e}_{cb}, \boldsymbol{u}_{cb}')' \sim MVN(0, \boldsymbol{\Omega})$.

We estimate model (11)-(16) with Bayesian techniques, i.e., Gibbs sampling. The benefit of this approach over classical approaches is that it, at the same time, (i) accommodates the multivariate nature of our dependent variable, (ii) allows for full variance-covariance between all decisions by brands within the same category, and (iii) estimates the moderator effects simultaneously with the other parameters rather than in a two-step approach. An overview of this procedure is given in appendix C.

To operationalize the Adstock and Ad-sensor variables, we need to know the brandspecific carry-over parameter λ_b (see equation 2). To estimate these carry-over parameters, we use the following traditional partial adjustment model (Hanssens et al., 2001 p147):

(17) $Sales_{b,t} = \alpha_b + \beta_b A dv_{b,t} + \lambda_b \ln Sales_{b,t-1} + \varepsilon_{b,t}$

Here as well, we use Gibbs sampling to obtain draws for λ_b . For more details, we refer to appendix D.

5. Data description

We estimate our model on 129 CPG categories in the UK. These categories cover nearly completely the assortment offered in a typical supermarket. An overview of the included product categories, along with the range of included brands, is given in table 1.

Product Fields	Examples	No. of Categories	Range of Brands
Assorted Foods	Breakfast cereals, dry pasta, flour	23	1-11
Beverages	Brandy, mineral waters, softdrinks	20	1-27
Cakes	Oatcakes, crumpets pickelets and muffins	4	1-9
Candy	Cereal bars, countline chocolate, fruit bars	7	1-15
Canned/bottled foods	Canned fish, canned fruit	8	1-9
Care products	Deodorants, shampoo, toilet tissue	22	1-27
Cleaning products	Descalers, scouring powders, drain care	14	1-14
Dairy products	Butter, cream, yoghurt	7	1-11
Frozen foods	Frozen fish, frozen vegetables	4	2-6
Household supplies	Batteries, car freshener	3	1-9
Pet products	Dog food, cat litter	3	2-21
Taste enhancers	Mustard, vinegar, Worcester sauce	14	1-30
Total		129	748

 Table 1. Overview of included product types

Four years of weekly advertising spending data were obtained from NielsenMedia, from which we use one year (52 weeks) as initialization period, and three years (156 weeks) as estimation period. All brands that (i) advertised at least once during the estimation period, and (ii) were present in the market during the whole estimation period, were included. As such, both small and large brands were considered, as reflected in the range of their average (across the three years estimation period) market-shares, which varied from a low of 0.0002% to a high of 97.76%. We focused on national brands, as private labels have a very different positioning and advertising strategy (Kumar and Steenkamp, 2007). Still, private labels were considered in the derivation of certain covariates, as the concentration level in the industry or the national brands' market shares. These decision rules resulted in a sample of 748 brands. Even though the focus of the current paper is to study the spending pattern of brands that advertise, it is interesting to note that 1855 brands never advertised during the considered three years, even though they were in the market for the entire period.

Among the 748 brands that did advertise at least once, considerable variability exists in their advertising behavior, as was already indicated in figure 2. First, the number of advertised weeks varies greatly. About one out of ten brands advertised only once, while a few brands (6 in total) advertised every week. However, the distribution is quite skewed, as nearly half of the brands advertised less than 10% of the time (see figure 6).

Figure 6. Advertising weeks.



Table 2 also provides evidence for the large variability in actual spending, even when conditioning on those weeks where a given brand advertises. Some brands typically spend large amounts (with an average level of £814,536 per week), while others spend only a limited amount per week (£19).

 Table 2. Advertising behavior of the Included Brands

	Range	Average	Standard deviation
Number of Advertising Weeks	1-156	37	42
Average Spending per Advertising Week (in £)	19-814,536	56,756	72,150

Combining both dimensions (using a median split on, respectively, the number of weeks of non-zero spending and the average value of such non-zero spending), we observe two main types of advertisers in Figure 7: Heavy advertisers, who spend large amounts per week for multiple weeks; and Light advertisers, limiting themselves to fewer actions and smaller amounts. Even though the heavy advertisers in the top-left cell only account for 39.57% of the brands, they represent 96.15% of the total advertising over the 3-year estimation period. Of these 296 brands, only 140 would comply with the Steenkamp et al. (2005) decision rules. These 140 brands account for 63.76% of all advertising expenditures. The remaining 11 of the 151 brands which would comply with these rules would account for only 0.16%. Limiting ourselves to these 151 brands would hence result in a loss of information on more than 1/3 of all advertising expenditures by branded products, in which especially the focus on large brands appears to have

strong consequences for the amount of advertising that is covered by the analyses. However, even though infrequent advertisers account for only a small part of total advertising expenditures, they still represent a large number of players, and hence in total as well a large number of advertising actions. By including them as well in our analyses, we can provide a better understanding of advertising in the market as a whole, thereby also providing evidence on how behavior may differ for smaller vs larger and more vs less frequently acting brands.





Numbers between brackets show the number of brands which comply with the Steenkamp et al. (2005) decision rules.

As a second main data source, we had access (through TNS) to consumer panel purchase data covering all purchases for over 17,000 families. These were used to calculate (i) the market shares of the different players, (ii) the category concentration and, (iii) the extent of category growth. Finally data on new product introductions (see e.g. Sorescu and Spanjol, 2008) were obtained through ProductScan[©].

Measurement

We now turn to the measurement of the different constructs. In section 4, we already provided an in-depth discussion of the *Adstock* and *Ad-sensor* concepts. The brand-specific lambdas were obtained by a Bayesian regression estimation of a partial adjustment sales model (Hanssens et al., 2001 p. 147), thereby allowing for correlations between brands' sales within the same category. An overview of the procedure can be found in appendix D. To account for the uncertainty in the estimated lambdas, we use each of the 60,000 draws of the lambdas obtained

after burn-in of 30,000 draws to calculate brand-specific Adstock and Ad-sensor series. These 60,000 series are subsequently used in the actual model estimation, with a burn in of 30,000 and sample of 30,000 draws for inference.

Advertising frequency equals the percentage of time the brand was advertised during the 52-week initialization period. In the choice of the data range to determine this factor, the aim is to match as closely as possible the estimation sample. However, as we analyze advertising decisions, Advertising frequency could suffer from endogeneity problems when estimated on the 156-week estimation period. This made us opt to determine the factor based on the 52-week initialization period. These endogeneity problems, however, are not as severe for *Brand market share,* which is consequently defined as the average market share over the 156-week estimation period.

The first four weeks a new product is on the market, *New product introduction* will equal 1; the other weeks 0. Similarly, the last four weeks of the year, the *End of year budget depletion* dummy variable will equal 1; the other weeks 0. *Competitive adstock* is defined as the weighted average of the competitors' Adstock values. The weights are dynamic, and based on the market share in volume terms over the previous 26 weeks (see e.g. Nijs et al., 2001; 2007). *Relative performance evolution* is expressed by the first difference of the logarithm of the brand volume shares (see e.g. Deleersnyder et al. 2004), also defined over a moving window of the previous 26 weeks. *Category growth* is measured as the first difference of the log-transformed category volume sales (cfr Franses and Koop, 1998). Finally, the Herfindahl index of volume shares over the previous 26 weeks period is used to quantify the *Category concentration*. All operationalizations are summarized in appendix E.

6. Empirical analysis

The coefficient estimates are shown in tables 3 and 4. They show the 95% posterior density intervals for the estimates, the latter printed in bold if zero is not included in the interval.

Adstock Management

Adstock. Table 3 shows the expected positive effect of Adstock on the likelihood of advertising actions ($\overline{\zeta}_{1,1,0} = 0.241$). The conclusions of the stream of normative literature that advertising in most instances can best be scheduled in pulses or campaigns, appear to be adopted

by the market. Brands show a clear state dependency, with periods of advertising during which the Adstock is rebuilt followed by periods during which it is allowed to deplete again. This state dependency, moreover, is also shown by the positive effect of Adstock on the magnitude of advertising actions ($\overline{\omega}_{1,1,0} = 0.011$). More intense actions during the previous weeks of an advertising campaign, resulting in higher Adstock levels, will be followed by higher spending in subsequent actions, implying that brands either opt for consistently high or low intensity campaigns.

	Timing							
		2.5th percentile	median	97.5 th percentile		2.5 th percentile	median	97.5 th percentile
Adstock	$\overline{\zeta}_{1,1,0}$	0.198	0.241	0.289	$\overline{\omega}_{\mathrm{l,l,0}}$	0.006	0.011	0.017
x Brand market share	$\overline{\zeta}_{1,1,1}$	0.512	1.00 9	1.529	$\overline{\omega}_{1,1,1}$	-0.028	0.004	0.035
x Advertising frequency	$\overline{\zeta}_{1,1,2}$	0.638	0.781	0.933	$\overline{\omega}_{\mathrm{l,l,2}}$	0.028	0.041	0.052
Ad-sensor	$\overline{\zeta}_{1,2,0}$	0.107	0.130	0.150	$\overline{\omega}_{1,2,0}$	0.066	0245	0.454
x Brand market share	$\overline{\zeta}_{1,2,1}$	0.417	0.663	0.915	$\overline{\omega}_{1,2,1}$	-0.002	0.019	0.041
x Advertising frequency	$\overline{\zeta}_{1,2,2}$	0.131	0.208	0.295	$\overline{\omega}_{\mathrm{l},2,2}$	0.009	0.017	0.025
Brand market share	$\overline{\zeta}_{1,0,1}$	-0.089	0.822	2.325	$\overline{\textit{\textit{0}}}_{\mathrm{l},0,1}$	0.066	0.245	0.454
Advertising frequency	$\overline{\zeta}_{1,0,2}$	4.735	5.192	5.586	$\overline{\omega}_{1,0,2}$	0.512	0.592	0.670

Table 3. Adstock Management: parameter estimates

Ad-sensor shows a significant positive effect in the starting and stopping of advertising campaigns ($\overline{\zeta}_{1,2,0} = 0.130$). When the Ad-sensor values become too high, a brand will start advertising. The closer to the desired maximum goodwill level in the campaign, the more likely the brand will stop advertising. Once the target has been reached, negative values of the Adsensor will increase the pressure to stop advertising even further. Over time, however, the pressure to start a new campaign will start building again. Moreover, as long as one is still far away from the goal level of Adstock for that specific campaign, i.e. when the Ad-sensor has relatively higher levels, it will also remain a source of pressure to spend more in order to reach the desired level faster ($\overline{\omega}_{1,2,0} = 0.006$). Once the desired target has been reached, it gives clear indications to no longer spend large amounts in case the brand would still decide to advertise. This also captures quite well the phenomenon of make goods: the target level was reached, but still some smaller remaining advertising can still be found at the end of the campaign.

