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Does Misery Love Company - Evidence from Pharmaceutical Markets before and after the Orphan Drug Act

Frank R. Lichtenberg
Columbia University Graduate School of Business

Joel Waldfogel
Wharton School of the University of Pennsylvania

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DOES MISERY LOVE COMPANY? EVIDENCE FROM PHARMACEUTICAL MARKETS BEFORE AND AFTER THE ORPHAN DRUG ACT

*Frank R. Lichtenberg**
*Joel Waldfogel***

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* Frank R. Lichtenberg is the Courtney C. Brown Professor of Business at the Columbia University Graduate School of Business, and a Research Associate of the National Bureau of Economic Research. He received a B.A. with Honors in history from the University of Chicago and an M.A. and a Ph.D. in economics from the University of Pennsylvania. He was awarded the 1998 Schumpeter Prize for his paper, *Pharmaceutical Innovation as a Process of Creative Destruction*, and a 2003 Milken Institute Award for Distinguished Economic Research for the paper, *Pharmaceutical Knowledge-Capital Accumulation and Longevity*.

** Joel Waldfogel is the Ehrenkranz Professor of Business and Public Policy at the Wharton School of the University of Pennsylvania. He received his A.B. in economics from Brandeis University in 1984 and a Ph.D. in economics from Stanford University in 1990. He does research on law and economics as well as on industrial economics. He also has a strong research interest in products with large fixed costs, such as pharmaceuticals and media products. He is the author of many articles in economics journals, as well as *THE TYRANNY OF THE MARKET* (Harvard Univ. Press 2007).

INTRODUCTION

When production entails fixed or sunk costs, the number of products developed can increase with the size of the market. A larger potential market provides greater reward for firms that can bring a new product to market. Additional products increase welfare because if products are differentiated, then additional products confer benefits by giving more consumers options that better suit their needs. In this way, consumers benefit each other via a mechanism one might term “preference externalities.” Of course, whether or not products are differentiated, additional products can place downward pressure on prices.¹

Although the relationship between market size and consumption and, by extension, welfare operating through product variety follows from theory in straightforward ways, corresponding empirical evidence is scarce.² Yet, the conditions giving rise to this phenomenon can appear whenever fixed costs are large relative to market size. Nowhere is this more likely to be true than in pharmaceutical markets. According to the Tufts Center for the Study of Drug Development, the average cost to develop a new molecular entity is \$802 million.³ The number of drugs available per condition bears out the claim that drug development costs are large, relative to market size, for many conditions. The median number of drugs labeled to treat a four-digit ICD9 condition is two.⁴ These

1. These mechanisms have been outlined in seminal papers. See Avinash K. Dixit & Joseph E. Stiglitz, *Monopolistic Competition and Optimum Product Diversity*, 67 AM. ECON. REV. 297 (1977); A. Michael Spence, *Product Differentiation and Welfare*, 66 AM. ECON. REV. 407 (1976) [hereinafter Spence, *Product Differentiation*]; A. Michael Spence, *Product Selection, Fixed Costs, and Monopolistic Competition*, 43 REV. ECON. STUD. 217 (1976) [hereinafter Spence, *Product Selection*]. See also N. Gregory Mankiw & Michael D. Whinston, *Free Entry and Social Inefficiency*, 17 RAND J. ECON. 48 (1977) (emphasizing possible inefficiencies of entry).

2. See Lisa George & Joel Waldfogel, *Who Affects Whom in Daily Newspaper Markets?*, 111 J. POL. ECON. 765 (2003); Joel Waldfogel, *Preference Externalities: An Empirical Study of Who Benefits Whom in Differentiated-Product Markets*, 34 RAND J. ECON. 557 (2003).

3. The Tufts study was based on detailed survey data obtained directly from ten drug companies. A similar study done by the Tufts Center a decade ago indicated that the average cost to develop a new drug then was about \$231 million, in 1987 dollars. Recent News, Tufts Center for the Study of Drug Development, *Tufts Center for the Study of Drug Development Pigs Cost of a New Prescription Medicine at \$802 Million* (Nov. 30, 2001), <http://csdd.tufts.edu/NewsEvents/RecentNews.asp?newsid=6>.

4. The *International Classification of Diseases* (ICD) was designed to promote international comparability in the collection, processing, classification, and presentation of mortality statistics. The ICD has been revised periodically to incorporate changes in the medical field. To date, there have been 10 revisions of the ICD. The ninth revision (ICD9) was used during the period 1979–1998. The ICD9 classification is hierarchical; a 4-digit ICD9 code is the most detailed code. See Nat’l Ctr. for Health Statistics, *International Classification of Diseases, Ninth Revision (ICD-9)* (Jan. 11, 2007), <http://www.cdc.gov/nchs/about/major/dvs/icd9des.htm>. The figure for the median number of drugs is based on proprietary data provided to Frank Lichten-

facts lead us to ask whether individuals are better off in their capacity as drug consumers if their condition is more common. More succinctly, we ask whether “misery loves company.”

Despite the novelty of the academic question of the welfare of small consumer groups in markets, concern about this issue is not new to policy makers. The possibility that small populations would see few medications developed for their conditions has already led the U.S. Congress to pass the 1983 Orphan Drug Act (“ODA”), giving firms special incentives to develop drugs for diseases afflicting fewer than 200,000 persons per year.⁵ The ODA contains provisions that reduce the cost, and raise the appropriability, of research on rare diseases. First, under the Act, drug makers receive seven years of exclusive marketing upon FDA approval of newly-developed drugs qualifying as “orphan drugs”—i.e., drugs for disorders affecting fewer than 200,000 persons.⁶ According to the FDA, this is the “most sought incentive.”⁷ For seven years following FDA approval, the FDA cannot approve another drug for the same indication without the sponsor’s consent.⁸ Second, drug makers qualify for a tax credit for clinical research expense of up to 50 percent of clinical testing expense.⁹ In addition, the FDA provides grant support for investigation of rare disease treatments.¹⁰ Together, these provisions (a) increase the effective market size; and (b) reduce fixed (sunk) costs. In doing so, the Act provides a natural experiment for measuring the impact of increased market size, relative to fixed costs, on product development, consumption, and welfare.

