

Method to detect gravitational waves from an ensemble of known pulsars

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Combining information from weak sources, such as known pulsars, for gravitational wave detection, is an attractive approach to improve detection efficiency. We propose an optimal statistic for a general ensemble of signals and apply it to an ensemble of known pulsars. Our method combines \mathcal{F} -statistic values from individual pulsars using weights proportional to each pulsar's expected optimal signal-to-noise ratio to improve the detection efficiency. We also point out that to detect at least one pulsar within an ensemble, different thresholds should be designed for each source based on the expected signal strength. The performance of our proposed detection statistic is demonstrated using simulated sources, with the assumption that all pulsar ellipticities belong to a common (yet unknown) distribution. Comparing with an equal-weight strategy and with individual source approaches, we show that the weighted combination of all known pulsars, where weights are assigned based on the pulsars' known information, such as sky location, frequency and distance, as well as the detector sensitivity, always provides a more sensitive detection statistic.

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I. INTRODUCTION

Pulsars are believed to be rapidly rotating neutron stars (NSs) that can emit continuous gravitational wave (GW) radiation if their mass distributions are asymmetric [1]. Observations from first-generation GW detectors have placed upper limits on the amplitude of these GWs from the known galactic millisecond pulsars. This in turn allows constraints to be placed on the ellipticities of these NSs [2]. With the advanced detector era having recently begun with Advanced LIGO [3] in operation and Advanced Virgo [4], and KAGRA [5] close behind, we will soon be able to make observations of these sources with significantly increased sensitivity.

For each pulsar with known sky location and assumed GW phasing (as inferred from arrival times of its radio pulses), time and frequency-domain matched-filtering approaches [6–10] are commonly applied. The former has been used within the LIGO-Virgo Collaboration for the known pulsar searches and applies a Bayesian marginalization strategy to the unknown system parameters [6]. The latter, frequency-domain approach, known as the \mathcal{F} statistic [8] performs an analytical maximization of the likelihood over the unknown parameters of each pulsar and it is this method that we make use of for the remainder of this paper.

Combining sources to improve detection probability is an attractive approach to weak signal detection (e.g. detecting NS ellipticity from analysis of the GW stochastic background [11] and detecting gravitational wave memory using binary black hole mergers [12]). Since GW detectors currently study ~ 200 known pulsars, the existing detection

strategy for this relatively large ensemble can be viewed as trying to detect each one separately, and then waiting for the first detection to appear. This is certainly the most obvious strategy to take, but not obviously the most optimal. Cutler and Schutz (CS) [13] proposed an alternative: first sum the \mathcal{F} -statistic from each pulsar, and then use that sum as a new detection statistic. In this initial study, CS used an equal weight for all the pulsars to be combined. One issue with this approach is that including pulsars which are likely to emit relatively weak GWs decreases the signal-to-noise ratio (SNR) of the combined statistic. As indicated in their paper, the SNR of the combined statistic decreases if the detection ensemble includes weak sources where the squared SNR is less than half of the average squared SNR for all observed pulsars. Therefore, to more efficiently detect GWs from an ensemble of all known pulsars, it seems sensible to investigate the effects of giving nonequal weights to the pulsars within the ensemble.

In this paper, we generalize the idea proposed by CS, by considering the prior distribution of GW strengths from the pulsars within the ensemble. After a brief introduction to pulsar GW emission and the \mathcal{F} statistic, we apply the general theory of hypothesis testing, and obtain a Neyman-Pearson criterion for detecting GWs from an ensemble of pulsars. This leads to an optimal detection statistic, which in idealized situations (i.e., when our prior knowledge of the signal and our model for the noise are an accurate representation of reality) provides the highest detection probability with a given false-alarm probability. As we show, this statistic can in some cases be approximated by linearly combining \mathcal{F} -statistic values from the ensemble of pulsars with appropriate weights.

We assume that the ellipticities of pulsars follow a common (yet unknown) intrinsic distribution and that the orientation of their rotation axes is isotropically distributed. We then draw on our knowledge of their sky location, distance from the Earth, and their rotation frequency to construct prior distributions on the expected GW amplitudes from our known pulsars. Since the intrinsic ellipticity distribution remains unknown, we model it as a simple exponential distribution, but perform tests using both exponential and Gaussian distributions.

This paper is organized as follows. In Sec. II, we briefly review the form of GW emission from individual pulsars and the \mathcal{F} statistic; in Sec. III, we introduce the optimal statistic for a general ensemble of pulsars and discuss how it may apply to a set of pulsars in idealized situations; in Sec. IV, we test our statistic on two possible intrinsic distributions of pulsar ellipticity. We summarize our main conclusions in Sec. VI.

II. BRIEF REVIEW OF GW FROM KNOWN PULSARS AND THE \mathcal{F} STATISTIC

In this section, we give a brief overview of the signal model and maximum-likelihood detection statistic for a single pulsar.

A. Gravitational waveform

For a single GW detector, the signal strain as a function of time, $h(t)$, from

$$h(t) = \frac{16\pi^2 \epsilon I f^2}{d} [\alpha_+ \tilde{F}_+(t) + \alpha_\times \tilde{F}_\times(t)] \cos[\Phi(t) + \Phi_0], \quad (1)$$

with

$$\alpha_+ = \frac{1 + \cos^2 \iota}{2} \cos 2\psi + \cos \iota \sin 2\psi, \quad (2)$$

$$\alpha_\times = -\frac{1 + \cos^2 \iota}{2} \sin 2\psi + \cos \iota \cos 2\psi. \quad (3)$$

Here we have assumed the pulsar, at distance d from the Earth, to be an triaxial ellipsoid rotating at frequency f around one of its minor axes, which stays constant in orientation. In [8] this is the case when the angle between the total angular momentum vector of the star and the star's axis of symmetry is $\pi/2$.

The pulsar is nearly spherical, with a moment of inertia I around its rotation axes; ϵ is its ellipticity, given by

$$\epsilon = \frac{I_1 - I_2}{I} \quad (4)$$

with I_1 and I_2 being the two moments of inertia around the two principal axes that are orthogonal to the rotation axis.

The above four quantities (d, f, ϵ, I) define the strength of the source as received at the detector.

In addition, $\tilde{F}_{+, \times}(t)$ are the (time-dependent, due to Earth's rotation) antenna patterns of the detector toward a source at the sky location of the pulsar, while $\Phi(t)$ defines the GW phase evolution inferred from its radio (or x-ray) pulsations—both are considered known. For the type of emission we are considering, GW radiation will be emitted at twice the rotation frequency, $2f$, with additional modulations due to the orbital motion of the pulsar and the motion of the detector due to the Earth's rotation and orbit.

Finally, we have the polarization angle ψ , the inclination angle ι that describes the pulsar's orientation, and Φ_0 , an additional unknown GW reference phase, all of which we consider as unknown.

