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ABSTRACT

The study compares estimates of life expectancy and active life expectancy across indicators of socioeconomic status for a cohort of older adults in Beijing Municipality. Our aim is to determine whether associations found are consistent across indicators and with those typically observed in Western industrialized countries. A multistate life table method is used to estimate expected years of total and active life, defined as life spent without limitation in functions necessary for performing daily tasks. We find that men of higher status experience advantages with respect to life and active life expectancies. Among women, only active life expectancy is significantly greater for those of higher status, but the difference by income is not statistically significant. With respect to the proportion of life spent in an active state, both men and women of higher status benefit in comparison to their lower status counterparts. Finally, we find that disparities by socioeconomic status generally increase with age. Despite several inconsistencies across socioeconomic status indicators by sex, findings generally confirm inequalities within a society that is organized very differently socially, economically, and politically from countries in the West.

Consistent and robust statistical associations between socioeconomic status and levels of mortality and morbidity have been well-documented for older adults, and reducing such inequality is an important challenge facing populations around the world (National Research Council 2001; Preston and Taubman 1994). Much of the research on these associations has taken place in Western developed countries. The current study examines socioeconomic differentials in life and active life expectancies among older adults in Beijing Municipality, an area encompassing the city of Beijing and its rural environs.

Active life expectancy is a useful measure for summarizing morbidity and mortality experiences and for comparing these across subgroups within populations (Robine and Ritchie 1991). It is obtained through decomposing life expectancy into estimates of time lived in active and inactive states of health—in other words, active life expectancy and inactive life expectancy (*ibid.*). Estimates can be viewed in absolute terms, that is, the number of years expected to be spent in an active state, or in relative terms, that is, active years as a proportion of total life expectancy (Nusselder 2003). For the current study, we compare these estimates across four indicators of socioeconomic status, namely, three conventional measures of education, income, and occupation and a household-level measure that refers to ownership of household possessions.

China provides a particularly interesting setting to assess the generality of an association between socioeconomic status and health. The country is characterized by a system of stratification that, in theory, promotes more egalitarian access to various resources, and by a history of economic development that differs greatly from that in Western industrialized societies. Despite historically lower levels of economic development and health care expenditures, China has attained a level of life expectancy that is almost comparable to the West, an achievement some have attributed to equalizing access to primary health care (Caldwell 1986). Furthermore, China's population is aging at an unprecedented rate because of drastic declines in fertility triggered in part by the strict family planning policies in place over the last several decades and in part by declining mortality at old age. This makes understanding quantity versus quality of life, that is, the total number of years versus active and healthy years of life among the elderly and across subgroups within the population an important topic for research (Hermalin 1995; Phillips 2000).

SOCIOECONOMIC STATUS, MORTALITY, AND MORBIDITY

Several classic studies on socioeconomic status differentials in health, such as Kitagawa and Hauser (1973) for the United States and Townsend and Davidson (1982) and Marmot et al. (1987) for the United Kingdom, established that substantial health benefits are associated with higher socioeconomic status. For older populations, it is likely that some selection effects influence the association. That is, those who are apt to be fit and healthy are more likely to survive to old age, potentially muting associations between socioeconomic status and health. Despite this possibility, the association has been confirmed for older populations in a series of more recent studies within several industrialized countries, particularly when health is measured with functional status items

(Berkman and Gurland 1998; Grundy and Sloggett 2002; Kaplan et al. 1993; Liang et al. 2002; Melzer et al. 2001; Olausson 1991; Ross and Wu 1996; Thorslund and Lundberg 1994; Valkonen 1989; Wolfson et al. 1993).

Previous research has also generally found links between socioeconomic status and life expectancy, active life expectancy, and inactive life expectancy in the West (Bossuyt et al. 2000; Cambois et al. 2001; Crimmins and Saito 2001; Crimmins et al. 1989; Doblhammer and Kytir 1999; Manton et al. 1997; Tu and Kuanjeng 1994; Valkonen et al. 1997). Higher status is associated not only with more years of expected life, but also with enjoying more years in an active state and fewer years in an inactive state. These studies have often shown greater differences by socioeconomic status in active life expectancy than in life expectancy, suggesting that inequalities are most pervasive with respect to quality of remaining life. These studies have also shown that associations persist and even strengthen into very old ages, although part of this may be due to selection. In contrast, no consistent differences in the strength of the associations have emerged by sex (Tu and Kuanjeng 1994). Findings have mostly been based on the Sullivan (1971) method of calculating active life expectancy, which uses cross-sectional data on prevalence rates of health problems.

The magnitude of socioeconomic status differences in life and active life expectancies appears to be influenced by the way in which status and health outcomes are measured (Crimmins and Cambois 2003; Guralnik and Kaplan 1989). This may be due to the fact that specific health outcomes are more readily influenced by factors associated with socioeconomic characteristics or that various socioeconomic status indicators are related to health outcomes through different mediating factors. Income may, for example, be more closely related to access to high-quality medical care or better nutrition, while occupation may be more closely related to health hazards in the work place. Other mediating factors that have been identified include psychosocial factors such as health-related behaviors, stress, social support, and self-efficacy (House et al. 1994; Williams 1990).

Although virtually all studies on the relationship between socioeconomic status and health discussed above have been conducted in Western industrialized societies, some research from Asia is beginning to emerge (Liang et al. 2001; Liang et al. 2000; Tu and Kuanjeng 1994; Zimmer and Amornsirisomboon 2000; Zimmer and Kwong, forthcoming; Zimmer et al. 1998). While these studies have substantiated some of the findings from Western societies, inconsistencies have also materialized. In particular, some indicators of socioeconomic status found to be significant predictors of health in the West are not consistently related to various health outcomes in Asian settings.

