## **[Journal of the Iowa Academy of Science: JIAS](https://scholarworks.uni.edu/jias?utm_source=scholarworks.uni.edu%2Fjias%2Fvol96%2Fiss1%2F5&utm_medium=PDF&utm_campaign=PDFCoverPages)**

[Volume 96](https://scholarworks.uni.edu/jias/vol96?utm_source=scholarworks.uni.edu%2Fjias%2Fvol96%2Fiss1%2F5&utm_medium=PDF&utm_campaign=PDFCoverPages) | [Number](https://scholarworks.uni.edu/jias/vol96/iss1?utm_source=scholarworks.uni.edu%2Fjias%2Fvol96%2Fiss1%2F5&utm_medium=PDF&utm_campaign=PDFCoverPages) [Article 5](https://scholarworks.uni.edu/jias/vol96/iss1/5?utm_source=scholarworks.uni.edu%2Fjias%2Fvol96%2Fiss1%2F5&utm_medium=PDF&utm_campaign=PDFCoverPages)

1989

# Estimates of Genetic Variability in F2 Maize Populations

Geng-Chen Han *Iowa State University*

A. R. Hallauer *Iowa State University*, hallauer@iastate.edu

Copyright © Copyright 1989 by the Iowa Academy of Science, Inc. Follow this and additional works at: [https://scholarworks.uni.edu/jias](https://scholarworks.uni.edu/jias?utm_source=scholarworks.uni.edu%2Fjias%2Fvol96%2Fiss1%2F5&utm_medium=PDF&utm_campaign=PDFCoverPages) Part of the [Anthropology Commons](http://network.bepress.com/hgg/discipline/318?utm_source=scholarworks.uni.edu%2Fjias%2Fvol96%2Fiss1%2F5&utm_medium=PDF&utm_campaign=PDFCoverPages), [Life Sciences Commons,](http://network.bepress.com/hgg/discipline/1016?utm_source=scholarworks.uni.edu%2Fjias%2Fvol96%2Fiss1%2F5&utm_medium=PDF&utm_campaign=PDFCoverPages) [Physical Sciences and](http://network.bepress.com/hgg/discipline/114?utm_source=scholarworks.uni.edu%2Fjias%2Fvol96%2Fiss1%2F5&utm_medium=PDF&utm_campaign=PDFCoverPages) [Mathematics Commons,](http://network.bepress.com/hgg/discipline/114?utm_source=scholarworks.uni.edu%2Fjias%2Fvol96%2Fiss1%2F5&utm_medium=PDF&utm_campaign=PDFCoverPages) and the [Science and Mathematics Education Commons](http://network.bepress.com/hgg/discipline/800?utm_source=scholarworks.uni.edu%2Fjias%2Fvol96%2Fiss1%2F5&utm_medium=PDF&utm_campaign=PDFCoverPages)

#### Recommended Citation

Han, Geng-Chen and Hallauer, A. R. (1989) "Estimates of Genetic Variability in F2 Maize Populations," *Journal of the Iowa Academy of Science: JIAS, 96(1),* 14-19. Available at: [https://scholarworks.uni.edu/jias/vol96/iss1/5](https://scholarworks.uni.edu/jias/vol96/iss1/5?utm_source=scholarworks.uni.edu%2Fjias%2Fvol96%2Fiss1%2F5&utm_medium=PDF&utm_campaign=PDFCoverPages)

This Research is brought to you for free and open access by UNI ScholarWorks. It has been accepted for inclusion in Journal of the Iowa Academy of Science: JIAS by an authorized editor of UNI ScholarWorks. For more information, please contact [scholarworks@uni.edu](mailto:scholarworks@uni.edu).

### **Estimates of Genetic Variability in F2 Maize Populations <sup>1</sup>**

#### GENG-CHEN HAN and A. R. HALLAUER

Department of Agronomy, Iowa State University, Ames, Iowa 50011

Maize (Zea mays L.) breeders emphasize selection within F<sub>2</sub> populations derived from crosses of inbred lines. Studies of the inheritance of quantitative traits in maize have been conducted primarily for genetically broad-based populations. Objectives of our study were to estimate the genetic variability in  $F_2$  populations developed from crosses of related and unrelated lines and to determine the effects of five generations of random intermating of plants within  $F_2$  populations on the estimates of genetic variability. Estimates of additive genetic variability were greater in the unrelated line crosses, but the estimates were not significantly different before and after random intermating within both crosses. Estimates of dominance variance decreased with random mating, suggesting that linkage effects were affecting the estimates. For applied breeding programs, it seems that adequate genetic variability was available in both types of crosses and that five generations of random intermating were not effective for increasing genetic variability.

INDEX DESCRIPTORS: Zea mays L., Corn, Quantitative traits, Intermating, Linkage effects, Breeding methods

Information on the genetic variability of breeding populations is important in the choice of effective breeding methods and the expected responses ro selection. Estimates of genetic components of variances and average levels of dominance of genes controlling the inheritance of quantitative traits also are useful in the theory of heterosis and its expression in hybrids. Mating designs were developed to provide estimates of genetic components of variance, assuming linkage equilibrium (Comstock and Robinson, 1948, 1952.) These authors emphasized that estimates of additive and dominance genetic variances were biased if the populations were in linkage disequilibrium, and that bias was reduced with genetic recombination in advanced random mating populations.

Estimates of genetic components of variance were summarized by Hallauer and Miranda ( 1981) for different types of maize *(Zea mays* L.) populations. Additive genetic variance estimates were of. greater importance than the dominance variance estimates for 11 traits with the average levels of dominance in the partial to complete dominance range. Linkage disequilibrium, however, biased estimates of genetic variances and average levels of dominance in  $F_2$  populations formed by crossing inbred lines. Few estimates of genetic parameters have been reported  $\ln$  F<sub>2</sub> populations developed from hybrids that exhibited different levels of heterosis. Moreno-Gonzalez and Dudley ( 1981) estimated genetic effects in crosses of related and unrelated maize lines. Heterosis was greater in crosses of unrelated lines, but the genetic effect estimates were not associated with crosses of either related or unrelated lines.

Objectives of our study were to estimate and compare genetic components of variance and average levels of dominance of genes in two maize populations that descended from a cross of two unrelated lines ( $B73 \times$  Mo17) and from a cross of two related lines ( $B73 \times B84$ ). B73 and Mo17 are elite lines that represent the 'Reid Yellow Dent' (873) and 'Lancaster Surecrop' (Mol7) heterotic pattern used in the U.S. Corn Belt. B73 and B84 were developed from BS13, a strain of Iowa Stiff Stalk Synthetic (BSSS) that has been under half-sib recurrent selection (Hallauer et al., 1983). BSSS includes germplasm that is considered representative of Reid Yellow Dent. Estimates of genetic components of variance were obtained for the  $F_2$  populations and after five generations of random intermating within each  $F_2$  population to determine the linkage effects on the estimates of genetic parameters.

