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Mortality Risk Reduction Differs According to Bisphosphonate Class: A 15-Year Observational Study

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ABSTRACT

Introduction/Purpose: Emerging evidence suggests that bisphosphonates (BP), first-line treatment of osteoporosis, are associated with reduced risks for all-cause mortality.

This study aimed to determine the association between different BP types and mortality risk in participants with or without a fracture.

Methods: A prospective cohort study of users of different BPs matched to non-users by propensity score (age, gender, co-morbidities, fragility fracture status) and time to starting the BP medication from the population-based Canadian Multicentre Osteoporosis Study from nine Canadian centres followed from 1995 to 2013.

Mortality risk for bisphosphonate users vs matched non-users was assessed using pairwise multivariable Cox proportional hazards models.

Results: There were 2048 women and 308 men on BP and 1970 women and 1794 men who did not receive medication for osteoporosis. The relationship between BP and mortality risk was explored in three separate 1:1 propensity score-matched cohorts of BP users and no treatment (etidronate, n=599, alendronate, n=498, and risedronate n=213). Nitrogen BP (n-BP) (alendronate and risedronate) was associated with lower mortality risks [pairwise HR, 0.66 (95% CI, 0.48-0.91)] while the less potent non-n-BP, etidronate, was not [pairwise HR: 0.89 (95% CI, 0.66-1.20)]. A direct comparison between n-BP and etidronate (n=340 pairs) also suggested a better survival for n-BP [paired HR, 0.47 (95% CI, 0.31-0.70)] for n-BP vs. etidronate].

Conclusion: Compared to no treatment, nitrogen but not non-nitrogen bisphosphonates appear to be associated with better survival.

Mini-abstract (50 words or less)

In this prospective cohort of 6120 participants aged 50+, nitrogen-bisphosphonates but not non-nitrogen bisphosphonates were associated with a significant 34% mortality risk reduction compared to non-treated propensity score matched controls. These findings open new avenues for research into mechanistic pathways.

INTRODUCTION

Osteoporotic fragility fracture is highly prevalent in the general population and is associated with serious consequences. From the age of 50, 40% of women and 25% of men will sustain a fragility fracture (trauma less than or equal to a fall from standing) during their remaining lifetimes¹. Men and women with a fracture have increased risk of further fractures²⁻⁴ and most importantly, premature mortality⁵⁻⁷. Despite the availability of effective medications, treatment rates continue to be low with <30% women and <20% men with fragility fractures on validated treatments. Bisphosphonates, first-line treatment for osteoporosis world-wide^{8,9}, are effective in reducing the relative risk of fracture by between 40 to 70%¹⁰ and also appear to confer a survival benefit among patients with a fracture^{11,12} based on a randomised controlled trial (RCT) of hip fracture patients¹³⁻¹⁵ and several cohort¹⁶⁻¹⁸, registry-based studies^{14,15}, and more recently in a Fracture Liaison Service setting¹⁹. In the RCT, hip fracture subjects given zoledronic acid had a 28% reduced mortality¹¹. A meta-analysis of anti-osteoporosis medications from eight RCTs found a pooled mortality risk benefit (~11 %) of these agents¹². More recently, zoledronic acid was reported to reduce mortality risk by 35% [OR 0.65 (95% CI, 0.40-1.05)] over 6 years in a RCT of women with osteopenia²⁰.

Despite multivariate adjustment, criticism persists that, at least in cohort studies, survival benefit may relate to healthy user bias. However, RCTs with mortality as the primary outcome will likely not be conducted due to necessary large numbers, expense and particularly ethical considerations. This issue is important to resolve as, if true, it may help to increase the acceptability and uptake of urgently needed treatments²¹. A scenario in which any potential bisphosphonate-related mortality benefit could be further explored, would therefore be a cohort study that examines the effect of bisphosphonates of different chemistries expected to have different effects on all-cause mortality. Thus if a difference were found, any healthy user bias would be avoided as indications for treatment would be similar.

There are two main classes of bisphosphonates: nitrogen bisphosphonates (n-BP, *e.g.* alendronate, risedronate) and non-n-BP (*e.g.* etidronate). The newer bisphosphonates (n-BP) have a different mechanism of action and are more potent than the non-nitrogen bisphosphonates^{22,23}. The higher potency of n-BP result in a greater reduction of bone loss that would limit the resorption-related release from bone of toxic substances (*e.g.* lead)^{24,25}. Furthermore, several studies have suggested that n-BP may have non-bone beneficial effects

1 such as on immune function^{26,27} (although adverse immune effects also occur²⁸), endothelial
2 function²⁹, systemic inflammation³⁰ and an antitumor effect³¹.

3 Our hypothesis was that participants on bisphosphonates would have a better survival than
4 those on no treatment and that the more potent nitrogen-bisphosphonate may have a greater
5 effect than the non-nitrogen bisphosphonates. This study therefore examined the association
6 between bisphosphonates of two different chemistries with all-cause mortality in a
7 population-based cohort of women and men aged 50 years and older.
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13 **METHODS**

14 **Subjects and setting**

15 The study population consisted of women and men participating in the Canadian Multicentre
16 Osteoporosis Study (CaMos), an ongoing prospective population-based study that started in
17 1995 with the aim to document the skeletal health of a randomly selected population of
18 women and men aged 25 and over. All non-institutionalised Canadians who resided within 50
19 km of a study centre, representing ~37% of all Canadians were eligible. Participants were
20 recruited using randomly generated telephone lists from the region surrounding nine urban
21 centres in Canada. A detailed description of the study design and population sampling has
22 been published previously³². CaMos was approved by the Ethics Committee of McGill
23 University and at each participating centre.
24

25 Of the 9,423 participants recruited, 7,689 aged 50+ were screened for medication uptake.
26 CaMOS is an observational study, thus all the medication was initiated by each participant's
27 physician without any intervention from the CaMOS investigators. Etidronate and
28 alendronate received Canadian regulatory approval for osteoporosis treatment within a year
29 of each other, and prior to the start of CaMOS. In most Canadian provincial drug plans,
30 access to alendronate (and risedronate) was restricted to patients who had already suffered an
31 osteoporotic fracture; or had either failed to respond to etidronate (had lost bone density or
32 suffered a new fracture) or were not able to tolerate etidronate. This is reflected in this
33 observational study by the large number of participants (~40%) who switched between
34 bisphosphonate types during the follow-up (Figure 1). To account for any potential immortal
35 time bias induced by this switch, the primary aim was investigated in the groups not treated
36 versus those treated with only one type of bisphosphonate for the entire follow-up.
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2 *Inclusion criteria:* Individuals who used bisphosphonates during the study follow-up
3 (etidronate, alendronate, and risedronate), and those who did not use any osteoporosis-related
4 medication (NoRx) were included.
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7 *Exclusion criteria:* A number of osteoporosis related medications were excluded due to small
8 number of users (clodronate, n=22, pamidronate, n=54, zoledronic acid, n=44, calcitonin,
9 n=14, denosumab, n=2, raloxifene, n=50), tamoxifen (n=100), and testosterone (n=39)
10 (Figure 1). A relatively large number of women reported hormone therapy (N=1268) at
11 baseline or throughout the study (Figure 1). This group of women had more favourable
12 characteristics than women who did not take any medication. They were younger, had higher
13 BMD, a higher proportion of distal compared to proximal fractures, and also had better
14 lifestyle habits (less smoking, more exercise, and more were taking vitamin D). Given the
15 unknown duration of prior exposure and the potential effect on cardiovascular risk, this group
16 was excluded from further analyses. However, after adjustment for baseline characteristics,
17 their survival was not significantly better than those on no treatment.
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28 **Outcomes and risk factors**

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30 A standardized interviewer administered questionnaire was obtained at baseline (1995-97).
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32 Information was obtained on lifestyle factors (i.e. smoking, physical activity), demographics,
33 education, co-morbidities and medication use. In addition to this structured questionnaire,
34 each participant had a clinical visit that included anthropometric measurements (i.e. height,
35 weight) and femoral neck areal bone mineral density (BMD). This information was
36 subsequently obtained in Years 3 (40 – 60 years of age only), 5 and 10. Yearly postal self-
37 administered questionnaires for incident fractures and medications were obtained between
38 clinical visits.
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46 *Bisphosphonate exposure*

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48 Bisphosphonate uptake was determined from yearly questionnaires and the inventory of
49 medications brought to each interview (baseline and Years 3, 5 and 10). Participants were
50 classified as bisphosphonate users based on yearly report of medication. Of the 2356 eligible
51 bisphosphonate users, 985 participants used more than one of type of bisphosphonate (Figure
52 1). Thus, 50% participants initiated on etidronate, 33% on alendronate and 14% on
53 risedronate switched during the study follow-up to another class of bisphosphonate. These
54 participants were included in a sensitivity analysis, classified according to the first
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1 bisphosphonate used. The treatment initiation date was taken as the year of the first reported
2 use of bisphosphonate.
3 The uptake of bisphosphonate during follow-up was much lower in men (~14%) than women
4 (~40%). Therefore we have performed two analyses: “any user” including both genders, and
5
6 women only.
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8 Adherence to bisphosphonates was not recorded. Participants who reported bisphosphonates
9 only once during the follow-up (n=251) were used as surrogate for non-adherence in a
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11 sensitivity analysis.
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16 Participants did not receive any formal fracture risk assessment or management suggestions
17 from the CaMOS investigators. They, and/or their primary care physician received a copy of
18 the BMD report performed at baseline and all subsequent visits.
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22 *Fracture ascertainment*

23 Self-reported incident clinical fractures were obtained yearly and at clinical visits.

24 Information on the date, site, circumstance of the fracture, and an x-ray report was obtained
25
26 by interview. Medical records were obtained and verified for 78% of fractures.
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28 This study included only incident fragility fractures. Skull, sternum, finger and toe fractures
29
30 were excluded.
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33 *Mortality ascertainment*

34 Mortality ascertainment was conducted annually throughout the study follow-up. All
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36 participants provided contact detail for next of kin. If a participant did not respond to the
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38 yearly questionnaire, the study co-ordinator contacted the next of kin. If this failed, obituaries
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40 were screened for death records. Although mortality data were not formally validated using
41
42 national figures or other external data sources, it was highly unlikely that these deaths were
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44 misclassified.
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51 **Statistical Analysis**

52 Baseline characteristics were examined for BP (alendronate, risedronate and etidronate) in
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54 comparison to NoRx (T-tests for continuous and Chi²-square tests for categorical variables).
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56 **Whole cohort**

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1 For the primary analysis participants who used only one type of bisphosphonate during the
2 follow-up (n=1371), classified as n-BP (alendronate and risedronate) and non-n-BP
3 etidronate were matched 1:1 to non-treated participants. Matching was by a propensity score,
4 including age, gender, fracture type, co-morbidities and life-style factors that predicted the
5 likelihood of being treated³³, and time to starting medication. Follow-up was calculated from
6 the time of medication start for both treated and non-treated. For non-treated this starting
7 point was obtained by the addition to baseline date his/her “pair’s” time of medication
8 commencement. This procedure ensured that all participants had similar baseline
9 characteristics, avoiding selection bias, and that a participant who started treatment later
10 during the follow-up was matched to a control still alive at that time point, avoiding immortal
11 time bias. Mortality risk was analysed using a paired Cox proportional hazards model.
12 Proportionality hazards assumption was tested by the inspection of Schoenfeld residuals over
13 time. Kaplan Meier survival curves for each bisphosphonate were also created.