MEASURING SOCIOECONOMIC STATUS

There are several issues to consider when measuring socioeconomic status among older populations generally and older populations in developing countries specifically. First, conventional indicators of status may not adequately describe the socioeconomic position of older adults. For instance, income often declines with retirement, and current income for older adults may mask the cumulative effect of economic status throughout

life (Kaplan et al. 1987). Assigning older adults to occupational categories can be problematic for those who are no longer employed or were never employed, or for those who held multiple jobs over their lifetimes. Use of these categories may be especially inappropriate among older women, who likely have been less attached to the labor force throughout their lifetimes than men.

The problems associated with measures of income and occupation may be amplified in developing countries. For instance, because income for older adults in developing countries typically comes from multiple sources, including monetary assistance from family members and various sideline economic activities, it is difficult to obtain accurate accounts (Hermalin et al. 2002). The common Western concept of occupational standing may be inappropriate in a setting where a large majority work in agricultural occupations, while only a small proportion are found across wide varieties of professional and other nonagricultural occupations.

A second problem is the difficulty in determining the correct causal order between socioeconomic status and health (Smith 1999). Poor health may be a result of poverty and occupational hazards, but it can also force individuals into early retirement and limit their income and occupational options. For this reason, education is often considered to be a more efficacious measure of socioeconomic status than is health status. Information on education is usually available for every adult in any society, and it is fixed relatively early in life and remains constant throughout adulthood (Freedman and Martin 1998; Winkleby et al. 1992). However, in settings with generally low educational attainment, narrow variation in levels of education can limit its power in predicting health outcomes. Moreover, in societies in earlier stages of socioeconomic development or in societies characterized by socialist economic systems, the relationship between education, income, and other indicators of status may be less clear than elsewhere.

Studies that have expanded upon the conventional indicators of socioeconomic status by including alternative indicators, such as financial assets, home or car ownership, and household goods, have shown these indicators to be highly effective in differentiating health status among older adults in the West (Arber and Ginn 1993; Marmot et al. 1987; Robert and House 1996). Studies incorporating these nontraditional measures in developing countries are, however, quite limited. Zimmer and Amornsirisomboon (2000) examined a measure of household amenities in Thailand and found it to be fairly predictive of health outcomes, although they found income to be a stronger predictor.

THE BEIJING SETTING

Before the Communist Party came to power in 1949, China was predominantly agricultural and, even in urban areas, wage labor was not widespread (Harrell 2000). In 1949, the Party took two important steps toward its goal of transforming China into a socialist society. First, it collectivized work and eliminated privately owned businesses, thereby moving production out of the household and into the public sphere (Goldstein and Ku 1993). Individuals either worked on allocated land in the rural communes or were employed in assigned work-units in urban areas and were paid according to a national wage scale (Davis-Friedmann 1991; Harrell 2000; Parish and Whyte 1978).

Second, the state demanded that both women and men take part in economic activities outside the household. This step was an attempt to remove preexisting sex differences in wage work and eliminate the norm that women worked inside the home while men worked outside (Harrell 2000).

Among other changes that accompanied the new economic system was the provision of health care. Workers and their dependents in urban areas were generally covered under government and labor insurance programs financed by state enterprises; those in rural areas were covered under a cooperative medical insurance program operated by local collectives (Olson 1993; Shi 1993). Through these programs, almost everyone in China had access to basic health care.

Following the death of the Communist Party Chairman Mao Zedong in 1976, China entered an era of economic reform with policies aimed at moving away from a planned economy (Davis-Friedmann 1991; Harrell 2000; Ikels 1990). Communes were decollectivized and privately owned businesses emerged. China's economy today is often described as a socialist market economy, still made up primarily of the state and collective sectors, yet with an increasing proportion of private-sector activities. Under the current system, China's economy has seen rapid growth. Yet unemployment, which was rare before the reforms, increased and so did the number of bankruptcies among enterprises. Wages, which were kept stable in the past regardless of supply and demand, have also become more unstable. Concurrent changes in health care have come through decentralization of health facilities and health care administration. There have also been declines in health insurance coverage and in the number of primary health care providers, especially in rural areas, and costs for patients have risen sharply (Liu et al. 1995; Shi 1996).

It is unclear how the economic changes of the last five decades have influenced individual-level socioeconomic characteristics and their associations with morbidity and mortality. One might hypothesize that such associations would be weaker among today's elderly, who have spent a large proportion of their adult years living within a system that discouraged differences in socioeconomic status and promoted equal access to health care. Given high rates of labor force participation for both sexes, any association between socioeconomic status and morbidity, and mortality may vary little by sex. However, growing inequalities during the reform era may be strengthening the widely observed associations between socioeconomic status, morbidity, and mortality. Finally, although there is little research on how changes related to health care since the early 1980s have affected the elderly's use of health services or their health status, older adults are likely to be particularly vulnerable to these changes.

METHODS

Data

Data for our analysis come from the Beijing Multidimensional Longitudinal Study on Aging, a project led by Zhe Tang and his colleagues at the Beijing Municipal Network for Health and Care of the Elderly at the Xuanwu Hospital of the Capital University of Medical Science in Beijing. We use data from the 1992, 1994, and 1997 waves of this

ongoing study. Three areas within Beijing Municipality are represented in the data collection. The first is Xuanwu district within the city of Beijing. The other two are Daxing, a suburban area, and Huairou, a mountainous area located in the rural suburbs of Beijing. The data are weighted by age and sex based on census figures, so that the weighted sample is representative of the population of these three districts. The sample size is 3,257, of which women make up 51.1 percent. The average age is 64.8 years for men and 65.2 years for women. Response rates for the three waves of data collection were 90.1 percent, 94.1 percent, and 88.4 percent. Further information on the data can be found in Jiang et al. (2002) and Tang et al. (1999).

