

W&M ScholarWorks

**VIMS Articles** 

1-2018

# Intensified environmental and density-dependent regulation of white perch recruitment after an ecosystem shift in the Hudson River Estuary

BK Gallagher Virginia Institute of Marine Science

DH Secor

Follow this and additional works at: https://scholarworks.wm.edu/vimsarticles

Part of the Aquaculture and Fisheries Commons

### **Recommended Citation**

Gallagher, BK and Secor, DH, "Intensified environmental and density-dependent regulation of white perch recruitment after an ecosystem shift in the Hudson River Estuary" (2018). *VIMS Articles*. 1293. https://scholarworks.wm.edu/vimsarticles/1293

This Article is brought to you for free and open access by W&M ScholarWorks. It has been accepted for inclusion in VIMS Articles by an authorized administrator of W&M ScholarWorks. For more information, please contact scholarworks@wm.edu.



Canadian Journal of Fisheries and Aquatic Sciences Journal canadien des sciences halieutiques et aquatiques

## Intensified environmental and density-dependent regulation of white perch recruitment after an ecosystem shift in the Hudson River Estuary

Journal:	Canadian Journal of Fisheries and Aquatic Sciences
Manuscript ID	cjfas-2016-0455.R1
Manuscript Type:	Article
Date Submitted by the Author:	10-Feb-2017
Complete List of Authors:	Gallagher, Brian; Virginia Institute of Marine Science, Department of Fisheries Science Secor, David H.; University of Maryland, Chesapeake Biological Laboratory; Chesapeake Biological Laboratory
Keyword:	RECRUITMENT < General, Freshwater flow, Hudson River, White perch, Ecosystem shift

SCHOLARONE<sup>™</sup> Manuscripts

1	Intensified environmental and density-dependent regulation of white
2	perch recruitment after an ecosystem shift in the Hudson River Estuary
3	By:
4	Brian K. Gallagher <sup>1,2</sup> and David H. Secor*
5 6	University of Maryland Center for Environmental Science, Chesapeake Biological Laboratory, P.O. Box 38, Solomons, Maryland, 20688, USA
7	
8	Footnotes:
9	<sup>1</sup> Corresponding author (e-mail: bkgallagher@vims.edu, phone: 804-684-7351)
10	<sup>2</sup> Present address: Virginia Institute of Marine Science, College of William and Mary, P.O. Box
11	1346, Gloucester Point, Virginia, 23062, USA
12	*e-mail: secor@umces.edu
13	
14	
15	
16	
17	
18	
19	
20	
21	
22	
23	

# 24 Abstract

Long-term monitoring data were used to test whether the invasion of zebra mussels in the 25 Hudson River Estuary (HRE) in 1991 altered the influence of density-dependence and 26 27 environmental conditions on life-stage transitions, growth and partial migration in white perch (Morone americana). During the post-invasion period (1992-2013), we estimated standing stocks 28 of white perch eggs, yolk-sac larvae (YSL), post yolk-sac larvae (PYSL), young-of-the-year 29 30 (YOY) and adults, as well as indices of YOY growth and spatial distribution. A series of linear and nonlinear functions were employed to model life-stage transitions, while the effects of six 31 environmental and density-dependent variables on YOY growth and partial migration were 32 33 quantified. Comparisons to pre-invasion observations (1974-1991), indicated that egg-YSL, PYSL-YOY and YOY-yearling transitions changed significantly after the invasion, while PYSL 34 abundance developed a stronger negative effect on YOY growth. The PYSL-YOY transition 35 became more sensitive to density-dependence and freshwater flow from 1992-2013, which is 36 consistent with diminished abundance and increased environmental sensitivity of the forage base 37 38 in the HRE reported after the zebra mussel invasion. 39 40 41

- 42
- 43
- 44

# 45 Introduction

Examining the potential of ecosystem shifts to intensify, dampen or reverse population 46 responses to long-term environmental change is an important and broadly relevant problem in 47 fisheries research. Shifts in aquatic ecosystems are characterized by altered food web structure, 48 species interactions, and population dynamics of individual species within the system (Hare and 49 Mantua 2000; Collie et al. 2004). Threshold changes can drive ecosystems into alternative stable 50 states, which control how populations and communities interact and respond to environmental 51 conditions and external forcing (Strayer et al. 2008; Planque et al. 2010), thereby confounding 52 53 traditional management strategies that depend on stationarity (Folke et al. 2004). Important consequences of abrupt ecosystem shifts and long-term environmental change 54 (i.e. gradual, but persistent trends) include alterations in primary and secondary production, 55 which influence the productivity and spatial distribution of fish populations. In particular, 56 reductions in prey availability have the potential to intensify density-dependence in growth and 57 recruitment. For example, grazing pressure from the invasive overbite clam (Potamocorbula 58 *amurensis*) played an important role in shifting the abundance and composition of the 59 60 zooplankton community in the San Francisco Estuary (Kimmerer et al. 1994), which contributed to food limitation and subsequent density-dependent reductions in carrying capacity for delta 61 smelt (Hypomesus transpacificus), threadfin shad (Dorosoma petenense) and striped bass 62 (Morone saxatilis; Kimmerer et al. 2000; Feyrer et al. 2007). In addition to changing density-63 dependent processes, shifts in population state brought about by ecosystem shifts, climate 64 65 oscillations or fishing pressure can potentially alter how populations respond to long-term temperature increases (Ottersen et al. 2006; Planque et al. 2010). Thus, determining whether 66 ecosystem shifts alter the state of a population and, subsequently, change the responses of one or 67

more life-stages to density-dependence or environmental conditions is a key step in incorporating
species responses to climate change into assessment and management frameworks (Collie et al.
2004).

71 The availability of extensive long-term monitoring data that overlaps with considerable 72 changes in the ecosystem makes the Hudson River Estuary (HRE) a model system to study the 73 ecological consequences of abrupt ecosystem shifts and long-term environmental change (Strayer et al. 2014a). The HRE is a large (243 km in length) estuary that receives most of its 74 75 freshwater input from the Troy Dam at the head of the estuary (Figure 1; Cooper et al. 1988). Through its effects on turbidity and advection, freshwater flow strongly controls primary and 76 77 secondary production in the HRE, with conditions of low flow resulting in increased water 78 clarity and residence time and generally leading to enhanced productivity (Gladden et al. 1988; Howarth et al. 2000). Freshwater flow and temperature have both increased in the HRE since 79 80 1950 (Seekell and Pace 2011; Strayer et al. 2014a) and are projected to continue increasing over the next century (Najjar et al. 2009). In addition to these long-term changes, the zebra mussel 81 (Dreissena polymorpha) invaded the tidal freshwater portion of the HRE in 1991, became highly 82 83 abundant in 1992, then triggered a rapid and extensive ecosystem shift (see reviews by Strayer et al. 2008; 2014a). Grazing pressure increased exponentially, greatly reducing the biomass of 84 phytoplankton, pelagic zooplankton and benthic invertebrates (Caraco et al. 1997; Pace et al. 85 86 1998; Strayer and Smith 2001). As a consequence of their role in reducing the forage base, zebra mussels had demonstrable impacts on many fishes, with pelagic species generally displaying 87 88 reduced abundance and growth rates after the invasion, as well as downriver shifts in their spatial 89 distribution (for details, see Strayer et al. 2004). However, increased water clarity after the invasion enhanced primary and secondary production in littoral habitats, resulting in increased 90

91	abundance, higher growth rates and upriver distributional shifts in many littoral fishes (e.g.
92	centrarchids; Strayer et al. 2004). Importantly, the zebra mussel invasion also made the
93	abundance of many organisms more sensitive to freshwater flow. For example, the abundance of
94	littoral benthic invertebrates and littoral fishes has developed a stronger inverse relationship with
95	freshwater flow than was observed before the invasion (Strayer et al. 2008). More recently, the
96	abundance and size of zebra mussels in the HRE have decreased (Strayer and Malcom 2006),
97	allowing some recovery of zooplankton, benthic invertebrate and fish stocks (Pace et al. 2010;
98	Strayer et al. 2011; Strayer et al. 2014b).

White perch (Morone americana) is a dominant estuarine fish species in North America 99 100 which completes its entire life-cycle within the estuary, and is therefore likely to be sensitive to 101 ecosystem shifts and long-term environmental change within these systems. Adult white perch undertake spring spawning migrations into tidal freshwater, where eggs develop into pelagic 102 103 larvae and subsequently metamorphose into fully-formed juveniles that select littoral habitats (Mansueti 1964; Klauda et al. 1988). Thus, different life-stages may respond to large-scale 104 ecosystem change in distinct ways. In addition, white perch life history characteristics are 105 106 sensitive to ecosystem attributes (e.g. Tuckett et al. 2013). For example, recruitment in several 107 Chesapeake Bay tributaries is positively associated with freshwater flow (Kraus and Secor 2005), likely due to a strengthening of the estuarine turbidity maximum in high flow years, 108 which enhances larval retention and feeding success (North and Houde 2003). Similarly, first-109 year growth in Oneida Lake white perch is positively influenced by summer water temperatures 110 111 and measures of primary and secondary productivity (VanDeValk et al. 2016). Before the zebra 112 mussel invasion, studies of young-of-the year (YOY) white perch growth and recruitment in the HRE suggested that growth was unrelated to YOY abundance, but positively influenced by water 113

114 temperature (Klauda et al. 1988), while recruitment was unrelated to larval abundance,

temperature or flow (Pace et al. 1993).

