Anthropometric Correlates of Human Anger

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Michael E. Price, Department of Psychology, Brunel University James Dunn, Department of Psychology, Brunel University Sian Hopkins, Department of Psychology, Brunel University Jinsheng Kang, School of Engineering and Design, Brunel University

Correspondence concerning this article should be addressed to:

Michael E. Price

Department of Psychology

School of Social Sciences

Brunel University

Uxbridge, Middlesex

UB8 3PH

United Kingdom

E-mail: michael.price@brunel.ac.uk

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1 Abstract

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3 The recalibrational theory of human anger predicts positive correlations between aggressive 4 formidability and anger levels in males, and between physical attractiveness and anger levels 5 in females. We tested these predictions by using a three dimensional body scanner to collect 6 anthropometric data about male aggressive formidability (measures of upper body muscularity and leg-body ratio) and female bodily attractiveness (waist-hip ratio, body mass 7 index, overall body shape femininity, and several other measures). Predictions were partially 8 supported: in males, two of three anger measures correlated significantly positively with 9 several muscularity measures; in females, self-perceived attractiveness correlated 10 11 significantly positively with two anger measures. However, most of these significant results 12 were observed only after excluding from the sample 27 participants who were older than 13 undergraduate age, leaving a subsample of 40 males and 51 females. Evidence for 14 relationships between anthropometric attractiveness indicators and anger measures was weak, but there was some evidence for relationships between anthropometric attractiveness 15 16 indicators and self-perceived attractiveness measures. While our results support the 17 recalibrational theory's prediction that anger usage and formidability are positively correlated 18 in males, and suggest that this formidability can be assessed via anthropometric measures alone, they also suggest that this prediction may not apply to populations older than 19 undergraduate age. Further, our results suggest that while female anger levels relate 20 21 positively to self-perceived attractiveness, they are unrelated to most anthropometric 22 measures of bodily attractiveness. *Keywords*: recalibrational theory; anger; aggression; self-perceived physical attractiveness; 23

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26 **1. Introduction**

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According to the recalibrational theory of anger (Sell et al., 2009), the human psychological 28 29 program that generates anger evolved to incentivize others to recalibrate upwards their 30 valuation of the angry individual's welfare. Ancestral individuals used anger to convince 31 others to treat them better, and the more power they had to harm and/or benefit others, the 32 more convincing they would have been in this regard: "anger is more likely to be triggered when an actor is positioned to make the price of resisting recalibration high. This price is 33 34 higher when the actor's formidability (ability to inflict costs on the target) or the actor's 35 ability to confer/withhold benefits is greater" (Sell et al., 2009: 15074).

The recalibrational theory predicts that certain traits would have been particularly 36 37 important influences in ancestral environments on an individual's ability to impose costs or 38 confer benefits. One of these influences, aggressive formidability, would have enhanced the ability to inflict costs. Because selection for traits promoting success in physical conflict is 39 40 stronger in males than in females (Trivers, 1972), traits that enhance aggressive formidability (e.g., upper body muscle mass) should be more important as aspects of cost-imposition ability 41 42 in males than in females. Another important influence on one's ability to impose costs/confer 43 benefits is physical attractiveness, which, as an indicator of health and mate value in both 44 sexes, would have enhanced the ability to confer benefits as an ally or mate. Attractiveness 45 would have probably been a more significant aspect of benefit-conferral ability in females 46 than in males, because it is relatively more important in females as an index of both fertility 47 and mate value (Grammer et al., 2003). It should also be noted that some traits may enhance 48 abilities to both inflict costs and confer benefits. For example, although Sell et al. focus on how strength could be used to impose costs, strength could also be used to provide benefits 49 50 such as protection from crime (Snyder et al., 2011). Further, moderately high male upper

body muscularity is itself perceived as attractive by females (Frederick & Haselton, 2007),
which may help explain why men who are more muscular report having had more sex
partners (Frederick & Haselton, 2007; Lassek & Gaulin, 2009).