#### **MATERIALS AND METHODS**

 $F_2$  populations from the B73  $\times$  Mo17 and B73  $\times$  B84 single-cross hybrids and  $F_2$  synthetic 5 ( $F_2$  Syn. 5) populations obtained after five generations of intermating the  $F_2$  populations were the source materials of the study. B73 and B84 were denved from Iowa Stiff Stalk Synthetic population after five and seven cycles of half-sib recurrent selection, respectively (Hallauer et al., 1983). Mo17 was derived by pedigree selection from the cross,  $C103 \times 187$ -2; C103 was developed from a strain of Lancaster Surecrop and 187-2 developed from Krug, a selected strain of Reid Yellow Dent. Mo17's performance in hybrids, however, is similar to lines derived from Lancaster Surecrop. Herein,  $B73 \times$  Mo17 and  $B73 \times B84$  are considered unrelated and related line crosses, respectively.

The B73  $\times$  Mo17 and B73  $\times$  B84 single crosses (F<sub>1</sub>'s) were selfpollinated to produce the respective  $F_2$  populations. For each  $F_2$ population, 250 plants were intermated. No tassel was used to pollinate more than two ears, and reciprocal crosses between the same plants were not made. Every plant, where possible, was used either as a male or as a female. All pollinated ears were harvested and equally sampled to form a 250-seed bulk for the next generation of intermating. These procedures were repeated for five generations to produce the  $F_2$  Syn. 5 populations.

The North Carolina Design III mating design was used to develop half-sib and full-sib progenies within each of the two  $F_2$  and two  $F_2$ Syn. 5 populations (Comstock and Robinson, 1952). Randomly chosen plants within each population were used as males and backcrossed to each of the respective parental lines (B73, B84, and Mo 17) of the respective  $F_2$  and  $F_2$  Syn. 5 populations. The plants used as males in the backcrosses also were self-pollinated to produce  $S_1$ progenies. Hence, for each randomly chosen plant, there was one pair of backcrosses and one  $S_1$  progeny. Adequate seed for testing was obtained for 40 pairs of backcross progenies and 40  $S<sub>1</sub>$  progenies for each population except for (B73  $\times$  Mo 17) F<sub>2</sub> Syn. 5, in which 38 pairs were available. The study included, therefore, three inbred lines, two single crosses, two  $F_2$  populations, 316 backcross progenies (158 pairs) from the  $F_2$  and  $F_2$  Syn. 5 populations, and 158  $S_1$  progenies from the  $F_2$  and  $F_2$  Syn. 5 populations.

Experiments were conducted at two locations (Agronomy Research Center, Ames, and Ankeny, 1985; Agronomy Research Center, Ames and another location near Ames, 1986) for two years. A randomized incomplete block design was used for each experiment. The genetic materials were assigned to four sets, with each set including an equal number of backcross and  $S_1$  progenies from each of the four populations. The first three sets included 80 backcross (20 from each population) and  $40 S<sub>1</sub>$  (10 from each population) progenies. The fourth set included 76 backcross and 38  $S<sub>1</sub>$  progenies, three inbred

<sup>&#</sup>x27;Joint contribution: USDA-ARS, and Journal Paper No. J-12956 of the Iowa Agric. and Home Econ. Exp. Stn., Ames, IA 50011. Project No. 2778. Part of dissertation submitted by Geng-Chen Han in pattial fulfillment of the requirements for Ph.D. degree at Iowa State *Univ.* 

lines (B73, B84, and Mo17), B73 $\times$ Mo17 and B73 $\times$ B84 single crosses, and the  $F_2$  populations of B73  $\times$  Mo17 and B73  $\times$  B84. Each set was partitioned to include subplots of  $S_1$  and backcross progenies to reduce competition effects between the two types of progenies. Each set was replicated twice for each location. Each plot was one row 5.5 m long with 76 cm between plots. Plots were overplanted and thinned to a uniform stand of 50,232 plants ha $^{\text{\tiny{-1}}}$  . All plots were handharvested except the experiment conducted at Ankeny (1985), which was discarded because of drought conditions. Harvested ears were placed in forced-air dryers and dried to a uniform moisture of 6% before data were recorded for yield and ear components.

Days from planting to 50% silk emergence were recorded for the two experiments conducted at the Agronomy Research Center. Ear height (cm) was measured from ground level to the node bearing the primary ear for five plants within each plot. Ear length (cm) and ear diameter (cm) were measured for 10 ears for each plot and expressed as the average of the 10 measured ears. Number of kernel rows were counted for 10 ears for each of the two experiments conducted in 1986 and expressed on an ear basis. Yield was recorded as the total shelled grain weight of all harvested ears and converted to t ha<sup>-1</sup>. Plot means (days to flower, ear height, ear length and diameter, and number of kernel rows), and plot totals (yield) were used in the analyses of variance.

A split-plot analysis of variance was calculated for each set within each environment. An analysis that included pooling over sets for the backcross and  $S_1$  progenies for each trait was calculated within each environment. Finally, the pooled analyses for each environment were combined over environments to permit estimation of genotype by environment interactions for the backcross and  $S<sub>1</sub>$  progenies.

The genetic model used for estimation of genetic components of variance for the backcross progenies was the Design III model developed by Comstock and Robinson (1952). Two components of variance are estimable from the backcross analyses: variation among males  $(\sigma_m^2)$ , which is the covariance of half sibs or  $(\frac{1}{4}) \sigma_A^2$  (additive genetic variance), and interaction of males with the two parent lines  $(\sigma_{\rm mp}^2)$  or  $\sigma_{\rm D}^2$  (variance due to dominance deviations). Average level of dominance (d) of genes controlling the expression of the trait was estimated as  $(\sigma_{mp}^2/2\sigma_{m}^2)^{1/2}$ . Analyses of the S<sub>1</sub> progenies also provide estimates of  $\sigma_A^2$  and  $\sigma_D^2$ . Because F<sub>2</sub> populations were sampled, the expected gene frequencies of the loci segregating for alleles affecting the traits are 0.5. Expected genetic components of variance among  $S_1$ progenies for gene frequencies of 0.5 are  $\sigma_A^2$  +  $(1/4)\sigma_D^2$ . The same expressions are expected for the  $F_2$  Syn. 5 populations, provided that sampling did not change gene frequencies after five generations of random mating. If linkages among loci controlling the traits are present, linkage would bias the estimates of  $\sigma_A^2$  positively with coupling-phase linkages and negatively with repulsion-phase linkages. Estimates of  $\sigma_{\rm D}^2$  are positively biased for both linkage phases. The effects of linkages are expected to be greatest in the  $F_2$  generation and gradually reduce with subsequent generations of genetic recombination. Estimates of average levels of dominance are expected to decrease from the  $F_2$  to the  $F_2$  Syn. 5 if repulsion-phase linkages are important.

