Comparative Causality Analyses between Hydrological Natural Inflow and Climate Variables in Brazil

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Abstract

5 Numbers of studies have proved the significant influence of climate variables on hydrological series. Considering the pivotal role of the hydroelectric power plants play in the electricity 6 production in Brazil this paper considers the natural hydrological inflow data from 15 major 7 basins and 8 climate variables containing 7 El Niño Southern Oscillation proxies and the 8 sunspot numbers. The causal relationships between hydrological natural inflows and climate 9 variables are investigated by adopting and comparing 5 different causality detection meth-10 ods (Granger Causality test, Frequency Domain Causality test, Convergent Cross Mapping 11 Causality test, Single Spectrum Analysis (SSA) Causality test and Periodic Autoregressive 12 Model Causality test) that cover both well established and novel empirical approaches. Both 13 time domain and frequency domain causality tests gain valid evidences of unidirectional 14 causality for a group of series; CCM achieved unidirectional causality for 18% of pairs and 15 overwhelmingly indicated the opposite direction of causality; a mixture of results are con-16 cluded by SSA causality test; PAR based causality test obtained six unidirectional causality, 17 but only one is really significant. 18

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Keywords: Hydrological Natural Inflow; Climate Variables; Causality Detection; Granger
 Causality; Frequency Domain Causality; Convergent Cross Mapping; Single Spectrum Anal ysis; Periodic Autoregressive Model.

23 Nomenclature

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	CCM	Convergent Cross Mapping.
	EDM	Empirical Dynamic Modeling.
	ENSO	El Niño-Southern Oscillation.
	GC	Granger Causality Test.
	NGDC	National Geophysical Data Center.
	NOAA	National Oceanic and Atmospheric Administration.
4	ONI	Oceanic Niño Index.
	PAR	Periodic Autoregressive Model.
	PARX	Periodic Autoregressive Model with One Exogenous Variable.
	RMSE	Root Mean Square Error.
	SOI	Southern Oscillation Index.
	SSA	Singular Spectrum Analysis.
	SST	Sea Surface Temperature.

25 1 Introduction

In Brazil there are 1268 hydroelectric power plants in operation, corresponding to 65% of total installed capacity and responsible for 73% of electricity production in 2016 [1]. This kind of power plant produces electricity by harnessing a river hydraulic potential so the electricity generation depends directly on hydrological regimes.

Since the 90s there are several studies showing that not only there is an influence of climate variables like El Niño-Southern Oscillation (ENSO) on hydrological series [2–6,8], but also that when correlation is taken into account there is improvement in the forecasting/modelling exercise of inflow time series [9–13], for instance, storm tides data at the Baltic Sea in [14] and streamflow data of the East River basin of China in [15] by adopting the significant Hurst exponent [16], which has also been applied in birth time series [17]. Another recent research considered Hurst exponent in analyzing hydrogeological series can be found in [18].

This paper aims to establish comprehensive causality analyses between natural inflow and climate variables in Brazil by embracing and comparing both well established and advanced causality detection methods, including time domain Granger causality (GC) test [19], frequency domain causality test [20], Convergent Cross Mapping (CCM) [21], Singular Spectrum Analysis (SSA) based causality test [22] and the Periodic Autoregressive model (PAR) based causality test [23,24].

Most of the works previously cited study the influence of ENSO events using the Sea Surface
Temperature (SST) variable for the Northeast region of Brazil, ignoring others geographic regions
and also other variables that possibly indicate a proxy for ENSO. In this paper, all the fifteen
Brazil major basins are considered to test the causality with more than seven ENSO proxies
and the Sunspot climate event.

The remainder of this paper is organized such that the background of this study is presented in Section 2; the causality detection techniques adopted in this paper are briefly summarized in Section 3; Section 4 introduces the data and summarizes the descriptive statistics along with
correlation analyses; the detailed causality test results by different methods are listed in Section
5; the paper concludes in Section 6 with proposals of future research.

53 2 Background

It is possible to find several studies that identify the influence of ENSO events in the Brazilian 54 river basins, but none of them apply any type of causality test. Amarasekera et al. [2] concludes 55 that the annual discharges of the Amazon river is weakly and negatively correlated with the 56 equatorial Pacific Sea Surface Temperature (SST) anomaly, while the Paraná river shows 57 a strong and positive correlation. Dettinger & Diaz [3] uses El Niño variations to characterize 58 geographic differences in the seasonality and year-to-year variability of stream flow from several 59 sites around the world, and shows that the Amazon basin is drier-than-normal in El Niño 60 years accordingly to Southern Oscillation Index (SOI) and North Atlantic Oscillation (NAO) 61 [25] index. Foley et al. [4] shows that during the El Niño there is a decrease in the Amazon 62 and Tocantins river discharge, and the opposite during the La Niña. Berri et al. [5] presents 63 that exactly the opposite happens in Paraná river, i.e., during El Niño the average inflow are 64 always larger than those observed during La Niña events. Garcia & Mechoso [6] concludes 65 that the Amazon, Tocantins, San Francisco, Paraguay, Paraná river streamflows shows El Niño-66 like periodicities. Soares et al. [8] notice that the sub-basins of the southern Brazilian regions 67 showed positive variations in water production during El Niño, while the Amazon basin showed 68 no response. 69

Souza Filho [9] shows that the correlation between climate and hydrological variables is 70 beneficial for the prediction of reservoirs inflows in Ceará. Cardoso & Silva Dias [10] use the 71 SST index to show that there is improvement in the reservoir inflow forecasting of Paraná River. 72 Lima & Damien [11] apply dynamic linear models to predict the inflow of the Brazil fifteen main 73 basins using precipitation and an El Niño index. Macaira et al. [12] developed a causal PAR 74 model to estimate the influence between several El Niño indices and the inflow time series of 75 some Brazilian locations. Silveira et al. [13] propose the Periodic Autoregressive model with 76 one exogenous variable (PARX) to simultaneously predict all natural inflows of the National 77 Interconnected System. 78

Apart from the significance of studying the causal links between natural inflow and climate factors, this paper has adopted 5 different causality detection techniques covering both well established and advanced time series analysis methods (note that the detailed introduction of these methods are summarized in section 3). It worth to be noted as another contribution of this paper that it comprehensively investigates the causal relationship with the most up to date time series analysis techniques to our knowledge.

The well established and widely applied GC approach enables researchers to evaluate dependence relationship, mostly linear, among factors in a complex system. It brings insights on

⁸⁷ whether the changes of one factor have relationship with the changes of another factor in the

current sequence or after specific lag of time. However, it assumes linearity and separability for 88 the selected variables in the model and the nonlinear applicability is limited. The frequency do-89 main causality test extends the GC approach to identify the causality for each frequency instead 90 of a single statistics for the whole time series, whilst the restricted assumptions and nonlinear 91 applicability maintain. In addition, by adopting the advanced time series analysis techniques 92 like SSA and CCM, this paper also explores the causality detection from the aspect of nonlin-93 earity and other complex dynamics. These advanced non-parametric techniques are relatively 94 new and have no assumptions of linear or restricted nonlinear model. They are designed to be 95 widely applicable and assumption free with straightforward way of thinking and implementing. 96 By adopting these advanced methods, this paper seeks to further distinguish possible causal 97 relationships that the empirical tests cannot achieve or fall short at. In general, to the best of 98 our knowledge, this paper is the first attempt that applies and compares all these five causality 99 detection methods to date. Moreover, for most of the advanced methods, it is also the initial 100 implementation study on the natural inflow and climate variables in Brazil. 101

¹⁰² 3 Causality Detection Methods

¹⁰³ 3.1 Time Domain Granger Causality Test

GC test [19] is the most generally accepted and significant method for causality analyses in various disciplines. The regression formulation of granger causality states that vector X_i is the cause of vector Y_i if the past values of X_i are helpful in predicting the future value of Y_i , two regressions are considered as follows:

$$Y_i = \sum_{t=1}^{T} \alpha_t Y_{i-t} + \varepsilon_{1i}, \qquad (1)$$

$$Y_i = \sum_{t=1}^T \alpha_t Y_{i-t} + \sum_{t=1}^T \beta_t X_{i-t} + \varepsilon_{2i}, \qquad (2)$$

where $i = 1, 2, \dots, N$ (N is the number of observations), T is the maximal time lag, α and β are vectors of coefficients, ε is the error term. The first regression is the model that predicts Y_i by using the history of Y_i only, while the second regression involves both X_i and Y_i . Therefore, the conclusion of existing causality is conducted if the second model is a significantly better model than the first one.

