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## The Complexity of Ergodic Games

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# The Complexity of Ergodic Games* 

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#### Abstract

We study finite-state two-player (zero-sum) concurrent mean-payoff games played on a graph. We focus on the important sub-class of ergodic games where all states are visited infinitely often with probability 1 . The algorithmic study of ergodic games was initiated in a seminal work of Hoffman and Karp in 1966, but all basic complexity questions have remained unresolved. Our main results for ergodic games are as follows: We establish (1) an optimal exponential bound on the patience of stationary strategies (where patience of a distribution is the inverse of the smallest positive probability and represents a complexity measure of a stationary strategy); (2) the approximation problem lie in FNP; (3) the approximation problem is at least as hard as the decision problem for simple stochastic games (for which $\mathrm{NP} \cap \operatorname{coNP}$ is the long-standing best known bound). We show that the exact value can be expressed in the existential theory of the reals, and also establish square-root sum hardness for a related class of games.


Keywords: Concurrent games; Mean-payoff objectives; Ergodic games; Approximation complexity.

## 1 Introduction

Concurrent games. Concurrent games are played over finite-state graphs by two players (Player 1 and Player 2) for an infinite number of rounds. In every round, both players simultaneously choose moves (or actions), and the current state and the joint moves determine a probability distribution over the successor states. The outcome of the game (or a play) is an infinite sequence of states and action pairs. Concurrent games were introduced in a seminal work by Shapley [29], and they are the most well-studied game models in stochastic graph games, with many important special cases.
Mean-payoff (limit-average) objectives. The most fundamental objective for concurrent games is the limitaverage (or mean-payoff) objective, where a reward is associated to every transition and the payoff of a play is the limit-inferior (or limit-superior) average of the rewards of the play. The original work of Shapley [29] considered discounted sum objectives (or games that stop with probability 1); and the class of concurrent games with limit-average objectives (or games that have zero stop probabilities) was introduced by Gillette in [17]. The Player-1 value $\operatorname{val}(s)$ of the game at a state $s$ is the supremum value of the expectation that Player 1 can guarantee for the limit-average objective against all strategies of Player 2. The games are

[^0]zero-sum, so the objective of Player 2 is the opposite. The study of concurrent mean-payoff games and its sub-classes have received huge attention over the last decades, both for mathematical results as well as algorithmic studies. Some key celebrated results are as follows: (1) the existence of values (or determinacy or equivalence of switching of strategy quantifiers for the players as in von-Neumann's min-max theorem) for concurrent discounted games was established in [29]; (2) the result of Blackwell-Ferguson established existence of values for the celebrated game of Big-Match [5]; and (3) developing on the results of [5] and Bewley-Kohlberg on Puisuex series [4] the existence of values for concurrent mean-payoff games was established by Mertens-Neyman [26].
Sub-classes. The general class of concurrent mean-payoff games is notoriously difficult for algorithmic analysis. The current best known solution for general concurrent mean-payoff games is achieved by a reduction to the theory of the reals over addition and multiplication with three quantifier alternations [7] (also see [19] for a better reduction for constant state spaces). The strategies that are required in general for concurrent mean-payoff games are infinite-memory strategies that depend in a complex way on the history of the game [26,5], and analysis of such strategies make the algorithmic study complicated. Hence several sub-classes of concurrent mean-payoff games have been studied algorithmically both in terms of restrictions of the graph structure and restrictions of the objective. The three prominent restrictions in terms of the graph structure are as follows: (1) Ergodic games (aka irreducible games) where every state is visited infinitely often almost-surely. (2) Turn-based stochastic games, where in each state at most one player can choose between multiple moves. (3) Deterministic games, where the transition functions are deterministic. The most well-studied restriction in terms of objective is the reachability objectives. A reachability objective consists of a set $U$ of terminal states (absorbing or sink states that are states with only self-loops), and the set $U$ is exactly the set of states where out-going transitions are assigned reward 1 and all other transitions are assigned reward 0 . For all these sub-classes, except deterministic mean-payoff games (that is ergodic mean-payoff games, concurrent reachability games, and turn-based stochastic mean-payoff games) stationary strategies are sufficient, where a stationary strategy is independent of the past history of the game and depends only on the current state.
Previous results. The decision problem of whether the value of the game at a state is at least a given threshold for turn-based stochastic reachability games (and also turn-based mean-payoff games with deterministic transition function) lie in $\mathrm{NP} \cap \operatorname{coNP}[8,32]$. They are among the rare and intriguing combinatorial problems that lie in NP $\cap$ coNP, but not known to be in PTIME. The existence of polynomial-time algorithms for the above decision questions are long-standing open problems. The algorithmic solution for turn-based games that is most efficient in practice is the strategy iteration algorithm, where the algorithm iterates over local improvement of strategies which is then established to converge to a globally optimal strategy. For ergodic games, Hoffman and Karp [21] presented a strategy iteration algorithm and also established that stationary strategies are sufficient for such games. For concurrent reachability games, again stationary strategies are sufficient (for $\epsilon$-optimal strategies, for all $\epsilon>0$ ) [12, 9]; the decision problem is in PSPACE and squareroot sum hard [10]. ${ }^{1}$
Key intriguing complexity questions. There are several key intriguing open questions related to the complexity of the various sub-classes of concurrent mean-payoff games. Some of them are as follows: (1) Does there exist a sub-class of concurrent games where the approximation problem is simpler than the exact decision problem, e.g., the decision problem is square-root sum hard, but the approximation problem can be solved in FNP? (2) There is no convergence result associated with the Hoffman-Karp algorithm for ergodic

[^1]games; and is it possible to establish a convergence for a strategy iteration algorithm for approximating the values of ergodic games. (3) The complexity of a stationary strategy is described by its patience which is the inverse of the minimum non-zero probability assigned to a move [12], and there is no bound known for the patience of stationary strategies for ergodic games.
Our results. The study of the ergodic games was initiated in the seminal work of Hoffman and Karp [21], and most of the complexity questions (related to computational, strategy, and algorithmic complexity) have remained open. In this work we focus on the complexity of simple generalizations of ergodic games (that will subsume ergodic games). Ergodic games form a very important sub-class of concurrent games subsuming the special cases of uni-chain Markov decision processes and uni-chain turn-based stochastic games (that have been studied in great depth in the literature with numerous applications, see [15, 27]). We consider generalizations of ergodic games called sure ergodic games where all plays are guaranteed to reach an ergodic component (a sub-game that is ergodic); and almost-sure ergodic games where with probability 1 an ergodic component is reached. Every ergodic game is sure ergodic, and every sure ergodic game is almostsure ergodic. Intuitively the generalizations allow to consider that after a finite prefix an ergodic component is reached.

1. (Strategy and approximation complexity). We show that for almost-sure ergodic games the optimal bound on patience required for $\epsilon$-optimal stationary strategies, for $\epsilon>0$, is exponential (we establish the upper bound for almost-sure ergodic games, and the lower bound for ergodic games). We then show that the approximation problem for turn-based stochastic ergodic mean-payoff games is at least as hard as solving the decision problem for turn-based stochastic reachability games (aka simple stochastic games); and finally show that the approximation problem belongs to FNP for almost-sure ergodic games. Observe that our results imply that improving our FNP-bound for the approximation problem to polynomial time would require solving the long-standing open question of whether the decision problem of turn-based stochastic reachability games can be solved in polynomial time.
2. (Algorithm). We present a variant of the Hoffman-Karp algorithm and show that for all $\epsilon$ approximation (for $\epsilon>0$ ) our algorithm converges with in exponential number of iterations for almost-sure ergodic games. Again our result is optimal, since even for turn-based stochastic reachability games the strategy iteration algorithms require exponential iterations [16, 13].
3. (Exact complexity). We show that the exact decision problem for almost-sure ergodic games can be expressed in the existential theory of the reals (in contrast to general concurrent mean-payoff games where quantifier alternations are required). Finally, we show that the exact decision problem for sure ergodic games is square-root sum hard.

Technical contribution and remarks. Our main result is establishing the optimal bound of exponential patience for $\epsilon$-optimal stationary strategies, for $\epsilon>0$, in almost-sure ergodic games. Our result is in sharp contrast to the optimal bound of double-exponential patience for concurrent reachability games [20], and also the double-exponential iterations required by the strategy iteration algorithm for concurrent reachability games [18]. Our upper bound on the exponential patience is achieved by a coupling argument. While coupling argument is a well-established tool in probability theory, to the best of our knowledge the argument has not been used for concurrent mean-payoff games before. Our lower bound example constructs a family of ergodic mean-payoff games where exponential patience is required. Our results provide a complete picture for almost-sure and sure ergodic games (subsuming ergodic games) in terms of strategy complexity, computational complexity, and algorithmic complexity; and present answers to some of the key intriguing open questions related to the computational complexity of concurrent mean-payoff games.

Comparison with results for Shapley games. For Shapley (concurrent discounted) games the fact that the approximation problem is in FNP is straight-forward to prove; and the more interesting and challenging question is whether the approximation problem can be solved in PPAD. The PPAD complexity for the approximation problem for Shapley games was established in [11]; and the PPAD complexity arguments use the existence of unique (Banach) fixpoint (due to contraction mapping) and the fact that weak approximation implies strong approximation. A PPAD complexity result for the class of ergodic games (in particular, whether weak approximation implies strong approximation) is a subject for future work.

## 2 Definitions

In this section we present the definitions of game structures, strategies, mean-payoff function, values, and other basic notions.
Probability distributions. For a finite set $A$, a probability distribution on $A$ is a function $\delta: A \rightarrow[0,1]$ such that $\sum_{a \in A} \delta(a)=1$. We denote the set of probability distributions on $A$ by $\mathcal{D}(A)$. Given a distribution $\delta \in \mathcal{D}(A)$, we denote by $\operatorname{Supp}(\delta)=\{x \in A \mid \delta(x)>0\}$ the support of the distribution $\delta$. We denote by $r$ the number of random states where the transition function is not deterministic, i.e., $r=\mid\left\{s \in S \mid \exists a_{1} \in\right.$ $\left.\Gamma_{1}(s), a_{2} \in \Gamma_{2}(s) .\left|\operatorname{Supp}\left(\delta\left(s, a_{1}, a_{2}\right)\right)\right| \geq 2\right\} \mid$.
Concurrent game structures. A concurrent stochastic game structure $G=\left(S, A, \Gamma_{1}, \Gamma_{2}, \delta\right)$ has the following components.

- A finite state space $S$ and a finite set $A$ of actions (or moves).
- Two move assignments $\Gamma_{1}, \Gamma_{2}: S \rightarrow 2^{A} \backslash \emptyset$. For $i \in\{1,2\}$, assignment $\Gamma_{i}$ associates with each state $s \in S$ the non-empty set $\Gamma_{i}(s) \subseteq A$ of moves available to Player $i$ at state $s$.
- A probabilistic transition function $\delta: S \times A \times A \rightarrow \mathcal{D}(S)$, which associates with every state $s \in S$ and moves $a_{1} \in \Gamma_{1}(s)$ and $a_{2} \in \Gamma_{2}(s)$, a probability distribution $\delta\left(s, a_{1}, a_{2}\right) \in \mathcal{D}(S)$ for the successor state.
We will denote by $\delta_{\min }$ the minimum non-zero transition probability, i.e., $\delta_{\min }=$ $\min _{s, t \in S} \min _{a_{1} \in \Gamma_{1}(s), a_{2} \in \Gamma_{2}(s)}\left\{\delta\left(s, a_{1}, a_{2}\right)(t) \mid \delta\left(s, a_{1}, a_{2}\right)(t)>0\right\}$. We will denote by $n$ the number of states (i.e., $n=|S|$ ), and by $m$ the maximal number of actions available for a player at a state (i.e., $\left.m=\max _{s \in S} \max \left\{\left|\Gamma_{1}(s)\right|,\left|\Gamma_{2}(s)\right|\right\}\right)$. We denote by $r$ the number of random states where the transition function is not deterministic, i.e., $r=\left|\left\{s \in S\left|\exists a_{1} \in \Gamma_{1}(s), a_{2} \in \Gamma_{2}(s) .\left|\operatorname{Supp}\left(\delta\left(s, a_{1}, a_{2}\right)\right)\right| \geq 2\right\} \mid\right.\right.$.
Plays. At every state $s \in S$, Player 1 chooses a move $a_{1} \in \Gamma_{1}(s)$, and simultaneously and independently Player 2 chooses a move $a_{2} \in \Gamma_{2}(s)$. The game then proceeds to the successor state $t$ with probability $\delta\left(s, a_{1}, a_{2}\right)(t)$, for all $t \in S$. A path or a play of $G$ is an infinite sequence $\pi=$ $\left(\left(s_{0}, a_{1}^{0}, a_{2}^{0}\right),\left(s_{1}, a_{1}^{1}, a_{2}^{1}\right),\left(s_{2}, a_{1}^{2}, a_{2}^{2}\right) \ldots\right)$ of states and action pairs such that for all $k \geq 0$ we have (i) $a_{1}^{k} \in \Gamma_{1}\left(s_{k}\right)$ and $a_{2}^{k} \in \Gamma_{2}\left(s_{k}\right)$; and (ii) $s_{k+1} \in \operatorname{Supp}\left(\delta\left(s_{k}, a_{1}^{k}, a_{2}^{k}\right)\right)$. We denote by $\Pi$ the set of all paths.
Strategies. A strategy for a player is a recipe that describes how to extend prefixes of a play. Formally, a strategy for Player $i \in\{1,2\}$ is a mapping $\sigma_{i}:(S \times A \times A)^{*} \times S \rightarrow \mathcal{D}(A)$ that associates with every finite sequence $x \in(S \times A \times A)^{*}$ of state and action pairs, and the current state $s$ in $S$, representing the past history of the game, a probability distribution $\sigma_{i}(x \cdot s)$ used to select the next move. The strategy $\sigma_{i}$ can prescribe only moves that are available to Player $i$; that is, for all sequences $x \in(S \times A \times A)^{*}$ and states $s \in S$, we require that $\operatorname{Supp}\left(\sigma_{i}(x \cdot s)\right) \subseteq \Gamma_{i}(s)$. We denote by $\Sigma_{i}$ the set of all strategies for Player $i \in\{1,2\}$. Once the starting state $s$ and the strategies $\sigma_{1}$ and $\sigma_{2}$ for the two players have been chosen, then we have a random walk $\pi_{s}^{\sigma_{1}, \sigma_{2}}$ for which the probabilities of events are uniquely defined [31], where an
event $\mathcal{A} \subseteq \Pi$ is a measurable set of paths. For an event $\mathcal{A} \subseteq \Pi$, we denote by $\operatorname{Pr}_{s}^{\sigma_{1}, \sigma_{2}}(\mathcal{A})$ the probability that a path belongs to $\mathcal{A}$ when the game starts from $s$ and the players use the strategies $\sigma_{1}$ and $\sigma_{2}$; and denote $\mathbb{E}_{s}^{\sigma_{1}, \sigma_{2}}[\cdot]$ as the associated expectation measure. We will consider in particular stationary and positional strategies. A strategy $\sigma_{i}$ is stationary (or memoryless) if it is independent of the history but only depends on the current state, i.e., for all $x, x^{\prime} \in(S \times A \times A)^{*}$ and all $s \in S$, we have $\sigma_{i}(x \cdot s)=\sigma_{i}\left(x^{\prime} \cdot s\right)$, and thus can be expressed as a function $\sigma_{i}: S \rightarrow \mathcal{D}(A)$. For stationary strategies, the complexity of the strategy is described by the patience of the strategy, which is the inverse of the minimum non-zero probability assigned to an action [12]. Formally, for a stationary strategy $\sigma_{i}: S \rightarrow \mathcal{D}(A)$ for Player $i$, the patience is $\max _{s \in S} \max _{a \in \Gamma_{i}(s)}\left\{\left.\frac{1}{\sigma_{i}(s)(a)} \right\rvert\, \sigma_{i}(s)(a)>0\right\}$. A strategy is pure (deterministic) if it does not use randomization, i.e., for any history there is always some unique action $a$ that is played with probability 1 . A pure stationary strategy $\sigma_{i}$ is also called a positional strategy, and represented as a function $\sigma_{i}: S \rightarrow A$. We call a pair of strategies $\left(\sigma_{1}, \sigma_{2}\right) \in \Sigma_{1} \times \Sigma_{2}$ a strategy profile.
The mean-payoff function. In this work we will consider limit-average (or mean-payoff) functions. We will consider concurrent games with a reward function $\mathrm{R}: S \times A \times A \rightarrow[0,1]$ that assigns a reward value $0 \leq \mathrm{R}\left(s, a_{1}, a_{2}\right) \leq 1$ for all $s \in S, a_{1} \in \Gamma_{1}(s)$, and $a_{2} \in \Gamma_{2}(s)$. For a path $\pi=\left(\left(s_{0}, a_{1}^{0}, a_{2}^{0}\right),\left(s_{1}, a_{1}^{1}, a_{2}^{1}\right), \ldots\right)$, the limit-inferior average (resp. limit-superior average) is defined as
 $\left.\sum_{i=0}^{n-1} \mathrm{R}\left(s_{i}, a_{1}^{i}, a_{2}^{i}\right)\right)$. For brevity we will denote concurrent games with mean-payoff functions as CMPGs (concurrent mean-payoff games).
Values and $\epsilon$-optimal strategies. Given a CMPG $G$ and a reward function R, the lower value $\underline{v}_{s}$ (resp. the upper value $\bar{v}_{s}$ ) at a state $s$ is defined as follows:

