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John Rice
Washington University School of Medicine in St. Louis

et al

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Review article

Interaction between the *FTO* gene, body mass index and depression: meta-analysis of 13701 individuals[†]

Margarita Rivera, Adam E. Locke, Tanguy Corre, Darina Czamara, Christiane Wolf, Ana Ching-Lopez, Yuri Milaneschi, Stefan Kloiber, Sara Cohen-Woods, James Rucker, Katherine J. Aitchison, Sven Bergmann, Dorret I. Boomsma, Nick Craddock, Michael Gill, Florian Holsboer, Jouke-Jan Hottenga, Ania Korszun, Zoltan Kutalik, Susanne Lucae, Wolfgang Maier, Ole Mors, Bertram Müller-Myhsok, Michael J. Owen, Brenda W. J. H. Penninx, Martin Preisig, John Rice, Marcella Rietschel, Federica Tozzi, Rudolf Uher, Peter Vollenweider, Gerard Waeber, Gonneke Willemsen, Ian W. Craig, Anne E. Farmer, Cathryn M. Lewis, Gerome Breen and Peter McGuffin

Background
Depression and obesity are highly prevalent, and major impacts on public health frequently co-occur. Recently, we reported that having depression moderates the effect of the *FTO* gene, suggesting its implication in the association between depression and obesity.

Aims
To confirm these findings by investigating the *FTO* polymorphism rs9939609 in new cohorts, and subsequently in a meta-analysis.

Method
The sample consists of 6902 individuals with depression and 6799 controls from three replication cohorts and two original discovery cohorts. Linear regression models were performed to test for association between rs9939609 and body mass index (BMI), and for the interaction between rs9939609 and depression status for an effect on BMI. Fixed and random effects meta-analyses were performed using METASOFT.

Results
In the replication cohorts, we observed a significant interaction between *FTO*, BMI and depression with fixed effects meta-analysis ($\beta = 0.12$, $P = 2.7 \times 10^{-4}$) and with the Han/Eskin random effects method ($P = 1.4 \times 10^{-7}$) but not with traditional random effects ($\beta = 0.1$, $P = 0.35$). When combined with the discovery cohorts, random effects meta-analysis also supports the interaction ($\beta = 0.12$, $P = 0.027$) being highly significant based on the Han/Eskin model ($P = 6.9 \times 10^{-8}$). On average, carriers of the risk allele who have depression have a 2.2% higher BMI for each risk allele, over and above the main effect of *FTO*.

Conclusions
This meta-analysis provides additional support for a significant interaction between *FTO*, depression and BMI, indicating that depression increases the effect of *FTO* on BMI. The findings provide a useful starting point in understanding the biological mechanism involved in the association between obesity and depression.

Declaration of interest
K.J.A., A.E.F. and P.M. have received consultancy fees and honoraria for participating in expert panels for pharmaceutical companies including GlaxoSmithKline (GSK). P.M. has received speaker's fees from Pfizer. K.J.A. has been on the advisory board for Bristol-Myers Squibb and Otsuka Pharmaceutical and in addition received consultancy fees including payment for lectures and educational presentations. She was previously a member of other advisory boards, receiving consultancy fees and honoraria, and has received research grants from various companies including Lundbeck and GSK. F.H. is co-founder of the biotech company Holsboer Maschmeyer Neuro Chemie GmbH (HMNC GmbH) in Germany. W.M. is a member of the advisory boards and has received fees for speaking from Lilly and Lundbeck. M.P. is part of advisory boards for Eli Lilly and Lundbeck.

Depression and obesity are leading causes of disease burden and disability, as well as major public health concerns worldwide.[1] Both conditions are highly prevalent and major risk factors for chronic physical diseases such as type 2 diabetes, cardiovascular disease and hypertension.[2] Shared aetiological factors, including genetic risk factors, between depression, obesity and physical disorders have been reported.[2,3]

The nature of the association between obesity and depression remains unclear. In a review and meta-analysis of longitudinal studies, obesity was found to increase the risk of depression, whereas depression was predictive of the development of overweight and obesity, suggesting a bidirectional relationship between depression and obesity.[4]

The *FTO* (fat mass and obesity associated) gene has consistently been associated with common forms of human obesity.[5] Since the 2007 discovery,[6,7] the role of *FTO* in body mass regulation and predisposition to obesity has been confirmed in multiple populations in many independent studies[8–10] (see reference 10 for a review), as well as in large genome-wide association studies (GWAS).[11–13]

Animal studies have shown that *FTO* is widely expressed in the brain, with high expression in hypothalamic nuclei, which are involved in regulating energy balance.[14]

The single-nucleotide polymorphism (SNP) rs9939609 is one of the most extensively studied *FTO* polymorphisms. The body mass index (BMI)-increasing 'A' allele has been associated with increased energy intake[15] and diminished satiety,[16] also implicating *FTO* in appetite regulation. Furthermore, the *FTO* gene, BMI and depression have been associated with structural brain differences in humans.[17,18]

In 2012, we reported the first study identifying an interaction effect between *FTO* genotype, depression and increased risk of obesity[19] in two independent samples of patients with depression and psychiatrically healthy controls, in which the effect of *FTO* was increased in those individuals who had experienced depression.

In the current study, we first replicate the *FTO* rs9939609 depression interaction in 3 independent cohorts and then combine results with the 2 original cohorts for a total combined meta-analysis of 6902 patients with depression and 6799 psychiatrically healthy controls.

Method

Samples

This meta-analysis includes data from five different studies: Radiant, PsyCoLaus, GSK, MARS and NESDA/NTR.

