



Review Article

Natural Regression of Frailty Among Community-Dwelling Older Adults: A Systematic Review and Meta-Analysis

Richard Ofori-Asenso, MSc,¹ Ken Lee Chin, PhD,^{1,2} Mohsen Mazidi, PhD,^{3,4} Ella Zomer, PhD,¹ Jenni Ilomaki, PhD,^{1,5} Zanfina Ademi, PhD,¹ J Simon Bell, PhD,^{1,5,6} and Danny Liew, PhD^{1,*}

¹Department of Epidemiology and Preventive Medicine, Monash University, Melbourne, Victoria, Australia. ²Department of Medical Education, Melbourne Medical School, The University of Melbourne, Victoria, Australia. ³Key State Laboratory of Molecular Developmental Biology, Institute of Genetics and Developmental Biology, Chinese Academy of Sciences, Chaoyang, China. ⁴Institute of Genetics and Developmental Biology, International College, University of Chinese Academy of Science (IC-UCAS), Chaoyang, China.⁵Centre for Medicine Use and Safety, Faculty of Pharmacy and Pharmaceutical Sciences, Monash University, Melbourne, Victoria, Australia. ⁶Centre for Research Excellence in Frailty and Healthy Ageing, University of Adelaide, South Australia, Australia.

*Address Correspondence to: Danny Liew, PhD, Department of Epidemiology and Preventive Medicine, Monash University, 553 St Kilda Road, VIC 3004, Melbourne, Victoria, Australia. E-mail: danny.liew@monash.edu

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Abstract

Background and Objectives: Frailty is a dynamic process with potential transitions over time. However, there is limited understanding of the patterns of frailty improvement. We conducted a systematic review and meta-analysis to estimate the natural rate of frailty regression among community-dwelling older adults aged at least 60 years.

Research Design and Methods: Systematic searches for studies reporting frailty improvement were performed in 5 databases (Medline, Embase, CINAHL plus, Web of Science, and PsycINFO) from inception until January 2019.

Results: Twenty-five studies from 26 countries were included. Among a baseline population of more than 50,000 individuals, the pooled prevalence of pre-frailty and frailty was 50.5% (95% confidence interval [CI] 47.8–53.3) and 12.8% (95% CI 9.1–17.0), respectively. During a median follow-up of 3.0 (range 1–10.0) years, 23.3% of surviving pre-frail individuals regressed to a robust state and 35.2% of surviving frail individuals reversed to a pre-frail or robust state. The pooled remission rates among people with pre-frailty and frailty were 80.4 (95% CI 61.7–104.6) and 135.3 (95% CI 98.1–186.5) per 1,000 person-years, respectively. Frailty and pre-frailty improvement rates varied by sex, diagnostic criteria, study region, and follow-up duration. The remission rates were significantly reduced when accounting for progressions to death. The heterogeneity of included studies was high which reflected considerable differences in methodological approach.

Discussion and Implications: Although frailty is highly prevalent in older people, natural remission is possible and common. Improved understanding of the factors that confer increased likelihood of frailty regression may support the design of interventions to reduce the burden of frailty.

Keywords: Pre-frailty, Physical phenotype, Deficit accumulation, Health status

Frailty, a complex phenomenon distinct from disability and comorbidity, which results from a decline in reserve capacity, affecting multiple physiological systems, is a major issue among older adults (Clegg, Young, Iliffe, Rikkert, & Rockwood, 2013; Xue, 2011). At present, there are no universally accepted criteria nor clinical definitions for frailty

(Dent, Kowal, & Hoogendijk, 2016). Nonetheless, two main theoretical concepts underpin the assessment of frailty: phenotype of frailty and accumulation of deficits. The frailty phenotype as described by Fried and colleagues, is based on five-predefined physical components: weight loss, exhaustion, weakness, slowness, and low physical activity, with frailty diagnosed as the presence of three or more of these components (Fried et al., 2001). Rockwood and colleagues also conceptualized frailty as an accumulation of deficits (symptoms, signs, functional impairment, and laboratory abnormalities), and demonstrated that more deficits confer greater risk (Rockwood et al., 1999). Both the Fried criteria and the accumulation of deficits approach have been extensively validated and are widely used (Cesari, Gambassi, van Kan, & Vellas, 2014; Hubbard, O'Mahony, & Woodhouse, 2009). The frailty phenotype is commonly used because it allows for better clinical operationalization in a variety of health care practice settings, whereas the deficit accumulation approach is also used because it has a strong relationship with mortality and allows for different combinations of deficits that all predict mortality equivalently well (Dent et al., 2016; Malmstrom, Miller, & Morley, 2014).

