# J2 optimality and Multi-level Minimum Aberration Criteria in fractional factorial design 

Salawu, I.S., Adeleke B.L. and Oyeyemi, G.M.<br>Department of Statistics University of Ilorin<br>salahus74@yahoo.com


#### Abstract

The desirable properties of fractional factorial design: Balance and orthogonal; was examined for near balance and near orthogonal using the balance coefficient and $\mathrm{J}_{2}$ optimality criteria respectively. Efficient orthogonal arrays with three factors having two, three and four levels were constructed with balance and orthogonal property for lowest common multiples of runs. The two forms of balance coefficient were used for classifying the designs into two and multi level minimum aberration criteria were used to determine designs with lesser aberration. It was observed that designs constructed using the maximum form of balance coefficient has the lesser aberration in both the generalized minimum aberration and minimum moment aberration criteria. The $\mathrm{J}_{2}-$ optimality criterion reveals that the higher the run of a design, the lesser it's optimality value.


Keywords: Balance Coefficient, fractional factorial, Generalized Minimum Aberration (GMA), $\mathrm{J}_{2}$ optimality and Minimum Moment Aberration (MMA).

## 1. Introduction

Factorial designs have broad applications in agricultural, engineering and scientific studies. In constructing and studying properties of factorial designs, traditional design theory treats all factors as nominal. However, this is not appropriate for experiments that involve quantitative factors. For designs with quantitative factors, level permutation of one or more factors in a design matrix could result in different geometric structures, and, thus, different design properties.

Basically two or three - levels factorial experiments are mostly used in the design of experimental research. In many situations, factors with more than two-three levels are desirable, when the factors are either qualitative or quantitative. As a result, designs with mixed-level factors have been used more often in designed experiments in modern industrial and agricultural trials, especially when only limited resources are allowed. Full factorial designs are test matrices that contain all possible combinations of the levels of the factors. For example, if a factor A has "a" levels, factor B has " $b$ " levels and factor C has " c " levels, then the full factorial design will contain "abc" combinations.

The two basic properties of factorial experiments are balance and orthogonal. Balance requires a level of particular factors replicated the same number of times as any other levels of this factor in an experiment. Orthogonal designs are pair wise linearly independent, useful for assessing factor significance.

As the number of factors or factor levels increases, the number of runs increases and maintaining the balance property requires too many runs in some situations. For examples, consider a design with four factors having 3, 5, 5, and 2 levels. To generate a balanced design, at least 150 runs are needed. Suppose an experimenter only has resources for 50 tests and the main objective is screening. Then, a mechanism for creating mixed-level designs which is capable of meeting desirable resources is required.

Optimality properties and ease of interpretation have popularized the use of orthogonal designs in applied research. Traditionally, the construction of orthogonal $\mathrm{S}^{\mathrm{n}}$ fractional factorial designs has been confined to using aliasing relations.

A two-level $2^{k-q}$ design is defined to be a fractional factorial design with k factors, each at two levels, consisting of $2^{k-q}$ runs. Therefore, it is a $2^{-q}$ fraction of the $2^{k}$ full factorial design in which the fraction is determined by q generators. The number of letters in a word is its word length and the word formed by the $q$ defining words is called the defining relation.

For a $2^{k-q}$ design, let $A_{K}(d)$ be the number of words of length $k$ in the defining contrast subgroup. The vector

$$
W(d)=\left(A_{1}(d), A_{2}(d), \ldots, A_{m}(d)\right)
$$

is called the word length pattern of the design d, (Fries and Hunter, 1980). The resolution of a $2^{m-q}$ design, R, is defined to be the smallest r such that $A_{r}(d) \geq 1$, that is, the length of the shortest word in the defining contrast subgroup for any two $2^{\mathrm{m}-\mathrm{q}}$ designs $\mathrm{d}_{1}$ and $\mathrm{d}_{2}$, let r be the smallest integer such that $\mathrm{A}_{\mathrm{r}}\left(\mathrm{d}_{1}\right) \neq \mathrm{A}_{\mathrm{r}}\left(\mathrm{d}_{2}\right)$. Then $d_{l}$ is said to have less aberration than $\mathrm{d}_{2}$ if $\mathrm{A}_{\mathrm{r}}\left(\mathrm{d}_{1}\right)<\mathrm{A}_{\mathrm{r}}\left(\mathrm{d}_{2}\right)$. If no design exists with less aberration than $d_{l}$, then $d_{l}$ has minimum aberration (Fries and Hunter, 1980).