Original cohorts

Extended Radiant

The depressive disorder sample included 2442 individuals sourced from several studies described in detail elsewhere: the Depression Case Control (DeCC) study,[20] Depression Network (DeNT) study[21,22] and the Genome-Based Therapeutic Drugs for Depression (GENDEP) study.[23] The DeCC is a case–control study that recruited individuals from three UK sites (London, Cardiff and Birmingham).[20] The DeNT sibling pair linkage study includes cases of recurrent unipolar depression collected at seven European sites and one US site.[21,22] All participants in the DeCC and DeNT studies had experienced at least two episodes of major depression of at least moderate severity. The GENDEP study includes individuals with one or more episodes of depression of at least moderate severity recruited from nine European centres.[23]

Diagnosis of major depressive disorder (MDD) was ascertained using the Schedules for Clinical Assessment in Neuropsychiatry (SCAN) interview in all three studies.[24] The control sample comprised 809 controls from the UK who were screened for lifetime absence of any psychiatric disorder using a modified version of the Past History Schedule.[25] All cases and controls were of White European ancestry.

PsyCoLaus study

The PsyCoLaus study focused on psychiatric symptoms in a population-based cohort randomly selected from the list of residents of the city of Lausanne (Switzerland) and originally assessed for cardiovascular risk factors (CoLaus study). All 35- to 66-year-old individuals of the CoLaus sample were invited to participate in the psychiatric evaluation for the PsyCoLaus substudy (see Firmann *et al*[26] and Preisig *et al*[27] for a detailed description). The PsyCoLaus sample included 1296 individuals who fulfilled lifetime criteria for MDD according to DSM-IV based on assessment using the Diagnostic Interview for Genetic Studies (DIGS).[28] The control sample included 1698 PsyCoLaus participants who had never fulfilled criteria for MDD. The PsyCoLaus study has been described in more detail elsewhere.[27]

Replication cohorts

GSK study

Participants from this study were recruited at the Max Planck Institute of Psychiatry in Munich, Germany, and at two satellite recruiting hospitals in the Munich area (BKH Augsburg and Klinikum Ingolstadt). A total of 821 White European individuals diagnosed with recurrent MDD and 856 White European age- and gender-matched unaffected controls were included in the meta-analysis. All patients with depression were evaluated using the SCAN interview and were included in the study if they had experienced at least two moderate to severe depression episodes according to DSM-IV. Controls were excluded if anxiety and mood disorders were present using the Composite International Diagnostic Screener (CIDS). The study has been described in detail elsewhere.[29]

MARS study

The Munich Antidepressant Response Signature (MARS) project at the Max Planck Institute of Psychiatry in Munich, Germany (www.mars-depression.de) is a naturalistic clinical study of in-patients with a major depressive episode. Individuals with a DSM-IV diagnosis of major depressive episode or recurrent depression ($n = 575$) were included in this meta-analysis. Five hundred and forty-one controls randomly selected from Munich community registries and screened for the absence of a lifetime history of DSM-IV Axis I disorders were included in the analyses. All patients and controls were of White European origin. Study details have been described previously.[30]

NESDA/NTR studies

This sample is part of the Netherlands Study of Depression and Anxiety (NESDA)[31] and the Netherlands Twin Register (NTR).[32] NESDA is a naturalistic multicentre and longitudinal cohort study designed to examine the course and consequences of individuals with depressive and anxiety disorders. Recruitment of participants was from the general population, general practices and mental health organisations.

The NTR project, in 1994, started collecting longitudinal data from twins and their families to create a resource for genetic studies on health, lifestyle and personality.

In both cohorts, similar inclusion and exclusion criteria were used to select MDD cases. The Composite Interview Diagnostic Interview (CIDI)[33] was used to diagnose depressive disorders according to the DSM-IV criteria. The control group had no lifetime diagnosis of depression or anxiety disorders. Controls were partly confirmed by the absence of lifetime diagnoses of psychiatric disorders, and partly by repeated measures of low genetic liability for MDD (determined by factor score derived from longitudinal measures of neuroticism, anxiety and depressive symptoms).[34]

Participants included in these studies were required to report western European ancestry. These studies have been previously described in more detail.[31,32]

The case sample comprised 1636 individuals from NESDA and 132 from NTR. The controls were mainly from NTR ($n = 2470$) with 424 additional controls from NESDA.

Phenotypic data

In all studies, BMI was defined as weight in kilograms divided by height in metres squared (kg/m^2). In Radiant, self-reported height and weight were obtained during the SCAN interview for cases and telephone interview for controls. The reliability of self-report of height and weight was assessed in the GENDEP data-set ($n = 811$) where we had also measured height and weight. The correlations for measured *v.* self-reported height, weight and BMI were 0.97, 0.95 and 0.95, respectively.

In the PsyCoLaus sample, weight and height were measured at the out-patient clinic at the Centre Hospitalier Universitaire Vaudois (CHUV).[26] In the GSK and MARS studies, anthropometric measures for patients and controls were taken at the Max Planck

Institute and associated study sites by trained technicians and study nurses.[29,30]

Weight and height were measured by medical examination at the study clinic during the visit for NESDA,[31] and during the home visit after blood sampling for NTR.[32]

In all studies the distribution of BMI was positively skewed. We therefore transformed the data to $\log_{10}$(BMI) to achieve a closer approximation to normal distribution.

Genotyping

The samples from the different studies were all genotyped with SNP arrays. If rs9939609 was not genotyped directly on the array, genotypes were imputed and best guess genotypes were used to perform the statistical analyses. For the Radiant study, we report results from a larger sample than previously reported for the UK subsample ($n = 2174$).[19] A thorough description of all genotyping and imputation is described for each study in more detail elsewhere.[29,30,32,35–37]

Inclusion criteria

As common inclusion criteria, we looked for studies with information available on a lifetime DSM-IV diagnosis of MDD, BMI and genotype data for the rs9939609 *FTO* polymorphism. Homogeneous ethnicity (White European) was also required for each study to be included in the meta-analysis to reduce the risk of population stratification. Demographic and clinical characteristics of the participants from the five studies included in the meta-analysis are summarised in Table 1.

