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First model-independent Dalitz analysis of $B^0 \rightarrow DK^{*0}, D \rightarrow K^0_S \pi^+ \pi^-$ decay

Belle Collaboration

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We report a measurement of the amplitude ratio r_S of $B^0 \to D^0 K^{*0}$ and $B^0 \to \overline{D^0} K^{*0}$ decays with a Dalitz analysis of $D \to K_S^0 \pi^+ \pi^-$ decays, for the first time using a modelindependent method. We set an upper limit $r_S < 0.87$ at the 68% confidence level, using the full data sample of 711 fb⁻¹ corresponding to 772 × 10⁶ $B\bar{B}$ pairs collected at the $\Upsilon(4S)$ resonance with the Belle detector at the KEKB e^+e^- collider. This result is obtained from observables $x_- = +0.4^{+1.0+0.0}_{-0.6-0.1} \pm 0.0$, $y_- = -0.6^{+0.8+0.1}_{-1.0-0.0} \pm 0.1$, $x_+ = +0.1^{+0.7+0.0}_{-0.4-0.1} \pm 0.1$, and $y_+ = +0.3^{+0.5+0.0}_{-0.8-0.1} \pm 0.1$, where $x_{\pm} = r_S \cos(\delta_S \pm \phi_3)$, $y_{\pm} = r_S \sin(\delta_S \pm \phi_3)$, and $\phi_3(\delta_S)$ is the weak (strong) phase difference between $B^0 \to D^0 K^{*0}$ and $B^0 \to \overline{D^0} K^{*0}$.

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Subject Index C02, C03

1. Introduction

Determination of parameters of the standard model (SM) plays an important role in the search for new physics. In the SM, the Cabibbo–Kobayashi–Maskawa (CKM) matrix [1,2] gives a successful description of all current measurements of *CP* violation. The *CP*-violating parameters ϕ_1 , ϕ_2 , and ϕ_3 are the three angles of the most equilateral of the CKM unitarity triangles, of which $\phi_3 \equiv \arg \left(-V_{ud}V_{ub}^*/V_{cd}V_{cb}^*\right)$ is the least accurately determined. In the usual quark-phase convention, where the complex phase is negligible in the CKM matrix elements other than V_{ub} and V_{td} [3], the measurement of ϕ_3 is equivalent to the extraction of the phase of V_{ub} . To date, ϕ_3 measurements have been performed mainly with *B* meson decays into $D^{(*)}K^{(*)}$ final states [4–13], all of which exploit the interference between the $\bar{D}^{(*)0}$ and $D^{(*)0}$ decaying into a common final state. In particular, Dalitz analyses of $B^{\pm} \rightarrow D^{(*)}K^{(*)\pm}$, $D \rightarrow K_S^0\pi^- + \pi^-$ provide the most precise determination of ϕ_3 . The Dalitz analysis technique for the measurement of ϕ_3 was proposed in Ref. [14]. Belle reported the first ϕ_3 measurement with the model-independent Dalitz analysis technique in Ref. [15], which exploits a set of measured strong phases instead of relying on a *D* decay model into a three-body final state.

In this paper, we present the first measurement of the amplitude ratio of $B^0 \to D^0 K^{*0}$ and $B^0 \to D^0 K^{*0}$ decays with a model-independent Dalitz analysis. We reconstruct $B^0 \to DK^{*0}$, with $K^{*0} \to K^+\pi^-$ (throughout the paper, charge-conjugate processes are implied; K^{*0} refers to $K^*(892)^0$ and D refers to either D^0 or \overline{D}^0 when the D^0 flavor is untagged). Here, the flavor of the B meson is identified by the kaon charge. Neutral D mesons are reconstructed in the $K_S^0\pi^+\pi^-$ decay mode. The reconstructed final states are accessible through $b \to c$ and $b \to u$ processes via the diagrams shown in Fig. 1.

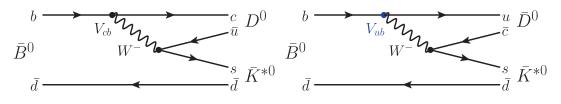


Fig. 1. Diagrams for the $\bar{B}^0 \rightarrow D\bar{K}^{*0}$ decay.

In this analysis, we use the variables r_S , k, and δ_S to parameterize the strong dynamics of the decay. These parameters are defined as [16]

$$r_S^2 \equiv \frac{\Gamma\left(B^0 \to D^0 K^+ \pi^-\right)}{\Gamma\left(B^0 \to \bar{D}^0 K^+ \pi^-\right)} = \frac{\int dp A_{b\to u}^2(p)}{\int dp A_{b\to c}^2(p)},\tag{1}$$

$$ke^{i\delta_S} \equiv \frac{\int dp A_{b\to c}(p) A_{b\to u}(p) e^{i\delta(p)}}{\sqrt{\int dp A_{b\to c}^2(p) \int dp A_{b\to u}^2(p)}},$$
(2)

where the integration is over the $B^0 \to DK^+\pi^-$ Dalitz distribution region corresponding to the K^{*0} resonance. Here, $A_{b\to c}(A_{b\to u})(p)$ is the magnitude of the amplitude for the $b \to c(u)$ transition and $\delta(p)$ is the relative strong phase, where the variable p indicates the position within the $DK^+\pi^-$ Dalitz distribution. If the B^0 decay can be considered as a DK^{*0} two-body decay, r_S becomes the ratio of the amplitudes for $b \to u$ and $b \to c$ and k becomes 1. According to a simulation study using a Dalitz model based on the measurements in Ref. [17], the value of k is 0.95 ± 0.03 within the phase space of the DK^{*0} resonance. The value of r_S is expected to be around 0.4, which corresponds naïvely to $|V_{ub}V_{cs}^*|/|V_{cb}V_{us}^*|$ but also depends on strong interaction effects. For r_S , the best experimental value is reported by LHCb [18] as $r_S = 0.240^{+0.055}_{-0.048}$ (different from zero by 2.7 σ) from $B^0 \to DK^{*0}$, $D \to K^+K^-$, $\pi^+\pi^-$, $K^{\pm}\pi^{\mp}$ decay.