Estuarine white perch display partial migration, characterized by a resident contingent 116 that remains in natal freshwater habitats throughout life, and a migratory contingent, which 117 118 exhibits active dispersal from the natal habitat and subsequently uses brackish water (Kraus and Secor 2004; Kerr and Secor 2011). Inter-annual variability in the numerical dominance of each 119 contingent is modulated by freshwater flow in Chesapeake Bay tributaries (Kraus and Secor 120 2004). White perch in the HRE also display partial migration, with a numerically dominant 121 122 freshwater resident contingent (Gallagher 2016). Likely owing to higher productivity in the 123 brackish portion of the HRE (Howarth et al. 2006), migratory white perch grow and mature 124 faster than residents (Gallagher 2016), and therefore may contribute disproportionately to the productivity of the population (Kerr et al. 2010). Although white perch recruitment, growth and 125 126 partial migration dynamics have been well-documented in the HRE and elsewhere, the influence of density-dependence, temperature and freshwater flow on these early-life characteristics may 127 have been altered by the zebra mussel invasion. Such changes, if present, could potentially affect 128 129 future population dynamics, as temperature and precipitation are both projected to increase (by 4°C and 7% on average, respectively) in the HRE over the next century (Najjar et al. 2009). 130

Our objective was to evaluate whether the zebra mussel invasion modified the relative influence of density-dependence and environmental conditions on the recruitment, growth, and migration dynamics of YOY white perch in the HRE. Due to the reduced abundance and increased environmental sensitivity of the forage base (Strayer et al. 2008), we hypothesized that, following the zebra mussel invasion: 1.) recruitment to the YOY stage became more sensitive to density-dependence and environmental conditions 2.) YOY growth developed stronger density137 dependence and 3.) partial migration became more strongly density-dependent, such that high YOY abundance produces a larger fraction of YOY in the migratory contingent. To test these 138 hypotheses, we applied a series of stage-structured models and exploratory statistical analyses to 139 an extensive set of monitoring data collected after the invasion of zebra mussels (1992-2013). 140 These findings were then compared to previous studies of YOY white perch recruitment, growth 141 142 and spatial distribution before the zebra mussel invasion (1971-1991), to test whether the ecosystem shift brought about by the invasion has altered these important processes through 143 changed functional relationships with abundance, flow, and other environmental variables. 144

# 145 Methods

### 146 Study area

The HRE is a large, partially mixed estuary characterized by unusually low sinuosity and 147 strong tidal influence compared to many estuaries, with a watershed spanning approximately 148 12,900 km<sup>2</sup> (Figure 1; Geyer and Chant 2006). Annual mean temperature and flow in the HRE 149 150 from 1951-2013 have positive, non-linear trends over this time period, with accelerated rates of 151 increase since 1990 (Figure 2). In addition, the position of the salt front in the HRE is strongly influenced by freshwater flow (Cooper et al. 1988) and can vary by 20-60 km on a seasonal basis 152 153 (Gever and Chant 2006). The position of the salt front controls the salinity distribution within the 154 river, which directly influences spatial variation in the abundance and composition of 155 zooplankton, benthic invertebrate and fish communities (Gladden et al. 1988; Strayer and Smith 156 2001; Daniels et al. 2005).

#### 157 **Data availability**

This analysis was shaped by data availability, which differed between pre-invasion
(1974-1991) and post-invasion (1992-2013) periods. Daily temperatures and salt front positions

160 in the HRE are only available during the post-invasion period since 1992 and 1991, respectively. We also lacked white perch growth and spatial distribution data before the zebra mussel invasion 161 (1974-1991) and therefore used relationships reported in previous studies by Klauda et al. (1988) 162 and Lawler, Matusky and Skelly Engineers (1989). Finally, two types of abundance data were 163 used. Standing stock estimates with high spatial and temporal resolution were only available for 164 165 the post-invasion period, and were thus preferred when describing life-stage transitions and spatial distribution indices during this period. A second set of annual abundance indices were 166 less resolved, but spanned the pre-invasion and post-invasion periods (1974-2013), and were 167 168 used to perform statistical comparisons of life-stage transitions before and after the invasion.

### 169 White perch abundance, growth and migration data

170 Standing stock abundance estimates were obtained from annual reports prepared by Applied Science Associates (2014) for white perch eggs, yolk-sac larvae (YSL), post yolk-sac 171 larvae (PYSL), young-of-the-year juveniles (YOY) and adults (age-1 yearlings and older). These 172 data are collected in "Year-class Reports" for the Hudson River Estuary Monitoring Program, an 173 extensive survey funded by four utility companies which has been in place since 1974. Data were 174 drawn from three separate surveys, which targeted different life-stages. Egg, YSL and PYSL 175 176 standing stocks were estimated from the Longitudinal River Survey; YOY standing stock was estimated from the Utilities Beach Seine Survey (seine survey hereafter) and adult white perch 177 standing stock was based on the Fall Juvenile Survey. Early life-stages are differentiated using 178 179 standardized laboratory protocols based upon developmental traits (e.g. yolk-sac absorption). 180 Each of these surveys follows a random sampling design, stratified by depth zones. A summary of the depth strata, time-span, gear and mesh sizes used in each of the three surveys can be found 181 in Table S.1. The spatial and temporal resolution of each survey is relatively high and consistent 182

(~1000 samples year<sup>-1</sup>), with fish sampling occurring in the same 13 river sections, spanning
from Albany to Manhattan (Figure 1), on weekly or bi-weekly intervals depending on the survey
and time of year.

For each life-stage in each survey, mean densities (number  $m^{-3}$ ) within each stratum and 186 river section were calculated each sampling week. These densities were subsequently converted 187 into river-wide standing stock estimates by multiplying the density by the stratum volume in 188 each river section, and summing across strata and river sections (Applied Science Associates 189 2014). Standing stocks were subsequently averaged over a different set of weeks for each life-190 191 stage in order to obtain annual means from 1992-2013. Weeks were selected if the life-stage of 192 interest was consistently present in the river and vulnerable to the sampling gear across the majority of years in our data set. For white perch early life-stages, the time periods selected were 193 weeks 17-25, 18-26 and 19-28 for eggs, YSL and PYSL, respectively, while weeks 28-40 and 194 195 27-41 were used in analyses of YOY and adults, respectively (Table S.1). The same set of weeks was also used to calculate the mean temperature and freshwater flow experienced by each life-196 stage during each year for use in subsequent analyses (see Statistical Analyses). The 197 198 Longitudinal River Survey tows the sampling gear against the prevailing current at 0.9-1.0 m s<sup>-1</sup>, so evasion of white perch YSL and PYSL (maximum swim speed ~ $0.1 \text{ m s}^{-1}$ ) is unlikely. 199 However, due to likely differences in sampling efficiency (e.g. size-selective sampling of larvae) 200 and stage duration between life-stages, standing stocks should be interpreted as indices of 201 abundance, rather than estimates of the absolute abundance of each stage in the HRE. 202 YOY white perch standing stocks were analyzed in relation to the position of the salt 203 204 front to separate resident and migratory contingents during each year. Daily salt front positions in the HRE have been computed by the US Geological Survey (USGS) since 1991 (USGS 205

206	2015a; <u>http://ny.water.usgs.gov/projects/dialer_plots/saltfront.html</u> ). Data were not available for
207	most of 2012, and in this year daily salt front positions were estimated using a multiple
208	regression model from Cooper et al. (1988), with lagged freshwater flow and tidal amplitude data
209	as the independent variables. The predicted daily salt front positions from the regression model
210	were proportional to observed values reported by the USGS from January-February in 2012
211	(slope = 0.97; $R^2 = 0.45$ ; data not shown). Based upon the mean salt front position each week,
212	river sections were designated as freshwater (above) or brackish water (below). Associated
213	freshwater and brackish water standing stocks were then averaged over weeks 28-40. The
214	proportion of YOY that were in brackish water was calculated each year (hereafter migrant
215	fraction) and used as an index of migratory contingent prevalence. In addition, mean YOY total
216	length at the end of the growing season (hereafter YOY length) was calculated each year as the
217	mean river-wide size reported during the last two sampling weeks of the seine survey (first and
218	third weeks of October; weeks 40 and 42), which served as an index of YOY white perch
219	somatic growth.

In addition to standing stock estimates, we also utilized a set of annual abundance indices 220 for white perch eggs, YSL, PYSL, YOY and yearlings reported from 1974-2013 in Year-class 221 Reports (indices were not reported for adults). These indices were calculated from the same 222 surveys as the standing stock estimates and are highly correlated and proportional to the annual 223 mean standing stocks from 1992-2013 ( $R^2 > 0.90$  in all cases; data not shown). However, in 224 order to account for differences in sampling duration between early (1974-1987) and recent 225 (1988-2013) survey time periods, the calculation of each index utilized a different set of 226 sampling weeks than the mean standing stocks (see above), with YOY and yearling indices 227 drawn from weeks 33-40; while egg, YSL and PYSL indices relied on a different set of 7-week 228

periods each year, depending on the week in which the cumulative density reached 5% of the
annual sum of densities over all sampling weeks (Applied Science Associates 2014). Despite
these discrepancies, the strong correlation of each annual abundance index to the annual mean
standing stock and the longer duration of the index data set supported quantitative comparisons
of white perch life-stage transitions before (1974-1991) and after (1992-2013) the zebra mussel
invasion (see Statistical Analyses).

