


SEVEN-DAY PROFILE PUBLICATION



# Acute hypoxemic respiratory failure in immunocompromised patients: the Efraim multinational prospective cohort study

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## Abstract

**Background:** In immunocompromised patients with acute hypoxemic respiratory failure (ARF), initial management aims primarily to avoid invasive mechanical ventilation (IMV).

**Methods:** To assess the impact of initial management on IMV and mortality rates, we performed a multinational observational prospective cohort study in 16 countries (68 centers).

**Results:** A total of 1611 patients were enrolled (hematological malignancies 51.9%, solid tumors 35.2%, systemic diseases 17.3%, and solid organ transplantation 8.8%). The main ARF etiologies were bacterial (29.5%), viral (15.4%), and fungal infections (14.7%), or undetermined (13.2%). On admission, 915 (56.8%) patients were not intubated. They received standard oxygen ( $N = 496$ , 53.9%), high-flow oxygen (HFNC,  $N = 187$ , 20.3%), noninvasive ventilation (NIV,  $N = 153$ , 17.2%), and NIV + HFNC ( $N = 79$ , 8.6%). Factors associated with IMV included age (hazard ratio = 0.92/year, 95% CI 0.86–0.99), day-1 SOFA (1.09/point, 1.06–1.13), day-1 PaO<sub>2</sub>/FiO<sub>2</sub> (1.47, 1.05–2.07), ARF etiology (*Pneumocystis jirovecii* pneumonia (2.11, 1.42–3.14), invasive pulmonary aspergillosis (1.85, 1.21–2.85), and undetermined cause (1.46, 1.09–1.98). After propensity score matching, HFNC, but not NIV, had an effect on IMV rate (HR = 0.77, 95% CI 0.59–1.00,  $p = 0.05$ ). ICU, hospital, and day-90 mortality rates were 32.4, 44.1, and 56.4%, respectively. Factors independently

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associated with hospital mortality included age (odds ratio = 1.18/year, 1.09–1.27), direct admission to the ICU (0.69, 0.54–0.87), day-1 SOFA excluding respiratory score (1.12/point, 1.08–1.16),  $\text{PaO}_2/\text{FiO}_2 < 100$  (1.60, 1.03–2.48), and undetermined ARF etiology (1.43, 1.04–1.97). Initial oxygenation strategy did not affect mortality; however, IMV was associated with mortality, the odds ratio depending on IMV conditions: NIV + HFNC failure (2.31, 1.09–4.91), first-line IMV (2.55, 1.94–3.29), NIV failure (3.65, 2.05–6.53), standard oxygen failure (4.16, 2.91–5.93), and HFNC failure (5.54, 3.27–9.38).

**Conclusion:** HFNC has an effect on intubation but not on mortality rates. Failure to identify ARF etiology is associated with higher rates of both intubation and mortality. This suggests that in addition to selecting the appropriate oxygenation device, clinicians should strive to identify the etiology of ARF.

**Keywords:** Noninvasive ventilation, High flow oxygen, Hematological malignancies, Transplantation, Systemic diseases, Pneumocystis

## Introduction

The number of living immunosuppressed adults is increasing [1]. This is the result of new indications and development of immunosuppressive treatments in auto-immune conditions and solid organ transplantation [2, 3], of greater life expectancy in patients with malignancies [4, 5] and the high number of averted deaths from cancer over the last decade [6]. Immunosuppression, either related to the underlying condition or to its treatments, increases the risk and severity of infections, primary bacterial, viral, or opportunistic infections, as well as reactivation of latent infections [7, 8]. Moreover, non-infectious causes of critical illness may also be triggered by specific infiltration from the underlying disease or by drug-related organ toxicity.

Acute hypoxemic respiratory failure (ARF) is the leading cause of ICU admission in immunocompromised patients [4, 9, 10]. It is also one of the most challenging conditions as initial management must optimize oxygenation, identify the ARF etiology, and support associated organ dysfunction at the same time [4]. Over the last decade, small and single-center studies, or post hoc analyses in limited numbers of patients provided conflicting information on what can be expected from high-flow oxygen administered through a nasal cannula (HFNC) in immunocompromised patients with ARF [11–13]. Moreover, a randomized clinical trial suggested no benefit from noninvasive ventilation (NIV) in immunocompromised patients with ARF [14]. Furthermore, NIV use alone or in association with HFNC was associated with a dramatic increase in both the need for invasive mechanical ventilation (IMV) and in hospital mortality, raising concerns about the selection of initial oxygenation strategy in these patients [15].

Mortality remains remarkably high, reaching 60% if ICU admission is delayed [16], in case of NIV failure, invasive fungal infection, or when the ARF etiology remains undetermined [17–22]. Thus, clinicians are left with a difficult dilemma of either preventing intubation by all means, or giving priority to identifying ARF etiology,

possibly including intubation for diagnostic procedures such as fiberoptic bronchoscopy and bronchoalveolar lavage (FO-BAL). Although bronchoscopy and noninvasive diagnostic tests perform equally [23], FO-BAL may be needed in several cases [18].

To guide clinical practice and inform the clinician on whether initial management should give priority to identifying ARF etiology or to avoiding IMV, we performed a large prospective multinational cohort study. Immunocompromised patients with hypoxemic ARF were recruited from 68 ICUs in 16 countries. We hypothesized that oxygenation strategy influences outcomes as much as recognizing the etiology of ARF. Need for IMV and hospital mortality were our primary endpoints.

## Patients and methods

### Main objective

The primary objective of this study was to assess whether initial management (i.e., oxygenation and diagnostic strategies) affects the proportion of patients requiring intubation and mechanical ventilation or hospital mortality.

### Study design and participants

Efrain was a multinational, observational prospective cohort study performed by the Nine-I (Caring for critically ill immunocompromised patients) study group. This group includes critical care physicians from 16 countries who have extensive experience in the management of various groups of critically ill immunocompromised patients. Physician participation was voluntary, without financial incentive. Participating providers obtained institutional review board (IRB) approval from their institutions in accordance with local ethics regulations.