24 *Subgroup and sensitivity analysis*

26 To determine the role of individual bisphosphonate type on mortality risk, alendronate and
27 risedronate users were separately compared to no treatment using a paired Cox proportional
28 hazards model adjusted for any variable which became unbalanced after stratification.

32 To test the hypothesis that n-BPs have a stronger association with mortality reduction than
33 the non-n-BPs, a head-to-head comparison between the 2 classes of bisphosphonates was
34 performed in a set of n-BP matched 1:1 to etidronate by propensity score, using a paired Cox
35 Proportional Hazards Model.

37 A sensitivity analysis including all bisphosphonate users (n=2356), with switchers classified
38 according to initial type of bisphosphonate was performed using inverse probability
39 weighting with treatment as a time dependent variable. The period of time prior to treatment
40 initiation contributed to no treatment, while the interval following the first bisphosphonate
41 uptake contributed to treatment, in an intention to treat analysis. Thus, all bisphosphonates
42 users were classified according to the initial type of bisphosphonate, regardless if they
43 continued on the same bisphosphonate, or switched to another type during the follow-up.

45 Individuals who reported bisphosphonate only at one visit during the follow-up (n=241),
46 were used as surrogate for non-adherence and excluded in a sensitivity analysis.

1
2 **Fracture cohort**

3 A subset analysis of the relationship between the bisphosphonate initiated at or following the
4 time of fracture and mortality was performed for individuals with incident fractures. n-BP
5 and etidronate users were matched 1:2 by age, gender, and fracture type to individuals who
6 did not use any treatment after the incident fracture. Fracture risk was assessed for all
7 individuals using the Garvan fracture risk calculator³⁴. The relationship between BP and
8 survival was assessed using a paired Cox Proportional Hazard Model.
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16 *Subgroup analysis*

17 Given the high mortality occurring immediately after the fracture event, this analysis was also
18 performed according to the time of BP initiation post-fracture (0-2, 2-5 and 5+ years in n-BP
19 group only, due to small number of etidronate users 2+ years post-fracture.
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24 In order to examine whether the mortality reduction could be mediated by a reduction in
25 subsequent fracture events, an additional Cox proportional hazards model with subsequent
26 fracture as the outcome was conducted.
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33 All statistical analyses were performed using SAS 9.4 and R statistical environment on a
34 Windows platform. There were no missing values for the main outcome measurement (i.e.
35 deaths). Missing variables were inputted using the R-Package Mice³⁵. The plausible values of
36 missing data for the covariates were imputed using multivariate imputation by chained
37 equations algorithm (MICE) which created 5 completed imputed datasets. Each variable has
38 its own imputation equation. The MICE method uses all variables in the dataset, including the
39 outcome of interest for imputation of missing data via chained regression equations
40 algorithm.
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49 **RESULTS**

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51 **Cohort characteristics**

52 This study included 4,018 women and 2,102 men aged 50+ followed for a median of 13.5
53 (IQR: 6.5-15.0) and 12.5 years (IQR: 5.4-15.0) for women and men, respectively.
54 During the follow-up 1,081 (27%) women and 284 men (14%) experienced an incident
55 fracture, 308 women and 53 in men experienced a further fracture and 899 women and 578
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1 men died. The length of follow-up post-fracture was 5.5 (IQR: 2.6-9.5) and 5.1 years (IQR:
2 2.3-9.9) for women and men, respectively.
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5 Approximately 65% of women and 23% of men had osteoporosis at baseline (femoral neck
6 T-score \leq -2.5 SD). Of those with baseline osteoporosis, 60% of women and 29% of men
7 received bisphosphonate medication during follow-up. Male gender, baseline diabetes and
8 cardio-vascular disease, smoking, physical inactivity and lower level of education were
9 associated with a higher likelihood of not receiving bisphosphonate therapy. A greater
10 number of medications at baseline did not represent a barrier to receiving bisphosphonate
11 therapy.
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14 15 16 17 18 19 20 **Treatment groups**

21 Reflecting Canadian practice at baseline, etidronate was the most frequently prescribed
22 bisphosphonate [1170 (57%) for women and 167 (54%) for men] followed by alendronate
23 [650 (32%) for women and 95 (31%) for men], and risedronate [228 (11%) for women and
24 46(15%) for men] (Table 1). Risedronate only became available in 1999 and was started on
25 average ~ 9 (\pm 3) years after baseline resulting in both the smaller number of risedronate users
26 and shorter follow-up: 5 (\pm 3) years compared to 8 (\pm 4) years for alendronate and 9 (\pm 4) years
27 for etidronate.
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36 Bisphosphonate users had significantly lower femoral neck BMD, weight, and more incident
37 fractures than (NoRx). They also had several factors associated with “healthy users” such as
38 better education lifestyle habits (less smoking, more exercise and more vitamin D use) and
39 less cardiovascular disease and diabetes. There were no substantive differences in
40 bisphosphonate uptake and year of initiation for the nine study centres across Canada (see
41 supplemental table (Table S1).
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48 49 **Bisphosphonate type and mortality for individuals with and without fracture**

50 51 **n-BP (alendronate and risedronate) vs NoRx**

52 Of the 735 n-BP users, 635 (83% women) were matched to NoRx (Table 2). After propensity
53 score matching there were no statistically significant differences in baseline characteristics
54 between treated and not treated. Mortality risk was reduced for the treated group [HR, 0.66
55 (95% CI, 0.48-0.91)], in particular for women [HR, 0.58 (95% CI, 0.39-0.84)].
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Subgroup analysis according to n-BP type

In order to determine whether the relationship between treatment and survival was similar for the two nitrogen bisphosphonates, a secondary analysis was performed separately for alendronate and risedronate. In these models, mortality risk was adjusted for the baseline characteristics unbalanced after stratification (i.e. cancer for alendronate group; weight and smoking for risedronate group).

Alendronate users were associated with mortality risk reduction [HR 0.62 (95% CI, 0.42-0.92) for any user and 0.60 (95% CI, 0.38-0.93) for women only] (Table 3 and Figure 2).

Risedronate was not associated with an overall mortality risk reduction [HR 0.97 (95% CI, 0.50-1.88)]. However, women who used risedronate appeared to have a mortality risk reduction compared to NoRx [HR, 0.52 (0.25-1.09)] (Table 3 and Figure 2), albeit not statistically significant due to low numbers.

Etidronate vs NoRx

Of the 663 etidronate users, 599 (83% women) were matched to NoRx. By contrast with n-BP users, mortality rates of etidronate users were similar to the matched NoRx [103 deaths/3535 person-years equating to 2.91 deaths/100 person-years (95% CI, 2.40-3.53) vs 110 deaths/3355 person-years equating to 3.28 deaths/100 person-years (95% CI, 2.72-3.95) for etidronate and matched NoRx, respectively; p=0.33] (Table 3 and Figure 2). Etidronate use was not associated with survival benefit in whole group [HR, 0.89 (95% CI, 0.66-1.20)] or in women only [HR 0.88 (95% CI, 0.63-1.25)] (Table 3).

The exclusion of participants who reported bisphosphonates (n-BP or etidronate) only once during the follow-up, did not change the findings.

Importantly, close inspection of the 2-year KM plots for both n-BP and etidronate matched sets revealed that there was no difference in survival during the first 6 months, suggesting that the groups were well matched for mortality risk prior to treatment initiation.

n-BP (Alendronate and Risedronate) vs Etidronate

Given the differences in the baseline characteristics between bisphosphonate types, reflecting different indication criteria, only a third of n-BP users (n=340) were successfully matched 1:1

1 to etidronate (n=340) by propensity score (Table 2). After matching, all characteristics were
2 balanced. However, n-BP users had a borderline higher bone mineral density [0.65 g/cm²
3 (0.10) and 0.64 g/cm² (0.09); p=0.07 for n-BP and etidronate, respectively] and a shorter
4
5 duration on medication [average 4.9 (±3.4) years and 5.5 (±3.4) years; p=0.03 for n-BP and
6 etidronate, respectively]. Mortality risk was significantly lower for n-BP users compared to
7 etidronate [paired HR, 0.47 (95% CI, 0.31-0.70)] (Figure 3).
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10 11 12 **Sensitivity analysis**

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14 In the analysis of all the bisphosphonate users, including those who switched to a different
15 type of bisphosphonate, the results were comparable to the single users analysis. n-BP use
16 was associated with 30-50% mortality risk reduction in the unadjusted [HR, 0.58 (95% CI,
17 0.48- 0.72)], and BMD-adjusted [HR, 0.66 (95% CI, 0.52-0.83)] analyses. Etidronate use
18 was not associated with mortality risk reduction in unadjusted [HR, 0.99 (95% CI, 0.84-
19 1.14)] or BMD adjusted [HR, 1.18, 95% CI, 0.99-1.40)] analyses. Low BMD was a
20 significant confounder in the model of etidronate and survival. A stratified analysis according
21 to BMD level, demonstrated that etidronate was associated with increased mortality risk for
22 osteoporosis group [HR, 1.28 (95% CI, 1.06-1.54)], and a non-significant survival benefit for
23 normal/osteopenia group [HR, 0.72 (95% CI, 0.46- 1.15)].
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26 However, when the models were adjusted for inverse treatment probability scores, both n-
27 BP and etidronate users were associated with survival benefit [HRs, 0.50 (95% CI, 0.39-0.63)
28 and 0.69 (95% CI, 0.59-0.82), for n-BP and etidronate, respectively].
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33 **Bisphosphonate type and mortality for women with incident fracture**