113 3.2 Frequency Domain Causality Test

The frequency domain causality test is the extension of time domain GC test that identifies the causality between different variables for each frequency. In order to briefly introduce the testing methodology, we mainly follow [20, 26].

It is assumed that two dimensional vector containing X_i and Y_i (where $i = 1, 2, \dots, N$ and Nis the number of observations) with a finite-order Vector Auto-regression Model representative 119 of order p,

$$\Theta(R) \begin{pmatrix} Y_i \\ X_i \end{pmatrix} = \begin{pmatrix} \Theta_{11}(R) & \Theta_{12}(R) \\ \Theta_{21}(R) & \Theta_{22}(R) \end{pmatrix} \begin{pmatrix} Y_i \\ X_i \end{pmatrix} + \mathcal{E}_i,$$
(3)

where $\Theta(R) = I - \Theta_1 R - ... - \Theta_p R_p$ is a 2×2 lag polynomial and $\Theta_1, ..., \Theta_p$ are 2×2 autoregressive parameter matrices, with $R^k X_i = X_{i-k}$ and $R^k Y_i = Y_{i-k}$. The error vector \mathcal{E} is white noise with zero mean, and $E(\mathcal{E}_i \mathcal{E}'_i) = \mathbf{Z}$, where \mathbf{Z} is positive definite matrix. The moving average representative of the system is

$$\begin{pmatrix} Y_i \\ X_i \end{pmatrix} = \Psi(R)\eta_i = \begin{pmatrix} \Psi_{11}(R) & \Psi_{12}(R) \\ \Psi_{21}(R) & \Psi_{22}(R) \end{pmatrix} \begin{pmatrix} \eta_{1i} \\ \eta_{2i} \end{pmatrix},$$
(4)

with $\Psi(R) = \Theta(R)^{-1} \mathbf{G}^{-1}$ and \mathbf{G} is the lower triangular matrix of the Cholesky decomposition $\mathbf{G'G} = \mathbf{Z}^{-1}$, such that $E(\eta_t \eta'_t) = I$ and $\eta_i = \mathbf{G}\mathcal{E}_i$. The causality test developed in [26] can be written as:

$$C_{X \Rightarrow Y}(\gamma) = \log \left[1 + \frac{|\Psi_{12}(e^{-i\gamma})|^2}{|\Psi_{11}(e^{-i\gamma})|^2} \right].$$
 (5)

However, according to this framework, no Granger causality from X_i to Y_i at frequency γ corresponds to the condition $|\Psi_{12}(e^{-i\gamma})| = 0$, this condition leads to

$$|\Theta_{12}(e^{-i\gamma})| = |\Sigma_{k=1}^{p}\Theta_{k,12}\cos(k\gamma) - i\Sigma_{k=1}^{p}\Theta_{k,12}\sin(k\gamma)| = 0,$$
(6)

where $\Theta_{k,1,2}$ is the (1,2)th element of Θ_k , such that a sufficient set of conditions for no causality is given by [20]

$$\Sigma_{k=1}^{p}\Theta_{k,1,2}\cos(k\gamma) = 0$$

$$\Sigma_{k=1}^{p}\Theta_{k,1,2}\sin(k\gamma) = 0$$
(7)

Hence, the null hypothesis of no Granger causality at frequency γ can be tested by using a standard F-test for the linear restrictions (7), which follows an F(2, B - 2p) distribution, for every γ between 0 and π , with B begin the number of observations in the series.

134 3.3 Convergent Cross Mapping

CCM is firstly introduced in [21] that aimed at detecting the causation among time series 135 and provide a better understanding of the dynamical systems that have not been covered by 136 other well established methods like GC. CCM has proven to be an advanced non-parametric 137 technique for distinguishing causations in a dynamical system that contains complex interactions 138 in ecosystems and climate studies [21,27], more details can be found in [28,29]. Some significant 139 rationales of embracing this advanced technique include: CCM is non-parametric approach 140 with no restrictions of assumptions for parametric methods; CCM can distinguish statistically 141 significant causality by considering only two key variables instead of building a complex model 142 by incorporating many possible influential variables based on regression modelling; CCM has 143 remarkable sensitivity at detecting causal links within complex systems whilst not being limited 144

to linearity or nonlinearity; the calculation itself is efficient and comparatively straight forward.
CCM is briefly introduced below by mainly following [21].

Assume there are two variables X_i and Y_i , for which X_i has a causal effect on Y_i . CCM test will test the causation by evaluating whether the historical record of Y_i can be used to get reliable estimates of X_i . Given a library set of n points (not necessarily to be the total number of observations N of two variables) and here set $i = 1, 2, \dots, n$, the lagged coordinates are adopted to generate an E-dimensional embedding state space [30,31], in which the points are the library vector X_i and prediction vector Y_i

$$X_i : \{x_i, x_{i-1}, x_{i-2}, \cdots, x_{i-(E-1)}\},$$
(8)

$$Y_i : \{y_i, y_{i-1}, y_{i-2}, \cdots, y_{i-(E-1)}\},$$
(9)

The E+1 neighbors of Y_i from the library set X_i will be selected, which actually form the smallest simplex that contains Y_i as an interior point. Accordingly, the forecast is then conducted by this process, which is the nearest-neighbour forecasting algorithm of simplex projection [31]. The optimal E will be evaluated and selected based on the forward performances of these nearby points in an embedding state space.

Therefore, by adopting the essential concept of Empirical Dynamic Modeling (EDM) and 158 generalized Takens' Theorem [30], two manifolds are conducted based on the lagged coordinates 159 of the two variables under evaluation, which are the attractor manifold M_Y constructed by Y_i 160 and respectively, the manifold M_X by X_i . The causation will then be identified accordingly 161 if the nearby points on M_Y can be employed for reconstructing observed X_i . Note that the 162 correlation coefficient ρ is used for the estimates of cross map skill due to its widely acceptance 163 and understanding, additionally, leave-one-out cross-validation is considered a more conservative 164 method and adopted for all evaluations in CCM. 165

¹⁶⁶ 3.4 Singular Spectrum Analysis based Causality Test

As GC formalized the causality concept and claimed causality if the elimination of one variable 167 from a system is harmful for explaining the other variable. Similarity, as can be seen in Figure 1, 168 the SSA based causality analysis is obtained by comparing the forecast values obtained by the 169 univariate procedure–SSA and multivariate process– multivariate SSA (MSSA). Consequently, 170 if the forecasting errors using MSSA are significantly smaller than those of univariate SSA, it is 171 concluded that there is a causal relationship detected between these series. As a nonparametric 172 technique, the SSA causality test is able to capture possible nonlinearities using a data-driven 173 approach without specifying any known functional nonlinear model to the relationship, which in 174 turn, could be incorrectly specified in the first place. Detailed introduction is presented below 175 which mainly follow [22, 32], where also summarize the details of SSA and MSSA formulation 176 and forecasting algorithms. 177

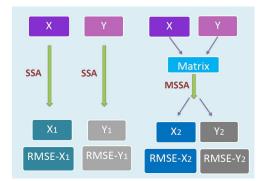


Figure 1: Flowchart of Cause Detection based on SSA Forecasting Accuracy.

