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EVALUATING LINEAR AND NONLINEAR MODELS FOR THE RESPIRATION RATE OF FOUR BREEDS OF HEAT STRESSED FEEDLOT HEIFERS

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ABSTRACT

Heat stress is a factor that causes loss of production and even death in cattle. Animals differ in vulnerability to heat stress. One reason for the difference may be the coat color associated with different breeds or genotypes. A good measure of the heat stress is respiration rate which increases in response to increasing ambient temperature. The objective of this study is to characterize the respiration rates of four genotypes of heat stressed feedlot heifers. Linear and nonlinear models will be compared to find an appropriate method of detecting differences among genotypes.

1. INTRODUCTION

Heat stress in cattle has a negative impact on animal performance and well-being. Frequently, losses in production are reported due to decreases in feed intake, growth and efficiency (Hahn, 1999). In extreme cases, heat stress can even result in death of vulnerable animals (Hahn and Mader, 1997). Coat color is one characteristic of vulnerable animals (Busby and Loy, 1996; Hungerford et al., 2000; Mader et al., 2001). The coat color associated with different breeds or genotypes may be one of the factors that causes differences in response to heat stress. Animals with dark hides are thought to be more vulnerable than light-hided animals due to the adsorption of solar radiation. In a study by da Silva et al. (2003), adsorption of solar radiation from a blackhided animal was 93%, while a light-hided animal was only 27%. Another characteristic of vulnerability to heat stress is the availability of shade. Busby and Loy (1996) reported results of a producer's survey taken to identify factors contributed by a 1995 heat wave in western Iowa. They found that producers with non-shaded lots had the highest death loss in dark-hided animals.

Respiration rate (RR) has been shown to be a good indicator of thermal stress (Brown-Brandl et al., 2002; Gaughan et al., 2000; Hahn et al., 1997). Several authors (Eigenberg et al., 2002; Hahn et al., 1997) have shown respiration rate increases in a non-linear fashion in response to increasing ambient temperature (Ta). One advantage of using respiration rate as an indication of stress is that it is readily observable in a production setting (Hahn et al., 1997).

In this study, we will attempt to characterize the non-linear relationship between RR and Ta by a quadratic regression and an exponential regression model. The quadratic regression is an empirical model that is linear in its parameters. We will examine the fixed effects and mixed model versions of this linear model. Similarly, we will examine fixed and mixed model versions of the nonlinear exponential model. We will also recommend a model that can address the question of whether or not dark-hided heifers become more heat stressed than light-hided ones; that is, does RR increase more rapidly for dark-hided heifers than for light-colored heifers during periods of heat stress? To strengthen our recommendation, after fitting the fixed and mixed

model versions of the quadratic and exponential, we will repeat the process incorporating the experiment treatment factor of genotype and shade.

2. MATERIALS AND METHODS

2.a Experimental Design

A study of four genotypes of heifers, Angus (A), Charolais (C), Gelbvieh(G), and MARC III crossbred (M), was conducted over two consecutive summers (2002 - 2003) in the USDA-ARS-MARC feedlot. The Angus heifers were all solid black. The Charolais heifers were solid white. The MARC III crossbred heifers were mostly dark red (three of 64 were black; some were solid, while others had white tailheads and/or white faces). The Gelbvieh heifers were solid tan. The

feedlot was a block of four adjacent pens. Shade was available in half of each pen. Each year heifers were assigned to one section of the eight pen-sections by genotype (16 heifers/pen-section). These arrangements resulted in a single level hierarchical design. Sixteen heifers of the same genotype are nested

Pen Environment	1	2	3	4
Shade	Α	Μ	G	С
No Shade	М	С	А	G

in one of the 8 pen-sections. Hence, the experiment treatment factor (shade-genotype) is unreplicated and is confounded with pen-section. Only data from the first summer (2002) was used in this study.

Measurements of RR were made twice daily (0800 and 1430) on ten heifers per genotype during six 5-day periods between June 24 and August 9 in 2002. On scheduled experimental days, two observers worked independently. Each randomly selected five animals per pen and recorded RR, by visual observation of flank movement, timing 10 breaths with a stopwatch. The average Ta was calculated based on weather data recorded prior to and immediately after animal observations.

