

🗓 CORE



IRT Assessment of Readiness for Interprofessional Learning Scale (RIPLS): Dimensionality, Reliability, and Item Function

Matthew J Kerry-Krause, Ph.D.

Background and Context

+ Translated into dozens of languages, available in public domain, and informs health education and care delivery policies

+ More recent, IPE researchers are singing a different RIPLS tune



Table 1

Journal of Interprofessional Care

ISSN: 1356-1820 (Print) 1469-9567 (Online) Journal homepage: <u>http://www.tandfonline.c</u>

The Readiness for Interprofessional Learning Scale (RIPLS): A problematic evaluative scale for the interprofessional field

Cornelia Mahler, Sarah Berger & Scott Reeves

And still others rightfully observe that measurement invalidity for RIPLS is only the 'tip of the iceberg'- "Unfortunately, these and other problems cited about the RIPLS... can probably be said of other measurement tools as well." (Schmitz & Brandt, 2015)





Research Question and Proposition

- •<u>Question</u> \rightarrow Can Mahler et al.'s (2015) "three key problems with the RIPLS", as follows: 1) psychometric invalidity as indexed by Cronbach's α , 2) unstable latent-factor structure, and 3) sensitivity or appropriateness for assessing attitude-change be successfully applied to empirical RIPLS data?
- <u>Proposition</u> → Evidence from modern psychometric methods will bear information as to RIPLS usefulness in IPE.

Summary Sample

- RIPLS was administered to N = 287 pre-licensure MD & BSN students participating a mixed-methods IPE session

Results

1) Cronbach's α as invalidity index

- RIPLS critics' follow-up study Abstract "Cronbach's alpha was used to examine internal consistency" . . .
 - Low subscale- α is evidence of poor psychometrics quality.
 - Spearman-Brown formula predicts lower α for short scales.
 - Current data verifies with $\alpha = .90$ (9-item), .84 (4-item), & .20 (3-item).

Comparativ	omparative Global Data-Fit Indices of RIPLS.									
Model		<u>-2InL (df)</u>		<u>AIC</u>		BIC		<u>RMSEA</u>		
1-Dim		8114.43 (56)		8274.43		8558.36		.02		
2-Dim _c		8098.10 (55)		8260.10		8547.57		.02		
3-Dim _c		8090.55 (53)		8256.55		8551.12		.02		
Bifact ₁		8044.15 (40)		8236.15		8576.86		.00		
Bifact ₂		8039.74 (40)		8231.74		8572.45		.02		

3) <u>Sensitivity for Change-Assessment</u>

- This is a content- and sample-determined criterion.
- Mean-level changes were indicated in our dataset
- Further inspection of item-difficulty parameters indicated that item #s 2, 6, &
 - 11 may be best candidates for pre-licensure IPE assessments

There is indeed something odd about the common practice of using factor analysis to establish the dimensionality of a scale but then ignoring the parameter estimates themselves when creating scale scores. Statements about the adequacy of a model from a factor analytic standpoint may not apply when the parameters from that model are ignored" (Edwards & Wirth, 2009; *p*. 84-85).

- Furthermore, estimated common variance (ECV) from a bifactor-model of RIPLS = 97% (total score attributable to general-latent factor)

2) <u>Unstable Latent-Factor Structure</u>

- This is related to issue #1.
 - Primarily, this is due to over-factorization (analogous to adding items)
 - This is empirically demonstrated with RIPLS (Table 1)
- The RIPLS has shortcomings, but every IPE self-report questionnaire is 'technically' insufficient from measurement-construction standards (see, Oates & Davidson, 2015)

Conclusion

- There are well-discriminating items and general support for RIPLS' unidimensional interpretations
- The only reason for SUBscores is meaningful SUBscale interpretation (as we showed, empirical justification of subscores based solely on global-fit is insufficient). Simply, broad constructs require 'broad', rather than 'narrowed' (over-factorized) measurement models.



: : : SANTIAGO DE CHILE