# Independent Range Sampling, Revisited Again 

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#### Abstract

We revisit the range sampling problem: the input is a set of points where each point is associated with a real-valued weight. The goal is to store them in a structure such that given a query range and an integer $k$, we can extract $k$ independent random samples from the points inside the query range, where the probability of sampling a point is proportional to its weight.

This line of work was initiated in 2014 by Hu, Qiao, and Tao and it was later followed up by Afshani and Wei. The first line of work mostly studied unweighted but dynamic version of the problem in one dimension whereas the second result considered the static weighted problem in one dimension as well as the unweighted problem in 3D for halfspace queries.

We offer three main results and some interesting insights that were missed by the previous work: We show that it is possible to build efficient data structures for range sampling queries if we allow the query time to hold in expectation (the first result), or obtain efficient worst-case query bounds by allowing the sampling probability to be approximately proportional to the weight (the second result). The third result is a conditional lower bound that shows essentially one of the previous two concessions is needed. For instance, for the 3D range sampling queries, the first two results give efficient data structures with near-linear space and polylogarithmic query time whereas the lower bound shows with near-linear space the worst-case query time must be close to $n^{2 / 3}$, ignoring polylogarithmic factors. Up to our knowledge, this is the first such major gap between the expected and worst-case query time of a range searching problem.


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## 1 Introduction

In range searching, the goal is store a set $P$ of points in a data structure such that given a query range, we can answer certain questions about the subset of points inside the query range. The difficulty of the range searching problem, thus depends primarily on the shape of the query as well as types of questions that the data structure is able to answer. These range searching questions have been studied extensively and we refer the reader to the survey by Agarwal and Erickson [3] for a deeper review of range searching problems.


Let us for a moment fix a particular query shape. For example, assume we are given a set $P \subset \mathbb{R}^{2}$ to preprocess and at the query time we will be given a query halfplane $h$. The simplest type of question is an emptiness query where we simply want to report to the user whether $h \cap P$ is empty or not. Within the classical literature of range searching, the most general (and thus the most difficult) variant of range searching is semigroup range searching where each point in $P$ is assigned a weight from a semigroup and at the query time the goal is to return the sum of the weights of the points of $h \cap P$. The restriction of the weights to be from a semigroup is to disallow subtraction. As a result, semigroup range searching data structures can answer a diverse set of questions. Other classical variants of range searching lie between the emptiness and the semigroup variant. In our example, the emptiness queries can be trivially solved with $O(n)$ space and $O(\log n)$ query time whereas the semigroup variant can only be solved with $O(\sqrt{n})$ query time using $O(n)$ space. Finally, the third important variant, range reporting, where the goal is to output all the points in $P \cap h$, is often close to the emptiness variant in terms of difficulty. E.g., halfplane queries can be answered in $O(\log n+k)$ time where $k$ is the size of the output, using $O(n)$ space.

Sampling queries. Let $P$ be a large set of points that we would like to preprocess for range searching queries. Consider a query range $h$. Classical range searching solutions can answer simple questions such as the list of points inside $h$, or the number of them. However, we immediately hit a barrier if we are interested in more complex questions, e.g., what if we want to know how a "typical" point in $h$ looks like? Or if we are curious about the distribution of the data in $h$. In general, doing more complex data analysis requires that we extract the list of all the points inside $h$ but this could be an expensive operation. For questions of this type, as well as many other similar questions, it is very useful to be able to extract a (relatively) small random sample from the subset of points inside $h$. In fact, range sampling queries were considered more than two decades ago within the context of database systems $[7,4,5,11,10]$. Indeed the entire field of sample complexity, which provides the basis for statistics and machine learning, argues how any statistical quantity can be understood from an iid sample of the data. Range sampling allows this literature to quickly be specified to data in a query range. However, many of these classical solutions fail to hold in the worst-case as they use R-trees or Quadtrees whose performance depends on the distribution of the coordinates of the input points (which could be pretty bad). Some don't even guarantee that the samples extracted in the future will be independent of the samples extracted in the past. For example, sometimes asking the same query twice would return the same samples.

The previous results. Given a set of $n$ weights, one can sample one weight with proportional probability using the well-known "alias method" by A. J. Walker [12]. This method uses linear space and can sample a weight in worst-case constant time.

The rigorous study of independent range sampling was initiated by Hu et al. [8]. They emphasized the requirement that any random sample extracted from the data structure should be independent of all the other samples. They studied the one-dimensional version and for an unweighted point set and presented a linear-sized data structure that could extract $k$ random samples in $O(\log n+k)$ time and it could be updated in $O(\log n)$ time as well. A few years later, Afshani and Wei [2] revisited the problem and solved the one-dimensional version of the problem for weighted points: they presented a linear-size data structure that could answer queries in $O(\operatorname{Pred}(n)+k)$ time where $\operatorname{Pred}(n)$ referred to the running time of a predecessor query (often $O(\log n)$ but sometimes it could be faster, e.g., if the input is indexed by an array then the predecessor query can be answered trivially in $O(1)$ time $)$. They also studied the 3D halfspace queries but for unweighted points. Their main result was an optimal data structure of linear-size that could answer queries in $O(\log n+k)$ time.

Our results. We provide general results for the independent range sampling problem on weighted input (wIRS), and design specific results for 3D halfspace queries. We show a strong link between the wIRS and the range max problem. Namely, we show that the range max problem is at least as hard as wIRS, and also we provide a general formulation to solve the wIRS problem using range max. For halfspace queries in 3D, our framework gives a structure that uses $O(n \log n)$ space and has $O\left(\log ^{2} n+k\right)$ query time. We improve the space complexity to $O(n)$ when the ratio of the weights is $n^{O(1)}$. This solution uses rejection sampling, so it only provides an expected query time bound. To compensate, we provide another solution that has worst-case query time, but allows the points to be sampled within a $(1 \pm \varepsilon)$ factor of their desired probability, and may ignore points that would be sampled with probability less than $\gamma / n$ for $\gamma<\varepsilon<1$. This structure requires $O(n \log n)$ space, and has a worst-case query time of $O\left(\log (n / \gamma)\left(\log n+1 / \varepsilon^{3}\right)+k\right)$.