$$
\underline{v}_{s}=\sup _{\sigma_{1} \in \Sigma_{1}} \inf _{\sigma_{2} \in \Sigma_{2}} \mathbb{E}_{s}^{\sigma_{1}, \sigma_{2}}[\text { LimInfAvg }] ; \quad \bar{v}_{s}=\inf _{\sigma_{2} \in \Sigma_{2}} \sup _{\sigma_{1} \in \Sigma_{1}} \mathbb{E}_{s}^{\sigma_{1}, \sigma_{2}}[\text { LimSupAvg }] .
$$

The celebrated result of Mertens and Neyman [26] shows that the upper and lower value coincide and gives the value of the game denoted as $v_{s}$. For $\epsilon \geq 0$, a strategy $\sigma_{1}$ for Player 1 is $\epsilon$-optimal if we have $v_{s}-\epsilon \leq \inf _{\sigma_{2} \in \Sigma_{2}} \mathbb{E}_{s}^{\sigma_{1}, \sigma_{2}}[$ LimInfAvg]. An optimal strategy is a 0 -optimal strategy.
Game classes. We consider the following special classes of CMPGs.

1. Variants of ergodic CMPGs. Given a CMPG $G$, a set $C$ of states in $G$ is called an ergodic component, if for all states $s, t \in C$, for all strategy profiles $\left(\sigma_{1}, \sigma_{2}\right)$, if we start at $s$, then $t$ is visited infinitely often with probability 1 in the random walk $\pi_{s}^{\sigma_{1}, \sigma_{2}}$. A CMPG is ergodic if the set $S$ of states is an ergodic component. A CMPG is sure ergodic if for all strategy profiles $\left(\sigma_{1}, \sigma_{2}\right)$ and for all start states $s$, ergodic components are reached certainly (all plays reach some ergodic component). A CMPG is almost-sure ergodic if for all strategy profiles $\left(\sigma_{1}, \sigma_{2}\right)$ and for all start states $s$, ergodic components are reached with probability 1 . Observe that every ergodic CMPG is also a sure ergodic CMPG, and every sure ergodic CMPG is also an almost-sure ergodic CMPG.
2. Turn-based stochastic games, MDPs and SSGs. A game structure $G$ is turn-based stochastic if at every state at most one player can choose among multiple moves; that is, for every state $s \in S$ there exists at most one $i \in\{1,2\}$ with $\left|\Gamma_{i}(s)\right|>1$. A game structure is a Player-2 Markov decision process (MDP) if for all $s \in S$ we have $\left|\Gamma_{1}(s)\right|=1$, i.e., only Player 2 has choice of actions in the game, and Player-1 MDPs are defined analogously. A simple stochastic game (SSG) [8] is an almost-sure ergodic turn-based stochastic game with two ergodic components, where both the ergodic components (called terminal states) are a single absorbing state (an absorbing state has only a selfloop transition); one terminal state $(T)$ has reward 1 and the other terminal state $(\perp)$ has reward 0 ;
and all positive transition probabilities are either $\frac{1}{2}$ or 1 . The almost-sure reachability property to the ergodic components for SSGs is referred to as the stopping property [8].

Remark 1. The results of Hoffman and Karp [21] established that for ergodic CMPGs optimal stationary strategies exist (for both players). Moreover, for an ergodic CMPG the value for every state is the same, which is called the value of the game. We argue that the result for existence of optimal stationary strategies also extends to almost-sure ergodic CMPGs. Consider an almost-sure ergodic CMPG G. Notice first that in the ergodic components, there exist optimal stationary strategies, as shown by Hoffman and Karp [21]. Notice also that we eventually will reach some ergodic component, with probability 1 after a finite number of steps and therefore that we can ignore the rewards we get while doing so (since mean-payoff functions are independent of finite prefixes). Hence, we get an almost-sure reachability game, in the states which are not in the ergodic components, by considering any ergodic component $C$ to be a terminal with reward equal to the value of $C$. In such games it is easy to see that there exist optimal stationary strategies.

Value and the approximation problem. Given a CMPG $G$, a state $s$ of $G$, and a rational threshold $\lambda$, the value problem is the decision problem that asks whether $v_{s}$ is at most $\lambda$. Given a CMPG $G$, a state $s$ of $G$, and a tolerance $\epsilon>0$, the approximation problem asks to compute an interval of length $\epsilon$ such that the value $v_{s}$ lies in the interval. We will present the formal definition of the decision version of the approximation problem in Section 3.3. In the following sections we will consider the value problem and the approximation problem for almost-sure ergodic, sure ergodic, and ergodic games.

## 3 Complexity of Approximation for Almost-sure Ergodic Games

In this section we will present three results for almost-sure ergodic games: (1) First we establish (in Section 3.1) an optimal exponential bound on the patience of $\epsilon$-optimal stationary strategies, for all $\epsilon>0$. (2) Second we show (in Section 3.2) that the approximation problem (even for turn-based stochastic ergodic mean-payoff games) is at least as hard as solving the value problem for SSGs. (3) Finally, we show (in Section 3.3) that the approximation problem lies in FNP.

### 3.1 Strategy complexity

In this section we will present results related to $\epsilon$-optimal stationary strategies for almost-sure ergodic CMPGs, that on one hand will establish an optimal exponential bound for patience, and on the other hand will be used to establish the complexity of approximation of values in the following subsection. The results of this section will also be used in the algorithmic analysis in Section 4. We start with the notion of $q$-rounded strategies.
The classes of $q$-rounded distributions and strategies. For $q \in \mathbb{N}$, a distribution $d$ over a finite set $Z$ is a $q$-rounded distribution if for all $z \in Z$ we have that $d(z)=\frac{p}{q}$ for some number $p \in \mathbb{N}$. A stationary strategy $\sigma$ is a $q$-rounded strategy, if for all states $s$ the distribution $\sigma(s)$ is a $q$-rounded distribution.
Patience. Observe that the patience of a $q$-rounded strategy is at most $q$. We will show that for almost-sure ergodic CMPGs for all $\epsilon>0$ there are $q$-rounded $\epsilon$-optimal strategies, where $q$ is as follows:

$$
\left\lceil 4 \cdot \epsilon^{-1} \cdot m \cdot n^{2} \cdot\left(\delta_{\min }\right)^{-r}\right\rceil .
$$

This will immediately imply an exponential upper bound on the patience. We start with a lemma related to the probability of reaching states that are guaranteed to be reached with positive probability.

Lemma 2. Given a CMPG G, let s be a state in $G$, and $T$ be a set of states such that for all strategy profiles the set $T$ is reachable (with positive probability) from s. For all strategy profiles the probability to reach $T$ from $s$ in $n$ steps is at least $\left(\delta_{\min }\right)^{r}$ (where $r$ is the number of random states).

Proof. The basic idea of the proof is to consider a turn-based deterministic game where one player is Player 1 and Player 2 combined, and the opponent makes the choice for the probabilistic transitions. (The formal description of the turn-based deterministic game is as follows: $\left(S \cup\left(S \times A_{1} \times A_{2}\right),\left(A_{1} \times\right.\right.$ $\left.\left.A_{2}\right) \cup S \cup\{\perp\}, \bar{\Gamma}_{1}, \bar{\Gamma}_{2}, \bar{\delta}\right)$; where for all $s \in S$ and $a_{1} \in \Gamma_{1}(s)$ and $a_{2} \in \Gamma_{2}(s)$ we have $\bar{\Gamma}_{1}(s)=$ $\left\{\left(a_{1}, a_{2}\right) \mid a_{1} \in \Gamma_{1}(s), a_{2} \in \Gamma_{2}(s)\right\}$ and $\bar{\Gamma}_{1}\left(\left(s, a_{1}, a_{2}\right)\right)=\{\perp\} ; \bar{\Gamma}_{2}\left(\left(s, a_{1}, a_{2}\right)\right)=\operatorname{Supp}\left(\delta\left(s, a_{1}, a_{2}\right)\right)$ and $\bar{\Gamma}_{2}(s)=\{\perp\}$. The transition function is as follows: for all $s \in S$ and $a_{1} \in \Gamma_{1}(s)$ and $a_{2} \in \Gamma_{2}(s)$ we have $\bar{\delta}\left(s,\left(a_{1}, a_{2}\right), \perp\right)\left(\left(s, a_{1}, a_{2}\right)\right)=1$ and $\bar{\delta}\left(\left(s, a_{1}, a_{2}\right), \perp, t\right)(t)=1$.) In the turn-based deterministic game, against any strategy of the combined players, there is a positional strategy of the player making the probabilistic choices such that $T$ is reached after being in each state at most once certainly (by positional determinacy for turn-based deterministic reachability games [30]), as otherwise there would exist a positional strategy profile such that $T$ is never reached. The probability that exactly the choices made by the positional strategy of the probabilistic player in the turn-based deterministic game is executed once in each state in the original game is at least $\left(\delta_{\min }\right)^{r}$. Hence the desired result follows.

Variation distance. We will use a coupling argument in our proofs and this requires the definition of variation distance of two probability distributions. Given a finite set $Z$, and two distributions $d_{1}$ and $d_{2}$ over $Z$, the variation distance of the distributions is

$$
\operatorname{var}\left(d_{1}, d_{2}\right)=\frac{1}{2} \cdot \sum_{z \in Z}\left|d_{1}(z)-d_{2}(z)\right| .
$$

Coupling and coupling lemma. Let $Z$ be a finite set. For distributions $d_{1}$ and $d_{2}$ over the finite set $Z$, a coupling $\omega$ is a distribution over $Z \times Z$, such that for all $z \in Z$ we have $\sum_{z^{\prime} \in Z} \omega\left(z, z^{\prime}\right)=d_{1}(z)$ and also for all $z^{\prime} \in Z$ we have $\sum_{z \in Z} \omega\left(z, z^{\prime}\right)=d_{2}\left(z^{\prime}\right)$. We will only use the second part of coupling lemma [1] which is stated as follows:

- (Coupling lemma). For a pair of distributions $d_{1}$ and $d_{2}$, there exists a coupling $\omega$ of $d_{1}$ and $d_{2}$, such that for a random variable $(X, Y)$ from the distribution $\omega$, we have that $\operatorname{var}\left(d_{1}, d_{2}\right)=\operatorname{Pr}[X \neq Y]$.

We will now show that in almost-sure ergodic CMPGs strategies that play actions with probabilities "close" to what is played by an optimal strategy also achieve values that are "close" to the values achieved by the optimal strategy.

Lemma 3. Consider an almost-sure ergodic CMPG and let $\epsilon>0$ be a real number. Let $\sigma_{1}$ be an optimal stationary strategy for Player 1. Let $\sigma_{1}^{\prime}$ be a stationary strategy for Player 1 s.t. $\sigma_{1}^{\prime}(s)(a) \in\left[\sigma_{1}(s)(a)-\right.$ $\left.\frac{1}{q} ; \sigma_{1}^{\prime}(s)(a)+\frac{1}{q}\right]$, where $q=4 \cdot \epsilon^{-1} \cdot m \cdot n^{2} \cdot\left(\delta_{\min }\right)^{-r}$, for all states $s$ and actions $a \in \Gamma_{1}(s)$. Then the strategy $\sigma_{1}^{\prime}$ is an $\epsilon$-optimal strategy.

Proof. First observe that we can consider $\epsilon \leq 1$, because as the rewards are in the interval $[0,1]$ any strategy is an $\epsilon$-optimal strategy for $\epsilon \geq 1$. The proof will be split up in two parts, and the second part will use the first. The first part is related to plays starting in an ergodic component; and the second part is the other case. In both cases we will show that $\sigma_{1}^{\prime}$ guarantees mean-payoff within $\epsilon$ of the mean-payoff guaranteed by $\sigma_{1}$, thus implying the statement. Let $\sigma_{2}$ be a positional best response strategy against $\sigma_{1}^{\prime}$. Our proof is based on a novel coupling argument. The precise nature of the coupling argument is different in the two parts, but
both will use the following: For any state $s$, it is clear that the variation distance between $\sigma_{1}^{\prime}(s)$ and $\sigma_{1}(s)$ is at most $\frac{\left|\Gamma_{1}(s)\right|}{2 \cdot q}$, by definition of $\sigma_{1}^{\prime}(s)$. For a state $s$, let $d_{1}^{s}$ be the distribution over states defined as follows: for $t \in S$ we have $d_{1}^{s}(t)=\sum_{a_{1} \in \Gamma_{1}(s)} \sum_{a_{2} \in \Gamma_{2}(s)} \delta\left(s, a_{1}, a_{2}\right)(t) \cdot \sigma_{1}(s)\left(a_{1}\right) \cdot \sigma_{2}(s)\left(a_{2}\right)$. Define $d_{2}^{s}$ similarly using $\sigma_{1}^{\prime}(s)$ instead of $\sigma_{1}(s)$. Then $d_{1}^{s}$ and $d_{2}^{s}$ also have a variation distance of at most $\frac{\left|\Gamma_{1}(s)\right|}{2 \cdot q} \leq \frac{m}{2 \cdot q}$. Let $s_{0}$ be the start state, and $P=\pi_{s_{0}}^{\sigma_{1}, \sigma_{2}}$ be the random walk from $s_{0}$, where Player 1 follows $\sigma_{1}$ and Player 2 follows $\sigma_{2}$. Also let $P^{\prime}=\pi_{s_{0}}^{\sigma_{1}^{\prime}, \sigma_{2}}$ be the similar defined walk, except that Player 1 follows $\sigma_{1}^{\prime}$ instead of $\sigma_{1}$. Let $X^{i}$ be the random variable indicating the $i$-th state of $P$, and let $Y^{i}$ be the similar defined random variable in $P^{\prime}$ instead of $P$.
The state $s_{0}$ is in an ergodic component. Consider first the case where $s_{0}$ is part of an ergodic component. Irrespective of the strategy profile, all states of the ergodic component are visited infinitely often almostsurely (by definition of an ergodic component). Hence, we can apply Lemma 2 and obtain that we require at most $n \cdot\left(\delta_{\min }\right)^{r}=\frac{\epsilon \cdot q}{4 \cdot n \cdot m}$ steps in expectation to get from one state of the component to any other state of the component.
Coupling argument. We will now construct a coupling argument. We will define the coupling using induction. First observe that $X^{0}=Y^{0}=s_{0}$ (the starting state). For $i, j \in \mathbb{N}$, let $a_{i, j} \geq 0$ be the smallest number such that $X^{i+1}=Y^{j+1+a_{i, j}}$. By the preceding we know that $a_{i, j}$ exists for all $i, j$ with probability 1 and $a_{i, j} \leq \frac{\epsilon \cdot q}{4 \cdot n \cdot m}$ in expectation. The coupling is done as follows: (1) (Base case): Couple $X^{0}$ and $Y^{0}$. We have that $X^{0}=Y^{0}$; (2) (Inductive case): (i) if $X^{i}$ is coupled to $Y^{j}$ and $X^{i}=Y^{j}=s_{i}$, then also couple $X^{i+1}$ and $Y^{j+1}$ such that $\operatorname{Pr}\left[X^{i+1} \neq Y^{j+1}\right]=\operatorname{var}\left(d_{1}^{s_{i}}, d_{2}^{s_{i}}\right.$ ) (using coupling lemma); (ii) if $X^{i}$ is coupled to $Y^{j}$, but $X^{i} \neq Y^{j}$, then $X^{i+1}=Y^{j+1+a_{i, j}}=s_{i+1}$ and $X^{i+1}$ is coupled to $Y^{j+1+a_{i, j}}$, and we couple $X^{i+2}$ and $Y^{j+2+a_{i, j}}$ such that $\operatorname{Pr}\left[X^{i+2} \neq Y^{j+2+a_{i, j}}\right]=\operatorname{var}\left(d_{1}^{s_{i+1}}, d_{2}^{s_{i+1}}\right)$ (using coupling lemma). Notice that all $X^{i}$ are coupled to some $Y^{j}$ almost-surely; and moreover in expectation $\frac{j}{i}$ is bounded as follows:

