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An Asymmetric Block Dynamic Conditional Correlation Multivariate GARCH Model¹

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ABSTRACT

The Block DCC model for determining dynamic correlations within and between groups of financial asset returns is extended to account for asymmetric effects. Simulation results show that the Asymmetric Block DCC model is competitive in in-sample forecasting and performs better than alternative DCC models in out-of-sample forecasting of conditional correlation in the presence of asymmetric effect between blocks of asset returns. Empirical results demonstrate that the model is able to capture the asymmetries in conditional correlations of some blocks of currencies in East Asia in the turbulent years of the late 1990s.

Keywords: asymmetric effect, block dynamic conditional correlation, multivariate GARCH
JEL Codes: C32, C5, G10

I. INTRODUCTION

By now, the GARCH model of Bollerslev (1986) has been extended to several classes of multivariate GARCH models (see Bauwens, Laurent and Rombouts (2006)). GARCH itself has come a long way since Robert Engle's (1982) pioneering paper on the ARCH. Multivariate GARCH (MGARCH) research focuses on ways of simplifying the variance-covariance matrix where the number of parameters to be estimated explodes for higher dimensions making estimation costly and computationally intractable. The approaches to simplify the estimation of the parameters of the variance-covariance matrix are now well-developed although suggestions have been made to come up with models that account for economic theory as a basis for simplifying this matrix, Diebold (2004).

It is widely held in financial econometrics that the returns of financial assets move together. The behavior of returns can be investigated using volatility models. Although models of MGARCH have accounted for the stylized fact that returns of similar assets move together there are other salient features of financial assets that need to be considered. Dissimilar assets have varying degrees of return correlation and one set of assets serves as a leading source of volatility for other sets of assets. Conditional correlation research has gained momentum only in recent years but its unconditional counterpart is the most common input to models being used by a typical investor.

The differences in types of financial assets result in certain groups of asset returns to be more correlated with each other than they are relative to others (see Kroner and Ng (1998), Billio,

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Caporin and Gobbo (2003)). Examples of these groups are the property, food and beverage, energy, banking and finance, mining sectors in a stock market. Stock price returns within each sector are highly correlated with each other and the degree and direction of correlations between sectors are known to vary. It has been observed that the impact of bad news on stock price return correlations is asymmetric between small and large firms. It is greater on small firms when bad news occur with large firms but not vice-versa, Kroner and Ng (1998). This is important in asset allocation where investors decide what proportion of different types of instruments should comprise their portfolio to lower the risk of their overall holdings while at the same time maximize their returns. This asymmetric behavior when applied to asset allocation modeling leads to economically significant gains in an investor's portfolio according to Patton (2004).

Asymmetric effect occurs when unexpected downward movements in the price of an asset raise the conditional volatility of returns more than when there are unexpected upward movements (see Nelson (1991) and Engle and Ng (1993)). The asymmetric effect was confirmed by empirical investigations of French, Schwert and Stambaugh (1987), Schwert (1990) and Nelson (1991) among others. In the stock market studies, Erb, Harvey and Viskanta (1994), Ang and Chen (2001), Longin and Solnik (2001) have shown that there is greater dependence between returns during market downturns.

Asymmetric effect in groups or blocks of financial asset returns were also observed by Lo and MacKinlay (1990) and Conrad, Gultekin and Kaul (1991). The lead-lag effect in portfolio returns by Lo and MacKinlay (1990) states that small firm portfolio returns lag large firm portfolio returns but not the other way around while the volatility spill-over hypothesis by Conrad, Gultekin and Kaul (1991) claims that volatility spills over large to medium and medium to small firms but not in the opposite way. Milunovich (2003) confirmed these asymmetric behaviors in stock prices of blocks of small, medium and large firms using a structural MGARCH model.

In the study of worldwide linkages in the dynamics of volatility and correlations of bonds and equity markets Capiello, Engle and Sheppard (2003) showed that there were strong asymmetries in conditional volatility of equity index returns while bond index returns have little evidence of this behavior. Capiello, Engle and Sheppard (2003) estimated the correlations of stock and bond indices of four major regions assuming the same dynamic condition for the correlations. On the other hand, Billio, Caporin and Gobbo (2003) introduced Block Dynamic Conditional Correlation (BDCC) which assumes different dynamic condition for correlation of assets within a certain block of assets. But Billo, Caporin and Gobbo's (2003) BDCC does not account for asymmetries between blocks while the Asymmetric DCC (ADCC) model of Capiello, Engle and Sheppard (2003) does not consider the asymmetric correlations between blocks of assets per se. Capiello, Engle and Sheppard (2003) only took the average dynamic correlations of individual indices to represent regional dynamic conditional correlations.

This paper will extend the BDCC by introducing asymmetric effects of conditional correlation between blocks of asset returns. Billio, Caporin and Gobbo (2003) noted the possible extension of their model that would include an asymmetric component but did not pursue this and thus the motivation for this study. In Section II of this paper the proposed model is presented including its relation to the relevant models in the literature and its estimation procedure. In Section III Monte Carlo simulation results are presented, in Section

IV the model is applied to East Asian currency returns and in the last section is the conclusion.

II. THE MODEL

This paper proposes a new model called Asymmetric Block Dynamic Conditional Correlation (ABDCC) MGARCH. A missing element in the DCC literature is the asymmetric effect in the conditional correlation between blocks of asset returns and this is the contribution of ABDCC. Actually, Cajigas and Urga (2005) have accounted for this kind of asymmetry in their Asymmetric Generalized DCC (AGDCC) but theirs is a special case of the proposed model. DCCs belong to the nonlinear class of MGARCH models that aims to reduce the dimensionality of the conditional variance-covariance matrix, H_t , by specifying the conditional correlation matrix. Bauwens, Laurent and Rombouts (2006) provide the most recent and comprehensive survey of the different classes of MGARCH models.

