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Jurassic shift from abiotic to biotic control on marine ecological

success

Kilian Eichenseer¹, Uwe Balthasar¹, Christopher W. Smart¹, Julian Stander², Kristian A. Haaga^{3,4,5}, Wolfgang Kiessling⁶

¹ School of Geography, Earth and Environmental Sciences, University of Plymouth, Drake Circus, Plymouth PL4 8AA, UK

² School of Computing, Electronics and Mathematics, University of Plymouth, Drake Circus, Plymouth PL4 8AA, UK

³ Department of Earth Science, University of Bergen, N-5020 Bergen, Norway

⁴ Bjerknes Centre for Climate Research, PO Box 7803, NO-5020 Bergen, Norway

⁵ K.G. Jebsen Centre for Deep Sea Research, P.O. Box 7803, N-5020, Bergen, Norway

⁶ GeoZentrum Nordbayern, Department of Geography and Geosciences, Universität Erlangen-Nürnberg, Loewenichstraße 28, 91054 Erlangen, Germany

1 Environmental change and biotic interactions both govern the evolution of the biosphere, 2 but the relative importance of these drivers over geological time remains largely unknown. 3 Previous work suggests that, unlike environmental parameters, diversity dynamics differ 4 profoundly between the Palaeozoic and post-Palaeozoic eras. Here we use the fossil record 5 to test the hypothesis that the influence of ocean chemistry and climate on the ecological 6 success of marine calcifiers decreased throughout the Phanerozoic eon. Marine calcifiers 7 build skeletons of calcite or aragonite, and the precipitation of these calcium carbonate 8 polymorphs is governed by the magnesium-to-calcium ratio and temperature in abiotic 9 systems. We developed an environmental forcing model based on secular changes of ocean 10 chemistry and temperature and assessed how well the model predicts the proliferation of 11 skeletal taxa with respect to calcium carbonate polymorphs. Abiotic forcing governs the 12 ecological success of aragonitic calcifiers from the Ordovician to the Middle Jurassic, but 13 not thereafter. This regime shift coincides with the proliferation of calcareous plankton in 14 the mid-Mesozoic. The deposition of biomineralizing plankton on the ocean floor buffers 15 CO₂ excursions and stabilizes Earth's biochemical cycle, and thus mitigates the evolutionary 16 impact of environmental change on the marine biota.

Dramatic shifts in the success of dominant animal groups in Earth history abound in the fossil record¹ and there are numerous Phanerozoic-scale macroevolutionary trends. Traits such as body size, the metabolic rate of dominant taxonomic groups, and physiological buffering capacity have increased over the course of the Phanerozoic²⁻⁵. As a consequence, biotic interactions may have increased as well⁶⁻⁹. In contrast, climate and seawater composition show a cyclical behaviour rather than Phanerozoic-scale trends^{10,11}. The biosphere has evolved to cushion some environmental variability: for example, the buffering of ocean

chemistry has increased due to the mid-Mesozoic ascent of calcifying plankton¹². Perturbations of the carbon cycle have fallen in amplitude, particularly since the mid-Mesozoic¹³ (Fig. 1d), and extinction rates have decreased towards the present¹⁴ (Fig. 1e). Accordingly, we hypothesize that the evolutionary importance of the abiotic environment, relative to intrinsic, biotic factors, has declined through geological time.

29 We tested this hypothesis using the vast fossil record of marine calcifiers. In the inorganic 30 formation of calcium carbonate (CaCO₃), high Mg/Ca ratios and high temperatures have been 31 shown to favour the precipitation of aragonite over calcite, and vice versa¹⁵. Across the Phanerozoic, tectonically driven changes in sea water chemistry and climate have caused 32 33 aragonite and calcite favouring conditions to alternate, giving rise to episodes of "aragonite 34 seas" and "calcite seas"^{10,16} (Fig. 1b). The skeletal mineralogy of calcifying organisms is 35 strongly tied to phylogenetic history, but the *de novo* acquisition of biominerals, skeletal composition, skeletal production, and growth rates of many marine calcifiers are affected by 36 the Mg/Ca ratio and temperature of the surrounding sea water¹⁷⁻²¹, analogous to inorganic 37 38 CaCO₃ formation. If aragonite and calcite seas were influential in the evolution of marine 39 calcifiers, we expect a correspondence of aragonite sea conditions with greater success of aragonitic taxa. 40