Functional limitations

In each survey wave, respondents were asked to report whether they can perform a series of functional tasks without any help. These included eating, dressing, getting on and off a bed, bathing, walking 300 meters, and walking up and down a flight of stairs. These are all tasks that are necessary for daily functioning and are sometimes referred to as activities of daily living or Nagi functional tasks (Katz et al. 1983; Nagi 1965). We define the “active” state as the ability to perform all six functional tasks without any help. The “inactive” state is defined as requiring some help in performing at least one task. Individuals are coded as being either functionally active or functionally inactive in each of the three waves. In the second and the third waves, individuals may also be coded as having died in the survey interval years, or as having missing data (not responding to the survey or to particular items). Active life expectancy indicates the length of expected life in which one is able to perform all six of these tasks. Inactive life expectancy indicates the length of expected life in which one is unable to perform at least one task.

Socioeconomic status indicators

We use four indicators of socioeconomic status: education, income, occupation, and household possessions. All are measured at baseline. Because of limited sample size, we dichotomized each indicator in a way that distinguishes those with relatively high status from those with relatively low status. Education is dichotomized as respondents with and without any formal schooling. For income, we first constructed a summary measure by adding income from different sources reported in the survey—for instance, work, pension, and rental income. We then dichotomized the total, with respondents having income above the median level being in the higher status group. Because the levels of income differ greatly by districts, this median level was determined separately for each of the three districts. For occupation, we recognize that the Western conceptualization of prestige is less appropriate to the Chinese setting. Instead, we focus on the physical demands of one’s occupation. We dichotomized reported occupation as those whose lifetime work was “white-collar” or light physical labor (considered as the higher status group), versus those in occupations requiring heavy physical labor. Because some adults in our sample report being retired, and some would have had multiple occupations, we used information on occupation that was reported to have been held the longest.

In contrast to the first three conventional indicators of socioeconomic status, the final indicator is a household-level measure of ownership of selected household possessions, namely, color television, refrigerator, and washing machine. Individuals from households that own all three are coded as being in the higher status group. While other household-level measures, such as home ownership or liquid assets, are more commonly used to indicate household status in other studies, these are not necessarily applicable to the Chinese context. For example, land in Beijing is state-owned (and leased to individuals); thus, there is little option for private home ownership.

We recognize that the division of each measure into higher versus lower socioeconomic status is arbitrary. Several points about these divisions should be made. First, we have attempted to make divisions that make sense within the Chinese context. Second, we have used cut-points that divide the sample fairly evenly. Table 1 shows the percent of individuals coded as having higher status according to each of the four indicators for the total sample and by sex. The fraction in the higher status category is around 40 to 50 percent across the four indicators. There are some appreciable differences by sex in education and income, where the percentages in the higher status groups are around 70 percent for men and around 30 percent for women, while occupation and household possessions show little difference by sex. Third, our goals are to describe general patterns of differences in active life expectancy across relative socioeconomic status categories rather than focus on values of estimates, and to determine whether associations are consistent with those that would be expected on the basis of previous studies.

Table 2 shows the extent to which the socioeconomic status indicators are correlated by presenting concurrence rates, that is, the proportion of individuals who are coded as having higher status or lower status given a combination of two indicators. For example, in our total sample, 66 percent are coded as either having both higher education and higher income or having both lower education and lower income. The results show that the socioeconomic status indicators are strongly correlated for both sexes, but they are also not completely dependent measures.

Analyses

We estimate life expectancy, active life expectancy, and inactive life expectancy separately by sex, using the IMACh (Interpolative Markov Chains) software program, a recent innovation in analysis of active life expectancy employing the multistate life table method (Brouard et al. 2002). The program, written by Nicholas Brouard and colleagues at the Institut National d'Etudes Démographiques in Paris, is based on methods for analyzing active life expectancy developed by Laditka and Wolf (1998). The program permits the use of longitudinal datasets when the periods between survey waves are not identical. In our case, first and second survey waves are separated by two years, while the second and third waves are separated by three years. The program also enables us to retain cases with missing data in some waves. IMACh also generates variance estimates for life expectancy, active life expectancy, and inactive life expectancy, which allows for

testing statistical significance for differences observed in these estimates across different groups.

The first step in the estimation of active life expectancy is a parameter estimation of transition probabilities between the survey waves using maximum likelihood. Our transition model consists of death as an absorbing state and the states of being functionally active and functionally inactive as two nonabsorbing states, with the possibility of movement between the two nonabsorbing states. There are, thus, three possible transitions from each nonabsorbing state for a total of six transitions between survey waves: from functionally active to functionally inactive, to functionally active, or dead; from functionally inactive to functionally active, to functionally inactive, or dead. These transition probabilities are calculated for all ages (between 55 and 80+ years), both sexes, and both higher and lower socioeconomic status groups across all socioeconomic status indicators. Then, based on these transition probabilities, we construct multistate life tables for each subgroup, from which we obtain estimates of life expectancy, active life expectancy, and inactive life expectancy.

Since we are primarily interested in comparisons of life expectancy, active life expectancy, and inactive life expectancy between individuals in higher and lower socioeconomic states, we present our main results in the form of a ratio of a given estimate for higher status over the corresponding estimate for lower status at selected ages. For instance, our results for men show that at age 55, those with some education are expected to live 21.9 years compared to 18.3 years for those with no formal education. The ratio of these life expectancies is 1.20, suggesting that men with some formal education at age 55 are expected to live 20 percent longer than their uneducated counterparts. A ratio above 1.00 indicates a positive relationship between socioeconomic status and the estimate under examination, while a ratio below 1.00 indicates a negative relationship. We focus our discussion on these ratios and on general patterns found in our results rather than highlighting single estimates.