Estimates of the  $\sigma_m^2$ ,  $\sigma_{mp}^2$ , and  $\sigma_{s1}^2$  components of variance were translated to the expected genetic ( $\sigma_A^2$  and  $\sigma_D^2$ ) components of variance to estimate  $\sigma_A^2$ ,  $\sigma_B^2$ , and  $\overline{d}$  for each population for the backcross and S<sub>I</sub> progenies. Because only one genetic parameter can be obtained from the S<sub>1</sub> analyses, the variation among S<sub>1</sub> progenies estimates  $\sigma_A^2$  +  $(1/4)\sigma_{\rm D}^2$ . Estimates of  $\sigma_{\rm A}^2$  and  $\sigma_{\rm D}^2$ , their interactions with environments, and experimental error also were determined by the method of least squares by using the mean squares from the backcross and  $S_1$ progenies analyses of variance. Expectations for each of the mean squares for both analyses were determined for  $\sigma_A^2$ ,  $\sigma_{AE}^2$ ,  $\sigma_{D}^2$ ,  $\sigma_{DE}^2$ , and  $\sigma^2$  (experimental error) and expressed in matrix notation that included the  $\sigma_{\rm m}^2$ ,  $\sigma_{\rm mp}^2$ ,  $\sigma_{\rm s1}^2$ ,  $\sigma_{\rm mee}^2$ ,  $\sigma_{\rm s1E}^2$ , and  $\sigma^2$  components of variance.

#### **RESULTS AND DISCUSSION**

Differences between backcross and  $S_1$  progenies and their interactions with environments were highly significant ( $P \le 0.01$ ) for each trait within each of the four populations (analyses not shown). The mean squares for backcrosses and  $S_1$  progenies, averaged for all traits, were 87.7 and 89.5% greater, respectively, than the backcross and  $S<sub>1</sub>$ progeny by environment mean squares. Differences between the average of the two sets of backcross progenies and the mean of the  $S_1$ progenies also were highly significant for all traits. Differences among males pooled over sets and combined over environments were highly significant for each trait for each population except for days to flower in the (B73  $\times$  B84) F<sub>2</sub> population. The male by parent line interaction was highly significant for each trait in each population except for the following instances: a significant ( $P \le 0.05$ ) difference was detected for ear diameter for the (B73  $\times$  B84)  $F_2$  population and number of kernel rows for the (B73  $\times$  Mo17) F<sub>2</sub> and (B73  $\times$  Mo17) F<sub>2</sub> Syn. 5 populations, and nonsignificant differences were detected for ear height for the (B73  $\times$  B84)  $F_2$  and (B73  $\times$  B84)  $F_2$  Syn. 5 populations, days to flower for the (B73  $\times$  B84) F<sub>2</sub> population, and number of kernel rows for the (B73  $\times$  B84) F<sub>2</sub> and (B73  $\times$  B84) F<sub>2</sub> Syn. 5 populations. Each of the four nonsignificant differences occurred in the populations derived from the related line cross. The analyses of variance indicated that significant genetic differences occurred among the progenies to permit the estimation of genetic components of variance.

Means for the pairs of backcross progenies and the  $S_1$  progenies for each trait for each population are listed in Table I. Orthogonal comparisons of the pairs of backcross progeny means for the F<sub>2</sub> vs. F<sub>2</sub> Syn. 5 populations were either significant or highly significant for each trait, and backcross progeny means for the  $(B73 \times M017)$ populations vs. the  $(B73 \times B84)$  populations were significant for all except plant height. Orthogonal comparisons between  $S_1$  progeny means were either significant or highly significant for all comparisons except grain yield for  $(B73 \times B84)$  F<sub>2</sub> vs.  $(B73 \times B84)$  F<sub>2</sub> Syn. 5. Comparisons of the means of the backcross progenies with those for the  $S_1$  progenies were used as an estimate of inbreeding depression (Table 1). Estimates of inbreeding depression were similar for the  $F_2$ and  $F_2$  Syn. 5 populations for both crosses, which would be expected if no significant changes in gene frequency occurred with random mating. S<sub>1</sub> progenies of the ( $\overline{B73} \times \text{Mo17}$ ) populations tended to have greater inbreeding depression than the  $S_1$  progenies for the B73  $\times$  B84 populations, particularly for grain yeld. Greater inbreeding depression would be expected in crosses of unrelated lines. The relative means and estimates of inbreeding depression agree with the a priori choice of inbred lines used in the crosses.

Estimates of components of variance from the combined analyses of variance of the backcross (Design III, Table 2) and  $S_1$  progenies (Table 3) for each trait measured within each population show that the estimates of the components of genetic variances were greater than their respective interactions with environments. The estimates of  $\sigma_A^2$ were numerically greater than the estimates of  $\sigma_{\rm D}^2$  in all instances except for grain yield for the (B73  $\times$  B84) F<sub>2</sub> population (Table 2). If one considers twice the standard errors of the estimates for judging significance between the estimates of  $\sigma_A^2$  and  $\sigma_D^2$ , the estimates of  $\sigma_A^2$ were not significantly greater than the estimates of  $\sigma_{\rm D}^2$  in all instances [e.g., estimates of  $\sigma_A^2$  and  $\sigma_D^2$  for grain yield for the (B73  $\times$  Mo I7) F<sub>2</sub> population, Table 2). Additive effects, however, were of greater importance than the dominance effects for each trait for each population. There were no indications that the five generations of random intermating of plants within the  $F_2$  populations increased genetic variability because of genetic recombination. It was expected a priori that the  $(B73 \times M017)$  F<sub>2</sub> population (cross of unrelated lines) would have greater genetic variability than the (B73  $\times$  B84) F<sub>2</sub> population (cross of related lines). Although not statistically significant in all