The general representation of the MGARCH model is

$$\begin{aligned}
 y_t &= \mu_t + \varepsilon_t \\
 \varepsilon_t &= H_t^{1/2} z_t \text{ where } z_t \sim N(0, I) \\
 \varepsilon_t | \mathfrak{I}_{t-1} &\sim N(0, H_t)
 \end{aligned}
 \tag{1}$$

where y_t is the $N \times 1$ vector of asset returns and \mathfrak{I}_{t-1} is a sigma algebra of information up to time $t-1$. Without loss of generality μ_t is assumed to be zero. The ABDCC model has the following specification for the $N \times N$ conditional variance-covariance matrix H_t :

$$H_t = D_t R_t D_t = \left(\rho_{ij,t} \sqrt{h_{ii,t} h_{jj,t}} \right)
 \tag{2}$$

as in Engle's (2002) DCC model where $D_t = \text{diag}(h_{11,t}^{1/2} \dots h_{NN,t}^{1/2})$. $R_t = Q_t^{*-1} Q_t Q_t^{*-1}$ is the $N \times N$ matrix of dynamic conditional correlation driven by the evolution process $Q_t = (q_{ij,t})$ where $Q_t^* = \text{diag}(\sqrt{q_{ii,t}})$. The ABDCC has the following specification for Q_t :

$$Q_t = (\bar{Q} - \alpha(L) \circ \bar{Q} - \beta(L) \circ \bar{Q} - \eta(L) \circ \bar{N}) + \alpha(L) \circ (\varepsilon_t^* \varepsilon_t^{*'}) + \beta(L) \circ Q_t + \eta(L) \circ (n_t n_t')
 \tag{3}$$

where \circ indicates the Hadamard product, $\bar{Q} = E[\varepsilon_t^* \varepsilon_t^{*'}]$ with sample equivalent

$\hat{\bar{Q}} = \frac{\sum_{t=1}^T \varepsilon_t^* \varepsilon_t^{*'}}{T}$ that serves as estimator of \bar{Q} , $\varepsilon_t^* \sim N(0, R_t)$ is an $N \times 1$ vector of residuals standardized by their conditional standard deviation, $\bar{N} = E[n_t n_t']$ with sample equivalent

$\hat{\bar{N}} = \frac{\sum_{t=1}^T n_t n_t'}{T}$ that serves as estimator of \bar{N} , $n_t = I[\varepsilon_t^* < \tau] \circ \varepsilon_t^*$ is the asymmetric component where τ is a time-varying threshold of asymmetry that is typically set to zero. The $\alpha(L)$,

$\beta(L)$ and $\eta(L)$ are $N \times N$ parameter matrices where the N asset returns series are grouped in w sets of m_1, m_2, \dots, m_w -dimensional vectors and

$$\alpha(L) = \sum_{i=1}^{\bar{q}} \alpha_i L^i, \quad \beta(L) = \sum_{j=1}^{\bar{p}} \beta_j L^j, \quad \eta(L) = \sum_{k=1}^{\bar{r}} \eta_k L^k.$$

Furthermore, α_i , β_j and η_k have the following structure

$$\alpha_i = \begin{bmatrix} \alpha_{i,11} i(m_1) i(m_1)' & \alpha_{i,12} i(m_1) i(m_2)' & \cdots & \alpha_{i,1w} i(m_1) i(m_w)' \\ \alpha_{i,21} i(m_2) i(m_1)' & \alpha_{i,22} i(m_2) i(m_2)' & \cdots & \alpha_{i,2w} i(m_2) i(m_w)' \\ \vdots & \vdots & \ddots & \vdots \\ \alpha_{i,w1} i(m_w) i(m_1)' & \alpha_{i,w2} i(m_w) i(m_2)' & \cdots & \alpha_{i,ww} i(m_w) i(m_w)' \end{bmatrix}$$

$$\beta_j = \begin{bmatrix} \beta_{j,11} i(m_1) i(m_1)' & \beta_{j,12} i(m_1) i(m_2)' & \cdots & \beta_{j,1w} i(m_1) i(m_w)' \\ \beta_{j,21} i(m_2) i(m_1)' & \beta_{j,22} i(m_2) i(m_2)' & \cdots & \beta_{j,2w} i(m_2) i(m_w)' \\ \vdots & \vdots & \ddots & \vdots \\ \beta_{j,w1} i(m_w) i(m_1)' & \beta_{j,w2} i(m_w) i(m_2)' & \cdots & \beta_{j,ww} i(m_w) i(m_w)' \end{bmatrix}$$

$$\eta_k = \begin{bmatrix} \eta_{k,11} i(m_1) i(m_1)' & \eta_{k,12} i(m_1) i(m_2)' & \cdots & \eta_{k,1w} i(m_1) i(m_w)' \\ \eta_{k,21} i(m_2) i(m_1)' & \eta_{k,22} i(m_2) i(m_2)' & \cdots & \eta_{k,2w} i(m_2) i(m_w)' \\ \vdots & \vdots & \ddots & \vdots \\ \eta_{k,w1} i(m_w) i(m_1)' & \eta_{k,w2} i(m_w) i(m_2)' & \cdots & \eta_{k,ww} i(m_w) i(m_w)' \end{bmatrix}$$

where $i(m_g)$ is a column vector of ones with dimension m_g and $g = \{1, 2, \dots, w\}$. L^i is the time lag of order i .

Parameter estimation of the ABDCC model invokes the concept of variance targeting introduced by Engle and Mezrich (1996). Variance targeting assumes that in the long run the Q_t approaches the sample variance-covariance matrix. In order to ensure that the estimation is conducted within a valid parameter space, ABDCC must be specified by maintaining the positive definiteness of the conditional correlation matrix, R_t . This can be done if Q_t is positive definite. The following conditions ensure the positive definiteness of Q_t :

1. $(\bar{Q} - \alpha(L) \circ \bar{Q} - \beta(L) \circ \bar{Q} - \eta(L) \circ \bar{N})$ must be positive definite;
2. $\alpha(L) \circ (\varepsilon_t^* \varepsilon_t^{*'})$ must be positive semidefinite;
3. $\beta(L) \circ Q_t$ must be positive definite; and
4. $\eta(L) \circ (n_t n_t')$ must be positive semidefinite.

In the optimization, these conditions make the location of the maximum of the likelihood function difficult. According to Engle and Sheppard (2001), a sufficient but not necessary condition for Q_t to be positive definite is for all the parameters to be positive; and, a minimum condition to ensure the positive definiteness of Q_t is for the first condition to be satisfied as shown by Engle and Mezrich (1996).

An alternative model of ABDCC is

$$Q_t = (I - \alpha(L) - \beta(L) - \eta(L)) \circ \bar{Q} + \alpha(L) \circ (\varepsilon_t^* \varepsilon_t^{*'}) + \beta(L) \circ Q_t + \eta(L) \circ (n_t n_t'), \quad (4)$$

but this specification no longer satisfies variance targeting. Conditions 1 to 4 above still apply to this alternative model where in the first condition $(\bar{Q} - \alpha(L) \circ \bar{Q} - \beta(L) \circ \bar{Q} - \eta(L) \circ \bar{N})$ is replaced by $(I - \alpha(L) - \beta(L) - \eta(L)) \circ \bar{Q}$.