In terms of notation, our ι and ψ are the same as used in [8], while $\tilde{F}_+(t)$ and $\tilde{F}_\times(t)$ are respectively equivalent to $a(t)$ and $b(t)$ of [8] where we have assumed that the angle between the two interferometer arms equals $\pi/2$.

B. The single-pulsar \mathcal{F} statistic

Under the assumption that the measured strain is a combination of a GW signal and additive detector noise n , with a single-sided noise spectral density $S_h(f)$, the “near optimal” statistic is given by the so-called \mathcal{F} statistic, derived by Jaranowski *et al.* [8]. For point hypotheses with no uncertain model parameters the maximum-likelihood approach of the \mathcal{F} statistic is optimal in the Neyman-Pearson sense whereby the detection probability P_{DE} is maximized at fixed false-alarm probability P_{FA} . However, even for individual pulsar detection the signal model does include additional unknown model parameters in which case the truly optimal approach is Bayesian and requires marginalization over those parameters [14]. Our investigation makes use of the \mathcal{F} statistic as our input data and hence by association also suffers from a lack of total optimality. However, as shown in [15] the reduction in sensitivity of the \mathcal{F} statistic over the fully optimal approach is slight.

For an observation time T_{obs} , the \mathcal{F} statistic satisfies a χ^2 distribution with four degrees of freedom (4-D) and has a noncentrality parameter equal to the squared optimal SNR ρ^2 , defined by

$$\rho^2 = \frac{256\pi^4 \epsilon^2 I^2 f^4 \mathcal{K}}{d^2 S_h(2f)} T_{\text{obs}} \quad (5)$$

(note that $2f$ is approximately the gravitational wave frequency) with

$$\mathcal{K} = \sum_{p,q=+,\times} \alpha_p \alpha_q F_{pq}, \quad (6)$$

$$F_{pq} = \frac{1}{T_{\text{obs}}} \int_0^{T_{\text{obs}}} \tilde{F}_p(t) \tilde{F}_q(t) dt.$$

The unknown quantities defining the optimal SNR are the ellipticity ϵ and the geometrical factors contained within $\alpha_{+, \times}$ describing the GW polarization and orientation of the pulsar. Note that $F_{++} \neq F_{\times \times}$ and that averaging over many sidereal days leads to $F_{+\times} \rightarrow 0$ and so such terms can be ignored.

For $\alpha_{+, \times}$ we assume that ι and ψ are distributed according to a random orientation of the pulsar's rotation axis. In this case, points with coordinates $(\alpha_+, \alpha_\times)$ are distributed on the two-dimensional plane axisymmetrically around the origin, with modulus

$$\zeta \equiv \alpha_+^2 + \alpha_\times^2 = \frac{1 + 6\cos^2\iota + \cos^4\iota}{4} \quad (7)$$

and $\cos \iota$ uniformly distributed between -1 and $+1$. We can write

$$\mathcal{K} = \zeta[F_{++}\cos^2(2\tilde{\psi}) + F_{\times \times}\sin^2(2\tilde{\psi})] \quad (8)$$

with $\tilde{\psi}$ related to ψ by an offset,

$$4\tilde{\psi} = 4\psi - \arctan \frac{4\cos\iota(1 + \cos^2\iota)}{\sin^4\iota}, \quad (9)$$

hence uniformly distributed between 0 and 2π . In this paper, we simply generate an ensemble of binaries using uniformly distributed $\cos \iota$ and uniformly distributed ψ . The average of \mathcal{K} over this ensemble is given by

$$\langle \mathcal{K} \rangle = \frac{\langle \zeta \rangle}{2}(F_{++} + F_{\times \times}) = \frac{2}{5}(F_{++} + F_{\times \times}). \quad (10)$$

It was shown by CS that for the detection of a single pulsar in a network of M detectors, the \mathcal{F} statistic still satisfies a 4-D χ^2 distribution with a noncentrality parameter $\rho_{\text{net}}^2 = \sum_i^M \rho_i^2$, where ρ_i^2 is the optimal single detector SNR as defined in Eq. (5).

C. Scaling of detectability with observation time

In our idealized treatment with Gaussian noise, the significance of detection only depends on the noncentrality parameter ρ^2 , which is proportional to the observation time T_{obs} . For this reason, the T_{obs} required for a detection with a particular confidence level is inversely proportional to \mathcal{K} and ϵ^2 , or

$$T_{\text{det}} = \frac{S_h(2f)d^2}{256\pi^4\epsilon^2 I^2 f^4 \mathcal{K}} \rho_*^2 \quad (11)$$

with ρ_* being a threshold (or a sensitivity level for ρ) determined by the desired false-alarm probability (P_{FA}) and detection probability (P_{DE}), as we discuss below.

Let us follow a frequentist approach of hypothesis testing. Suppose X is our detection statistic, which is either a 4-D χ^2

distribution or a 4-D noncentral χ^2 distribution with noncentrality parameter ρ^2 . Let us first impose a detection threshold X_{th} on X , so that $P[X > X_{\text{th}} | \rho^2 = 0] = P_{\text{FA}}$, which leads to

$$\left(1 + \frac{X_{\text{th}}}{2}\right) e^{-\frac{X_{\text{th}}}{2}} = P_{\text{FA}}, \quad (12)$$

where the threshold X_{th} is determined implicitly from P_{FA} . If X now has a nonzero ρ^2 , its probability of overcoming the threshold becomes the detection probability, or

$$P_{\text{DE}} = P[X > X_{\text{th}} | \rho^2]. \quad (13)$$

The threshold ρ_*^2 is determined by requiring that when $\rho^2 \geq \rho_*^2$, Eq. (13) provides a significant P_{DE} .

III. THE DETECTION STATISTIC OF MULTIPLE PULSARS

In this section we extend the single-pulsar analysis approach of Sec. II C to apply to the detection of GWs from an ensemble of pulsars.

A. General theory

To formulate how we might detect a combination of n nearby sources, let us consider the general problem of distinguishing the distribution of n random variables, (X_1, \dots, X_n) , between two probability densities p_A and p_B . Suppose we have a region \mathcal{V} , and we claim A if $(X_1, \dots, X_n) \in \mathcal{V}$, and B otherwise. In the context of GW detection A is without signal, while B is detection. In this way, our false-alarm probability is

$$P_{\text{FA}} = \int_{\bar{\mathcal{V}}} p_A(x_1, \dots, x_n) dx_1 \dots dx_n, \quad (14)$$

where $\bar{\mathcal{V}}$ represents not being within the region \mathcal{V} , and our detection probability is

$$P_{\text{DE}} = \int_{\bar{\mathcal{V}}} p_B(x_1, \dots, x_n) dx_1 \dots dx_n. \quad (15)$$

We then have to find the region \mathcal{V} for which P_{DE} is maximized given P_{FA} . It is possible to find that the boundary of \mathcal{V} should be given by

$$\mu p_A(x_1, \dots, x_n) = p_B(x_1, \dots, x_n). \quad (16)$$

This is an implicit formula: given different values of the Lagrange multiplier μ , we arrive at regions that have particular pairs of $(P_{\text{FA}}, P_{\text{DE}})$. For each pair, the detection probability is the maximum possible value given P_{FA} . Operationally, the boundaries of all these \mathcal{V} 's are given by surfaces specified by Eq. (16). In other words, for data $X_{1, \dots, n}$, if we define the likelihood ratio

$$L = \frac{p_B(X_1, \dots, X_n)}{p_A(X_1, \dots, X_n)} \quad (17)$$

as a detection statistic, and by imposing a threshold, we obtain the best P_{DE} with given P_{FA} .