Minimum aberration has been widely recognized as a useful criterion for selecting regular fractional factorials. Recent work on minimum aberration designs includes Chen and Wu (2001), Tang and Wu (1996), Chen and Hedayat (1996), and Cheng et al. (1999). Minimum aberration mixed-level designs are also balanced, Cheng et al, (1999), Deng and Tang (1999), Mukerjee and Wu (2001), Xu and Wu (2001). For unbalanced mixed-level fractional factorial designs, the degree of balance was evaluated using a balance coefficient (Bashir (2003)).

As an extension of two level fractional factorial designs, Franklin (1984) and Suen, Chen and Wu (1997) discuss the construction of multi-level minimum aberration designs. Xu and Wu (2001) proposed a generalized minimum aberration for mixed -level fractional factorial designs. Wu and Zhang (1993) and Ankenman (1999) used minimum aberration designs in two-level and four - level designs. Murkerjee and Wu (2001) developed minimum aberration designs for mixed-level fractional factorial designs involving factors with two or three distinct levels.

Montgomery (2005) gives a slightly different formatted word length pattern from Wu and Zhang (1993), instead of using numbers of words of length k in the defining contrast subgroup, Montgomery (2005) directly shows the length of each word in the defining contrast group.

The objective of this paper is to screen designs using the measure of balance and orthogonal in an efficient fractional factorial design and to compare the designs using multi - level minimum aberration criteria at various run sizes.

## 2. Balance Coefficient

### 2.1 Form I

In form I of the balnced coefficient as defined by Guo (2009), the motivation behind the definition of the balance coefficient is a simple optimization problem. The balance coefficient of design matrices will be derived from the optimization problem stated below:

$$
\begin{aligned}
& \text { Max } \quad G=\prod_{K=1}^{m} X_{K} \\
& \text { Subject to } \quad \sum_{k=1}^{m} X_{K}=C,
\end{aligned}
$$

## Where C is a constant

The balance coefficient for design matrix k, $F(k)$, is defined as the combination of the balanced coefficient of each column, $F_{j}$,

$$
F(k)=\sum_{j=1}^{m} w_{j} F_{J}=\sum_{j=1}^{m}\left(\prod_{i=1}^{l_{j}} l_{i j}\right) w_{j}
$$

Where $w_{j}$ are the weights for the corresponding column $j$. This balance coefficient depends on the runs. A standard number of levels will be used to standardize the balance coefficient. The notations $f_{\mathrm{ij}}$ is used instead of $l_{i j}$.

### 2.2 Form II

In form II as defined by Guo (2009), the definition of balance coefficient employs the concept of the distance function. Consider a distance function

$$
H_{J}=\sum_{i=1}^{l j}\left(l_{i j}-T\right)^{2}
$$

where $T=n / l_{j}$, is a fixed value.
The balance coefficient under this definition becomes

$$
H_{j}=\sum_{i=1}^{l_{j}}\left(l_{i j}-\frac{n}{l_{j}}\right)^{2}
$$

and

$$
H=\sum_{j=1}^{m} H_{j}=\sum_{j=1}^{m} w_{j} \sum_{i=1}^{l_{j}}\left(l_{i j}-\frac{n}{l_{j}}\right)^{2}
$$

If $f_{i j}$ are used instead of $l_{i j}$, then standardized $H_{j}$ and H can be given by

$$
\begin{aligned}
& \hat{H}_{j}=\sum_{i=1}^{l_{j}}\left(f_{i j}-\frac{1}{l_{j}}\right)^{2}, \text { and } \\
& \hat{H}=\sum_{j=1}^{m} w_{j} H_{j}=\sum_{j=1}^{m} w_{j} \sum_{i=1}^{l_{j}}\left(f_{i j}-\frac{1}{l_{j}}\right)^{2}
\end{aligned}
$$