Inclusion criteria were 18 years or older; acute hypoxemic respiratory failure ( $\text{PaO}_2 < 60$  mmHg or  $\text{SpO}_2 < 90\%$  on room air, or tachypnea  $> 30/\text{min}$ , or labored breathing or respiratory distress or dyspnea at rest or cyanosis); need for more than 6 L/min oxygen; respiratory symptom duration less than 72 h;

and non-AIDS-related immune deficiency defined as hematologic malignancy or solid tumor (active or in remission for less than 5 years, including recipients of autologous or allogeneic stem cell transplantation), solid organ transplant, long-term (> 30 days) or high-dose (> 1 mg/kg/day) steroids, or any immunosuppressive drug for more than 30 days. Exclusion criteria included postoperative acute respiratory failure (within 6 days of surgery), admission after a cardiac arrest, ICU admission only to secure bronchoscopy, or refusal of the patient or family to participate in the study. Primary endpoints were the need for IMV in patients not intubated on ICU admission and all-cause hospital mortality.

After IRB approval, participating ICUs enrolled patients between 5 November 2015 and 1 July 2016. A standardized paper case report was prepared by investigators and tested in 35 patients. After feedback and corrections, the case report form was sent to participating ICUs and, once completed, it was sent back to the coordinating center in Paris for data entry by specialized technicians used to handling data on critically ill immunocompromised patients. The study was funded by the Groupe de Recherche en Réanimation Respiratoire Onco-Hématologique (GRRR-OH), an academic not-for-profit French organization.

### Study treatments

All management decisions were independently made by the attending physicians according to standard practice in each ICU. Diagnostic tests to identify the cause of respiratory failure were invasive (FO-BAL), noninvasive (blood and sputa cultures for bacteria and fungi, serum and urine antigens, PCR in blood, serum and nasopharyngeal aspirates, high resolution CT scan, echocardiography, serology, and specific tests according to each situation), or both [14, 18–20, 23]. All diagnoses were reviewed by two study investigators for coherence and for alignment with established definitions. Oxygenation modalities and the use of NIV or HFNC were at the discretion of the primary team. Management of associated organ dysfunction and handling of immunosuppressive drugs or chemotherapy were done as per local preferences. The decision to intubate was not controlled by the study.

Patient's code status on ICU admission was recorded as full code management, time-limited trial, do not intubate (DNI), do not resuscitate (DNR), or palliative.

### Study outcomes

The two primary study outcomes were the need for IMV and all-cause mortality before hospital discharge. Patients were enrolled immediately at ICU admission and the data in the tables and figures were collected prospectively

using the paper case report form. The sequential organ failure assessment (SOFA) score was recorded within 24 h of admission [24]. We did not use the term acute respiratory distress syndrome (ARDS) as defined on the basis of the 2012 definition because most patients were not ventilated at day 1 and were not receiving positive end-expiratory pressure (PEEP) [25].

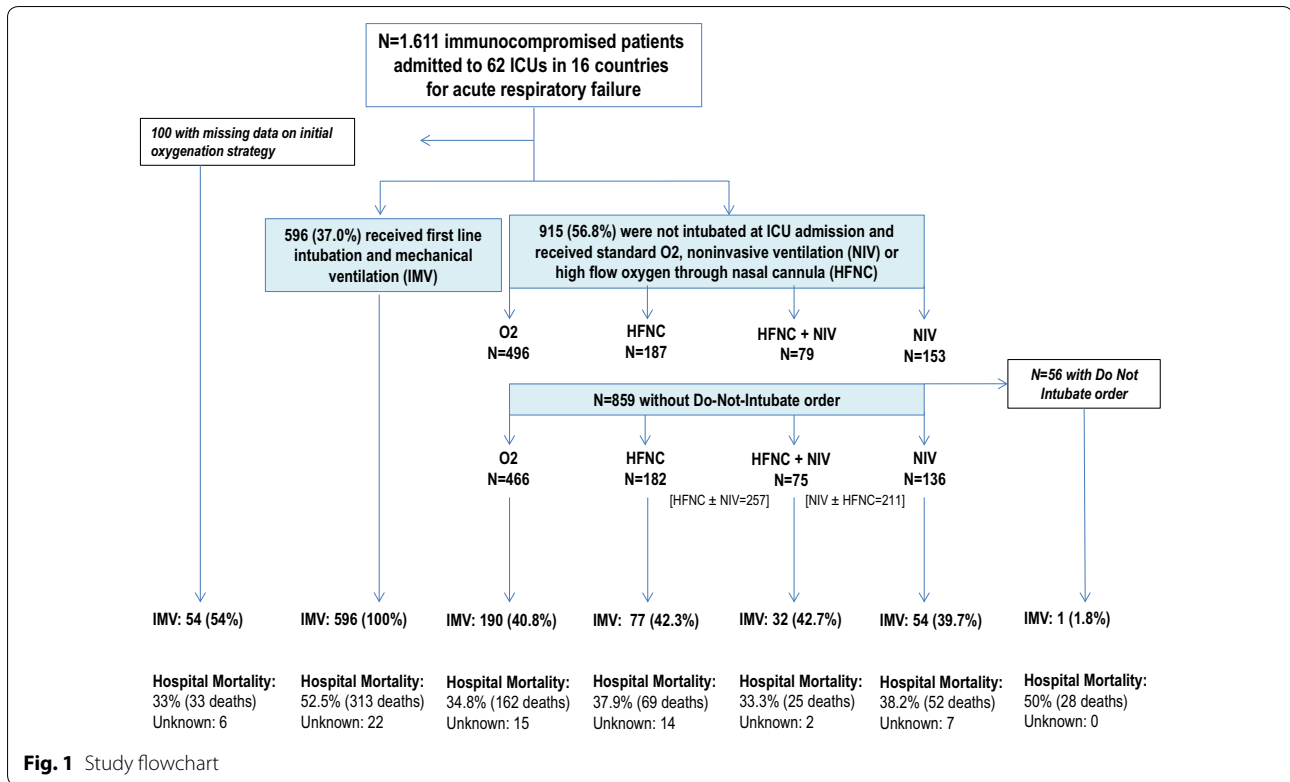
### Statistical analysis

We based our computation on the potential effect size demonstrated using logistic regression modelling, with fixed sample size, and prevalence of hospital mortality ranging from 0.15 to 0.5. It showed that a sample size of 1500 will reach an 80% power to detect an odds ratio (OR) of 1.3 with balanced groups, and OR of at least 1.5 with imbalanced groups of 1:4.

