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RESEARCH ARTICLE

Morbidity after surgical management of cervical cancer in low and middle income countries: A systematic review and metaanalysis

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Abstract

Objective

To investigate morbidity for patients after the primary surgical management of cervical cancer in low and middle-income countries (LMIC).

Methods

The Pubmed, Cochrane, the Cochrane Central Register of Controlled Trials, Embase, LILACS and CINAHL were searched for published studies from 1st Jan 2000 to 30th June 2017 reporting outcomes of surgical management of cervical cancer in LMIC. Random-effects meta-analytical models were used to calculate pooled estimates of surgical complications including blood transfusions, ureteric, bladder, bowel, vascular and nerve injury, fistulae and thromboembolic events. Secondary outcomes included five-year progression free (PFS) and overall survival (OS).

Findings

Data were available for 46 studies, including 10,847 patients from 11 middle income countries. Pooled estimates were: blood transfusion 29% (95%Cl 0.19–0.41, P = 0.00, I^2 = 97.81), nerve injury 1% (95%Cl 0.00–0.03, I^2 77.80, P = 0.00), bowel injury, 0.5% (95%Cl 0.01–0.01, I^2 = 0.00, P = 0.77), bladder injury 1% (95%Cl 0.01–0.02, P = 0.10, I^2 = 32.2), ureteric injury 1% (95%Cl 0.01–0.01, I^2 0.00, P = 0.64), vascular injury 2% (95% Cl 0.01–0.03, I^2 60.22, P = 0.00), fistula 2% (95%Cl 0.01–0.03, I^2 = 77.32, P = 0.00,), pulmonary embolism 0.4% (95%Cl 0.00–0.01, I^2 26.69, P = 0.25), and infection 8% (95%Cl 0.04–0.12,

 I^2 95.72, P = 0.00). 5-year PFS was 83% for laparotomy, 84% for laparoscopy and OS was 85% for laparotomy cases and 80% for laparoscopy.

Conclusion

This is the first systematic review and meta-analysis of surgical morbidity in cervical cancer in LMIC, which highlights the limitations of the current data and provides a benchmark for future health services research and policy implementation.

Introduction

Cervical cancer is the third most common malignancy in women worldwide and performs poorly in all objective measurements of outcomes in less developed countries[1]. The disease is a notable example of an extreme global health disparity with almost all cervical cancers, and the deaths caused by them, occurring in low and middle-income countries (LMIC)[2]. Contributors to this inequity are complex and multifaceted and include insufficient access to HPV vaccines and screening, and lack of trained health care professionals, radiation services and infrastructure, that prohibit reductions in cervical cancer incidence and mortality within these countries [3].

Opportunities to prevent and control cervical cancer arise from the human Papillomavirus (HPV) vaccine and adopting an organised approach to cervical screening. Whilst there have been global efforts towards this, for the half a million women with cervical cancer in LMIC, access to appropriate surgical, radiotherapy, and chemotherapy treatments are critical to reducing the burden of disease in LMIC and should remain a focus of global healthcare [4–7].

Surgical care should be a fundamental component of all health systems regardless of development level. This is particularly relevant in LMIC, where surgeons may be the sole physician involved in delivery of cancer care[3]. The delivery of high-quality surgical services that treat cervical cancer will prevent deaths, limit disability and suffering and promote economic growth (15). Research investigating surgical outcomes is critical to improve patient health outcomes. Practitioner adherence to protocols should be tracked and adverse events openly reviewed (15). Currently patients in LMIC are over-represented in reports of global rates of adverse surgical outcomes[8], with women's cancers having been long neglected in these settings[9]. Allowing for variations in reporting, overall complication rates for radical hysterectomy are, often from single centres[10], reported at up to 37% [11, 12], and rates of individual complications, e.g. voiding dysfunction, are reported at up to 42%[13, 14]. However, these reports largely come from high-income countries.

Data regarding cervical cancer surgical outcomes in LMIC are lacking and a baseline measure is necessary in order to advocate for ongoing quality care in these countries. Therefore, this systematic review and meta-analysis aims to report the morbidity of cervical cancer surgery in LMIC and provides a benchmark for future research initiatives.

Materials and methods

This systematic review and meta-analysis was performed in accordance with the Preferred Reporting Items for Systematic reviews and Meta-Analysis (PRISMA) guidelines [15]. We searched Pubmed, Cochrane, the Cochrane Central Register of Controlled Trials, Embase, LILACS, and CINAHL to identify all relevant articles published January 1, 2000, to June 30,

2017, without language restriction. The year 2000 was chosen as the limit of the search for both pragmatic reasons and in order that the findings were reflective of current practice.

We applied a search strategy combining relevant terms, including: Uterine Cervical Neoplasms/surgery AND ((Hysterectomy/adverse effects*) OR (Laparoscopy/adverse effects) OR (Postoperative Complications*) OR (Operative Time) OR (Perioperative Period*) OR (Blood Loss, Surgical) OR (Intraoperative Complications) OR (Ureter) OR (Cystotomy) OR (Intestines/surgery) OR (Urinary Bladder, Neurogenic) OR (Lymphocele) OR (Fistula)) OR (Hystere* OR Laparo* OR Complication* OR Operati* OR Blood OR Morbidity OR Bowel injury OR Pelvic abscess OR Vascular injury OR Lymph*)). The full search strategy can be found in S1 Table. Reference lists from full text articles identified following the title and abstract screen were hand searched for any additional references relevant to the review question.

Inclusion and exclusion criteria

Included participants were women in a LMIC [16], with a diagnosis of cervical cancer undergoing surgical intervention. The exception to this was for studies from Taiwan, for which the data (for most indicators) is added to high-income country aggregates [17]. Surgical management was defined as a simple or radical hysterectomy (+/- pelvic lymphadenectomy or pelvic lymph node dissection). All studies were required to have a minimum of 20 cervical cancer patients (at a minimum rate of 5 patients per year included, to avoid overestimation of complications, such as from small case series) and report on at least two surgical complications. Studies were excluded if there were no data available on the individual surgical complications.

Two investigators independently screened the titles and abstracts of articles retrieved from the literature search, and the full texts of potentially eligible articles were obtained and further assessed for final inclusion (Refer to Fig 1). Disagreements were resolved through consensus.

Outcome measures

The primary measure was to evaluate the reported prevalence of surgical morbidity for cervical cancer patients treated with surgery in LMIC. Secondary outcome measures included progression free survival (PFS) (the time from surgery to the first relapse, progression of existing disease or the last contact date) and OS (the interval from surgery to death (from any cause).