Frailty is usually considered as a three-staged process: robust, pre-frail, and frail (Xue, 2011). In the case of the Fried five-phenotype criteria, individuals are characterized as robust, pre-frail, and frail if 0, 1–2, and at least 3 of the criteria are met, respectively (Fried et al., 2001). For the accumulation of deficits approach, variable cut-offs have been used to define frailty in the literature (Aguayo et al., 2017). A systematic review by Collard and colleagues, reported a prevalence of 9.9% (95% confidence interval [CI] 9.6-10.2) and 13.6% (95% CI 13.2-14.0) for frailty in community-dwelling adults aged 65 years and older in high-income countries based on the Fried's definition and the frailty index, respectively (Collard, Boter, Schoevers, & Oude Voshaar, 2012). Increasingly, frailty has been associated with adverse health outcomes such as falls (Cheng & Chang, 2017; Kojima, 2015), delirium (Persico et al., 2018), hospitalization (Chang, Lin, & Cheng, 2018; Kojima, 2016), institutionalization (Kojima, 2018), incident disability (Kojima, 2017; Vermeiren et al., 2016), and death (Chang & Lin, 2015; Kojima, Iliffe, & Walters, 2018), as well as being associated with higher health care costs (Bock et al., 2016). Despite the increased recognition of the importance of early identification and management of frailty (Dent et al., 2017; Turner, Clegg, British Geriatrics, Age, & Royal College of General, 2014), there is considerable debate about frailty screening particularly related to who should be screened and where and when (Ambagtsheer et al., 2019).

Frailty is considered a dynamic process, with potential fluctuations over time (Lang, Michel, & Zekry, 2009; Xue, 2011). Nonetheless, the epidemiological data on frailty remain dominated by a focus on its burden and progression with less attention to the patterns of frailty improvement. Recently, there has been increased interest in understanding the various transition patterns among people with frailty (Michel,

Cruz-Jentoft, & Cederholm, 2015; O'Caoimh et al., 2018). Kojima and colleagues, for example, reported a systematic review and meta-analysis describing the patterns of frailty transitions among adults aged 50 years and over based on data from 16 studies (Kojima, Taniguchi, Iliffe, Jivraj, & Walters, 2019). However, their analysis focused only on studies using the Frailty phenotype and reported the pooled proportion of people who regressed with no indication of how quickly remission occurs (i.e., remission rate). In addition, there has been limited comparison of the rates of frailty improvement across different regions or country income levels.

Greater insight of the extent of frailty improvement may benefit clinical decision making as well as informing the design and targeting of public health strategies (Bandeen-Roche & Espinoza, 2017). Thus, we conducted a systematic review to summarize the available epidemiological data on the rate of frailty improvement among communitydwelling older adults aged 60 years and older.

Methods

We undertook a systematic review in accordance with the PRISMA (Preferred Reporting Items for Systematic Reviews and Meta-Analyses) guidelines (Moher, Liberati, Tetzlaff, Altman, & Group, 2009) (Supplementary Table 1 for checklist). The review protocol was registered in PROSPERO (International Prospective Register of Systematic Reviews); CRD42019121303.

Search Strategy and Selection Criteria

We searched Medline, Embase, CINAHL Plus, Web of Science, and PsycINFO for studies reporting on frailty remission. The searches were first performed on December 10, 2018 and updated on a weekly basis until January 29, 2019 so as to enable the tracking of new publications. The keywords used in the search were "frailty *or* frailty syndrome or geriatric disorders" *AND* "elderly *or* older *or* senior" *AND* "transition *or* trajectory *or* improvement *or* remission". No language restrictions were imposed. Supplementary Table 2 presents the search strategies used for the various databases. Reference lists of the relevant articles were also hand searched for additional studies.

Cohort studies were considered potentially eligible if they examined frailty regression (transition from more to less severe frailty states) among community-dwelling older adults aged 60 years and older at baseline in the absence of interventions (i.e., natural regression of frailty). We selected the 60 years cut-off to be consistent with the United Nations' definition of older persons (Department of Economic and Social Affairs Population Division, 2013). Studies must have measured frailty using a validated method, such as the Fried phenotype criteria or the frailty index (Fried et al., 2001; Rockwood et al., 1999).

Studies were excluded if they did not assess or report data on frailty improvement or involved participants selected on the basis of an index disease. Studies involving participants across the age spectrum were excluded unless there were data reported specifically for people aged 60 years and older at baseline. Studies involving hospitalized or institutionalized people were excluded. Similarly, studies reporting only changes in mean frailty scores or those using statistical methods such as group-based trajectory modeling were excluded because these studies were designed primarily to identify groups of individuals following approximately the same frailty trajectories. Randomized controlled trials, reviews, conference abstracts, editorials, and commentaries were also excluded.

All potentially eligible studies identified were searched for duplicates using Endnote X7, followed by title, abstract, and full-text assessments. In the case of multiple studies reporting on the same cohort, the study with the more detailed information or sample size was selected. Title and abstract screening were performed independently by two reviewers (R. Ofori-Asenso and K. L. Chin) and the list of studies selected for full text assessment were cross-checked for consistency. All articles selected for full text evaluation were also screened by two reviewers (R. Ofori-Asenso and K. L. Chin) independently. A prespecified procedure involving adjudication by a third reviewer (D. Liew) was in place to address potential disagreements (related to study selection) between the reviewers. However, this mechanism was never used as no disagreements occurred.

