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10-1-2013

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Accepted version. *Journal of Counseling Psychology*, Vol. 60, No. 4 (October 2013): 593-600. DOI. © 2013 American Psychological Association. Used with permission

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# Dose–Effect Relationship in Routine Outpatient Psychotherapy: Does Treatment Duration Matter? / BRIEF REPORT

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## Acknowledgement:

There is an ongoing debate on the rate and shape of change in outcome variables during the course of psychotherapy ([Baldwin, Berkeljon, Atkins, Olsen, & Nielsen, 2009](#); [Barkham et al., 2006](#); [Kopta, Howard, Lowry, & Beutler, 1994](#); [Lambert, Hansen, & Finch, 2001](#); [Reese, Toland, & Hopkins, 2011](#)). The dose–effect model of change was proposed more than 25 years ago by [Howard, Kopta, Krause, and Orlinsky \(1986\)](#). In their seminal study, they demonstrated a negatively accelerated relationship between the number of treatment sessions (i.e., dose) and the normalized probability of measurable client improvement (i.e., effect) using a variety of global outcome variables. For example, after two sessions, 30% of clients had shown some measurable improvement. This figure increased to 41% after four sessions, 53% after eight sessions, and 75% after 26 sessions, consistent with a negatively accelerated dose–effect curve.

[Howard et al.'s \(1986\)](#) study is considered a milestone toward empirically estimating how much psychotherapy is enough ([Kopta, 2003](#)). Specifically, [Howard et al. \(1986\)](#) suggested that 26 sessions—the dose at which 75% of the clients clearly benefited from treatment—would be a rational limit of the number of sessions in light of the diminishing cost-effectiveness. However, using a stricter criterion for clinically significant change, [Lambert et al.'s \(2001\)](#) survival analysis showed that 21 sessions were necessary for 50% of clients starting within the dysfunctional range to achieve clinically significant change and that 35 sessions were necessary for 70% of clients. Additional research showed differential (negatively accelerated) change patterns by diagnosis and symptoms ([Barkham et al., 1996](#); [Hansen, Lambert, & Forman, 2002](#); [Kopta et al., 1994](#); [Maling, Gurtman, & Howard, 1995](#)). Although the specific number of sessions are often left to administrative purview, monitoring of client progress and feedback to therapists are often based on comparing clients' individual progress against negatively accelerating trajectories (i.e., expected treatment response curves) modeled using a log-linear function of the session number (e.g., [Lambert, 2010](#); [Lambert et al., 2003](#); [Lueger et al., 2001](#); [Lutz, Lambert, et al., 2006](#); [Lutz, Lowry, Kopta, Einstein, & Howard, 2001](#); [Lutz, Martinovich, & Howard, 1999](#); [Lutz, Saunders, et al., 2006](#); [Lutz, Stulz, Martinovich, Leon, & Saunders, 2009](#); [Stulz, Lutz, Leach, Lucock, & Barkham, 2007](#)).

However, criticisms have been raised against [Howard et al.'s \(1986\)](#) and others' argument for a negatively accelerating dose–effect curve, as it was modeled using the percentage of clients reaching pre-to-post treatment improvement as a function of treatment length (i.e., aggregated across clients). For example, [Barkham et al. \(1996\)](#) reported that session-by-session plots of individual clients' improvement in time-limited psychotherapies tended to look linear. Further, in their “good enough level” (GEL) model, [Barkham et al. \(2006\)](#) demonstrated that even with an assumption of linear change, aggregating data across clients with varying lengths of treatment could misleadingly produce negatively accelerating aggregate curves, as clients responding more quickly to treatment tend to terminate earlier. Therefore, to model clients' trajectory during treatment, data must be modeled longitudinally rather than cross-sectionally.

