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Abstract

This paper considers a model of household demand for water in a theoretical framework consistent with funtamendal principles of comsumer behaviour. It applies this model to individual household data to estimate the price and income elasticities of residential demand for water in Cyprus and evaluate the welfare effects associated with changes in the water pricing system. We &nd that the current regionally heterogeneous increasing block pricing system in the island introduces gross price distortions that are not justi&ed either on efficiency or equity grounds. A shift towards uniform marginal cost pricing will eliminate the deadweight loss of the current system. However, its bene&ts will be distributed in favour of the better off households. Overall, price can be an effective tool for residential water demand management, however, it may also lead to socially undesirable distributional effects on households.

JEL Classi&cation:

Keywords: Residential water demand, increasing block pricing, consumer welfare, water management.

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1 Introduction

The use of price as a water consumption management tool has been an issue of growing concern among private and public utilities in Europe and the United States. Water utilities often choose among three types of pricing schemes, uniform, decreasing and increasing block rates (or some combination of them) to in! uence water use. Economists have attempted to shed some light on the consequences of this choice by paying attention to demand estimation. However, opinions concerning the appropriate methodology for estimating water demand models differ. This paper considers demand for water in the context of a theoretical framework consistent with fundamental principles of consumer behaviour. It then applies this model to individual household data drawn from the Cyprus Family Expenditure Survey (1996-97) to estimate the price and income elasticities of residential demand for water and evaluate the welfare effects associated with potential changes in the current water pricing system.

Given that consumers in Cyprus are metered, there are strong efficiency arguments for prices to re! ect the (long-run) marginal social cost of water scarcity. At the moment water pricing in Cyprus follows an increasing block structure. In general, increasing block (progressive) tariffs are becoming more common in developed as well as developing economies. The rational for the popularity of this pricing system relates to the perception that it can be used as a tool for social justice and conservation of a scarce natural resource. Strictly speaking, there are no obvious efficiency arguments for an increasing block water tariff structure and the evidence in favour of the argument that this pricing structure has a psychological effect helping water demand management is ambiguous: OECD (1987) reports evidence in favour of this argument in Japan, Italy, Denmark and Switzerland (Zurich) while other studies (e.g. the UK metering trials in the Thames and Yorkshire water authorities) prove inconclusive.

Consumer theory provides a useful and convenient framework for residential water demand analysis and for the investigation of the efficiency and distributional effects of alternative water pricing systems. It is, therefore, not surprising that so many investigators use consumer demand analysis tools to estimate the effects of price on water demand, most mostly in the United States (Billings 1982, Schefter and David 1985, Chicoine and Ramamurthy 1986, Nieswiadomy and Molina 1989, and Renwick and Archibald 1998, among others). Studies that use European data include Hanke and de

Maré (1982), Hansen (1996), Höglund (1999), Nauges and Thomas (2000) and Martinez-Espiñeira (2000). A summary of the findings of some of these studies is provided by Herrington (1987). Overall the existing empirical results suggest that the price elasticity of demand for water is significantly different from zero and lie somewhere under -0.3.

Earlier studies of water demand ignore the peculiar features of alternative water pricing policies, such as the presence of block rates, and perform empirical estimation using ex post-calculated average prices (Gottlieb 1963; Young 1973; Foster and Beattie 1979, 1981a, 1981b). Taylor (1975), studying electricity demand, suggests that under a block pricing scheme the explanatory variables should include marginal and average price. Subsequently, Nordin (1976) demonstrated that Taylor s specification should be modified to include a "difference" variable represents the income effect imposed by the tariff structure by accounting for the effects of fixed and intra-marginal rates. In the case of multiple tariffs (and those cases where fixed quota and/or a free allowance is used) the difference variable is the difference between the total bill and what the user would have paid if all units were charged at the marginal price. More recently, investigators combine marginal price and Nordin s difference variable in empirical models of residential water demand.[1]

In this paper we follow a different approach to modelling residential demand for water. The novelty of our approach is mainly in terms of exploiting the cost and indirect utility functions underlying the consumer theory framework to derive a most general (rank-3) integrable demand model (Lewbel 1991). This enable us to obtain empirical results that conform to the fundamentals of consumer theory (such as adding-up, price homogeneity and symmetry) and have meaningful behavioural and welfare interpretation. We consider the ability to evaluate the welfare implications of alternative water pricing policies particularly important, given the significance attached to equity and the strong political objections to water price reform based on political economy arguments (Dinar 2000). In the empirical analysis we treat the measurement error problem arising from the increasing block pricing structure by using an instrumental

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[1] Estimating water demand under a block pricing structure requires an appropriate modelling to account for the choice of both within and between block consumption. Hewitt and Hanemann (1995), Corral et al. (1998) and Pint (1999), apply a two-stage model in which the choice of the block is modelled firstly in a discrete-choice fashion (using Probit analysis) and then the quantity within that block is chosen in a continuous way. The latter choice is modelled using simple regression analysis where an additional variable (the so called Mill s ratio) is used as additional explanatory variable to account for the first stage (block) choice..

variables (IV) estimation method. Moreover, unlike most other studies, here we use individual household data and this allows us to study the behavioural and welfare effects of alternative pricing policy on households grouped by income and other interesting policy characteristics. The empirical analysis in the paper focuses on the comparison between the effects of the increasing block and the uniform pricing systems on household demand and welfare. The former is the prevailing structure of water pricing in Cyprus and the latter is the pricing structure generally considered to be most efficient on the basis of the standard marginal cost pricing criterion.
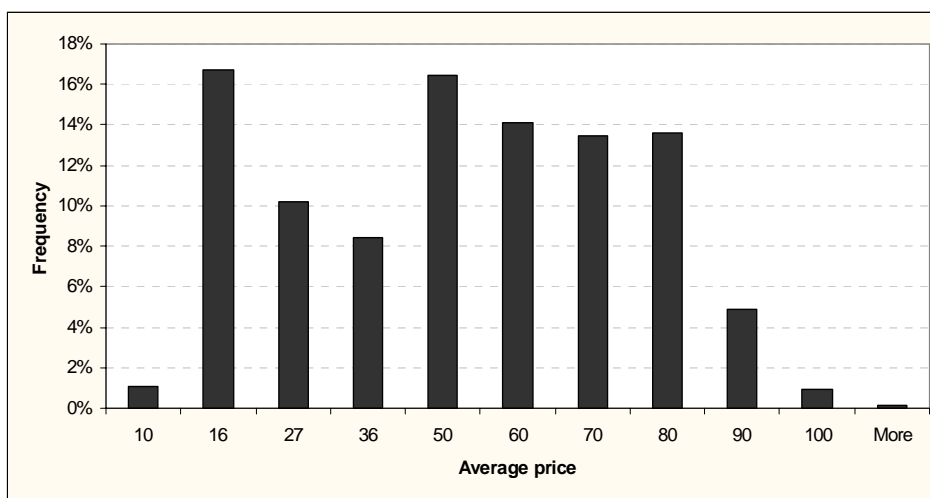
In the next section we describe the current water tariff structure applied to domestic users and examine its regional and income distribution aspects. In the same section we also compare the existing water tariff structure with a hypothetical situation where all water authorities in the island adopt a uniform &xed price policy. All calculations in this section are based on the assumption that water demand for domestic use does not respond to change in prices. This assumption is relaxed in the subsequent two sections where &rst an integrable demand for water model is derived (Section 3) and then applied to individual household data for the estimation of income and price elasticities and for the calculation of the welfare of a switch to a uniform &xed price policy (Section 4). The main conclusions of the paper are summarised in Section 5.