## 3. $\mathbf{J}_{\mathbf{2}}$-Optimality

Consider a n x m, matrix $\mathrm{D}=\left[\mathrm{x}_{\mathrm{ij}}\right]$, with weights $w \mathrm{k}>0$, and $\mathrm{S}_{\mathrm{k}}$ levels for each factor as defined by Guo (2009). The weights indicate the relative importance of each factor and if $\mathrm{w}_{\mathrm{k}}=1$ are chosen, $\delta_{\mathrm{i} j}(\mathrm{~d})$ is the number of coincidences between the $i^{\text {th }}$ and $j^{\text {th }}$ rows. For $1 \leq i, j \leq n$,

$$
\delta_{i, j}(d)=\sum_{k=1}^{m} w_{k} \delta\left(x_{i k}, x_{j k}\right)
$$

Where $\delta(\mathrm{x}, \mathrm{y})=1$ if $\mathrm{x}=\mathrm{y}$ and 0 otherwise.
Then the $J_{2}$-optimality of this design matrix $d$ is defined by Guo (2009) as

$$
J_{2}(d)=\sum_{1 \leq i<j \leq n}\left[\delta_{i, j}(d)\right]^{2}
$$

### 3.1 Lower bound of J2-optimality for balanced designs

When the $\mathrm{J}_{2}$ optimality of a balanced design matrix reaches the lower bound, then, the design is orthogonal. This conclusion was proved by Xu (2003). The lower bound, $\mathrm{L}(\mathrm{n})$, is given as follows,

$$
L(n)=2^{-1}\left[\left(\sum_{k=1}^{m} n s_{k}^{-1} w_{k}\right)^{2}+\left(\sum_{k=1}^{m}\left(s_{k}-1\right)\left(n s_{k}^{-1} w_{k}\right)^{2}\right)-n\left(\sum_{k=1}^{m} w_{k}\right)^{2}\right]
$$

## 4. Minimum Aberrations

Minimum aberration has been widely recognized as a useful criterion for selecting regular fractional factorials. Recent work on minimum aberration designs includes Chen and Wu (2001), Tang and Wu (1996), Chen and Hedayat (1996), and Cheng et al. (1999).

### 4.1 Generalized Minimum Aberration Criterion

Xu and Wu (2001) proposed a generalized minimum aberration (GMA) criterion for multi-level and mixed-level designs. For a design $d$, the ANOVA model has the following form

$$
Y=X_{0} \alpha_{0}+X_{1} \alpha_{1}+\ldots+X_{m} \alpha_{m}+\varepsilon
$$

where $Y$ is the response, $\alpha_{k}$ is the vector of all $k$-factor interactions and $X_{k}=\left[x_{i j}^{(k)}\right]$ is the matrix of contrast coefficients for $\alpha_{k}$. Let

$$
A_{k}(d)=n^{-2} \sum_{j}\left|\sum_{i=1}^{n} x_{i j}^{(k)}\right|^{2} .
$$

The $A_{k}(d)$ are invariant with respect to the choice of orthogonal contrasts. The vector $\left(A_{1}(d), A_{2}(d), \ldots A_{m}(d)\right)$ is called the generalized word length pattern. Then the generalized minimum aberration criterion is to sequentially minimize $A_{k}(d)$ for $k=1, \ldots, m$.

### 4.2 Minimum Moment Aberration Criterion

The Minimum Generalized Aberration ( $M G A$ ), Minimum $\mathrm{G}_{2}$ Abberation $\left(M G_{2} A\right.$ ), and Generalized Minimum Aberration (GMA) criteria all require contrast coefficients of factors. Xu (2003) developed a Minimum Moment Aberration criterion (MMA), which does not need contrast coefficients. For a design matrix $d$, with $d_{i j}$ as the elements of $i^{\text {th }}$ row and $j^{\text {th }}$ column. The coincidence between two elements $d_{i j}$ and $d_{l j}$ is defined by $\delta\left(d_{i j}, d_{l j}\right)$, where $\delta\left(d_{i j}, d_{l j}\right)=1$ if $d_{i j}=d_{l j}$ and 0 otherwise. The value of $\sum_{j=1}^{m} \delta\left(d_{i j}, d_{l j}\right)$ measures the coincidence between $i^{t h}$ and $j^{t h}$ rows of $d$. The $k^{t h}$ power moment is defined by Xu (2003) as

$$
K_{k}(d)=[n(n-1) / 2]^{-1} \sum_{1 \leq i \leq l \leq n}\left[\sum_{j=1}^{m}\left(d_{i j}, d_{l j}\right)\right]^{k}
$$

For two designs $d_{1}$ and $d_{2}, d_{1}$ is said to have less moment aberration than $d_{2}$ if there exists an $r$ such that $K_{r}\left(d_{l}\right)<K_{r}\left(d_{2}\right)$ and $K_{t}\left(d_{l}\right)=K_{t}\left(d_{2}\right)$ for all $t=1, \ldots, r-1$. Therefore, $d_{l}$ is said to have minimum moment aberration if there is no other design with less moment aberration than $d_{l}$.