To capture clients' typical patterns of improvement during treatment, progress should ideally be measured session-by-session using a continuous outcome variable and modeled against the amount of treatment that the individual clients received ([Krause, Howard, & Lutz, 1998](#)). In this study, we used session-by-session outcome assessments during psychotherapy conducted under routine outpatient

conditions to examine further the relationship between the number of sessions and progress ([Krause et al., 1998](#)). The large database allowed us to examine the pattern of change in client subgroups with different treatment lengths. Specifically, we addressed the following research questions: (a) Does individual change on a global clinical symptom scale measured during routine outpatient treatment follow a linear or a negatively accelerated pattern of change? (b) Does the pattern of change vary according to treatment duration? (c) Does the rate of change vary according to treatment duration?

## Method

### Participants

Our sample consisted of 6,375 clients who were treated for at least three sessions at 26 centers, which included 20 college counseling centers (94% of the clients), four primary care medical centers (5.8%), and two private mental health centers (0.2%). Data from shorter treatments (one or two sessions) were excluded as it did not allow for differentiation between linear and negatively accelerated change in outcome variables during treatment. In addition, clients must have begun treatment in the dysfunctional range of the measure used, according to the clinical significance cutoff criterion established by [Jacobson and Truax \(1991\)](#). All clients gave informed consent to allow for their data to be analyzed in aggregate during the welcome screen message presented to them prior to completing the electronically administered questionnaire.

The clients were treated between June 2006 and December 2011. All clients were 18 years and older, and 64% were female. Regarding ethnicity, 62% indicated that they were European American, 7% Asian American, 6% other, 5% Latino/Hispanic, 5% African American, 1% Native American, 1% Middle Eastern, and 1% multiracial (13% did not report).

### Treatment and Therapists

There were no external criteria for treatment termination; discharge from treatment was at the therapist's and client's discretion. We considered a treatment episode to have ended if there was no additional treatment session for 3 months or more. If a client had several treatment episodes, only the first episode was included in the analyses. Treatment durations varied between three and 86 sessions ( $M = 8.0$ ,  $SD = 7.2$ ,  $Mdn = 5$ ).

Over 500 therapists provided treatment. Most were female (69%) and European American (61%). Therapists included psychologists, psychiatrists, clinical social workers, and trainees. Most had a master's degree (42%) or a doctorate (52%). The therapists varied in professional background and theoretical orientation, and, although many were familiar with published treatment manuals, they were not required to follow a manualized treatment protocol. Treatment duration was variable and not subject to strict time limits, although some therapists set time limits as part of their treatment strategy.

### Measure

The Behavioral Health Measure (BHM; [Kopta & Lowry, 2002](#)) is a reliable and valid 21-item self-report measure that assesses well-being (three items), psychological symptoms (13 items), and life functioning (four items). An overall suicide risk (one item) appears only when clients indicate self harm on one of the psychological symptom items. Respondents rate the items regarding how they have been

feeling over the past 2 weeks on a Likert-type scale ranging from 0 (*extreme distress/poor functioning*) to 4 (*no distress/excellent functioning*). The Global Mental Health (GMH) subscale used in this study is calculated by averaging the 20 items (i.e., excludes the overall suicide risk item). For our analyses, if the client answered less than 80% of the items, the GMH score was treated as a missing value. A GMH score below 2.94 indicates that clients are more likely in the clinical population than in the nonclinical population ([Kopta & Lowry, 2002](#)). Analyses have established the internal consistency (Cronbach's  $\alpha = .91$ ) and test-retest reliability ( $r = .80$ ) of the GMH subscale, as well as its validity ([Kopta & Lowry, 2002](#)).

## Procedure

Clients were administered the BHM at each psychotherapy session. The instrument was administered through a computer-based system using an iPad, desktop, or Netbook computer. Clients answered the items before the session. The therapist then immediately received a clinical report on his or her computer based on the client's responses.

## Data Analysis

Clients were classified into groups based on the length of the treatment (3, 4, 5, 6, 7, 8, 9, 10, 11–12, 13–14, 15–17, 18–20, 21–25, and 26 or more sessions). We then modeled the pattern of GMH scores (i.e., the dose-effect relationship) up to Session 20 for each of these groups using latent growth curve modeling (LGCM; [Bollen & Curran, 2006](#)). LGCM enables analysis of longitudinal data by relating an observed outcome variable (e.g., treatment progress) to time or time-related variables (e.g., number of treatment sessions). Specifically, individual variation in treatment progress was captured by latent coefficients on initial severity (i.e., intercept) and trajectory of outcome (i.e., slope for linear or log-linear). Drawing the log of the session number enabled modeling a negatively accelerated relationship between the number of sessions and treatment progress on the GMH. We thus compared the linear and log-linear LGCM for every treatment length category. Both the linear and log-linear models were identical in terms of the number of parameters to estimate.