2 The structure and distribution of water tariffs in Cyprus

The government controlled part of Cyprus is divided into 37 water authorities each having its own tariff structure, as shown in Table A1 in the Appendix. The adoption of an increasing block tariff structure and differences in the application of this pricing policy across water authorities give rise to a substantial water price heterogeneity in the island. The effects of differences in water tariffs on households are examined in this paper using information on water consumption contained in the Family Expenditure Survey (FES) 1996/97. For each of over 2700 households randomly sampled, the FES reports its annual water bill together with its expenditure on a large number of other items, the level and sources of its income and many other characteristics such as demographic composition, economic/employment position of its members, housing variables, ownership of durable goods etc.

Using the FES 1996/97 standard geographical code we were able to allocate households to water authorities areas and calculate the level of annual water consumption for each household and the average price paid per cubic meter (pcm) of water.[2] As shown in Diagram 1, there is substantial variation in the average price paid pcm of water for domestic consumption in Cyprus. At the extreme bottom and top ends of the distribution there is a small number of households paying 10 cents and over one Cyprus pound pcm of water, respectively. Most households pay between 20 and 90 cents per cubic meter for their water consumption.[3]

Diagram 1: Distribution of the average water price

We have also used the FES 1996/97 data to calculate the marginal price pcm of water for each household, an important piece of information as it shows how households would be affected by a water price reform. Given the increasing block tariff structure, the marginal price pcm of water is higher than the average one in all water authorities. As seen in Diagram 2, the frequency distribution of the marginal price pcm of water is skewed to the left, indicating that moving up the price scale the proportion of households paying the corresponding price for an additional cubic meter of water increases.

Diagram 2: Distribution of the marginal water price (cents)
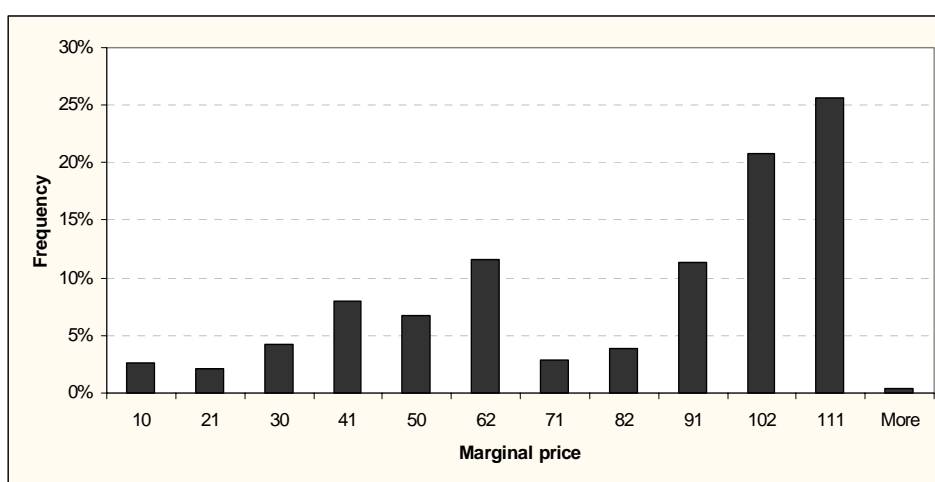
According to the Water Development Department (WDD), the average cost of supplying one cubic meter of water for domestic consumption (including distribution costs) is currently around 70 cents in Cyprus. Looking at the above frequency distributions, one can conclude that over 20% of the households are paying above average cost for their domestic water consumption. Furthermore, as seen from Diagram 2, more households are burdened rather than subsidised. at the margin of their water consumption.

2.1 Price heterogeneity between regions

Table 1 shows how the level of consumption and prices vary across water authorities. Under the heading Consumption is the average amount in cubic meters purchased by the households in the corresponding water authority and under the heading Price pcm are the average and marginal prices paid pcm of water purchased.

6

Table 1: Consumption and price pcm of water by region (in cents)

Water Authority Area	Consumption	Price (cents pcm)	
		Average	Marginal
Greater Urban SE of Pafos	170	26	33
Episkopi	167	20	38
Agios Theodoros	156	69	90
Alethriko	155	54	65
Pyla	145	54	59
Greater Urban E of Lemesos	140	20	45
Lemesos Town	140	24	40
Larnaka Greater Urban	136	60	93
Klirou	132	43	55
Kornos	119	50	70
Athienou	119	56	94
Kellaki	116	63	100
Katokopia	115	61	96
Greater Urban W of Lemesos	110	22	33
Paralimni	109	63	88
Nisou	101	56	92
Lefkosia Town	99	74	106
Lefkosia Suburbs	99	74	107
Lefkosia Greater Urban	98	70	105
Akrotiri	98	55	100
Solea	97	35	89
Pafos Town	97	40	80
Marathasa	92	20	42
Kiti	91	61	80
Larnaka Town	90	57	86
Pachna	90	46	88
Greater Urban N of Pafos	82	46	100
Parekklisia	82	45	98
Pegeia	80	38	100
Kokkinotrimithia	77	51	72
Paramytha	70	41	100
Gioulou	65	54	100
Omodos	62	26	100
Pano Panayia	56	62	100
Agros	56	22	78
Troodos	51	22	38
Salamiou	39	38	50

In relation to consumption, a notable feature of the &gures in Table 1 are the substantial regional differences in the annual amount of water purchased by households ranging from 170 cubic meters per household in the Greater Urban Area South-East of Paphos down to 39 cubic meters per household in the Salamiou water authority area. To some extend these differences re! ect regional differences in life-style, for example household living in high altitude areas can rely more on rainfall and have access to

more and better quality groundwater resources for their domestic needs. Therefore, their water consumption may not be as low as suggested by the &gures in Table 1.

Regarding differences in average and marginal price, the range of variation shown in Table 1 is striking considering the size of (the non-occupied part of) Cyprus. Some households purchase water from their local authority at an average price of 20 cents pcm (e.g. Episkopi and Greater Urban Area East of Lemesos) whereas in other water authority areas pay an average price of 74 cents pcm of water (Nicosia Town and its suburbs). Even more striking are differences in the price paid by households for the last cubic meter of water purchased from their local water authority area: from 33 cents in Greater Urban Area South-East of Paphos and Greater Urban Area West of Lemesos going up to over 105 cents in Nicosia Town and its suburbs.