Data Extraction

The data extraction was performed by R. Ofori-Asenso and cross-checked by K. L. Chin. For each included study, the following information was collected: author details, country, study cohort name if any, publication year, sample size, age range at baseline, frailty assessment method, duration of follow-up, and data on frailty improvement. If available, age- and sex-specific remission data were collected. To highlight the burden of frailty and pre-frailty, we also collected data on the baseline prevalence of frailty and pre-frailty where available. Study authors were contacted for additional data where necessary.

Methodological Quality Assessment

Two reviewers (R. Ofori-Asenso and K. L. Chin) evaluated the methodological quality of studies using the nine items of the Joanna Briggs Institute (JBI)'s critical assessment tool for prevalence and incidence studies (Munn, Moola, Lisy, Riitano, & Tufanaru, 2015). A study was considered ineligible if it met less than five of the criteria.

Statistical Analysis

For each study, we extracted or calculated the rate of frailty improvement (transition from severe to less severe state) per 1,000 person-years based on the event rates and the mean duration of follow-up (Jager, Zoccali, Kramar, & Dekker, 2007; Yousef et al., 2008). Some studies used a 100% survivor cohort (i.e., excluded people who died) over the duration of the study. As such, for studies that included those who died in the estimation of remission rate, we recalculated the transition rate by restricting the sample to the fully surviving cohort with frailty assessment data at follow-up. This approach was intended to improve comparability across studies so as to minimize the impact of survivorship bias (Delgado-Rodriguez & Llorca, 2004). To derive 95% CIs of event rates, the exact methods, based on the Poisson distribution was adopted. If studies reported zero number of remissions a correction of 0.5 was applied to the observed events and person-years to allow for the calculation of CIs (Sutton, Abrams, Jones, Sheldon, & Song, 2000).

A meta-analysis was performed using the log-transformed transition rates and corresponding 95% CIs. The metaanalysis used the random effects (DerSimonian and Laird) model because of the anticipated heterogeneity across studies. Cochran's Q test and the I^2 statistic were used to quantify the presence of statistical heterogeneity (Higgins, Thompson, Deeks, & Altman, 2003). I^2 values of 25%, 50%, and 75% were considered to be low, moderate, and high degrees of heterogeneity, respectively (Higgins et al., 2003).

To explore potential sources of heterogeneity, subgroup analyses were undertaken based on the following characteristics: sex, study region (North America vs other), frailty measurement approach (physical phenotype vs other), sample size (\leq 1,000 vs >1,000), duration of follow-up (\leq 2 years vs >2 years), and country income level (highincome vs other). A high-income country (HIC) was defined using the World Bank criteria as any country with a gross national income per capita of USD 12,056 or more in 2017 (The World Bank, 2019). Differences between subgroups were assessed via a chi-square test (Sedgwick, 2015). We also pooled results specifically for studies that accounted for progression to death to examine its impact on the frailty and pre-frailty remission rates.

Publication bias was assessed using visual inspection of funnel plots and quantified via Egger's test (Egger, Davey Smith, Schneider, & Minder, 1997).

For the baseline prevalence of frailty and pre-frailty, the meta-analysis was performed using the Freeman–Tukey double arcsine transformed proportions to stabilize the variance (Barendregt, Doi, Lee, Norman, & Vos, 2013). If studies did not report baseline prevalence or lacked data to allow for calculations, they were excluded from this analysis.

All analyses were conducted using Stata 15/IC (StataCorp LP, College Station, Texas, USA). p < .05 was considered statistically significant.

Results

Selection Processes

A PRISMA flowchart illustrating the study selection process is presented in Figure 1. The electronic database search yielded 16,945 citations, of which 4,782 duplicates were excluded. A total of 12,048 articles were also excluded after title and abstract review, leaving 115 studies for potential inclusion. Following full-text assessment, 22 studies were found to meet the eligibility criteria. Three additional studies were retrieved from reference screening, resulting in a total of 25 studies included for the final review. No study was excluded on the basis of methodological (quality) evaluation using the JBI critical assessment tool for prevalence and incidence studies (Munn et al., 2015).

Study Characteristics

The characteristics of the 25 studies included in the review are summarized in Table 1. The studies were conducted among a total of 31,336 older adults with frailty or prefrailty from 26 countries. In terms of the regional distribution of studies; six were from Asia, six from North America, three from South America, seven from Europe, two from Australia, and one cross-regional study. The study sample size varied from 50 to 8,913 (median = 463). There were considerable differences in the duration of follow-up, with the median being 3.0 (range 1–10.0) years.

Twenty studies measured frailty according to physical phenotype (using the Fried criteria or a modified version), three studies used the frailty index, one study used the Vulnerable Elders Survey-13 screening tool, and one study used both the Fried's criteria and frailty index. Among the included studies that used the deficit accumulation approach, the number of deficits used ranged from 20 to 44.



Figure 1. Flow chart of studies selection process.