The fit indices to evaluate the LGCMs (i.e., to see how well a hypothesized model fits the observed data) included  $\chi^2$ , the comparative fit index (CFI), the Tucker-Lewis index (TLI), the root-mean-square error of approximation (RMSEA), the standardized root-mean-square residual (SRMR), and the Akaike information criterion (AIC). As the  $\chi^2$  test is oversensitive with large samples, most researchers regard a model as acceptable if the CFI and TLI—which are both based on the  $\chi^2$  statistic and compare the null (i.e., uncorrelated) model against the hypothesized model—are greater than 0.9. The RMSEA is a measure of the discrepancy between the model and the data per degree of freedom, with values closer to 0 indicating better fit. There is general agreement that a value of less than 0.05 indicates a “good” fit, and a value of less than 0.08 indicates an “acceptable” fit. The SRMR represents the absolute difference between the observed and the model-implied correlations with values closer to 0 indicating better fit; a value of less than 0.10 indicates “acceptable” fit ([McDonald & Ho, 2002](#); [Stull, 2008](#)). For the AIC, no absolute cutoffs are proposed and the model with the smallest AIC is deemed best ([Akaike, 1974](#)). Using the *Mplus* software (Version 6.1; [Muthén & Muthén, 2010](#)), the LGCMs were estimated with full information maximum likelihood to handle missing values in outcome variables, which were assumed to be missing at random ([Rubin, 1976](#); [Schafer & Graham, 2002](#)).

## Results

The mean GMH score was 2.24 ( $SD = 0.46$ ) at intake and 2.82 ( $SD = 0.61$ ) at posttherapy (difference = 0.58; 95% confidence interval = 0.56 to 0.59). The corresponding average pre-post effect size was  $d = 1.26$ . Based on [Martinovich, Saunders, and Howard's \(1996\)](#) reliable change index (RCI; see also [Jacobson & Truax, 1991](#)), the magnitude of the pre-post difference  $\Delta_{RC}$  that is considered to be reliable (in the sense of being large enough to be not attributable to measurement error) is calculated as

$$\begin{aligned}\Delta_{RC} &= z_{RC} \cdot \sqrt{2(SD\sqrt{1-r})^2} \\ &= 1.96 \cdot \sqrt{2(0.47 \cdot \sqrt{1-.91})^2} = 0.39, \\ \Delta_{RC} &= Z_{RC} \cdot \sqrt{2(SD\sqrt{1-r})^2} = 1.96 \cdot \sqrt{2(0.47 \cdot \sqrt{1-.91})^2} = 0.39.\end{aligned}$$

where  $z_{RC}$  is the desired reliability of the change (i.e., 95% two-tailed),  $SD$  is the standard deviation of the GMH score in a community adult sample ([Kopta & Lowry, 2002](#)), and  $r$  is the reliability (internal consistency) of the BHM. Of the 6,375 participants in the data set,  $n = 44$  clients did not answer at least 80% of the BHM items at posttreatment and thus were excluded. Of the remaining 6,331 clients, 3,965 (62.6%) clients reliably improved, including 2,657 (42.0%) clients who also achieved clinically significant improvement (i.e., scored above the GMH cutoff of 2.94 at termination). The proportion of clients who achieved clinically significant improvement (42.0%) was within the range usually found in clinical trials (40–70%, [Lambert, 2013](#)).

The visual inspection of [Figures 1](#) and [2](#) suggests that the patterns of GMH mean scores over the first 20 psychotherapy sessions appear negatively accelerated rather than linear irrespective of the treatment length. Statistical analyses supported this impression, as the log-linear LGCMs consistently outperformed the linear models on all model fit criteria irrespective of the total amount of treatment (cf. online Supplement 1).

