

# Distribution-based fuzzy clustering of electrical resistivity tomography images for interface detection

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## SUMMARY

A novel method for the effective identification of bedrock subsurface elevation from electrical resistivity tomography images is described. Identifying subsurface boundaries in the topographic data can be difficult due to smoothness constraints used in inversion, so a statistical population-based approach is used that extends previous work in calculating iso-resistivity surfaces. The analysis framework involves a procedure for guiding a clustering approach based on the fuzzy *c*-means algorithm. An approximation of resistivity distributions, found using kernel density estimation, was utilized as a means of guiding the cluster centroids used to classify data. A fuzzy method was chosen over hard clustering due to uncertainty in hard edges in the topography data, and a measure of clustering uncertainty was identified based on the reciprocal of cluster membership. The algorithm was validated using a direct comparison of known observed bedrock depths at two 3-D survey sites, using real-time GPS information of exposed bedrock by quarrying on one site, and borehole logs at the other. Results show similarly accurate detection as a leading isosurface estimation method, and the proposed algorithm requires significantly less user input and prior site knowledge. Furthermore, the method is effectively dimension-independent and will scale to data of increased spatial dimensions without a significant effect on the runtime. A discussion on the results by automated versus supervised analysis is also presented.

**Key words:** Image processing; Neural networks, fuzzy logic; Tomography.

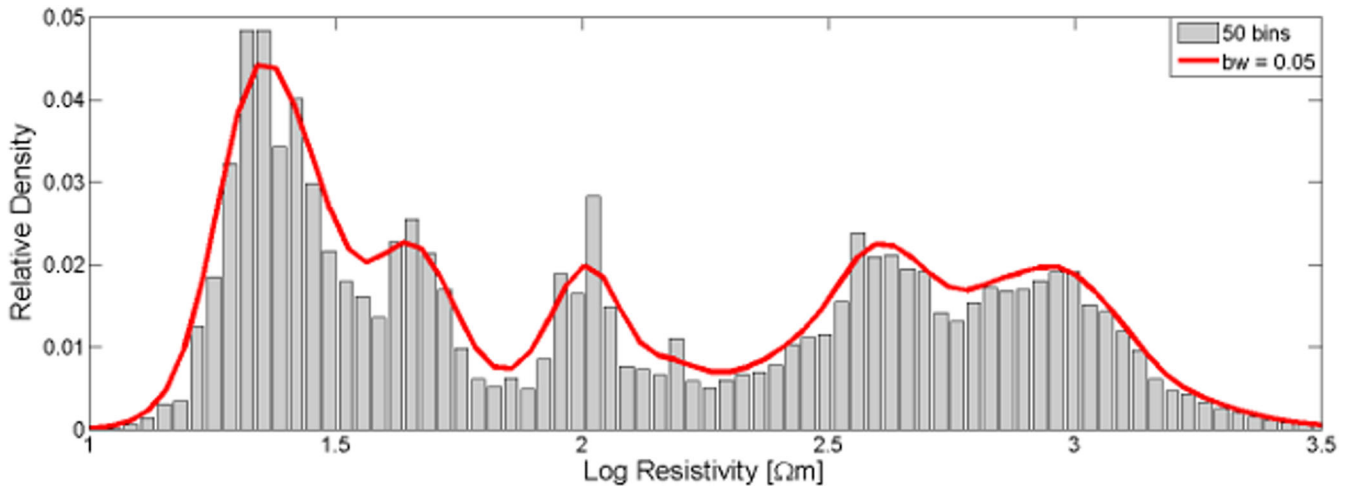
## 1 INTRODUCTION

Intrusive investigation, especially drilling, is the most significant and common method by which to analyse shallow soft-rock aggregate mineral resources in unconsolidated superficial geological deposits (e.g. sand and gravel). A complementary method, involving minimal intrusion, is electrical resistivity tomography (ERT), which has been demonstrated as a viable means of mineral deposit characterisation. Known benefits of ERT imaging over direct intrusion include the provision of spatial information and rapid non-invasive survey coverage. Although ERT imaging of the subsurface is not yet routinely used for soft-rock aggregate exploration, research has been undertaken in recent years to develop ERT for this application (e.g. Hirsch *et al.* 2008; Hickin *et al.* 2009; Hsu *et al.* 2010; Chambers *et al.* 2012, 2013; Loke *et al.* 2013). The purpose of such studies is to provide the evidence base needed to validate ERT for this application, and to establish a good practice framework covering survey design, data processing and interpretation, and the integrated use of ERT (e.g. Böhm *et al.* 2013) alongside conventional intrusive techniques (i.e. drilling and trial pitting).

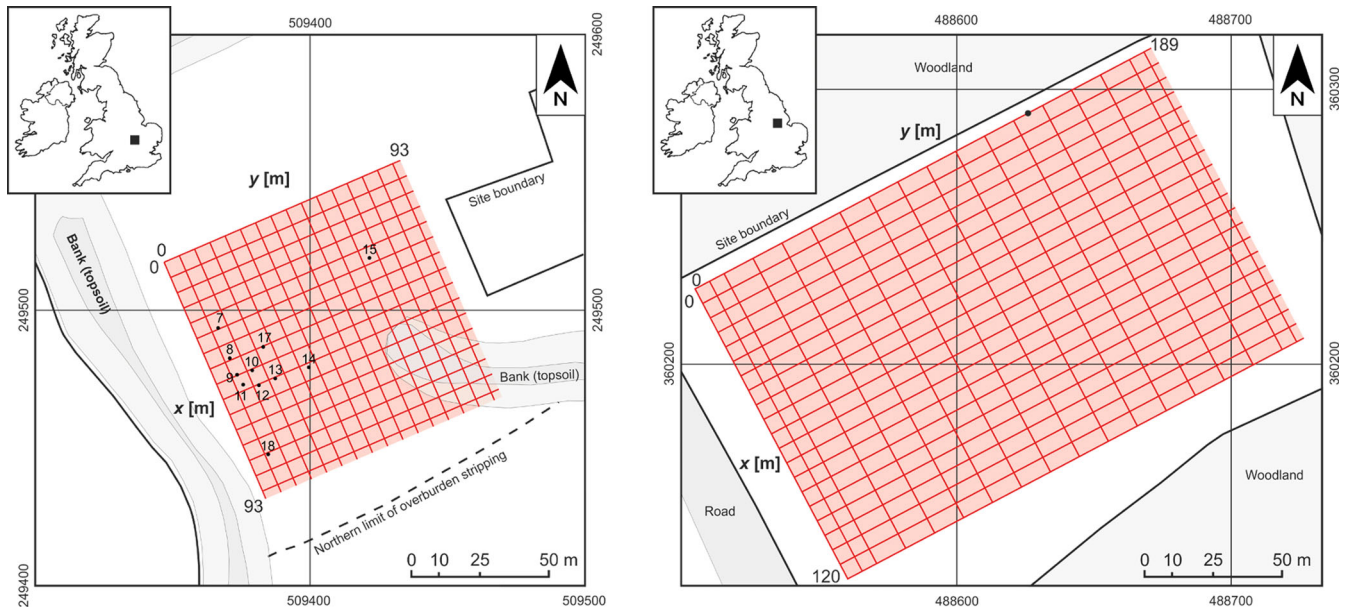
Accurate delineation of subsurface boundaries or edges is essential to achieve reliable estimates of overburden volumes and

minerals reserves. Common image processing approaches to edge detection typically involve gradients in the image. Such attempts on ERT images are detailed in Hsu *et al.* (2010) and Chambers *et al.* (2012, 2013). Problems occur, however, if the steepest gradients in the image do not coincide with the locations of the mineral interfaces, which can occur due to the nature of smoothness-constrained inversion and the fundamental lack of resolution at depth, even when the true interface is sharp. However, in certain cases where deposits are relatively homogeneous, resistivity isosurfaces can be used instead to identify interfaces (Chambers *et al.* 2013).

In this study, the aim was to develop a reliable method for the analysis of 3-D ERT images generated using standard 3-D ERT survey and inversion approaches to delineate mineral volumes and thereby estimate yields. Due to the gradational transitions in the ERT images, a fuzzy algorithm was chosen. It involves edge detection based on a machine learning approach, incorporating clustering methods guided by exploiting the probability density properties of the resistivity image. A framework was developed to automatically determine the density function. It was found that a probability density function (pdf) provided a suitable means of guiding cluster initialization that both increased accuracy and significantly reduced the runtime. The accuracy of the method was improved by choosing



**Figure 1.** Density distribution by histogram and KDE approximation on the Norton Disney site showing similarities in structures displayed. Control parameters were 50 equal width histogram bins and kernel bandwidth  $bw = 0.05$ .