Let us consider the procedure for constructing vectors of forecasting error for out-of-sample 178 tests in a two variable case X_N and Y_N by both univariate and multivariate SSA techniques 179 respectively. Firstly, the series $X_N = (x_1, ..., x_N)$ is divided into two separate subseries X_R 180 and X_F that satisfy $X_N = (X_R, X_F)$, where $X_R = (x_1, ..., x_R)$ and $X_F = (x_{R+1}, ..., x_N)$. Same 181 procedure is also conducted for Y_N . The subseries X_R and Y_R are used in the reconstruction step 182 to provide the noise-free series X_R and Y_R . The noise-free series are then used for forecasting the 183 subseries X_F and Y_F with the help of the forecasting algorithms (see Appendix A) of SSA and 184 MSSA respectively. For variable X_N , two different forecasting values of $X_F = (\hat{x}_{R+1}, ..., \hat{x}_N)$ 185 by SSA and MSSA are then used for computing the forecasting errors accordingly, which will 186 be the same process in terms of variable Y_N . Therefore, in a multivariate system like this, the 187 vectors of forecasts obtained can be used in computing the forecasting accuracy and therefore 188 conducting the causality analysis between the two variables. 189

The length of out-of-sample does not have specific limitation, generally considering the sim-190 ulation scenario, the length of time series for reconstruction will take 2/3 of the whole series 191 and the rest 1/3 is considered as out-of-sample for constructing forecasting error. The separate 192 point to define the out-of-sample size for different series can be chosen respectively, whilst it 193 is important that when it goes to comparing the performances of different techniques based on 194 constructed forecasting error of one specific series, the sizes of reconstruction and out-of-sample 195 for all techniques should be identical. In addition, the choices of window length L and the 196 referring options of numbers of eigenvalues r should also be carefully evaluated in practice of 197 SSA causality test respectively. Considering this as the first attempt of application, also in 198 order to conduct the most accurate results, all the possibilities of L and its referring choices of r199 should be applied for both univariate SSA and MSSA processes, then the optimal ones with best 200 performance of forecasting will be chosen to construct the finally cause detection procedure. 201

²⁰² Consequently, define the criterion $F_{X|Y} = \Delta X_{F|Y} / \Delta X_F$ corresponding to the forecast of the ²⁰³ series X_N in the presence of the series Y_N . Specifically, if $F_{X|Y}$ is small, then having information ²⁰⁴ obtained from the series Y can help to achieve better forecasts of the series X. If $F_{X|Y} < 1$, it is ²⁰⁵ concluded that the information provided by the series Y can be regarded as useful or supportive ²⁰⁶ for forecasting the series X. Alternatively, if the values of $F_{X|Y} \geq 1$, then either there is no detectable causality between X and Y or the performance of the univariate SSA is better than of the MSSA (this may happen, for example, when the series Y has structural breaks misdirecting the forecasts of X).

210 3.5 Periodic Autoregressive Model based Causality Test

To perform monthly forecasts and simulation of hydrological series, the classical PAR model has 211 been widely used [23]. This type of model adjusts the series using the estimated parameters of 212 the historical data [33], and does not consider any exogenous information that could affect the 213 hydrological regimes in equation (10). To consider any exogenous variable in the PAR model, 214 there is the Periodic Autoregressive model with one exogenous variable (PARX) as presented in 215 equation (11). PAR models fit for each season an autoregressive term being able to capture the 216 monthly variability of hydrological regimes, this is the main reason for its success for this type 217 of data. The mathematical details of PAR and PARX can be found in [12, 13, 24]. 218

$$\left(\frac{Y_i - \mu_m}{\sigma_m}\right) = \sum_{t=1}^{p_m} \varphi_t^m \left(\frac{Y_{i-t} - \mu_{m-t}}{\sigma_{m-t}}\right) + \varepsilon_t,\tag{10}$$

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$$\left(\frac{Y_i - \mu_m}{\sigma_m}\right) = \sum_{t=1}^{p_m} \varphi_t^m \left(\frac{Y_{i-t} - \mu_{m-t}}{\sigma_{m-t}}\right) + \sum_{t=0}^{v_m} \vartheta_t^m X_{i-t} + \varepsilon_t, \tag{11}$$

Where μ_m and σ_m are the average and the standard deviation of season m, respectively; φ_t^m is the *t*-th autoregressive coefficient of season m, p_m is the order of the autoregressive operator of season m. In (11), X_i is the predictor variable, ϑ_t^m is the autoregressive coefficient and v_m is the order of the autoregressive operator of season m for the predictor variable.

Similar to the SSA based Causality Test, it was developed the PAR based Causality Test that compares the forecasts values obtained by the univariate process PAR and the PARX. If the forecasting errors using PARX are significantly smaller than those of PAR, it is conclude that there is a causality detected among the variables.

228 4 Data

229 4.1 The Natural Inflow Series in Brazil

According to the Brazilian Electricity Regulatory Agency (ANEEL) there are fifteen major river basins in Brazil, with an installed capacity of approximately 90 GigaWatts [GW] in 2016, representing 66% of the total installed capacity in the country (Figure 2). The Parana river basin has the highest hydroelectric potential, around 43 GW, which represents 48% of total hydroelectric capacity. It can be further subdivided into six minor basins based on its major rivers - Paranaiba, Grande, Tiete, Parana, Paranapanema and Iguacu.



Figure 2: Major rivers basins in Brazil.(Source: [11])

The historical data available is the natural inflow¹ for each generator, on a monthly basis, starting in January 1931 and ending in December 2015, measured in cubic meters per second $[m^3/s]$. For generators built after 1931, the National Electric System Operator performs a backward forecasting in order to standardize the records for the hydrothermal dispatch optimization process.

In the major rivers there are around 164 hydroelectric power plants currently in operation [34], and these plants operate in a cascade scheme, see in Figure 3 this cascade scheme for Paranaíba and Grande basins with 19 generators with reservoirs, represented by triangles, and 15 generators with no reservoir (circles). This way decisions taken at the upstream reservoirs will impact the inflow of the downstream reservoirs.

¹The natural inflow is the average incoming water per unit of time at each generator's reservoir from affluent rivers, lakes and its own drainage area.