The effects of Adstock and Ad-sensor for our numerical example, as well as their combined effect, are represented in figure 8. This is an illustration for an 'average' brand, with zero-effects of the mean-centered moderators. The first panel of figure 8 shows the effects on the timing decision, in which the period during which advertising is taking place is indicated by the grey zone. The second panel subsequently shows the effects on the magnitude decision



Figure 8. Effects of Adstock and Ad-sensor on Timing and Magnitude decision for our example

Moderators of Adstock and Ad-sensor

Brand market share. The larger brands in our set are likely to spend more on individual advertising actions ($\overline{\omega}_{1,0,1} = 0.245$). As marketing budgets are often set as a percentage of sales of previous periods (e.g. Allenby and Hanssens, 2005), more powerful brands will have larger budgets at their disposal, leading to more intense behavior. These brands, however, do not show differential baseline preferences in relation to their timing decisions.

Larger brands are better able to respond to their internal advertising pressure relative to smaller brands. Our results show that this is certainly the case in the timing decision ($\overline{\zeta}_{1,1,1} = 1.009$, $\overline{\zeta}_{1,2,1} = 0.663$). However, in the magnitude decision, no significant effects could be found.

Advertising frequency. Brands advertising more frequently in the past will continue to do so ($\overline{\zeta}_{1,0,2} = 5.192$). This is in line with previous research showing that current brand actions are often strongly influenced by previous behavior due to inertia effects (e.g. Frederickson and Iaquinto, 1989; Nijs et al., 2007). In addition, these brands will also spend more when advertising ($\overline{\omega}_{1,0,2} = 0.592$), confirming that basically two types of advertisers can be found: high-intensity (advertising often, spending more per decision) and low-intensity (advertising less often, spending less on single actions), as was already argued in the data section.

Our findings confirm the hypotheses that the effects of both Adstock and Ad-sensor are stronger for more experienced brands, and this for both the timing and magnitude decisions $(\overline{\zeta}_{1,1,2} = 0.781 \text{ and } \overline{\omega}_{1,1,2} = 0.041, \text{ respectively}; \overline{\zeta}_{1,2,2} = 0.663 \text{ and } \overline{\omega}_{1,2,2} = 0.017, \text{ respectively}).$

The resulting effects for both Adstock and Ad-sensor are depicted in figures 9 and 10. Figures are based on low (10) vs high (50) Adstock and low (-5) vs high (10) Ad-sensor values. The graphs in figure 9 show the effect of Adstock and Ad-sensor in the Timing decision for low vs high values of Brand market share and Advertising frequency. The graphs in figure 10 show the effects for the Magnitude decision. Moderating effects are reasonably strong for the Adstock effect. Larger and more experienced brands show much stronger campaigning behavior (panels 9a and 9c). They have more and better means and capabilities to pursue longer campaigns. More experienced brands will also be more likely adopters of consistently high or low advertising campaigns, as they as well show higher state dependence in their spending decision (panel 10c). Similarly, larger and more experienced brands can better track their Ad-sensor, and will more strongly react to its evolution in their decision on whether or not to start/stop a campaign (panels 9b and 9d). In the beginning of a campaign, Ad-sensor is high. Combined with a strong response coefficient for more experienced brands, this will cause a faster and stronger build-up at the beginning. The combination of a strong response coefficient and negative Ad-sensor values after the target level was reached, in turn, will result in a strong tendency of more experienced brands to refrain from investing relatively large amounts after the target level was attained (panel 10d).

Figure 9. Effect of Adstock and Ad-sensor on timing decision as a function of Brand market share and Advertising frequency



Figure 10. Effect of Adstock and Ad-sensor on magnitude decision as a function of Brand market share and Advertising frequency



Impact of baseline advertising preference drivers and covariates

Table 4 reports the parameter estimates of time-invariant baseline advertising preference drivers and the set of time-varying covariates we included in our analysis. We do not discuss the effects of Brand market share and Advertising frequency as they were already reported above, but

include them for reasons of completeness. For similar reasons, we also report the effects of the product class variables which were added to the model as control variables.

	Timing			Timing							Magnitud	e
		2.5 th		97.5 th		2.5 th		97.5 th				
		percentile	median	percentile		percentile	median	percentile				
Brand market share	$\overline{\zeta}_{1,0,1}$	-0.089	0.822	2.325	$\overline{\omega}_{1,0,1}$	0.066	0.245	0.454				
Advertising frequency	$\overline{\zeta}_{1,0,2}$	4.735	5.192	5.586	$\overline{\omega}_{\mathrm{l},0,2}$	0.512	0.592	0.670				
Food	$\overline{\zeta}_{1,0,3}$	-0.641	-0.319	-0.022	$\overline{\omega}_{\mathrm{l},\mathrm{0,3}}$	-0.073	-0.016	0.042				
Drinks	$\overline{\zeta}_{1,0,4}$	-0.028	0.273	0.587	$\overline{\omega}_{\mathrm{l},0,4}$	-0.069	-0.020	0.026				
Cosmetics	$\overline{\zeta}_{1,0,5}$	-0.081	0.319	0.666	$\overline{\omega}_{\mathrm{l},0,5}$	0.005	0.062	0.121				
Competitive Adstock	$\overline{\zeta}_{1,3,0}$	-0.033	-0.011	0.016	$\overline{\omega}_{1,3,0}$	-0.006	-0.002	0.002				
Relative Performance Evolution	$\overline{\zeta}_{1,4,0}$	-0.152	-0.066	0.075	$\overline{\omega}_{\mathrm{l},4,0}$	-0.050	-0.005	0.026				
New Product Introduction	$\overline{\zeta}_{1,5,0}$	-0.138	0.412	0.807	$\overline{\omega}_{\mathrm{l},\mathrm{5},\mathrm{0}}$	-0.013	0.039	0.082				
End of Year Remaining Budget	$\overline{\zeta}_{1,6,0}$	-0.324	-0.252	-0.177	$\overline{\omega}_{\mathrm{l},6,0}$	-0.030	-0.017	-0.008				
Category Growth	$\overline{\zeta}_{1,7,0}$	7.583	8.310	8.652	$\overline{\omega}_{\mathrm{l},\mathrm{7,0}}$	0.736	0.908	0.983				
Market Concentration	$\overline{\zeta}_{1,8,0}$	-3.653	-3.340	-3.159	$\overline{\omega}_{\!\scriptscriptstyle 1,8,0}$	-0.148	-0.073	0.005				
Market Concentration	€ _{1,8,0}	-3.653	-3.340	-3.159	$\omega_{\scriptscriptstyle 1,8,0}$	-0.148	-0.073	0.005				

Table 4. Baseline advertising drivers and covariates: parameter estimates

Competitive Adstock. Neither in timing nor size of their advertising actions, brands seem to be guided by their competitors. Although it has been shown that advertising clutter lowers advertising effectiveness (e.g. Villas-Boas, 1993; Danaher et al., 2008), brands do not refrain from spending when competitive advertising actions are likely, neither do they engage in a competitive escalation as was argued by Metwally (1978). Although perhaps surprising to some, this is in line with research by Steenkamp et al. (2005), indicating that little evidence of competitive reactions in advertising could be found to sudden competitive advertising shocks. Brands appear not to retaliate with a new own advertising action to those of competitors, nor will they spend more when advertising.

Relative performance evolution. Brands that perform well and gain market share have better means to compete. Larger marketing budgets become available (e.g. Allenby and Hanssens, 2005), enabling them to advertise more often and spend more. Conversely, the idea prevails that brands act in order to correct for a negative performance evolution relative to competitors (e.g. Metwally, 1978; Armstrong and Collopy, 1996). These theories are countered by our findings, as brands do not react with increased spending to make up for short term negative sales evolutions. Advertising budgets, on the other hand, are not adjusted immediately according to increased sales, leading to overall insignificance of short term performance evolution on advertising behavior. Advertising thus proves to be a strategic means of competition rather than a short term tactic means.

New product introduction. Advertising has been shown to be most effective for new products (Lodish et al., 1995) as it, e.g., increases trial probability of such products (Steenkamp and Gielens, 2003). New products still need to persuade customers into buying them. As advertising is an effective means to build awareness and convey product information, new products should be advertised more heavily (Tellis, 2004; Kotler and Armstrong, 2005). However, although our findings point in that direction, no such significant effects could be found on the actual advertising decisions.

End of year budget depletion. Advertising budgets are mostly set on a yearly basis, based on rules of thumb (e.g. percentage of sales of the previous year), formal advertising response modeling and management judgments (Farris and West, 2007). During the year, these budgets are used for advertising campaigns, driven by a wide set of factors (e.g. Montgomery et al., 2005). Thus, these financial resources become depleted. Our results indicate that managers tend to spend relatively more in the beginning of the year, making them advertise less often and spending less on single actions at the end ($\overline{\zeta}_{1,6,0} = -0.252$ and $\overline{\omega}_{1,6,0} = -0.017$), after accounting for Holiday season spending. Having money seemingly leads to spending it. In the beginning of the year, resources are still large, so one can more easily engage in more intense actions. As resources get depleted, one has to become more careful in when and what to spend, especially if one has spent relatively more in the beginning and hence has already used 'too much' of the resources.

Category growth. When category growth is higher, brands will be more inclined to advertise ($\overline{\zeta}_{1,7,0} = 8.310$), and their subsequent actions will as well be more intense ($\overline{\omega}_{1,7,0} = 0.908$). High growth categories are often younger product categories, requiring more advertising to inform and convince new customers (Narayanan et al., 2005). Consumers in more mature markets, with lower to zero growth, relay mostly on own experiences, and pay less attention to

advertising (Chandy et al., 2001). Higher growth, in addition, can be considered an indicator of potential future profits, leading brands to defend their position even more fiercely (Gatignon et al., 1990).