As said earlier in this section, the observed regional differences in the price of water can be due to differences in the water pricing policies followed by the various water authorities, and the application of an increasing block tariff system resulting in large water users paying a higher average price than small water users. Looking at the &gures reported in Table 1, it appears that regional differences in water pricing policies is more likely to the cause of the observed price differences. For instance, households is Lemesos and its suburbs consume, on average, between 40% and 70% more water per annum yet, on average, they pay around 60% less pcm of water than households in Nicosia Town and its suburbs.

To consider the implications of the regional heterogeneity in water prices we have calculated the extent to which water consumption is subsidised. (or burdened) by the current water pricing system.[4] More speci&cally, we have compared the amount paid by each household under the present water tariff system with the amount which the same household would pay for the same water consumption under a system of uniform price where all households in all water authorities were charged 70 cents pcm of water, i.e. the amount corresponding to the average cost of supplying one cubic meter of water for domestic consumption. The results of this calculation are given in Table 2.

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[4]The subsidy (or burden) associated with water consumption is calculated for each household as the difference between the average supply cost pcm of water (estimated by the WDD to be 70 pence pcm of water) times the quantity of water consumed, minus the water bill paid. The average subsidy is the total subsidy divided by the cubic meters of water consumed by the household; and the marginal subsidy the difference between the estimated average supply cost pcm of water and the marginal price paid pcm of water by the household.

Table 2: Subsidies by region and consumption level (cents)[5]

Water Authority Area	All consumers		Bottom 20%		Top 20%	
	Average	Marginal	Average	Marginal	Average	Marginal
Lefkosia Suburbs	-4	-37	14	-25	-20	-40
Lefkosia Town	-4	-36	14	-22	-20	-40
Lefkosia Greater Urban	-3	-35	16	-15	-20	-40
Agios Theodoros	0	-20	10	-20	-8	-20
Kellaki	7	-30	.	.	.	.
Paralimni	7	-19	24	14	-11	-30
Pano Panayia	8	-30	.	.	.	.
Katokopia	9	-26	31	-10	-12	-30
Kiti	9	-11	.	.	.	.
Larnaka Greater Urban	9	-23	28	6	-10	-30
Larnaka Town	12	-16	27	-2	-2	-20
Athienou	13	-24	.	.	.	.
Nisou	14	-23	32	5	-4	-30
Akrotiri	15	-30	.	.	.	.
Gioulou	16	-30	.	.	.	.
Pyla	16	11	21	16	12	1
Alethriko	16	4	.	.	.	.
Kokkinotrimithia	19	-2	33	28	-4	-30
Kornos	20	5	31	13	1	-10
Greater Urban N of Pafos	24	-30	31	-30	11	-30
Pachna	24	-18	.	.	.	.
Parekklisia	24	-29	38	-24	1	-30
Klirou	27	15	38	30	18	10
Paramytha	29	-30	.	.	.	.
Pafos Town	30	-11	42	10	13	-20
Pegeia	31	-30	37	-30	26	-30
Salamiou	32	20	.	.	.	.
Solea	35	-19	49	22	11	-30
Omodos	44	-30	.	.	.	.
Greater Urban SE of Pafos	44	37	46	40	41	28
Lemesos Town	46	30	46	56	38	10
Greater Urban W of Lemesos	47	36	50	57	39	20
Troodos	48	32	.	.	.	.
Marathasa	48	28	.	.	.	.
Agros	48	-8	.	.	.	.
Greater Urban E of Lemesos	49	24	58	47	38	10
Episkopi	50	31	58	50	32	10

Looking at the &rst column, under the heading All consumers , we can see that, on the basis of the average price paid pcm of water, only households in the Lefkosia and its suburbs have annual water bills above what they would have had under the uniform (70 cents pcm of water) pricing policy. Furthermore, even in these few water authority where households would bene&t from a move to uniform water pricing, the gain is very small, around 4 cents pcm of water. In contrast, households in all other water authority

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[5]A dot indicates that the number of obsrvations in the corresponding cell were less than 5 and no average was computed.

areas would be worse off under the uniform pricing system, especially those living in water authority areas where the average price pcm of water is very low, e.g. the Lemesos and its suburbs and some Troodos mountain areas.

The picture, however, is completely different when we look at the marginal benefit to households from switching to the uniform price policy. In this case under the uniform water pricing system households in most water authority areas would pay less for an extra cubic meter of water than the amount they pay under the current system. In some areas the gain derived from switching to the uniform price system is very high, 35-37 pcm of water for households living in Lefkosia and its suburbs. At the same time, however, even at high levels of water consumption, households living in Lemesos and its suburbs would loose from switching from the current to the uniform price system. This is seen more clearly in the columns under the heading 'Bottom 20%' and 'Top 20%', showing the change in subsidy (or burden) which would result from switching from the current to a uniform price system for the 20% of households with the lowest and highest water consumption, respectively.

Assuming that the uniform price system considered here represents the true cost of supplying a cubic meter of water to households, the results above suggest that under the present water tariff system in Cyprus domestic water consumption in some water authorities (Lemesos and its suburbs in particular) is heavily subsidised. even in the case of the very large water users. In contrast, large users in other areas (especially in Nicosia and its suburbs) are heavily burdened by the current pricing system.

2.2 Price heterogeneity between income groups

The observed regional heterogeneity in the water price structure in Cyprus is primarily due to historical reasons, most probably the fact that areas blessed with more water resources resist economic and social arguments for a more equitable sharing of these resources. The increasing block tariff system, however, is based on both efficiency and equity considerations. The efficiency aspect relates to the negative externality of depleting a scarce natural resource, hence a system where the price of water increases with consumption can be seen as a measure to reduce this externality. The equity aspect relates to the fact that water is an essential item, therefore every household should be able to consume a subsistence amount of it. Indeed, one may also argue that since a

minimum amount of water consumption is required for hygienic purposes, subsidising water up to a certain level can be justi&ed on efficiency grounds too.

Table 3 shows the average annual consumption (cubic meters) and the average and marginal price paid pcm of water by households in different income groups. Water is a *normal good*, i.e. its level of consumption increases with income (second column in Table 3) and a *necessity*, i.e. its share in income decreases with income (third column in Table 3).[6] The average price pcm of water increases as we move from lower to higher income groups, indicating that the current water pricing structure in Cyprus is progressive. For example, households in the lowest 10% of income distribution pay 42 cents whereas those at the top 10% of income distribution pay 59 cents pcm of water. Also progressive, albeit by a lower rate, is the marginal price pcm of water, rising from 70 cents for households in the lowest 10% of income distribution to 85 cents pcm of water for households in the top 10% of income distribution.