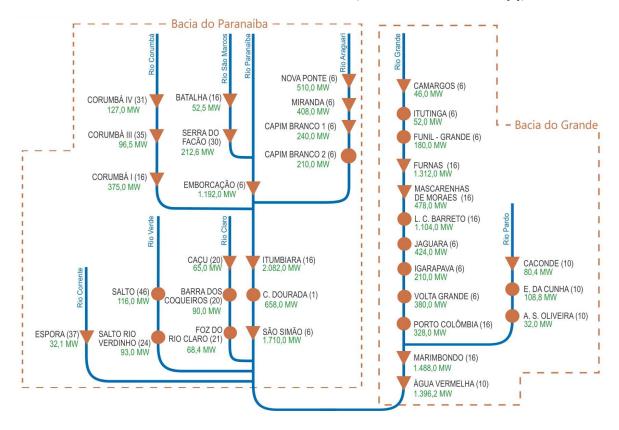


Figure 3: Example of a cascade scheme. (Source: Adapted from [1])

Since there is a cascade scheme, the natural inflow of each generator has to be calculated 246 based on the concept of incremental inflows. For exemplifications reason, assume that Camargos 247 is Generator number 1 and Itutinga is Generator number 2 in Figure 3. If Generator 1 comes 248 first in the cascade, the incremental inflow will be equal to its natural inflow. But, if Generator 249 2 has 1 upstream, so its incremental inflow will be given by the difference between its natural 250 inflow and the natural inflow of Generator 1. The generators will be grouped by basin creating 251 an equivalent generator with natural inflow equal to the sum of the incremental inflows of all 252 reservoirs belonging to the basin (Figure 4). It is of note that all natural inflow data analyzed 253 in the following sections have been adjusted accordingly considering the cascade scheme. 254

255 4.2 Climate Variables

The climate variables were selected trough a literature search. The selected variables are related to El Niño and the Sunspots numbers; the variables representing El Niño/La Niña phenomenon are: Southern Oscillation Index (SOI), Equatorial SOI, Niño variations and Oceanic Niño Index (ONI).

The SOI is calculated based on the difference between the atmospheric pressure at sea level in the regions of Tahiti (in the Western Pacific) and Darwin (Australia, Western Pacific) [35]. The Equatorial SOI measures the average difference of atmospheric pressure at sea level between two regions centered on the equator: Indonesia and East Pacific. The range to indicate the presence

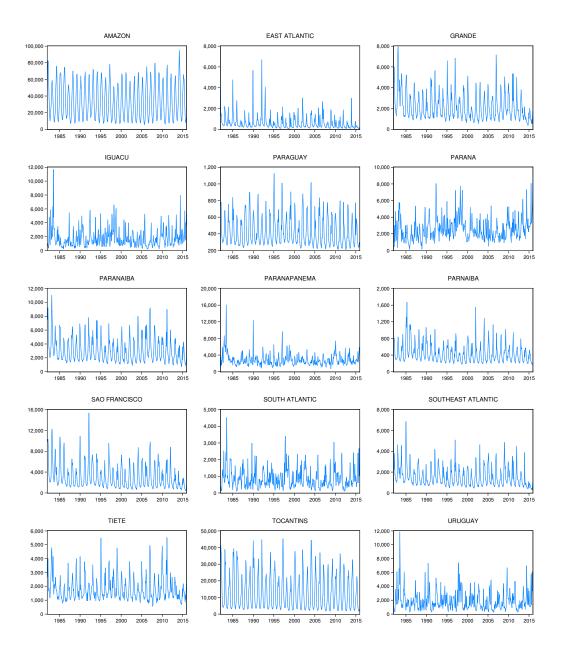


Figure 4: The natural inflow series in Brazil.

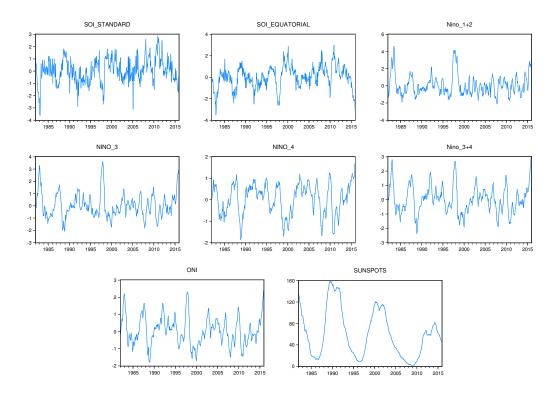


Figure 5: The climate variables in Brazil.

or absence of El Niño/La Niña is the same for both the SOI index and Equatorial SOI. It is also of note that the influence of El Niño in North America indicate a nearly 30 year long cycle due to the different geographical zone [36]. Consecutive periods of negative figures indicate El Niño phenomenon occurrence; meanwhile consecutive positive figures denote the presence of La Niña and values close to zero indicate a normal situation, where none of the two phenomenons occur. The official historical monthly series of these indices are provided by the National Oceanic and Atmospheric Administration (NOAA).

The Sea Surface Temperature (SST) is the water temperature close to the ocean's surface. 271 The SST anomaly, that is, the temperature variation by month, is a proxy for El Niño and La 272 Niña. Thus, this index is used to classify and quantify such phenomena in four Niño regions: 273 Niño 1+2, Niño 3, Niño 4 and Niño 3.4, defined as follows by NOAA in 2014. Through the 274 location of the Niño regions it is possible to conclude that regions Niño 1+2 and Niño 3 better 275 identify temperature anomalies for the Eastern Pacific Ocean sea surface and region Niño 4 for 276 the Western Pacific. The Niño 3.4 region is centralized in the Pacific, which allows a better 277 understanding of anomalies across it. Therefore, currently the Niño 3.4 region is the official 278 measure used to represent SST. However, depending on the study, other regions may be a better 279 alternative. The threshold for the normal state of this index is between $-0.5^{\circ}C$ and $+0.5^{\circ}C$. The 280 criteria commonly used to define an El Niño phenomenon consists of five consecutive averages 281 of SST anomalies above $+0.5^{\circ}C$. Similarly, for La Niña, this criterion remains, but now the 282

SST anomaly should be below $-0.5^{\circ}C$. The monthly time series for all regions are provided by NOAA.

The ONI measures the average sea surface temperature anomalies for the region Niño 3.4, removing the existing warming trend on it. According to the NOAA website, multiple centered 30-year based periods are adopted for obtaining ONI values of five successive years. For instance, the 1956-1960 ONI values are based on the 1941-1970 period, while 1936-1965 base period produces the 1950-1955 ONI values. The El Niño and La Niña are indicated in the same manner as the SST index, the time series is monthly and is provided by NOAA.

Sunspots comprehend solar surface regions of high magnetic field, which have considerably lower temperature than its surroundings and thus appear as a dark area. The magnetic flux amount on the sun surface varies over eleven year periods, known as sunspot and solar cycles. During this cycle there is a minimum and a maximum magnetic flux, which is not only difficult to identify the sunspots and but also they appear almost all the time. The cycle reaches its maximum approximately every eleven years, therefore the observed cycle duration corresponds to eleven years.

The number of sunspots calculation is accomplished with the Relative Index American number of sunspots. This index indicates the solar phenomenon occurrence taking into account their relationship with the Earth, including geomagnetic variations and ionosphere effects. The Solar Division from American Association of Variable Star Observers coordinates the data collection program and the analysis of this phenomenon. Thus, the National Geophysical Data Center (NGDC), provides the historical data from the number of sunspots per month since 1749.

Considering the availability of all series and since the SST is only available after 1982, the data used for this paper are at monthly frequency from January 1982 to December 2015.

A brief summary table is listed below in Table 1 and the descriptive statistics can be found in Table 2.