Category concentration. Economic theory tells that more concentrated markets show higher profits as such oligopolistic markets are often characterized by barriers to entry (e.g. Bain, 1951; Modigliani, 1958; Karakaya and Stahl, 1989). Combined with the easy monitoring of competitors' actions in such markets, this may lead to collusive behavior and the use of non-price forms of competition such as advertising in order not to compete away these attractive margins (Ramaswamy et al., 1994; Lipczynski and Wilson, 2001). Our findings, however, indicate that brands appear to be less inclined to advertise in more concentrated categories ($\overline{\zeta}_{1,8,0} = -3.340$). This is in line with previous research by Steenkamp et al. (2005) showing less overall competitive interaction behavior in such categories. Brands thus advertise less often, although no effect on their actual spending when advertising could be found.

Validation

To find guidance on the relative performance of our model, we compare the proposed model (Model 0) to two other plausible specifications. In the first competing model (Model 1), we restrict all covariances between error terms to zero. Model 2 is specified with only the Adstock level but without the Ad-sensor, thus only accounting for the state dependency and not for the pressure to start or stop advertising.

To compare these models, we assess to what extent they are capable of predicting both timing and magnitude of advertising actions. We compare the models regarding their performance on four different prediction statistics. The first statistic we consider, is the Mean Squared Error, one of the most widely used loss functions in statistics. Theil U allows us to judge the performance of the models relative to a naïve no-change model. The closer the value to zero, the better the model performance over the naïve no-change model in which $y_{cbt} = y_{cbt-1}$. We implement the so-called U2 specification (Theil, 1966), which allows to make a distinction between models performing better or worse, as it allows values beyond 1. The hit rate provides evidence on the percentage correctly predicted Timing (0/1) decisions. Finally, we report the correlation between the predicted and observed advertising expenditures. Higher values for the

latter two statistics prove better fit of the model. The best values on each summary statistic are underlined

The model performance statistics are considered both in- and out-of-sample. Indeed, as Van Heerde and Bijmolt (2005) argue, such time-based split provides us with a model robustness check as the estimation and validation samples may differ in a systematic way. Parameter estimates are based on a 130 week ($2\frac{1}{2}$ year) estimation period. The remaining 26 weeks ($\frac{1}{2}$ year) are used as a time hold-out sample. In the spirit of e.g. Brodie and de Kluyver (1987), we include *observed* Competitive Adstock, i.e. we assume that competitive actions are known. An overview of these statistics is given in table 5.

Table 5. Model Comparison

Model	Description		In-Sa	mple	
		MSE	Theil's U	Hit Rate	Correlation
Model 0	Proposed model	<u>0.078</u>	<u>0.552</u>	0.891	<u>0.779</u>
Model 1	No correlations	0.120	0.684	<u>0.893</u>	0.675
Model 2	No Ad-sensor	0.094	0.608	0.870	0.739
Model	Description		Out-of-	Sample	
		MSE	Theil's U	Hit Rate	Correlation
Model 0	Proposed model	<u>0.141</u>	<u>0.749</u>	0.829	<u>0.588</u>
Model 1	No correlations	1.144	2.130	<u>0.839</u>	0.214
Model 2	No Ad-sensor	0.152	0.770	0.810	0.540

In-sample performance

Our specification (Model 0) outperforms the alternative specifications on three diagnostics with regard to in-sample estimations. Although the hit rate (89%) is slightly lower than in model 1, our result is still impressive, especially when compared to the expected result for a random model. Given 24% of the observations showing advertising actions, a random model would show an overall hit rate of no more than 64% (Morrison, 1969). By means of such random model, we would a priori choose to classify $\alpha = 24\%$ of the observations as actions. The

observations, in turn, have an a priori probability of being an advertising action of p = 24%. The resulting hit rate would then equal $p*\alpha + (1-p)*(1-\alpha) = 0.24*0.24+0.76*0.76 = 0.64$. In addition, when we decompose our hit rate into the percentage of correctly predicted actions and no-actions, we obtain results of 71% and 95% respectively, far beyond the expected values of 24% and 76%. Besides a good predictor of the timing decisions, our model as well proves high in-sample fit when correlating the predicted with the actually observed expenditures, with a correlation equal to 0.779.

Out-of-sample performance

The second part of table 5 summarizes the out-of-sample performance statistics of our validation models. Here as well, our focal model 0 outperforms the other models on 3 out of the 4 statistics. The hit rate (about 83%) and correlation (0.588) are still impressive. Decomposition of the hit rate shows correct predictions of action in 60% of the cases, and a correct prediction of non-actions in 90% of the cases. These are still far beyond the aforementioned expected values of 24% and 76% of a random model (Morrison, 1969). Overall, this provides evidence of the excellent predictive validity of our model specification.

Timing and Magnitude

In our study, we allowed for differential decisions processes in the timing and magnitude decisions. We already provided evidence of the good performance of our model in predicting the timing decisions by means of an impressive hit rate. In addition, we reported the overall correlation between observed and predicted advertising values. To more specifically assess the relative value of the magnitude part, we now restrict ourselves to those observations for which the outcome of the timing decisions was positive. Thus we obtain the following in-sample (out-of-sample) statistics: MSE = 0.199 (0.330); Theil U = 0.431 (0.548); correlation = 0.512 (0.407). These results indicate that the large share of correctly predicted zero advertising in the overall sample – and thus the good performance of the timing part – may inflate the overall correlation. Our model consequently performs better in predicting the timing of the advertising actions relative to predicting the actual amounts spent when advertising.

Based on these analyses, we can conclude that our model provides an encouraging mimicking of advertising decisions in the marketplace. Not accounting for the correlations between the error terms generally worsens the performance of our model. Finally, the inclusion of our Ad-sensor, capturing the process leading to the start and stop of an advertising campaign, clearly adds value compared to models including only the Adstock level.

8. Discussion

Summary

Notwithstanding the fact that advertising is one of the most used marketing tools, little is known about what is driving the timing and magnitude of advertising actions. Building on normative theory, we developed a parsimonious model that captures this dual investment process. We explained advertising spending patterns as observed in the market, and investigated the impact of company, competitive, and category-related factors on these decisions. Analyses were based on a combination of (i) weekly advertising data from a wide variety of CPG brands from the UK, (ii) household panel purchase data, and (iii) data on new product introductions. We included both large and small brands, both frequent and infrequent advertisers in order to avoid data pruning biases, and to obtain a more complete overview of the market.

The empirical findings provide broad support for our conceptual framework. Adstock and Ad-sensor have a positive effect on both timing and magnitude decisions. Advertising spending patterns can hence be considered a result of the application of campaigning strategies, based on dynamic Adstock management systems. The extent to which such strategies are more or less the outcome of such systems, however, varies across brands as a function of their size and Advertising frequency, a clear indication that inclusion of all types of brands in our dataset was warranted.

Discussion

In this study, we explained observed advertising spending patterns. Even strategies with same amounts of advertising in every single week, as were advocated by early normative studies (e.g. Zielske, 1959; Sasieni, 1971; 1989), are not observed in real world data. Pulsing strategies on the other hand, characterized by an alternation of periods with higher advertising and periods

with lower to zero advertising (e.g. Mahajan and Muller, 1986) emerge as a dominant strategy. Differing from previous advertising research, we argued these advertising investments to be the outcome of a dual decision process, thereby distinguishing between two conceptually different but at the same time closely linked decisions: whether to advertise or not (timing), and conditional upon the choice to advertise, how much to spend (magnitude). As our results indicated, differentiating between these two decisions is warranted. They are partly driven by different factors, and common drivers have clearly different weights in both decisions. Accounting for differential processes thus allows us to obtain a much richer view as well as a more correct understanding of what is driving observed behavior.

Both decisions were subsequently embedded in an advertising goodwill stock management system. Advertising investments are scheduled in campaigns of several weeks, during which brands will build advertising goodwill among consumers. Carry-over effects of advertising allow for longer periods without advertising, during which the goodwill will gradually decay. As such, the created advertising goodwill is managed in a dynamic way. Similar to strategies in other inventory management, this process can be described as an (s,S) inventory management system, by means of which brands in a systematic way monitor and adjust their advertising goodwill stock. Brands will decide to launch a new advertising action as soon as this stock falls below a certain minimum desired level s, and stop advertising once the desired (higher) target level S was reached. The Ad-sensor concept we developed in this study captures this pressure, not only to launch and stop a new campaign, but also to spend more when still far away from the target level in order to reach it faster. As such, it provides evidence of the dynamics associated with advertising campaigns. The state dependencies of such campaigns, i.e. the fact that they last for several weeks and that brands prefer consistent spending strategies within each campaign, in turn, are well represented by the Adstock concept. Together, Ad-sensor and Adstock allow us to explain observed advertising spending patterns as real-life applications of the normative literature, in which advertising and advertising goodwill management are embedded in dynamic (s,S) inventory systems.

However, the extent to which such systems affect the actual advertising decisions clearly differs among different types of brands. Larger brands, as well as more frequently advertising brands, show a greater reactance to changes in their advertising goodwill. The former have better means to do so, whereas the latter may have learned from previous experience, resulting in a

closer advertising goodwill stock management. Both findings, taken together, provide clear evidence that limiting the analyses to only the frequently acting large brands would have biased our results. Our flexible model specification, in contrast to previous work, allowed us to cover nearly all players in the market, thus avoiding data pruning biases. Given the number of brands that otherwise would have been excluded from our analyses, this certainly gives extra weight to our empirical generalizations.

Although advertising is often argued to be driven by competitive reasoning and reaction, no such tendencies could be found. Not only did we not find any direct effects of competitive advertising in the advertising decision making process, indirect effects through the relative performance evolution of the brands did not show significant impact either. This may seem even more surprising, as annual advertising budgets are often determined on a percentage of sales basis (e.g. Allenby and Hanssens, 2005), and given the well-known argument that companies advertise in order to retain market shares (e.g. Metwally, 1978; Armstrong and Collopy, 1996). Advertising, characterized by small short run elasticities (e.g. Assmus et al., 1984; Sethuraman and Tellis, 1991), takes time to build an effect in the mindset of customers (e.g. Assael, 1998; Kardes, 2004). Price promotions, on the other hand, show much higher short run elasiticities (Neslin, 2002; Bijmolt et al., 2005), making it a more interesting instrument to make up for negative sales evolutions in the short run. This is also in line with the findings by Steenkamp et al. (2005) who show that, in contrast to price promotions, advertising will hardly be used as a means to react to competitors. Advertising thus appears not to be a tactic means to counter negative evolutions in the short run, but rather a strategic means which builds goodwill that lasts. Anecdotal evidence from practitioners adds to our findings, as several advertising agency account managers confirmed that brands, in general, focus on their own internal advertising utility calculi, and much less on what their competitors are doing in their actual advertising decisions.

