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## Primary Language Spoken at Home and Children's Dental Service Utilization in the United States

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#### Abstract

Objective: Language barriers have been well documented as a contributing factor to disparities in the receipt of medical services, especially for Hispanic children. However, there is a paucity of information on the effect of language barriers on children's dental service utilization. We examined the association of primary language spoken at home with the receipt of preventive and routine dental care for children in the United States. Methods: We analyzed data from the Medical Expenditure Panel Survey (2002-2004), which contains data on 21,049 children weighted to represent 75.8 million children nationally. Results: Among children aged 118 years, 13 percent spoke a language other than English at home. Whites, females, children between the ages of 7 and 12 years, and those whose parents spoke English at home had the highest marginal rates of preventive and routine dental visits. However, the large marginal effect of language, even among Hispanics, was not significant after adjusting for other covariates. Parental education and having a primary provider were the strongest predictors of preventive and routine dental visits. Conclusion: Children that did not speak English at home were less likely to receive preventive or routine dental care. However, after adjusting for other socioeconomic factors, our study suggests that language barriers may not play as pronounced a role in the receipt of dental care as that documented for medical services.


## Introduction

An increasing number of people in the United States speak a language other than English at home, with Spanish being the second most prevalent language spoken. Between the years 1990 and 2000, the number of Spanish speakers in the United States rose by 62 percent, increasing from 17.3 to 28.1 million individuals ${ }^{1}$. Language barriers have been consistently linked to the under-utilization and nonreceipt of health-care services by racial/ethnic minorities. For example, Flores and Tomany-Korman reported that children from non-English primary language (NEPL) households experience a multitude of health-related disparities, including reduced access to dental and medical care ${ }^{2}$. Additionally, children of non-English-speaking parents are more likely to be from disadvantaged families and to experience barriers to health care when compared with children of Englishspeaking parents ${ }^{3}$. Furthermore, language barriers have been shown to affect safety in health care, result in low quality of care, and cause dissatisfaction with care among recipients ${ }^{4-7}$.

Disparities in the receipt of oral health care are greatest among racial and ethnic minority children, a large proportion of which speak a language other than English at home, and lack dental insurance ${ }^{3,8}$. Cohen and Christakis ${ }^{9}$ reported disparities based on primary language in the receipt of recommended pediatric preventive care by White, Hispanic, and African-American infants enrolled in Medicaid. Graham et al. ${ }^{10}$ examined the extent to which perceived social status and primary language are associated with having a dental home among four Hispanic groups in Florida. They concluded that those respondents who spoke Spanish at home were less likely to have a dental home, compared with those who spoke English.

Preventive and routine dental care is essential in achieving and maintaining good oral health ${ }^{11}$. Preventive dental interventions, such as fluoride varnish and dental sealant applications are cost-effective ways to reduce dental caries and their associated complications and expenses ${ }^{12}$. Studies have reported that under-utilization of preventive dental services are associated with the following factors: poverty or lower income, male gender, lack of regular dental care, feelings of anxiety toward dental care, lack of insurance, lack of education, and race/ethnicity ${ }^{\mathbf{1 3}, 14}$. Lack of access to preventive and routine dental services commonly results in emergency treatment as the only option for many poor and minority children, resulting in significant financial burdens for their families as well as society through higher health-care costs and inappropriate use of scarce resources. This creates a self-sustaining cycle of unmet treatment need ${ }^{15,16}$.

Although language barriers have been indirectly implicated in the observed oral health disparities for racial/ethnic minority children in the said reports, there has been limited evaluation of the role language barriers play in the receipt of dental services. Flores and Tomany-Korman² reported that children from NEPL households were more likely than children who spoke English at home to lack dental insurance, have teeth in fair/poor condition, have unmet dental care needs, and have not seen a dentist in the last year. However, one limitation of that study (which used The National Survey of Childhood Health) is the unusual high percentage of dental utilization exhibited in the study sample. For example, within children living in NEPL households (comprised of $86.7 \%$ Hispanic children), $85.8 \%$ had a routine preventive dental visit in the last 12 months. This level of utilization is considerably higher than other published reports on the use of dental services among Hispanic populations ${ }^{17-19}$, and even considerably higher than similar reports using the same database ${ }^{20}$. Therefore, it is unclear whether the study of Flores and Tomany-Korman truly generalizes to its intended population in terms of dental services and outcomes. Because of this uncertainty, this study sought to examine whether primary language spoken at home was associated with lower utilization of preventive and routine dental care among racial and ethnic minority children on a national level.