Statistical analyses

In each study, linear regression models were performed to test for association between rs9939609 polymorphism and $\log_{10}$(BMI) assuming an additive genetic model. Models were tested separately in the cases and controls and in the combined sample. We then tested the interaction effect between rs9939609 variant and depression status on $\log_{10}$(BMI). Gender and age were included as covariates in the regression analyses. Genotype-based principal components were used to control for possible population stratification within each study. Standardised beta coefficients were obtained in each study to allow direct comparison between studies. Statistical analyses were performed using PLINK v1.07.[38]

Meta-analyses of main association effects

Fixed effects meta-analyses of the association between rs9939609 variant and $\log_{10}$(BMI) were performed in the whole sample and in cases and controls separately, using PLINK v1.07.[38]

Heterogeneity across studies was assessed using Cochran's Q statistic and I^2 heterogeneity index.

Meta-analysis of the interaction effect

Fixed effects and random effects meta-analyses based on inverse-variance-weighted effect size of the interaction effects were performed using METASOFT,[39] (http://genetics.cs.ucla.edu/meta/index.html). Heterogeneity across studies was assessed using Cochran's Q statistic and I^2 heterogeneity index.

Results

Main effects of the association between *FTO* and BMI

Original cohorts: Extended Radiant and PsyCoLaus

In Radiant, as previously reported in a subset of the same data,[19] there was a significant association, although strengthened in significance here, between the rs9939609 A variant and $\log_{10}$(BMI) ($\beta = 0.08$, $P = 0.001$) in the whole sample.

Linear regression analysis in PsyCoLaus also showed that the rs9939609 A variant was significantly associated with $\log_{10}$(BMI) ($\beta = 0.07$, $P = 0.006$).

Moreover, in both studies the association was strengthened when analysing the cases alone (Radiant: $\beta = 0.12$, $P = 6.15 \times 10^{-5}$; PsyCoLaus: $\beta = 0.12$, $P = 8.52 \times 10^{-4}$). The analyses in the control groups alone showed no significant association with $\log_{10}$(BMI) (Table 2).

Replication cohorts: GSK, MARS and NESDA/NTR

In the GSK and MARS studies, there were no statistically significant associations between the *FTO* rs9939609 A variant and $\log_{10}$(BMI) (Table 2). In NESDA/NTR, the rs9939609 A allele was associated with $\log_{10}$(BMI) in the whole sample ($\beta = 0.09$, $P = 1.24 \times 10^{-5}$) and showed a stronger effect in the depression group alone than in the combined sample ($\beta = 0.2$, $P = 2.26 \times 10^{-8}$), whereas the association in controls was not significant (Table 2).

In all studies, depression status, gender, age and principal components were included as covariates when the analyses were performed in the combined sample; gender, age and principal components were included when cases and controls were analysed separately. The results for the association analyses between the rs9939609 polymorphism and $\log_{10}$(BMI) in the whole sample, and in cases and controls separately, for each individual study are shown in Table 2.

We also explored the association between *FTO* gene and depression in all the studies and found no association between the rs9939609 A risk variant and depression (data not shown).

Table 1	Demographic and clinical characteristics of the participants from the studies included in the meta-analysis									
	Radiant		PsyCoLaus		GSK		MARS		NESDA/NTR	
	Cases	Controls	Cases	Controls	Cases	Controls	Cases	Controls	Cases	Controls
Sample size, *n*	2442	809	1296	1698	821	856	575	541	1768	2895
Gender, %										
Male	30.3	38.7	33.49	57.36	33.74	32.48	47.13	44.92	31.4	38.7
Female	69.7	61.3	66.51	42.64	66.26	67.52	52.87	55.08	68.6	61.3
Mean age (s.d.), years	45.25 (12.15)	39.90 (13.71)	49.69 (8.68)	50.59 (8.94)	50.94 (13.74)	51.92 (13.26)	48.09 (13.95)	47.42 (13.50)	42.68 (12.41)	42.83 (14.96)
Mean body mass index (s.d.), kg/m²	26.62 (4.32)	25.18 (4.31)	26.04 (3.84)	26.39 (3.80)	26.25 (4.89)	24.62 (4.08)	25.28 (4.52)	26.04 (4.72)	25.74 (5.13)	24.52 (4.00)

GSK, GlaxoSmithKline study; MARS, Munich Antidepressant Response Signature project; NESDA/NTR, Netherlands Study of Depression and Anxiety/Netherlands Twin Register.

Table 2 Association results between the rs9939609 polymorphism and standardised $\log_{10}$(BMI) and fixed effects meta-analyses in the whole sample and in cases and controls separately

Study	n	Whole sample			Cases			Controls		
		β	s.e.	P	β	s.e.	P	β	s.e.	P
Radiant	3251	0.08	0.02	0.001	0.12	0.03	6.2×10^{-5}	−0.04	0.04	0.3
PsyCoLaus	2994	0.07	0.02	0.006	0.12	0.04	8.5×10^{-4}	0.03	0.03	0.44
GSK	1677	0.04	0.03	0.193	0.01	0.05	0.89	0.09	0.04	0.04
MARS	1116	0.06	0.04	0.119	0.08	0.06	0.16	0.04	0.05	0.45
NESDA/NTR	4663	0.09	0.02	1.2×10^{-5}	0.2	0.04	2.3×10^{-8}	0.01	0.02	0.56
Meta-analysis: fixed effects	13701	0.07	0.01	1.3×10^{-12}	0.12	0.02	6.9×10^{-11}	0.02	0.01	0.15

BMI, body mass index; GSK, GlaxoSmithKline study; MARS, Munich Antidepressant Response Signature project; NESDA/NTR, Netherlands Study of Depression and Anxiety/Netherlands Twin Register.