		<b>Traits</b>									
				Ear	Days to	Kernel-row					
Population	Progeny <sup>a</sup>	Yield	Height	Length	Diameter	flower <sup>b</sup>	number				
		$tha^{-1}$		-cm--							
$(B73 \times Mo17) F2$	BC1	7.70	104	16.4	4.7	85.3	16.4				
	BC2	7.51	97	19.2	4.3	85.3	13.3				
	S1	5.40	87	16.0	4.3	85.9	14.5				
	I $(\%)^c$	$-29.0$	$-13.4$	$-10.1$	$-4.4$	0.7	$-2.4$				
$(B73 \times M017)$ F <sub>2</sub> Syn. 5	BC <sub>1</sub>	7.75	102	16.5	4.7	85.1	16.5				
	BC2	7.72	98	19.2	4.3	84.8	13.5				
	S <sub>1</sub>	5.50	87	16.3	4.3	85.3	14.7				
	I $(\%)^c$	$-28.9$	$-13.0$	$-8.7$	$-4.4$	0.4	$-2.0$				
$(B73 \times B84) F_2$	BC <sub>1</sub>	6.28	104	15.0	4.7	87.2	17.6				
	BC2	6.48	104	15.8	4.6	87.9	16.2				
	S <sub>1</sub>	5.26	93	14.6	4.6	87.8	16.5				
	$I(\mathcal{C})^c$	$-17.6$	$-10.6$	$-5.2$	$-1.1$	0.3	$-2.4$				
$(B73 \times B84)$ F <sub>2</sub> Syn. 5	BC <sub>1</sub>	6.80	102	15.1	4.7	86.7	17.7				
	BC <sub>2</sub>	6.66	100	16.1	4.6	87.0	16.5				
	S <sub>1</sub>	5.43	88	14.7	4.4	87.0	16.9				
	I $(\%)^c$	$-19.3$	$-12.9$	$-5.8$	$-5.4$	$-0.2$	$-1.2$				
SE (BC means)		0.04	0.28	0.04	0.006	0.07	0.03				
SE $(S_1 \text{ means})$		0.05	0.41	0.06	0.009	0.11	0.05				
C.V. BC $(\%)$		10.9	6.1	5.1	2.8	5.7	3.6				
C.V. $S_1$ (%)		13.2	7.1	5.9	3.0	4.1	4.1				

Table 1. Progeny (BCl, BC2, and Sl) means of six traits for four maize populations averaged over replications, sets, and environments.

<sup>a</sup>BC1 is the backcross to B73, BC2 is the backcross to either Mo17 or B84, and S1 is the progenies after one generation of selfing. bDays from planting to 50% silk emergence.

'Estimate of inbreeding depression calculated as  $1-[S1/(BC1+BC2)/2] \times 100$ .

instances, genetic variability was greater in the population developed from the cross of unrelated lines.

Specific comparisons of the relative magnitude of the variance component estimates and levels of dominance will be discussed for grain yield and plant height for the two crosses. Estimates of  $\sigma_A^2$ , averaged for the  $F_2$  and  $F_2$  Syn. 5 populations, for yield and plant height were 2.4 and 2.1 times greater, respectively, for the  $B73 \times$  Mo17 cross than for the B73  $\times$  B84 cross. Estimates of  $\sigma_{\rm D}^2$  were 2.7 times greater in the  $B73 \times M017$  cross for plant height than the  $B73 \times B84$  cross. Thus, the genetic variability in the unrelated line cross for additive and dominance effects was at least twice the genetic variability in the related line cross. If repulsion-phase linkages were important, they would be expected to cause a negative bias in the estimates of  $\sigma_A^2$ . Random intermating would be expected to reduce the frequency of repulsion phase linkages and, subsequently, reduce the negative bias in the estimates of  $\sigma_A^2$ . On the basis of the estimates obtained for the  $F_2$  and  $F_2$  Syn. 5 populations for both crosses, it does not seem that repulsion-phase linkages had a large affect on the estimates of  $\sigma_A^2$ . If linkage occurs equally often in coupling and repulsion, the bias in  $\sigma_A^2$  also is likely to be small. Linkages have a positive bias on the estimates of  $\sigma_{\rm D}^2$  regardless of the linkage phases. The greatest changes in the estimates of  $\sigma_D^2$  occurred for grain yield in both crosses, suggesting dominance effects were greater than for the other traits (Table 2).

Estimates of the levels of dominance were less in the  $F_2$  Syn. 5 populations relative to the  $F_2$  populations in both crosses for all traits except number of kernel rows for the B73  $\times$  Mo 17 cross and ear height for the  $B73 \times B84$  cross (Table 2). The effects of dominance were greater for grain yield than for the other traits, and the bias due to linkages also was greater. In both crosses, the estimates of levels of dominance were in the overdominance range for the  $F_2$  populations.

After five generations of intermating, the estimates of average levels of dominance were in the partial to complete dominance range. Dominance effects were of lesser importance for the expression of ear height in both crosses, and the effects of random intermating had less effect in changing the estimates.

Estimates of levels of dominance for the other traits were intermediate to those for grain yield and ear height. Estimates of levels of dominance and the biases due to linkage effects for the two crosses were similar to those reported for other  $F_2$  populations of maize (Hallauer and Miranda, 1981).

Estimates of genetic variability among the  $S_1$  progenies had a trend similar to those obtained from the Design III analyses (Table 3). Estimates of variability among  $S_1$  progenies for the B73  $\times$  Mo 17 cross were greater than those for the  $B73 \times B84$  cross, and five generations of random intermating did not have a consistenr effect on the estimates. The trends in the relative magnitude of the estimates for  $S_1$ progenies and the estimates of  $\sigma_A^2$  from the Design III analyses were similar except for grain yield for the (B73  $\times$  B84) F<sub>2</sub> and F<sub>2</sub> Syn. 5 populations, which were reversed (Tables 2 and 3). The estimates of variation among  $S_1$  progenies tended to be greater than the estimates of  $\sigma_A^2$  from the Design III analyses, but this would be expected because variance among  $S_1(\sigma_{51}^2)$  progenies is  $\sigma_A^2 + (1/4)\sigma_D^2$  for  $F_2$  populations. If dominance effects were not important,  $\sigma_{S1}^2 = \sigma_A^2$ . The estimates of  $\sigma_{S1}^2$ , averaged for the  $F_2$  and  $F_2$  Syn. 5 populations, were 135.1 for  $B73 \times M017$  and 49.1 for  $B73 \times B84$  for grain yield (Table 3). The estimates of  $\sigma_A^2$  and  $\sigma_D^2$ , averaged for the F<sub>2</sub> and F<sub>2</sub> Syn. 5 populations, were 106.2 for  $\sigma_A^2$  and 66.9 for  $\sigma_D^2$  for B73  $\times$  Mo17 (Table 2) and 44.2 for  $\sigma_A^2$  and 24.5  $\sigma_D^2$  for B73 × B84;  $\sigma_A^2$  +(1/4) $\sigma_D^2$  is 122.9 for  $B73 \times M017$  and 50.3 for  $B73 \times B84$ , which are similar to the estimates of  $\sigma_{51}^2$  for B73 × Mo17 (135.1, Table 3) and B73 × B84 (49.1, Table 3). The estimates of genetic variability for grain yield

were consistent for both crosses. Assuming only additive and dominance effects, the total genetic variability in the  $B73 \times$  Mo17 populations was 2.5 times greater than in the  $B73 \times B84$  populations, for both the estimates from Design III (Table 2) and the  $S_1$  progenies (Table 3). Dominance effects were important in both crosses, accounting for 38.6 (B73 $\times$ Mo17) and 35.7% (B73 $\times$ B84) of the total genetic variability.