## Methods

We analyzed data from the Medical Expenditure Panel Survey (MEPS) Household Component for 2002-2004. Conducted by the Agency for Healthcare Research and Quality, MEPS is a national survey of health-care use and expenditures that can be generalized to the civilian non-institutionalized population of the United States ${ }^{21,22}$. Response rates were 64.7 percent, 64.5 percent, and 63.1 percent for 2002, 2003, and 2004 MEPS surveys, respectively. Each year a "panel," a subset of the National Health Interview Survey (NHIS) households, is chosen and is followed up over 2 years by MEPS interviewers. The 2002-2004 period contains information from four panels: the second year of panel 6 which was recruited in 2001, both years of panels 7 and 8 recruited in 2002 and 2003, respectively, and the first year of panel 9 recruited in 2004. To avoid dependencies introduced by having multiple data points per respondent, we used only 1 year's worth of data for each panel. Specifically, 2002 data were used for panel 6, 2003 data for panels 7 and 8, and 2004 data for panel 9 . Extensive details on MEPS 2002-2004 methodologies, exclusion and inclusion criteria, as well as the panel structure are provided elsewhere ${ }^{23}$. The two primary outcome variables were preventive dental visits (defined as having had a dental visit during the index year for any of the following procedures: general exam/checkup or consultation, cleaning or prophylaxis/polishing, radiographs/x-rays or bitewings, fluoride treatment, sealant application), and routine dental visits (defined as having at least one event during the index year which may be preventive, restorative, surgical, or orthodontic care classified as dental care by the respondent).

Explanatory variables included were sex, age (categorized as 1-3, 4-6, 7-12, 13-15, and 16-18 years of age), race/ethnicity with Hispanic ethnicity defined as a separate category from non-Hispanics, which was categorized as African-American, Asian, and White/other. Primary language spoken at home was available as English, Spanish, and Other. A combination of the three different languages spoken at home with the four racial/ethnic groups leads to 12 race/ethnic-language categories. Because certain race/ethnicity-language combinations were very rare (for example, Spanish was spoken almost exclusively by Hispanics), these two variables were combined into one and we were left with seven categories. Two percent of the respondents did not belong to any of these seven categories, and were omitted from further analysis. Dental insurance and having a primary medical provider were collected as yes/no variables. Parental education was classified as less than high school, high school degree, or more than high school. When information from both parents was available, the higher level was taken. Poverty status was based on the relationship of household income to the federal poverty level, with levels defined as poor, near poor, low income, middle income, and high income.

## Statistical Analysis.

All analyses took into account the complex survey design of MEPS using Proc SurveyFreq and Proc SurveyLogistic in SAS 9.1.3 (SAS Institute, Cary, NC, USA). We followed the recommendations laid out by the NHIS for pooling data from multiple years ${ }^{24}$, and report weighted frequencies for each of the separate demographic characteristics. In univariate analyses, rates of preventive and routine dental visits were compared across levels of the demographic variables using the Rao-Scott design-adjusted chi-square test for difference in proportions. Odds ratios (OR) and $95 \%$ confidence intervals (CI) were calculated to assess the association of preventive and routine dental visits with each covariate, both marginally and in a multivariable logistic regression model. We fit separate models for preventive and routine dental visits.

In addition, we examined potential interactions between the combined language-race/ethnicity variable and the other predictors. Only one statistically significant interaction was noted: the effect of having a primary provider was higher among racial/ethnic minorities. The inclusion of the interactions did not alter the main significant findings of the multivariable models. We therefore chose to omit the findings of multivariable models that included interaction terms. An alpha level of 0.05 was used throughout to denote statistical significance. This study was approved by the Institutional Review Board of Marquette University.