Table 3: Consumption level and prices by income group

Income group	Consumption (Pounds)	Share in income	Average price (cents)	Marginal price (cents)
0%-10%	81	1.6%	42	70
11%-25%	95	0.9%	47	76
26%-50%	110	0.6%	50	79
51%-75%	116	0.4%	54	82
76%-90%	123	0.3%	54	83
Top 10%	130	0.2%	59	85

It follows from the points above that making the same assumption about water tariff reform as in the previous subsection (i.e. switching from the regionally heterogeneous increasing block pricing system to a homogeneous one where everyone pays a ! at rate of 70 cents pcm of water) will not be advantageous to the low income households. The effects of this water price reform on households grouped by income are shown in Diagram 3.

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[6]It is worth noting here that the decline in the share of water in consumer expenditure decreases at high income levels. This phenomenon is behind the increasing with income elasticity of demand for water which we as &nd in the empirical analysis in Section 4.

On average, households in all income groups would end up paying more for their current water consumption if the existing increasing block pricing system is replaced by a uniform pricing system. Again, assuming that the uniform price system considered here represents the true cost of supplying a cubic meter of water to households, this means that the water consumption of all income groups is subsidised. The largest subsidy (28 cents pcm of water) is enjoyed by households in the lowest 10% and the smallest (11 cents) by the top 10% of income distribution.

The picture, again, is very different in the case of the marginal benefits from switching to a uniform pricing system: all income groups would pay either the same (lowest income group) or less (all other income groups) for an additional cubic meter of water consumption under the uniform price system compared to the existing system. This means that marginal water consumption is not subsidised for any income group under the existing water tariff system. Those at the lowest 10% of income distribution, on average, purchase their last cubic meter of water at supply cost (70 cents pcm). As we move up the income scale households pay more for their marginal water consumption, with those in the top 10% of income distribution paying 15 cents pcm of water above the supply cost.

Diagram 3: Average and marginal subsidy by income group

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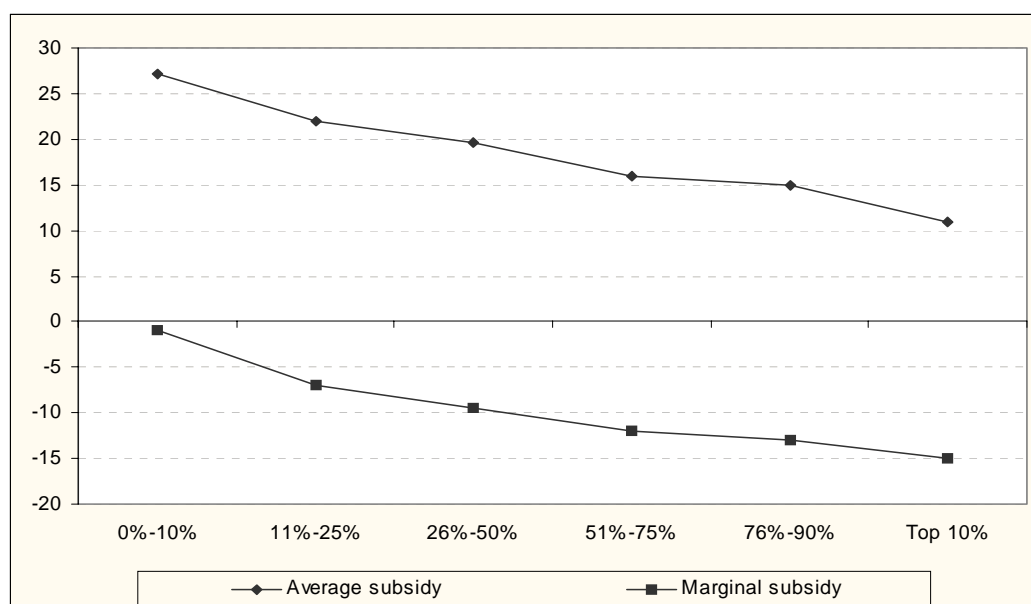
The tables and diagrams above compare the current consumption and increasing block tariff system with the hypothetical uniform pricing system on the assumption that water demand for domestic use is the same under the two price regimes. This assumption however is not realistic because household demand for water is likely to change in opposite direction to the change in price. The extent to which this will happen is determined by the price elasticity. In the next section we consider consumer demand for water in the context of a theoretical model satisfying the fundamental principles of consumer theory. This enables us to not only estimate the price elasticity of demand for water in a theoretically consistent manner but also integrate this demand back to the underlying cost and (indirect) utility functions to evaluate the welfare implications of the two price regimes.

3   Modelling tariff effects on consumer demand and welfare

We assume that preferences over goods are separable from leisure and public goods and can be expressed in terms of the Quadratic Logarithmic cost function (Lewbel 1990)

$$lnC\left(\mathsf{p},u_h\right)=a\left(\mathsf{p}\right)+\frac{b\left(\mathsf{p}\right)u_h}{1-g\left(\mathsf{p}\right)u_h}, \tag{1}$$

where u_h is the utility of the hth household (consumer) and p the vector of market prices of goods $(p_1, p_2, ..., p_I)$. Also $a\left(\mathsf{p}\right), b\left(\mathsf{p}\right)$ and $g\left(\mathsf{p}\right)$ are linearly independent and homogeneous functions the parameters of which can be allowed to vary with observable household characteristics (family size and composition, the education and occupation of household members etc) as shown in the empirical analysis described the next section.

The three price indices in (1) are assumed to have the Quadratic Almost Ideal Demand System (QUAIDS) form proposed by Banks, Blundell and Lewbel (1997),

$$
\begin{aligned}
a\left(\mathsf{p}\right) &= a_o + \Sigma_i a_i lnp_i + .5\Sigma_i\Sigma_j\gamma_{ij}lnp_jlnp_i,\\
b\left(\mathsf{p}\right) &= \Pi_i p_i^{\beta_i},\\
g\left(\mathsf{p}\right) &= \Sigma_i\lambda_i lnp_i,
\end{aligned}
\tag{2}
$$

yielding Marshallian demands in budget shares form,

$$w_{ih}=a_i+\Sigma_j\gamma_{ij}lnp_j+\beta_i\left[lny_h-a\left(\mathsf{p}\right)\right]+\frac{\lambda_i}{\Pi_i p_i^{\beta_i}}\left[lny_h-a\left(\mathsf{p}\right)\right]^2, \tag{3}$$

where y_h is the budget of the hth household.[7] The parameters obey the restrictions: $\Sigma_i a_i = 1$, $\Sigma_i \beta_i = \Sigma_i \lambda_i = 0$ and $\Sigma_i \gamma_{ij} = 0$ for adding-up; $\Sigma_j \gamma_{ij} = 0$ for homogeneity; and $\gamma_{ij} = \gamma_{ji}$ for symmetry.