We use variance estimates generated by IMACh to test for statistical significance of observed socioeconomic status differentials and report them in our tables. To this end, we first calculate 95 percent and 99 percent confidence intervals around each estimate, then report the socioeconomic status differentials as being significant where confidence intervals do not overlap. Finally, we present the proportion of expected life spent in an active state for each higher and lower status group across all socioeconomic status indicators. This allows for the examination of socioeconomic status differentials in active life expectancy in relative sense.

RESULTS

Table 3 shows the percent of respondents who are unable to perform each of the six functional tasks without any help as well as the percent unable to perform at least one of the six at baseline by age and sex. In total, 10.5 percent report the inability to perform at least one of the six tasks without help, although the figure differs greatly by age and sex. The percent reporting limitation with at least one task increases with age, as expected, as does the percent reporting limitation with any specific task. For instance,

about 50 percent of those aged 80 and older compared to only about 3 percent of those aged 55 to 59 report a limitation with at least one task. Women are almost twice as likely as men to report at least one limitation, and this disadvantage exists within virtually all age groups and for all tasks. Looking at individual tasks, more people have difficulty walking up and down stairs than performing other tasks, followed by walking 300 meters, bathing, getting on and off a bed, dressing, and eating. This sequence is quite similar across age groups and for both sexes.

Before introducing socioeconomic status indicators, we present Figure 1 to describe overall patterns of life expectancy, active life expectancy, and inactive life expectancy. The figure graphs years of expected life, active life, and inactive life, as well as the proportion of expected life spent in an active state. For men, both life expectancy and active life expectancy decline steadily with age, as would be expected. In contrast, inactive life expectancy fluctuates little across ages. The proportion of expected life in an active state declines slightly with age, from 0.91 for those aged 55 to 0.75 for those aged 80.

Patterns are similar for women. Both life expectancy and inactive life expectancy are higher for women than for men at all ages. For instance, at age 55 women can expect to live another 23 years, about four of which are inactive. Men at the same age expect to live 20 years with only about two years in inactive health.

We turn now to the main focus of the current study, namely, the ratios that compare these estimates among respondents of higher status to respondents of lower status across different indicators of socioeconomic status. Table 4 shows ratios for life expectancy, active life expectancy, and inactive life expectancy for each status indicator separately by sex for ages 55, 60, 65, 70, 75, and 80.

Life expectancy ratios are all greater than 1.00 across all socioeconomic status indicators for all ages and both sexes, indicating that respondents in the higher status groups live longer than do those in the lower status groups. For instance, the first row of the table indicates that, compared to their counterparts in the corresponding lower status groups men aged 55 in the higher education group can expect to live 20 percent longer, those in the higher income group can expect to live 37 percent longer, those in the white-collar and light physical labor occupation group can expect to live 27 percent longer, and those in the higher household possessions group can expect to live 30 percent longer.

Despite consistency in the direction of the association, the size and the statistical significance of the ratios differ by age and gender. Among men, socioeconomic status differentials in life expectancy ratios are mostly statistically significant and, as noted above, are quite substantial. Ratios increase fairly steadily by age so that by age 80 those of higher status can expect to live 40 percent to 52 percent longer than those of lower status. The life expectancy ratios among women are noticeably smaller across all socioeconomic status indicators and are not generally statistically significant, with the exception of ratios for the household possession measure. Depending upon age and indicator, except for household possessions, women categorized as being of higher socioeconomic status live only 4 percent to 16 percent longer than those of lower status.

For household possessions, the ratio exceeds 1.20 regardless of age. In contrast to men, no clear age patterns are observed in the life expectancy ratios among women.

Differences in active life expectancy by socioeconomic status are more consistent with expectations for both men and women, although, again, men of higher status are more advantaged than are their female counterparts. All ratios are above 1.00 and, except for income among women, differences tend to be significant, meaning that those in the higher status group can expect to live more active years. Active life expectancy ratios also increase with age for both sexes. For men, depending upon age and status indicator, the advantage ranges from 30 percent to 77 percent. Among women, the ratios are smaller. Except for income, women categorized as being of higher status can expect to live between 20 percent and 73 percent more active years.

The last set of estimates shows the ratios for inactive life expectancy. Patterns here are less clear. While ratios are mostly below 1.00, as expected, many are close to 1.00, particularly at older ages. Some of the ratios for men are also above 1.00. The exceptions are ratios among women for education and household possessions, which are substantially below 1.00. None of the ratios, however, is statistically significant, which is likely partly due to the fact that the actual values for inactive life expectancy are quite small and therefore the variances, in relation to the estimates, are large.

Figures 2 and 3 show the proportion of remaining life expected to be lived in an active state for higher and lower status groups. Figure 2 shows the proportions for men and Figure 3 for women. Looking first at Figure 2, the proportion of active life among men is higher for the higher status group across all socioeconomic status indicators. The amount of difference ranges from about 5 to 10 percentage points depending upon the indicator and age. In all cases, the proportion of life spent in an active state decreases with age. At age 55, it is about 90 to 95 percent for those of higher status and 85 to 90 percent for those of lower status. By age 80, it has decreased to about 80 percent or less for those in the higher status categories and to 70 percent or less for those in the lower categories.

Figure 3 for women shows some consistencies with the results for men. For all indicators, women of higher socioeconomic status can expect to spend more time in an active state regardless of age. But differences are less consistent across indicators, with higher status individuals being relatively more advantaged with respect to education and household possessions, especially compared to the corresponding advantage with respect to income. The proportion of time expected to be lived in an active state is somewhat lower for women than for men, particularly among those in lower socioeconomic status categories, where the proportion of time active can fall to 50 percent at very old ages.

