Is sea-ice-driven Eurasian cooling too weak in models?

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9 In a recent Letter, Mori et al.¹ examined connections in observations and climate models 10 between reduced Arctic sea ice and the 'warm Arctic and cold Eurasia' (WACE) pattern. They concluded that models systemically underestimate Eurasian cooling in response to sea-ice 11 loss. relative to observations. If correct, their result implies that up to half of the observed 12 13 Eurasian cooling from 1995-2014 is attributable to sea-ice loss¹, whereas previous studies have found a negligible contribution from sea-ice loss²⁻⁴. Here, we highlight that their 14 comparisons between observations and models are not like-for-like and when fair 15 16 comparisons are made, modelled and observed estimates are consistent with each other. The upward adjustment of the contribution of sea-ice loss to observed Eurasian cooling in Mori et 17 18 al.¹ is therefore unjustified. 19 20 An essential first step in model evaluation is to derive an observational benchmark against 21 which models can be assessed. Mori et al.¹ seek an observational estimate of the WACE 22 response to Barents-Kara sea-ice loss. Their approach is to correlate time-series of the WACE 23 and Barents-Kara sea ice. They interpret the squared correlation coefficient multiplied by the

- 24 WACE variance (hereafter $r^2\sigma^2$) as the fraction of WACE variance *driven* by sea ice. In doing
- so, they make an assumption about causality. Their apparent justification for this causal
 inference is that their analysis detects the WACE as the main pattern of sea-ice-driven
- inference is that their analysis detects the WACE as the main pattern of sea-ice-driven
 temperature variability. However, the WACE also varies irrespective of sea ice and as a result,
- their observed time-series of the WACE (EC_{ERA}) contains variability driven by sea ice and
- 29 variability not driven by sea ice. Importantly, even the non-sea-ice-driven component will
- 30 correlate with Barents-Kara sea ice because a warm Arctic causes reduced sea ice, and vice
- 31 versa. Sea ice and WACE variability will correlate because Arctic temperature and sea ice are
- 32 strongly and physically connected, even if Eurasian temperature is minimally affected by sea
- 33 ice, as recent work suggests⁵. For the reasons just given, the interpretation of $r^2\sigma^2$ as the
- fraction of WACE variance *driven* by sea ice is questionable and, as we will demonstrate, it is
 likely an overestimation.
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- 37 Mori et al.¹ also calculate $r^2 \sigma^2$ from atmospheric general circulation model (AGCM)
- 38 experiments in which sea ice is specified. Interaction between ice and atmosphere is one-way
- in this experimental setup: the atmospheric response to varying sea ice is represented but the
- 40 response of sea ice to atmospheric variation is not. Mori et al.¹ report that the observed $r^2\sigma^2$
- is roughly twice as high as those obtained from seven different AGCMs, leading them to
 conclude that models systemically underestimate the WACE response to sea-ice loss.
- 42 conclude that models systemically underestimate the WACE response to sea-ice loss.
 43 However, their comparison between observations and AGCMs is misleading because the
- 45 nowever, their comparison between observations and AGLMS is misleading because the 44 observed $r^2\sigma^2$ reflects two-way interaction between sea ice and atmosphere whereas AGCM-
- 45 derived $r^2\sigma^2$ only reflects the one-way influence of sea ice on the atmosphere.
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- 47 Here, we test how much $r^2 \sigma^2$ is suppressed by a lack of two-way interaction in AGCM
- 48 simulations. We first reproduced the results of Mori et al.¹ (with minor methodological
- 49 differences; see Methods), which confirms that none of the seven AGCMs reproduce the

