**Title:** An 8-item short form of the *Eating Disorder Examination-Questionnaire* adapted for children (*ChEDE-Q8*).

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#### Abstract

*Objective:* Eating disturbances are common in children placing a vulnerable group of them at risk for full-syndrome eating disorders and adverse health outcomes. In order to provide a valid self-report assessment of eating disorder psychopathology in children, a short form of the child version of the *Eating Disorder Examination (ChEDE-Q)* was psychometrically evaluated. Similar to the *EDE-Q*, the *ChEDE-Q* provides assessment of eating disorder psychopathology related to anorexia nervosa, bulimia nervosa, and binge-eating disorder. However, it does not assess symptoms of avoidant/restrictive food intake disorder, pica, or rumination disorder.

*Method:* In 1836 participants ages 7 to 18 years, recruited from two independent populationbased samples, the factor structure of the recently established 8-item short form *EDE-Q8* for adults was examined, including measurement invariance analyses on age, gender, and weight status derived from objectively measured weight and height. For convergent validity, the *ChEDE-Q* global score, Body Esteem Scale, Strengths and Difficulties Questionnaire, and sociodemographic characteristics were used. Item characteristics and age- and gender-specific norms were calculated.

*Results:* Confirmatory factor analysis revealed good model fit for the 8-item *ChEDE-Q*. Measurement invariance analyses indicated strict invariance for all analyzed subgroups. Convergent validity was provided through associations with well-established questionnaires and age, gender, and weight status, in expected directions.

*Discussion:* The newly developed *ChEDE-Q8* proved to be a psychometrically sound and economical self-report assessment tool of eating disorder psychopathology in children. Further validation studies are needed, particularly concerning discriminant and predictive validity.

# An 8-item short form of the Eating Disorder Examination-Questionnaire adapted for children (*ChEDE-Q8*)

Eating disturbances are relatively common in childhood and adolescence,<sup>1–7</sup> including a wide range of eating or weight-related symptoms (i.e., loss of control eating, binge eating, avoidant or restrictive eating), compensatory strategies of weight control (i.e., self-induced vomiting or misuse of laxatives), and associated general eating disorder psychopathology, including dietary restraint and concerns about body image and weight.<sup>8</sup> Although the tendency to remit is high in children<sup>9</sup>, an early manifestation of eating disturbance is a frequently reported risk factor for later full-blown eating disorders,<sup>10–14</sup> increased psychopathology,<sup>12,15–18</sup> obesity,<sup>12,15,16,18</sup> and impairment in quality of life<sup>19</sup>.

Given the lack of psychometrically sound instruments for children and youth, the child version of the *Eating Disorder Examination-Questionnaire* (*EDE-Q*),<sup>20</sup> the *ChEDE-Q*, was developed for children aged 8-14 years<sup>21</sup>, and evaluated in samples of children and adolescents aged 10-16<sup>22</sup>, 12-17<sup>23</sup> and 8-13<sup>24</sup> years. This self-report instrument contains, like the adult version, 22 items allocated to four subscales (restraint, eating concern, weight concern, shape concern) which can be integrated into one global score of eating disorder psychopathology. Six additional key behavioral items measure diagnostically relevant information, for example, binge eating or self-induced vomiting.

With the *EDE-Q8*<sup>25</sup> an 8-item short form for the assessment of eating disorder psychopathology was developed and evaluated using two representative population surveys in Germany, for use in adults and adolescents aged  $\geq 14$  years. The *EDE-Q8* had excellent item characteristics ( $r_{it} > .30$ ), very good reliability ( $\alpha > .90$ ), and a very good model fit for the postulated second-order factorial structure. Furthermore, a strong correlation (r=.75) between the *EDE-Q8* and a 13-item short form of the Eating Attitudes Test<sup>26</sup> was observed.

Along the lines of the *EDE-Q*, the *ChEDE-Q* provides assessment of eating disorder psychopathology related to anorexia nervosa (AN), bulimia nervosa (BN), and binge-eating

disorder (BED), but it does not assess symptoms of avoidant/restrictive food intake disorder (ARFID), pica, or rumination disorder. The aim of this study was to evaluate a comparable short form of the *ChEDE-Q*, based on an item selection of the *EDE-Q8*, for an economical assessment of eating disorder psychopathology related to AN, BN, and BED in children and youth.