2. The model-independent Dalitz analysis technique

The amplitude of the $B^0 \to DK^{*0}$, $D \to K_S^0 \pi^+ \pi^-$ decay is a superposition of the $B^0 \to \overline{D}^0 K^{*0}$ and $B^0 \to D^0 K^{*0}$ amplitudes

$$A_B\left(m_{+}^2, m_{-}^2\right) = \bar{A} + r_S e^{i(\delta_S + \phi_3)} A,$$
(3)

where m_{\pm}^2 and m_{-}^2 are the squared invariant masses of the $K_S^0\pi^+$ and $K_S^0\pi^-$ combinations, respectively, $\bar{A} = \bar{A}(m_{\pm}^2, m_{-}^2)$ is the amplitude of the $B^0 \to \bar{D}^0 K^{*0}$, $\bar{D}^0 \to K_S^0 \pi^+ \pi^-$ decay, and $A = A(m_{\pm}^2, m_{-}^2)$ is the amplitude of the $B^0 \to D^0 K^{*0}$, $D^0 \to K_S^0 \pi^+ \pi^-$ decay. In the case of *CP* conservation in the *D* decay, we have $A(m_{\pm}^2, m_{-}^2) = \bar{A}(m_{\pm}^2, m_{\pm}^2)$ as a *CP* transformation changes $\pi^{\pm} \to \pi^{\mp}$, thus $m_{\pm}^2 \to m_{\mp}^2$. The Dalitz distribution density of the *D* decay from $B^0 \to$ DK^{*0} is given by

$$P_B = |A_B|^2 = |\bar{A} + r_S e^{i(\delta_S + \phi_3)} A|^2 = \bar{P} + r_S^2 P + 2k\sqrt{P\bar{P}} (x_+ C + y_+ S),$$
(4)

where $P = P(m_+^2, m_-^2) = |A|^2$, $\bar{P} = \bar{P}(m_+^2, m_-^2) = |\bar{A}|^2$, and

$$x_{+} = r_{S}\cos(\delta_{S} + \phi_{3}), \quad y_{+} = r_{S}\sin(\delta_{S} + \phi_{3}).$$
 (5)

The functions $C(m_+^2, m_-^2)$ and $S(m_+^2, m_-^2)$ are the cosine and sine of the strong-phase difference $\delta_D(m_+^2, m_-^2) = \arg \bar{A} - \arg A$ between the $\bar{D}^0 \to K_S^0 \pi^+ \pi^-$ and $D^0 \to K_S^0 \pi^+ \pi^-$ amplitudes.

Here, we have used the definition of k given in Eq. (2). The equations for the charge-conjugate mode $\bar{B}^0 \to D\bar{K}^{*0}$ are obtained with the substitution $-\phi_3 \to \phi_3$ and $A \leftrightarrow \bar{A}$; the corresponding parameters that depend on the \bar{B}^0 decay amplitude are

$$x_{-} = r_S \cos(\delta_S - \phi_3), \quad y_{-} = r_S \sin(\delta_S - \phi_3).$$
 (6)

If *P*, \bar{P} , *C*, *S*, and *k* are known, one can obtain (x_+, y_+) from B^0 and (x_-, y_-) from \bar{B}^0 decays. Combining both B^0 and \bar{B}^0 measurements, r_S , ϕ_3 , and δ_S can be extracted.

In the model-dependent analysis, one deals directly with the Dalitz distribution density and the functions *C* and *S* are obtained from a model based upon a fit to the $D^0 \rightarrow K_S^0 \pi^+ \pi^-$ amplitude. On the other hand, in the model-independent approach [19,20], where the assumption of a model for $D^0 \rightarrow K_S^0 \pi^+ \pi^-$ decay is not necessary, the Dalitz plot is divided into 2*N* bins symmetric under the exchange $m^2_- \leftrightarrow m^2_+$. The bin index *i* ranges from -N to N (excluding 0); the exchange $m^2_- \leftrightarrow m^2_+$ corresponds to the exchange $i \leftrightarrow -i$. The expected number of signal events in bin *i* of the Dalitz distribution of the *D* mesons from $B^0 \rightarrow DK^{*0}$ is

$$N_{i}^{\pm} = h_{B} \left[K_{\pm i} + r_{S}^{2} K_{\mp i} + 2k \sqrt{K_{i} K_{-i}} \left(x_{\pm} c_{i} \pm y_{\pm} s_{i} \right) \right], \tag{7}$$