If we have various versions of B parametrized by a set of parameters θ , we can further average over these possibilities with their prior probability distributions $w(\theta)$, such that

$$P_{\text{DE}} = \int d\theta w(\theta) \int_{\mathcal{V}} p_B(x_1, \dots, x_n; \theta) dx_1 \dots dx_n. \quad (18)$$

This simply arrives at modified boundaries of \mathcal{V} given by

$$\mu p_A(x_1, \dots, x_n) = \int w(\theta) p_B(x_1, \dots, x_n; \theta) d\theta \quad (19)$$

meaning that

$$L = \frac{\int w(\theta) p_B(X_1, \dots, X_n; \theta) d\theta}{p_A(X_1, \dots, X_n)}. \quad (20)$$

This is in fact the same as the marginal likelihood ratio (the Bayes factor in a Bayesian approach) for obtaining the data $X_{1,\dots,n}$; therefore we have simply established the optimality of the Neyman-Pearson approach in our case.

B. Multiple pulsars

In the detection of multiple pulsars, let us consider A to be n independent 4-D χ^2 distributions, and B to be n independent 4-D noncentral χ^2 distributions, with non-centrality parameter $\lambda_1, \dots, \lambda_n$ (for simplicity, we use λ rather than the optimal SNR ρ^2). Recall that for a k -D noncentral χ^2 distribution, we have

$$p_{(k,\lambda)}(x) = \frac{1}{2} e^{-(x+\lambda)/2} \left(\frac{x}{\lambda}\right)^{k/4-1/2} I_{k/2-1}(\sqrt{\lambda x}), \quad x > 0, \quad (21)$$

where I is the modified Bessel function of the first kind. We can then write

$$p_A(x_1, \dots, x_n) = p_{(4,0)}(x_1) \cdots p_{(4,n)}(x_n) \quad (22)$$

and

$$p_B(x_1, \dots, x_n) = p_{(4,\lambda_1)}(x_1) \cdots p_{(4,\lambda_n)}(x_n). \quad (23)$$

Following Eq. (16), for fixed values of $\lambda_1, \dots, \lambda_n$, we have

$$\prod_{j=1}^n \frac{2e^{-\lambda_j/2} I_1(\sqrt{\lambda_j x_j})}{\sqrt{\lambda_j x_j}} = \mu \quad (24)$$

as optimal boundaries of \mathcal{V} which can also be written as

$$\sum_j \log \left[\frac{I_1(\sqrt{\lambda_j x_j})}{\sqrt{\lambda_j x_j}} \right] = \text{const.} \quad (25)$$

This shows how signals should be combined resulting in our combined detection statistic

$$L_{\text{fix}}^{\text{opt}} = \sum_j \log \left[\frac{I_1(\sqrt{\lambda_j X_j})}{\sqrt{\lambda_j X_j}} \right] \quad (26)$$

where $X_{1,\dots,n}$ are the n observables.

If each λ depends on a set parameters θ , and for each j there is a corresponding prior distribution $w_j(\theta)$, then from Eq. (19), we can write

$$L^{\text{opt}} = \sum_j \log \left[\int w_j(\theta) e^{-\lambda(\theta)/2} \frac{I_1(\sqrt{\lambda(\theta) X_j})}{\sqrt{\lambda(\theta) X_j}} d\theta \right]. \quad (27)$$

As a sanity check, if $w_j(\lambda) = \delta(\lambda - \bar{\lambda}_j)$, we recover the previous result.

C. Special case: exponential distribution

We can further simplify the construction of the optimal statistic, simply and arbitrarily assuming that each λ_j value is drawn from an exponential distribution, or

$$w_j(\lambda) = \frac{1}{\bar{\lambda}_j} e^{-\lambda/\bar{\lambda}_j}, \quad \lambda > 0, \quad (28)$$

where $\bar{\lambda}$ is the mean value of the prior distribution on λ for each pulsar. In this case, we obtain the following closed-form expression:

$$\mathcal{L}_{\text{exp}}^{\text{opt}} = \sum_j \log \left[\frac{e^{Y_j} - 1}{Y_j} \right] \quad (29)$$

with

$$Y_j = \frac{\bar{\lambda}_j}{\bar{\lambda}_j + 2} \frac{X_j}{2}. \quad (30)$$

This is quite interesting: those sources with $\bar{\lambda}_j \gg 2$ (already quite detectable individually) should be combined with a similar weight, while those much less than unity should be combined according to the expectation value of the non-centrality parameter, $\bar{\lambda}_j$. The latter case is discussed further below.

D. Special case: weak-signal limit

A different way to obtain an optimal statistic is to directly assume that we should linearly combine the \mathcal{F} statistic according to

$$\mathcal{L}_{\text{lin}}^{\text{wopt}} = \sum_j \alpha_j X_j, \quad (31)$$

and optimize the signal-to-noise ratio, which is given by the increase of $\langle \mathcal{L}_{\text{lin}}^{\text{wopt}} \rangle$ due to nonzero λ_j divided by the variance of $\mathcal{L}_{\text{lin}}^{\text{wopt}}$ in the absence of signal. This leads to

$$\begin{aligned} \alpha_j &\propto \bar{\lambda}_j \\ &= \frac{f_j^4 \langle \mathcal{K}_j \rangle}{d_j^2 \mathcal{S}_h(2f_j)}, \end{aligned} \quad (32)$$

where the second line is valid for the known pulsars case, if we assume the intrinsic parameter of pulsars follows the same distribution (see the discussion in Sec. V). This can be derived from the optimal statistic, if we assume that we are interested in the low signal amplitude limit where the λ_j are small. In this case, we can Taylor expand Eq. (27) and obtain, at leading order,

$$\mathcal{L}_{\text{lin}}^{\text{wopt}} \approx \sum_j \frac{(X_j - 2)\bar{\lambda}_j}{4} \quad (33)$$

which is equivalent to using

$$\mathcal{L}_{\text{lin}}^{\text{wopt}} \approx \sum_j \alpha_j X_j, \quad (34)$$

which is also consistent with Eq. (29) when $\bar{\lambda}_j$ is small.