Meta-analyses of the main effects of the association between *FTO* and BMI

Fixed effects meta-analysis supports a significant association between rs9939609 polymorphism and $\log_{10}$(BMI) in the whole sample and in cases (whole sample: β = 0.07, P = 1.2910^{-12}; cases: β = 0.12, P = 6.9210^{-12}). There was no association between rs9939609 and $\log_{10}$(BMI) in controls (β = 0.02, P = 0.15).

No significant heterogeneity was detected among studies (whole sample: Q = 0.7, I^2 = 0; cases: Q > 0.05, I^2 = 56.88; controls: Q = 0.21, I^2 = 31.43). The results for the fixed effects meta-analyses in each group are shown in Table 2.

Interaction between *FTO*, BMI and depression

Original studies: Extended Radiant and PsyCoLaus

In the analysis of updated data from the original studies, we confirmed the significant interaction effect on $\log_{10}$(BMI) between

rs9939609 genotype and depression that we had previously published (Radiant: β = 0.18, P = 0.002; PsyCoLaus: β = 0.12, P = 0.034) (Table 3). The P-value for the interaction results in the Radiant extended sample is lower than previously reported (P = 0.005).[19]

Replication studies: GSK, MARS and NESDA/NTR

There was also no significant interaction effect in GSK (β = −0.09, P = 0.168) (Table 3). The interaction effect in MARS was consistent with NESDA/NTR (β = 0.26), but was not significant (P = 0.083), likely reflecting the smaller sample size in MARS. The NESDA/NTR studies showed a significant interaction between rs9939609 genotype and depression status in relation to $\log_{10}$(BMI) (NESDA/NTR: β = 0.19, P = 3.22×10^{-6}), replicating our previous findings.[19]

Meta-analysis of the interaction effect between rs9939609 and depression in the three replication cohorts was also consistent with our earlier finding. In the fixed effects analysis there was

Table 3 Interaction results between rs9939609 risk allele and depression on $\log_{10}$(BMI) in the five independent studies and fixed effects, random effects and Han/Eskin model meta-analyses

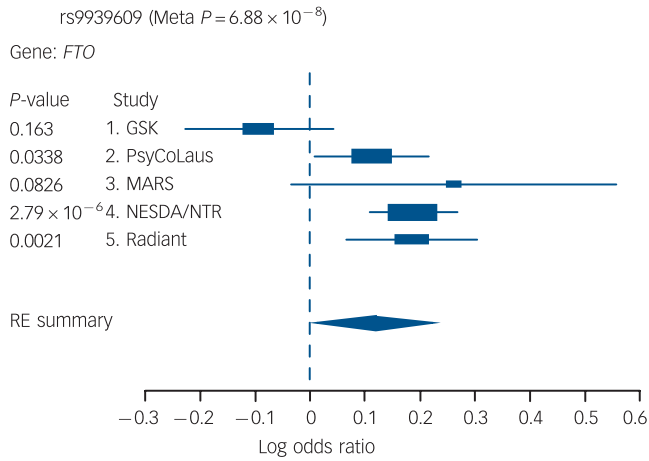|	n	Interaction		
		β	s.e.	P
Study				
Radiant	3251	0.18	0.06	0.002
PsyCoLaus	2994	0.12	0.05	0.034
GSK	1677	−0.09	0.07	0.168
MARS	1116	0.26	0.15	0.083
NESDA/NTR	4663	0.19	0.04	3.2×10^{-6}
Meta-analysis				
Fixed effects				
Replication studies[a]	7456	0.12	0.03	2.7×10^{-4}
Replication studies (no GSK)	5779	0.19	0.04	6.6×10^{-7}
All studies	13701	0.13	0.03	3.1×10^{-7}
All studies (no GSK)	12024	0.17	0.03	1.1×10^{-9}
Random effects				
Replication studies[a]	7456	0.10	0.11	0.35
Replication studies (no GSK)	5779	0.19	0.04	6.6×10^{-7}
All studies	13701	0.12	0.05	0.02
All studies (no GSK)	12024	0.17	0.03	1.110^{-9}
Han/Eskin model				
Replication studies[a]	7456	0.10	0.11	1.410^{-5}
Replication studies (no GSK)	5779	0.19	0.04	7.910^{-7}
All studies	13701	0.12	0.05	6.910^{-8}
All studies (no GSK)	12024	0.17	0.03	1.710^{-9}
Meta-analysis				
Replication studies	7456	85.3[b]	13.614[c]	0.001[d]
Replication studies (no GSK)	5779	0	0.221[c]	0.638[d]
All studies	13701	72.6[b]	14.608[c]	0.006[d]
All studies (no GSK)	12024	0	1.828[c]	0.609[d]

BMI, body mass index; GSK, GlaxoSmithKline study; MARS, Munich Antidepressant Response Signature project; NESDA/NTR, Netherlands Study of Depression and Anxiety/Netherlands Twin Register.
a. Replication studies: GSK, MARS and NESDA/NTR.
b. I^2 statistic.
c. Q statistic.
d. P-value of Q.

Fig. 1 Forest plot showing interactions between *FTO*, depression and body mass index.