**Figure 2.** Three-dimensional ERT survey designs for the Willington (left) and Norton Disney (right) sites, showing survey areas (red shading), lines (red lines) and borehole positions (black dots).

the number of clusters to match the expected number of formations under investigation.

## 2 TECHNIQUES/METHODOLOGY

### 2.1 Kernel density estimation (KDE)

The application of ERT can provide fully 3-D volumetric models of subsurface resistivity distributions. Features of contrasting resistivity can be located and characterized using KDE, a method for estimating the pdf of random variables. KDE is similar to creating a histogram to represent the distribution of data, except that it sums a symmetric weighting function, called a kernel, applied to each point in the data, rather than assigning each point to an interval. This provides large responses at areas of high frequency, that is, common values in the data (Botev *et al.* 2010). An example of this is shown in Fig. 1, which demonstrates the similarities in shape between KDE and histogram estimation.

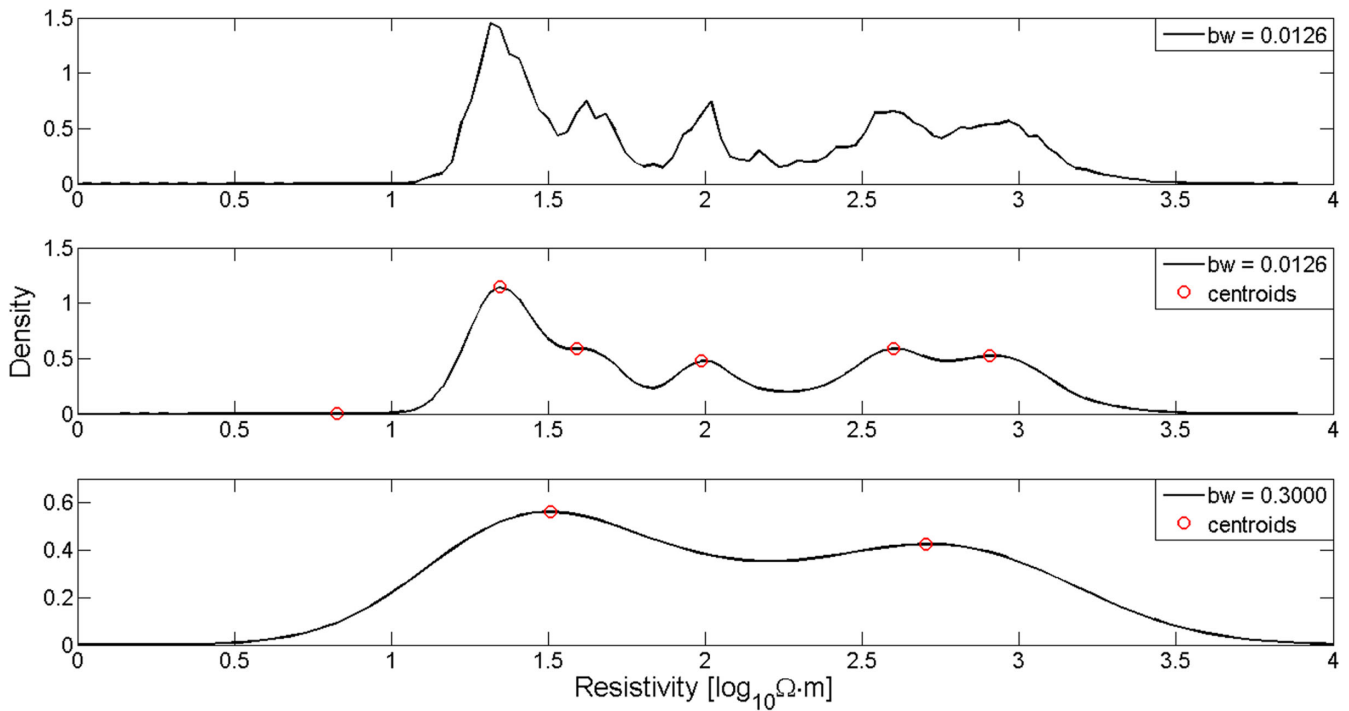
Given  $N$  sample points of a random variable  $X = \{x_1, x_2, \dots, x_N\}$  of an unknown continuous pdf,  $f$ , the KDE of  $f$  at  $x \in \mathbb{R}$ ,  $\hat{f}_\sigma(x)$ , is defined by

$$\hat{f}_\sigma(x) = \frac{1}{N} \sum_{i=1}^N K_\sigma(x - x_i),$$

where  $K_\sigma(\cdot)$  is the kernel function and  $\sigma$  is the bandwidth, a pre-defined smoothing parameter of  $K$  (Lanhan 1990). Like the choice of interval size in a histogram, the bandwidth is an important consideration that strongly influences the density estimate. A small bandwidth will give a distribution containing many small peaks, whereas choosing a large bandwidth will give wide responses and return a very smooth curve with few, wider variations.

The kernel used on the data in this research is the Gaussian function,

$$K_\sigma(x) = \phi(x; \sigma) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{x^2}{2\sigma^2}},$$



**Figure 3.** Probability density functions of Norton Disney ERT data. The top is the pdf result using kernel density estimation with automated bandwidth detection (Sheather & Jones 1991). The detected bandwidth for this is  $bw = 0.0126$ . The middle plot shows smoothing applied to the top plot, with resulting peaks identified by red circles. The bottom shows KDE applied with a wider, manually selected bandwidth of  $bw = 0.3$ , showing two detected peaks representing cluster centroids.

where each sample point  $x_i$  represents the mean of the kernel distribution in  $K_\sigma(x - x_i)$  and  $\sigma$  is the kernel bandwidth.

There exist multiple methods for selection and analysis of bandwidth suitability (Sheather & Jones 1991; Botev *et al.* 2010), and an automated approach is considered in this research to contrast with manual selection is the improved Sheather–Jones method proposed by Botev *et al.* (2010). This method involves a completely data-driven iterative scheme based on sample variance: for a number of estimated bandwidths, functionals are calculated, each being used to approximate the next until an optimal bandwidth for the data is found. An optimal bandwidth is one which minimizes the mean square error between the estimation and the true density function. This is approximated as a solution to a differential equation based on the assumed asymptotic behaviour of the error (Rosenblatt 1956).

## 2.2 Fuzzy clustering

The process of clustering data is the task of grouping a set of objects within a data set in such a way that objects of the same group are more similar to each other in a particular way than those in other groups. It has many applications, such as pattern recognition, image analysis and machine learning (Estivill-Castro 2002). Clustering techniques may be classified in terms of how they handle data and rate object similarities: the major types are hierarchical; distribution-based; density-based and centroid-based clustering. Because of the nature of data in this study, the method used belongs to the centroid-based clustering family. It is largely based on fuzzy  $c$ -means (FCM) clustering, which, in turn, takes its theory from the commonly used  $k$ -means clustering method (MacQueen 1967). The  $k$ -means algorithm is an unsupervised method for statistically classifying data. For some specified number of clusters, the method assigns each datum based on the minimized distance