Finally, we show a conditional lower bound when we enforce worst-case query time and exact sampling probabilities, in what we call the separated algebraic decision tree (SAD) model. This model allows any decision tree structure that compares random bits to algebraic functions of a set of input weights. In this model, we show wIRS is as hard as the range sum problem, which is conjectured to be hard. In particular, if the best known solution to the range sum problem for halfspaces in 3D is optimal, then the wIRS problem would require $\Omega\left(n^{2 / 3-o(1)}\right)$ query time if it uses near-linear space. This provides the first such separation between expected $O\left(\log ^{2} n+k\right)$ and worst-case $\Omega\left(n^{2 / 3-o(1)}\right)$ query time for a range searching problem that we are aware of.

## 2 A Randomized Data Structure

In this section, we show that if we allow for the query bound to hold in expectation, then the range sampling problem can be solved under some general circumstances. Intuitively, we show that we need two ingredients: one, a data structure for range maximum queries, and two, a data structure that can sample from a weighted set of points under the assumptions that the weights are within a constant factor of each other. Furthermore, with an easy observation, we can also show that the range sampling problem is at least as hard as range maximum problem. We consider the input as a general set system $(X, \mathcal{R})$. We assume the input is a set $X$ of $n$ data elements (e.g., points) and we consider queries to be elements of a set of ranges $\mathcal{R}$ where each $R \in \mathcal{R}$ is a subset of $X$. The set $\mathcal{R}$ is often given implicitly, and for example, if $X$ is a set of points in $\mathbb{R}^{3}, \mathcal{R}$ could be the set of all $h \cap X$ where $h$ is a halfspace in 3D. Note that our model of computation is the real RAM model.

- Definition 1 (The Range Maximum Problem). Let $X$ be a data set, s.t., each element $x \in X$ is assigned a real-value weight $w(x)$. The goal is to store $X$ in a structure, s.t., given a range $R \in \mathcal{R}$, we can find the element in $R$ with the maximum weight.

Given a weighted set $X$, for any subset $Y \subset X$, we denote by $w(Y)$ the sum $\sum_{x \in Y} w(x)$. First we observe that range sampling is at least as hard as the range maximum problem.

- Lemma 2. Assume we can solve range sampling queries on input $(X, \mathcal{R})$ and for any weight assignment $w: X \rightarrow \mathbb{R}$ using $S(|X|)$ space and $Q(|X|)$ query time where the query only returns one sample. Then, given the set $X$ and $a$ weight function $w^{\prime}: X \rightarrow \mathbb{R}$, we can store $X$ in a data structure of size $S(|X|)$ such that given a query $R$, we can find the data element in $R$ with the maximum weight in $Q(|X|)$ time, with high probability.

Proof. See the full paper for the proof [1].

Next, we show that weighted range sampling can be obtained from a combination of a range maximum data structure and a particular form of weighted range sampling data structure for almost uniform weights.

- Lemma 3. Let $(X, \mathcal{R})$ be an input to the range sampling problem. Assume, we have a structure for the range maximum queries that uses $O\left(S_{m}(|X|)\right)$ space and with query time of $O\left(Q_{m}(|X|)\right)$. Furthermore, assume for any subset $X^{\prime} \subset X$ we can build a structure $\mathcal{D}_{s}\left(X^{\prime}\right)$ that uses $O\left(S_{s}\left(\left|X^{\prime}\right|\right)\right)$ space and given a query $R \in \mathcal{R}$, it does the following: it can return $w(Y)$ for a subset $Y \subset X^{\prime}$ with the property that $R \cap X^{\prime} \subset Y$, and $|Y|=O\left(\left|R \cap X^{\prime}\right|\right)$ and furthermore, the structure can extract $k$ random samples from $Y$ in $O\left(Q_{s}(|X|)+k\right)$ time.

Then, we can answer range sampling queries using $O\left(S_{m}(|X|)+S_{s}(|X|)\right)$ space and with expected query time of $O\left(Q_{m}(|X|)+Q_{s}(|X|) \log |X|+k\right)$.

Proof. Let $n=|X|$. We store $X$ in a data structure for the range maximum queries. We partition $X$ into subsets $X_{i} \subset X$ in the following way. We place the element $x_{1}$ with the largest weight in $X_{1}$ and then we add to $X_{1}$ any element whose weight is at least $w\left(x_{1}\right) / 2$ and then recurse on the remaining elements. Observe that for all $x, x^{\prime} \in X_{i}$ we have $w(x) / w\left(x^{\prime}\right) \in(1 / 2,2]$. Thus, the weight function $w$ is almost uniform on each $X_{i}$. We store $X_{i}$ in a data structure $\mathcal{D}_{s}\left(X_{i}\right)$.

Next, we build a subset sum information over the total weight $W\left(X_{i}\right)$ of the (disjoint) union of all subsets $X_{i^{\prime}}$ with $i^{\prime} \geq i$, that is, $W\left(X_{i}\right)=\sum_{j \geq i} w\left(X_{j}\right)$. Consider a query $R \in \mathcal{R}$.

Step 1: Use the Range Maximum Structure. Issue a single range-max query on $X$ for the query range $R \in \mathcal{R}$. Let $x$ be the answer to the range-max query, assume $x \in X_{i}$.

Step 2: top-level alias structure. Having found $i$, we identify the smallest index $i^{\prime} \geq i$ such that the maximum weight $w\left(x^{\prime}\right)$ for $x^{\prime} \in X_{i^{\prime}}$ and the minimum weight $w(x)$ for $x \in X_{i}$ satisfy $w(x) / w\left(x^{\prime}\right)>n^{2}$. As the weights in the sets $X_{i}$ decreases geometrically, we have $i^{\prime}-i=O(\log n)$. Then, for each $X_{j}$ with $j \in\left[i, i^{\prime}\right)$, we use the data structure $\mathcal{D}_{s}\left(X_{j}\right)$ to identify the set $Y_{j}$ such that $Y_{j}$ contains the set $R \cap X_{j}$. This returns the values $w\left(Y_{j}\right)$, and the total running time of this step is $O\left(Q_{s}(n) \log n\right)$.

We build a top-level alias structure on $i^{\prime}-i+1$ values: all the $i^{\prime}-i$ values $w\left(Y_{j}\right)$, $i \leq j<i^{\prime}$, as well as the value $W\left(X_{i^{\prime}}\right)$. Let $T=W\left(X_{i^{\prime}}\right)+\sum_{j=i}^{i^{\prime}-1} w\left(Y_{j}\right)$. This can be done in $O(\log n)$ time as $i^{\prime}-i=O(\log n)$.

