

## Validation of a Spanish version of the Brief Experiential Avoidance Questionnaire (BEAQ) in clinical population

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### Abstract

**Background:** The Brief Experiential Avoidance Questionnaire (BEAQ) has been suggested as the most appropriate instrument for measuring experiential avoidance. However, no Spanish validation has been published. The aim of this study was to validate a Spanish version of the BEAQ in a clinical sample treated at a community mental health unit. **Methods:** Participants ( $N = 332$ ) completed the BEAQ as well as other self-report measures of experiential avoidance and psychopathology. **Results:** Internal consistency was satisfactory ( $\alpha = .82$ ). No statistically significant gender differences were found in the BEAQ scores. The data also showed high test-retest reliability after four to six weeks, acceptable concurrent validity with another experiential avoidance measure and acceptable convergent validity with the psychopathology measure. The principal component analysis, forcing the one factor solution proposed in the original scale, produced indicators similar to the English version of the BEAQ. **Conclusions:** These results firmly support the reliability and validity of this Spanish validation, stressing its usefulness as a measure of experiential avoidance in clinical populations.

**Keywords:** Experiential avoidance, Brief Experiential Avoidance Questionnaire, psychometric characteristics, assessment, clinical sample.

### Resumen

**Validación de la versión española del Cuestionario Breve de Evitación Experiential (BEAQ) en población clínica. Antecedentes:** el Cuestionario Breve de Evitación Experiential (BEAQ) ha sido propuesto como el instrumento más adecuado para medir la evitación experiential. Sin embargo, todavía no ha sido publicada ninguna validación en español. Por lo tanto, el objetivo del presente estudio ha sido la validación de una versión española del BEAQ, en una muestra clínica atendida en un Centro de Salud Mental Comunitario. **Método:** los participantes ( $N = 332$ ) completaron el BEAQ, así como otras medidas de autoinforme de evitación experiential y psicopatología. **Resultados:** la consistencia interna fue satisfactoria ( $\alpha = .82$ ). No se encontraron diferencias de género estadísticamente significativas en las puntuaciones del BEAQ. Los datos también mostraron una alta fiabilidad test-retest en un intervalo de cuatro a seis semanas, validez concurrente aceptable con otra medida de evitación experiential y validez convergente aceptable con la medida de psicopatología. El análisis de componentes principales forzando a la solución de un factor, como se propuso en la escala original, obtuvo unos indicadores similares a los obtenidos en dicha versión inglesa. **Conclusiones:** los resultados confirman la adecuada fiabilidad y validez de la presente versión española del BEAQ, destacando su utilidad como medida de la evitación experiential en población clínica.

**Palabras clave:** evitación experiential, Cuestionario Breve de Evitación Experiential, características psicométricas, evaluación, muestra clínica.

Experiential avoidance has been defined as “the phenomenon that occurs when a person does not want to remain in contact with particular private experiences (e.g. bodily sensations, emotions, thoughts, memories, images, behavioral predispositions) and takes steps to alter the form or frequency of those experiences or the contexts that produce them” (Hayes, Wilson, Gifford, Follette, & Strosahl, 1996, p. 1154).

The importance of this construct is supported by research finding experiential avoidance to be a major factor in the development and maintenance of psychopathology. Several studies have pointed to

the association between more experiential avoidance and greater interpersonal problems, as well as more severe psychopathology (García Montes, Luciano Soriano, Hernández López, & Zaldívar Basurto, 2004; Gerhart, Baker, Hoerger, & Ronan, 2014; Hershenberg, Mavandadi, Wright, & Thase, 2017; Kroska, Miller, Roche, Kroska, & O’Hara, 2018; Mellick, Vanwoerden, & Sharp, 2017; Skinner, Rojas, & Veilleux, 2017; Spinhoven, Drost, de Rooij, van Hemert, & Penninx, 2015; Wagener, Baeyens, & Blairy, 2016). In addition, less experiential avoidance is associated with better well-being and quality of life (Álvarez Díaz & Esteve Zarazaga, 2009; Hayes, Luoma, Bond, Masuda, & Lillis, 2006; Jacob, Ower, & Buchholz, 2013; Kashdan, Barrios, Forsyth, & Steger, 2006).

The first instrument developed to measure experiential avoidance was the Acceptance and Action Questionnaire (AAQ; Hayes et al., 2004). A second version of this questionnaire, the Acceptance and Action Questionnaire II (AAQII; Bond et al., 2011), was developed after the AAQ received criticism, especially

concerning its low internal consistency (Cronbach's alpha <.70), inadequate distinction between process and result (Chawla & Ostafin, 2007), and poor discriminant validity with measures of negative affect and neuroticism.