41 Aragonite - calcite seas and the success of marine calcifiers

42 We combine a model of past Mg/Ca ratios²² (Fig. 1b) with δ^{18} O temperature 43 reconstructions²³ (Fig. 1c) to quantify *aragonite sea intensity* (*ASI*) in 85 post-Cambrian stages. 44 The *ASI* is parametrized from experimental data¹⁵ via multiple regression. We contrast *ASI* 45 with a measure of the environmental occupancy, or success, of aragonitic genera relative to 46 all calcifying genera (*SCOR_{ara}*), calculated with the Summed Common species Occurrence Rate 47 (SCOR)²⁴. Given that novel acquisitions of calcium carbonate skeletons are rare and that switches in skeletal mineralogy are largely restricted to a few clades^{20,21,25}, changes in SCOR_{ara} 48 49 reflect predominantly the differential success of aragonitic taxa, rather than changing 50 mineralogies within calcifying clades (Supplementary Materials S6). Although the lower 51 preservation potential of aragonitic taxa may lead to underestimate the abundance of 52 aragonitic taxa²⁶, we find that the completeness of the record of aragonitic and calcitic genera 53 is not significantly different (Wilcoxon signed-rank test: p = 0.28, Supplementary Materials 54 S3). The strength of abiotic controls on marine calcifiers is assessed by estimating linear models of SCOR_{ara} against ASI using generalised least squares (GLS) to account for temporal 55 autocorrelation, and by convergent cross mapping (CCM) to detect causal coupling²⁷ 56 (Methods). 57

58 Visual inspection of the SCOR_{ara} and ASI time series suggests an association in the Ordovician - Carboniferous and again in the early Mesozoic, but not for most of the Mesozoic and 59 60 Cenozoic (Fig. 1a). A linear model of SCOR_{ara} against ASI is significant in the Palaeozoic $(R^2 = 0.15, p = 0.017, Table 1)$, and not in the Mesozoic – Cenozoic $(R^2 < 0.01, p = 0.68)$, 61 62 suggesting a decreasing dependence of ecological success on relevant environmental conditions in the Mesozoic. In the entire Ordovician – Pleistocene data, the linear relationship 63 is not significant ($R^2 < 0.01$, p = 0.70), however the sharp rise of SCOR_{ara} across the Permian-64 65 Triassic makes estimating linear models across this boundary problematic.

We used a Bayesian approach to identify the timing of the changes in the relationship between *SCOR_{ara}* and *ASI* (see *Methods*). This analysis identifies the Permian-Triassic boundary as the strongest change point of the entire Phanerozoic time series (supported by 100 % posterior probability; Fig. 2a, purple bar). When evaluating only the Palaeozoic time

series, the Carboniferous – Permian boundary shows the strongest change (73% probability;
Fig. 2a, red bars), whereas no unambiguous single change point is found for the Mesozoic –
Cenozoic time series (all probabilities <25%; Fig. 2a, green bars).

73 To evaluate the possibility of a gradual change, we fitted linear models of SCOR_{ara} against ASI 74 from the Triassic onwards, adding successively more stages towards the present, and 75 repeated the same process with the Palaeozoic data, separately. The relationship of SCORara 76 and ASI is consistently strong in the Palaeozoic, with a weakening around the Carboniferous 77 – Permian boundary (Figure 2b). In the Mesozoic, a positive linear relationship persists up to 78 the Middle Jurassic, although with lower statistical support. No relationship is supported from 79 the Late Jurassic onwards. The relationship of SCOR_{ara} and ASI evolves in a similar way when 80 SCOR_{ara} is calculated only in the subset of organisms which are considered especially 81 responsive to physiochemical changes in the environment (Supplementary Fig. S7)^{2,25,28}.

82 These results are reinforced by using CCM, a technique developed to detect causal coupling 83 between time series by quantifying the extent to which a putative response time series can 84 be used to predict a driver time series²⁷ (see *Methods*). If this prediction is successful beyond 85 some appropriate null hypothesis (*Methods*), we take it as evidence of dynamical coupling. 86 Applying CCM in expanding time windows, we find a significant dynamical influence of ASI on 87 SCOR_{ara} in the Ordovician - Jurassic (Fig. 2c), implying a causal link between aragonite-calcite 88 sea conditions and SCOR_{ara} in this part of the record. Following a maximum in the early 89 Jurassic (Sinemurian), the CCM prediction skill decreases gradually until the Early Cretaceous 90 and remains low thereafter. This decline indicates a weakening influence of ASI on the success 91 of marine calcifiers towards the present.