where $N^{+(-)}$ stands for the number of $B^0(\bar{B}^0)$ meson decays, $h_B^{+(-)}$ is the normalization constant, and K_{+i} is the number of events in the *i*th bin of flavor-tagged $D^0 \to K_S \pi^+ \pi^-$ decays measured with a sample of inclusively reconstructed $D^{*+} \to D^0 \pi^+$ decays. Equation (7) can be obtained by integrating Eq. (4) over the *i*th bin region. Here, $K_i \propto \int_{\mathcal{D}_i} |A|^2 d\mathcal{D}$, and \mathcal{D} represents the Dalitz plane and \mathcal{D}_i is the bin over which the integration is performed. The values of K_i are measured from a sample of flavor-tagged D^0 mesons obtained by reconstructing $D^{*\pm} \to D\pi^{\pm}$ decays. The terms c_i and s_i are the amplitude-weighted averages of the functions *C* and *S* over the bin:

$$c_{i} = \frac{\int_{\mathcal{D}_{i}} |A| \left| \bar{A} \right| C d\mathcal{D}}{\sqrt{\int_{\mathcal{D}_{i}} |A|^{2} d\mathcal{D} \int_{\mathcal{D}_{i}} \left| \bar{A} \right|^{2} d\mathcal{D}}}.$$
(8)

The terms s_i are defined similarly with *C* substituted by *S*. The absence of *CP* violation in the *D* decay implies $c_i = c_{-i}$ and $s_i = -s_{-i}$. The values of c_i and s_i can be measured using quantum-correlated *D* pairs produced at charm-factory experiments operating at the threshold of $D\bar{D}$ pair production. The CLEO Collaboration has reported c_i and s_i values from *CP*-tagged and flavor-tagged $D\bar{D}$ events data, and this analysis is performed with the optimal binning in Refs. [21,22], as shown in Fig. 2. Given that c_i and s_i are measured and K_i and *k* are known, Eq. (7) has only three free parameters (x, y,and h_B) for each of B^0 and \bar{B}^0 , and can be solved. We use the values of (c_i, s_i) for the "optimal $D^0 \rightarrow K_S^0 \pi^+ \pi^-$ binning" reported in Table XVI of Ref. [22], "the optimal binning" K_i values reported in Table II of Ref. [15], and $k = 0.95 \pm 0.03$ [17]. We have neglected charm-mixing effects in *D* decays from both the $B^0 \rightarrow DK^{*0}$ process and in the quantum-correlated $D\bar{D}$ production [23].

3. Event reconstruction and selection

This analysis is based on a data sample that contains 711 fb⁻¹ corresponding to $772 \times 10^6 B\bar{B}$ pairs, collected with the Belle detector at the KEKB asymmetric-energy e^+e^- (3.5 on 8 GeV) collider [24] operating at the $\Upsilon(4S)$ resonance. The Belle detector is a large-solid-angle magnetic spectrometer that consists of a silicon vertex detector, a 50-layer central drift chamber (CDC), an array of aerogel threshold Cherenkov counters (ACC), a barrel-like arrangement of time-of-flight scintillation

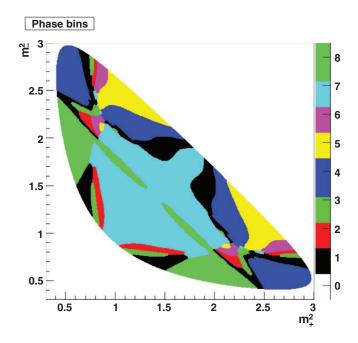


Fig. 2. The optimal binning in Ref. [21,22].

counters (TOF), and an electromagnetic calorimeter comprised of CsI(Tl) crystals located inside a superconducting solenoid coil that provides a 1.5 T magnetic field. An iron flux-return located outside of the coil is instrumented to detect K_L^0 mesons and to identify muons. The detector is described in detail elsewhere [25].

We reconstruct $B^0 \to DK^{*0}$ events with $K^{*0} \to K^+\pi^-$ and $D \to K^0_S \pi^+\pi^-$. The event selection described below is developed from studies of continuum data taken at center-of-mass energies just below the $\Upsilon(4S)$ resonance and Monte Carlo (MC) simulated events.

The K_S^0 candidates are identified using the output of a neural network. Inputs to the network for a pair of oppositely charged pions are the invariant mass, 20 kinematic parameters, and particle identification (PID) information from the ACC, TOF, and the ionization energy loss in the CDC. The K_S^0 selection has a simulated purity of 92.2% and an efficiency of 75.1%. Charged kaon and pion candidates are identified using PID information. The efficiency is 80%–90% and the probability of misidentification is 6%–10%, depending upon the momentum of hadrons and obtained using dedicated data control samples. We reconstruct neutral *D* mesons by combining a K_S^0 candidate with a pair of oppositely charged pion candidates. We require that the invariant mass be within $\pm 15 \text{ MeV}/c^2 (\pm 3 \sigma)$ of the nominal D^0 mass. K^{*0} candidates are reconstructed from $K^+\pi^-$ pairs. We require that the invariant mass be within $\pm 50 \text{ MeV}/c^2$ of the nominal K^{*0} mass. We combine D and K^{*0} candidates to form B^0 mesons. Candidate events are identified by the energy difference $\Delta E \equiv \sum_i E_i - E_b$ and the beam-constrained mass $M_{bc}c^2 \equiv \sqrt{E_b^2 - |c\sum_i \vec{p}_i|^2}$, where E_b is the beam energy and \vec{p}_i and E_i are the momenta and energies, respectively, of the B^0 meson decay products in the e^+e^- center-of-mass (CM) frame. We select events with 5.21 GeV/c^2 < $M_{bc} < 5.29 \text{ GeV}/c^2$ and $-0.10 \text{ GeV} < \Delta E < 0.15 \text{ GeV}$.