This implies that if we could tolerate a high false-alarm probability by setting our threshold low, it is plausible that combining the observables proportional to the (prior) expectation value of noncentrality parameters would be optimal. However, as shown in Sec. IV, in the situations we encounter, this approximation is not quite valid.

E. Comparison with individual pulsar detection

Before we compare our strategy with existing strategies that do not combine signals from multiple pulsars, let us first clarify what it means to not combine signals. A careful examination provides two possible variants.

1. Assigning equal false-alarm probability to each pulsar

The first approach regards treating each pulsar as truly independent, and by setting the same false-alarm probability for each pulsar—even though each pulsar is not equally likely to provide detection. In this procedure, we therefore set the same threshold X^{th} for each pulsar, requiring

$$1 - P^n(X < X^{\text{th}} | \rho^2 = 0) = P_{\text{FA}} \quad (35)$$

and leading to the following total detection probability,

$$P_{\text{DE}} = 1 - \prod_{j=1}^n P(X < X^{\text{th}} | \lambda_j), \quad (36)$$

of detecting at least one pulsar within this ensemble.

2. Assigning false-alarm probability according to signal strength

This is clearly problematic since we have potentially ~ 200 pulsars—assigning the same false-alarm value to pulsars with dramatically different potential signal strength is clearly wasteful. If a different threshold is set for each pulsar, in such a way that the detection probability of an ensemble is maximum, we then require

$$\begin{aligned} \mu &= \int w_j(\boldsymbol{\theta}) \frac{P^{(4, \lambda_j)}}{P^{(4, 0)}} d\boldsymbol{\theta} \\ &= \int w_j(\boldsymbol{\theta}) \frac{e^{-\lambda(\boldsymbol{\theta})/2} I_1\left(\sqrt{\lambda(\boldsymbol{\theta}) X_j^{\text{th}}}\right)}{\sqrt{\lambda(\boldsymbol{\theta}) X_j^{\text{th}}}} d\boldsymbol{\theta} \end{aligned} \quad (37)$$

where μ is a constant independent of j . As we vary μ , we obtain a varying set of X_j^{th} that would provide us with the optimal thresholds for each X_j , such that the total detection probability of detecting a GW signal within this ensemble is maximum given the false-alarm probability [as defined in Eq. (35) with different X^{th}].

IV. MONTE CARLO SIMULATIONS OF SIMPLE MODELS

In this section, we perform numerical investigations of two simple models. In particular, we study the case of constant λ_j first, then the case where the λ_j values follow exponential distributions. This provides important basic understanding before we move on to the known pulsars.

A. Constant λ_j

In this section, we perform Monte Carlo simulations for signals with fixed λ_j , the simplest case. We compare four strategies: (i) imposing a constant threshold on all X_j (Sec. III E 1), (ii) imposing a variable threshold on X_j , according to Eq. (37), (iii) using a linear-combination statistic

$$\mathcal{L}_{\text{lin}} = \sum_{j=1}^n \alpha_j^\beta X_j \quad (38)$$

with various values of β , and (iv) using the optimal statistic, according to Eq. (26).

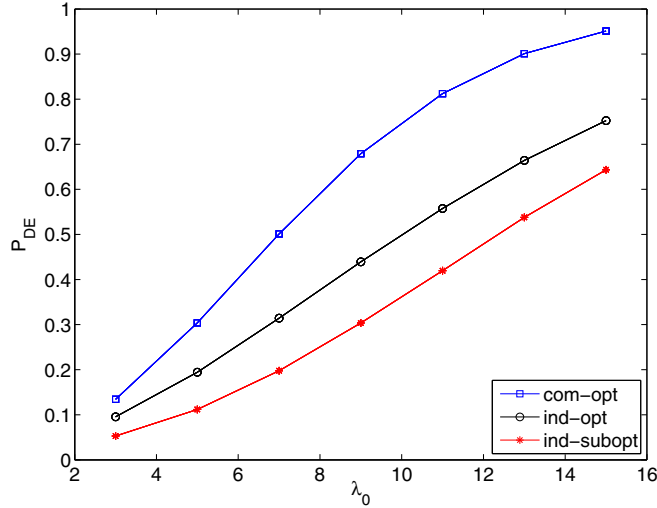


FIG. 1. Detection probability for the model given by Eq. (39), using the optimal combined statistic (com-opt) and individual thresholds (ind-subopt for equal thresholds, see Sec. III E 1, and ind-opt for optimal thresholding, see Sec. III E 2). We have fixed $P_{FA} = 0.01$.

We have chosen

$$\alpha_j = \lambda_j = \frac{\lambda_0}{j}, \quad j = 1, \dots, 8, \quad (39)$$

which is designed to simulate an ensemble of sources that are distributed on a two-dimensional plane. If, within each disk with radius r , the number of sources is proportional to r^2 , then for the N th source, its distance should be $\sim\sqrt{N}$; therefore the noncentrality parameter should be $\sim 1/N$.

As we vary λ_0 from 3 to 15, and fixing $P_{FA} = 0.01$, we compare the detection probability. As is shown by Fig. 1, the optimal strategy is substantially better than strategies (i) and (ii). In particular, in order for (i) and (ii) to achieve 50% detection probability, the noncentrality parameter must be a factor of ~ 2 stronger.

In Fig. 2, we investigate the performance of the linear-combination statistics. For $P_{FA} = 0.01$, we plot P_{DE} as a function of the index β . It seems here that $\beta \sim 0.5$ performs slightly better than $\beta \sim 1$, although the optimal β value depends on λ_0 , and is located somewhere between 0.5 and 0.8.

B. Exponential distributions for λ_j

Let us now consider λ_j values that have simple exponential prior distributions for which we have analytical formulas derived in Sec. III C. This is also important because we can test whether having the correct prior information in constructing the detection statistic can significantly affect detection efficiency. In particular, while the optimal statistic seems highly dependent on the prior

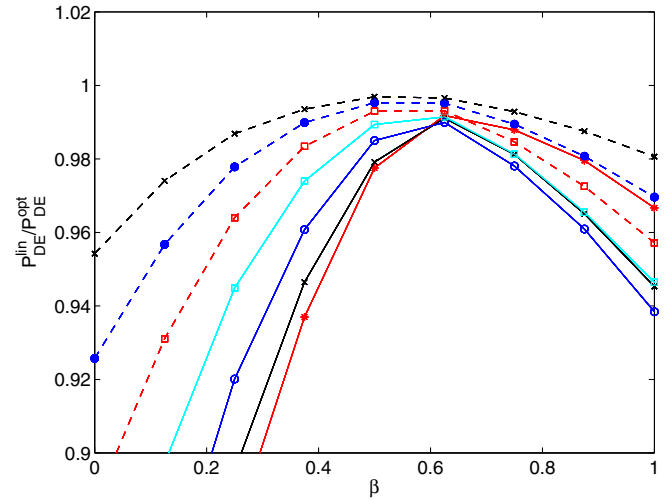


FIG. 2. Detection probability of linear combination statistics [Eq. (38)] compared to the optimal detection probability. Different traces correspond to λ_0 ranging from 3 to 10, and we have fixed $P_{FA} = 0.01$.

distribution of λ , the linear statistic \mathcal{L}_{lin} is robust against a rescaling of the distributions of all λ_j .