34 Of the 1081 women with incident fracture, 659 received bisphosphonates at the time or after
35 the fracture and 412 used only one type of bisphosphonate. n-BP (n=260 alendronate or
36 risedronate) and etidronate users (n=114) were matched 1:2 to women who never used
37 osteoporosis treatment following fracture. Treated and not treated participants had similar
38 baseline characteristics. However, individuals on treatment had a higher estimated 5-year
39 fracture risk than those not treated (p-value<0.0001 for both n-BP and etidronate pairs).
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55 In women with incident fracture, use of n-BP was associated with better survival [HR, 0.49
56 (0.29-0.80)] while etidronate was not associated with survival benefit [HR, 0.82 (95% CI,
57 0.46-1.49)] (Figure 4).
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1 The relationship between n-BP and mortality was similar regardless of the time of its
2 initiation: HR, 0.48 (95% CI, 0.27-0.85), 0.41 (95% CI, 0.11-1.57) and 0.53 (95% CI, 0.10-
3 2.76) for 0-2, 2-5 and 5+ years post-fracture, respectively).
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7 The risk of a further fragility fracture was similar for BP and matched NoRx [HR, 1.20 (95%
8 CI, 0.74-1.94)] and etidronate [HR, 1.55 (95% CI, 0.78-3.11)].
9

10 11 12 13 **DISCUSSION**

14 Individuals with osteoporotic fracture are at increased risk of death. Emerging evidence
15 suggests that bisphosphonate treatment of those with an osteoporotic fracture is associated
16 with a reduction of all-cause mortality. However, the mechanism for this association is
17 unknown. In this observational study, we found that participants on the nitrogen
18 bisphosphonates, alendronate and risedronate, experienced a 40% survival benefit,
19 particularly in women. Participants on non-n-BP, etidronate had no overall survival benefit.
20 These findings were further supported by a head to head comparison which demonstrated that
21 nitrogen bisphosphonate users had ~50% better survival compared to etidronate users.
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31 These findings suggest that the benefit seen with the nitrogen bisphosphonates either lies in
32 their greater anti-resorptive effect or via a non-bone effect that may be related to their
33 disruption of the mevalonate kinase pathway³⁶. However, most importantly, this analysis in a
34 population-based cohort of two different bisphosphonate biochemistries meant that user bias
35 played a less significant role in the different outcomes observed, thus increasing the
36 likelihood that there is a true decrease in mortality associated with use of nitrogen
37 bisphosphonates.
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45 The magnitude of mortality reduction associated with nitrogen bisphosphonates [HR 0.66
46 (95% CI, 0.48-0.91] in this study, is comparable to previous studies on all-cause mortality
47 risk^{15,16,18,37}. In the Dubbo Osteoporosis Epidemiology Study, bisphosphonate use in women
48 was associated with a 69% reduction in mortality risk compared to no treatment¹⁶. Another
49 study reported 27% lower mortality in institutionalised older people³⁸, and two other studies
50 reported a survival benefit of bisphosphonates in critically ill people^{18,37}. In a Danish database
51 study, the association between bisphosphonates and mortality following hip fracture was
52 similar to this study¹⁵, while two other previous prospective cohort studies reported a stronger
53 relationship (~ 63-66%) with mortality risk reduction^{13,17}. The differences between these
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1 results are most likely related to differences in baseline characteristics between the treated
2 and non-treated populations. Importantly, these findings are also consistent with the 28%
3 mortality risk reduction observed in the zoledronic acid RCT¹¹.
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7 The role of etidronate on mortality risk reduction was less clear. Our primary analysis, with
8 participants who only used etidronate during the follow-up, showed no survival benefit over
9 the follow-up period. By contrast, in a sensitivity analysis, including a large number of
10 participants who switched during follow-up to either alendronate or risedronate, etidronate
11 was associated with ~ 31% mortality risk reduction. However, this finding has to be
12 interpreted in the light of the immortal time bias inherently induced by longer follow-up with
13 participants being alive to switch treatments. In addition, based on the previous analyses, the
14 benefit in etidronate ‘switchers’ could also be attributed to the n-BP to which they were
15 switched.
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19 The uptake of bisphosphonate by men in this study was very low in comparison to women.
20 The association in men between bisphosphonate use and mortality risk was in a similar
21 direction to women, particularly for alendronate and etidronate. The effect of risedronate in
22 men is most likely unreliable, due to the joint effects of a very small sample size (n<50) and a
23 shorter follow-up time. The gender discrepancy in both use of bisphosphonates as well as
24 survival benefit has been reported previously^{14,16}. In an Austrian study, only 12% of men
25 compared to 30% of women reported initiation of bisphosphonate therapy following hip
26 fracture. The association between bisphosphonates and survival was lower than that for
27 women, perhaps driven by the smaller number of men¹⁴.
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42 The mechanism through which bisphosphonates may reduce mortality risk is likely to be
43 multifactorial. The most obvious mechanism would be through a reduction in subsequent
44 fracture risk. However, the RCT of zoledronic acid in women with hip fracture showed that
45 only 8% of the mortality risk reduction in the treatment group was attributable to a reduction
46 in the subsequent fracture rate³⁹. Similarly, reduction in subsequent fracture risk in the
47 current study did not account for the observed mortality risk reduction. The lack of a
48 significantly lower subsequent fracture risk reduction in the treated groups is probably due to
49 their higher baseline fracture risk and possibly the survival advantage providing more time to
50 sustain a fracture.
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1 It is also possible, that the relationship between bisphosphonates and mortality risk could be
2 mediated through a reduction in the rate of bone loss, a marker of poor health and increased
3 mortality⁴⁰ in both individuals with⁴¹ and without fractures^{42,43}. This would also be consistent
4 with the current finding that mortality reduction was greater with n-BP than with etidronate;
5 parallel with their greater antiresorptive effects. On the other hand, there is emerging
6 evidence that nitrogen bisphosphonates may have anti-inflammatory³⁰ and anti-cancer
7 effects³¹.

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14 This study has several strengths. The large number of bisphosphonates users permitted a
15 detailed analysis of bisphosphonates by biochemistry and mortality as well as adjusting for a
16 large set of risk factors not available in registry-based studies. However, there are some
17 limitations. Treatment was a decision made in clinical care and not randomly allocated, thus
18 part of the observed association could be related to confounding. In order to counteract this
19 potential confounding bias, this study employed propensity score matching which is currently
20 recognised as a valid method to account for bias in observational studies⁴⁴. Although this
21 procedure cannot account for unmeasured confounding, the resultant treatment groups had
22 equal baseline risks for all measured variables. Furthermore, matching by the time of
23 medication commencement ensured that the treatment groups started follow-up around the
24 same calendar time. Furthermore, Kaplan Meier survival for matched treated versus no
25 treated did not diverge until after 6 months, suggesting that pairs were well matched and had
26 similar mortality risk at initiation of treatment.

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34 It is possible that there was residual unmeasured confounding such as number of medications,
35 severity of co-morbidities and socio-economic status that could not be accounted for in this
36 observational study. The cost of the different included bisphosphonates was not directly
37 addressed in this study, however in the main analysis adjustment was made for education as a
38 surrogate of socio-economic status. Furthermore, medication number did not predict
39 likelihood of receiving bisphosphonate treatment.

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53 Etidronate was not approved worldwide for the treatment of osteoporosis. It is a weaker anti-
54 resorptive than nitrogen-bisphosphonates and is given in a cyclical regime for two weeks
55 every three months. Didrocal was a formulation of etidronate (Didronel) for ease of managing
56 the two weeks on each 3 months with calcium provided for the other days. Etidronate was
57 available before the nBPs and could have been prescribed to sicker patients. However, this is
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1 unlikely to have been the case as in many jurisdictions, etidronate had to have ‘failed’ or not
2 been tolerated before n-BPs could be prescribed. Thus n-BPs were often prescribed to sicker
3 patients. Any such differences could not be identified or excluded.
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7 Given the yearly collection of data, this study could not address the issue of misclassification
8 of exposure due to differences in treatment adherence. It is possible that some participants
9 may have been prescribed bisphosphonates between yearly questionnaires but did not adhere
10 to treatment for the full year. In this situation, participants would have been classified as non-
11 treatment despite a “window” of treatment exposure. These participants would have only
12 under-estimated the true effect. However, the exclusion of participants who reported
13 bisphosphonates only once during the study follow-up (~10%) and thus more likely to be
14 non-adherent did not impact the findings.
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24 In summary, compared to no treatment, nitrogen bisphosphonate use, particularly in women,
25 was associated with better survival in this long term prospective population-based cohort
26 study irrespective of incident fracture status while etidronate either lacked or had a minor
27 mortality benefit. This observation is important as it points toward mechanistic hypotheses
28 that need to be confirmed in further studies. Importantly, this study suggests that nitrogen
29 bisphosphonate treatment for osteoporosis, whether or not a fragility fracture has occurred,
30 improves survival irrespective of fracture risk prevention.
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39 **Acknowledgement**

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45

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Disclosure Summary:

DB, TT, TvG, CB, DG, RGJ, LL, CSK, SMK and JCP have no competing interests to declare. JDA has received research grants and/or personal fees from Amgen, Eli Lilly, Merck, Actavis and AgNovos. JvdB has received grants and/or personal fees from Amgen, MSD, and Eli Lilly. JAE has consulted for and/or received research funding from Amgen, deCode, Merck Sharp and Dohme, and Sanofi-Aventis. PG was advisory member for Amgen, has received speaker fee and/or research grants from Amgen, Pfizer, MSD, UCB, Abbott, Lilly, BMS, Novartis, Roche and Will Pharma. DAH has consulted for and/or received speaker fee and /or research funding from Amgen, Merck and Eli Lilly. TVN has received honoraria for consulting and symposia from Merck Sharp and Dohme, Roche, Servier, Sanofi-Aventis, and Novartis. JRC has consulted for and/or given educational talks for Merck Sharp and Dohme, Amgen, Actavis and Sanofi-Aventis.