92 Our findings corroborate the hypothesis of a decreasing environmental influence on marine 93 calcifiers' success. The correlation of aragonite-calcite seas and the success of aragonitic calcifiers decreased chiefly in two episodes: Around the Carboniferous-Permian boundary, 94 95 and in a second episode centred in the Middle Jurassic (Fig 2). Permian seawater had an 96 unusually high CaCO₃ saturation^{12,29}. All else being equal, higher CaCO₃ saturation facilitates calcification³⁰ and tends to shift the CaCO₃ polymorph balance in favour of calcite³¹. 97 98 Consequently, calcifiers with a calcitic skeleton performed better than predicted by ASI when 99 the CaCO₃ saturation state rose in the Permian. Calcitic taxa, particularly brachiopods, 100 became more successful across the Carboniferous – Permian boundary (Supplementary Fig. S8, a detailed discussion of the relative success of the major taxonomic groups of marine 101 102 calcifiers is provided in the Supplementary Materials). The influence of ASI on SCOR_{ara} 103 decreased but their correlation remained positive within the Permian (Fig. 2b, Table 1). 104 Although CaCO₃ saturation state probably remained high during the Triassic³², ASI continued 105 to affect SCOR_{ara} in the early Mesozoic, but not thereafter.

106 The role of calcifying plankton

107 In the mid-Mesozoic, the Earth-Life system was revolutionised by the rise of calcifying 108 plankton¹². Before the widespread occurrence of planktonic calcifiers, CaCO₃ precipitation 109 was largely confined to the continental shelves and linked to the success of benthic calcifiers 110 such as corals and brachiopods. The evolutionary success of calcifying plankton, especially of coccolithophores³³, shifted the carbonate factory from the shelves to the open ocean. 111 112 Calcareous tests sink to the ocean floor and either dissolve or accumulate, depending on the local CaCO₃ saturation state. Since the proliferation of planktonic calcifiers, changes in the 113 114 atmospheric and oceanic CO₂ content have been compensated on geologically short time scales by increased CaCO₃ deposition or dissolution in the deep sea³⁴. The evolutionary impact
 of episodes of severe climatic change with associated ocean acidification such as the
 Palaeocene-Eocene Thermal Maximum has been less severe than comparable events in the
 Palaeozoic and early Mesozoic^{35,36}, possibly a consequence of the increased ocean
 buffering^{12,37}

120 Producing a skeleton out of *sync* with aragonite-calcite sea conditions may be costly especially 121 when CaCO₃ secretion is impeded by ocean acidification. Enhanced ocean buffering after the 122 proliferation of calcifying plankton can explain the diminished response of marine calcifiers to changing aragonite-calcite sea conditions after the mid-Jurassic. Several key events in the 123 124 evolution of calcifying plankton fall into the Middle – Late Jurassic. Planktonic foraminifera 125 first appear in the fossil record during the Lower Jurassic, but the first known deep water carbonate oozes composed of planktonic foraminifera date back to the Middle Jurassic^{38,39}. 126 The thick-walled coccolithophore genus Watznaueria diversified in the early Middle Jurassic, 127 128 resulting in an increase of coccolith flux to the sediment by two orders of magnitude⁴⁰. 129 Nannofossil deposits from the Tethys ocean show that coccolithophores colonised the open 130 ocean during the Late Jurassic and became abundant enough to affect the marine carbonate system⁴¹. 131

132 A new evolutionary regime

Beyond skeletal mineralogy, there is evidence for a wider regulatory change of evolutionary patterns and environmental state shifts in the mid-Phanerozoic. Devastating extinctions in the Late Permian – early Mesozoic overturned the taxonomic composition of marine calcifiers^{36,42} and favoured the survival of active and physiologically buffered animals^{28,43}. As a consequence, the dominant, modern marine biota are less vulnerable to abiotic stressors

than their Palaeozoic counterparts⁴³, and individual energy budgets of bivalves and gastropods increased throughout the early Mesozoic^{3,5}. This rise in available metabolic energy may have helped absorb the cost of secreting a shell out of *sync* with aragonite-calcite sea conditions.