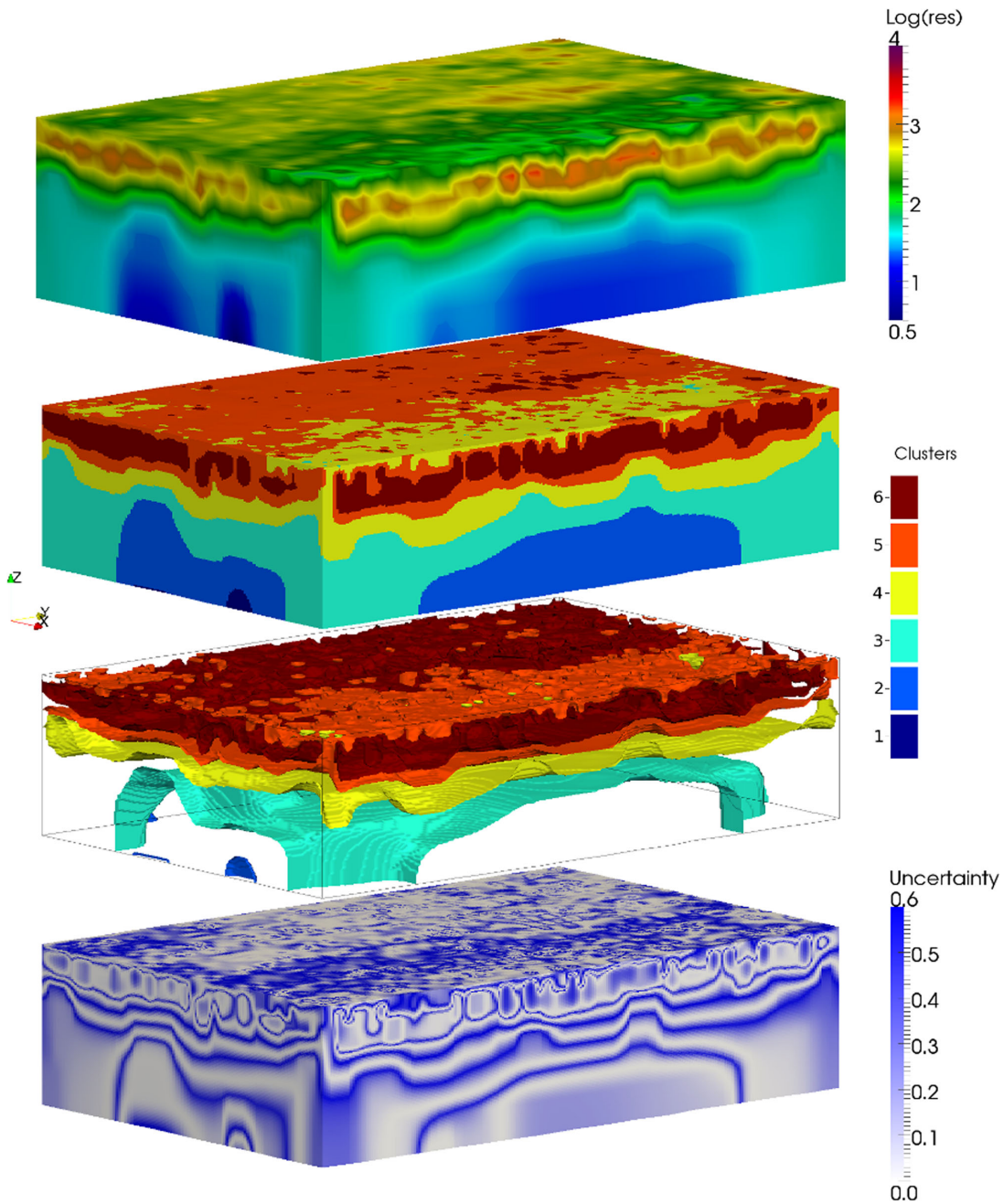
to the cluster's geometric centroid. The clusters are updated based on the new members, new centroids are found and the points are reclassified. This continues until the method reaches convergence between iterations.

An alternative to assigning data to specific clusters with an absolute in or out value is for a fuzzy subset to represent the point in relation to each cluster. This set assigns a fuzzy value to each datum for each cluster, similar to a probability value, based on the likelihood of membership of the datum into that cluster. FCM clustering makes use of this concept, assigning fuzzy membership values based on some measure of the distances of the data from the cluster centroids (Cannon *et al.* 1986). For  $n$  data points  $X = \{x_1, x_2, \dots, x_n\}$  and  $c$  clusters, a fuzzy partition of a data set can be described by a  $c \times n$  real matrix  $U$ . The entries of  $U$  must satisfy the following conditions, with  $u : X \rightarrow [0, 1]$  being a function to assign each  $x \in X$  its grade of membership to each cluster in the fuzzy set:

1.  $U_i = \{u_i(x_1), u_i(x_2), \dots, u_i(x_n)\}$  is the  $i$ th fuzzy subset of  $X$ , that is, the  $i$ th membership function.
2.  $U^j = \{u_1(x_j), u_2(x_j), \dots, u_c(x_j)\}$  are the values of the  $c$  membership values of the  $j$ th data point in  $X$ .
3.  $\sum_i u_i(x_k) = 1, \forall k$ , that is, the sum of membership values for a data point is equal to 1.
4.  $0 < \sum_k u_i(x_k) < n, \forall i$ , that is, no fuzzy subset is empty or contains all of  $X$ .

A fuzzy partition  $U^{(0)}$  is randomly generated based on the above criteria, and this is used to initialise the FCM method. A step function  $b = 0, 1, 2, \dots$  is initialized and the  $c$  cluster centroids, contained in the set  $\mathbf{v}$ , of  $U^{(b)}$  are calculated using the weighted membership function for the  $i$ th cluster centroid:

$$v_i = \frac{\sum_{k=1}^n [u_i(x_k)]^m \cdot x_k}{\sum_{k=1}^n u_i(x_k)^m}.$$



**Figure 4.** ERT model of Norton Disney site (top) with 6-clustering guided with a smoothed pdf estimated using automated bandwidth  $bw = 0.0126$ , and calculated interfaces of deposits. The bottom model shows the uncertainty of the clustering as the reciprocal of fuzzy membership used to assign clusters.

An updated fuzzy subset  $U^{(b+1)}$  can then be found using the weighting cluster assignment operation:

$$u_i(x_k) = \left[ \sum_{j=1}^c \left( \frac{d_{ik}}{d_{jk}} \right)^{\frac{2}{m-1}} \right]^{-1} \quad \forall i \in \{1, \dots, c\}, \forall k \in \{1, \dots, n\}.$$

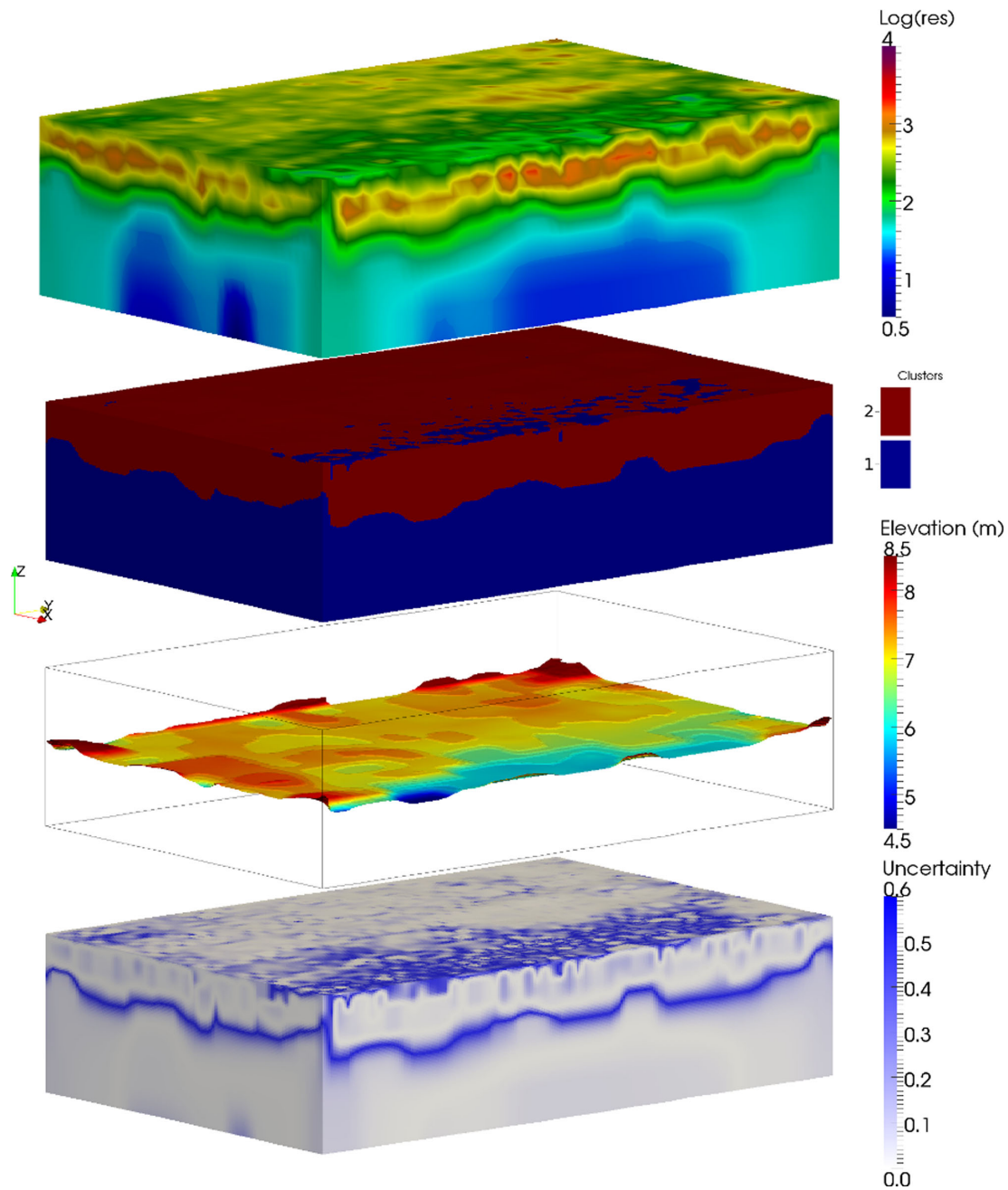
The distance function,  $d_{ik} = \|x_k - v_i\|$ , is based on some arbitrary inner product norm, and  $m \in [1, \infty)$  is a fixed weighting exponent. The FCM algorithm uses an iterative optimization to approach minima of  $U$ , and thus requires convergence less than some chosen error term. This can take significant computational effort to cluster sufficiently complex data, and due to the random nature of the initial step, may take varying lengths to reach convergence that may also be different in repeat applications due

to the dependence on  $U^{(0)}$ . It is typical to bound the number of iterations such that  $b < b_{\max}$  and if convergence is not reached, the final clustering partition is taken to be  $U = U^{(b_{\max})}$ . Furthermore, the method needs to be initialized with some value for the number of clusters,  $c$ .