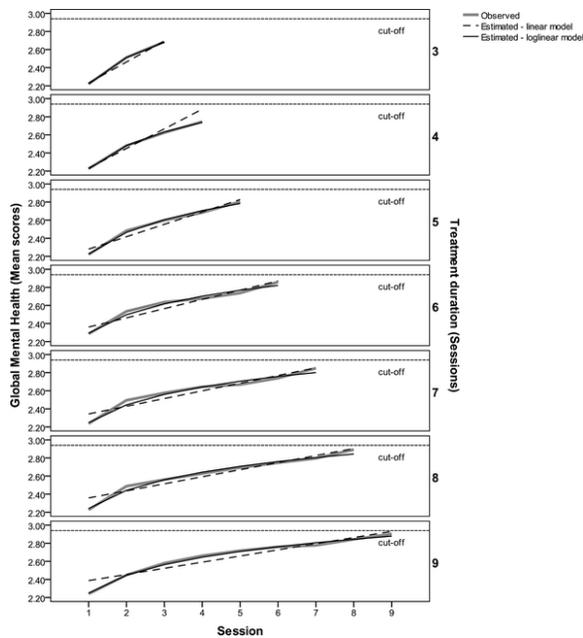


Figure 1. Dose–outcome relationships for varying treatment durations (three to nine sessions): observed and estimated global mental health (GMH) mean scores.

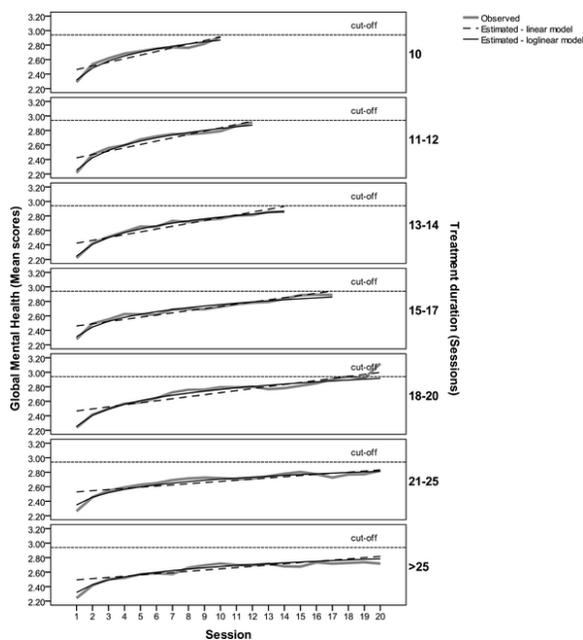


Figure 2. Dose–outcome relationships for varying treatment durations (more than nine sessions): observed and estimated global mental health (GMH) mean scores.

We then examined the development of the percentage of reliably improved clients as a function of the number of sessions. Although these analyses are subject to loss of information (due to dichotomizing continuous outcomes into reliably improved or not), they allow us to compare our findings to previous studies reporting percentage of clients with reliable change ([Howard et al., 1986](#); [Lambert et al., 2001](#)). As can be seen by visual inspection of the Figures in online Supplement 2, treatment progress also

suggests negative acceleration of percentage of clients obtaining reliable change, and this occurred irrespective of the total amount of treatment.

Our analyses also revealed that faster rates of change were related to shorter treatments (Table 1). The Spearman rank correlations between the mean rates of change and the categorized treatment durations were  $r = -.996$  ( $p < .001$ ) for the slopes derived from linear LGCMs and  $r = -.974$  ( $p < .001$ ) for the slopes derived from the log-linear growth models. Table 1 furthermore presents the proportion of clients who achieved clinically significant improvement by the end of therapy stratified by the total number of sessions attended. The percentage of clients attaining clinically significant improvement ranged between 32.1% and 63.1%, depending on the duration of treatment (Spearman rank correlation between percentage attaining clinical significance and the categorized treatment durations:  $r = .714$ ;  $p = .004$ ).