2.b Statistical Models

When choosing a regression model to characterize how the RR response (Y) varies with Ta (X), we have the option of choosing a polynomial which is linear in its parameters. The quadratic polynomial we choose is an empirical model, devoid of theoretical assumptions, based only on the data. The physical meaning of the parameters is irrelevant. We also choose a mechanistic model which is based on theoretical considerations. These considerations endow the parameters with a natural physical interpretation. The nonlinear exponential regression is derived from a differential equation which is based on the assumption that the rate of change in RR is proportional to the current RR, i.e., $\partial Y/\partial X = kappa^*Y$. For the resulting exponential regression, $y = alpha * e^{-kappa^*X}$, the parameter alpha is the respiration rate when Ta is 0°C (32°F), and kappa is the rate constant in units of reciprocal Ta. Other issues to consider when using nonlinear models are the intrinsic and parameter-effects nonlinearity (Bates and Watts, 1988). Preliminary analyses to check for nonlinear behavior showed the percent excess variance for all parameters was greater than 1%. Thus, we re-parameterized the model using *log (alpha) = beta*. The exponential model $y = e^{beta-kappa^*x}$ passed the checks for parameter-effects nonlinearity. In fact, this parameterization is described by Ratkowsky (1990) as having close-to-linear behavior and is called the exponential offset regression, with beta being referred to as the offset. Note:

Another parameterization with poorer close-to-linear behavior but with an interesting interpretation is obtained if we let beta=tau/kappa. In this case, tau is the lag in Ta, that is, the setback in Ta to reach 0C.

Since animal studies are notable for substantial variability in animals, we group the observations by heifer and employ mixed model methods to incorporate both fixed-effects terms and random-effects terms. The fixed-effects terms describe the behavior of the entire population of heifers in a pen-section. The random-effects terms describe the behavior of the individual heifers in the pen-section. Thus, we use the linear mixed models to fit the quadratic regression to the data and the nonlinear mixed model to fit the exponential regression. The analyses were preformed in R, a free software environment for statistical computing and graphics. The linear and nonlinear mixed models were fit using the lme and nlme functions, respectively, in the nlme library in version 2-1.1 of R. Additional powerful diagnostic graphics come from the trellis library.

2.b.i Linear Mixed Model: Quadratic Regression

The following linear mixed effects covariance model was considered.

$$\mathbf{Y}_{i} = \mathbf{betaI} + \mathbf{betaL}^{*} \mathbf{X}_{i} + \mathbf{betaQ}^{*} \mathbf{X}^{2} + \varepsilon_{i} \quad i = 1, ..., 8$$
where $\varepsilon_{i} \sim N(0, \mathbf{I\sigma}^{2})$,
$$\mathbf{betaI} = \mathbf{B}_{10} + \mathbf{B}_{11} \cdot \mathbf{C}_{1} + \mathbf{B}_{12} \cdot \mathbf{C}_{2} + \cdots + \mathbf{B}_{17} \cdot \mathbf{C}_{7} + \mathbf{b}_{10}$$

$$betaI = \beta_{10} + \beta_{11} \cdot C_1 + \beta_{12} \cdot C_2 + \cdots + \beta_{17} \cdot C_7 + b_{i,0}$$

$$betaL = \beta_{20} + \beta_{21} \cdot C_1 + \beta_{22} \cdot C_2 + \cdots + \beta_{27} \cdot C_7 + b_{i,Ta}$$

$$betaQ = \beta_{30} + \beta_{31} \cdot C_1 + \beta_{32} \cdot C_2 + \cdots + \beta_{37} \cdot C_7 + b_{i,Ta}^2$$

$$\boldsymbol{b}_{i} = \begin{pmatrix} \boldsymbol{b}_{i,0} \\ \boldsymbol{b}_{i,Ta} \\ \boldsymbol{b}_{i,Ta}^{2} \end{pmatrix} \sim N(0, \Psi), \text{ where } \Psi = \begin{pmatrix} \boldsymbol{\sigma}^{2}_{0} & \boldsymbol{\sigma}_{0 Ta} & \boldsymbol{\sigma}_{0 Ta}^{2} \\ \boldsymbol{\sigma}_{Ta 0} & \boldsymbol{\sigma}^{2}_{Ta} & \boldsymbol{\sigma}_{Ta Ta}^{2} \\ \boldsymbol{\sigma}_{Ta}^{2}_{0} & \boldsymbol{\sigma}_{Ta}^{2}_{Ta} & \boldsymbol{\sigma}^{2}_{Ta}^{2} \end{pmatrix}$$