Again, to be concrete, we chose to have λ_j 's follow exponential distributions, with mean values given by Eq. (39). The detection probability with $P_{FA} = 0.01$ is shown in Fig. 3 for λ_0 ranging from 3 to 30. Here, we see again that the optimal statistic is substantially better than

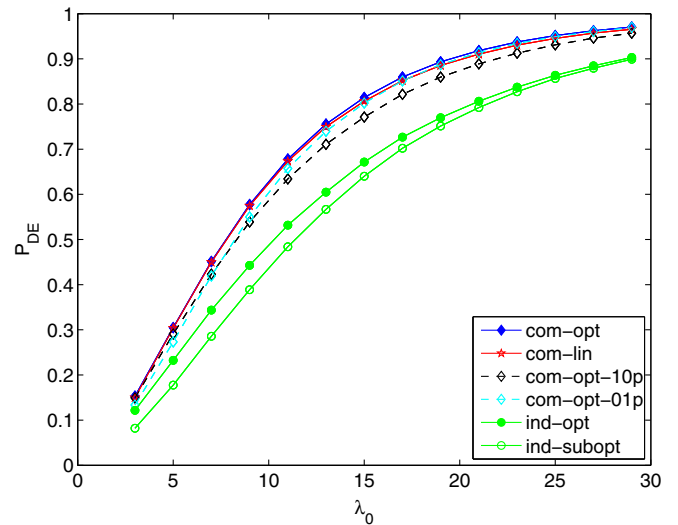


FIG. 3. Detection probability for models with λ_j following exponential distributions with mean value given by Eq. (39). We have fixed $P_{FA} = 0.01$. Shown here are from the optimal statistic (com-opt), optimal statistic scaling prior distributions by 10 (com-opt-10p) and by 1/10 (com-opt-01p), linear-combination statistic with $\beta = 1/2$ (com-lin), individual pulsar detection with optimal thresholding on each X_j (ind-opt) and individual pulsar detection with common threshold (ind-subopt).

individually detecting the pulsars—while a more strategic thresholding allows some improvement.

In this case, we can see the potential benefits of the linear statistic: when the wrong prior distributions are used (with $\bar{\lambda}_j \rightarrow 10\bar{\lambda}_j$ and $\bar{\lambda}_j \rightarrow \bar{\lambda}_j/10$) to compute the optimal statistic, the detection efficiency drops to a level worse than using the linear statistic, which is independent of an overall rescaling of all λ_j values.

C. Scaling with the number of sources

Let us now consider how the detection probabilities of the various schemes scale with the number of sources. We do this by simply extending Eq. (39) to include a variable number of sources N .

In Fig. 4, we can see that as the number of sources increases the detection probability of the optimal and linear combination statistics with $\beta = 0.5$ also increases. The detection probability of individual pulsars using a common threshold decreases, while the individual detection with optimal thresholding also keeps increasing, but stops increasing at a relatively low number of sources. This can be explained as being due to the combined statistics' ability to incorporate weaker sources without sacrificing sensitivity.

Numerically, we can see that a substantially larger signal strength has to be present for the individual detection strategies. In addition, we emphasize that the linear-combination statistic, here shown to be very close to being

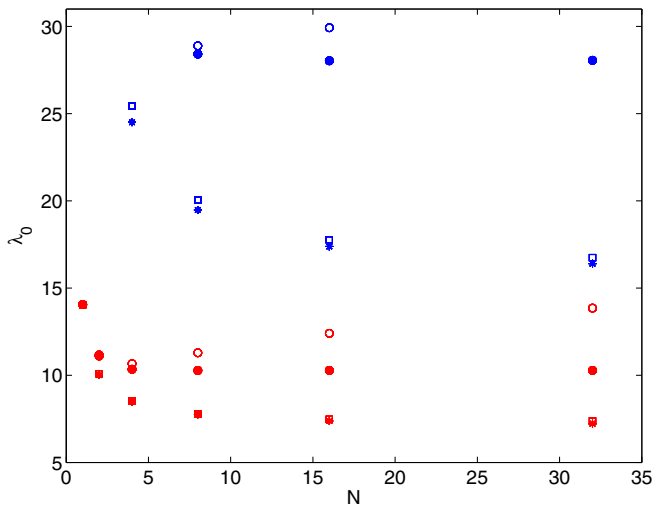


FIG. 4. Values of signal strength (λ_0) at which each detection strategy can yield detection probability of 50% (red symbols) and 90% (blue symbols), respectively, as a function of the number of sources, N . We have used the optimal statistic (stars), linear-combination statistic with $\beta = 0.5$ (square), uniform threshold for all sources (hollow circles) and optimal thresholding (solid circles). Increase of λ with N in the uniform threshold case indicates that including more sources introduces contamination from weaker sources. We assumed exponential distributions for λ_j in this plot and have fixed $P_{FA} = 0.01$.

optimal, is independent from an overall rescaling of the distribution of λ_j 's. The optimal thresholding, on the other hand, does depend on the particular model of λ .

V. MONTE CARLO SIMULATIONS FOR KNOWN PULSARS

We now discuss the case of detecting GWs from multiple known pulsars. We start by describing the known and unknown aspects of these sources, and then present the setup and conclusions of our numerical simulations.

A. Known pulsars: prior distributions for λ_j

For the case of multiple pulsars, the noncentrality parameter λ_j for each pulsar in a single detector is simply equal to ρ_j^2 , as given by Eq. (5). We now discuss in detail all factors contributing to our prior knowledge of ρ_j^2 .

The ellipticity ϵ_j crucially defines the level of quadrupole deformation of the NS. At present, we have only theoretical constraints based on the internal structure of NSs, which span a wide range, and observational upper limits from previous GW searches, which span the range $\sim 10^{-7} - 10^{-2}$ [2]. Our baseline assumption is that the ϵ of all pulsars follows a common (yet unknown) distribution; this could be motivated as arising from the belief that all these eccentricities were generated by the same physical mechanism. We note that it is plausible for Advanced LIGO to detect at the level of $\bar{\epsilon} \sim \text{few} \times 10^{-8}$.

The geometrical factor \mathcal{K} depends on the inclination angle ι , polarization angle $\tilde{\psi}$, and antenna patterns F_{++} and $F_{\times\times}$; see Sec. II B. We assume no knowledge concerning the orientation of the pulsar, therefore uniformly distributing $\cos \iota$ between -1 and $+1$, and uniformly distributing $\tilde{\psi}$ between 0 and 2π . As for F_{++} and $F_{\times\times}$, they further depend on the geographical location and orientation of the detector, as well as the source's declination angle (the right ascension dependence is averaged away after many sidereal days observation).