$$
\frac{j}{i} \leq 1+\frac{m}{2 \cdot q} \cdot \frac{\epsilon \cdot q}{4 \cdot n \cdot m}=1+\frac{\epsilon}{8 \cdot n}
$$

The expression can be understood as follows: consider $X^{i}$ being coupled to $Y^{j}$. With probability at most $\frac{m}{2 \cdot q}$ they differ. In that case $X^{i+1}$ is coupled to $Y^{j+1+a_{i, j}}$. Otherwise $X^{i+1}$ is coupled to $Y^{j+1}$. By using our bound on $a_{i, j}$ we get the desired expression. For a state $s$, let $f_{s}$ (resp. $f_{s}^{\prime}$ ) denote the limit-average frequency of $s$ given $\sigma_{1}$ (resp. $\sigma_{1}^{\prime}$ ) and $\sigma_{2}$. Then it follows easily that for every state $s$, we have $\left|f_{s}-f_{s}^{\prime}\right| \leq \frac{\epsilon}{8 \cdot n}$. The formal argument is as follows: for every state $s$, consider the reward function $\mathrm{R}_{s}$ that assigns reward 1 to all transitions from $s$ and 0 otherwise; and then it is clear that the difference of the mean-payoffs of $P$ and $P^{\prime}$ is maximized if the mean-payoff of $P$ is 1 under $\mathrm{R}_{s}$ and the rewards of the steps of $P^{\prime}$ that are not coupled to $P$ are 0 . In that case the mean-payoff of $P^{\prime}$ under $\mathrm{R}_{s}$ is at least $\frac{1}{1+\frac{\epsilon}{8 \cdot n}}>1-\frac{\epsilon}{8 \cdot n}$ (since $\left.1>1-\left(\frac{\epsilon}{8 \cdot n}\right)^{2}=\left(1+\frac{\epsilon}{8 \cdot n}\right)\left(1-\frac{\epsilon}{8 \cdot n}\right)\right)$ in expectation and thus the difference between the mean-payoff of $P$ and the mean-payoff of $P^{\prime}$ under $\mathrm{R}_{s}$ is at most $\frac{\epsilon}{8 \cdot n}$ in expectation. The mean-payoff value if Player 1 follows a stationary strategy $\sigma_{1}^{1}$ and Player 2 follows a stationary strategy $\sigma_{2}^{1}$, such that the frequencies of the states encountered is $f_{s}^{1}$, is $\sum_{s \in S} \sum_{a_{1} \in \Gamma_{1}(s)} \sum_{a_{2} \in \Gamma_{2}(s)} f_{s}^{1} \cdot \sigma_{1}^{1}(s)\left(a_{1}\right) \cdot \sigma_{2}^{1}(s)\left(a_{2}\right) \cdot \mathrm{R}\left(s, a_{1}, a_{2}\right)$. Thus the differences in mean-payoff value when Player 1 follows $\sigma_{1}$ (resp. $\sigma_{1}^{\prime}$ ) and Player 2 follows the positional strategy $\sigma_{2}$, which plays action $a_{2}^{s}$ in state $s$, is

$$
\sum_{s \in S} \sum_{a_{1} \in \Gamma_{1}(s)}\left(f_{s} \cdot \sigma_{1}(s)\left(a_{1}\right)-f_{s}^{\prime} \cdot \sigma_{1}^{\prime}(s)\left(a_{1}\right)\right) \cdot \mathrm{R}\left(s, a_{1}, a_{2}^{s}\right)
$$

Since $\left|f_{s}-f_{s}^{\prime}\right| \leq \frac{\epsilon}{8 \cdot n}$ (by the preceding argument) and $\left|\sigma_{1}(s)\left(a_{1}\right)-\sigma_{1}^{\prime}(s)\left(a_{1}\right)\right| \leq \frac{1}{q}$ for all $s \in S$ and
$a_{1} \in \Gamma_{1}(s)$ (by definition), we have the following inequality

$$
\begin{aligned}
& \sum_{s \in S} \sum_{a_{1} \in \Gamma_{1}(s)}\left(f_{s} \cdot \sigma_{1}(s)\left(a_{1}\right)-f_{s}^{\prime} \cdot \sigma_{1}^{\prime}(s)\left(a_{1}\right)\right) \cdot \mathrm{R}\left(s, a_{1}, a_{2}^{s}\right) \\
\leq & \sum_{s \in S} \sum_{a_{1} \in \Gamma_{1}(s)}\left|f_{s} \cdot \sigma_{1}(s)\left(a_{1}\right)-\left(f_{s}-\frac{\epsilon}{8 \cdot n}\right) \cdot\left(\sigma_{1}(s)\left(a_{1}\right)-\frac{1}{q}\right)\right| \\
= & \sum_{s \in S} \sum_{a_{1} \in \Gamma_{1}(s)}\left|\frac{\epsilon}{8 \cdot n} \cdot \sigma_{1}(s)\left(a_{1}\right)+f_{s} \cdot \frac{1}{q}-\frac{\epsilon}{8 \cdot n \cdot q}\right| \\
\leq & \sum_{s \in S}\left(\frac{\epsilon}{8 \cdot n}+\frac{f_{s} \cdot m}{q}+\frac{\epsilon \cdot m}{8 \cdot n \cdot q}\right) \\
= & \frac{\epsilon}{8}+\frac{m}{q}+\frac{\epsilon \cdot m}{8 \cdot q} \leq \frac{\epsilon}{8}+\frac{\epsilon}{4}+\frac{\epsilon}{8}=\frac{\epsilon}{2}
\end{aligned}
$$

The first inequality uses that $\mathrm{R}\left(s, a_{1}, a_{2}^{s}\right) \leq 1$ and the preceding comments on the differences. The second inequality uses that (a) when we sum over $\sigma_{1}(s)\left(a_{1}\right)$ for all $a_{1}$, for a fixed $s \in S$, we get 1 ; (b) $\left|\Gamma_{1}(s)\right| \leq m$. The following equality uses that $\sum_{s \in S} f_{s}=1$ since they represent frequencies. Finally since $4 \cdot m \cdot n \cdot \epsilon^{-1} \leq$ $q, \epsilon \leq 1$, and $n \geq 1$ we have $\frac{m}{q} \leq \frac{\epsilon}{4}$ and $\frac{\epsilon \cdot m}{8 \cdot q} \leq \frac{\epsilon}{32} \leq \frac{\epsilon}{8}$. The desired inequality is established.

The state $s_{0}$ is not in an ergodic component. Now consider the case where the start state $s_{0}$ is not part of an ergodic component. We will divide the walks $P$ and $P^{\prime}$ into two parts. The part inside some ergodic component and the part outside all ergodic components. If $P$ and $P^{\prime}$ ends up in the same ergodic component, then the mean-payoff differs by at most $\frac{\epsilon}{2}$ in expectation, by the first part. For any pair of strategies the random walk defined from them almost-surely reaches some ergodic component (since we consider almostsure ergodic CMPGs). Hence, we can apply Lemma 2 and see that we require at most $n \cdot\left(\delta_{\min }\right)^{r}=\frac{\epsilon \cdot q}{4 \cdot n \cdot m}$ steps in expectation before we reach an ergodic component.
Coupling argument. To find the probability that they end up in the same component we will again make a coupling argument. Notice that $X^{0}=Y^{0}=s_{0}$. We will now make the coupling using induction. (1) (Base case): Make a coupling between $X^{1}$ and $Y^{1}$, such that $\operatorname{Pr}\left[X^{1} \neq Y^{1}\right]=\operatorname{var}\left(d_{1}^{s_{0}}, d_{2}^{s_{0}}\right) \leq \frac{\left|\Gamma_{1}\left(s_{0}\right)\right|}{2 \cdot q} \leq \frac{m}{2 \cdot q}$ (such a coupling exists by the coupling lemma). (2) (Inductive case): Also, if there is a coupling between $X^{i}$ and $Y^{i}$ and $X^{i}=Y^{i}=s_{i}$, then also make a coupling between $X^{i+1}$ and $Y^{i+1}$, such that $\operatorname{Pr}\left[X^{i+1} \neq Y^{i+1}\right]=$ $\operatorname{var}\left(d_{1}^{s_{i}}, d_{2}^{s_{i}}\right) \leq \frac{\left|\Gamma_{1}\left(s_{i}\right)\right|}{2 \cdot q} \leq \frac{m}{2 \cdot q}$ (such a coupling exists by the coupling lemma). Let $\ell$ be the smallest number such that $X^{\ell}$ is some state in an ergodic component. In expectation, $\ell$ is at most $\frac{\ell \cdot q}{4 \cdot n \cdot m}$. The probability that $X^{i} \neq Y^{i}$ for some $0 \leq i \leq \ell$ is by union bound at most $\frac{m}{2 \cdot q} \cdot \frac{\epsilon \cdot q}{4 \cdot n \cdot m} \leq \frac{\epsilon}{8 \cdot n} \leq \frac{\epsilon}{2}$ in expectation. If that is not the case, then $P$ and $P^{\prime}$ do end up in the same ergodic component. In the worst case, the component the walk $P$ ends up in has value 1 and the component that the walk $P^{\prime}$ ends up in (if they differ) has value 0 . Therefore, with probability at most $\frac{\epsilon}{2}$ the walk $P^{\prime}$ ends up in an ergodic component of value 0 (and hence has mean-payoff 0 ); and otherwise it ends up in the same component as $P$ does and thus gets the same mean-payoff as $P$, except for at most $\frac{\epsilon}{2}$, as we established in the first part. Thus $P^{\prime}$ must ensure the same mean-payoff as $P$ except for $\frac{2 \epsilon}{2}=\epsilon$. We therefore get that $\sigma_{1}^{\prime}$ is an $\epsilon$-optimal strategy (since $\sigma_{1}$ is optimal).

We show that for every integer $q^{\prime} \geq \ell$, for every distribution over $\ell$ elements, there exists a $q^{\prime}$-rounded
distribution "close" to it. Together with Lemma 3 it will show the existence of $q$ '-rounded $\epsilon$-optimal strategies, for every integer $q^{\prime}$ greater than the $q$ defined in Lemma 3.

Lemma 4. Let $d_{1}$ be a distribution over a finite set $Z$ of size $\ell$. Then for all integers $q \geq \ell$ there exists a $q$-rounded distribution $d_{2}$ over $Z$, such that $\left|d_{1}(z)-d_{2}(z)\right|<\frac{1}{q}$.

Proof. WLOG we consider that $\ell \geq 2$ (since the unique distribution over a singleton set clearly have the desired properties for all integers $q \geq 1$ ). Given distribution $d_{1}$ we will construct a witness distribution $d_{2}$. There are two cases. Either (i) there is an element $z \in Z$ such that $\frac{1}{q} \leq d_{1}(z) \leq 1-\frac{1}{q}$, or (ii) no such element exists.

- We will first consider case (ii), i.e., there exists no element $z$ such that $\frac{1}{q} \leq d_{1}(z) \leq 1-\frac{1}{q}$. Consider an element $z^{*} \in Z$ such that $1-\frac{1}{q}<d_{1}\left(z^{*}\right)$. Precisely only one such element exists in this case since not all $\ell$ elements can have probability strictly less than $\frac{1}{q} \leq \frac{1}{\ell}$, and no more than one element can have probability strictly more than $1-\frac{1}{q} \geq \frac{1}{2}$. Then let $d_{2}\left(z^{*}\right)=1$ and $d_{2}(z)=0$ for all other elements in $Z$. This clearly ensures that $\left|d_{1}(z)-d_{2}(z)\right|<\frac{1}{q}$ for all $z \in Z$ and that $d_{2}$ is a $q$-rounded distribution.
- Now we consider case (i). Let $z^{\ell}$ be an arbitrary element in $Z$ such that $\frac{1}{q} \leq d_{1}\left(z^{\ell}\right) \leq 1-\frac{1}{q}$. Let $\left\{z^{1}, \ldots, z^{\ell-1}\right\}$ be an arbitrary ordering of the remaining elements. We will now construct $d_{2}$ iteratively such that in step $k$ we have assigned probability to $\left\{z^{1}, \ldots, z^{k}\right\}$. We will establish the following iterative property: in step $k$ we have that $\sum_{c=1}^{k}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right) \in\left(-\frac{1}{q} ; \frac{1}{q}\right)$. The iteration stops when $k=\ell-1$, and then we will assign $d_{2}\left(z^{\ell}\right)$ the probability $1-\sum_{c=1}^{\ell-1} d_{2}\left(z^{c}\right)$. For all $1 \leq k \leq \ell-1$, the iterative definition of $d_{2}\left(z^{k}\right)$ is as follows:

$$
d_{2}\left(z^{k}\right)= \begin{cases}\frac{\left\lfloor q \cdot d_{1}\left(z^{k}\right)\right\rfloor}{q} & \text { if } \sum_{c=1}^{k-1}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right)<0 \\ \frac{\left\lceil q \cdot d_{1}\left(z^{k}\right)\right\rceil}{q} & \text { if } \sum_{c=1}^{k-1}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right) \geq 0\end{cases}
$$