Continuous variables were described as medians [interquartile ranges (IQRs)] and categorical variables as proportions. Data management allowed checking for data inconsistencies that were solved by consensus.

Comparisons of proportions between groups were made using the  $\chi^2$  test. Comparisons of continuous variables between groups were made using the Wilcoxon rank-sum test.

To analyze the predictive factors of intubation, the analysis has to deal with death prior to intubation that acts as a competing event for intubation. Thus, cumulative incidence curves were plotted, and compared across baseline groups using the Gray test. Predictive factors were assessed through the use of cause-specific Cox models; univariate models were fitted, and then those associated with the outcome at the 5% level were introduced into a multivariable model, with variable selection based on the Akaike criterion. The analysis of hospital mortality was based on logistic models, with model selection similar to that described above for intubation outcome. In this model, SOFA score was introduced without items of the respiratory score to assess the impact of PaO<sub>2</sub>/FiO<sub>2</sub> ratio in immunocompromised patients [20]. To assess the benefit of the use on ICU admission of NIV or HFNC in non-mechanically ventilated patients, we used a propensity score approach to control for observed confounding factors that might influence both group assignment and outcome. Each model was performed separately as follows. Briefly, we estimated the propensity score of being administered NIV (HFNC, resp.) at the time of ICU admission using a logistic regression model including potential predictors as selected by prognostic analyses described above. To handle missing values in confounders, multiple imputation with chained equation was used, where propensity score for each patient was averaged across 30 completed datasets while propensity score matching used these averaged scores to estimate



the treatment effect. A 1:1 matching algorithm without replacement within a caliper of 0.15 standard deviation of the logit of the propensity score was used; balances in confounders before and after matching were checked using standardized mean differences. Estimates of the effect of NIV on outcomes, accounting for paired data, were based on either a cause-specific frailty Cox model or logistic regression model by generalized estimating equations, respectively.

Primary analyses were performed on the complete cases, assuming missing completely at random covariates. Then, sensitivity analyses for such assumptions were performed, based on multiple imputation with chained equation. To incorporate potential center effects, models with centers introduced as clusters or random variables were also fitted. Last, we performed exploratory subset analyses, restricting ourselves to patients with hematological malignancies or full code status.

All reported *P* values are 2-sided; *P* < 0.05 was considered statistically significant. All analyses were performed using R version 3.3.2 (<http://www.R-project.org/>).

## Results

Over the 8-month study period, 1611 patients [60% male, median age 63 years old (IQR 54–71)] were enrolled in the 68 participating ICUs. Supplemental Fig. 1 shows the number of patients recruited per country.

Immunosuppression was related to hematological malignancy [837 (51.9%), including 103 allogeneic and 152 autologous hematopoietic stem cell transplantations], solid tumor [567 (35.2%)], systemic rheumatic disease [278 (17.3%)], and solid organ transplant [142 (8.8%)]. Fourteen (0.9%) patients received immunosuppressive drugs for rare diseases. More than one condition was reported in 199 patients. Half the patients were newly diagnosed, 25% were in second-line therapy, 15% were in remission, and 15% had received elements of palliative care. The resuscitation status was full code in 1292 (80.2%), time-limited trial in 69 (4.3%), DNI in 58 (3.6%), DNR in 56 (3.5%), and the code status was not established in 136 (8.4%) cases.

Most patients were admitted from the wards (57%); however, 348 (21.6%) were directly admitted from the emergency department or by ambulances. On admission, respiratory rate was 30/min (25–36), SpO<sub>2</sub> was 92% through 8 L/min (4–15) O<sub>2</sub> or a FiO<sub>2</sub> of 50% (40–80). PaO<sub>2</sub>/FiO<sub>2</sub> ratio was 151 (100–233) and SOFA score was 7 (4–10) on day 1. IMV was needed in 1004 (62.3%) patients throughout the ICU stay. ICU, hospital, and day-90 mortality rates were 32.4% (522 deaths), 44.1% (682 deaths, 66 missing information), and 56.4% (771 deaths, 245 missing), respectively.

Figure 1 depicts oxygenation and ventilation modalities following ICU admission. First-line IMV was offered