#### Method

#### **Sampling Procedure**

Children and adolescents were recruited for two different research projects at the Universities of Marburg (Sample 1) and Leipzig (Sample 2): Sample 1 included an unselected community-based sample, recruited for a research project on loss of control eating between 2005-2007 (for methodological detail see Hilbert and Czaja<sup>27</sup>). N=594 8-13 year old children were assessed with the *ChEDE-Q*. Sample 2 is part of the ongoing 'Leipzig Research Centre for Civilization Diseases (LIFE)' Child study which started in 2011. One aim of this prospective population-based cohort study is to identify risk factors of childhood obesity and associated mental disorders (for methodological detail see Quante et al.<sup>28</sup>). N=1242 7-18 year old youth were investigated with the *ChEDE-Q*. Broadening of the age range was allowed to evaluate suitability for higher (and lower) ages than postulated.<sup>20</sup> Using multigroup confirmatory factor analysis, we found strict measurement invariance (see Statistical Analyses) between both samples (see Appendix A). Thus, pooling both samples was considered to be adequate.

#### **Participants**

Sample characteristics of the total sample (N=1836) are displayed in Table 1. The two projects were approved by the Ethics Committees of the German Psychological Society and the Medical Faculty of the University of Leipzig, Germany (Reg. No. 264-10-19042010). Informed consent and assent were obtained from at least one parent of the under-aged participants.

#### Measures

In sample 2, the *appearance* subscale of the child version of the *Body Esteem Scale* (BES)<sup>29</sup>, containing 12-items on a 4-point Likert scale (1 = "strongly disagree" to 4 = "strongly agree"), and the *emotional problems* subscale of the *Strengths and Difficulties Questionnaire* (SDQ)<sup>30</sup>, containing five items on a 3-point scale (0 = "not true" to 2 = "certainly true") were included. Higher scores on the *BES* indicate greater body satisfaction; higher scores on the *SDQ* subscale indicate more emotional problems. The observed internal consistencies of the subscales of the *BES* ( $\alpha$ =.92) and *SDQ* ( $\alpha$ =.69) were excellent to acceptable. The body mass index (BMI, kg/m<sup>2</sup>) of children was calculated from objectively measured weight and height in sample 2 (height and weight were not measured for sample 1) and subsequently standardized regarding age and sex using German reference data (i.e., BMI standard deviation scores [SDS]).<sup>31</sup>

#### **Statistical Analyses**

The German version of the *ChEDE-Q8* was used, embedded in the *ChEDE-Q*<sup>32</sup> (i.e., children completed the *ChEDE-Q*, but only eight items were considered in the data analyses; for items see Table 2). At the item level, means and standard deviations were determined. Item selectivity indicating if one item varies in line with other items was computed as the correlation of the respective item with the sum of all other items ( $r_{it}$ , item-rest correlation coefficients). In accordance with Field<sup>33</sup>, item selectivity values less than 0.30 can be classified as unsatisfactory. Item difficulty coefficients were calculated as quotients of the sum of the item values that were obtained, and the sum of the maximum achievable item values, multiplied by 100 (possible range: 0 [none of the participants answered the item positively] to 100 [all of the participants answered the item positively]). Differences in *ChEDE-Q8* global scores were analyzed using two-factorial (age and gender) analysis of variance (ANOVA). Missing data (0.1–0.7% per item) were imputed using chained equation

modeling,<sup>34</sup> based on gender and age. Predictive mean matching was used for imputation (i.e., only realistic values were computed). Construct validity of the ChEDE-Q8 was examined through positive correlations with the full ChEDE-Q (including and excluding ChEDE-Q8 items), the emotional problems subscale of the SDQ, and weight (i.e., BMI-SDS). Furthermore, we expected a negative correlation with the appearance subscale of the child version of the BES. Based on the EDE-Q8,<sup>25</sup> factorial validity was tested using confirmatory factor analysis (CFA) by a second-order general factor model with four first-order factors comprising the postulated subscales. We used robust maximum likelihood estimation with the mean-adjusted Satorra-Bentler  $\chi^2$  test statistic.<sup>35</sup> To evaluate goodness-of-fit, the following four criteria were considered: The Standardized Root Mean Square Residual (SRMR) and the Root Mean Square of Approximation (RMSEA) with its 90% confidence interval (CI) were used to assess absolute model fit. The Comparative Fit Index (CFI) and Tucker Lewis Index (TLI) measure relative model fit compared to the "null" model. RMSEA and SRMR values < .050 represent a close fit, values between .050 and .080 a reasonably close fit, and values > .080 an unacceptable model.<sup>30</sup> Regarding CFI and TLI, Hu and Bentler<sup>36</sup> suggested a CFI and TLI >.900 for an adequate fit and a CFI and TLI >.950 for a good model fit.