	Abbreviation	Variable
	AMZ	Amazon
	EAT	East Atlantic
	GRA	Grande
	IGU	Iguacu
	P1	Paranaiba
	P2	Paranapanema
	P3	Parana
Natural Inflow Series	P4	Paraguay
	P5	Parnaiba
	SAT	South Atlantic
	SEAT	Southeast Atlantic
	\mathbf{SF}	Sao Francisco
	TIE	Tiete
	TOC	Tocantins
	URU	Uruguay
	SOLSt	Southern Oscillation Index Standard
	SOLEq	The Equatorial Southern Oscillation Index
	NN12	Sea Surface Temperature of Niño 1+2 Region
Climate Variables	NN3	Sea Surface Temperature of Niño 3 Region
Climate variables	NN4	Sea Surface Temperature of Niño 4 Region
	NN34	Sea Surface Temperature of Niño 3.4 Region
	ONI	The Oceanic Niño Index
	SS	Sunspots Number

Table 1: Summary of tested series.

Table 2: Descriptive statistics of data.

Series	Mean	Median	Maximum	Minimum	Std. Dev.	Skewness	Kurtosis	Obs
AMZ	33553.30	27753.5	95088	5600	21564.07	0.47	1.97	408
EAT	622.75	356	6690	45	746.51	3.67	22.33	408
\mathbf{GRA}	2155.66	1721	7938	364	1323.23	1.36	4.79	408
IGU	1802.17	1414	11670	206	1368.56	2.12	10.77	408
P4	425.12	352	1124	212	178.78	1.07	3.42	408
P3	2587.07	2309	8911	130	1368.81	1.35	5.69	408
P1	3094.21	2496	11025	705	1796.74	1.29	4.53	408
P2	2857.46	2504	16004	699	1521.65	3.12	20.81	408
P5	438.33	366.5	1668	178	232.30	1.70	6.71	408
SF	3152.20	2229	15360	406	2481.43	1.55	5.49	408
SAT	868.39	689.5	4524	110	616.66	1.64	7.01	408
SEAT	1535.49	1226	6862	319	975.27	1.56	5.97	408
TIE	1787.83	1549.5	5519	548	849.09	1.64	6.03	408
TOC	13064.57	8229	45317	1772	10943.19	0.93	2.81	408
URU	1939.78	1472	11834	262	1472.04	2.00	9.32	408
SOI_St	0.03	0	2.9	-3.6	1.01	-0.21	3.60	408
SOI_Eq	0.02	0.1	3	-3.5	1.03	-0.37	3.51	408
NN12	0.08	-0.18	4.62	-2.1	1.21	1.32	5.00	408
NN3	0.05	-0.13	3.62	-2.07	0.99	0.95	4.46	408
NN34	0.03	0.005	2.95	-2.38	0.97	0.41	3.25	408
NN4	0.04	0.19	1.67	-1.87	0.72	-0.50	2.49	408
ONI	0.05	-0.01	2.37	-1.78	0.84	0.36	3.00	408
SS	60.12	56.6	158.5	1.7	44.31	0.57	2.24	408

308 4.3 Correlation Analysis

Prior to the comparison of causality analyses by different methods, the correlation analysis results are here summarized in Table 3 and Table 4. Note that the results are Pearson correlation coefficients respectively considering the empirical status of Pearson approach and ** indicates

significance at the 0.01 level whilst * reflects the 0.05 level.

As can be seen in Table 3, the correlation between natural inflow and climate variables are overwhelmingly weak, except a few weak correlations detected among NN12, NN3, P2, P3, SAT and URU. The correlations between the climate series are also evaluated and summarized in Table 4. Similar conclusions are obtained as expected following the results in [12]: SOI indices hold negative correlation with the others, whilst the El Niño and ONI series indicate strong positive values.

Table 3: Correlation between natural inflow and climate variables.

	\mathbf{AMZ}	EAT	GRA	IGU	$\mathbf{P1}$	P2	P3	$\mathbf{P4}$	P5	SAT	SEAT	\mathbf{SF}	TIE	TOC	URU
SOI_St	0.06	-0.06	0.00	-0.07	-0.03	-0.09	-0.09	0.09	0.09	-0.13**	0.01	-0.03	0.01	0.06	-0.11*
SOI_Eq	0.03	-0.04	-0.05	-0.22**	-0.05	-0.24**	-0.23**	0.06	0.09	-0.27**	0.01	-0.04	-0.06	0.05	-0.26**
NN12	-0.03	-0.05	0.08	0.37^{**}	0.04	0.36^{**}	0.38^{**}	-0.03	-0.13**	0.33^{**}	-0.03	0.02	0.14^{**}	-0.07	0.37^{**}
NN3	-0.06	-0.03	0.03	0.28^{**}	0.04	0.26^{**}	0.31^{**}	-0.08	-0.14**	0.32^{**}	-0.05	-0.01	0.09	-0.09	0.33^{**}
NN4	-0.03	0.02	-0.06	0.11*	0.01	0.07	0.19^{**}	-0.05	-0.08	0.16^{**}	-0.07	-0.02	-0.02	-0.03	0.16^{**}
NN34	-0.06	0.01	0.01	0.20**	0.03	0.18^{**}	0.25^{**}	-0.08	-0.12*	0.26^{**}	-0.05	-0.01	0.05	-0.07	0.25^{**}
ONI	-0.07	0.01	0.01	0.19^{**}	0.04	0.18^{**}	0.23^{**}	-0.09	-0.11*	0.28^{**}	-0.04	0.01	0.04	-0.07	0.27^{**}
SS	0.00	0.10*	-0.02	0.04	0.02	0.06	0.02	0.05	0.00	0.01	-0.06	0.03	0.05	0.05	0.00

Table 4: Correlation between climate variables.

	SOI_St	SOI_Eq	NN12	NN3	NN4	NN34	ONI	\mathbf{SS}
SOI_St	1.00							
SOI_Eq	0.80**	1.00						
NN12	-0.47**	-0.65**	1.00					
NN3	-0.67**	-0.83**	0.82^{**}	1.00				
NN4	-0.69**	-0.75**	0.41^{**}	0.73^{**}	1.00			
NN34	-0.75**	-0.85**	0.64^{**}	0.94^{**}	0.88^{**}	1.00		
ONI	-0.74**	-0.85**	0.63**	0.92^{**}	0.88**	0.99^{**}	1.00	
SS	-0.02	-0.04	-0.02	-0.02	-0.03	-0.03	-0.03	1.00

³¹⁹ 5 Causality Analyses Comparison

The causality detections between natural inflow and climate variables in Brazil are here evaluated and compared by implying different causality detection methods summarized in section 2. It is of note that all the results were obtained using R.

323 5.1 Time Domain Granger Causality Test

Given the significant and empirical role of GC causality test, the GC test results are firstly con-324 ducted and summarized as follows in Table 5. It is of note that the preconditions of time domain 325 GC test are satisfied for all tests across various combinations of variables and the corresponding 326 optimal lag is determined respectively by a group of information criteria. Specifically, the re-327 sults that are highlighted in red represent that the valid evidence is obtained for unidirectional 328 causality from corresponding climate variable to the natural inflow. Note that these valid cases 329 have no conflicts of causality for the reverse direction and all shows significance level less than 330 10%. 331

It is observed that the GC test shows relatively promising performance for NN12 and ONI across climate variables, for AMZ, URU and SAT among all natural inflow series. However, there are general misleading results of the reverse direction and many cases of mutual directionalcausality with high significant levels.

336 5.2 Frequency Domain Causality Test

The frequency domain test extends the GC test and further investigates into the causal links by each particular frequency. Note that the preconditions are stratified and the optimal lagstructures are maintained for all tests. As can be seen in Table 6, the valid cases are again highlighted in red, which indicates unidirectional causality from climate variable to natural inflow without the evidence of causality for the other direction.