Here we concentrate on consumer demand for a single commodity, water for domestic use, and assume that all other goods can be grouped in a Hicksian composite good. Therefore, the ith subscript is dropped for convenience so that the QUAIDS budget share of water is written,

$$w_{hr} = a + \gamma ln(p/P) + \beta lnY_h + \frac{\lambda}{P(p/P)^\beta}\left(lnY_h\right)^2, \tag{4}$$

where P is the price (index) summarising the cost of all the goods other than water and $lnY_h = lny_h - a_o - .5\gamma ln(p/P)^2 - lnP - aln(p/P)$.

We de&ne the market price of water as $p_r = p^* s_r$ where p^* is the producer s price, the subscript $r = 1, ..., R$ denotes the water tariff area and $s_r = (1 + t_r)$, where $-1 < t_r < 0$ is the tax paid and $0 < t_r < 1$ the subsidy received by consumers in the rth tariff area as proportion of the producer s price. When the latter and the prices of all goods other than water for domestic consumption are &xed we may normalise to $p_r = s_r$ and $P = 1$ so that the (4) becomes

$$
\begin{aligned}
w_{hr} &= a + \gamma lns_r + \beta \left[lnx_{hr} - alns_r - .5\gamma \left(lns\right)_r^2\right] \\
&\quad + \frac{\lambda}{s_r^\beta}\left[lnx_h - alns_r - .5\gamma \left(lns_r\right)^2\right]^2,
\end{aligned} \tag{5}
$$

where lnx_h is consumer expenditure measured from some minimum a_o.[8] The variable s_r in (5) is a price index re!ecting the tax paid (if $s_r > 1$) or subsidy received (if $s_r < 1$) by the consumer.

Knowledge of the parameters in (5) enable one to consider the implications of alternative water pricing policies on consumer behaviour and welfare. The behavioural

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[7]The QUAIDS demand system belongs to the family of Rank-3 demand systems, the most general empirical representation of consumer preferences that satis&es integrability (Gorman 1980 and Lewbell 1991). In the context of our analysis integrability is vital for considering the welfare implications of alternative water pricing policies.

[8]The parameter a_o is difficult to identify in empirical application and is set at a value corresponding to the minimum log expenditure in the sample. In the empirical analysis below we take this to be the average log expenditure of the households in our sample which are in the lowest percentile of the expenditure distribution.

effects can be summarised by the budget elasticity

$$\frac{1}{w_{hr}}\left(\beta + \frac{2\lambda}{s_r^{\beta}}\left[lnx_h - alns_r - .5\gamma\left(lns\right)^2_{hr}\right]\right) + 1 \tag{6}$$

and the (compensated) elasticity with respect to s_r

$$\frac{1}{w_{hr}}\left\{ \begin{array}{c} \gamma - \beta\left(a - \gamma lns\right) + \frac{2\lambda}{s_r}\left(a - \gamma lns\right)\left[lnx_h - alns_r - .5\gamma\left(lns\right)^2_{hr}\right] \\ -\frac{\lambda}{s_r^{\beta+1}}\left[lnx_h - alns_r - .5\gamma\left(lns\right)^2_{hr}\right]^2 \end{array} \right\} - 1 \tag{7}$$

of consumer demand for water. Evaluated at $s_r = 1$ the budget and tax (or subsidy) elasticity formulas simplify to

$$(1/w_{hr})\left(\beta + 2\lambda lnx_h\right) + 1 \tag{8}$$

$$\text{and} \quad (1/w_{hr})\left[\gamma - \beta a - \left(2\lambda a + \lambda\beta\right)lnx_h\right] - 1, \tag{9}$$

respectively. The parameters estimates required for the evaluation of these elasticities are obtained from &tting the water demand equation (5) to the data.

The effects of alternative water pricing policies on consumer welfare can be evaluated using the indirect utility function corresponding to the quadratic logarithmic cost function (1)

$$V(x_h, \mathsf{p}) = \frac{lnx_h - a\left(\mathsf{p}\right)}{g\left(\mathsf{p}\right)\left[lnx_h - a\left(\mathsf{p}\right)\right] + b\left(\mathsf{p}\right)}. \tag{10}$$

For the hth household to obtain the same utility level under the reference price vector p^* and some other price regime p, the following equality must hold,

$$\frac{lnx_h - a\left(\mathsf{p}\right)}{g\left(\mathsf{p}\right)\left[lnx_h - a\left(\mathsf{p}\right)\right] + b\left(\mathsf{p}\right)} = \frac{lnx_h^* - a\left(\mathsf{p}^*\right)}{g\left(\mathsf{p}^*\right)\left[lnx_h^* - a\left(\mathsf{p}^*\right)\right] + b\left(\mathsf{p}^*\right)}, \tag{11}$$

where lnx_h^* is the household expenditure under the reference price regime. At $\mathsf{p}^* = 1$ (11) can be written as the log expenditure index

$$\ln X_h \equiv ln\left(\frac{x_h}{x_h^*}\right) = a\left(\mathsf{p}\right) + \frac{\left[b\left(\mathsf{p}\right) - 1\right]lnx_h^*}{1 - lnx_h^* g\left(\mathsf{p}\right)}, \tag{12}$$

showing the change in log expenditure required by the household facing prices p to obtain the same level of utility as at reference prices p^*. Assuming p are the increasing block tariff and p^* the uniform price system, $\ln X_h$ can be interpreted as the consumer s willingness to pay a proportion of her/his income to avoid (or the compensation required to accept) the former in place of the latter pricing system.

Given the functional form of the QUAIDS demand system and the assumption that goods other than water can be grouped as one composite item, at $p_i^* = 1$ all i, (12) becomes

$$lnX_{hr} = (a + .5\gamma lns_r)lns_r + \frac{\left(s_r^\beta - 1\right)lnx_h^*}{1 - lnx_h^*\lambda lns_r} \tag{13}$$

and, like the budget and tax (or subsidy) elasticity, can be computed for each household in the sample using the parameter estimates obtained from &tting the water demand equation (5) to the data as described in the next section. The only difficulty here is that the expenditure index depends on lnx_h^*, a well known problem in the consumer welfare literature arising from the dependence of the expenditure index on the base utility level. In the empirical analysis below we compute the expenditure index at $x_h^* = 1$. This is equivalent to evaluating the cost of the price change at subsistence utility level, in which case $w_{hr} = a + \gamma lns_r$ and (5) can be written as

$$lnX_{hr} = (w_{hr} - .5\gamma lns_r)lns_r. \tag{14}$$

Furthermore we compute the deadweight loss index associated with a change in the price regime. We de&ned this index as the change in (indirect) log utility caused by a change in price regime with total expenditure remaining at its base period level,

$$lnW_h = \frac{ln\left[x_h^*\left(\mathsf{p}\right)\right] - a\left(\mathsf{p}\right)}{g\left(\mathsf{p}\right)\left[lnx_h^*\left(\mathsf{p}\right) - a\left(\mathsf{p}\right)\right] + b\left(\mathsf{p}\right)} - \frac{lnx_h^* - a\left(\mathsf{p}^*\right)}{g\left(\mathsf{p}^*\right)\left[lnx_h^* - a\left(\mathsf{p}^*\right)\right] + b\left(\mathsf{p}^*\right)}, \tag{15}$$

where, using q_{ih} to denote quantities,

$$ln\left[x_h^*(\mathsf{p})\right] = ln(\Sigma_i q_{ih}^* p_i) = ln\left[x_h^*\Sigma_i w_{ih}^*(p_i/p_i^*)\right] \simeq lnx_h^* + \Sigma_i w_{ih}^* ln(p_i/p_i^*)$$

is the expenditure in the base period spent under the current price regime.