**Table 1 Characteristics of the 915 patients who were not intubated on ICU admission**

Numbers (%) or median (IQR)	Standard oxygen (O <sub>2</sub> , N = 496)	High-flow oxygen (HFNC, N = 187)	Noninvasive ventilation (NIV, N = 153)	HFNC + NIV (N = 79)	P value
Age (years)	63.7 (54.4–70.8)	62.1 (52.7–69.8)	65.1 (56.6–71.9)	64.1 (52–71)	0.40
Female gender	184 (37.1%)	63 (33.7%)	63 (41.2%)	41 (51.9%)	0.03
Chronic pulmonary disease	62 (12.5%)	23 (12.3%)	42 (27.4%)	11 (13.9%)	0.0002
Chronic cardiac failure	107 (21.6%)	31 (16.6%)	41 (26.8%)	23 (29.1%)	0.09
Days since symptoms onset	1 (0–4)	2 (0–8)	1 (0–3)	2 (1–5)	0.0001
<b>Location before ICU admission</b>					<b>0.05</b>
Emergency department	112 (22.6%)	36 (19.3%)	48 (31.4%)	9 (11.4%)	
Hospital wards	307 (61.9%)	116 (62%)	89 (58.2%)	55 (69.6%)	
Other ICUs	22 (4.4%)	2 (1.1%)	3 (1.9%)	1 (1.3%)	
Other locations	55 (11.1%)	33 (17.6%)	13 (8.5%)	14 (17.7%)	
Nightly or weekend admissions	304 (61.3%)	118 (63.1%)	82 (53.6%)	57 (72.1%)	0.06
<b>Underlying condition</b>					
Hematological malignancy	274 (55.2%)	110 (58.8%)	73 (47.7%)	46 (58.2%)	0.11
Solid tumor	158 (31.8%)	54 (28.9%)	49 (32%)	23 (29.1%)	0.84
Systemic diseases	84 (16.9%)	36 (19.3%)	32 (20.9%)	14 (17.7%)	0.84
Solid organ transplant	42 (8.5%)	23 (12.3%)	20 (13.1%)	6 (7.6%)	0.21
Long-term/high-dose steroids	384 (77.4%)	148 (79.1%)	114 (74.5%)	47 (59.4%)	0.13
Neutropenia	78 (15.7%)	34 (18.2%)	24 (15.7%)	20 (25.3%)	0.18
<b>Hematopoietic stem cell transplantation</b>					
Autologous	36 (7.3%)	13 (6.9%)	7 (4.6%)	6 (7.6%)	0.13
Allogeneic	36 (7.3%)	26 (13.9%)	15 (9.8%)	12 (15.2%)	
Days since diagnosis	142 (15–672)	204 (26–725)	118 (24–671)	81 (1–424)	0.20
Newly diagnosed disease	150 (30.2%)	51 (27.3%)	58 (37.9%)	20 (25.3%)	0.28
SOFA score on ICU admission	5 (3–8)	6 (4–8)	5 (3–8)	5 (4–8)	0.04
PaO <sub>2</sub> /FiO <sub>2</sub> on day 1	173 (136–220)	150 (100–173)	168.5 (112–188)	128 (100–173)	< 0.0001
<b>Acute respiratory failure (ARF) etiology</b>					
Bacterial infection	131 (26.4%)	27 (14.4%)	63 (41.2%)	21 (26.7%)	0.001
Viral infection	70 (14.1%)	38 (20.3%)	30 (19.6%)	14 (17.7%)	0.18
Pneumocystis pneumonia	21 (4.2%)	18 (9.6%)	6 (3.9%)	10 (12.6%)	0.005
Invasive pulmonary aspergillosis	20 (4.0%)	7 (3.7%)	7 (4.6%)	3 (3.8%)	1
Invasive fungal infection	65 (13.1%)	39 (20.8%)	21 (13.7%)	13 (16.4%)	0.05
Disease-related infiltrates	52 (10.5%)	20 (10.7%)	12 (7.8%)	9 (11.4%)	0.71
Cardiogenic pulmonary edema	44 (8.9%)	8 (4.3%)	16 (10.4%)	8 (10.1%)	0.20
Undetermined	63 (12.7%)	28 (15%)	14 (9.1%)	12 (15.2%)	0.53
More than one ARF etiology	71 (14.3%)	34 (18.2%)	30 (19.6%)	13 (16.4%)	0.39
<b>Code status on ICU admission</b>					<b>0.62</b>
Full code	442 (89.1%)	173 (92.5%)	124 (81.1%)	65 (82.3%)	
Time-limited trial	12 (2.4%)	7 (3.7%)	3 (1.9%)	8 (10.1%)	
DNI/DNR	42 (8.5%)	7 (3.7%)	26 (17%)	6 (7.6%)	
Need for intubation (overall)	202 (40.7%)	77 (41.2%)	61 (39.9%)	34 (43.0%)	0.95
Intubation ≥ day 3	99/202 (49%)	45/77 (58.4%)	29/61 (47.5%)	17/34 (50%)	0.74
ICU mortality	107 (21.6%)	55 (29.4%)	39 (25.5%)	20 (25.3%)	0.28
Hospital mortality (38 missing)	162 (32.7%)	69 (36.9%)	62 (40.5%)	27 (34.2%)	0.54
Day-90 mortality (140 missing)	198 (39.9%)	78 (41.7%)	74 (48.4%)	36 (45.6%)	0.85

to 596 (37%) patients who were intubated on ICU admission (Fig. 1). Among the 915 patients not intubated on ICU admission, 496 (54.2%) received standard oxygen

only, 187 (20.4%) HFNC only, 153 (16.7%) NIV only, and 79 (8.6%) both HFNC + NIV. Table 1 shows patients' characteristics in these four groups. Patients receiving