54 Support for the predictions of the recalibrational theory is provided by Sell et al. 55 (2009), who found significant positive correlations in males, but not in females, between 56 upper body strength and the likelihoods of getting angry, of getting into physical fights, of 57 believing in the utility of personal and political aggression, and of succeeding in conflict. (A related finding is that women perceive men with more muscular bodies as being more 58 "volatile" [Frederick & Haselton, 2007]). Sell et al. also found significant positive 59 correlations in females between self-perceived physical attractiveness and all of these anger 60 61 measures, with the exception of likelihood of getting into physical fights. The correlations between attractiveness and anger measures were weaker in males, and most of these zero-62 63 order correlations became marginal or non-significant after controlling for the effects of 64 strength.

Previous research suggests, then, that there are significant sex differences in the traits 65 66 associated with anger-related outcome variables: upper body strength is important in males 67 but unimportant in females, whereas self-assessed physical attractiveness is relatively more 68 important in females than in males. The main purpose of the current study was to test these predictions using predictor variables that were based on anthropometric data, as opposed to 69 the kinds of data (mainly strength tests and self-report) collected in prior research. Sell et al. 70 71 (2009) assessed male aggressive formidability via several measures of strength: weight-lifting 72 ability, self-perceived strength, strength as perceived by others, flexed bicep circumference, and strength scores as recorded by a hand dynamometer. Their only anthropometric measure, 73 flexed bicep circumference, was included as one item in a four-item composite variable, so 74

they did not investigate the extent to which male anger measures could be predicted based onanthropometric variables alone, which is an issue that we examined in depth.

77 Further, Sell et al. used a measure of attractiveness that was based on self-report data, rather 78 than on any anthropometric data that would be informative about bodily attractiveness. 79 Because they did not collect this kind of anthropometric data, they could not test the 80 hypotheses that females with more attractive bodies (as assessed anthropometrically) tend to 81 use anger more, and that self-perceived attractiveness accurately reflects anthropometrically-82 measured attractiveness. Both of these hypotheses are relevant to the recalibrational theory, because this theory assumes that attractive people (especially, females) have greater power to 83 confer benefits because other people think they are attractive. While Sell et al. are correct 84 85 that human ancestors needed to perceive their own attractiveness in order to assess the extent 86 to which they could leverage it to their own advantage, it is also true that their perceptions 87 needed to have been rooted in reality in order to be used adaptively (otherwise, they would 88 have miscalculated the actual extent of their benefit-conferral advantage, and would thus have used anger either less or more frequently than would have been optimal for them). 89 90 However, good evidence for a positive relationship between anthropometrically-measured 91 attractiveness and female anger levels has apparently not yet been produced, and some 92 evidence suggests that there is no relationship between self-perceived attractiveness and anthropometric attractiveness in females (Brewer et al., 2007). We tested for both of these 93 types of relationships. 94

We used two kinds of variables as indicators of male formidability. First, as
explained above, we predicted that measures of upper-body muscularity such as chest,
shoulder and bicep circumference would correlate positively with anger use in males.
Second, we expected leg-body ratio (LBR) to correlate negatively with anger use in males.
LBR appears to correlate inversely with male fighting ability in a variety of primate species,

perhaps because males with lower LBR are harder to knock down (Carrier, 2006), and/or
because as a sexually dimorphic trait (Brown et al., 2008), LBR is associated with other
testosterone-dependent traits that influence formidability.

103 We used a variety of variables as indicators of female attractiveness. First, we 104 predicted that LBR would relate positively to anger in females: LBR appears to be associated 105 positively with body-shape femininity and attractiveness in females (Brown et al., 2008; 106 Rilling et al., 2009; Swami et al., 2006), although some research suggests that a moderate 107 LBR is most attractive in females (Frederick et al., 2010; Swami et al., 2007). Second, a 108 lower female waist-hip ratio (WHR) is regarded as more attractive by males in a wide variety 109 of cultures (Singh, 1993, 2002), so we expected WHR to relate negatively to anger in 110 females. (However, some evidence suggests cross-cultural variation in WHR preferences; for 111 reviews see Sugiyama [2005] and Swami & Salem [2011]). Third, waist circumference was 112 found by Rilling et al. (2009) to be a particularly significant inverse correlate of female 113 attractiveness, so we predicted it would relate negatively to female anger. Fourth and fifth, 114 we predicted that body mass index (BMI) and volume height index (VHI) would relate 115 negatively to female anger: Tovee et al. (2002) emphasized the importance of BMI as a 116 predictor of female attractiveness, while Fan et al. (2004) found VHI to be a better predictor. 117 Sixth, we predicted that bust-underbust ratio (BUR) would relate positively to female 118 attractiveness, based on work by Brown et al. (2008) which identified this relationship. 119 Finally, because a more sex-typical (i.e., feminine) female body shape is regarded as more 120 attractive (Brown et al., 2008), we derived a measure of overall "body shape femininity" 121 using principal component analysis and predicted that it would associate positively with 122 female anger.

