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Causality deficit-inflation: Wavelet Transform

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Abstract

Few researchers have addressed the problem of the causality inflation-deficits, and even Previous work treated this causality they have only focused on the Granger causality technique in which the results of this approach raises many doubts. This investigation aim to study the relationship between inflation and the primary deficit using the wavelet transform for different euro area member countries from 1980 until 2014 with quarterly data. We have characterized the inflation-primary deficit relationship in a time-frequency scale. First we found that the deficit causes inflation in log term. We detected also, that inflation-primary deficit relationship is highly consistent during the post-financial crisis in the euro area in the long run. Finally we can judge through the wavelet transform that this relationship is non linear.

Keywords: deficit, inflation, wavelet.

1 Introduction

We bring down the curtain that, after the 2008 financial crisis, the inflation rate and debt ratio sharply increase. To that end, it will be substantive to shed light on these economic phenomena. Indeed, the primary balance is an important indicator of a State's economic situation and gives a conspicuous idea of the unbounded evolution of public debt. Whilst copious theoretical studies have shown that primary deficits are inflationary Catao & Terrones (2005), Lin & Chu (2013). Notwithstanding, to date, no empirical study started the ball rolling concerning a statistically significant link between the deficit-inflation relationship in developed countries, albeit in developing countries, De Haan & Zelhorst (1990), Metin (1998), Domaç & Yücel (2005), have found a chief causality of budget deficits on inflation in high-inflation countries. Hence the importance of our research. In this investigation our objective is to study the relationship between inflation, and the primary deficit using the wavelet transform for different euro area member countries. We have characterized the primary inflation-deficit relationship in a time-frequency scale in an attempt to disentangle four conclusions. i) find out if there is a relationship between these target variables. ii) know if this causality is a long-term causality or a short-term causality. iii) know the intention of causality, i. e. if inflation causes the deficit or reciprocally. iv) know the type of causality, i. e. linear or non-linear causality. The originality of this research is to know what inflation and the primary deficit are dependent on before or after the 2008 financial crisis? Did inflation drive the deficit or reciprocally? And are euro area member countries converging or diverging? These findings have important implications for risk management, and for monetary policies to control inflationary pressures as well as the debt ratio. In this article to tackle our problem and answer all these questions, we will start in the first section with a literature review. The second section will present the wavelet approach and then we will finish with the results and the conclusion.

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2 literature review

Manifold studies, have agitated the linkage between inflation and deficit. In their research Lin & Chu (2013) bring to pass that the budget deficit has a high strike on inflation in high inflation episodes and a low impact in low inflation episodes. The results intimate that fiscal consolidation would be the more efficacious in firming prices in the higher the inflation rate. According to Sargent et al. (1981), it is true that in a monetarist economy, the money supply is exogenous and determined independently by the central bank to control inflation, but this control is limited. Putting it differently, Assuming that the monetary authority is controlled by the fiscal authority, this leads the latter to independently announce all current and future deficits, which leads to deficits with signorage and therefore the emergence of inflation. Against this background, it can be said that the deficit and inflation are dynamically correlated. Given the past theoretical research (Catao & Terrones (2005) Blanchard et al. (1989)), until today we can not conclude that there is a study that is interested in mentoring the exact relationship between inflation and deficit in spite of Kliem et al. (2016)'s recherche where he has studied the co-movement between these two target variables and who can only show that this relationship is non-linear. In fact, Kliem et al. (2016) studied the relationship between the government deficit and inflation using American data from 1900 to 2011 using the low-frequency approach of Lucas (1980) and Sargent and Surio (2011) to assess the low-frequency relationship of these two filtered time series (filtering inflation and deficit and deficit by the filter proposed by Lucas (1980)), allowing them to focus on the low-frequency variation. They found that any increase in the primary deficit leads to an increase in inflation. In this context, the empirical studies regarding the united states (Hamburger & Zwick (1981), Dwyer Jr (1982), Darrat (1985), Ahking & Miller (1985), King & Plosser (1985)) and those in others developed countries (King & Plosser (1985), Giannaros & Kolluri (1985), Protopapadakis & Siegel (1987), Barnhart & Darrat (1988)) could not find results that explain the relationship between these two target variables. for developing countries, most empirical studies can draw a conclusion on these two variables. indeed, during the period of high inflation, they find that there are significant causal effects between inflation and the budget deficit. (Metin (1998), De Haan & Zelhorst (1990), Domaç & Yücel (2005), and Lin & Chu (2013)). Since before, several studies have used different methodology to explain the relationship betwixt inflation and deficit.