## Results

Table 1 displays weighted population demographics for children in the United States between the ages of 1 and 18 years. Children that had a primary provider and those that had dental insurance had almost twice the odds of reporting the use of preventive or routine dental services than children without a primary provider and without dental insurance, respectively (Table 2). Compared with English-speaking White children, racial/ethnic minority groups, regardless of language spoken at home, were less likely to report preventive or routine dental care. Within Hispanics, a higher proportion of those who spoke English at home reported preventive or routine dental visit compared with those who spoke Spanish at home ( $P<0.01$ ). Asians who spoke English at home also were more likely to report a preventive dental care compared with those who spoke other languages at home.

Table 1. Unweighted and Weighted Characteristics of US Children Based on MEPS: 2002-2004

| Description | Characteristics | Unweighted <br> frequency | Weighted <br> frequency (in <br> millions) | Weighted \% <br> (SE) |
| :--- | :--- | :--- | :--- | :--- |
| Sex | Male | 10,777 | 38.9 | $51.3(0.5)$ |
|  | Female | 10,272 | 36.9 | $48.7(0.5)$ |
| Age (years) | $1-3$ | 3,331 | 12.0 | $15.9(0.3)$ |
|  | $4-6$ | 3,474 | 12.2 | $16.2(0.3)$ |
|  | $7-12$ | 7,246 | 25.3 | $33.4(0.4)$ |
|  | $13-15$ | 3,666 | 13.4 | $17.6(0.3)$ |
| Primary provider | $16-18$ | 3,332 | 12.8 | $16.9(0.3)$ |
|  | Yes | 18,191 | 67.4 | $89.5(0.4)$ |
| Primary language at home | No | 2,756 | 7.8 | $10.5(0.4)$ |
|  | English | 16,106 | 65.8 | $86.9(0.6)$ |
|  | Spanish | 4,221 | 7.3 | $9.6(0.5)$ |
| Dental insurance | Other | 689 | 2.6 | $3.5(0.4)$ |
|  | Yes | 7,539 | 34.9 | $46.1(0.7)$ |
|  | No | 13,497 | 40.8 | $53.9(0.7)$ |


| Race/Ethnicity | White/Other | 9,444 | 47.6 | 62.8 (1.0) |
| :---: | :---: | :---: | :---: | :---: |
|  | Hispanic | 7,157 | 14.1 | 18.7 (0.9) |
|  | Black | 3,740 | 11.3 | 14.9 (0.7) |
|  | Asian | 708 | 2.7 | 3.6 (0.3) |
| Poverty category | Poor | 5,633 | 12.9 | 17.0 (0.6) |
|  | Near poor | 1,745 | 4.3 | 5.7 (0.3) |
|  | Low income | 4,141 | 12.3 | 16.3 (0.4) |
|  | Middle income | 5,747 | 24.9 | 32.9 (0.7) |
|  | High income | 3,783 | 21.3 | 28.1 (0.7) |
| Perceived health status | Excellent | 9,652 | 38.4 | 50.8 (0.7) |
|  | Very good | 6,609 | 23.2 | 30.7 (0.5) |
|  | Good | 4,147 | 12.2 | 16.2 (0.5) |
|  | Fair | 559 | 1.5 | 2.0 (0.1) |
|  | Poor | 65 | 0.2 | 0.3 (0.0) |
| Education | Less than HS degree | 5,526 | 12.0 | 15.8 (0.5) |
|  | HS degree | 8,446 | 30.3 | 40.1 (0.7) |
|  | More than HS degree | 6,127 | 30.5 | 40.2 (0.8) |
|  | No information | 950 | 3.0 | 3.9 (0.2) |
| Year | 2002 | 5,984 | 20.3 | 26.7 (0.5) |
|  | 2003 | 10,029 | 36.7 | 48.5 (0.6) |
|  | 2004 | 5,036 | 18.8 | 24.8 (0.6) |
| Dental visit in last year | Yes | 8,798 | 36.3 | 47.9 (0.7) |
|  | No | 12,251 | 39.6 | 52.1 (0.7) |
| Preventive visit in last year | Yes | 8,116 | 33.6 | 44.4 (0.7) |
|  | No | 12,933 | 42.1 | 55.6 (0.7) |