For the purpose of all analysis in this paper, the weighting exponent  $m = 2$ , and the distance norm used in cluster assignment is the Euclidean norm,  $\|x\| = \sqrt{\sum x_i^2}$ .

### 2.3 Guided fuzzy clustering

The random element of the application of FCM to a data set leads to an analysis tool that does not provide identical results upon repeat applications and, in cases of some randomly generated choices of



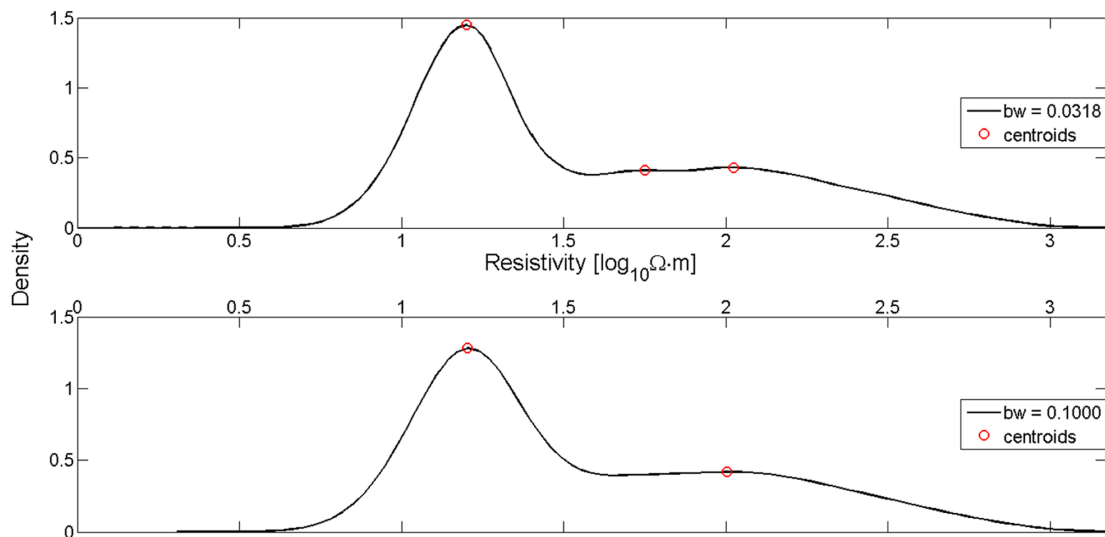
**Figure 5.** ERT model of Norton Disney site (top) with 2-clustering guided with a pdf estimated using the manually selected bandwidth  $bw = 0.30$  (second). The interface between the two clusters is taken to be the bedrock surface and an elevation map is shown. The bottom model shows the uncertainty of the clustering as the reciprocal of fuzzy membership used to assign clusters.

$U^{(0)}$ , will not give adequate clustering detail, whereas a different choice of  $U^{(0)}$  would under the same conditions. In order to remove this inconsistency in cluster membership, a method to guide the fuzzy clustering is introduced. The method also identifies a desirable cluster number as part of the generation of  $U^{(0)}$ .

For the data set  $X$  representing the random variable (in this study, resistivity values at each datum), the fuzzy subset  $U^{(0)}$  of  $X$  may be calculated using the weighted cluster assignment of FCM on some pre-defined set of cluster centroids,  $\mathbf{v}^{(-1)}$ . Here, the centroids are pre-calculated by first approximating the density distribution of the data set. Applying the KDE method to the data and finding an approximation  $\hat{f}(x)$  provides a statistically grounded analysis of the data set. Each peak in the pdf shows an estimation of individual

populations in the data. Using the Gaussian kernel in the calculation of  $\hat{f}$  means that the location of the maximum of each data peak is approximately the mean of its corresponding data population. Therefore, the number of peaks can be assumed to represent the appropriate number of clusters required to group the data. Each cluster has a centroid equal to the value of its population density maxima, and  $\mathbf{v}^{(-1)}$  is calculated such that  $\mathbf{v} = \{\bar{x}_i, i = 1, 2, \dots\}$ , where  $c = |\mathbf{v}|$  is the number of population means,  $\bar{x}_i$ , in the set  $\mathbf{v}$ .

Using this statistical approach to cluster initialisation removes the necessity of iterative optimization. Assuming that the bandwidth,  $\sigma$ , in the kernel  $K_\sigma$  is appropriate for the data set, the peaks themselves represent adequate finalized centroids for clusters. This means that using this distribution-guided fuzzy clustering,  $U \equiv U^{(0)}$  such that



**Figure 6.** Probability density functions of Willington ERT data, using kernel density estimation of automated (top) and manually selected bandwidths. The detected population peaks that correspond to cluster centroids are identified by red circles. The automatically detected bandwidth was calculated as  $bw = 0.0318$ , while  $bw = 0.10$  was chosen to approximate only two populations.

$U$  represents the final fuzzy partition containing membership values for  $X$ .

To obtain a final  $c$ -clustering of the data set, at each data point  $x_i$ , the maximum membership value in  $U$  is taken as the absolute membership. The additional fuzzy information can be used to classify uncertainty in the final cluster model.

## 2.4 Geological population segmentation

While the most common and popular approaches to edge detection use gradient information in multiple spatial directions to identify interfaces between objects in the model or image (Chambers *et al.* 2012), this approach may be limited by the gradational nature of smoothness constrained ERT images and the fundamental decrease in resolution with increasing distance from the electrodes (Wilkinson *et al.* 2012). While most approaches have been tested on 2-D ERT data (Sass 2007; Hsu *et al.* 2010), some extensions of existing methods have been used on 3-D data sets (Chambers *et al.* 2012, 2013). The method proposed here ignores the spatial properties of the resistivity image and analyses the data solely on the statistical distribution of the resistivities. Using some mapping function,  $\chi : \mathbb{R}^3 \rightarrow \mathbb{R}$ , that converts 3-D coordinate structural data to an ordered 1-D set of resistivity values, distribution-guided fuzzy clustering can be applied. The resulting ordered fuzzy membership set is then used to assign each resistivity values to the cluster for which it has maximum membership. Applying the inverse mapping  $\chi^{-1} : \mathbb{R} \rightarrow \mathbb{R}^3$ , to the 1-D ordered cluster detail vector gives a 3-D clustered set, corresponding to the original resistivity image. The interfaces between the clusters are then assumed to represent the geological boundaries.