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Contributors

DB, TT, TvG, JAE, and JRC designed the study. DB, TT, JAE, and JRC analysed the data. DB, TT, and JRC drafted the manuscript. All authors contributed to the interpretation of the data and revision of the manuscript. JRC had primary responsibility for final content and acts as guarantor. All authors read and approved the final manuscript. The funding sources had no role in the design and conduct of the study; collection, management, analysis, and interpretation of the data; nor did they have a role in preparation, review, or approval of the manuscript and the decision to submit for publication.

Figures and legends:

- 1
- 2 **Figure 1** Flow chart of participants aged 50+ from Canadian Multicentre Osteoporosis Study
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- 4 **Figure 2** Kaplan-Meier survival curves for alendronate, risedronate and etidronate and matched not
- 5 treated
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- 7 **Figure 3** Kaplan Meier survival curves for nitrogen bisphosphonate (n-BP: alendronate or risedronate)
- 8 and etidronate versus matched not treated in individuals with incident fracture
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- 10 **Figure 4** Kaplan Meier survival curves for nitrogen bisphosphonate and etidronate and matched not
- 11 treated for individuals with incident fractures
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- 14 This includes only bisphosphonate initiated at the time of or after the initial incident fragility
- 15 fracture
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References

1. Nguyen ND, Ahlborg HG, Center JR, Eisman JA, Nguyen TV. Residual lifetime risk of fractures in women and men. *J Bone Miner Res.* Jun 2007;22(6):781-788.
2. Center JR, Bliuc D, Nguyen TV, Eisman JA. Risk of subsequent fracture after low-trauma fracture in men and women. *Jama.* Jan 24 2007;297(4):387-394.
3. Kanis JA, Johnell O, De Laet C, et al. A meta-analysis of previous fracture and subsequent fracture risk. *Bone.* Aug 2004;35(2):375-382.
4. Klotzbuecher CM, Ross PD, Landsman PB, Abbott TA, 3rd, Berger M. Patients with prior fractures have an increased risk of future fractures: a summary of the literature and statistical synthesis. *Journal of bone and mineral research : the official journal of the American Society for Bone and Mineral Research.* Apr 2000;15(4):721-739.
5. Bliuc D, Nguyen ND, Milch VE, Nguyen TV, Eisman JA, Center JR. Mortality risk associated with low-trauma osteoporotic fracture and subsequent fracture in men and women. *Jama.* 2009;301(5):513-521.
6. Center JR, Nguyen TV, Schneider D, Sambrook PN, Eisman JA. Mortality after all major types of osteoporotic fracture in men and women: an observational study. *Lancet.* Mar 13 1999;353(9156):878-882.
7. Cauley JA, Thompson DE, Ensrud KC, Scott JC, Black D. Risk of mortality following clinical fractures. *Osteoporos Int.* 2000;11(7):556-561.
8. Chen JS, Sambrook PN. Antiresorptive therapies for osteoporosis: a clinical overview. *Nature reviews. Endocrinology.* Feb 2012;8(2):81-91.
9. Reginster JY. Antifracture efficacy of currently available therapies for postmenopausal osteoporosis. *Drugs.* Jan 1 2011;71(1):65-78.
10. Zhou J, Ma X, Wang T, Zhai S. Comparative efficacy of bisphosphonates in short-term fracture prevention for primary osteoporosis: a systematic review with network meta-analyses. *Osteoporos Int.* Nov 2016;27(11):3289-3300.
11. Lyles KW, Colon-Emeric CS, Magaziner JS, et al. Zoledronic acid and clinical fractures and mortality after hip fracture. *N Engl J Med.* Nov 01 2007;357(18):1799-1809.
12. Bolland MJ, Grey AB, Gamble GD, Reid IR. Effect of osteoporosis treatment on mortality: a meta-analysis. *The Journal of clinical endocrinology and metabolism.* Mar 2010;95(3):1174-1181.
13. Beaupre LA, Morrish DW, Hanley DA, et al. Oral bisphosphonates are associated with reduced mortality after hip fracture. *Osteoporos Int.* Nov 4 2010.
14. Brozek W, Reichardt B, Zwerina J, Dimai HP, Klaushofer K, Zwettler E. Antiresorptive therapy and risk of mortality and refracture in osteoporosis-related hip fracture: a nationwide study. *Osteoporos Int.* Jan 2016;27(1):387-396.
15. Bondo L, Eiken P, Abrahamsen B. Analysis of the association between bisphosphonate treatment survival in Danish hip fracture patients-a nationwide register-based open cohort study. *Osteoporos Int.* Jan 2013;24(1):245-252.
16. Center JR, Bliuc D, Nguyen ND, Nguyen TV, Eisman JA. Osteoporosis Medication and Reduced Mortality Risk in Elderly Women and Men. *The Journal of clinical endocrinology and metabolism.* Feb 2 2011.
17. Cree MW, Juby AG, Carriere KC. Mortality and morbidity associated with osteoporosis drug treatment following hip fracture. *Osteoporos Int.* Sep 2003;14(9):722-727.
18. Lee P, Ng C, Slattery A, Nair P, Eisman JA, Center JR. Preadmission bisphosphonate and mortality in critically ill patients. *The Journal of clinical endocrinology and metabolism.* Jan 18 2016;jc20153467.

19. van Geel T, Bliuc D, Geusens PPM, et al. Reduced mortality and subsequent fracture risk associated with oral bisphosphonate recommendation in a fracture liaison service setting: A prospective cohort study. *PLoS one*. 2018;13(6):e0198006.
20. Ian Reid HA, Mihov Borislav, Stewart Angela, Garratt Liz, Bolland Mark, Bastin Sonja, Gamble Greg. Abstracts of the ECTS Congress 2018. *Calcified tissue international*. May 01 2018;102(1):1-159.
21. Brennan MB, Huang ES, Lobo JM, et al. Longitudinal trends and predictors of statin use among patients with diabetes. *Journal of Diabetes and its Complications*. 2018/01/01/ 2018;32(1):27-33.
22. Russell RG, Watts NB, Ebetino FH, Rogers MJ. Mechanisms of action of bisphosphonates: similarities and differences and their potential influence on clinical efficacy. *Osteoporos Int*. Jun 2008;19(6):733-759.
23. Rogers MJ. New insights into the molecular mechanisms of action of bisphosphonates. *Current pharmaceutical design*. 2003;9(32):2643-2658.
24. Gulson B, Mizon K, Smith H, et al. Skeletal lead release during bone resorption: effect of bisphosphonate treatment in a pilot study. *Environmental health perspectives*. Oct 2002;110(10):1017-1023.
25. Khalil N, Wilson JW, Talbott EO, et al. Association of blood lead concentrations with mortality in older women: a prospective cohort study. *Environ Health*. 2009;8:15.
26. Rossini M, Adami S, Viapiana O, et al. Long-Term Effects of Amino-Bisphosphonates on Circulating $\gamma\delta$ T Cells. *Calcified tissue international*. 2012;91(6):395-399.
27. Musso A, Catellani S, Canevali P, et al. Aminobisphosphonates prevent the inhibitory effects exerted by lymph node stromal cells on anti-tumor Vdelta 2 T lymphocytes in non-Hodgkin lymphomas. *Haematologica*. Jan 2014;99(1):131-139.
28. Kalyan S, Quabius ES, Wiltfang J, Monig H, Kabelitz D. Can peripheral blood gammadelta T cells predict osteonecrosis of the jaw? An immunological perspective on the adverse drug effects of aminobisphosphonate therapy. *Journal of bone and mineral research : the official journal of the American Society for Bone and Mineral Research*. Apr 2013;28(4):728-735.
29. Ugur Ural A, Avcu F, Ozturk K. Bisphosphonates may retrieve endothelial function in vascular diseases similar to statins' effects. *European journal of haematology*. Jul 2008;81(1):77-78.
30. Conti L, Casetti R, Cardone M, et al. Reciprocal activating interaction between dendritic cells and pamidronate-stimulated gammadelta T cells: role of CD86 and inflammatory cytokines. *Journal of immunology (Baltimore, Md. : 1950)*. Jan 1 2005;174(1):252-260.
31. Clezardin P, Ebetino FH, Fournier PG. Bisphosphonates and cancer-induced bone disease: beyond their antiresorptive activity. *Cancer research*. Jun 15 2005;65(12):4971-4974.
32. Kreiger N, Tenenhouse A, Joseph L, et al. The Canadian Multicentre Osteoporosis Study (CaMos): Background, Rationale, Methods. *Canadian Journal on Aging*. 1999;18(3):376-387.
33. Austin PC. An Introduction to Propensity Score Methods for Reducing the Effects of Confounding in Observational Studies. *Multivariate Behavioral Research*. 06/08 2011;46(3):399-424.
34. Nguyen ND, Frost SA, Center JR, Eisman JA, Nguyen TV. Development of prognostic nomograms for individualizing 5-year and 10-year fracture risks. *Osteoporos Int*. Oct 2008;19(10):1431-1444.
35. Azur MJ, Stuart EA, Frangakis C, Leaf PJ. Multiple Imputation by Chained Equations: What is it and how does it work? *International journal of methods in psychiatric research*. 2011;20(1):40-49.
36. Rogers MJ, Gordon S, Benford HL, et al. Cellular and molecular mechanisms of action of bisphosphonates. *Cancer*. 2000;88(S12):2961-2978.
37. Schulman RC, Moshier EL, Rho L, et al. INTRAVENOUS PAMIDRONATE IS ASSOCIATED WITH REDUCED MORTALITY IN PATIENTS WITH CHRONIC CRITICAL ILLNESS. *Endocrine practice* :

official journal of the American College of Endocrinology and the American Association of Clinical Endocrinologists. Feb 26 2016.