Recent research has shown that new product introductions, combined with heavier advertising, can result in higher sales and increased shareholder value (see e.g. Pauwels et al., 2004; Srinivasan et al., 2009), combined with a rejuvenated brand (Slotegraaf and Pauwels, 2008). This may as well be a consequence of the fact that advertising is especially effective for those new products (e.g. Lodish et al., 1995; Steenkamp and Gielens, 2003). However, no significant link between such introductions and advertising decisions could be found, although there were some weak indications in that direction. Overall, adopted strategies appear to be

continued in a similar fashion, without granting much extra support to new products relative to those that are already present in the portfolio. Brands which already have intense advertising strategies in place, are likely to continue these strategies for the new products as well, whereas those with lower intensity strategies may simply not have the means to increase their efforts.

Directions for future research

In this work, we investigated and explained advertising spending patterns as observed in the market. We did not aim at modeling the particular decision process of individual managers, but mimic the advertising decisions as observed in the market by means of a paramorphic model (e.g. Hoffman, 1960; Slovic and Lichtenstein, 1971; Steenkamp, 1989). Past research has proven that such models, although showing deviations when applied to specific individual reasoning processes, perform very well in capturing judgment and decision processes at a higher level (e.g. Einhorn et al., 1979). Thus we are able to capture the phenomenon that is taking place in the market, abstracting from short term individual deviations. However, this provides an avenue for future research focusing on how individual managers in practice decide on their advertising actions, and this by means of in-depth personal interactions with individual managers.

A second limitation of our work, is the fact that we limit ourselves to the dual decision of timing and magnitude, thereby aggregating over all media. Media choice as such is not investigated. This, however, would represent an interesting area for future research, as not all media show the same effectiveness for different product categories, and as synergy effects can be present in multimedia communications (see e.g. Naik and Raman, 2003). Moreover, media effectiveness is not static, but does evolve as well as a consequence of the appearance of new media. These phenomena can have far-reaching implications for the advertising decision process and render it even more complicated.

Computing limitations (with the present runtime of our model on a fast Dell Precision Workstation equaling 30 days) made us opt to include only time-invariant moderators in our model. However, the inclusion of time-varying moderators like price-promotions would enrich our analyses even further, and would also enable us to capture the interplay between Adstock and advertising management and other instruments of the marketing mix.

Finally, we showed that commonly applied data pruning rules may engender results which are only valid for that specific subsample of the market, raising questions to the validity of

the corresponding conclusions for the market as a whole. New econometric techniques enable us to include all different types of players, even those that would have been excluded by these pruning rules. An investigation as to what extent findings of previous research are valid for the market as a whole can therefore be suggested.

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Appendix A: Mathematical derivation of Ad-sensor

Ad-sensor captures the dynamics in the evolution of the brand's Adstock. This Adstock is defined as (Broadbent, 1984):

(A1)
$$Adstock_{b,t} = (1 - \lambda_b)Advert_{b,t} + \lambda_bAdstock_{b,t-1}$$

The first order condition for the optimum, i.e. the maximum Adstock level during a campaign, given that we use discrete time observations and we can thus only observe up to time t-1, is:

(A2)
$$\frac{\Delta A dstock_{b,t-1}}{\Delta t} = 0$$

As we are analyzing discrete time data, this yields

(A3)
$$\frac{Adstock_{b,t-1} - Adstock_{b,t-2}}{\Delta t} = 0$$

(A4)
$$\Leftrightarrow Adstock_{b,t-1} - Adstock_{b,t-2} = 0.$$

Given (A1), this is satisfied if:

(A5)
$$(1 - \lambda_b) A dvert_{b,t-1} + \lambda_b A dstock_{b,t-2} - A dstock_{b,t-2} = 0$$

(A6)
$$\Leftrightarrow (1 - \lambda_b) A dvert_{b,t-1} = (1 - \lambda_b) A dstock_{b,t-2}$$

(A7)
$$\Leftrightarrow Advert_{b,t-1} = Adstock_{b,t-2}$$

The second order condition for this maximum at t-1 then requires:

(A8)
$$Advert_{b,t-2} > Adstock_{b,t-3}$$
.

because Adstock will increase as long as Advertising is larger than Adstock:

(A9)
$$Adstock_{b,t-1} > Adstock_{b,t-2}$$

(A10)
$$\Leftrightarrow (1 - \lambda_b) A dvert_{b,t-1} + \lambda_b A dstock_{b,t-2} > A dstock_{b,t-2}$$

(A11)
$$\Leftrightarrow (1 - \lambda_b) A dvert_{b,t-1} > (1 - \lambda_b) A dstock_{b,t-2}$$

(A12)
$$\Leftrightarrow Advert_{b,t-1} > Adstock_{b,t-2}$$

The starting point for our Ad-sensor at time t is hence:

(A13)
$$Advert_{b,t-1} - Adstock_{b,t-2} > 0$$
 if the maximum is reached after time t-1;

= 0 if the maximum is reached in t-1;

< 0 if the maximum was reached before t-1.

We subsequently rewrite (A1) in function of $Advert_{b,t-1}$:

(A14)
$$Advert_{b,t-1} = \frac{Adstock_{b,t-1} - \lambda_b Adstock_{b,t-2}}{(1 - \lambda_b)},$$

which yields:

(A15)
$$Advert_{b,t-1} - Adstock_{b,t-2} = \frac{Adstock_{b,t-1} - \lambda_b Adstock_{b,t-2}}{(1 - \lambda_b)} - Adstock_{b,t-2}$$

(A16)
$$\approx Adstock_{b,t-1} - \lambda_b Adstock_{b,t-2} - (1 - \lambda_b) Adstock_{b,t-2}$$

or

(A13)
$$Advert_{b,t-1} - Adstock_{b,t-2} \approx Adstock_{b,t-1} - Adstock_{b,t-2}$$

We therefore define our Ad-sensor variable as the difference between Adstock in time *t*-1 and Adstock in *t*-2:

(A17)
$$Adsensor_{b,t} = Adstock_{b,t-1} - Adstock_{b,t-2}$$

Ad-sensor thus captures the dynamics in the evolution of a brand's Adstock. In the beginning of a campaign, Adstock will increase very fast, causing high values of Ad-sensor. Closer to the maximum, increases will become smaller as Adstock approaches the Advertising values. As a consequence, the value of Ad-sensor starts to decrease. Once beyond the maximum, Adstock starts to decline, and Ad-sensor takes relatively strong negative values. Adstock decays at a constant rate λ , but not in constant absolute terms. When Adstock levels are still high, decay will be large in absolute terms, causing strong negative Ad-sensor values. Over time, the Adstock level becomes smaller, and decay will be smaller in absolute terms. The Ad-sensor values will become less negative, indicating a growing pressure to start advertising again.

Wook	Advorticing	Adetock	Ad concor	C	6
VVEEK	Auventising	AUSIOCK	70-2611201	3	3
1	0.00	0.00	0.00	5.00	197.65
2	0.00	0.00	0.00	5.00	197.65
3	0.00	0.00	0.00	5.00	197.65
4	120.00	36.00	0.00	5.00	197.65
5	200.00	85.20	36.00	5.00	197.65
6	240.00	131.64	49.20	5.00	197.65
7	230.00	161.15	46.44	5.00	197.65
8	240.00	184.80	29.51	5.00	197.65
9	210.00	192.36	23.66	5.00	197.65
10	210.00	197.65	7.56	5.00	197.65
11	190.00	195.36	5.29	5.00	197.65
12	110.00	169.75	-2.31	5.00	197.65
13	80.00	142.83	-25.61	5.00	197.65
14	50.00	114.98	-26.93	5.00	197.65
15	0.00	80.48	-27.85	5.00	197.65
16	0.00	56.34	-34.49	5.00	197.65
17	0.00	39.44	-24.15	5.00	197.65
18	0.00	27.61	-16.90	5.00	197.65
19	0.00	19.32	-11.83	5.00	197.65
20	0.00	13.53	-8.29	5.00	197.65
21	0.00	9.47	-5.80	5.00	197.65
22	0.00	6.63	-4.06	5.00	197.65
23	0.00	4.64	-2.84	5.00	197.65
24	0.00	3.25	-1.99	5.00	197.65
25	0.00	2.27	-1.39	5.00	197.65

Appendix B: Numerical example of the Adstock and Ad-sensor development

We first simulate an average advertising campaign. These advertising expenditures are given in the second column. We subsequently calculate the Adstock (with imposed carry-over $\lambda = 0.70$) and Ad-sensor series.

For period 6, these are:

Adstock6 = (1-0.70)*240.00-0.70*85.20 = 131.64 Ad-sensor6 = 85.20-36.00 = 49.20

For period 11, these are:

Adstock11 = (1-0.70)*190.00-0.70*197.65 = 195.36 Ad-sensor11 = 197.65-192.36 = 5.29 For period 12, these are:

Adstock12 = (1-0.70)*110.00-0.70*195.36= 169.75 Ad-sensor12 = 195.36-197.65 = -2.31

For period 15, these are:

Adstock15 = (1-0.70)*0.00-0.70*114.98 = 80.48 Ad-sensor15 = 142.83-114.98 = -27.85

For period 16, these are:

Adstock16 = (1-0.70)*0.00-0.70*80.48 = 56.34 Ad-sensor16 = 114.98-80.48 = -34.49

For period 20, these are:

Adstock20 = (1-0.70)*0.00-0.70*19.32 = 13.53 Ad-sensor20 = 19.32-27.61 = -8.29

The highest observed Adstock level in this series equals 197.65. We consequently assume that this was the desired maximum level S. In this simulation, we impose s=5. In practice, we do not know s and S, and derive these from the observed patterns. In addition, we estimate λ for each individual brand.