In general, NN34 and ONI obtain overall valid evidences of unidirectional causality without misleading results, whilst only AMZ out of all the natural inflow series shows identical valid results with NN12, NN4 and NN34. Even P2, P3, IGU, URU and SAT indicate a few valid unidirectional causality cases, however, it is not consistent and solid enough considering the amount of misleading results showing causality in reverse direction or mutual direction.

		SOL_St		SO	SOI_Eq		Ξ.	5		NN3			NN4			NN34			INO			\mathbf{ss}	
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Table 5: Time domain GC test results.

→ causality from climate variable to natural influe.
 → : causality from natural inflow to climate variable.
 → : causality from natural inflow to climate variable.
 empty cell: no causality detected.
 ,,***: indicate 10%, 5% and 1% significant level respectively.

		SOLSt		SOI_Eq		NN12			NN3			NN4		Z	NN34		0	INO		\mathbf{SS}	
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Table 6: Frequency domain causality test results.

←: causality from climate variable to natural inflow.
 →: causality from natural inflow to climate variable.
 empty cell: no causality detected.
 ,,***: indicate 10%, 5% and 1% significant level respectively.

347 5.3 CCM

The CCM causality test has the significant advantage of no prior linear model assumptions are 348 made and this technique is designed for better understanding of causal relationships in complex 349 dynamical system. The results of CCM tests are briefly summarized in Table 7, Table 8 and 350 Table 9 and organized by each pair of tested variables. Moreover, the time lag has been involved 351 to the evaluation, where lag 1 to 6 are considered coving 6 months of lag effect. It is of note 352 that all test results are obtained by the optimal embedding dimension respectively, which is 353 determined by the nearest neighbor forecasting performance using simplex projection and leave-354 one-out cross validation is applied for the best choice on library size with optimal performance. 355 The results overwhelmingly indicate causality from natural inflow to climate variable², whilst 356 only 18% of the pairs get positive evidence on unidirectional causality from climate variable to 357 natural inflow. However, even among those 18% pairs, there are misleading results of no clear 358 causality detected for specific time lag options. In general, SAT and SF along with NN3 and 359 NN4 obtain relatively more positive results. 360

 $^{^{2}}$ This is possibly due to the oversensitivity of CCM on noise, however, it is of note that the cross mapping skills of both directions are significantly high, indicating the strong link between natural inflow and climate variables.

		01	SOL_St		Š	SOL-Eq	F	Z	NN12		Z	N3		2	JN4		Z	NN34			INC			$\mathbf{s}^{\mathbf{s}}_{\mathbf{s}}$	
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Table 7: CCM causality test results (1).

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_	ю	4	¢	ю	9	¢	5	7	¢	ъ	7	¢	ю	2	Ŷ	ъ	7	¢	ъ	ŝ	Ŷ	ю	
_	9	x	Ļ	9	5	Ŷ	9	x	¢	9	7	¢	9	2	¢	9	7	Ŷ	9	2	¢	9	6 2
	-	ъ	¢	1	9	¢	1	7	¢	1	7	¢	-	2	Ļ	1	7	¢	1	7	¢		
_	2	4	1	7	9	1	2	7	¢	2	2	1	7	7	¢	2	ŝ	¢	2	0	¢	2	
COL	e	ŋ	1	0	ŋ	¢	с	с	¢	с	5	1	e	0	¢	З	7	¢	e	с	¢	e	
	4	ŋ	1	4	7	¢	4	7	¢	4	7	1	4	7	¢	4	7	¢	4	0	¢	4	
_	ъ	4	1	ъ	9	¢	ŋ	7	¢	ŋ	2	1	S	7	¢	S	7	1	ŋ	ŝ	¢	n	5 2
	9	8	¢	9	5	1	9	ø	¢	9	2	1	9	2	¢	9	2	¢	9	2	¢	9	
	1	IJ	¢	1	9	↑	1	7	NO	1	2	¢	1	2	↑	1	7	¢	1	2	Ŷ	1	
_	2	7	¢	7	9	¢	2	7	¢	2	7	¢	2	2	¢	2	ŝ	Ŷ	7	7	Ŷ	2	
ълт	e	IJ	Ŷ	e	5	¢	e	e	¢	ŝ	7	1	n	7	¢	ŝ	7	Ŷ	e	ę	Ŷ	n	3 2
	4	ŋ	ON	4	7	¢	4	7	ON	4	7	1	4	7	¢	4	7	¢	4	0	¢	4	
_	ъ	4	¢	ъ	9	¢	ŋ	7	ON	ŋ	7	1	ъ	7	¢	S	7	¢	5 2	ŝ	¢	n	
	Ś	x	1	ď	ц	1	u u	0	,	a a	c	CIV.	¢	c		ç	c	,	u	c			

Table 8: CCM causality test results (2).

		s_0	حد	10	SOL_Eq			NN12			SZ3 Σ		1	ZZ4			NN34			INO				\mathbf{SS}
	Lag	ы	dir.	Lag	ы	dir.	Lag	ы	dir.	Lag	ы	dir.	Lag	ы	dir.	Lag	ы	dir.	Lag	ы	dir.	Г	ag	
	1	ъ	1	1	9	\uparrow	1	0	1	-	0	1	1	5	↑	1	7	¢	1	0	↑		-	
	2	7	1	2	9	1	2	2	1	2	2	¢	2	2	¢	2	3	¢	2	0	¢		\sim	
51.5	ñ	ŋ	Ŷ	n	ŋ	Ŷ	ŝ	с	Ŷ	ŝ	2	Ŷ	ŝ	2	¢	ę	2	Ŷ	c,	ę	Ŷ		3	
111	4	Ŋ	1	4	7	¢	4	7	Ļ	4	7	¢	4	7	¢	4	7	¢	4	7	Ŷ		4	
	ю	4	1	ъ	9	1	ъ	2	1	ю	2	¢	ŋ	2	¢	ŋ	2	¢	ю	с	¢		ъ	
	9	x	Ŷ	9	ŋ	¢	9	x	Ŷ	9	2	Ŷ	9	2	¢	9	2	Ŷ	9	2	Ŷ	_	.0	
	1	ŋ	¢	1	9	1	1	2	¢	1	2	¢	1	2	¢	1	2	¢	1	7	¢			
	2	7	1	2	9	¢	2	7	ON	2	2	Ļ	2	2	Ļ	2	ŝ	¢	2	7	Ļ		~	
	ŝ	ŋ	1	e	ŋ	¢	ŝ	e	1	ŝ	2	¢	e	7	Ļ	ĉ	2	¢	ŝ	ę	¢		e	
זינ	4	ŋ	1	4	7	1	4	0	ON	4	7	ON	4	2	1	4	2	¢	4	0	¢		4	
	ъ	4	1	ъ	9	¢	ъ	7	1	ъ	2	¢	ю	2	¢	ŋ	2	¢	5 C	ę	Ŷ		ы	
	9	x	Ŷ	9	5 C	Ŷ	9	80	Ŷ	9	2	Ŷ	9	2	¢	9	2	¢	9	7	Ŷ		9	
	1	ъ	¢	-1	9	¢	1	7	1	1	2	NO	1	2	Ļ	1	2	¢	1	7	↓			
	7	7	1	2	9	Ŷ	2	7	1	2	2	¢	2	2	Ļ	2	ŝ	¢	2	7	Ŷ		0	
TIDII	e	ŋ	¢	n	5 C	Ŷ	ŝ	ი	Ŷ	n	2	Ŷ	ŝ	2	¢	ŝ	2	¢	ĉ	ŝ	Ŷ		~	
	4	ŋ	1	4	7	Ŷ	4	0	1	4	7	ON	4	7	ON	4	7	1	4	7	ON	4.		
	n	4	1	ъ	9	Ŷ	ъ	7	1	ъ	2	¢	5	2	¢	ŋ	2	¢	ŋ	ŝ	Ŷ	ŝ		
	9	8	1	9	S	1	9	×	1	9	2	¢	9	2	¢	9	2	1	9	2	¢	_		
	1	ъ	1	1	9	1	1	2	1	1	7	Ļ	1	5	Ļ	1	5	Ļ	1	2	Ļ	-		
	2	7	1	7	9	1	2	2	1	2	7	Ļ	7	7	t	7	ĉ	¢	2	7	Ļ	7		
сAТ	с С	Ŋ	1	e	ŋ	¢	e	с	1	e	7	Ļ	С	7	¢	3	7	¢	с	с	¢	3		
	4	ŋ	1	4	7	1	4	7	1	4	7	Ļ	4	7	Ļ	4	7	¢	4	0	Ļ	4		
	ŋ	4	1	ъ	9	1	ŋ	2	1	ŋ	7	ON	Ŋ	7	Ļ	ю	7	ON	ŋ	ĉ	¢	ŋ		
	9	8	1	9	S	1	9	×	1	9	2	¢	9	2	¢	9	2	1	9	2	¢	9		
	-	ъ	¢	-1	9	1	1	7	1	1	7	¢	-	2	¢	1	2	¢	1	7	¢			
	2	7	1	2	9	Ŷ	2	7	1	2	2	¢	2	2	¢	2	ŝ	¢	2	7	Ŷ	2		
Д	e	ŋ	Ŷ	n	5 C	Ŷ	ŝ	ი	Ŷ	n	2	Ŷ	ŝ	2	¢	ŝ	2	¢	ĉ	ŝ	Ŷ	ŝ		
2	4	Ŋ	1	4	7	¢	4	7	1	4	7	¢	4	7	¢	4	7	¢	4	7	Ŷ	~		
	n	4	1	ъ	9	1	S	0	1	ъ	2	¢	S	0	¢	ŋ	0	¢	ų	ŝ	¢	_	ы	
	g	x	1	ç	Ľ	1	u u	x	,	a	c		¢	c		a	c	,	U	c		0		