Step 3: Extracting samples. To generate $k$ random samples from $R$, we first sample a value using the top-level alias structure. This can result in two different cases:
Case 1. It returns the value $W\left(X_{i^{\prime}}\right)$. Let $X_{i^{\prime}+}=\cup_{j \geq i^{\prime}} X_{j}$. In this case, we sample an element from the set $R \cap X_{i^{\prime}+}$, by building an alias structure on $X_{i+}$ in $O(n)$ time. The probability of sampling an element $x_{j} \in X_{i^{\prime}}+$ is set exactly to $\frac{w\left(x_{j}\right)}{W\left(X_{i^{\prime}}\right)}$. Note that these probabilities of $x_{j} \in X_{i^{\prime}+} \cap R$ do not add up to one which means the sampling might fail and we might not return any element. If this happens, we go back to the top-level alias structure and try again. Notice that $R$ contains at least one element $x_{i}$ from $X_{i}$, and that for any $x \in X_{j}, j \geq i^{\prime}$ have $w(x) \leq w\left(x_{i}\right) / n^{2}$ which implies $W\left(X_{i^{\prime}}\right) \leq w\left(x_{i}\right) / n$. Thus, this case can happen with probability at most $1 / n$, meaning, even if we spend $O(n)$ time to answer the query, the expected query time is $O(1)$.
Case 2. It returns a value $w\left(Y_{j}\right)$, for $i \leq j<i^{\prime}$. We place $Y_{j}$ into a list for now. At some later point (to be described) we will extract a sample $z$ from $Y_{i}$. If $z$ happens to be inside $R$, then we return $z$, otherwise, the sampling fails and we go back to the top-level structure.

We iterate until $k$ queries have been pooled. Then, we issue them in $O(\log n)$ batches to data structure $\mathcal{D}_{s}\left(X_{j}\right), i \leq j<i^{\prime}$. Ignoring the failure events in case (2), processing a batch of $k$ queries will take $O\left(Q_{s}(n) \log n+k\right)$ time. Notice that each iteration of the above procedure will succeed with a constant probability: case (1) is very unlikely (happens with probability less than $1 / n)$ and for case (2) observe that we have $w\left(Y_{i}\right)=O\left(w\left(R \cap X_{i}\right)\right)$ and thus each query will succeed with constant probability. As a result, in expectation we only issue a constant number of batches of size $k$ to extract $k$ random samples.

It remains to show that we sample each element with the correct probability. The probability of reaching case (1) is equal to $\frac{W\left(X_{i^{\prime}}\right)}{T}$. Thus, the probability of sampling an element $x_{j} \in X_{i^{\prime}+}$ is equal to $\frac{W\left(X_{i^{\prime}}\right)}{T} \cdot \frac{w\left(x_{j}\right)}{W\left(X_{\left.i^{\prime}\right)}\right)}=\frac{w\left(x_{j}\right)}{T}$. The same holds in case (2) and the probability of sampling an element $x_{j} \in X_{j}, i \leq j<i^{\prime}$ is $\frac{w\left(x_{j}\right)}{T}$. Thus, conditioned on the event that the sampling succeeds, each element is sampled with the correct probability.

## 3 3D Halfspace Sampling

### 3.1 Preliminaries

In this section, we consider random sampling queries for 3D halfspaces. But we first need to review some preliminaries.

- Lemma 4. Let $P$ be a set of $n$ points in $\mathbb{R}^{3}$. Let $P=P_{1} \cup \cdots \cup P_{t}$ be a partition of $P$ into $t$ subsets. We can store $P$ in a data structure of size $O(n \log t)$ such that given a query halfspace $h$, we can find the smallest index $i$ such that $P_{i} \cap h \neq \emptyset$ in $O(\log n \log t)$ time.

Proof. See the full paper for the proof [1].

The following folklore result is a special case of the above lemma.

- Corollary 5. Let $P$ be a set of $n$ points in $\mathbb{R}^{3}$ where each point $p \in P$ is assigned a real-valued weight $w(p)$. We can store $P$ in a data structure of size $O(n \log n)$ such that given a query halfspace $h$, we can find the point with maximum weight in $O\left(\log ^{2} n\right)$ time.

We also need the following preliminaries. Given a set $H$ of $n$ hyperplanes in $\mathbb{R}^{3}$, the level of a point $p$ is the number hyperplanes that pass below $p$. The $(\leq k)$-level of $H$ (resp. $k$-level of $H$ ) is the closure of the subset of $\mathbb{R}^{3}$ containing points with level at most $k$ (resp. exactly $k$ ). An approximate $k$-level of $H$ is a surface composed of triangles (possibly infinite triangles) that lies above $k$-level of $H$ but below $(c k)$-level of $H$ for a fixed constant $c$. For a point $q \in \mathbb{R}^{3}$, we define the conflict list of $q$ with respect to $H$ as the subset of hyperplanes in $H$ that pass below $q$ and we denote this with $\Delta(H, q)$. Similarly, for a triangle $\tau$ with vertices $v_{1}, v_{2}$, and $v_{3}$, we define $\Delta(H, \tau)=\Delta\left(H, v_{1}\right) \cup \Delta\left(H, v_{2}\right) \cup \Delta\left(H, v_{3}\right)$. One of the main tools that we will use is the existence of small approximate levels. This follows from the existence of shallow cuttings together with some geometric observations.

- Lemma 6. For any set $H$ of $n$ hyperplanes in $\mathbb{R}^{3}$, and any parameter $1 \leq k \leq n / 2$, there exists an approximate $k$-level which is a convex surface consisting of $O(n / k)$ triangles.

Furthermore, we can construct a hierarchy of approximate $k_{i}$-levels $\mathcal{L}_{i}$, for $k_{i}=2^{i}$ and $i=0, \cdots, \log n$, together with the list $\Delta(H, \tau)$ for every triangle $\tau \in \mathcal{L}_{i}$ in $O(n \log n)$ time and $O(n)$ space. Given a query point $q$, we can find an index $i$ in $O(\log n)$ time such that there exists a triangle $\tau \in \mathcal{L}_{i}$ that lies above $q$ such that $|\Delta(H, q)|=O(|\Delta(H, \tau)|)$.

Proof. See the full paper for the proof [1].

We will also use the following results.