Table 1  
Latent Growth Parameters and Clinically Significant Pre-Post-Treatment Improvement

Treatment duration (sessions)	Clients <i>n</i>	Linear model				Log-linear model				Clinically significant improved <sup>a</sup>	
		Intercept		Slope		Intercept		Slope		<i>n</i>	%
		<i>M</i>	<i>SD</i>	<i>M</i>	<i>SD</i>	<i>M</i>	<i>SD</i>	<i>M</i>	<i>SD</i>		
3	1,517	2.231	0.406	0.232	0.235	2.220	0.423	0.970	0.925	486	32.1
4	958	2.233	0.404	0.217	0.210	2.225	0.414	0.853	0.802	366	38.4
5	719	2.280	0.396	0.137	0.105	2.225	0.406	0.799	0.628	299	41.8
6	505	2.361	0.365	0.102	0.089	2.294	0.385	0.678	0.604	222	44.3
7	426	2.344	0.434	0.085	0.089	2.247	0.429	0.656	0.622	182	43.2
8	336	2.358	0.417	0.078	0.071	2.241	0.391	0.667	0.554	168	50.1
9	301	2.387	0.410	0.068	0.055	2.249	0.392	0.659	0.510	145	48.3
10	243	2.462	0.392	0.050	0.063	2.317	0.382	0.552	0.528	115	47.3
11–12	344	2.423	0.446	0.046	0.055	2.252	0.422	0.576	0.509	176	51.9
13–14	237	2.424	0.451	0.039	0.045	2.245	0.431	0.539	0.475	107	45.3
15–17	252	2.461	0.411	0.030	0.032	2.312	0.416	0.445	0.415	136	54.6
18–20	159	2.467	0.430	0.028	0.032	2.253	0.417	0.510	0.449	99	63.1
21–25	163	2.529	0.439	0.016	0.032	2.349	0.452	0.356	0.469	74	46.3
≥26	215	2.468	0.453	0.017	0.032	2.300	0.447	0.357	0.396	100	48.1

<sup>a</sup>  $n = 44$  clients did not answer at least 80% of the Behavioral Health Measure (Kopta & Lowry, 2002) items at posttreatment. These clients were excluded from the analyses, as they were not considered to have a valid Global Mental Health score at posttreatment.



Table 1 Latent Growth Parameters and Clinically Significant Pre-Post-Treatment Improvement

		Linear model				Log-linear model				Clinically improved <sup>a</sup>	significant	
		Intercept		Slope		Intercept		Slope				
Treatment duration (sessions)	Clients <i>n</i>	<i>M</i>	<i>SD</i>	<i>M</i>	<i>SD</i>	<i>M</i>	<i>SD</i>	<i>M</i>	<i>SD</i>	<i>n</i>	%	
3	1,517	2.231	0.406	0.232	0.235	2.220	0.423	0.970	0.925	486	32.1	
4	958	2.233	0.404	0.217	0.210	2.225	0.414	0.853	0.802	366	38.4	
5	719	2.280	0.396	0.137	0.105	2.225	0.406	0.799	0.628	299	41.8	
6	505	2.361	0.365	0.102	0.089	2.294	0.385	0.678	0.604	222	44.3	
7	426	2.344	0.434	0.085	0.089	2.247	0.429	0.656	0.622	182	43.2	
8	336	2.358	0.417	0.078	0.071	2.241	0.391	0.667	0.554	168	50.1	
9	301	2.387	0.410	0.068	0.055	2.249	0.392	0.659	0.510	145	48.3	
10	243	2.462	0.392	0.050	0.063	2.317	0.382	0.552	0.528	115	47.3	
11–12	344	2.423	0.446	0.046	0.055	2.252	0.422	0.576	0.509	176	51.9	
13–14	237	2.424	0.451	0.039	0.045	2.245	0.431	0.539	0.475	107	45.3	
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<sup>a</sup> *n* = 44 clients did not answer at least 80% of the Behavioral Health Measure (Kopta & Lowry, 2002) items at posttreatment. These clients were excluded from the analyses, as they were not considered to have a valid Global Mental Health score at posttreatment.