### 235 Environmental and biological variables

236 Daily temperature and freshwater flow records were obtained respectively from a

pumping station in Poughkeepsie at river km 120, (USGS 2015b;

238 <u>http://waterdata.usgs.gov/nwis/uv?site\_no=01372058</u>) and the Green Island monitoring site at

the head of the estuary in Troy, NY (USGS 2015c;

240 <u>http://waterdata.usgs.gov/nwis/uv?site\_no=01358000</u>). In addition to being the richest data set,

freshwater flow at Green Island represents approximately 70% of freshwater input into the HRE

242 (Howarth et al. 2006). Daily temperature and flow values were averaged during spring (April-

June) and summer (July-September) months to calculate seasonal means for each year. In

addition, estimates of annual mean zebra mussel filtration rates during the growing season (May-

September) from 1992-2013 were obtained from the Cary Institute of Ecosystem Studies (D.

246 Strayer, personal communication; see Strayer and Malcom 2006 for details) and used to examine

- 247 possible effects of zebra mussels on white perch early life-stages. Descriptive statistics for all
- white perch early life-history, biological and environmental variables can be found in Table 1.

249

#### 251 Statistical analyses

252 A life-cycle analysis was employed to model the transitions across early life-history stages of HRE white perch. We used the methods developed by Paulik (1973) to decompose the 253 stock-recruitment relationship of HRE white perch (in this study, the relationship between adult 254 and YOY standing stocks) into a series of life-stage transitions leading up to the YOY stage. 255 256 Similar life-cycle analyses have been successfully applied to describe the life-stage transitions in North Sea herring (Nash and Dickey-Collas 2005) and assess changes in these transitions over 257 time (Payne et al. 2009). Transitions were modeled in two possible ways, depending on whether 258 the relationship was linear or non-linear. For linear transitions, the mean standing stock of life-259 260 stage s  $(N_s)$  was modeled as a proportion of the mean standing stock of one or more previous life-stages  $(N_{s-1})$  by the equation: 261

262 (1) 
$$N_s = \alpha(N_{s-1})$$

where  $\alpha$  is a density-independent multiplier. We estimated  $\alpha$  using simple linear regression with the intercept fixed at 0, reflecting traditional assumptions in stock-recruit theory (Quinn and Deriso 1999). Preliminary analyses indicated that environmental variables had negligible impacts on linear transitions (data not shown), and were therefore not included. Similarly, non-linear transitions were modeled using a Ricker stock-recruitment function by the equation:

268 (2) 
$$N_s = \alpha(N_{s-1})e^{-\beta(N_{s-1})}$$

- 269 where  $\alpha$  is once again the density-independent multiplier, and  $\beta$  is the density-dependent
- 270 parameter, equal to the inverse of the value of  $N_{s-1}$  which corresponds to the maximum value of
- 271  $N_s$  (Quinn and Deriso 1999). To test for environmental effects on non-linear transitions,
- additional Ricker models including each combination of temperature, freshwater flow (both from

**.** .

the time-period over which  $N_{s-1}$  was averaged) and annual mean zebra mussel filtration rates as covariates were constructed by the equation:

275 (3) 
$$N_s = \alpha(N_{s-1})e^{-\beta(N_{s-1})-\sum \gamma_E(E)}$$

where  $\Sigma$  is a summation term for all environmental variables (*E*) in a given model, and  $\gamma_E$ represent the coefficients corresponding to each environmental variable. All models (n=7) were then compared and the best model was selected using Akaike's information criterion corrected for small sample sizes (AICc; Burnham and Anderson 2002). Ricker models were linearized to the form  $\log_e(N_s \times N_{s-1}^{-1}) = \log_e(\alpha) - \beta(N_{s-1})$  (equation 2) or  $\log_e(N_s \times N_{s-1}^{-1}) = \log_e(\alpha) - \beta(N_{s-1}) - \Sigma \gamma_E(E)$  (equation 3) and the parameters were estimated using simple linear regression (two parameters) or multiple regression (three or more parameters).

The effects of environmental and density-dependent factors on YOY white perch growth 283 and partial migration from 1992-2013 were explored using a Pearson correlation analysis. The 284 response variables of interest were YOY length and the migrant fraction, while the predictors in 285 both analyses were the mean summer temperature and flow, annual mean zebra mussel filtration 286 287 rates (environmental factors), as well as the mean standing stocks of white perch YOY and PYSL (density-dependent factors). In addition, because the availability of brackish habitat in the 288 HRE may influence white perch growth and migration, the mean proportion of the total HRE 289 shoal volume in the brackish portion of the estuary during the summer was calculated and used 290 291 as a predictor (hereafter termed brackish shoal proportion). This proportion was calculated by 292 assigning the 13 river sections (Figure 1) into freshwater and brackish regions based on the mean summer salt front position, summing the shoal volumes of all brackish sections (depth < 6 m) 293 and dividing by the total shoal volume of the HRE ( $\sim$ 345 million m<sup>3</sup>). For comparison, 294

correlations reported in previous studies of YOY white perch growth and spatial distribution in the 1970s and 80s (Klauda et al. 1988; Lawler, Matusky and Skelly Engineers 1989) were compiled and compared to estimated correlations from 1992-2013. These previous studies selected a Type-I error rate of 0.05 and did not adjust p-values of their correlations for multiple comparisons. However, for the 12 correlations analyzed in the more recent period, we applied a Bonferroni correction  $(0.05 \times 12^{-1} = 4.17 \times 10^{-3})$  to adjust the Type-I error rate.

Indices of white perch egg, YSL, PYSL, YOY and yearling abundance were compared 301 between pre-invasion (1974-1991) and post-invasion (1992-2013) periods. Two-sample t-tests 302 assuming unequal variance were performed to test for differences in the mean index of 303 304 abundance between these two time periods. Subsequently, life-stage transitions during each time 305 period were modeled using the functional relationships described previously (equations 1, 2 and 3). Differences in parameter estimates  $(\alpha, \beta, \gamma)$  between pre-invasion and post-invasion periods 306 307 were analyzed using analysis of covariance (ANCOVA) and post-hoc contrasts, with invasion period (pre- or post-invasion) used as a categorical variable. The relationships between the egg 308 index and the YOY index, as well as the yearling index and the YOY index from the previous 309 year were also modeled for both periods. 310

# 311 **Results**

#### 312 Life cycle analysis

Based on standing stock data, the stock-recruitment (adult-YOY) relationship for white perch in the HRE from 1992-2013 (Figure 3) was described by a Ricker model with freshwater flow during the PYSL period as a covariate ( $\log_e(\alpha) = 0.91$ ;  $\beta = 5.61 \times 10^{-4}$ ;  $\gamma_{flow} = 2.18 \times 10^{-3}$ ; Table 2). Paulik diagrams decomposing the stock-recruitment relationship into a series of life317 cycle transitions (Figure 3) illustrated that the shape of the stock-recruitment relationship was primarily driven by the transition from PYSL to YOY (Figure 3;  $\log_{e}(\alpha) = -4.72$ ;  $\beta = 2.21 \times 10^{-6}$ ; 318  $\gamma_{flow} = 1.89 \times 10^{-3}$ ; Table 2). The standing stock of eggs, YSL and PYSL were proportional to the 319 320 standing stock of their previous life-stages in a density-independent manner (p < 0.001; Table 2). The standing stocks of white perch YSL and PYSL were also proportional to the adult standing 321 stock. While each  $\alpha$  estimate for linear transitions was positive and statistically significant, it is 322 important to emphasize that the standing stocks of each life-stage are indices of abundance, and 323 therefore  $\alpha$  estimates do not represent the actual proportionality between life-stage abundances. 324 Comparisons of seven linearized Ricker model structures describing adult-YOY and 325 PYSL-YOY relationships from 1992-2013 indicated that the most likely model included 326 freshwater flow during the PYSL period for both transitions ( $\Delta AICc > 2$  for all other models; 327 Table 3). Flow during the PYSL period had a strong negative effect on  $log_{e}(YOY \times adult^{-1})$  and 328  $log_e(YOY \times PYSL^{-1})$ , which was consistent across models (p < 0.05; Table 3), while the effects of 329 temperature during the PYSL period and zebra mussel filtration rates were comparatively weak 330 (p > 0.25; Table 3). Similarly, the density-dependent term in both relationships (i.e. adult 331 standing stock for  $log_e(YOY \times adult^{-1})$  and PYSL standing stock for  $log_e(YOY \times PYSL^{-1})$ 332 consistently had a strong negative effect on transition responses in each model (p < 0.001; Table 333 334 3).