From a theoretical point of view and taking into consideration the factors that can influence inflation Kaur (2018) focused on the empirical study. Indeed, in his work he tried to increase the investigation on the causality between money supply, exchange rates, budget deficit and inflation during the period 1970-1971 to 2014-2015 in India. To do this he used two techniques. The first is Granger causality to test short-term causality and the second is the Johansen's co-integration technique to test the existence and effectiveness of several of the co-integration vectors. Kaur (2018) argued that there is a long-term relationship between budget deficit, money supply, exchange rate volatility and inflation, but he failed to provide adequate evidence to confirm the causality from budget deficit to inflation through Granger's causality tests. One of the major drawbacks to adopting this technique (Granger causality) is that fiscal theory of the price level is unable to find empirical support in the case of India. then, This is something of a pitfall if once accepts that the effectiveness of the budget deficit as an instrument for price stabilization in the short term. Vast amount of documentation on the relationship between budget deficits and inflation, using cross-sectional and panel data. Cutting edge paper of Karras (1994) tested the relationship using panel estimation. He claimed that deficits are not inflationary in 32 countries. Despite Cottarelli (1998) demonstrate a significant results concerning the impact of budget deficits on inflation in industrial economies and transition economies using the dynamic panel model.

Using the data set of 94 developed and developing countries from 1960 to 1995, Fischer et al. (2002) traces the advances in relationship between budget deficits and inflation. In their seminal, this relationship is

low in countries with low-inflation and countries with high-inflation during low-inflation episodes. However this connection is strong only in high-inflation countries during high inflation episodes. Catao & Terrones (2005) They draw our attention to the impact of deficits on inflation and this influence is stronger either in high inflation or in developing countries. this empirically investigation using the group average estimation model, showed that, developing countries in which there is political instability, limited access to external borrowing, and less effective tax collection , have a propensity to have a higher inflation tax. is known about the regimes of "monetary dominance" and "fiscal dominance" examine the connection between inflation and budget deficits which are largely treated in the well-known article of Sargent and Wallace (1981). On the first hand, as we know, the the budget deficit is generated by the issuance of bonds as well as the seignorage created by the central bank whose monetary policy is independent. On the second hand, the State representing the budgetary authority faces a budgetary constraint imposed by the central bank, which in turn represents the monetary authority . in this case both the money supply and the budget deficit can be controlled by the central bank without inflation occurring. if not, if there is budgetary domination, then in these circumstances the central bank can not control neither money supply nor budget deficits because the latter lead to inflation in this case. There is still considerable ambiguity concerning the relationship between primary deficit and inflation. in this context. several experts found that the effect of primary deficit on inflation is negligible In a major advance in 1988, Barnhart and Darrat surveyed that the deficit don't causes an increase in monetary growth and Reciprocally monetary growth don't causes deficit. Giannaros and Kolluri (1986) has already noted a consistency with Barnhart and Darrat (1988) and claimed that the impact of budget deficits on money supply and inflation is insignificant for the data from 10 industrialized countries. The set up used by previous authors can be found with Komulainen & Pirttilä (2002) how use data from Russia, Bulgaria and Romania and find that the deficit does not play an inflationary role. However, Metin (1998) concludes that deficits causes directly inflation in Turkey during the period 1954-1986. Using the co-integration analysis. De Haan and Zelhorst (1990) underlines that there is no no evidence to affirm "the fiscal dominance hypothesis" and concluded that the correlation betwixt inflation and deficits are only in the high inflation periods. Using the coi-ntegration analysis, Also Loungani and Swagel (2003) comes to the conclusion that the fiscal balance is weakly correlated with inflation in 53 developing countries between 1964 and 1998, and in a high average inflation the correlation becomes stronger. add to that, they find a non-linear relationship between these tow variables, and when the deficit-to-GDP ratio is above 5 % the effect of deficits on inflation is significant. In the same reflection, that budget deficit has a significant impact on inflation in high inflation periods, the investigation of DomaŃğ and YŃijcel (2005) intervenes using a poled probit estimations in 15 emerging markets from 1980 to 2001.

3 Methodology : wavelet approach

The method that will be presented throughout this section is abundant in the field of electrical engineering to filter signals. The idea is to borrow all these techniques from the macroeconomic domain in order to study the behaviour of a non-stationary process in order to find results explaining whether there is a causality between the macroeconomic aggregates in question not only in a time domain but rather in a time-frequency-scale domain. To do this. Indeed, the choice of this technique is explained by economic intuition, whose knowledge of the behaviour of the series from a time-frequency point of view provides information on the nature of instabilities and the velocity of the bi-variance between inflation and the deficit.