As noted by CS, for the network of M detectors case, the noncentrality parameter of each pulsar is simply

$$\lambda_j^{\text{net}} = \sum_i^M \lambda_{ji}, \quad (40)$$

B. Simulations and results

Our simulations assume one year of observation using the network of Advanced LIGO and Virgo at design sensitivity [16,17]. The positions and orientations of the detectors are taken from Table I of [8]. The known pulsar parameters (distance, sky location and frequency) used to compute λ_j^{net} are taken from the 195 known pulsars analyzed in the initial detection era, and we assume that the moment of inertia $I = 10^{38} \text{ kg m}^2$ [10]. To provide a proof of principle of our proposed method, we assume (i) e^2 values follow exponential distributions with two different

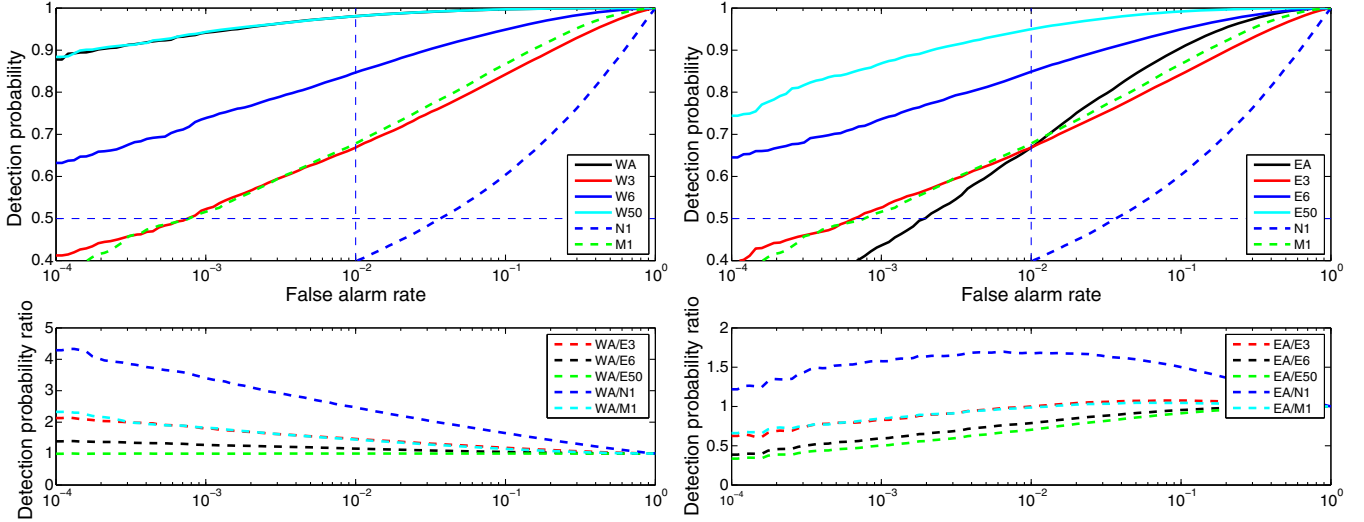


FIG. 5. ROC curves for different detection methods. WA (EA), W3 (E3), W6 (E6), W50 (E50), N1 and M1 correspond to weighted (equal) combinations of all, the expected brightest three, the expected brightest six, the expected brightest fifty, the expected brightest and the measured brightest pulsar(s) respectively. The ϵ^2 in injections were drawn from an exponential distribution with rate parameter 2×10^{-16} .

rate parameters 4×10^{-16} and 9×10^{-16} , and (ii) ϵ values follow normal distributions with mean values of 1.5×10^{-8} , and 2×10^{-8} with standard deviations equal to half of their respective mean values.

We have performed simulations to test the detection efficiency of our proposed robust statistic L_{lin} [Eq. (38); see discussion in IV] via the receiver operating characteristic (ROC) curve, which is a parametric plot of the probability of false alarm versus the probability of detection.

The ROC curve is constructed using 10^5 simulations of X_j with noncentrality parameter λ_j^{net} and 10^5 noise only simulations. With the assumption of ϵ^2 following the same distribution, the α_j^{net} is used in Eq. (38) to compute every

simulated L_{lin} in place of α_j for the M detectors case, defined as

$$\alpha_j^{\text{net}} = \sum_i^M \alpha_{ji} = \sum_i^M \frac{f_j^4 \langle \mathcal{K}_j \rangle}{d_j^2 S_{hi}(2f_j)}. \quad (41)$$

We compare the detection efficiencies of ensemble based strategies including the weighted-combination ($\beta = 0.5$) and equal-combination ($\beta = 0$, the CS case) method, with the individual pulsar detection strategy including the expected brightest [the largest value of $\frac{f^4 \langle \mathcal{K} \rangle}{d^2 S_{hi}(2f)}$], measured brightest (the maximum ρ_j^2 in each simulation) case.

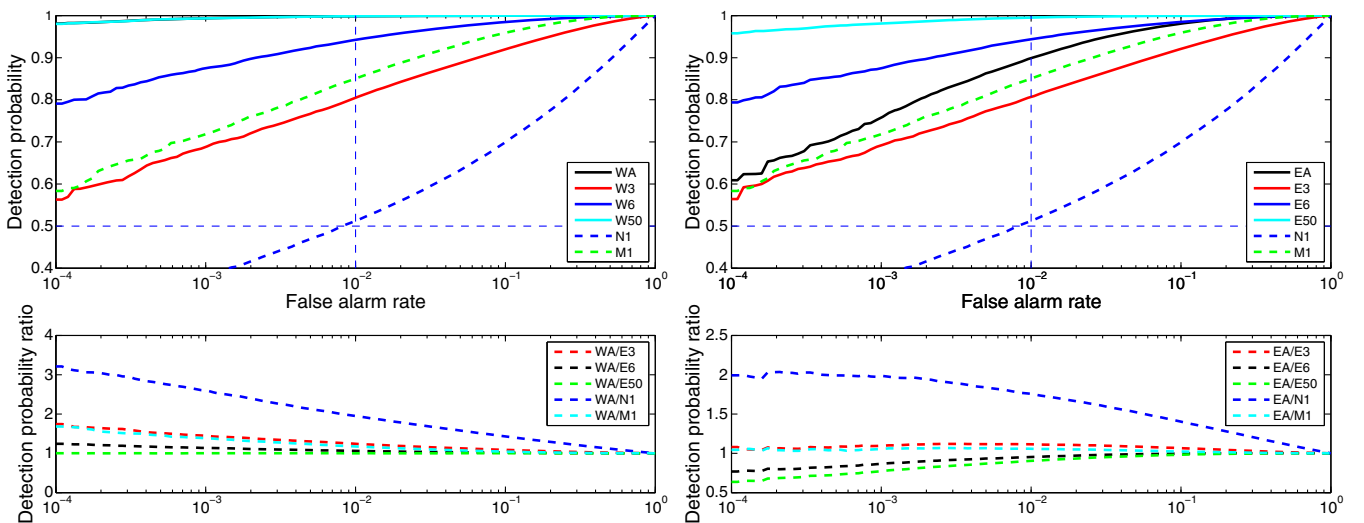


FIG. 6. Same as Fig. 5 but with ϵ following the Gaussian distribution with mean value of 1.5×10^{-8} , and half of mean values as its variance value.