### **Somatic growth and contingent structure**

For the period 1992-2013, YOY length at the end of the growth season was positively associated with mean summer water temperature (r = 0.70; p < 0.01) and negatively correlated to summer flow (r = -0.50; p < 0.05) and the PYSL standing stock (r = -0.46; p < 0.05; Table 4;

https://mc06.manuscriptcentral.com/cjfas-pubs

339	Figure 4). In addition, the abundance of the migratory contingent of YOY was closely related to
340	the abundance of the resident contingent from 1992-2013 ( $R^2 = 0.81$ ), with the migrant fraction
341	ranging from 0.12 to 0.60, (mean = 0.42; Table 1). The migrant fraction was unrelated to any of
342	the environmental or density-dependent variables analyzed ( $p > 0.15$ ), but was positively
343	correlated to the brackish shoal proportion from July-September ( $r = 0.67$ ; $p < 0.01$ ; Table 4),
344	although this trend was largely driven by two years (2004 and 2011) when the migrant fraction
345	and brackish shoal proportion were both relatively low (Figure 4d). However, after applying a
346	Bonferroni correction for multiple comparisons, the post-invasion correlations between YOY
347	length and summer flow and PYSL abundance, as well as between migrant fraction and brackish
348	shoal proportion became non-significant ( $p > 0.01$ ; Table 4). The effects of temperature, flow
349	and salt front position on YOY white perch growth and partial migration were generally similar
350	to correlations reported prior to the zebra mussel invasion, although the effect of PYSL
351	abundance on YOY growth was weaker during this period (Table S.2).

352

### 2 **Pre- vs. post-invasion indices of abundance**

Post-invasion (1992-2013) indices of white perch abundance were significantly lower 353 than pre-invasion (1974-1991) indices for eggs (p < 0.02), YOY (p < 0.01) and yearlings (p < 0.02) 354 355 0.01; Table 5). In the post-invasion time period, abundance indices for eggs, YOY and yearlings declined by approximately 50%, 40% and 60%, respectively, of their pre-invasion means (Table 356 5). Each life-stage transition differed between the two time periods, although differences varied 357 in magnitude (Figures 5 and 6). The egg-YSL transition exhibited a significantly steeper slope (p 358 < 0.001) in the post-invasion time period ( $\alpha = 0.94$ ) than before the invasion ( $\alpha = 0.41$ ) (Table 6). 359 Change in the YSL-PYSL transition between the two time-periods was not statistically 360 significant (p > 0.15), but showed a slightly flatter slope after the invasion ( $\alpha = 5.62$ ) relative to 361

the pre-invasion period ( $\alpha = 6.53$ ) (Table 6). The PYSL-YOY transition during the post-invasion 362 period exhibited a reduced YOY maximum and increased influence of freshwater flow during 363 the PYSL period ( $log_e(\alpha) = 2.59$ ;  $\beta = 0.57$ ;  $\gamma_{flow} = 1.43 \times 10^{-3}$ ) relative to estimates before the 364 zebra mussel invasion ( $log_e(\alpha) = 2.08; \beta = 0.32; \gamma_{flow} = 2.68 \times 10^{-4}$ ) (Table 6; Figure 5), although 365 contrasts performed on the linearized model coefficients indicated that none of these changes 366 were statistically significant (p > 0.05). Similarly, contrasts indicated that the egg-YOY 367 transition showed significantly enhanced density-dependence after the invasion ( $\beta_{post} = 3.61$ ;  $\beta_{pre}$ 368 = 1.34; p > 0.001), as well as a strong, but non-significant, increase in the effect of freshwater 369 flow ( $\gamma_{post} = 1.57 \times 10^{-3}$ ;  $\gamma_{pre} = -4.86 \times 10^{-4}$ ; p = 0.097) (Table 6). Finally, the linear relationship 370 between the yearling index of abundance and the YOY index from the previous year exhibited a 371 significantly flatter slope (p < 0.05) after the zebra mussel invasion ( $\alpha = 0.21$ ) than was observed 372 during the pre-invasion time period ( $\alpha = 0.32$ ) (Table 6; Figure 6). 373

# 374 **Discussion**

375 In accordance with expected effects of the zebra mussel invasion on the early life history of white perch in the HRE, our analyses revealed substantial alterations in life-stage transitions 376 after the zebra mussel invasion. Following the invasion, the PYSL-YOY transition became more 377 378 sensitive to PYSL abundance (i.e. density-dependence) and freshwater flow experienced during the PYSL period. In addition, declines in yearling abundance and egg production after the zebra 379 mussel invasion suggest that the adult population has been negatively impacted, which has not 380 been reported in previous studies (Strayer et al. 2004). These shifts in white perch abundance and 381 382 life-stage transitions are generally consistent with broader changes in the abundance and flow sensitivity of the forage base in the HRE after the zebra mussel invasion (Strayer et al. 2004; 383

2008). In contrast, similar patterns in pre-invasion and post-invasion correlations suggested that
the effects of density-dependence and environmental conditions on YOY white perch growth and
partial migration changed little after the zebra mussel invasion.

387

### **Environmental and density-dependent effects**

The reduced abundance of YOY white perch observed in years with high PYSL 388 abundance (Figure 3) is consistent with density-dependent mortality of YOY after settlement in 389 390 littoral habitats. This density-dependent mortality is probably most intense shortly after peak settlement, when the abundance of newly metamorphosed YOY is highest. Upon settlement, 391 YOY white perch likely experience high levels of mortality due to predation, which is in turn 392 393 regulated by density and habitat type (Juanes 2007). Density-dependent prey limitation could 394 also cause mortality either directly through starvation or indirectly by exposing juveniles to a prolonged period of size-dependent predation (Walters and Juanes 1993). Lower survival 395 between the PYSL-YOY transition during high flow years could relate to the findings of Strayer 396 397 et al. (2008), who observed an inverse relationship between flow and the abundance of littoral benthic invertebrates (prey items for YOY white perch) in the HRE. Reduced pelagic forage 398 399 availability during the PYSL stage may also be contributing to lower YOY abundances during high flow years. Densities of principal zooplankton prey are known to affect white perch 400 401 recruitment (Limburg et al. 1999), while cladocerans and (to a lesser degree) copepods have been observed to be negatively influenced by flow (Strayer et al. 2008). Contrary to this study, North 402 and Houde (2003) reported that the stock-recruitment relationships of white perch and 403 congeneric striped bass in the Upper Chesapeake Bay were best described when incorporating a 404 405 positive effect of freshwater discharge. Interestingly, differences in the direction of the flow effect on the white perch stock-recruitment function in the Hudson River (negative) and 406

407 Chesapeake Bay (positive) generally match the effects of freshwater flow on primary and
408 secondary production in these contrasting systems (Howarth et al. 2000; Strayer et al. 2008;
409 Testa et al. 2008).

The positive effect of temperature on YOY white perch growth is not surprising, and has 410 been observed in other freshwater and estuarine systems (Kerr and Secor 2010; VanDeValk et al. 411 412 2016). The negative influence of freshwater flow on white perch growth is consistent with previous reports that low flows coincide with increased water clarity and residence times in the 413 HRE (Gladden et al. 1988, Howarth et al. 2000), which may increase primary and secondary 414 production, particularly in littoral habitats used by juvenile white perch (Strayer et al. 2008). The 415 416 negative effect of the PYSL standing stock on YOY growth is more difficult to interpret, 417 especially since growth was not significantly related to YOY abundance, in contrast to other studies that have documented density-dependent growth in white perch (VanDeValk et al. 2016) 418 419 and striped bass (Martino and Houde 2010). The effect of larval abundance on YOY growth 420 (Figure 4c) may be related to the density-dependent mortality observed during the PYSL-YOY transition (see above); such that competition for food in newly settled YOY may increase in 421 422 years of high larval abundance, potentially reducing growth rates early in the YOY period, which carry over to influence length at the end of the growing season. In addition, the relatively 423 constant proportionality between migrant abundance and total YOY abundance implies that 424 years with high recruitment (i.e. moderate PYSL abundance and low flow; Figure 3) will 425 produce higher numbers of migrants with associated higher growth rates (Gallagher 2016), 426 427 which may further enhance population productivity (Kerr et al. 2010).

#### 429 Impacts of the zebra mussel invasion

430 Indices of abundance from pre-invasion (1974-1991) and post-invasion (1992-2013) time periods demonstrated, for the first time, that the abundance of white perch eggs and yearlings 431 declined after zebra mussels invaded the HRE (Table 5), while also confirming the reduction of 432 YOY standing stocks reported in previous studies (Strayer et al. 2004, 2014b). Our analysis of 433 434 life-stage transitions from egg to YSL to PYSL prior to the zebra mussel invasion largely mirrored previous research. Pace et al. (1993) analyzed the egg-YSL, YSL-PYSL and PYSL-435 YOY life-stage transitions during the pre-invasion period from 1974-1990, and found that YSL 436 437 and PYSL abundances were related to the abundance of their previous life-stages in a density-438 independent manner, similar to our analyses of the egg-YSL and YSL-PYSL transitions from 439 1992-2013 (Figure 3; Table 2). However, our analysis indicated that the slope of egg-YSL transition became steeper during the post-invasion time period (Table 6), resulting in a similar 440 441 range of YSL abundance between the two time periods despite declines in egg production (Figure 5; Table 5). This pattern suggests that a compensatory increase in hatching success and 442 early larval survival may have occurred after the invasion. 443

In contrast to our findings, Pace et al. (1993) reported that the relationship between PYSL 444 and YOY abundance was relatively flat, and unrelated to annual variations in temperature and 445 flow, whereas the PYSL-YOY relationship reported here (1992-2013) exhibited strong density-446 dependence and a significant negative effect of freshwater flow during the PYSL period (Figure 447 3; Table 3a). This change in the PYSL-YOY transition after the zebra mussel invasion was 448 corroborated to an extent by the index-based Ricker models describing the transition, which 449 showed a 1.7-fold and 5-fold increase in  $\beta$  ( $\beta_{post} = 0.57$ ;  $\beta_{pre} = 0.32$ ) and  $\gamma_{flow}$  ( $\gamma_{post} = 1.43 \times 10^{-3}$ ; 450  $\gamma_{\rm pre} = 2.68 \times 10^{-4}$ ), respectively, during the post-invasion time-period. Although contrasts 451

indicated that the differences between Ricker model parameters in the pre-invasion and postinvasion time periods were not statistically significant, we argue that the significant reduction in
YOY abundance after the invasion, the consistent negative effect of flow on the post-invasion
adult-YOY and PYSL-YOY transitions, and substantial differences in the shape of the Egg-YOY
and PYSL-YOY relationships between time periods suggest that these differences are
ecologically significant.