GSK, GlaxoSmithKline study; MARS, Munich Antidepressant Response Signature project; NESDA/NTR, Netherlands Study of Depression and Anxiety/Netherlands Twin Register; RE, risk estimate.

directional consistency and a significant association, with a greater increase in BMI for individuals with depression carrying rs9939609 alleles ($\beta = 0.12$, $P = 2.72 \times 10^{-4}$). Owing to the opposite but non-significant interaction effect observed in GSK, a traditional random effects meta-analysis was not significant ($\beta = 0.10$, $P = 0.35$). However, the Han/Eskin random effects model,[39] optimised to detect an effect in the presence of heterogeneity, was able to detect the interaction effect ($P = 1.35 \times 10^{-7}$) (Table 3).

Five study meta-analysis of the interaction effect

Fixed effects meta-analysis in all five studies further confirmed the interaction between *FTO*, BMI and depression ($\beta = 0.13$, $P = 3.1 \times 10^{-7}$) (Table 3). A forest plot showing the interaction effect in each study as well as the fixed effects meta-analysis is shown in Fig. 1. The effect corresponds to an additional increase of 2.2% in BMI for subjects with depression for each risk allele, over and above the main effect of *FTO*. This translates to a 1.76 kg weight increase in an individual whose height and weight were 180 cm and 8 kg. Similar to meta-analysis in the three new cohorts, traditional random effects meta-analysis was only nominally significant ($\beta = 0.12$, $P = 0.02$), but highly significant based on the Han/Eskin model ($P = 6.89 \times 10^{-8}$) (Table 3).

Heterogeneity and sensitivity analyses

Both the three-study replication meta-analysis and the full meta-analysis showed significant heterogeneity among studies, using both Cochran's Q statistic ($P_{3\text{-way}} = 0.001$, $P_{5\text{-way}} = 0.006$) and I^2 (3-way index 85.3%; 5-way index 72.6%). GSK accounted for the full extent of the heterogeneity. In sensitivity analyses of both the replication and full meta-analysis, removing GSK increased the significance of both fixed effects and random effects associations and eliminated all evidence of heterogeneity (Table 3).

Discussion

Main findings

The aim of this study was to confirm our previously reported interaction effect between *FTO* polymorphism and depression on BMI, conducting a replication meta-analysis in three new studies and a combined meta-analysis including 13 701 individuals from 5 studies.[19]

The main effect meta-analysis showed a significant association between *FTO* and BMI in the whole sample. This association was attributed to the case group alone, with no association found in controls, replicating our previous findings.[19]

In the Radiant, GSK and NESDA/NTR studies, patients with depression had higher BMI than controls. This could be attributable to a side-effect of antidepressant treatment or because individuals with depression are less physically active and/or have an increased food intake. In MARS, BMI in controls was higher than in patients with depression, and this could be because this study includes patients with untreated, first-episode depression. In the PsyCoLaus sample, BMI was not different between cases and controls. These differences possibly reflect the fact that PsyCoLaus participants were recruited from the community and were an arguably less severely affected group (18% were experiencing a current depressive episode and 82% were in remission), where only one episode was required for inclusion and only 37.5% of cases had ever received treatment with antidepressants.

In addition to this, the control samples were screened to have no history of psychiatric disorders. Previous studies investigating BMI, obesity and *FTO* have not taken this into account and it could further explain why there is no association observed in the control samples.

Another plausible explanation for the BMI differences across studies could be the clinical heterogeneity of depression. Evidence suggests that metabolic dysregulation may be more involved in one subtype of depression (atypical) than in another (melancholic).[40] Recently, it has been shown that obesity and increased appetite are more prevalent in atypical depression, whereas rates of obesity are similar or even lower compared with controls in melancholic (typical) depression.[40,41] Therefore, when considering the overall depression diagnosis these differential effects are blurred.

Unfortunately, we could not include medication as a covariate in the analyses as this information was not available for all the studies. Therefore, we cannot exclude the possibility that the FTO effect that we have found is at least partly reflecting an increased susceptibility to the weight-inducing effects of medications. We also cannot exclude the possibility that psychiatric conditions that are frequently comorbid with depression play a part in modulating the effect of *FTO*.

The involvement of overlapping physiological mechanisms and shared genes between depression and obesity could support the hypothesis that the two disorders have shared genetic vulnerability.[3] In 2010, a systematic review and meta-analysis on the longitudinal relationship between depression, overweight and obesity confirmed a bidirectional association between depression and obesity.[4] Several lines of evidence support the possibility of a biological pathway. Metabolic, immune–inflammatory and hypothalamic–pituitary–adrenal (HPA) axis dysregulations could be mediators of the reported association as they have a role in both depression and obesity.[40]

Psychological factors such as body dissatisfaction, low self-esteem, stigmatisation and eating patterns should also be considered in addition to the biological mechanisms and could further contribute to the observed association between depression and obesity.[42,43] Reduced physical activity, sedentary lifestyle and/or unhealthy dietary choices as well as antidepressant treatment could be additional risk factors that induce weight gain in individuals with depression who are genetically predisposed to the disorder.

Unfortunately, measures of important confounding factors such as smoking, alcohol consumption or socioeconomic status, which might influence the association between higher BMI and MDD, were not available for all the studies.

The results from the meta-analysis of the interaction effect suggest a genetic mechanism by which individuals who have depression are at increased risk for obesity. Our results demonstrate that depression enhances the effect of *FTO* variants on BMI, such that individuals with depression have an additional 2.2% increase in BMI for each rs9939609 risk allele (A) compared with psychiatrically healthy controls.

Limitations and conclusions

The main limitation of this study is that the inclusion criteria for participants, study design, recruitment and sample composition vary across the studies. This could explain the significant heterogeneity found between studies.

To our knowledge this is the largest and most comprehensive study and meta-analysis investigating the interaction between *FTO*, BMI and depression concurrently. The overall interaction meta-analysis results suggest that having depression moderates the effect of *FTO* on BMI, such that the BMI-increasing effect is significantly enlarged. This meta-analysis demonstrates a modest but a consistent effect of the interaction between *FTO*, depression and BMI.