Prevalence of Frailty and Pre-frailty Among Baseline Source Population

From a baseline source population of 52,312 older adults across 24 included studies, the pooled prevalence of frailty was 12.8% (95% CI 9.1–17.0; $I^2 = 99.4\%$, p < .001) (Supplementary Figure 1). The pooled prevalence of frailty was lower among studies that measured frailty according to physical phenotype (9.2%, 95% CI 7.8–10.6; $I^2 = 95.1$, p < .001) than studies using other methods (37.0%, 95% CI 27.3–47.3; $I^2 = 98.4$, p < .001) (p-value for difference <.001). The prevalence of frailty was also significantly lower among studies from HICs (12.1%, 95% CI 9.6–14.8; $I^2 = 98.0\%$, p < .001) compared to those from non-HICs (16.5%, 95% CI 6.1–30.7; $I^2 = 99.6$, p < .001) (p-value for difference <.001).

Data on baseline prevalence of pre-frailty were obtained from 21 studies. In these studies, involving 50,633 older adults at baseline, the pooled prevalence of pre-frailty was 50.5% (95% CI 47.8–53.3; $I^2 = 97.1\%$, p < .001) (Supplementary Figure 2). The pooled prevalence of pre-frailty was higher among studies that measured frailty according to physical phenotype (51.3%, 95% CI 47.1–53.5; $I^2 = 97.1\%$, p < .001) than studies using other criteria (46.0%, 95% CI 34.7–57.5, $I^2 = 98.5\%$, p < .001) (p-value for difference <.001). There was no significant difference in the pooled prevalence of prefrailty between HICs (50.2%, 95% CI 47.1–54.3; $I^2 = 95.3$, p < .001) and other countries (51.1%, 95% CI 45.1–57.1, $I^2 = 97.4\%$, p < .001) (p-value for difference =.4120).

Transition From Pre-frailty to Robust

A total of 22 studies reported data on transitions among 23,869 people with pre-frailty at baseline. Of the 20,281 people with pre-frailty at baseline who were alive at the end of a median follow-up duration of 3.0 years, the proportion that regressed from pre-frailty to a robust state was 23.3%. The pooled transition rate from pre-frailty to robust was 80.4 (95% CI 61.7–104.6; $I^2 = 98.6$, p < .001) per 1,000 person-years (Figure 2). Visual inspection of funnel plots did not reveal any patterns of publication bias (Supplementary Figure 3), and this was confirmed via Egger's regression test (p = .873).

Fourteen studies reported transitions to death among 19,749 pre-frail people. The proportion of participants who progressed from pre-frailty to death was 15.6% over a median follow-up of 3.5 years. When factoring in death, the pooled rate of regression from pre-frailty to robust was 51.9 (95% CI 37.3–72.1; $I^2 = 98.9$, p < .001) per 1,000 person-years (Supplementary Figure 4). Restricting the analysis to survivors in these 14 studies resulted in a pooled transition rate from pre-frailty to robust of 61.8 (95% CI 47.5–80.3; $I^2 = 98.2$, p < .001) per 1,000 person-years (Supplementary Figure 5).

Transition From Frailty to Pre-frailty or Robust

A total of 24 studies reported data on transitions among 7,467 people with frailty at baseline. Of the 4,180 people with frailty who were alive at the end of a median follow-up

			Age range	Sample si	ze $(n)^a$		Mean follow-up
Study reference	Country(s)	Name of study cohort	at baseline	Pre-frail	Frail	Frailty assessment method	duration (years)
Ahmad et al., 2018	Malaysia		≥60	1,072	178	FFI	1.0
Alencar, Dias, Figueiredo, &	Brazil		≥65	108	43	FFI	1.0
Dias, 2015							
Alves, Duarte, & Santos, 2018	Brazil	SABE Study	≥60	581	119	FFI	4.0
Bentur, Sternberg, & Shuldiner, 2016	Israel		≥65		447	VES-13	6.0
Borrat-Besson, Ryser, & Wernli, 2013	Sweden, Denmark, Germany, Netherlands, Belgium, France, Switzerland, Austria, Spain, Italy, Poland, Czech Republic	SHARE survey	≥60	4,109	818	FFI	4.3
Doi et al., 2018	Japan	Obu Study of Health Promotion for the Elderly	≥65	2,344	354	FFI	4.0
Espinoza, Jung, & Hazuda, 2012	United States	The San Antonio Longitudinal Study of Aging cohort	≥65	298	52	FFI	7.0
Gill, Gahbauer, Allore, & Han, 2006	United States	The Precipitating Events Project	≥70	369	183	FFI	1.5^{b}
Gomes et al., 2018	Colombia, Albania, Brazil and Canada	International Mobility in Aging Study	65–74	804	104	FFI	2.0
Gruenewald, Seeman, Karlamangla, & Sarkisian, 2009	United States	MacArthur Study of Successful Aging	70-79	363	I	FFI	3.0
Hyde et al., 2016	Australia		≥60	I	106	FI (20 deficits used: on a scale of 0 to 1, frailty defined as ≥0.2 deficits)	7.0
Lanziotti Azevedo da Silva et al., 2015	Brazil	FIBRA study	≥65	111	27	FFI	1.1
Lee, Auyeung, Leung, Kwok, & Woo, 2014	Hong Kong	Mr and Ms OS Study	≥65	1,734	271	FFI	2.0
Liu et al., 2018	China	The Chinese Longitudinal Healthy Longevity Survey	65-99	5,349	3,564	FI (44 deficits were used: on a scale of 0 to 1, robust, pre- frail and frail were defined as ≤0.1, 0.1–0.21 and >0.21 respectivelv)	3.0
Lorenzo-López et al., 2019	Spain	VERISAÚDE project	≥65	379	18	FFI	1.0