The C_1, C_2, \ldots, C_7 are dummy variables defined to incorporate the shade-genotype effects. Both the fixed and random effects were considered for *betaI*, *betaL* and *betaQ*, the regression coefficients for a second-order polynomial regression of RR, the response variable (y), on Ta, the covariate (x). The meaning of the fixed-effect coefficients is as follows: β_{10}, β_{20} and β_{30} are the means of the first shade-genotype treatment for *betaI*, *betaL* and *betaQ* respectively; β_{11}, β_{21} and β_{31} are the differences of the means between the second and first shade-genotype treatment for *betaI*, *betaL* and *betaQ* respectively; and similarly up to the eighth shade-genotype treatment level. The random effects $b_{i,0}$ represent the deviation from the population mean associated with the ith pen and *betaI*, similarly for the linear, and quadratic regression coefficients. The b_i 's and ε_i are assumed independent of each other.

2.b.ii Nonlinear Mixed Model: Exponential Regression

When respiration was assumed to increase at an exponential rate, the following nonlinear mix model was considered.

 $\mathbf{Y}_{i} = \boldsymbol{beta} * \mathbf{e}^{-kappa*X}_{i} + \varepsilon_{i} \qquad i = 1, ..., 8 \qquad \text{Eq 2.}$ where $\varepsilon_{i} \sim N(0, \mathbf{I\sigma}^{2})$, $\boldsymbol{beta} = \beta_{10} + \beta_{11} \cdot C_{1} + \beta_{12} \cdot C_{2} + \dots + \beta_{17} \cdot C_{7} + \boldsymbol{b}_{i,beta}$ $\boldsymbol{kappa} = \beta_{20} + \beta_{21} \cdot C_{1} + \beta_{22} \cdot C_{2} + \dots + \beta_{27} \cdot C_{7} + \boldsymbol{b}_{i,kappa}$ $\boldsymbol{b}_{i} = \begin{pmatrix} \boldsymbol{b}_{i,beta} \\ \boldsymbol{b}_{i,beta} \end{pmatrix} \sim N(0, \Psi), \qquad \text{where } \Psi = \begin{pmatrix} \boldsymbol{\sigma}_{beta}^{2} & \boldsymbol{\sigma}_{beta | kappa} \\ \boldsymbol{\sigma}_{kappa | beta} & \boldsymbol{\sigma}_{kappa}^{2} \end{pmatrix}$

The $C_1, C_2, ..., C_7$ are dummy variables defined to incorporate the shade-genotype effects. Both the fixed and random effects were considered for *beta* and *kappa*. The meaning of the fixedeffect coefficients is as follows: β_{10} and β_{20} are the means of the first shade-genotype treatment for beta and kappa respectively; β_{11} , and β_{21} are the differences of the means between the second and first shade-genotype treatment for beta and kappa respectively; and similarly up to the eight shade-genotype treatment levels. The random effects $b_{i, beta}$ represents the deviation from the population mean associated with the ith pen and beta. The b_i 's and ε are assumed independent of each other.

3. RESULTS AND DISCUSSION

3. a. Relationship between Respiration Rate and Dry-bulb Temperature

A popular approach to modeling the relationship between RR and Ta is to fit a linear fixedeffects model to the entire data set (Figure 1 and Table 1). If the primary focus is on the effects of genotype, it is tempting to fit a separate linear regression for each genotype, Figure 2. A slightly more sophisticated approach is to perform an analysis of covariance which incorporates the genotype treatment effects (ignoring availability of shade) with the covariate Ta. The inclusion of genotype effect (model not shown) significantly reduces the residual standard deviation of the linear fixed-effect model (p < 0.0001) from 26.02 to 24.26. None of these approaches take into account the curvature in the data.

In an attempt to capture the curvature, two regression fixed-effects models, quadratic and exponential, were fit to the data. The coefficients and residual standard error are given in Table 1. Both models have statistically significant parameter estimates (alpha=0.05). Thus, both models support the assumption of curvature.