Although our analyses cannot infer causality or directionality about the relationship between obesity and depression, this study provides additional evidence that shared genetic factors between depression and obesity do exist. Furthermore, it is evident that *FTO* gene, BMI and depression influence brain structure.[17,18] Altogether, our results indicate that depression-related alterations in key biological processes may interact with the *FTO* risk allele to increase BMI or obesity risk. Future studies that include samples followed longitudinally will be crucial to better understand the nature and direction of this association.

Overall, our findings provide evidence that *FTO* is involved in the association between obesity and depression. Although *FTO* genotyping has modest implications for predicting which patients with depression are at risk of BMI-related disorders, the findings provide a useful starting point in understanding the biological mechanism involved in the association between obesity and depression. The identification of such mechanisms should in turn lead to better understanding of the development of comorbid states and eventually contribute to prevention of obesity-related disorders that are currently overrepresented among patients with depression.[2]

Margarita Rivera, PhD, Department of Biochemistry and Molecular Biology II and Institute of Neurosciences, Center for Biomedical Research, University of Granada, Granada, Spain, and MRC Social Genetic and Developmental Psychiatry Centre, Institute of Psychiatry, Psychology & Neuroscience, Kings College London, UK; **Adam E. Locke**, PhD, Department of Biostatistics and Center for Statistical Genetics, University of Michigan, Ann Arbor, Michigan, USA; **Tanguy Corre**, PhD, Department of Medical Genetics, University of Lausanne, Lausanne, and Swiss Institute of Bioinformatics, Lausanne, Switzerland; **Darina Czamara**, PhD, **Christiane Wolf**, PhD, Max Planck Institute of Psychiatry, Munich, Germany; **Ana Ching-Lopez**, Department of Psychiatry, School of Medicine, University of Granada, and Institute of Neurosciences Federico Olóriz, Centro de Investigación Biomédica, University of Granada, Spain; **Yuri Milaneschi**, PhD, Department of Psychiatry and EMGO Institute for Health and Care Research, VU University Medical Center/GGZ in Geest, Amsterdam, The Netherlands; **Stefan Kloiber**, MD, Max Planck Institute of Psychiatry, Munich, Germany; **Sara Cohen-Woods**, PhD, School of Psychology, Flinders University, Adelaide, South Australia, Australia; **James Rucker**, MD, PhD, MRC Social Genetic and Developmental Psychiatry Centre, Institute of Psychiatry, Psychology & Neuroscience, King's College London, UK; **Katherine J. Aitchison**, MD, PhD, Department of Psychiatry, University of Alberta, Alberta, Canada; **Sven Bergmann**, PhD, Department of Medical Genetics, University of Lausanne, Lausanne, and Swiss Institute of Bioinformatics, Lausanne, Switzerland; **Dorret I. Boomsma**, PhD, Department of Biological Psychology, VU University Amsterdam, Amsterdam, The Netherlands; **Nick Craddock**, MB, PhD, FMedSci, Department of Psychological Medicine and Neurology, Cardiff University School of Medicine, Henry Wellcome Building, Cardiff, UK; **Michael Gill**, MD, Department of Psychiatry, Trinity Centre for Health Sciences, Dublin 8, Ireland; **Florian Holsboer**, MD, PhD, Max Planck Institute of Psychiatry, Munich, Germany; **Jouke-Jan Hottenga**, PhD, Department of Psychiatry, University of Alberta, Alberta, Canada; **Ania Korszun**, PhD, Barts and The London School of Medicine and Dentistry, Queen Mary University of London, London, UK; **Zoltan Kutalik**, PhD, Department of Medical Genetics, University of Lausanne, Lausanne, and Swiss Institute of Bioinformatics, Lausanne, Switzerland; **Susanne Lucae**, MD, PhD, Max Planck Institute of Psychiatry, Munich, Germany; **Wolfgang Maier**, MD, Department of Psychiatry, University of Bonn, Bonn, Germany; **Ole Mors**, MD, PhD, Research Department P, Aarhus University Hospital, Risskov, Denmark; **Bertram Müller-Myhsok**, MD, Max Planck Institute of Psychiatry, Munich, Germany; **Michael J. Owen**, MB, PhD, FMedSci, MRC Centre for Neuropsychiatric Genetics and Genomics, Department of Psychological Medicine and Neurology, School of Medicine, Cardiff University, Cardiff, UK; **Brenda W. J. H. Penninx**, PhD, Department of Psychiatry and EMGO Institute for Health and Care Research, VU University Medical Center/GGZ in Geest, Amsterdam, The Netherlands; **Martin Preisig**, MD, Department of Psychiatry, Lausanne University Hospital, 1008 Prilly-Lausanne, Switzerland; **John Rice**, PhD, Department of Psychiatry, Washington University School of Medicine, St Louis, Missouri, USA; **Marcella Rietschel**, MD, Central Institute of Mental Health, Mannheim, Germany; **Federica Tozzi**, MD, Genetics Division, Drug Discovery, GlaxoSmithKline Research and Development, Verona, Italy; **Rudolf Uher**, MD, PhD, MRC Social Genetic and Developmental Psychiatry Centre, Institute of Psychiatry, Psychology & Neuroscience, King's College London, UK, and Department of Psychiatry, Dalhousie University, Halifax, Nova Scotia, Canada; **Peter Vollenweider**, MD, PhD, **Gerard Waeber**, MD, PhD, Division of Internal Medicine, CHUV, Lausanne, Switzerland; **Gonneke Willemsen**, PhD, Department of Psychiatry, University of Alberta, Alberta, Canada; **Ian W. Craig**, PhD, **Anne E. Farmer**, MD, MRC Social Genetic and Developmental Psychiatry Centre, Institute of Psychiatry, Psychology & Neuroscience, King's College London, UK; **Cathryn M. Lewis,** PhD, MRC Social, Genetic and Developmental Psychiatry Centre, Institute of Psychiatry, Psychology & Neuroscience, King's College London, and Department of Medical and Molecular Genetics, School of Medicine, King's College London, UK; **Gerome Breen**, PhD, **Peter McGuffin**, MB, PhD, FMedSci, MRC Social Genetic and Developmental Psychiatry Centre, Institute of Psychiatry, Psychology & Neuroscience, King's College London, UK