The ABDCC, Eq. (3), regresses to the following models:

1. Engle's (2002) DCC if the elements of $\alpha(L)$, $\beta(L)$ and $\eta(L)$ have $\alpha_{ij} = a$ in $\alpha(L)$, $\beta_{ij} = b$ in $\beta(L)$, and $\eta(L) = 0$, for all $i, j \in \{1, \dots, N\}$;
2. Billio, Caporin and Gobbo's (2003) BDCC if $\eta(L) = 0$;
3. Capiello, Engle and Sheppard's (2003) ADCC if the elements of $\alpha(L)$, $\beta(L)$ and $\eta(L)$ have $\alpha_{ij} = (\sqrt{a})^2$ in $\alpha(L)$, $\beta_{ij} = (\sqrt{b})^2$ in $\beta(L)$, and $\eta_{ij} = (\sqrt{g})^2$ in $\eta(L)$, for all $i, j \in \{1, \dots, N\}$;
4. Cajigas and Urga's (2005) AGDCC, if the elements of $\alpha(L)$, $\beta(L)$ and $\eta(L)$ have $\alpha_{ii} = a_{ii}^2$ and $\alpha_{ij} = a_{ii} a_{jj}$ in $\alpha(L)$, $\beta_{ii} = b_{ii}^2$ and $\beta_{ij} = b_{ii} b_{jj}$ in $\beta(L)$, $\eta_{ii} = g_{ii}^2$ and $\eta_{ij} = g_{ii} g_{jj}$ in $\eta(L)$, for all $i, j \in \{1, \dots, N\}$ where $i \neq j$.

for time lag of order 1. These can be easily generalized to higher orders of time lag. While the alternative ABDCC, Eq. (4), regresses to Cajigas and Urga's (2005) alternative AGDCC if the same conditions in No.4 are satisfied.

The likelihood function of ABDCC is derived under the assumption that the underlying distribution of the vector of returns is multivariate normal. Limited information maximum likelihood (LIML) is utilized here. In a two-stage LIML parameter estimation procedure the set of parameters is divided into two. The likelihood function is maximized with respect to the first set of parameters then the parameter estimates of the first set serve as inputs to the second stage where the next set of parameters are then estimated. The likelihood function of the model is

$$L(\vartheta, \phi | y_t) = \prod_{t=1}^T \left[\frac{1}{(\sqrt{2\pi})^N |H_t|^{1/2}} e^{-\frac{1}{2} y_t H_t^{-1} y_t'} \right]. \quad (5)$$

The vector ϑ consists of univariate GARCH parameters for each element of the N -dimensional y_t vector of asset returns and vector ϕ consists of the parameters of Q_t . Engle and Sheppard (2001) have shown the consistency and asymptotic normality of this two-stage procedure. Now, taking the natural logarithm of the likelihood function

$$\begin{aligned}\log L(\vartheta, \varphi | y_t) &= -\frac{1}{2} \sum_{t=1}^T \left(N \log(2\pi) + \log |H_t| + y_t' H_t^{-1} y_t \right) \\ &= -\frac{1}{2} \sum_{t=1}^T \left(N \log(2\pi) + \log |R_t| + 2 \log |D_t| + y_t' D_t^{-1} R_t^{-1} D_t^{-1} y_t \right)\end{aligned}\quad (6)$$

where $H_t = D_t R_t D_t$ and $D_t = \text{diag}(h_{11,t}^{1/2} \dots h_{NN,t}^{1/2})$. As shown by Engle and Sheppard (2001), if R_t is assumed to be an identity matrix in the first stage estimation,

$$\log L(\vartheta | y_t) = -\frac{1}{2} \sum_{t=1}^T \left(N \log(2\pi) + \log |I_N| + 2 \log |D_t| + y_t' D_t^{-1} I_N^{-1} D_t^{-1} y_t \right) \quad (7)$$

$$\hat{\vartheta} = \arg \max[\log L(\vartheta | y_t)] \quad (8)$$

turns out to be the univariate estimation of the individual GARCH models of asset returns of y_t . The second stage estimation will have

$$\log L(\varphi | \hat{\vartheta}, y_t) = -\frac{1}{2} \sum_{t=1}^T \left(N \log(2\pi) + \log |R_t| + 2 \log |\hat{D}_t| + y_t' \hat{D}_t^{-1} R_t^{-1} \hat{D}_t^{-1} y_t \right) \quad (9)$$

where $\varepsilon_t^* = \hat{D}_t^{-1} y_t$. And since $R_t = Q_t^{*-1} Q_t Q_t^{*-1}$ where $Q_t^* = \text{diag}(\sqrt{q_{iit}})$

$$\log L(\varphi | \hat{\vartheta}, y_t) = -\frac{1}{2} \sum_{t=1}^T \left(N \log(2\pi) + \log |Q_t^{*-1} Q_t Q_t^{*-1}| + 2 \log |\hat{D}_t| + \varepsilon_t^{*'} (Q_t^{*-1} Q_t Q_t^{*-1})^{-1} \varepsilon_t^* \right).$$

Excluding the constant terms in the loglikelihood function we have

$$\log L'(\varphi | \hat{\vartheta}, y_t) = -\frac{1}{2} \sum_{t=1}^T \left(\log |Q_t^{*-1} Q_t Q_t^{*-1}| + \varepsilon_t^{*'} (Q_t^{*-1} Q_t Q_t^{*-1})^{-1} \varepsilon_t^* \right) \quad (10)$$

$$\hat{\varphi} = \arg \max[\log L'(\varphi | \hat{\vartheta}, y_t)]. \quad (11)$$

Expanding the 2nd stage loglikelihood function

$$\log L'(\varphi | \hat{\vartheta}, y_t) \quad (12)$$

$$\begin{aligned}&= -\frac{1}{2} \sum_{t=1}^T \left\{ \log |Q_t^{*-1} \left((\bar{Q} - \alpha(L) \circ \bar{Q} - \beta(L) \circ \bar{Q} - \eta(L) \circ \bar{N}) + \alpha(L) \circ (\varepsilon_t^* \varepsilon_t^{*'}) + \beta(L) \circ Q_t + \eta(L) \circ (n_t n_t') \right) Q_t^{*-1}| \right. \\ &\left. + \varepsilon_t^{*'} \left(Q_t^{*-1} \left((\bar{Q} - \alpha(L) \circ \bar{Q} - \beta(L) \circ \bar{Q} - \eta(L) \circ \bar{N}) + \alpha(L) \circ (\varepsilon_t^* \varepsilon_t^{*'}) + \beta(L) \circ Q_t + \eta(L) \circ (n_t n_t') \right) Q_t^{*-1} \right)^{-1} \varepsilon_t^* \right\}\end{aligned}$$

where $Q_t = (\bar{Q} - \alpha(L) \circ \bar{Q} - \beta(L) \circ \bar{Q} - \eta(L) \circ \bar{N}) + \alpha(L) \circ (\varepsilon_t^* \varepsilon_t^{*'}) + \beta(L) \circ Q_t + \eta(L) \circ (n_t n_t')$, is a long expression and tedious for large dimensional vector of asset returns. In the optimization process the variance targeting constraint of Engle and Mezrich (1996) is invoked.