## 3 STUDY SITES

### 3.1 Willington

The first site used in this study is located in the valley of the Great Ouse, around 4 km east of Bedford, UK, near the village of Willington (Fig. 2a). The Great Ouse is an important part of the Wash

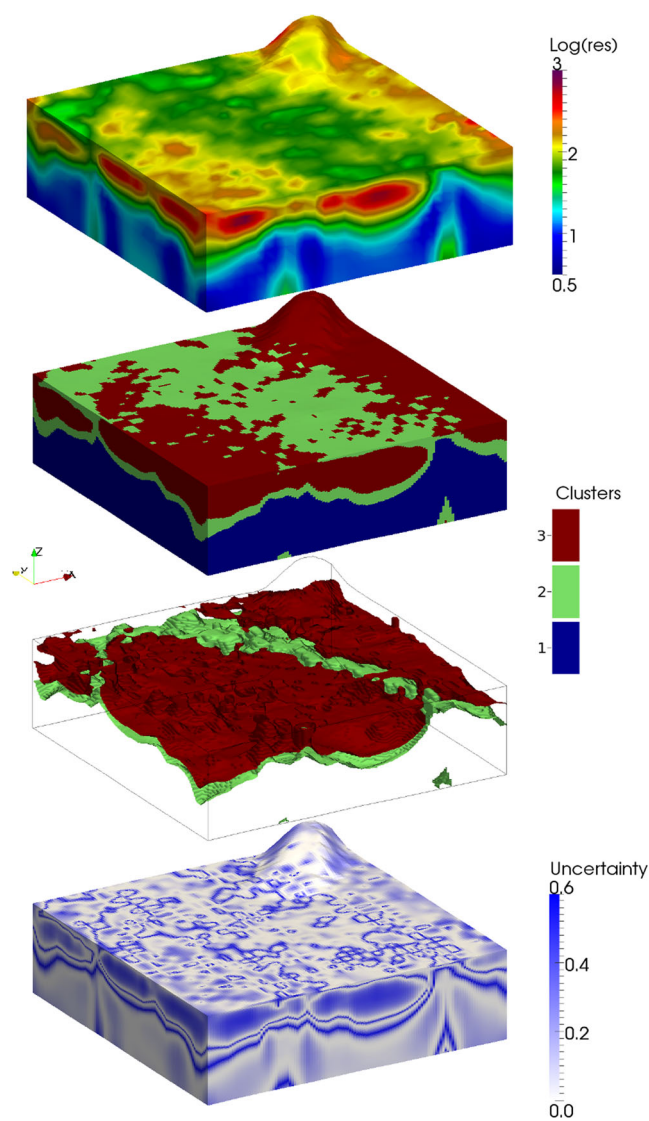
fluvial network, preserving a record of late Quaternary uplift and climate variation. It also contains records of Palaeolithic human activity.

In terms of geology, the site is composed of Quaternary alluvium and river terrace sand and gravel overlying Oxford Clay formation bedrock (Jurassic—Boreham *et al.* 2010). The Oxford Clay bedrock consists of the Peterborough member, a brownish grey, fissile mudstone. It crops out to both the southeast and northwest of the survey area and has an approximate thickness of 20 m, partly exposed by extractive activities in the river valley. The river terrace deposits here are of Ouse Valley formation, likely to have been formed by braided rivers under periglacial conditions during different Quaternary cold stages (Rogerson *et al.* 1992; Green *et al.* 1996; Bridgland 2006). There are three principle deposits observed in the area (Horton 1970; Barron *et al.* 2010; Boreham *et al.* 2010): the first is approximately 3 m thick, overlies Felmersham member and has a surface that is 0.6–2 m above the floodplain. The next terrace, with a surface 2–7 m above the floodplain, overlies Stoke Goldington member. The third terrace overlies Biddenham member, and is up to 7 m thick, its surface lying between 11 and 13 m above the floodplain. Sands and gravel of these three terraces display a similar composition, and are composed of a planar-bedded brownish yellow sand and gravel, which is mainly made up of flint and limestone.

The present day floodplain at the Willington site is covered by a brown clay and silt alluvium, which is up to 4 m thick and overlies Ouse Valley formation. In places, this may occupy channels that were cut in the Felmersham member by meandering rivers under temperate climate condition (Barron *et al.* 2010).

There has been extensive removal and reworking of superficial deposits that have occurred from mineral extraction in this area, particularly quarrying of sand and gravel from river terrace deposits. In many places, there has been exposure of bedrock as a result of the removal of sand and gravel.

The study site is situated on terrace deposits of undifferentiated Felmersham and Stoke Goldington members, overlying Oxford Clay bedrock. The terrace deposits are the focus of long-standing sand and gravel operations, and at the time of study, the topsoil was stripped and banked, exposing alluvium at surface.



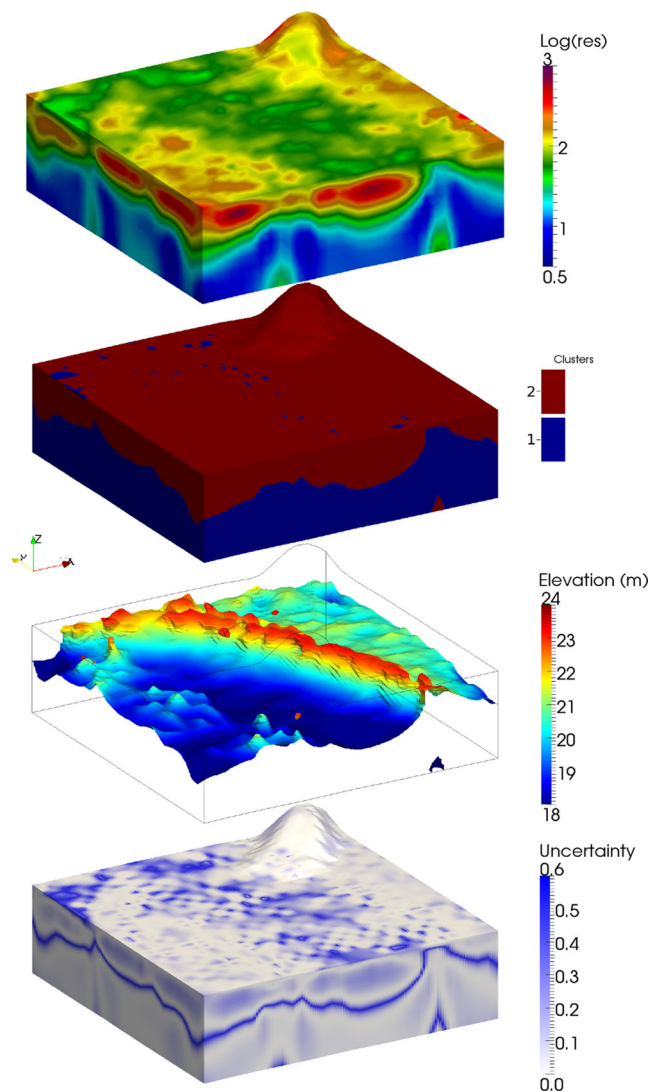
**Figure 7.** ERT model of Willington site (top) with 3-clustering guided with a smoothed pdf estimated using automated bandwidth  $bw = 0.0318$ , and calculated interfaces of deposits. The bottom model shows the uncertainty of the clustering as the reciprocal of fuzzy membership used to assign clusters.

This area was selected mainly because of the availability of good subsurface data in the form of borehole logs, which can be used to interpret and calibrate the geophysical results. Deposits are unsaturated due to dewatering in mineral working immediately south of the study site.

### 3.2 Norton Disney

The second site detailed is a sand and gravel quarry near Norton Disney, Lincolnshire, approximately 10 km northeast of Newark and the River Trent (Fig. 2b). At the time of the survey, the site was a grassed field bounded by woodland, and the land immediately surrounding the area had been worked for sand and gravel for many years. After the ERT survey was completed, the site was quarried revealing much of the bedrock across the survey area.

The geology of the Norton Disney site consists of Quaternary river terrace deposits of Balderton Sand and Gravel Member and a thin layer of topsoil, overlying flat lying Lower Lias mudstone bedrock (Jurassic—Berridge *et al.* 1999). The Lias Group is