Among other *B* decays, the most serious background is from \bar{B}^0 decaying to the same final state as $B^0 \rightarrow DK^{*0}$. To suppress this background, we exclude candidates for which the invariant mass of the $K^{*0}\pi^+$ system is within $\pm 4 \text{ MeV}/c^2$ of the nominal D^+ mass. This criterion leads to negligible contamination from this mode and a relative loss of 0.6% in the signal efficiency.

1	Fisher discriminants based on modified Fox-Wolfram moments.*
2	The angle in the CM frame between the thrust axes of the <i>B</i> decay and that of the remaining particles.
3	The signed difference of the vertices between the <i>B</i> candidate and the remaining charged tracks.
4	The distance of closest approach between the trajectories of the K^* and D candidates.
5	The expected flavor dilution factor described in Ref. [29].
6	The angle θ between the <i>B</i> meson momentum direction and the beam axis in the CM frame.
7	The angle between the D and $\Upsilon(4S)$ directions in the rest frame of the B candidate.
8	The projection of the sphericity vector with the largest eigenvalue onto the e^+e^- beam direction.
9	The angle of the sphericity vector with the largest eigenvalue with respect to that of the remaining particles.
10	The angle of the sphericity vector with the second largest eigenvalue.
11	The angle of the sphericity vector with the smallest eigenvalue.
12	The magnitude of the thrust of the particles not used to reconstruct the signal.

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Table	1.	Variables	used	tor	qq	suppression.

*The Fox–Wolfram moments were introduced in [26]. The Fisher discriminant used by Belle, based on modified Fox–Wolfram moments, is described in Refs. [27,28].

The large combinatorial background of true D^0 and random K^+ and π^- combinations from the $e^+e^- \rightarrow c\bar{c}$ process and other $B\bar{B}$ decays is reduced if D^0 candidates that are a decay product of $D^{*+} \rightarrow D^0\pi^+$ are eliminated. We use the mass difference ΔM between the $[K_S^0\pi^+\pi^-]_D\pi^+$ and $[K_S^0\pi^+\pi^-]_D$ systems for this purpose: if $\Delta M > 0.15 \text{ GeV}/c^2$ for any additional π^+ candidate not used in the *B* candidate reconstruction, the event is retained. This requirement removes 19% of $c\bar{c}$ background and 11% of $B\bar{B}$ background according to MC simulation. The relative loss in signal efficiency is 5.5%.

In the rare case where there are multiple candidates in an event, the candidate with M_{bc} closest to the nominal value is chosen. The relative loss in signal efficiency is 0.8%.

To discriminate signal events from the large combinatorial background dominated by the twojet-like $e^+e^- \rightarrow q\bar{q}$ continuum process, where q indicates u, d, s, or c, a multivariate analysis is performed using the 12 variables introduced in Table 1. To effectively combine these 12 variables, we employ the NeuroBayes neural network package [30]. The NeuroBayes output is denoted as $C_{\rm NB}$ and lies within the range [-1, 1]; events with $C_{\rm NB} \sim 1$ are signal-like and events with $C_{\rm NB} \sim -1$ are $q\bar{q}$ -like. Training of the neural network is performed using signal and $q\bar{q}$ MC samples. The $C_{\rm NB}$ distribution of signal events peaks at $C_{\rm NB} \sim 1$ and is therefore difficult to represent with a simple analytic function. However, the transformed variable

$$C'_{\rm NB} = \ln \frac{C_{\rm NB} - C_{\rm NB, low}}{C_{\rm NB, high} - C_{\rm NB}},\tag{9}$$

where $C_{\text{NB,low}} = -0.6$ and $C_{\text{NB,high}} = 0.9992$, has a distribution that can be modeled by a Gaussian for signal as well as background. The events with $C_{\text{NB}} < C_{\text{NB,low}}$ are rejected; the relative loss in signal efficiency is 7.4%.

4. Analysis procedure

In this section we describe the fit to determine the physics parameters. In Sect. 4.1 we describe the signal and background shape parametrization. In Sect. 4.2 we describe how we correct for the effect of migration and acceptance variations between bins. In Sect. 4.3 the fit to extract the values of (x, y) is described.

4.1. Signal and background parametrization

The number of signal events is obtained by fitting the three-dimensional distribution of variables $M_{\rm bc}$, ΔE , and $C'_{\rm NB}$ using the extended maximum likelihood method. We form three-dimensional probability density functions (PDFs) for each component as the product of one-dimensional PDFs for ΔE , $M_{\rm bc}$, and $C'_{\rm NB}$, since the correlations among the variables are found to be small. The fit region is defined as $\Delta E \in [-0.1, 0.15]$ GeV and $M_{\rm bc} > 5.21$ GeV/ c^2 .

Backgrounds are divided into the following components:

- \circ Continuum background from $q\bar{q}$ events.
- $B\bar{B}$ background, in which the tracks forming the $B^0 \to DK^{*0}$ candidate come from decays of both *B* mesons in the event. The number of possible *B* decay combinations that contribute to this background is large; therefore, both the Dalitz distribution and the distribution of the fit parameters are quite smooth. $B\bar{B}$ backgrounds are further subdivided into two components: events reconstructed with a true $D \to K_S^0 \pi^+ \pi^-$ decay, referred to as $D_{\text{true}} B\bar{B}$ background, and those reconstructed with a combinatorial *D* candidate, referred to as $D_{\text{fake}} B\bar{B}$ background.
- Peaking $B\bar{B}$ background, in which all tracks forming the $B^0 \to DK^{*0}$ candidate arise from the same *B* meson. This background has two types: events with one pion misidentified as a kaon, such as $D^0[\pi^+\pi^-]_{\rho^0}$, and one pion misidentified as a kaon and one pion not reconstructed, such as $D^0[\pi^+\pi^+\pi^-]_{a_1^+}$. The backgrounds come from individual *B* decays and are well separated from the signal.