III. SIMULATION RESULTS

Monte Carlo simulations were conducted to determine the consistency of the MLE estimator as the sample size is increased for the ABDCC in the case of a 4-dimensional vector return series, partitioned into two 2-dimensional vectors where lag orders \bar{q} , \bar{p} and \bar{r} of L are all equal to 1 or ABDCC(2,2;1,1,1). The general notation of the model is ABDCC($m_1, \dots, m_w; \bar{p}, \bar{q}, \bar{r}$). Three cases were considered: nonpersistent-nonpersistent, persistent-nonpersistent and persistent-persistent blocks of volatility series where each block consists of two series. A GARCH(1,1) process, $h_{ii,t} = \omega_i + \alpha_i \varepsilon_{i,t-1}^2 + \beta_i h_{ii,t-1}$, is considered persistent if $\alpha_i + \beta_i \geq 0.80$, the parameters α_i and β_i are the impact and persistence of the process. Similarly, the evolution process, Q_t , of the conditional correlation consists of impact, persistence and asymmetric effect parameters α_{ij} , β_{ij} and η_{ij} , respectively. Conditional correlation is considered persistent within a block if $\alpha_{ii} + \beta_{ii} \geq 0.80$ and between blocks if $\alpha_{ij} + \beta_{ij} \geq 0.80$. The specifications of the model are given below.

Nonpersistent-Nonpersistent case:

$$\begin{array}{llll}
 h_{11,t} = 0.07 + 0.10\varepsilon_{1,t-1}^2 + 0.40h_{11,t-1} & \alpha_{11} = 0.08 & \alpha_{22} = 0.05 & \alpha_{12} = \alpha_{21} = 0.06 \\
 h_{22,t} = 0.08 + 0.12\varepsilon_{2,t-1}^2 + 0.45h_{22,t-1} & \beta_{11} = 0.43 & \beta_{22} = 0.40 & \beta_{12} = \beta_{21} = 0.47 \\
 h_{33,t} = 0.08 + 0.08\varepsilon_{3,t-1}^2 + 0.37h_{33,t-1} & \eta_{11} = 0.04 & \eta_{22} = 0.03 & \eta_{12} = \eta_{21} = 0.02 \\
 h_{44,t} = 0.11 + 0.07\varepsilon_{4,t-1}^2 + 0.42h_{44,t-1} & & &
 \end{array}$$

Persistent-Nonpersistent case:

$$\begin{array}{llll}
 h_{11,t} = 0.10 + 0.15\varepsilon_{1,t-1}^2 + 0.70h_{11,t-1} & \alpha_{11} = 0.14 & \alpha_{22} = 0.07 & \alpha_{12} = \alpha_{21} = 0.11 \\
 h_{22,t} = 0.12 + 0.12\varepsilon_{2,t-1}^2 + 0.80h_{22,t-1} & \beta_{11} = 0.73 & \beta_{22} = 0.48 & \beta_{12} = \beta_{21} = 0.54 \\
 h_{33,t} = 0.07 + 0.09\varepsilon_{3,t-1}^2 + 0.44h_{33,t-1} & \eta_{11} = 0.08 & \eta_{22} = 0.04 & \eta_{12} = \eta_{21} = 0.07 \\
 h_{44,t} = 0.09 + 0.04\varepsilon_{4,t-1}^2 + 0.51h_{44,t-1} & & &
 \end{array}$$

Persistent-Persistent case:

$$\begin{array}{llll}
 h_{11,t} = 0.12 + 0.17\varepsilon_{1,t-1}^2 + 0.73h_{11,t-1} & \alpha_{11} = 0.15 & \alpha_{22} = 0.19 & \alpha_{12} = \alpha_{21} = 0.12 \\
 h_{22,t} = 0.11 + 0.14\varepsilon_{2,t-1}^2 + 0.81h_{22,t-1} & \beta_{11} = 0.79 & \beta_{22} = 0.65 & \beta_{12} = \beta_{21} = 0.61 \\
 h_{33,t} = 0.09 + 0.20\varepsilon_{3,t-1}^2 + 0.67h_{33,t-1} & \eta_{11} = 0.08 & \eta_{22} = 0.09 & \eta_{12} = \eta_{21} = 0.13 \\
 h_{44,t} = 0.08 + 0.22\varepsilon_{4,t-1}^2 + 0.64h_{44,t-1} & & &
 \end{array}$$

Table 1 summarizes the bias and MSE of the estimates in 500 Monte Carlo runs of all the parameters for series lengths of 500, 1,000, and 1,500 in the three different cases. As the sample size is increased the resulting bias of the estimates is mixed for the different parameters although the magnitudes of the bias for all the parameters are tolerable. Tse and Tsui (2002) considered such bias small in their paper where bias is typically 10% or less of the actual value. The bias is expected for parameters that are constrained in the parameter space Θ^9 where $\Theta \in (0,1)$. Engle and Sheppard (2001) placed such constraint as a sufficient

condition for the evolution process, Q_t , to remain positive definite. As the sample size is increased, the MSE dropped for all the parameters in all cases considered and so confirms the consistency of the maximum likelihood estimator for the model.

In order to assess how well the ABDCC model performs in in-sample and out-of-sample forecasting of conditional correlation between blocks of asset returns, in the absence and presence of asymmetric effects within and between conditional correlation of blocks, its performance vis-à-vis the other dynamic conditional correlations models DCC, BDCC, ADCC and AGDCC is compared. Below are specifications for three main cases: Case A nonpersistent-nonpersistent, Case B persistent-nonpersistent and Case C persistent-persistent conditional correlations between blocks of assets for a series of 1,000 observations. There were nine cases considered in these simulations and for each 500 Monte Carlo runs were conducted. Cases 1, 2 and 3 consists of cases A, B and C, respectively, but with zero asymmetric effects, that is, $\eta_{11} = \eta_{22} = \eta_{12} = 0$. Cases 4, 5 and 6 follow the exact specifications of A, B and C which have small between block asymmetric effects, η_{12} 's. While cases 7, 8 and 9, on the other hand, have the same specification as A, B and C but with larger η_{12} 's of 0.12, 0.15 and 0.15, respectively.