Normalising at $x_h^* = 1$ and $p_i^* = 1$ all i, and using the same assumptions as before (i.e. the functional form of the QUAIDS demand system with goods other than water grouped as one composite item),

$$
\begin{aligned}
lnW_h &= \frac{w_{hr}^* lns_r - (a + .5\gamma lns_r)lns_r}{\lambda lns_r\left[w_{hr}^* lns_r - (a + .5\gamma lns_r)lns_r\right] + s_r^\beta} \\
&= \frac{-.5\gamma\left(lns_r\right)^2}{-.5\gamma\lambda\left(lns_r\right)^3 + s_r^\beta}
\end{aligned} \tag{16}
$$

This effectively re!ects the substitution effect caused by changes in relative prices, as de&ned by $-.5\gamma\left(lns_r\right)^2$ in (16); the denominator simply adjusts this effect to account for the non-homotheticity of preferences and should not deviate from unity.

4   Empirical analysis

The analysis in the previous section ignores two fundamental issues pertaining to empirical application based on individual household data: (i) demand for water is affected not just by the consumer s budget and prices but also by many demographic, housing and other characteristics and (ii) under increasing block pricing the subsidy (or burden) depends on the amount of water used by the individual household.

We take account of household heterogeneity by allowing the intercept in the demand for water equation (5) to depend on a large number of characteristics drawn from the FES 1996/97 and found to affect consumer behaviour in other consumer demand studies based on individual household data (e.g. Blundell, Pashardes and Weber 1993). These household characteristics are shown in Table A2 in the Appendix.

The dependence of tax/subsidy on the amount of water used is accounted for by replacing s_r in (5), calculated on the basis of the actual quantity of water used (as explained in Section 2) with $\widehat{s}_{hr}$, its value corresponding to the predicted quantity of water used. The latter is obtained from the reduced form equation,

$$q_{hr} = \varepsilon_o + \Sigma_m \varepsilon_m Z_{mhr} + v_{hr} \tag{17}$$

where Z_{mhr} is the mth exogenous variable corresponding to the hth household in the rth region and v_{hr} a random error. Among the Z, s are all the household characteristics and other variables included in the demand for water equation below (except s_r) plus dummies for the water tariff area, capturing differences in unit cost and other differences in water consumption between water tariff areas in the island. The full list of variables included in (17) is reported in Table A3 in the Appendix. The same table also reports the parameter estimates and diagnostic statistics obtained from &tting (17) to data.

Incorporating in (5) the modi&cations required to account for household heterogeneity and increasing block pricing (i.e. the endogeneity of s_r) we obtain

$$
\begin{aligned}
w_{hr} \;=\; & a + \Sigma_k \delta_k lnz_{kh} + \beta \left[lnx_h - (a + \Sigma_k \delta_k lnz_{kh}) \, ln\widehat{s}_{hr} - .5\gamma \left(ln\widehat{s} \right)^2_{hr} \right] \\
& + \gamma ln\widehat{s}_{hr} + \frac{\lambda}{\widehat{s}^{\beta}_{hr}} \left[lnx_h - (a + \Sigma_k \delta_k lnz_{kh}) \, ln\widehat{s}_{hr} - .5\gamma \left(ln\widehat{s} \right)^2_{hr} \right]^2 + u_{hr},
\end{aligned}
\tag{18}
$$

where z_{kh} is the kth characteristic of the hth household and u_{hr} a random error. The

dependent variable is de&ned as the share of water in household expenditure on non-durable goods.

The parameters in (18), estimated by Maximum Likelihood, are reported together with their standard errors (and some system diagnostic statistics) in Table A2 in the Appendix. The price and budget (income) elasticities corresponding to these parameter estimates have been calculated for each household in the sample using the formulas (8) and (9) in Section 3. The elasticities, grouped by household income, are shown in Table 4.

As one would expect, water appears to be a necessity, with a budget elasticity ranging between 0.25 for the lowest income group to 0.48 for the highest income group. The increasing budget elasticity with income, suggesting that water is more of a necessity to households with lower rather than higher income, is a puzzle to us. A possible explanation of this phenomenon can be the fact that water is a complement to large houses with lawned gardens, swimming pools, jacuzzis and other luxury goods purchased by the rich.[9] Another possible explanation for the increase in budget elasticity with income is the fact that under the current increasing block tariff system rich people tend to be large water users and pay more for a given quantity of water. Therefore a given proportional change in their demand means more to them in terms of income than the same proportional change in the consumption of low income households. The average budget elasticity of demand for water for the population as a whole is 0.32.

Table 4: Price and budget elasticities by income group

	Income group					
Elasticity	0%-10%	11%-25%	26%-50%	51%-75%	75%-90%	Top 10%
Budget	0.25	0.22	0.23	0.30	0.35	0.48
Price	-0.79	-0.69	-0.60	-0.56	-0.50	-0.39

---

[9]As seen in Table A2 in the Appendix, although we conditon demand for water on certain durable goods (washing machines, dishwashers etc) and household characteristics, we have not been able to do this for all such goods and characteristics affecting water consumption due to lack of information. One particularly important conditioning variable missing from our empirical water demand equation, for the same reason, is access to groundater.

The price elasticity of demand for water declines with income: it starts from -0.79 for the worse off and decreases (in absolute size) to -0.39 for the better off households. This means that low income households are more sensitive to changes in the price of water than high income households. One possible reason for this result is that, as said earlier, high income households tend to consume water in conjunction with expensive durables goods; therefore, their demand is less responsive to water price changes.

One could point to the decline in the price elasticity of demand for water as income increases, as evidence against the argument in favour of using price as a water demand management tool. At the same time, however, the same argument is strengthened by the empirical evidence here, in the sense that the price elasticities of demand for water reported in Table 4 are probably among the highest ever estimated.

Turning to the welfare effects we calculate the expenditure index (14) and the deadweight loss index (16) by region and income group, by comparing the present water price system with the uniform price one. Furthermore, we express the results in Cyprus pounds per annum (in 1997 prices) to show (i) in the case of the expenditure index the amount the various household groups are willing to pay (be paid) to accept a switch from the current to the uniform price system and (ii) in the case of the deadweight loss index the loss in pounds due the price distortion created by the regional variation and the increasing block tariff system.