Appendix C: MCMC Estimation of a Hierarchical Multivariate Type-2 Tobit Model

We will first briefly repeat the model specification.

An advertising decision in category c by brand b in week t (z_{cbt}) is described by a multivariate probit model:

(C1)
$$z_{cbt} = \begin{cases} 1 & if \ z_{cbt}^* > 0 \\ 0 & otherwise \end{cases}$$

The latent variable z_{cbt}^* is modeled through a linear model:

(C2)
$$z_{cbt}^* = \mathbf{x}_{1cbt}^{'} \boldsymbol{\zeta}_{1cb} + \mathbf{x}_{2cbt}^{'} \boldsymbol{\zeta}_{2cb} + \boldsymbol{\mu}_{cbt}$$

Conditional on the decision to advertise ($z_{cbt} = 1$), we model y_{cbt} , the ln of spending by brand b in category c during week t as:

(C3)
$$y_{cbt} = \mathbf{v}_{1cbt}\mathbf{\omega}_{1cb} + \mathbf{v}_{2cbt}\mathbf{\omega}_{2cb} + \varepsilon_{cbt}$$

where

(C4)
$$(\mathbf{\epsilon}'_{ct}, \mathbf{\mu}'_{ct})' \sim MVN(0, \mathbf{\Sigma})$$

We relate the response parameters ζ_{1cb} and ω_{1cb} to a set of second stage variables.

(C5)
$$\boldsymbol{\zeta}_{1cb} = \mathbf{Q}_{cb} \overline{\boldsymbol{\zeta}} + \mathbf{u}_{cb}$$

(C6)
$$\boldsymbol{\omega}_{1cb} = \mathbf{R}_{cb}\overline{\boldsymbol{\omega}} + \mathbf{e}_{cb}$$

where

(C7)
$$(\mathbf{e}_{cb},\mathbf{u}_{cb})' \sim MVN(0,\mathbf{\Omega})$$
.

We stack the dependent variables of equations (C2) and (C3) for all brands *b* in category *c* and time periods *t* so that the vector of ln expenditures is $\mathbf{y}_c = [y_{c11}, y_{c12}, ..., y_{cB_cT}]'$ and the vector of advertising action indicator variables is $\mathbf{z}_c^* = [z_{c11}^*, z_{c12}^*, ..., z_{cB_cT}^*]'$, (b) the predictor variables for the advertising action equation, $\mathbf{v}_{1cbt}' = [v_{1cb1b}v_{1cb2b}..., v_{1cbMt}]'$ and $\mathbf{v}_{2cbt}' = [v_{2cb1b}v_{2cb2b}..., v_{2cbMt}]'$; the predictor variables for the ln expenditures equation, $\mathbf{x}_{1cbt}' = [x_{1cb1b}x_{1cb2b}..., x_{1cbMt}]'$ and $\mathbf{x}_{2cbt}' = [x_{2cb1b}x_{2cb2b}..., x_{1cbMt}]'$ and $\mathbf{x}_{2cbt}' = [x_{2cb1b}x_{2cb2b}..., x_{1cbMt}]'$ and $\mathbf{x}_{2cbt}' = [x_{2cb1b}x_{2cb2b}..., x_{2cbMt}]'$; and (c) the error terms of these two equations for all brands b and

time periods *t* so that $[\mathbf{\epsilon}'_{ct}, \mathbf{\mu}'_{ct}]' = [\varepsilon_{c_{1b}} \varepsilon_{c_{2b}} \dots, \varepsilon_{c_{B_{c}t}}, \mu_{c_{1t}}, \mu_{c_{2t}}, \dots, \mu_{c_{B_{c}t}}]'$ follows a $(2B_c)$ -variate normal distribution with zero mean and full covariance matrix

$$\boldsymbol{\Sigma}_{\mathbf{c}} = \begin{bmatrix} \boldsymbol{\Sigma}_{c11} & \boldsymbol{\Sigma}_{c12} \\ \boldsymbol{\Sigma}_{c12}' & \boldsymbol{\Sigma}_{c22} \end{bmatrix}$$

where $\Sigma_{c11} = E(\varepsilon_{ct}\varepsilon'_{ct})\forall c$, $\Sigma_{c12} = E(\varepsilon_{ct}\mu'_{ct})\forall c$, and $\Sigma_{c22} = E(\mu_{ct}\mu'_{ct})\forall c$, which has ones on the diagonal since each selectivity mechanism is binary.

Next we specify the hierarchies associated with the two advertising decisions. We stack (i) the parameter coefficients per category and per brand across equations (C2) and (C3) and (ii) the error terms of the hierarchical equations for all brands b in a similar way. We model the response parameters as follows:

$$\begin{bmatrix} \boldsymbol{\omega}_{1c1} \\ \boldsymbol{\zeta}_{1c1} \\ \boldsymbol{\omega}_{1c2} \\ \boldsymbol{\zeta}_{1c2} \\ \vdots \\ \boldsymbol{\omega}_{1cB_c} \\ \boldsymbol{\zeta}_{1cB_c} \end{bmatrix} = \begin{bmatrix} \mathbf{Q}_{c1} \\ \mathbf{R}_{c1} \\ \mathbf{Q}_{c2} \\ \vdots \\ \mathbf{Q}_{c2} \\ \vdots \\ \mathbf{Q}_{cB_c} \\ \mathbf{R}_{cB_c} \end{bmatrix} \begin{bmatrix} \overline{\boldsymbol{\omega}} \\ \overline{\boldsymbol{\zeta}} \end{bmatrix} + \begin{bmatrix} \mathbf{e}_{c1} \\ \mathbf{u}_{c1} \\ \mathbf{e}_{c2} \\ \mathbf{u}_{c2} \\ \vdots \\ \mathbf{e}_{cB_c} \\ \mathbf{u}_{cB_c} \end{bmatrix} \sim N_{|\mathbf{u}_c|+|\mathbf{e}_c|} (0, \mathbf{I}_{B_c} \otimes \boldsymbol{\Omega}),$$

with $\mathbf{Q}_{cb} = \mathbf{R}_{cb} = \mathbf{I}_M \otimes \mathbf{q}'_{cb}$, where \mathbf{I}_M is an (M x M) identity matrix and \mathbf{q}'_{cb} is a 1 x (N/M) vector of covariates. The hyperparameters relating these covariates to the actual first level response parameters are stacked in $[\overline{\boldsymbol{\omega}}' \overline{\boldsymbol{\zeta}}']'=$

$$\left[\omega_{1I,I},\omega_{1I,2}\ldots\omega_{1I,N/M},\omega_{12,I}\ldots\omega_{1M,N/M},\zeta_{1I,I},\zeta_{1I,2}\ldots\zeta_{1I,N/M},\zeta_{12,I}\ldots\zeta_{1M,N/M}\right], \text{ with } \overline{\boldsymbol{\omega}} \text{ and } \overline{\boldsymbol{\zeta}} \text{ both being (N x 1) vectors.}$$

$$\mathbf{\Omega} = \begin{bmatrix} \mathbf{\Omega}_{11} & \mathbf{\Omega}_{12} \\ \mathbf{\Omega}_{12}' & \mathbf{\Omega}_{22} \end{bmatrix} \text{ is a full covariance matrix where } \mathbf{\Omega}_{11} = E(\mathbf{e}_{cb}\mathbf{e}_{cb}') \forall c, b, \ \mathbf{\Omega}_{12} = E(\mathbf{e}_{cb}\mathbf{u}_{cb}') \forall c, b, \ \mathbf{\Omega}_{12} =$$

and $\mathbf{\Omega}_{22} = E(\mathbf{u}_{cb}\mathbf{u}'_{cb}) \forall c, b$.

We use an MCMC approach to estimate the marginal distributions of the latent dependent variables, parameters and covariances. The MCMC algorithm involves sampling sequentially from the relevant conditional distributions over a large number of iterations. These draws can be shown to converge to the marginal posterior distributions. Our implementation of the MCMC algorithm has 6 steps that are described below.

Conditional distributions

The first implementation step requires that we specify conditional distributions of the relevant variables. The solutions of these distributions follow from the normality assumption of the disturbances terms. We employ natural conjugate priors. Specifications of the conditional distributions are as follows:

1. y_{cbt}^* is y_{cbt} if $z_{cbt} = 1$, otherwise y_{cbt}^* is drawn from a normal distribution:

$$y_{cbt}^{*} | \mathbf{y}_{cj\neq b,t}^{*}, \mathbf{z}_{ct}^{*}, \mathbf{\omega}_{cb}, \alpha_{cb}, t_{cb} \Sigma$$

$$\approx \begin{cases} y_{cbt} | z_{cbt} = 1 \\ N(\mathbf{v}_{1cbt}' \mathbf{\omega}_{1cb} + \mathbf{v}_{2cbt}' \mathbf{\omega}_{2cb} + \mathbf{\sigma}_{cb,cj} \widetilde{\boldsymbol{\Sigma}}_{cj,cj}^{-1} \left(\mathbf{y}_{cj\neq b,t}^{*} - \mathrm{E}(\mathbf{y}_{cj\neq b,t}^{*}) \right), \sigma_{cb,cb}^{(11)} - \mathbf{\sigma}_{cb,cj}^{*} \widetilde{\boldsymbol{\Sigma}}_{cj,cj}^{-1} \mathbf{\sigma}_{cj,cb}^{*}) | | z_{cbt} = 0 \end{cases}$$
where $\mathbf{y}_{ct}^{*} = \begin{bmatrix} y_{cbt}^{*} \\ -- \\ \mathbf{y}_{cj\neq b,t}^{*} \end{bmatrix}$ and $\mathbf{\Sigma}_{c} = \begin{bmatrix} \sigma_{cb,cb}^{(11)} & | \mathbf{\sigma}_{cb,cj}^{*} \\ -- & -- \\ \mathbf{\sigma}_{cj,cb}^{*} & | \widetilde{\boldsymbol{\Sigma}}_{cj,cj}^{*} \end{bmatrix}$

As the notation suggests, the \mathbf{y}_{ct}^* vector and $\sum_{\mathbf{c}}$ matrix are partitioned between the brand of interest, *cb*, and all other brands *cj*, $j \neq b$ (the entries in $\sum_{\mathbf{c}}$ corresponding to \mathbf{z} are not shuffled). Without loss of generality, we have assumed the brand of interest to be the first. Each brand is then drawn in succession for category c, conditioning on $\mathbf{y}_{c,j\neq b,t}^*$, \mathbf{z}_{ct}^* , and $\sum_{\mathbf{c}}$.