Our collection of anthropometric data was aided through use of a three dimensionalbody scanner, which uses white-light to generate a point cloud display of the body, which it

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can then use to extract hundreds of accurate measurements. According to the manufacturer,

126 the scanner's point accuracy is < 1 mm, and its circumferential accuracy is < 3 mm (TC²,

127 2010).

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129 **2. Method**

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131 *2.1. Participants and procedure*

One-hundred and eighteen participants (56 males, 62 females, mean age 21.95 ± 4.53 132 years), mostly undergraduates at an English University, participated in exchange for 133 134 participation pool credit and/or a copy of their 3D body scan (for reasons discussed below, 135 this sample size was ultimately reduced to 91). After completing the questionnaire portion of 136 the study, participants' height to top of head and weight were recorded by stadiometer and 137 digital scale. Using methods similar to those used by Brown et al. (2008), participants were then body-scanned with an NX12 scanner, manufactured by TC^2 (Cary, North Carolina, 138 139 USA). During the scan, participants wore scanner-appropriate clothing (tight-fitting briefs 140 and for females, a sports bra), and stood erect in a standardized pose, without flexing any 141 muscles, with arms straightened and held slightly away from the sides of the body. Two 142 high-quality scans were obtained from each participant, and the 23 trait measurements used in 143 this study were extracted from each scan. The two measurements of each trait were first used 144 to assess repeatabilities, and were then averaged to produce the single measurement used to 145 generate predictor variables. Complete scans were obtained for all participants, with the 146 exception of two very dark-skinned males whose scans were incomplete below the elbow (the 147 NX12 sometimes has difficulty scanning very dark skin). These participants were excluded 148 from analyses that required the missing data.

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151 With the exception of height and weight (described above), all other anthropometric 152 measurements were extracted, in centimeters, by the NX12 scanner. The following 153 measurements were used to assess upper body muscularity: horizontal shoulder 154 circumference, left and right vertical shoulder circumference (measured from the underarm to 155 the top of the shoulder), chest circumference, left and right bicep circumference, left and right 156 elbow circumference, left and right forearm circumference, and left and right wrist 157 circumference. In order to produce a general measure of upper body muscularity, we created 158 a composite measure out of the most important components of upper body muscle mass: we 159 summed the z-scores of horizontal shoulder circumference, chest circumference, and mean 160 bicep circumference, and called it "upper body size". In addition, so that we could compare the effects of upper and lower body muscularity, we extracted measurements of left and right 161 162 thigh circumference and left and right calf circumference. To measure waist circumference, 163 we took the minimum circumference between the lower ribs and top of pelvis, and to measure 164 hip circumference we took the widest circumference between crotch and waist. We 165 determined WHR by dividing waist circumference by hip circumference. We calculated BMI 166 by dividing weight in kilograms by the square of height (to top of head) in meters, and VHI by dividing body volume by the square of height (to top of head) in meters. To calculate 167 168 BUR, we divided bust circumference by underbust circumference. Finally, to determine LBR, we measured circumference at the base of the lower torso, passing through the small of the 169 170 lower back, and divided the frontal height of this circumference by height (to top of head). 171 Repeatabilities (intraclass correlation coefficients) for all trait measurements extracted by the NX12 were high, ranging from .893 to .999. 172

To create the overall body shape femininity variable, our technique was similar to that described in Brown et al. (2008). First, independent *t*-tests were conducted on a variety of