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Table

			Age range	Sample siz	$e(n)^a$		Mean follow-un
Study reference	Country(s)	Name of study cohort	at baseline	Pre-frail	Frail	Frailty assessment method	duration (years)
Ottenbacher et al., 2009	United States	Hispanic Established Populations for the Epidemiological Study of the Elderly	≥65	975	156	FFI	10.0
Pollack et al., 2017	United States	Osteoporotic Fractures in Men Study	≥65	2342	422	FFI	4.6
Potier et al., 2018	Belgium		≥70	44	9	FFI	1.33
Swiecicka et al., 2018	Italy, Belgium, Poland, Sweden, United Kingdom, Spain, Hungary, Estonia	European Male Ageing Study	≥60	256	15	FFI	4.3
Serra-Prat et al., 2017	Spain		≥75	174	22°	FFI	1.0
Stephan et al., 2017	Germany		≥65	522	165	FI (30 items used: on a	3.0
:			2			scale of 0 to 1 robust, pre-frail and frailty were defined as ≤0.08, 0.08 > to < 0.25 and ≥0.25, respectively)	
Thompson, Theou, Adams, Tucker, & Visvanathan, 2018	Australia	North West Adelaide Health Study	≥65	357 (FFI) 219 (FI)	106 (FFI) 302 (FI)	FFI and FI (30 items used: on a scale of 0 to 1, robust, pre-frail and frailty were defined as ≤ 0.08 , $0.08 > to < 0.25, \geq 0.25$, respectively)	4.5
Trevisan, et al., 2017	Italy	The Progetto Veneto Anziani cohort	≥65	1,441	223	FFI	4.4
Wang et al., 2019	Taiwan	Taichung Community Health Study for Elders	65–99	I	56	FFI	1.0
Xue, 2011	United States	Women Health and Aging Study II	70-79	137	12	FFI	1.5 ^b

Notes: FFI = Fried frailty index; FI = frailty index; VES-13 = Vulnerable Elders Survey-13. "Sample size excludes those lost to follow-up but includes participants who died were repoi

"Sample size excludes those lost to follow-up but includes participants who died were reported. bWe selected transitions from baseline to 18 months because it had the most comprehensive data. "Sample derived from transitions between year 1 and 2.

Study	Cases/Person Years	ES (95% CI)	% Weight
Ahmad (2018)	313/1072	> 292.0 (260.5, 326.2)	4.77
Alencar (2015)	12/104	115.4 (59.6, 201.6)	3.83
Alves (2018)	137/2324		4.72
Borrat-Besson (2013)	1129/16524.9	68.3 (64.4, 72.4)	4.80
Doi (2018)	701/8928	* 78.5 (72.8, 84.6)	4.79
Espinoza (2012)	37/1372	27.0 (19.0, 37.2)	4.46
Gill (2006)	44/526.5	83.6 (60.7, 112.2)	4.52
Gomes (2018)	265/1608.0	🛨 164.8 (145.6, 185.9)	4.76
Grueneweld (2009)	103/1089.0	94.6 (77.2, 114.7)	4.69
Lanziotti (2015)	31/122.1	→ 253.9 (172.5, 360.4)	4.40
Lee (2014)	434/3016	🛨 143.9 (130.7, 158.1)	4.78
Liu (2018)	618/11976	➡ 51.6 (47.6, 55.8)	4.79
Lorenzo lopez (2019)	33/379	87.1 (59.9, 122.3)	4.42
Ottenbacher (2009)	69/3270.0	21.1 (16.4, 26.7)	4.62
Pollack (2017)	352/9457.6	★ 37.2 (33.4, 41.3)	4.78
Potier (2018)	10/58.52	170.9 (81.9, 314.3)	3.67
Serra-Prat (2017)	20/158.0	126.6 (77.3, 195.5)	4.19
Stephan (2017)	32/1428.0	22.4 (15.3, 31.6)	4.41
Swiecicka (2018)	80/1100.8	72.7 (57.6, 90.4)	4.65
Thompson (2018)	83/1336.5	62.1 (49.5, 77.0)	4.66
Trevisan (2018)	177/4395.6	40.3 (34.6, 46.7)	4.74
Xue (2011)	48/198.0	> 242.4 (178.7, 321.4)	4.54
Overall (I-squared = 9	8.6%, p = 0.000)	80.4 (61.7, 104.6)	100.00
NOTE: Weights are fro	m random effects analysis		

Figure 2. Forest plot of the rate of regression (per 1,000 person-years) from pre-frailty to robust state.

of 3.0 years, the proportion of people that transitioned from frailty to pre-frailty or a robust state was 35.2%. Among the frail population, the pooled rate of remission was 135.3 (95% CI 98.1–186.5; l^2 =96.7, p < .001) per 1,000 person-years (Figure 3). Visual inspection of funnel plots did not reveal any patterns of publication bias (Supplementary Figure 6) and this was confirmed via Egger's regression test (p = .686).