The Mesozoic rise of plankton such as foraminifera³⁸, coccolithophores, and dinoflagellates⁴⁴ 142 143 had an additional effect on the bio-geosphere: tests of phytoplankton act as ballast, 144 increasing its sinking velocity and increasing the depth at which organic carbon is oxidised, 145 which in turn can explain the much lower prevalence of anoxia on Mesozoic and Cenozoic shelves⁴⁵. Increasing oxygenation of shallow ocean water is indicated from the Jurassic 146 147 onwards by iodine-to-calcium ratios⁴⁶ (Fig. 1f). Well-oxygenated shelves stabilise the carbon 148 cycle by reducing the impact of sea level changes on the burial capacity of organic carbon¹³, 149 thus decreasing the potential for catastrophic environmental change. Increasing oxygen availability also allows for higher metabolic rates and more active modes of life in the shelf 150 biota, as has been inferred for the mid-Mesozoic⁴⁷ and may have increased the pace of 151 152 escalation in evolution⁴⁸. The onset of a persistent diversity rise in the Middle Jurassic agrees 153 with this interpretation (Fig. 1g).

Our results specify the long-held notion that "the evolutionary milieu in which taxa find themselves changed substantially" from the Palaeozoic to the modern world⁴⁹: We found a prominent decrease in environmental influence on the ecological success of marine calcifiers, although some 80 million years after the end of the Palaeozoic. This regime shift was caused by a number of abiotic and biotic revolutions in the Earth-Life system. Of all the factors contributing to this pattern, the onset of the modern carbon cycle via deep-sea CaCO₃ sedimentation was likely the most consequential for marine calcifying organisms. The high-

level taxonomic composition of marine life changed towards a "modern" biota after the endPermian catastrophe¹, but the Palaeozoic evolutionary regime may have persisted well into
the Mesozoic.

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282 **Corresponding author**

283 Correspondence and requests for materials should be addressed to Kilian Eichenseer284 (kilian.eichenseer@plymouth.ac.uk).

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293 Author contributions

- 294 K.E., U.B., W.K. and C.W.S. designed the study. J.S. developed and implemented the Bayesian
- changepoint regression analysis. K.A.H. performed the convergent cross mapping analysis.
- 296 K.E. carried out all other data analysis and wrote the initial manuscript draft, and all authors
- 297 contributed substantially to its improvement.

298 Competing interests

299 The authors declare no competing interests.

300 Figures

Figure 1



301



303 (a) ASI (aragonite sea intensity, blue) and SCORara (relative Summed Common species 304 Occurrence Rate of aragonitic genera, black), in 85 Ordovician – Pleistocene stages. represent 305 Shaded areas 2 standard errors around the mean with 306 the ASI error envelope being based on the temperature component (Methods). Stages 307 with only one measurement are drawn as solid dots, stages without observations308 (circles) have been averaged from the neighboring stages.

309 Legend: O = Ordovician, S = Silurian, D = Devonian, C = Carboniferous, P = Permian, T

310 = Triassic, J = Jurassic, K = Cretaceous, Pg = Palaeogene, N = Neogene and Quaternary,

- 311 Ma = Million years ago. The blue red transitions and the vertical bar mark the time
- 312 when the relationship between *ASI* and *SCOR*_{ara} decreased most strongly (See Fig. 2b,
- 313

c).

(b) Modelled Mg/Ca ratio from ref. 22 (blue line) and a compilation of Mg/Ca proxy data
(black dots, see Supplementary Materials S1). The bar at the top delineates calcite and
aragonite sea intervals as predicted by ref. 16.

- (c) Mean stage-level tropical shallow water temperatures calculated from oxygen isotope
 measurements compiled in ref. 23. Stages with only one measurement are drawn as
 solid dots, stages without observations have been averaged from the neighbouring
 stages and are shown as circles. Shaded areas represent 2 standard errors around the
 mean.
- 322 (d) Periodic changes in the envelope of third-order δ^{13} C variations, reprinted from ref. 13 323 with permission of the American Journal of Science. The grey area highlights the 324 variability.
- (e) Genus-level, sampling-standardised extinction proportions (circles, see *Methods*) with
 long-term trend line (brown; LOESS regression with a smoothing span of 0.1).
- (f) Box plots showing the variability of iodine-to-calcium (I/Ca) ratios from shallow water
 carbonates within sampling localities, reprinted from ref. 46 with permission from

- 329 AAAS. I/Ca ratios are considered a proxy for oxygenation, with higher I/Ca ratios330 indicating better oxygenation.
- 331 (g) Sampling-standardised marine genus-level diversity (circles, see *Methods*) with long-
- term trend line (brown; LOESS regression with a smoothing span of 0.1).





334

335 Figure 2: Changing relationship of SCOR_{ara} and ASI.

(a) Bayesian posterior probabilities for changes in the linear regression of SCOR_{ara} against ASI
for the entire time series (purple bar), the Palaeozoic (red) and the Triassic – Pleistocene

(green). The strength of correlation between SCOR_{ara} and ASI changed around the
 Carboniferous – Permian boundary and at the Permian – Triassic boundary, whereas no
 distinct point of change is found in the Mesozoic – Cenozoic time series.