175	traits known to be sexually dimorphic, in order to see which traits would be associated with
176	the greatest between-sex variation. The four traits which generated the largest t values were
177	WHR ($t = 11.44$), horizontal shoulder circumference ($t = 10.97$), BUR ($t = -9.63$), and LBR
178	(t = -9.14). Next, a principal component analysis on these four traits produced one
179	component with an eigenvalue of 2.56 (the only eigenvalue greater than 1.00), which
180	accounted for 63.98% of the total variance. Variable loadings (WHR = .89, horizontal
181	shoulder circumference = $.81$, BUR = 74 , LBR = 76) indicated that this component
182	captured body shape sex typicality such that higher values indicated a more masculine shape,
183	and lower values indicated a more feminine shape. We reverse coded this component so that
184	we could label it "body shape femininity".
185	Finally, so that we could examine the relationship between bodily and self-perceived
186	attractiveness, we measured three kinds of self-perceived attractiveness via three different
187	items. With two of these items, male and female participants responded on a nine-point scale
188	ranging from "Very unattractive" to "Very attractive" to the following items: "Please tick the
189	box indicating how physically attractive you think you are, in general" ("S-P attract
190	[general]"), and "Please tick the box indicating how physically attractive you think your body
191	is" ("S-P attract [body]"). All participants provided responses to both of these items, with the
192	exception of three males who responded to neither. The third item, administered to female
193	participants only, was the measure of self-perceived attractiveness used by Sell et al. (2009):
194	"Please fill in the blank: 'I am more attractive than% of other women'" ("S-P attract
195	[percentile]").

2.3. Outcome variables

Our three anger-related outcome variables were composed of items used in theoriginal measures designed by Sell et al. (2009). Participants responded to all items on a 9-

200	point scale from "Strongly disagree" to "Strongly agree". First, our "proneness to anger"
201	measure, which assesses one's general likelihood to become angry in everyday life, was the
202	mean response to items 3, 6, 7, 12, 15, 18 and 21 of the original measure (Cronbach's α =
203	.75). Sample items are "People often irritate me" and "It is harder to get me angry than other
204	people" (reverse-coded). Second, our "history of fighting" measure, which indicates one's
205	past frequency of engagement in physical conflict, was the mean response to all five items of
206	the original measure (Cronbach's $\alpha = .83$). Sample items are "I have physically intimidated
207	someone who had it coming" and "I have physically defended myself against attack".
208	Finally, our "utility of political aggression" measure was the mean response to items 8, 9, 11,
209	12 and 14 of the original measure (Cronbach's $\alpha = .74$). Sample items are "To deter
210	violence, a country needs a strong military" and "When it comes to international conflicts,
211	violence never solves anything" (reverse-coded). One male participant failed to respond to
212	most of the history of fighting and utility of political aggression items and so could not be
213	included in analyses involving these variables.

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- 215 **3. Results**
- 216
- 217 *3.1. Exclusion of older participants*

The initial analysis indicated that while few of the kinds of effects reported by Sell et al. (2009) could be observed when analyzing the entire 118-participant sample, many of them could be observed among the undergraduate-aged participants only. The age structure of our sample differed from that of the samples used by Sell et al., in that we had a larger percentage of participants who were older than undergraduate age, which gave our sample a higher age mean and standard deviation. The age means and standard deviations of the samples used by Sell et al. were 21.13 ± 2.38 (males, sample one), 19.94 ± 1.97 (males, sample two) and 18.99 225 \pm 1.24 (females, sample two) (A. Sell, personal communication, May 25, 2010). Figures for 226 our original sample were 21.95 \pm 4.53, but by excluding all participants who were older than 227 the typical undergraduate age range of 18-23 years, we reduced these figures to 19.93 \pm 1.44, 228 which were more in line with the Sell at al. samples. Excluding the 27 participants who were 229 older than 23 years reduced our sample size to 91 (51 females, 40 males). The below 230 analysis was conducted on this undergraduate-aged subsample only.