Correspondence: Margarita Rivera, PhD, Center for Biomedical Research, University of Granada, Avda. del Conocimiento, s/n, Armilla, Granada 18016, Spain. Email: margarita.rivera_sanchez@kcl.ac.uk

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References

1 Mathers CD, Loncar D. Projections of global mortality and burden of disease from 2002 to 2030. *PLoS Med* 2006; **3**: e442.

2 Farmer A, Korszun A, Owen MJ, Craddock N, Jones L, Jones I, et al. Medical disorders in people with recurrent depression. *Br J Psychiatry* 2008; **192**: 351–5.

3 Afari N, Noonan C, Goldberg J, Roy-Byrne P, Schur E, Golnari G, et al. Depression and obesity: do shared genes explain the relationship? *Depress Anxiety* 2010; **27**: 799–806.

4 Luppino FS, de Wit LM, Bouvy PF, Stijnen T, Cuijpers P, Penninx BW, et al. Overweight, obesity, and depression: a systematic review and meta-analysis of longitudinal studies. *Arch Gen Psychiatry* 2010; **67**: 220–9.

5 Loos RJ, Bouchard C. FTO: the first gene contributing to common forms of human obesity. *Obes Rev* 2008; **9**: 246–50.

6 Frayling TM, Timpson NJ, Weedon MN, Zeqqini E, Frealthy RM, Lindgren CM, et al. A common variant in the FTO gene is associated with body mass index and predisposes to childhood and adult obesity. *Science* 2007; **316**: 889–94.

7 Dina C, Meyre D, Gallina S, Durand E, Körner A, Jacobson P, et al. Variation in FTO contributes to childhood obesity and severe adult obesity. *Nat Genet* 2007; **39**: 724–6.

8 Cotsapas C, Speliotes EK, Hatoum IJ, Greenawalt DM, Dobrin T, Lum PY, et al. Common body mass index-associated variants confer risk of extreme obesity. *Hum Mol Genet* 2009; **18**: 3502–7.

9 Hinney A, Nguyen TT, Scherag A, Friedel S, Brönner G, Müller TD, et al. Genome wide association (GWA) study for early onset extreme obesity supports the role of fat mass and obesity associated gene (FTO) variants. *PLoS ONE* 2007; **2**: e1361.

10 Fawcett KA, Barroso I. The genetics of obesity: FTO leads the way. *Trends Genet* 2010; **26**: 266–74.

11 Scuteri A, Sanna S, Chen WM, Uda M, Albai G, Strait J, et al. Genome-wide association scan shows genetic variants in the FTO gene are associated with obesity-related traits. *PLoS Genet* 2007; **3**: e115.

12 Speliotes EK, Willer CJ, Berndt SI, Monda KL, Thorleifsson G, Jackson AU, et al. Association analyses of 249,796 individuals reveal 18 new loci associated with body mass index. *Nat Genet* 2010; **42**: 937–48.

13 Thorleifsson G, Walters GB, Gudbjartsson DF, Steinthorsdottir V, Sulem P, Helgadottir A, et al. Genome-wide association yields new sequence variants at seven loci that associate with measures of obesity. *Nat Genet* 2009; **41**: 18–24.

14 Gerken T, Girard CA, Tung YC, Webby CJ, Saudek V, Hewitson KS, et al. The obesity-associated FTO gene encodes a 2-oxoglutarate-dependent nucleic acid demethylase. *Science* 2007; **318**: 1469–72.

15 Cecil JE, Tavendale R, Watt P, Hetherington MM, Palmer CN. An obesity-associated FTO gene variant and increased energy intake in children. *N Engl J Med* 2008; **359**: 2558–66.

16 Wardle J, Carnell S, Haworth CM, Farooqi IS, O'Rahilly S, Plomin R. Obesity associated genetic variation in FTO is associated with diminished satiety. *J Clin Endocrinol Metab* 2008; **93**: 3640–3.

17 Cole JH, Boyle CP, Simmons A, Cohen-Woods S, Rivera M, McGuffin P, et al. Body mass index, but not FTO genotype or major depressive disorder, influences brain structure. *Neuroscience* 2013; **252**: 109–17.

18 Ho AJ, Stein JL, Hua X, Lee S, Hibar DP, Leow AD, et al. A commonly carried allele of the obesity-related FTO gene is associated with reduced brain volume in the healthy elderly. *Proc Natl Acad Sci* 2010; **107**: 8404–9.

19 Rivera M, Cohen-Woods S, Kapur K, Breen G, Ng MY, Butler AW, et al. Depressive disorder moderates the effect of the FTO gene on body mass index. *Mol Psychiatry* 2012; **17**: 604–11.

20 Cohen-Woods S, Gaysina D, Craddock N, Farmer A, Gray J, Gunasinghe C, et al. Depression Case Control (DeCC) Study fails to support involvement of the muscarinic acetylcholine receptor M2 (CHRM2) gene in recurrent major depressive disorder. *Hum Mol Genet* 2009; **18**: 1504–9.