However, this second version has also been criticized for insufficient discriminant validity with measures of negative affect and neuroticism, especially in clinical samples (Gámez, Chmielewski, Kotov, Ruggero, Suzuki, & Watson, 2014). Thus, the AAQII is more correlated with measures of neuroticism and negative affect than with the AAQ or other avoidance measures (Gámez, Chmielewski, Kotov, Ruggero, & Watson, 2011). According to these authors, this is observed in the item content, which focuses, for example, on failed attempts at controlling discomfort and its consequences (e.g., "emotions cause problems in my life"; "worries get in the way of my success"). This can lead to criterion contamination (Gámez et al., 2011), since it confuses process with result.

Another criticism of both scales (AAQ and AAQII) is that the authors initially conceptualized them as measures of experiential avoidance, but later redefined their scope as psychological inflexibility. This construct includes some aspects of experiential avoidance, but also other concepts, such as cognitive fusion, lack of contact with values, lack of committed action and attention to the present (Bond et al., 2011), thereby exceeding the content of the experiential avoidance variable (Gámez et al., 2011).

The Multidimensional Experiential Avoidance Questionnaire (MEAQ; Gámez et al., 2011) was designed to overcome the limitations of both versions of the AAQ. It was intended as a comprehensible and reliable measure of experiential avoidance, differentiated from neuroticism. However, its length (62 items) made its applicability difficult in contexts where time is limited. The Brief Experiential Avoidance Questionnaire (BEAQ; Gámez et al., 2014) emerged as a brief, but still effective questionnaire. Its authors pointed out that this instrument manages to reduce the application time from 12 minutes on the full scale (MEAQ) to 3 minutes. In addition, it covers content similar to the MEAQ and both scales show practically the same convergent and discriminant association. Because of the limitations of those already mentioned above (AAQ, AAQII and MEAQ), it has been suggested that this brief version, the BEAQ, is the most appropriate instrument for measuring experiential avoidance (Wolgast, 2014). However, no Spanish validation has yet been published.

Due to the relevance of the experiential avoidance construct and its negative association with wellbeing and health measures, we consider the validation of this instrument in Spanish population to be of great importance. The purpose of the present research was to analyze the psychometric properties of a Spanish version of the BEAQ in a clinical sample.

#### Method

##### Participants

The sample consisted of 332 patients treated in a community mental health unit. It included 228 women (69%) and 104 men (31%). All the participants were Caucasian and were aged between 18 and 78 years ( $M = 45.97$ ,  $SD = 12.44$ ). The inclusion criteria were: 1) have been diagnosed with a mental disorder according to ICD-10 (World Health Organization, 1992), specifically those corresponding to Depressive Disorder (F32.0, F32.1, F33.0 and F33.1), Dysthymia (F34.1), Anxiety Disorder (F40 and F41),

Obsessive-Compulsive Disorder (F42), Adjustment disorders (F43), Eating Disorder (F50) and Personality Disorder (F60); 2) be of legal age; 3) have no cognitive impairment; 4) give their voluntary consent to participation in the study. Other sociodemographic and clinical data appear in Table 1.

##### Instruments

*Brief Experiential Avoidance Questionnaire* (Gámez et al., 2014). This is a self-report questionnaire with 15 items that measure experiential avoidance. The items are rated on a 6-point Likert scale (1="strongly disagree" to 6="strongly agree"). The scores range from 15 to 90 points (item 6 must be inverted), with higher scores showing greater experiential avoidance. In its original version, the BEAQ shows association with measures of avoidance, psychopathology and quality of life, as well as it shows strong convergence with each of the six dimensions of the MEAQ (Gámez et al., 2011), the original scale which the items for this short version were extracted from. The English version has good reliability, with a Cronbach's alpha of .80 to .89 for the various samples.