According to the FDA, the ODA has had a large effect on drug development: “ODA has been very successful—more than 200 drugs and biological products for rare diseases have been brought to market since 1983. In contrast, the decade prior to 1983 saw fewer than ten such products come to market.”¹¹ A complete list of the drugs that have been granted orphan drug status by the FDA is provided in Appendix

berg by First DataBank. See First DataBank, Drug Indications Master Table (First DataBank’s National Drug Data File CD-ROM, 1999).

5. Orphan Drug Act, Pub. L. No. 97-414, 96 Stat. 2049 (1983) (codified as amended at 21 U.S.C. §§ 360aa–360ee (1994)).

6. *Id.* § 526(a)(2).

7. Marlene E. Haffner, Director, FDA Office of Orphan Drug Prods., From Bench to Bedside to Practice: A Practical Course—Genetic Alliance Annual Conference (July 28, 2006), slide 6, available at http://www.geneticalliance.org/ksc_assets/pdfs/conf06/incentives_to_drug_development.ppt.

8. Orphan Drug Act § 527(a)(3).

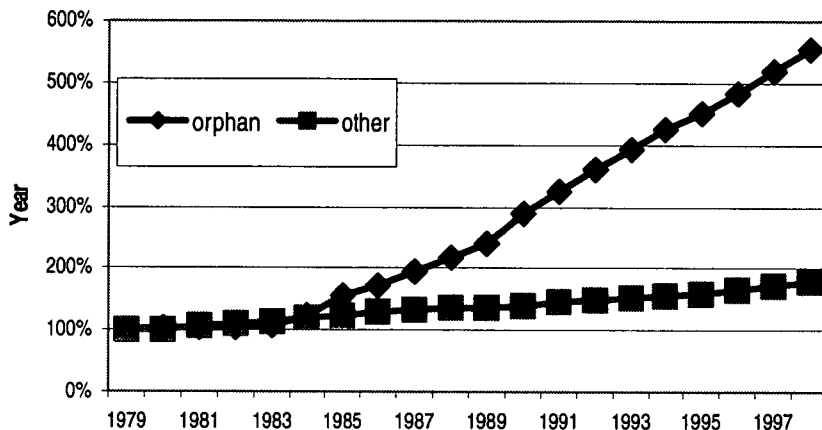
9. See 26 U.S.C. § 45C (2008).

10. See Information on the OOPD Grant Program, <http://www.fda.gov/orphan/grants/info.htm> (last visited Feb. 23, 2009).

11. Welcome to the Office of Orphan Products Development, <http://www.fda.gov/orphan/history.htm> (last visited Feb. 23, 2009).

Table 1.¹² Figure 1 shows the cumulative number of orphan and non-orphan drugs approved, 1979-1998, as a percent of the cumulative number of drugs approved in 1979.

FIGURE I
CUMULATIVE NUMBER OF DRUGS APPROVED, AS PERCENTAGE
OF CUMULATIVE NUMBER OF DRUGS APPROVED IN 1979:
ORPHAN VS. OTHER DRUGS



Between 1979 and 1983, the number of orphan drugs increased at about the same rate as the number of other drugs. By 1998, there were more than five times as many orphan drugs as there had been in 1979, and fewer than twice as many non-orphan drugs.

In light of the apparent effect of the ODA on drug development, we examine its effect on two measures related to welfare: consumption and mortality. First, we ask whether there is evidence, in the pharmaceutical context, that misery loves company. We compare across conditions with different levels of prevalence (“market size”), asking whether physicians are more likely to prescribe drugs for common diseases, and whether people with common diseases are likely to live longer. Results from this approach are highly suggestive: more prevalent conditions have substantially more products available, and we document both that larger affected populations are much more likely to take a drug than smaller affected patient populations, and that mortality rates are lower for persons with more common conditions. A shortcoming of this approach, however, is the possibility of unobserved heterogeneities leading both to large mar-

12. Some of these drugs also have non-orphan indications: i.e., they may be used to treat common diseases.

kets and many drugs. Putting this differently, the cross-sectional measurement strategy may not reflect a clean source of exogenous variation in market size.

Conveniently, the passage of the Orphan Drug Act provides a source of exogenous variation in market size, relative to fixed costs, for drugs targeting small populations. This motivates our second measurement approach for documenting the effect of market size on drug consumption and, by extension, welfare. We document growth in consumption and increases in longevity for individuals with less common conditions, relative to those with more common conditions. Moreover, we document that these effects on consumption and longevity are significantly related to orphan drug use for the condition.

This Article proceeds in four parts. Part I provides background by outlining the mechanism underlying preference externalities. We also review relevant literature in this Part. Part II describes the data used in this study. Part III presents our empirical strategy and results. We find clear cross-sectional evidence that misery loves company, both before and after the passage of the Orphan Drug Act. But the Act appears to have weakened the link between market size and welfare: conditions with substantial orphan drug use have larger increases in consumption and longevity than others. In the conclusion, we consider our results in both narrow and broad contexts.

I. MARKET SIZE, ENTRY, AND WELFARE: WHY WOULD MISERY LOVE COMPANY?

This Article is mainly concerned with the positive question of how market size affects drug development, consumption, and other measures of welfare. It is helpful to locate this problem in its normative context, which we briefly do below.