- Theorem 7 (The Partition Theorem, [9]). Given a set $P$ of $n$ points in 3D and an integer $0<r \leq n / 2$, there exists a partition of $P$ into $r$ subsets $P_{1}, \cdots, P_{r}$, each of size $\Theta(n / r)$, where each $P_{i}$ is enclosed by a tetrahedron $T_{i}$, s.t., any hyperplane crosses $O\left(r^{2 / 3}\right)$ tetrahedra.
- Lemma 8. Let $T$ be tree of size $n$ where each leaf $v$ stores a real-valued non-negative weight $w(v)$. We can build a data structure s.t., given an internal node $u \in T$ at the query time, we can independently sample a leaf with probability proportional to its weight in the subtree of $u$. The data structure uses $O(n)$ space and it can answer queries in $O(1)$ worst-case time.

Proof. See the full paper for the proof [1].

Afshani and Wei [2] consider unweighted sampling for 3D halfspace queries. We next use the following technical result of theirs.

- Lemma 9. Let $H$ be a set of n hyperplanes in 3D. Let $f(n)=(\log n)^{c \log \log n}$ where $c$ is a large enough constant. We can build a tree $T_{\text {global }}$ with $n$ leafs where each hyperplane is stored in one leaf such that the following holds: Given a point $q \in \mathbb{R}^{3}$ with level $k$ where $k \geq f(n)$, we can find $k^{\prime}=O\left(k / \log ^{2} n\right)$ internal nodes $u_{1}, \cdots, u_{k^{\prime}}$ in $T_{\text {global }}$ such that $\Delta(H, q)=T_{\text {global }}\left(u_{1}\right) \cup \cdots \cup T_{\text {global }}\left(u_{k^{\prime}}\right)$ where $T_{\text {global }}\left(u_{i}\right)$ is the set of hyperplanes stored in the subtree of $u_{i}$.

Unfortunately, the above lemma is not stated explicitly by Afshani and Wei, however, $T_{\text {global }}$ is the "Global Structure" that is described in [2] under the same notation.

### 3.2 A Solution with Expected Query Time

We now observe that we can use Lemma 3 to give a data structure for weighted halfspace range sampling queries in 3D. We first note that using Corollary 5 and by building a hierarchy of shallow cuttings, i.e., building approximate $k_{i}$-levels for $k_{i}=2^{i}, i=0, \cdots, \log n$, we can get a data structure with $O(n \log n)$ space that can answer queries in $O\left(\log ^{2} n+k\right)$ query time. Furthermore, as Lemma 2 shows, our problem is at least as hard as the halfspace range maximum problem which currently has no better solution than $O(n \log n)$ space and $O\left(\log ^{2} n\right)$ expected query time. Thus, it seems we cannot do better unless we can do better for range maximum queries, a problem that seems very difficult.

However, the reduction given by Lemma 2 is not completely satisfying since we need to create a set of weights that are exponentially distributed. As a result, it does not capture a more "natural" setting where the ratio between the largest and the smallest weight is bounded by a parameter $U$ that is polynomial in $n$. Our improved solution is the following which shows when $U=n^{O(1)}$ we can in fact reduce the space to linear.

- Theorem 10. Let $P$ be a set of $n$ weighted points in $\mathbb{R}^{3}$, where the smallest weight is 1 and the largest weight is $U$. We can store $P$ in a data structure of size $O\left(n \min \left\{\log n, \log \log _{n} U\right\}\right)$ such that given a query halfspace and an integer $k$, we can extract $k$ independent random samples from the points inside $h$ in $O\left(\log ^{2} n+k\right)$ time.

Proof summary. For the complete proof see the full paper [1]. The basic idea is to partition the input into sets where the ratio of weights is within $n^{2}$. Then, we use Lemmas 9 and 8 .

### 3.3 Worst-Case Time with Approximate Weights

All of the previous data structures sample items exactly proportional to their weight, but rely on rejection sampling. Hence, their query time is expected, and not worst case. With small probability these structures may repeatedly sample items which are not in $h$, and then need to reset and try again with no bound on the worst-case time. To achieve worst-case query time, we need some modifications. We allow for items to be sampled "almost" proportional to their weights, i.e., we introduce a notion of approximation. As we shall see in the next chapter, without some kind of approximation, our task is very likely impossible.

Problem definition. We consider an input set $X$ of $n$ points where $x_{i}$ has weight $w_{i}$ and we would like to store $X$ in a data structure. At the query time, we are given a halfspace $h$ and a value $k$ and we would like to extract $k$ random samples from $X \cap h$. Let $w(h)=\sum_{x_{i} \in X \cap h} w_{i}$, and set two parameters $0<\gamma<\varepsilon<1$. Ideally, we would like to sample each $x_{i}$ with probability $w_{i} / w(h)$. Instead we sample $x_{i}$ with probability $\rho_{i}$, if $w_{i} / w(h) \geq \gamma / n$ then

$$
\begin{equation*}
(1-\varepsilon) \frac{w_{i}}{w(h)} \leq \rho_{i} \leq(1+\varepsilon) \frac{w_{i}}{w(h)}, \tag{1}
\end{equation*}
$$

and if $w_{i} / w(h)<\gamma / n$ then we must have $\rho_{i} \leq(1+\varepsilon) \frac{w_{i}}{w(h)}$. That is, we sample all items within a $(1 \pm \varepsilon)$ factor of their desired probability, except for items with very small weight, which could be ignored and not sampled. The smaller items are such that the sum of their desired probabilities is at most $\gamma$.

Overview of modifications. We start in the same framework as Section 3.2 and Lemma 3, i.e., we partition $X$ into subsets $X_{\ell}$ by weight. Then, we need to make the following modifications: (1) We can now ignore small enough weights; (2) We can no longer use rejection sampling to probe into each set $X_{\ell}$, rather we need to collect a bounded number (a function $f(\varepsilon)$ ) of candidates from all $X_{\ell}$ which is guaranteed to contain some point in the query $h$. (3) We can also no longer use rejection sampling within each $X_{\ell}$ to get candidates, instead we build a stratified sample via the partition theorem.

Change (1) is trivial to implement. Remember that given query $h$, we first identify indices $i$ and $i^{\prime}$ such that $X_{i}$ contains the largest weight in $h$ and the weights in $X_{i^{\prime}}$ are a factor $n^{2}$ smaller. We now require weights in $X_{i^{\prime}}$ to be a factor $n / \gamma$ smaller, and let $i^{\prime}=i+t$ where $t=O(\log (n / \gamma))$. We can now ignore all the remaining sets: the sets $X_{j}$ with $j \geq i+t$ will have weights so small that even if there are $\Omega(n)$ points within, the sum of their weights will be at most $\gamma$ times the largest weight. Since $\gamma<\varepsilon$, this implicitly increases all of the other weights in each $X_{\ell}$ (for $\ell \in[i, i+t)$ ) by at most a factor $(1+\varepsilon)$.