We use the standard convention that the empty sum is 0 . For $1 \leq k \leq \ell-1$, observe that (a) $\mid d_{1}\left(z^{k}\right)-$ $d_{2}\left(z^{k}\right) \left\lvert\,<\frac{1}{q}\right.$; and (b) since $d_{1}\left(z^{k}\right) \in[0,1]$ also $d_{2}\left(z^{k}\right)$ is in $[0 ; 1]$. Moreover, there exists an integer $p$ such that $d_{2}\left(z^{k}\right)=\frac{p}{q}$. We have that

$$
d_{1}\left(z^{k}\right)-\frac{1}{q}<\frac{\left\lfloor q \cdot d_{1}\left(z^{k}\right)\right\rfloor}{q} \leq d_{1}\left(z^{k}\right) \leq \frac{\left\lceil q \cdot d_{1}\left(z^{k}\right)\right\rceil}{q}<d_{1}\left(z^{k}\right)+\frac{1}{q}
$$

Thus, if the sum $\sum_{c=1}^{k-1}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right)$ is negative, then we have that

$$
\frac{-1}{q}<\sum_{c=1}^{k-1}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right) \leq \sum_{c=1}^{k}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right)<\sum_{c=1}^{k-1}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right)+\frac{1}{q}<\frac{1}{q},
$$

where the first inequality is the iterative property (by induction for $k-1$ ); the second inequality follows because in this case we have $d_{2}\left(z^{k}\right)=\frac{\left\lfloor q \cdot d_{1}\left(z^{k}\right)\right\rfloor}{q} \leq d_{1}\left(z^{k}\right)$ by ( $\ddagger$ ); the third inequality follows since $d_{1}\left(z^{k}\right)-d_{2}\left(z^{k}\right)=d_{1}\left(z^{k}\right)-\frac{\left\lfloor q \cdot d_{1}\left(z^{k}\right)\right\rfloor}{q}<\frac{1}{q}$ by ( $\ddagger$ ); the final inequality follows since
$\sum_{c=1}^{k-1}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right)$ is negative. Symmetrically, if the sum $\sum_{c=1}^{k-1}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right)$ is not negative, then we have that

$$
\frac{1}{q}>\sum_{c=1}^{k-1}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right) \geq \sum_{c=1}^{k}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right)>\sum_{c=1}^{k-1}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right)-\frac{1}{q} \geq \frac{-1}{q}
$$

using the iterative property (by induction) and the inequalities of $(\ddagger)$ as in the previous case. Thus, in either case, we have that $\frac{-1}{q}<\sum_{c=1}^{k}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right)<\frac{1}{q}$, establishing the iterative property by induction.
Finally we need to consider $z^{\ell}$. First, we will show that $\left|d_{1}\left(z^{\ell}\right)-d_{2}\left(z^{\ell}\right)\right|<\frac{1}{q}$. We have that

$$
d_{2}\left(z^{\ell}\right)=1-\sum_{c=1}^{\ell-1} d_{2}\left(z^{c}\right)=\sum_{c=1}^{\ell}\left(d_{1}\left(z^{c}\right)\right)-\sum_{c=1}^{\ell-1}\left(d_{2}\left(z^{c}\right)\right)=d_{1}\left(z^{\ell}\right)+\sum_{c=1}^{\ell-1}\left(d_{1}\left(z^{c}\right)-d_{2}\left(z^{c}\right)\right)
$$

Hence $\left|d_{1}\left(z^{\ell}\right)-d_{2}\left(z^{\ell}\right)\right|<\frac{1}{q}$, by our iterative property. This also ensures that $d_{2}\left(z^{\ell}\right) \in[0 ; 1]$, since $d_{1}\left(z^{\ell}\right) \in\left[\frac{1}{q} ; 1-\frac{1}{q}\right]$, by definition. Thus, $d_{2}$ is a distribution over $Z$ (since it is clear that $\sum_{z \in Z} d_{2}(z)=$ 1 , because of the definition of $d_{2}\left(z^{\ell}\right)$ and we have shown for all $z \in Z$ that $\left.d_{2}(z) \in[0 ; 1]\right)$. Since we have ensured that for each $z \in\left(Z \backslash\left\{z^{\ell}\right\}\right)$ that $d_{2}(z)=\frac{p}{q}$ for some integer $p$, it follows that $d_{2}\left(z^{\ell}\right)=\frac{p^{\prime}}{q}$ for some integer $p^{\prime}$ (since $q$ is an integer). This implies that $d_{2}$ is a $q$-rounded distribution. We also have $\left|d_{1}(z)-d_{2}(z)\right|<\frac{1}{q}$ for all $z \in Z$ (by $(\ddagger)$ ) and thus all the desired properties have been established.

This completes the proof.
Corollary 5. For all almost-sure ergodic CMPGs, for all $\epsilon>0$, there exists an $\epsilon$-optimal, $q^{\prime}$-rounded strategy $\sigma_{1}$ for Player 1, for all integers $q^{\prime} \geq q$, where

$$
q=4 \cdot \epsilon^{-1} \cdot m \cdot n^{2} \cdot\left(\delta_{\min }\right)^{-r}
$$

Proof. Notice that the $q$ defined here is the same $q$ as is defined in Lemma 3. Let the integer $q^{\prime} \geq q$ be given. Consider an almost-sure ergodic CMPG $G$. Let $\sigma_{1}^{\prime}$ be a optimal stationary strategy in $G$ for Player 1. For each state $s$, pick a $q^{\prime}$-rounded distribution $d^{s}$ over $\Gamma_{1}(s)$, such that $\left|\sigma_{1}^{\prime}(s)\left(a_{1}\right)-d^{s}\left(a_{1}\right)\right|<\frac{1}{q^{\prime}} \leq \frac{1}{q}$ for all $a_{1} \in \Gamma_{1}(s)$. Such a distribution exists by Lemma 4 , since $q^{\prime} \geq q \geq m \geq\left|\Gamma_{1}(s)\right|$. Let the strategy $\sigma_{1}$ be defined as follows: $\sigma_{1}(s)=d^{s}$ for each state $s \in S$. Hence $\sigma_{1}$ is a $q^{\prime}$-rounded strategy. By Lemma 3, the strategy $\sigma_{1}$ is also an $\epsilon$-optimal strategy.

Exponential lower bound on patience. We will now present a family of ergodic CMPGs where the lower bound on patience is exponential in $r$. We will present the lower bound on a special class of ergodic CMPGs, namely, skew-symmetric ergodic CMPGs which we define below.
Skew-symmetric CMPGs. A CMPG $G$ is skew-symmetric ${ }^{2}$, if there is a bijective map $f: S \rightarrow S$, where $f(f(s))=s$, (for all $s$ we will use $\bar{s}$ to denote $f(s)$ ) where the following holds: For each state $s$, there is a bijective map $f_{1}^{s}: \Gamma_{1}(s) \rightarrow \Gamma_{2}(\bar{s})$ (for all $i \in \Gamma_{1}(s)$ we will use $\bar{i}$ to denote $f_{1}^{s}(i)$ ) and a bijective map $f_{2}^{s}: \Gamma_{2}(s) \rightarrow \Gamma_{1}(\bar{s})$ (similarly to the first map, for all $j \in \Gamma_{2}(s)$ we will use $\bar{j}$ to denote $f_{2}^{s}(j)$ ), such that for

[^2]all $i \in \Gamma_{1}(s)$ and all $j \in \Gamma_{2}(s)$, the following conditions hold: (1) we have $\mathrm{R}(s, i, j)=1-\mathrm{R}(\bar{s}, \bar{j}, \bar{i})$; (2) for all $s^{\prime}$ such that $\delta(s, i, j)\left(s^{\prime}\right)>0$, we have $\delta(\bar{s}, \bar{j}, \bar{i})\left(\bar{s}^{\prime}\right)=\delta(s, i, j)\left(s^{\prime}\right)$; and (3) we have $f_{2}^{\bar{s}}\left(f_{1}^{s}(i)\right)=i$ and that $f_{1}^{\bar{s}}\left(f_{2}^{s}(j)\right)=j$.

Lemma 6. Consider a skew-symmetric CMPG G. Then for all s we have $v_{s}=1-v_{\bar{s}}$.
Proof. Let $s$ be a state. For a stationary strategy $\sigma_{k}$ for Player $k, k \in\{1,2\}$, let $\bar{\sigma}_{k}$ be a stationary strategy for the other player defined as follows: For each state $s$ and action $i \in \Gamma_{k}(s)$, let $\bar{\sigma}_{k}(\bar{s})(\bar{i})=\sigma_{k}(s)(i)$. For a stationary strategy $\sigma_{1}$ for Player 1 , consider the stationary strategy profile $\left(\sigma_{1}, \bar{\sigma}_{1}\right)$. For the random walk $P=\pi_{s}^{\sigma_{1}, \bar{\sigma}_{1}}$, where the players follows $\left(\sigma_{1}, \bar{\sigma}_{1}\right)$, starting in $s$ corresponds to the random walk $\bar{P}=\pi_{\bar{s}}^{\sigma_{1}, \bar{\sigma}_{1}}$, where the players follows $\left(\sigma_{1}, \bar{\sigma}_{1}\right)$, starting in $\bar{s}$, in the obvious way (that is: if $P$ is in state $s_{i}$ in the $i$-th step and the reward is $\lambda$, then $P^{\prime}$ is in $\bar{s}_{i}$, in the $i$-th step and the reward is $1-\lambda$ ). The two random walks, $P$ and $P^{\prime}$, are equally likely. This implies that $v_{s}=1-v_{\bar{s}}$.

Corollary 7. For all skew-symmetric ergodic CMPGs the value is $\frac{1}{2}$.
Family $G_{\eta}^{k}$. We will now provide a lower bound for patience of $\epsilon$-optimal strategies in skew-symmetric ergodic CMPGs. More precisely, we will give a family of games $\left\{G_{\eta}^{k} \left\lvert\, k \geq 2 \vee 0<\eta<\frac{1}{4 \cdot k+4}\right.\right\}$, such that $G_{\eta}^{k}$ consists of $2 \cdot k+5$ states and such that $\delta_{\min }$ for $G_{\eta}^{k}$ is $\eta$. The game $G_{\eta}^{k}$ will be such that all $\frac{1}{48}$-optimal stationary strategies require patience at least $\frac{1}{2 \cdot \eta^{k / 2}}$.
Construction of the family $G_{\eta}^{k}$. For a given $k \geq 2$ and $\eta$, such that $0<\eta<\frac{1}{4 \cdot k+4}$, let the game $G_{\eta}^{k}$ be as follows: The game consists of $2 \cdot k+5$ states, $S=\left\{a, b, \bar{b}, c, \bar{c}, s_{1}, \bar{s}_{1}, s_{2}, \bar{s}_{2}, \ldots, s_{k}, \bar{s}_{k}\right\}$. For $s \in(S \backslash\{c, \bar{c}\})$, we have that $\left|\Gamma_{1}(s)\right|=\left|\Gamma_{2}(s)\right|=1$. For $s^{\prime} \in\{c, \bar{c}\}$, we have that $\left|\Gamma_{1}\left(s^{\prime}\right)\right|=\left|\Gamma_{2}\left(s^{\prime}\right)\right|=2$, and let $\Gamma_{1}\left(s^{\prime}\right)=\left\{i_{1}^{s^{\prime}}, i_{2}^{s^{\prime}}\right\}$ and $\Gamma_{2}\left(s^{\prime}\right)=\left\{j_{1}^{s^{\prime}}, j_{2}^{s^{\prime}}\right\}$. For $y \geq 2$ we have that $s_{y}$ (resp. $\bar{s}_{y}$ ) has a transition to $s_{k}$ (resp. $\bar{s}_{k}$ ) of probability $1-\eta$; to $s_{y-1}$ (resp. $\bar{s}_{y-1}$ ), where $s_{0}=\bar{s}_{0}=a$, with probability $\eta$; and also the reward of the transition is 0 (resp. 1). The state $b$ (resp. $\bar{b}$ ) is deterministic and has a transition to $a$ of reward 0 (resp. 1). The transition function at state $c$ is deterministic, and thus for each pair $(i, j)$ of actions we define the unique successor of $c$.

1. For $\left(i_{1}^{c}, j_{1}^{c}\right)$ and $\left(i_{2}^{c}, j_{2}^{c}\right)$ the successor is $\bar{b}$.
2. For $\left(i_{1}^{c}, j_{2}^{c}\right)$ the successor is $b$.
3. For $\left(i_{2}^{c}, j_{1}^{c}\right)$ the successor is $s_{k}$.

The reward of the transitions from $c$ is 0 . Intuitively, the transitions and rewards from $\bar{c}$ will be defined from skew-symmetry. Formally, we have:

1. For $\left(i_{1}^{\bar{c}}, j_{1}^{\bar{c}}\right)$ and $\left(i_{2}^{\bar{c}}, j_{2}^{\bar{c}}\right)$ the successor is $b$.
2. For $\left(i_{2}^{\bar{c}}, j_{1}^{\bar{c}}\right)$ the successor is $\bar{b}$.
3. For $\left(i_{1}^{\bar{c}}, j_{2}^{\bar{c}}\right)$ the successor is $\bar{s}_{k}$.

The reward of the transitions from $\bar{c}$ is 1 . There is a transition from $a$ to each other state. The probability to go to $c$ and the probability to go to $\bar{c}$ are both $\frac{1}{4}$. For each other state $s^{\prime}$ (other than $c, \bar{c}$ and $a$ ), the probability to go to $s^{\prime}$ from $a$ is $\frac{1}{4 \cdot k+4}$. The transitions from $a$ have reward $\frac{1}{2}$. There is an illustration of $G_{\eta}^{k}$ in Figure 1.

Lemma 8. For any given $k$ and $\eta$, such that $0<\eta<\frac{1}{4 \cdot k+4}$, the $C M P G G_{\eta}^{k}$ is both skew-symmetric and ergodic. Thus $G_{\eta}^{k}$ has value $\frac{1}{2}$.

Proof. We first argue about ergodicity: from any starting state $s$, the state $a$ is reached almost-surely; and from $a$ there is a transition to all other states with positive probability. This ensures that $G_{\eta}^{k}$ is ergodic.

The following mappings implies that CMPG $G_{\eta}^{k}$ is skew-symmetric: (i) $f\left(s_{i}\right)=\bar{s}_{i}$ for all $i$; and (ii) $f(a)=a$; and (iii) $f(b)=\bar{b}$; and (iv) $f(c)=\bar{c}$. The bijective map $f_{1}^{c}$ between $\Gamma_{1}(c)$ and $\Gamma_{2}(\bar{c})$ is such that $\bar{i}_{1}^{c}=j_{1}^{\bar{c}}$ (and thus also $\bar{i}_{2}^{c}=j_{2}^{\bar{c}}$ ). The bijective map $f_{2}^{c}$ is such that $\bar{j}_{1}^{c}=i_{1}^{\bar{c}}$ (and thus also $\left.\bar{j}_{2}^{c}=i_{2}^{\bar{c}}\right)$.

Lemma 9. For any given $k$ and $\eta$, such that $0<\eta<\frac{1}{4 \cdot k+4}$, consider the set $\mathcal{C}_{p}$ of stationary strategies for Player 1 in $G_{\eta}^{k}$, with patience at most $\frac{1}{p}$, where $p=2 \cdot \eta^{k / 2}$. Consider the stationary strategy $\sigma_{1}^{*}$ defined as: $(i) \sigma_{1}^{*}(c)\left(i_{2}^{c}\right)=p\left(\right.$ and $\left.\sigma_{1}^{*}(c)\left(i_{1}^{c}\right)=1-p\right)$; and (ii) $\sigma_{1}^{*}(\bar{c})\left(i_{2}^{\bar{c}}\right)=1-p\left(\right.$ and $\left.\sigma_{1}^{*}(\bar{c})\left(i_{1}^{\bar{c}}\right)=p\right)$. Then the strategy $\sigma_{1}^{*}$ ensures the maximal value among all strategies in $\mathcal{C}_{p}$.