Because five generations of random intermating did not have a consistent effect on the estimates of genetic variability within the  $F_2$ and  $F<sub>2</sub>$  Syn. 5 populations, the best estimates of genetic variability in the  $F<sub>2</sub>$  populations would be those that included all sources (Table 4). The estimates of  $\sigma_A^2$ ,  $\sigma_D^2$ , and their interactions with environments were similar to those for the respective backcross and  $S_1$  analyses. Intermating within the  $F_2$  populations was not effective in changing the estimates of  $\sigma_A^2$ , but the estimates of  $\sigma_D^2$  and the ratio ( $\sigma_D^2/\sigma_A^2$ ) tended to decrease with intermating. Greater genetic variability was present in the unrelated cross and did not change with intermating.

The parental lines used in the crosses to produce the  $F<sub>2</sub>$  populations were selected on the basis of their presumed genetic divergence. The parent lines and their respective  $F_1$ ,  $F_2$ , and  $F_3$  (bulk of  $S_1$  progenies) generations were included in set 4 to compare the means of the different generations for the two crosses (Table 5). The estimates of hcterosis (H) and inbreeding depression (I) for grain yield were greater

for the cross ( $B73 \times M017$ ) of unrelated lines, indicating a greater frequency of heterozygous loci. Except for number of kernel rows, the estimates of H and I agree with the expected genetic differences among the lines used for the two crosses. The estimates of genetic variability were greater in the unrelated line cross, suggesting greater potential responses to selection (Tables 2, 3, and 4). But these types of crosses may not be the more desirable because of the confounding effects of germplasm for maintaining the heterotic pattern (Reid Yellow Dent  $\times$  Lancaster Surecrop) emphasized in the U.S. Corn Belt. Use of unrelated line crosses for developing recovered lines may require more extensive testing to determine the crosses that maximize the heterotic response in hybrids.

Random intermating plants within  $F<sub>2</sub>$  populations did not increase the genetic variability in either related or unrelated crosses. The current practice of sampling  $F_2$  populations without additional genetic recombination by random mating has been used successfully, and our data suggest that the additional time and effort of random intermating are not necessary. The effects of random intermating within  $F<sub>2</sub>$  populations have been examined theoretically and empirically. Bos ( 1977) and Pederson ( 1974) determined theoretically that there were no advantages to intermating  $F_2$  plants in autogamous crop species. Hanson (1959) suggested that  $3$  to  $5$  generations of intermating were necessary to establish linkage equilibrium before initiating

Table 2. Estimates of additive  $(\sigma_A^2)$  and dominance  $(\sigma_B^2)$  components of variance, their interactions  $(\sigma_{AF}^2$  and  $\sigma_{DF}^2)$  with environments, experimental errors, and levels of dominance for six traits obtained from the Design III analysis pooled over sets and combined over environments for four maize populations.

			Variance component estimates		Level of			
	Population Generation	Trait	$\sigma^2_A$	$\sigma^2_{AE}$	$\sigma^2$ <sub>D</sub>	$\sigma^2$ <sub>DE</sub>	$\sigma^2$	dominance
$B73 \times Mo17$ F <sub>2</sub>	$F_2$ Syn. 5	Yield (t $ha^{-1}$ ) <sup>a</sup>	$104.3 \pm 30.8$ $108.0 \pm 32.6$	$17.1 \pm 9.0$ $17.6 \pm 9.2$	$85.0 \pm 23.0$ $48.8 \pm 14.2$	$13.9 \pm 5.2$ $2.3 \pm 4.0$	60.1 $\pm$ 2.8 $60.1 \pm 2.8$	1.28 0.95
	F <sub>2</sub> $F_2$ Syn. 5	Ear height (cm)	$85.1 \pm 22.9$ $127.1 \pm 33.1$	$2.9\pm 4.6$ $0.3 \pm 4.2$	$12.9 \pm 4.3$ $14.7 \pm 4.7$	$-2.6 \pm 2.0$ $-4.4 \pm 1.9$	$38.3 \pm 1.8$ $38.3 \pm 1.8$	0.55 0.48
	F <sub>2</sub> $F_2$ Syn. 5	Ear length (cm)	$11.5 \pm 3.4$ $11.5 \pm 3.5$	$1.4 \pm 1.0$ $1.3 \pm 1.0$	$3.8 \pm 1.3$ $3.5 \pm 1.0$	$1.4 \pm 1.1$ $0.6 \pm 0.5$	$7.3 \pm 0.3$ $7.3 \pm 0.3$	0.81 0.78
	F <sub>2</sub> $F_2$ Syn. 5	Ear diameter (cm) <sup>a</sup>	$3.4\pm 0.9$ $2.5 \pm 0.7$	$-0.1\pm 0.2$ $-0.1\pm 0.2$	$1.0 \pm 0.3$ $0.6\pm 0.2$	$0.0 \pm 0.1$ $-0.1 \pm 0.1$	$1.6 \pm 0.1$ $1.6 \pm 0.1$	0.78 0.72
	F <sub>2</sub> $F_2$ Syn. 5	Days to flower (no.)	$26.2 \pm 5.8$ $19.4 \pm 3.8$	$7.4 \pm 3.6$ $-3.3\pm 3.3$	$4.7 \pm 1.3$ $3.0 \pm 0.7$	$0.1 \pm 1.2$ $-4.5 \pm 0.5$	$14.3 \pm 0.8$ $14.3 \pm 0.8$	$0.60**$ $0.56**$
	F <sub>2</sub> $F_2$ Syn. 5	Kernels rows $(no.)^c$	$20.8 \pm 3.4$ 6.0 $\pm$ 1.4	$-0.1\pm 0.3$ $2.0 \pm 1.8$	$0.8\pm 0.3^c$ $0.1 \pm 0.0$	$0.8 \pm 0.4$ $0.1 \pm 0.0$	$3.3 \pm 0.2$ $3.3 \pm 0.2$	$0.28***$ $0.53*†$
$B73 \times B84$	$F_2$ and $F_3$ $F_2$ Syn. 5	Yield $(t ha^{-1})^a$	$32.9 \pm 11.6$ $55.4 \pm 17.9$	$-7.7\pm 5.3$ $2.5 \pm 6.8$	$38.3 \pm 10.8$ $10.7 \pm 5.2$	$-4.6 \pm 3.1$ $1.6 \pm 3.8$	60.1 $\pm$ 2.8 60.1 $\pm$ 2.8	$1.53*$ $0.62*$
	F <sub>2</sub> $F_2$ Syn. 5	Ear height (cm)	$55.1 \pm 15.9$ $45.8 \pm 14.6$	$1.2 \pm 4.3$ $8.1 \pm 5.3$	$1.9\pm 3.0$ $3.2 \pm 2.6$	$9.0 \pm 3.3$ $1.2 \pm 2.4$	$38.3 \pm 1.8$ $38.3 \pm 1.8$	$0.26***$ $0.38**$
	F <sub>2</sub> $F_2$ Syn. 5	Ear length $(cm)^b$	$6.2 \pm 2.1$ $8.1 \pm 2.6$	$0.9\pm 0.9$ $1.1 \pm 1.0$	$1.9\pm 0.7$ $1.3 \pm 0.6$	$-0.6 \pm 0.4$ $-0.6 \pm 0.4$	$7.3 \pm 0.3$ $7.3 \pm 0.3$	0.78 $0.57***$
	F <sub>2</sub> $F_2$ Syn. 5	Ear diameter (cm) <sup>a</sup>	$0.7 \pm 0.4$ $1.5 \pm 0.4$	$0.5 \pm 0.2$ $-0.2\pm 0.1$	$0.4\pm 0.1$ $0.3 \pm 0.1$	$0.4 \pm 0.1$ $-0.3 \pm 0.1$	$1.6 \pm 0.1$ $1.6 \pm 0.1$	1.06 $0.64*$
	F <sub>2</sub> $F_2$ Syn. 5	Days to flower $(no.)^b$	$6.0 \pm 4.6$ $12.5 \pm 2.8$	$26.6 \pm 13.3$ $-3.1\pm 3.8$	$1.8 \pm 1.5$ $3.3 \pm 0.9$	$6.0 \pm 2.1$ $-2.3 \pm 0.8$	$14.3 \pm 0.8$ $14.3 \pm 0.8$	0.77 0.73
	F <sub>2</sub> $F_2$ Syn. 5	Kernel rows $(no.)^c$	$10.6 \pm 2.0$ $9.1 \pm 1.7$	$1.1 \pm 1.5$ $0.1 \pm 1.1$	$0.8 \pm 0.3$ $0.3\pm 0.2$	$0.7 \pm 0.4$ $-0.3 \pm 0.2$	$3.3 \pm 0.2$ $3.3 \pm 0.2$	$0.38**$ $0.27***$