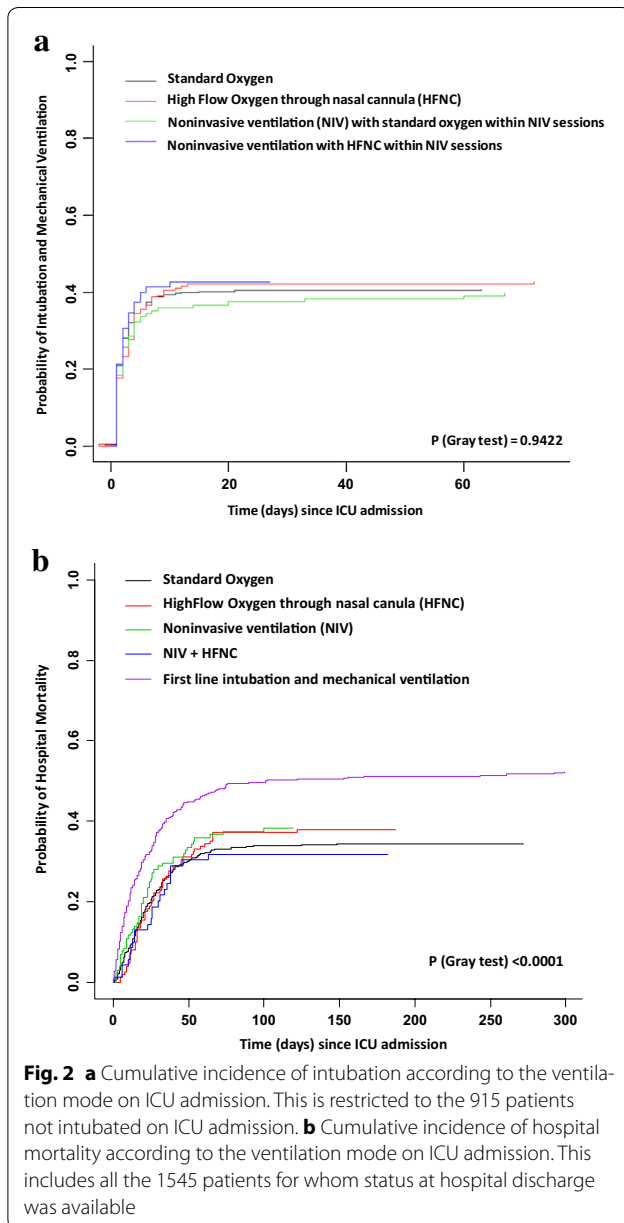


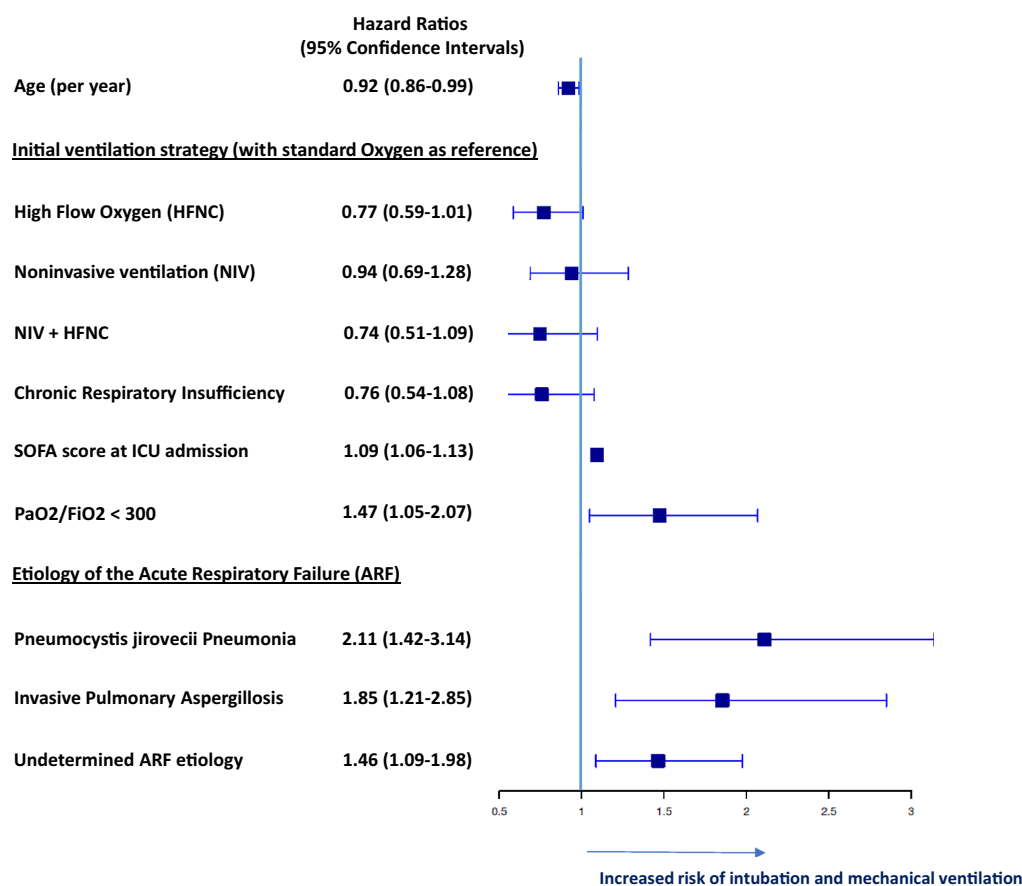
HFNC at day 1 were more severely hypoxemic than other patients ( $\text{PaO}_2/\text{FiO}_2 = 122$  (88–183) vs. 163 (103–250),  $P < 0.0001$ ). In patients receiving continuous standard oxygen, oxygen flow was 8 L/min (4–15), 5 L/min (3–8), and 4 L/min (2–7), at day 1, day 2, and day 3, respectively. In patients receiving NIV, NIV duration was 8 h/day (5–14), 8 h/day (4–15), and 7 h/day (3–15), at day 1, day 2, and day 3, respectively.  $\text{FiO}_2$ , PEEP, and pressure support at day 1 were 60% (40–80), 8  $\text{cmH}_2\text{O}$  (5–10), and 14  $\text{cmH}_2\text{O}$  (9–16). In patients receiving continuous HFNC, oxygen flow was 50 L/min (40–50) on day 1 and day 2, and 40 L/min (30–50) on day 3. HFNC- $\text{FiO}_2$  was

80% (60–100), 70% (50–100), and 50% (40–75) at day 1, day 2, and day 3, respectively. Among patients who were intubated, initial settings were tidal volume of 6.7 mL/kg (6–7.5) ideal body weight and PEEP of 8  $\text{cmH}_2\text{O}$  (6–12).

Cumulative incidence of intubation and hospital mortality are reported in Fig. 2a, b. By multivariable analysis (Fig. 3), factors associated with the cause-specific hazard of intubation and ventilation were age (hazard ratio = 0.92/year, 95% CI 0.86–0.99), day-1 SOFA score (1.09/point, 1.06–1.13),  $\text{PaO}_2/\text{FiO}_2 < 300$  at day 1 (1.47, 1.05–2.07), and ARF etiology [*Pneumocystis jirovecii* pneumonia, 2.11 (1.42–3.14); invasive pulmonary aspergillosis, 1.85 (1.21–2.85); and undetermined etiology, 1.46 (1.09–1.98)]. The cause-specific hazard of intubation and ventilation associated with HFNC use was 0.77 (0.59–1.01),  $P = 0.06$ . After matching on propensity score of being administered HFNC (Supplemental Fig. 2), there was an effect from HFNC on the cause-specific hazard of intubation (HR = 0.77, 95% CI 0.59–1.00,  $P = 0.05$ ). After matching on propensity score of being administered NIV (Supplemental Fig. 2), there was no effect of NIV on the cause-specific hazard of intubation (HR = 0.90, 95% CI 0.67–1.21,  $P = 0.48$ ).