3.b Linear Model: Quadratic Regression

3.b.i Linear Fixed Model: Quadratic Regression

When we compare the quadratic and linear regression fixed models, discussed above, we see the residual standard error is reduced as we add the quadratic term. The F-ratio for reduction in sum of squares is significant, p=0.0002961, indicating the quadratic effect makes a contribution to

fitting this data. When the experiment shade-genotype treatment factor is included in the model, all the regression coefficients are significant (not shown). Once again, we compare the quadratic and linear regression fixed model, this time including the treatment factor. For the fixed case including treatments, the quadratic is superior to the linear (model not shown). The reduction sum of squares F-ratio is significant (p=0.001091) and residual standard error is reduced from 24.46 to 23.76 for the linear and quadratic, respectively.

Recall the primary motivation for this study is to see if there are differences in how genotypes react to heat challenges (increases in Ta). We begin by including the shade and genotype treatment factor in the fixed-effects model. Results for the quadratic linear fixed-effects model are listed in Table 2 Model 6. We compare these results to those for model 1 (quadratic fixed effect model without treatment, discussed above) using the SS-reduction F-test. We find inclusion of the treatment effects significantly reduces the residual error (p=0.0006).

3.b.ii Linear Mixed Model: Quadratic - Ignoring Experiment Treatment Factors

For the mixed models, we consider four random-effects structures: intercept, linear Ta, quadratic Ta, and the diagonal (intercept, Ta, Ta²) which assumes the random-effects are independent. The residual standard error, information criteria and log-likelihood are provided in Table 2-Models 2-5, respectively. To establish the importance of including random effects in the model, the fixed-effects for mixed model 2 (quadratic regression with random intercept), are reported in Table 1 and compared to those for model 1 (quadratic regression: fixed effect model without treatment). The standard errors are smaller for all estimates. The likelihood ratio constructed from Table 2 (Models 1&2), is significant (p<0.0001). Thus, the inclusion of random-effects significantly improves the fit.

To identify which random effects to incorporate in the model and what their covariance structure should be, we use the results from 127 heifer individual regressions. We examine the 95% confidence intervals for each heifer, Figure 3 Quadratic, to see how the parameter estimates vary among heifers. While it is clear that random effects are needed to account for heifer to heifer variability, the parameter estimates appear to be highly correlated. The scatter plot of the correlation among the random effects, Figure 4, suggests the need for a covariance structure.

Plots of the individual and mixed model (3) fixed-effect estimates are displayed in Figure 5. The mixed model estimates are frequently called "shrinkage estimates" (Pinheiro and Bates, 2000) because they are "pulled toward" the individual estimates. They represent a compromise between the individual fits and the fixed-effects of the mixed model which are associated with the population average. Again we see substantial variation across heifers.

The random-effects of the four mixed models (2-5) are reported in Table 3. There appears to be little variability associated with Ta², while there is a strong relationship between the intercept and Ta. In fact, comparing these models using the information criteria and likelihood ratio tests we find model 3 (linear random-effects structure) provides the best fit. We assess the quality of this fit by examining plots of the predicted values from the linear mixed model with the predicted values from the individual heifer regressions. A subset of these plots for 24 heifers is given in

Figure 6 Quadratic. Model 3 does appear to account for variability among heifers; but, it fails to capture the full extend of the curvature in the data.

3.b.iii Linear Mixed Model: Quadratic Including Experiment Treatment Factors

For the mixed model with treatment factor, the four random-effects structures described above are fit to the data. The information criteria and log likelihood for these models are listed in Table 2 (Models 7-10). The random effects estimates are given in Table 3. Comparing random-effects estimates to those from models without the treatment factor (models 2-5), we see the quadratic effect, Ta^2 , becomes even smaller and the correlation between intercept and Ta is reduced, although it is still ample. Comparing the models with the treatment factor (models 7-10) using the information criteria and likelihood ratio tests, we find model 8 (linear random effects) provides the best fit. The random-effects structure did not change when the treatment factor was included.

Examination of the random-effects produced by model 8, Figure 7 Quadratic, shade is an important factor. Heifers without access to shade have a different pattern then those that do have shade available.

Next we examine the within-group (i.e. heifer) residuals. From the box-plots in Figure 8, we see the standardized residuals are centered at zero but there is large within-group variance, that is, variability changes with animal. A few of the intervals do not cover zero and several have an outlying observation. Treatment factor plots of the residuals versus fitted (or predicted) values, Figure 9 Quadratic, show only Charolais with/without shade and Gelbvieh-without shade are free of outliers. Variability in RR differs by both genotype and shade. The assumption of normality for the within-group residuals can be assessed with normal probability plots, Figure 10. The tails are heavier than expected under normality. But, the picture changes when viewed by treatment factors, Figure 11 Quadratic. Although there are outliers, the deviations from normality do not appear so worrisome.