Proof. First, observe that from $s_{k}$, the probability to reach $a$ in $k$ steps is $\eta^{k}$. If $a$ is not reached in $k$ steps, then in these $k$ steps $s_{k}$ is reached again. Similarly for $\bar{s}_{k}$. Thus, the expected length $L_{s_{k}}$ of a run from $s_{k}$ (or $\bar{s}_{k}$ ) to $a$, is (strictly) more than $\eta^{-k}$, but (strictly) less ${ }^{3}$ than $k \cdot \eta^{-k}$.

The proof will be split in three parts. The first part will consider strategies in $\mathcal{C}_{p}$ that plays $i_{2}^{c}$ with probability greater than $p$; the second part will consider strategies in $\mathcal{C}_{p}$ that plays $i_{2}^{c}$ with probability 0 ; and the third part will show that the optimal distribution for the actions in $\bar{c}$ is to play as $\sigma_{1}^{*}$.

1. Consider some stationary strategy $\sigma_{1}^{\prime} \in \mathcal{C}_{p}$ such that $\sigma_{1}^{\prime}(c)\left(i_{2}^{c}\right)=p^{\prime}>p$. Consider the strategy $\sigma_{1}$ such that $\sigma_{1}(c)=\sigma_{1}^{*}(c)$ and $\sigma_{1}(\bar{c})=\sigma_{1}^{\prime}(\bar{c})$. We will show that $\sigma_{1}$ guarantees a higher expected mean-payoff value for the run between $a$ and $c$ than $\sigma_{1}^{\prime}$, and thus $\sigma_{1}$ will ensure greater mean-payoff value than $\sigma_{1}^{\prime}$.
For $\ell \in\{1,2\}$, let $\sigma_{2}^{\ell}$ be an arbitrary stationary strategy which plays $j_{\ell}^{c}$ with probability 1 . Let $m_{\ell}$ be the mean-payoff of the run from $c$ to $a$, when Player 1 plays $\sigma_{1}^{*}$ and Player 2 plays $\sigma_{2}^{\ell}$. Define $m_{\ell}^{\prime}$ similarly, except that Player 1 plays $\sigma_{1}^{\prime}$ instead of $\sigma_{1}^{*}$. Then, $m_{1}=\frac{1-p}{p \cdot\left(L_{s_{k}}+1\right)+(1-p) \cdot 2}$ and $m_{1}^{\prime}=$ $\frac{1-p^{\prime}}{p^{\prime} \cdot\left(L_{s_{k}}+1\right)+\left(1-p^{\prime}\right) \cdot 2}$ (the expected length of the run is $p^{\prime} \cdot\left(L_{s_{k}}+1\right)+\left(1-p^{\prime}\right) \cdot 2$ and it gets reward 1 only once and only with probability $1-p^{\prime}$ ). We will now argue that $m_{1}>m_{1}^{\prime}$. Consider $m_{1}-m_{1}^{\prime}$ :

$$
\begin{aligned}
m_{1}-m_{1}^{\prime} & =\frac{1-p}{p \cdot\left(L_{s_{k}}+1\right)+(1-p) \cdot 2}-\frac{1-p^{\prime}}{p^{\prime} \cdot\left(L_{s_{k}}+1\right)+\left(1-p^{\prime}\right) \cdot 2} \\
& =\frac{(1-p) \cdot\left(p^{\prime} \cdot\left(L_{s_{k}}+1\right)+\left(1-p^{\prime}\right) \cdot 2\right)-\left(1-p^{\prime}\right) \cdot\left(p \cdot\left(L_{s_{k}}+1\right)+(1-p) \cdot 2\right)}{\left(p \cdot\left(L_{s_{k}}+1\right)+(1-p) \cdot 2\right) \cdot\left(p^{\prime} \cdot\left(L_{s_{k}}+1\right)+\left(1-p^{\prime}\right) \cdot 2\right)}
\end{aligned}
$$

Hence, see that the numerator of the above expression is

$$
\begin{gathered}
(1-p) \cdot\left(p^{\prime} \cdot\left(L_{s_{k}}+1\right)+\left(1-p^{\prime}\right) \cdot 2\right)-\left(1-p^{\prime}\right) \cdot\left(p \cdot\left(L_{s_{k}}+1\right)+(1-p) \cdot 2\right) \\
=\left(p^{\prime}-p\right) \cdot\left(L_{s_{k}}+1\right)>0
\end{gathered}
$$

and therefore $m_{1}>m_{1}^{\prime}$.
We will now argue that $m_{1}<m_{2}$ and $m_{1}^{\prime}<m_{2}^{\prime}$ (and thus Player 2 will play $j_{1}^{c}$ in $c$ against both $\sigma_{1}$ (and thus also $\sigma_{1}^{*}$ ) and $\sigma_{1}^{\prime}$ ). We have that $m_{2}=\frac{p}{2}$ and (repeated for convenience)

[^3]$m_{1}=\frac{1-p}{p \cdot\left(L_{s_{k}}+1\right)+(1-p) \cdot 2}<\frac{1}{p \cdot\left(L_{s_{k}}+1\right)}<\frac{1}{p \cdot \eta^{-k}}$. But $p=2 \cdot \eta^{k / 2}$ and therefore $\frac{1}{p \cdot \eta^{-k}} \leq \frac{1}{2 \cdot \eta^{k / 2} \cdot \eta^{-k}}=$ $\frac{1}{2 \cdot \eta^{-k / 2}}<\frac{1}{\eta^{-k / 2}} \leq \frac{p}{2}$. Similar for $m_{1}^{\prime}<m_{2}^{\prime}$, and hence we have the desired result.
2. Consider some stationary strategy $\sigma_{1}^{0} \in \mathcal{C}_{p}$ such that $\sigma_{1}^{0}(c)\left(i_{2}^{c}\right)=0$. Now consider the strategy $\sigma_{1}$ such that $\sigma_{1}(c)=\sigma_{1}^{*}(c)$ and $\sigma_{1}(\bar{c})=\sigma_{1}^{0}(\bar{c})$. Then, the best response $\sigma_{2}^{0}$ for Player 2 against $\sigma_{1}^{0}$ will play $j_{2}^{c}$ with probability 1 . We see that if Player 1 follows $\sigma_{1}^{0}$ and Player 2 follows $\sigma_{2}^{0}$, then the mean-payoff of the run from $c$ to $a$ is 0 . Thus $\sigma_{1}$ will ensure greater mean-payoff value than $\sigma_{1}^{0}$.
3. Similar to the first two parts, it follows that a strategy that plays like $\sigma_{1}^{*}$ in $\bar{c}$ ensures at least the meanpayoff value of any other stationary strategy in $\mathcal{C}_{p}$ for the play between $\bar{c}$ and $a$. (In this case, the best response for Player 2 will play $j_{1}^{\bar{c}}$ with probability 1 and therefore the mean-payoff for the run from $\bar{c}$ to $a$ is $\frac{2-p}{2}$ as the length of the run is 2 ; and with probability $1-p$ both rewards are 1 , otherwise the first reward is 1 and the second reward is 0 ).

It follows from above that $\sigma_{1}^{*}$ ensures the maximal mean-payoff value among all strategies in $\mathcal{C}_{p}$.
Lemma 10. For any given $k$ and $\eta$, such that $0<\eta<\frac{1}{4 \cdot k+4}$, consider the set $\mathcal{C}_{p}$ of stationary strategies for Player 1 in $G_{\eta}^{k}$, with patience at most $\frac{1}{p}$, where $p=2 \cdot \eta^{k / 2}$. For all strategies in $\mathcal{C}_{p}$, the mean-payoff value is at most $\frac{23}{48}$; and hence no strategy in $\mathcal{C}_{p}$ is $\frac{1}{48}$-optimal.

Proof. By Lemma 9 we only need to consider $\sigma_{1}^{*}$ as defined in Lemma 9. Now we calculate the expected mean-payoff value for a run from $a$ to $a$ given $\sigma_{1}^{*}$ and a positional best-response strategy $\sigma_{2}$ for Player 2, (which is then the expected mean-payoff value of the strategies in $G_{\eta}^{k}$ ) as follows:

1. With probability $\frac{1}{2}$ in the first step, the run goes to some state which is neither $c$ nor $\bar{c}$. Since the probability is equally large to go to some state $s$ or to the corresponding skew-symmetric state $\bar{s}$ and no state $s$ can be reached such that $\left|\Gamma_{1}(s)\right|$ or $\left|\Gamma_{2}(s)\right|$ is more than 1 , such runs has mean-payoff $\frac{1}{2}$.
2. Otherwise with probability $\frac{1}{2}$ in the first step we get reward $\frac{1}{2}$ and go to either $c$ or $\bar{c}$ with equal probability (that is: the probability to go to $c$ or $\bar{c}$ is $\frac{1}{4}$ each). As shown in Lemma 9, (i) the length of the run from $\bar{c}$ to $a$ is 2 ; and with probability $1-p$ both rewards are 1 , otherwise the first reward is 1 and the second reward is 0 ; (ii) the expected length of the run from $c$ to $a$ is $p \cdot\left(L_{s_{k}}+1\right)+(1-p) \cdot 2$ and it gets reward 1 only once and only with probability $1-p$ (where $L_{s_{k}}$ is as defined in Lemma 9).

From the above case analysis we conclude that the mean-payoff of the run from $a$ to $a$ is

$$
\begin{aligned}
& \frac{1}{2} \cdot \frac{1}{2}+\frac{1}{4} \cdot\left(\frac{\frac{1}{2}+1+(1-p)}{3}+\frac{\frac{1}{2}+1-p}{1+p \cdot\left(L_{s_{k}}+1\right)+(1-p) \cdot 2}\right) \\
& =\frac{1}{4}+\frac{\frac{1}{2}+1+(1-p)}{12}+\frac{\frac{1}{2}+1-p}{4 \cdot\left(1+p \cdot\left(L_{s_{k}}+1\right)+(1-p) \cdot 2\right)} \\
& <\frac{1}{4}+\frac{\frac{1}{2}+2}{12}+\frac{2}{4 \cdot p \cdot L_{s_{k}}}<\frac{1}{4}+\frac{5}{24}+\frac{1}{2 \cdot 2 \cdot \eta^{k / 2} \cdot \eta^{-k}} \\
& =\frac{1}{4}+\frac{5}{24}+\frac{1}{4 \cdot \eta^{-k / 2}}<\frac{1}{4}+\frac{5}{24}+\frac{1}{48}=\frac{23}{48}
\end{aligned}
$$



Figure 1: For a given $k \in \mathbb{N}$ and $0<\eta<\frac{1}{4 \cdot k+4}$, the skew-symmetric ergodic game $G_{\eta}^{k}$, except that the transitions from $a$ are not drawn. There is a transition from $a$ to each other state. The probability to go to $c$ from $a$ and the probability to go to $\bar{c}$ from $a$ are both $\frac{1}{4}$. For each other state $s^{\prime}$ (other than $c, \bar{c}$ and $a)$, the probability to go to $s^{\prime}$ from $a$ is $\frac{1}{4 \cdot k+4}$. The transition from $a$ has reward $\frac{1}{2}$. Dashed edges have reward 1 and non-dashed edges have reward 0 . Actions are annotated with probabilities if the successor is not deterministic.

In the first inequality we use that $1+p \cdot\left(L_{s_{k}}+1\right)+(1-p) \cdot 2>p \cdot L_{s_{k}}$ and that $p>0$. In the second inequality we use that $p=2 \cdot \eta^{k / 2}$ and that $\eta^{-k}<L_{s_{k}}$. In the third we use that $\eta^{-k / 2}>12$, which comes from $k \geq 2$ and $\eta<\frac{1}{4 \cdot k+4} \leq \frac{1}{12}$. Therefore, we see that there is no $\frac{1}{48}$-optimal strategy with patience at most $\frac{\eta^{-k / 2}}{2}$ in the game $G_{\eta}^{k}$.

Theorem 11 (Strategy complexity). The following assertions hold:

1. (Upper bound). For almost-sure ergodic CMPGs, for all $\epsilon>0$, there exists an $\epsilon$-optimal strategy of patience at most $\left\lceil 4 \cdot \epsilon^{-1} \cdot m \cdot n^{2} \cdot\left(\delta_{\text {min }}\right)^{-r}\right\rceil$.
2. (Lower bound). There exists a family of ergodic CMPGs $G_{n}^{\delta_{\min }}$, for each odd $n \geq 9$ and $0<\delta_{\min }<$ $\frac{1}{2 \cdot n}$ and $n=r+5$, such that any $\frac{1}{48}$-optimal strategy in $G_{n}^{\delta_{\min }}$ has patience at least $\frac{1}{2} \cdot\left(\delta_{\min }\right)^{-r / 4}$.

Proof. The upper bound comes from Corollary 5, since all $q$-rounded strategies have patience at most $q$; and the lower bound follows from Lemma 10 .

### 3.2 Hardness of approximation

We present a polynomial reduction from the value problem for SSGs to the problem of approximation of values for turn-based stochastic ergodic mean-payoff games (TEMPGs).
The reduction. Consider an SSG $G$ with $n$ non-terminal states, and two terminal states ( $\top$ and $\perp$ ). Given a state $s$ in $G$ we construct a TEMPG $G^{\prime}=\operatorname{Red}(G, s)$ that has the same states as $G$ (including the terminal states) and one additional state $s^{\prime}$. For every transition in $G$, there is a corresponding transition in $G^{\prime}$, with
reward 0 . The 1 terminal $\top$ (resp. 0 terminal $\perp$ ) instead of the self-loop, has two outgoing transitions that go to $\top$ (resp. $\perp$ ) with probability $1-\frac{1}{2^{9 n}}$ and to $s^{\prime}$ with probability $\frac{1}{2^{9 n}}$. The reward of the transitions are 1 (resp. 0 ) for $T$ (resp. $\perp$ ). The additional state $s^{\prime}$ goes to $s$ with probability $1-\frac{1}{2^{7 n}}$ and to each other state (including the terminals, but not $s$ and $s^{\prime}$ ) with probability $\frac{1}{(n+1) \cdot 2^{7 n}}$. The rewards of the transitions from $s^{\prime}$ are 0 . We first observe that the game $G^{\prime}$ is ergodic: since the SSG $G$ is stopping, from all states and for all strategies in $G$, the terminal states are reached with probability 1 ; and hence in $G^{\prime}$, from all states and for all strategies, the state $s^{\prime}$ is reached with probability 1 ; and from $s^{\prime}$ there exists a positive transition probability to every state other than $s^{\prime}$. It follows that under all strategy profiles, from all starting states, the state $s^{\prime}$ is visited infinitely often almost-surely, and hence every other state is visited infinitely often almost-surely. Hence $G^{\prime}$ is ergodic. We will now show that the value $v$ of $G^{\prime}$ is "close" to the value $v_{s}$ of $s$ in $G$. We then argue that we can obtain $v_{s}$ from $v$ in polynomial time by rounding.
Lemma 12. Let $G$ be an $\operatorname{SSG}$, and consider a state $s$ in $G$ with value $v_{s}$. The value $v$ of $\operatorname{Red}(G, s)$ is in the interval $\left[v_{s}-2^{-7 n+1} ; v_{s}+2^{-7 n+1}\right]$.
Proof. We show that the value of $G^{\prime}$ is at least $v_{s}-2^{-7 n+1}$; and the other part of the proof is symmetric. Notice that since $G$ is stopping, we reach a terminal in $n$ steps with probability at least $\frac{1}{2^{n}}$, from every starting state. The expected number of steps required to reach the terminal states is at most $n \cdot 2^{n}$ (one can also use a more refined argument similar to [22] to show that the expected number of steps is at most $2^{n+1}$ ). By construction this is also the case in $G^{\prime}$. From a terminal state in $G^{\prime}$ the expected number of steps required to reach $s^{\prime}$ is $2^{9 n}$. Consider an optimal strategy $\sigma_{1}$ in $G$ for Player 1 . Since $G$ and $G^{\prime}$ have the same set of states where Player 1 has a choice (and the same choices in those states), we can also use $\sigma_{1}$ in $G^{\prime}$. Now consider the best response strategy $\sigma_{2}$ against $\sigma_{1}$ for Player 2 in $G^{\prime}$. We will now estimate the value of $G^{\prime}$. The best $\sigma_{2}$ can ensure for Player 2 is the following:

- By the argument above, for the plays from any starting state in $G$, the expected number of steps required to reach a terminal state is (at most) $n \cdot 2^{n}$.
- For a state $t$ different from $s$, the plays from $t$ reach the 0 terminal with probability 1 .
- The plays from $s$ reach the 0 terminal with probability $1-v_{s}$ and the 1 terminal with probability $v_{s}$.