In a multivariate model, to study the relationship between the primary deficit and inflation, it is first necessary to study the causality between the target variables. Indeed, one of the best ideas in the world of research is the conceptualization that a variable could have predictive information about the behaviour of another variable

Granger (1969). This notion, often referred to as the Granger causality, has been applied in many fields of study, such as finance and economics, signal processing, neuroscience, image processing, geophysics, etc. Despite Granger's causality has succeeded in affirming whether there is really a relationship between the variables, it has limitations, first of all, regarding the causality of variance, the causality of risk, the causality of quantiles, the causality of the mean as well as the distribution. In addition to these limitations, this Granger causality method is unable to distinguish between long and short-term causal effects. Once causal effects are recorded and based on Granger causality, location in the frequency domain regardless of whether it is low (short term) or high frequency (long term) will be difficult. However, the causal relationship occurs over different horizons, which was developed in Geweke's study (1982) to help us understand this difficulty. Indeed, Geweke (1982) innovative idea was to transform causality into a frequency domain. In this context Hosoya (2001), Breitung & Candelon (2006) and Yao & Hosoya (2000) align with Geweke (1982). At this stage the research studies have found that the frequency representation of the so-called "Fourier transform" makes it possible to distinguish long-term causal effects from short-term causal effects, however, the fact that causality cannot be followed over time has caused limitations regarding "Granger-Geweke causality", this has generated other innovative ideas such as "continuous wavelet transforms" to analyse time-frequency causality links Dhamala et al. (2008). Wavelet analysis has become widely used in empirical studies to understand the relationship between variables and more precisely the temporal fluctuations between them over different horizons. The ideas of Crowley (2007), Yogo (2008), Dhamala et al (2008), Gallegati et al. (2011), Gençay et al. (2001), Fan & Gençay (2010) and Gallegati et al (2011) support this approach. However, what is more interesting in this chapter are the causal effects between the budget deficit and inflation. Indeed, the originality of this investigation is reflected in the fact that it records how the causal effects between the deficit and inflation change over time and over time. Let's look at the Fourier transform (frequency representation), one of its advantages is that not only can long-term causal effects be isolated from short-term causal effects, but also it allows us to see if it is the short or long term that modulates the correlation between the variables. Unfortunately, the measurement of causal effects using the continuous wavelet transform was particularly problematic because such a measure only encompasses the amplitude between variables while the direction information needed to collect causal relationships is not available (see diagram). However, useful information about the delay relationship (i.e. who starts first and causes who) is coded in the phase difference (or even phase circle). In few studies that are likely to undertake non-parametric wavelet causality (see Dhamala et al. 2008), the problem lies in calculating the spectral matrix factors in order to deduce the minimum phase. This process involves an inverse Fourier to communicate between the time domain and the frequency domain. At this stage, Owolabi et al. (2014) his study proposed an alternative method that abandons spectral matrix factorization but instead uses the information contained in both amplitude and phase difference to derive causal information between the variables of interest encoded in the phase difference. He authenticated the results based on this method by developing a data generation process (DGP) and found that the method is able to identify the causal period in synthetic data.

3.1 Causality of Geweke-Granger

To illustrate the two-variable (x and y) p-order VAR model is assumed as follows (Breitung and Candelon (2006), Dhamala et al (2008) or Olayeni (2015)):

$$\boxed{\Lambda(L)X_t = \epsilon_t} \quad (1)$$

With, $\Lambda(L)$, X_t and ϵ_t denote respectively, the delay polynomial, the endogenous variables of target variables and the variance-covariance matrix error term noted Ω which represents the mysterious part of the model. The matrix writing of the previous equation is given by :

$$\begin{bmatrix} I - \Lambda_{xx}(L) & \Lambda_{yx}(L) \\ \Lambda_{xy}(L) & I - \Lambda_{yy}(L) \end{bmatrix} \times \begin{bmatrix} x_t \\ y_t \end{bmatrix} = \begin{bmatrix} \epsilon_t^x \\ \epsilon_t^y \end{bmatrix} \quad (2)$$

$$\Omega = \begin{bmatrix} \Omega_{xx} & \Omega_{yx} \\ \Omega_{xy} & \Omega_{yy} \end{bmatrix} \quad (3)$$

knowing that $\Lambda(L)$ is invertible, then we note its inverse by $\Theta(L)$ then we have :

$$X_t = \Theta(L)\epsilon_t \quad (4)$$

The Fourier transform of the previous equation gives us:

$$X(\omega) = \Theta(\omega)\Omega\tilde{\Theta}(\omega) \quad (5)$$

With, $X(\omega)$ denotes the spectral power of variable X in ω (it denotes the frequency); $\Theta(\omega)$ denotes the transfer function between variables x and y and $\tilde{\Theta}(\omega)$ the conjugate (complex part) of $\Theta(\omega)$.