*Acceptance and Action Questionnaire II* (Bond et al., 2011). This is a self-report scale with seven items measuring experiential avoidance. The items evaluate unwillingness to experience

Table 1  
Sociodemographic information (N = 332)

Variables	N	%
<b>Marital Status</b>		
Single	109	32.8
Married/domestic partner	147	44.3
Separated/Divorced	45	13.6
Widowed	6	1.8
Unknown	25	7.5
<b>Total</b>	<b>332</b>	<b>100</b>
<b>Educative level</b>		
Primary school	37	11.1
Secondary School	44	13.3
Secondary school certificate/Professional training	92	27.7
University	126	38
Unknown	33	9.9
<b>Total</b>	<b>332</b>	<b>100</b>
<b>Employment</b>		
Working	169	50.9
Unemployed	67	20.3
Pensioned	20	6
Retired	6	1.8
Student	27	8.1
Homemaker	20	6
Other	23	6.9
<b>Total</b>	<b>332</b>	<b>100</b>
<b>Primary Diagnostics</b>		
Depression	41	12.3
Anxiety	117	35.3
Adjustment Disorder	125	37.7
Eating Disorder	30	9
Personality Disorder	19	5.7
<b>Total</b>	<b>332</b>	<b>100</b>

unwanted emotions, thoughts and feelings and the inability to act according to values on a 7-point Likert scale (1="never true" to 7="always true"). Scores range from 7 to 49, with higher scores showing higher experiential avoidance. The Spanish version used (Ruiz, Langer Herrera, Luciano, Cangas, & Beltrán, 2013) has acceptable internal consistency, with a Cronbach's alpha of .75 for a sample of clinical patients.

*Symptoms Checklist-90-Revised* (SCL-90-R; Derogatis, 1977). This self-report scale is composed of 90 items covering a wide variety of psychopathological manifestations. The nine dimensions evaluated are somatization, obsessive-compulsive, interpersonal sensitivity, depression, anxiety, hostility, phobic anxiety, paranoid ideation and psychoticism. The results offer a symptomatic profile as well as three global indexes: global severity index, total positive symptoms and positive symptom distress index. The items are scored on a scale of 5-point Likert scale (0 = "not at all" to 4 = "very or extremely"), based on the intensity with which a series of symptoms were experienced during the previous week. The Spanish adaptation used (González de Rivera, de las Cuevas, Rodríguez-Abuin, & Rodríguez-Pulido, 2002) has a Cronbach's alpha coefficient of .94 for the 90 elements included in the test.

#### Procedure

First, following guidelines for test validation (Muñiz, Elosua, & Hambleton, 2013), authorization was requested from the main author of the BEAQ scale for its translation and validation in a Spanish sample. With his permission, it was translated into Spanish by the reverse translation method. This scale was applied to a pilot sample of six patients to make sure they understood it, and it was unnecessary to make any changes in the items. Due to the length of the questionnaire, the SCL-90-R was applied only to the first 112 participants. This research was approved by the Andalusian Biomedical Research Ethics Committee.

#### Data analysis

First, the Kaiser-Meyer-Olkin and the Barlett's sphericity test analysis were performed to find out whether the items were correlated and adequate for factor analysis. To determine the factor loadings, an analysis of principal component analysis was carried out, specifying a one-factor solution as proposed by the authors of the original scale.

Furthermore, a descriptive analysis of the items on this Spanish version of the BEAQ was carried out by calculating measures of central tendency and dispersion. Skewness and kurtosis were calculated as well to determine the behavior of the data found from the questionnaire. Next, the normal distribution of data was checked with the Kolmogorov-Smirnov test of normality. The Student's *t* test was also calculated for any possible gender differences in the scores.

Internal consistency of the BEAQ was then calculated using the Cronbach's alpha coefficient, and Omega coefficient was used to examine its factorial reliability (Dunn, Baguley, & Brunson, 2014; Elosua & Zumbo, 2008; Ventura-León & Caycho-Rodríguez, 2017). In addition, temporal reliability was analyzed with the Pearson's *r* correlation of the measure at two different times four to six weeks apart. Finally, the instrument's sources of validity were analyzed (Muñiz, 2018). First, the concurrent validity of the BEAQ with the AAQII was calculated, using the

Pearson's *r* correlation coefficient. An analysis of convergent validity of experiential avoidance with the SCL-90-R, a general psychopathology measure, was performed using the Pearson's correlation. Specific correlation with the global severity index, as well as with five of its subscales (anxiety, depression, phobic anxiety, somatization and obsessive-compulsive), were also explored. Statistical analyses were performed using SPSS v.21.

## Results

### Sources of validity evidence of internal structure

The Kaiser-Meyer-Olkin test for sample adequacy ( $KMO = .83$ ) and the Barlett's sphericity test ( $X^2 = 1251.87, p < .01$ ) both confirmed adequacy of the sample data for factor analysis. A principal component analysis was done, forcing a one-factor solution, as proposed in the original scale. The first five eigenvalues were 4.52, 1.57, 1.11, 1.08 and .94. The component extracted explained 30.12% of the variance. Table 2 shows the saturation of each item on this single component.