## Discussion

Consistent with the dose–effect model ([Howard et al., 1986](#)), our results suggest that treatment responses modeled longitudinally under routine outpatient conditions proceed in a negatively accelerated manner. Using a global mental health measure, this negatively accelerating curve of improvement was shown to model treatment progress better than a linear model irrespective of the length of treatment.

These findings suggest that the negatively accelerating pattern of change is not an artifact of disregarding the varying treatment lengths and cross-sectionally aggregating the percentage of clients reaching a criterion for improvement, as was suggested by [Barkham et al. \(2006\)](#). While Barkham et al. examined time-limited therapies, our findings suggest that, at least in unrestricted therapies, the average rate of change during psychotherapy can be expected to diminish with each additional session irrespective of the length of treatment. These results are consistent with previous studies (e.g., [Howard et al., 1986](#); [Kopta et al., 1994](#)), and the results may also corroborate [Reese et al.'s \(2011\)](#) findings had they retained the (statistically significant) curvilinear parameters in their growth models. While their argument toward discarding curvilinear trends in favor of parsimony is understandable, the linear and log-linear LGCMs used in our study were equivalent in model complexity given that the number of parameters to estimate was identical in both models. Further, as the negatively accelerated LGCMs were shown to estimate treatment progress consistently better than linear models, improvement of clients' distress cannot be expected to be linear in outpatient psychotherapy. On average, clients appear to experience relatively rapid improvement in early treatment stages, whereas the effect of treatment seems to decrease in later stages. Our findings are therefore in line with the dose–effect model.

Our results, however, also corroborate the GEL model ([Barkham et al., 2006](#)). Although treatment progress typically proceeded in a negatively accelerated manner, the rate of change was inversely related to the total amount of psychotherapy. Consistent with the GEL model, rapid improvement was associated with shorter treatment duration and vice versa ([Baldwin et al., 2009](#); [Barkham et al., 2006](#); [Reese et al., 2011](#)). This may suggest that clients decide to terminate when they feel that they have reached an acceptable level of improvement. Concurrently, our findings are at odds with [Barkham et al.'s \(2006\)](#) in that they found the length of treatment to be negatively related to the proportion of clinically significant improvement at discharge ( $r = -.91$ ). Consistent with [Baldwin et al. \(2009\)](#), we found the rate of clinically significant improved clients to increase with the total number of sessions attended ( $r = .71$ ).

Taking these findings together, our study provides support for both the dose–effect and the GEL models of change in psychotherapy. As proposed by the dose–effect model, the expected course of improvement in psychotherapy follows a negatively accelerated pattern, irrespective of the duration of the treatment. However, the dose–effect model must not be understood to imply that the rate of the negative acceleration is constant across all treatment lengths. Given that the rate of change varies as a function of treatment duration, we can no longer argue for a consistent relationship between the expected effect of treatment and number of sessions across clients (e.g., a 75% chance of improvement after 26 sessions of treatment for every client). The amount of improvement must be evaluated in light of whether the treatment has concluded or is ongoing. As proposed by the GEL

model, people seem to remain in therapy until they reach a GEL, leading to termination. This level of improvement, however, is typically reached by a log-linear rather than by a linear and steady pattern of change.

### Limitations

Our study has several limitations. First, the study was based on naturalistic data. While this may suggest higher external validity, it implies lower internal validity compared to experimentally controlled studies. The current study shares typical limitations of studies conducted under routine care conditions, including limited information on clients, diagnoses, therapists, and therapies. An additional consequence of a naturalistic design is the possibility that there are unobserved variables associated with treatment progress that were not controlled. In particular, the treatment response curves presented in this study can be argued as a combination of both the effects of psychotherapy and of “spontaneous” recovery, as the data do not allow for modeling a nontreatment control group. However, given that the majority of controlled studies have shown that psychotherapy is far superior to spontaneous recovery ([Lambert, 2013](#)) and in light of the observed absolute pre-post treatment effect size, we believe that what was modeled using this data reflects actual treatment progress.

