Title: Single-Subject Structural Cortical Networks in Clinically Isolated Syndrome.

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ABSTRACT

Background: Structural cortical networks (SCNs) represent patterns of coordinated morphological modifications in cortical areas and they present the advantage of being extracted from previously acquired clinical MRI scans. SCNs have shown pathophysiological changes in many brain disorders, including multiple sclerosis.

Objective: To investigate alterations of SCNs at the individual level in patients with clinically isolated syndrome (CIS), thereby assessing their clinical relevance.

Methods: We analyzed baseline data collected in a prospective multi-center (MAGNIMS) study. CIS patients (n=60) and healthy controls (n=38) underwent high resolution 3T MRI. Measures of disability and cognitive processing were obtained for patients. Single-subject structural cortical networks were extracted from brain 3D-T1 weighted sequences; global and local network parameters were computed.

Results: Compared to healthy controls, CIS patients showed altered small-world topology, an efficient network organization combining dense local clustering with relatively few long-distance connections. These disruptions were worse for patients with higher lesion load and worse cognitive processing speed. Alterations of centrality measures and clustering of connections were observed in specific cortical areas in CIS patients when compared with healthy controls.

Conclusion: Our study indicates that SCNs can be used to demonstrate clinically relevant alterations of connectivity in clinically isolated syndrome.

INTRODUCTION

Multiple sclerosis (MS) is a chronic inflammatory demyelinating and degenerative disease of the central nervous system. For most patients, MS starts with an acute neurological episode known as a clinically isolated syndrome (CIS)¹. Conventional MRI findings are the best predictors of conversion to MS, although measures such as the WM lesion volume only partially correlate with the variable course of the disease². MRI abnormalities in the normal-appearing white and gray matter, and connectivity alterations may contribute to clinical outcome of patients with CIS³. Functional changes in the brain can lead to related morphological modification in cortical areas^{4,5} and, recently, it has become possible to describe these coordinated patterns of cortical morphology with network parameters⁶. In structural cortical networks (SCNs), the nodes correspond to cortical areas considered connected by edges when they have structural similarity in thickness or volume across subjects⁷ or within single subjects⁸. In comparison with other brain graphs (e.g., structural graphs derived from diffusion tensor imaging or functional graphs derived from synchronic activation of brain areas), SCNs present the unique advantage of being derived from anatomical MRI acquisitions, such as a 3D T1-weighted sequence, frequently available in acquired datasets and even in routine clinical protocols.

SCNs have shown clinically relevant changes in many brain diseases^{9–11}. However, few studies so far have investigated SCNs in multiple sclerosis (MS)^{12–15}.

SCN analysis has revealed in MS patients a disruption of physiological small-world network topology of the brain: an efficient organization, shared by most of the complex systems in nature, including brain networks, that is neither completely regular nor completely random, because it combines a dense local clustering with relatively few long-distance connections¹⁶. Whilst MS patients can shift in either direction, either a more regular^{13,14} or a more random network^{12,15}, the only study including a small cohort of CIS patients¹⁴ demonstrated an increased clustered organization of SCNs in CIS subjects compared with healthy controls (HCs). Regular networks, such as lattices, are highly clustered, but with a long path length: nodes are densely connected with

their clique of neighbors, whereas they are linked to distant nodes only through several intermediate steps, causing a decrease in the global network efficiency. However, the SCNs were extracted at group level in the CIS cohort, and so it remains unclear whether and how these disruptions are associated with individual measures of disability.

Our aim was to investigate SCNs alterations in a larger CIS patient cohort by constructing the network at the individual level, thereby assessing the clinical relevance of the SCN alterations. Moreover, since SCNs have been studied exclusively in single-center studies before, our secondary aim was to evaluate the impact of across-scanner variations on single-subject SCNs analysis.

MATERIALS AND METHODS

Participants

This study represents a re-analysis of the baseline data which were acquired as part of a prospective MAGNIMS (<u>www.magnims.eu</u>) 3-year multi-center, multi-vendor project, conducted at the following six MS Centers (Table 1). Previous studies using the same cohort addressed a different issue and are unrelated to this analysis^{17,18}.

Sixty patients experiencing symptoms suggestive of a CIS (age 18 to 59 years at baseline) were recruited from the six participating centers within six months of symptom onset. Thirty-eight age and gender-matched healthy controls (HCs) were also recruited (Table 1). Exclusion criteria for all participants were a history of other known medical conditions that could have affected the brain (vascular, malignant, neurological) and MRI related contra-indications.

Clinical examination

At baseline, CIS patients underwent neurological examination and Expanded Disability Status Scale (EDSS) was assessed by a trained physician. Forty-seven patients were also tested for information processing speed with the Symbol Digit Modality Test (SDMT).