341 (b) Linear models of SCOR_{ara} against ASI in windows of increasing length for the Palaeozoic 342 and for the Mesozoic – Cenozoic. Expanding windows start with the first six stages and all 343 data are plotted at the last stage of the respective window. Black line = slope; blue area = R^2 ; diamonds = p-values; only p-values < 0.1 are shown. The boxes at the bottom 344 345 of the graph indicate the gap for the first five stages (hatched pattern), and whether linear 346 models were generated using ordinary least squares (OLS), or generalised least squares 347 (GLS, see Methods). These results demonstrate a strong positive correlation between 348 SCOR_{ara} and ASI for the Palaeozoic time series, and a strong, although less robust positive 349 correlation in the early Mesozoic that weakens with the inclusion of data from the Middle 350 Jurassic onwards-

351 (c) Predicting ASI from SCOR_{ara} with convergent cross mapping (CCM) for expanding time 352 series. All included time series start with the first Ordovician stage and all data are plotted 353 at the last stage of their respective window. The CCM prediction skills (solid black line) 354 can be interpreted as the strength of dynamical influence of ASI on $SCOR_{ara}$. The dashed red line shows the 95th percentile of 500 random surrogate time series, which we take as 355 356 a significance criterion. Time windows shorter than the Ordovician – Middle Triassic did 357 not pass the CCM convergence test and were not included (see Methods). The CCM skill 358 shows a sustained drop when Jurassic – Early Cretaceous stages are added, which implies 359 a continuously weakening dynamical influence of ASI on SCORara.

360 Tables

Table 1: Linear models with GLS

(d) Generalised least squares linear models of $SCOR_{ara}$ against *ASI* in the entire data set and in temporal subsets. N denotes the number of observations, ϵ denotes the error term of the regression, with AR indicating autoregressive errors and *ind*. indicating independent errors. ϕ denotes the autocorrelation of the error at lag = 1. In models with autocorrelated error terms, R² is calculated from the log-likelihoods of the model and the corresponding null model (*Methods*). Models with independent errors are equivalent to an ordinary least squares linear model.

	Ν	ε	ф	Intercept	Slope	R ²	p
entire data set	85	AR	0.93	0.45	0.05	0.00	0.70
Palaeozoic	38	AR	0.57	0.13	0.25	0.15	0.017
Ordovician - Carboniferous	29	ind.	-	0.07	0.60	0.59	< 0.001
Permian	9	ind.	-	0.05	0.28	0.68	0.006
Mesozoic - Cenozoic	47	AR	0.68	0.65	0.04	0.00	0.68

361

363 Methods

Fossil and palaeoenvironmental data were grouped into 85 Ordovician to Pleistocene geological stages. To achieve more uniform stage durations, we combined stages shorter than 1 million years (myr) with neighbouring stages (Table S4). Consequently, the two Early Triassic stages were combined, as were the four Pleistocene stages. All analyses were carried out in R, version 3.4.1⁵⁰

368 Fossil data. We used the Paleobiology Database (PBDB, https://paleobiodb.org/) to assess the global 369 fossil record of marine calcifiers. The PBDB records occurrences of fossil organisms and the geological 370 setting in which they were preserved. For our analysis, we relied on internal PBDB information on 371 stratigraphy, taxonomy, mineralogy, life habits, preservation, lithification, and palaeocoordinates 372 (Table S3). All Phanerozoic occurrence data were downloaded on 24 January 2017 with standard 373 settings. Cambrian occurrences were later omitted as ASI could only be calculated in the Ordovician -374 Pleistocene due to the insufficient Cambrian palaeotemperature record. We also excluded 375 occurrences from non-marine settings and occurrences that could not be assigned with confidence to 376 a geological stage. We only included occurrences that were identified to genus level and that could be 377 reliably assigned to an invertebrate animal phylum, Foraminifera, Chlorophyta, Rhodophyta, or 378 calcifying "Problematica". Occurrences of plankton (coccolithophores, planktonic foraminifera, 379 planktonic gastropods or planktonic tentaculites) were omitted. We followed the classification of 380 skeletal mineralogy in the PBDB and considered calcifiers as "aragonitic" if their dominant mineralogy 381 was aragonite and no secondary mineral was listed. The skeletal mineralogies recorded in the PBDB 382 are based on the protocol described in ref. 25, from which we deviated in few exceptions – labechild 383 and Palaeozoic chaetetid sponges were classified as possessing a high Mg calcite skeleton (Balthasar 384 et al., unpublished).