The *ChEDE-Q8* is a complex instrument and it could be considered challenging to complete for children at the younger end of the age range included in this study. For this reason, we investigated age-related effects on the measurement model of the *ChEDE-Q8* using analysis of measurement invariance (group 1: 7 to 10 years; group 2: >10 to 12 years; group 3: >12 to 14 years; group 4: >14 to 18 years, see Table 1). Furthermore, we conducted measurement invariance tests across gender and weight status (obesity and overweight vs. normal weight). Underweight was not analyzed in measurement invariance testing because of few underweight cases

In accordance with the sequential strategy developed by Meredith and Teresi<sup>37</sup>, the weak invariance model was first tested by constraining the estimate factor loadings to be

equal across groups. If empirical support for weak invariance is provided, it allows for the comparison of structural relationships between latent constructs in groups. Second, the strong invariance model was tested by constraining both loadings and intercepts to be equal across groups. This level of invariance allows for the comparison of means of the latent construct between groups. Finally, we tested the strict invariance model by constraining the loadings, intercepts, and item error variances to be equal across groups. Lack of strict invariance can affect decisions in screening processes that depend on the expression of a construct, resulting in different error rates (i.e., sensitivity, specificity) for different groups.<sup>38</sup>

As recommended by Chen,<sup>39</sup> a change of  $\Delta CFI \leq -.010$  in CFI, supplemented by a change of  $\Delta RMSEA \geq .015$ , was considered as indicative of non-invariance. Data analysis was carried out using the R packages lavaan<sup>40</sup> and mice.<sup>41</sup>

#### Results

#### **Item Characteristics**

Means and standard deviations, and item characteristics of the *ChEDE-Q8* are depicted in Table 2. The global mean score of the *ChEDE-Q8* was 1.05 (*SD*=1.33), with an observed range from 0 to 6. Item difficulty values ( $p_i$ ) ranged between  $p_i$ =9 (preoccupation with food) and  $p_i$ =27 (desire to lose weight). Regarding item selectivity, all item-rest correlation coefficients exceeded the relevant cut-off (range:  $r_{it}$ =.43 [preoccupation with food] to  $r_{it}$ =.80 [feelings of fatness]). For some items, there were statistically significant differences between boys and girls, however, with negligible effect sizes (d=0.04 to 0.12).

### Norms

A two-factorial ANOVA with the factors gender and age resulted in a significant interaction effect (F(3, 1836)=6.50; p<.001). Thus, age- and gender-specific norms are provided in Table 3. For practical reasons, we determined severity categories in accordance with the *ChEDE-Q8* percentile ranks. The level of eating disorder psychopathology was

classified as critical if the percentiles were between the 90th and 95th percentile, as high if the percentiles were between the 95th and 98th, and as very high if the percentiles were above the 98th percentile.

#### **Internal Consistency**

With regard to the global score of the *ChEDE-Q8*, the internal consistency for the total sample was  $\alpha$ =.89 (boys:  $\alpha$ =.87; girls:  $\alpha$ =.90).

### **Factorial Validity**

CFA revealed good fit parameters for the second-order general factor model (SRMR=.042; RMSEA=.077, 90% CI [.070; .083]; CFI=.946; TLI=.916). Second-order corrected factor loadings were medium to high (.41 to .87). Detailed results are displayed in Appendix B. Measurement invariance analyses indicated strict invariance for all of the analyzed subgroups. Detailed results can be found in Appendix A of the supporting information.

#### **Construct Validity**

Correlations of the *ChEDE-Q8* global score with the *ChEDE-Q* global score (r=.97, p<.001) as well as with the *ChEDE-Q* global score not including the items of the short form (r=.91; p<.001) were very high. The correlation coefficients were as follows: *BES*: r=-.65, p<.001; *SDQ emotional problems*: r=.37, p<.001 BMI-SDS: r=.51, p<.001.

#### Discussion

Based on a large-scale population-based sample of more than 1800 children, the present study provided a psychometrically sound 8-item short form of the commonly used *ChEDE-Q* for assessing children's eating disorder psychopathology. The newly developed *ChEDE-Q8* showed good internal consistency, a clear second-order one-factorial structure, and satisfactory convergent validity with well-established measures. Age- and gender-specific population-based norms of the *ChEDE-Q8* allow for directly evaluating the severity of