$$\Theta(\omega) = \begin{bmatrix} \Theta_{xx}(\omega) & \tilde{\Theta}_{yx}(\omega) \\ \Theta_{yx}(\omega) & \Theta_{yy}(\omega) \end{bmatrix} \quad (6)$$

Granger-Geweke causality in the frequency domain is given by:

$$G_{y \rightarrow x}(\omega) = \log \left[\frac{X_{xx}(\omega)}{X_{xx}(\omega) - \{\Omega_{xx} - \Omega_{xy}^2/\Omega_{xx}\}|\Theta_{xy}(\omega)|^2} \right] \quad (7)$$

With, the numerator represents the spectral power of the variable x at the frequency ω and the denominator presents the total power minus the causal contribution which represents the intrinsic power. Causality in the frequency domain allows us to reduce the time dimension to a single time point, resulting in a loss of information on time variation. However, the causality content in the frequency domain is interesting insofar as it informs us about fluctuations in causality in the time domain by decomposing the variables into different scales (frequency). This decomposition uses the transform into a Discrete Wavelet. This approach does not indicate the temporal content of causal effects. The extension of Granger-Geweke causality to non-parametric time-frequency domain modeling, as well as the analysis of the power distribution of Granger causality requires the factorized spectral matrix. The factorization of the spectral matrix is obtained through the use of Wilson's algorithm (Wilson 1972). The necessary condition to factorize a spectral matrix is given by :

$$\int_{-\pi}^{+\pi} \log \det X(\omega) d\omega > -\infty \quad (8)$$

If we note Γ the minimum phase of the spectral density factor matrix phase :

$$\Gamma = \sum_{t=0}^{\infty} \Lambda_t \exp(-2i\pi\omega t) \quad (9)$$

Thus, if we note by Ω and Θ respectively the noise variance-covariance matrix and the minimum phase of the spectral transfer function :

$$\begin{cases} \Omega = \Lambda_0 \Lambda_0^T \\ \Theta = \Gamma \Lambda_0^{-1} \end{cases} \quad (10)$$

Then, the minimum phase of the spectral density factor matrix phase can be rewritten as follows:

$$\Gamma = \Theta \Sigma \Theta^H \quad (11)$$

With, Θ^H transposed it from the Hermitian matrix Θ . The factorization of the matrix $X(\omega)$ is given by :

$$X = \Gamma \tilde{\Gamma} \quad (12)$$

With, $\tilde{\Gamma}$ refers to the complex conjugate of the transposed Γ .

3.2 Causality in the Geweke-Granger sense proposed by Dhamala et al (2008)

The calculation of the spectral matrix and the minimum phase transfer function allow us to evaluate the Geweke-Granger time-frequency causality presented by equation (13) However, the calculation of the transfer function requires the determination of the spectral matrix and the convergence of the Wilson algorithm “La factorisation des densit s spectrales matricielles” (n.d.) which have disadvantages of this approach. These disadvantages are addressed by Dhamala et al (2008), who introduced the Wavelet transformation to the Geweke-Granger analysis. The new causality formula is parallel to the formula presented in the equation (7) in the following form:

$$G_{y \rightarrow x}(s, \tau) = \log \left[\frac{W_{xx}(s, \tau)}{W_{xx}(\omega, \tau) - \{\Omega_{xx} - \Omega_{xy}^2 / \Omega_{xx}\} |\Theta_{xy}(s, \tau)|^2} \right] \quad (13)$$

With, $W_{xx}(\omega, \tau)$ indicates the spectral power in Wavelet.

3.2.1 wavelet transform

The wavelet is a function defined on $(L^2(\mathbb{R}))$, noted $(\psi(t))$. It must check the following analytical properties such as mean and integral equal to zero and normalized :

$$\begin{cases} \int_{-\infty}^{+\infty} \psi(t) dt = 0 \\ \int_{-\infty}^{+\infty} |\psi(t)|^2 dt < \infty \end{cases} \quad (14)$$

In addition to the continuity that has been verified in the properties explained above, the wavelet must also satisfy the following admissibility status:

$$\boxed{\int_{-\infty}^{+\infty} \frac{|\Psi(\omega)|^2}{|\omega|} d\omega < \infty} \quad (15)$$

where, $(\Psi(\omega))$ is considered by the Fourier transform of $(\psi(t))$.