Tables 2 and 3 compare patients' characteristics according to status at hospital discharge. By multivariable analysis (Fig. 4), determinants of hospital mortality were age [odds ratio/year = 1.18 (1.09–1.27)], direct admission to the ICU [0.69 (0.54–0.87)], day-1 SOFA without respiratory score [1.12/point (1.08–1.16)],  $\text{PaO}_2/\text{FiO}_2 < 100$  [1.60 (1.03–2.48)], and undetermined ARF etiology [1.43 (1.04–1.97)]. The need for intubation was also associated with hospital mortality; however, the magnitude of the effect depended on the timing of intubation and initial oxygenation strategy. For instance, odds for mortality were 2.55 (95% confidence interval, 1.94–3.29) after first-line IMV, 2.31 (1.09–4.91) after failure of HFNC + NIV, 3.65 (2.05–6.53) after failure of NIV alone, 4.16 (2.91–5.93) after failure of standard oxygen, and 5.54 (3.27–9.38) after failure of HFNC alone. On the basis of matched samples from the eligible population of patients with available mode of oxygenation or ventilation at ICU admission and status at hospital discharge, there was no evidence of any survival benefit from NIV (OR = 0.98, 95% CI 0.9–1.1,  $P = 0.69$ ), nor from HFNC (OR = 1.0, 95% CI 0.9–1.1,  $P = 0.86$ ). When restricting these analyses to the population of 859 patients not intubated at ICU admission and free of any DNI order, and after excluding patients with no available status at hospital discharge, results were not markedly modified, with still no survival benefit from NIV (OR = 1.0, 95% CI 0.9–1.1,  $P = 0.68$ ) or from HFNC (OR = 1.0, 95% CI 0.9–1.1,  $P = 0.57$ ).





**Fig. 3** Multivariate model of the cause-specific hazard of intubation. This analysis is restricted to the 915 patients not intubated on ICU admission. Plots report variables independently associated with the need for intubation in the final model, with their 95% confidence intervals

Sensitivity analyses performed after multiple imputation of missing data, or in the subgroup of patients with hematological malignancies, or restricted to the group of patients with full code status led to the same results. When the center was introduced as a cluster effect in the regression models, it was not significantly associated with outcomes (hospital mortality,  $P = 0.30$ ; intubation,  $P = 0.76$ , Supplemental Fig. 3).

Compared to patients with established causes of ARF, those with undetermined ARF etiology had longer symptoms prior to ICU admission [1 day (0–4) vs. 2 days (0–5),  $P = 0.008$ ], were more frequently receiving antifungal prophylaxis (50.3% vs. 41.8%,  $P = 0.04$ ), had a higher SOFA score excluding respiratory score on day 1 [5 (3–8) vs. 4 (2–8),  $P = 0.03$ ], were more frequently intubated on day 1 or 2 (56.4% vs. 47.9%,  $P = 0.05$ ), and had higher hospital mortality rate (52.8% vs. 42.8%,  $P = 0.009$ ). Interestingly, the number of FO-BAL performed were not different (62% vs. 59%,  $P = 0.79$ ). Also, the type and number of noninvasive diagnostic tests including sputum samples, serum biomarkers,

and computed tomography scan were not significantly different between the two groups.

## Discussion

We report a large, prospective multicenter multinational cohort study on critically ill immunocompromised patients with ARF. HFNC has an effect on intubation rate but did not influence all-cause hospital mortality. Use of NIV did not impact outcomes. The need for intubation was associated with mortality with higher odds for mortality in case of NIV or HFNC failure. Most importantly, the cause of ARF influenced IMV and mortality rates, particularly when ARF etiology remained unknown despite comprehensive diagnostic workup, suggesting that any effort towards a diagnosis may trump avoiding intubation.

The hospital mortality of 44.1% stresses the need to maintain vigilance to improve outcomes in this group of patients that is growing and will develop more complications in the future [26]. Of those patients who were not intubated on admission, 39% needed IMV, which is consistent with recent studies [11, 14].

**Table 2 Patient's characteristics according to status at hospital discharge**

Numbers (%) or median (IQR)	Patients discharged alive, N = 863	Patients died during hospitalization, N = 682	P value
Age (years)	63.0 (52.8–70.6)	64.8 (56.6–72.8)	0.001
Female gender	346 (40.1%)	267 (39.1%)	0.78
Chronic pulmonary disease	62 (16.9%)	23 (3.4%)	0.30
Chronic cardiac failure	107 (20.0%)	31 (4.5%)	0.005
Poor performance status (ECOG > 2)	311 (36.0%)	364 (53.4)	< 0.0001
Symptoms duration prior ICU	1 (0–4)	1 (0–4)	0.96
<b>Location before ICU admission</b>			<b>0.08</b>
Emergency department	223 (25.8%)	142 (20.8%)	
Hospital wards	467 (54.1%)	417 (61.1%)	
Other ICUs	56 (6.5%)	48 (7.0%)	
Other locations	117 (13.6%)	75 (11%)	
Direct admission to the ICU	324 (37.6%)	188 (27.6%)	< 0.0001
Nightly/weekend admissions	524 (60.7%)	393 (57.6%)	0.24
<b>Underlying conditions</b>			
Hematological malignancy	438 (50.75%)	362 (53.2%)	0.39
Solid tumor	295 (34.2%)	249 (36.5%)	0.37
Systemic diseases	163 (18.9%)	106 (15.5%)	0.09
Solid organ transplant	97 (11.2%)	59 (8.6%)	0.13
Drug-related immunosuppression	10 (1.2%)	3 (0.4%)	0.21
Neutropenia	122 (14.1%)	130 (19.1%)	0.01
<b>Hematopoietic stem cell transplantation</b>			<b>0.16</b>
Autologous	50 (5.8%)	45 (6.6%)	
Allogeneic	71 (8.2%)	74 (10.9%)	
Days since diagnosis	156 (18–734)	120 (11–578)	0.26
Newly diagnosed disease	241 (27.9%)	216 (31.7%)	< 0.0001
Day-1 SOFA score	6 (3–9)	8 (5–11)	< 0.0001
Day-1 SOFA without respiratory items	4 (1–7)	6 (3–9)	< 0.0001
PaO <sub>2</sub> /FiO <sub>2</sub> on day 1	173 (114–210)	158 (100–195)	< 0.0001
<b>PaO<sub>2</sub>/FiO<sub>2</sub> classes</b>			<b>0.02</b>
≥ 300/200–299	91 (10.5%)/156 (18.1%)	52 (7.6%)/159 (23.3%)	
100–199/< 100	469 (54.3%)/147 (17.0%)	363 (53.2%)/108 (15.8%)	
<b>Goals of care on ICU admission</b>			<b>&lt; 0.0001</b>
Full code status	802 (92.9%)	566 (83.0%)	
Time-limited trial	13 (1.5%)	51 (7.5%)	
DNR/DNI	20/28 (5.6%)	30/35 (9.5%)	
<b>Need for intubation and ventilation</b>			<b>&lt; 0.0001</b>
No	461 (53.4%)	163 (23.9%)	
Within 2 h of ICU admission	261 (30.2%)	261 (30.2%)	
After failure of standard oxygen	75 (8.7%)	112 (16.4%)	
After noninvasive ventilation (NIV) failure	24 (2.8%)	30 (4.4%)	
After high-flow oxygen (HFNC) failure	26 (3.0%)	50 (7.3%)	
After failure of NIV + HFNC	16 (1.8%)	14 (2.0%)	