We next describe how we can build a data structure on each $X_{\ell}$ to generate $f(1 / \varepsilon)$ candidate points. Once we have these $t \cdot f(1 / \varepsilon)$ points, we can build two alias structures on them (they will come with weights proportional to the probability they should be selected), and select points until we find one in $h$. As a first pass, to generate $k$ samples, we can repeat this $k$ times, or bring $k f(1 / \varepsilon)$ samples from each $X_{\ell}$. We will return to this subproblem to improve the dependence on $k$ and $\varepsilon$ by short-circuiting a low-probability event and reloading these points dynamically.

### 3.3.1 Generating Candidate Points

Here we will focus on sampling our set of candidates. We will do this for every set $X_{\ell}$, $i \leq \ell<i^{\prime}$. However, to simplify the presentation, we will assume that the input is a set $X$ of $n$ points such that the weights of the points in $X$ are within factor 2 of each other. We will
sample $f(1 / \varepsilon)$ candidate points from $X$ (representing a subset $X_{\ell}$ ) s.t., the set of candidates intersects with the query halfspace $h$. Each candidate will be sampled with a probability that is almost uniform, i.e., it fits within our framework captured by Eq. 1.

Let $H$ be the set of hyperplanes dual to points in $X$. We maintain a hierarchical shallowcutting of approximate levels (Lemma 6) on $H$. By Lemma 6, we get the following in the primal setting (on $X$ ), given a query halfspace $h$ : We can build $O(|X|)$ subsets of $X$ where the subsets have in total $O(|X| \log |X|)$ points. Given a query halfspace $h$, in $O(\log n)$ time, we can find a subset $X^{\prime}$ so that $X \cap h \subset X^{\prime}$ and $\left|X^{\prime}\right|=O(|X \cap h|)$. We now augment this structure with the following information on each such subset $X^{\prime}$, without increasing the space complexity. We maintain an $r$-partition $\left(Z_{1}, \Delta_{1}\right),\left(Z_{2}, \Delta_{2}\right), \ldots\left(Z_{r^{\prime}}, \Delta_{r^{\prime}}\right)$ on $X^{\prime}$. (That is, so $r^{\prime}=\Theta(r)$, each subset $Z_{j} \subset \Delta_{j}$ and has size bound $\left|X^{\prime}\right| / r \leq\left|Z_{j}\right| \leq 2\left|X^{\prime}\right| / r$, and the boundary of any halfspace $h$ intersects at most $O\left(r^{2 / 3}\right)$ cells $\Delta_{j}$.) For each $Z_{j}$ we maintain an alias structure so in $O(1)$ time we can generate a random $s_{j}$ from the points within. It is given a weight $W_{j}=\sum_{x \in Z_{j}} w(x)$. In $O(r)$ time we can generate a weighted set $S=\left\{s_{1}, s_{2}, \ldots, s_{r^{\prime}}\right\}$; this will be the candidate set.

The sum of all weights of points within $h$ is $w(h)=\sum_{x \in X \cap h} w(x)$, and so we would like to approximately sample each $x \in X^{\prime} \cap h$ with probability $w(x) / w(h)$.

- Lemma 11. Let $r=\Omega\left(1 / \varepsilon^{3}\right)$ and consider a candidate set $S$, and sample one point proportional to their weights. For a point $x \in X^{\prime} \cap h$, the probability $\rho_{x}$ that it is selected satisfies

$$
(1-\varepsilon) \frac{w(x)}{w(h)} \leq \rho_{x} \leq(1+\varepsilon) \frac{w(x)}{w(h)}
$$

Proof. Of the $r^{\prime}$ cells in $S$, classify them in sets as inside if $\Delta_{j} \in h$, as outside if $\Delta_{j} \cap h=\emptyset$, and as straddling otherwise. We can ignore the outside sets. There are $O(r)$ inside sets, and $O\left(r^{2 / 3}\right)$ straddling sets.

For point $x \in S_{j}$ from an inside set, it ideally should be selected with probability $\frac{w(x)}{W_{j}} \cdot \frac{W_{j}}{w(h)}=\frac{w(x)}{w(h)}$. Indeed it is the representative of $S_{j}$ with probability $\frac{w(x)}{W_{j}}$ and is give weight proportional to $W_{j}$ in the alias structure. We now examine two cases, that all representative points in the straddling sets are in $h$, and that none are; the probability $x$ is selected will be between the probability of these two cases, and the desired probability it is selected will also be between these two cases. Let $W_{i n}=\sum_{S_{j} \text { is inside }} W_{j}$ and $W_{s t r}=\sum_{S_{j} \text { is straddling }} W_{j}$. The probability $x$ is selected if it is the representative of $S_{j}$ is then in the range $\left[\frac{W_{j}}{W_{\text {in }}+W_{\text {str }}}, \frac{W_{j}}{W_{i n}}\right]$. The ratio of these probabilities is $\frac{W_{j}}{W_{i n}} \cdot \frac{W_{i n}+W_{s t r}}{W_{j}}=\frac{W_{i n}+W_{s t r}}{W_{i n}}=1+\frac{O\left(r^{2 / 3}\right)}{\Theta(r)}=1+O\left(r^{1 / 3}\right)$. Setting $r=\Omega\left(1 / \varepsilon^{3}\right)$ ensures that these probabilities are within a $(1+\varepsilon)$-factor of each other, and on all points from an inside set, are chosen with approximately the correct probability.

For a point $x$ in a straddling set $S_{j}$, it should be selected with probability $\frac{w(x)}{W_{j}}$ and is selected with probability between $L_{x}=\frac{w(x)}{W_{j}} \frac{W_{j}}{W_{\text {in }}+W_{\text {str }}}$ and $U_{x}=\frac{w(x)}{W_{j}} \frac{W_{j}}{W_{\text {in }}+W_{j}}$. As with points from an inside set, these are within a $(1+\varepsilon)$-factor of each other if $r=\Omega\left(1 / \varepsilon^{3}\right)$. And indeed since $W_{i n} \leq w(h) \leq W_{i n}+W_{\text {str }}$ then $L_{x} \leq \frac{w(x)}{w(h)} \leq U_{x}$, and the desired probability of sampling straddling point $x$ is in that range.