The projection of a series $(x(t))$ through a mother wavelet function is given by the following complex coefficients:

$$\boxed{W(s, \tau) = \int_{-\infty}^{+\infty} x(t) \psi_{s,\tau}^*(t) dt} \quad (16)$$

with,

$$\boxed{\psi_{s,\tau}^*(t) = \frac{1}{\sqrt{|s|}} \psi\left(\frac{t-\tau}{s}\right)} \quad (17)$$

Where $(x(t))$ and $(\psi_{s,\tau}^*(t))$ denote, respectively, a series and the atom of the mother wavelet transform. We note by (s) a scale parameter (dilation) ($s \in \mathbb{R}_+^*$) and by (τ) a time location parameter (translation) ($\tau \in \mathbb{R}$). Translation and expansion allow us to determine the atoms of the wavelet. Dilation is a time shift, while translation is a time location, Translation and expansion allow us to determine the atoms of the wavelet. Dilation is a time shift, while translation is a time location Torrence & Compo (1998) have shown that it is possible to reconstruct the signal through the following formula:

$$\boxed{x(t) = \frac{1}{C} \int_{s=-\infty}^{+\infty} \int_{\tau=-\infty}^{+\infty} \frac{1}{|s|^2} W(s, \tau) \psi_{s,\tau}(t) ds d\tau} \quad (18)$$

3.3 The Geweke-Granger causality proposed by Rua (2013) and Olayeni (2015)

Rua (2013) proposed a measure of the correlation by the Continuous Wavelet Transform noted $(\rho_{xy}(s,\tau))$ ¹ given by:

$$\boxed{\rho_{xy}(s, \tau) = \frac{\gamma\{s^{-1}|\Re W_{xy}^m(s, \tau)|\}}{\gamma\{s^{-1}\sqrt{|W_x^m(s, \tau)|^2}\} \gamma\{s^{-1}\sqrt{|W_y^m(s, \tau)|^2}\}}}$$

with γ designates a time-scale smoothing operator. The correlation by the Continuous Wavelet Transform differs from the Wavelet Coherence noted $(R_{xy}(s, \tau))$ ², which is given by

$$\boxed{R_{xy}(s, \tau) = \frac{\gamma\{s^{-1}|W_{xy}^m(s, \tau)|\}}{\gamma\{s^{-1}\sqrt{|W_x^m(s, \tau)|^2}\} \gamma\{s^{-1}\sqrt{|W_y^m(s, \tau)|^2}\}}}$$

However, the analysis proposed by Rua (2013) shows these limitations insofar as it does not integrate information on the direction between the variables. Olayeni (2015) proposed a modification on the correlation of Rua (2013) by integrating the concept of phase difference between the variables. He proposed an indicator function (see equation.....) that takes the value one if the variables are in phase and the value zero if not while basing

¹ $-1 < \rho_{xy}(s,\tau) < 1$.

² $0 < R_{xy}(s, \tau) < 1$.

their analysis on the phase difference circle proposed by Aguirra-Contraria and Soares (2011) (see figure 1).

The concept of phase difference between two variables x and y is given by :

$$\boxed{\Phi_{xy}(s, \tau) = \Phi_x(s, \tau) - \Phi_y(s, \tau)} \quad (21)$$

with, $-\pi \leq \Phi_{xy} \leq \pi$.

The interval $[-\pi, \pi]$ can be divided into four intervals, as indicated (see Table 1 figure ??):

- If $\Phi_{xy} \in [0, \pi/2], [-\pi/2, 0]$: the two variables are in phase, they are moving in the same direction (direction). We say that x leads to y , i. e. there is predictable information on x in the sense of Granger when $\Phi_{xy} \in [0, \pi/2]$ and vice versa.
- If $\Phi_{xy} \in [\pi/2, \pi], [-\pi, -\pi/2]$: the two variables are anti-phase, they move in the opposite direction. It is said that y leads to x , i. e. x has predictable information about y when $\Phi_{xy} \in [\pi/2, \pi]$ and vice versa.

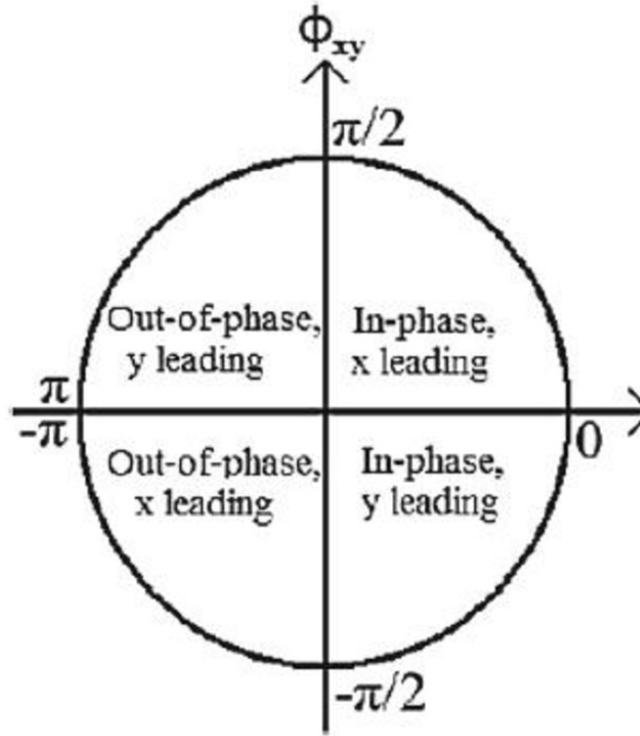


Figure 1: Phase difference circle Contraria & Soares (2011)