The results of these calculations are shown in Diagrams 4 and 5. Commending &rst on the results by region (Diagram 4)[10], as one would expect from the analysis in Section 2, those willing to pay the highest amount, households living in the Greater Urban area SE of Pafos, Lemesos and its suburbs are the most willing to pay (up to 47 Cyprus pounds) to avoid a change from the current to the uniform pricing system. Notably, the Greater Urban area SE of Pafos fares badly here due to the large price subsidies maintained at high levels of water consumption in this water authority area. In contrast households in Nicosia and its suburbs and in Larnaca are the most willing to pay (up to 33 Cyprus pounds) for the replacement of the current increasing bloc tariff with a uniform water pricing system.

---

[10]In Diagram 4 we show only regions where the water price reform causes a change in water above ten Cyprus pounds per annum.

Diagram 4: Welfare effects by region (in Cyprus pounds p.a.)
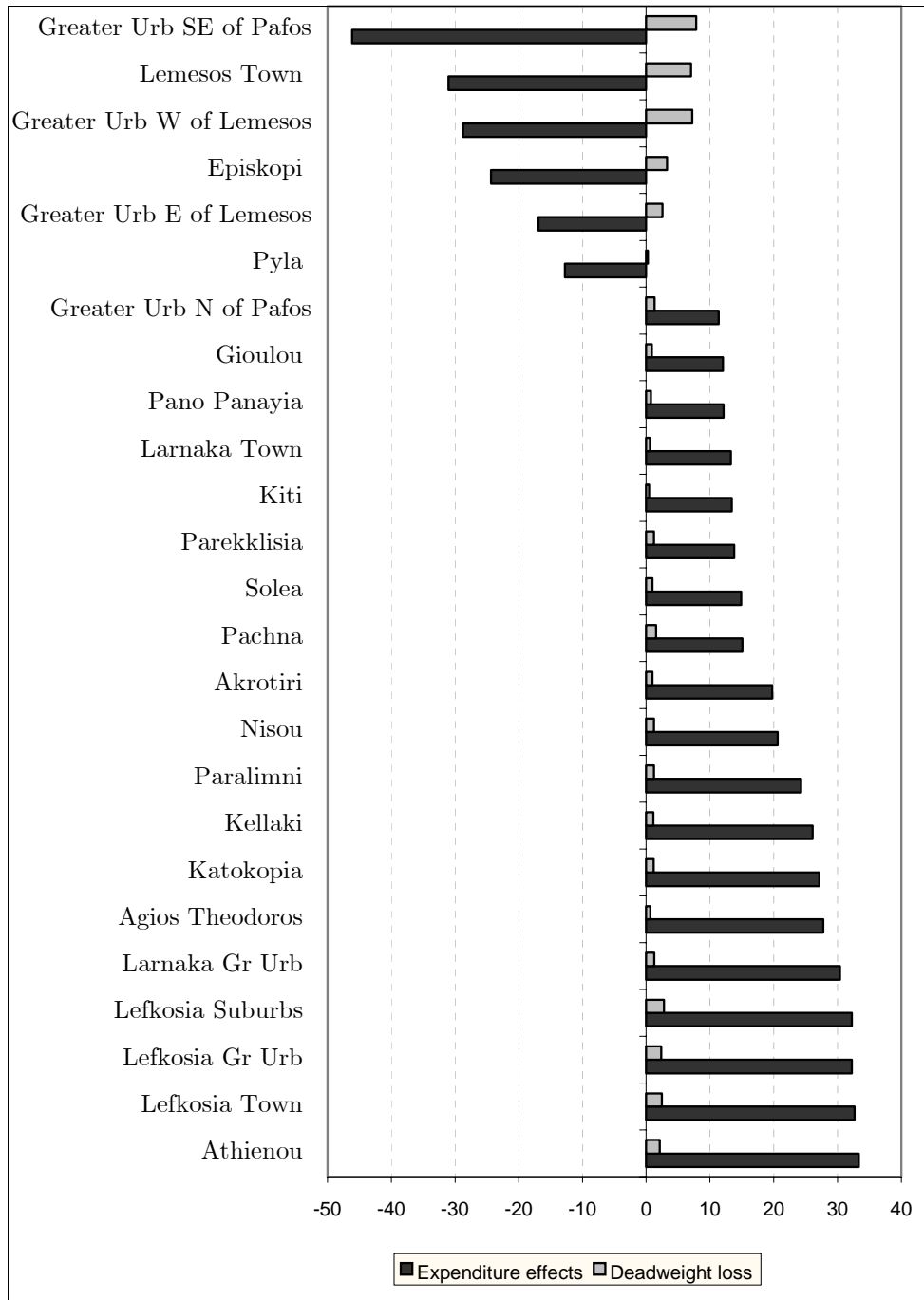
Turning to the deadweight loss, again, this is most pronounced (up to 8 Cyprus pounds or more per annum) in areas with large price distortions among large water users: Greater Urban area SE of Pafos, Lemesos and its suburbs. In interpreting the

deadweight loss here one can say that it re!ects the amount households would be able
to forgo if hey were able to replace the water subsidy with a cash payment. It therefore
shows the waste associated with the current pricing system.

Diagram 5: Welfare effects by income (in Cyprus pounds p.a.)
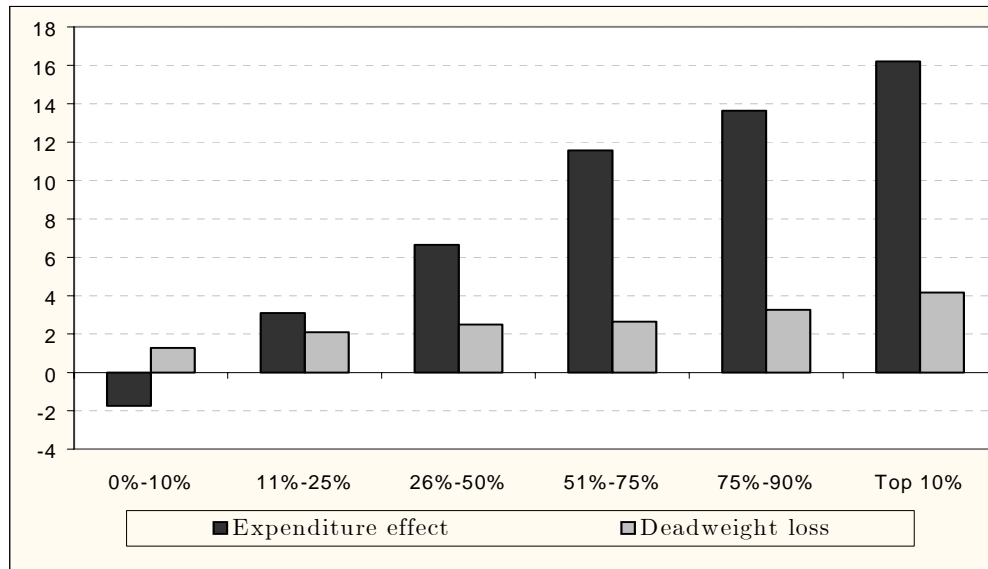
The welfare effects by income in Diagram 5 show that all household groups, on
average, would be willing to pay to move to the uniform water pricing system except
those in the lowest income group which would pay to avoid such a move. Furthermore,
given the inherent progressivity of the current increasing block pricing system, the
amount households would be willing to pay for a change from the current to the uniform
pricing system increases with income, up to 16 Cyprus pounds p.a for those in the top
10% of income distribution. Again, there is a deadweight loss associated with the current
pricing system measured by the amount households would be willing to forgo as bene&t
received through water subsidies if they had the option to replace this bene&t with cash.
The deadweight loss also increases with income but very slightly so, indicating that the
price distortions are not highly correlated with household income.

