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# Effect of Uncertainty in Localized Imperfection on the Ultimate Compressive Strength of Cold-Formed Stainless Steel Hollow Sections

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**Abstract:** Stainless steel has excellent corrosion resistance properties, considerable long-term durability, and good mechanical strength. Hollow sections are a versatile and efficient form for construction applications. The use of cold-formed stainless steel rectangular hollow section (RHS) and square hollow section (SHS) in construction industry grasps the attention of designers conceiving long-term, cost-effective structures. For cold-formed RHS and SHS, localized imperfection ( $\omega$ ) resulting from rolling and fabrication process is inevitable.  $\omega$  has inherent variability and has no definitive characterization. In this paper, statistical analysis of the maximum value of  $\omega$  collected from available experimental data is conducted. A new approach utilizing Fourier series to generate the three-dimensional (3D) models of members with random  $\omega$  is proposed. Probabilistic studies based on the proposed 3D models are then carried out to evaluate the effect of uncertainty in  $\omega$  on the ultimate compressive strength of stainless steel columns with cold-formed RHS and SHS. A total of 21 columns that are prone to local buckling reduction are studied. The results show that uncertainty in  $\omega$  has a considerable influence on the columns with relatively higher cross-sectional slenderness.

**Keywords:** stainless steel; cold-formed; uncertainty; localized imperfection

# 1. Introduction

Stainless steel is a steel alloy that contains a minimum of 10.5% chromium (Cr) content by mass. There is a wide range of stainless steels with varied levels of corrosion resistance and mechanical strength. More than 200 standardized stainless steel grades have hitherto been developed. The basic alloying elements of stainless steel grades are chromium (Cr) and nickel (Ni). According to Cr–Ni content by mass, stainless steels can be classified into five basic groups [1,2], as shown in Figure 1.

The five types are austenitic, ferritic, austenitic–ferritic (duplex), martensitic, and precipitation hardening stainless steels. The first three types have a wide and diverse application in the construction industry, such as building exteriors, facades, and pedestrian bridges. The last two types are for specialist applications [3]. For instance, martensitic grades are commonly used for bearings and turbine blades, and precipitation hardening grades are mostly used in the nuclear and aerospace industry.

Steel hollow sections are a versatile and efficient form for construction applications [4–9]. Buildings that use steel hollow sections have high strength-to-weight ratios. The efficient use of steel hollow sections reduces material usage, resulting in lightweight structures. It allows for large spans, and thus is a choice to achieve the optimal economic benefit. As stainless steels have excellent corrosion resistance properties, considerable long-term durability, and good mechanical strength [10–14], the use of cold-formed stainless steel rectangular hollow section (RHS) and square hollow section (SHS) (shown in Figure 2a) in construction industry has attracted considerable attention. For example, these sections

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have been used in the facade of the building of department of Chemistry, ETH Zürich (Switzerland), shown in Figure 2b, and the support frame of Marqués de Riscal Vineyard (Spain), shown in Figure 2c. Also, they can be used for the main frame structure of residential buildings.

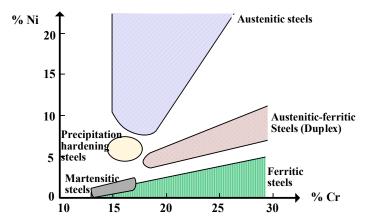
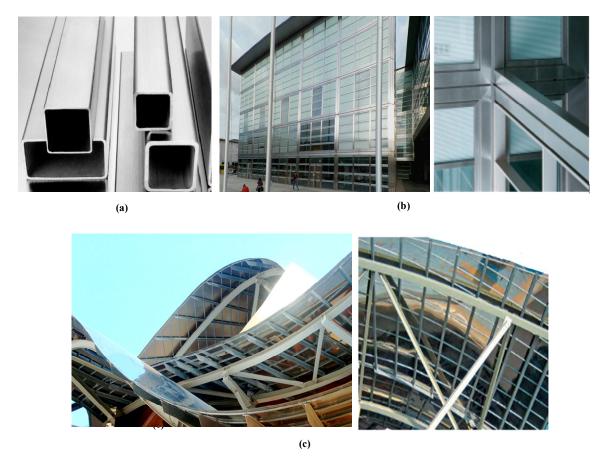


Figure 1. Classification of stainless steels groups based on the content of chromium (Cr) and nickel (Ni).

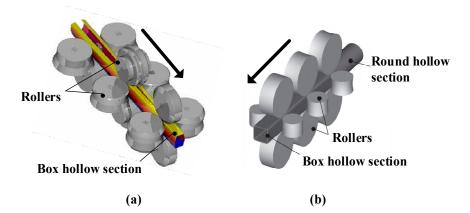


**Figure 2.** Cold-formed stainless steel rectangular hollow section (RHS) and square hollow section (SHS) and their applications in construction industry. (a) Cold-formed stainless steel RHS and SHS; (b) facade of the building of department of Chemistry, ETH Zürich (Switzerland); (c) support frame of Marqués de Riscal Vineyard (Spain).