2. We next draw the latent dependent variable values for the probit component of the model. If the indicator variable $z_{cbt} = 1$, then z_{cbt}^* is drawn from a normal distribution, truncated below at 0. Otherwise, z_{cbt}^* is drawn from a normal distribution, truncated above at 0.

$$\mathbf{z}_{cbt}^{*} \left| \mathbf{z}_{c,j\neq b,t}^{*}, \mathbf{y}_{ct}^{*}, \boldsymbol{\zeta}_{cb}, \boldsymbol{\alpha}_{cb}, \boldsymbol{\iota}_{cb}, \boldsymbol{\Sigma} \right|$$

$$\sim N_{T} \left(\mathbf{x}_{c1bt}^{\prime} \boldsymbol{\zeta}_{1cb} + \mathbf{x}_{2cbt}^{\prime} \boldsymbol{\zeta}_{2cb} + \boldsymbol{\sigma}_{cb,cj}^{\ast} \widetilde{\boldsymbol{\Sigma}}_{cj,cj}^{-1} \left(\mathbf{z}_{c,j\neq b,t}^{*} - \mathbf{E}(\mathbf{z}_{c,j\neq b,t}^{*}) \right), \boldsymbol{\sigma}_{cb,cb}^{\ast} - \boldsymbol{\sigma}_{cb,cj}^{\ast} \widetilde{\boldsymbol{\Sigma}}_{cj,cj}^{-1} \boldsymbol{\sigma}_{cj,cb}^{\ast} \right)$$
where: $\mathbf{z}_{ct}^{*} = \begin{bmatrix} \mathbf{z}_{cbt}^{*} \\ -- \\ \mathbf{z}_{cj\neq b,t}^{*} \end{bmatrix}$ and $\mathbf{\Sigma}_{c}^{\ast} = \begin{bmatrix} \boldsymbol{\sigma}_{cb,cb}^{(22)} & | \mathbf{\sigma}_{cb,cj}^{\ast} \\ --- & -- \\ \mathbf{\sigma}_{cj,cb}^{\ast} & | \mathbf{\Sigma}_{cj,cj}^{\ast} \end{bmatrix}$

The latent probit dependent variables are drawn using the inverse cdf method.

3. The parameters in $[\omega_{1c1}, \zeta_{1c1}, \omega_{1c2}, \zeta_{1c2}, ..., \omega_{1cB_c}, \zeta_{1cB_c}]'$ are drawn from a SUR model with variance/covariance matrix of disturbances Σ_c :

$$\begin{bmatrix} \mathbf{\omega}_{1c1} \\ \boldsymbol{\zeta}_{1c1} \\ \mathbf{\omega}_{1c2} \\ \boldsymbol{\zeta}_{1c2} \\ \vdots \\ \mathbf{\omega}_{1cB_c} \\ \boldsymbol{\zeta}_{1cB_c} \end{bmatrix} \mathbf{y}^{*(t)}, \mathbf{z}^{*(t)}, \mathbf{\Sigma}_{c}^{(t-1)} \sim N \left(\mathbf{O}_{c} \left(\mathbf{P}_{1c}^{'} \left(\mathbf{I}_{T_{c}} \otimes \mathbf{\Sigma}_{c}^{-1} \right) \right) \right) \left(\mathbf{I}_{T_{c}}^{*(t)} - \mathbf{v}_{2cl1}^{'} \mathbf{u}_{2c2} \\ \mathbf{z}_{cl1}^{*(t)} - \mathbf{v}_{2cl2}^{'} \mathbf{u}_{2c2} \\ \mathbf{z}_{cl2}^{*(t)} - \mathbf{v}_{2cl2}^{'} \mathbf{u}_{2c2} \\ \vdots \\ \mathbf{y}_{cB_{c}l}^{*(t)} - \mathbf{v}_{2cB_{c}l}^{'} \mathbf{u}_{2cB_{c}} \\ \mathbf{z}_{cB_{c}l}^{*(t)} - \mathbf{v}_{2cB_{c}l}^{'} \mathbf{u}_{2cB_{c}} \\ \mathbf{z}_{cB_{c}r}^{*(t)} - \mathbf{v}_{2cB_{c}l}^{'} \mathbf{u}_{2cB_{c}} \\ \mathbf{z}_{cB_{c}r}^{*(t)} - \mathbf{v}_{2cB_{c}r}^{'} \mathbf{u}_{2cB_{c}} \\ \mathbf{z}_{cB_{c}r_{c}}^{*(t)} - \mathbf{v}_{2cB_{c}r_{c}}^{'} \mathbf{u}_{2cB_{c}} \\ \mathbf{z}_{cB_{c}r_{c}}^{*(t)} - \mathbf{z}_{2cB_{c}r_{c}}^{'} \mathbf{u}_{2cB_{c}} \\ \mathbf{z}_{cB_{c}r_{c}}^{*(t)} - \mathbf{z}_{2cB_{c}r_{c}}^{'} \mathbf{u}_{2cB_{c}} \\ \mathbf{z}_{cB_{c}r_{c}}^{*(t)} - \mathbf{z}_{2cB_{c}r_{c}}^{'} \mathbf{u}_{2cB_{c}}^{'} \mathbf{u}_{2cB_{c}$$

where
$$\Sigma_{c} = \begin{bmatrix} \sigma_{y_{c1}y_{c1}} & \sigma_{y_{c1}z_{c1}} & \sigma_{y_{c1}y_{c2}} & \sigma_{y_{c1}z_{c2}} & \cdots & \sigma_{y_{c1}y_{cB_{c}}} & \sigma_{y_{c1}z_{cB_{c}}} \\ \sigma_{z_{c1}y_{c1}} & \sigma_{z_{c1}z_{c1}} & \sigma_{z_{c1}y_{c2}} & \sigma_{z_{c1}z_{c2}} & \cdots & \sigma_{z_{c1}y_{cB_{c}}} & \sigma_{z_{c1}z_{cB_{c}}} \\ \sigma_{y_{c2}y_{c1}} & \sigma_{y_{c2}z_{c1}} & \sigma_{y_{c2}y_{c2}} & \sigma_{y_{c2}z_{c2}} \\ \sigma_{z_{c2}y_{c1}} & \sigma_{z_{c2}z_{c1}} & \sigma_{z_{c2}y_{c2}} & \sigma_{z_{c2}z_{c2}} \\ \vdots & & & \ddots & \\ \sigma_{y_{cB_{c}}y_{c1}} & \sigma_{y_{cB_{c}}z_{c1}} & & & \sigma_{y_{cB_{c}}y_{cB_{c}}} & \sigma_{y_{cB_{c}}z_{cB_{c}}} \\ \sigma_{z_{cB_{c}}y_{c1}} & \sigma_{z_{cB_{c}}z_{c1}} & & & \sigma_{z_{cB_{c}}y_{cB_{c}}} & \sigma_{z_{cB_{c}}z_{cB_{c}}} \end{bmatrix},$$

and
$$\mathbf{O}_{c} = \left(\mathbf{P}_{1c}^{T}\left(\mathbf{I}_{T_{c}} \otimes \boldsymbol{\Sigma}_{c}^{-1(c-1)}\right) \mathbf{P}_{1c} + \mathbf{I}_{B_{c}} \otimes \boldsymbol{\Omega}^{-1}\right)^{-1}$$
,

$$\left| \begin{array}{c} \mathbf{v}_{1c11} & & & & & \\ & \mathbf{v}_{1c21} & & & & \\ & & \mathbf{x}_{1c21} & & & \\ & & & \mathbf{v}_{1cB_{c}1} & & \\ & & & & \mathbf{v}_{1cB_{c}1} \\ & & & & \mathbf{v}_{1cB_{c}2} \\ & & & & \mathbf{x}_{1cB_{c}2} \\ & & & & & \mathbf{v}_{1cB_{c}2} \\ & & & & & \mathbf{v}_{1cB_{c}2} \\ & & & & & & \mathbf{v}_{1cB_{c}7_{c}} \\ & & & & & & & \mathbf{v}_{1cB_{c}7_{c}} \\ & & & & & & & \mathbf{v}_{1cB_{c}7_{c}} \\ & & & & & & & & \mathbf{v}_{1cB_{c}7_{c}} \end{array} \right]$$

4. The vector of hyper-parameters, $\left[\overline{\omega}', \overline{\zeta}'\right]'$, is drawn from a SUR model with variance/covariance matrix of disturbances

$$\mathbf{\Omega} : \begin{bmatrix} \overline{\mathbf{\omega}}^{(t)} \\ \overline{\boldsymbol{\zeta}}^{(t)} \end{bmatrix} | \mathbf{\omega}^{(t)}, \mathbf{\zeta}^{(t)}, \mathbf{\Omega} \sim N \left[\mathbf{J} \left[\mathbf{S}'(\mathbf{I}_{B} \otimes \mathbf{\Omega}^{-1}) \right] \left[\begin{array}{c} \mathbf{\omega}_{111} \\ \mathbf{\zeta}_{111} \\ \mathbf{\omega}_{112} \\ \mathbf{\zeta}_{121} \\ \vdots \\ \mathbf{\omega}_{121} \\ \mathbf{\zeta}_{121} \\ \vdots \\ \mathbf{\omega}_{1CB_{c}} \\ \mathbf{\zeta}_{1CB_{c}} \end{bmatrix} + \overline{\mathbf{\Omega}}^{-1} \left[\begin{array}{c} \overline{\mathbf{\omega}} \\ \overline{\mathbf{z}} \\ \overline{\mathbf{z}} \\ \overline{\mathbf{z}} \\ \mathbf{\zeta}_{1CB_{c}} \\ \mathbf{\zeta}_{1CB_{c}} \end{bmatrix} + \overline{\mathbf{z}}^{-1} \left[\begin{array}{c} \overline{\mathbf{\omega}} \\ \overline{\mathbf{z}} \\ \overline{\mathbf{z}} \\ \mathbf{\zeta}_{1CB_{c}} \\ \mathbf{\zeta}_{1CB_{c}} \\ \mathbf{\zeta}_{1CB_{c}} \end{bmatrix} + \overline{\mathbf{z}}^{-1} \left[\begin{array}{c} \overline{\mathbf{\omega}} \\ \overline{\mathbf{z}} \\ \overline{\mathbf{z}} \\ \mathbf{\zeta}_{1B_{1}} \\ \mathbf{\zeta}_{1B$$

with B the sum of all B_c and hence the total number of brands.