When development carries sunk costs and products are imperfect substitutes, markets can fail to achieve optimal outcomes.¹³ First, if sellers cannot capture the entire consumer valuation of their product, some products with consumer valuation in excess of their production cost will not be developed. That is, inefficient under-provision is possible. At the same time, because products are substitutes, the private benefit of entry can exceed the resulting social benefit if some of a product's business is

13. These problems are the subject of important theoretical papers. See Dixit & Stiglitz, *supra* note 1; Spence, *Product Differentiation*, *supra* note 1; Spence, *Product Selection*, *supra* note 1.

diverted from other products.¹⁴ For illustration, consider an additional identical product. It imposes its fixed cost on society, but adds no consumer benefit (except, possibly, reduced prices). It is possible, as a result, for markets to support inefficient overprovision of products with sufficient total demand to cover the costs of multiple products. Spence terms the process by which the market determines what to produce the “product selection” problem.¹⁵

Some products that the market selects not to produce are candidates for the “inefficient under-provision” designation. Indeed, one can view the ODA as an attempt to remedy inefficient under-provision. In this case, the reason the allocation may be inefficient is presumably the inability to price discriminate.

We envision firms introducing competing products as long as it is profitable to do so. Competing products are imperfect substitutes for one another. Different products in a category work better for different patients, so that additional products in a category may draw additional persons to consumption. A sufficient, although not necessary, condition for additional products to increase welfare is that additional products raise the tendency for patients to consume a drug in the category corresponding to their condition. We assume that drug development carries only fixed (sunk) costs.¹⁶ The presence of more products creates greater potential for consumers to find a product closer to their ideal. Unless pricing extracts all surplus, consumer welfare is greater.¹⁷

The “business stealing vs. market expansion” distinction provides a helpful framework for viewing the relationship between consumption and welfare.¹⁸ If a new drug is substantially differentiated, it may draw new customers into the market rather than simply diverting business away from existing products. In this case, the share of affected people

14. Lichtenberg and Philipson provide evidence that the present discounted value of a pharmaceutical innovator’s returns is reduced more by competition from other brands (“creative destruction”) prior to patent expiration than it is by competition from generic manufacturers after patent expiration. Frank R. Lichtenberg & Tomas J. Philipson, *The Dual Effects of Intellectual Property Regulations: Within- and Between-Patent Competition in the U.S. Pharmaceuticals Industry*, 45 J.L. & ECON. 643 (2002).

15. Spence, *Product Differentiation*, *supra* note 1.

16. Drugs cost a great deal to develop, but this development cost does not vary with the number of doses produced or sold. Hence the costs are fixed and not variable.

17. We recognize that a higher tendency to consume in a cross section does not necessarily reflect higher welfare. Welfare is not higher if 80 percent of people are barely willing to consume than if 79 percent of persons consume and derive substantial surplus. On the other hand, if the arrival of a new product—without withdrawal of existing products—raises the tendency to consume, then by revealed preference, welfare would be higher. We will treat consumption tendencies as suggestive evidence about welfare in this Article, paying particular attention to results from longitudinal measurement approaches.

18. See Mankiw & Whinston, *supra* note 1.

consuming a drug will increase with entry. On the other hand, an undifferentiated product may draw all of its business from existing products and will therefore not increase the tendency to consume. Of course, additional products can put downward pressure on prices, and this pressure is presumably more acute, as the products are less differentiated.¹⁹

In this scheme, it is easy to see how misery loves company. An increase in market size raises the amount of revenue available to a product category, possibly justifying the development of an additional product. An additional product may attract a new customer (who values the product above its price) whose use of the product generates some combination of consumer surplus and greater longevity. Furthermore, additional products may reduce the price paid by all customers.

The passage of the ODA increases the effective size of the market, relative to fixed costs, for drugs targeting uncommon conditions. This may give rise to more products in those categories, as well as a greater tendency to consume. Because rare conditions are targeted by few products, especially prior to the ODA, new products spurred by the ODA are likely to be strongly differentiated products—that is, their entry provides *some* product, as opposed to *no* product.

The foregoing suggests the following questions: do larger markets attract more products? Is there a greater tendency to consume in markets with either more products or lower prices, or both? Do additional products promote longevity? We now turn to the empirical analysis of these questions.

II. DATA

The basic data for this study are information on disease prevalence, prescription drug consumption, and longevity, by 3-digit ICD-9 disease code, in 1979 and 1998. These observations occur before and fairly long after the 1983 ODA, respectively. In addition, we have information on the fraction of prescriptions written for orphan drugs between 1995 and 2000. Our data are drawn from two sources, which we describe below.

A. Physician Survey Data on Drug Consumption and Condition Prevalence

Our primary data on drug consumption and prevalence are drawn from a physician survey, the National Ambulatory Medical Care Surveys

19. This mechanism has been documented indirectly based on the relationship between market size and entry. See Timothy F. Bresnahan & Peter C. Reiss, *Entry and Competition in Concentrated Markets*, 99 J. POL. ECON. 977 (1991); Timothy F. Bresnahan & Peter C. Reiss, *Entry in Monopoly Markets*, 57 REV. ECON. STUD. 531 (1990).

(“NAMCS”). The NAMCS surveys offer information on patients’ visits to a national sample of office-based physicians. The universe consists of office visits to non-federally employed physicians classified by the American Medical Association (“AMA”) or the American Osteopathic Association (“AOA”) as “office-based, patient care”—excluding specialties of anesthesiology, pathology, and radiology—from 112 Primary Sampling Units (“PSUs”) in the United States.