children's eating disorder symptoms. Because children and adolescents between 7 and 18 years of age were analyzed, the ChEDE-Q8 was proved to qualify for use over a wide age range. There is an ongoing debate about difficulties around the use of a questionnaire version of the *EDE* with younger individuals. For example, Goldschmidt and colleagues<sup>23</sup> argue that younger children may not be able to understand complex items of ChEDE-Q without assistance. We agree that in younger children the interview-based assessment of eating disorder psychopathology is preferable. On the other hand, face-to-face interviews cannot be applied in all areas (e.g., epidemiological studies or longitudinal studies with many assessment points) due to a lack of time or for financial reasons. There are thus some arguments in favor of conducting the ChEDE-Q8 as a self-report assessment of eating disorder psychopathology in children. First, the questionnaire is relatively short and therefore appears suitable for young children. Second, the instrument allows for a reliable assessment in younger age groups (<12 years:  $\alpha$ =.87). Third and most importantly, confirmation of strict measurement invariance indicated that the ChEDE-Q8 measures the same construct across different children's age groups, which is a prerequisite for meaningful cross-group comparisons and emphasizes high measurement quality. However, in future studies, results obtained using the ChEDE-Q8 should be compared to the results of a diagnostic interview (e.g., *ChEDE*).

#### Limitations

In order to achieve sufficient sample size, two different samples were combined for analysis. Although both samples were unselected, included boys and girls, and covered a broad age and socioeconomic range, measurement invariance analyses were conducted to account for these heterogeneities. Furthermore, it is unclear whether the psychometric properties of the *ChEDE-Q8* would also apply to child or adolescent samples with more severe eating disorder psychopathology or to samples from other countries. This is certainly an important area for future research, along with the need for examining predictive and

discriminant validity of the *ChEDE-Q8*. Because the *ChEDE-Q8* was derived from the adult *EDE-Q8*,<sup>12</sup> there may be debate about whether age-specific features of eating disorder psychopathology are adequately represented. However, additional exploratory factor analyses produced an item selection identical with that found for the adult *EDE-Q8* (data not shown). Finally, although the *ChEDE-Q8* targets children ages 8-14 years, in accordance with the *ChEDE-Q*, the present study included somewhat younger children and older adolescents. The fact that the *ChEDE-Q8* showed favorable psychometrics for these broadened age ranges, however, highlighted its suitability and comparability across childhood and adolescence, so that methodological variances (e.g., based on changes in assessments) can be reduced. While the *ChEDE-Q8* targets at the assessment of eating disorder psychopathology related to AN, BN, and BED, just as the *ChEDE-Q*, alternative measures are needed to identify symptoms of ARFID, pica, or rumination disorder, e.g. the *Eating Disturbances in Youth-Questionnaire.*<sup>42</sup>

#### Conclusions

Although further validation studies are needed, particularly with respect to discriminant and predictive validity, the results suggested that the *ChEDE-Q8* has a clear unidimensional factor structure and favorable psychometrics, and thus represents an evidence-based, economical self-report assessment of eating disorder psychopathology in children. While the *ChEDE-Q8* appears to be particularly suitable within an epidemiological framework, the assessment of global eating disorder psychopathology, unifying restraint as well as eating, weight and shape concern, provides clinically relevant information on the level of eating disturbances in youth.

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## Table 1

## Sociodemographic characteristics

| Characteristics |            | Sample 1         |       |        | Sample 2          |       |                   | Total Sample | :     |  |
|-----------------|------------|------------------|-------|--------|-------------------|-------|-------------------|--------------|-------|--|
|                 |            | ( <i>n</i> =594) |       |        | ( <i>n</i> =1242) |       | ( <i>N</i> =1836) |              |       |  |
| Gender          |            | n (%)            |       |        | n (%)             |       | n (%)             |              |       |  |
| Boys            | 290 (48.8) |                  |       |        | 637 (51.3)        |       | 927 (50.5)        |              |       |  |
| Girls           |            | 304 (51.2)       |       |        |                   |       | 909 (49.5)        |              |       |  |
| Age             | M (SD)     | median           | range | M (SD) | median            | range | M(SD)             | median       | range |  |
|                 | 10.81      | 11.00            | 8.00- | 11.95  | 11.00             | 7.12- | 11.58             | 11.20        | 7.12- |  |
|                 | (1.48)     | 11.00            | 16.81 | (2.25) | 11.80             | 17.99 | (2.10)            | 11.50        | 17.99 |  |
| Age group       |            | n (%)            |       |        | n (%)             |       |                   | n (%)        |       |  |
| 7 to 10 y       |            | 117 (19.7)       |       |        | 265 (21.3)        |       |                   | 382 (20.8)   |       |  |
| >10 to 12 y     |            | 292 (49.2)       |       |        | 387 (31.2)        |       |                   | 676 (37.0)   |       |  |
| >12 to 14 y     |            | 179 (30.1)       |       |        | 348 (28.0)        |       |                   | 527 (28.7)   |       |  |
| >14 to 18 y     |            | 6 (1.0)          |       |        | 242 (19.5)        |       |                   | 248 (13.5)   |       |  |
|                 |            |                  |       |        |                   |       |                   |              |       |  |