Table 1: The lead-lag relationship

	x leads y	y leads x
In-phase	$\Phi_{xy}(s, \tau) \in (0, \pi/2)$	$\Phi_{xy}(s, \tau) \in (-\pi/2, 0)$
Out-of-phase	$\Phi_{xy}(s, \tau) \in (-\pi, -\pi/2)$	$\Phi_{xy}(s, \tau) \in (\pi/2, \pi)$
Total phase	$\Phi_{xy}(s, \tau) \in (0, \pi/2) \cup (-\pi, -\pi/2)$	$\Phi_{xy}(s, \tau) \in (-\pi/2, 0) \cup (\pi/2, \pi)$

$$\boxed{I_{y \rightarrow x}(s, \tau) = \begin{cases} 1, & \text{si } \Phi_{xy}(s, \tau) \in J \\ 0, & \text{si non} \end{cases}} \quad (22)$$

With, $\Phi_{xy}(s, \tau)$ denote the phase difference function :

$$\Phi_{xy}(s, \tau) = \tan^{-1} \left(\frac{\Im \{ W_{xy}^m(s, \tau) \}}{\Re \{ W_{xy}^m(s, \tau) \}} \right) \quad (23)$$

And J denotes intervals : $[0, \frac{\pi}{2}]$, $[-\pi, -\frac{\pi}{2}]$, and $[0, \frac{\pi}{2}] \cup [-\pi, -\frac{\pi}{2}]$,

4 Data

We have built a quarterly database associated with the different macroeconomic variables. The various data are collected from data Stream. Indeed, this database of quarterly variables covers the period from 1980Q1 to 2014Q4 in the euro area. In our investigation we selected 10 countries from the area in addition to the euro area as classified into two categories: core countries (Germany, France, Austria, the Netherlands, Belgium, and Finland), and peripheral countries (Greece, Ireland, Spain, Italy). In our estimate we considered three main variables such as inflation, primary deficit and economic growth in order to study the relationship between inflation and deficit while considering the economic growth of the area. Based on the article by ? , inflation (π) is measured as the first annual differences in the logarithmic GDP deflator, where data are taken from the Oxford Economics database (Price Deflator, Gross Domestic Product, Index, 2010 =100). Oxford Economics defines inflation as a measure that tracks the price of all goods and services in a net of tax recovery. This measure differs from the consumer price index, which is taxed on goods and services purchased by households, including indirect taxes. Regarding the primary deficit (d) it was measured by the variable ' general government primary balance, euro ' from the Oxford Economics database in Billions of euros at constant price which was defined by central and local government revenues less expenditure, excluding the payment of interest on consolidated government liabilities expressed in local currency. If the balance is negative, then the primary deficit is negative.

Table 2: Descriptive Statistics

		GRM	FRA	ITA	SPA	NTH	FIN	BEL	IRL	GRE	AUS	ZOE
*Min	Inflation	-0,71	-0,30	-0,03	-0,09	-2,52	-0,68	-0,18	-6,99	-3,78	0,08	0,59
	Primary deficit	-31,37	-24,53	-6,43	-22,25	-6,73	-0,87	-2,72	-13,80	-31,37	-2,28	-94,05
*Max	Inflation	6,22	13,57	20,21	14,21	7,22	12,75	8,14	18,49	29,33	13,63	11,46
	Primary deficit	19,74	5,00	14,84	12,63	5,15	3,34	3,35	2,04	19,74	1,16	69,98
*Moyenne	Inflation	1,90	2,93	5,17	5,03	2,00	3,44	2,73	3,88	9,64	3,00	3,51
	Primary deficit	0,33	-4,27	2,34	-0,85	-0,09	0,90	0,57	-0,28	0,33	-0,17	-4,86
*Variance	Inflation	2,10	9,08	21,00	13,16	2,28	8,88	2,60	23,57	67,53	7,50	7,00
	Primary deficit	83,35	41,78	39,90	70,67	7,02	1,34	3,16	6,65	83,35	0,49	1215,83
*Ecart-type	Inflation	1,45	3,01	4,58	3,63	1,51	2,98	1,61	4,85	8,22	2,74	2,64
	Primary deficit	9,13	6,46	6,32	8,41	2,65	1,16	1,78	2,58	9,13	0,70	34,87
*C.V	Inflation	0,76	1,03	0,89	0,72	0,76	0,87	0,59	1,25	0,85	0,91	0,75
	Primary deficit	27,98	-1,51	2,70	-9,84	-28,54	1,28	3,11	-9,11	27,98	-4,09	-7,18