<sup>a</sup> Estimates were multiplied by  $10<sup>2</sup>$ .

bEstimates were multiplied by 10.

'Days from planting to  $50\%$  silk emergence.

\* and \*\* indicate that levels of dominance were diffferent from complete dominance at 0.01 and 0.05 probability levels, respectively. t and § indicate that derivations from no dominance were different from no dominance at 0.01 and 0.05 probability levels, respectively.

		Component of variance estimates								
Population		Trait	$\sigma_A^2$	$\sigma_{AE}^2$	$\sigma_{\rm D}^2$ .	$\sigma_{DE}^2$	$\sigma^2$	$\sigma_{D}/\sigma_A^2$		
	$F_2$ Syn. 5	$B73 \times M017$ F <sub>2</sub> Yield (t ha <sup>-1</sup> ) <sup>a</sup>	$112.9 \pm 3.6$ $118.8 \pm 3.4$	$34.8 \pm 9.2$ $22.0 \pm 8.3$	$85.5 \pm 3.9$ $49.5 \pm 3.7$	$5.3 \pm 9.2$	$19.5 \pm 9.6 \quad 48.6 \pm 10.3$ $52.2 \pm 9.8$	0.76 0.42		
	$F_2$ $F_2$ Syn. 5	Ear height (cm)	$69.6 \pm 6.0$ $102.2 \pm 9.2$	$-0.3 \pm 15.4$ $-2.4 \pm 23.7$	$11.9\pm 6.5$ $13.2 \pm 10.0$	$-7.8 \pm 15.9$ $-11.5 \pm 24.5$	$49.5 \pm 17.1$ $54.2 \pm 26.3$	0.17 0.13		
	$F_2$ $F_2$ Syn. 5	Ear length (cm) <sup>b</sup>	$15.1 \pm 1.5$ $12.2 \pm 0.3$	$3.2 \pm 3.7$ $1.3 \pm 0.7$	$4.0 \pm 1.6$ $3.6\pm 0.3$	$2.4 \pm 3.9$ $0.5 \pm 0.7$	$5.0 \pm 4.2$ 7.3 $\pm$ 0.8	0.26 0.30		
	F <sub>2</sub> $F_2$ Syn. 5	Ear diameter (cm)ª	$4.0 \pm 0.3$ $1.4 \pm 0.4$	$0.6 \pm 0.8$ $-0.2\pm 1.1$	$1.1 \pm 0.3$ $0.6\pm 0.4$	$0.3 \pm 0.8$ $-0.4\pm 1.7$	$1.0 \pm 0.9$ $2.3 \pm 1.2$	0.28 0.43		
	F <sub>2</sub> $F_2$ Syn. 5	Days to flower (no.) <sup>bc</sup>	$11.6 \pm 0.2$ $13.1 \pm 2.8$	$3.5 \pm 5.8$ $12.0 \pm 7.3$	$2.8 \pm 2.4$ $2.0 \pm 3.1$	$-3.0\pm 6.1$ $-1.3\pm 7.5$	$20.9 \pm 6.5$ $9.4 \pm 8.1$	0.24 0.15		
$(B73 \times B84)$	$F_2$ $F_2$ Syn. 5 $F_2$ Syn. 5	Kernels rows $(no.)^b$ $F_2$ Yield (t ha <sup>-1</sup> ) <sup>a</sup>	$12.4 \pm 0.4$ $51.4 \pm 8.1$ $35.0 \pm 7.5$	$1.2 \pm 1.1$ $3.7 \pm 0.4$ - 0.7 $\pm$ 1.0 $7.5 \pm 20.8$ $8.2 \pm 19.3$	$0.5 \pm 0.5$ $0.5 \pm 0.4$ $39.4 \pm 8.8$ $9.5 \pm 8.1$	$0.1 \pm 1.1$ $-0.1\pm 1.1$ $5.8 \pm 21.6$ 39.0 $\pm 23.2$ $-1.1 \pm 20.0$	$3.9 \pm 1.2$ $4.8 \pm 1.2$ $67.1 \pm 21.5$	0.04 0.14 0.77 0.27		
	$F_2$ $F_2$ Syn. 5	Ear height (cm)	$59.3 \pm 1.5$ $46.4 \pm 1.7$	$0.4\pm 3.8$ $-1.6\pm 4.4$	$2.2 \pm 1.6$ $3.3 \pm 1.8$	$9.5 \pm 3.9$ 36.9 $\pm 4.2$ $-2.0\pm 4.5$	$43.7 \pm 4.8$	0.04 0.07		
	$F_2$ $F_2$ Syn. 5	Ear length (cm) <sup>b</sup>	$10.3 \pm 1.7$ $8.8 \pm 0.3$	$2.9\pm 4.3$ $1.1 \pm 0.6$	$2.0 \pm 1.8$ $1.3 \pm 0.3$	$0.7 \pm 4.4$ $-0.7\pm 0.7$	$4.6 \pm 4.8$ 7.4 $\pm$ 0.7	0.19 0.15		
	F <sub>2</sub> $F_2$ Syn. 5	Ear diameter (cm) <sup>a</sup>	$1.0 \pm 0.1$	$0.1 \pm 0.2$ $0.3 \pm 0.4$ - $0.4 \pm 1.1$	$0.4 \pm 0.1$ $0.2 \pm 0.5$	$0.3 \pm 0.2$ $-0.7\pm 1.2$	$1.8 \pm 0.2$ $2.4 \pm 1.3$	0.40 0.67		
	$F_2$ $F_2$ Syn. 5	Days to flower (no.) <sup>bc</sup>		$6.1\pm3.2$ - $2.1\pm8.4$ $7.5\pm0.6$ - $2.1\pm1.7$	$1.3 \pm 3.6$ $2.2 \pm 0.7$	$-0.9\pm$ 8.7 25.8 $\pm$ 9.4 $-3.0\pm 1.7$	$16.1 \pm 1.8$	0.21 0.29		
	$F_2$ $F_2$ Syn. 5	Kernel rows (no.) <sup>b</sup>	$5.5 \pm 2.5$	$4.9\pm0.9$ - 0.2 $\pm$ 2.3 $0.0 \pm 0.5$	$0.4\pm 1.0$ $0.2 \pm 0.2$	$-0.3\pm 2.4$ $-0.6\pm 0.5$	$5.4 \pm 2.4$ $3.9\pm 0.5$	0.08 0.04		