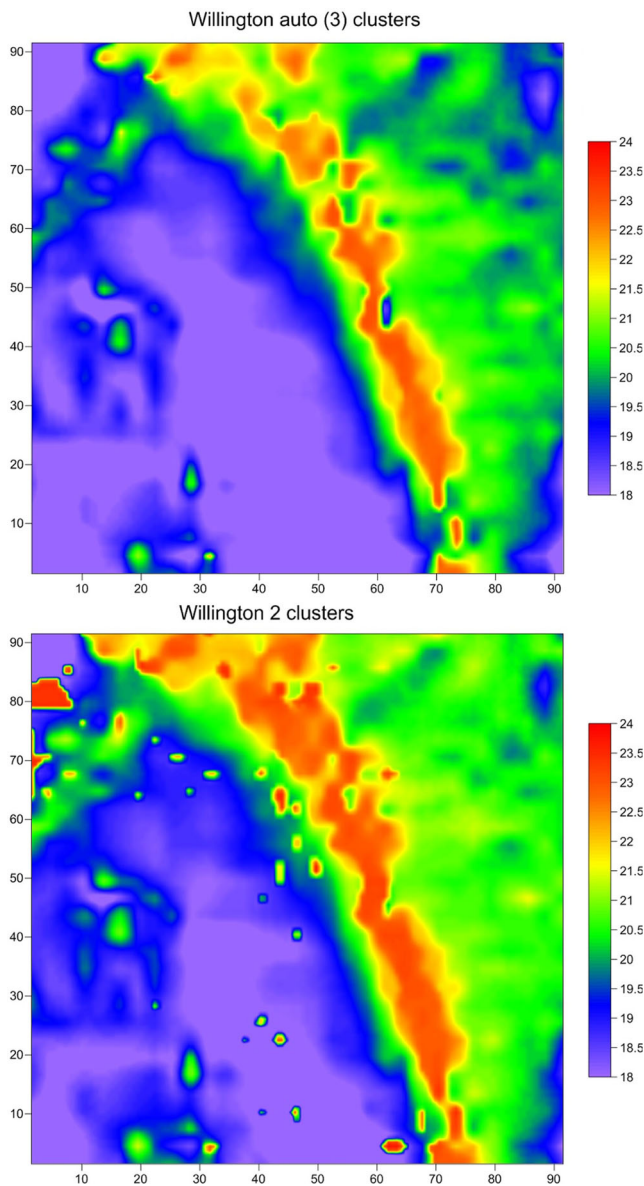


**Figure 8.** ERT model of Willington site (top) with 2-clustering guided with a pdf estimated using the manually selected bandwidth  $bw = 0.10$  (second). The interface between the two clusters is taken to be the bedrock surface and an elevation map is shown. The bottom model shows the uncertainty of the clustering as the reciprocal of fuzzy membership used to assign clusters.

composed mainly of grey shaly mudstone, with minor limestone, sandstone and ironstone beds. The site itself lies in the Scunthorpe Mudstone Formation, in the lower Lias Group, the formation of which is characterized by grey, variably calcareous, silty mudstone with numerous thin limestones. These limestones are typically around 0.1–0.3 m thick and can be well cemented and laterally persistent.

The Balderton Sand and Gravel Member is an early River Trent deposit, with a surface level around 14 to 15 m above Ordnance Datum at the Norton Disney site. The deposit at the site has a thickness of between 7.8 and 9.8 m, and is brown and yellow-brown according to borehole logs. The bulk of the deposit is slightly silty fine to coarse grained gravelly to very gravelly sand, and very sandy gravel. The deposit has poorly bedded gravels at the base, with sandier gravels further up and brown to orange-brown sandy, gravelly soil at the surface.

Borehole data were available for this site, the most recent being from 2005, including holes drilled close to the ERT survey area. Records for water levels close to the site indicate that they were



**Figure 9.** Elevation maps of detected bedrock for Willington site using three clusters (top) and two clusters.

likely to have been approximately 4 m below ground level. After quarrying, a real-time kinematic GPS survey of the exposed bedrock surface was conducted (Chambers *et al.* 2013) to provide ground truth with which to compare the results of ERT interface detection.

### 3.3 ERT data collection & inversion

Methodological descriptions of ERT deployment and image generation at the study sites are given by Chambers *et al.* (2012, 2013), so only brief descriptions of data collection and inversion are provided here.

For the Willington site, the survey was carried out in an area of 93 m  $\times$  93 m, using 16 survey lines positioned at 6 m intervals in both the  $x$ - and  $y$ -directions. Data were collected using dipole–dipole configurations with dipole lengths  $a = 3$  and 6 m and separations of  $1a$  to  $8a$ . For the Norton Disney site, the survey area dimensions were 120 m  $\times$  189 m, using 21 lines at 6 m intervals in the  $x$ -direction and 16 lines at 12 m intervals in  $y$ -direction. Dipole

lengths were 3, 6, 9 and 12 m with separations of  $1a$  to  $8a$ . Full sets of normal and reciprocal measurements were collected at both sites. Here, the line separation in  $y$  was chosen to be twice that in  $x$  to maximize survey coverage rate while avoiding too great a degree of undersampling. Bias that would normally result from the use of a single-line direction in the ERT inversion is minimized by the use of orthogonal lines (Gharibi & Bentley 2005). Dipole–dipole arrays were used as they provide a relatively high level of resolution, and can be efficiently acquired with multichannel ERT instruments in both normal and reciprocal configurations (Dahlin & Zhou 2004). The reciprocal configuration is found by exchanging current and potential dipoles, and gives the same result as the normal configuration in the absence of non-linear effects. The difference between the normal and reciprocal configuration can therefore be used to assess random and certain systematic sources of error (Wilkinson *et al.* 2012). Here, the reciprocal error is defined as percentage standard error in the mean of the forward and reciprocal measurements.

The data sets from the Willington and Norton Disney sites comprised 11 270 and 46 196 pairs of normal and reciprocal measurements, respectively. Pairs with reciprocal errors greater than 5 per cent were removed from the data set. This data removal accounted for only 2 per cent of the Willington data but approximately 13 per cent of data were removed for the Norton Disney site. This relatively high level of reciprocal errors can be accounted for by the presence of high contact resistances recorded during the field survey, which limits the current that can be injected into the subsurface (Chambers *et al.* 2013).

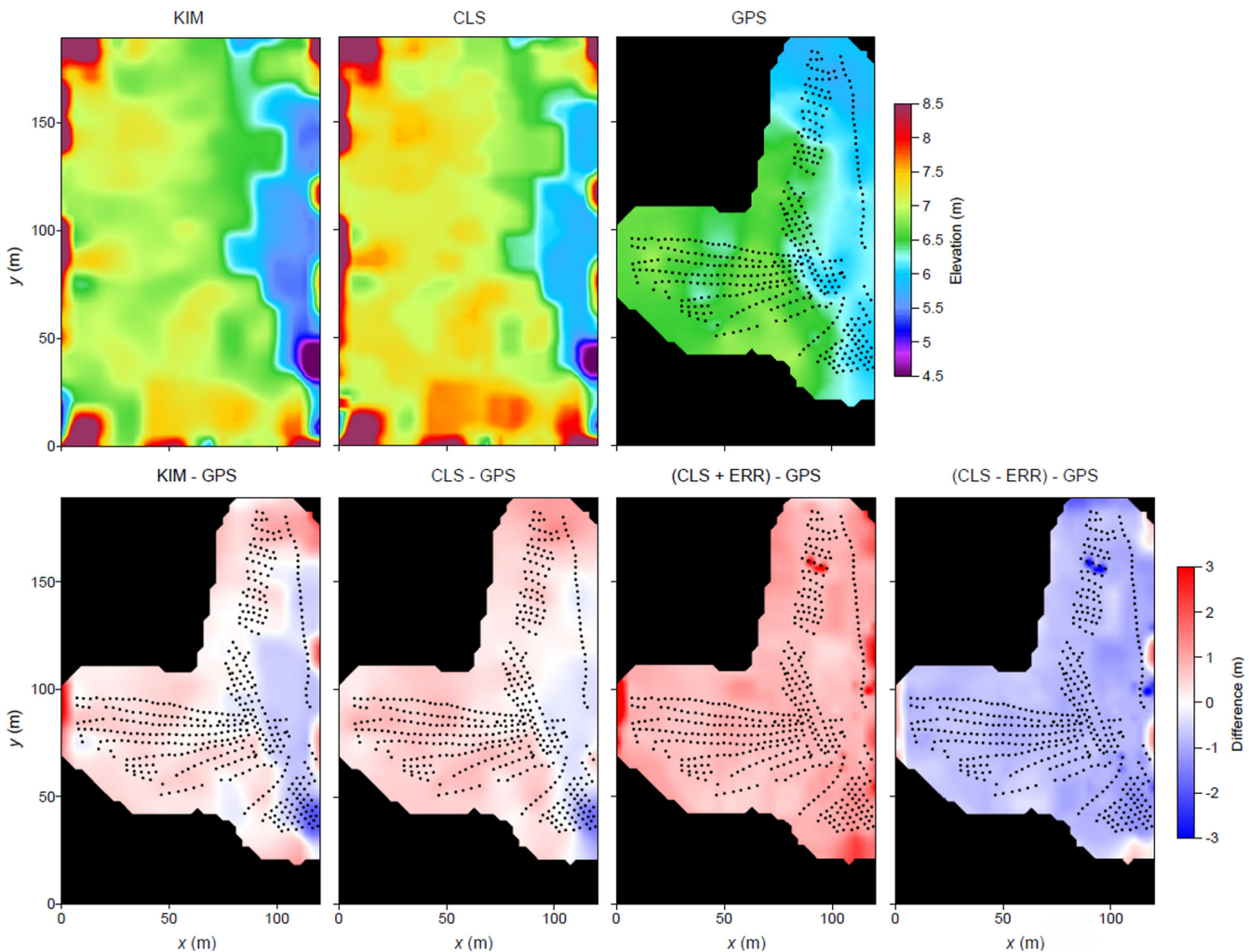
Field data were inverted using a 3-D regularized least-squares optimization algorithm (Loke & Barker 1996) and the resulting forward problem was solved using the finite-element method. After inversion, the resulting model for the Willington site contained 10 571 model cells of dimension 31  $\times$  31  $\times$  11 ( $x \times y \times z$ ), and 20 160 cells for Norton Disney, with shape dimensions 40  $\times$  63  $\times$  8.