Each NAMCS office visit record reports the following: the physician’s diagnoses (usually only one), any drugs ordered, administered, or provided, and a sampling weight. We measure condition i ’s prevalence in a year based on the number of visits with primary diagnosis i . In particular, we define:

$N_VISIT_PRE_i$ = the estimated annual number of office-based physician visits in which 3-digit ICD9 diagnosis i was recorded in the pre-ODA period (1980–1981)²⁰; and

$N_VISIT_POST_i$ = the estimated annual number of office-based physician visits in which 3-digit ICD9 diagnosis i was recorded in the post-ODA period (1997–1998).

Thus, the NAMCS-based prevalence measure is based only on physician visits. The advantage of this sampling condition is that physician diagnoses are more likely to be correct than self-diagnoses. At the same time, this sampling has the possible disadvantage of excluding persons who are ill but who do not seek medical care.

We measure drug consumption tendencies from prescription information in the NAMCS in two ways. Our first measure is whether patients diagnosed with a condition have one or more drugs prescribed for them. The “consumption” measure is therefore based not literally on consumption but rather on whether the physician believes beneficial drugs exist for the individual’s circumstance. In particular, the fraction of visits with primary diagnosis i in which one or more drugs were prescribed were defined as follows:

$RX\%_PRE_i$ = visits in which any medications were prescribed as a fraction of total visits in which 3-digit ICD9 diagnosis i was recorded in the pre-ODA period (1980–1981); and

$RX\%_POST_i$ = visits in which any medications were prescribed as a fraction of total visits in which 3-digit ICD9 diagnosis i was recorded in the post-ODA period (1997–1998).

20. NAMCS was conducted in 1980, 1981, 1985, and annually since 1989.

Our second consumption measure is the average number of drugs prescribed per visit, by condition.

B. Longevity and Prevalence Data from Vital Statistics

Our data on longevity, as well as a second measure of prevalence, are drawn from Vital Statistics-Multiple Cause of Death files.²¹ Two items that are recorded on death certificates are the *cause of death*, and the *age at death*. The number of non-infant deaths due to a condition is our second measure of prevalence. We measure longevity using the percentage of non-infant deaths occurring before age 55 due to a condition.

C. Orphan Drug Use

The third piece of information for this study is a measure of orphan drug use. We use the percentage of 1995–2000 prescriptions, by 3-digit ICD-9 condition, that are for the orphan drugs listed in Appendix Table 1. These data are drawn from NAMCS.

Table 1 provides summary statistics on prevalence and consumption measures from the NAMCS survey and the Vital Statistics mortality data. We restrict attention to the 479 3-digit ICD-9 codes for which all variables are available. The first two columns report unweighted means, while the latter two columns report means weighted by contemporaneous MD visit prevalence measures. All measures of prevalence increase over time as do all of the measures of drug consumption. The share of deaths occurring among the young declines over time in both weighted and unweighted measures. We do not observe orphan drug use for the early period, although we can safely assume it is close to zero. By contrast, roughly 5 percent of prescriptions written between 1995 and 2000 are for orphan drugs. The median is 3.3 percent, while the 25th and 75th percentiles of the distribution are 0.3 percent and 8.1 percent, respectively.

III. EMPIRICAL STRATEGY AND RESULTS

A. Empirical Strategy

Our goal in this Article is to measure the effect of market size on consumer welfare in drug markets, and we employ two empirical strategies. First, we exploit cross-sectional comparisons across conditions with different levels of prevalence—i.e., “market size”—asking whether

21. Nat'l Bureau of Econ. Research, Mortality Data—Vital Statistics NCHS's Multiple Cause of Death Data, 1959–2005 (Mar. 6, 2008), <http://www.nber.org/data/vital-statistics-mortality-data-multiple-cause-of-death.html>.

physicians are more likely to prescribe drugs to patients with more prevalent conditions. The inherent difficulty with this approach, however, is the possibility of unobserved heterogeneity leading both to large markets and many drugs.

Fortunately, the passage of the Orphan Drug Act provides a source of exogenous variation in market size for drugs targeting small populations. Using panel data at two points in time, together with a measure of orphan drug use, we can exploit this policy change to provide more compelling evidence of the effects of market size on consumption and mortality than one might find using cross-sectional comparisons across medical conditions alone. As a useful byproduct of this approach, we can also simply examine the effectiveness of the Orphan Drug Act.

B. *Prevalence and Consumption Using Physician Survey Data*

Do persons with more common conditions have a greater tendency to take a drug? First, we estimate cross-condition relationships between the tendency to take a drug and condition prevalence, via the following equations:

$$RX\%_PRE_i = \alpha_0 + \beta_0 \ln(N_VISIT_PRE_i) + \varepsilon_{i0} \quad (1)$$

$$RX\%_POST_i = \alpha_1 + \beta_1 \ln(N_VISIT_POST_i) + \varepsilon_{i1} \quad (2)$$

where equation (1) characterizes the pre-ODA period and equation (2) characterizes this relationship in the post-ODA period. We recognize that these are very parsimonious specifications of what are, essentially, demand equations. It would be natural to also include drug price as an explanatory variable. We experimented with a number of price measures and found little sensitivity of consumption to prices, perhaps owing to the role of insurance in financing prescription drug expenditures. We also include specifications using the mortality-based prevalence measure, as well as both measures. All regressions are weighted by the contemporaneous *MD visit prevalence* measure.

The estimates are reported in Appendix Table 2. Consistent with our expectations, probability of drug use is higher for more prevalent diseases both before and after enactment of the ODA in all specifications. That is, misery loves company in the sense that persons with more prevalent conditions are more likely to find a suitable prescription drug. The dependence of drug use on the *MD visit prevalence* measure declines after the enactment of the ODA. Using the pre-ODA estimates in the first column, 45 percent of persons with a condition in the 25th percentile of prevalence would take a drug, compared with 62 percent of persons with a condition in the 75th prevalence percentile. By contrast, the column 4

post-ODA estimates indicate that 45 percent of persons in the 25th prevalence percentile receive a prescription, compared with 59 percent in the 75th percentile.