The ΔE PDFs are parameterized by a double Gaussian for the signal, an exponential function for the $D_{\text{true}} B\bar{B}$ background, an exponential function for the $D_{\text{fake}} B\bar{B}$ background, a linear function for the $q\bar{q}$ background, a double Gaussian for the $\bar{D}^0 \rho^0$ background, and a Gaussian for the $\bar{D}^0 a_1^+$ background. The M_{bc} PDFs are a Gaussian for signal, a Crystal Ball function [31] for the D_{true} $B\bar{B}$ background, an ARGUS function [32] for the $D_{\text{fake}} B\bar{B}$ background, an ARGUS function for the $q\bar{q}$ background, a sum of a Gaussian and ARGUS functions for the $\bar{D}^0 \rho^0$ background, and a Gaussian for the $\bar{D}^0 a_1^+$ background. For each component, the C'_{NB} PDF is the sum of a Gaussian and a bifurcated Gaussian. The shape parameters of the PDFs are fixed from MC samples.

The numbers of events in each bin are free parameters in the fit. This procedure has been justified for background that is either well separated from the signal (such as peaking $B\bar{B}$ background) or is constrained by a much larger number of events than the signal (such as $q\bar{q}$ background). The results of the fit to the full Dalitz plot are shown in Fig. 3. We obtain a total of $44.2^{+13.3}_{-12.1}$ signal events. The statistical significance is 2.8σ relative to the no-signal hypothesis. Simultaneously, we obtain $695.8^{+177.6}_{-175.6}$ for $D_{\text{true}} B\bar{B}$, $1963.2^{+228.1}_{-227.5}$ for $D_{\text{fake}} B\bar{B}$, $11075.7^{+156.6}_{-135.5}$ for $q\bar{q}$, $16.6^{+16.7}_{-13.6}$ for $\bar{D}^0 \rho^0$, and $59.3^{+22.3}_{-20.8}$ for $\bar{D}^0 a_1^+$ background events.

4.2. Corrections to the bin-by-bin yields

There are further effects that must be accounted for before the values of (x, y) can be determined from a binned fit. Equation 7 only holds if there is no migration between bins and the Dalitz acceptance is uniform. Here, we consider the crossfeed for bin-by-bin yields as the migration and the acceptance as the event reconstruction efficiency.

First, we discuss migration due to momentum resolution and flavor misidentification. Momentum resolution leads to migration of events among the bins. In the binned approach, this effect can be corrected in a non-parametric way. The migration can be described by a linear transformation of the

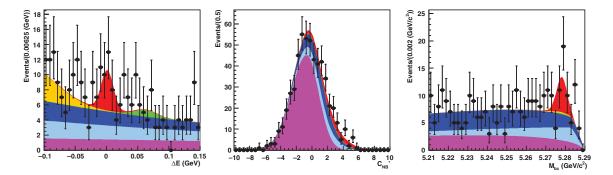


Fig. 3. Projection of the fit to real data using the full Dalitz plot. Left: ΔE distribution with $M_{\rm bc} > 5.27 \,{\rm GeV}/c^2$ and $C'_{\rm NB} > 2$ requirements. Middle: $C'_{\rm NB}$ distribution with $|\Delta E| < 0.03 \,{\rm GeV}$ and $M_{\rm bc} > 5.27 \,{\rm GeV}/c^2$ requirements. Right: $M_{\rm bc}$ distribution with $|\Delta E| < 0.03 \,{\rm GeV}$ and $C'_{\rm NB} > 2$ requirements. Curves show the fitted signal and background contributions (red is signal, yellow is $D^0 a_1^+$, green is $D^0 \rho^0$, blue is $D_{\rm fake} B\bar{B}$, light blue is $D_{\rm true} B\bar{B}$, and magenta is $q\bar{q}$), and points with error bars are the data.

number of events in each bin,

$$N_{\text{obs},i} = \sum \alpha_{ik} N'_k,\tag{10}$$

where N'_i is the number of events that bin *i* would contain without the migration with acceptance and $N_{\text{obs},i}$ is the reconstructed number of events in bin *i*. The migration matrix α_{ik} is nearly the unit matrix; it is obtained from a signal MC simulation generated with the amplitude model reported in Ref. [6]. Most of the off-diagonal elements are null; only a few have values $|\alpha_{ik}| \le 0.04$. In the case of a $D \to K_S^0 \pi^+ \pi^-$ decay from a *B*, the migration depends on the parameters *x* and *y*. However, this is a minor correction to an already small effect and so is neglected.

The second migration effect to be considered is due to misidentification of the *B* flavor. Double misidentification in K^{*0} reconstruction from $K^+\pi^-$ where K^- is misidentified as π^- and π^+ is misidentified as K^+ at the same time leads to migration of events between $N_i^+ \leftrightarrow N_{-i}^-$ due to assignment of the wrong flavor to the *B* candidate. If the fraction of doubly misidentified events is β , the number of events in each bin can be written as

$$N_i^{\prime\pm} = N_{\text{obs},i}^{\pm} + \beta N_{\text{obs},-i}^{\mp}.$$
(11)

The value of β is obtained from MC simulation and is found to be (0.12 ± 0.01) %. Therefore, the effect of flavor misidentification is neglected.