As aragonite is thermodynamically instable at Earth surface conditions and eventually dissolves or recrystallizes to calcite, aragonite has a lower preservation potential than calcite^{26,51}. Most occurrences of aragonite preservation are therefore concentrated in the youngest stages. We

minimized this time-dependant bias by excluding all occurrences from collections with aragonite preservation and unlithified sediments, but kept the 19 % of data which had no information on their preservation recorded to avoid excessive loss of fossil data (see Supplementary Text S3). Barring some potential undeclared occurrences of aragonite preservation, the record of aragonitic taxa used herein thus consists of specimens that have been recrystallized to the stable calcite polymorph or that have been preserved in another way, e.g. by silicification.

Aragonite sea intensity (ASI). We developed a proxy for the degree to which aragonite precipitation is favoured relative to calcite precipitation in the non-biogenic environment based on the joint influence of temperature (*T*) and the Mg/Ca ratio on CaCO₃ formation in experiments⁵². We conducted a multiple linear regression of the mole percent of aragonite present in every experiment against the temperature and Mg/Ca ratio under which the experiments were conducted (see figure 1 in ref. 15). We only used experiments that produced > 1 % calcite and > 1 % aragonite. This yields the equation 400

401 (1)
$$ASI = -119.61 + 46.57 \times \frac{Mg}{Ca} + 4.30 \times T$$

To infer past ASI for the Ordovician - Pleistocene, palaeotemperatures were calculated from a 402 403 Phanerozoic δ^{18} O compilation²³, including only measurements from fossil brachiopods, bivalves and 404 planktonic foraminifera from 35° south to 35° north, because measurements from higher latitudes are 405 unavailable for most of the Phanerozoic. For the δ^{18} O (∞ PDB) to T($^{\circ}$ C) transfer function, we calculated 406 palaeotemperatures assuming a Phanerozoic trend of increasing δ^{18} O as in equation (2) from ref. 23. 407 The Mg/Ca ratio ratios were taken from a Phanerozoic model of seawater composition and digitised 408 from fig. 2.A of ref. 22 in steps of 2 million years using the R package digitize. For the Mg/Ca data, the 409 mean from all observations falling into a geological stage was taken. ASI was calculated with average Mg/Ca data and with every individual temperature observation using equation (1). Mean and standard 410 411 error of all ASI were calculated for each stage, and the resulting ASI was normalised to values between 412 0 and 1. No temperature data was available for the Hettangian and for the Hauterivian stage. We calculated *ASI* for these stages with the mean temperature of the two neighbouring stages,respectively.

SCOR_{ara}. Evolutionary success is widely assessed by diversity. Although species diversity and abundance of higher taxonomic ranks or ecological groups are tightly coupled⁵³, we prefer a direct measure of occupancy to assess the ecological success of individual genera. We apply the Summed Common species Occurrence Rate (SCOR), which is driven by the most widespread and common taxa²⁴. SCOR reflects the actual abundance of a group of taxa with good accuracy⁵⁴. The cumulative SCOR of a set of *m* genera is calculated as

421 (2) SCOR =
$$\sum_{i=1}^{m} -\ln(1 - \frac{y_i}{k})$$

where genus *i* is present in *y_i* subsets out of a total of *k* occupied localities. We defined a locality as a 422 423 cell in a global penta-hexagonal grid with 6240 hexagonal and 12 pentagonal grid cells, with an area 424 of ca. 40,800 km². The grid was generated using the hexagrid() function in the icosa package⁵⁵. The 425 contribution of a genus to SCOR thus depends solely on the number of cells it was sampled in. 426 Repeated sampling within a cell does not increase SCOR. SCOR_{ara} reflects the ecological occupancy of 427 aragonitic taxa, relative to all calcifiers occurring at k_{all} localities. SCOR_{ara} is generated by dividing the 428 SCOR of aragonitc taxa with $k = k_{all}$ by the SCOR of all calcifying genera, again with $k = k_{all}$. To get the 429 relative success of major taxonomic groups, we divided the group SCOR by the SCOR of all calcifiers. 430 For every stage, we calculated SCOR using all genera recorded in the PBDB fossil data recorded in that 431 respective stage.