The increased sensitivity of the PYSL-YOY transition to density-dependence and 458 freshwater flow may be related to reductions in the food supply for white perch. This change is 459 460 generally consistent with riverwide declines in phytoplankton (Caraco et al. 1997), pelagic 461 zooplankton (Pace et al. 1998) and benthic invertebrate (Strayer and Smith 2001) densities, as 462 well as the strong negative relationship of littoral benthic invertebrate density with flow that Strayer et al. (2008) observed in the first 10-15 years after the zebra mussel invasion. In addition, 463 464 the increased sensitivity to freshwater flow has probably contributed to the diminished postinvasion production of YOY, as flow during the PYSL period (from weeks 19-28) increased at a 465 rate of 1.9% year<sup>-1</sup> on average in the HRE from 1992 to 2013 (Figure 7). The mechanistic basis 466 467 for intensified density-dependence observed after the zebra mussel invasion in the HRE merits further research in white perch and other fishes that reside in rivers, lakes and estuaries affected 468 by invasive bivalves (Kimmerer et al. 2000; Higgins and Vander Zanden 2010). 469

White perch in the HRE begin to mature as yearlings (Klauda et al. 1988), which suggests that reduced yearling abundance after the zebra mussel invasion may have depleted the overall spawning stock abundance of the population. Declines in yearling abundance after the zebra mussel invasion were likely driven by reduced YOY abundance and a flattened slope in the transition between YOY and yearlings, both of which would be expected to decrease the number 475 of yearlings. If the diminished abundance of yearling white perch in the HRE carries over to subsequent age-classes, this would explain the significant reduction in post-invasion egg 476 production (Table 5), and provides one explanation for the negative temporal trends observed in 477 adult, YSL and PYSL standing stocks from 1992-2013 (Table 1). Overall, these comparative 478 479 analyses suggest that the zebra mussel invasion has forced the HRE white perch population into 480 an alternate state, characterized by lower spawning stock abundance and reduced juvenile production that is more sensitive to environmental variation, which can in turn be attributed to 481 differences in life-stage transitions before and after the invasion. 482

### 483 Interactions between ecosystem shifts and climate change

While our study was correlative in nature, the weight of evidence suggests the HRE white 484 485 perch population has been perturbed to an alternate state through the complex interactions of two large-scale phenomena: an abrupt ecosystem shift due to the invasion of zebra mussels (Strayer 486 et al. 2008) and long-term increases in flow associated with climate change (Seekell and Pace 487 488 2011). Precipitation in the HRE watershed is projected to increase by 7% on average over the next century, which should result in a 10-15% increase in freshwater flow (Najjar et al. 2009), 489 490 and this long-term change will likely be overlain by decadal-scale variability that will affect the productivity of white perch and the HRE ecosystem (Strayer et al. 2014a). The response of HRE 491 492 white perch to long-term changes in flow will be further complicated by concurrent increases in 493 water temperature, which Najjar et al. (2009) projected to increase by approximately 4°C on average by 2100. The effects of rising temperatures on the HRE ecosystem are likely to be 494 495 modest over decadal timescales (Strayer et al. 2014a), but the increased biological rates (e.g. zooplankton productivity, white perch growth) that would accompany temperatures at the end of 496

the century could potentially interact with higher freshwater flow (which favors lower primaryproduction; Howarth et al. 2000) in unpredictable ways.

Where adequate environmental and biological monitoring data are available, careful 499 assessment of temporal shifts in life-history transitions may generate a better understanding of 500 501 the interactive effects of ecosystem shifts and climate change on fish populations, and the 502 diversity of environmental drivers involved. Ecosystem shifts are typically abrupt (1-3 years), and can be triggered by overfishing (Daskalov et al. 2007), invasive species (Shiganova et al. 503 1998), eutrophication (Österblom et al. 2007), climate oscillations (Francis and Hare 1994), 504 gradual changes in temperature and precipitation (Smol et al. 2005), or some combination of 505 506 these factors (e.g. Weijerman et al. 2005; Möllmann et al. 2008). Regardless of the cause, abrupt 507 aquatic ecosystem shifts will lead to changes in the abundance and population dynamics of many 508 constituent species (Collie et al. 2004). Further comparisons of fish population characteristics 509 across various stages of ecosystem disturbance and recovery in multiple systems (e.g. nutrient remediation; Kemp et al. 2009) can potentially uncover more diverse mechanisms by which 510 populations have responded to changes in ecosystem structure in the past. 511

This research highlights the importance of abrupt and long-term changes in ecosystem 512 characteristics to the dynamics of the HRE white perch population. Alterations in the abundance 513 514 and environmental sensitivity of white perch in the HRE after the zebra mussel invasion will likely influence how the population responds to long-term increases in temperature and 515 precipitation in the HRE. Similar changes in the shape of early life-stage transitions may be 516 detectable in other fish species in the HRE (e.g. striped bass, river herring, centrarchids) and 517 518 other areas subjected to ecosystem shifts, especially those characterized by intense modifications of the forage base. Improved knowledge of the causes and consequences of state shifts in fish 519

### **Canadian Journal of Fisheries and Aquatic Sciences**

- 520 populations, such as those highlighted here, can enhance efforts to adequately detect and account
- 521 for such changes in conservation and fisheries management practices.

538

Acknowledgements

539	We thank J. Young, the Hudson River Estuary Monitoring Program and Applied Science
540	Associates for allowing us to access data from Year-Class Reports, as well as D. Strayer and the
541	Cary Institute of Ecosystem Studies for sharing zebra mussel filtration rates. We acknowledge J.
542	Testa and H. Bailey for providing insights and comments which improved a previous draft of this
543	manuscript. This research was funded by the Hudson River Foundation under grant number
544	009/13A.
545	
546	
547	
548	
549	
550	
551	
552	
553	
554	
555	
556	
557	
558	

# 559 **References**

- Applied Science Associates, Inc. (ASA). 2014. 2013 Year class report for the Hudson River
   Estuary monitoring program and Central Hudson Gas and Electrical Corporation. New
   Hampton, NY.
- Burnham, K. P. and D. R. Anderson. 2002. Model selection and multimodel inference:
   Apractical information-theoretic approach. Springer-Verlag:New York.
- Caraco, NF, JJ Cole, PA Raymond, and DL Strayer. 1997. Zebra mussel invasion in a large,
   turbid river: phytoplankton response to increased grazing. Ecology 78(2): 588-602.
- Collie, J., K. Richardson, and J. Steele. 2004. Regime shifts: Can ecological theory illuminate the
   mechanisms? Progress in Oceanography 60: 281-302.
- Cooper, J., F. Cantelmo, and C. Newton. 1988. Overview of the Hudson River Estuary. in: L.W.
  Barnthouse, R.J. Klauda, D.S. Vaughan and R.L. Kendall (eds). Science, Law and
  Hudson River Power Plants. American Fisheries Society Monograph 4: 11-24.
- Daniels, R.A., Limburg, K.E., and Schmidt, R.E. 2005. Changes in fish assemblages in the
   tidal Hudson River, New York. American Fisheries Society Symposium 45: 471-503.
- Daskalov G.M., A.N. Grishin, S. Rodionov, and V. Mihneva. 2007. Trophic cascades triggered
   by overfishing reveal possible mechanisms of ecosystem regime shifts. Proceedings of
   the National Academy of Sciences 104(25): 10518-10523.
- Feyrer, F., M. Nobriga, and T. Sommer. 2007. Multidecadal trends for three declining fish
  species: habitat patterns and mechanisms in the San Francisco Estuary, California, USA.
  Canadian Journal of Fisheries and Aquatic Sciences 64: 723–734.
- Folke, C., S. Carpenter, B. Walker, M. Scheffer, T. Elmqvist, L. Gunderson and C.S. Holling.
  2004. Regime shifts, resilience, and biodiversity in ecosystem management. Annual
  Review of Ecology, Evolution, and Systematics 35: 557-581.
- Francis, R.C. and S.R. Hare. 1994. Decadal-scale regime shifts in the large marine ecosystems of
  the North-east Pacific: A case for historical science. Fisheries Oceanography 3(4): 279291.
- Gallagher, B.K. 2016. Influence of partial migration and environmental change on the population
  dynamics of white perch (*Morone americana*) within the Hudson River Estuary. M.S.
  Thesis, University of Maryland, College Park, MD.
- Geyer, W.R. and R. Chant. 2006. The physical oceanography processes in the Hudson River
  Estuary. in: J.S. Levinton and J.R. Waldman (eds). The Hudson River Estuary.
  Cambridge University Press, New York, NY, 121-136.