Cold-formed hollow sections are manufactured by cold working and welding. There are two different forming methods for cold-formed stainless steel RHS and SHS—direct forming and round to square forming. In direct forming (shown in Figure 3a), the steel strip is transformed into a square or

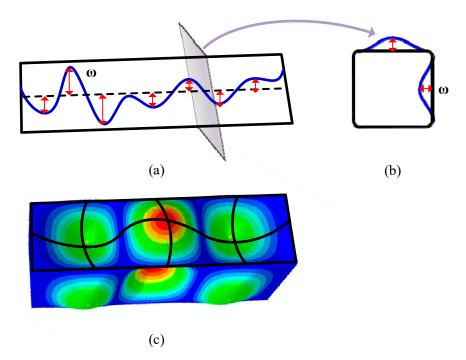
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rectangular hollow section by bending it through rollers, and welding the seam after that. In round to square forming (shown in Figure 3b), the steel strip is first formed into a circular hollow section and then it is welded. After forming the circular hollow section, square or rectangular shapes are created using profiling rollers.



**Figure 3.** Two forming methods for cold-formed RHS and SHS. (a) Direct forming [15]; (b) round to square forming [16].

For the two forming methods, localized imperfection ( $\omega$ ) induced by rolling and fabrication process is inevitable (shown in Figure 4). It has sufficient variability and has no definitive characterization [17]. The study of Wang et al. [18] and Zhao et al. [19,20] showed that both the shape and magnitude of  $\omega$  varied randomly in the longitudinal direction for cold-formed members with box section, as shown in Figure 4a. The shape of  $\omega$  in the transverse direction (cross-sectional) was convexity/concavity [21,22], which can be modelled by a half-sine wave, as shown in Figure 4b. In numerical studies,  $\omega$  is typically modelled by local buckling mode (shown in Figure 4c) obtained from linear perturbation buckle analysis [23]. The local buckling mode gives idealized  $\omega$  and neglects uncertainty in localized imperfection.



**Figure 4.** Localized imperfection ( $\omega$ ). (a) Random  $\omega$  along longitudinal centerline of the surface; (b)  $\omega$  in transverse direction (convexity/concavity); (c) idealized  $\omega$  obtained from buckle analysis.

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The cross-sectional resistance of cold-formed hollow sections, which contain slender thin-walled elements, is sensitive to initial localized imperfection [24]. These sections undergo local buckling reduction in advance of failure and their ultimate capacity may be influenced by the uncertainty in localized imperfection. Nevertheless, research related to the effect of uncertainty in localized imperfection on the ultimate compressive strength of cold-formed stainless steel RHS and SHS has not been explicitly reported, even though they have been studied for structural applications at the material, member, and system levels [1,17,19–23].

For this purpose, a mathematical approach utilizing Fourier series to generate the 3D models of members with random  $\omega$  is proposed in this paper. On the basis of the proposed 3D models, probabilistic studies are carried out to evaluate the effect of uncertainty in  $\omega$  on the ultimate compressive strength of the columns.

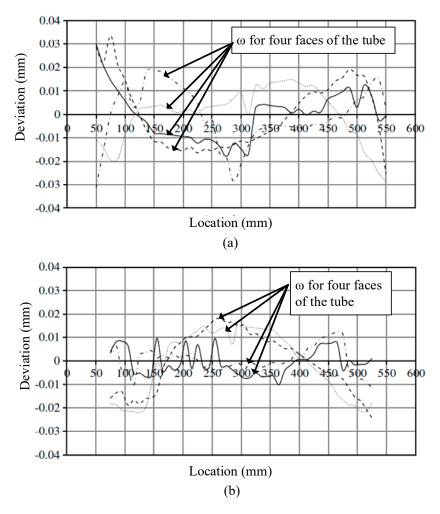
# 2. Statistical Analysis of the Maximum Localized Imperfection ( $\omega$ )

Although there is a considerable uncertainty when characterizing localized imperfection ( $\omega$ ) in cold-formed hollow sections, experimental data on the maximum localized imperfection ( $\omega_{max}$ ) are available for the particular case of RHS and SHS stainless steel specimens. A statistical analysis of experimental results of the  $\omega_{max}$  from the literature [19–22,25–31] is carried out in this section. A total of 161 cold-formed stainless steel RHS and SHS samples are collected. A summary of the samples is shown in Table 1. The studied samples refer to the stainless steel grades commonly used in construction. In these references, some studies [21,22] provided the pattern of  $\omega$  in the transverse direction (cross-sectional), in which all the reported patterns are very close to a half-sine wave. Few of them reported the variation of localized imperfection in the longitudinal direction. The distribution of  $\omega$  in the longitudinal direction for two tubes reported in the work of [19] is shown in Figure 5. It is observed that  $\omega$  in the longitudinal direction has a considerable variability and its characterization in a definite closed-form is not feasible.