MEPS, Medical Expenditure Panel Survey; HS, high school; SE, standard error.
Table 2. Univariate Relationship with Rates of Preventive and Routine Dental Visits

|  | Percent with <br> preventive visit (SE) | OR (95\% CI) | Percent with any <br> dental visit (SE) | OR (95\% CI) |
| :--- | :--- | :--- | :--- | :--- |
| Sex |  |  |  |  |
| Male | $43.1(0.8)$ | $0.90(0.84,0.96)$ | $46.6(0.8)$ | $0.90(0.84,0.96)$ |
| Female | $45.8(0.8)$ | 1.0 | $49.3(0.8)$ | 1.0 |
| Age (years) |  |  |  |  |
| $1-3$ | $13.0(0.8)$ | $0.19(0.16,0.23)$ | $13.5(0.8)$ | $0.16(0.14,0.19)$ |
| $4-6$ | $49.1(1.2)$ | $1.24(1.09,1.40)$ | $50.7(1.2)$ | $1.06(0.94,1.20)$ |
| $7-12$ | $54.3(1.0)$ | $1.52(1.38,1.69)$ | $57.3(0.9)$ | $1.39(1.25,1.54)$ |
| $13-15$ | $50.2(1.2)$ | $1.29(1.17,1.43)$ | $57.2(1.1)$ | $1.38(1.24,1.53)$ |
| $16-18$ | $43.8(1.2)$ | 1.0 | $49.2(1.2)$ | 1.0 |
| Has primary provider |  | $2.38(2.04,2.77)$ | $50.0(0.7)$ | $2.19(1.90,2.53)$ |
| Yes | $46.7(0.7)$ | 1.0 | $31.4(1.5)$ | 1.0 |
| No | $26.9(1.5)$ | $0.33(0.29,0.38)$ | $29.6(1.3)$ | $0.33(0.29,0.38)$ |
| Race/Ethnicity-language |  | $0.48(0.42,0.55)$ | $37.8(1.5)$ | $0.48(0.42,0.55)$ |
| Hispanic-Spanish | $26.5(1.2)$ | $0.45(0.40,0.52)$ | $35.6(1.2)$ | $0.44(0.38,0.49)$ |
| Hispanic-English | $34.5(1.5)$ |  |  |  |
| Black-English | $33.2(1.3)$ |  |  |  |