Classification of surgical outcomes

Complications were extracted as reported by the authors of each study (S1 File). Four authors (EA, AP, PC, YL) reviewed these in a collaborative session and agreed on the grouping of complications for analysis. Some complications were not considered surgical morbidity (for example, "IUD withdrawal difficulty") and so were excluded from any further analysis. A pragmatic approach was adopted, and complications that were only reported once and/or unlikely to require acute surgical management (e.g. lymphocyst) were excluded from the meta-analysis.

We have presented results for blood transfusions, bladder injury, ureteric injury, bowel injury, nerve injury, vascular injury, conversion to laparotomy (for planned laparoscopic cases), fistula, pulmonary embolism, combined thromboembolic events (pulmonary embolism and venous thromboembolism), and infectious morbidity. Infectious morbidity included wound infection, post-operative infection, urinary tract infection, urosepsis, vaginal cuff infection, febrile morbidity, necrotising fasciitis, pelvic cellulitis, pneumonia, pyelonephritis, and pelvic abscess.

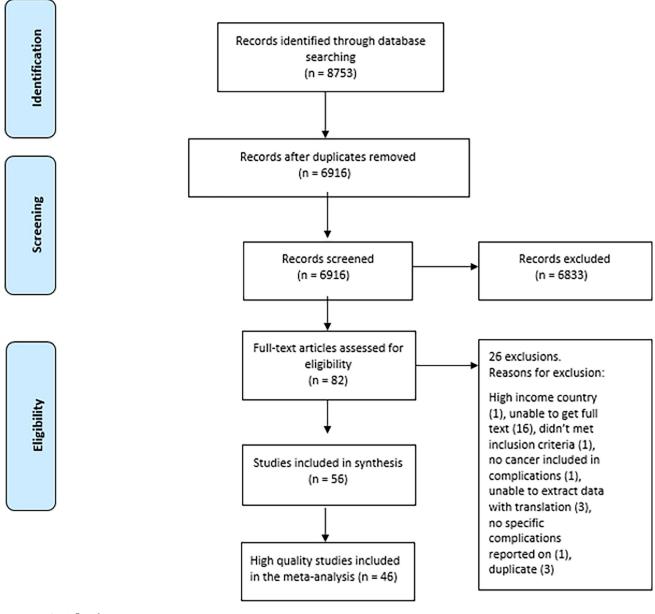


Fig 1. PRISMA flowchart.

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Assessment of methodological and reporting quality

Methodological quality was assessed using the Cochrane Risk of Bias tool for randomized trials. For non-randomised studies, the Newcastle-Ottawa Scale (NOS) was adapted and used for the cohort studies [18]. We also included three elements (all outcome measures reported, methods of assessment for outcome provided, authors discuss potential sources of bias) from the Strengthening the Reporting of Observational Studies in Epidemiology (STROBE) quality assessment tool when assessing cohort studies [19–21].

Meta-analysis

Studies were included in the meta-analysis if they were assessed as high quality. Randomised control trials were excluded from the analysis if they had a high risk of bias in any domain [22]. Cohort studies were excluded if they had a Newcastle-Ottawa score of less than 7 [23].

Statistical analysis

Summary estimates and 95% confidence intervals (CIs) for various complication types were obtained with a random effects model by pooling proportions from each study[24]. A random effects model for pooling proportions was used to account for variability in the effect estimates and subgroup meta-analyses were used to distinguish the respective treatments performed (i.e. laparotomy vs. laparoscopic surgery)[25]. The percentage of total variation across studies due to heterogeneity was evaluated by the I2 measure[26, 27]. Forrest plots were drawn showing the variation of the specific surgical complication rate among all studies together with the pooled estimate measure[28, 29]. Given the large number of studies included from China, a sensitivity analysis was conducted that excluded all Chinese studies. Funnel plots were used to assess the potential role of publication bias. Egger's test was used as a formal test of funnel plot asymmetry and publication bias[30]. Stata, version 15.0 (Stata Corp., College Station, TX, USA), was used for statistical analysis[31].

Registration

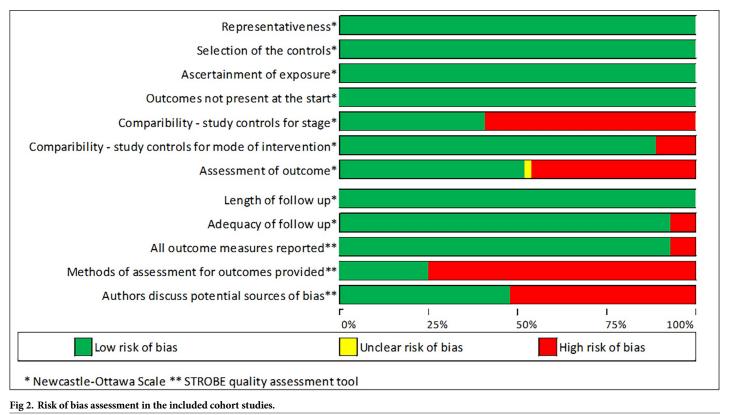
The protocol for this review was registered PROSPERO registration number CRD420170572 05.

Results

Fig 1 shows the selection of studies. A total number of 56 articles met the study inclusion criteria. Two studies which reported on trachelectomy only were excluded from the meta-analysis [32, 33]. One high quality study included stage 4 cancers (the remainder included only early stage disease 1A-2B) and was excluded from the analysis. Figs 2 and 3 present the risk of bias assessment in the included high quality studies. Studies were high quality in most domains; however, it was common that stage of disease was not controlled for, and that the assessment of the outcome and the method of assessment were not adequately described.

After excluding low-quality studies, data were available for 46 studies [34–79], for 10,847 patients (from 2 randomized trials, 9 prospective and 35 retrospective studies) from 11 middle-income countries (Fig 4). Low income country studies from Asia, Africa, South America and Russia were excluded studies as they were assessed as low quality studies. Characteristics of included studies are shown in Table 1. The median age of participants was in the 5th or 6th decade in all but one study[40]. The study authors variably described surgical approaches, and while inclusion criteria specified the need for confirmed malignancy, this was rarely confirmed with post-operative histology. 10814 patients had a radical hysterectomy. 33 had a simple hysterectomy. 2,187 patients had a laparoscopic procedure, and 8,660 patients had a laparotomy. Surgery was performed in a wide variety of centres, from general hospitals to specialised cancer units. 23 high quality studies included 1663/4572 patients (36.4%) that had neoadjuvant treatment, both chemotherapy and / or radiotherapy. The neoadjuvant treatment regimens were varied and are not reported here.

Sixty-six separate complications in all studies (high and low quality) were reported on (the full list is available in <u>S1 File</u>).