**Table 1.** Summary of the samples collected from the literature.

Reference	Stainless Steel Groups	Grade	No. of Samples with Measured ω	
B.F. Zheng et al., 2016 [25]	Austenitic	EN1.4301	4	
I. Arrayago. et al., 2016 [26]	Ferritic	EN1.4003	12	
B. Young and W.M. Lui, 2005 [21]	Duplex	EN1.4162	5	
	Austenitic	EN1.4301	10	
	Austenitic	EN1.4571	6	
O. Zhao et al., 2015 [20]	Austenitic	EN1.4307	6	
	Austenitic	EN1.4404	6	
	Duplex	EN1.4162	6	
M. Theofanous and L. Gardner, 2009 [27]	Duplex	EN1.4162	8	
W.M. Lui et al., 2014 [22]	Duplex	EN1.4462	10	
Y. Huang and B. Young, 2013 [28]	Duplex	EN1.4162	22	
S. Afshan and L. Gardner,	Ferritic	EN1.4003	6	
2013 [29]	Ferritic	EN1.4509	2	
M. Bock et al., 2015 [30]	Ferritic	EN1.4003	8	
I. Arrayago and E. Real, 2015 [31]	Ferritic	EN1.4003	26	
O. Zhao et al., 2016 [19]	Ferritic	EN1.4003	24	
			Total: 161	

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**Figure 5.** Distribution of  $\omega$  in the longitudinal direction for two tubes reported in the work of [19]. (a) RHS  $100 \times 40 \times 2$ ; (b) SHS  $60 \times 60 \times 3$ .

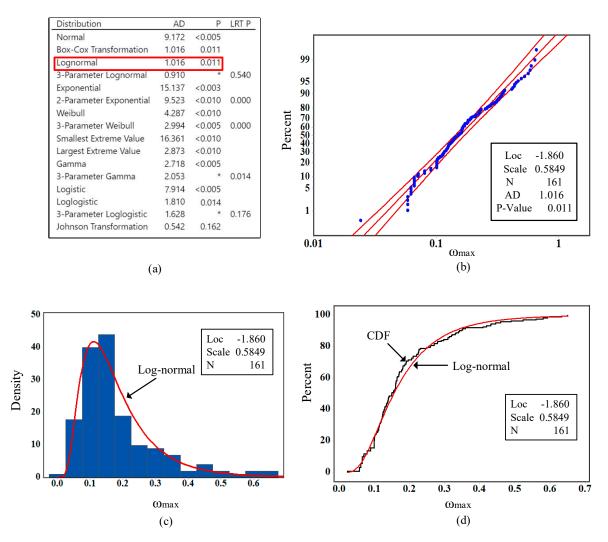
The probability distribution for  $\omega_{max}$  among the samples collected in the literature was identified by statistical distribution tests (Anderson–Darling method), as well as from probability plots. Both distribution tests and probability plots were performed using the statistical software Minitab 18 [32].

For the case of distribution tests, Anderson–Darling statistics (AD) and p-values measure how well the specified distributions fit the data. For a given sample data and distribution, the smaller the AD, the better the distribution fits the data. Higher p-values indicate a better fit, and p-values less than 0.05 typically indicate that the data do not follow the specified distribution. The indicator LRT P is for three-parameter distributions only. A lower LRT P indicates that the related two-parameter distribution can be significantly improved by a third parameter. Goodness of fit test results for 16 different distribution tests are shown in Figure 6a. The Box–Cox transformation and the Johnson transformation are disregarded as the target is to identify probability distribution rather than to perform any transformation. It is found that the log-normal distribution (AD = 1.016, p-value = 0.011) represents the best fit for the data of  $\omega_{max}$ .

Probability plots are another efficient way to determine whether the specified distribution fits the sample data. The closer the data to the middle straight line, the better the distribution fits the data. The probability plot of the sample data is shown in Figure 6b. It is observed that the data points are in close agreement with the center straight line. It again demonstrates that the sample data follow the log-normal distribution. The histogram of  $\omega_{max}$  is shown in Figure 6c. The log-normal distribution is fitted to the histogram. Comparison of the cumulative probability (CDF) curve against the log-normal

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distribution is shown in Figure 6d, in which CDF determines the probability that an observation will be less than or equal to a certain value.



**Figure 6.** Identifying probability distribution for  $\omega_{max}$ . (a) Goodness of fit test results for 16 different distribution tests; (b) probability plot of  $\omega_{max}$ ; (c) histogram of  $\omega_{max}$ ; (d) comparison of cumulative probability (CDF) curve against the log-normal distribution. AD, Anderson–Darling.

Some standards provide the maximum allowable value of localized imperfection in the transverse direction [33,34]. The tolerance for convexity/concavity of a cross-section specified in EN-10219-2:2006 [34] is  $\min\{\omega_{max}/b, 0.5 \text{ mm}\}$ , where  $\omega_{max}/b \leq 0.008$ ; b is the side (straight side of the cross-section) length; and  $\omega_{max}$  represents the maximum deviation from the straight side.