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232 *3.2. Formidability–anger correlations*

Table 1 presents descriptive statistics for all study variables, separately by sex, as well 233 234 as the results of *t*-tests that reveal whether each trait displayed significant sexual dimorphism. 235 Table 2 presents the sex-specific correlations between each predictor variable and each 236 outcome variable. Results in Table 2 show that as predicted, measures of upper body 237 muscularity correlated significantly positively with proneness to anger and political 238 aggression in males. Four distinct upper body traits (not including the composite variable, 239 upper body size) correlated significantly positively with both of these anger variables in 240 males: horizontal shoulder circumference, mean vertical shoulder circumference, chest 241 circumference, and mean bicep circumference. Overall, the anger measure that correlated 242 most strongly positively with the upper body muscularity measures was proneness to anger. 243 Surprisingly, none of these upper body muscularity measures correlated significantly 244 positively with history of fighting, although most of these correlations were positive and several approached significance. As expected, no significant correlations occurred between 245 246 upper body muscularity measures and anger measures among females. Also as expected, 247 lower body muscularity measures were relatively weak predictors of anger measures in 248 males, compared to upper body muscularity measures: the only significant positive

correlation between a lower body muscularity trait and an anger measure was between meancalf circumference and political aggression in males.

251 The predicted negative correlations between LBR and anger measures in males were 252 not significant, although in the case of proneness to anger this correlation was marginally 253 significant (p = .065). Because LBR and upper body size were both highly correlated with 254 proneness to anger in males, but not with each other (the correlation between upper body size and LBR in males was r = -.14, p = .189), the independent effects of these two predictors on 255 256 male proneness to anger were investigated via a multiple regression model. Upper body size 257 and LBR were entered together into this model with male proneness to anger as the outcome variable. The overall model was significant ($R^2 = .22$, p = .005), but only upper body size (β 258 = .40, p = .005), and not LBR ($\beta = -.19$, p = .106), explained significant unique variance in 259 260 male proneness to anger.

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262 *3.3. Attractiveness–anger correlations*

263 In females, the only significant correlations between any of the seven anthropometric 264 attractiveness measures and any of the three anger measures were between LBR and 265 proneness to anger (in the expected positive direction) and between WHR and political 266 aggression (in the expected negative direction). The only other indexes of attractiveness that 267 did correlate significantly with any anger measure in females were two non-anthropometric 268 ones, S-P attract [body], which correlated significantly positively with political aggression, 269 and S-P attract [percentile], which correlated significantly positively with both proneness to 270 anger and political aggression.

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272 *3.4. Correlations between self-perceived and anthropometric attractiveness*

The three self-perceived attractiveness measures varied in the extent to which they related to the female anthropometric attractiveness measures. As Table 3 shows, while S-P attract [general] and S-P attract [percentile] correlated significantly in the expected direction with only one of the seven anthropometric attractiveness measures (body shape femininity), S-P attract [body] performed relatively well, correlating significantly in the expected direction with six of these seven measures.

279 The one anthropometric attractiveness measure which failed to correlate significantly 280 in the expected direction with any self-perceived attractiveness measure in females was LBR. 281 Although the one-tailed correlations between LBR and S-P attract [body] were significant 282 among both males and females, in neither case was this correlation in the expected direction 283 or significant as a two-tailed correlation. Excessive speculation about the reasons for these correlations is thus not warranted. However, we should note that in the female case, the 284 correlation is due to LBR relating positively to body fat measures (WC, WHR, BMI and 285 286 VHI) which themselves relate negatively to S-P attract [body].

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288 **4. Discussion**

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Two types of predictions of the recalibrational theory of anger (Sell et al., 2009) were 290 291 supported in the current study. First, we found that indicators of aggressive formidability were significantly positively related to proneness to anger and political aggression in males. 292 293 While Sell at al. used predictors based mainly on upper body strength to reveal these 294 relationships, we used only anthropometric predictors (especially measures of upper body muscularity). Second, we replicated the findings that among females, self-perceived physical 295 296 attractiveness is significantly positively correlated with proneness to anger and political 297 aggression. We did not find indicators of aggressive formidability to be significantly

298 positively correlated with history of fighting in males, but some of these correlations were299 close to significantly positive.