(continued)

## Table 1 (continued)

## Sociodemographic characteristics

| Characteristics      | Sample 1         | Sample 2          | Total Sample      |
|----------------------|------------------|-------------------|-------------------|
|                      | ( <i>n</i> =594) | ( <i>n</i> =1242) | ( <i>N</i> =1836) |
| Weight status        | n (%)            | n (%)             | n (%)             |
| Underweight          |                  |                   | 07 (5.2)          |
| (BMI-SDS≤-1.28)      |                  | 97 (7.8)          | 97 (5.3)          |
| Normal weight        |                  | 791 (63 7)        | 791 (43-1)        |
| (-1.28>BMI-SDS<1.28) |                  | (05.7)            | (7)1 (43.1)       |
| Overweight           |                  | 107 (9.6)         | 107 (5.8)         |
| (1.28≤BMI-SDS<1.88)  |                  | 107 (8.0)         | 107 (3.8)         |
| Obesity              |                  | 218(17.6)         | 219(110)          |
| (BMI-SDS≥1.88)       |                  | 210 (17.0)        | 210 (11.7)        |
| Missing              | 594 (100)        | 29 (2.3)          | 623 (34.0)        |

*Note.* Sample 1 = University of Marburg; Sample 2 = University of Leipzig; BMI-SDS = body mass index-standard deviation score. BMI-SDS is a measure of relative weight adjusted for child's age and sex.<sup>35</sup> BMI-SD scores for Sample 2 were determined based on objectively measured weight and height. As for Sample 1 only subjective weight and height were available, BMI-SD scores were not calculated.

## Table 2

|                             | Т    | otal Sa | mple  | ;                      | Boys Girls |      |       |                 | Group Differences |      |       |                        |       |      |       |      |
|-----------------------------|------|---------|-------|------------------------|------------|------|-------|-----------------|-------------------|------|-------|------------------------|-------|------|-------|------|
|                             | М    | SD      | $p_i$ | <i>r</i> <sub>it</sub> | М          | SD   | $p_i$ | r <sub>it</sub> | М                 | SD   | $p_i$ | <i>r</i> <sub>it</sub> | t     | df   | р     | d    |
| Restraint over Eating       | 0.97 | 1.67    | 16    | .57                    | 0.91       | 1.65 | 15    | 0.55            | 1.04              | 1.69 | 17    | .59                    | -1.73 | 1834 | .085  | 0.04 |
| Food Avoidance              | 0.63 | 1.33    | 11    | .55                    | 0.58       | 1.31 | 10    | 0.52            | 0.68              | 1.36 | 11    | .57                    | -1.53 | 1834 | .126  | 0.04 |
| Preoccupation with Food     | 0.56 | 1.34    | 9     | .43                    | 0.59       | 1.40 | 10    | 0.43            | 0.53              | 1.27 | 9     | .49                    | 0.85  | 1834 | .394  | 0.02 |
| Feelings of Fatness         | 1.46 | 2.15    | 24    | .80                    | 1.24       | 2.02 | 21    | 0.77            | 1.69              | 2.26 | 28    | .82                    | -4.45 | 1834 | <.001 | 0.10 |
| Desire to Lose Weight       | 1.59 | 2.33    | 27    | .79                    | 1.41       | 2.25 | 24    | 0.75            | 1.78              | 2.40 | 30    | .82                    | -3.40 | 1834 | .001  | 0.08 |
| Guilt about Eating          | 0.68 | 1.32    | 11    | .65                    | 0.54       | 1.17 | 9     | 0.62            | 0.83              | 1.44 | 14    | .67                    | -4.71 | 1834 | <.001 | 0.11 |
| Dissatisfaction with Weight | 1.43 | 1.99    | 24    | .78                    | 1.19       | 1.89 | 20    | 0.75            | 1.68              | 2.07 | 28    | .80                    | -5.28 | 1834 | <.001 | 0.12 |
| Discomfort Seeing Body      | 1.06 | 1.74    | 18    | .75                    | 0.86       | 1.58 | 14    | 0.71            | 1.25              | 1.87 | 21    | .78                    | -4.86 | 1834 | <.001 | 0.11 |
| Global score                | 1.05 | 1.33    | 18    | -                      | 0.92       | 1.23 | 15    | -               | 1.19              | 1.41 | 20    | -                      | -4.37 | 1834 | <.001 | 0.10 |