Case A:

$$\begin{aligned}
 h_{11,t} &= 0.08 + 0.12\varepsilon_{1,t-1}^2 + 0.55h_{11,t-1} \\
 h_{22,t} &= 0.11 + 0.10\varepsilon_{2,t-1}^2 + 0.59h_{22,t-1} \\
 h_{33,t} &= 0.15 + 0.08\varepsilon_{3,t-1}^2 + 0.64h_{33,t-1} \\
 h_{44,t} &= 0.12 + 0.11\varepsilon_{4,t-1}^2 + 0.52h_{44,t-1} \\
 \alpha_{11} &= 0.07 & \alpha_{22} &= 0.06 & \alpha_{12} &= \alpha_{21} &= 0.04 \\
 \beta_{11} &= 0.51 & \beta_{22} &= 0.58 & \beta_{12} &= \beta_{21} &= 0.45 \\
 \eta_{11} &= 0.04 & \eta_{22} &= 0.04 & \eta_{12} &= \eta_{21} &= 0.03
 \end{aligned}$$

Case B:

$$\begin{aligned}
 h_{11,t} &= 0.05 + 0.05\varepsilon_{1,t-1}^2 + 0.80h_{11,t-1} \\
 h_{22,t} &= 0.10 + 0.07\varepsilon_{2,t-1}^2 + 0.75h_{22,t-1} \\
 h_{33,t} &= 0.30 + 0.20\varepsilon_{3,t-1}^2 + 0.55h_{33,t-1} \\
 h_{44,t} &= 0.40 + 0.17\varepsilon_{4,t-1}^2 + 0.60h_{44,t-1} \\
 \alpha_{11} &= 0.15 & \alpha_{22} &= 0.12 & \alpha_{12} &= \alpha_{21} &= 0.05 \\
 \beta_{11} &= 0.60 & \beta_{22} &= 0.50 & \beta_{12} &= \beta_{21} &= 0.40 \\
 \eta_{11} &= 0.05 & \eta_{22} &= 0.05 & \eta_{12} &= \eta_{21} &= 0.05
 \end{aligned}$$

Case C:

$$h_{11,t} = 0.05 + 0.06\varepsilon_{1,t-1}^2 + 0.90h_{11,t-1}$$

$$h_{22,t} = 0.10 + 0.08\varepsilon_{2,t-1}^2 + 0.85h_{22,t-1}$$

$$h_{33,t} = 0.08 + 0.20\varepsilon_{3,t-1}^2 + 0.68h_{33,t-1}$$

$$h_{44,t} = 0.11 + 0.17\varepsilon_{4,t-1}^2 + 0.72h_{44,t-1}$$

$$\alpha_{11} = 0.05 \quad \alpha_{22} = 0.15 \quad \alpha_{12} = \alpha_{21} = 0.05$$

$$\beta_{11} = 0.82 \quad \beta_{22} = 0.65 \quad \beta_{12} = \beta_{21} = 0.55$$

$$\eta_{11} = 0.07 \quad \eta_{22} = 0.08 \quad \eta_{12} = \eta_{21} = 0.06$$

Table 2 shows the in-sample forecast performance of the different models in capturing simulated conditional correlations between two blocks of asset returns. The results show that the ADCC of Cappiello, Engle and Sheppard (2003) outperformed all the other DCC models for cases 4 to 9 in terms of the lowest average root mean squared error (RMSE) criterion in forecasting each of the 500 generated series. ADCC is best in in-sampling forecast performance for those with persistent and nonpersistent conditional correlations with small and large asymmetric effects. The ABDCC follows next for cases 5 to 9. However, the simple DCC is competitive compared to the other models with regard to nonpersistent blocks with small asymmetric effects, as seen in case 4; and it outperforms more parameterized models in nonpersistent and persistent cases without asymmetric effects as evidenced by cases 1 to 3.

For out-of-sample forecast evaluation, 500 Monte Carlo runs of 1,200 series lengths of between block conditional correlations were generated for the nine cases. The last 200 observations serve as the hold-out sample that all the models will forecast 200-step-ahead out-of-sample. Table 3 presents the average mean squared errors (MSEs) for the 500 forecast series of each of the model for each case. The DCC and BDCC models dominate those cases without asymmetric effects, cases 1 to 3; and, these can compete fairly with the others for the case of nonpersistent-nonpersistent and persistent-persistent blocks with small asymmetric effects, respectively, as seen in cases 4 and 6. But clearly, in the presence of a persistent block or blocks with small and large asymmetric effects ABDCC outperforms the rest with ADCC coming in next. These results lend support to the ABDCC model as the model of choice in out-of-sample forecasting of persistent conditional correlation between blocks of asset returns with asymmetric effects. This behavior is widely observed in portfolios of stocks providing a venue for practical application of ABDCC in asset allocation models.

To test whether there is a significant difference in forecast accuracy of the two models in Table 3 with the lowest average MSEs, the Diebold and Mariano (1995) test or DM test for forecasting accuracy and the modified DM test (MDM) of Harvey, Leybourne and Newbold (1997) were used to evaluate the performance of these two tests. Table 4 shows the proportion of rejection of the null hypothesis of equal forecast accuracy in each test conducted on the 500 corresponding MSEs of the two models. The DM and MDM tests rejected the null hypothesis in 81% to 85% of the 500 comparisons in favor of the DCC

model compared to BDCC and ABDCC for cases 2 and 4, respectively. While the BDCC performed 90% of the time better than DCC in case 1 but did poorly for case 3 with respect to ABDCC at 65%. The ABDCC edged out ADCC in cases 5 and 7 to 9 by 83% to 97% of the time, while it did not do as well with BDCC at 73% in case 6.

The DM and MDM tests of forecast accuracy confirmed that, in general, when persistence of conditional correlation and asymmetric effect exist within and between blocks of asset returns the ABDCC model is expected to perform better than the rest in out-of-sample forecasting. At the same time, the test results favor the parsimonious DCC and also the BDCC in cases where asymmetric effect is absent. This is expected since DCC and BDCC models assume symmetry in the conditional correlations.

IV. APPLICATION TO EAST ASIAN FOREX RETURNS

The data was derived from FXHistory in oanda.com. These are 1,279 daily average exchange rates with respect to the U.S. dollar from 1996 to 2000 of the Philippine peso (PHP), Thailand baht (THB), Malaysian ringgit (MYR), Singapore dollar (SGD), Hong Kong dollar (HKD) and China yuan (CNY). Adjustments have been made to account for holidays. An ABDCC(2,2;1,1,1) was used to model the following blocks: PHP-THB, MYR-SGD and HKD-CNY. The Philippine peso and Thailand baht are considered as a block because these two economies are regarded as *twins* in the ASEAN. On the other hand, Malaysian ringgit and Singaporean dollar are in another block given that these two economies are relatively larger compared to Philippines and Thailand. Hong Kong dollar and Chinese yuan belong to a block for these two economies are reasonably linked since Hong Kong became a special administrative region of China in July of 1997. Asset return here is defined as the negative of the difference in daily logarithmic forex rate.