300 Other findings from our study suggest that the predictions of the recalibrational theory 301 should be accompanied by some important caveats. First, most of our significant results were 302 significant only among the younger participants in our sample. This was especially true with 303 the correlations between muscularity and anger usage in males. Among the 40 males that 304 were included in the younger subsample (i.e., the male participants analyzed in the above 305 results section), there were a total of 12 significantly positive correlations found between the 306 eight upper body muscularity traits and the three anger measures (Table 2), but when the full 307 56-male sample was analyzed, only two such correlations were found. The correlations 308 between self-perceived attractiveness and anger measures in females held up comparatively 309 well in the full sample: of the three positive correlations between the self-perceived 310 attractiveness and anger measures that were significant among the 51 females in the younger 311 subsample, two were also significant in the full 62-female sample (the exception was the 312 correlation between S-P attract [percentile] and proneness to anger). 313 Our lack of full-sample results does not contrast with any findings from Sell et al., 314 because our younger subsample was more closely matched, compared to our full sample, 315 with their samples. However, this lack does suggest that the predictions of the recalibrational

theory – especially, that of a positive relationship between aggressive formidability and anger

in males – might apply best to populations of undergraduate age. Male aggression tends to

peak around the undergraduate-age years and to decline thereafter (Daly & Wilson, 1988),

and in our full male sample (n = 56), age was significantly negatively correlated with political

aggression (r = -.30, p = .014), and nearly so with proneness to anger and history of fighting

321 (both r's = -.21, both p's = .06). It may be that as anger levels fade with age, so does the

322 relationship between muscularity and anger. Testosterone levels probably affect some of this

study's most important sexually dimorphic variables (e.g., muscularity, proneness to anger),
and as age advances further past the pubertal stages when most sexual differentiation occurs,
associations between these variables may weaken. Further, as males age beyond their
undergraduate years, their formidability may become more dependent on social power and
achievement (e.g., on having a high income or an influential role in their organization) as
opposed to physique and strength, which may also contribute to the weakening of the
muscularity-anger correlation.

330 The second caveat is that although the recalibrational theory predicts a positive 331 correlation between female attractiveness and anger usage, this prediction may not be 332 accurate if attractiveness is assessed in terms of standard anthropometric measures of bodily 333 attractiveness. Of the seven such anthropometric predictors that we looked at, leg-body ratio 334 and waist-hip ratio were the only two that correlated significantly with any of the three anger 335 measures in females, and each predictor did so with only one anger measure. The 336 recalibrational theory does predict that the effect of 'objective' (other-perceived) 337 attractiveness on anger usage will be mediated by self-perceived attractiveness, so perhaps 338 anthropometric attractiveness influences anger only indirectly. This argument receives some 339 support from the fact that some self-perceived attractiveness measures correlated significantly 340 positively both with some anger measures, and with some anthropometric attractiveness 341 measures, in females. On the other hand, the self-perceived attractiveness measure that best 342 predicted female anger, S-P attract [percentile] – which was also the self-perceived 343 attractiveness measure used in Sell et al. (2009) – correlated significantly with only one of the 344 seven anthropometric predicators of attractiveness (body shape femininity). Another selfperceived attractiveness measure, S-P attract [body], performed much better as a predictor of 345 346 anthropometric attractiveness by correlating significantly in the expected direction with six of 347 the seven anthropometric predicators; however, it performed worse than S-P attract

348 [percentile] as a predictor of female anger usage.

349 The lack of strong correlations between anthropometric attractiveness measures and 350 female anger measures, and the lack of strong evidence that the measures of self-perceived 351 attractiveness that best predict anthropometric attractiveness are the same ones that best 352 predict female anger usage, raise a question: if females modulate their anger usage based on 353 their self-perceived attractiveness, then what information are they using to assess their own 354 attractiveness? Perhaps their assessments are based on facial attractiveness, which is an 355 important element of overall physical attractiveness that the current study did not measure. 356 However, it is possible that anger is modulated based on self-perceived attractiveness 357 assessments that are largely inaccurate. Previous research suggests that people are poor to 358 moderately good assessors of their own attractiveness (Brewer et al., 2007; Marcus & Miller, 359 2003; Mulford et al., 1998). In a study that was relatively similar to the current one in terms 360 of variables examined, females' self-perceived attractiveness ratings did not correlate 361 significantly with anthropometric measures of their bodily attractiveness (WHR, BMI) or 362 with other people's ratings of their facial attractiveness (Brewer et al., 2007). The authors of 363 that study explained this inaccuracy by noting that their female participants, particularly the 364 most attractive ones, tended to overestimate their own attractiveness, perhaps as the result of 365 adaptive self-deception (Trivers, 1999). However it is not clear that female participants in the 366 current study were overestimating their own physical attractiveness; in fact, as our average 367 female participant stated that she was more attractive than only about 41% of other women 368 (i.e., the mean of S-P attract [percentile] was 41.39), it is plausible that they tended to err on the side of humility. Moreover, we found that when female participants were asked 369 370 specifically about their self-perceived bodily attractiveness (S-P attract [body]), their 371 perceptions became reasonably accurate (although as noted, S-P attract [body] was not the