#### 3. Fourier Series-Based 3D Models with Random $\omega$

On the basis of a superposition of Fourier series expansion of different functions, a 3D model with random localized imperfection ( $\omega$ ) is proposed in this section. The Fourier series technique has been widely used for 3D surface modeling [35,36]. For a function f(x) that is periodic on an interval [-L, L], it can be expressed as the Fourier series, given by

$$f(x) = \frac{a_0}{2} + \sum_{k=1}^{\infty} \left[ a_k \cos\left(\frac{k\pi x}{L}\right) + b_k \sin\left(\frac{k\pi x}{L}\right) \right],\tag{1}$$

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where

$$a_0 = \frac{1}{L} \int_{-L}^{L} f(x) dx,\tag{2}$$

$$a_k = \frac{1}{L} \int_{-L}^{L} f(x) \cos\left(\frac{k\pi x}{L}\right) dx \ (k = 0, 1, 2, 3, \dots),$$
 (3)

$$b_k = \frac{1}{L} \int_{-L}^{L} f(x) \sin\left(\frac{k\pi x}{L}\right) dx \ (k = 1, 2, 3, \cdots).$$
 (4)

Assume a surface consists of  $n \times m$  points in a 3D coordinate system, where X coordinate represents the longitudinal (length) direction, Y represents the transverse (width) direction, and Z represents deviation from the flat surface parallel to the XY plane. For a point  $(x_i, y_j, z_{ij})$  (I = 1, 2, ..., m; j = 1, 2, ..., n) on the surface,  $z_{ij}$  governs localized imperfection ( $\omega$ ). All  $z_{ij}$  elements comprise an  $n \times m$  matrix [**Z**], which can be determined by

$$[\mathbf{Z}] = [\mathbf{F}_1] + ([\mathbf{F}_2] - [\mathbf{F}_1])[\mathbf{S}],$$
 (5)

$$[\mathbf{F}_{1}] = \begin{bmatrix} f_{1}(\mathbf{x}_{1}) & f_{1}(\mathbf{x}_{2}) & \cdots & f_{1}(\mathbf{x}_{i}) & \cdots & f_{1}(\mathbf{x}_{m}) \\ f_{1}(\mathbf{x}_{1}) & f_{1}(\mathbf{x}_{2}) & \cdots & f_{1}(\mathbf{x}_{i}) & \cdots & f_{1}(\mathbf{x}_{m}) \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ f_{1}(\mathbf{x}_{1}) & f_{1}(\mathbf{x}_{2}) & \cdots & f_{1}(\mathbf{x}_{i}) & \cdots & f_{1}(\mathbf{x}_{m}) \end{bmatrix}_{n \times m},$$

$$(6)$$

$$[\mathbf{F}_{2}] = \begin{bmatrix} f_{2}(\mathbf{x}_{1}) & f_{2}(\mathbf{x}_{2}) & \cdots & f_{2}(\mathbf{x}_{i}) & \cdots & f_{2}(\mathbf{x}_{m}) \\ f_{2}(\mathbf{x}_{1}) & f_{2}(\mathbf{x}_{2}) & \cdots & f_{2}(\mathbf{x}_{i}) & \cdots & f_{2}(\mathbf{x}_{m}) \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ f_{2}(\mathbf{x}_{1}) & f_{2}(\mathbf{x}_{2}) & \cdots & f_{2}(\mathbf{x}_{i}) & \cdots & f_{2}(\mathbf{x}_{m}) \end{bmatrix},$$

$$(7)$$

where  $f_1(x_i)$  and  $f_2(x_i)$  are functions that are decomposed into Fourier series.

[S] is an  $m \times m$  diagonal matrix,

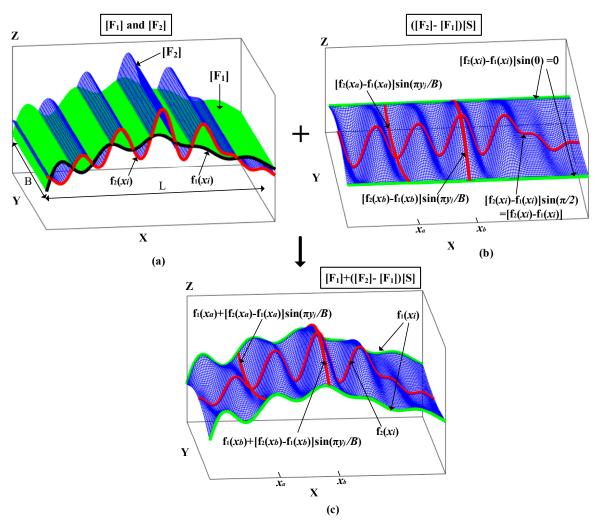
$$[\mathbf{S}] = \begin{bmatrix} \sin(\pi y_j/B) & 0 & \cdots & 0 \\ 0 & \sin(\pi y_j/B) & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \sin(\pi y_j/B) \end{bmatrix}_{m \times m}$$
(8)

$$\begin{aligned}
&([\mathbf{F}_{2}] - [\mathbf{F}_{1}])[\mathbf{S}] = \\
&\begin{bmatrix} [f_{2}(x_{1}) - f_{1}(x_{1})] \sin\left(\frac{\pi y_{1}}{B}\right) & \cdots & [f_{2}(x_{i}) - f_{1}(x_{i})] \sin\left(\frac{\pi y_{1}}{B}\right) & \cdots & [f_{2}(x_{m}) - f_{1}(x_{m})] \sin\left(\frac{\pi y_{1}}{B}\right) \\
&[f_{2}(x_{1}) - f_{1}(x_{1})] \sin\left(\frac{\pi y_{2}}{B}\right) & \cdots & [f_{2}(x_{i}) - f_{1}(x_{i})] \sin\left(\frac{\pi y_{2}}{B}\right) & \cdots & [f_{2}(x_{m}) - f_{1}(x_{m})] \sin\left(\frac{\pi y_{2}}{B}\right) \\
& \vdots & \vdots & \vdots & \vdots & \vdots \\
&[f_{2}(x_{1}) - f_{1}(x_{1})] \sin\left(\frac{\pi y_{m}}{B}\right) & \cdots & [f_{2}(x_{i}) - f_{1}(x_{i})] \sin\left(\frac{\pi y_{m}}{B}\right) & \cdots & [f_{2}(x_{m}) - f_{1}(x_{m})] \sin\left(\frac{\pi y_{m}}{B}\right) \\
\end{bmatrix} 
\end{aligned} (9)$$