Table 4.1: Design selected using balance coefficient and their $\mathbf{J}_{\mathbf{2}}$ - optimality in $O A\left(n, 2^{1} 3^{1} 4^{1}\right)$

| Runs | Balance coefficient methods | Column 1 | Column 2 | Column 3 | $\mathrm{J}_{2}$ optimality |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 6 | Distance Function | $\left[\begin{array}{ll}1 & 5\end{array}\right]$ | $\left[\begin{array}{lll}1 & 4 & 1\end{array}\right]$ | $\left[\begin{array}{llll}1 & 1 & 1 & 3\end{array}\right]$ | 39 |
|  | Optimization Procedure | $\left[\begin{array}{ll}3 & 3\end{array}\right]$ | [2, 2, 2] | [1, 1, 1, 3] | 24 |
| 7 | Distance Function | [1, 6] | [1, 5, 1] | [1, 4, 1, 1] | 77 |
|  | Optimization Procedure | [3, 4] | [1, 3, 3] | [1, 2, 2, 2] | 36 |
| 8 | Distance Function | [1, 7] | [1, 6, 1] | [1, 5, 1, 1] | 116 |
|  | Optimization Procedure | [4, 4] | [3, 3, 2] | [2, 2, 2, 2] | 47 |
| 9 | Distance Function | [1, 8] | [1, 7, 1] | [1, 1, 6, 1] | 166 |
|  | Optimization Procedure | $[4,5]$ | [3, 3, 3] | [1, 2, 3, 3] | 70 |
| 10 | Distance Function | [1, 9] | [1, 8, 1] | [7, 1, 1, 1] | 201 |
|  | Optimization Procedure | [5,5] | [3, 3, 4] | [2, 2, 3, 3] | 82 |
| 11 | Distance Function | [2, 9] | [1, 9, 1] | [8, 1, 1, 1] | 231 |
|  | Optimization Procedure | [5, 6] | [2, 4, 5] | [3, 2, 3, 3] | 112 |
| 12 | Distance Function | [3, 9] | [1, 9, 2] | [1, 9, 1, 1] | 276 |
|  | Optimization Procedure | $[6,6]$ | [4, 4, 4] | [3, 3, 3, 3] | 128 |
| 13 | Distance Function | [9, 4] | [3, 9, 1] | [9, 1, 2, 1] | 272 |
|  | Optimization Procedure | [7, 6] | [3, 5, 5] | [3, 3, 3, 4] | 162 |
| 14 | Distance Function | [9, 5] | [1, 9, 4] | [1, 3, 1, 9] | 271 |
|  | Optimization Procedure | [7,7] | [5, 5, 4] | [2, 3, 4, 5] | 220 |
| 15 | Distance Function | [9, 6] | [1, 9, 5] | [1, 4, 9, 1] | 312 |
|  | Optimization Procedure | $[8,7]$ | [4, 5, 6] | [2, 3, 4, 6] | 249 |
| 16 | Distance Function | [7, 9] | [2, 9, 5] | [1, 9, 1, 5] | 336 |
|  | Optimization Procedure | [8, 8] | [3, 6, 7] | [1, 8, 3, 4] | 313 |
| 17 | Distance Function | [9, 8] | [1, 9, 7] | [1, 7, 1, 8] | 839 |
|  | Optimization Procedure | [9, 8] | [4, 6, 7] | [1, 7, 1, 8] | 670 |
| 18 | Distance Function | [9, 9] | [1, 9, 8] | [1, 7, 1, 9] | 515 |
|  | Optimization Procedure | [9, 9] | [6, 6, 6] | [3, 4, 5, 6] | 341 |