observed $r^2\sigma^2$ (Fig. 1a). It is important to note that this discrepancy is seen in both the winter 50 51 months collectively (Fig. 1a) and in winter averages (c.f. Fig. 4a in Mori et al.¹), which strongly 52 suggests that the origin of the model-observation disparity is not time-scale dependent, and 53 justifies our use of monthly averages in what follows. Next, we compared output from an 54 atmosphere-ocean general circulation model (AOGCM) and an AGCM prescribed with sea ice 55 and sea surface temperatures taken from the parent AOGCM (see Methods). The $r^2\sigma^2$ in the 56 AOGCM experiment is roughly twice as high as in the AGCM (Fig. 1b), which we attribute to 57 the simulation of two-way atmosphere-ice interaction in the AOGCM but not in the AGCM. The 58 smaller $r^2 \sigma^2$ in AGCMs compared to observations shown originally by Mori et al.¹ and 59 reproduced here in Fig. 1a need not imply model error. Instead, it could be largely explained by the lack of two-way interaction in AGCM experiments and the inability of the $r^2\sigma^2$ method 60 to extract the sea-ice-driven fraction of observed WACE variability. We caution not to directly 61 compare the observed (Fig. 1a) and AOGCM-derived (Fig. 1b) $r^2\sigma^2$ as they correspond to 62 different climate states. Although the observed $r^2\sigma^2$ is higher than in the AOGCM, this likely 63 reflects greater sea-ice variability (by \sim 50%; not shown) in the recent past compared to the 64

65 preindustrial climate, rather than model error.

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67 We note that the higher $r^2\sigma^2$ in the AOGCM compared to the AGCM (Fig. 1b) could arise due 68 to one of two reasons: because coupling is necessary to simulate the effects of WACE

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variability on sea ice or because ocean feedbacks could amplify the WACE response to sea-ice 70 loss. Regarding the former, coupled models simulate the effects of WACE-related atmospheric

71 circulation variability on Barents-Kara sea ice similarly to the real world⁵. Regarding the

72 latter, it is clear that ocean feedbacks would strengthen the local warming response to sea-ice

73 loss. However, this need not imply any change in Eurasian cooling. Indeed, Deser et al.⁶ 74 showed that ocean coupling enhanced the Arctic warming and atmospheric circulation

75 responses to sea-ice loss, but suppressed Eurasian cooling. Surveying the literature, there is

76 no evidence that coupled models simulate stronger Eurasian cooling in response to sea-ice

77 loss than uncoupled models⁶⁻¹⁰.

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79 To gain further insight, and building upon Blackport et al.⁵, we propose a refined approach to 80 estimate the WACE variance driven by sea ice, which can be applied equally to observations and AGCMs. We hypothesise that a better estimate can be obtained from the correlation 81 82 between the WACE and Barents-Kara sea ice in the preceding month, rather than using the 83 contemporaneous correlation. Lead-lag correlation is a common first step in causal discovery 84 and is physically justified here because the upward surface heat flux anomalies that might 85 cause or reinforce the WACE tend to lag reductions in Barents-Kara sea ice^{5,11}. Conversely, WACE-induced (and more generally, circulation-induced) downward heat flux anomalies tend 86 to precede reductions in Barents-Kara sea ice^{5,11}. Blackport et al.⁵ presented evidence that this 87 88 1-month lead or 1-month lag approach can effectively distinguish between regimes of 'ice 89 driving atmosphere' and 'atmosphere driving ice'. To test our hypothesis, we repeated the analysis but calculating $r^2\sigma^2$ with Barents-Kara sea ice one month ahead of the WACE (see 90 Methods). The AOGCM and AGCM gave almost identical estimates of the WACE variance 91 92 driven by sea ice when using the refined method (Fig. 1d), lending support to our hypothesis 93 and suggesting that ocean coupling has little effect on the WACE response to sea-ice loss. 94

95 Applying the refined approach to the observations and AGCMs (Fig. 1c) leads to a rather

different conclusion on model performance to that in Mori et al.¹. Now, the estimates from all 96

97 seven AGCMs lie close to that from observations. Only two AGCMs have ensemble member

98 ranges that do not span the observed estimate, and only by very small margins (likely within

99 observational uncertainty, which has not been accounted for here). We conclude that the

modelled and observed estimates of the WACE variance driven by sea ice are consistent with 100