*Means (M), standard deviation (SD), item difficulty (p\_i), selectivity (r\_{it}), and gender differences for the ChEDE-Q8 items and global score* 

## Table 3

|              | Total<br>Sample |            |          | Boys             |          |                  |   |         | Girls            |          |                  |
|--------------|-----------------|------------|----------|------------------|----------|------------------|---|---------|------------------|----------|------------------|
| ChEDE-08     | 7-18v           | 7-18v      | 7-10v    | >10-             | >12-     | >14-             | 7-18v                                   | 7-10v   | >10-             | >12-     | >14-             |
| global score | (n=1836)        | (n=927)    | (n=185)  | 12v              | 14v      | 18v              | (n=909)                                 | (n=197) | 12v              | 14v      | 18v              |
| 8            |                 |            |          | ( <i>n</i> =366) | (n=265)  | ( <i>n</i> =111) | ( ) ) ) ) ) ) ) ) ) ) ) ) ) ) ) ) ) ) ) |         | ( <i>n</i> =313) | (n=262)  | ( <i>n</i> =137) |
| 0.00         | 30              | 33         | 33       | 32               | 31       | 40               | 27                                      | 29      | 33               | 24       | 18               |
| 0.13         | 39              | 43         | 42       | 42               | 42       | 49               | 35                                      | 41      | 43               | 29       | 22               |
| 0.25         | 45              | 49         | 49       | 48               | 49       | 55               | 41                                      | 48      | 49               | 36       | 26               |
| 0.38         | 50              | 54         | 54       | 53               | 53       | 60               | 47                                      | 55      | 54               | 40       | 29               |
| 0.50         | 54              | 58         | 58       | 56               | 56       | 64               | 51                                      | 59      | 58               | 45       | 33               |
| 0.63         | 58              | 61         | 60       | 61               | 59       | 67               | 54                                      | 62      | 63               | 49       | 36               |
| 0.75         | 61              | 65         | 66       | 65               | 62       | 67               | 57                                      | 65      | 66               | 50       | 39               |
| 0.88         | 64              | 68         | 70       | 69               | 64       | 69               | 60                                      | 67      | 69               | 53       | 43               |
| 1.00         | 66              | 69         | 71       | 71               | 66       | 71               | 62                                      | 70      | 70               | 57       | 45               |
| 1.13         | 68              | 72         | 74       | 72               | 69       | 75               | 65                                      | 73      | 71               | 58       | 50               |
| 1.25         | 71              | 74         | 77       | 75               | 71       | 76               | 67                                      | 75      | 74               | 58       | 53               |
| 1.38         | 72              | 76         | 80       | 76               | 73       | 77               | 67                                      | 76      | 74               | 59       | 54               |
| 1.50         | 74              | 78         | 81       | 78               | 75       | 77               | 70                                      | 80      | 75               | 61       | 57               |
| 1.63         | 75              | 79         | 82       | 81               | 76       | 78               | 71                                      | 83      | 76               | 62       | 58               |
| 1.75         | 77              | 81         | 83       | 82               | 79       | 78               | 72                                      | 84      | 78               | 62       | 61               |
| 1.88         | /8              | 82         | 84       | 83               | 80       | 80               | 74                                      | 85      | /9               | 64       | 61               |
| 2.00         | /9              | 83         | 85       | 84               | 80       | 81               | /5                                      | 86      | 80               | 66       | 65               |
| 2.13         | 80              | 84         | 85       | 85               | 81       | 82               | /6                                      | 8/      | 80               | 68<br>71 | 66               |
| 2.25         | 82<br>82        | 80<br>97   | 88       | 8/               | 83<br>95 | 82               | /8                                      | 89      | 82               | /1       | 09<br>70         |
| 2.38         | 83              | 8/         | 90<br>01 | 88               | 85<br>97 | 82<br>82         | /9<br>81                                | 90      | 83<br>85         | 71       | 70               |
| 2.50         | 04<br>85        | 00<br>80   | 91       | 00<br>80         | 07       | 03<br>95         | 81                                      | 90      | 0J<br>86         | 74       | 74               |