self-perceived attractiveness measure that best predicted female anger usage). The lack of
relationships in females between anthropometric attractiveness measures and anger, and
between anthropometric attractiveness measures and the self-perceived attractiveness
measures that best predict anger, are puzzles meriting further investigation.

376 In addition to these caveats, some study limitations bear mentioning. First, we (like 377 Sell et al.) have been interpreting correlations between formidability and anger as evidence 378 that formidability causes anger. However, anger could also lead to formidability, if anger-379 prone men spent more time working out in the gym. Either of these causal relationships 380 would be consistent with the correlations we found, and if both existed, there could be a 381 recursive relationship between formidability and anger: as anger-prone people became 382 stronger, they would become even more anger-prone. A second limitation is that our sample 383 size was relatively small, due both to time constraints imposed by our use of the body 384 scanner, and to our reliance on an age-restricted subsample. With a larger sample, some 385 observed non-significant relationships (such as those between history of fighting and 386 muscularity in males) would likely have been significant.

387 In conclusion, our results provide support for the recalibrational theory's predictions 388 that aggressive formidability will correlate positively with male anger usage, and that self-389 perceived attractiveness will correlate positively with female anger usage, with the caveat that these correlations may be stronger in populations of undergraduate age. Further, our 390 391 results suggest that the relationship between aggressive formidability and anger in males can 392 be observed even when formidability is assessed only via anthropometric measurements. Our 393 results also suggest that while anger usage correlates positively with female self-perceived 394 attractiveness, it in general does not do so with relatively objective measures of female bodily 395 attractiveness. While females may be basing their anger usage on their perceptions of their

396	own attractiveness, their perceptions do not seem to be related in a straightforward way to
397	information about their attractiveness as assessed anthropometrically.
398	
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402	
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Table 1

Descriptive statistics of all study variables, with results of t-tests for sex differences

	Males				Females				Sex difference	
Variable	п	Mean	S.D.	ľ	ı	Mean	S.D.	t	р	
Horiz. shoulder circ. (cm)	40	111.98	4.83	5	1	101.41	5.31	9.81	< .00	
Vert. shoulder circ. (cm)	40	44.20	2.35	5	1	39.57	3.59	7.05	< .00	
Chest circumference (cm)	40	101.61	6.08	5	1	91.65	6.58	7.42	< .00	
Bicep circumference (cm)	40	30.44	2.63	5	1	27.75	3.36	4.15	<.00	
Elbow circumference (cm)	39	26.48	1.43	5	1	24.54	2.21	4.79	<.00	
Forearm circ. (cm)	39	27.06	1.59	5	1	24.68	2.01	6.06	<.00	
Wrist circumference (cm)	39	17.84	0.82	5	1	16.86	0.81	5.70	<.00	
Upper body size	40	1.63	2.05	5	1	-1.80	2.38	7.23	<.00	
Thigh circumference (cm)	40	54.84	4.10	5	1	57.10	6.65	-1.89	.06	
Calf circumference (cm)	40	36.50	2.35	5	1	35.45	3.35	1.69	.09	
Leg to body ratio	40	0.58	0.03	5	1	0.62	0.02	-7.65	<.00	
Waist circumference (cm)	40	78.88	7.12	5	1	71.93	8.18	4.25	<.00	
Waist to hip ratio	40	0.80	0.04	5	1	0.71	0.05	9.77	<.00	
Body mass index	40	22.86	2.59	5	1	22.29	3.60	0.85	.40	
Volume-height index	40	20.63	2.61	5	1	20.51	3.37	0.19	.852	
Bust to underbust ratio	40	1.11	0.03	5	1	1.19	0.07	-7.47	<.00	
Body shape femininity	40	-0.83	0.55	5	1	0.83	0.44	-16.12	<.00	
S-P attract [general]	38	6.08	0.91	5	1	5.92	1.02	0.75	.45	
S-P attract [body]	38	5.87	1.34	5	1	5.43	1.60	1.36	.17	
S-P attract [percentile]	_	_		5	1	41.39	15.90			
Proneness to anger	40	5.29	1.19	5	1	5.27	1.33	0.06	.95	
History of fighting	39	4.96	2.26	5	1	3.60	1.92	3.10	.00	
Political aggression	39	4.15	1.54	5	1	3.25	1.28	3.04	.00	