CONCLUSIONS

We have extended previous research on the association between socioeconomic status and health by examining life expectancy and active life expectancy among older adults living in Beijing Municipality. We focused on ratios of life expectancy, active life expectancy, and inactive life expectancy, contrasting each estimate across categories of higher versus lower status. These ratios were calculated for four measures of

socioeconomic status, three of which are standard indicators—education, income, and occupation—and one of which is less often used—ownership of household possessions.

Despite large differences in the stage of socioeconomic development and political and economic systems between China and Western industrialized countries, our findings are mostly consistent with findings from the West. We generally find advantages in life expectancy and active life expectancy for older adults categorized as having higher levels of socioeconomic status, especially among men. Consistent with previous findings, we also find starker differences between socioeconomic status groups in active years of life than in total years of life and we find a widening of the gap between the two socioeconomic status groups with increasing age.

Contrary to our expectation, we find none of the socioeconomic status differences in inactive life expectancy to be statistically significant. Although the associations between socioeconomic status and inactive life expectancy are all in the expected direction among women, they are not consistent among men. This lack of significance reflects, in part, the small number of years expected to be spent in an inactive state. However, some of the ratios are close to 1.00, indicating, in part, little difference in the number of inactive years of life expected between those of higher and lower status. Hence, the result may also reflect circumstances found in China that differ from those found elsewhere, where individuals with higher socioeconomic status tend to live fewer years in an inactive state. It is difficult to speculate on why socioeconomic status characteristics would distinguish active but not inactive years of life. One factor to consider is that older adults in China today, even those of higher status, have spent their childhoods within a relatively unhealthy environment, characterized by high rates of communicable diseases, malnourishment, and other difficult circumstances. Although these older adults may now have an opportunity to extend their life, their earlier disadvantages may result not only in extra years in active states but also some extra years in an inactive state. This situation may change as new cohorts of Chinese, who have lived in relatively healthier environments throughout their lives, begin to age.

Another set of results contrary to expectations involves gender differentials. Despite socialist policies promoting greater gender equality in work, we find that socioeconomic status differences in life expectancy and active life expectancy are greater for men than for women. Furthermore, the specific status indicator made a difference in the results among women while it made little difference among men. Women with the three household possessions had a great advantage over those without, whereas there were smaller differences when using other, individual-level indicators. Income, in particular, was not able to distinguish those women living longer as well as healthier lives. One reason for the gender difference in our findings may be that despite the strong attachment Chinese women have had to the labor market under the communist regime, socioeconomic status indicators such as income do not adequately capture the cumulative effects of economic status for women. Our findings hint at the possibility that women's socioeconomic status may be more closely related to the socioeconomic status of their households than to their own socioeconomic characteristics. Another possibility is that our results reflect gender differences in behaviors that are associated both with

socioeconomic status and with health outcomes. As one example, prevalence rates for smoking and immoderate drinking in our sample are much lower among women than among men and are also associated with socioeconomic status.

We advise caution when interpreting our results on socioeconomic status differences for several reasons. First, some of these differences may reflect the ways in which specific socioeconomic status indicators were operationalized. Because of sample size limitations, we dichotomized each indicator into a logical division between higher and lower status. This results in admittedly crude divisions. Still, our findings with respect to life and active life expectancies are strong and the patterns that emerge are clear. Indeed, the division of socioeconomic status into basic categories better lends itself to examination of general patterns rather than to estimates for specific years. (We conducted several sensitivity tests by dividing our measures in different ways and the general conclusions remain unchanged.)

Second, we caution against misreading the causal direction of some of these associations. In particular, it may be that those who have better functional health are able to remain employed longer and earn higher incomes at older ages and also are in a better position to purchase household possessions. Yet, when it comes to education and occupation, socioeconomic status is likely to have been determined earlier in life, and the direction of causation from socioeconomic status to health can be more easily assumed.

Third, associations we have presented here do not adjust for possible covariates that might intervene in the association between socioeconomic status and health. Indeed, it is difficult to include many covariates in analyses of active life expectancy, especially if sample sizes are no larger than they are in the current study. We suggest that future research begin to disentangle the mediating factors that provide advantages in life expectancy and active life expectancy for older adults of higher status.

Despite the limitations of our study and several inconsistencies in the results according to socioeconomic status indicators and by sex, the findings generally confirm status inequalities within a society that is organized very differently socially, economically, and politically from countries in the West. In a time of rapid socioeconomic and demographic change in China, it would be highly valuable to monitor both active life expectancy of older adults and its relationship to socioeconomic status over time. Given the population aging taking place, this monitoring will be useful in assessing the needs of older adults within a context of rapid changes in social stratification, social organization, and patterns of old-age mortality. Further examinations of long-term panel data in China would also be useful for determining how trends in morbidity and mortality are being affected by the various social and economic changes taking place.

Table 1 Percent of Respondents Categorized as Having High Socioeconomic Status by Indicator and Sex

	Education	Income	Occupation	Household possessions
Total	50.0	51.2	42.9	42.8
Men	69.6	71.1	40.2	43.4
Women	30.7	31.5	45.6	42.1

Table 2 Concurrence Rates across Socioeconomic Indicators for the Total Sample, Men, and Women

	Education	Income	Occupation
<u>Total Sample</u>			
Education	---	---	---
Income	.657	---	---
Occupation	.679	.532	---
Household possessions	.727	.547	.782
<u>Men</u>			
Education	---	---	---
Income	.653	---	---
Occupation	.656	.540	---
Household possessions	.669	.529	.799
<u>Women</u>			
Education	---	---	---
Income	.661	---	---
Occupation	.701	.525	---
Household possessions	.784	.566	.765

Table 3 Percent Reporting Specific Functional Limitations at Baseline (1992) by Age and Sex