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38. Sambrook PN, Cameron ID, Chen JS, et al. Oral bisphosphonates are associated with reduced mortality in frail older people: a prospective five-year study. *Osteoporos Int*. Oct 20 2010.
39. Colon-Emeric CS, Mesenbrink P, Lyles KW, et al. Potential mediators of the mortality reduction with zoledronic acid after hip fracture. *Journal of bone and mineral research : the official journal of the American Society for Bone and Mineral Research*. Jan 2010;25(1):91-97.
40. Cauley JA, Lui LY, Barnes D, et al. Successful skeletal aging: a marker of low fracture risk and longevity. The Study of Osteoporotic Fractures (SOF). *Journal of bone and mineral research : the official journal of the American Society for Bone and Mineral Research*. Jan 2009;24(1):134-143.
41. Bliuc D, Nguyen ND, Alarkawi D, Nguyen TV, Eisman JA, Center JR. Accelerated bone loss and increased post-fracture mortality in elderly women and men. *Osteoporos Int*. Apr 2015;26(4):1331-1339.
42. Kado DM, Browner WS, Blackwell T, Gore R, Cummings SR. Rate of bone loss is associated with mortality in older women: a prospective study. *Journal of bone and mineral research : the official journal of the American Society for Bone and Mineral Research*. Oct 2000;15(10):1974-1980.
43. Nguyen ND, Center JR, Eisman JA, Nguyen TV. Bone loss, weight loss, and weight fluctuation predict mortality risk in elderly men and women. *J Bone Miner Res*. Aug 2007;22(8):1147-1154.
44. Freemantle N, Marston L, Walters K, Wood J, Reynolds MR, Petersen I. Making inferences on treatment effects from real world data: propensity scores, confounding by indication, and other perils for the unwary in observational research. *BMJ : British Medical Journal*. 2013;347.

Table 1 Characteristics of participants according to medication-groups

| | Bisphosphonate type | | | No Treatment |
|--|---------------------|--------------------|--------------------|--------------|
| | Alendronate | Risedronate | Etidronate | |
| | Women | | | |
| Number | 650 | 228 | 1170 | 1970 |
| Age ^a , yrs | 66 (8) | 65 (8) | 68 (8) | 70 (10) |
| Death ^b | 92 (14) | 22 (10) | 242 (21) | 543 (28) |
| Weight ^a , kg | 64 (12) | 65 (11) | 65 (12) | 71 (15) |
| Higher education ^b | 174 (27) | 50 (22) | 236 (20) | 335 (17) |
| Year of initiation | 1996 | 1999 | 1995 | |
| Years on medication ^a | 7 (4) | 5 (3) | 7 (4) | N/A |
| Fractures ^{1,b} | 243 (37) | 72 (32) | 414 (35) | 352 (18) |
| Hip | 25 (10) | 6 (8) | 52 (13) | 50 (14) |
| Vertebral | 47 (19) | 9 (13) | 50 (12) | 25 (7) |
| Proximal ² | 81 (33) | 19 (26) | 131 (32) | 102 (29) |
| Distal ³ | 90 (37) | 38 (53) | 181 (44) | 175 (50) |
| BMD ^{4,a} , g/cm ² | 0.63 (0.10) | 0.67 (0.09) | 0.63 (0.09) | 0.71 (0.12) |
| Co-morbidities ^b | | | | |
| Heart disease | 40 (6) | 34 (15) | 199 (17) | 233 (12) |
| Diabetes | 46 (7) | 23 (10) | 85 (7) | 320 (16) |
| Hypertension | 251 (39) | 104 (46) | 537 (46) | 993 (51) |
| Neurological | 21 (3) | 9 (4) | 53 (5) | 64 (3) |
| Respiratory | 98 (15) | 28 (13) | 186 (16) | 264 (14) |
| Cancer | 115 (18) | 34 (15) | 213 (18) | 297 (15) |
| Life style factors ^b | | | | |
| Exercise | 409 (63) | 139 (69) | 679 (58) | 972 (49) |
| Smoking | 76 (12) | 27 (12) | 131 (11) | 302 (15) |
| Vitamin D | 256 (39) | 75 (33) | 453 (39) | 546 (28) |
| | Men | | | |
| Number | 95 | 46 | 167 | 1794 |
| Age ^a , yrs | 64 (9) | 65 (8) | 69 (9) | 66 (10) |
| Death ^b | 22 (23) | 11 (24) | 54 (32) | 491 (27) |
| Weight ^a , kg | 79 (12) | 79 (13) | 76 (12) | 82 (14) |
| Higher education ^b | 39 (41) | 17 (37) | 49 (29) | 595 (33) |
| Year of initiation | 1997 | 1999 | 1996 | |
| Years on medication ^a | 5 (4) | 4 (3) | 6 (4) | N/A |
| Fractures ^{1,b} | 24 (25) | 8 (17) | 39 (22) | 213 (12) |
| Hip | 6 (25) | 0 (0) | 10 (25) | 32 (15) |
| Vertebral | 1 (4) | 0 (0) | 4 (10) | 16 (8) |
| Proximal ² | 10 (42) | 7 (88) | 11 (28) | 88 (41) |
| Distal ³ | 7 (29) | 1 (12) | 14 (36) | 77 (36) |
| BMD ^{4,a} , g/cm ² | 0.71 (0.11) | 0.72 (0.11) | 0.69 (0.11) | 0.81 (0.12) |
| Co-morbidities ^b | | | | |
| Heart disease | 14 (15) | 5 (11) | 18 (11) | 293 (16) |
| Diabetes | 10 (11) | 6 (13) | 18 (11) | 297 (17) |
| Hypertension | 35 (36) | 22 (48) | 67 (40) | 728 (41) |
| Neurological | 1 (1) | 1 (2) | 4 (2) | 48 (3) |

| | | | | |
|---------------------------------|----------------|---------|----------------|-----------|
| Respiratory | 13 (14) | 9 (20) | 21 (13) | 201 (11) |
| Cancer | 17 (18) | 8 (17) | 39 (23) | 265 (15) |
| Life style factors ^b | | | | |
| Exercise | 52 (55) | 23 (50) | 87 (52) | 1002 (56) |
| Smoking | 12 (13) | 9 (20) | 23 (14) | 336 (19) |
| Vitamin D | 27 (28) | 9 (20) | 52 (31) | 358 (20) |

¹Incident fragility fractures; ²Proximal fractures: humerus, elbow, pelvis, femur; ³Distal fractures: forearm, carpal, metacarpal, tibia/fibula, ankle, tarsal, metatarsal; ⁴Femoral neck BMD

missing values: weight 3%, BMD (13%), heart disease 0.001%; diabetes (0.001%), hypertension 4%),

respiratory (13%)

^a-mean (sd); ^b-number (%) Bold face corresponds to a global p-value<0.05 for the comparison between treated (alendronate, risedronate, etidronate) and No treatment);

Table 2 Characteristics of the treated and not-treated matched pairs of women by use of a single specific bisphosphonate

| | n-BP vs Not treated pairs | | Etidronate vs Not treated pairs | | n-BP vs Etidronate pairs | |
|---------------------------------------|------------------------------|-----------------|------------------------------------|-----------------|-----------------------------|--------------------|
| | Women | | | | | |
| | Treated | Not Treated | Treated | Not Treated | n-BP | Etidronate |
| Number | 530 | 530 | 496 | 496 | 340 | 340 |
| Age ^a , yrs | 66 (8) | 66 (8) | 68.9 (8.2) | 68.9 (8.2) | 68 (7) | 68 (7) |
| Death ^b | 64 (12) | 87 (16) | 103 (21) | 110 (22) | 39 (11) | 62 (18) |
| Weight ^a , kg | 64 (11) | 70 (14) | 66.0 (12.9) | 67.5 (13.5) | 65 (11) | 65 (11) |
| Higher education ^b | 126 (24) | 122 (23) | 93 (19) | 93 (19) | 61 (18) | 55 (16) |
| Years on medication | 6 (4) | - | 5 (4) | - | 4.9 (3.4) | 5.5 (3.4) |
| Fractures ^{1,b} | 140 (26) | 140 (26) | 105 (21) | 105 (21) | 113 (33) | 95 (28) |
| Hip | 10 (7) | 10 (7) | 11 (10) | 11 (10) | 13 (4) | 11 (3) |
| Vertebral | 6 (4) | 6 (4) | 5 (5) | 5 (5) | 17 (5) | 5 (1) |
| Proximal ² | 45 (38) | 45 (38) | 37 (35) | 37 (35) | 31 (9) | 33 (10) |
| Distal ³ | 79 (51) | 79 (51) | 52 (50) | 52 (50) | 52 (15) | 46 (14) |
| BMD ^{4a} , g/cm ² | 0.66 (0.10) | 0.68 (0.10) | 0.64 (0.09) | 0.65 (0.09) | 0.65 (0.10) | 0.64 (0.09) |
| Co-morbidities^b | | | | | | |
| Heart disease | 27 (5) | 32 (6) | 51 (10) | 44 (9) | 32 (9) | 29 (9) |
| Diabetes | 18 (3) | 27 (5) | 41 (8) | 49 (10) | 28 (8) | 26 (8) |
| Neurological | 13 (2) | 20 (4) | 14 (3) | 17 (3) | 12 (4) | 14 (4) |
| Respiratory | 38 (8) | 52 (11) | 79 (16) | 72 (15) | 44 (13) | 46 (14) |
| Cancer ^a | 92 (17) | 112 (21) | 88 (17) | 93 (19) | 58 (17) | 62 (18) |
| Life style factors^b | | | | | | |
| Exercise | 319 (60) | 326 (62) | 280 (56) | 291 (59) | 211 (62) | 201 (59) |
| Smoking | 67 (13) | 83 (16) | 65 (13) | 60 (12) | 40 (12) | 45 (13) |
| Vitamin D ^a | 188 (35) | 156 (29) | 174 (35) | 137 (28) | 123 (36) | 124 (36) |