Sixteen studies reported transitions to death among 6,744 frail people. The proportion of participants who progressed from frailty to death was 45.8% over a median follow-up of 4.2 years. When factoring in death, the pooled rate of regression from frailty to pre-frailty or robust was 57.2 (95% CI 37.0–88.4; $I^2 = 97.8$, p < .001) per 1,000 person-years (Supplementary Figure 7). Restricting the analyses to those who survived in these 15 studies resulted in a pooled transition rate from frailty to pre-frailty or pre-frailty or robust of 98.5 (95% CI 72.4–133.9; $I^2 = 95.2$, p < .001) per 1,000 person-years (Supplementary Figure 8).

Subgroup and Sensitivity Analyses

The pooled remission rates of frailty and pre-frailty were significantly higher across studies that measured frailty according to physical phenotype compared to those using other criteria (Table 2). Frailty and pre-frailty remission rates were significantly higher in women than men. The pooled regression rates of frailty and pre-frailty were also significantly higher in studies with shorter follow-up duration (≤ 2 years). Pooled rates of frailty and pre-frailty remission were lower among studies involving participants from North America compared to studies involving participants from rest of the world. There was also variability in the pooled rates of frailty remission as per country income levels and study sample size. Nonetheless, the subgroup analyses did not reveal any consistent patterns in the source of heterogeneity as evidenced by the change in the I^2 statistic.

Discussion

In this systematic review and meta-analysis, we found that nearly one in four people who were pre-frail and more than one in three people who were frail and who survived over a median follow-up of 3.0 years regressed to a less severe frailty state. The remission rates among pre-frail and frail individuals were estimated as 80 and 135 cases per 1,000

Study	Cases/Person Years		ES (95% CI)	% Weigh
Ahmad (2018)	111/178	>	623.60 (513.00, 750.97)	4.72
Alencar (2015)	11/35		314.29 (156.89, 562.34)	4.03
Alves(2018)	38/476	— •	79.83 (56.49, 109.58)	4.56
Bentur (2016)	6/1020		5.88 (2.16, 12.80)	3.50
Borrat-Besson (2013	283/2829.4	+	100.02 (88.71, 112.38)	4.77
Doi (2018)	141/1200	•	117.50 (98.91, 138.57)	4.73
Espinoza (2012)	7/119	•	58.82 (23.65, 121.20)	3.66
Gill (2006)	42/238.5	· •	176.10 (126.92, 238.04)	4.59
Gomes (2018)	54/208		259.62 (195.03, 338.74)	4.63
Hyde (2016)	9/350		26.79 (12.25, 50.85)	3.88
Lanziotti (2015)	15/29.7	\rightarrow	505.05 (282.67, 833.00)	4.22
Lee (2014)	121/376	-	321.81 (267.03, 384.52)	4.72
Liu (2018)	406/4395	•	92.38 (83.61, 101.82)	4.78
Lorenzo lopez (2019)	6/18	· · · · · · · · · · · · · · · · · · ·	333.33 (122.33, 725.53)	3.50
Ottenbacher (2009)	11/140		78.57 (39.22, 140.59)	4.03
Pollack (2017)	71/1200.6		59.14 (46.19, 74.59)	4.67
Potier (2018)	2/7.98		250.63 (30.35, 905.35)	2.05
Serra-Prat (2017)	9/18	\longrightarrow	500.00 (228.63, 949.16)	3.88
Stephan (2017)	14/345	•	40.58 (22.19, 68.09)	4.18
Swiecicka (2013)	12/73.1		164.16 (84.82, 286.75)	4.09
Thompson (2018)	9/279	•	32.26 (14.75, 61.24)	3.88
Trevisan (2017)	63/589.6		106.85 (82.11, 136.71)	4.66
Wang (2019)	20/56		357.14 (218.15, 551.58)	4.36
Kue (2011)	9/16.5	>	545.45 (249.42, 1035.44) 3.88
Overall (I-squared =	96.7%, p = 0.000)	\diamond	135.30 (98.15, 186.51)	100.00
NOTE: Weights are f	om random effects analysis			
	1	I I 50 200 50	0	

Figure 3. Forest plot of the rate of regression (per 1,000 person-years) from frailty to robust or pre-frail state.

person-years, respectively. These rates varied by gender, frailty assessment method, as well as by study characteristics such as the duration of follow-up.