It is possible that the cross sectional relationship between consumption and prevalence arises because of unobserved heterogeneity. Some factors determining consumption may be correlated with prevalence for reasons outside our explanation.²² Because we have consumption data at two points in time, we can eliminate the fixed unobservable by differencing. We can then test whether the *change* in consumption is larger for the conditions for which patients take orphan drugs. It is also possible that drug consumption is growing at different rates for conditions with different levels of prevalence. To avoid attributing a general prevalence effect to orphan drug use, we also include measures of condition prevalence in the regressions as follows:

$$\Delta RX\%_i = \alpha_{\Delta} + \beta_{\Delta} \ln(N_VISIT_PRE_i) + \gamma_{\Delta} (\%orphan_i) + \varepsilon_{i,\Delta} \quad (3)$$

where

$$\Delta RX\%_i = RX\%_POST_i - RX\%_PRE_i$$

Finally, we include regressions controlling for prevalence with dummies for prevalence quintiles according to 1979 MD visits. Appendix Table 3 reports results of these regressions. The change in the tendency to have a drug prescribed bears a positive and generally significant relationship to the orphan drug measure. This relationship survives the inclusion of controls for condition prevalence. The range of point estimates falls between 0.16 and 0.3. Between 1979 and 1998, the tendency to have some drug prescribed increases from 71 percent to 73 percent. At the mean level of orphan drug use, approximately 5 percent, orphan drugs raise the tendency to have a drug by between 0.8 percent and 1.5 percent points beyond the baseline increase with time. At the 90th percentile of orphan drug use (16 percent), the effect is between 2.5 percent and 5 percent points. The orphan effects on whether one has a drug prescribed are large relative to the overall increase in this measure.

Appendix Tables 4 and 5 revisit the relationships in Tables 2 and 3 with a different measure of consumption: the number of drugs prescribed rather than the tendency to have any prescriptions. In Appendix Table 4, as in Appendix Table 2, drug use increases in market size. Using this measure of consumption, misery loves company. In contrast with the

22. For example, the measure of prevalence used in these regressions, doctor visits where a condition is diagnosed, may be driven by the known availability of particular medications, e.g., Viagra. Other measurement approaches we employ, including both longitudinal data and mortality-based prevalence measures, avoid these problems. Death is not endogenous in the same way as doctor visits.

consumption results based on the share consuming, here, the dependence of consumption on market size is higher after the ODA than before.

Similarly, in Appendix Table 5, as in Appendix Table 3, the number of drugs prescribed increases more quickly as our orphan drug measure is higher. The mean number of drugs taken increases by 0.26 between 1979 and 1998 (see Appendix Table 1). At the mean of orphan drug use (0.05), the orphan effect adds between 0.02 and 0.05 to the baseline increase in drugs taken; at the 90th percentile of orphan drug use, the effect is between 0.06 and 0.17. In contrast with results on whether one takes a drug, the orphan effects on number of drugs taken are small.

C. Mortality and Prevalence

Although product consumption is the usual economic measure underlying welfare inferences, the medical context provides other intuitive measures of welfare. We can use mortality data to examine the relationship between prevalence and longevity, as measured by the percent of persons, among those dying of a condition, dying before age 55. Our empirical approaches are analogous to those above.

Appendix Table 6 shows cross sectional regressions of our longevity measure on the prevalence measures. In all cases, conditions that are more prevalent have lower fractions of their deaths occurring young. A disease at the 25th prevalence percentile (by *MD visits*) in 1979 has 21 percent of its deaths occurring young, compared with 13 percent dying young for conditions in the 75th prevalence percentile. In 1998, the percentage of deaths occurring young for conditions at the 25th prevalence percentile had fallen by 6 percentage points to 16 percent, while the percentage of deaths occurring young for more common conditions—in the 75th percentile—had fallen only two percentage points, to 11 percent. Both before and after the ODA, misery loves company in the sense that more prevalent conditions have greater longevity. Furthermore, the dependence of longevity on prevalence declines following the ODA.

Finally, Appendix Table 7 shows that the change over time in longevity is larger—the percent dying young declines more—for conditions with more orphan drug consumption.²³ Overall, the percentage dying before the age of fifty five falls by 6.7 percentage points (from 25.6 percent to 18.9 percent). For a disease with the mean orphan drug use, the additional orphan-related decline is 0.2–0.4 percentage points; for a con-

23. This is consistent with evidence that, in general, medical conditions with greater increases in the number of drugs available exhibit greater increases in longevity—as measured by the mean age at death. See Frank Lichtenberg, *Pharmaceutical Innovation, Mortality Reduction, and Economic Growth*, in *MEASURING THE GAINS FROM MEDICAL RESEARCH: AN ECONOMIC APPROACH* 74 (Kevin M. Murphy & Robert H. Topel eds., 2003).

dition at the 90th percentile of orphan drug use, the additional decline is 0.8–1.3 percentage points.

D. Discussion: *The ODA's Effects and Context*

The effects of the ODA are visible in a variety of ways in our results. Prior to the ODA, drug availability—and ensuing welfare—were more sensitive to market size. We see this primarily in the contrast between the pre- and post-ODA estimates of the relationship between share consuming and prevalence. The ODA increased the incentive for firms to develop drugs for small populations, relative to the incentive for larger populations. As a result, there was a sharper growth in the drug consumption tendency in low-prevalence conditions than in more common conditions. Similarly, there was a large decrease in mortality for low-prevalence conditions relative to higher-prevalence conditions. The ODA decreased the extent to which misery loves company. It is not clear whether these effects are efficient, although if the Act simply allows more complete appropriation of drug benefits, then there would be no reason to suspect inefficiency.