5. The vector of parameters, $[\omega_{2c1}, \zeta_{2c1}, \omega_{2c2}, \zeta_{2c2}, ..., \omega_{2cB_c}, \zeta_{2cB_{c}}]'$, is drawn from a Bayesian regression model:

$$\begin{bmatrix} \boldsymbol{\omega}_{2c1} \\ \boldsymbol{\zeta}_{2c1} \\ \boldsymbol{\omega}_{2c2} \\ \boldsymbol{\zeta}_{2c2} \\ \boldsymbol$$

where $\mathbf{J}_{2c} = \left(\mathbf{P}_{2c}'(\mathbf{I}_{T_c} \otimes \boldsymbol{\Sigma}_c^{-1})\mathbf{P}_{2c} + \overline{\boldsymbol{\Omega}}_2^{-1}\right)^{-1}$,

,



6. Σ_c is drawn from an inverted Wishart distribution with $T_c + \nu_{\Sigma}$ degrees of freedom:

$$\boldsymbol{\Sigma}_{c}^{-l(t)} | \boldsymbol{\omega}^{(t)}, \boldsymbol{\zeta}^{(t)}, \boldsymbol{y}^{*(t)}, \boldsymbol{z}^{*(t)}, \boldsymbol{\delta}^{(t)}, \boldsymbol{\psi}^{(t)}, \boldsymbol{V}^{(t)}, \boldsymbol{V}_{\Sigma_{c}}, \boldsymbol{v}_{\Sigma_{c}} \sim Wish \left(T_{c} + \boldsymbol{v}_{\Sigma_{c}}, \left(\boldsymbol{V}_{\Sigma_{c}} + \begin{bmatrix} \boldsymbol{\varepsilon}_{c} \\ \boldsymbol{\mu}_{c} \end{bmatrix} \begin{bmatrix} \boldsymbol{\varepsilon}_{c} & \boldsymbol{\mu}_{c} \end{bmatrix} \right)^{-1} \right)$$

7. Ω is drawn from an inverted Wishart distribution with $B + v_{\Omega}$ degrees of freedom.

$$\mathbf{\Omega}^{-1(t)} \Big| \mathbf{\omega}^{(t)}, \overline{\mathbf{\omega}}^{(t)}, \zeta^{(t)}, \overline{\zeta}^{(t)}, \mathbf{V}_{\Omega}, \mathbf{v}_{\Omega} \sim Wish \left(B + \mathbf{v}_{\Omega}, \left(\mathbf{V}_{\Omega} + \begin{bmatrix} \mathbf{e} \\ \mathbf{u} \end{bmatrix} \begin{bmatrix} \mathbf{e}' & \mathbf{u'} \end{bmatrix} \right)^{-1} \right).$$

For identification purposes we need ones on the diagonal of incidence equation error matrix $\Sigma_{c22} = E(\mu_{ct}\mu'_{ct})\forall c$ (while Σ_c is positive definite), and we need to rescale the parameters from the incidence equation relative to Σ_{c22} . To achieve this, we follow the procedure proposed by Edwards and Allenby (2003) and Rossi, Allenby, and McCulloch (2005, p. 108). That is, we do not impose any restrictions when drawing Σ_c . Instead, we postprocess the draws using the following diagonal matrices:

$$\begin{split} \mathbf{C}_{c1} &= \begin{bmatrix} \mathbf{I}_{B_{c}} & \mathbf{0} \\ \mathbf{0} & \operatorname{diag}(\mathbf{\Sigma}_{c22})^{-\frac{1}{2}} \end{bmatrix}, \forall c; \\ \mathbf{C}_{1c2} &= \begin{bmatrix} \mathbf{I}_{M} & \mathbf{0} \\ \mathbf{0} & (1/\sqrt{\mathbf{\Sigma}_{c22}^{(1,1)}}) \cdot \mathbf{I}_{M} \\ & \mathbf{I}_{M} & \mathbf{0} \\ \mathbf{0} & (1/\sqrt{\mathbf{\Sigma}_{c22}^{(2,2)}}) \cdot \mathbf{I}_{M} \\ & & \mathbf{0} & (1/\sqrt{\mathbf{\Sigma}_{c22}^{(2,2)}}) \cdot \mathbf{I}_{M} \end{bmatrix}, \forall c, \\ \mathbf{C}_{1c2} &= \begin{bmatrix} \mathbf{I}_{L} & \mathbf{0} \\ \mathbf{0} & (1/\sqrt{\mathbf{\Sigma}_{c22}^{(1,1)}}) \cdot \mathbf{I}_{L} \\ & & \mathbf{0} & (1/\sqrt{\mathbf{\Sigma}_{c22}^{(2,2)}}) \cdot \mathbf{I}_{L} \\ n & & \mathbf{0} & (1/\sqrt{\mathbf{\Sigma}_{c22}^{(2,2)}}) \cdot \mathbf{I}_{L} \\ n & & \mathbf{0} & (1/\sqrt{\mathbf{\Sigma}_{c22}^{(2,2)}}) \cdot \mathbf{I}_{L} \\ & & & \mathbf{0} & (1/\sqrt{\mathbf{\Sigma}_{c22}^{(2,2)}}) \cdot \mathbf{I}_{L} \\ \mathbf{C}_{2c2} &= \begin{bmatrix} \mathbf{I}_{N} & \mathbf{0} \\ \mathbf{0} & \operatorname{mean}(\operatorname{diag}(\mathbf{\Sigma}_{c22})^{-\frac{1}{2}}) \cdot \mathbf{I}_{N} \end{bmatrix} \text{ and } \mathbf{C}_{4} = \begin{bmatrix} \mathbf{I}_{M} & \mathbf{0} \\ \mathbf{0} & \operatorname{mean}(\operatorname{diag}(\mathbf{\Sigma}_{c22})^{-\frac{1}{2}}) \cdot \mathbf{I}_{M} \end{bmatrix} \end{split}$$

After completing the Gibbs chain, we calculate $\Sigma_c^* = \mathbf{C}_{c1}\Sigma_c \mathbf{C}'_{c1}$ for each saved Gibbs draw, monitor its convergence, and use it for inference purposes. Analogously, for inferences we use $(\mathbf{\omega}'_{1c}, \zeta'_{1c})^{*'} = \mathbf{C}_{1c2} \cdot (\mathbf{\omega}'_{1c}, \zeta'_{1c})', \forall c; (\mathbf{\omega}'_{2c}, \zeta'_{2c})^{*'} = \mathbf{C}_{2c2} \cdot (\mathbf{\omega}'_{2c}, \zeta'_{2c})', \forall c; \qquad (\overline{\mathbf{\omega}}', \overline{\zeta}')^{*'} = \mathbf{C}_3.(\overline{\mathbf{\omega}}', \overline{\zeta}')',$ and $\mathbf{\Omega}^* = \mathbf{C}_4 \mathbf{\Omega} \mathbf{C}'_4.$

Prior distributions

The second implementation step is to specify prior distributions for the parameters of interest. Note that the priors are set to be non-informative so that inferences are driven by the data.

The prior distribution of $[\overline{\omega}', \overline{\zeta}']$ is N($[\overline{\omega}', \overline{\zeta}']$, $\overline{\Omega}$), where $[\overline{\omega}', \overline{\zeta}']$ =0 and $\overline{\Omega}$ = diag(10³). The prior distribution of Σ_c^{-1} is Wishart: W($v_{\Sigma c}, V_{\Sigma c}$), where $v_{\Sigma} = 2 B_c + 2$ and $V_{\Sigma c} = \text{diag}(10^{-3})$. The prior distribution of Ω^{-1} is Wishart: W(v_{ω}, V_{Ω}), where $v_{\omega} = 2M+2$ and $V_{\Omega} = \text{diag}(10^{-3})$.

Initial values

The third implementation step is to set initial values for the parameters of the marginal distributions. The starting values for $\boldsymbol{\omega}$ and $\boldsymbol{\delta}$ are computed by OLS, using $\ln(y_{cbt})$ as the dependent variable of the regression. The covariance matrix, $\boldsymbol{\Sigma}_{11}$, is initiated by computing the sample covariances of this regression's residuals. In a similar fashion, the starting values for the patronage equation parameters, $\boldsymbol{\zeta}$, are computed by OLS, using z_{cbt} as the dependent variable, and the residuals from this regression, μ_{cbt} , are used to compute the sample correlations, which serve as the initial value for $\boldsymbol{\Sigma}_{22}$.

The final step is to generate N_1+N_2 random draws from the conditional distributions. We use a "burn in" of $N_1 = 30.000$ iterations. To reduce autocorrelation in the MCMC draws, we "thin the line," using every 50th draw in the final $N_2 = 30.000$ draws for our estimation. In this way, 600 draws are used to estimate marginal posterior distributions of the parameters of interest. Test runs of our Gauss implementation of the MCMC draws show that we can retrieve parameters used to simulate artificial data.

Appendix D: MCMC estimation of brand-specific carry-over parameter lambda

Basic sales model for each brand:

(D1) $Sales_{cb,t} = \alpha_{cb} + \beta_{cb} Adstock_{cb,t} + \varepsilon_{cb,t}$

We transform this model:

(D2)
$$Sales_{cb,t} - \lambda_{cb}Sales_{cb,t-1} = (1 - \lambda_{cb})\alpha_{cb} + \beta_{cb}Adstock_{cb,t} - \lambda_{cb}\beta_{cb}Adstock_{cb,t-1} + \varepsilon_{cb,t}^*$$

Adstock definition:

(D3)
$$Adstock_{cb,t} = (1 - \lambda_{cb})Adv_{cb,t} + \lambda_{cb}(Adstock_{cb,t-1})$$

Inserting this in the transformed model:

(D4)
$$Sales_{cb,t} - \lambda_{cb}Sales_{cb,t-1} = (1 - \lambda_{cb})\alpha_{cb} + \beta_{cb}(1 - \lambda_{cb})Adv_{cb,t} + \varepsilon_{cb,t}^{*}$$
or

(D5)
$$Sales_{cb,t} - \lambda_{cb}Sales_{cb,t-1} = \alpha_{cb}^* + \beta_{cb}^*Adv_{cb,t} + \varepsilon_{cb,t}^*$$

or

(D6)
$$Sales_{cb,t} = \alpha_{cb}^* + \beta_{cb}^* Adv_{cb,t} + \lambda_{cb} Sales_{cb,t-1} + \varepsilon_{cb,t}^*$$

which is the well-known Partial Adjustment Model (Hanssens et al., 2001 p147). From this model, we can estimate the brand-specific lambda's.