- 101 each other. This suggests that AGCMs are able to realistically simulate the WACE response to
- sea-ice loss, effectively ruling out either AGCM error or lack of ocean coupling as the main
- 103 reason for the stronger Eurasian cooling trends in observations compared to AGCMs.
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- Lastly, we repeated the analysis but calculating $r^2\sigma^2$ with Barents-Kara sea ice lagging one month behind the WACE. Now, $r^2\sigma^2$ provides an estimate of the WACE variance related to
- 107 atmospheric driving of sea ice. $r^2\sigma^2$ is non-zero in the AGCMs partly because the imprint of
- 108 observed WACE variability is contained in the sea ice conditions specified in the AGCMs. Also,
- 109 due to serial correlation, the $r^2\sigma^2$ at 1-month lead (Fig. 1d,e) and 1-month lag (Fig. 1c,d) do
- 110 not, and should not be expected to, sum to the contemporaneous $r^2\sigma^2$ (Fig. 1a,b).
- 111 Nevertheless, when sea ice lags the WACE, we find a clear discrepancy between observations
- and AGCMs (Fig. 1e) and between the AOGCM and AGCM (Fig. 1f), in stark contrast to the
- 113 consistency found when sea ice leads the WACE (Fig. 1c,d). This provides further evidence
- that the apparent divergence between observations and models reported by Mori et al.¹ stems
- from the inability of AGCM experiments to simulate the effects of the WACE on sea ice, and the failure of the original $r^2\sigma^2$ method to extract the sea-ice-driven fraction of observed WACE
- failure of the original $r^2\sigma^2$ method to extract the sea-ice-driven fraction of observed WACE variability. This reasoning is likely valid across timescales, at least qualitatively, given that
- 117 Variability. This reasoning is fixely valuacross timescales, at least qualitatively, given that 118 WACE-related weather patterns drive warming in the Arctic, thereby reducing sea ice, on sub-
- monthly^{5,11-13}, monthly^{5,14,15}, seasonal^{5,16} and multidecadal^{2-5,7,17} timescales.
- 120
- 121 In summary, here we have shown that the two main conclusions of Mori et al.¹ that models
- 122 systematically underestimate Eurasian cooling in response to Arctic sea-ice loss and that
- 123 ~44% of observed Eurasian cooling is attributable to sea-ice loss were based upon a
- 124 misleading comparison of observations and models. When fair comparisons are made,
- observations and models agree on the fraction of WACE variance driven by sea ice. There is
- 126 therefore, no justification for the adjustment to the model output that leads Mori et al.¹ to
- 127 conclude that 32-51% of observed Eurasian winter cooling from 1995-2014 is attributable to
- sea-ice loss. Without this misleading adjustment, models suggest that sea-ice loss has
- 129 contributed little to colder Eurasian winters, which can instead be largely explained as a
- 130 manifestation of internal climate varaibility²⁻⁵.
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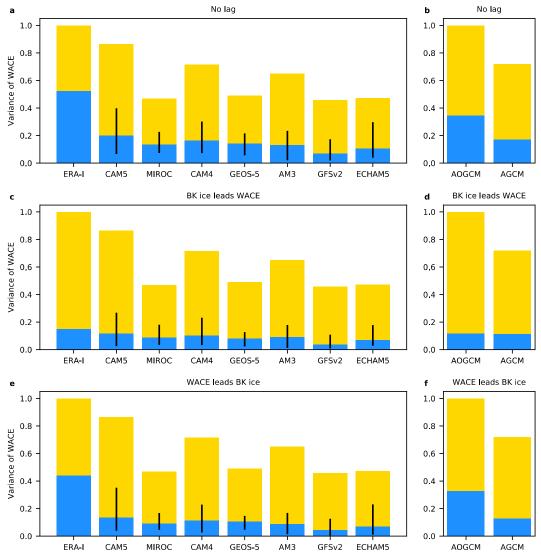
184 Methods