³ CV (ariation Coefficient v= Moyenne/Ecart-type)

5 Results

In our study we will focus on the types of causality (linear or non-linear causality) for this reason we will interpret coherence by transforming it into a wavelet. When ($R_{xy}(s, \tau) = 1$), therefore we can conclude that causality is linear. Before starting to interpret our results it is essential to do the scalogram decoding: First the

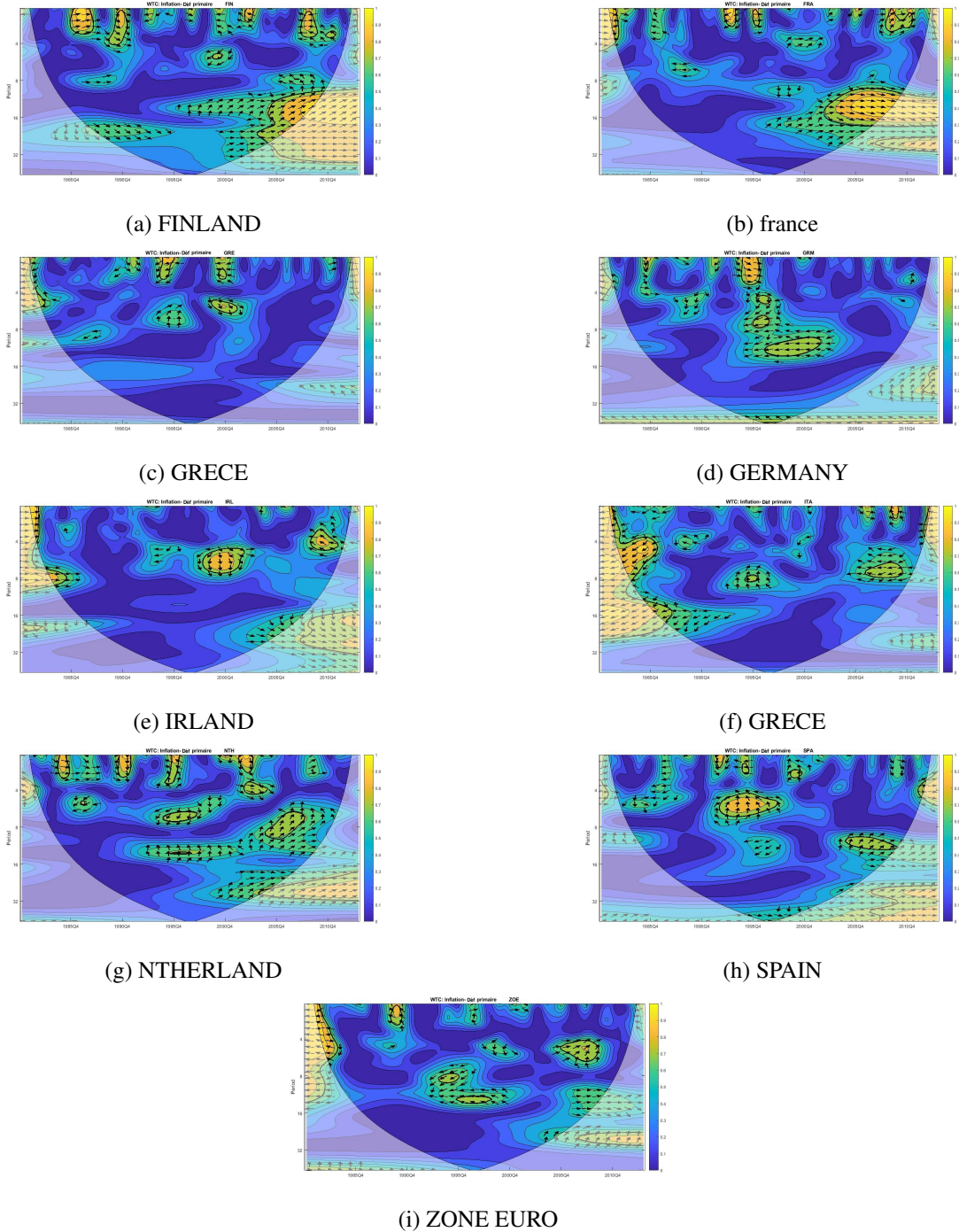


Figure 2: Wavelet Transform Coherence

olor (shows type of causality).The degradation of the blue colour means that the causality between inflation and the primary deficit is non-linear ($R_{xy}(s, \tau) = 0$) and that the dependence between these two variables is low. The degradation of the yellow colour means that the causality between inflation and the primary deficit is linear ($R_{xy}(s, \tau) = 1$) and that the dependence between these two variables is high. Second the Arrows (shows the direction of causality).The arrows to the right indicate that inflation is the cause of the primary deficit. The arrows to the left indicate that the primary deficit causes inflation.The upward arrows indicate an antiphase between the variables and inflation ahead (leading indicator) of the primary deficit. The arrows at the bottom indicate an antiphase between the variables and lagging inflation (lagging indicator) relative to the primary deficit. And finally Y-axis (quarter period). [0.4] is the short-term. [4.16] is the medium term. [16.32] is the long-term.

Austria : In the short term: it can be seen that there is a high dependence between inflation and the deficit (2004-2005). This causal relationship is linear during this same period. Indeed, this year we see that the primary deficit caused inflation in the short term. However, inflation-deficit consistency is low and non-linear between (1981-2003) and (2006-2014). Long-term: throughout the period it has been noted that there is a low dependence between inflation and deficit and throughout the same period have a non-linear relationship.

Germany : It should be noted that the Berlin Wall crisis (1991) had a negative impact on the country's economic situation, particularly in terms of inflation. These effects spread until (1995-2000). There is a strong interdependence in the scalogram during this short term to medium term period. In the long term, it can be seen that before the 2008 financial crisis, there was little consistency between the deficit and inflation. This consistency is accentuated from 2008 onwards, when it can be seen that the two variables in question are in antiphase (arrow upwards, so inflation is ahead of the deficit). Concerning the type of causality, we note that in LT the causality between inflation and deficit, until 2011, is non-linear and then becomes linear.

France : In the short term: The wavelet transformed coherence scalogram shows that the inflation-primary deficit relationship is highly consistent over the periods (1980-1981), (1993-1994), (2009-2011). Economically speaking, there is a dependency relationship between the variables, i.e. inflation drives the primary deficit in France during this short term period. In the long term: it can be seen that there is no interdependence between inflation and the primary deficit before the financial crisis (blue colour). The blue color also shows us that the causality between these two variables is non-linear in this period and then becomes linear beyond 2008. After the financial crisis, a strong coherence between the variables in question appears (yellow colour). According to our results, we can conclude that in the long term inflation leads to the primary deficit after the 2008 crisis in France.