The final-state radiation also causes migration between bins. The measured values of c_i and s_i by CLEO are not corrected for the radiation and the effect upon our analysis is found to be negligible [15].

Second, we consider the effect of the variation of the efficiency profile over the Dalitz plane. We note that Eq. (4) does not change under the transformation $P \rightarrow \epsilon P$ when the efficiency profile $\epsilon(m_+^2, m_-^2)$ is symmetric: $\epsilon(m_+^2, m_-^2) = \epsilon(m_-^2, m_+^2)$. The effect of non-uniform efficiency over the Dalitz plane cancels when using a flavor-tagged *D* sample with kinematic properties that are similar to the sample from the signal *B* decay. This approach allows for the removal of the systematic uncertainty associated with the possible inaccuracy of the detector acceptance description in the MC simulation. With the efficiency taken into account (that is, in general non-uniform across the bin region), the number of events reconstructed is

$$N' = \int p(\mathcal{D})\epsilon(\mathcal{D})d\mathcal{D}.$$
 (12)

Here, p is the probability density on the Dalitz plane and D is the position on the Dalitz plane. Clearly, the efficiency does not factorize. One can use an efficiency averaged over the bin, then correct for it in the analysis:

$$\bar{\epsilon}_i = \frac{N'_i}{N_i} = \frac{\int p(\mathcal{D})\epsilon(\mathcal{D})d\mathcal{D}}{\int p(\mathcal{D})d\mathcal{D}}.$$
(13)

Here, N_i are the number of events corrected for variations in acceptance and migration, which should be used for (x_{\pm}, y_{\pm}) extraction. The averaged efficiency $\bar{\epsilon}_i$ can be determined from MC. The assumption that the efficiency profile depends only on the *D* momentum is tested using MC simulation and the residual difference is treated as a systematic uncertainty. The correction for c_i and s_i due to efficiency variation within a bin cannot be calculated in a completely model-independent way, since the correction terms include the amplitude variation inside the bin. Calculations using the Belle $D \rightarrow K_S^0 \pi^+ \pi^-$ model [6] show that this correction is negligible even for very large non-uniformity of the efficiency profile.

4.3. Fit to determine (x, y)

If N_i in each bin is measured, x_{\pm} and y_{\pm} can be obtained according to Eq. (7) by minimizing

$$-2\log \mathcal{L}(x, y) = -2\sum_{i} \log p(\langle N_i \rangle (x, y), N_i, \sigma_{N_i}), \qquad (14)$$

where $\langle N_i \rangle$ are the expected number of signal events in bin *i* obtained from Eq. (7). Here, N_i and σ_{N_i} are the observed number of events in the data and the uncertainty on N_i , respectively.

The procedure described above does not make any assumptions about the Dalitz distribution of the background events, since the fits in each bin are independent. Thus, there is no uncertainty related to the Dalitz model. However, in our case, where there are a small number of events and many background components, such independent fits are not feasible. Therefore, we obtain (x_{\pm}, y_{\pm}) from a combined fit with a common likelihood for all bins. The relative numbers of background events in each bin are constrained to the numbers found in the MC. The amount of the $D_{\text{true}} B\bar{B}$ background in bins from the ratio of $D^0(K_i)$ and $\bar{D}^0(K_{-i})$ from MC and the amount of the $D_{\text{fake}}, B\bar{B}, q\bar{q}$, and the background from individual B decays from the MC. The yields integrated over the Dalitz plot of the background components are additional free parameters. Thus, the variables (x_{\pm}, y_{\pm}) become free parameters of the combined likelihood fit and the assumption that the signal yield obeys a Gaussian distribution is not needed. While the normalization parameter h_B is also a free parameter of the fit, we do not mention it in the following as it is not a quantity of interest.

5. Combined fits to data

The results of the combined fit in each bin of the B^0 and \overline{B}^0 are shown in Figs. 4 and 5, respectively. The plots show the projections of the data and the fitting model on the ΔE variable, with the additional requirements $M_{\rm bc} > 5.27 \,{\rm GeV}/c^2$ and $C'_{\rm NB} > 2$. The values of the (x, y) parameters and their statistical correlations, obtained from the combined fit for the signal sample, are given in Table 2. In this study, these (x, y) values from the likelihood distribution of the combined fit are corrected using the frequentist approach with Feldman–Cousins ordering [33], which is described in Sect. 7.

6. Systematic uncertainties

The systematic uncertainties of (x, y) are obtained by taking deviation from the default procedure under various assumptions. The systematic uncertainties are summarized in Table 3; most are

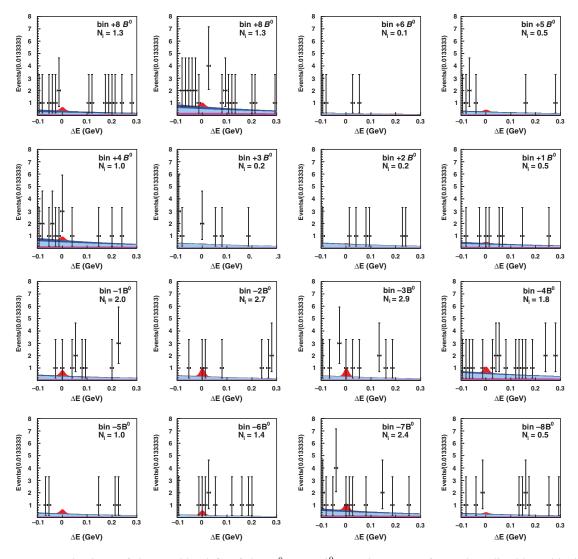


Fig. 4. Projections of the combined fit of the $B^0 \rightarrow DK^{*0}$ sample on ΔE for each Dalitz bin, with the $M_{\rm bc} > 5.27 \,{\rm GeV}/c^2$ and $C'_{\rm NB} > 2$ requirements. The fill styles for the signal and background components are the same as in Fig. 3.