| Asian-Other | 22.8 (3.2) | 0.27 (0.19, 0.39) | 28.6 (3.4) | 0.32 (0.23, 0.44) |
| :---: | :---: | :---: | :---: | :---: |
| Asian-English | 46.1 (4.2) | 0.78 (0.56, 1.09) | 48.3 (4.2) | 0.74 (0.53, 1.03) |
| White/Other-Other | 39.2 (5.6) | 0.59 (0.37, 0.94) | 45.2 (5.7) | 0.65 (0.41, 1.02) |
| White/Other-English | 52.2 (0.9) | 1.0 | 56.0 (0.9) | 1.0 |
| Dental insurance |  |  |  |  |
| Yes | 53.7 (1.1) | 2.01 (1.82, 2.23) | 57.4 (1.0) | 2.04 (1.85, 2.25) |
| No | 36.5 (0.8) | 1.0 | 39.8 (0.8) | 1.0 |
| Poverty category |  |  |  |  |
| Poor | 30.3 (1.1) | 0.28 (0.24, 0.32) | 33.0 (1.1) | 0.26 (0.23, 0.30) |
| Near poor | 34.1 (1.7) | 0.33 (0.28, 0.40) | 36.7 (1.8) | 0.31 (0.26, 0.37) |
| Low income | 32.5 (1.2) | 0.31 (0.27, 0.35) | 35.2 (1.2) | 0.29 (0.25, 0.33) |
| Middle income | 45.3 (1.2) | 0.53 (0.47, 0.60) | 48.9 (1.2) | 0.50 (0.44, 0.57) |
| High income | 60.8 (1.1) | 1.0 | 65.5 (1.1) | 1.0 |
| Perceived health status |  |  |  |  |
| Poor | 31.2 (6.1) | 0.51 (0.29, 0.88) | 33.6 (6.2) | 0.49 (0.28, 0.84) |
| Fair | 39.0 (2.8) | 0.71 (0.56, 0.90) | 43.4 (2.8) | 0.74 (0.59, 0.93) |
| Good | 37.2 (1.2) | 0.66 (0.59, 0.74) | 40.8 (1.2) | 0.67 (0.60, 0.74) |
| Very good | 44.2 (1.0) | 0.88 (0.81, 0.97) | 47.3 (1.0) | 0.87 (0.80, 0.95) |
| Excellent | 47.2 (0.9) | 1.0 | 50.8 (0.9) | 1.0 |
| Parental education |  |  |  |  |
| More than HS degree | 57.3 (1.0) | 4.15 (3.53, 4.88) | 61.1 (1.0) | 4.21 (3.58, 4.96) |
| HS degree | 40.1 (0.9) | 1.99 (1.74, 2.27) | 43.6 (0.9) | 2.07 (1.82, 2.35) |
| Less than HS degree | 26.9 (1.1) | 1.0 | 29.4 (1.1) | 1.0 |
| No information | 27.1 (2.0) | 1.11(0.88, 1.40) | 30.5 (2.1) | 1.17 (0.93, 1.47) |
| Hispanic - English versus Spanish |  | 1.34 (1.13, 1.60) |  | 1.29 (1.08, 1.54) |
| Asian - English versus Other |  | 2.04 (1.14, 3.66) |  | 1.66 (0.96, 2.88) |

Cl , confidence interval; HS, high school; OR, odds ratio; SE, standard error.
Compared with those who perceived their health status as excellent, the proportion of children that reported the use of preventive or routine dental care was lowest in children with perceived health status identified as poor. Children in the 7-12 age groups had the highest proportion of preventive and routine dental care, with slightly lower proportions for 4 - to 6 -year-olds and 13 - to 15 -year-old children, and lowest among 1 - to 3 -yearolds. Children of parents with more than a high school degree had almost four times the odds of reporting use of preventive or routine dental care compared with those with less than a high school degree. A direct relationship between family income and the probability of reporting use of preventive or routine dental care was identified, with poor children being the least likely to report receiving either type of care.

## Multivariable Analysis of Factors Associated with Dental Service Utilization.

Several factors were significantly associated with preventive and routine dental care in the multivariable regression models (Table 3). Male children were somewhat less likely to report a preventive or routine dental care ( $O R=0.88,95$ percent $\mathrm{Cl}: 0.78-1.0$ ) and ( $\mathrm{OR}=0.86,95$ percent $\mathrm{Cl}: 0.76-0.98$ ) than female children, respectively. Parental education [i.e., having more than high school degree ( $\mathrm{OR}=2.29$, 95 percent Cl : 1.772.96)] and having a primary provider ( $\mathrm{OR}=1.64,95$ percent $\mathrm{Cl}: 1.30-2.07$ ) were the strongest predictors of preventive or routine dental care. Children with dental insurance had increased probability of preventive and routine dental care.