Ireland : In the short term: there is a dependency relationship between inflation and the primary deficit during the period 1980- 1982. Indeed, according to the scalogram, it can be concluded that during this period inflation led to the primary deficit in Ireland. This causality is therefore of a linear type (yellow colour). This interdependence during this period is explained by the Irish hunger strike in 1981. However, beyond 1983 there is little coherence between these two variables (a non-linear relationship). In the long term: from the 2008 financial crisis onwards, the coherence between inflation and the primary deficit is beginning to become increasingly strong. This causality is of a linear type. Whereas before 2008 the causality was non-linear. Indeed, according to the chart, inflation generates the primary deficit at both CT and LT during the period (1981-1982) and (2008-2014) respectively. This strong coherence is expressed by the subprime crisis and then by the 2010 eurozone debt crisis (the colour has become increasingly warm).

Greece : In the long term: throughout the period up to the financial crisis, there is little consistency between inflation and the deficit (blue colour). This coherence this coherence this coherence is accentuated only slightly after the financial crisis and from 2010 the coherence has become slightly strong (green color). Economically speaking, these results can be interpreted as follows: from the financial crisis onwards, the deficit has a slight influence on inflation (right arrow with green colour). But from the Greek debt crisis of 2010 the influence of the primary deficit on inflation becomes more noticeable. The causal relationship between these two variables is non-linear from 1981 to 2008.

Spain : In the long term, Spain's result does not differ too much from the other results of the other euro zone member countries, where it can be seen that there is an influence between inflation and the deficit after the

financial crisis. It can therefore be concluded that inflation leads the deficit in the long term after the subprime crisis

6 Conclusion

many studies can be found in the existing literature on the relationship between the deficit and inflation, the exact consistency between these two variables remains undetermined. The analysis is based on the wavelet transform approach through a rich European database from 1980 to 2014. This study contributes to the literature by exploring the so-called time-frequency-scale-relationship. indeed, on the horizon it has been concluded that the relationship between the deficit and inflation continues to strengthen after the 2008 financial crisis. So if the European central bank wants to control inflation it should give full importance to the primary deficit.

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