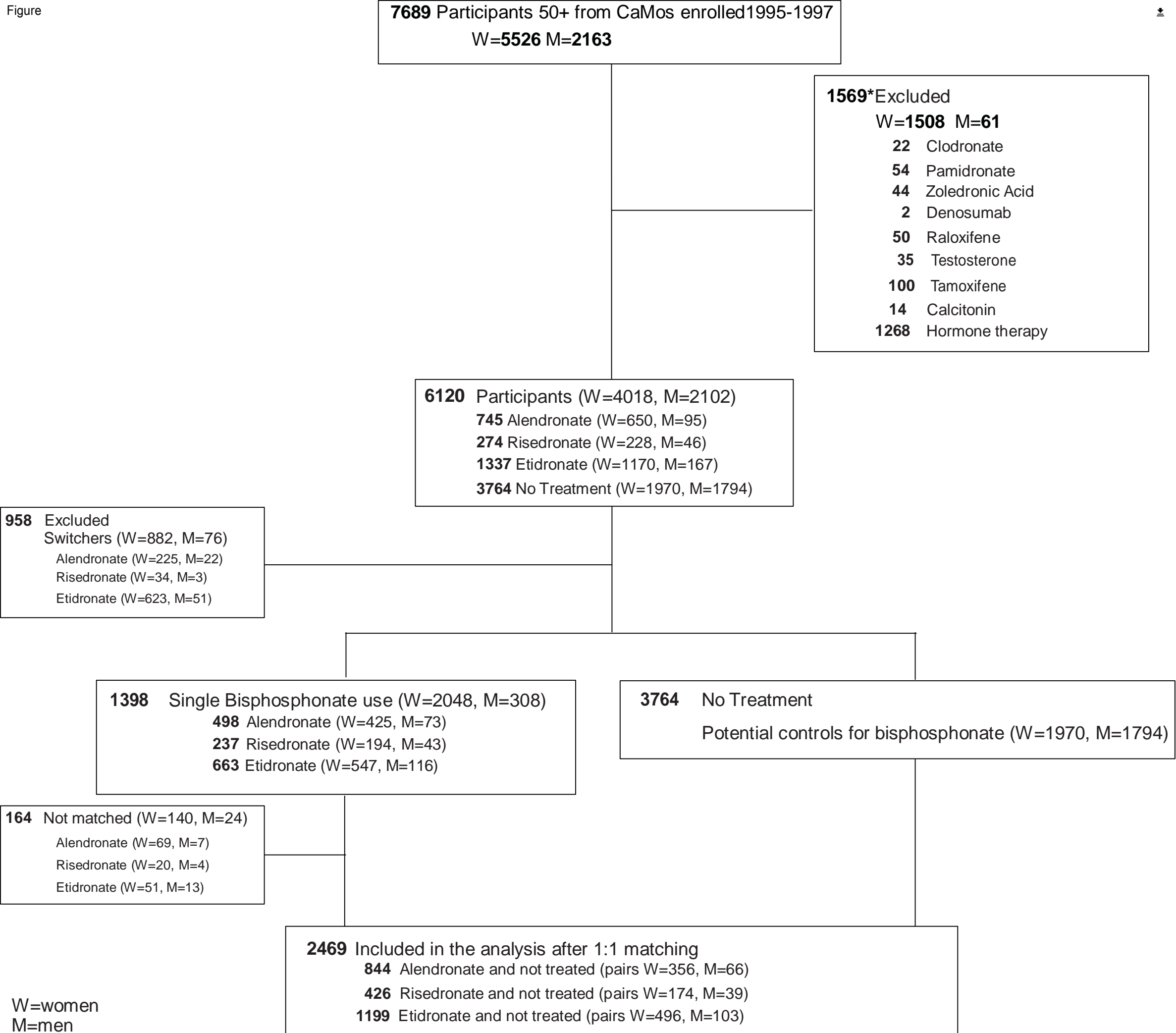
^a-mean (sd); ^b-number (%); Bold face corresponds to a global p-value<0.05 for comparison between treated and not treated within each pair.

¹Incident fragility fractures; ²Proximal fractures: humerus, elbow, pelvis, femur; ³Distal fractures: forearm, carpal, metacarpal, tibia/fibula, ankle, tarsal, metatarsal; ⁴Femoral neck BMD

Table 3 Mortality rates and hazard ratio for pairs of participants treated with different bisphosphonates propensity matched 1:1 to those who were not treated

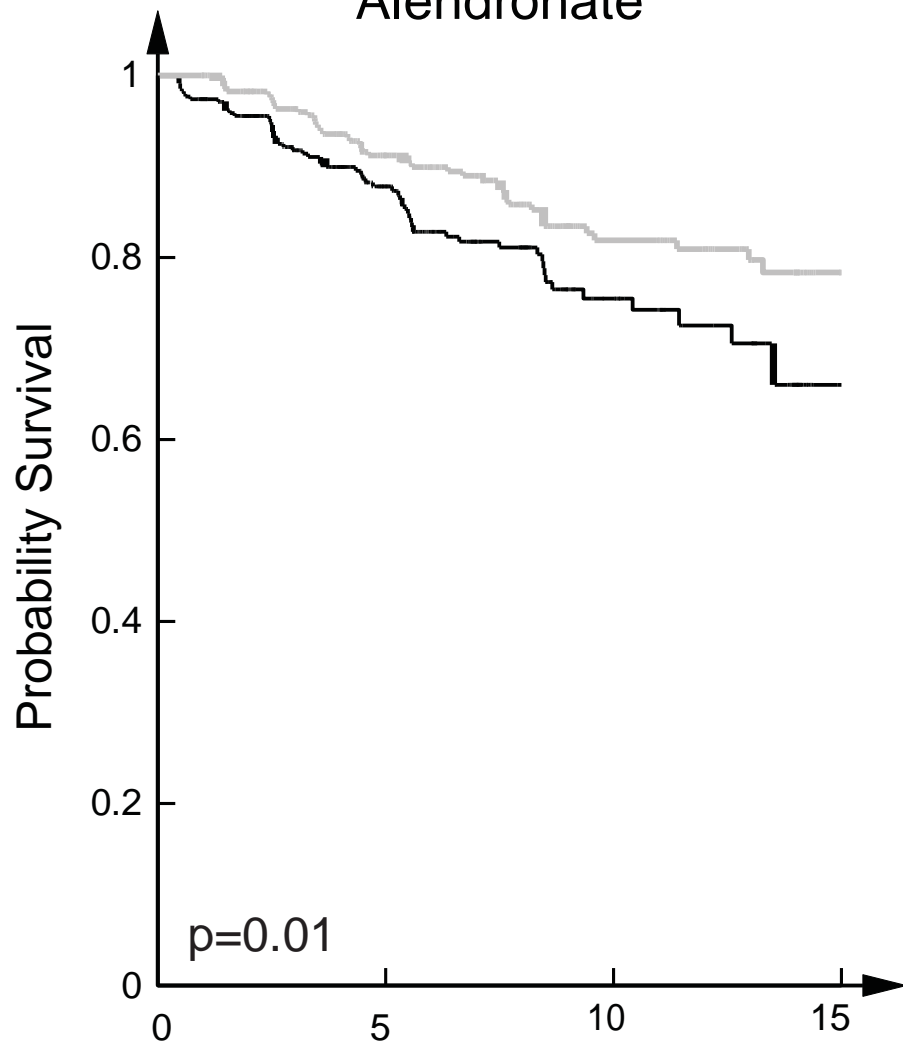
| | N | Treated | | Not Treated | | Paired HR ^a (95% CI) (Treated vs Not treated) |
|-------------|-------|----------|---|-------------|---|---|
| | Pairs | Deaths | Mortality rates (/100 person-yr) (95% CI) | Deaths | Mortality rates (/100 person-yr) (95% CI) | |
| n-BP | 530 | 64 (12) | 1.80 (1.41-2.30) | 87 (16) | 2.94 (2.39-3.63) | 0.58 (0.39-0.84) |
| Alendronate | 356 | 46 (13) | 1.73 (1.29-2.31) | 61 (17) | 2.78 (2.16-3.57) | 0.60 (0.38-0.94) |
| Risedronate | 174 | 18 (10) | 2.03 (1.28-3.22) | 26 (15) | 3.42 (2.32-5.01) | 0.67 (0.30-1.49) |
| Etidronate | 496 | 103 (21) | 2.91 (2.40-3.53) | 110 (22) | 3.28 (2.72-3.95) | 0.88 (0.63-1.25) |

^aHRs were adjusted for all the baseline variables still unbalanced after matching (cancer for alendronate group, weight and smoking for risedronate group)

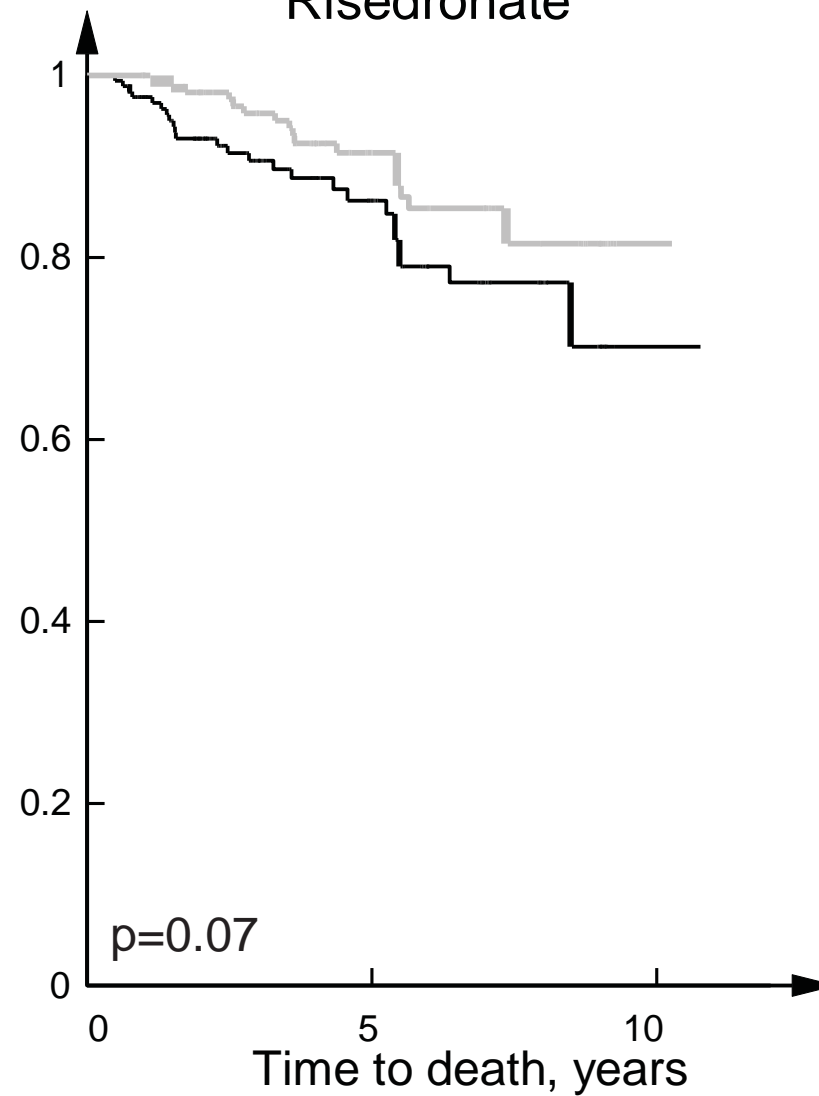


* The individual numbers do not add up to the total as some individual used more than one medication and thus listed more than once