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186 In Fig. 1a,c,e, we use the exact same data as Mori et al.¹. Briefly, observational results come 187 from the ERA-Interim reanalysis for the period 1979-2014 and modelled results come from 188 seven AGCM simulations in which observed sea surface temperatures and sea ice have been 189 specified. Further details on these model simulations can be found in Mori et al.¹. Additionally 190 in Fig. 1b,d,f, we analyse two experiments performed as part of the Community Earth System Model (CESM) Large Ensemble project¹⁸. The first is a 200-year section (years 401-600) of a 191 192 preindustrial control run of the CESM configured with the Community Atmosphere Model 193 version 5 (CAM5). CESM-CAM5 is a global coupled climate model at approximately 1° 194 horizontal resolution in all model components. The second additional experiment is a 200-195 year simulation with CAM5 in which sea surface temperatures and sea ice were specified from 196 years 401-600 of the parent CESM-CAM5 simulation. External forcing is the same in both 197 simulations. Initially in Fig. 1a, we employed the same methodologies as Mori et al.¹ to 198 calculate the WACE pattern, the WACE time-series and the Barents-Kara sea-ice index (see 199 Mori et al.¹ for details), with one exception. We computed the WACE and Barents-Kara sea-ice time-series from monthly averages whereas Mori et al.¹ used winter averages. The purpose of 200

201 this modification is to facilitate subseasonal lead-lag correlations. Projecting near-surface 202 temperatures for December, January and February separately onto the winter-mean WACE 203 pattern produced monthly WACE time-series. Fig. 1a shows the total and sea-ice-driven 204 variance for the three winter months combined. To construct Fig. 1b, we performed an 205 analogous analysis but substituted data from ERA-Interim with that from CESM-CAM5. More 206 specifically, we derived the WACE pattern as the leading mode of covariability from singular 207 value decomposition applied to the CESM-CAM5 and CAM5 simulations. The r^2 was then 208 calculated by correlating the corresponding WACE time-series with the Barents-Kara sea-ice 209 index from CESM-CAM5. In this so-called "perfect model" comparison^{19,20}, the coupled model 210 simulation is an analogue for the observations and therefore, any difference between the 211 CESM-CAM5 and CAM5 results can be attributed solely to ocean coupling. To construct Fig. 1c-212 f, we adapted the Mori et al.¹ approach by introducing a 1-month lead or lag time between the 213 WACE and Barents-Kara sea-ice time-series. In Fig1c,d we show the combination of three 214 cases where the Barents-Kara sea-ice index leads the WACE time-series by one month: 215 November sea ice correlated with December WACE, December sea ice correlated with January 216 WACE, and January sea ice correlated with February WACE. In Fig1c,d we show the 217 combination of three cases where the Barents-Kara sea-ice index lags the WACE time-series 218 by one month: January sea ice correlated with December WACE, February sea ice correlated 219 with January WACE, and March sea ice correlated with February WACE. 220 221 **Acknowledgements.** The authors thank Masato Mori for providing data from the MIROC 222 simulations and for useful discussions. We acknowledge the individuals and modelling groups 223 that contributed to the Facility for Climate Assessments (FACTS) multimodel data set and the 224 **CESM Large Ensemble Project.** 225 226 Data availability. The FACTS and CESM simulations are freely available and were obtained 227 from the following repositories: https://www.esrl.noaa.gov/psd/repository/facts and 228 https://www.cesm.ucar.edu/projects/community-projects/LENS/. 229 230 **Contributions.** J.A.S and R.B. jointly conceived the analysis. R.B. created Fig. 1. J.A.S. wrote the 231 manuscript with input from R.B. 232 233 **Competing interests.** The authors declare no competing interests. 234

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236 237 Figure 1. Observed and modelled estimates of the total and sea-ice-driven WACE 238 variance. a, Total variance of WACE for ERA-Interim and seven AGCM experiments (blue plus 239 yellow bars). The vertical axis in **a** is scaled by the total WACE variance in ERA-Interim. The 240 blue bars show the variance explained by Barents-Kara sea ice estimated from $r^2\sigma^2$. For the 241 AGCMs, the bars show results for all ensemble members (concatenated in series) and the 242 vertical black lines provide the ensemble ranges. **b**, As (**a**) but for a single AGCM (CAM5) and it's parent AOGCM (CESM-CAM5). The vertical axis in **b** is scaled by the total WACE variance in 243 the coupled model. c,d, As (a,b) but calculating $r^2\sigma^2$ with Barents-Kara sea ice one month 244 ahead of the WACE. **e**,**f**, As (**a**,**b**) but calculating $r^2\sigma^2$ with Barents-Kara sea ice one month 245 246 behind the WACE. 247