For the Willington data, the model was produced using an L2-norm constraint. This was chosen because the site has significant gradational lithological variations that are observed in drift deposits and undulating topography of bedrock. In contrast, an L1-norm constraint was found desirable for the Norton Disney site. This method of inversion minimizes the sum of absolute values of the changes in model resistivity (Loke *et al.* 2003), leading to sharper changes within the inverted model. This was suitable since the Norton Disney deposit is dominated by sharp boundaries in the interface of sand and gravel and more conductive clay bedrock (Chambers *et al.* 2013).

## 4 RESULTS

A comparison of results at the Norton Disney site with its ground truth was undertaken. After the ERT survey, the site was excavated and bedrock details for a large proportion of the site are known. Using the guided clustering method, two levels of cluster detail were identified: the first used the pdf from an approximation with automatically detected bandwidth (Botev *et al.* 2010) that were then smoothed by a moving average method to remove superficial local maxima. This identified six populations which can be seen in the pdf in Fig. 3. However, this did not reflect the broad geological divisions observed at the site (i.e. river terrace sands and gravels overlying mudstone bedrock) and so a manually selected bandwidth was also used that gave two distinct populations. This was chosen by reviewing the generated density function and adjusting the bandwidth accordingly to give the desired resolution. Figs 4 and 5 show





**Figure 10.** Elevation maps showing the bedrock surfaces detected by the known interface method (KIM) and guided clustering method (CLS) using  $bw = 0.30$ . The GPS measured bedrock surface (taken as the ground truth) is also shown. The bottom row shows the corresponding differences between the detected and ground truth bedrock for the two methods, and the limits of uncertainty of the CLS surface, where the error ERR is taken to be half the FWHM of the fuzzy uncertainty distribution around the interface.

the models produced by the two different clustering approaches. Additionally, a model displaying the uncertainty of final chosen fuzzy subset for each clustering is given. The uncertainty was taken to be the complement of the membership value to the final cluster of each data point.

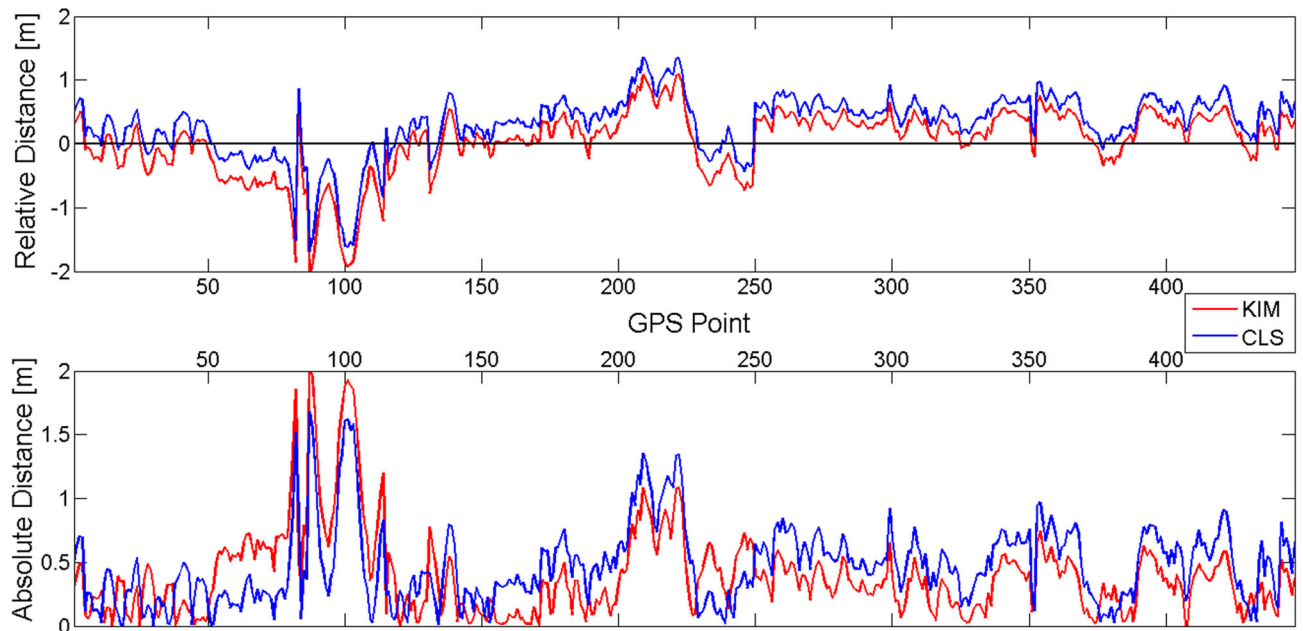
The same process of clustering was applied to the Willington site. In this case, the automated bandwidth selection produced a pdf with three clusters, but again a second bandwidth was manually determined to match the number of clusters to the major formations at the site (terrace deposits overlying clay bedrock). The difference between the bandwidths is smaller in this case, as can be seen in Fig. 6 which shows the pdfs from the two approximations. Peaks have been identified on the plots and correspond with the centroids used to guide the clustering. Figs 7 and 8 show models with the 2- and 3-clusterings and their respective fuzzy uncertainty.

In both Figs 4 and 7, where automated bandwidth is used, more than one interface is present. These are all shown as clustering surfaces, but none of them are selected to represent the bedrock. Conversely, for the 2-clustering results (Figs 5 and 8), there is only a single surface separating the clusters, which is assumed to represent the interface between the deposit and the bedrock. In these figures, the single surface is coloured based on its elevation to highlight the

topography. For the more detailed analysis of Willington results, the only continuous clustering interface for the automated bandwidth model is assumed to represent the bedrock and is compared to the interface from the manual bandwidth results in Fig. 9.

Smoothing of the automated bandwidth density estimations is necessary due to the nature of the algorithm used. This typically selects a relatively small bandwidth that gives a good global fit to the data but picks out an increased number of potentially insignificant populations. A simple smoothing algorithm (moving average) has given positive results for removing these fluctuations while leaving larger significant populations present. A similar issue arises in the choice of bin sizes for the data histograms, where too small a bin range can lead to unwanted detail. In both cases, matching the number of distributions to the expected number of major formations has produced better results for estimating the bedrock surface. However, in cases where less ground truth is available or where the deposit is known or suspected to be highly variable, the automated estimates will provide a useful ‘first look’ analysis of the images.

In Fig. 10, a comparison of the detected bedrock for the manual bandwidth cluster model with the GPS bedrock surface ground truth for Norton Disney is shown. The results are also compared to the known interface method (KIM; Chambers *et al.* 2013), which uses



**Figure 11.** Plots showing detected relative distance (top) and absolute distance from bedrock of KIM (red) and clustering results at multiple known elevations at the Norton Disney site.