Although not necessarily an inevitable consequence of aging, frailty is a problematic issue which is highly prevalent among older adults (Clegg et al., 2013; Xue, 2011). Across the studies included in this review, we found the prevalence of frailty and pre-frailty among the baseline source population to be 12.8% and 50.5%, respectively. These estimates are higher than that reported from a previous systematic review by Collard and colleagues (2012). Nonetheless, the review by Collard and colleagues pooled data mainly from HICs, whereas our analyses combined data from countries with different income levels. A recent meta-analysis observed a much higher prevalence of pre-frailty (49.3%) and frailty (17.5%) in low and middle income countries (Siriwardhana, Hardoon, Rait, Weerasinghe, & Walters, 2018). Indeed, among the studies included in this review, the pooled prevalence of frailty was significantly lower among studies from HICs than across studies from non-HICs.

The dynamic nature of frailty and the corresponding opportunity to transition between frailty states has been recognized (Lang et al., 2009; Michel et al., 2015; Xue, 2011). In particular, Kojima and colleagues, recently characterized various transition patterns among people aged at least 50 years with frailty based on the frailty phenotype (Kojima et al., 2019). They found that about 46% of frail individuals regressed to robust or pre-frail state over a median follow-up of 3.9 years. Our analysis showed that around 35.2% of frail older adults aged 60 years and older regressed if they survived a period of 3.0 years. The higher proportion of frail people experiencing remission in the analysis by Kojima and colleagues, may be partly due to their inclusion of younger people, because frailty remission has been found to decrease with increasing age (Trevisan et al., 2017). The current review has also estimated the rate of remission (per 1,000 person-years) that allows for better understanding of how quickly remission occurs (Jager et al., 2007). Furthermore, our study has characterized the frailty and pre-frailty remission rates

Table 2. Subgroup Analyses of the Rate of Frailty and Pre-frailty Regression

	No. of		Pooled rate (95% CI)		χ^2 , <i>p</i> -value for
	studies ^a	Sample size ^b	per 1,000 person-years	I^2 (<i>p</i> -value)	difference ^c
Pre-frailty to robust					
Method of assessment					
Physical phenotype	20	15,813	87.6 (65.9-116.4)	98.6 (<.001)	16.2 (<.001)
Other	3	4,662	41.8 (27.0-64.8)	90.3 (<.001)	
Sex					
Male	6	5,136	68.0 (43.5-106.3)	97.7(<.001)	4.2 (.040)
Female	5	4,037	94.1 (53.5-165.7)	98.3 (<.001)	
Study region					
North America	6	3,425	58.5 (31.1-110.1)	98.0 (<.001)	5.5 (.019)
Other	15	1,6052	86.8 (65.5-116.9)	98.6 (<.001)	
Mean follow-up period					
≤2 years	10	4663	157.2 (119.9-206.0)	93.8 (<.001)	57.3 (<.001)
>2 years	12	15,618	48.6 (39.8-59.4)	96.7 (<.001)	
Sample size					
≤1,000	16	5,578	78.2 (54.2-112.9)	97.2 (<.001)	0.45 (.4983)
>1,000	6	14,703	86.9 (53.7-140.6)	99.5 (<.001)	
Country income level					
HIC	16	13,617	67.4 (52.2-87.1)	97.8 (<.001)	15.1 (<.001)
Other	5	5,860	120.7 (48.9-297.8)	99.4 (<.001)	
Frailty to pre-frailty or robust					
Method of assessment					
Physical phenotype	20	2,382	160.2 (150.4–170.5)	96.3 (<.001)	26.4 (<.001)
Other	5	2,009	80.5 (73.3-88.3)	95.2 (<.001)	
Sex					
Male	6	591	134.9 (79.8-231.7)	98.1 (<.001)	6.7 (.01)
Female	5	686	181.1 (82.3-398.0)	96.9 (<.001)	
Study region					
North America	5	462	120.2 (57.3-252.1)	92.7 (<.001)	2.2 (.1399)
Other	18	3,614	144.2 (91.8-196.2)	97.2 (<.001)	
Mean follow-up period					
≤2 years	11	800	358.3 (265.2-484.1)	84.0 (<.001)	195.1 (<.001)
>2 years	13	3,380	69.4 (55.8-86.4)	87.3 (<.001)	
Sample size					
≤1,000	23	2,715	137.4 (95.9–196.9)	96.4 (<.001)	d
>1,000	1	1,465	92.4 (83.6-101.8)	_	
Country income level					
HIC	18	2,252	110.5 (78.1-156.5)	94.6 (<.001)	44.1 (<.001)
Other	5	1,824	233.7 (83.9–651.2)	98.8 (<.001)	

Notes: CI = confidence interval; HIC = high-income economy.

^aSome studies were excluded from subgroup or may fit under more than one subgroup and thus the total number of studies in a subgroup analysis may exceed the total number of included studies.

^bThe results presented in this table are per the fully surviving cohort in each study.

^c*p*-values derived via chi-square test.

^dStatistical significance not assessed.

across different diagnostic criteria and shown that the rates of remission are significantly higher when using the frailty phenotype than other methods. This observation further supports the growing calls for a harmonized approach to frailty assessment to improve the precision in frailty data. However, overall, the study results are in accord with those of prior studies (Kojima et al., 2019; O'Caoimh et al., 2018), demonstrating that natural remission of frailty is possible and that individuals classified as frail are not necessarily in an irreversible health condition.