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Blood transfusion

The need for blood transfusion was reported in fifteen studies (3,108 patients) with a pooled estimate of 29% (95% CI 0.19–0.41, P = 0.00, I² = 97.81). Pooled estimates varied greatly between laparotomy cases (42%, 95% CI 0.32–0.53, I² = 95.75, P = 0.00) and laparoscopic cases (10%, 95% CI 0.04–0.16, P = 0.00, I² = 85.75). High heterogeneity was present for both analyses (refer to Fig 5).

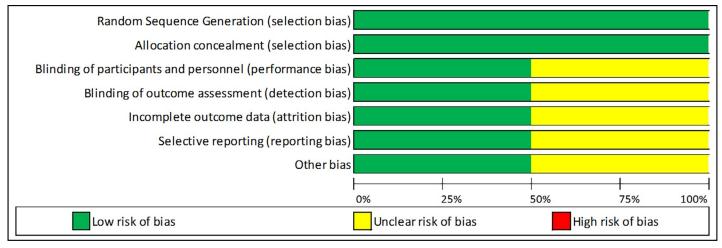
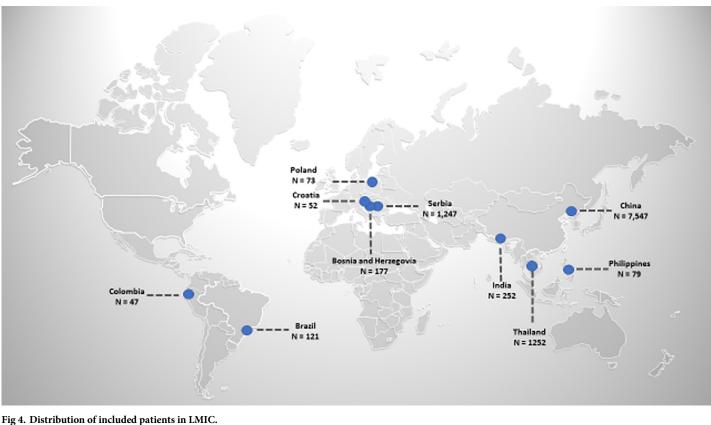


Fig 3. Risk of bias assessment in the included randomised studies.



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Nerve injury

Four studies (995 patients) reported nerve injury as an intraoperative complication. The overall pooled estimated was 1% (95% CI 0.00–0.03, I² 77.80, P = 0.00). Pooled estimates were higher for laparotomy cases (3%, 95% CI 0.05–0.14) than for laparoscopic cases (0.3%, 95% CI 0.00–0.10). Overall high heterogeneity was identified for both analyses (I² 77.8, P = 0.00).

Bowel injury

Six studies (1,957 patients) reported bowel injury, with a pooled estimate of 0.5% (95% CI 0.01–0.01, $I^2 = 0.00$, P = 0.77). Pooled estimates were slightly higher for laparotomy cases (0.7%, 95% CI 0.01–0.01) than in laparoscopic cases (0.4%, 95% CI 0.00–0.01). Low heterogeneity was identified for both analyses.

Bladder injury

Bladder injury was reported in sixteen studies (4,643 patients) and had a pooled estimate of 1% (95% CI 0.01–0.02, P = 0.10, I^2 = 32.2). Pooled estimates varied slightly between laparotomy cases 1% (95% CI 0.01–0.02, I^2 = 0.00, P = 0.92) and laparoscopic cases (2%, 95% CI 0.01–0.02, P = 0.07, I^2 36.67). Low heterogeneity was identified for laparotomy and moderate heterogeneity for laparoscopy (refer to Fig 6).

Table 1. Characteristics of included high quality studies.

Reference	Country	Study design	Total study cohort	FIGO stage	Type of surgery*	Mode(s) of surgery
Srisomboon, 2002 (24)	Thailand	RCT	100	1A2-2A	RH and PLND	Laparotomy
Cai, 2006 (25)	China	RCT	106	1B	RH and PLND	Laparotomy
Cao, 2006 [<u>36</u>]	China	Retrospective cohort	139	1A-2B	RH and PLND	Laparotomy
Coruasic, 2007 [<u>37</u>]	Croatia	Cohort**	52	1A-1B	RH	Laparoscopy, laparotomy
Li, 2007 [<u>38</u>]	China	Retrospective cohort	125	1B-2A	RH and PLND	Laparoscopy, laparotomy
Likic-Ladevic, 2007 [<u>39</u>]	Serbia	Prospective cohort	536	1B-2B	RH	Laparotomy
Puntambekar, 2007 [<u>40</u>]	India	Retrospective cohort	252	1A2-1B1	RH and PLND	Laparoscopy
Xu, 2007 [<u>41</u>]	China	Retrospective cohort	317	1B-2A	RH and PLND +/- aortic LND	Laparoscopy
Chen, 2008 [<u>42</u>]	China	Retrospective cohort	295	1A2-2B	RH and PLND +/- aortic LND	Laparoscopy
Kietpeerakool, 2008 [43]	Thailand	Retrospective cohort	357	1B2-2B	RH and PLND	Laparotomy
Likic, 2008 [<u>44</u>]	Serbia	Prospective cohort	536	1B	RH	Laparotomy
Liu, 2008[<u>96</u>]	China	Retrospective cohort	143	1A-2B	RH and PLND	Laparotomy
Cai, 2009 [<u>45</u>]	China	Prospective cohort	480	1B-2A	RH	Laparotomy
u, 2009 [<u>46</u>]	China	Retrospective cohort	93	1A-2B	RH	Laparotomy
Manchana, 2009 [<u>47</u>]	Thailand	Retrospective cohort	281	1B-2A	RH	Laparotomy
Cai, 2010 [<u>48</u>]	China	Retrospective cohort	372	1B-2A	RH	Laparotomy
Espino-Strebel, 2010 [<u>49</u>]	Philippines	Retrospective cohort	79	1-2A	RH	Laparotomy
Zhu, 2010 [<u>50</u>]	China	Retrospective cohort	132	IB-2A	RH	Laparotomy
Bezerra, 2011 [<u>51</u>]	Brazil	Retrospective cohort	88	1A-2A	RH and PLND	Laparotomy
Hou, 2011 [52]	China	Prospective cohort	63	1A-2B	RH and PLND	Laparoscopy, laparotomy
Li, 2011 [<u>53</u>]	China	Prospective cohort	73	1B2-2A	RH	Laparotomy
Lucic, 2011 [<u>54</u>]	Bosnia and Herzegovina	Retrospective cohort	177	1B1-2B	RH	Laparotomy
Yan, 2011 [<u>55]</u>	China	Retrospective cohort	240	1A2-2B	RH and PLND	Laparoscopy
Zheng, 2011 [<u>56</u>]	China	Retrospective cohort	960	1B-2B	RH	Laparotomy
Zhou, 2011 [<u>57</u>]	China	Retrospective cohort	80	1B2-2A	RH and PLND	Laparotomy
Zhu, 2011 [<u>58</u>]	China	Cohort**	61	1B1-2A	RH	Laparotomy
Achavanuntakul, 2012 [59]	Thailand	Retrospective cohort	456	1A2-2A	RH and PLND	Laparotomy
Li, 2012 [<u>60</u>]	China	Retrospective cohort	391	1A2-2B	RH and PLND +/- aortic LND	Laparotomy
Pareja, 2012 [<u>61</u>]	Colombia	Retrospective cohort	47	1A2-1B2	RH	Laparoscopy
Yan, 2012 [62]	China	Retrospective cohort	148	1B1	RH and PLND	Laparoscopy