Notice that for plays starting from any state $t \neq s^{\prime}$, the expected number of steps to reach $s^{\prime}$ is at most $n \cdot 2^{n}+2^{9 n}$. Hence the expected number of steps required to reach $s^{\prime}$ again from itself is at most $n \cdot 2^{n}+$ $2^{9 n}+1$. We now argue that the mean-payoff value is at least $v_{s}-2^{-7 n+1}$. With probability $1-\frac{1}{2^{7 n}}$, the successor of $s^{\prime}$ is $s$. From $s$ the play reaches $s^{\prime}$ after being in the 1 terminal for $v_{s} \cdot 2^{9 n}$ steps in expectation. Each reward obtained in the 1 terminal is 1 . All remaining rewards are 0 . Hence, the mean-payoff value is at least

$$
\begin{aligned}
\frac{v_{s} \cdot 2^{9 n} \cdot\left(1-\frac{1}{2^{7 n}}\right)}{n \cdot 2^{n}+2^{9 n}+1} & =\frac{v_{s} \cdot 2^{9 n}}{n \cdot 2^{n}+2^{9 n}+1}-\frac{v_{s} \cdot 2^{9 n} \cdot \frac{1}{2^{7 n}}}{n \cdot 2^{n}+2^{9 n}+1} \\
& \geq \frac{v_{s} \cdot 2^{9 n}}{\left(1+2^{-7 n}\right) 2^{9 n}}-\frac{v_{s} \cdot 2^{9 n} \cdot 2^{-7 n}}{2^{9 n}} \\
& >\left(1-2^{-7 n}\right) \cdot v_{s}-v_{s} \cdot 2^{-7 n} \\
& =v_{s}-v_{s} \cdot 2^{-7 n+1} \\
& \geq v_{s}-2^{-7 n+1} .
\end{aligned}
$$

The first inequality comes from $n \cdot 2^{n}=2^{n+\log n}<2^{2 n}$; the second inequality comes from $1-2^{-14 n}=$ $\left(1-2^{-7 n}\right)\left(1+2^{-7 n}\right)<1 \Rightarrow 1-2^{-7 n}<\frac{1}{1+2^{-7 n}}$; and the last inequality comes from $v_{s} \leq 1$.

Using a similar argument for Player 2, we obtain that the mean-payoff value is at most $v_{s}+2^{-7 n+1}$, by using that the expected path-length from a state $t$ in $G$ to a terminal is at least 0 . Therefore $v$, the value of $G^{\prime}$, is in the interval $\left[v_{s}-2^{-7 n+1} ; v_{s}+2^{-7 n+1}\right]$.

Observe that if the value $v$ of $G^{\prime}$ can be approximated within $2^{-6 n}$, then Lemma 12 implies that the approximation $a$ is in $\left[v_{s}-2^{-7 n+1}-2^{-6 n} ; v_{s}+2^{-7 n+1}+2^{-6 n}\right]$; which shows that $a$ is in $\left[v_{s}-2^{-5 n}\right.$; $v_{s}+$ $\left.2^{-5 n}\right]$. Hence we see that $a-2^{-5 n}$ is in $\left[v_{s}-2^{-4 n} ; v_{s}\right]$. As observed by Ibsen-Jensen and Miltersen [22], if the value of a state of an SSG can be approximated from below within $2^{-4 n}$, then one can use the KwekMehlhorn algorithm [24] to round the approximated value to obtain the correct value, in polynomial time. We therefore get the following lemma.

Lemma 13. The problem of finding the value of a state in an SSG is polynomial time Turing reducible to the problem of approximating the value of a TEMPG (turn-based stochastic ergodic mean-payoff game) within $2^{-6 n}$.

### 3.3 Approximation complexity

In this section we will establish the approximation complexity for almost-sure ergodic CMPGs. We first recall the definition of the decision problem for approximation.
Approximation decision problem. Given an almost-sure ergodic CMPG $G$ (with rational transition probabilities given in binary), an $\epsilon>0$ (in binary), and a rational number $\lambda$ (in binary), the promise problem PromValerg (i) accepts if the value of $G$ is at least $\lambda$, (ii) rejects if the value of $G$ is at most $\lambda-\epsilon$, and (iii) if the value is in the interval $(\lambda-\epsilon ; \lambda)$, then it may both accept or reject.

Theorem 14 (Approximation complexity). For almost-sure ergodic CMPGs, the following assertions hold:

1. (Upper bound). The problem PromValErg is in FNP.
2. (Hardness). The problem of finding the value of a state in an SSG is polynomial time Turing reducible to the problem PRomVALERG, even for the special case of turn-based stochastic ergodic mean-payoff games (TEMPGs).

Proof. We present the proof for both the items.

1. We first present an FNP algorithm for PromValErg as follows: Guess an $\frac{\epsilon}{4}$-optimal, $q^{\prime}$-rounded strategy $\sigma_{1}$ for Player 1, where $q^{\prime}=\lceil q\rceil$ such that $q$ is as in Corollary 5 (also such a strategy exists by Corollary 5). The strategy is then described using at most $O\left(n \cdot m \cdot \log q^{\prime}\right)$ many bits. Since $\epsilon$ and $\delta_{\text {min }}$ is given in binary, $\log q^{\prime}$ uses at most polynomial many bits. Now compute the best response strategy for Player 2. Since $\sigma_{1}$ is a stationary strategy (because it is $q^{\prime}$-rounded), when Player 1 restricted to follow $\sigma_{1}$, the game becomes an MDP for Player 2, and the size of the MDP is also polynomial in the size of $G$ and $\log q^{\prime}$. Hence there exists a positional best response strategy $\sigma_{2}$, which we can find in polynomial time using linear programming [15, 27, 23]. When Player 1 follows $\sigma_{1}$ and Player 2 follows $\sigma_{2}$ some expected mean-payoff val is achieved. Similarly guess an $\frac{\epsilon}{4}$-optimal, $q$-rounded strategy $\sigma_{2}^{\prime}$ for Player 2. Again there exists a positional best response strategy $\sigma_{1}^{\prime}$ for Player 1 which can again be computed in polynomial time. When Player 1 follows $\sigma_{1}^{\prime}$ and Player 2 follows $\sigma_{2}^{\prime}$ some expected mean-payoff $\mathrm{val}^{\prime}$ is achieved. If $\mathrm{val}^{\prime}-\mathrm{val}>\frac{\epsilon}{2}$, then reject, because then not both $\sigma_{1}$ and $\sigma_{2}^{\prime}$
can be $\frac{\epsilon}{4}$ optimal. Clearly the value of $G$ must be in $\left[\mathrm{val}^{\prime} ; \mathrm{val}^{\prime}\right]$. Notice that both $\lambda-\epsilon$ and $\lambda$ cannot be in $\left[\mathrm{val} ; \mathrm{val}^{\prime}\right]$, since $\mathrm{val}^{\prime}-\mathrm{val} \leq \frac{\epsilon}{2}$. Therefore if $\lambda \leq \mathrm{val}^{\prime}$, then accept, otherwise reject. This establishes that PromValErg is in FNP.
2. We will now show that the problem of finding the value of a state in an SSG is polynomial time Turing reducible to the problem PromValErg for TEMPGs. By Lemma 13, we just need to approximate the value $v$ of a TEMPG $G$ within $2^{-6 n}$. For any number $0<a<1$ and integer $b$, let $\operatorname{Proc}^{b}$ be a procedure, that takes $\frac{p}{q}$ as an input and returns if $a \geq \frac{p}{q}$, where $0 \leq p \leq q \leq b$. For any integer $b$, given procedure Proc ${ }^{b}$, the Kwek-Mehlhorn algorithm [24], finds integers $0 \leq p \leq q \leq b$, such that $a-\frac{p}{q}<\frac{1}{b}$ in $O(\log b)$ time and $O(\log b)$ calls to $\mathrm{Proc}^{b}$. We will argue how to use the Kwek-Mehlhorn algorithm [24] to find the value of $G$ within $2^{-6 n}$ using polynomially many calls to PromValErg. Let $b$ be $2^{8 n}$. Let Proc ${ }^{b}$ be PromValErg with $\epsilon=2^{-16 n}$. Notice that the choice of $\epsilon$ ensures that there can be at most one pair $p, q$ such that $\frac{p}{q} \in[v-\epsilon ; v]$, where $0 \leq p \leq q \leq 2^{8 n}$, because all such numbers are at least $2^{-16 n}$ apart. On such an input PromVaLErg will answer arbitrarily, but on all other inputs it will accurately answer if $\frac{p}{q} \geq v$. The Kwek-Mehlhorn algorithm queries a pair of variables only once, and will find a fraction $\frac{p}{q}$ such that $0 \leq p \leq q \leq 2^{8 n}$. But the four best such fractions must be within $2^{-6 n}$ of $v$.

The desired result follows.

## 4 Strategy Iteration Algorithm for Almost-sure Ergodic CMPGs

The classic algorithm for solving ergodic CMPGs was given by Hoffman and Karp [21]. We will present a variant of the algorithm, and show that for every $\epsilon>0$ it runs in exponential time for $\epsilon$ approximation. Also observe that even for the value problem for SSGs the strategy iteration algorithms require exponential time [16, 14], and hence our exponential upper bound is optimal (given our reduction of the value problem of SSGs to the approximation problem for TEMPGs).
The variant of Hoffman-Karp algorithm. For an almost-sure ergodic CMPG $G$, an $\epsilon>0$, and a state $t$, we will present an algorithm to compute a $q$-rounded $\epsilon$-optimal strategy in $O\left(q^{n \cdot m}\right)$ iterations, and each iteration will require $O\left(2^{\operatorname{POLY}(m)} \cdot \operatorname{POLY}\left(n, \log \left(\epsilon^{-1}\right), \log \left(\delta_{\text {min }}^{-1}\right)\right)\right)$ time, where

$$
q=\left\lceil 4 \cdot \epsilon^{-1} \cdot m \cdot n^{2} \cdot\left(\delta_{\min }\right)^{-r}\right\rceil
$$

Note that in all typical cases, $n$ is large and $m$ is constant, and every iteration takes polynomial time if $m$ is constant. The basic informal description of the algorithm is as follows. In every iteration $i$, the algorithm considers a $q$-rounded strategy $\sigma_{1}^{i}$, and then improves the strategy locally as follows: first it computes the potential $v_{s}^{\sigma_{1}^{2}}$ given $\sigma_{1}^{i}$ as in the Hoffman-Karp algorithm, and then for every state $s$, the algorithm locally computes the best $q$-rounded distribution at $s$ to improve the potential. The intuitive description of the potential is as follows: Fix the specific state $t$ as a target state (where the potential must be 0 ); and given a stationary strategy $\sigma$, consider a modified reward function that assigns the original reward minus the value ensured by $\sigma$. Then the potential for every state $s$ other than the specified state $t$ is the expected sum of rewards under the modified reward function for the random walk from $s$ to $t$. The local improvement step is achieved by playing a matrix game with potentials. Our variant differs from the Hoffman-Karp algorithm that while solving the matrix game we restrict Player 1 to only $q$-rounded distributions. The formal description of the algorithm is given in Figure 2, and the formal definition of the expected one-step reward

```
Function VarHoffmanKarp \((G, \epsilon, t)\)
    Let \(q \leftarrow\left\lceil 4 \cdot \epsilon^{-1} \cdot m \cdot n^{2} \cdot\left(\delta_{\text {min }}\right)^{-r}\right\rceil\);
    Let \(\sigma_{1}^{0}\) be a \(q\)-rounded strategy;
    for \(\left(i \in \mathbb{Z}_{+}\right)\)do
        Compute \(g^{i},\left(v_{s}^{i}\right)_{s \in S}\) as the unique solution of
\[
\begin{aligned}
\forall s \in S: g^{i}+v_{s}^{i} & =\min _{a_{2} \in \Gamma_{2}(s)}\left(\operatorname{ExpRew}\left(s, \sigma_{1}^{i-1}(s), a_{2}\right)+\sum_{s^{\prime} \in S} \delta\left(s, \sigma_{1}^{i-1}(s), a_{2}\right)\left(s^{\prime}\right) \cdot v_{s^{\prime}}^{i}\right) \\
v_{t}^{i} & =0 ;
\end{aligned}
\]
for \((s \in S)\) do
Let \(M_{s}\) be the matrix game defined as follows:
\(M_{s}\left[a_{1}, a_{2}\right] \leftarrow \mathrm{R}\left(s, a_{1}, a_{2}\right)+\sum_{s^{\prime} \in S} \delta\left(s, a_{1}, a_{2}\right)\left(s^{\prime}\right) \cdot v_{s^{\prime}}^{i}\), for all \(a_{1} \in \Gamma_{1}(s)\) and \(a_{2} \in \Gamma_{2}(s)\);
if \(\left(\sigma_{1}^{i-1}(s)\right.\) is a best \(q\) rounded distribution for the matrix game \(M_{s}\) ) then
Let \(\sigma_{1}^{i}(s) \leftarrow \sigma_{1}^{i-1}(s)\);
else
Let \(\sigma_{1}^{i}(s)\) be an arbitrary best \(q\)-rounded distribution over \(\Gamma_{1}(s)\) for the matrix game \(M_{s}\);
if \(\left(\sigma_{1}^{i}=\sigma_{1}^{i-1}\right)\) then
return \(\sigma_{1}^{i}\);
```