The fundamental principles of generating 3D surfaces with random  $\omega$  are illustrated by the following.  $f_1(x_i)$  and  $f_2(x_i)$  are two functions that are decomposed into Fourier series with random coefficients. [F<sub>1</sub>] and [F<sub>2</sub>] govern two curved surfaces, as shown in Figure 7a, where L and B are the length and width of the surface, respectively. Localized imperfection ( $\omega$ ) is determined by matrix ([F<sub>2</sub>] - [F<sub>1</sub>]) [S]. It consists of two components: the transverse variation and longitudinal variation, as shown in Figure 7b. The shape and magnitude of  $\omega$  in the longitudinal direction depends on the curve along the longitudinal centerline. It is determined by the function [ $f_2(x_i) - f_1(x_i)$ ] sin ( $\pi$ /2). The shape of  $\omega$  in the transverse direction is modelled by a half-sine-wave, as its shape in the transverse direction reported in most literatures is convexity/concavity. The half-sine-wave is determined by the function [ $f_2(x_i) - f_1(x_i)$ ] sin ( $\pi$ y<sub>i</sub>/B), as shown in Figure 7b, where the two half-sine waves correspond to ( $x_a$ ,  $y_i$ )

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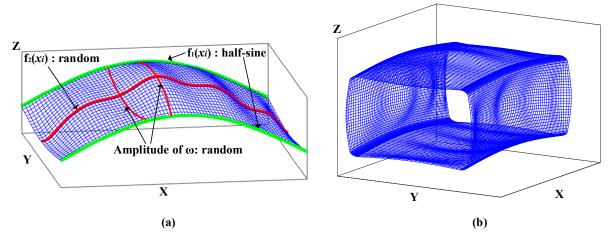
and  $(x_b, y_j)$  (j = 0, 1, ..., m). The generated surface with random  $\omega$  is determined by  $[\mathbf{F_1}] + ([\mathbf{F_2}] - [\mathbf{F_1}])$  [S], as shown in Figure 7c.



**Figure 7.** Development of 3D surface with random  $\omega$ . (a) Two surfaces determined by  $[F_1]$  and  $[F_2]$ ; (b) surface determined by  $[F_1] - [F_1]$ ) [S]; (c) surface determined by  $[F_1] + ([F_2] - [F_1])$  [S].

The 3D model of a member with random  $\omega$  is finally assembled by four surfaces. It should be mentioned that in order to fit the four faces together, relevant coordinate transformation should be conducted. Coordinate transformation depends on assembling order and the position of the surface in a 3D space. For developing a 3D model of RHS and SHS with round corners, additional curved surfaces representing round corners need to be modelled. The generated 3D model for a typical surface with random  $\omega$  and half-sine edges is shown in Figure 8a, and for a typical member with random  $\omega$  and half-sine edges, it is shown in Figure 8b.

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**Figure 8.** Generated 3D model. (a) Surface with random  $\omega$  and half-sine edges; (b) member with random  $\omega$  and half-sine edges.

# 4. Case Study of Stainless Steel Columns with Cold-Formed RHS and SHS

A total of 21 stainless steel columns with cold-formed RHS and SHS are selected among the tested specimens reported in the literature [21,28,29,37–39]. The studied columns have cross-sectional slenderness ( $\lambda_l$ ) higher than 0.776. This is to ensure that the columns undergo cross-sectional local buckling reduction before they reach the ultimate compressive strength. According to the work of [24], the nominal compressive strength of a column with RHS or SHS is determined by min{P<sub>ne</sub>, P<sub>nl</sub>}, where P<sub>ne</sub> and P<sub>nl</sub> are the nominal global buckling strength and local buckling strength, respectively. The interaction between global and local buckling depends on  $\lambda_l$  and is determined by

When 
$$\lambda_l \le 0.776 P_{nl} = P_{ne}$$
, (10)

When 
$$\lambda_l > 0.776 \frac{P_{nl}}{P_{ne}} = \lambda_l^{-0.8} - 0.15\lambda_l^{-1.6}$$
. (11)

A plot of  $P_{nl}/P_{ne}$  versus  $\lambda_l$  is shown in Figure 9. It should be pointed out that, although Equations (10) and (11) are developed based on carbon steel members, the two equations are applicable to stainless steel members with cold-formed RHS and SHS and give an accurate prediction [40,41].