21 Farmer A, Breen G, Brewster S, Craddock N, Gill M, Korszun A, et al. The Depression Network (DeNT) Study: methodology and sociodemographic characteristics of the first 470 affected sibling pairs from a large multi-site linkage genetic study. *BMC Psychiatry* 2004; **4**: 42.

22 McGuffin P, Knight J, Breen G, Brewster S, Boyd PR, Craddock N, et al. Whole genome linkage scan of recurrent depressive disorder from the depression network study. *Hum Mol Genet* 2005; **14**: 3337–45.

23 Uher R, Huezo-Diaz P, Perroud N, Smith R, Rietschel M, Mors O, et al. Genetic predictors of response to antidepressants in the GENDEP project. *Pharmacogenomics J* 2009; **9**: 225–33.

24 Wing JK, Babor T, Brugha T, Burke J, Cooper JE, Giel R, et al. SCAN. Schedules for Clinical Assessment in Neuropsychiatry. *Arch Gen Psychiatry* 1990; **47**: 589–93.

25 McGuffin P, Katz R, Aldrich J. Past and present state examination: the assessment of 'lifetime ever' psychopathology. *Psychol Med* 1986; **16**: 461–5.

26 Firmann M, Mayor V, Vidal PM, Bochud M, Pécoud A, Hayoz D, et al. The CoLaus study: a population-based study to investigate the epidemiology and genetic determinants of cardiovascular risk factors and metabolic syndrome. *BMC Cardiovasc Disord* 2008; **8**: 6.

27 Preisig M, Waeber G, Vollenweider P, Bovet P, Rothen S, Vandeleur C, et al. The PsyCoLaus study: methodology and characteristics of the sample of a population-based survey on psychiatric disorders and their association with genetic and cardiovascular risk factors. *BMC Psychiatry* 2009; **9**: 9.

28 Nurnberger Jr JI, Blehar MC, Kaufmann CA, York-Cooler C, Simpson SG, Harkavy-Friedman J, et al. Diagnostic interview for genetic studies. Rationale, unique features, and training. NIMH Genetics Initiative. *Arch Gen Psychiatry* 1994; **51**: 849–59.

29 Lucae S, Salyakina D, Barden N, Harvey M, Gagné B, Labbé M, et al. P2RX7, a gene coding for a purinergic ligand-gated ion channel, is associated with major depressive disorder. *Hum Mol Genet* 2006; **15**: 2438–45.

30 Hennings JM, Owashi T, Binder EB, Horstmann S, Menke A, Kloiber S, et al. Clinical characteristics and treatment outcome in a representative sample of depressed inpatients – findings from the Munich Antidepressant Response Signature (MARS) project. *J Psychiatr Res* 2009; **43**: 215–29.

31 Penninx BW, Beekman AT, Smit JH, Zitman FG, Nolen WA, Spinhoven P, et al. The Netherlands Study of Depression and Anxiety (NESDA): rationale, objectives and methods. *Int J Methods Psychiatr Res* 2008; **17**: 121–40.

32 Willemsen G, de Geus EJ, Bartels M, van Beijsterveldt CE, Brooks AI, Estourgie-van Burk GF, et al. The Netherlands Twin Register biobank: a resource for genetic epidemiological studies. *Twin Res Hum Genet* 2010; **13**: 231–45.

33 Ter Smitten MH, Smeets RMW, Van der Brink W. *Composite International Diagnostic Interview (CIDI), Version 2.1, 12-Months*. World Health Organization, 1998.

34 Major Depressive Disorder Working Group of the Psychiatric GWAS Consortium. A mega-analysis of genome-wide association studies for major depressive disorder. *Mol Psychiatr* 2013; **18**: 497–511.

35 Lewis CM, Ng MY, Butler AW, Cohen-Woods S, Uher R, Pirlo K, et al. Genome-wide association study of major recurrent depression in the UK population. *Am J Psychiatry* 2010; **167**: 949–57.

36 Ising M, Lucae S, Binder EB, Bettecken T, Uhr M, Ripke S, et al. A genome-wide association study points to multiple loci that predict antidepressant drug treatment outcome in depression. *Arch Gen Psychiatry* 2009; **66**: 966–75.

37 Nivard MG, Mbarek H, Hottenga JJ, Smit JH, Jansen R, Penninx BW, et al. Further confirmation of the association between anxiety and CTNND2: replication in humans. *Genes Brain Behav* 2014; **13**: 195–201.

38 Purcell S, Neale B, Todd-Brown K, Thomas L, Ferreira MA, Bender D, et al. PLINK: a tool set for whole-genome association and population-based linkage analyses. *Am J Hum Genet* 2007; **81**: 559–75.

39 Han B, Eskin E. Random-effects model aimed at discovering associations in meta-analysis of genome-wide association studies. *Am J Hum Gen* 2011; **88**: 586–98.

40 Penninx BW, Milaneschi Y, Lamers F, Vogelzangs N. Understanding the somatic consequences of depression: biological mechanisms and the role of depression symptom profile. *BMC Med* 2013; **11**: 129.

41 Lamers F, Vogelzangs N, Merikangas KR, de Jonge P, Beekman AT, Penninx BW. Evidence for a differential role of HPA-axis function, inflammation and metabolic syndrome in melancholic versus atypical depression. *Mol Psychiatry* 2013; **18**: 692–9.

42 Green MA, Scott NA, Cross SE, Liao KY, Hallengren JJ, Davids CM, et al. Eating disorder behaviors and depression: a minimal relationship beyond social comparison, self-esteem, and body dissatisfaction. *J Clin Psychol* 2009; **65**: 989–99.

43 Friedman KE, Reichmann SK, Costanzo PR, Zelli A, Ashmore JA, Musante GJ. Weight stigmatization and ideological beliefs: relation to psychological functioning in obese adults. *Obes Res* 2005; **13**: 907–16.