Table 9: CCM causality test results (3).

³⁶¹ 5.4 SSA based Causality Test

Follow the brief introduction of SSA based causality test in section 3, the test results of natural 362 inflow and climate variable in Brazil are summarized in Table 10^3 . It is of note that both recur-363 rent and vector forecasting algorithms are evaluated respectively; the out-of-sample is defined as 364 the last 1/3 of the total observation for both SSA and MSSA forecastings; the root mean square 365 error (RMSE) of forecasting for SSA and MSSA are the optimal outcome obtained respectively 366 with the optimal window length (L) and numbers of eigenvalues (r) that are also listed in the 367 table; causality is detected if the corresponding F statistics is smaller than 1 and the significant 368 level of causality increases while the value of F statistics decreases. 369 In general, the results are again a mixture of different unidirectional causality, mutual direc-370

tional causality and no causality, and no significant pattern can be identified, except that NN34 and NN4 work slightly better among all series. Moreover, the F statistics are very close to 1, which indicate that the forecasting of MSSA by involving the other variable is improved by a very limited amount comparing to the performance of univariate SSA.

³It is of note that the listed pairs are part of all combinations that cover almost all tested series and types of results. The complete details of these results are available upon request from the authors.

		SS	SA			MS	\mathbf{SA}			S	SA Causality	
	I	Rec	7	Vec		Rec		Vec	Rec	Vec		
	L,r	RMSE	L,r	RMSE	L,r	RMSE	L,r	RMSE	F stat	F stat	Direction	Decisior
ONI	3,2	0.21	$16,\!14$	0.20	9,6	0.20	3,2	0.21	0.92	1.09	URU caus ONI	XX /
URU	12,3	1349.63	3,1	1357.71	2,1	1426.68	2,1	1426.68	1.06	1.05	ONI caus URU	Wrong
NN34	3,2	0.28	3,2	0.28	3,2	0.28	3,2	0.29	1.02	1.02	SAT caus NN34	No
SAT	20,5	541.68	20,10	544.38	11,3	556.97	12,3	559.69	1.03	1.03	NN34 caus SAT	INO
SS	16,14	1.00	16,11	1.22	8,6	1.19	8,6	1.47	1.20	1.21	P4 caus SS	37
$\mathbf{P4}$	19,4	100.14	17,5	106.68	19,4	85.11	15,4	91.61	0.85	0.86	SS caus P4	Yes
SOI_St	2,1	0.74	3,1	0.73	2,1	0.69	2,1	0.69	0.93	0.94	SAT caus SOLSt	117
SAT	20,5	541.68	20,10	544.38	11,3	556.97	12,3	559.69	1.03	1.03	SOLSt caus SAT	Wrong
NN3	3,2	0.34	3,2	0.35	3,2	0.33	3,2	0.34	0.96	0.96	IGU caus NN3	
IGU	18,1	1292.42	20,1	1282.83	17,1	1338.35	19,1	1315.74	1.04	1.03	NN3 caus IGU	Wrong
NN4	6,3	0.20	17,8	0.21	3,2	0.22	3,2	0.22	1.06	1.07	GRA caus NN4	
GRA	19,5	829.02	13,3	821.12	19,5	844.93	17,5	854.13	1.02	1.04	NN4 caus GRA	No
NN12	3,2	0.57	3,2	0.57	8,5	0.49	9,6	0.48	0.86	0.83	SF caus NN12	
SF	12,5	1529.22	13,3	1509.66	20,3	1395.22	16,6	1425.42	0.91	0.94	NN12 caus SF	Mutual
NN34	3,2	0.28	3,2	0.28	3,2	0.28	3,2	0.29	1.01	1.02	P2 caus NN34	
P2	20,7	1223.53	20,7	1224.24	16,7	1111.85	12,4	1088.23	0.91	0.89	NN34 caus P2	Yes
ONI	3,2	0.21	16,14	0.20	5,4	0.19	5,4	0.20	0.88	1.03	P3 caus ONI	
P3	20,5	1330.56	5,1	1320.23	20,6	1307.70	15,4	1306.01	0.98	0.99	ONI caus P3	Mutua
NN34	3,2	0.28	3,2	0.28	3,2	0.28	3,2	0.29	1.01	1.02	AMZ caus NN34	
AMZ	13,6	5052.20	20,12	4762.11	10,7	4904.91	13,9	4678.68	0.97	0.98	NN34 caus AMZ	Yes
NN4	6,3	0.20	17,8	0.21	3,2	0.22	3,2	0.22	1.06	1.06	SAT caus NN4	
SAT	20,5	541.68	20,10	544.38	11.3	556.97	12,3	559.69	1.03	1.03	NN4 caus SAT	No
NN3	3,2	0.34	3,2	0.35	3,2	0.33	3,2	0.34	0.96	0.97	P1 caus NN3	
P1	20,5	1049.30	13,3	1029.57	20,5	1047.73	20,5	1053.18	1.00	1.02	NN3 caus P1	Wrong
SOI_Eq	20,0	0.52	2,1	0.52	7,3	0.54	7,3	0.55	1.05	1.02	URU caus SOLEq	
URU	12,1	1349.63	3,1	1357.71	$^{1,0}_{2,1}$	1426.68	17,10	1385.08	1.06	1.02	SOLEq caus URU	No
NN4	6,3	0.20	17,8	0.21	3,2	0.22	8,4	0.23	1.10	1.12	TOC caus NN4	
TOC	11,5	5595.27	20,12	5229.04	15,10	4538.45	15,10	4358.17	0.81	0.83	NN4 caus TOC	Yes
NN12	3,2	0.57	3,2	0.57	10,10	0.49	10,10	0.50	0.86	0.87	TIE caus NN12	
TIE	16,4	642.53	16,5	633.07	12,6	729.93	10,0 19,14	702.37	1.14	1.11	NN12 caus TIE	Wrong
SOLSt	2,1	0.74	3,1	0.73	8,3	0.69	8,3	0.68	0.93	0.93	IGU caus SOLSt	
IGU	18,1	1292.42	20,1	1282.83	17,1	1338.35	18,1	1315.74	1.04	1.03	SOLSt caus IGU	Wrong
ONI	3,2	0.21	16,14	0.20	8,6	0.17	8,6	0.18	0.82	0.92	SF caus ONI	
SF	$^{3,2}_{12,5}$	1529.22	13,3	1509.66	20,3	1395.22	16,6	1425.42	0.82	0.92	ONI caus SF	Mutua
NN12	3,2	0.57	3,2	0.57	14,9	0.48	10,0	0.48	0.91	0.94	P3 caus NN12	
P3	$^{3,2}_{20,5}$	1330.56	$^{3,2}_{5,1}$	1320.23	$^{14,9}_{20,6}$	0.48 1307.70		0.48 1306.01	0.83	0.84	NN12 caus P3	Mutua
SOI_Eq	20,5	0.52	2,1	0.52	5,2	0.56	15,4 5,2	0.55	1.07	1.06	P2 caus SOLEq	
P2	$^{2,1}_{20,7}$	0.52 1223.53	$^{2,1}_{20,7}$	0.52 1224.24	$^{5,2}_{16,7}$	0.56	$^{5,2}_{12,4}$	0.55 1088.23	0.91	0.89	SOI_Eq caus P2	Yes
	,		,	1224.24 le to natura	/		12,4	1000.23	0.91	0.69	SOLEq caus P2	