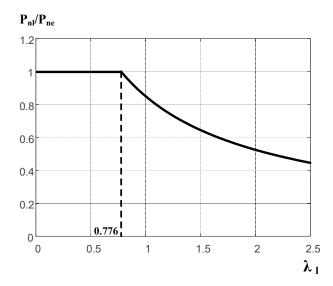


Figure 9. The interaction between global and local buckling strength.

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Details of the studied columns are shown in Table 2. Namely,  $b_1$ ,  $b_2$ , and t are the depth, width, and thickness of the hollow cross-section, respectively; R is external radius of the round corner; L is the length of the column;  $\lambda_c$  and  $\lambda_l$  are member slenderness and cross-sectional slenderness, respectively;  $\omega_g$  is the amplitude of global member imperfection (out-of-straightness).  $\omega_g$  is not reported for some cases of stub columns ( $\lambda_c \leq 0.2$ ), while the shape of  $\omega_g$  is adopted as a half-sine wave for other columns.

Reference	Specimen	b <sub>1</sub> (mm)	b <sub>2</sub> (mm)	t (mm)	R (mm)	L (mm)	$\lambda_{\rm c}$	$\lambda_{\mathrm{l}}$	$\omega_{g}$
[37]	SHS2L300	50.1	50.3	1.58	2.8	300	0.14	0.8	-
	RHS1L3000	140.1	79.9	3.01	10.0	3000	0.71	0.9	0.927
	C5L200	100.1	50.1	2.5	3.7	200	0.18	1.0	-
	C6L200	150.0	50.1	2.5	4.3	200	0.18	1.5	-
	C6L550	150.1	50.2	2.49	4.5	550	0.48	1.4	0.5
[28]	C5L900R	100.1	50.4	2.49	3.5	900	0.79	0.8	0.857
	C6L900	150.4	50.3	2.47	4.7	900	0.79	1.3	0.857
	C6L1200	149.9	50.5	2.46	4.5	1200	1.04	1.2	1.143
	C6L1550	150.5	50.3	2.49	4.5	1550	1.35	1.0	1.476
[20]	RHS 120 × 80 × 3-SC2	120.0	80.0	2.83	6.7	362	0.16	0.82	-
[29]	RHS 120 × 80 × 3-1077	120.0	79.9	2.87	6.8	1077	0.35	0.8	0.95
	RHS 120 × 80 × 3-1577	120.0	79.9	2.81	6.4	1577	0.51	0.8	0.96
	R1L1200	120.1	40.1	1.94	5.0	1199	0.48	1.1	0.254
[38]	R1L2000	120.2	40.0	1.95	5.1	2000	0.80	1.0	0.444
	R3L2000	120.0	80.0	2.80	6.7	2000	0.44	0.8	0.381
	SHS100 × 100 × 2-LC-2 m	99.8	99.9	1.86	3.2	2000	0.73	1.0	0.1
[39]	RHS100 $\times$ 50 $\times$ 2-LCJ-2 m	99.8	49.8	1.83	3.7	2000	0.80	0.9	0.6
	RHS100 $\times$ 50 $\times$ 2-LC-1 m	99.8	50.0	1.82	3.6	1000	0.69	1.0	0.1
	RHS120 × 80 × 3-LC-1 m	120.0	80.2	2.86	5.7	1001	0.45	0.8	1
[21]	$160 \times 80 \times 3$	160.1	80.8	2.87	9.0	600	0.09	1.2	-
[41]	$200\times110\times4$	196.2	108.5	4.01	13.0	600	0.07	1.1	-

**Table 2.** Details of the selected stainless steel columns for probabilistic studies.

# 5. Generation of 3D Models of the Studied Columns and Finite Element (FE) Analysis

The structural behavior of stainless steel columns with random  $\omega$  was studied using finite element (FE) software Abaqus 6.13 [42]. The Fourier series-based 3D model of the columns with random  $\omega$  was generated by Matlab 2017b [43]. Then, the generated models were imported into Abaqus to conduct the FE analysis. The input file of ABAQUS is generated by MATLAB script.

# 5.1. Generation of 3D Model with Random $\omega$ Using MATLAB

The development of the coefficient of Fourier series terms of function  $f_2(x)$  and  $f_2(x)$  was performed in Matlab. For the stub columns ( $\lambda_c \leq 0.2$ ), Fourier series expansion of function  $f_1(x)$  generated a straight line. For other columns,  $f_1(x)$  generated half-sine-waves, where the magnitude of the half-sine wave was taken as the corresponding  $\omega_g$ , shown in the above Table 2. For all columns, coefficients of Fourier series terms of function  $f_2(x)$  were defined as random. The maximum amplitude of the modelled  $\omega$  for each column was limited to min{0.008b, 0.5}. For each column, 50 models with random values of localized imperfection  $\omega$  were produced. The developed Matlab program automatically created a Python script associated with an Input file operated in Abaqus. It is worth pointing out that the distribution of the generated random  $\omega_{max}$  followed a log-normal distribution as the experimental data of  $\omega_{max}$ . This was explicitly set in the developed Matlab program.