Estimation.

$$Sales_{cb,t} = \alpha_{cb}^{*} + \beta_{cb}^{*} A dv_{cb,t} + \lambda_{cb} Sales_{cb,t-1} + \varepsilon_{cb,t}^{*}, \text{ with } \varepsilon_{ct} \sim MVN(\mathbf{0}, \Sigma_{c}^{\lambda})$$

and $\alpha_{cb}^{*} = \overline{\alpha} + e_{cb}^{\alpha}$
 $\beta_{cb}^{*} = \overline{\beta} + e_{cb}^{\beta}$
 $\lambda_{cb} = \overline{\lambda} + e_{cb}^{\lambda} \text{ with } (e_{cb}^{\alpha}, e_{cb}^{\beta}, e_{cb}^{\lambda})' \sim MVN(\mathbf{0}, \Omega_{\lambda})$

1. We draw the vector of parameters $\left[\alpha_{c1}^*, \beta_{c1}^*, \lambda_{c1}, \alpha_{c2}^*, \beta_{c2}^*, \lambda_{c2}, \dots, \alpha_{cB_c}^*, \beta_{cB_c}^*, \lambda_{cB_c}\right]$ from a SUR model with variance/covariance matrix of disturbances Σ_c^{λ} :

$$\begin{bmatrix} \alpha_{c1}^{*} \\ \beta_{c1}^{*} \\ \lambda_{c1}^{*} \\ \alpha_{c2}^{*} \\ \beta_{c2}^{*} \\ \vdots \\ \alpha_{cB_c}^{*} \end{bmatrix} \mathbf{Sales}^{(t)}, \mathbf{\Sigma}_{c}^{\lambda(t-1)} \qquad \mathbf{N} \mathbf{K}_{c} \left(\mathbf{L}_{c}^{*} (\mathbf{I}_{T_{c}} \otimes \mathbf{\Sigma}_{c}^{\lambda-1}) \begin{bmatrix} \mathbf{Sales}_{c1,1} \\ \mathbf{Sales}_{c2,1} \\ \mathbf{Sales}_{c1,2} \\ \mathbf{Sales}_{c2,2} \\ \vdots \\ \mathbf{Sales}_{c2,2} \\ \vdots \\ \mathbf{Sales}_{cB_{c},2} \\ \vdots \\ \mathbf{Sales}_{cB_{c},2} \\ \vdots \\ \mathbf{Sales}_{c2,T_{c}} \\ \vdots \\ \mathbf{Sales}_{cB_{c},T_{c}} \end{bmatrix} + \left[\mathbf{I}_{B_{c}} \otimes \mathbf{\Omega}_{\lambda}^{-1} \left[\mathbf{I}_{B_{c}} \otimes \mathbf{I}_{3} \right] \left[\frac{\overline{\alpha}}{\overline{\beta}} \\ \overline{\lambda} \right] \right], \mathbf{K}_{c} \right], \mathbf{K}_{c} \left[\frac{\overline{\alpha}}{\overline{\beta}} \\ \mathbf{Sales}_{c1,T_{c}} \\ \mathbf{Sales}_{c2,T_{c}} \\ \vdots \\ \mathbf{Sales}_{cB_{c},T_{c}} \end{bmatrix} + \mathbf{Sales}_{cB_{c},T_{c}} \right]$$

where
$$\mathbf{K}_{c} = \left(\mathbf{L}_{c}' \left(\mathbf{I}_{T_{c}} \otimes \boldsymbol{\Sigma}_{c}^{\lambda - 1} \right) \mathbf{L}_{c} + \mathbf{I}_{B_{c}} \otimes \overline{\boldsymbol{\Omega}}_{\lambda}^{-1} \right)^{-1},$$



2. The vector of hyper-parameters, $[\overline{\alpha}, \overline{\beta}, \overline{\lambda}]'$ is drawn from a SUR model with variance/covariance matrix of disturbances Ω_{λ} :

$$\begin{bmatrix} \overline{\boldsymbol{\alpha}}^{(t)} \\ \overline{\boldsymbol{\beta}}^{(t)} \\ \overline{\boldsymbol{\lambda}}^{(t)} \end{bmatrix} \boldsymbol{\alpha}^{*(t)}, \boldsymbol{\beta}^{*(t)}, \boldsymbol{\lambda}^{*(t)}, \boldsymbol{\Omega}_{\boldsymbol{\lambda}} \sim N \left(\mathbf{F} \left(\mathbf{G}'(\mathbf{I}_{B} \otimes \boldsymbol{\Omega}_{\boldsymbol{\lambda}}^{-1}) \begin{vmatrix} \boldsymbol{\alpha}_{11} \\ \boldsymbol{\beta}_{11} \\ \boldsymbol{\lambda}_{11} \\ \vdots \\ \boldsymbol{\alpha}_{1B_{1}} \\ \boldsymbol{\beta}_{1B_{1}} \\ \boldsymbol{\lambda}_{1B_{1}} \\ \boldsymbol{\alpha}_{21} \\ \boldsymbol{\beta}_{21}^{*} \\ \boldsymbol{\lambda}_{21} \\ \vdots \\ \boldsymbol{\alpha}_{CB_{C}} \\ \boldsymbol{\beta}_{CB_{C}}^{*} \\ \boldsymbol{\beta}_{CB_{C}} \\ \boldsymbol{\lambda}_{CB_{C}} \end{bmatrix} + \overline{\mathbf{\Omega}}_{\boldsymbol{\lambda}}^{-1} \begin{bmatrix} \overline{\boldsymbol{\alpha}} \\ \overline{\boldsymbol{\beta}} \\ \overline{\boldsymbol{\beta}} \\ \overline{\boldsymbol{\lambda}} \end{bmatrix} \right), \mathbf{F} \right],$$

where $\mathbf{F} = (\mathbf{G}'(\mathbf{I}_B \otimes \mathbf{\Omega}_{\lambda}^{-1})\mathbf{G} + \overline{\mathbf{\Omega}}_{\lambda}^{-1})^{-1}$, and $\mathbf{G} = \mathbf{1}_B \otimes \mathbf{I}_3$, with B the sum of all B_c and hence the total number of brands.

3. We subsequently draw Σ_c^{λ} from an inverted Wishart distribution with $T_c + v_{\Sigma_c^{\lambda}}$ degrees of freedom:

$$\boldsymbol{\Sigma}_{c}^{\lambda-1(t)} | \boldsymbol{\alpha}^{(t)}, \boldsymbol{\beta}^{(t)}, \boldsymbol{\lambda}^{(t)}, \mathbf{Sales}, \mathbf{V}_{\Sigma_{c}^{\lambda}}, \mathbf{v}_{\Sigma_{c}^{\lambda}} \sim Wish\left(T_{c} + \boldsymbol{v}_{\Sigma_{c}^{\lambda}}, \left(\mathbf{V}_{\Sigma_{c}^{\lambda}} + \boldsymbol{\varepsilon}_{c} \cdot \boldsymbol{\varepsilon}_{c}\right)^{-1}\right)$$

4. Ω_{λ} is drawn from an inverted Wishart distribution with $B + v_{\Omega_{\lambda}}$ degrees of freedom:

$$\mathbf{\Omega}_{\lambda}^{-1(t)} | \boldsymbol{\alpha}^{(t)}, \boldsymbol{\beta}^{(t)}, \boldsymbol{\lambda}^{(t)}, \overline{\boldsymbol{\alpha}}^{(t)}, \overline{\boldsymbol{\beta}}^{(t)}, \overline{\boldsymbol{\lambda}}^{(t)}, \mathbf{V}_{\Omega_{\lambda}}, \mathbf{v}_{\Omega_{\lambda}} \sim Wish\left(B + \nu_{\Omega_{\lambda}}, \left(\mathbf{V}_{\Omega_{\lambda}} + \mathbf{e'e}\right)^{-1}\right)\right)$$

5. We specify prior distributions for the parameters of interest. These are set to be non-informative so that inferences are driven by the data. The prior distribution of Σ_c^{λ-1} is Wishart: W (v_{Σ_c^λ}, V_{Σ_c^λ}), where v_{Σ_c^λ} = 2B_c+2 and V_{Σ_c^λ} = diag(10⁻³). The prior distribution of Ω_λ⁻¹ is Wishart: W (v_{Ω_λ}, V_{Ω_λ}), where $v_{\Omega_{\lambda}} = 3+2$ and $\mathbf{V}_{\Omega_{\lambda}} = \text{diag}(10^{-3})$.

Appendix E: Operationalization of the variables

/ariable	Measurement
Adstock management	
Adstock	$Adstock_{cb,t} = (1 - \lambda_{cb})Adv_{cb,t} + \lambda_{cb}(Adstock_{cb,t-1})$
Ad-sensor	$Ad - sensor_{b,t} = Adstock_{b,t-1} - Adstock_{b,t-2}$
Moderators	
Brand market share	Average volume share over the 156 weeks estimation period (cfr. Gatignon et al., 1990)
Advertising frequency	Percentage of time the brand was advertised during the 52 weeks initialization period
Company factors	
New product introduction	Dummy variable; 1 = if within four weeks after product introduction, 0 = otherwise
End of year budget depletion	Dummy variable; 1 = if within last four weeks of the year, 0 = otherwise
Competitive factors	
Competitive Adstock	$CompAdstock_{b,t} = \sum_{i=1}^{k} \left(ms_{i,t} * Adstock_{i,t} \right)$
Relative performance evolution	First difference of the log-transformed volume share (cfr. Franses and Koop, 1998)
Category factors	
Category growth	First difference of the log-transformed category volume sales (cfr. Franses and Koop, 1998)
Category concentration	Herfindahl index of volume shares of the brand over a moving window of previous 26 weeks