	Eating	Dressing	On/off Bed	Bathing	Walking	Climbing Stairs	At Least One Limitation	Total N
<u>Both Sexes</u>								
All ages	1.3	1.7	1.8	4.9	7.7	9.5	10.5	3,257
55-59	0.0	0.4	0.5	1.3	2.5	2.3	2.8	1,036
60-69	1.4	1.5	1.4	3.5	5.0	5.7	6.7	1,397
70-79	1.7	3.1	2.7	8.2	13.6	18.5	20.2	644
80+	6.5	6.0	8.6	24.2	37.5	47.1	50.1	180
<u>Men</u>								
All ages	1.1	1.8	1.8	4.1	5.4	6.8	7.3	1,593
55-59	0.0	0.7	1.0	1.7	2.4	2.4	2.4	523
60-69	1.4	1.9	1.7	3.3	3.9	4.4	4.8	687
70-79	1.6	2.8	2.1	6.9	8.8	13.3	14.5	312
80+	5.2	4.3	7.5	17.6	26.8	34.5	37.4	71
<u>Women</u>								
All ages	1.4	1.6	1.8	5.7	10.0	12.1	13.7	1,664
55-59	0.0	0.0	0.0	1.0	2.6	2.3	3.2	513
60-69	1.3	1.0	1.2	3.8	6.2	7.1	8.6	710
70-79	1.9	3.3	3.4	9.5	18.4	23.7	25.8	332
80+	7.3	7.2	9.4	28.6	44.7	55.6	58.7	109

Table 4 Ratios for Life Expectancy, Active Life Expectancy, and Inactive Life Expectancy, Comparing Higher to Lower Status Respondents, by Sex and Selected Ages

		Education		Income		Occupation		Possessions	
Life Expectancy									
<u>Men</u>									
	55	1.20		1.37	**	1.27	**	1.30	**
	60	1.28		1.41	**	1.30	**	1.34	**
	65	1.35	**	1.45	**	1.34	**	1.38	**
	70	1.41	**	1.49	**	1.37	**	1.42	**
	75	1.44	**	1.51	**	1.39	*	1.45	**
	80	1.47	*	1.52	*	1.40	*	1.47	*
<u>Women</u>									
	55	1.11		1.04		1.14	*	1.21	**
	60	1.12		1.04		1.15	*	1.23	**
	65	1.14		1.06		1.16		1.25	*
	70	1.14		1.07		1.16		1.27	*
	75	1.14		1.09		1.15		1.27	*
	80	1.12		1.11		1.11		1.25	
Active Life Expectancy									
<u>Men</u>									
	55	1.30	**	1.44	**	1.32	**	1.34	**
	60	1.37	**	1.51	**	1.37	**	1.40	**
	65	1.44	**	1.57	**	1.42	**	1.45	**
	70	1.50	**	1.64	**	1.48	**	1.52	**
	75	1.57	**	1.71	**	1.55	**	1.58	**
	80	1.63	*	1.77	*	1.61	*	1.64	*
<u>Women</u>									
	55	1.23	*	1.05		1.20	**	1.33	**
	60	1.28	*	1.07		1.24	**	1.39	**
	65	1.33	*	1.09		1.28	**	1.47	**
	70	1.40	*	1.11		1.32	*	1.56	**
	75	1.48	*	1.14		1.36	*	1.65	**
	80	1.56	*	1.19		1.35		1.73	*
Inactive Live Expectancy									
<u>Men</u>									
	55	0.69		0.89		0.84		0.97	
	60	0.84		0.91		0.85		0.99	
	65	0.96		0.93		0.87		1.00	
	70	1.04		0.95		0.88		1.02	
	75	1.08		0.96		0.90		1.03	
	80	1.11		0.97		0.92		1.04	
<u>Women</u>									
	55	0.57		0.95		0.84		0.69	
	60	0.57		0.96		0.84		0.70	
	65	0.58		0.97		0.84		0.71	
	70	0.58		0.98		0.84		0.72	
	75	0.60		1.00		0.85		0.74	
	80	0.62		1.02		0.87		0.76	

*p<.05 **p<.01

Figure 1 Expected Length of Total Life (LE), Active Life (ALE), and Inactive Life (IALE) and Ratio of Active Life Expectancy to Life Expectancy, by Age and Sex