### Descriptive statistics

The descriptive characteristics of the BEAQ in the sample are presented in Table 3. The Kolmogorov-Smirnov normality test confirmed that the distribution of the BEAQ scores followed a normal distribution, at a significance level over .05 ( $Z = .864, p = .445$ ).

The Student *t* test for independent samples showed absence of significant differences in the scores on experiential avoidance ( $t = .103, p = .92$ ) between women ( $M = 55.54, SD = 13.40$ ) and men ( $M = 55.72, SD = 15.04$ ). The effect size, calculated by Cohen's *d*, was insignificant ( $d = 0.013$ ).

### Reliability of the scores

*Internal consistency.* The Cronbach's alpha was .82, which shows high internal consistency of the instrument. Table 4 shows the means and variance in the scale if the item was eliminated, as

Table 2  
Saturation for the items on the BEAQ after principal component analysis for a one-factor solution (N = 332)

Item N°	Factorial loadings
1	.55
2	.56
3	.33
4	.43
5	.55
6	.26
7	.57
8	.61
9	.44
10	.68
11	.61
12	.71
13	.57
14	.63
15	.56

well as the corrected item-total correlation and Cronbach's alpha when each item of the scale is eliminated. The corrected item-total correlation was over .35 for 13 of the 15 items, which is adequate and significant according to Cohen and Manion (1989). The two exceptions were Items 3 and 6, which had a correlation with the total scale of  $r = .24$  and  $r = .21$ , respectively. Because of these results, the Cronbach's alpha was calculated for the scale after eliminating these items. By eliminating Item 3, Cronbach's alpha remained at .82 and only rose slightly when Item 6 was eliminated, from .82 for the total items to .83. In no case would eliminating any of the remaining items improve the Cronbach's alpha for the total scale. Omega index was .86, also supporting the adequate factorial reliability of the scale.

**Temporal reliability.** The results showed a high positive correlation ( $r = .70, p < .01$ ) between the BEAQ before and after

an interval of 4-6 weeks ( $N = 48$ ). This shows high temporal reliability of the Spanish version of the instrument with a clinical population.

*Sources of validity evidence in relation with other variables*

The concurrent validity between the BEAQ and the AAQII, analyzed using the Pearson's correlation, was significant with  $r = .67 (p < .01)$ . To analyze the convergent validity of the BEAQ experiential avoidance scale with other psychopathology measures, both the symptom inventory SCL-90-R (Derogatis, 1977) global severity index, and five of that instrument's subscales (anxiety, depression, phobic anxiety, somatization and obsession-compulsion) were used. The Pearson correlation analyses for the indicators above showed significant correlations for the global BEAQ score in all cases. The results are shown in Table 5.

*Table 3*  
Descriptive statistics ( $N = 332$ )

Item Nº BEAQ	Mean	SD	Skewness	Kurtosis
1	3.24	1.84	0.10	-1.48
2	3.55	1.70	-0.08	-1.30
3	4.00	1.64	-0.36	-1.09
4	2.78	1.75	0.47	-1.24
5	3.64	1.85	-0.12	-1.46
6	3.45	1.85	0.05	-1.47
7	4.12	1.74	-0.54	-1.07
8	3.32	1.70	0.11	-1.31
9	3.43	1.84	0.01	-1.45
10	4.09	1.74	-0.50	-1.10
11	4.46	1.56	-0.82	-0.44
12	4.03	1.71	-0.45	-1.10
13	4.14	1.53	-0.49	-0.82
14	3.17	1.67	0.28	-1.16
15	4.17	1.81	-0.50	-1.21
Total	55.60	13.92	-0.24	-0.27

*Table 5*  
Convergent validity. Pearson correlations between BEAQ and SCL-90-R global severity index and subscales

SCL-90-R	N	BEAQ
Anxiety	112	.34**
Depression	110	.46**
Phobic anxiety	110	.35**
Somatization	109	.27**
Obsessive-Compulsive	109	.43**
Global Severity Index	96	.48**

\*\*  $p < .01$

*Table 4*  
Internal consistency, item-total statistics ( $N = 332$ )