Table 4. Estimates of additive ( $\sigma_A^2$ ) and dominance ( $\sigma_D^2$ ) variance components from S<sub>1</sub> and Design III combined analyses Upooled over sets and combined over environments for six traits for four maize populations.

<sup>a</sup>Estimates were multiplied by  $10^2$ .

bEstimates were multiplied by 10.

<Days from planting to 50% silk emergence.

Table 5. Means and estimates of heterosis and inbreeding depression of six traits for inbred lines and their  $F_1$ ,  $F_2$  and  $F_3$ generations measured in three environments.

			Generations					I $(\%)^b$			Generations				I $(\%)^b$	
Trait	<b>B</b> 73	Mol7	${\bf F}$			H $(\%)^a$	$F_2/F_1$ $F_3/F_1$		<b>B</b> 73	<b>B84</b>	${\bf F}$	F,		$H(\%)^a$ $F_2/F_1$ $F_3/F_1$		
Yield (t ha <sup><math>-1</math></sup> )	3.44		4.14 10.87	9.01	5.76	187	$-17$	$-47$	3.43	4.45	7.21	6.49	4.85	83	$-10^{-}$	$-33$
Ear height (cm)	98	87	109	110	90	17		$-17$	98	83	103	101	89	14	$-2$	$-14$
Ear length (cm)	12.3	13.5	20.8	17.3	16.0		$61 - 17$	$-23$	12.3	17.4	18.2	15.1	14.3	23	$-17$	$-21$
Ear diameter																
(cm)	4.3	4.1	4.9	4.9	4.4	17	$\theta$	$-10$	4.3	3.6	4.4	4.5	4.4	11	$\mathcal{P}$	$\Omega$
Days to																
flower $(no.)^c$	89	91	83	86	86	-8	4	4	-89	89	86	87	88			2
Kernel rows																
(no.)	19.7	14.2	15.3	17.6	14.6	$-10$	15	$-5$	19.7	-11.1	- 14.7	16.9	16.4		15	12

<sup>a</sup>H is heterosis estimate calculated as  $(F_1 - P)P^{-1} \times 100$  where P is average of parent lines.

MlllER, P.A., and J. 0. RAWLINGS. 1967. Breakup of initial linkage blocks through inrermating in a corron breeding population. Crop Sci. 7: 199-204.

populations. Crop Sci. 13:132.

- PEDERSON, D. G. 1974. Arguments against intermating before selection in a self-ferrilizing species. Theor. Appl. Genet. 45: 157-162. VERNA, M. M., S. KOCHHAR, and W. R. KAPOOR. 1979. The
- MORENO-GONZALEZ, M., and J. W. DUDLEY. 1981. Epistasis in related and unrelated maize hybrids determined by three methods. Crop Sci. 21:644-651.
- NORDQUIST, P. T., 0.]. WEBSTER, C. 0. GARDNER, and W. M. ROSS. 1973. Registrarion of three sorghum germplasm random mating
- 82:174-181. YUNUS, M., and R. S. PARODA. 1982. Impact of biparental matings on correlation coefficients of bread wheat. Theor. Appl. Genet. 62:337-343.