### 5.2. FE Analysis Using ABAQUS

Abaqus/Standard (implicit solver) was employed for FE analysis. A four-node shell element with reduced integration (S4R) was used. It allows transverse shear deformation, and accounts for finite membrane strains and arbitrarily large rotations. The number of integration points through the thickness is five (Simpson's rule). The load–displacement response was predicted using an incremental procedure based on arc-length methods. The modified Riks method [44], which is available in Abaqus, was used. On the basis of a mesh convergence study, at least ten elements across the plate widths were used. Stainless steels have considerable non-linear stress–strain response. To accurately predict the structural behavior of the studied columns, the adopted material property for each column was obtained from a corresponding uniaxial tensile stress–strain coupon test. Details of the parameters that describe the stress–strain curves can be found in the literature [21,28,29,37–39]. For all the models, edge elements at both ends (top and bottom) were kinematically coupled and connected to two control points where the relevant degrees of freedom were constrained.

Spread of plasticity through cross-section and along member length was traced by distributed plasticity approach. In the FE analysis, residual stresses have to be considered as they may have negative effects on the ultimate capacity of a structure. As the effect of through-thickness longitudinal bending residual stresses on the global behavior of stainless steel members with box sections is dominant [45], only longitudinal bending residual stresses were considered, and they were implicitly included in the stress–strain curves obtained from the tensile coupons test. For each model of the column with random  $\omega_{max}$ , geometrically and materially nonlinear analysis with imperfections (GMNIA) was carried out to determine the ultimate compressive strength of the column.

# 6. Effect of Uncertainty in $\omega$ on the Ultimate Compressive Strength of the Studied Columns

A probabilistic study was carried out to investigate the effect on the uncertainty in localized imperfection ( $\omega$ ) on the distribution of the ultimate strength of the stainless steel columns with cold-formed RHS and SHS. The experimental results and predicted results of the columns with random localized imperfection are shown in Table 3. Namely,  $P_{u\text{-EXP}}$  is the ultimate compressive strength obtained from experiment;  $P_{u\text{-rand}}$  is the predicted ultimate compressive strength for each model (each column have 50 models);  $\mu$  and COV are the mean value and coefficients of variation, respectively; and  $|\varepsilon_{max}|$  is the maximum value of relative error for each set of 50 models.

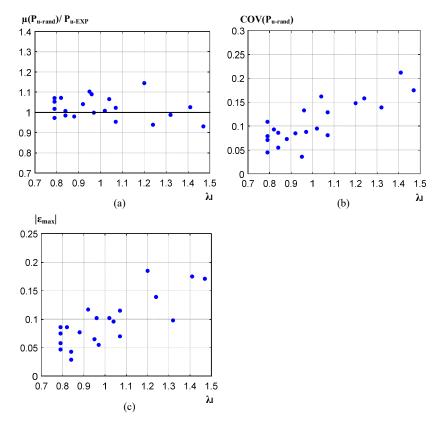
**Table 3.** Experimental results and predicted results for the studied columns. COV, coefficients of variation.

Specimen	$\lambda_{l}$	P <sub>u-EXP</sub> (kN)	μ(P <sub>u-rand</sub> ) (kN)	μ(P <sub>u-rand</sub> ) /P <sub>u-EXP</sub>	COV(P <sub>u-rand</sub> )	$ \epsilon_{max} $
SHS2L300	0.84	175.7	177.8	0.985	0.086	0.043
RHS1L3000	0.88	513.5	454.7	0.980	0.073	0.077
C5L200	0.95	370.1	387.5	1.103	0.036	0.065
C6L200	1.47	404.1	413.2	0.931	0.175	0.171
C6L550	1.41	353.2	388.1	1.026	0.212	0.175
C5L900R	0.84	336.0	326.0	1.007	0.055	0.029
C6L900	1.32	333.5	345.2	0.988	0.139	0.098
C6L1200	1.20	284.5	300.7	1.142	0.108	0.185
C6L1550	1.02	230.0	249.2	1.008	0.095	0.102
RHS 120 × 80 × 3-SC2	0.82	441.0	434.2	1.072	0.093	0.086
RHS $120 \times 80 \times 3-1077$	0.79	463.0	432.1	1.018	0.109	0.075
RHS $120 \times 80 \times 3-1577$	0.79	382.0	401.5	0.973	0.045	0.058

Tab	1 ~ 2	Cont

Specimen	$\lambda_{l}$	P <sub>u-EXP</sub> (kN)	μ(P <sub>u-rand</sub> ) (kN)	μ(P <sub>u-rand</sub> ) /P <sub>u-EXP</sub>	COV(P <sub>u-rand</sub> )	[ε <sub>max</sub> ]
R1L1200	1.07	167.0	153.5	1.02	0.129	0.115
R1L2000	0.97	141.3	137.9	0.999	0.088	0.055
R3L2000	0.79	394.0	355.7	1.071	0.071	0.047
SHS100 × 100 × 2-LC-2 m	1.04	176.0	183.0	1.066	0.162	0.096
RHS100 $\times$ 50 $\times$ 2-LCJ-2 m	0.92	157.0	145.3	1.041	0.085	0.117
RHS100 $\times$ 50 $\times$ 2-LC-1 m	0.96	163.0	151.4	1.090	0.133	0.102
RHS120 $\times$ 80 $\times$ 3-LC-1 m	0.79	448.0	415.5	1.053	0.079	0.086
$160 \times 80 \times 3$	1.24	537.3	505.0	0.939	0.158	0.139
$200 \times 110 \times 4$	1.07	957.0	928.0	0.958	0.081	0.070