Item Nº BEAQ	Scale mean if item is deleted	Scale variance if item is eliminated	Corrected total-item correlation	Cronbach's alpha if item is eliminated
1	52.36	168.91	.45	.81
2	52.05	170.17	.46	.81
3	51.60	180.22	.24	.82
4	52.82	173.54	.37	.82
5	51.96	167.94	.47	.81
6	52.15	179.64	.21	.83
7	51.48	169.55	.47	.81
8	52.28	168.84	.50	.81
9	52.17	172.99	.36	.82
10	51.51	164.95	.58	.80
11	51.14	171.08	.50	.81
12	51.57	164.22	.61	.80
13	51.46	173.40	.45	.81
14	52.43	167.94	.53	.81
15	51.42	168.77	.46	.81

*Appendix*  
Brief Experiential Avoidance Questionnaire

Please indicate the extent to which you agree or disagree with each of the following statements (1= Strongly disagree; 6= Strongly agree)

- The key to a good life is never feeling any pain [La clave para vivir bien es no sentir nunca ningún dolor]
- I'm quick to leave any situation that makes me feel uneasy [Rápidamente dejo cualquier situación que me haga sentir mal]
- When unpleasant memories come to me, I try to put them out of my mind [Cuando se me vienen a la mente recuerdos desagradables, trato de desecharlos]
- I feel disconnected from my emotions [Me siento desconectado/a de mis emociones]
- I won't do something until I absolutely have to [No me pongo a hacer algo hasta que no me veo totalmente obligado/a a hacerlo]
- Fear or anxiety won't stop me from doing something important [El miedo o la ansiedad no me impedirán hacer las cosas importantes]
- I would give up a lot not to feel bad [Renunciaría a muchas cosas con tal de no sentirme mal]
- I rarely do something if there is a chance that it will upset me [Difícilmente hago algo si hay alguna posibilidad de que me desagrade]
- It's hard for me to know what I'm feeling [Es difícil para mí saber lo que estoy sintiendo]
- I try to put off unpleasant tasks for as long as possible [Trato de retrasar todo lo posible las tareas que son desagradables]
- I go out of my way to avoid uncomfortable situations [Me esfuerzo en evitar situaciones desagradables]
- One of my big goals is to be free from painful emotions [Uno de mis mayores objetivos es estar libre de cualquier emoción dolorosa]
- I work hard to keep out upsetting feelings [Me esfuerzo mucho para evitar las sensaciones desagradables]
- If I have any doubts about doing something, I just won't do it [Si tengo cualquier duda al hacer una cosa, dejo de hacerla]
- Pain always leads to suffering [El dolor siempre lleva al sufrimiento]

## Discussion

The results of this study show adequate psychometric properties of the Spanish adaptation of the BEAQ in a clinical population and advise its use as a measure of experiential avoidance in this context. Furthermore, in general terms, it could be said that this adaptation achieves psychometric characteristics similar to those reflected by the authors in the original English version of the BEAQ (Gámez et al., 2014).

The mean of the total score in the Spanish sample is 55.60. This is very similar to the two clinical samples analyzed with the original English instrument ( $M = 56.41$  and  $M = 52.03$ , respectively). The standard deviation ( $SD = 13.92$ ) is also similar to the original instrument ( $SD = 12.64$  and  $SD = 12.23$ , respectively).

This Spanish adaptation has good internal consistency ( $\alpha = .82$ ), which in turn is similar to the  $\alpha = .83$  found by Gámez et al. (2014) in their clinical population. This result supports the homogeneity of the scale and the contribution of all the items to the measurement of the experiential avoidance construct in the Spanish version analyzed.

However, if the correlation of each item with the total scale is analyzed in detail, lower scores are observed on Items 3 and 6, with correlations of .24 and .21, respectively, when .35 is considered advisable for significant correlation with the total scale (Cohen & Manion, 1989). This suggests lower contribution of the items to the construct to be measured. Item 3 contains the statement “When unpleasant memories come to me, I try to put them out of my mind”. Because of this non-significant correlation, the internal consistency of the scale was tested again without Item 3, but the Cronbach’s alpha for the scale did not increase, remaining at .82. So the item was kept to maintain the greatest possible similarity with the original scale.

The Cronbach’s alpha for the scale was also analyzed without Item 6, “Fear or anxiety won’t stop me from doing something important”. However, as it only increased slightly from .82 to .83, it was decided to keep Item 6 for its content representation, as did the authors of the BEAQ.