3.c.i Nonlinear Fixed Model: Exponential Regression

The fixed nonlinear exponential regression is another model that captures the curvature in the data, Table1. The information criteria and log-likelihood are reported in Table 2 (Model 11). When the data are grouped by heifer, 127 individual regressions can be fit. The estimated within-group standard error 25.98 is slightly larger than the residual standard error in the individual regressions, 22.47. Plots of the individual heifer confidence intervals for beta and kappa, Figure 3 Exponential, indicate substantial variation among heifers; and, even though the parameter estimates are highly correlated (0.98), we elect to keep both parameters in the model and proceed to fit the nonlinear mixed-effects version of the exponential model.

3.c.ii Nonlinear Mixed Model: Exponential Ignoring Experiment Treatment Factors

We consider three mixed models; one with a random effect for beta, another with a random effect for kappa and another with random effects for both parameters. All three models indicate a very high correlation (~ 0.92) between fixed effects, beta and kappa. The residual standard error,

information criteria and log likelihood are provided in Table 2 (Models 12 - 14), respectively. The mixed models offer significant improvement over the fixed model 11; the p-values for the likelihood ratio test were all less than 0.0001. The estimates of the random effects for the three models are given in Table 4. The low correlation (0.368) between random effects, beta and kappa suggests we may not need both parameters in the random-effects structure to fit this data. If we restrict our attention to a single random effect, beta is by far the largest. The likelihood ratio test indicates that the model with two parameters (model 14) improves the fit compared to the model with one random-effect if that effect is kappa (p<0.000); but it offers no significant improvement over using beta alone (p=0.1624). Using the principle of parsimony, we conclude model 12, the nonlinear mixed exponential regression with random effect beta is the best choice to fit the data ignoring the experiment treatment factor.

An assessment of the quality of the fitted model is provided by the plot of the augmented predictions, Figure 6 Exponential. The fitted values, representing the average of the heifer population, track the curvature over the observed Ta range. Also, the individual heifer predictions are in good agreement with the observed variability in RR.

3.c.iii Nonlinear Mixed Model: Exponential Including Experiment Treatment Factors

Using the same three mixed models with the treatment factor further reduces the residual standard error, Table 2-Models 15 - 17. The question is which random effects are important. A plot of the random effects for both parameters by shade and genotype, Figure 7 Exponential, indicates the random effects parameters are correlated. The random effect estimates are given in Table 4. The standard deviation for kappa-alone (model 16) is reduced to 0.00323 and the correlation between beta and kappa (model 17) is increased to 0.884. The reduction in kappa implies part of the heifer-to-heifer variation may be explained by the treatment factor. The increase in the correlation between beta and kappa suggests the random effects may be overparameterized. To test these implications, we check the information criteria and likelihood ratios. The likelihood test shows two parameters (model 17) are better than kappa alone (model 16), p<0.0001, but, that may be because beta is one of the random effects. The two parameter structure (model 17) offers no significant improvement to using beta alone (model 15), p=0.2042. Once again, the principle of parsimony is invoked. We conclude that model 15 offers a good fit to the data. The same random effects structure is chosen whether we include or ignore the treatment factors.

Model diagnostics were performed by examining the plot of the standardized residual versus the fitted values by shade and genotype, Figure 9 Exponential. The residuals are distributed symmetrically around zero, with roughly uniform variance. Only Charolais-with/without shade and Gelbvieh-without shade are free of outliers. The normal plot of the within-group residuals by shade and genotype, Figure 11 Exponential, supports the assumption of normality.

3.d Comparing Quadratic and Exponential Regressions

Practical differences favor the quadratic. It is easier to fit and understand. The exponential requires good starting values for convergence and that can be a difficult task; but, the payoff is the parameters have a natural physical interpretation; beta is the offset and kappa is the rate

constant. In other respects, the regressions are similar. Both the quadratic and exponential use data grouped by heifer to describe RR as a function of Ta taking into account the correlation among observations for a heifer. Both the quadratic and exponential provide for curvature in the relationship between RR and Ta. For this data, both needed significant random-effects structures to represent within-heifer dependence and both affirmed the validity of assumptions on the random-effects and residual error. The main difference is in the quality of fit as assessed by Figure 6 Quadratic and Exponential. The nonlinear exponential model appears to be far superior to the quadratic regression.