For all the studied columns,  $\mu(P_{u-rand})/P_{u-EXP}$  and  $COV\ (P_{u-rand})$  versus  $\lambda_l$  are plotted in Figure 10a,b, respectively. A plot of  $|\epsilon_{max}|$  against  $\lambda_l$  is shown in Figure 10c. It is observed that the value of  $\mu(P_{u-rand})/P_{u-EXP}$  ranges between 0.931 and 1.103 for all the columns, except the column with  $\lambda_l=1.20$ . Compared with  $P_{u-EXP}$ , the value of  $\mu(P_{u-rand})$  for most columns with relatively lower cross-sectional slenderness ( $\lambda_l<1.0$ ) is overestimated, while the value of  $\mu(P_{u-rand})$  for columns with higher cross-sectional slenderness seems to be underestimated. For the column with  $\lambda_l=1.20$ , the value of  $\mu(P_{u-rand})/P_{u-EXP}$  is 1.142. It indicates that most predicted results (from the 50 models with random  $\omega$ ) significantly overestimate the experimental result. This may be because the value of the actual maximum localized imperfection for this column, which is not reported in the work of [28], is relatively larger compared with the modelled  $\omega$ , whose maximum value is min{0.008b, 0.5}.

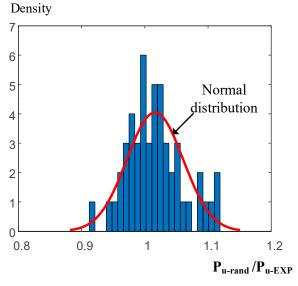


**Figure 10.** Predicted results against cross-sectional slenderness. (a)  $\mu(P_{u\text{-rand}})/P_{u\text{-EXP}}$  versus  $\lambda_l$ ; (b) coefficients of variation (COV)  $(P_{u\text{-rand}})$  versus  $\lambda_l$ ; (c)  $|\epsilon_{max}|$  versus  $\lambda_l$ .

On the other hand, both COV ( $P_{u\text{-rand}}$ ) and  $|\epsilon_{max}|$  increase as  $\lambda_l$  increases. One explanation is that, the larger the cross-sectional slenderness, the more sensitive the column is to initial localized imperfection. Consequently, the change in the value of modelled localized imperfection can result in a larger discrepancy in the ultimate compressive strength.

For the columns with relatively lower cross-sectional slenderness ( $\lambda_l < 1.0$ ), the values of COV ( $P_{u\text{-rand}}$ ) and  $|\epsilon_{max}|$  are less than 0.13 and 0.12, respectively. The result indicates that uncertainty in  $\omega$  has no considerable influence on the ultimate compressive strength of these columns. This may be because the columns with relatively lower cross-sectional slenderness are still not sensitive to initial localized imperfection. Besides, the result indicates that  $\omega$  can statistically be modelled as deterministic for these columns, such as using measured  $\omega$  in the experimental study.

For the columns with  $\lambda_l \geq 1.2$ , COV ( $P_{u\text{-rand}}$ ) are around 0.139–0.238 and the maximum value of  $|\epsilon_{max}|$  is 17.5%. It demonstrates that random  $\omega$  results in largely scattered ultimate compressive strength for the columns with larger cross-sectional slenderness, and it is important to consider the effect of uncertainty in  $\omega$  on these columns. The distribution of  $P_{u\text{-rand}}$  for a typical column (R1L1200) is shown in Figure 11. In the figure,  $P_{u\text{-rand}}$  is normalized by  $P_{u\text{-EXP}}$ . It is found that the distribution of  $P_{u\text{-rand}}/P_{u\text{-EXP}}$  can be fitted by a normal distribution.



**Figure 11.** Histograms of P<sub>u-rand</sub>/P<sub>u-EXP</sub> for R1L1200.

#### 7. Conclusions

In this paper, a new approach for modelling initial localized imperfection ( $\omega$ ) is introduced. The proposed approach considers uncertainty in  $\omega$  and is based on superposition of Fourier series expansion of different functions. The Fourier series-based 3D models for the columns with random  $\omega$  were generated by MATLAB, and were then imported into ABAQUS to conduct the FE analysis. The studied stainless steel columns with cold-formed RHS and SHS are selected among the tested specimens reported in the literature. The selected columns have relatively higher cross-sectional slenderness to ensure that they undergo cross-sectional local buckling reduction prior to failure. On the basis of this study, it is found that both the coefficients of variation and the maximum value of absolute error for the predicted results increase as cross-sectional slenderness increases. This is because columns with larger cross-sectional slenderness are sensitive to initial localized imperfection and, consequently, the change in the value of modelled localized imperfection can lead to much discrepancy in the ultimate compressive strength. Therefore, the effect of uncertainty in  $\omega$  on the columns with larger cross-sectional slenderness requires a careful probabilistic consideration that should be considered in practical design, in which deterministic methods are often employed.

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