Image acquisition and post-processing

All subjects underwent a baseline MRI scan of the brain and spinal cord at a 3T scanner, and CIS patients additionally underwent administration of intravenous contrast. The scanning parameters were based on local protocols and followed the MAGNIMS guidelines¹⁹. The acquisition protocols for the different vendors has been previously described¹⁸: it included brain 3D T1-weighted (1mm isotropic) and 3D fluid-attenuated inversion recovery (FLAIR); proton density/T2 and post-contrast T1-weighted sequences for brain and spinal cord.

White matter lesion number and location (periventricular, juxtacortical, deep WM, infratentorial and spinal) were recorded. Subsequently, dissemination in space and time, as well as MS diagnosis, were determined according to the revised 2017 McDonald criteria²⁰.

White matter lesion segmentation was performed with a semi-automated process using the 3D FLAIR and 3D T1-weighted MRI sequences²¹ All lesion masks were checked, corrected manually and lesion volumes subsequently obtained. These lesion masks were registered to the 3D T1-weighted sequence and the registered T1 lesion masks were used for lesion filling²², after which cortical gray matter (CGM) was automatically segmented and parcellated using GIF²³. All segmentations were quality checked. Ninety-eight cortical areas were identified in the native-space bilateral CGM according to the Desikan-Killiany-Tourville protocol

(http://braincolor.mindboggle.info)²⁴.

For computational reasons, CGM segmentation on the 3D T1-weighted images was resliced to 2.0 mm isotropic voxels to reduce the dimensionality of the data⁸.

Single-subject gray matter graphs

SCNs were extracted at subject level using software developed by T.B.M

(https://github.com/bettytijms/Single_Subject_Grey_Matter_Networks) in Matlab v7.12.0.635. Briefly, first, the SCNs' nodes were defined as regions-of-interest in the native space CGM segmentation, corresponding to 3x3x3 voxel cubes. This was chosen because a 3-voxel size can capture the thickness and folding of the cortex²⁵ without generating large matrices of cubes. The similarity between all the cubes (nodes) in the network was determined with the correlation coefficient, which was collected in a matrix that was subsequently thresholded and binarized, so that the chance of having spurious correlations for all single-subject SCNs was < or = to 5%. The obtained graphs were undirected with nodes connected each other by edges, only if their similarity matrix survived this threshold (Figure 1).

Graph properties

First, we determined the following global graph-defining properties: the size of the SCNs (i.e., the number of cubes), the average degree (i.e., the number of edges) and the connectivity density (i.e., the proportion of existing connections to the maximum number of possible connections). Then, we computed the global graph parameters averaged across all cubes: the characteristic path length²⁶, the clustering coefficient¹⁶, the betweenness centrality²⁷ and the eigenvector centrality²⁸. In addition, the clustering coefficient and characteristic path length of each graph were normalised with those of 20 randomised reference graphs of identical size and degree distribution⁸. The smallworld coefficient was then obtained from the ratio between the normalised clustering coefficient (γ) and the normalised characteristic path length (λ). A network shows a small-world properties when its nodes are highly clustered, as in regular lattices, but they can also be reached from every other node in a small number of steps, e.g. in random graphs, resulting in a ratio between γ and λ greater than 1²⁹.

At regional level, we averaged the local graph parameters across the nodes within each of the 98 cortical areas, which had been parcellated with GIF in each individual subject, in order to reduce the dimensionality of the SCNs and increase the comparability between subjects, across centers and possibly with previous studies³⁰.

Statistical analysis

Group differences in demographic characteristics and CGM volume were assessed using twosample t-tests for continuous variables and chi-square tests for categorical variables. We initially applied a multivariate linear regression model to compare graph properties between groups adjusting for age and gender and CGM volume. Since non-normalised global graph parameters (betweenness centrality, eigenvector centrality, clustering coefficient and path length) are dependent on graph-defining properties (size, connectivity density and degree), we included any of these graph-defining properties when they were significantly different between groups as additional independent variables to the model. We tested possible nonlinear effects of age with quadratic regression analysis imputing graph parameters as dependent variables. Since the analysis of the global graph parameters was meant to be exploratory, we did not perform multiple comparison adjustments.

Group differences in local graph parameters in the 98 CGM areas were assessed with linear regression, correcting for age, gender and local CGM volume). Due to concerns about residual variance heterogeneity across centers, the regression was estimated using the Hubert-White^{31,32} standard error, which is known to be robust to residual heterogeneity. Significant P values were identified correcting for multiple comparisons using a False Discovery Rate (FDR) of 5% for multiple hypothesis testing³³.