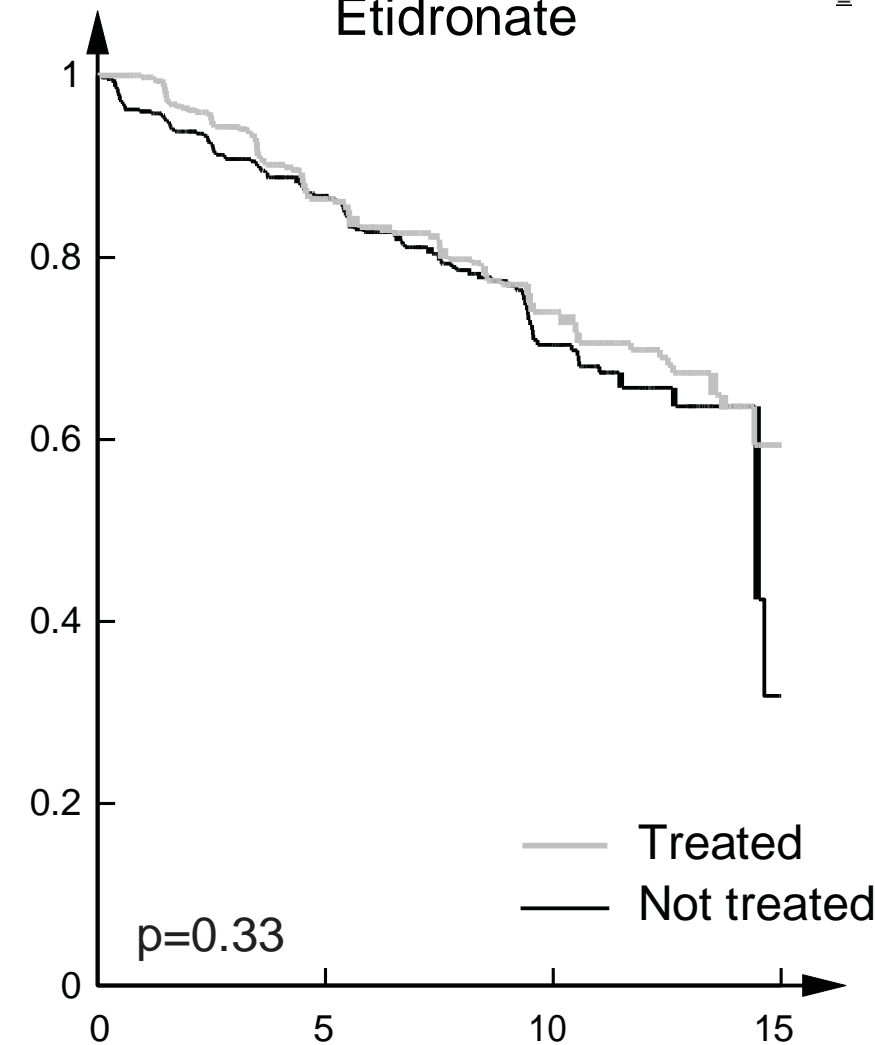
Alendronate



Risedronate



Etidronate

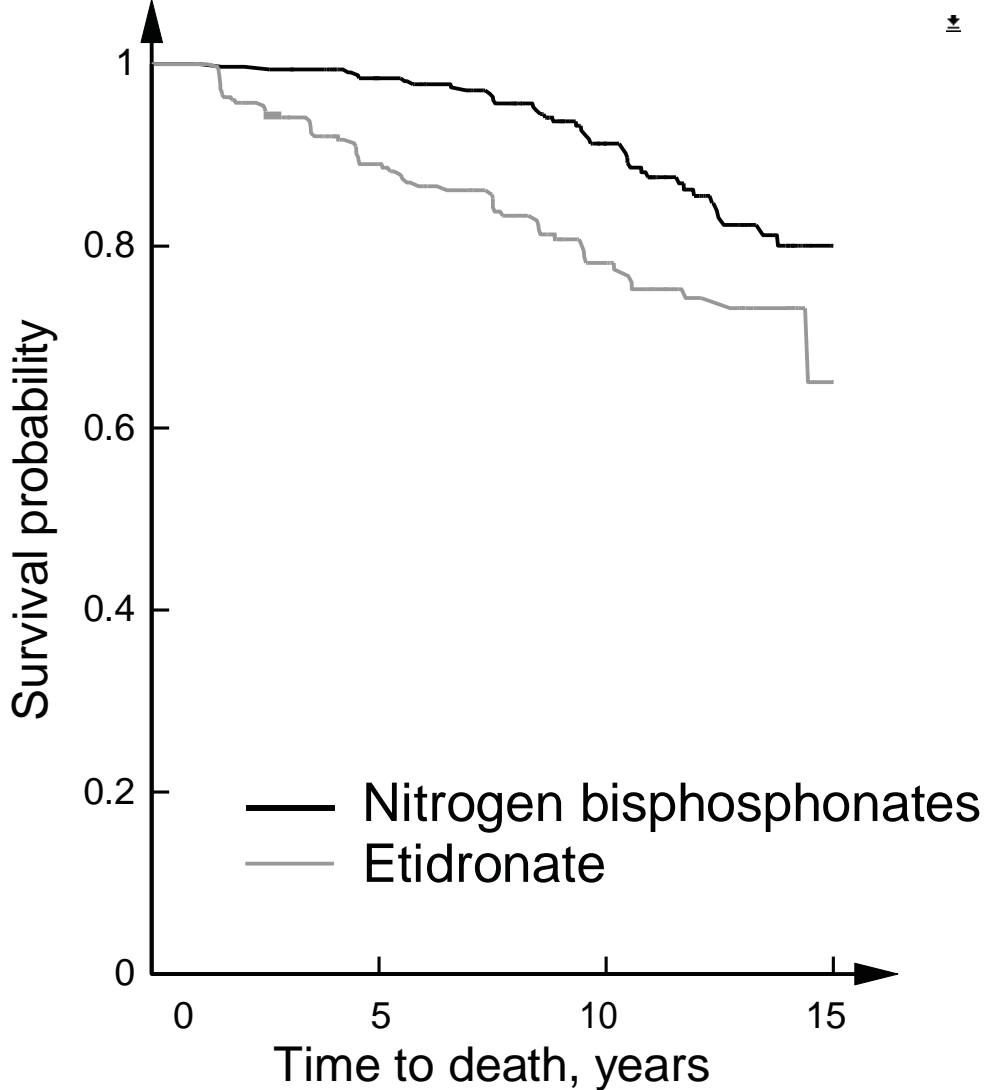


Number at risk
Ratio treated: Not Treated 1:1

| | | | | | | | | | | | |
|-------------|-----|-----|-----|----|-----|----|---|-----|-----|-----|---|
| Treated | 356 | 220 | 103 | 13 | 174 | 79 | 8 | 496 | 291 | 138 | 4 |
| Not treated | 356 | 189 | 70 | 1 | 174 | 64 | 5 | 496 | 293 | 127 | 1 |

Women only

Figure 3



Number at risk

Ratio Nitrogen bisphosphonate: Etidronate 1:1

| | | | | |
|-------------------------|-----|-----|-----|---|
| Nitrogen Bisphosphonate | 340 | 304 | 178 | 7 |
| Etidronate | 340 | 228 | 114 | 2 |