The analysis did not include 66 patients in whom hospital mortality was missing

NIV use in critically ill immunocompromised patients with ARF was believed to be the gold standard at a time where mortality associated with standard oxygen was above 80%, but only 50% with NIV [27]. A post hoc analysis of a HFNC trial reported that NIV was independently

associated with increased mortality [11]. However, this was not confirmed in a randomized trial that found no benefit and no harm from NIV [14]. Similarly, in the present study, NIV did not influence the need for IMV or mortality rates: among the 237 patients receiving NIV,



**Table 3 Association between acute respiratory failure etiologies and hospital mortality**

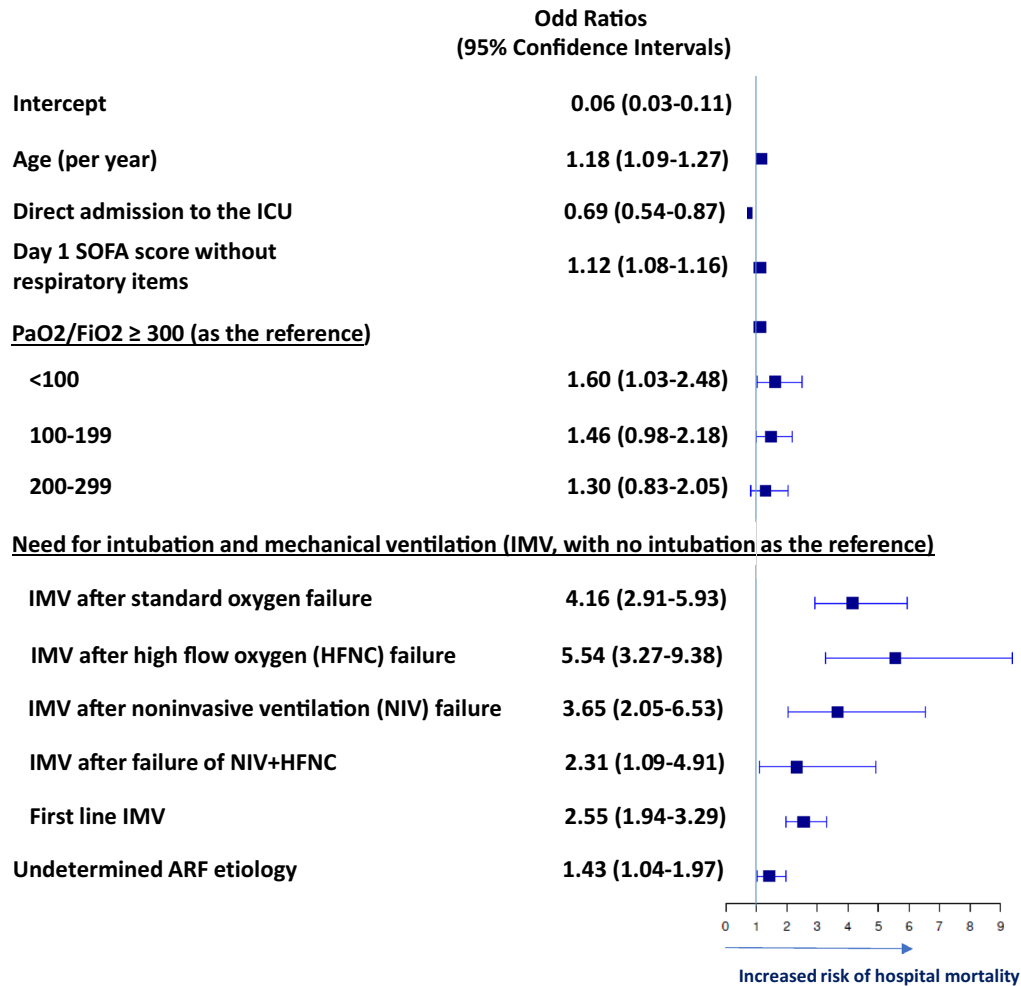
Numbers (%) or median (IQR)	Patients discharged alive from the hospital, N = 863 (55.8%)	Patients who died before hospital discharge, N = 682 (44.2%)	P value
Bacterial infection			0.27
Clinically documented	125 (14.5%)	84 (12.3%)	
Microbiologically documented	129 (14.9%)	118 (17.3%)	
Viral infection			
Influenza	86 (10.0%)	66 (9.7%)	0.92
Other viruses	138 (16.0%)	106 (15.5%)	0.87
Septic shock from extrathoracic source	81 (9.4%)	68 (10.0%)	0.76
Invasive fungal infection (IFI)			
Proven or probable invasive pulmonary aspergillosis (IPA)	30 (3.5%)	31 (4.5%)	0.35
<i>Pneumocystis jirovecii</i> pneumonia	34 (3.9%)	35 (5.1%)	0.32
Candidemia with septic shock	15 (1.7)	21 (3.1%)	0.12
All IFI cases including other fungi and possible cases of IPA	105 (12.2%)	121 (17.7%)	0.003
Aspiration pneumonia	38 (4.4%)	31 (4.5%)	0.99
Airway-related disorders	38 (4.4%)	16 (2.3%)	0.04
Drug-related pulmonary toxicity	34 (3.9%)	19 (2.8%)	0.27
Disease-related infiltrates	69 (8.0%)	73 (10.7%)	0.08
Cardiogenic pulmonary edema	66 (7.6%)	33 (4.8%)	0.03
Undetermined	94 (10.9%)	105 (15.4%)	0.01
More than one ARF etiology	125 (14.5%)	105 (15.4%)	0.67