WM lesion volumes were log transformed to allow parametric testing³⁴ and the relationships between these and global/local graph metrics as well as clinical parameters (EDSS, SDMT, MS diagnosis according to the 2017 McDonald criteria) were assessed using linear regression adjusting for age and gender. Associations between local/global graph metrics changes and clinical parameters were assessed with linear regression adjusting for age and gender. The WM lesion volume was included as an additional independent variable to the model if associated with the clinical parameters.

Since CIS patients are heterogeneous in terms of the number of WM lesions, an established risk factor for conversion to MS, we also divided patients into distinct groups according to the number of intracranial WM lesions: 0-1, 2-9 and >10 lesions (Table 1). This allowed further comparisons between HCs and patients at different risk stage of developing MS^{35} .

Not all centers provided both HCs and patients, hence we investigated and controlled for center effects in the following manner. In addition to entering the center as a covariate in a linear regression including all centers, with graph metrics as dependent variables and age, gender and CGM volume as the other covariates, we performed a further analysis restricted to those centers contributing both patients and HCs (Basel, Bochum and London). If the between-group graph parameter differences or p-values were materially altered by inclusion of center in either the full or the restricted analyses, center was retained as covariate in the unrestricted model over all centers. Statistical analyses were performed with Stata v. 14.1 (Stata Corporation, College Station, Texas, USA).

RESULTS

Subject Characteristics

The clinical and radiological characteristics of the participants (60 patients and 38 healthy controls) are provided in Table 1. CIS patients and HCs did not show significant differences in age, gender and CGM volumes. CIS patients had low disability and were cognitively preserved. Out of the entire sample of 60 patients, eight fulfilled the 2017 McDonald criteria for MS.

Participants' characteristics across centers are detailed in eTable1 in Supporting Information.

Global graph metrics

All single-subject SCNs did not exhibit disconnected nodes.

Both CIS and HC groups retained small-world topology ($\gamma/\lambda > 1$), but in CIS patients the multivariate regression analysis (R²=0.06, F(4, 92)=2.38, p<0.05) showed a higher γ than HCs (B= 0.01, 95% CIs=0.003-0.02, p=0.012). CIS patients showed also (R²=0.11, F(4, 92)=3.41, p<0.05) a higher small-world coefficient than HCs (B= 0.01, 95% CIs=0.004-0.02, p=0.003). Age and age squared did not have any effects on the global graph metrics.

After correction for center effect, the differences between CIS patients and HCs for small-world coefficient remained significant (p=0.024) but not for γ (Table 2).

Center/scanner effect was also significant for connectivity density (p<0.0001) and for all global graph parameters (p<0.0001). Therefore, center was retained as covariate in the multivariate regression analysis. See eTable2 in Supplemental Material for the SCNs parameters across centers. For global graph-defining properties, CIS patients had smaller network sizes compared with HCs (Table 2); therefore, in order to compare SCNs between CIS and HC groups, we adjusted the analysis of the non-normalised global graph parameters for the network size. CIS patients initially showed higher eigenvector centrality than controls (B=0.07, 95% CIs=0.01-0.1, p=0.02), but this difference lost significance after adjusting for network size (Table 2).

In CIS patients, a higher small-world coefficient was associated ($R^2=0.17$, F(3, 55)=3.64, p<0.05) with higher WM lesion volume ($\beta=0.36$, p=0.007), after adjusting for age ($\beta=0.03$, p=0.832) and gender ($\beta=-0.18$, p=0.162); additionally, CIS patients with higher lesion counts (>10 lesions) showed higher gamma and greater small-world coefficients compared with HCs and with CIS patients with lower lesion numbers (Figure 2). The small-world coefficient in CIS patients and the WM lesion location were not significantly associated.

Amongst the 47 CIS patients with SDMT scores available, a higher small-world coefficient was associated ($R^2=0.21$, F(3,43)=3.86, p<0.05) with a lower SDMT score (i.e. slower information processing speed) (β = -0.39, p=0.008) after adjusting for age (β = -0.81, p=0.42) and gender (β = 0.85, p=0.4) (Figure 3). There were too few MS patients (n=9) in our cohort to allow comparisons with HCs and other CIS patients not diagnosed with MS.

Local graph metrics

We investigated whether differences between CIS patients and HCs were specific to particular cortical areas. Patients and controls showed similar volumes in the 98 CGM areas and there were no significant differences in degree. Several cortical areas showed significant differences for betweenness centrality, clustering coefficient and path length (p<0.05), but only three survived correction for multiple testing (see eTable3 in Supplemental Material). Patients showed lower betweenness centrality in the right superior frontal gyrus (0.30 ± 0.01 vs 0.33 ± 0.01 , p<0.0001;

 p_{FDR} <0.05) and lower clustering coefficient in the right postcentral gyrus (0.51±0.04 vs 0.54±0.03, p<0.0001; p_{FDR} <0.05) in comparison to HCs (Figure 4). Among CIS patients, these alterations were not correlated with clinical parameters or WM lesion volume.