Figure 2: Algorithm for solving ergodic games
$\operatorname{ExpRew}\left(s, d_{1}, d_{2}\right)$ for distributions $d_{1}$ over $\Gamma_{1}(s)$ and $d_{2}$ over $\Gamma_{2}(s)$ is as follows: $\operatorname{ExpRew}\left(s, d_{1}, d_{2}\right)=$ $\sum_{a_{1} \in \Gamma_{1}(s), a_{2} \in \Gamma_{2}(s)} \mathrm{R}\left(s, a_{1}, a_{2}\right) \cdot d_{1}\left(a_{1}\right) \cdot d_{2}\left(a_{2}\right)$.
Computation of every iteration. The computation of every iteration is as follows. The computation of the unique solution $g^{i}$ and $\left(v_{s}^{i}\right)_{s \in S}$ is obtained in polynomial time using linear programming. The fact that the solution is unique follows from the fact that once a strategy for Player 1 is fixed, we obtain an MDP for Player 2, and then the MDP solution is unique. For a state $s$, let $\mathcal{D}^{q}(s)$ denote the set of all $q$-rounded distributions over $\Gamma_{1}(s)$. A $q$-rounded distribution $d$ is best for the matrix game $M_{s}$ iff $d \in$ $\arg \max _{d_{1} \in \mathcal{D}^{q}(s)} \min _{a_{2} \in \Gamma_{2}(s)} \sum_{a_{1} \in \Gamma_{1}(s)} d_{1}\left(a_{1}\right) \cdot M_{s}\left[a_{1}, a_{2}\right]$. The computation of a best $q$-rounded strategy is achieved as follows: given an $\left(m_{1} \times m_{2}\right)$-matrix game $M$, solve the following integer linear program for $v$ and $\left(x_{i}\right)_{1 \leq i \leq m_{1}}$ :
$\max v$

$$
\begin{array}{ll}
\text { subject to } & v \leq \sum_{i=1}^{m_{1}} M[i, j] \cdot x_{i} ; \quad 1 \leq j \leq m_{2}, \\
& \sum_{i=1}^{m_{1}} x_{i}=1 ; \\
& x_{i} \cdot q \in \mathbb{N} ; \quad 1 \leq i \leq m_{1} \\
& v \cdot q \cdot \ell \in \mathbb{Z} ;
\end{array}
$$

where $\ell$ is the gcd of all the entries of $M$. It was shown by Lenstra [25], that any integer linear programming problem on an integer $\left(m_{1} \times m_{2}\right)$-matrix (that is, with $m_{1}$ variables) can be solved in time
${ }_{2} \operatorname{POLY}\left(m_{1}\right) \cdot \operatorname{POLY}\left(m_{2}, \log a\right)$, where $a$ is an upper bound on the greatest integer in the matrix and associated vectors. Notice that we can simply scale our matrix with $q \cdot \ell$ and obtain our optimization problem in the required form. Since the entries in the original game was defined from a solution to an MDP (which can be represented using polynomially many bits, because the Player-1 strategy is $q$-rounded), we know that only polynomially many bits are needed to represent $M_{s}$ (also after scaling). Thus, such an integer linear programming problem can be solved in time $O\left(2^{\operatorname{POLY}(m)} \cdot \operatorname{POLY}\left(n, \log \left(\epsilon^{-1}\right), \log \left(\delta_{\min }^{-1}\right)\right)\right)$. This gives us the desired time bound for every iteration.
Turn-based game for correctness. For the correctness analysis, we consider a turn-based stochastic version of the game (which is not ergodic), and refer to the turn-based game as $G^{\prime}=\mathrm{TB}(G)$. The game $G^{\prime}=\mathrm{TB}(G)$ will be a bipartite game of exponential size. For a state $s$ in $G$, let $S_{s}^{q}=\left\{\left(s \times d_{1}\right) \mid d_{1} \in \mathcal{D}_{s}^{q}\right\}$. The state space in $G^{\prime}$ is $S^{\prime}=\left(\bigcup_{s \in S} S_{s}^{q}\right) \cup S$. Whenever we mention $S$ in the rest of this paragraph it should be clear from the context if we refer to $S$ as a part of $G$ or $G^{\prime}$. In $G^{\prime}$, Player 1 controls the states in $S$ and Player 2 the ones in $\bigcup_{s \in S} S_{s}^{q}$. From state $s \in S$, for every $d_{1} \in \mathcal{D}_{s}^{q}$, there is a transition from $s$ to $\left(s, d_{1}\right) \in S_{s}^{q}$ with reward 0 ; and from each state $\left(s, d_{1}\right) \in S_{s}^{q}$ there are $\left|\Gamma_{2}(s)\right|$ actions. For an action $a_{2} \in \Gamma_{2}(s)$, the probability distribution over the next state is given by $\delta\left(s, d_{1}, a_{2}\right)$, and the reward is given by $\operatorname{ExpRew}\left(s, d_{1}, a_{2}\right)$. Given a $q$-rounded strategy $\sigma_{1}$ for Player 1 in $G$ and a positional strategy $\sigma_{2}$, if we interpret the strategies in $G^{\prime}$, then the mean-payoff value in $G^{\prime}$ is exactly half of the mean-payoff value in $G$.
Correctness analysis and bound on iterations. We now present the correctness analysis, and the bound on the number of iterations will follow. The classic strategy iteration algorithm computes the same series of strategies for Player 1 on $\mathrm{TB}(G)$ as our modified Hoffman-Karp algorithm does on the original game ${ }^{4}$. This is because, if we consider a fixed strategy for Player 1 in $\mathrm{TB}(G)$ and the corresponding strategy in $G$, then the best response positional strategy for Player 2 in $\mathrm{TB}(G)$ and $G$ resp. must correspond to each other. Then, by the way the potentials are calculated by the two algorithms, we get the same potential for a given state $s \in S$ for Player 1 in $\mathrm{TB}(G)$ as we do for the corresponding state in $G$ (they are precisely the same, since the value in $\operatorname{TB}(G)$, for any given strategy profile, is half the value of $G$, and thus, when we have taken two steps in $\operatorname{TB}(G)$ we have subtracted precisely the value of $G$ ). For $d_{1} \in \mathcal{D}_{s}^{q}$, the potential of state $\left(s, d_{1}\right)$ in $\mathrm{TB}(G)$ is the same as the value ensured for Player 1 in $M_{s}$, if Player 1 plays $d_{1}$. Thus, also the next strategy for Player 1 is the same. Thus, since the turn-based algorithm correctly finds the optimal strategy for Player 1, our modified Hoffman-Karp algorithm also correctly finds the $q$-rounded strategy that guarantees the highest value in $G$ for all states, among all $q$-rounded strategies. Since the best $q$-rounded strategy in $G$ is $\epsilon$-optimal for $G$ (by Corollary 5), we have thus found an $\epsilon$-optimal strategy. It is well known that the classic strategy iteration algorithm only considers each strategy for Player 1 once (because the potential of the strategies picked by Player 1 are monotonically increasing in every iteration of the loop). Therefore our VarHoffmanKarp algorithm requires at most $q^{m \cdot n}$ iterations, since there are most $q^{m \cdot n}$ strategies that are $q$-rounded.
Inefficiency in reduction to $\mathrm{TB}(G)$. Observe that we only use $\mathrm{TB}(G)$ for the correctness analysis, and do not explicitly construct $\mathrm{TB}(G)$ in our algorithm. Constructing $\mathrm{TB}(G)$ and then solving $\mathrm{TB}(G)$ using strategy iteration could also be used to compute $\epsilon$-optimal $q$-rounded strategies. However, as compared to our algorithm there are two drawbacks in constructing $\operatorname{TB}(G)$ explicitly. First, then every iteration would take time polynomial in $q$ (which is exponential in $n$ ), whereas every iteration of our algorithm requires

[^4]only polynomial time in $n$ and $\log q$. Second, our algorithm only requires polynomial space, whereas the construction of $\mathrm{TB}(G)$ would require space polynomial in $q$ (which is exponential in the input size).

Theorem 15. For an almost-sure ergodic CMPG, for all $\epsilon>0$, VarHoffmanKarp correctly computes an $\epsilon$-optimal strategy, and (i) requires at most $O\left(\left(\epsilon^{-1} \cdot m \cdot n^{2} \cdot\left(\delta_{\min }\right)^{-r}\right)^{n \cdot m}\right)$ iterations, and each iteration requires at most $O\left(2^{\operatorname{POLY}(m)} \cdot \operatorname{POLY}\left(n, \log \left(\epsilon^{-1}\right), \log \left(\delta_{\text {min }}^{-1}\right)\right)\right)$ time; and (ii) requires polynomial space.

## 5 Exact Value Problem for Almost-sure Ergodic Games

We present two results related to the exact value problem: (1) First we show that for almost-sure ergodic CMPGs the exact value can be expressed in the existential theory of the reals; and (2) we establish that the value problem for sure ergodic CMPGs is square-root sum hard.

### 5.1 Value problem in existential theory of the reals

We show how to express the value problem for almost-sure ergodic CMPGs in the existential theory of the reals (with addition and multiplication) in three steps (for details about the existential theory of the reals see $[6,3]$ ).
Step 1: Ergodic decomposition computation. First we compute the ergodic decomposition of an almost-sure ergodic CMPG in polynomial time, and let $C_{1}, C_{2}, \ldots, C_{\ell}$, be the $\ell$ ergodic components. The polynomial time algorithm is as follows: construct a graph with state space $S$, and put an edge ( $s, t$ ) iff $t$ is reachable from $s$ in the CMPG. The bottom scc's of the graph are the ergodic components, where a bottom scc is an scc with no out-going edges leaving the scc.
Step 2: Existential theory of the reals sentence for an ergodic component. For an ergodic CMPG $G$, Hoffman-Karp [21] shows that the value is the unique fixpoint of the strategy iteration algorithm. The algorithm iteratively takes a strategy $\sigma_{1}$ for Player 1 , computes the optimal best response strategy $\sigma_{2}$ for Player 2, and computes the potentials of each state $v_{s}^{\sigma_{1}}$ and the value $g^{\sigma_{1}}$ guaranteed by $\sigma_{1}$. A strategy for Player 1 that ensures a higher value than $g^{\sigma_{1}}$ is then, for every state $u$, to use an optimal distribution in the matrix game defined by $M\left[a_{1}, a_{2}\right]=\mathrm{R}\left(u, a_{1}, a_{2}\right)+\sum_{s \in S} \delta\left(u, a_{1}, a_{2}\right)(s) \cdot v_{s}^{\sigma_{1}}$. We will quantify over stationary strategies in the existential theory of the reals, and use the following notation: for a set $\left\{x_{1}, x_{2}, \ldots, x_{k}\right\}$ of variables we write $\operatorname{ProbDist}\left(x_{1}, x_{2}, \ldots, x_{k}\right)$ to denote the constraints (i) $x_{i} \geq 0$ for $1 \leq i \leq k$, and (ii) $\sum_{i=1}^{k} x_{i}=1$; which specifies that the set of variables forms a probability distribution. We can formulate the fixpoint of the Hoffman-Karp algorithm (and thus the value $g$ ) using existential first order theory as follows. Fix a specific state $s^{*}$, and then consider the following sentence where we quantify
existentially over the variables $g,\left(x_{s, i}\right)_{s \in S, i \in \Gamma_{1}(s)},\left(y_{s, j}\right)_{s \in S, j \in \Gamma_{2}(s)},\left(v_{s}\right)_{s \in S}$, have the following constraints:

$$
\begin{align*}
& \Phi\left(g,\left(x_{s, i}\right)_{s \in S, i \in \Gamma_{1}(s)},\left(y_{s, j}\right)_{s \in S, j \in \Gamma_{2}(s)},\left(v_{s}\right)_{s \in S}\right)=  \tag{1}\\
& \bigwedge_{s \in S} \bigwedge_{j \in \Gamma_{2}(s)}\left(g+v_{s} \leq \sum_{i \in \Gamma_{1}(s)}\left(x_{s, i} \cdot\left(\mathrm{R}(s, i, j)+\sum_{t \in S}\left(\delta(s, i, j)(t) \cdot v_{t}\right)\right)\right) \wedge\right.  \tag{2}\\
& \bigwedge_{s \in S} \bigwedge_{i \in \Gamma_{1}(s)}\left(g+v_{s} \geq \sum_{j \in \Gamma_{2}(s)}\left(y_{s, j} \cdot\left(\mathrm{R}(s, i, j)+\sum_{t \in S}\left(\delta(s, i, j)(t) \cdot v_{t}\right)\right)\right) \wedge\right.  \tag{3}\\
& \bigwedge_{s \in S} \operatorname{ProbDist}\left(x_{s, 1}, x_{s, 2} \ldots, x_{s,\left|\Gamma_{1}(s)\right|}\right) \wedge \bigwedge_{s \in S} \operatorname{ProbDist}\left(y_{s, 1}, y_{s, 2} \ldots, y_{s,\left|\Gamma_{2}(s)\right|}\right) \wedge  \tag{4}\\
& \left(v_{s^{*}}=0\right) . \tag{5}
\end{align*}
$$

Notice that (2) and the fact that the variables $x_{s, i}$ gives a probability distribution, ensures that $x_{s, i}$ gives an optimal strategy in the matrix game of potentials, similar for (3) and $y_{s, j}$. Also, (2) and (3) implies that

$$
\bigwedge_{s \in S}\left(g+v_{s}=\sum_{j \in \Gamma_{2}(s)} \sum_{i \in \Gamma_{1}(s)}\left(y_{s, j} \cdot x_{s, i} \cdot\left(\mathrm{R}(s, i, j)+\sum_{t \in S}\left(\delta(s, i, j)(t) \cdot v_{t}\right)\right)\right)\right.
$$

which together with (2) ensures that

$$
\forall s: g+v_{s}=\max _{j \in \Gamma_{2}(s)} \sum_{i \in \Gamma_{1}(s)}\left(x_{s, i} \cdot\left(\mathrm{R}(s, i, j)+\sum_{t \in S}\left(\delta(s, i, j)(t) \cdot v_{t}\right)\right)\right) .
$$

The preceding equality together with $\left(v_{s^{*}}=0\right)$ ensures that $\left(v_{s}\right)_{s \in S}$ is the potential associated with the stationary strategy $x$, and hence, $g$ is the value of the game. The sentence $\Phi$ in the existential theory of the reals for the value is

$$
\exists g,\left(x_{s, i}\right)_{s \in S, i \in \Gamma_{1}(s)},\left(y_{s, j}\right)_{s \in S, j \in \Gamma_{2}(s)},\left(v_{s}\right)_{s \in S}: \Phi\left(g,\left(x_{s, i}\right)_{s \in S, i \in \Gamma_{1}(s)},\left(y_{s, j}\right)_{s \in S, j \in \Gamma_{2}(s)},\left(v_{s}\right)_{s \in S}\right) ;
$$

and $g$ denotes the value of the component.
Step 3: Existential theory of the reals sentence for an almost-sure ergodic CMPG. Given a real number $\lambda$ and an almost-sure ergodic CMPG $G$, we will now give an existential theory of the reals sentence, which can be satisfied iff $G$ has value at most $\lambda$. Let $C_{1}, C_{2}, \ldots, C_{\ell}$ be the ergodic components, and let $C=\bigcup_{i=1}^{\ell} C_{i}$. We denote by $\Phi_{C_{i}}$ the existential theory of the reals sentence for the value in component $C_{i}$ (as described in Step 2) and the variable $g_{i}$ is the value. The existential theory sentence for other states is given using the formula for reachability games. We quantify existential over the variables $\left(\left(z_{s}\right)_{s \in S},\left(x_{i, s}\right)_{s \in(S \backslash C), i \in \Gamma_{1}(s)},\left(y_{s, j}\right)_{s \in(S \backslash C), j \in \Gamma_{2}(s)}\right)$ and have the following constraints:
$\wedge_{1 \leq i \leq \ell} \quad \Phi_{C_{i}} \wedge$
$\bigwedge_{s \in(S \backslash C)} \quad \bigwedge_{j \in \Gamma_{2}(s)} \quad\left(z_{s} \leq \sum_{i \in \Gamma_{1}(s)} \sum_{t \in S} x_{s, i} \cdot \delta(s, i, j)(t) \cdot z_{t}\right) \wedge$
$\bigwedge_{s \in(S \backslash C)} \quad \bigwedge_{i \in \Gamma_{1}(s)} \quad\left(z_{s} \geq \sum_{j \in \Gamma_{2}(s)} \sum_{t \in S} y_{s, j} \cdot \delta(s, i, j)(t) \cdot z_{t}\right) \wedge$
$\bigwedge_{1 \leq i \leq \ell} \quad \bigwedge_{s \in C_{i}} \quad\left(z_{s}=g_{i}\right) \wedge$

$$
\begin{aligned}
\bigwedge_{s \in(S \backslash C)} & \operatorname{ProbDist}\left(x_{s, 1}, x_{s, 2} \ldots, x_{s,\left|\Gamma_{1}(s)\right|}\right) \wedge \bigwedge_{s \in S} \operatorname{ProbDist}\left(y_{s, 1}, y_{s, 2} \ldots, y_{s,\left|\Gamma_{2}(s)\right|}\right) \wedge \\
& \left(z_{s} \leq \lambda\right) .
\end{aligned}
$$