### 3.3.2 Constructing the $k$ Samples

To select $k$ random samples, the simplest way is to run this procedure $k$ times sequentially, generating $O\left(k / \varepsilon^{3}\right)$ candidate points; this is on top of $O(\log n)$ to identify the proper subset $X^{\prime}$ from the shallow cutting, applied to all $t=O(\log (n / \gamma))$ weight partitions $X_{\ell}$.

We can do better by first generating $O\left(1 / \varepsilon^{3}\right)$ candidate points per $X_{\ell}$, enough for a single random point in $h$. Now we place these candidates in two separate alias structures along with the candidate points from the other $t$ structures. There are $O\left(t / \varepsilon^{3}\right)$ candidate points placed in an inside alias structure, and $O\left(t / \varepsilon^{2}\right)$ points placed in a straddling alias structure. Now to generate one point, we flip a coin proportional to the total weights in the two structures. If we go to the inside structure, we always draw a point in $h$, we are done. If we go to the straddling structure, we may or may not select a point in $h$. If we do, we are done; if not we flip another coin to decide to go to one of the two structures, and repeat.

It is easy to see this samples points with the correct probability, but it does not yet have a worst case time. We could use a dynamic aliasing structure [6] on the straddling set, so we sample those without replacement, and update the coin weight. However, this adds a $O\left(\log \left(\frac{1}{\varepsilon} \log (n / \gamma)\right)\right)$ factor to each of $O\left(t / \varepsilon^{2}\right)$ steps which might be needed. A better solution is to only allow the coin to direct the sampler to the straddling sets at most once; if it goes there and fails, then it automatically redirects to the alias structure on the inside sets which must succeed. This distorts the sampling probabilities, but not by too much since in the rejection sampling scheme, the probability of going to the straddling set even once is $O(\varepsilon)$.

- Lemma 12. For any candidate point s let $\rho_{s}$ be the probability it should be sampled, and $\rho_{s}^{\prime}$ the probability it is sampled with the one-shot deterministic scheme. Then $\rho_{s} \leq \rho_{s}^{\prime} \leq(1+\varepsilon) \rho_{s}$ if $s$ is an inside point and $(1-\varepsilon) \rho_{s} \leq \rho_{s}^{\prime} \leq \rho_{s}$ if $s$ is from a straddling set.

Proof. The probability that the coin directs to the inside set is $\pi_{i n}=\frac{W_{\text {in }}}{W_{\text {in }}+W_{\text {str }}}=\frac{1}{1+W_{s t r} / W_{\text {in }}}=$ $\frac{1}{1+\Omega(\varepsilon)}=1-O(\varepsilon)$. Let $w_{\text {str }}$ be the probability of selecting a point inside $h$, given that the coin has directed to the straddling set; we only need that $W_{\text {str }} \in[0,1]$.

For a candidate point in the inside set $s$, the probability it is selected in the rejection sampling scheme is $\rho_{s}=\frac{w(s)}{W_{\text {in }}}(1-O(\varepsilon))$, and in the deterministic scheme is $\rho_{s}^{\prime}=\frac{w(s)}{W_{\text {in }}}\left(\pi_{i n}+\right.$ $\left.\left(1-\pi_{i n}\right)\left(1-w_{s t r}\right)\right)$ which is in the range $\left[\rho_{s}, \frac{w(s)}{W_{i n}}\right]$, and these have a ratio $1+O(\varepsilon)$.

Similarly, for a candidate point in the straddling set $s$, the probability it is selected in the rejection sampling scheme is

$$
\rho_{s}=\frac{w(s)}{W_{s t r}}\left(1-\pi_{i n}\right)\left(1+W_{s t r}\left(1-\pi_{i n}\right)(1+\ldots)\right)=\frac{w(s)}{W_{s t r}}\left(1-\pi_{i n}\right)(1+O(\varepsilon))
$$

and in the deterministic scheme is $\rho_{s}^{\prime}=\frac{w(s)}{W_{s t r}}\left(1-\pi_{i n}\right)$. Thus $\rho_{s}^{\prime}$ is in the range $\left[\rho_{s}(1-O(\varepsilon)), \rho_{s}\right]$. Adjusting the constant coefficients in $\varepsilon$ elsewhere in the algorithm completes the proof.

Now to generate the next independent point (which we need $k$ of), we do not need to re-query with $h$ or rebuild the alias structures. In particular, each candidate point can have a pointer back to the alias structure within its partition cell, so it can replace its representative candidate point. Moreover, since the points have weight proportional to their cell in the partition $W_{j}$, these weights do not change on a replacement. And more importantly, the points we never inspected to see if they belong to $h$ do not need to be replaced. This deterministic process only inspects at most 2 points, and these can be replaced in $O(1)$ time. Hence extending to $k$ samples, only increases the total runtime by an additive term $O(k)$.

Final bound. We now have the ingredients for our final bound. In general the argument follows that of Lemma 3 except for a few changes. First, we allow items with probability total less than $\gamma n$ to be ignored. This replaces a $\log n$ factor in query time to be replaced with $t=\log (n / \gamma)$ term. Second, we require $O(t \log n)$ time to identify the relevant subset $X^{\prime}$ in each of $t$ shallow-cutting structures. Then we select $O\left(1 / \varepsilon^{3}\right)$ candidate points from each of $t$ weight partitions, in $O\left(t / \varepsilon^{3}\right)$ time. Then pulling $k$ independent samples, and refilling the candidates takes $O(k)$ time. This results in the following final bound:

- Theorem 13. Let $X$ be a set of $n$ weighted points in $\mathbb{R}^{3}$, where the smallest weight is 1 and the largest weight is $U$. Choose $0<\gamma<\varepsilon \leq 1 / 2$. We can store $X$ in a data structure of size $O\left(n \min \left\{\log n, \log \log _{n} U\right\}\right)$ such that for any integer $k$, we can extract $k$ independent random samples from the points $(\varepsilon, \gamma)$-approximately proportional to their weights in worst case time $O\left(\log (n / \gamma)\left(\log n+1 / \varepsilon^{3}\right)+k\right)$

In particular, if $\varepsilon, \gamma=\Omega(1)$, then the worst case time is $O\left(\log ^{2} n+k\right)$ matching the expected time algorithm to sample by the exact weights.