(Continued)

Table 1. (Continued)

Reference	Country	Study design	Total study cohort	FIGO stage	Type of surgery*	Mode(s) of surgery
Ma, 2013 [<u>97]</u>	China	Retrospective cohort	50	1B2-2A	RH and PLND	Laparotomy
Mahawerawat, 2013 [<u>63</u>]	Thailand	Retrospective cohort	58	1A2	RH and PLND	Laparotomy
Chai, 2014 [<u>64</u>]	China	Retrospective cohort	438	2B	RH and PLND	Laparotomy
Favero, 2014 [<u>65</u>]	Brazil	Prospective cohort	33	1B2-2B	Extra-fascial hysterectomy	Laparoscopy
Li, 2014 [<u>66</u>]	China	Retrospective cohort	134	1B1-2A2	RH and PLND	Laparoscopy, laparotomy
Makowski, 2014 [<u>67</u>]	Poland	Retrospective cohort	73	1A2-2A1	RH	Laparotomy
Zhang, 2014 [<u>68</u>]	China	Prospective cohort	126	1B1	RH and PLND	Laparotomy
Chen, 2015[<u>98]</u>	China	Retrospective cohort	137	1B-2A	RH	Laparotomy
Durdevic, 2015 [<u>69</u>]	Serbia	Retrospective cohort	175	1B-2B	RH	Laparotomy
Xie, 2015 [70]	China	Retrospective cohort	86	1B1-2A1	RH	Laparoscopy
Yang, 2015 [71]	China	Retrospective cohort	403	1A1-2B	RH and PLND	Laparoscopy
Yang, 2105 [<u>72</u>]	China	Retrospective cohort	120	1B2-2B	RH and PLND + aortic LND	Laparotomy
Gong, 2016 [73]	China	Retrospective cohort	800	1B2-2B	RH	Laparotomy
Liu, 2016 [<u>74</u>]	China	Prospective cohort	120	1B2-2A2	RH	Laparoscopy
Wu, 2016 [75]	China	Retrospective cohort	839	1B1-2A2	RH	Laparotomy
Yang, 2016 [<u>76</u>]	China	Prospective cohort	76	1B1-2A2	RH	Laparoscopy

* RH = radical hysterectomy, PLND = pelvic lymph node dissection, LND = lymph node dissection

** Not reported as retrospective or prospective

https://doi.org/10.1371/journal.pone.0217775.t001

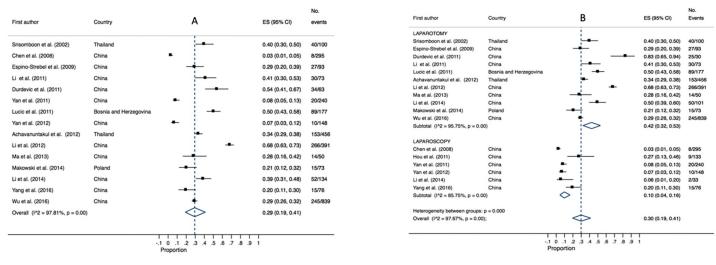


Fig 5. Annotated Forrest plot of overall prevalence of blood transfusion (A) and prevalence by surgical mode (B).

First author	Country	А	ES (95% CI)	No. events	First author	Country	В	ES (95% CI)	No. Events
Srisomboon et al. (2002) Li et al. (2007) Likic-Ladevic et al. (2007) Puntambekar et al. (2007)	Thailand China Serbia India	-	0.01 (0.00, 0.05) 0.03 (0.01, 0.08) 0.01 (0.01, 0.03) 0.01 (0.00, 0.03)	1/100 4/125 8/536 3/252	LAPAROTOMY Srisomboon et al. (2002) Likic-Ladevic et al. (2007) Likic et al. (2008) Cai et al. (2009) Lucic et al. (2011)	Thailand Serbia Serbia China Bosnia and Herzegovina		0.01 (0.00, 0.05) 0.01 (0.01, 0.03) 0.01 (0.01, 0.03) 0.01 (0.00, 0.02) 0.01 (0.00, 0.04)	8/536 8/536 4/480 2/177
Xu et al. (2007) Chen et al. (2008) Likic et al. (2008)	China China Serbia		0.02 (0.01, 0.04) 0.02 (0.01, 0.04) 0.01 (0.01, 0.03)	5/317 5/295 8/536	Makowski et al. (2014) Durdevic et al. (2015) Wu et al. (2016) Subtotal (I^2 = 0.00%, p =	Poland Serbia China 0.92)	•	0.03 (0.00, 0.10) 0.01 (0.00, 0.04) 0.01 (0.01, 0.02) 0.01 (0.01, 0.02)	2/175 10/839
Cai et al. (2009) Lucic et al. (2011) Yan et al. (2011)	China Bosnia and Herzegovina China	-	0.01 (0.00, 0.02) 0.01 (0.00, 0.04) 0.04 (0.02, 0.07)	4/480 2/177 9/240	LAPAROSCOPY Li et al. (2007) Puntambekar et al. (2007) Xu et al. (2007)	China India China	•	0.04 (0.01, 0.11) 0.01 (0.00, 0.03) 0.02 (0.01, 0.04)	3/252
Pareja et al. (2012) Yan et al. (2012) Makowski et al. (2014) Durdevic et al. (2015)	Colombia China Poland Serbia	-	0.06 (0.01, 0.18) 0.01 (0.00, 0.04) 0.03 (0.00, 0.10) 0.01 (0.00, 0.04)	3/47 1/148 2/73 2/175	Chen et al. (2008) Yan et al. (2011) Pareja et al. (2012) Yan et al. (2012)	China China Colombia China		0.02 (0.01, 0.04) 0.02 (0.01, 0.04) 0.04 (0.02, 0.07) 0.06 (0.01, 0.18) 0.01 (0.00, 0.04)	5/295 9/240 3/47
Yang et al. (2015) Wu et al. (2016) Overall (I^2 = 32.20%, p = 0.1	China China		0.01 (0.00, 0.04) 0.00 (0.00, 0.01) 0.01 (0.01, 0.02) 0.01 (0.01, 0.02)	1/403 10/839	Yang et al. (2015) Subtotal (I ^A 2 = 65.71%, p Heterogeneity between gro		•	0.00 (0.00, 0.01) 0.02 (0.01, 0.03)	
	I 1 Proporti		1.3		Overall (I^2 = 36.37%, p =		0 .1 .2 Nortion	0.01 (0.01, 0.02) 1 .3	

Fig 6. Annotated Forrest plot of overall prevalence of bladder injury (A) and prevalence by surgical mode (B).