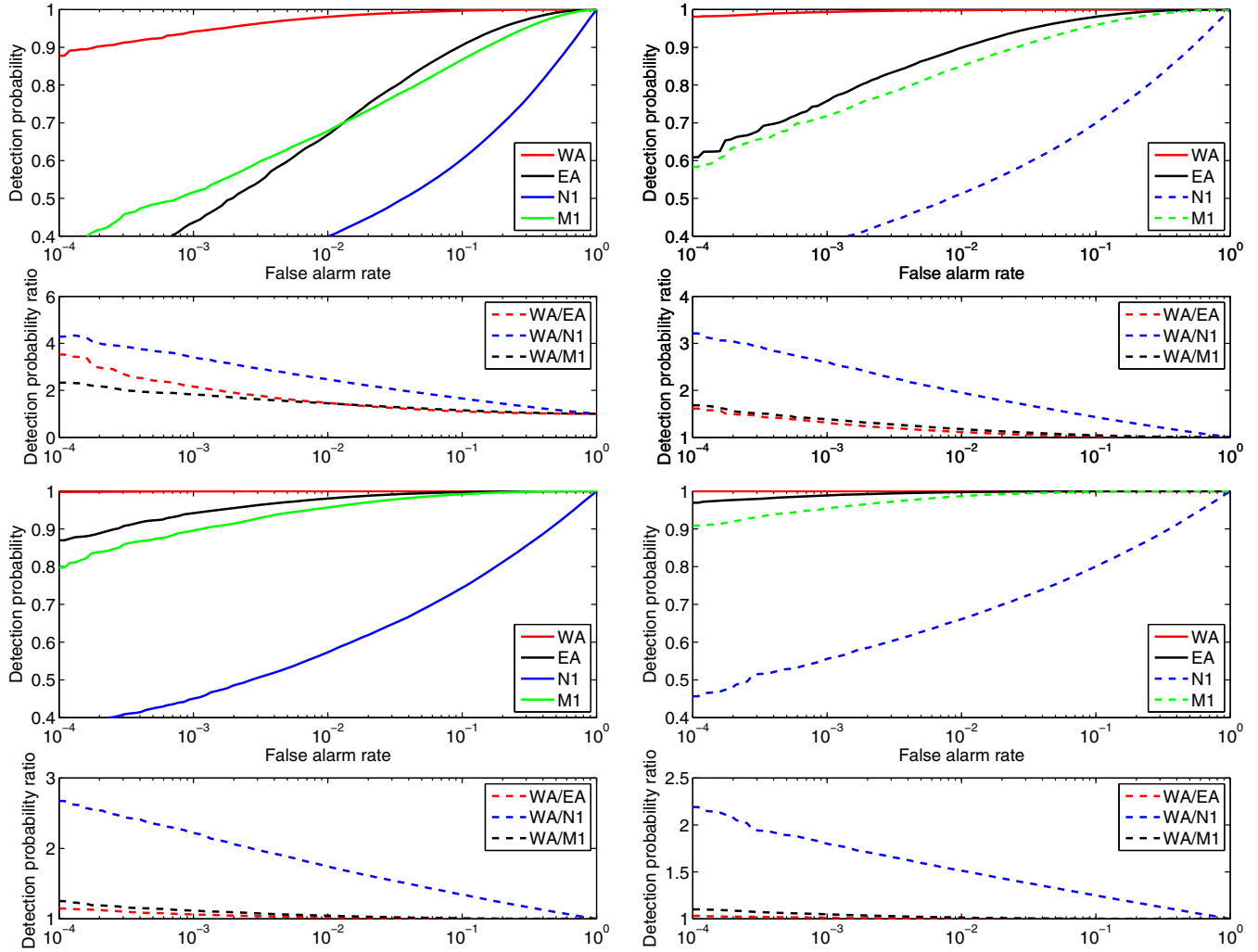


FIG. 7. ROC curves of the weighted-combination (WA), equal-combination (EA), the expected brightest (N1) and the measured brightest (M1) detection methods for 10^5 signal injection simulations and 10^5 noise only simulations. WA and EA correspond to weighted and equal combinations of all pulsars, respectively. The intrinsic parameter in injections was drawn from two distributions: (i) ϵ^2 follows an exponential distribution with rate parameter 2×10^{-16} (top-left panel) and 4×10^{-16} (bottom-left panel), and (ii) ϵ follows a Gaussian distribution with mean value of 1.5×10^{-8} (top right) and 2×10^{-8} (bottom right) with standard deviations equal to half of the mean value. Detection probability ratios (dashed lines) are shown in each panel.

The results are first presented in terms of answers to the following two questions: Will collecting more pulsars return higher P_{DE} values than an individual detection? How many sources should be combined to obtain the maximum P_{DE} at given P_{FA} ? As shown in Figs. 5 and 6 for various ellipticity distributions, the more sources are combined, the higher P_{DE} is for our proposed robust statistic L_{lin} with $\beta = 0.5$, although combining the weakest part of the population (e.g. the weakest 50 sources) will not greatly contribute to P_{DE} . As expected, we find that the P_{DE} increases when combining the first few high amplitude sources, and then decreases for the equal-weight method ($\beta = 0$) as more and more weak sources are added to the combination. These results are consistent with the simple test in Sec. IV. Since we do not know the true values of all

pulsar parameters, it is interesting to ask whether the measured brightest source or the expected brightest source would be more detectable than any other ensemble of sources.

Since the weighted-combination method is optimized for the whole population of GW signals, neither the measured brightest source nor the expected brightest one is more detectable than the whole population. This is not the case for the equal-weight combination method (see Figs. 5 and 6).