90 (38%) subsequently required IMV, so did patients receiving standard oxygen or HFNC. The same applied to hospital mortality. Regarding HFNC, a post hoc analysis restricted to immunocompromised patients with ARF reported that ORs for intubation were higher in patients treated with NIV than in those treated with HFNC [11]. However, this was not confirmed by the post hoc analysis of the NIV trial in which HFNC was neither associated with a lower intubation rate nor day-28 mortality [14]. In the present study, HFNC, which was used in more severely hypoxemic patients, has an effect of IMV rate. Last, the association of HFNC + NIV was not associated with increased intubation or hospital mortality.

Another major finding from this study is the association between the cause of ARF and outcomes. More particularly, patients with ARF from unknown cause in spite of extensive workup were at high risk of IMV and death. The prevalence of 13.2% of unknown diagnoses is in agreement with previous reports [18, 23, 27, 28]. Also, this study confirms that these patients underwent as much diagnostic efforts as patients with identified ARF etiology. That patients with undetermined cause of ARF have poor outcomes could be the result of delay and failure to recognize and treat a pulmonary complication. However, more research is needed to understand whether this clinical entity does actually refer to unidentified pathogens, unknown toxicity, or atypical organ

involvement by the disease. For now, this result suggests that more data should be gathered on ARF from undetermined etiology to guide clinicians towards a more aggressive diagnostic strategy, including cytokine concentrations, as well as molecular diagnostics in blood and BAL, and minimally invasive CT-guided lung biopsies, cryobiopsies, or surgical biopsies. Meanwhile, empirical steroids, antifungals, or antiviral agents may also be considered despite the lack of evidence. Furthermore, this study suggests that in addition to adjusting oxygenation strategy to a patient's severity, clinicians should also focus on how the selected oxygenation device enables an appropriate diagnostic workup.

Strengths of our study include the multicenter and multinational design and a large number of patients, making this study the largest to date. Also, follow-up and hospital mortality were missing in only 4% of the patients. The profile of infectious diseases in our population indicates severe immunologic impairment. These results also have a high degree of external validity, since the centers belong to a large multinational study group including university and nonuniversity hospitals. We found no evidence that oxygenation modalities influenced any of the endpoints or were beneficial in some subgroups. Moreover, NIV and HFNC were administered optimally in terms of duration and settings. Overall, now that NIV use is either neutral [14]



**Fig. 4** Multivariate model of the prevalence of hospital death. This analysis is restricted to the 1545 patients with available status at hospital discharge. Plots report variables independently associated with hospital mortality in the final model, with their 95% confidence intervals

or associated with increased mortality [11], NIV can hardly be part of the standard of care in this setting until trials adequately powered and with appropriate control groups demonstrate that a better NIV administration (continuous and prolonged NIV or helmet-based continuous positive airway pressure) translates into improved survival. Also, in immunocompromised patients with hypoxemic ARF requiring intubation, no trial has assessed benefits from NIV or HFNC as demonstrated in non-immunocompromised patients [29]. Regarding HFNC, no impact was found regarding mortality. Specific trials of oxygen therapy in immunocompromised patients with hypoxemic ARF are needed before any conclusion can be drawn.

The present study has several limitations. First, the observational design precludes strong conclusions; however, the use of propensity score matching, which allowed us to erase main imbalances in potential

confounders across oxygenation groups, reached conclusions that were similar to those from regression modelling. Second, criteria for intubation, management decisions, or diagnostic and therapeutic strategies were not standardized across centers. However, use of NIV, HFNC, and IMV rates were not significantly different across centers, and no center effect was found. Similarly, the number of diagnostic procedures and the proportion of patients left with undetermined ARF etiologies were not significantly different across centers; nevertheless, such a multicenter design argues for the good external validity of our results. Third, debates about ventilation and oxygenation modalities often arise in studies reporting no benefit. However, the 8-h median NIV duration per day and HFNC settings observed in the present study are consistent with those in other recent reports [11–13]. Studies comparing HFNC use versus standard oxygen are needed. Last, earlier and more

aggressive diagnostic and therapeutic strategies should be developed and validated for patients with ARF from undetermined etiology. Lung biopsies are difficult to perform in patients with refractory hypoxemia and frequent hemostatic disorders and thrombocytopenia. Also, the value of rescue ECMO to increase time to treatment response or to support lung biopsy is unclear, as reported outcomes in immunocompromised patients are inconsistent, however, grim in specific subgroups [30].

## Conclusions

Among immunocompromised patients admitted to ICU with hypoxemic acute respiratory failure, initial oxygenation strategy has an impact on IMV rate but not on mortality. However, patients with ARF from unknown etiology are more likely to be intubated and have a higher hospital mortality. Clinicians managing these patients should increase their ability to identify and treat the cause of pulmonary involvement.

### Electronic supplementary material

The online version of this article (doi:10.1007/s00134-017-4947-1) contains supplementary material, which is available to authorized users.

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### Compliance with ethical standards

### Conflicts of interest

The authors declare no conflict of interest in relation to this study.

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