592 593 594 595	<ul> <li>Gladden, J., F. Cantelmo, J. Croom and R. Shapot. 1988. Evaluation of the Hudson River ecosystem in relation to the dynamics of fish populations. in: L.W. Barnthouse, R.J. Klauda, D.S. Vaughan and R.L. Kendall (eds). Science, Law and Hudson River Power Plants. American Fisheries Society Monograph 4: 69-88.</li> </ul>
596 597	Hare, S.R. and Mantua, N.J. 2000. Empirical evidence for North Pacific regime shifts in 1977 and 1989. Progress in Oceanography, 47(2): 103-145.
598 599 600	Higgins, S.N. and M.J. Vander Zanden. 2010. What a difference a species makes: a meta- analysis of <i>Dreissenid</i> mussel impacts on freshwater ecosystems. Ecological Monographs 80:179-196
601 602	Howarth, R., D. Swaney, T. Butler, and R. Marino. 2000. Rapid Communication: Climatic control on eutrophication of the Hudson River Estuary. Ecosystems 3: 210-215.
603 604 605 606	<ul><li>Howarth, R, R. Marino, D. Swaney and E. Boyer. 2006. Wasterwater and watershed influences on primary productivity and oxygen dynamics in the Lower Hudson River Estuary. in: J.S. Levinton and J.R. Waldman (eds). The Hudson River Estuary. Cambridge University Press, New York, NY, 121-136.</li></ul>
607 608	Juanes, F. 2007. Role of habitat in mediating mortality during the post-settlement transition phase of temperate marine fishes. Journal of Fish Biology 70(3): 661-677.
609 610 611	Kemp, W. M., J.M. Testa, D.J. Conley, D. Gilbert, and J.D. Hagy. 2009. Temporal responses of coastal hypoxia to nutrient loading and physical controls. Biogeosciences 6(12): 2985- 3008.
612 613	Kerr, L., and D. Secor. 2010. Latent effects of early life history on partial migration for an estuarine-dependent fish. Environmental Biology of Fishes 89: 479–492.
614 615 616	Kerr, L., and D. Secor. 2011. Partial migration across populations of white perch ( <i>Morone americana</i> ): A flexible life history strategy in a variable estuarine environment. Estuaries and Coasts 35(1): 227-236.
617 618 619	Kerr, L.A., S.X. Cadrin, and D.H. Secor. 2010. The role of spatial dynamics in the stability, resilience, and productivity of an estuarine fish population. Ecological Applications 20(2): 497-507.
620 621 622	Kimmerer, W., E. Gartside, and J. Orsi. 1994. Predation by an introduced clam as the likely cause of substantial declines in zooplankton of San Francisco Bay. Marine Ecology Progress Series 113: 81-93.
623	

624	Kimmerer, W., J. Cowan, L. Miller, and K. Rose. 2000. Analysis of an estuarine striped bass
625	( <i>Morone saxatilis</i> ) population: influence of density-dependent mortality between
626	metamorphosis and recruitment. Canadian Journal of Fisheries and Aquatic Sciences 57:
627	478–486.
628	<ul> <li>Klauda, R.J., J.B. McLaren, R.E. Schmidt and W.P. Dey. 1988. Life history of white perch in the</li></ul>
629	Hudson River Estuary. in: L.W. Barnthouse, R.J. Klauda, D.S. Vaughan and R.L.
630	Kendall (eds). Science, Law and Hudson River Power Plants. American Fisheries Society
631	Monograph 4: 69-88.
632 633 634	Kraus, R., and D. Secor. 2004a. Dynamics of white perch <i>Morone americana</i> population contingents in the Patuxent River estuary, Maryland, USA. Marine Ecology Progress Series 279: 247–259.
635	Kraus, R., and D. Secor. 2005. Connectivity in estuarine white perch populations of Chesapeake
636	Bay: evidence from historical fisheries data. Estuarine, Coastal and Shelf Science 64(1):
637	108-118.
638	Limburg, K.E., M.L. Pace, D. Fischer and K.K. Arend. 1997. Consumption, selectivity, and use
639	of zooplankton by larval striped bass and white perch in a seasonally pulsed estuary.
640	Transactions of the American Fisheries Society 126(4): 607-621.
641	Limburg, K.E., M.L. Pace, and K.K. Arend. 1998. Growth, mortality, and recruitment of larval
642	Morone spp. in relation to food availability and temperature in the Hudson River. Fishery
643	Bulletin 97: 80–91.
644	Lawler, Matusky and Skelly Engineers, Inc (LMS). 1989. Hudson River Estuary white perch
645	adult and subadult stock assessment study: fall 1988. Final report to: Orange and
646	Rockland Utilities, Inc. Pearl River, NY.
647 648	Mansueti, R.J. 1964. Eggs, larvae, and young of the white perch, <i>Roccus americanus</i> , with comments on its ecology in the estuary. Chesapeake Science. 5(1-2): 3-45.
649	Martino, E.J., and E.D. Houde. 2012. Density-dependent regulation of year-class strength in age-
650	0 juvenile striped bass ( <i>Morone saxatilis</i> ). Canadian Journal of Fisheries and Aquatic
651	Sciences 69(3): 430-446.
652	Möllmann, C., B. Müller-Karulis, G. Kornilovs, and M.A. St John. 2008. Effects of climate and
653	overfishing on zooplankton dynamics and ecosystem structure: regime shifts, trophic
654	cascade, and feedback loops in a simple ecosystem. ICES Journal of Marine Science
655	65(3): 302-310.
656	Najjar, R., L. Patterson, and S. Graham. 2009. Climate simulations of major estuarine watersheds
657	in the Mid-Atlantic region of the US. Climatic Change 95: 139–168.

658	Nash, R., and M. Dickey-Collas. 2005. The influence of life history dynamics and environment
659	on the determination of year class strength in North Sea herring ( <i>Clupea harengus L.</i> ).
660	Fisheries Oceanography 14: 279–291.
661	New York State Department of Environmental Conservation (NYSDEC). 2016. New York State
662	GIS Clearinghouse. Accessed July 2, 2016. Available:
663	<u>https://gis.ny.gov/gisdata/inventories/member.cfm?organizationID=529</u>
664	North, E. W., and E. D. Houde. 2003. Linking ETM physics, zooplankton prey, and fish early-
665	life histories to striped bass <i>Morone saxatilis</i> and white perch <i>M. americana</i> recruitment.
666	Marine Ecology Progress Series 260: 219–236.
667	Österblom H, Hansson S, Larsson U, Hjerne O, F. Wulff, R. Elmgren, and C. Folke. 2007.
668	Human-induced trophic cascades and ecological regime shifts in the Baltic Sea.
669	Ecosystems 10(6): 877-89.
670 671 672	Ottersen, G., D.O. Hjermann, and N.C. Stenseth. 2010. Changes in spawning stock structure strengthen the link between climate and recruitment in a heavily fished cod ( <i>Gadus morhua</i> ) stock. Fisheries Oceanography 15(3): 230–243.
673	Pace, M., S. Baines, H. Cyr, and J. Downing. 1993. Relationships among early life stages of
674	<i>Morone americana</i> and <i>Morone saxatilis</i> from long-term monitoring of the Hudson River
675	Estuary. Canadian Journal of Fisheries and Aquatic Sciences 50: 1976-1985.
676 677	Pace, M., S. Findlay, and D. Fischer. 1998. Effects of an invasive bivalve on the zooplankton community of the Hudson River. Freshwater Biology 39: 103-116.
678 679	Pace, M., D. Strayer, D. Fischer, and H. Malcom. 2010. Recovery of native zooplankton associated with increased mortality of an invasive mussel. Ecosphere 1: 1-10.
680	Paulik, G. J. 1973. Studies of the possible form of the stock-recruitment curve. Rapports et
681	Procès-Verbaux des Réunions du Conseil International pour l'Exploration de la Mer 164:
682	302–315.
683 684 685 686	Payne, M., E. Hatfield, M. Dickey-Collas, T. Falkenhaug, A. Gallego, J. Gröger, P. Licandro, M. Llope, P. Munk, C. Röckmann, J. Schmidt, and R. Nash. 2009. Recruitment in a changing environment: the 2000s North Sea herring recruitment failure. ICES Journal of Marine Science 66: 272-277.
687	Planque, B., J. Fromentin, P. Cury, K.F. Drinkwater, S. Jennings, R.I. Perry, and S. Kifani. 2010.
688	How does fishing alter marine populations and ecosystems sensitivity to climate? Journal
689	of Marine Systems 79: 403–417.
690	Quinn J.T. and R.B. Deriso. 1999. Quantitative fish dynamics. Oxford University Press.