As shown in Fig. 7, our proposed weighted-combination method includes the known information of all sources and detectors; therefore combining all sources should yield a higher P_{DE} compared to other methods. In the case of $\overline{\epsilon^2} = 2 \times 10^{-16}$, given $P_{FA} = 0.0001$, the $P_{DE} \sim 0.9$ for the

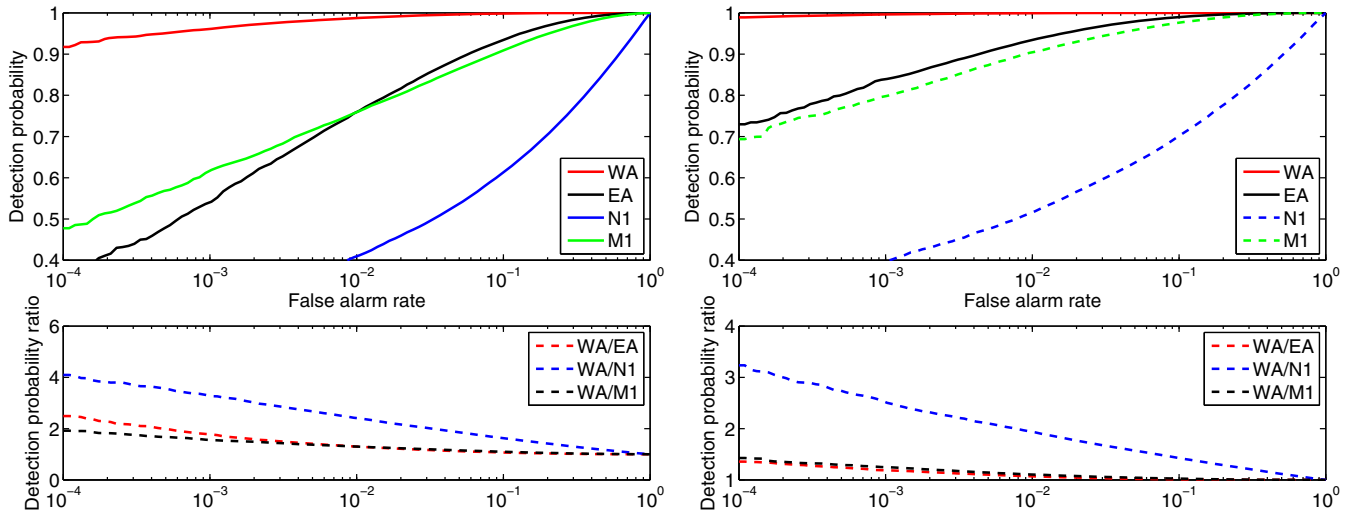


FIG. 8. The same as for the top panels of Fig. 7 but taking into account the measured distance error. The measured distance error effect is included by drawing “true” pulsar distances from a normal distribution with mean value equal to their current best estimates and a standard deviation of 20% of those mean values.

weighted-combination method is a factor of ~ 2 to 4 more sensitive than other methods (see the top-left panel of Fig. 7). The improved performance of the weighted-combination method over other methods appears to be independent of the ellipticity distribution types and distribution parameter values used in our simulations.

The typical pulsar distance measurement error is $\sim 20\%$ but could be up to a factor of 2–3 larger (e.g. [18]). To test the robustness of our proposed method, we test our sensitivity to distance uncertainty by drawing our pulsar distances from Gaussian distributions with mean values equal to the best estimated distance and with a standard deviation equal to 20% of the mean. As shown in Fig. 8, the distance uncertainties do not change the general performance of all methods: our proposed weighted-combination method still is the most efficient method and improves the P_{DE} by a factor of ~ 1.5 to 4 compared to different methods and given $P_{FA} = 0.0001$. The level of improvement decreases when the GW signals become stronger.

VI. DISCUSSION

We have proposed a novel weighted-combination detection statistic for GWs from an ensemble of known pulsars. The aim of this approach is to improve the detection efficiency of GWs over that of individual pulsar detection based on the \mathcal{F} statistic applied to single pulsars. The general argument behind the combination detection strategy is that a group of sources should be more detectable than an individual one if they share certain characteristics. We have shown that our general optimal statistic for the weighted combination of GW signals outperforms all other approaches.

We have shown that to more efficiently detect GW signals emitted from an ensemble of pulsars, each source

within the ensemble could be assigned a different detection statistic threshold based on the expected signal strength. Furthermore, by assuming that the SNRs of all sources are constant or follow exponential distributions, we have shown that the linearly weighted-combination statistic is very close to being optimal and is robust to the choice of prior SNR distributions. These analytic and simple Monte Carlo test predictions are consistent with results obtained from simulations of known pulsars.

We have also used the ROC function to determine the sensitivity of a range of possible search strategies where the detection probability between approaches is compared as a function of false-alarm probability. To demonstrate the performance of the new weighted-combination detection method for the advanced detectors era, we have compared the detection efficiency of the linearly weighted-combination method versus the equal-combination and individual detection method. We have done this by simulating GW signals emitted from the 195 known pulsars within the sensitive frequency band of Advanced LIGO and Virgo. We assume that the intrinsic pulsar parameter ellipticity ϵ^2 follows a common distribution in these simulations. The true form of the ellipticity distribution and its associated parameters are unknown. We have chosen to use both exponential and Gaussian distributions with mean values corresponding to ellipticities $\epsilon \sim 10^{-8}$, a value consistent with the initial GW era nondetection of pulsar signals and a possible advanced era detection. In general, the combination methods return better detection efficiency than a method that simply considers the closest or brightest pulsar. Being consistent with results of simple Monte Carlo tests, the most efficient method in simulations for known pulsars involves combining all known pulsars with weights $\propto \bar{\rho}$, the expected value of the optimal SNR of each pulsar. For the specific case where $\bar{\epsilon} \sim 1.5 \times 10^{-8}$, for one year

observation of the advanced detector network, we find that $P_{\text{DE}} \sim 0.95$ given $P_{\text{FA}} = 0.0001$. In this case, the improvement by our proposed combined method could be up to a factor of ~ 4 compared with other methods. These results are consistent with the case of taking into account the measurement errors of pulsar distances.

An important feature of the proposed combination method is that it is very flexible. Using the new method it is simple to include more observed pulsars or updated source information (e.g. distance or orientation parameters), without recalculating any individual detection \mathcal{F} -statistic values. However, we would expect that a fully Bayesian approach for combining all known pulsars may be more sensitive albeit at an increased computational cost.

The flagship known pulsar analysis within the GW community is a Bayesian approach [6,7,15,19]. We note that it is likely that a comprehensive Bayesian approach to combining all known pulsars into a single analysis may produce a truly optimal result. Besides all of the information discussed above, one could also consider the uncertainty of the major assumption (model) of this work: that all pulsars' ellipticity values follow a common but unknown distribution. A hierarchical Bayesian approach would allow us to naturally investigate the true priors governing the distribution. In this case the form of the prior would be

represented as a possible model and the parameters governing that distribution would be the ‘‘hype’’ parameters of that model. We could also apply Bayesian model selection to distinguish between different prior distributions e.g. exponential vs Gaussian or power law, etc. However, it is unclear how constraining such an analysis would be and we hope to tackle this problem in future studies. Beyond the detection of GWs emitted by an ensemble of pulsars, the posterior probability of all parameters could be output from a Bayesian approach. In future studies we hope to investigate such a Bayesian application to the detection of GWs from the ensemble of known pulsars.

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