5   Conclusions

In this paper we investigate the effects of switching from the current regionally het-
erogeneous increasing block water pricing system to a regionally homogeneous uniform
pricing one. We find that the current system is progressive but inefficient in the sense
that it introduces gross price distortions resulting in deadweight loss. The regional
differences, in particular, introduce a substantial price heterogeneity that cannot be
justified on the basis of efficiency or equity criteria. It cannot be justified on efficiency
grounds because it is difficult to imagine that in a small island like Cyprus such large
regional differences in price can reflect differences in supply costs. The regional price
heterogeneity cannot also be justified on equity grounds because we found that users
of large quantities of water pay substantially less per cubic meter of water than users
consuming much smaller amounts of water.

Our empirical analysis suggests that the price elasticity of demand for water ranges
between -.4 for households in the lowest and -.8 for households in the highest 10% of
income distribution. This means that, in the case of residential water use, price can play
a role in the context of a demand management scheme designed to tackle the growing
fresh water problems in Cyprus. Such an approach, however, should take into account
the distributional impact of alternative price regimes. Any major water price reform is
bound to have effects on the welfare of individual consumers. In other words there will
be winners and losers, and therefore there will also be a need to consider how to deal
with potential hardship caused by the water price reform.

We believe that inevitably there will be a move towards a more uniform marginal cost
pricing system in Cyprus. Our empirical results show that such a system is desirable
in terms of reducing the deadweight loss associated with the price distortions of the
present system. At the same time, however, its benefits will be distributed in favour
of the better off households, while households in the bottom 10% of the distribution
of income will be net losers. Furthermore, we believe that for a water pricing policy
to be capable of ensuring sustainable development of water resources, must go beyond
the purely quantitative aspect of water demand to also take into account qualitative
considerations. This is an important area in need of more research in the water demand
management literature.

References

Banks J., Blundell R. and Lewbel A., 1997, Quadratic Engel Curves and Consumer Demand, Review of Economics and Statistics, 4,527-539.

Blundell, R., Pashardes, P., and Weber, G., 1993, What do we Learn about Consumer Demand Patterns from Micro-data, American Economic Review, 83, 570-597.

Billings, B., 1982. Speci&cation of block rate price variables in demand models. Land Economics, 58(3), 386-393.

Chicoine, D. and Ramamurthy, G., 1986. Evidence on the speci&cation of price in the study of domestic water demand. Land Economics, 62(1), 26-32.

Dinar, A., 2000, The political economy of water pricing reforms, in Dinar A (ed) The political economy of water pricing reforms, Oxford University Press.

Dandy, G., Nguyen, T. and Davies, C., 1997. Estimating residential water demand in the presence of free allowances. Land Economics, 73(1), 125-139.

Foster, H. S. Jr., and Beattie B., 1979. Urban residential demand for water in the United States. Land Economics, 55 (Feb.), 43-58.

Foster and Beattie, 1981a. Urban residential demand for water in the United States: Reply. Land Economics, 57 (May), 257-265.

Foster and Beattie, 1981b. On the speci&cation of price in studies of consumer demand under block price scheduling. Land Economics, 57 (Nov.), 624-29.

Gibbons, D. C., 1986. The economic value of water. Washington, DC, Resources for the Future.

Gorman, T., 1981, Some Engel curves, in Essays in the Theory and Measurement of Consumer Behaviour in Honor of Richard Stone, ed. A. Deaton, Cambridge University Press.

Gottlieb, M., 1963. Urban domestic demand for water: A Kansas case study. Land Economics, 39 (May): 204-210.

Hanke, S., and de Maré, L., 1982. Residential water demand: a pooled time series cross-section study of Malmö, Sweden. Water Resources Bulletin, 18(4), 621-625.

Hansen, L., 1996. Water and energy price impacts on residential water demand. Copenaghen. Land Economics, 72(1), 66-79.

Herrington, P. R., 1987. Pricing of water services. OECD Publication.

Höglund, L., 1999. Household demand for water in Sweden with implications of a potential tax on water use. Water Resources Research, 35(12), 3853-3863.

Lewbel, A., 1990, Full Rank Demand Systems, International Economic Review, 31, 289-300.

Lewbel, A., 1991, The Rank of Demand Systems: Theory and Nonparametric Estimation, Econometrica, 59, 711-730.

Martinez-Espiñeira, 2000. Residential water demand in the Northwest of Spain. Paper presented in the 10th annual conference of the European Association of Environmental and Resource Economists (EAERE 2000).

Nauges, C., and Thomas, A., 2000. Privately-operated water utilities, municipal price

negotiation, and estimation of residential water demand: The case of France. Forthcoming in Land Economics, 76(1).

Nieswiadomy, M., and Cobb, S., 1993. Impact of pricing structure selectivity on urban water demand. Contemporary Policy Issues, 11(6), 101-113.

Nieswiadomy, M. L., and Molina, D. G., 1989. comparing residential water estimates under decreasing and increasing block rates using household data.

Nordin, J., 1976. A proposed modi&cation on Taylor s demand supply analysis: comment. Bell Journal of Economic Management and Science, 7(2), 719-721.

OECD, 1987. Pricing of water services, organization of economic co-operation and development, Paris.

Renwick, M., and Archibald, S., 1998. Demand side management policies for residential water use: Who bears the conservation burden? Land Economics, 74(3), 343-359.

Schefter, J., and David, E., 1985. Estimating residential water demand under multi-tariffs using aggregate data. Land Economics, 61(3), 272-80.

Taylor, L. D., 1975. The demand elasticity a survey. The Bell Journal of Economics, 6 (Spring), 74-110.

Young, R. A., 1973. Price elasticity of demand for municipal water: Case study of Tucson, Arizona. Water Resources Research, 9 (Dec.), 1068-1072.