## 4 Lower Bound for Worst-Case Time with Exact Weights

In this section, we focus on proving a (conditional) lower bound for a data structure that can extract one random sample in $Q(n)$ worst-case time using $S(n)$ space. Our main result is that under a reasonably restricted model, the data structure must essentially solve an equivalent range searching problem in the "group model". As a result, we get a conditional lower bound as this latter problem is conjectured to be difficult. As an example, our conditional lower bound suggests that halfspace range sampling queries would require that $S(n) Q^{3}(n)=\Omega\left(n^{3}\right)$, i.e., with near-linear space we can only expect to get close to $O\left(n^{2 / 3}\right)$ query time. This is in contrast to the case when we allow expected query time or approximate weights. As already shown, we can solve the same problem with $O(n \log n)$ space and $O\left(\log ^{2} n\right)$ query time which reveals a large polynomial gap between the expected and the worst-case variants of the problem. To our knowledge, this is the first time such a large gap appears in the range searching literature between the worst-case and expected query times.

The Model of Computation. We assume the input is a list $X$ of $n$ elements, $x_{1}, \cdots, x_{n}$, where each element $x_{i}$ is associated with a real-valued weight $w\left(x_{i}\right)$. We use the decision tree model of computation. We assume the data structure has three components: a preprocessing algorithm that builds the data structure, a data structure which is a set of $S(n)$ stored real values, and finally a query algorithm that given a query $q$ it returns an element sampled with the correct probability in $Q(n)$ worst-case time. We allow no approximation: the query algorithm should return an element $x$ with exactly $w(x) / w(X)$ probability.

The main bottleneck. The challenge in obtaining our lower bound was understanding where the main computational bottleneck lied and formulating a plan of attack to exploit it. This turned out to be in the query algorithm. As a result, we place no restrictions on the storage, or the preprocessing part of the algorithm, an idea that initially sounds very counter intuitive. After giving the algorithm the input $X$ together with the weight assignment $w: X \rightarrow \mathbb{R}$, the algorithm stores some $S(n)$ values in its storage. Afterwards, we move to the query phase and this is where we would like to put reasonable limits. We give the query algorithm the query $q$. We allow the query algorithm access to a set of $t$ real random numbers, $R=\left\{r_{1}, \cdots, r_{t}\right\}$, generated uniformly in $[0,1]$. We restrict the query algorithm to be a "binary decision tree" $T$ but in a specialized format: each node $v$ of $T$ involves a comparison between some random number $r_{v} \in R$ and a rational function $f_{v}=g_{v} / G_{v}$ where $g_{v}$ and $G_{v}$ are $n$-variate polynomials of the input weights $w_{1}, \cdots, w_{n}$. To be more precise, we assume the query algorithm can compute polynomials $g_{v}$ and $G_{v}$ (either using polynomials stored at the ancestors of $v$ or using the values stored by the data structure). The query algorithm at node $v$ computes the ratio $g_{v}$ and $G_{v}$ and compares it to the random value $r_{v} \in\left\{r_{1}, \cdots, r_{t}\right\}$. If $v$ is an internal node, then $v$ will have two children $u_{1}$ and $u_{2}$ and the algorithm will proceed to $u_{1}$ if $f_{v}\left(w\left(x_{1}\right), \cdots, w\left(x_{n}\right)\right)<r_{v}$ and to $u_{2}$ if otherwise. If $v$ is a
leaf node, then $v$ will return (a fixed) element $x_{v} \in X$. Note that there is no restriction on the rational function $f_{v}$; it could be of an arbitrary size or complexity. We call this the separated Algebraic decision tree (SAD) model since at each node, we have "separated" the random numbers from the rational function that involves the weights of the input elements; a more general model would be to allow a rational function of the weights $w_{i}$ and the random numbers $r_{j}$. If we insist the polynomials $g_{v}$ and $G_{v}$ be linear (i.e., degree one), then we call the model separated linear decision tree model (SLD).

- Theorem 14. Consider an algorithm for range sampling in the SAD model where the input is a list of $n$ of elements, $x_{1}, \cdots, x_{n}$ together with a weight assignment $w\left(x_{i}\right) \in \mathbb{R}$, for each $x_{i}$. Assume that the query algorithm has a worst-case bound, i.e., the maximum depth of its decision tree $T$ is bounded by a function $d(n)$. Then, for every query $q$, there exists a node $v \in T$ such that $G_{v}$ is divisible by the polynomial $\sum_{p \in q} w(p)$.

Proof. Consider the query decision tree $T$. By our assumptions, $T$ is a finite tree that involves some $N(n)$ nodes and it has the maximum depth of $d(n)$. Each node $v$ involves a comparison between some random number $r_{i}$ and a rational function $f_{v}\left(w\left(x_{1}\right), \cdots, w\left(x_{t}\right)\right)$. To reduce clutter, we will simply write the rational function as $f_{v}$. Let $\mathcal{P}$ be the set of all the rational functions stored at the nodes of $T$. Note that each unique rational function appears only once in $\mathcal{P}$. Consider the set $\Delta_{\mathcal{P}}=\left\{f_{1}-f_{2} \mid f_{1}, f_{2} \in \mathcal{P}\right\}$ which is the set of pairwise differences. As each unique rational function appears only once in $\mathcal{P}$, it follows that none of the rational functions in $\Delta_{\mathcal{P}}$ is identical to zero. This in particular implies that we can find real values $w_{1}, \cdots, w_{n}$ such that none of the rational functions in $\Delta_{\mathcal{P}}$ are zero on $w_{1}, \cdots, w_{n}$. Thus, as these functions are continuous at the points $\left(w_{1}, \cdots, w_{n}\right)$, it follows that we can find a real value $\varepsilon>0$ such that for every weight assignment $w\left(x_{i}\right) \in\left[w_{i}, w_{i}+\varepsilon\right]$, the rational functions in $\Delta_{\mathcal{P}}$ have the same (none zero) sign. In the rest of the proof, we only focus on the weight assignment functions $w$ with the property that $w\left(x_{i}\right) \in\left[w_{i}, w_{i}+\varepsilon\right]$.