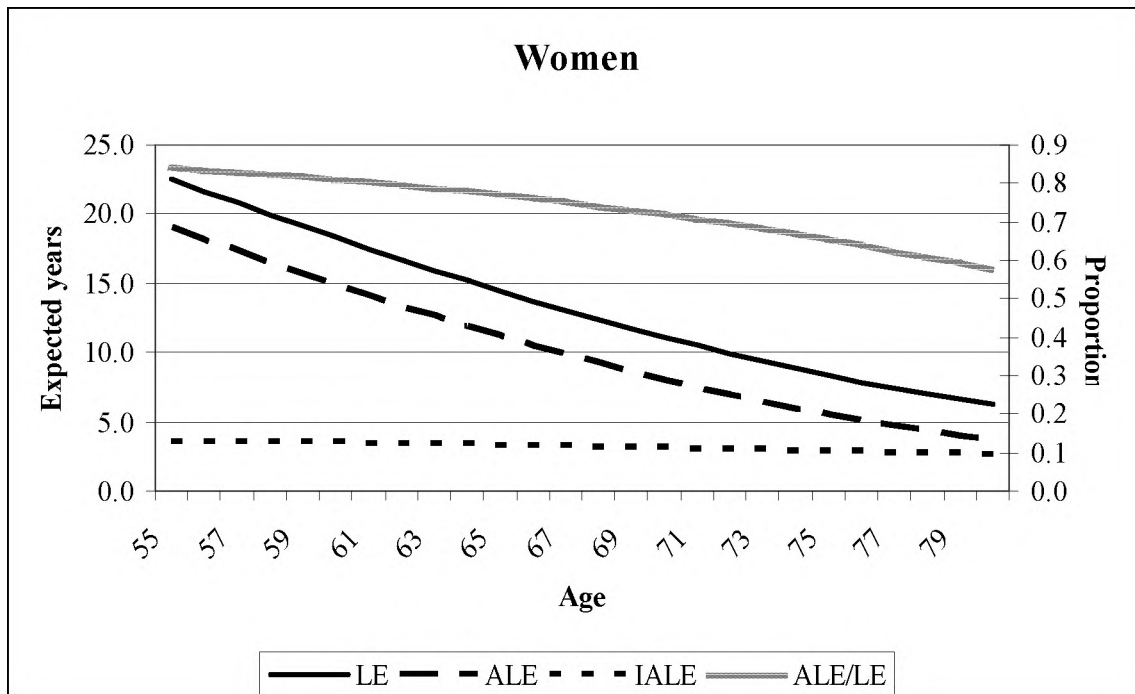
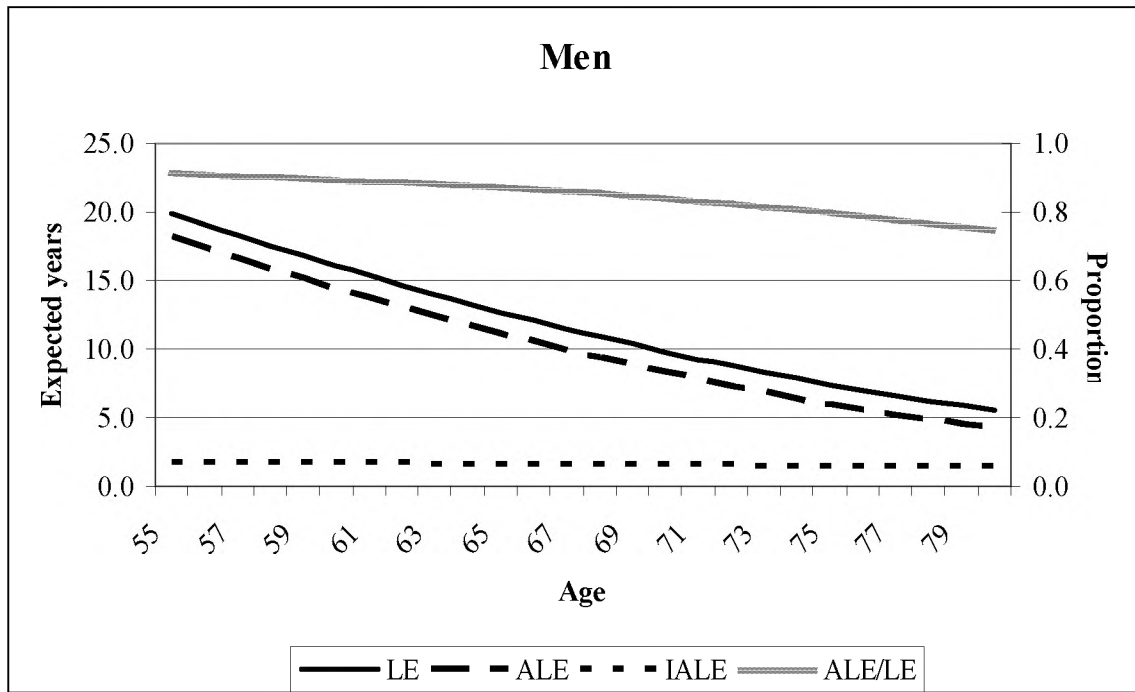


Figure 2 Proportion of Life Expected to be Spent in an Active State, by Men in Higher and Lower Status Groups According to Four Indicators of Socioeconomic Status, by Age