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Ureteric injury

Fourteen studies (3,063 patients) reported ureteric injury, with a pooled estimate of 1% (95% CI 0.01–0.01, I^2 0.00, P = 0.64). Pooled estimates were equivalent for laparotomy cases 1% (95% CI 0.01–0.01, I^2 0.00, P = 0.53) and laparoscopic cases (1%, 95% CI 0.00–0.01, I^2 0.00, P = 0.78). The included studies had low heterogeneity in both analyses (refer to Fig 7).

Vascular injury

Fourteen studies reported vascular injury (2,758 patients) and had a pooled estimate of 2% (95% CI 0.01–0.03, I² 60.22, P = 0.00). Pooled estimates were lower for laparotomy cases 1% (95% CI 0.01–0.04, I² 57.24, P = 0.03) than for laparoscopic cases 2% (95% CI 0.01–0.04, I² 42.10, P = 0.10). Overall high heterogeneity was identified for both analyses (I² 58.06, P = 0.00) (refer to Fig 8).

First author	Country	A	ES (95% CI)	No. events	First author	Country	В	ES (95% CI)	No. Events
Likic-Ladevic et al. (2007)	Serbia		0.01 (0.01, 0.03)	7/536	LAPAROTOMY Li et al. (2007)	China	-	0.06 (0.01, 0.19)	2/35
Li et al. (2007)	China	÷	0.02 (0.00, 0.06)	2/125	Likic-Ladevic et al. (2007) Likic et al. (2008)	Serbia Serbia	±	0.01 (0.01, 0.03) 0.01 (0.01, 0.03)	
Puntambekar et al. (2007)	India	+	0.00 (0.00, 0.02)	1/252	Cai et al. (2009)	China		0.02 (0.01, 0.03)	
Likic et al. (2008)	Serbia	÷	0.01 (0.01, 0.03)	7/536	Lucic et al. (2011)	Bosnia and Herzegovina	<u>+</u>	0.01 (0.00, 0.04)	
Chen et al. (2008)	China	+	0.00 (0.00, 0.02)	1/295	Zheng et al. (2011) Makowski et al. (2014)	China Poland	<u> </u>	0.01 (0.00, 0.02) 0.01 (0.00, 0.07)	
Cai et al. (2009)	China	H	0.02 (0.01, 0.04)	9/480	Durdevic et al. (2015)	Serbia	÷	0.01 (0.00, 0.03)	
Lucic et al. (2011)	Bosnia and Herzegovina	÷-	0.01 (0.00, 0.04)	2/177	Wu et al. (2016)	China	F	0.02 (0.01, 0.03)	
Zheng et al. (2011)	China	÷	0.01 (0.00, 0.02)	9/960	Subtotal (I^2 = 0.00%, p	= 0.53)		0.01 (0.01, 0.02)	
Yan et al. (2011)	China	+	0.01 (0.00, 0.04)	3/240	LAPAROSCOPY				
Yan et al. (2012)	China	÷	0.01 (0.00, 0.04)	1/148	Puntambekar et al. (2007) Chen et al. (2008)) India China		0.00 (0.00, 0.02) 0.00 (0.00, 0.02)	
Makowski et al. (2014)	Poland	÷	0.01 (0.00, 0.07)	1/73	Yan et al. (2008)	China	-	0.01 (0.00, 0.02)	
Yang et al. (2015)	China	+	0.00 (0.00, 0.02)	2/403	Yan et al. (2012)	China	-	0.01 (0.00, 0.04)	
Durdevic et al. (2015)	Serbia	÷	0.01 (0.00, 0.03)	1/175	Yang et al. (2015) Subtotal (I ² = 0.00%, p	China - 0.78)		0.00 (0.00, 0.02) 0.01 (0.00, 0.01)	
Wu et al. (2016)	China	•	0.02 (0.01, 0.03)	15/839		- 0.70)		0.01 (0.00, 0.01)	
Overall (I ^A 2 = 0.00%, p = 0.64)			0.01 (0.01, 0.01)		Heterogeneity between gr Overall (I ² = 4.72%, p =			0.01 (0.01, 0.01)	
	ا 1. Proportion	0.1	.2			I 1 Proportior		.2	

Fig 7. Annotated Forrest plot of overall prevalence of ureteric injury (A) and prevalence by surgical mode (B).