Table 3.
Weighted Logistic Regression Analysis of Factors Associated with Preventive and Routine Dental Visits

| Characteristic | Adjusted odds ratio (95\% <br> Cl ) for preventive visits | $P$ value | Adjusted odds ratio (95\% <br> Cl ) for routine dental visits | $P$ value |
| :---: | :---: | :---: | :---: | :---: |
| Sex |  |  |  |  |
| Male | 0.88 (0.78, 1.0) | 0.05 | 0.86 (0.76, 0.98) | 0.02 |
| Female | Referent |  | Referent |  |
| Age (years) |  |  |  |  |
| 1-3 | 0.17 (0.13, 0.23) | <0.01 | 0.15 (0.11, 0.19) | <0.01 |
| 4-6 | 1.18 (0.95, 1.46) | 0.14 | 0.99 (1.1, 1.57) | 0.87 |
| 7-12 | 1.47 (1.23, 1.75) | <0.01 | 1.31 (0.80, 1.22) | <0.01 |
| 13-15 | 1.30 (1.09, 1.55) | <0.01 | 1.29 (1.07, 1.54) | <0.01 |
| 16-18 | Referent |  | Referent |  |
| Has primary provider |  |  |  |  |
| Yes | 1.77 (1.40, 2.23) | <0.01 | 1.64 (1.30, 2.07) | <0.01 |
| No | Referent |  | Referent |  |
| Race/Ethnicity-language |  |  |  |  |
| Hispanic-Spanish | 0.71 (0.56, 0.89) | <0.01 | 0.72 (0.57, 0.89) | <0.01 |
| Hispanic-English | 0.72 (0.58, 0.90) | <0.01 | 0.72 (0.58, 0.88) | <0.01 |
| African American-English | 0.59 (0.50, 0.71) | <0.01 | $0.57(0.48,0.68)$ | <0.01 |
| Asian-Other | 0.44 (0.25, 0.79) | <0.01 | 0.54 (0.40, 1.37) | 0.02 |
| Asian-English | 0.75 (0.39, 1.45) | 0.39 | 0.74 (0.32, 0.92) | 0.34 |
| White/Other-Other | 1.58 (0.87, 2.88) | 0.13 | 1.59 (0.86, 2.94) | 0.14 |
| White/Other-English | Referent |  | Referent |  |
| Dental insurance |  |  |  |  |
| Yes | 1.29 (1.09, 1.54) | <0.01 | 1.29 (1.08, 1.54) | <0.01 |
| No | Referent |  | Referent |  |
| Poverty category |  |  |  |  |
| Poor-low income | 0.95 (0.79, 1.14) | 0.59 | 0.97 (0.81, 1.17) | 0.77 |
| Middle-high income | Referent |  | Referent |  |
| Perceived general health |  |  |  |  |
| Poor | 1.00 (0.37, 2.75) | 0.99 | 1.22 (0.49, 3.04) | 0.67 |
| Fair | 1.16 (0.84, 1.60) | 0.37 | 1.17 (0.85, 1.61) | 0.34 |
| Good | 0.91 (0.75, 1.11) | 0.35 | 0.88 (0.73, 1.07) | 0.20 |
| Very good | 1.06 (0.90, 1.25) | 0.48 | 1.04 (0.88, 1.22) | 0.65 |
| Excellent | Referent |  | Referent |  |
| Parental education |  |  |  |  |
| More than HS degree | 2.16 (1.66, 2.80) | <0.01 | 2.29 (1.77, 2.96) | <0.01 |
| HS degree | 0.95 (0.71, 1.27) | 0.72 | 1.3 (1.07, 1.58) | <0.01 |
| No information | 1.24 (1.01, 1.51) | 0.04 | 0.99 (0.74, 1.33) | 0.96 |
| Less than HS degree | Referent |  | Referent |  |
| Hispanic-English versus Spanish | 1.0 (0.81, 1.28) | 0.88 | 1.0 (0.8, 1.24) | 0.98 |
| Asian-English versus Other | 1.7 (0.71, 4.05) | 0.24 | 1.36 (0.6, 3.07) | 0.46 |

Cl , confidence interval; HS , high school.

Children from racial and ethnic minority groups had significantly lower odds of preventive and routine dental visits compared with Whites (OR ranged from 0.44 to 0.75 ). Among the minority groups, African Americans (who spoke English at home) and Asians (who spoke other languages at home) had the lowest odds of preventive and routine dental care compared with Whites. Hispanic children who spoke Spanish at home and those who spoke English had the same lower odds of preventive and routine dental care. Although the comparison between English-speaking Asians and Asians who spoke another language at home is not statistically significant ( $O R=$ 1.7, 95 percent $\mathrm{Cl}: 0.71-4.05$ for preventive care), we cannot rule out a potential difference because of the small sample size and wide confidence interval.