Most observers of the ODA applaud this policy precisely for its effect of reducing the dependence of welfare on market size. Intuitively, in the context of disease, it is not hard to understand the popularity of this policy. Yet, the conditions facing would-be consumers of drugs for unpopular conditions are not unique to pharmaceutical markets. These conditions arise, generically, whenever there are large fixed costs and preferences that differ across consumers.

The process by which markets select which products to make causes markets to deliver more welfare to persons with common preferences than to persons with uncommon ones. As Spence has emphasized, there is no reason to expect the market to select the right mix of products in contexts of this sort.²⁴ As we consider the sense of the ODA, we might also ask whether other policies aimed at raising the welfare of small consumer groups are also justified.

Some commentators believe that investment is not too sensitive to incentives (e.g., patent enforcement, price controls).²⁵ They doubt that weakening patent protection or imposing price controls would significantly reduce investment in new drug development.²⁶ Our evidence supports the hypothesis that at least one type of incentive, the extent of

24. Spence, *Product Differentiation*, *supra* note 1.

25. See, e.g., Public Citizen, *Would Lower Prescription Drug Prices Curb Drug Company Research & Development?*, http://www.citizen.org/congress/reform/drug_industry/r_d/articles.cfm?ID=7909 (last visited Feb. 23, 2009).

26. *Id.*

the market, has an important effect on the amount of investment. It may shed light on the effect of changes in other incentives on investment. For example, a government-mandated 25 percent price reduction may have a similar effect on investment as a (“market-mandated”) 25 percent reduction in prevalence.

CONCLUSION

The results show two things. First, the results show that in this market, as in some others, supply-side non-convexities give rise to an important relationship between market size and consumption and, arguably, welfare. In this context, misery loves company. This has broad implications. First, market size matters in providing incentives for product development.

Second, the results show that the prevailing, and generally implicit, view is that market allocation, unlike allocation through collective choice, gives each consumer whatever she wants, regardless of her fellow consumers’ preferences. Given the large drug development costs, however, consumers see drugs developed for their conditions only as they make up large potential markets. Our results are, frankly, not surprising; but they do provide some evidence about how the mix of differentiated products selected in a market depends on the distribution of product-preferring types in the market.

Third, the results show that the Orphan Drug Act “works,” in the sense that it has induced increased development of drugs targeted at small populations and that these populations are now more likely to take drugs. The policy has had the effect of reducing disparities in wellbeing between large and small populations. Other government policies also promote this objective. Perhaps most notably, the U.S. Postal Service has an explicit policy of charging the same rates for postage regardless of letter origin or destination within the United States. If mail pricing were left entirely to the market, postage rates would presumably be lower for letters sent to and from densely populated areas. Under government provision, by contrast, administered rates are the same for consumers with substantially different costs of service, in densely and sparsely populated areas.

It is becoming increasingly clear that in large-fixed cost contexts where preferences differ across individuals, markets deliver fewer products and perhaps less satisfaction to small groups. In the pharmaceutical market, this is deemed a bad feature of market outcomes; and policies have been devised to remedy the situation. Yet, there is no clear distinction between the economic circumstances of pharmaceutical markets and

other large-fixed-cost markets. How widely such a policy rationale should be applied is remains an important question for policymakers.

APPENDIX

TABLE 1: SUMMARY STATISTICS

	Unweighted		Weighted	
	Pre-ODA (1978)	Post-ODA (1998)	Pre-ODA (1978)	Post-ODA (1998)
MD Visits (mil)	0.76	1.04	5.69	7.16
Deaths	3601.34	4478.90	8551.9	11087.7
% Dying Young	31.23%	27.60%	25.59%	18.87%
Share w/ Rx	0.59	0.64	0.71	0.73
Mean Rxs	1.16	1.51	1.40	1.66
Orphan % of Rx (1995–2000)		5.94%		4.45%
N (3-digit ICD-9 codes)	479	479	479	479

Notes: weighted means are weighted by contemporaneous MD visits.

TABLE 2: SHARE CONSUMING A DRUG AND PREVALENCE

	(1)	(2)	(3)	(4)	(5)	(6)
	Share Getting Rx Pre	Share Getting Rx Pre	Share Getting Rx Pre	Share Getting Rx Post	Share Getting Rx Post	Share Getting Rx Post
Log MD Visits (pre)	0.0657		0.0638			
	(0.0063)**		(0.0060)**			
Log Deaths 1979		0.0176	0.0079			
		(0.0033)**	(0.0030)**			
Log MD Visits (post)				0.0573		0.0537
				(0.0051)**		(0.0052)**
Log Deaths 1998					0.0127	0.0083
					(0.0024)**	(0.0022)**
Constant	-0.2570	0.8630	-0.1591	-0.1406	0.8288	-0.0182
	(0.0940)**	(0.0296)**	(0.0971)	(0.0782)	(0.0212)**	(0.0839)
Observations	479	479	479	479	479	479
R-squared	0.18	0.06	0.22	0.21	0.05	0.23

Notes: Standard errors in parentheses. * significant at 5%; ** significant at 1%.
All regressions weighted by contemporaneous MD visits.