691 692	Seekell, D., and M. Pace. 2011. Climate change drives warming in the Hudson River Estuary, New York (USA). Journal of Environmental Monitoring 13: 2321–2327.
693 694	Shiganova, T. A. 1998. Invasion of the Black Sea by the ctenophore <i>Mnemiopsis leidyi</i> and recent changes in pelagic community structure. Fisheries Oceanography 7(3-4): 305-310.
695 696 697	Smol, J.P., A.P. Wolfe, H.J.B. Birks, M.S. Douglas, V.J. Jones, and 21 others. 2005. Climate- driven regime shifts in the biological communities of arctic lakes. Proceedings of the National Academy of Sciences of the United States of America 102(12): 4397-4402.
698 699 700	Strayer, D.L., and L.C. Smith. 2001. The zoobenthos of the freshwater tidal Hudson River and its response to the zebra mussel ( <i>Dreissena polymorpha</i> ) invasion. Archiv Fur Hydrobiologie-Supplement 139: 1-52.
701 702 703	Strayer, D., K. Hattala, and A. Kahnle. 2004. Effects of an invasive bivalve ( <i>Dreissena polymorpha</i> ) on fish in the Hudson River estuary. Canadian Journal of Fisheries and Aquatic Sciences 61: 924-941.
704 705	Strayer, D., and H. Malcom. 2006. Long-term demography of a zebra mussel ( <i>Dreissena polymorpha</i> ) population. Freshwater Biology 51: 117-130.
706 707	Strayer, DL, ML Pace, NF Caraco, JJ Cole, and S. Findlay. 2008. Hydrology and grazing jointly control a large-river food web. Ecology 89(1): 12-18.
708 709	Strayer, D. L., N. Cid, and H.M. Malcom. 2011. Long-term changes in a population of an invasive bivalve and its effects. Oecologia 165(4): 1063-1072.
710 711 712	Strayer, D., J. Cole, S. Findlay, D. Fischer, J. Gephart, H. Malcom, M. Pace, and E. Rosi- Marshall. 2014a. Decadal-Scale Change in a Large-River Ecosystem. BioScience 64: 496-510.
713 714 715	Strayer, D., K. Hattala, A. Kahnle, R. Adams, and A. Fisk. 2014b. Has the Hudson River fish community recovered from the zebra mussel invasion along with its forage base? Canadian Journal of Fisheries and Aquatic Sciences 71: 1146-1157.
716 717 718	Testa, J., M. Kemp, W. Boynton, and J. Hagy. 2008. Long-term changes in water quality and productivity in the Patuxent River Estuary: 1985 to 2003. Estuaries and Coasts 31(6): 1021-1037.
719 720 721	Tuckett, Q., K. Simon, J. Saros, D. Halliwell, and M. Kinnison. 2013. Fish trophic divergence along a lake productivity gradient revealed by historic patterns of invasion and eutrophication. Freshwater Biology 58: 2517-2531.

722 723 724	United States Geological Survey (USGS). 2015a. National Water Information System data available on the World Wide Web (USGS Water Data for the Nation). Accessed June 9, 2015. Available: <u>http://ny.water.usgs.gov/projects/dialer_plots/saltfront.html</u>
725 726 727	United States Geological Survey (USGS). 2015b. National Water Information System data available on the World Wide Web (USGS Water Data for the Nation). Accessed June 9, 2015. Available: <u>http://waterdata.usgs.gov/nwis/uv?site_no=01372058</u>
728 729 730	United States Geological Survey (USGS). 2015c. National Water Information System data available on the World Wide Web (USGS Water Data for the Nation). Accessed June 9, 2015. Available: <u>http://waterdata.usgs.gov/nwis/uv?site_no=01358000</u>
731 732 733 734	VanDeValk, A.J., J.L. Forney, T.E. Brooking, J.R. Jackson, and L.G. Rudstam. 2016. First-year density and growth as they relate to recruitment of white perch to the adult stock in Oneida Lake, New York, 1968–2011. Transactions of the American Fisheries Society 145(2): 416-426.
735 736 737	Walters, C.J., and F. Juanes. 1993. Recruitment limitation as a consequence of natural selection for use of restricted feeding habitats and predation risk taking by juvenile fishes. Canadian Journal of Fisheries and Aquatic Sciences 50(10): 2058-2070.
738 739	Weijerman, M., H. Lindeboom, and A.F. Zuur. 2005. Regime shifts in marine ecosystems of the North Sea and Wadden Sea. Marine Ecology Progress Series 298: 21-39.
740	
741	
742	
743	
744	
745	
746	
747	
748	
749	
750	

# 752 **Tables**

**Table 1.** List of 13 primary variables analyzed in this study, with descriptive statistics (mean, minimum, maximum and correlation

coefficient with year) for each variable from 1992-2013. Statistically significant temporal correlations (p < 0.05) are marked with an

asterisk (\*). Sources are abbreviated as ASA (Applied Science Associates) and USGS (United States Geological Survey).

Variable	Units	Туре	Source	Mean	Min	Max	$\mathbf{r}_{year}$
Mean adult standing stock	Millions	White Perch	ASA (2014)	2.54	0.89	5.89	-0.30
Mean YOY standing stock	Millions	White Perch	ASA (2014)	0.71	0.05	2.35	0.13
Mean PYSL standing stock	Millions	White Perch	ASA (2014)	709	340	1529	-0.65*
Mean YSL standing stock	Millions	White Perch	ASA (2014)	105	32	257	-0.47*
Mean Egg standing stock	Millions	White Perch	ASA (2014)	123	22	351	-0.17
Proportion of migratory YOY	Unitless	White Perch	ASA (2014)	0.42	0.12	0.60	0.07
Mean YOY length in October	mm	White Perch	ASA (2014)	73.9	62.8	82.3	0.14
Mean zebra mussel filtration rate	m <sup>3</sup> m <sup>-2</sup> day <sup>-1</sup>	Biological	D. Strayer	3.76	0.06	8.28	-0.59*
Mean summer salt front position	km from Battery	Environmental	USGS (2015a)	99	74	118	0.16
Mean spring temperature	°Celsius	Environmental	USGS (2015b)	14.7	13.2	17.2	0.22
Mean summer temperature	°Celsius	Environmental	USGS (2015b)	24.5	23.2	26.4	0.32
Mean spring flow	m <sup>3</sup> sec <sup>-1</sup>	Environmental	USGS (2015c)	587	181	857	-0.01
Mean summer flow	m <sup>3</sup> sec <sup>-1</sup>	Environmental	USGS (2015c)	234	98	546	0.43*

756 757

<b>Table 2.</b> Parameter estimates (with standard errors in parentheses) for each wh	ite perch
---	-----------

life-stage transition based on mean standing stocks during the post-invasion period (1992-2013). The  $\alpha$  estimates, model R<sup>2</sup> and p-values for Ricker models (denoted by <sup>†</sup>) are for the linearized equations, where  $\alpha = \log_e(\alpha)$  (see Methods). 759

760

	Life-Stage Transition	α	β	Yflow	Model R <sup>2</sup>	Model p-value
	Adult-Egg	46.40 (4.56)	-	-	0.38	<0.001
	Adult-YSL	40.41 (3.19)	-	-	0.53	<0.001
	Adult-PYSL	258.25 (21.59)	-	-	0.39	<0.001
	Egg-YSL	0.76 (0.08)	-	-	0.36	<0.001
	Egg-PYSL	4.87 (0.53)	-	-	0.28	<0.001
	YSL-PYSL	6.07 (0.47)	-	-	0.55	<0.001
	Adult-YOY <sup>†</sup>	0.91 (0.49)	5.61 10 <sup>-4</sup> (1.24 10 <sup>-4</sup> )	2.18 10 <sup>-3</sup> (8.61 10 <sup>-4</sup> )	0.59	<0.001
	PYSL-YOY <sup>†</sup>	-4.72 (0.47)	2.21 10 <sup>-6</sup> (4.44 10 <sup>-7</sup> )	1.89 10 <sup>-3</sup> (8.07 10 <sup>-4</sup> )	0.62	<0.001
762						
763						
764						
765						
766						
767						
768						
769						
770						
771						
772						
773						
774						
775						
775						
//6						

777	Table 3. Parameter	• estimates, $\mathbb{R}^2$ ,	, AICc and	$\Delta$ AICc values	for seven	linearized Ricke	er
-----	--------------------	-------------------------------	------------	----------------------	-----------	------------------	----

778 model configurations for white perch describing the PYSL-YOY transition (a) and the

adult-YOY transition (b) with each combination of temperature, flow (both during the

780 PYSL period) and zebra mussel filtration rates. Parameters that were significantly

- different than 0 (p < 0.05) are marked with an asterisk (\*). Model 3 had the lowest AICc
- in both cases (bold italic text).