Note. Results are for the university-aged subsample only (ages 18-23). Values of p are two-tailed. For bilateral traits the mean of left-right measurements are presented.

Table 2

Correlations between predictors and anger measures

	Pronene	ess to anger	History	of fighting	Political aggression		
Trait	Male <i>r</i>	Female <i>r</i>	Male <i>r</i>	Female <i>r</i>	Male r	Female <i>i</i>	
Horizontal shoulder circ.	.41**	.01	.24	.09	.31*	18	
Vertical shoulder circ.	.35*	.03	.22	.08	.28*	.01	
Chest circumference	.42**	06	.18	.20	.32*	14	
Bicep circumference	.39**	.04	.21	.23	.34*	12	
Elbow circumference	.43**	.01	.06	.12	.15	08	
Forearm circumference	.37**	.02	.09	.16	.21	08	
Wrist circumference	.26	07	03	.11	.04	15	
Upper body size	.43**	<01	.22	.19	.34*	15	
Thigh circumference	.11	04	.16	.14	.23	15	
Calf circumference	.05	08	02	.14	.27*	15	
Leg to body ratio	24	.31*	13	.19	04	13	
Waist circumference	.20	.03	08	.05	.23	22	
Waist to hip ratio	.13	.18	21	.05	.13	28*	
Body mass index	.21	02	02	.13	.26	16	
Volume-height index	.31*	<01	.07	.14	.24	18	
Bust to underbust ratio	01	02	.27*	02	.05	01	
Body shape femininity	33*	.02	<01	01	18	.19	
S-P attract [general]	.22	.20	.05	09	.03	.11	
S-P attract [body]	.19	.08	.19	18	.10	.27*	
S-P attract [percentile]		.26*	_	12	_	.29*	

Note. Results are for the university-aged subsample only (ages 18-23, 40 males, 51 females). Values of p are one-tailed. For bilateral traits the mean of left-right measurements are presented. *p < .05, **p < .01.

Table 3

Measure	1	2	3	4	5	6	7	8	9
1. LBR		32*	26	41**	24	.18	.67***	.08	.30*
2. WC	.41**		.84***	.83***	.88***	34*	78***	06	16
3. WHR	.28*	.75***	—	.66***	.62***	50**	79***	.05	02
4. BMI	.34**	.89***	.55***	—	.84***	37*	80***	05	11
5. VHI	.37**	.90***	.56***	.98***		18	66***	15	18
6. BUR	.12	.22	.05	.36**	.28*		.56***	14	13
7. BSF	.15	51***	65***	31*	35**	.54***		08	.04
8. S-P [g]	.03	10	10	12	14	.21	.28*		.74***
9. S-P [b]	25*	34**	45***	27*	32*	.29*	.44**	.56***	
10. S-P [p]	.03	13	17	14	15	.12	.26*	.60***	.57***

Intercorrelations between attractiveness measures

Note. Results are for the university-aged subsample only (ages 18-23, 40 males, 51 females). Correlations for males are presented above the diagonal, and correlations for females are presented below it. Values of *p* are one-tailed. WC = waist circumference, BSF = body shape femininity, S-P [g] = self-perceived attractiveness [general], S-P [b] = self-perceived attractiveness [body], S-P [p] = self-perceived attractiveness [percentile]. *p < .05, **p < .01, ***p < .001.