Tables 5, 6 and 7 show the conditional correlations of the blocks of currencies. All the parameters are significant at the 0.05 level except for the asymmetric effect parameter between the PHP-THB and MYR-SGD blocks as shown in Table 5. This implies that there is no significant asymmetric correlation in the volatilities of the peso-baht and ringgit-SG dollar although the conditional correlations are highly persistent between these two blocks. Asymmetry in conditional correlations is present for these two blocks when compared with the HK dollar-yuan block as presented in Tables 6 and 7. Table 7 further shows that the correlation of volatility is persistent between HK dollar-yuan and ringgit-SG dollar owing largely to the level of development and relative stability of the financial markets in Hong Kong and Singapore while HK dollar-yuan and peso-baht correlation of volatility are not as persistent. There is also a higher asymmetry in the conditional correlation of HK dollar-yuan and peso-baht.

To determine the conditional correlation between blocks, the estimated conditional variances, $\hat{q}_{ij,t}$'s, within a block were pooled to get the estimated pooled conditional variance for that block. The estimated pooled conditional covariance was also computed between blocks. Figures 1, 2 and 3 show the conditional correlations between blocks. The asymmetric block DCC is clearly positive for the peso-baht and ringgit-SG dollar in Figure 1 for the period considered which indicate the linkage of the economies of these four ASEAN countries. At

the same time, there was a sharp rise in the conditional correlation of the two blocks during the 1997-98 Asian Financial Crisis. The drop in the conditional correlation occurred in September of 1998 when Malaysia imposed capital controls to discourage speculators against its currency. In Figures 2 and 3, the conditional correlations of peso-baht and ringgit-SG dollar against HK dollar-yuan is fluctuating around zero which indicate periods of positive and negative conditional correlations. But in Figure 3, the 1997-98 period resulted in higher volatility of the conditional correlation between ringgit-SG dollar and HK dollar-yuan and this volatility in conditional correlation relatively mellowed after 1998 when the ringgit exchange rate was pegged and the SG-dollar has stabilized. This is dissimilar to the pattern between peso-baht and HK dollar-yuan where the relatively stable exchange rates of the HK dollar and the yuan are fluctuating similarly across the period against the peso-baht.

The significant asymmetries of the ASEAN blocks with respect to the HK dollar-yuan block may be explained by the relatively stable movement of these two currencies compared to their ASEAN counterparts in the period considered. Hong Kong operates a form of currency board system which composes a pegged exchange rate against the U.S. dollar. The Chinese yuan is also pegged to a tight band since 1994. It follows that whatever pressure to depreciation on non-pegged currencies will not impact the pegged ones. And this is evident in the asymmetries that are observed here with respect to the HKD-CNY block. The insignificant asymmetry between the ASEAN blocks implies that although the correlation of the volatilities of the peso-baht and the ringgit-SG dollar is persistent there is no asymmetric effect in the dynamic conditional correlation of the returns between the two blocks.

V. CONCLUSION

The Asymmetric Block Dynamic Conditional Correlation model, a new model which extends the Block DCC of Billio, Caporin and Gobbo (2003), introduces the asymmetric effect in the evolution process between blocks of asset returns. The two-stage maximum likelihood estimation procedure utilizing the limited information maximum likelihood approach is shown to be consistent in estimating the parameters of the ABDCC model. Furthermore, simulation results support the ABDCC as a better model compared to other DCCs in out-of-sample forecasting of conditional correlations in the presence of asymmetric effect in dynamic conditional correlation between blocks of asset returns.

The model was able to capture the behavior of six Asian currency returns considered to reflect the turbulent times of the Asian Financial Crisis in 1997-98. The peso-baht and ringgit-SG dollar blocks are positively conditionally correlated, this conditional correlation rose sharply during the 1997-98 Asian Financial Crisis. The conditional correlation of peso-baht and ringgit-SG dollar blocks volatilities is highly persistent but not asymmetric which confirms the linkage of these ASEAN currencies. On the other hand, the HK dollar-yuan block saw periods of highly positive and highly negative conditional correlation with respect to the peso-baht and ringgit-SG dollar blocks and the model has shown the presence of asymmetry in the block volatilities of the HK dollar-yuan with the ASEAN blocks due to the generally pegged HK dollar-yuan rates with the U.S. dollar. Also, the volatility of the HK dollar-yuan is highly asymmetric against the peso-baht and is highly persistent with the ringgit-SG dollar. The conditional correlation between the HK dollar-yuan and ringgit-SG dollar blocks are especially high during the crisis years.

**Table 1.1. Monte Carlo Simulation Results for Varying Sample Sizes of ABDCC(2,2;1,1,1)
Model: Nonpersistent-Nonpersistent Blocks**

Block	Actual Value	Series Length	Bias	MSE
Within Block 1				
α_{11}	0.08	500	-0.003809	0.000166
		1000	-0.004972	0.000084
		1500	-0.005246	0.000065
β_{11}	0.43	500	0.004801	0.004376
		1000	-0.001105	0.001664
		1500	-0.005583	0.001024
η_{11}	0.04	500	-0.000749	0.000262
		1000	-0.000026	0.000110
		1500	-0.000043	0.000067
Within Block 2				
α_{22}	0.05	500	0.003376	0.000114
		1000	0.002025	0.000047
		1500	0.001844	0.000026
β_{22}	0.40	500	0.054681	0.010422
		1000	0.031651	0.003922
		1500	0.019816	0.002329
η_{22}	0.03	500	-0.002382	0.000143
		1000	-0.002701	0.000069
		1500	-0.002479	0.000045
Between Blocks 1 & 2				
α_{12}	0.06	500	-0.003043	0.001135
		1000	-0.004683	0.000529
		1500	-0.004768	0.000332
β_{12}	0.47	500	-0.068321	0.050844
		1000	-0.043553	0.024045
		1500	-0.030205	0.014946
η_{12}	0.02	500	-0.007086	0.001877
		1000	-0.004606	0.000804
		1500	-0.005783	0.000558

Table 1.2. Monte Carlo Simulation Results for Varying Sample Sizes of ABDCC(2,2;1,1,1) Model: Persistent-Nonpersistent Blocks