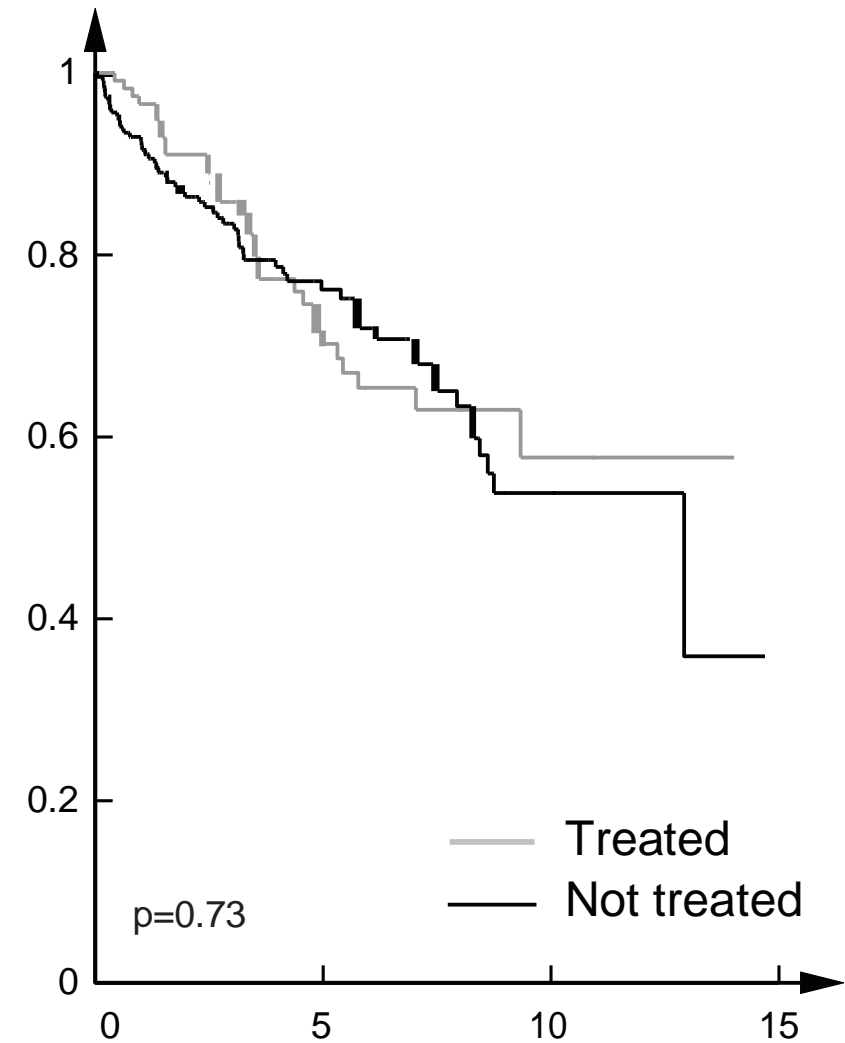
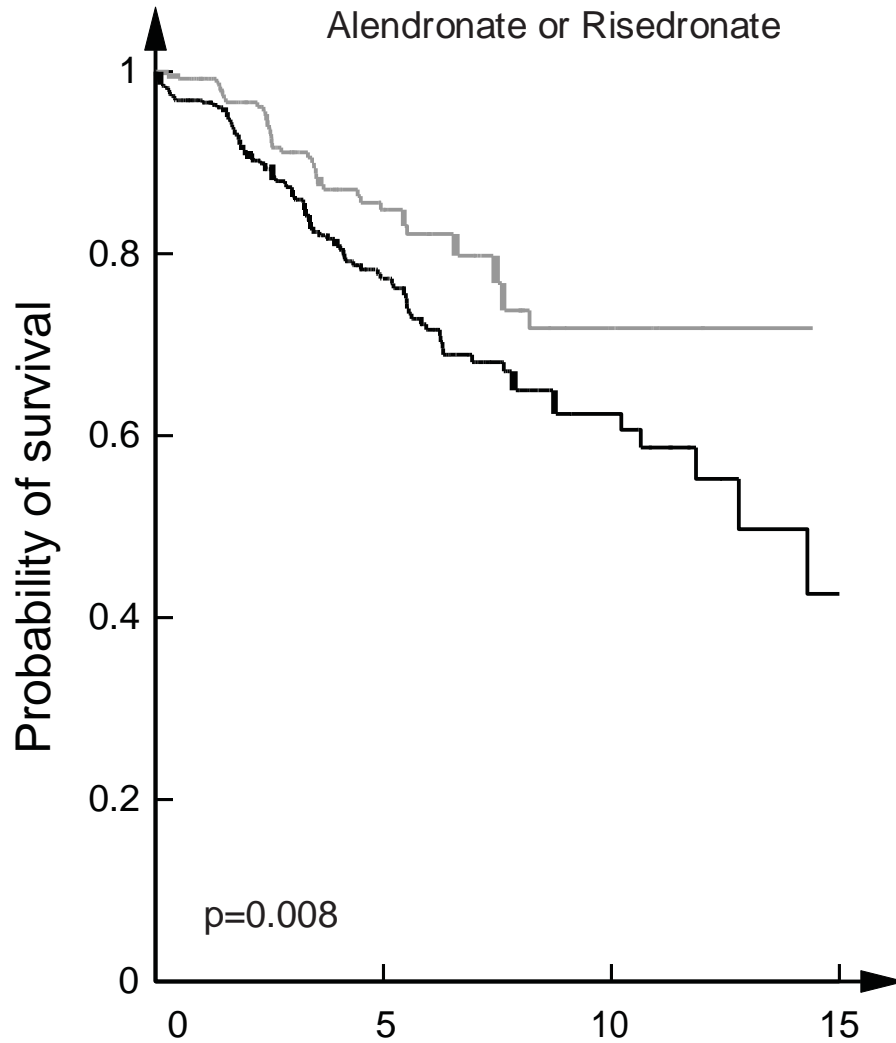
Women only

Fracture group

Nitrogen-Bisphosphonates

Etidronate

Alendronate or Risedronate



Number at risk

Ratio treated: Not Treated 1:2

| | | | | | | | | |
|-------------|-----|-----|----|---|-----|----|----|---|
| Treated | 265 | 108 | 22 | 1 | 149 | 63 | 10 | 0 |
| Not treated | 447 | 151 | 38 | 1 | 213 | 81 | 17 | 0 |

Women only

Propensity Score Matching

Propensity scores for the three sets of matched nBp vs not Treated, etidronate vs not treated and nBP vs etidronate were obtained from three separate logistic regression models. The dependent variable in each of these models was binary, taking value 1 for cases: nBP (first and last model) or etidronate (second model) and value 0 for controls (not treated first 2 models) and etidronate (third models). The independent variables in all three models were all baseline characteristics including age, gender, anthropometric measurements (i.e. height, weight), bone mineral density, baseline co-morbidities, education, life-style factors (i.e. smoking, physical activity), incident fracture type (i.e. hip, vertebral, proximal and distal fracture) and region. The models were well calibrated for most of propensity scores range (Table S1).

Matching was performed using the SAS macro %Gmatch, which uses a greedy algorithm to select a control with the smallest difference in the absolute weighted sums of the matching parameters. We have performed a 1:1 case control matching, based on a propensity score difference ≤ 0.01 and with the condition that the control is alive at the time the case started treatment.

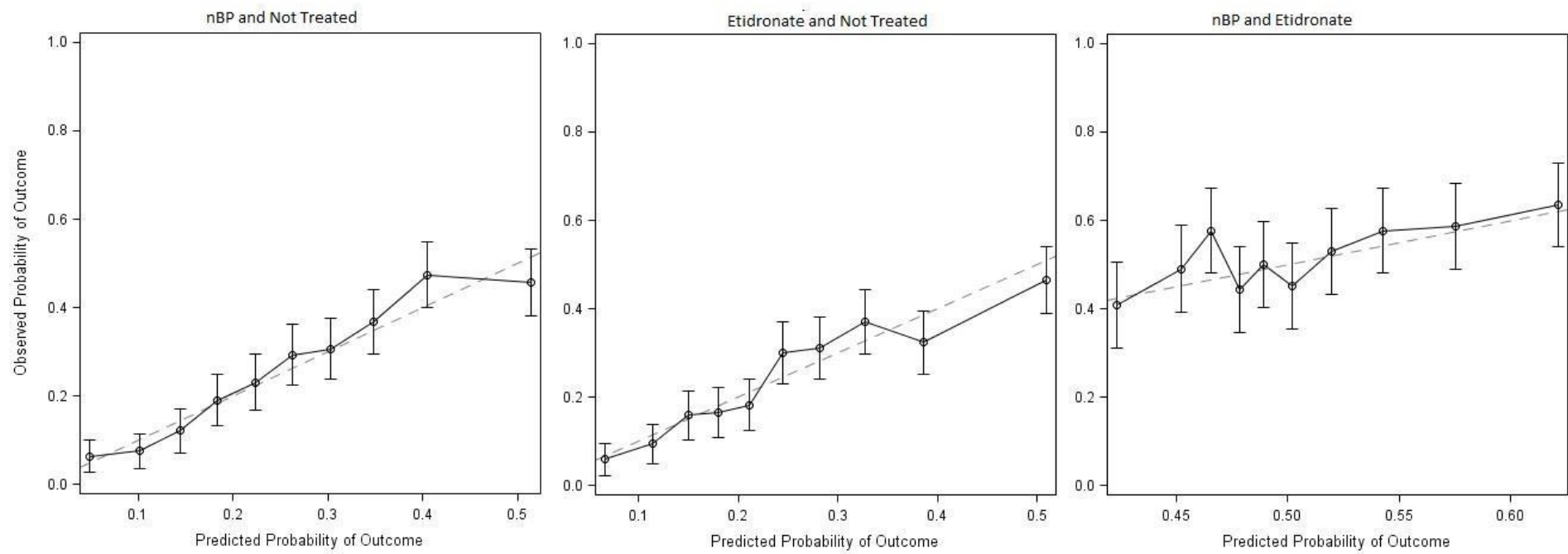


Figure S1 Calibration graphs for logistic models of n-BP and not treated, etidronate and not treated and n-BP and etidronate

Table S2 Bisphosphonates according to province and year of initiation

| | Women | | | Men | | |
|--------------------|------------------------|------------------------|------------------------|-----------------------|-----------------------|-----------------------|
| | Alendronate (n=650) | Risedronate (n=228) | Etidronate (n=1170) | Alendronate (n=95) | Risedronate (n=46) | Etidronate (n=167) |
| Calgary | | | | | | |
| Year of initiation | 1997 | 1999 | 1996 | 1997 | 2002 | 1997 |
| Number (%) | 75 (12) | 16 (7) | 150 (13) | 10 (11) | 8 (17) | 22 (13) |
| Hamilton | | | | | | |
| Year of initiation | 1996 | 2001 | 1996 | 1997 | 1999 | 1996 |
| Number | 64 (10) | 34 (15) | 172 (15) | 13 (14) | 6 (13) | 28 (17) |
| Halifax | | | | | | |
| Year of initiation | 1996 | 2001 | 1996 | 1997 | 2000 | 1997 |
| Number (%) | 80 (12) | 30 (13) | 89 (8) | 11 (12) | 2 (4) | 4 (2) |
| Kingston | | | | | | |
| Year of initiation | 1996 | 2000 | 1996 | 2000 | 2002 | 1996 |
| Number (%) | 61 (9) | 48 (21) | 88 (8) | 7 (7) | 6 (13) | 11 (7) |
| Quebec city | | | | | | |
| Year of initiation | 1996 | 2001 | 1995 | 1998 | 2002 | 1997 |
| Number (%) | 82 (13) | 44 (19) | 71 (6) | 11 (12) | 10 (22) | 1 (0.6) |
| Saskatoon | | | | | | |
| Year of initiation | 1996 | 2002 | 1996 | 1997 | 2002 | 1997 |
| Number (%) | 61 (9) | 15 (7) | 170 (15) | 7 (7) | 2 (4) | 25 (15) |
| St John's | | | | | | |
| Year of initiation | 1996 | 1999 | 1996 | 1997 | 1999 | 1997 |
| Number (%) | 80 (12) | 12 (5) | 101 (9) | 10 (11) | 4 (8) | 10 (6) |
| Toronto | | | | | | |
| Year of initiation | 1996 | 2001 | 1996 | 1997 | 1999 | 1997 |
| Number (%) | 71 (11) | 20 (9) | 143 (12) | 16 (17) | 8 (17) | 21 (13) |
| Vancouver | | | | | | |
| Year of initiation | 1996 | 2003 | 1996 | 1997 | * | 1996 |
| Number | 76 (12) | 9 (4) | 186 (16) | 10 (11) | * | 45 (27) |