TABLE 3
CHANGE IN SHARE CONSUMING A DRUG, PREVALENCE,
AND ORPHAN DRUG USE

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	<i>Change in Share Consuming Rx</i>	<i>Change in Share Consuming Rx</i>	<i>Change in Share Consuming Rx</i>	<i>Change in Share Consuming Rx</i>	<i>Change in Share Consuming Rx</i>	<i>Change in Share Consuming Rx</i>	<i>Change in Share Consuming Rx</i>	<i>Change in Share Consuming Rx</i>
Log MD Visits (pre)	-0.0172			-0.0156		-0.0158		
	(0.0034)**			(0.0036)**		(0.0037)**		
Log Deaths 1979		-0.0008			-0.0016	0.0004		
		(0.0017)			(0.0017)	(0.0018)		
Orphan % of Rxs			0.3095	0.1692	0.3247	0.1637		0.2534
			(0.1065)**	(0.1094)	(0.1078)**	(0.1124)		(0.1085)*
Prevalence Quintile 2							-0.0487	-0.0538
							(0.0769)	(0.0765)
Quintile 3							-0.0004	-0.0019
							(0.0720)	(0.0717)
Quintile 4							-0.0417	-0.0418
							(0.0703)	(0.0700)
Quintile 5							-0.0664	-0.0613
							(0.0693)	(0.0690)
Constant	0.2712	0.0117	0.0047	0.2398	-0.0097	0.2464	0.0769	0.0619
	(0.0504)**	(0.0156)	(0.0069)	(0.0543)**	(0.0171)	(0.0624)**	(0.0690)	(0.0690)
Observations	479	479	479	479	479	479	479	479
R-squared	0.05	0.00	0.02	0.06	0.02	0.04	0.01	0.03

Notes: Standard errors in parentheses. * significant at 5%; ** significant at 1%.
 All regressions weighted by 1998 MD visits.

TABLE 4
MEAN DRUGS CONSUMED AND PREVALENCE

	(1)	(2)	(3)	(4)	(5)	(6)
	<i>Mean # Rx Pre</i>	<i>Mean # Rx Pre</i>	<i>Mean # Rx Pre</i>	<i>Mean # Rx Post</i>	<i>Mean # Rx Post</i>	<i>Mean # Rx Post</i>
Log MD Visits (pre)	0.1390		0.1138			
	(0.0177)**		(0.0155)**			
Log Deaths 1979		0.1034	0.0798			
		(0.0079)**	(0.0077)**			
Log MD Visits (post)				0.1502		0.0991
				(0.0201)**		(0.0163)**
Log Deaths 1998					0.1248	0.1167
					(0.0072)**	(0.0071)**
Constant	-0.6560	2.2681	0.3959	-0.6139	2.6730	1.1090
	(0.2628)*	(0.0702)**	(0.2502)	(0.3051)*	(0.0625)**	(0.2646)**
Observations	479	479	479	479	479	479
R-squared	0.11	0.26	0.30	0.10	0.39	0.43

Notes: Standard errors in parentheses. * significant at 5%; ** significant at 1%.
All regressions weighted by contemporaneous MD visits.

TABLE 5
CHANGE IN MEAN DRUGS, PREVALANCE, AND ORPHAN DRUG USE

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Change in Mean # Rx	Change in Mean # Rx	Change in Mean # Rx	Change in Mean # Rx	Change in Mean # Rx	Change in Mean # Rx	Change in Mean # Rx	Change in Mean # Rx
Log MD Visits (pre)	-0.0160			-0.0062		-0.0286		
	(0.0115)			(0.0120)		(0.0118)*		
Log Deaths 1979		0.0371			0.0371	0.0407		
		(0.0055)**			(0.0055)**	(0.0056)**		
Orphan % of Rxs		0.7250	1.0773	1.0218	0.7250	0.4340		1.0772
		(0.3379)*	(0.3493)**	(0.3658)**	(0.3379)*	(0.3570)		(0.3587)**
Prevalence Quintile 2							-0.1216	-0.1434
							(0.2550)	(0.2530)
Quintile 3							-0.0810	-0.0873
							(0.2389)	(0.2369)
Quintile 4							-0.1962	-0.1967
							(0.2333)	(0.2314)
Quintile 5							-0.1726	-0.1512
							(0.2298)	(0.2280)
Constant	0.5102	0.5601	0.2277	0.3206	0.5601	1.0227	0.4441	0.3804
	(0.1697)**	(0.0535)**	(0.0225)**	(0.1817)	(0.0535)**	(0.1981)**	(0.2290)	(0.2281)
Observations	479	479	479	479	479	479	479	479
R-squared	0.00	0.11	0.02	0.02	0.11	0.12	0.01	0.02

Notes: Standard errors in parentheses. * significant at 5%; ** significant at 1%. All regressions weighted by 1998 MD visits in the condition.

TABLE 6
PERCENT DYING YOUNG AND PREVALENCE

	(1)	(2)	(3)	(4)	(5)	(6)
	% Dying Young, 1979	% Dying Young, 1979	% Dying Young, 1979	% Dying Young, 1998	% Dying Young, 1998	% Dying Young, 1998
Log MD Visits (pre)	-0.0324		-0.0143			
	(0.0037)**		(0.0043)**			
Log Deaths 1979		-0.0353	-0.0278			
		(0.0032)**	(0.0039)**			
Log MD Visits (post)				-0.0218		-0.0069
				(0.0032)**		(0.0035)*
Log Deaths 1998					-0.0339	-0.0301
					(0.0031)**	(0.0037)**
Constant	0.5601	0.0116	0.2270	0.4017	-0.0058	0.1003
	(0.0497)**	(0.0119)	(0.0665)**	(0.0432)**	(0.0115)	(0.0547)
Observations	479	479	479	479	479	479
R-squared	0.14	0.20	0.22	0.09	0.20	0.20

Notes: Standard errors in parentheses. * significant at 5%; ** significant at 1%. All regressions weighted by contemporaneous deaths in the condition.