In the above Table 4.1, the balance coefficient criteria are used to select the designs of columns 1, 2, and 3 and their $\mathrm{J}_{2}$ optimality obtained. It was observed that the designs selected using the optimization procedure method of balance coefficient has a lower $\mathrm{J}_{2}$ optimality in all the designs selected.
The two basic criteria of fractional factorial design: balance and orthogonality are measured using the balance coefficient and $\mathrm{J}_{2}$-optimality was used to select designs between Optimization Procedure and Distance Function pattern of balance coefficient. Minimum aberration criterion is used to select the best among efficient designs. The selected designs were subjected to comparison using multi-level aberration criteria such as minimum moment aberration criteria (MMAC) and generalized minimum aberration criteria (GMAC). These criteria were selected because; the constructed design involves more than two levels i.e. mixed level design that does not make use of design generator.

Table 4.2: Design Comparison using Generalized Minimum Aberration Criterion (GMAC) in $\left(n, 2^{1} 3^{1} 4^{1}\right)$

|  |  | Generalized Minimum Aberration Criteria |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Runs | Designs | Sum of Squares | $\mathrm{A}_{\mathrm{i}}\left(\mathrm{d}_{\mathrm{i}}\right)$ | $\mathrm{d}_{2}$ |
| 6 | Distance Function | 16.99 | 0.472 |  |
|  | Optimization Procedure | 2.99 | 0.083 |  |
| 7 | Distance Function | 29.89 | 0.61 | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | 3.92 | 0.08 |  |
| 8 | Distance Function | 46.04 | 0.72 | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | 0.67 | 0.01 |  |
| 9 | Distance Function | 67.21 | 0.82 | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | 3.24 | 0.04 |  |
| 10 | Distance Function | 91.62 | 0.92 | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | 1.67 | 0.02 |  |
| 11 | Distance Function | 103.86 | 0.86 | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | 5.91 | 0.05 |  |
| 12 | Distance Function | 103.95 | 0.72 | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | 0 | 0 |  |
| 13 | Distance Function | 91.87 | 0.54 | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | 3.92 | 0.02 |  |
| 14 | Distance Function | 83.63 | 0.43 | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | 5.67 | 0.03 |  |
| 15 | Distance Function | 79.22 | 0.35 | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | 11.25 | 0.05 |  |
| 16 | Distance Function | 70.64 | 0.28 | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | 34.66 | 0.14 |  |
| 17 | Distance Function | 77.89 | 0.27 | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | 5.91 | 0.02 |  |
| 18 | Distance Function | 88.97 | 0.27 | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | 4.99 | 0.02 |  |

The generalized minimum aberration criteria in the comparison of the designs, $O A\left(n, 2^{1} 3^{1} 4^{1}\right)$, using both

Optimization Procedure and Distance Function methods of balance coefficient for $6 \leq n \leq 18$, it was shown that at $6 \leq n \leq 18, A_{1}\left(\mathrm{~d}_{2}\right)<A_{1}\left(\mathrm{~d}_{1}\right)$, i.e. design $\mathrm{d}_{2}$ has less aberration than $\mathrm{d}_{1}$. Therefore, $\mathrm{d}_{2}$ is better than $\mathrm{d}_{1}$ by the GMAC.

Table 4.3: Designs using Minimum Moment Aberration Criteria (MMA) in $\left(n, 2^{1} 3^{1} 4^{1}\right)$