negligible compared to the statistical uncertainty. There is an uncertainty due to the Dalitz efficiency variation because of the difference in average efficiency over each bin for the flavor-tagged D and $B^0 \rightarrow DK^{*0}$ samples. A maximum difference of 1.5% is obtained in an MC study. The uncertainty is taken as the maximum of two quantities:

- the root mean square of x and y from smearing the numbers of events in the flavor-tagged sample K_i by 1.5%, or
- the bias in x and y between the fits with and without efficiency correction for K_i obtained from signal MC.

The uncertainty due to migration of events between bins is estimated by taking the bias between the fits with and without the migration correction. The uncertainties due to the fixed parameterization of the signal and background PDFs are estimated by varying them by $\pm 1 \sigma$. The uncertainty due to the $C'_{\rm NB}$ PDF distributions for $B\bar{B}$ is estimated by replacing them with the signal $C'_{\rm NB}$ PDF. The uncertainty due to the $D_{\rm true}$ and $D_{\rm fake} B\bar{B}$ fractions is estimated by varying them between 0 and 1.

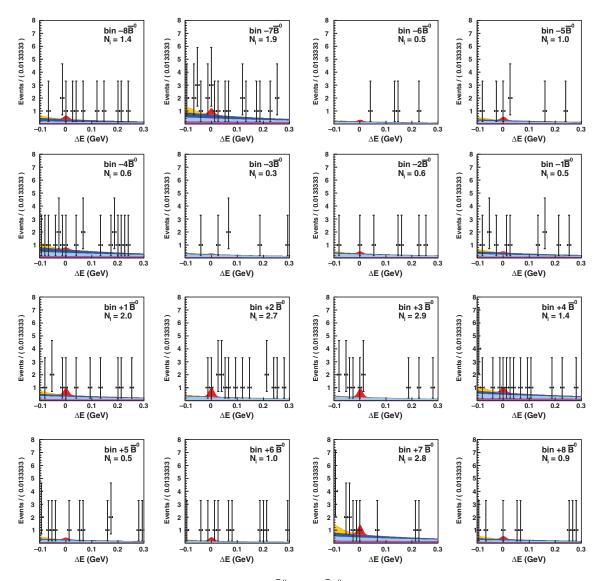


Fig. 5. Projections of the combined fit of the $\bar{B}^0 \rightarrow D\bar{K}^{*0}$ sample on ΔE for each Dalitz bin, with the $M_{\rm bc} > 5.27 \,{\rm GeV}/c^2$ and $C'_{\rm NB} > 2$ requirements. The fill styles for the signal and background components are the same as in Fig. 3.

The uncertainty arising from the finite sample of flavor-tagged $D \rightarrow K_S^0 \pi \pi$ decays is evaluated by varying the values of K_i within their statistical uncertainties. The uncertainty due to the limited precision of c_i and s_i parameters is obtained by smearing the c_i and s_i values within their total errors and repeating the fits for the same experimental data. The uncertainty due to k in Eq. (2) is evaluated by varying the value of $k (= 0.95 \pm 0.03)$ within its error [17]. Total systematic uncertainties in (x, y) are obtained by summing all uncertainties in quadrature and are listed in Table 3.

7. Result

We use the frequentist approach with Feldman–Cousins ordering [33] to obtain the physical parameters $\mu = (\phi_3, r_S, \delta_S)$ [or true parameters $\mu = z_{\text{true}} = (x_-, y_-, x_+, y_+)$] from the measured parameters $z = z_{\text{meas}} = (x_-, y_-, x_+, y_+)$ taken from the likelihood distribution. In

Table 2. (x, y) parameters and their statistical correlations from the combined fit of the $B^0 \rightarrow DK^{*0}$ sample. The error is statistical. The values and errors are obtained from the likelihood distribution.

Parameter	
<i>x</i>	$+0.29 \pm 0.32$
<i>Y</i> _	-0.33 ± 0.41
$\operatorname{corr.}(x, y)$	+7.0%
x_+	$+0.07 \pm 0.42$
y_+	$+0.05 \pm 0.45$
$corr.(x_+, y_+)$	-7.5%

Table 3. Systematic uncertainties in the (x, y) measurement for the $B^0 \rightarrow DK^{*0}$ mode. Values are rounded to two significant digits and those less than 0.005 are quoted as 0.00.