Block	Actual Value	Series Length	Bias	MSE
Within Block 1				
α_{11}	0.14	500	-0.008456	0.000151
		1000	-0.007377	0.000096
		1500	-0.007061	0.000080
β_{11}	0.73	500	-0.017922	0.000522
		1000	-0.016023	0.000338
		1500	-0.014865	0.000286
η_{11}	0.08	500	0.005639	0.000195
		1000	0.005305	0.000119
		1500	0.005888	0.000107
Within Block 2				
α_{22}	0.07	500	0.005026	0.000101
		1000	0.004404	0.000052
		1500	0.004175	0.000036
β_{22}	0.48	500	0.028977	0.003126
		1000	0.014804	0.001173
		1500	0.007542	0.000703
η_{22}	0.04	500	-0.003247	0.000123
		1000	-0.003380	0.000059
		1500	-0.003330	0.000043
Between Blocks 1 & 2				
α_{12}	0.11	500	-0.007472	0.000482
		1000	-0.008521	0.000255
		1500	-0.008505	0.000194
β_{12}	0.54	500	-0.041871	0.005670
		1000	-0.038906	0.003129
		1500	-0.036490	0.002362
η_{12}	0.07	500	-0.002862	0.000574
		1000	-0.000277	0.000260
		1500	-0.000180	0.000186

**Table 1.3. Monte Carlo Simulation Results for Varying Sample Sizes of ABDCC(2,2;1,1,1)
Model: Persistent-Persistent Blocks**

Block	Actual Value	Series Length	Bias	MSE
Within Block 1				
α_{11}	0.15	500	-0.002340	0.000040
		1000	-0.001537	0.000016
		1500	-0.001145	0.000010
β_{11}	0.79	500	-0.000861	0.000036
		1000	0.000217	0.000013
		1500	0.000200	0.000007
η_{11}	0.08	500	0.001164	0.000077
		1000	0.000701	0.000029
		1500	0.000827	0.000019
Within Block 2				
α_{22}	0.19	500	0.005067	0.000077
		1000	0.003018	0.000033
		1500	0.002200	0.000019
β_{22}	0.65	500	-0.005330	0.000133
		1000	-0.004166	0.000062
		1500	-0.003992	0.000045
η_{22}	0.09	500	-0.002042	0.000077
		1000	-0.002772	0.000040
		1500	-0.002817	0.000027
Between Blocks 1 & 2				
α_{12}	0.12	500	-0.010866	0.000289
		1000	-0.009514	0.000185
		1500	-0.008090	0.000134
β_{12}	0.61	500	-0.019136	0.000873
		1000	-0.015233	0.000510
		1500	-0.012096	0.000328
η_{12}	0.13	500	0.007745	0.000279
		1000	0.008206	0.000173
		1500	0.007510	0.000131

Table 2. Average RMSEs of the Different DCC Models in 500 Monte Carlo Simulations of In-Sample Forecasts in Estimating Simulated Block Dynamic Conditional Correlations

Case	DCC(1,1)	BDCC(2,2;1,1)	ADCC(1,1,1)	AGDCC(2,2;1,1,1)	ABDCC(2,2;1,1,1)
1	0.01458	0.01775	0.01795	0.09336	0.01867
2	0.01509	0.01836	0.01882	0.09850	0.01926
3	0.01547	0.01897	0.01951	0.08000	0.01929
4	0.01823	0.02086	0.01578	0.08666	0.02047
5	0.02144	0.02518	0.01745	0.08775	0.02130
6	0.02575	0.02800	0.01971	0.09084	0.02298
7	0.04366	0.04145	0.02281	0.07186	0.02448
8	0.04893	0.04537	0.02602	0.07537	0.02701
9	0.05353	0.04913	0.02926	0.08230	0.02932

Highlighted Average RMSEs are the two lowest in each case.

Table 3. Average MSEs of the Different DCC Models in 500 Monte Carlo Simulations of 200-Step-Ahead Out-of-Sample Forecasts in Estimating Simulated Block Dynamic Conditional Correlations

Case	DCC(1,1)	BDCC(2,2;1,1)	ADCC(1,1,1)	AGDCC(2,2;1,1,1)	ABDCC(2,2;1,1,1)
1	0.000345	0.000337	0.000514	0.004908	0.000787
2	0.000328	0.000367	0.000414	0.004156	0.000399
3	0.000469	0.000395	0.000509	0.001558	0.000413
4	0.000440	0.000506	0.000455	0.003909	0.000443
5	0.000604	0.000594	0.000478	0.002840	0.000477
6	0.001146	0.000769	0.000878	0.001138	0.000535
7	0.001968	0.001609	0.001021	0.002114	0.000596
8	0.002699	0.002067	0.001112	0.001746	0.000719
9	0.003879	0.002367	0.002071	0.002301	0.000830

Highlighted Average MSEs are the two lowest in each case.

Table 4. Proportion of H_0 Rejections in the Diebold-Mariano and the Modified Diebold-Mariano Tests in 500 Monte Carlo Simulations of 200-Step-Ahead Out-of-Sample Forecasts of the Top 2 Models from Table 3 in Estimating Simulated Block Dynamic Conditional Correlations

Case	MSE				DM Test	MDM Test
	DCC(1,1)	BDCC(2,2;1,1)	ADCC(1,1,1)	ABDCC(2,2;1,1,1)		
1	0.000345	0.000337	–	–	0.898	0.900
2	0.000328	0.000367	–	–	0.852	0.854
3	–	0.000395	–	0.000413	0.646	0.648
4	0.000440	–	–	0.000443	0.812	0.814
5	–	–	0.000478	0.000477	0.900	0.902
6	–	0.000769	–	0.000535	0.726	0.728
7	–	–	0.001021	0.000596	0.966	0.968
8	–	–	0.001112	0.000719	0.942	0.944
9	–	–	0.002071	0.000830	0.830	0.832

Highlighted Average MSE is the smaller between the two in each case.

Table 5. ABDCC(2,2;1,1,1) Model of PHP-THB and MYR-SGD Forex Returns

Block $\hat{\alpha}$ (s.e.) $\hat{\beta}$ (s.e.) $\hat{\eta}$ (s.e.)	PHP-THB	MYR-SGD
PHP-THB	0.164128 (0.002960) 0.523677 (0.003629) 0.046345 (0.003323)	0.041819 (0.000753) 0.824190 (0.002214) 0.002132 (0.003567)
MYR-SGD		0.129147 (0.001858) 0.508595 (0.004072) 0.087879 (0.002228)

All highlighted parameter estimates are significant at the 0.05 level.

Table 6. ABDCC(2,2;1,1,1) Model of PHP-THB and HKD-CNY Forex Returns

Block $\hat{\alpha}$ (s.e.) $\hat{\beta}$ (s.e.) $\hat{\eta}$ (s.e.)	PHP-THB	HKD-CNY
PHP-THB	0.126281 (0.002443) 0.510580 (0.008824) 0.124693 (0.000132)	0.207637 (0.002027) 0.501463 (0.001917) 0.219513 (0.000512)
HKD-CNY		0.195552 (0.002426) 0.482574 (0.003070) 0.291877 (0.003016)

All highlighted parameter estimates are significant at the 0.05 level.