Several factors such as low physical activity (Ahmad et al., 2018; Trevisan et al., 2017), polypharmacy (including the use of medications such as sedatives and anticholinergic agents) (Lorenzo-López et al., 2019; Thompson et al., 2018), obesity (Thompson et al., 2018), underweight (Trevisan et al., 2017), presence of comorbidities such as stroke, diabetes (Espinoza et al., 2012; Lee et al., 2014;

Trevisan et al., 2017), smoking, vision or hearing loss (Trevisan et al., 2017), and cognitive impairment (Trevisan et al., 2017), have been identified to reduce the likelihood of frailty regression. This highlights that efforts to address these comorbidities may help to reduce the burden of frailty.

Previous studies have shown higher burden of frailty among females than males across all age groups (Collard et al., 2012; Gordon et al., 2017). Nonetheless, our pooled data suggested that frailty and pre-frailty remission is more common in females than in males. The contributory factors to the higher rate of frailty and pre-frailty remission in females deserves further exploration.

Owing to the adverse health outcomes and costs associated with frailty (Bock et al., 2016; Cheng & Chang, 2017; Kojima, 2015), better understanding of the patterns and determinants of frailty is important. By characterizing the rate of frailty remission, our study contributes to improving the understanding of the natural trajectory of frailty among community dwelling older adults. Such information is likely to be useful to professionals involved in the delivery of health care to the geriatric population who may want to know about possible outcomes of frailty in the absence of interventions.

As frailty improvement has been shown to be possible, the timely implementation of appropriate interventions could facilitate remission among pre-frail and frail individuals and potentially avert adverse-related consequences. For example, several interventions incorporating physical exercise, health education, nutritional supplements, home visits, hormone supplements, and counselling have been evaluated for their potential to delay or reverse frailty (Apostolo et al., 2018; Chin, van Uffelen, Riphagen, & van Mechelen, 2008; de Labra, Guimaraes-Pinheiro, Maseda, Lorenzo, & Millan-Calenti, 2015; Giné-Garriga, Roque-Figuls, Coll-Planas, Sitja-Rabert, & Salva, 2014; Puts et al., 2017), and highlighted in contemporary guidelines (Dent et al., 2017).

Recently, Travers, Romero-Ortuno, Bailey, and Cooney (2019) also compared the relative effectiveness and ease of implementation of the various interventions to delay or reverse frailty specifically within primary care settings. They found that interventions with both strength training and protein supplementation ranked highest in terms of relative effectiveness and ease of implementation. Although relatively easy to implement, interventions with mild-intensity mixed exercises or singular exercises such as walking ranked in the mid-zone for relative effectiveness (Travers et al., 2019). Similarly, educational or health promotion activities typically placed in the mid-zone for both relative effectiveness as well as ease of implementation. On the other hand, interventions targeting behavioral change ranked low in relative effectiveness and the mid-zone for ease of implementation. Whereas comprehensive geriatric assessments and home visits ranked mid-low for both relative effectiveness as well as ease of implementation (Travers et al., 2019).

Our review has some strengths. We included a large number of studies from multiple countries, with no

evidence of publication bias. By not restricting our analyses to any specific frailty assessment method, the results provide a broader context for understanding the patterns of remission among people with frailty. In addition, for the main analysis, we recalculated the remission rate in studies that incorporated deaths to allow for comparison to studies that used fully surviving cohorts. This approach was important to reduce the impact of survivorship bias (Delgado-Rodriguez & Llorca, 2004). Indeed, a subsequent subgroup analyses showed that incorporation of transition to deaths had significant impact on the remission rates. This was particularly pronounced for frailty than pre-frailty given that transition to deaths are more frequent in frail than prefrail individuals (Chang & Lin, 2015; Kojima et al., 2018; Vermeiren et al., 2016).

In terms of the limitations of our study, there was substantial heterogeneity unexplained by subgroup analyses and largely reflecting methodological differences. In addition, we could not stratify results according to age groups because of limited number of studies reporting age-stratified data. Moreover, this review was focused on the proportion of people who experienced remission rather than the number of transitions. Because an individual may undergo multiple intermediary transitions (Jamsen et al., 2016; Kojima, Taniguchi, Iliffe, Jivraj, & Walters, 2019; O'Caoimh et al., 2018), our analytic framework did not fully capture the full extent of frailty transitions.

Conclusions

Although frailty is highly prevalent among older adults, it is a dynamic syndrome, and natural remission is not only possible, but common. Improved understanding of the factors that confer increased likelihood of remission may support the design of interventions to reduce the burden of frailty in an era of population aging.

Supplementary Material

Supplementary data are available at *The Gerontologist* online.

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Conflict of Interest

The authors declare no potential conflicts of interest.

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Ofori-Asenso, R; Chin, KL; Mazidi, M; Zomer, E; Ilomaki, J; Ademi, Z; Bell, JS; Liew, D

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