Table 10: SSA based causality test results.

Yes: only causality from climate variable to natural inflow is detected. Wrong: only causality from natural inflow to climate variable is detected.

No: no causality detected.

Mutual: mutual directional causality between climate variable and natural inflow.

375 5.5 Periodic Autoregressive Model based Causality Test

The PAR causality test results are summarized in Table 11⁴ where both PAR and PARX RMSE 376 are present when forecasting the last 1/3 of the total observation. Following the procedure 377 describe in SSA causality test, if the corresponding F statistics is smaller than 1 then there is 378 causality. When causality is detected in both directions the causality is not computed, and when 379 the direction of causality is from the natural inflow in the climate variable then is computed as 380 wrong. The causality is computed in the right decision only in six cases, but the F statistics 381 is very close to one in most cases, showing that even when causality can be considered, the 382 improvements of considering a climate variable it's on the edge. The only case where can be 383 clearly found a causality is between SOI Equatorial and Paranapanema basin. 384

⁴It is of note that the listed pairs are part of all combinations that cover almost all tested series and types of results. The complete details of these results are available upon request from the authors.

	RM	ISE		PAR Causality	
	PAR	PARX	F stat	Direction	Decision
GRA	1129.17	1076.456	0.953	NN4T caus GRA	Mutual
NN4	1.128	0.789	0.7	GRA caus NN4	Mutual
P2	2056.97	1235.528	0.601	SOLEq caus P2	Yes
SOI_Eq	1.69	1.754	1.038	No	res
P3	1618.171	1369.67	0.846	ONI caus P3	Yes
ONI	1.306	1.955	1.497	No	res
P3	1618.171	1587.472	0.981	NN3 caus P3	37
NN3	1.56	2.199	1.409	No	Yes
P3	1618.171	1434.836	0.887	NN4 caus P3	Yes
NN4	1.128	1.34	1.188	No	res
P3	1618.171	1583.034	0.978	NN34 caus P3	Yes
NN34	1.574	2.143	1.362	No	res
AMZ	8519.513	8581.495	1.007	No	337
NN4	1.128	0.91	0.807	AMZ caus NN4	Wrong
TOC	6359.823	5489.439	0.863	ONI caus TOC	Mutual
ONI	1.306	1.006	0.77	TOC caus ONI	Mutual
EAT	806.574	472.349	0.586	SS caus EAT	Mutual
ss	64.101	24.251	0.378	EAT caus SS	Mutual
TIE	827.817	747.594	0.903	NN3 caus TIE	Yes
NN3	1.56	1.624	1.041	No	res
IGU	1629.477	1348.128	0.827	ONI caus IGU	Mutual
ONI	1.306	0.847	0.649	IGU caus ONI	Mutual
URU	1827.213	1497.444	0.82	SOLSt caus URU	Mutual
SOI_St	1.537	1.152	0.75	URU caus SOLSt	mutual
SAT	712.875	550.122	0.772	NN12 caus SAT	Maturi
NN12	1.606	1.406	0.876	SAT caus NN12T	Mutual
P5	193.369	196.162	1.014	No	XX 7
SOI_Eq	1.69	1.544	0.914	P5 caus SOLEq	Wrong

Table 11: PAR based causality test results.

Yes: only causality from climate variable to natural inflow is detected.

Wrong: only causality from natural inflow to climate variable is detected.

No: no causality detected.

Mutual: mutual directional causality between climate variable and natural inflow.

³⁸⁵ 6 Final Discussion and Conclusion

In general, this paper successfully obtains comprehensive investigation of the causality relation-386 ship between natural inflow and climate variables in Brazil by analyzing the data of 15 major 387 basins and 8 different climate series. For the first time to the best of our knowledge, it in-388 corporates and compares five different causality detection methods for the causality study on 389 hydrological series. In specific, GC test shows relatively promising performance for AMZ, URU 390 and SAT among all natural inflow series, NN12 and ONI across climate variables; frequency 391 domain causality test indicates generally valid evidences of unidirectional causality for AMZ, 392 NN34 and ONI; CCM overwhelmingly obtains significant unidirectional causality from the op-393 posite direction (natural inflow to climate variables), whilst SAT, SF, NN3 and NN4 relatively 394 give more positive results of the valid direction; SSA based causality test shows that NN34 and 395 NN4 work slightly better, and the forecasting improvements by involving the other variable are 396 generally very limited; PAR based causality test computed six unidirectional causality, but only 397 one is really significant (P2 and SOLEq). 398

The overall results indicate that there is no single method which stands out and outperforms the others. The conclusions are a mixture of different unidirectional, mutual directional, and no causality. There is no obvious pattern that can be clearly identified across 15 natural inflow series and 8 climate variables. However, it is noticed that the overwhelming evidences of opposite direction of causality are obtained by CCM, which is the most concurrent outcome of all five different tests. It is frankly interesting discovery that is possibly caused by significant noises that generally exist in those series, which will be one of the main focuses for future research.

The works presented in the Background section showed improvements when using information from the climate variables in the inflow prediction procedure, so even if the tests applied here did not present favourable results, a natural continuation of this study will be the application of different models that incorporate exogenous variables to verify the significance of the climate variables in the prediction of each of the inflow series studied.

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