|  |  | Minimum Moment Aberration Criteria |  |
| :---: | :---: | :---: | :---: |
| N | Designs | $\left(K_{1}(\mathrm{~d}), \mathrm{K}_{2}(\mathrm{~d}), \mathrm{K}_{3}(\mathrm{~d}), \mathrm{K}_{4}(\mathrm{~d})\right)$ |  |
| 6 | Distance function | (1.267, 2.6, 5.667, 13) | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | (0.8, 1.6, 3.6, 8.8) |  |
| 7 | Distance function | (1.429, 3.667, 9.476, 26.43) | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | (0.857, 1.714, 4.809, 10.286) |  |
| 8 | Distance function | (1.643, 4.143, 11.286, 32) | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | (0.821, 1.679, 3.964, 10.607) |  |
| 9 | Distance function | (1.75, 4.611, 12.778, 36.611) | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | (0.88, 1.94, 4.88, 13.28) |  |
| 10 | Distance function | (1.889, 4.467, 11.622, 31.933) | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | (0.88, 1.82, 3.67, 10.62) |  |
| 11 | Distance function | (1.727, 4.2, 10.545, 28.2) |  |
|  | Optimization Procedure | (0.872, 2.036, 4.873, 12.509) | $\mathrm{d}_{2}$ |
| 12 | Distance function | (1.697, 4.181, 11.060, 30.879) | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | (0.909, 1.939, 4.727, 12.485) |  |
| 13 | Distance function | (1.513, 3.487, 8.897, 24.103) | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | (1.025, 2.077, 5.103, 13.462) |  |
| 14 | Distance function | (2.494, 2.978, 7.132, 17.703) | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | (1, 2.418, 10.099, 16.923) |  |
| 15 | Distance function | (1.352, 2.971, 7.038, 18.371) | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | (0.96, 2.37, 6.314, 17.457) |  |
| 16 | Distance function | (1.267, 2.75, 7.05, 19) | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | (1.1, 2.608, 6.55, 17.483) |  |
| 17 | Distance function | (1.235, 6.169, 9.279, 26.779) | $\mathrm{d}_{2}$ |
|  | Optimization Procedure | (0.98, 4.93, 5.78, 15.93) |  |
| 18 | Distance function | (1.261, 3.366, 10.189, 26.634) |  |
|  | Optimization Procedure | (0.987, 2.229, 5.693, 15.562) | $\mathrm{d}_{2}$ |

The minimum aberration criteria for two selected designs using Optimization Procedure and Distance Function method of balance coefficient, for $6 \leq n \leq 18$.
The observation shows that at $6 \leq n \leq 18$

$$
K_{1}\left(\mathrm{~d}_{2}\right)<K_{1}\left(\mathrm{~d}_{1}\right)
$$

This indicated that in all the runs mentioned, $K_{1}\left(\mathrm{~d}_{2}\right)$ has a lesser aberration than $K_{1}\left(\mathrm{~d}_{1}\right)$, that is, the design $\mathrm{d}_{2}$ is a better fractional factorial of all possible designs in the runs considers.

Table 4.4: Summary of designs evaluated in Table 4.2 and 4.3

|  |  | $\left(n, 2^{1} 3^{1} 4^{1}\right)$ |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\mathbf{n}$ | MMAC | GMAC |  | $\mathbf{n}$ | MMAC | GMAC |
| Optimization procedure | $\mathbf{6}$ | $\mathrm{d}_{2}$ | $\mathrm{~d}_{2}$ |  | $\mathbf{1 3}$ | $\mathrm{~d}_{2}$ | $\mathrm{~d}_{2}$ |
| Optimization Procedure | $\mathbf{7}$ | $\mathrm{d}_{2}$ | $\mathrm{~d}_{2}$ |  | $\mathbf{1 4}$ | $\mathrm{~d}_{2}$ | $\mathrm{~d}_{2}$ |
| Optimization Procedure | $\mathbf{8}$ | $\mathrm{d}_{2}$ | $\mathrm{~d}_{2}$ |  | $\mathbf{1 5}$ | $\mathrm{~d}_{2}$ | $\mathrm{~d}_{2}$ |
| Optimization Procedure | $\mathbf{9}$ | $\mathrm{d}_{2}$ | $\mathrm{~d}_{2}$ |  | $\mathbf{1 6}$ | $\mathrm{~d}_{2}$ | $\mathrm{~d}_{2}$ |
| Optimization Procedure | $\mathbf{1 0}$ | $\mathrm{d}_{2}$ | $\mathrm{~d}_{2}$ |  | $\mathbf{1 7}$ | $\mathrm{~d}_{2}$ | $\mathrm{~d}_{2}$ |
| Optimization Procedure | $\mathbf{1 1}$ | $\mathrm{d}_{2}$ | $\mathrm{~d}_{2}$ |  | $\mathbf{1 8}$ | $\mathrm{~d}_{2}$ | $\mathrm{~d}_{2}$ |
| Optimization Procedure | $\mathbf{1 2}$ | $\mathrm{d}_{2}$ | $\mathrm{~d}_{2}$ |  |  |  |  |

### 5.0 Conclusion

In this paper, constructed efficient fractional factorial designs with balance coefficient and $\mathrm{J}_{2}$ optimality criteria were used to compare the two forms of balanced coefficient methods using the generalized minimum aberration and minimum moment aberration criteria. It was observed that designs constructed using the maximum form of balance coefficient has the less aberration in both the generalized minimum aberration and minimum moment aberration criteria.
The major contributions of this paper include:

The screening of efficient mixed level designs associated with optimal near-balance and near orthogonal properties. The efficient mixed-level fractional designs are the solution to the excessive run requirements associated with many balanced mixed-level designs.
Balance coefficient of form I using maximum methods produces an efficient mixed fractional factorial designs as shown by the minimum aberrations
$\mathrm{J}_{2}$ optimality values are lower in efficient fractional factorial designs
The relationship between the balance coefficient and $\mathrm{J}_{2}$ optimality is inverse, that is, as one increase the other decreases

## References

Ankenman, B. E. (1999) "Design of Experiments with Two-Level and Four-Level Factors", Journal of Quality Technology, 31, 368-375.
Bashir,A. (2003): "A New Approach for Constructing and Analyzing Supersaturated Design for Experiments involving Numerous Factors".
Cheng, C.S., Steinberg, D.M. \& Sun, L.X. (1999): Minimum aberration and model robustness for two-level fractional factorial design. J. Roy Statis. Soc. Ser. B, 61, 85-93.
Chen, H. \& Hedayat, A.S. (1996): 2n-l designs with weak minimum aberration. Ann. Statist. 24 2536-2548.
Cheng, S.W. and Wu, C.F.J. (2001): Factor Screening and response surface exploration (with discussion) Statist. Sinica. 11 553-604.
Deng, L.Y. \& Tang, B. (1999): Minimum G2- aberration for non-regular fractional factorial designs. Statist. Sinica 5 235-250.
Franklin, M.F. (1984): Constructing tables of minimum aberration Pn-m designs. Technometrics 26 225-232
Montgomery, D.C. (2005). Design and Analysis of Experiments, 5th ed., John Wiley \& sons, Inc. New York, NY. Mukerjee, R. \& Wu, C.F.J. (2001): A Modern Theory of Factional Designs. New York: Springer.
Suen, C., Chen, H. \& Wu, C.F.J. (1997): Some identities on designs with application to minimum aberration designs. Ann. Statist. 25 1176-1188.
Tang, B. \& Wu, C.F.J. (1996): Characterization of minimum aberration $2 \mathrm{n}-\mathrm{m}$ designs in terms of their complementary designs. Ann Statist. 242549 - 2559.
Wu, C.F.J. \& Zhang, R.(1993): Minimum aberration designs with two - level and four - level factors Biometrika 80 203-209
Xu, H. (2003): "An Algorithm for Construction Orthogonal and Nearly-Orthogonal Designs with Mixed Levels and Small Runs" Technometrices, 44, 356-368.
Xu, H. \& Wu, C.F.J. (2001): Generalized minimum aberration for asymmetrical fractional factorial designs. Ann. Statist. 29 1066-1077
Yong Guo (2009): 'Deciphering all those Minimum Aberration Criteria for Experimental Designs'. Quality Engineering 10:2009

This academic article was published by The International Institute for Science, Technology and Education (IISTE). The IISTE is a pioneer in the Open Access Publishing service based in the U.S. and Europe. The aim of the institute is Accelerating Global Knowledge Sharing.

More information about the publisher can be found in the IISTE's homepage: http://www.iiste.org

## CALL FOR PAPERS

The IISTE is currently hosting more than 30 peer-reviewed academic journals and collaborating with academic institutions around the world. There's no deadline for submission. Prospective authors of IISTE journals can find the submission instruction on the following page: http://www.iiste.org/Journals/

The IISTE editorial team promises to the review and publish all the qualified submissions in a fast manner. All the journals articles are available online to the readers all over the world without financial, legal, or technical barriers other than those inseparable from gaining access to the internet itself. Printed version of the journals is also available upon request from readers and authors.

## IISTE Knowledge Sharing Partners

EBSCO, Index Copernicus, Ulrich's Periodicals Directory, JournalTOCS, PKP Open Archives Harvester, Bielefeld Academic Search Engine, Elektronische Zeitschriftenbibliothek EZB, Open J-Gate, OCLC WorldCat, Universe Digtial Library, NewJour, Google Scholar

(1) ULRICHSWEB

JournalTOCs
PKP | public knowledge project