The idea is as follows: First note that the constraint $z_{s}=g_{i}$, for $s \in C$, ensures that for all states in the ergodic component the variable $z_{s}$ denotes the value of $s$ (by the correctness of the formula $\Phi_{C_{i}}$ for an ergodic component $C_{i}$ ). If the value of state $s \in(S \backslash C)$ in $G$ is $z_{s}$, for all $s$, then

$$
\bigwedge_{s \in(S \backslash C)} \bigwedge_{j \in \Gamma_{2}(s)}\left(z_{s} \leq \sum_{i \in \Gamma_{1}(s)} \sum_{t \in S} x_{s, i} \cdot \delta(s, i, j)(t) \cdot z_{t}\right)
$$

ensures that $x$ is an optimal strategy in the game. Also, similar to the ergodic part,

$$
\bigwedge_{s \in(S \backslash C)} \bigwedge_{j \in \Gamma_{2}(s)}\left(z_{s} \leq \sum_{i \in \Gamma_{1}(s)} \sum_{t \in S} x_{s, i} \cdot \delta(s, i, j)(t) \cdot z_{t}\right) ; \bigwedge_{s \in(S \backslash C)} \bigwedge_{i \in \Gamma_{1}(s)}\left(z_{s} \geq \sum_{j \in \Gamma_{2}(s)} \sum_{t \in S} y_{s, j} \cdot \delta(s, i, j)(t) \cdot z_{t}\right)
$$

implies that for all $s$ :

$$
\left(z_{s}=\max _{j \in \Gamma_{2}(s)} \sum_{i \in \Gamma_{1}(s)} \sum_{t \in S} x_{s, i} \cdot \delta(s, i, j)(t) \cdot z_{t}\right) .
$$

Therefore, the vector $\bar{z}$, such that $\bar{z}_{s}=z_{s}$ is a fixpoint for the value iteration algorithm. Hence, the fact that $z_{s} \leq \lambda$, implies that the least fixpoint $\tilde{z}$ of the value iteration algorithm (which is the value of the game) is such that $\tilde{z}_{s} \leq \lambda$. Thus, we get the following theorem.

Theorem 16. The value problem for almost-sure ergodic CMPGs can be expressed in the existential theory of the reals.

### 5.2 Square-root sum hardness

In this section we show that the value problem for sure ergodic CMPGs is at least as hard as the square-root sum problem.
Square-root sum problem. The square-root sum problem is the following decision problem: Given a positive integer $v$ and a set of positive integers $\left\{n_{1}, \ldots, n_{\ell}\right\}$, is $\sum_{i=1}^{\ell} \sqrt{n_{i}} \geq v$ ? The problem is known to be in the the fourth level of the counting hierarchy [2], but it is a long-standing open problem if it is in NP.
Reduction to sure ergodic CMPGs. The reduction is similar to [10, 11]. First we will define a family of ergodic CMPGs $\left\{G_{b} \mid b \in \mathbb{N}\right\}$, such that $G_{b}$ has value $\sqrt{b}$. Given an instance of the square-root sum problem, $\left(v,\left\{n_{1}, \ldots, n_{\ell}\right\}\right)$, we use our family to get an ergodic CMPG $G_{n_{i}}$ for each number $n_{i}$. We use one more state $s^{*}$, with one action for each player. The successor of $s^{*}$ is $G_{n_{i}}$ with probability $\frac{1}{\ell}$ for every $i$. This will ensure that the value of $s^{*}$ is $\frac{\sum_{i} \sqrt{n_{i}}}{\ell}$. Thus, the value of $s^{*}$ is at least $\frac{v}{\ell}$ iff $\sum_{i} \sqrt{n_{i}} \geq v$. Notice that we reach an ergodic component in precisely one step from $s^{*}$, and thus the game is sure ergodic.
The numbers $k_{b}$ and $d_{b}$. First we will define $G_{b}$, for $b \notin\{1,2,4\}$. We will define $G_{b}$ for $b \in\{1,2,4\}$ afterwards. To define $G_{b}$ for $b \notin\{1,2,4\}$, we will use two numbers $k_{b}$ and $d_{b}$, such that $k_{b}>d_{b}>0$, defined as follows: Let $k_{b}$ be the smallest positive integer such that $k_{b}^{2}>b$. Let $d_{b}=2 \cdot k_{b}-\frac{2 \cdot b}{k_{b}}$, implying that $b=k_{b}^{2}-\frac{d_{b} \cdot k_{b}}{2}$. This gives us directly that $d_{b}>0$ (and hence also $\frac{d_{b} \cdot k_{b}}{2} \in \mathbb{N}$ ). We will show that $k_{b}>d_{b}$. First, for $b=3$, we see that $k_{3}$ is 3 and $3=2^{2}-\frac{1 \cdot 2}{2}$ and thus $d_{3}=1$, implying that $k_{3}>d_{3}$. For $9>b \geq 5$, we see that $k_{b}=3$ and $d_{b} \in\left[\frac{2}{3} ; \frac{8}{3}\right]$ and again have that $k_{b}>d_{b}$. For $b \geq 9$, we will show the statement using contradiction. Assume therefore that $d_{b} \geq k_{b}$. We then get that $b=k_{b}^{2}-\frac{d_{b} \cdot k_{b}}{2} \Rightarrow b \leq \frac{k_{b}^{2}}{2}$. By definition of $k_{b}$ we know that $b \geq\left(k_{b}-1\right)^{2}=k_{b}^{2}+1-2 \cdot k_{b} \geq k_{b}^{2}+1-\frac{k_{b}}{2} \cdot k_{b}>\frac{k_{b}^{2}}{2}$. That is a contradiction. The second to last inequality is because for $b \geq 9$, we have that $k_{b} \geq 4$. Thus, $k_{b}>d_{b}$ for $b \notin\{1,2,4\}$.


Figure 3: The game $G_{b}$, such that $b=k_{b}^{2}-\frac{k_{b} \cdot d_{b}}{2}$. Dashed edges has reward $k_{b}-d_{b}$ and non-dashed edges has reward $k_{b}$. Actions are annotated with probabilities if the probability is not 1 .

Construction of $G_{b}$. For a positive integer $b \notin\{1,2,4\}$, we define $G_{b}$ as follows. There are two states in $G_{b}$, $u$ and $w$. The state $w$ has a single action for Player 1 and a single action for Player $2, a_{w}$ and $b_{w}$ respectively, and the successor of $w$ is always $u$. Also $\mathrm{R}\left(w, a_{w}, b_{w}\right)=k_{b}$. The state $u$ has two actions for each of the two players. Player 1 has actions $a_{u}^{1}$ and $a_{u}^{2}$. Player 2 has actions $b_{u}^{1}$ and $b_{u}^{2}$. For any pair of actions $a_{u}^{i}$ and $b_{u}^{j}$ we have that the successor, $\delta\left(u, a_{u}^{i}, b_{u}^{j}\right)$ is $w$, except for $a_{u}^{1}$ and $b_{u}^{1}$ for which the successor is $u$ with probability $\frac{d_{b}}{k_{b}}$ and $w$ with probability $1-\frac{d_{b}}{k_{b}}$. Note that $\frac{d_{b}}{k_{b}}$ is a number in $(0,1)$, since $k_{b}>d_{b}>0$. The rewards $\mathrm{R}\left(u, a_{u}^{1}, b_{u}^{2}\right)=\mathrm{R}\left(u, a_{u}^{2}, b_{u}^{1}\right)$ are $k_{b}-d_{b}$. The rewards $\mathrm{R}\left(u, a_{u}^{1}, b_{u}^{1}\right)=\mathrm{R}\left(u, a_{u}^{2}, b_{u}^{2}\right)$ are $k_{b}$. The game is ergodic, since $\frac{d_{b}}{k_{b}}<1$, and thus there is a positive probability to change to the other state in every step, no matter the choice of the players. There is an illustration of $G_{b}$ in Figure 3.

Remark 17. For $b \notin\{1,2,4\}$, the numbers $k_{b}$ and $d_{b}$ have short binary descriptions. The number $k_{b}>0$ cannot be larger than $\sqrt{2 \cdot b}$, because otherwise $k_{b}^{2}-\frac{d_{b} \cdot k_{b}}{2} \geq \frac{k_{b}^{2}}{2}>b$. It must also be a positive integer and thus has a binary representation of length at most $\frac{1+\log b}{2}$. Also $k_{b}>d_{b}>0$ and $\frac{d_{b} \cdot k_{b}}{2}$ is a positive integer and thus, $d_{b}$ has a binary representation of length at most $\frac{1+\log b}{2}+\frac{1+\log b}{2}=1+\log b$.
$G_{b}$ for $b \in\{1,2,4\}$. One can, using the preceding, define $G_{b}$ for all positive integers $b$ which is not in $\{1,2,4\}$. It is also easy to construct games, which has value $\sqrt{1}$ and $\sqrt{4}$, since they are integers. Let $G_{1}$ be an arbitrary ergodic CMPG of value 1 and $G_{4}$ be an arbitrary ergodic CMPG of value 2 . One can also construct a ergodic CMPG, which has value $\sqrt{2}$, similar to our construction of $G_{b}$ for $b \notin\{1,2,4\}$, using fractional ${ }^{5} k_{2}$ and $d_{2}$. We see that $k_{2}=\frac{3}{2}$ and $d_{2}=\frac{1}{3}$ gives us that $2=k_{2}^{2}-\frac{d_{2} \cdot k_{2}}{2}$, while ensuring that $k_{2}>d_{2}>0$. Let $G_{2}$ be the game defined analogous to $G_{b}$ for $b \notin\{1,2,4\}$ using $k_{2}=\frac{3}{2}$ and $d_{2}=\frac{1}{3}$.
The value in $G_{b}$ is $\sqrt{b}$. We will now argue that for a fixed $b \notin\{1,4\}$, the game $G_{b}$ has value $\sqrt{b}$ (by definition, the CMPGs $G_{1}$ and $G_{4}$ had value 1 and 2 resp.). We will use that $b=k_{b}^{2}-\frac{d_{b} \cdot k_{b}}{2}$ and that $k_{b}>d_{b}>0$. Let $\sigma_{1}$ be some arbitrary stationary optimal strategy for Player 1 . Let $p$ be the probability that $\sigma_{1}$ plays $a_{u}^{1}$. Let $a$ be the optimal potential of state $u$, then the potential of $w$ is 0 . Let $v$ be the value of $G_{b}$.

[^5]Then as shown by Hoffman-Karp [21] the strategy $\sigma_{1}$ must satisfy the equation system

$$
\begin{aligned}
& a=p \cdot\left(k_{b}-d_{b}\right)+(1-p) \cdot k_{b}-v \\
& a=(1-p) \cdot\left(k_{b}-d_{b}\right)+p \cdot k_{b}+\frac{d_{b} \cdot p \cdot a}{k_{b}}-v \\
& 0=a+k_{b}-v
\end{aligned}
$$

From the third equation we obtain $a=v-k_{b}$, and substituting in the first equation we obtain that

$$
2 \cdot k_{b}=p \cdot d_{b}+2 \cdot v \quad \Rightarrow p=\frac{2 \cdot k_{b}-2 \cdot v}{d_{b}}
$$

Substituting $a$ and $p$ from above into the second equation we obtain

$$
\begin{aligned}
0 & =2 \cdot k_{b}-2 \cdot v-d_{b}+2 \cdot k_{b}-2 \cdot v+\frac{d_{b} \cdot\left(2 \cdot k_{b}-2 \cdot v\right) \cdot\left(v-k_{b}\right)}{k_{b} \cdot d_{b}} \\
\Rightarrow 0 & =2 \cdot k_{b}-d_{b}-\frac{2 \cdot v^{2}}{k_{b}} \\
\Rightarrow 0 & \left.=\frac{k_{b}^{2}}{2}-\frac{d_{b} \cdot k_{b}}{4}-\frac{v^{2}}{2} \quad \text { (Multiply by } k_{b} \text { and divide by } 4\right) .
\end{aligned}
$$

Solving the above second degree equation for $v$ we obtain that

$$
v=\frac{-0 \pm \sqrt{-4 \cdot\left(\frac{k_{b}^{2}}{2}-\frac{d_{b} \cdot k_{b}}{4}\right) \cdot \frac{-1}{2}}}{2 \cdot \frac{-1}{2}} \quad \Rightarrow v= \pm \sqrt{b}
$$

Since we know that the value is positive (since all rewards are positive, because $k_{b}>d_{b}>0$ ), we see that $v=\sqrt{b}$. Thus the desired property is established.

Theorem 18. The value problem for sure ergodic CMPGs is square-root sum hard.

## 6 Conclusion

In this work we established the strategy complexity and the approximation complexity for ergodic, sure ergodic, and almost-sure ergodic mean-payoff games. Our results also show that the approximation problem for turn-based stochastic ergodic mean-payoff games is at least as hard as the value problem for SSGs. In contrast, for concurrent deterministic almost-sure ergodic games, the value problem can be solved in polynomial time. In concurrent deterministic games, in every ergodic component all states have an unique successor, and hence an optimal strategy and the value can be computed in polynomial time. In any given concurrent deterministic almost-sure ergodic game, once the values of the ergodic components have been computed, the value iteration algorithm computes the values for the remaining states in $n$ iterations. Moreover, we established that the value problem for sure ergodic games is square-root sum hard. Note that for sure ergodic games with reachability objectives, the values can be computed in polynomial time by value iteration for $n$ iterations. This shows informally that the hardness of sure ergodic games is due to mean-payoff objectives. Since we have shown that values of ergodic games can be irrational, we conjecture that the value problem for ergodic games itself is sqaure-root sum hard, but an explicit reduction will be cumbersome.

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[^1]:    ${ }^{1}$ The square-root sum problem is an important problem from computational geometry, where given a set of natural numbers $n_{1}, n_{2}, \ldots, n_{k}$, the question is whether the sum of the square roots exceed an integer $b$. The square root sum problem is not known to be in NP.

[^2]:    ${ }^{2}$ For the special case of matrix games (that is; the case where $n=1$ ), this definition of skew-symmetry exactly corresponds to the notion of skew-symmetry for such.

[^3]:    ${ }^{3}$ It is also less than $2 \cdot \eta^{-k}+k$, since for any state $s_{i}$, for $i \geq 1$, there is a probability of more than $\frac{1}{2}$ to go to $s_{k}$ and whenever the play is in $s_{k}$ there is a probability of $\eta^{k}$ that it is the last time.

[^4]:    ${ }^{4}$ The proof that the strategy iteration algorithm works for turn-based mean-payoff games seems to be folk-lore, and also see [28] for the related class of discounted games. Moreover, though $\mathrm{TB}(G)$ is not almost-sure ergodic, if we consider an ergodic component $C$ in $G$, and consider the corresponding set of states in $\mathrm{TB}(G)$, then from all states in $C$ every other state in $C$ is visited infinitely often with probability 1 in $\operatorname{TB}(G)$. Thus for the concrete game $\mathrm{TB}(G)$, the proof can also be done similarly to the proof by Hoffman and Karp [21] for ergodic games, by picking $t$ as a state in $C$ in $\operatorname{TB}(G)$.

[^5]:    ${ }^{5}$ We do not use fractional $k_{b}$ in general only because it becomes harder to argue that the games has a polynomial length binary representation.