Let $\mathcal{U}$ be the unit cube in $\mathbb{R}^{t} . \mathcal{U}$ denotes the total probability space that corresponds to the random variables $r_{1}, \cdots, r_{t}$. Every point in $\mathcal{U}$ corresponds to a possible value for the random variables $r_{1}, \cdots, r_{t}$. Now consider one rational function $f_{v}$ stored at a node $v$ of $T$, and assume $v$ involves a comparison between $r_{i}$ and $f_{v}$ and let $u_{1}$ and $u_{2}$ be the left and the right child of $v$. By our assumption, we follow the path to $u_{1}$ if $f_{v} \leq r_{i}$ but to $u_{2}$ if otherwise. Observe that this is equivalent to partitioning $\mathcal{U}$ into two regions by a hyperplane perpendicular to the $i$-th axis at point $f_{v}$. As a result, for every node $v \in T$, we can assign a rectangular subset of $\mathcal{U}$ that denotes its region and it includes all the points $\left(r_{1}, \cdots, r_{t}\right)$ of $\mathcal{U}$ such that we would reach node $v$ if our random variables were sampled to be $\left(r_{1}, \cdots, r_{t}\right)$.

The next observation is that we can assume the region of each node is defined by fixed rational functions. Consider the list of rational functions encountered on the path from $v$ to the root of $T$. Assume among this list, the rational functions $f_{1}, \cdots, f_{m}$ are involved in comparisons with $r_{i}$. Clearly, the lower boundary of the region of $v$ along the $i$-th dimension is $\min \left\{f_{1}, \cdots, f_{m}\right\}$ whereas its upper boundary is $\max \left\{f_{1}, \cdots, f_{m}\right\}$. Now, observe that since we have assumed that each $f_{i}-f_{j}$ has a fixed sign, it follows that these evaluate to a fixed rational function. Thus, let $f_{i, v}$ (resp. $F_{i, v}$ ) be the rational function that describes the lower (resp. upper) boundary of the region of $v$ along the $i$-th dimension. Let $f_{i, v}=a_{i, v} / b_{i, v}$ and $F_{i, v}=A_{i, v} / B_{i, v}$ where $a_{i, v}, b_{i, v}, A_{i, v}, B_{i, v}$ are some polynomials of $w\left(x_{1}\right), \cdots, w\left(x_{n}\right)$ stored in tree $T$ (e.g., $b_{i, v}$ is equal to some $G_{u}$ for some ancestor $u$ of $v$ and the same holds for $B_{i, v}$ ).

The Lebesgue measure of the region of $v, \operatorname{Vol}(v)$, is thus defined by the rational function

$$
\operatorname{Vol}(v)=\prod_{i=1}^{t}\left(F_{i, v}-f_{i, v}\right)=\prod_{i=1}^{t}\left(\frac{A_{i, v}}{B_{i, v}}-\frac{a_{i, v}}{b_{i, v}}\right)=\prod_{i=1}^{t} \frac{A_{i, v} b_{i, v}-a_{i, v} B_{i, v}}{B_{i, v} b_{i, v}}
$$

$\operatorname{Vol}(v)$ is the probability of reaching $v$. Consider a query $q$ that contains $k$ elements. W.l.o.g, let $x_{1}, \cdots, x_{k}$ be these $k$ elements. Let $v_{1}, \cdots, v_{\ell}$ be the set of all the leaf nodes that return $x_{1}$. For $x_{1}$ to have been sampled with correct probability we must have the following identity

$$
\sum_{i=1}^{\ell} \operatorname{Vol}\left(v_{i}\right)=\frac{w\left(x_{1}\right)}{w\left(x_{1}\right)+\cdots+w\left(x_{k}\right)}
$$

Observe that since the polynomial $w\left(x_{1}\right)+\cdots+w\left(x_{k}\right)$ is irreducible, it follows that at least one of the polynomials $b_{i}$ or $B_{i}$ for some $i$ has this polynomial as a factor.

Conditional Lower Bound. Intuitively, our above theorem suggests that if we can perform range sampling in finite worst-case time, then we should also be able to find the total weight of the points in the query range - since the total weight $\sum_{p \in q} w(p)$ is encoded in some rational function. This latter problem is conjectured to be difficult but obtaining provable good lower bounds remains elusive. This suggests that short of a breakthrough, we can only hope for a conditional lower bound. This is reinforced by this observation that an efficient data structure for range sum queries leads to an efficient data structure for range sampling.

- Observation 15. Assume for any set of $X$ of $n$ elements each associated with a real-valued weight, we can build a data structure that uses $S(n)$ space such that given a query range $q$, it can output the total weight of the elements in $q$ in $Q(n)$ time.

Then, we can extract $k$ random samples from $q$ by a data structure that uses $O(S(n) \log n)$ space and has the query time of $O(k Q(n) \log n)$.

Proof. Partition $X$ into two equal-sized sets $X_{1}$ and $X_{2}$ and build the data structure for range sum on each. Then recurse on $X_{1}$ and $X_{2}$. The total space complexity is $O(S(n) \log n)$. Given a query $q$, it suffices to show how to extract one sample. Using the range sum query $q$ on $X_{1}$ and $X_{2}$, we can know the exact value of $w_{1}=\sum_{x \in q \cap X_{1}} w(x)$ and $w_{2}=\sum_{x \in q \cap X_{2}} w(x)$. Thus, we can recurse into $X_{1}$ with probability $w_{1} /\left(w_{1}+w_{2}\right)$ and into $X_{2}$ with $w_{2} /\left(w_{1}+w_{2}\right)$. In $O(\log n)$ recursion steps we find a random sample with query time $O(Q(n) \log n)$.

This implies that in the SLD model, the range sampling problem in the worst-case is equivalent to the range sum problem, ignoring polylog factors. This has consequences for query ranges like halfspaces where the latter problem is conjectured to be hard.

- Conjecture 16. For every integer n, there exists a set $P$ of $\Theta(n)$ points in $\mathbb{R}^{d}$ with the following property: If for any function $w: P \rightarrow \mathbb{R}$ given as input, we can build a data structure of size $S(n)$ such that for any query halfspace $h$, it can return the value $\sum_{p \in P \cap h} w(p)$ in $Q(n)$ time, then, we must have $S(n) Q^{d}(n)=\Omega\left(n^{d-o(1)}\right)$.
- Corollary 17. Assume Conjecture 1 holds for $d=3$. Then, there exists an input set of $n$ points in $\mathbb{R}^{3}$ such that for any data structure that uses $S(n)$ space and solves the range sampling problem where the query algorithm is a decision tree $T$ with worst-case query time $Q(n)$, we must have $S(n) Q^{3}(n)=\Omega\left(n^{3-o(1)}\right)$. Thus if space $S(n)$ is near-linear, query time $Q(n)=\Omega\left(n^{2 / 3-o(1)}\right)$.


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