First author	Country	A	ES (95% CI)	No. events	First author	Country	В	ES (95% CI)	No. Events
					LAPAROTOMY Srisomboon et al. (2002)	Thailand	_	0.01 (0.00, 0.05)	1/100
Srisomboon et al. (2002)	Thailand		0.01 (0.00, 0.05)	1/100	Li et al. (2007)	China	-	0.03 (0.00, 0.15)	
Li et al. (2007)	China	÷	0.04 (0.01, 0.09)	5/125	Cai et al. (2009)	China		0.00 (0.00, 0.01)	
Puntambekar et al. (2007)	India	-	0.04 (0.02, 0.07)	9/252	Lucic et al. (2011)	Bosnia and Herzegovina	i -	0.02 (0.00, 0.05)	
		-			Zheng et al. (2011)	China	+	0.02 (0.01, 0.03)	
Chen et al. (2008)	China	÷	0.02 (0.01, 0.05)	7/295	Durdevic et al. (2015)	Serbia	—	0.03 (0.01, 0.07)	6/175
Cai et al. (2009)	China		0.00 (0.00, 0.01)	1/480	Wu et al. (2016)	China	4	0.01 (0.00, 0.02)	9/839
Zheng et al. (2011)	China	÷	0.02 (0.01, 0.03)	16/960	Subtotal (1 ² = 57.24%,)	0 = 0.03)	¢.	0.01 (0.00, 0.02)	ł
Yan et al. (2011)	China	-	0.03 (0.01, 0.06)	7/240	LAPAROSCOPY				
Lucic et al. (2011)	Bosnia and Herzegovina	÷	0.02 (0.00, 0.05)	3/177	Li et al. (2007)	China		0.04 (0.01, 0.11)	4/90
		1_			Puntambekar et al. (2007		-	0.04 (0.02, 0.07)	
Yan et al. (2012)	China	-	0.04 (0.02, 0.09)	6/148	Chen et al. (2008)	China		0.02 (0.01, 0.05)	
Pareja et al. (2012)	Colombia	+	0.02 (0.00, 0.11)	17168	Yan et al. (2011)	China		0.03 (0.01, 0.06)	
Durdevic et al. (2015)	Serbia	.	0.03 (0.01, 0.07)	6/175	Pareja et al. (2012)	Colombia	-	0.02 (0.00, 0.11)	
		1			Yan et al. (2012)	China	-	0.04 (0.02, 0.09)	
Yang et al. (2015)	China		0.01 (0.00, 0.02)	3/403	Xie et al. (2015)	China	5	0.01 (0.00, 0.06)	
Xie et al. (2015)	China	+	0.01 (0.00, 0.06)	1/86	Yang et al. (2015)	China		0.01 (0.00, 0.02)	
Wu et al. (2016)	China		0.01 (0.00, 0.02)	9/839	Subtotal (I ² = 42.10%, I	5 = 0.10)	ř	0.02 (0.01, 0.04)	
Overall (I*2 = 60.22%, p = 0.00)		•	0.02 (0.01, 0.03)		Heterogeneity between g Overall (I^2 = 58.06%, p		0	0.02 (0.01, 0.02)	
	l 1 Proportic		.3			l 1 Proportia	0 .1 .2	.3	

Fig 8. Annotated Forrest plot of overall prevalence of vascular injury (A) and prevalence by surgical mode (B).

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Fistula

Fourteen studies (4,429 patients) reported fistula as a surgical complication. Fistula was defined by the study authors and not necessarily definitive as to type (e.g. vesicovaginal or rectovaginal). The overall pooled estimated was 2% (95% CI 0.01–0.03, $I^2 = 77.32$, P = 0.00,). Pooled estimates were comparable between laparotomy cases (2%, 95% CI 0.01–0.04, P = 0.00, $I^2 = 86.19$) and laparoscopic cases (1%, 95% CI 0.01–0.01, $I^2 = 22.70$, P = 0.26). Overall high heterogeneity was identified for laparotomy and moderate heterogeneity for laparoscopy (refer to Fig 9).

First author	Country	A	ES (95% CI)	No. events	First author	Country	в	ES (95% CI)	No. Events
One at al. (2000)	China	1	0.00 (0.00, 0.00)	0/100	LAPAROTOMY		1		
Cao et al. (2006)		T	0.02 (0.00, 0.06)	3/139	Cao et al. (2006) Likic-Ladevic et al. (2007	China) Serbia	T	0.02 (0.00, 0.06) 0.05 (0.03, 0.07)	3/139 27/536
Li et al. (2007)	China		0.02 (0.00, 0.06)	2/125	Likic et al. (2008)	Serbia		0.05 (0.03, 0.07)	27/536
Xu et al. (2007)	China	+	0.02 (0.01, 0.04)	5/317	Cai et al. (2009)	China		0.01 (0.00, 0.02)	5/480
Likic-Ladevic et al. (2007)	Serbia		0.05 (0.03, 0.07)	27/536	Chai et al. (2014)	China		0.01 (0.00, 0.02)	4/438
Likic et al. (2008)	Serbia		0.05 (0.03, 0.07)	27/536	Makowski et al. (2014)	Poland	+	0.03 (0.00, 0.10)	2/73
Cai et al. (2009)	China		0.01 (0.00, 0.02)	5/480	Durdevic et al. (2015)	Serbia	-	0.01 (0.00, 0.04)	2/175
		1			Wu et al. (2016) Subtotal (I^2 = 86.19%,	China	~	0.01 (0.00, 0.01) 0.02 (0.01, 0.04)	5/839
Yan et al. (2011)	China	- F	0.03 (0.01, 0.05)	6/240	Subiotal (1.2 = 86.19%,	p = 0.00)	Y	0.02 (0.01, 0.04)	
Pareja et al. (2012)	Colombia	-	0.04 (0.01, 0.15)	2/47	LAPAROSCOPY				
Makowski et al. (2014)	Poland	÷	0.03 (0.00, 0.10)	2/73	Li et al. (2007)	China	÷	0.02 (0.00, 0.08)	2/90
Favero et al. (2014)	Brazil	+	0.03 (0.00, 0.16)	1/33	Xu et al. (2007)	China	+	0.02 (0.01, 0.04)	5/317
Chai et al. (2014)	China		0.01 (0.00, 0.02)	4/438	Yan et al. (2011)	China	†	0.03 (0.01, 0.05)	6/240
Yang et al. (2015)	China		0.01 (0.00, 0.02)	3/403	Pareja et al. (2012) Favero et al. (2014)	Colombia Brazil	-	0.04 (0.01, 0.15) 0.03 (0.00, 0.16)	2/47 1/33
Durdevic et al. (2015)	Serbia	1	0.01 (0.00, 0.04)	2/175	Yang et al. (2015)	China		0.01 (0.00, 0.02)	3/403
		3			Subtotal (1^2 = 22.70%,	p = 0.26)	Q	0.01 (0.01, 0.02)	
Wu et al. (2016)	China	1	0.01 (0.00, 0.01)	5/839			- 1		
Overall (I ² = 77.32%, p =)	0.00)	Ŷ	0.02 (0.01, 0.03)		Heterogeneity between g				
					Overall (I ² = 77.40%, p	= 0.00);	Ŷ	0.02 (0.01, 0.03)	
		1 0 .1 .2 . Proportion	3.4			1	0 .1 .2 Proportion	.3 .4	

Fig 9. Annotated Forrest plot of overall prevalence of fistula injury (A) and prevalence by surgical mode (B).

Conversion to open surgery

There were three studies (1015 patients) that reported the requirement to convert to open surgery. The overall pooled estimate was 1% (95% CI 0.00–0.02, $I^2 = 21.29$, P = 0.28) and the included studies showed moderate heterogeneity.

Thromboembolic events

Pulmonary embolism was reported in five studies (that included 1,270 patients) as a postoperative complication. The pooled estimate was 0.4% (95% CI 0.00–0.01, I² 26.69, P = 0.25). Seven studies reported on all thromboembolic events (that included 1,501 patients) and had a pooled estimate of 1% (95% CI 0.01–0.02, I² 0.00, P = 0.73). Pooled estimates were comparable for RH (laparotomy) cases (1%, 95% CI 0.00–0.02) vs. RH (laparoscopic) cases (1%, 95% CI 0.00– 0.05). Overall the included studies showed low heterogeneity for both analyses (I² 0.00, P = 0.73).