4. CONCLUSIONS

Descriptions of the relationship between RR and Ta improve when curvature is built-into the model. Since curvature is a prominent feature of both the quadratic and exponential regressions, they showed improvement when compared to the linear regression.

Variability in RR differs by both genotype and shade. Each regression shows considerable improvement when the shade-genotype treatment factor is incorporated into the model, attesting to the presence of shade-genotype effects. Future studies conducted using a 2-way factorial treatment design with replication would allow us to examine the interaction of shade and genotype.

Both regressions also show improvement when random effects are used to account for the random variation in heifers attesting to the usefulness of mixed models. The plots of the augmented predictions provide a way to assess the quality of fit and a way to compare mixed models. For this data, the exponential regression is the model of choice. It is clearly superior to the quadratic regression both in terms of quality of fit and insightfulness of interpretation.

5. SUMMARY

In an attempt to characterize heat stress in four breeds of heifers, respiration rate, RR, was measured twice daily and recorded along with the computed ambient temperature. The data was gathered from a single level hierarchical design conducted in the USDA-ARS-MARC feedlot. Each of four pens was sectioned into shade and no shade areas. Sixteen heifers from each of four genotypes, two dark-hided (Angus and MARC-III) and two light-colored (Charolais and Gelbvieh) heifers were assigned to one of the eight pen-sections. The shade-genotype treatment factor was un-replicated and confounded with the pen-section. A linear mixed model analysis was performed for a quadratic regression; while, a nonlinear mixed model analysis was performed for the two-parameter exponential regression, RR=exp (beta-kappa*Ta). Both regressions verified the assumption of curvature in the RR-Ta relationship. Both mixed model regressions showed improvement over their fixed model counterpart and they both improved when the experimental shade-genotype treatment factor was included in the analysis. The best fit for the quadratic came from a mixed model with a linear Ta random-effects structure. The best fit for the exponential came from a mixed model with singe random-effect on beta. Plots of the augmented predictions show the exponential regression is far superior. The parameters of that model have a natural interpretation useful for comparing levels of the experiment treatment factor with beta being the offset constant and kappa, the RR rate constant. Thus, when

investigating shade and genotype effects under the conditions of the motivating experiment, we recommend using the nonlinear exponential mixed model with random-effect beta, to analyze the data.

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LINEAR REGRESSION MODELS							NONLINEAR REGRESSION MODELS		
Linear Fixed		Quadratic Fixed		Quadratic Mixed Random Intercept		neter	Expor Fix	nential ked	
Param	Estimates	Std Error	Estimates	Std Error	Estimates	Std Error	Param	Estimates	Std Error
Intercept	-39.348*	2.9865	11.577	14.363	14.238	12.961	Beta	3.069*	0.0362
Temp	4.712*	0.1053	0.924	1.050	0.759	0.946	Kappa	-0.509 *	0.0010
Temp ²			0.068*	0.019	0.071*	0.017			
RSE	26.02		25.95		22.755		RSE	25	.98

Table 1. Fixed Estimates and Residual Standard Error for Models Ignoring Experiment Treatment Factors

* Estimate is significant, α =0.05.

Table 2. Residual Standard Error, Information Criteria and Log Likelihood for Linear and Nonlinear Models