## Effect of Language Spoken at Home, Parental Education, and Dental Insurance on Preventive Visits among Hispanic Children.

The disappearance of the large marginal effect of home language among Hispanics in the multivariable analysis was very surprising, so we conducted further analysis to investigate this change. Hispanic children who spoke Spanish at home differed from their English-speaking counterparts in a variety of ways: for example, they had lower rates of dental insurance ( 17 percent versus 38 percent for English speakers) and their parents had lower levels of education ( 59 percent had less than high school education versus 24 percent for English speakers).

Table 4 shows the observed rates of preventive dental visits among Hispanics broken down by language spoken at home, dental insurance, and parental education. There is no consistent trend of children with one primary language having more preventive care than those that speak a different language at home, especially among parents with a high school education or lower. There is some evidence that children of parents with more than a high school education are disadvantaged if Spanish is spoken at home, but the relatively low number of such families (<5 percent of the Hispanic children) results in large standard errors and low power to detect an effect in this subgroup.

Table 4. Weighted Percent of Hispanic Children with Preventive Visits by Primary Language, Parental Education, and Dental Insurance

| Parental education | Dental insurance | Parental language spoken <br> at home <br> \% of Spanish speakers <br> (standard error) | \% of English speakers <br> (standard error) | P value |
| :--- | :--- | :--- | :--- | :--- |
|  |  | $22.7(1.7)$ | $26.4(2.8)$ | 0.26 |
| Less than high school | No | $34.3(4.5)$ | $26.8(6.2)$ | 0.33 |
|  | Yes | $31.0(2.3)$ | $27.1(2.3)$ | 0.23 |
| High school | No | $36.8(5.1)$ | $41.6(3.1)$ | 0.42 |
|  | Yes | $30.6(6.8)$ | $50.4(7.9)$ | 0.06 |
| Above high school | No | $35.2(11.0)$ | $51.1(6.1)$ | 0.21 |
|  | Yes |  |  |  |

## Discussion

Findings from this study support similar reports that have documented disparities in the receipt of dental care for racial and ethnic minority children ${ }^{2,7,12,16}$. Regardless of the language spoken at home, African-American, Hispanic, and Asian children were generally less likely to receive either preventive or routine dental services compared to their White counterparts. We found that almost 1 in 7 children between the ages of 1 and 18 years spoke a language other than English at home. There was little indication of a language effect within racial and ethnic minority children's dental service utilization after considering other factors such as parental education
and having a primary provider. This was observed even among Hispanic children, the ethnic group most likely to speak a language other than English at home.

At first glance, our results appear in stark contrast to other studies from the medical community that have examined the effect of language barriers on health service utilization. Cheng et al. ${ }^{25}$, reported that speaking a language other than English at home identifies Hispanics as being at risk for not receiving recommended healthcare services. Cohen and Christakis ${ }^{9}$ reported disparities in the receipt of recommended pediatric preventive care among infants enrolled in Medicaid based on primary language spoken at home. Flores et al. ${ }^{6}$ also documented language and cultural barriers to health care for Hispanic children. Upon careful examination of the above studies, our results are not as discrepant as they first appear. For example, Cheng et al. based their conclusion on a comparison of Spanish-speaking Hispanics to English-speaking Whites. Their presented multivariable analysis, however, indicates little to no effect of language spoken at home on the receipt of healthcare services within Hispanics (after considering socio-economic differences). While language issues were shown to affect safety and quality of care in the study by Flores et al., the most frequently cited access barriers were primarily income driven, such as transportation difficulties and not being able to afford care.