Another potential limitation is that we chose the number of sessions as the dose of treatment to determine the dose–effect relationship. The underlying assumption is that the number of sessions is related to the amount of exposure to the active ingredients in psychotherapy ([Howard et al., 1986](#)). Although some findings suggest that the time between sessions (i.e., session frequency) is also related to the rate of change ([Reese et al., 2011](#)), it is unlikely that session frequency would affect the trajectory of change (i.e., negative acceleration) even if it may contribute to interindividual variability in growth rates.

Third, our analyses are limited to the first 20 sessions of treatment. Thus, there is no evidence that our findings generalize to treatments beyond 20 sessions.

Fourth, our clients were treated at different types of centers (college counseling centers, primary care medical centers, and private mental health centers), and not all of them had complete session-by-session outcome data. However, the majority of the clients in our sample were treated at university counseling centers (94%) and most of the clients had no missing data points (97%). Thus, it is unlikely that our findings were affected by missing data or by varying types of treatment centers. When rerunning the LGCM analyses for the subsample of those 5,836 (92%) clients who were treated at university counseling centers and who had no missing values in the outcome variables, the log-linear LGCMs again consistently outperformed the linear models irrespective of treatment duration.

Fifth, therapists were provided regular feedback regarding their clients’ BHM scores. Thus, our findings may not generalize to settings where no feedback is given to therapists. As Lambert and others (e.g., [Lambert, 2010](#)) have shown, feedback to therapists over the course of treatment tends to improve the outcomes of clients.

Sixth, we used a single, client-reported general clinical symptom measure as the outcome variable. Whether our findings generalize to different measures, domains, and perspectives remains an empirical issue.

Last, it is important to note that while the log-linear growth models explained change in outcome consistently better than the linear models, both types of models are nevertheless aggregated approximations of treatment progress. As individual client's progress varies substantially from the average course of treatment, these models should not be considered trajectories that every client should follow. Research has shown that large between-session changes occur during the course of treatment (so-called "sudden gains" and "sudden losses"; e.g., [Tang & DeRubeis, 1999](#)). Thus, not surprisingly, the proportion of unexplained variance is considerably large even when modeled with log-linear LGCMs, especially for longer treatments. Nevertheless, we believe that modeling the negatively accelerating rate of change has benefits over and above linear models in predicting treatment progress.

## Implications

These findings have important implications for psychotherapy research, practice, and policy. Consistent with previous research, our results suggest that client-focused research should model the course of treatment assuming a negatively accelerating rather than a linear rate of change ([Lueger, Lutz, & Howard, 2000](#); [Lutz et al., 2001](#); [Lutz et al., 1999](#); [Stulz, Gallop, Lutz, Wrenn, & Crits-Christoph, 2010](#); [Stulz & Lutz, 2007](#); [Stulz et al., 2007](#)).

For clinicians, it is important to bear in mind that the impact of each session may diminish as treatment length increases. While the rate of change may differ considerably across clients and is typically related to the total number of sessions that the client attends, on an intraindividual level, a negatively accelerating rate of improvement may be expected. However, this does not mean that treatments should be brief. When the curve starts accelerating more slowly, this does not necessarily indicate termination or that clients will not improve with more treatment, as the percentage of clinically significant improvement was shown to increase with the total number of sessions attended. Furthermore, our data indicate considerable variability with the clients' individual slopes, particularly for those terminating treatments early. Some clients may have terminated early because they felt better, whereas others may have terminated early because therapy was not working well or for other reasons.

Policy makers should be aware that there is no prespecified dose of therapy that could be recommended across all clients. Despite the recent surge of time-limited psychotherapy, these session limits have been determined arbitrarily. To achieve optimal treatment outcomes from the perspectives of both individual clients *and* organizations (e.g., third-party payers), resource allocation should be based not on a predetermined number of sessions that a particular client should be offered but rather on the client's actual degree of improvement based on psychometric progress monitoring.

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*Submitted:* January 15, 2013 *Revised:* May 7, 2013 *Accepted:* May 8, 2013

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**Source:** Journal of Counseling Psychology. Vol. 60. (4), Oct, 2013 pp. 593-600)

**Accession Number:** 2013-23513-001

**Digital Object Identifier:** 10.1037/a0033589