| 2.03         | 85              | 80         | 91       | 09               | 00       | 85<br>86         | 82<br>83                                | 91      | 80<br>87         | 74       | 74               |
| 2.75         | 87              | 00         | 01       | 90               | 88       | 88               | 85                                      | 02      | 88               | 70       | 70               |
| 3.00         | 88              | 90<br>01   | 02       | 90<br>01         | 80       | 80               | 86                                      | 92      | 90               | 70       | 80               |
| 3.13         | 89              | 91         | 92       | 93               | 89       | 90               | 80                                      | 93      | 91               | 82       | 81               |
| 3 25         | 91              | 92         | 93       | 94               | 90       | 91               | 89                                      | 95      | 93               | 82       | 83               |
| 3.38         | 92              | 9 <u>4</u> | 94       | 95               | 93       | 92               | 90                                      | 95      | 94               | 84       | 85               |
| 3.50         | 93              | 94         | 95       | 95               | 94       | 92               | 91                                      | 95      | 94               | 86       | 85               |
| 3.63         | 94              | 95         | 95       | 95               | 95       | 93               | 92                                      | 97      | 96               | 88       | 87               |
| 3.75         | 95              | 96         | 96       | 96               | 96       | 95               | 93                                      | 97      | 96               | 90       | 88               |
| 3.88         | 95              | 96         | 97       | 97               | 97       | 95               | 94                                      | 97      | 96               | 92       | 90               |
| 4.00         | 96              | 97         | 97       | 97               | 97       | 96               | 95                                      | 98      | 96               | 94       | 90               |
| 4.13         | 97              | 98         | 99       | 98               | 97       | 96               | 96                                      | 98      | 97               | 95       | 90               |
| 4.25         | 97              | 98         | >99      | 98               | 97       | 97               | 97                                      | 99      | 97               | 97       | 92               |
| 4.38         | 98              | 98         | >99      | <b>98</b>        | 98       | 97               | 97                                      | 99      | 97               | 97       | 92               |
| 4.50         | 98              | 99         | >99      | <b>98</b>        | 99       | 98               | 97                                      | 99      | <b>98</b>        | 97       | 93               |
| 4.63         | 98              | 99         | >99      | 99               | 99       | 98               | 98                                      | 99      | 98               | 98       | 95               |
| 4.75         | 99              | 99         | >99      | 99               | 99       | 98               | 98                                      | 99      | 99               | 99       | 96               |
| 4.88         | 99              | 99         | >99      | 99               | 99       | 99               | 99                                      | >99     | 99               | 99       | 96               |
| 5.00         | 99              | 99         | >99      | 99               | 99       | 99               | 99                                      | >99     | >99              | 99       | 96               |
| 5.13         | 99              | 99         | >99      | 99               | 99       | 99               | 99                                      | >99     | >99              | 99       | 96               |
| 5.25         | 99              | >99        | >99      | >99              | 99       | >99              | 99                                      | >99     | >99              | 99       | 97               |
| 5.38         | >99             | >99        | >99      | >99              | >99      | >99              | >99                                     | >99     | >99              | >99      | 99               |
| 5.50         | >99             | >99        | >99      | >99              | >99      | >99              | >99                                     | >99     | >99              | >99      | 99               |
| 5.63         | >99             | >99        | >99      | >99              | >99      | >99              | >99                                     | >99     | >99              | >99      | >99              |
| 5./5         | >99             | >99        | >99      | >99              | >99      | >99              | >99                                     | >99     | >99              | >99      | >99              |
| 5.88         | >99             | >99        | >99      | >99              | >99      | >99              | >99                                     | >99     | >99              | >99      | >99              |
| 0.00         | >99             | >99        | >99      | >99              | >99      | >99              | >99                                     | >99     | >99              | >99      | >99              |