Model	Intercept	Adult	Temperature	Flow	Filtration	R-square	AICc	ΔΑΙϹϲ
1	0.02	-0.57*	-	-	-	0.46	57.5	3.4
2	-1.95	-0.56*	0.10	-	-	0.46	60.2	6.1
3	0.91	-0.56*	-	-2.19 10 <sup>-3</sup> *	-	0.59	54.1	0.0
4	-0.03	-0.58*	-	-	0.02	0.46	60.5	6.3
5	1.68	-0.56*	-0.04	-2.26 10 <sup>-3</sup> *	-	0.59	57.5	3.3
6	1.13	-0.54*	-	-2.40 10 <sup>-3</sup> *	-0.05	0.60	57.0	2.9
7	2.90	-0.54*	-0.09	-2.40 10 <sup>-3</sup> *	-0.06	0.61	60.6	6.5

**a** 

#### 

b

Model	Intercept	PYSL	Temperature	Flow	Filtration	R-square	AICc	ΔAICc
1	-5.48*	-2.29 10 <sup>-3</sup> *	-	-	-	0.51	53.8	2.6
2	-5.41	-2.28 10 <sup>-3</sup> *	0.00	-	-	0.51	56.8	5.6
3	-4.72*	-2.20 10 <sup>-3</sup> *	-	-1.91 10 <sup>-3</sup> *	-	0.62	51.2	0.0
4	-5.51*	-2.31 10 <sup>-3</sup> *	-	-	0.01	0.51	56.8	5.6
5	-1.51	-2.39 10 <sup>-3</sup> *	-0.16	-2.15 10 <sup>-3</sup> *	-	0.64	53.7	2.5
6	-4.49*	-2.11 10 <sup>-3</sup> *	-	-2.15 10 <sup>-3</sup> *	-0.06	0.64	53.9	2.7
7	-0.30	-2.30 10 <sup>-3</sup> *	-0.20	-2.58 10 <sup>-3</sup> *	-0.07	0.66	56.3	5.0

- **Table 4.** Correlations between white perch migrant fraction and YOY length and six
- 797 predictor variables of interest from the post-invasion period (1992-2013). Correlations
- that were statistically significant (p < 0.05) before multiple comparison adjustment are
- 799 marked with an asterisk (\*), and those that remained significant after applying a
- 800 Bonferroni correction (p <  $4.17 \times 10^{-3}$ ) are marked with two asterisks (\*\*).

				Respon	se
	Predictor	Туре	Time Period	<b>Migrant Fraction</b>	YOY Length
_	PYSL Standing Stock	Density-dependent	May-Jul	-0.16	-0.46*
	YOY Standing Stock	Density-dependent	Jul-Sep	-0.05	0.19
	Summer Temperature	Environmental	Jul-Sep	0.24	0.70**
	Summer Flow	Environmental	Jul-Sep	-0.30	-0.50*
	Zebra Mussel Filtration	Biological	May-Oct	-0.03	-0.06
-	Brackish Shoal Proportion	Habitat availability	Jul-Sep	0.67*	0.37
301					
302					
202					
303					
304					
305					
306					
07					
08					
09					
10					
11					
12					
13					
14					
15					
16					
17					
18					
19					

820	Table 5. Statistical	comparisons	of mean	post-invasion	(1992-2013)	and pre-invasion
-----	----------------------	-------------	---------	---------------	-------------	------------------

821 (1974-1991) indices of abundance for white perch eggs, YSL, PYSL, YOY and yearlings

822 (age-1). The % Difference metric is calculated as (Mean<sub>post</sub>-Mean<sub>pre</sub>) × Mean<sub>pre</sub><sup>-1</sup> × 100.

823 Statistically significant (p < 0.05) p-values are denoted by an asterisk (\*).

Life-Stage	Post-Invasion Mean	Pre-Invasion Mean	t-Statistic	% Difference	p-value
Egg	0.41	0.89	-2.66	-54	0.02*
YSL	0.45	0.46	-0.06	-1	0.95
PYSL	2.85	3.50	-1.30	-19	0.20
YOY	4.82	8.08	-2.83	-40	0.01*
Yearling	1.21	3.26	-3.95	-63	<0.01*

**Table 6.** Parameter estimates (with standard errors in parentheses) and p-values for life-stage transitions based on white perch indices

of abundance shown in Figures 6 and 7. The % Difference metric is calculated as (Estimate<sub>post</sub>-Estimate<sub>pre</sub>) × Estimate<sub>pre</sub><sup>-1</sup> × 100. The contrast t-statistics and p-values are for (Estimate<sub>post</sub>-Estimate<sub>pre</sub>) contrasts performed on each parameter using ANCOVA (estimates

for PYSL-YOY and egg-YOY transitions are from linearized Ricker models where  $\log_e(N_s \times N_{s-l}) = \log_e(\alpha) - \beta(N_{s-l}) - \gamma_{flow}(flow)$ ; see

Methods). Statistically significant (p < 0.05) p-values are denoted by an asterisk (\*).

Life-Stage Transition	Parameter	Post-Invasion Estimate	Post-Invasion p-value	Pre-Invasion Estimate	Pre-Invasion p-value	Contrast t-Statistic	% Difference	Contrast p-value
Egg-YSL	α	0.94 (0.12)	< 0.001*	0.41 (0.06)	< 0.001*	3.81	129	< 0.001*
YSL-PYSL	α	5.62 (0.48)	< 0.001*	6.53 (0.49)	< 0.001*	-1.31	-14	0.199
YOY-Yearling	α	0.21 (0.04)	< 0.001*	0.32 (0.03)	< 0.001*	-2.33	-34	0.026*
PYSL-YOY	$\log_{e}(\alpha)$	2.59 (0.46)	< 0.001*	2.08 (0.88)	0.025*	0.68	25	0.503
	β	0.57 (0.10)	< 0.001*	0.32 (0.17)	0.077	1.84	78	0.074
	γflow	1.43 10 <sup>-3</sup> (9.31 10 <sup>-4</sup> )	0.070	2.68 10 <sup>-4</sup> (7.18 10 <sup>-4</sup> )	0.872	0.97	508	0.34
Egg-YOY	$\log_{e}(\alpha)$	4.47 (0.35)	< 0.001*	3.49 (0.37)	< 0.001*	1.88	28	0.069
	β	3.61 (0.60)	< 0.001*	1.34 (0.20)	< 0.001*	3.60	169	< 0.001*
	γ <sub>flow</sub>	1.57 10 <sup>-3</sup> (7.94 10 <sup>-4</sup> )	0.056	-4.86 10 <sup>-4</sup> (9.08 10 <sup>-4</sup> )	0.596	1.71	-423	0.097

848

849

850

851

852

# 853 Figures

- **Figure 1.** Map of the Hudson River Estuary (HRE), with the 13 river sections sampled by the
- 855 Hudson River Estuary Monitoring Program outlined by river kilometer (rkm), and markers
- 856 denoting approximate locations of monitoring stations for temperature (Poughkeepsie) and
- 857 freshwater flow (Troy; see text). Manhattan and Albany are shown for reference. Shape files for
- the map were obtained from the New York State Department of Environmental Conservation
- 859 GIS Clearinghouse (NYSDEC 2016).



Figure 2. Plots of annual mean temperature (a) and freshwater flow (b) in the HRE from 1951-

866 2013, each plotted with a locally weighted regression (thin black line; LOESS quadratic

smoother with a span of 0.5) to show long-term patterns. The dashed grey lines on each plot

denote the year 1991, when zebra mussels first appeared in the HRE (see text). Temperature data

869 were reported in Applied Science Associates (2014) and freshwater flow data were acquired from USCS (2015a)

870 from USGS (2015c).



- 875
- 876

https://mc06.manuscriptcentral.com/cjfas-pubs

- **Figure 3.** Paulik diagram depicting (clockwise from top right) adult-YOY, adult-YSL, YSL-
- 878 PYSL and PYSL-YOY life-stage transitions based on mean standing stocks during the post-
- invasion period (1992-2013). Note that all four plots within the diagram have a common origin.
- 880 Ricker models are plotted both with (dashed lines) and without (solid lines) freshwater flow
- during the PYSL period as a covariate (see Methods).



**Figure 4.** Scatter plots and regression lines for the four strongest post-invasion correlations

(1992-2013; p < 0.05 before multiple comparison adjustments) between YOY white perch mean

length in October and summer temperature (a), summer flow (b) and PYSL standing stock (c)

and between migrant fraction and the brackish shoal proportion (d). Note that relationships in b, c and d were no longer statistically significant after applying a Bonferroni correction for multiple

901 comparisons (p >  $4.17 \times 10^{-3}$ ).





904

905

906

- 907
- 908
- 909

- 911
- 912

- 913 **Figure 5.** Paulik diagram depicting (clockwise from top right) egg-YOY, egg-YSL, YSL-PYSL
- and PYSL-YOY life-stage transitions based on indices of abundance during pre-invasion (1974-
- 1991; black points; fitted by solid lines) and post-invasion (1992-2013; white points; fitted by
- dashed lines) time periods. Note that all four plots within the diagram have a common origin.
- 917 The egg-YOY and PYSL-YOY relationships are plotted with freshwater flow during the PYSL
- 918 period as a covariate (see Methods).



- **Figure 6.** Relationship between white perch yearling (age-1) abundance and YOY abundance
- from the previous year based on indices of abundance for each life-stage during pre-invasion
- 932 (1974-1991; black points; fitted by solid line) and post-invasion (1992-2013; white points; fitted
- 933 by dashed line) time periods.



**Figure 7.** Plot of mean freshwater flow at Green Island in Troy, NY (USGS 2015c) during

weeks 19-28, when white perch post yolk-sac larvae (PYSL) are most abundant, from 1992-

2013. The regression equation is displayed in the top right corner, and the slope (7.99 year<sup>-1</sup>)
corresponds to an increase of 1.9% year<sup>-1</sup>.