The temporal reliability of the BEAQ, after 4 to 6 weeks was high ( $r = .70$ ,  $p < .01$ ), showing that its measure of the experiential avoidance construct is stable over time. This contribution of temporal reliability is of special interest because there is no information on it in the original English version.

Concurrent validity of the BEAQ is shown by its significant correlation with another avoidance measure, the AAQII ( $r = .67$ ,  $p < .01$ ). In the original English version, they found a significant correlation between these two scales in their sample of students ( $r = .61$ ,  $p < .01$ ), and a non significant correlation in the two samples of patients ( $r = .62$  and  $r = .63$ , respectively). However, despite the fact that the correlation is not significant in any of these two samples, the authors of the original version concluded that the BEAQ’s correlation with other avoidance measures was stronger than its correlation with other measures of negative emotionality. It should be recalled that this was one of their main objectives, specifically, that the instrument should be a reliable and valid measure of experiential avoidance differentiated from other associated constructs, such as neuroticism. Thus, it may be concluded that in this study, the data also support the validity of the BEAQ as a measure of experiential avoidance by confirming its correlation with another measure of experiential avoidance, the AAQII.

In relation to the convergent validity of the BEAQ with the measures of psychopathology measured by the SCL-90-R, it is noteworthy that all were significant at  $p < .01$ . This is also consistent with the result for the original scale. In their study, Gámez et al. (2014) pointed out robust relationships between the BEAQ and various scales of psychopathology, among which are measures of depression, situational phobia and obsession-compulsion.

Based on these results, it may be observed, as in the original instrument, that the BEAQ showed moderate correlations with indicators of negative emotionality (mean  $r = .39$ ), but stronger correlation with a measure of experiential avoidance ( $r = .67$ ). Thus, the Spanish adaptation also achieves the objective of the authors of the original instrument of measuring experiential avoidance as discriminated from negative emotionality. These results therefore support the validity of the BEAQ as a measure of the experiential avoidance construct.

The results of factor analysis were similar to those for the English version. Using principal component analysis and forcing one factor like the original scale, 30.12% of the variance was explained. This percentage is very similar to the one presented by Gámez et al. (2014), who reported 28.98% as the variance explained by the factor. The first five eigenvalues in the Spanish sample were 4.52, 1.57, 1.11, 1.08 and .94, respectively.

Similar to this study, in a first sample of patients who were administered the English version, Item 6 had the lowest factor loading, which was  $-.33$ . The rest of the factor loadings were .38 to .63. It might be thought that because Item 6 was included to cover specific content in the MEAQ questionnaire, despite having an initial factor loading of less than .40, this means that in our sample the same trend is maintained, and that therefore, the loading on the factor is lower. On the other hand, Item 3, “When unpleasant memories come to me, I try to put them out of my mind”, also had a lower factor loading (.33) than the rest of the 13 items with significant loadings. This coincides with the findings by Gámez et al. (2014), who also found this item to have one of the lowest factor loadings. But in this case, it was significant. In view of the results above, we can affirm that the scale fits to the one-factor solution, similar to the one proposed by the authors of the original version.

Regarding limitations, it should be noted that the data was collected exclusively through self-reported measures, so they are vulnerable to possible biases associated with the use of a single method (Podsakoff, MacKenzie, Lee, & Podsakoff, 2003). Thus, it would be important to validate the BEAQ data with those found by other methods, such as professional evaluation of the patient concerning the experiential avoidance construct. Another limitation was the non-randomized sample. Therefore, the sample could not be representative of the Spanish population with clinical disorders. Moreover, the diagnosis is not based on a structured interview, but on the clinical judgement of a clinical psychologist or psychiatrist after a clinical interview based on the ICD-10 diagnostic criteria. Besides, the study does not use an infrequency response scale to identify dishonest or non-compliant participants.

Future research should confirm the factor structure with confirmatory factor analysis in a larger sample. In addition, the psychometric characteristics of the present scale should be studied in a Spanish clinical population with other psychopathological disorders, as well as in the general population.

In view of the results of this study, we can conclude by highlighting the adequate psychometric characteristics of this Spanish adaptation of the BEAQ. These good results support the use of the BEAQ as an instrument for measuring experiential avoidance in a Spanish clinical population, and is therefore a relevant contribution providing a short form (approximately 3

minutes) of the instrument considered the most appropriate at present for measuring experiential avoidance (Wolgast, 2014) in various clinical contexts. This enables its administration in care contexts where time management and limited resources prioritize instruments that take a short time for their application.

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