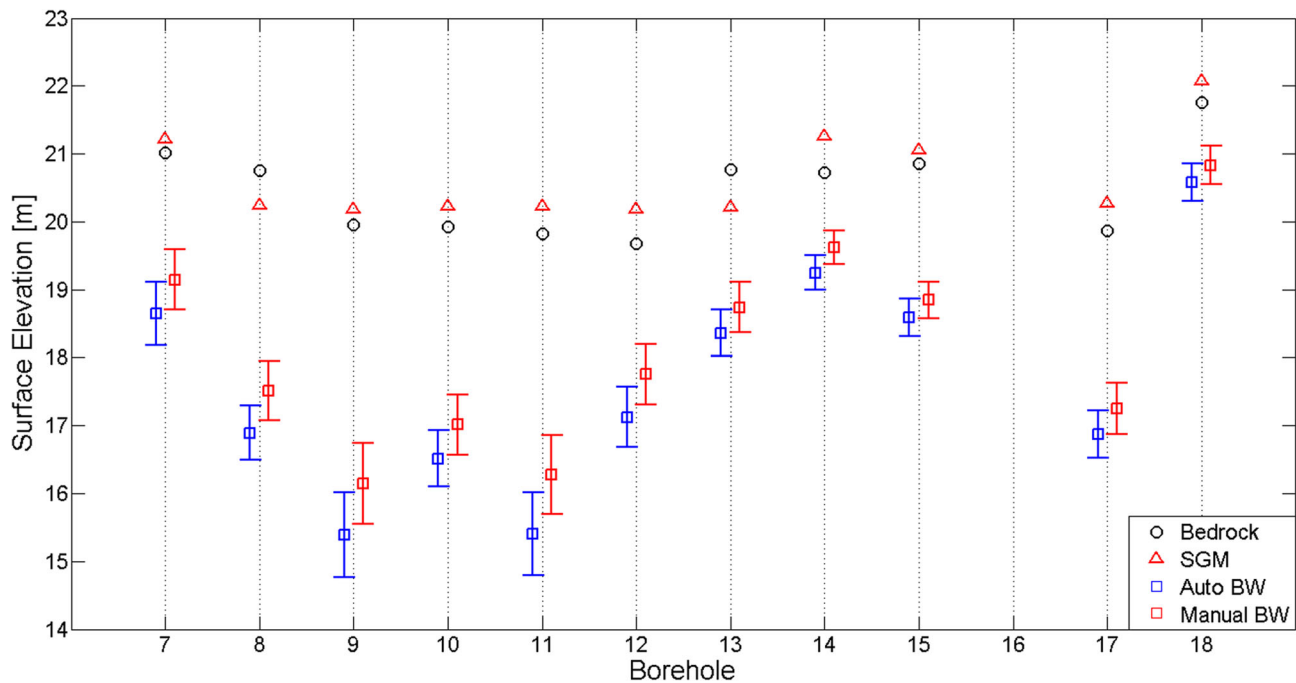
**Table 1.** Willington borehole information with bedrock elevation [m] and comparative values of predicted surface using SGM and guided FCM using kernel bandwidths 0.0318 (automated) and 0.1000 (manually selected). Additionally, for the results from clustering, the full width at half maximum (FWHM) of the uncertainty distribution around the interface is included.

| Borehole | $x$   | $y$   | Bedrock | SGM   | $bw = 0.03$ | FWHM | $bw = 0.10$ | FWHM |
|----------|-------|-------|---------|-------|-------------|------|-------------|------|
| 7        | 29.50 | 8.60  | 21.02   | 21.22 | 18.64       | 0.92 | 19.15       | 0.89 |
| 8        | 41.20 | 8.10  | 20.76   | 20.25 | 16.89       | 0.80 | 17.51       | 0.87 |
| 9        | 49.30 | 8.44  | 19.95   | 20.19 | 15.39       | 1.25 | 16.14       | 1.20 |
| 10       | 48.50 | 13.90 | 19.93   | 20.23 | 16.51       | 0.83 | 17.01       | 0.88 |
| 11       | 51.90 | 8.80  | 19.83   | 20.23 | 15.40       | 1.22 | 16.28       | 1.17 |
| 12       | 54.40 | 14.00 | 19.68   | 20.18 | 17.13       | 0.88 | 17.75       | 0.89 |
| 13       | 54.50 | 20.40 | 20.77   | 20.21 | 18.36       | 0.68 | 18.73       | 0.74 |
| 14       | 55.60 | 33.20 | 20.73   | 21.26 | 19.24       | 0.51 | 19.62       | 0.50 |
| 15       | 27.80 | 69.20 | 20.86   | 21.06 | 18.59       | 0.55 | 18.84       | 0.54 |
| 17       | 42.20 | 21.00 | 19.87   | 20.27 | 16.88       | 0.70 | 17.25       | 0.76 |
| 18       | 78.80 | 7.20  | 21.76   | 22.07 | 20.58       | 0.55 | 20.83       | 0.56 |

a resistivity isosurface known to intersect the bedrock surface at a chosen point. The distances between the detected surfaces and the GPS surface are shown in Fig. 11. The average absolute distances for the clustering and known interface methods are 0.47 and 0.40 m, respectively. An error estimate was derived from the fuzzy uncertainty distribution in the vicinity of the interface (Fig. 10). This was taken to be half the full width at half maximum (FWHM) of the uncertainty distribution in the vertical direction. This error estimate had a mean of 0.80 m and a standard deviation of 0.26 m across the model space. In the bottom rightmost two images in Fig. 10, it can be seen that the error limits effectively bracket the bedrock interface, showing that the bedrock has been detected to within the limits of accuracy of the clustering method.

For the Willington site, it was known that the KIM would not accurately detect the interface due to the variability of the deposit (Hsu *et al.* 2010; Chambers *et al.* 2012). However, a steepest gradient method (SGM) was found to be applicable in this case

(Chambers *et al.* 2012). The SGM results are compared with the guided clustering method surfaces for both two and three clusters. This comparison does not cover the whole site as ground truth is only known from borehole logs at 11 locations. Table 1 and Fig. 12 give the results and the FWHM for the clustering surface estimates (these are shown as error bars in Fig. 12). For this site, the SGM produced results very close to the interfaces detected in the boreholes but the clustering results did not agree to within their error estimates. This is because the guided clustering method produces resistivity isosurfaces which, like the KIM, do not work well if the deposit is variable (Chambers *et al.* 2012 found that the inverted model resistivity values at the 11 drilling interface locations ranged from 42 to 520  $\Omega\text{m}$ ). Due to the heterogeneity of the model, there are some discrepancies, similar to the issues identified in Chambers *et al.* (2013) when using the KIM algorithm. It was found that a resistivity isosurface cannot be expected to delineate the mineral in such data, suggesting that the error does not lie in the method.



**Figure 12.** Plot showing detected elevations of bedrock (circle) and three methods on the Willington data at borehole sites. The methods shown are steepest gradient method (triangle) and guided clustering using automated bandwidth  $bw = 0.0318$  (blue square) and manually selected  $bw = 0.10$  (red square). The error bars are given as half of the FWTM of uncertainty about the interface at each point.

This is highlighted in Fig. 12, and shows that even within the fuzzy uncertainty of clustering, the results do not coincide with borehole data.

## 5 DISCUSSION AND CONCLUSION

Guided fuzzy clustering has been applied to 3-D ERT images of sand and gravel deposits to detect the interface between the deposit and the bedrock. The method is, however, independent of the spatial dimension of the data and could equally be applied to similar 2.5-D images (e.g. Hsu *et al.* 2010).

The use of fuzzy clustering addresses uncertainty in the models, and reduces the impact of gradational interfaces that can cause problems with gradient-based interface detection on certain sites (Chambers *et al.* 2013). When compared to other methods that attempt to assign resistivity isosurfaces to formation interfaces (Chambers *et al.* 2013), considerably fewer assumptions were needed in the guided fuzzy clustering approach. Knowledge of the expected number of major formations was sufficient to achieve good results; otherwise, there was little user intervention. Even when manually selecting bandwidths, a plot of the pdf provides an easily accessible visualization of the feature space regardless of the size and dimension of the input data.

In terms of runtime efficiency, the original FCM algorithm runs at  $O(cbN)$ , where  $N$  is the number of data,  $c$  is the number of clusters and  $b$  is the number of iterations required to converge:  $b$  increases in magnitude with the increase of dimension. The iteration limit,  $b_{\max}$ , is typically chosen to be 1000 for a 3-D data set, although convergence does not often occur. With the distribution-guided clustering introduced in this paper, the framework runs at  $O(cN)$ , with KDE and local maxima detection running at  $O(N)$ . For the given examples, the algorithm takes approximately 1 min to run on a dual Intel Xeon E5620 system.

Limitations of this approach fall largely into two categories: limitations of resistivity imaging and limitations of the edge detection methods. The former include the exponential decrease of resolution with distance from the electrodes and the gradational nature of interfaces produced by smoothness-constrained inversion. The edge detection algorithm involved KDE and FCM to identify the interfaces, which discard any spatial information. It may be possible to improve the results by incorporating summation of the membership function over a defined neighbourhood of each cell under consideration (Chuang *et al.* 2006). While the means taken from KDE were identified as population centroids, no other distribution information was used. Using the trough locations of the pdf to approximate, the standard deviation of each population could give a means of further guiding the clustering approach by incorporating it into the FCM weighting function.

An extension of the methods presented in this paper could follow multiple directions. One of the more simple possibilities includes applying the methods to higher dimensional data, such as 4-D (time-lapse) ERT monitoring of sites. A further improvement of the methods could be made by incorporating data from other geophysical survey methods (Ellefsen *et al.* 1998). This could be achieved by adapting multivariate versions of both fuzzy  $c$ -means and KDE (Gustafson & Kessel 1978; Silverman 1986; Simonoff 1996). This may further isolate deposits and would perhaps increase the capability of classification.

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