In our study, parental education was the strongest predictor of utilization of either preventive or routine dental visits. Children of parents with more than a high school degree had twice the odds of utilizing dental services when compared with those with less than a high school degree. This result supports the hypothesis that people with higher levels of education have higher health literacy, which has been shown to have a positive correlation with the utilization of dental services ${ }^{26,27}$. People with higher levels of education also tend to have higher levels of income which has been shown to be a predictor of dental service utilization. Additionally, there was some evidence in our study that children of parents with more than a high school education are disadvantaged if Spanish is spoken at home. However, we likely had low power to detect such an effect given the relatively low number ( $<5$ percent of the Hispanic children) of such families in our study. Further research will therefore be required to investigate whether or not language differentially affects oral health-care usage in educated Hispanic families. Furthermore, children who reported having a primary provider were almost twice as likely to report a dental visit as those who did not. Children who have primary providers are more likely to visit the dentist because of the potential for subsequent referrals following a routine health check ${ }^{28}$.

It is important to acknowledge, given the observational nature of the present study, that it is impossible to completely separate the effect of language from related factors such as parental education, family income, or insurance status. For example, among Hispanic children, those that spoke Spanish at home had lower rates of dental insurance coverage and their parents tended to be less educated than those that did not. While the present study suggests that families that do not speak English at home experience barriers to oral health care because they are likely to be economically and educationally disadvantaged, one can argue that language barriers at least indirectly contribute to this lack of educational and economic attainment. To address the language barrier issue, studies conducted within the medical community have advocated for the use of interpreters and training of bilingual providers as a means of facilitating the receipt of health-care services. While we agree that this will help ensure safe and quality care, it is unclear how this would impact access to dental care because of the effect of socio-economic barriers.

Children between the ages of 7 and 15 years were more likely to have preventive and routine dental visits compared with children 1-3 years old. Even though the American Academy of Pediatric Dentistry and the American Dental Association have recommended that children have two dental visits per year beginning at age $1^{29-31}$, children aged 1-3 years were significantly less likely to visit the dentist for preventive and routine care compared with all other age groups in our study. Despite all efforts within the dental community to educate dentists and parents on the importance of seeking dental care for their children starting from age 1, our results indicate that there is still a lot of work to be done in this area.

## Limitations.

The primary limitation of the present study is that it is observational in nature, and so it is not possible to evaluate the causal effect of language on dental service utilization. However, a number of similar studies using national databases and comparable analytic methods have been used to document language-related disparities in the use of health-care services ${ }^{2,26,32}$. Therefore, we feel that our conclusion regarding the lack of an association between language barriers and preventive and routine dental visits is warranted. However, our hope is that this result may spur future monitoring and data collection concerning language barriers and the receipt of dental care. In addition, MEPS does not contain information related to limited English proficiency (LEP), which has been conjectured as a better metric for assessing the link between language barriers and children's health outcomes and service usage ${ }^{32}$. While the inclusion of both LEP individuals and English-proficient individuals that do not speak English at home has the potential to bias our results, we believe that this bias is likely rather small. Based on calculations described in Flores and Tomany-Korman ${ }^{2}$, the proportion of individuals that are LEP and speak a language other than English at home is on the order of 7 percent of all NEPL households. Therefore, the vast majority of NEPL households have at least one member that is English proficient. However, future research should investigate the relationship between LEP and utilization of dental services. Finally, although MEPS contains a nationally representative sample, the survey contains only a small number of non-English-speaking Asians. We therefore had limited power to detect a language effect within this population group and to contrast this effect with that observed among Hispanics.

## Conclusions.

Children whose primary language spoken at home was not English exhibited lower rates of preventive and routine dental visits. After accounting for socio-economic factors such as family income and parental education, we did not observe an association between lower dental service utilization and families that did not speak English at home. However, we did observe an indication that language barriers may play a role in the use of dental services for educated Hispanic families (at least one parent with greater than a high school education). While racial/ethnic children experience barriers to dental care, our results indicate that language may not play as pronounced a role in the receipt of dental care as research has documented for medical care. Nonetheless, further research is required to elucidate any potential relationship between primary language spoken at home and other factors associated with disparities in access to oral health care, such as parental education, family income, and dental insurance coverage.

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