Infectious morbidity

There were twenty studies (3,826 patients) that reported infectious morbidity as a postoperative complication. The overall pooled estimated was 8% (95% CI 0.04–0.12, I² 95.72, P = 0.00). Pooled estimates were higher for RH (laparotomy) cases 9% (95% CI 0.05–0.14, I² = 95.82, P = 0.00) vs. RH (laparoscopic) cases 5% (95% CI 0.00–0.12, I² 88.88, P = 0.00). High heterogeneity was identified for both analyses (refer to Fig 10).

Assessment of complications excluding studies from China

For the complications of blood transfusion, bladder injury, ureteric injury, vascular injury and fistula, there was less than 0.08 difference in the random pooled estimates when studies from China were excluded. (S2 Table).

Progression free survival (PFS) and overall survival (OS)

The 5-year PFS was reported for 4,731 patients from 17 studies. Overall PFS was 81.15%, 83% for laparotomy (n = 3,917; 95% CI:0.75–0.89) and 84% for laparoscopy (n = 664; 95% CI:0.81–0.87). The 5-year OS was reported for 5130 patients from 16 studies. OS was 83.5%, 85% for laparotomy cases (n = 4,466; 95% CI: 0.78–0.91) and 80% for laparoscopy (n = 664; 95% CI:0.72–0.88).

Assessment of publication bias

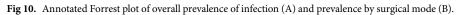
There was no apparent publication bias for all complications included in the meta-analysis. Funnel plots and p values of Egger's test for each complication are shown in <u>S2 File</u>

Discussion

To our knowledge, this is the first systematic review and meta-analysis of morbidity after primary surgical management of cervical cancer specifically in LMIC, although only data from middle-income countries (MIC) were included in the meta-analysis. The incidence of complications in this meta-analysis is largely comparable to those reported in high-income countries (HIC). Blood transfusions are reported in HIC at up to 25% for minimally invasive surgical approaches for cervical cancer and up to 75% for laparotomy[80], and so our findings of 10% and 42% respectively are consistent with published studies that have been conducted in well resourced settings. This may however also reflect differing resources and thresholds for

Espino-Strebel et al. (2009) Chin	sa aa aa aa aa aa aa aa aa aa aa aa aa a	╌╎┽╴┤╎╌╷╴╷╴			0.11 (0.06, 0.19) 0.02 (0.00, 0.07) 0.04 (0.02, 0.09) 0.07 (0.00, 0.02) 0.08 (0.02, 0.09) 0.08 (0.02, 0.09) 0.02 (0.01, 0.04) 0.02 (0.01, 0.07) 0.00 (0.00, 0.01) 0.00 (0.00, 0.01)	2/10 6/13 4/53 10/3 4/52 9/29 6/35 4/14
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ku et al. (2008) Chin kike et al. (2008) Sett sai et al. (2009) Chin fanchara et al. (2009) Thai spino-Strebei et al. (2009) Chin spino-Strebei et al. (2009) Chin spino-Strebei et al. (2009) Chin ai et al. (2009) Chin ai et al. (2009) Chin	na bia Na Jiand ppines	-+-			0.03 (0.01, 0.07) 0.00 (0.00, 0.01)	4/14
kic et al. (2008) Sert ai et al. (2009) Chin tarchana et al. (2009) Thai spino-Strebel et al. (2009) Chin spino-Strebel et al. (2010) Phil u et al. (2010) Chin ai et al. (2009) Chin	ba na iland ha ippines				0.00 (0.00, 0.01)	
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spino-Strebel et al. (2009) Chin spino-Strebel et al. (2010) Phili hu et al. (2010) Chin al et al. (2009) Chin	na ippines	- +_			0.01 (0.00, 0.03)	
spino-Strebel et al. (2010) Phili hu et al. (2010) Chin cai et al. (2009) Chin	ippines	· · · ·			0.09 (0.04, 0.16)	
hu et al. (2010) Chin ai et al. (2009) Chin					0.11 (0.05, 0.21)	
ai et al. (2009) Chin			_		0.04 (0.01, 0.09)	
					0.01 (0.00, 0.03)	
ezerra et all(2011) Braz						
					0.03 (0.01, 0.10)	
urdevic et al. (2011) Chin	-		-		0.06 (0.02, 0.15)	
nu et al. (2011) Chin			_		0.13 (0.06, 0.24)	
heng et al. (2011) Chin					0.01 (0.00, 0.01)	
et al. (2011) Chin	-		-		0.05 (0.02, 0.13)	
hou et al. (2011) Chin	-		-		0.11 (0.05, 0.20)	
an et al. (2011) Chin			_		0.06 (0.03, 0.10)	
	nia and Herzegovina				0.08 (0.04, 0.13)	
et al. (2012) Chin			_		0.02 (0.01, 0.04)	
an et al. (2012) Chin	-		_		0.09 (0.05, 0.15)	
	iland		_		0.00 (0.00, 0.02)	
ahawerawat et al. (2013) Thai					0.12 (0.05, 0.23)	
la et al. (2013) Chin			-		0.14 (0.06, 0.27)	
hang et al. (2014) Chin	-	-	-		0.03 (0.01, 0.08)	
akowski et al. (2014) Pola	and				0.11 (0.05, 0.20)	8/73
avero et al. (2014) Braz	zil .			-	0.18 (0.07, 0.35)	6/3
et al. (2014) Chin	na				0.04 (0.02, 0.09)	6/13
urdevic et al. (2015) Sert	bia		-		0.08 (0.04, 0.13)	14/
e et al. (2015) Chin	na l	1			0.19 (0.11, 0.28)	164
ang et al. (2015) Chin	na l	-			0.02 (0.00, 0.06)	2/13
hen. (2015) Chin	na l		_		0.04 (0.01, 0.08)	5/13
ing et al. (2015) Chin	na l				0.03 (0.02, 0.06)	14
ang et al. (2016) Chin	na				0.07 (0.02, 0.15)	5/7
u et al. (2016) Chin	na l				0.07 (0.03, 0.13)	8/1
lu et al. (2016) Chin	na l				0.02 (0.01, 0.03)	14/
verall (I*2 = 85.73%, p = 0.00)		\$			0.04 (0.03, 0.06)	1

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transfusions in LMIC. In HIC 1.2% and 2.8% of women develop fistula (ureterovaginal and vesicovaginal) and pulmonary embolism respectively[81] following surgery for cervical cancer. Ureteric injuries following laparotomy and radical hysterectomy for cervical cancer were reported to occur in 2.48% of cases in a review of nearly 400,000 hysterectomies in the United Kingdom (UK)[82]. The reported incidence of complications is similar or lower in this meta-analysis and may represent publication bias. However, our findings suggest that, morbidity following surgery for cervical cancer in MIC appears to compare favourably to the reported incidence in HIC.