432 We calculated the variance of any SCOR metric with the delta method^{24,56}

433 (3) Var(SCOR) =
$$\sum_{i=1}^{m} \frac{\frac{y_i}{k}}{(1-\frac{y_i}{k})^{*k}}$$

and can approximate the variance of SCOR_{ara} using the means and variance of aragonitic SCOR and the
SCOR of all calcifiers, assuming they are independent:

436 (4)
$$\operatorname{Var}(\operatorname{SCOR}_{ara}) = \left(\frac{\operatorname{aragonitic SCOR}}{\operatorname{all calcifier SCOR}}\right)^2 \times \left(\frac{\operatorname{Var}(\operatorname{aragonitic SCOR})}{\operatorname{aragonitic SCOR}} + \frac{\operatorname{Var}(\operatorname{all calcifier SCOR})}{\operatorname{all calcifier SCOR}}\right).$$

Bayesian change point regression analysis. The relationship of *ASI* and *SCOR*_{ara} changed through time. We developed statistical methodology to identify change points. In particular, we performed inference in the Bayesian framework about the unknown parameters of the model $y_i \sim N(\mu_i, \sigma_i^2)$, i = 1,...,n, independently, in which *n* is the overall sample size and

441 (5)
$$\mu_i = \begin{cases} \alpha_1 + \beta_1 x_i & i = 1, ..., n_1 \\ \sum_{j=1}^2 \alpha_j + \sum_{j=1}^2 \beta_j x_i & i = n_1 + 1, ..., n_n \end{cases}$$

This model allows a distinct linear relationship between y_i and the covariate x_i in a first and second part of the time series. The parameter α_2 and β_2 represent the additional intercept and slope in the second part, added to α_1 and β_1 of the first part. log σ_i is defined in a similar way, allowing for a different relationship between the standard deviation σ_i and the covariate x_i in each time series part.

446 In the Bayesian framework, it is necessary to specify prior distributions for all unknown parameters. 447 We adopted normal priors with very high variances for all intercept and slope parameters. For the 448 change point n_1 , we adopted a discrete uniform prior across integer values from 5 to n-5, implying 449 that the change point divides the time series into two sections with at least five data points each. This 450 prior distribution expresses considerable uncertainty about the position of the change point before 451 seeing the data. As it is impossible to handle the posterior distribution of all these parameters 452 analytically, we followed the standard approach of sampling from this distribution using a Markov chain Monte Carlo (MCMC) algorithm⁵⁷. To do this we used the jags program⁵⁸, accessed in R through 453 454 the R2jags package⁵⁹. Our posterior inference is based on 100,000 iterations of the MCMC algorithm, 455 half of which were discarded as burn-in.

After inferring a first change point at the Permian-Triassic boundary, we split the data set into a
Palaeozoic part and a Mesozoic – Cenozoic part. In each of the two parts, we estimated additional
change points with the method described above.

459

460

(6)
$$y = \alpha + \beta x + \varepsilon$$

461 assumes that the errors ε are independent between observations. For our data, the residuals of an 462 OLS linear model of *SCOR*_{ara} against *ASI* are autocorrelated, which suggests that this assumption of 463 independent errors does not hold. Generalised least square regression (GLS) can incorporate 464 autoregressive errors and thus allows us to test for linear relationships between autocorrelated time 465 series⁶⁰. Autocorrelated errors ε of order *p* can be modelled as

Generalised least squares. Linear regression with ordinary least squares (OLS) of the form

466 (7)
$$\varepsilon_i = \sum_{j=1}^p \phi_j \varepsilon_{i-j} + \delta_i$$

467 with $\delta_i \sim N(0, \sigma^2)$ independently, in which σ is the standard deviation. We created linear models with 468 independent and with autocorrelated error terms of the first order using the gls() function of the nlme 469 R package⁶¹, performing maximum likelihood estimation by specifying gls(..., method = "ML"). For 470 model selection we compared pairs of models with and without autoregressive errors using a 471 likelihood ratio test⁶², implemented using the anova.gls() function. We selected the more complicated 472 model only if the associated p-value was < 0.05 and the likelihood ratio was > 1. In a few instances, 473 models with autocorrelated errors estimated $\phi < 0$, which we attributed to model overfitting. In these 474 cases, we chose the model without autocorrelation.

475 As a goodness-of-fit measure for GLS models, we calculated the likelihood ratio test R² as

476 (8)
$$R^2 = 1 - \exp\left(-\frac{2}{m}(\log L_M - \log L_0)\right)^{63,64}$$
,

477 with *m* being the number of observations, $\log L_M$ being the log-likelihood of the model, and $\log L_0$ being 478 the log-likelihood of the null model of the form $y = 1 + \varepsilon$, with ε being the error as in equation (7).