Model		Model	Name	Random-effects	RSE	Information Criteria		Log-
		#				AIC	BIC	Likelihood
dels	dratic hout riment ctor	1	Fixed lm	None	25.95	20948.31	20971.16	-10470.16
		2	Mixed lme	Intercept	22.755	20580.85	20609.41	-10285.42
		3	Mixed Ime	Linear Temp	22.357	20527.49	20567.47	-10256.74
	Quac Wit Fac Fac	4	Mixed lme	QuadraticTemp	22.364	20535.46	20592.59	-10257.73
	EX 4 O	5	Mixed lme	Diag	22.391	20529.73	20569.72	-10257.87
Ŭ				(Intercept, Temp, Temp ²)				
ear	Quadratic With Experiment Factor	6	Fixed lm-trt	None	23.76	20649.22	20746.23	-10307.61
Line		7	Mixed lme-trt	Intercept	21.689	20427.45	20530.16	-10195.72
		8	Mixed lme-trt	Linear Temp	21.560	20411.51	20525.64	-10185.75
		9	Mixed lme-trt	QuadraticTemp	21.563	20417.63	20548.89	-10185.82
		10	Mixed lme-trt	Diag	21.565	20412.01	20526.14	-10186
				(Intercept, Temp, Temp ²)				
	tial t ent	11	Fixed nl	None	25.98	20942.83	20954.26	-10469.42
s	Exponent Withou Experime Factor	12	Mixed nlme	beta	22.475	20530.42	20553.27	-10261.21
ode		13	Mixed nlme	kappa	22.542	20543.18	20566.03	-10267.59
Nonlinear M		14	Mixed nlme	beta + kappa	22.448	20534.27	20568.55	-10261.13
	t I	15	Mixed	beta	21.667	20350.56	20453.41	-10157.28
	ntia nen ner		nlme-trt					
	Exponer With Experin Facto	16	Mixed nlme-trt	kappa	21.784	20368.92	20471.77	-10166.46
		17	Mixed nlme-trt	beta + kappa	21.6	20350.96	20465.23	-10155.48

	Quadrat	tic Regress Experime	sion Model w ent Factors	ithout	Quadratic Regression Models with Experiment Factors			
	Mixed	Mixed	Mixed	Mixed	Mixed	Mixed	Mixed	Mixed
	lme 2	lme 3	lme 4	lme 5	lme-trt 7	lme-trt 8	lme-trt 9	lme-trt 10
	Intercept	linear	Quadratic	Diag	Intercept	Linear	Quadratic	Diag
Doromotors		Temp	Temp			Temp	Temp	
Tarameters	Std	Std	Std	Std	Std	Std	Std	Std
	Dev	Dev	Dev	Dev	Dev	Dev	Dev	Dev
Intercept	12.21638	10.3375	3.7209	0.7603	9.90583	2.9843	2.50345	0.18186
Temp	-	0.7607	0.4233	0.2486	-	0.4144	0.39958	0.33637
Temp ²	-	-	0.0122	0.0125	-	-	0.00084	0.00397
Intercept & Temp	-	-0.92 ^c	-0.651 ^c	-	-	-0.636 ^c	-0.565 ^c	-
Intercept & Temp ²	-	-	-0.115 ^c	-	-	-	0.007 ^c	-
Temp & Temp ²	-	-	-0.128 ^c	-	-	-	-0.04 ^c	-
RSE	22.755	22.357	22.364	22.391	21.68898	21.5604	21.5625	21.5648

Table 3 Random Effects and Residual Standard Error for Linear Model Quadratic Regression and Nonlinear Models Exponential Regression

^c denotes correlation.

Table 4 Random Effects and Residual Standard Error for Nonlinear Models Exponential Regression

	Nonlinear I Ex	Regression Moc speriment Facto	lels without ors	Nonlinear Regression Models with Experiment Factors			
Parameter	ter Mixed M nlme 12 nlm random ra beta k		Mixed nlme 14 random both	Mixed nlme-trt 15 random beta	Mixed nlme-trt 16 random kappa	Mixed nlme-trt 17 random both	
Beta	0.013369	Std Dev	0.14303	0.0103366	- Sta Dev	0.18466	
Kappa	-	0.00434	0.00225	-	0.00323	0.00346	
Beta & Kappa	pa 0.368 ^c		0.368 ^c	-	-	0.884 ^c	
RSE 22.47477		22.542	22.448	21.66688	21.7837	21.59978	

^c denotes correlation.

Figure 2. Linear Regression Lines for



Figure 1. Plot of Respiration Rate vs. Ta with Linear Regression





Figure 4. Scatter Plot of Correlation between Random Effects from Mixed Model Quadratic (random: linear Ta)



Figure 5. Individual and Mixed Model Fixed-effects:Quadratic(random: linear Ta)



Figure 6 Predictions for population (fixed) and within-group (Individual Heifer) for observed RR (circles) versus Ta





Figure 7. Parameter Estimates of Heifer Random Effects by Shade and Genotype from Mixed Model

Figure 8. Box-plots of Standardized Residuals for Each Heifer Linear Mixed Model Quadratic (Random Linear)





Figure 9. Plot of Standardized Residuals vs. Fitted Values by Shade and Genotype for Mixed Model

Fig 10. Normal Probability Plot of Standardized Residuals from Mixed Model Quadratic