TABLE 7
CHANGE IN PERCENT DYING YOUNG, PREVALENCE,
AND ORPHAN DRUG CONSUMPTION

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Change in % Dying Young	Change in % Dying Young	Change in % Dying Young	Change in % Dying Young	Change in % Dying Young	Change in % Dying Young	Change in % Dying Young	Change in % Dying Young
Log MD Visits (pre)	0.0068		0.0067		0.0030	0.0029		
	(0.0013)**		(0.0013)**		(0.0015)*	(0.0015)		
Log Deaths 1979		0.0078		0.0078	0.0063	0.0063		
		(0.0011)**		(0.0011)**	(0.0014)**	(0.0014)**		
Orphan % of Rxs			-0.0782	-0.0824		-0.0793		-0.0515
			(0.0373)*	(0.0366)*		(0.0365)*		(0.0376)
Prevalence Quintile 2							0.0081	0.0098
							(0.0113)	(0.0113)
Quintile 3							-0.0228	-0.0197
							(0.0105)*	(0.0107)
Quintile 4							0.0114	0.0135
							(0.0094)	(0.0095)
Quintile 5							0.0256	0.0270
							(0.0094)**	(0.0094)**
Constant	-0.1152	0.0051	-0.1076	0.0110	-0.0409	-0.0330	-0.0351	-0.0330
	(0.0170)**	(0.0047)	(0.0173)**	(0.0054)*	(0.0232)	(0.0234)	(0.0087)**	(0.0089)**
Observations	479	479	479	479	479	479	479	479
R-squared	0.06	0.09	0.07	0.10	0.10	0.11	0.10	0.11

Notes: Standard errors in parentheses. * significant at 5%; ** significant at 1%. All regressions weighted by the 1998 deaths in the condition.

TABLE 8
 DRUGS THAT HAVE BEEN GRANTED ORPHAN DRUG
 STATUS BY THE FDA²⁷

Albendazole	Etanercept	Naltrexone Hydrochloride
Aldesleukin	Ethanolamine Oleate	Nitric Oxide
Alglucerase	Etidronate Disodium	Octreotide Acetate
Alitretinoin	Exemestane	Ofloxacin
Allopurinol	Factor IX (Human)	Oprelvekin
Altretamine	Felbamate	Paclitaxel
Amifostine	Filgrastim	Pegademase Bovine
Aminosalicic Acid	Fludarabine Phosphate	Pegaspargase
Amiodarone Hydrochloride	Follitropin Alfa	Pentamidine Isethionate
Amphotericin B	Fomepizole	Pentastarch
Amphotericin B Lipid Complex	Fosphenytoin Sodium	Pentosan Polysulfate Sodium
Anagrelide Hydrochloride	Ganciclovir Sodium	Pentostatin
Antihemophilic Factor (Recombinant)	Gemtuzumab Ozogamicin	Pilocarpine
Antithrombin III (Human)	Glatiramer Acetate	Poractant Alfa
Aprotinin Bovine	Gonadorelin Acetate	Porfimer Sodium
Atovaquone	Halofantrine Hydrochloride	Potassium Citrate
Baclofen	Hemin	Proteinase Inhibitor (Human), Alpha 1
Basiliximab	Histrelin Acetate	Respiratory Syncytial Virus Immune Globulin
Beractant	Hydroxyurea	Rho (D) Immune Globulin
Betaine, Anhydrous	Idarubicin Hydrochloride	Rifabutin
Bexarotene	Ifosfamide	Rifampin
Bleomycin Sulfate	Imiglucerase	Rifapentine
Botulinum Toxin	Immune Globulin (Human)	Riluzole
Busulfan	Infliximab	Rituximab
Caffeine Citrate	Interferon Alfa-2a, Recombinant	Sacrosidase
Calcium Acetate	Interferon Alfa-2b, Recombinant	Sargramostim
Calfactant	Interferon beta-1a	Satumomab Pendetide
Cetyl Alcohol; Colfosceril Palmitate; Tyloxapol	Interferon Beta-1b, Recombinant	Selegiline Hydrochloride
Chenodiol	Interferon Gamma-1b, Recombinant	Sermorelin Acetate
Citric Acid; Glucono-Delta-Lactone; Magnesium Carbonate	Iobenguane Sulfate I 131	Sodium Benzoate; Sodium Phenylacetate
Cladribine	Isoniazid; Pyrazinamide; Rifampin	Somatrem
Clofazimine	Lamotrigine	Somatropin, Biosynthetic
Clonidine Hydrochloride	Lepirudin (rdNA)	Sotalol Hydrochloride
Coagulation Factor VIIa (Recombinant)	Leucovorin Calcium	Succimer

27. Data obtained from Mosby, Mosby's Drug Consult for Health Professions (2002).

Corticotrelin Ovine Triflutate	Leuprolide Acetate	Sulfadiazine
Cromolyn Sodium	Levocarnitine	Temozolomide
Cysteamine Bitartrate	Levomethadyl Acetate Hydrochloride	Teniposide
Cytarabine Liposome	Lidocaine	Teriparatide Acetate
Cytomegalovirus Immune Globulin	Liothyronine Sodium	Thalidomide
Daclizumab	Lodoxamide Tromethamine	Thyrotropin
Daunorubicin Citrate Liposome	Mafenide Acetate	Tiopronin
Denileukin Diftitox	Mefloquine Hydrochloride	Tobramycin
Desmopressin Acetate	Megestrol Acetate	Toremifene Citrate
Dexrazoxane Hydrochloride	Melphalan	Tretinoin
Diazepam	Mesna	Trientine Hydrochloride
Digoxin Immune Fab (Ovine)	Methotrexate Sodium	Trimetrexate Glucuronate
Dornase Alfa	Metronidazole	Urofollitropin
Doxorubicin, Liposomal	Midodrine Hydrochloride	Ursodiol
Dronabinol	Mitoxantrone Hydrochloride	Valrubicin
Eflornithine Hydrochloride	Modafinil	Zalcitabine
Epirubicin Hydrochloride	Monoctanoin	Zidovudine
Epoetin Alfa	Morphine Sulfate	Zinc Acetate
Epoprostenol Sodium	Nafarelin Acetate	