Source of uncertainty	Δx_{-}	Δy_{-}	Δx_+	Δy_+
Dalitz efficiency	± 0.00	$+0.01 \\ -0.00$	± 0.01	$^{+0.00}_{-0.01}$
Migration between bins	± 0.00	$+0.01 \\ -0.00$	$+0.01 \\ -0.00$	± 0.00
PDF parameterization	$+0.01 \\ -0.07$	$+0.07 \\ -0.01$	$^{+0.01}_{-0.10}$	$^{+0.04}_{-0.06}$
Flavor-tag statistics	± 0.00	± 0.00	± 0.00	$+0.00 \\ -0.01$
c_i, s_i precision	± 0.03	$^{+0.09}_{-0.08}$	± 0.05	$^{+0.08}_{-0.10}$
k precision	± 0.00	± 0.01	± 0.00	± 0.00
Total without c_i , s_i precision	$+0.01 \\ -0.07$	+0.07 -0.02	$+0.02 \\ -0.10$	$^{+0.04}_{-0.06}$
Total	$+0.03 \\ -0.08$	$+0.12 \\ -0.08$	+0.05 -0.11	$+0.09 \\ -0.12$

essence, the confidence level α for a set of physical parameters μ is calculated as

$$\alpha(\mu) = \frac{\int_{\mathcal{D}(\mu)} p(z|\mu) \, dz}{\int_{\infty} p(z|\mu) \, dz},\tag{15}$$

where $p(z|\mu)$ is the probability density to obtain the measurement z given by the set of true parameters μ . The integration domain $\mathcal{D}(\mu)$ is given by the likelihood ratio (Feldman–Cousins) ordering:

$$\frac{p(z|\mu)}{p(z|\mu_{\text{best}}(z))} > \frac{p(z_0|\mu)}{p(z_0|\mu_{\text{best}}(z_0))},\tag{16}$$

where $\mu_{\text{best}}(z)$ is the μ that maximizes $p(z|\mu)$ for the given z, and z_0 is the result of the data fit. This PDF is taken from MC pseudo-experiments.

Systematic uncertainties in μ are obtained by varying the measured parameters z within their systematic uncertainties assuming a nominal distribution. In this calculation, we ignore the correlations of uncertainties between the B^0 and \bar{B}^0 as the two samples are independent.

As a result of this procedure, we obtain the confidence levels (C.L.) for (x, y) and the physical parameter r_S . The C.L. contours on (x, y) are shown in Fig. 6. 1 – C.L. as a function of r_S is shown

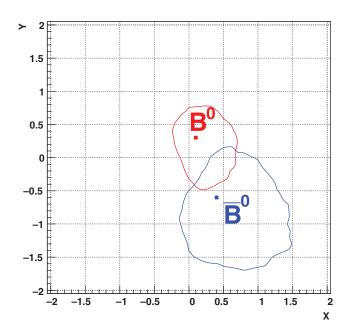


Fig. 6. C.L. contours for (x_-, y_-) (blue) and (x_+, y_+) (red). The dots show the most probable (x, y) values; the lines show the 68% contours. The fluctuations arise from the statistics of the pseudo-experiments and the C.L. step used.

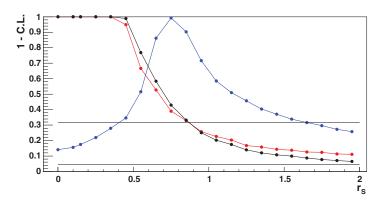


Fig. 7. Likelihood profile for r_s . The blue points are for $\bar{B}^0(x_-, y_-)$, red are for $B^0(x_+, y_+)$, and black are \bar{B}^0 and B^0 combined. The two horizontal lines show 68% and 95% C.L.

in Fig. 7. The final results are:

$$x_{-} = +0.4^{+1.0+0.0}_{-0.6-0.1} \pm 0.0, \tag{17}$$

$$y_{-} = -0.6^{+0.8+0.1}_{-1.0-0.0} \pm 0.1, \tag{18}$$

$$x_{+} = +0.1^{+0.7+0.0}_{-0.4-0.1} \pm 0.1, \tag{19}$$

$$y_{+} = +0.3^{+0.5+0.0}_{-0.8-0.1} \pm 0.1,$$
(20)

$$r_S < 0.87$$
 at 68% C.L., (21)

where the first error is statistical, the second is systematic without uncertainties in (c_i, s_i) , and the third is from the (c_i, s_i) precision from CLEO.

8. Conclusion

We report the first measurement of the amplitude ratio r_S using a model-independent Dalitz analysis of $D \to K_S \pi^+ \pi^-$ decays in the process $B^0 \to DK^{*0}$ with the full data sample of 711 fb⁻¹ corresponding to $772 \times 10^6 B\bar{B}$ pairs collected by the Belle detector at the $\Upsilon(4S)$ resonance. Model independence is achieved by binning the Dalitz plot of the $D \to K_s^0 \pi^+ \pi^-$ decay and using the strong-phase coefficients with binning as in the CLEO experiment [22]. We obtain the value $r_{s} < 0.87$ at 68% C.L. This measurement results in lower statistical precision than the model-dependent measurement from BaBar with the $B^0 \rightarrow DK^0$ mode [9] despite the larger data sample due to the smaller $B^0 \to DK^{*0}$ signal observed. The result is consistent with the most precise r_S measurement reported by the LHCb Collaboration [18] of $r_S = 0.240^{+0.055}_{-0.048}$ that uses $B^0 \to [K^+K^-, K^\pm \pi^\mp, \pi^+\pi^-]_{DK}^{*0}$ decays. We have confirmed the feasibility of the modelindependent Dalitz analysis method with neutral $B \rightarrow DK^*$. The value of r_S indicates the sensitivity of the neutral $B \to DK^*$ decay to ϕ_3 because the statistical uncertainty is proportional to $1/r_S$. In future high statistics experiments such as Belle II and the LHCb upgrade, this method will give a precise and model-independent determination of ϕ_3 . A more advanced double-Dalitz-plot analysis of $B^0 \to DK^+\pi^-$, $D \to K_S\pi^+\pi^-$ [34] has been proposed; this result can be considered as that from one bin of such an analysis.

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