Percentiles for the ChEDE-Q8 by age and gender

*Note*. Normative data are presented as *ChEDE-Q8* global mean scores with corresponding percentiles. Percentiles are shown for the total sample and for subsamples based on age and gender; light grey area indicate critical severity of eating disorder psychopathology; grey area indicate high severity of eating disorder psychopathology; dark grey area indicate very high severity of eating disorder psychopathology.

# Appendix A

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|   | 2        | 10  | <u>CEI</u> |       |       |        | Measurement                  |
|---|----------|-----|------------|-------|-------|--------|------------------------------|
|   | $\chi^2$ | df  | CFI        | ΔCFI  | RMSEA | ΔRMSEA | Invariance Test <sup>a</sup> |
| Gender (boys; girls)  |          |     |            |       |       |        |                              |
| Configural invariance                                       | 531.49   | 36  | .941       | -     | .079  | -      |                              |
| Weak invariance (equal loadings)                            | 544.12   | 43  | .940       | 001   | .073  | 006    |                              |
| Strong invariance (equal loadings + intercepts)             | 564.48   | 46  | .936       | 04    | .073  | .000   | $\checkmark$                 |
| Strict invariance (equal loadings + intercepts + residuals) | 606.06   | 54  | .938       | +.002 | .066  | 007    | $\checkmark$                 |
| Age (7-10y; >10-12y; >12-14y; >14-18y)                      |          |     |            |       |       |        |                              |
| Configural invariance                                       | 555.31   | 72  | .947       | -     | .075  | -      |                              |
| Weak invariance (equal loadings)                            | 611.65   | 93  | .943       | 004   | .069  | 006    | $\checkmark$                 |
| Strong invariance (equal loadings + intercepts)             | 661.23   | 102 | .935       | 008   | .070  | +.001  | $\checkmark$                 |
| Strict invariance (equal loadings + intercepts + residuals) | 848.36   | 126 | .929       | 006   | .066  | 004    | $\checkmark$                 |
|   |          |     |            |       |       |        |                              |

Analysis of factorial invariance using multigroup confirmatory factor analyses

(continued)

## Appendix A (continued)

| A 1 ·            | CC · 1   | • •               | •     | 1          | C•         | C ·          | 1        |
|------------------|--|-------------------|-------|------------|------------|--------------|----------|
| $\Delta nalvere$ | of factorial                                       | invariance        | usina | multioroun | contirmato | ry tactor    | analyses |
|                  | $o_i$ incrotin                                     | <i>invariance</i> | usuig | manistond  | communu    | i v $iacior$ | unuivses |
|                  | · <b>J J</b> · · · · · · · · · · · · · · · · · · · |                   |       |            |            | J J          |          |

| x<br>401.60<br>417.72<br>419.62 | <i>ay</i><br>36<br>43<br>46                                 | .924<br>.916  | 008  | .090<br>.087  | 003   | Invariance Test <sup>a</sup>  |
|---------------------------------|---|---|--|---|---|---|
| 401.60<br>417.72<br>419.62      | 36<br>43<br>46  | .924<br>.916  | 008  | .090<br>.087  | 003   | √   |
| 401.60<br>417.72<br>419.62      | 36<br>43<br>46  | .924<br>.916<br>916   | -<br>008   | .090<br>.087  | 003   | $\checkmark$  |
| 417.72<br>419.62                | 43<br>46  | .916  | 008  | .087  | 003   | $\checkmark$  |
| 419.62                          | 46  | 016   |  |   |   |   |
|                                 |   | .910  | 000  | .084  | 003   | $\checkmark$  |
| 688.75                          | 54  | .901  | 015  | .084  | .000  | $\checkmark$  |
|                                 |   |   |  |   |   |   |
| 506.97                          | 36  | .945  | -  | .078  | -   | $\checkmark$  |
| 539.16                          | 43  | .941  | 004  | .074  | 004   | $\checkmark$  |
| 552.00                          | 46  | .938  | 003  | .073  | 001   | $\checkmark$  |
| 613.79                          | 54  | .937  | 001  | .069  | 004   | $\checkmark$  |
|                                 | 688.75<br>506.97<br>539.16<br>552.00<br>613.79<br>models (1 | 688.75 54<br>506.97 36<br>539.16 43<br>552.00 46<br>613.79 54<br>models (1 and 2: | 688.75 54 .901<br>506.97 36 .945<br>539.16 43 .941<br>552.00 46 .938<br>613.79 54 .937<br>models (1 and 2; 2 and 3 | 688.75 54 .901015<br>506.97 36 .945 -<br>539.16 43 .941004<br>552.00 46 .938003<br>613.79 54 .937001<br>models (1 and 2; 2 and 3; 3 and 4 | 688.75       54       .901      015       .084         506.97       36       .945       -       .078         539.16       43       .941      004       .074         552.00       46       .938      003       .073         613.79       54       .937      001       .069         models (1 and 2; 2 and 3; 3 and 4); RMSEA=r | 688.75       54       .901      015       .084       .000         506.97       36       .945       -       .078       -         539.16       43       .941      004       .074      004         552.00       46       .938      003       .073      001         613.79       54       .937      001       .069      004 |

 $\Delta$ RMSEA=differences between models (1 and 2; 2 and 3; 3 and 4); <sup>a</sup> $\Delta$ *CFI* $\leq$ -.010 supplemented by  $\Delta$ *RMSEA* $\geq$ .015 indicates non-invariance.  $\sqrt{}$  indicates invariance

## Appendix B.

Confirmatory factor analysis of the higher order general factor model.



*Note:* The model was estimated using MLM estimation. Loadings are standardized. Grey loading indicates the second-order corrected factor loadings (item loading \* one-order factor loading).

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