Table 7. ABDCC(2,2;1,1,1) Model of HKD-CNY and MYR-SGD Forex Returns

Block		
$\hat{\alpha}$ (s.e.)		
$\hat{\beta}$ (s.e.)	HKD-CNY	MYR-SGD
$\hat{\eta}$ (s.e.)		
HKD-CNY	0.150419 (0.000233) 0.574708 (0.000624) 0.194977 (0.001349)	0.238101 (0.000199) 0.683436 (0.000378) 0.118465 (0.000325)
MYR-SGD		0.119230 (0.000485) 0.570860 (0.000473) 0.087742 (0.000376)

All highlighted parameter estimates are significant at the 0.05 level.

Figure 1. Asymmetric Block DCC of PHP-THB and MYR-SGD Blocks of Forex Returns

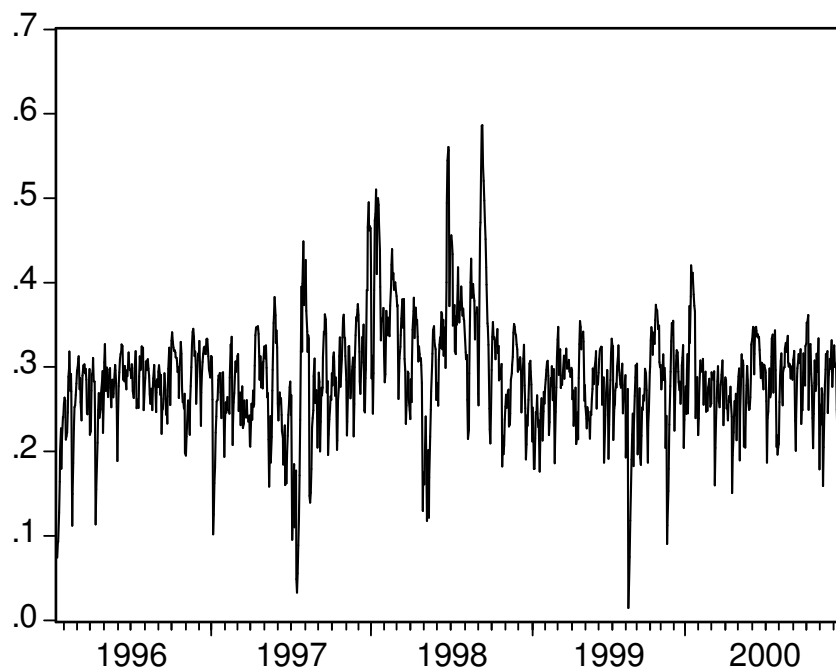


Figure 2. Asymmetric Block DCC of PHP-THB and HKD-CNY Blocks of Forex Returns

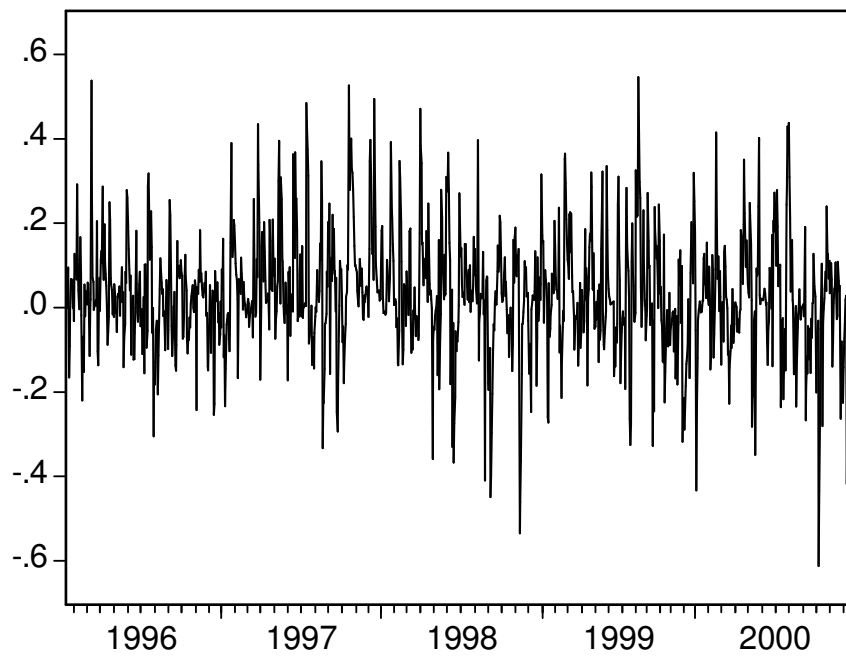
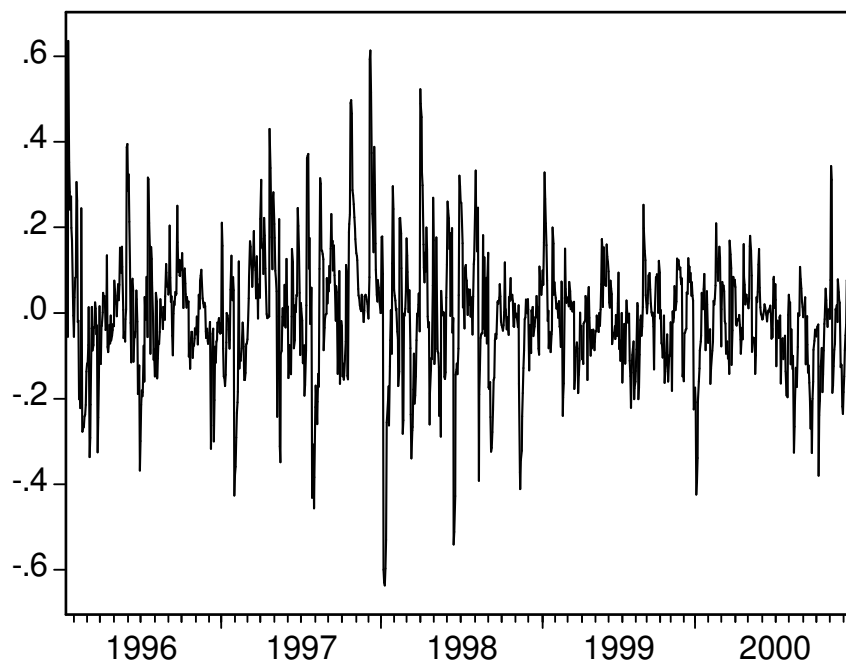


Figure 3. Asymmetric Block DCC of HKD-CNY and MYR-SGD Blocks of Forex Returns



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