The comparative survival between laparoscopy and laparotomy cases in this review is clearly different to that of a recently published retrospective cohort study [83] and prospective RCT in HIC [84]. However, our data is neither registry based nor largely prospective so comparisons should be made with caution. It is reassuring that the risks of complications reported in this meta-analysis are comparable to HICs. Accepting that some settings within LMIC have individual survival rates more comparable to HIC[85], overall long-term outcomes do not compare favourably for cervical cancer patients in LMIC where overall 5-year survival rate is 50%, which is markedly different to HIC, which report survival rates up to 80%. Reasons for this difference may include the stages of tumour (e.g. 2B), being offered operative care which may not happen in HIC, high rates of HIV positivity in these settings in women who are younger at diagnosis than non-HIV positive women, although in many settings their outcomes appear comparable[86]. Equally, surgeons are frequently the mainstay of treatment for women in LMIC and the lack of adjuvant treatment likely impacts on outcomes[3, 6].

The strengths of the current study include the rigorous inclusion criteria and a search that was not limited by language. Moreover, the broad search criteria were not limited to LMIC (which are not necessarily linked in MeSH terms or titles and abstracts) and so this increased the likelihood of capturing studies reporting on outcomes in LMIC. A large number of both articles and patients were included in the meta-analysis with a surgical treatment period since 2000, which is likely reflective of current practice. Each complication has a large denominator, although we accept that the definition of morbidity (in both high income and LMIC studies) is not necessarily standardised, and subject to variations in regional practices and definitions [87]. However, we have only included high quality studies in the meta-analysis, adding to its robustness.

There are several limitations to this study. As the main body of data are derived from retrospective studies with high heterogeneity, a random-effects model was applied to decrease the possibility of overestimating surgical morbidity. Nevertheless, the retrospective and unblinded study designs of the majority of included studies may introduce biases that cannot be evaluated. Furthermore, there is variability in surgical approaches, including lymphadenectomy, that potentially affect the meta-analysis; however this is reflective of current surgical practice, and the issue is similar to that encountered in other meta-analyses of surgical complications [88].

Operative morbidity could also be influenced by neoadjuvant chemotherapy. A recent randomized trial from India reported inferior disease-free survival (DFS) with neoadjuvant-chemotherapy followed by surgery compared to concurrent chemoradiation. This study has also reported operative morbidity. Patients randomised to neoadjuvant chemotherapy had a significantly higher rate of grade 3 or 4 thrombocytopenia compared to those treated by concurrent chemoradiation, but there was no significant difference between the two groups in grade 3 or 4 gastrointestinal and bladder toxicities. In the neoadjuvant chemotherapy plus surgery group, perioperative hemorrhage with 1000 mL blood loss occurred in 7.9% of patients[89]. In view of the superior DFS with chemoradiation it is conceivable that the neoadjuvant modality might decline in the future. Due to the heterogeneity of the regimens of the included studies, it was not possible to analyse the impact of neoadjuvant and adjuvant therapies on surgical morbidity in this meta-analysis.

Only one high quality study^[65] reported on simple hysterectomy rather than radical hysterectomy. However, this study included only 33 patients and so is unlikely to impact on the outcome of the meta-analyses. Equally patient co-morbidities that confound complication results are not reported. No high-quality studies from LICs were included and there is little representation from Africa in this analysis (2 studies from this region that met inclusion criteria were excluded at the quality assessment stage) and furthermore, data from China represents 65% of the included studies. This may account for the reported incidences of surgical complications in the current meta-analysis, which are comparable to those in high-income countries, as many parts of China are increasingly becoming higher income settings[90]. However, the similar pooled estimates for reported complications in our sensitivity analysis that excluded studies from China do not support this. It is also plausible that the surgical case load and experience in caring for cervical cancer cases in LMIC mitigates the limitations in these settings and results in comparable surgical complications with HIC. An important limitation is that included data were not population or registry based. This may explain the relatively high progression free and overall survival reported in our study, which is markedly different to that reported by Allemani and colleagues, with 5 year OS in LMIC of around 50% [91], based on population registry data that included patients with all stages of cervical cancer. Finally, it was not possible to stratify surgical complications by stage of disease or timing of the complication (e.g. immediately postoperative or delayed complication) as this was infrequently reported in the included studies.

Given these limitations, there is a clear lack of globally comprehensive and representative data on surgical outcomes in LMIC[8, 92], which, by virtue of inadequate assessment of deficiencies where resources can be targeted, and by limiting knowledge on the natural history and variations in cancer outcomes in these settings, worsens cancer care[3]. Data are particularly lacking from LICs. Developing registries to meet this gap is critical in addressing the global burden of non-communicable diseases including cancer[93], particularly with increasing access to adjuvant therapies in these settings. Inclusive access to the most basic cervical cancer management is lacking in LMIC and system strengthening is critical to changing this [94]. The Lancet Commission on Global Surgery clearly outlines that scale-up of universal access to safe surgical care includes the need for strong system development grounded in data collection, including surgical indicators, such that system issues and deficiencies can be identified and changes advocated for. If we are to meet these goals then, similar to global efforts to count every newborn[95], the oncology community urgently needs to work with LMIC policy makers to ensure longitudinal data collection and the counting of every cancer case[91, 92].

Conclusion

This is the first systematic review and meta-analysis of surgical morbidity in cervical cancer in LMIC and provides a benchmark for future health services research and policy implementation. Surgery is the cornerstone of management in early stage cervical cancer and is likely to remain a major treatment modality in LMIC for many years; we must as a global oncology community mobilise to ensure the foundations for health advocacy and program development are in place, including surgical morbidity registries.

Supporting information

S1 Table. Complete search strategy. (DOCX)

S2 Table. Sensitivity analysis for surgical complications reported in the meta-analysis. (DOCX)

S1 File. All reported complications in included studies. (DOCX)

S2 File. Funnel plots reporting surgical complications for all studies included in the metaanalysis.

(DOCX)

S3 File. PRISMA checklist. (PDF)

S4 File. Surgical complications in high quality studies. (XLSX)

S5 File. Surgical complications (laparoscopy and laparotomy) in high quality studies. (XLSX)

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