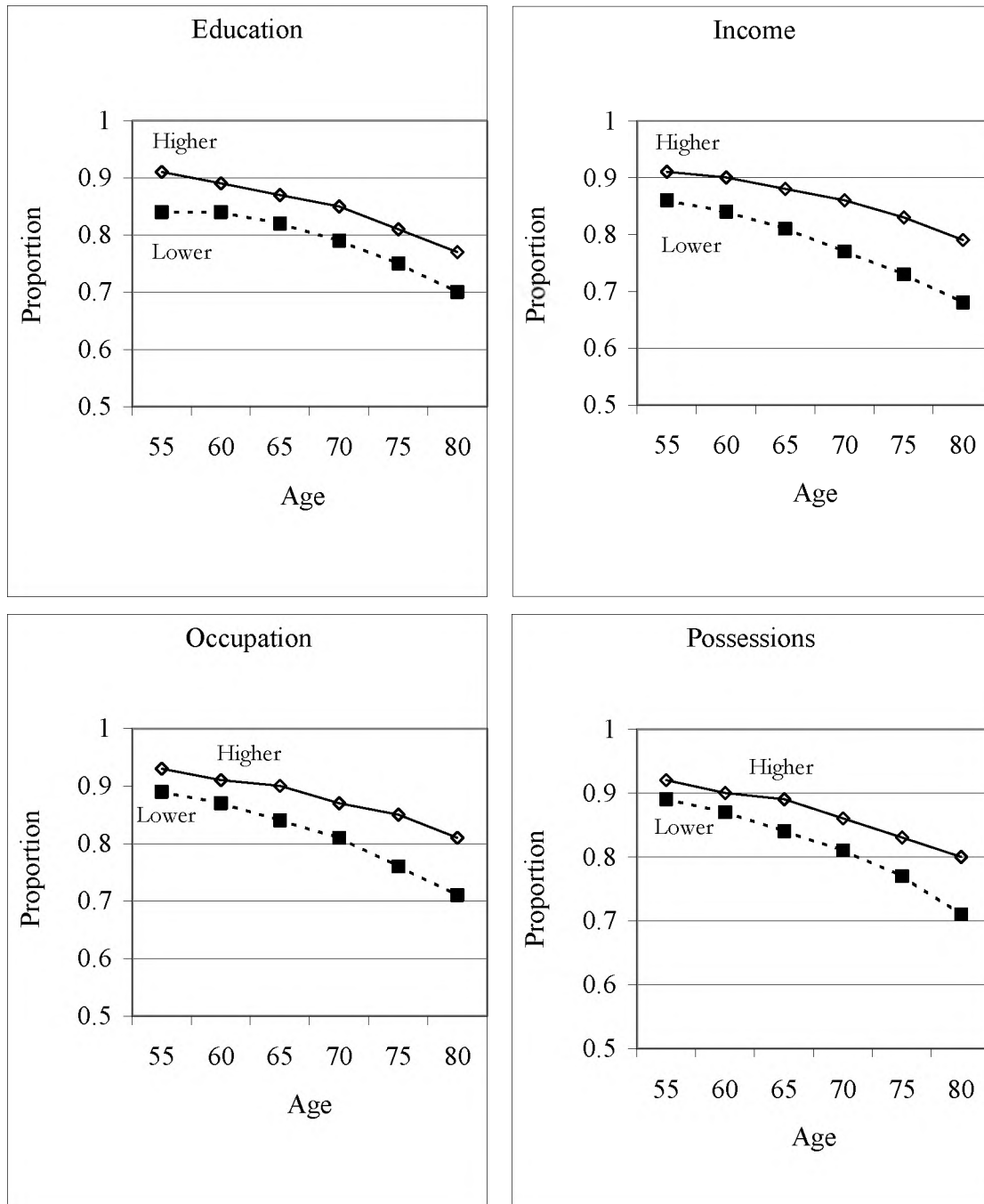
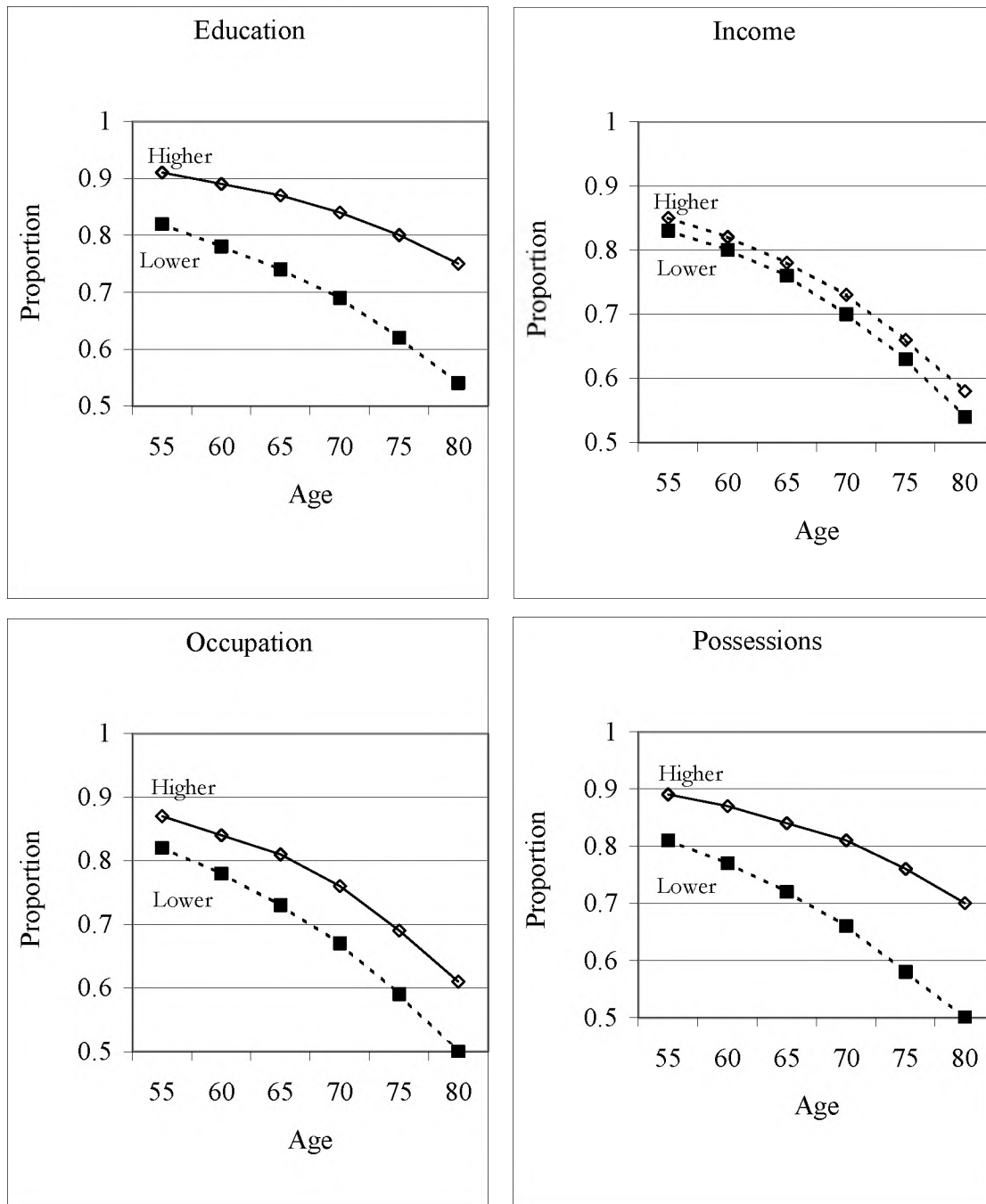


Figure 3 Proportion of Life Expected to be Spent in an Active State, by Women in Higher and Lower Status Groups, According to Four Indicators of Socioeconomic Status, by Age



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