Finally, CIS patients with >10 WM lesions had lower betweenness centrality in the right postcentral gyrus ($0.29\pm0.01\ 0.33\pm0.01$, $p_{FDR}<0.05$) and lower clustering coefficient in the right medial orbital gyrus (0.51 ± 0.04 vs 0.53 ± 0.03 , $p_{FDR}<0.05$), compared with HCs.

There was significant (p_{FDR} <0.05) variability across centers in the following local graph properties: characteristic path length in the bilateral anterior insula, left inferior and middle temporal gyrus, left parahippocampal gyrus; clustering coefficient in the right inferior and middle occipital gyrus; betweenness centrality in the left posterior cingulate cortex and eigenvector centrality in the right middle occipital gyrus.

DISCUSSION

In this study, using single-subject SCN analysis, we showed that in CIS patients, graph alterations are related to disease burden. At the global level, CIS patients possess altered small-world topology characterized by a shift towards a more regular network. These findings are in line with previous studies on SCNs in CIS patients¹⁴. A regular network tends to possess dense local connections (i.e. high clustering) between nodes at the expenses of the long-distance ones (i.e. low path length), thus compromising the efficient balance between short and long-range information transfer, distinctive of small-world networks¹⁶.

When comparing the whole CIS cohort with HCs, the difference in small-world coefficient was significant, but small. This small difference could be because CIS patients were at the very onset of the disease, some of them with little WM damage. When patients were stratified according to the number of WM lesions, this difference in the small-world coefficient increased: subjects with higher lesion load at onset showed higher small-world coefficient values compared not only with HCs, but also with the group of CIS patients with 0-1 lesion.

This altered topology was correlated with a high lesion load. Moreover, we did not find any significant differences in cortical volumes between groups: in CIS patients the tendency to lose long-distance connections could be a consequence of altered connectivity due to lesional disruption of WM tracts more than to alterations in cortical areas.

The increase in the small-world coefficient, in our study, was associated with worse cognitive performance. Noteworthy, the analysis of single-subject SCNs in established MS¹⁵, similar to studies conducted with the same methodology in dementia¹⁰, showed that cognitively impaired patients tend to possess random SCN topology. However, unlike individuals in our cohort, those patients already had evidence of definite gray matter atrophy. This suggests that small-world topology alterations of the SCNs may have different characteristics at different stages of the disease, being more regular in the early phases and more random when neurodegeneration starts to occur.

At local level, CIS patients had altered network properties in several cortical region, including low clustering coefficient in areas relevant to cognition. These alterations do not seem to affect the global SCNs properties and the clinical phenotype. However, studies with further cognitive assessments and longitudinal follow-up could investigate the clinical relevance of these findings. In our study, SCNs differed in size between CIS patients and HCs, although the two groups did not show any difference in CGM volume. It is possible that the smaller size of patients' SCNs could result from the different MRI pre-processing, particularly the WM lesion segmentation in patients, or from the presence of intra-cortical lesions. Most studies use methods that constrain networks into identical sizes (i.e. parcellated cortical areas), whereas, here, we chose to extract SCNs from native space segmentations, thus preserving the inter-individual variability. However, it is still not clear how to compare graphs with different sizes and/or connectivity densities³⁰. Here, we adjusted for SCN size in our statistical models, but after this, we could not detect any differences in non-normalised global graph-parameters between CIS patients and HCs.

One final consideration is the multi-center setting of our study giving rise to inter-center variability. Inter-scanner variations had significant effects on the global graph parameters. This may have been due to differences in WM-GM interface contrast across different scanners that influenced the CGM segmentation outcomes. Our findings suggest that it is important to account for inter-center variability for single-subject SCN extraction from multi-center neuroimaging datasets. The advantage of multi-center settings though is the putative increase in statistical power and, although variance heterogeneity across centers may be a potential problem for our analysis, the relative robustness of the small-world result suggests this may not in our case be a major limitation.

CONCLUSIONS

In this study, we applied, for the first time, a recently developed method to extract single-subject SCNs in CIS. Using clinical scans, we were able to demonstrate altered small-world topology in CIS patients SCNs that was associated with cognitive processing. This suggests that SCNs are affected early in MS and can represent a potential biomarker. Future longitudinal analyses will assess if the observed changes in SCNs can predict conversion to MS during the follow-up.

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CONFLICT OF INTEREST STATEMENT

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