479 We take 0.05 as the alpha level for the statistical significance of linear regressions.

480 Regression in expanding windows. We assessed the changing strength of a relationship between ASI
481 and SCOR_{ara} through time by calculating linear models in windows of expanding length. Due to the

482 severe increase of SCOR_{ara} across the Permian-Triassic boundary, calculating the linear relationship 483 across the entire data set may be misleading. Instead, linear models were formulated separately in 484 the Palaeozoic and Mesozoic-Cenozoic. In both cases, the shortest window considered comprised the 485 first six data points, while the longest window had 38 points in the Palaeozoic and 47 points in the 486 Mesozoic – Cenozoic. In every window, an OSL and a GLS model was fitted, with autoregressive errors 487 of the first order incorporated into the GLS model. GLS models were used from the first window 488 onwards in which ϕ was positive and a likelihood ratio test comparing the OLS and the GLS model 489 produced a *p*-value < 0.05.

490 **Convergent cross mapping.** The success of marine calcifiers may be influenced by environmental 491 parameters other than climate and ocean chemistry, as well as biotic interactions and chance. It is 492 therefore possible that a causal connection between ASI and SCOR_{ara} exists even when no linear 493 relationship is detected. We test for this possibility using CCM, a model-free time series analysis 494 method based in dynamical systems theory that can detect causal coupling in nonlinear and even chaotic systems²⁷. It asserts that if two processes are causally linked, then information about the 495 driver variable can be recovered from the response variable²⁷. CCM indirectly measures the dynamical 496 497 influence of the driver variable on the response variable by quantifying the extent to which a state 498 space reconstruction (time delay embedding) of the response variable can be used to predict the 499 driver time series. A description of the algorithm can be found in the Supplementary Materials (S2).

To test for a temporally variable influence of *ASI* on *SCOR*_{ara}, we performed CCM analysis on expanding time windows on the stage level data, under the assumption that these coarse-grained data contain sufficient dynamical information about Phanerozoic Earth system dynamics (see Supplementary Materials S2). We used the rEDM R package⁶⁵ to perform CCM analyses for the main paper. Because of the limited number of time series points, we used embedding dimension 2 and embedding lag 1, with zero temporal exclusion radius in the predictions due to the coarse temporal resolution of the data. For a given time window, the CCM analysis is convergent if prediction skill

507 increases with increasing library size (Supplementary Materials S2). If convergence is achieved, for 508 each time window of length L, we report the median CCM skill for 500 bootstrapped samples at a 509 library size L. In addition, to assess the significance of the results, we used surrogate testing with randomly shuffled surrogates⁶⁶. The analysis for a given time window was considered significant if the 510 511 median prediction skill at the largest library size exceeded the 95th percentile of the median 512 prediction skills obtained for an ensemble of 500 surrogate CCM analyses (dashed, red line in Fig. 2c), 513 where each surrogate realization, the driver time series is replaced by a randomly shuffled version of 514 itself.

515 **Diversity dynamics.** We calculated second-for-third extinction proportions⁶⁷ using classical 516 rarefaction^{14,68} with a sampling quota of 500 occurrences per stage and took the mean extinction 517 proportions over 100 subsampling trials. A sampling-standardised diversity curve was generated with 518 shareholder-quorum subsampling¹ by taking the mean of 100 subsampling iterations, each with a 519 quorum of 0.7. Following the recommendations in ref. 1, we relied on the reference-based singleton 520 count, excluded the dominant genus from frequency calculations, and excluded the largest collection 521 from the single-publication occurrence count. To control for short-term sampling variation, we used the corrected sampled-in-bin richness metric⁶⁹, except for the first and the last stage, in which no 522 523 sampling correction could be made. A locally estimated scatterplot smoothing (LOESS) regression⁷⁰ 524 has been calculated from these results using a smoothing span of 0.1. Extinction and diversity computations were performed using the divDyn R package 71 . 525

526 Data availability

527 The data used to calculate *SCOR*_{ara} are available from the Paleobiology Database at 528 https://paleobiodb.org/. The data used to calculate aragonite sea intensity were taken from ref. 15, 529 22, and 23.

530 Code availability

531	The	code	used	to	generate	the	results	can	be	accessed	at
532	https://figshare.com/articles/R_scripts_and_protocols/7199561.										
533											
534	Methods only References										
535	50	R Core	Team.	R: A	Language	and	Environmen	t for S	statistic	al Comput	ing.
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