assessment of the biparental approach in a wheat cross. Z. Pflanzenzucht

			Variance component estimates				
Population	Generation	Trait	$\sigma_{S1}^2$	$\sigma_{S1E}^2$	$\sigma^2$	$h^{2^e}$	
$B73 \times Mo17$	F <sub>2</sub> $F_2$ Syn. 5	Yield $(t ha^{-1})^a$	$136.4 \pm 36.3$ $133.8 \pm 35.3$	$38.3 \pm 10.7$ $21.6 \pm 8.0$	$50.1 \pm 3.3$ 50.1 $\pm$ 3.3	$0.87 \pm 0.23$ $0.90 \pm 0.24$	
	F <sub>2</sub> $F_2$ Syn. 5	Ear height (cm)	$68.7 \pm 17.7$ $99.3 \pm 25.4$	$4.2 \pm 4.2$ $5.2 \pm 4.5$	$39.8 \pm 2.6$ $39.8 \pm 2.6$	$0.90 \pm 0.23$ $0.92 \pm 0.24$	
	F <sub>2</sub> $F_2$ Syn. 5	Ear length (cm)	$17.0 \pm 4.4$ $13.3 \pm 3.5$	$2.1 \pm 1.0$ $1.0 \pm 0.9$	$8.1 \pm 0.5$ $8.1 \pm 0.5$	$0.89 \pm 0.21$ $0.89 \pm 0.27$	
	F <sub>2</sub> $F_2$ Syn. 5	Ear diameter (cm) <sup>a</sup>	$4.5 \pm 1.1$ $1.2 \pm 0.4$	$-0.3\pm 0.2$ $0.4 \pm 0.2$	$1.8 \pm 0.1$ $1.8 \pm 0.1$	$0.88 \pm 0.22$ $0.74 \pm 0.17$	
	F <sub>2</sub> $F_2$ Syn. 5	Days to flower $(no.)^{bc}$	$17.5 \pm 3.6$ $20.5 \pm 4.9$	$4.0 \pm 3.3$ $9.5 \pm 4.6$	$19.8 \pm 1.6$ $19.8 \pm 1.6$	$0.72 \pm 0.16$ $0.67 \pm 0.17$	
	F <sub>2</sub> $F_2$ Syn. 5	Kernels rows $(no.)^b$	$18.3 \pm 3.1$ $5.6 \pm 1.0$	$1.5 \pm 0.8$ $0.8 \pm 0.4$	$4.2 \pm 0.3$ $4.2 \pm 0.3$	$0.92 \pm 0.15$ $0.80 \pm 0.13$	
$B73 \times B84$	F <sub>2</sub> $F_2$ Syn. 5	Yield $(t \text{ ha}^{-1})^a$	$65.9 \pm 17.3$ $32.3 \pm 11.1$	$2.9 \pm 4.9$ $19.2 \pm 7.5$	$50.1 \pm 3.3$ $50.1 \pm 3.3$	$0.88 \pm 0.23$ $0.69 \pm 0.24$	
	F <sub>2</sub> $F_2$ Syn. 5	Ear height (cm)	$60.9 \pm 15.6$ $47.3 \pm 12.2$	$0.6 \pm 3.6$ $-2.2 \pm 3.2$	$39.8 \pm 2.6$ $39.8 \pm 2.6$	$0.90 \pm 0.23$ $0.88 \pm 0.23$	
	F <sub>2</sub> $F_2$ Syn. 5	Ear length $(cm)^b$	$11.9\pm 3.1$ $9.3 \pm 2.5$	$1.1 \pm 0.9$ $0.5 \pm 1.1$	$8.1 \pm 0.5$ $8.1 \pm 0.5$	$0.88 \pm 0.23$ $0.85 \pm 0.23$	
	F <sub>2</sub> $F_2$ Syn. 5	Ear diameter (cm) <sup>a</sup>	$1.2 \pm 0.3$ $1.2 \pm 0.3$	$0.0 \pm 0.2$ $-0.1\pm 0.1$	$1.8 \pm 0.1$ $1.8 \pm 0.1$	$0.67 \pm 0.20$ $0.67 \pm 0.20$	
	F <sub>2</sub> $F_2$ Syn. 5	Days to flower $(no.)^{bc}$	$10.5 \pm 2.1$ $11.8 \pm 2.3$	$-3.6\pm 1.6$ $-4.1 \pm 1.6$	$19.8 \pm 1.6$ $19.8 \pm 0.19$	$0.67 \pm 0.19$ $0.70 \pm 0.19$	
	F <sub>2</sub> $F_2$ Syn. 5	Kernel rows $(no.)^b$	$6.6 \pm 1.2$ $8.2 \pm 1.4$	$-0.5 \pm 0.6$ $-0.2 \pm 0.5$	$4.2 \pm 0.3$ $4.2 \pm 0.3$	$0.88 \pm 0.15$ $0.89 \pm 0.16$	

Table 3. Estimates of variance components among S<sub>1</sub> progenies ( $\sigma_{S1}^2$ ), their interactions with environments ( $\sigma_{S1E}^2$ ), and experimental error for six traits for four maize populations evaluated in three environments.

<sup>a</sup>Estimates were multiplied by  $10^2$ .

bEstimates were multiplied by 10.

"Days from planting to 50% silk emergence.

 ${}^{4}\sigma_{51}^{2}$  has an expected value of  $\sigma_{A}^{2} + (1/4) \sigma_{D}^{2}$  for allele frequency of 0.05.

 ${}^{\text{c}}$ h<sup>2</sup> is the heritability estimate for S<sub>1</sub> progenies calculated as  $\sigma_{\text{S1}}^2/(\sigma_{\text{S1E}}^2/e + \sigma_{\text{S1}}^2)$ .

selection. Intermating followed by selection could increase the REFERENCES chances of obtaining useful recombinants compared with selection without intermating. Empirical studies in cotton *( Gossypium hirsutum*  1.) tobacco *(Nicotiana tabacum* L.), sorghum *[Sorghum bicolor* (1.) Moench), and wheat *(Triticum aertivum* L.) suggested that increased genetic recombination resulted from random intermating (Miller and Rawlings, 1967; Humphrey et al., 1969; Meredith and Bridge, 1971; Verna et al., 1979; Yunus and Paroda, 1982; Nordquist et al., 1973). Altman and Busch ( 1984), however, concluded from a study in spring wheat that the benefits of recombination were inadequate to justify random intermating of  $F_2$  populations of single crosses as a primary breeding procedure before selection. Similar studies have not been reported in allogamous crop species except to determine the linkage biases in the estimation of average levels of dominance of genes controlling the expression of quantitative traits (Hallauer and Miranda, 1981).

Random intermating within two  $F_2$  maize populations included in our study agree with theoretical conclusions for autogamous crop species. Although the two  $F_2$  populations were developed from lines based on level of genetic relationships, the effects of random intermating were similar; random mating did not increase genetic variability. Linkage effects, both coupling and repulsion, were evident in the  $F_2$ populations because the estimates of average levels of dominance generally were reduced after five generations of random intermating.

- ALTMAN, D. W., and R. H. BUSCH. 1984. Random intermating before selection in spring wheat. Crop Sci. 24: 1085-1089.
- BOS, I. 1977. More arguments against intermating  $F_2$  plants of a selffertilizing crop. Euphytica 26:33-46.
- COMSTOCK, R. E., and H.F. ROBINSON. 1948. The components of genetic variance in populations of biparencal progenies and their use in estimating the average degree of dominance. Biometrics 4:254-266.
- COMSTOCK, R. E., and H.F. ROBINSON. 1952. Estimation of average dominance of genes. p. 494-516. In J. W. Gowen (ed.) Heterosis. Iowa State University Press, Ames, IA.
- HALLAUER, A. R., and J. B. MIRANDA, Fo. 1981. Quantitative genetics in maize breeding. Iowa State University Press, Ames, IA.
- HALLAUER, A. R., W. A. RUSSELL, and O. S. SMITH. 1983. Ouantitative analysis of Iowa Stiff Stalk Synthetic. p. 83-104. In J. P. Gustafson (ed.) 15th Annu. Stadler Genetics Symposia. University Missouri, Columbia, MO.
- HANSON, W. D. 1959. The breakage of initial linkage blocks under selected mating systems. Genetics 44:857-868.
- HUMPHREY, A. B., D. F. MATZINGER, and C. C. COCKERHAM. l 969. Effects of random intercrossing in a naturally self-fertilizing species, *Nicotiana tabacum* L. Crop Sci. 9:495-498.
- MEREDITH, W. R., and R.R. BRIDGE. 1971. Breakup of linkage blocks in cotton, *Gossypium hirsutum* L. Crop Sci. 11:695-698.