

Decomposition of Bivariate Inequality Indices by Attributes

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Abstract

We provide, for the class of relative bidimensional inequality indices, a decomposition of inequality into two univariate Atkinson-Kolm-Sen indices and a third statistic which depends on the joint distribution of resources.

Keywords: multidimensional inequality, relative indices, correlation increasing transfers, copulas.

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When measuring inequality over two attributes, it may be of practical interest to express inequality as a function of the underlying level of inequality in each attribute. This note provides, for the class of relative bidimensional inequality indices, a decomposition of inequality into two univariate Atkinson-Kolm-Sen [AKS] indices ¹ and a third statistic which depends on the joint distribution of resources.

1 Notation and definitions

A Population consists of n individuals. Individual i has resources $x_i \doteq (x_{i1}, x_{i2})$, where $x_i \in R_{++}^2$. The joint distribution is a matrix $X \doteq \begin{bmatrix} x_1 \\ \vdots \\ x_n \end{bmatrix} \in$

M^n , the set of all $n \times 2$ matrices with strictly positive elements. Accordingly, the j th column of X , X_j , gives a distribution (an $n \times 1$ vector) for the j th attribute in the population, and we also write $X = [X_1 \ X_2]$.

A bidimensional inequality index is a real valued function $I(X) : M^n \rightarrow R_+$. Underlying I is a social welfare function $W(X) : M^n \rightarrow R$. We take $W(\cdot)$ to be the average of individual welfare levels:

ADD : $W(X) = \frac{1}{n} \sum_i u(x_i)$ for any X (additive separability).

CON : $W(\cdot)$ is continuous in X .

PAR : $W(\cdot)$ is strictly increasing in all its arguments (Pareto principle).

EQUAL : $W(\cdot)$ is a strictly quasi-concave function (preference for equality).

RSINV : If Λ is an 2×2 diagonal matrix with strictly positive elements, then $W(X^A) = W(X^B) \iff W(X^A \Lambda) = W(X^B \Lambda)$, (ratio-scale invariance).

The above axioms on W entail that $u(\cdot)$ is continuous, increasing, concave and ratio-scale invariant. In what follows, we shall refer to the set of axioms *ADD*, *CON*, *PAR*, *EQUAL* and *RSINV* as the *basic axioms*.

Let $\mu \doteq (\mu_1, \mu_2)$ denote the vector of sample means and let $w^o \doteq W(X)$, be the level of welfare attained by X . Then, if $W(\cdot)$ satisfies the basic axioms, we may define a scalar $\theta(X)$ in the unit interval, such that $u(\theta\mu) = \frac{1}{n} \sum_i u(x_i) = w^o$. Starting from a distribution X , a fraction θ of the sum-total of *each* attribute would lead to the same level of welfare as X , provided each attribute were equally distributed in the population. Therefore, θ is a

¹See Atkinson (1970), Sen (1973), Kolm (1977) and Tsui (1995).

measure of *equality* in X , and

$$I(X) = 1 - \theta(X) \tag{1}$$

is the corresponding relative inequality index (Kolm, 1977). Likewise, below we shall define equality indices θ_j for each separate attribute. Our aim is to decompose θ as a function of statistics θ_1, θ_2 , pertaining to the marginal distributions and a term κ which depends on the correlation between the two attributes.

Denote by F_{12} the cumulative distribution function of X , and by F_j the marginal distribution of the j -th attribute X_j . By Sklar's theorem (Sklar, 1959), there exists a copula function $c : [0, 1] \times [0, 1] \rightarrow [0, 1]$ such that for any (x_1, x_2) , $F_{12}(x_1, x_2) = c[F_1(x_1), F_2(x_2)]$. The copula captures the association between the two variables, and the joint distribution F_{12} may be identified with its two marginal distributions and its copula function.

In general terms, we shall say that $\theta(X)$ is decomposable into two indices θ_1, θ_2 , and a measure of association κ if there exist four $R \rightarrow R$ functions $h(\cdot), g^1(\cdot), g^2(\cdot)$, and $\chi(\cdot)$, where h, g^1 , and g^2 are monotonically increasing, such that:

$$h[\theta(F_{12})] = g^1[\theta_1(F_1)] + g^2[\theta_2(F_2)] + \chi[\kappa(c(F_1, F_2))] \tag{2}$$

where $\kappa(\cdot)$ measures the association between attributes via the copula function.

Note that there are different ways to define equality in a given dimension, in particular with regard to whether the distribution over the other attribute has been equalized. Two polar cases may thus be defined. First, consider a situation where only x_2 is redistributed in the population. Then, we may define a scalar $\rho_2(X)$ such that $\frac{1}{n} \sum_i u(x_{i1}, \rho_2 \mu_2) = w^o$. Hence, ρ_2 is a measure of equality in X_2 holding the distribution X_1 constant. Next, there exists a scalar $\gamma_1(X)$, such that $u(\gamma_1 \mu_1, \rho_2 \mu_2) = \frac{1}{n} \sum_i u(x_{i1}, \rho_2 \mu_2)$. The scalar γ_1 is the fraction of μ_1 that, if equally distributed in the population, would attain the same level of welfare as X , *conditional* on everyone receiving an endowment $\rho_2 \mu_2$ of x_2 . Both γ_1 and ρ_1 provide measures of equality in the first attribute, and we derive below inequalities relating γ_j and ρ_j .

When $u(\cdot)$ is additively separable in x_1 and x_2 we show that $\gamma_j = \rho_j$, and θ may be decomposed in terms of γ_1 and γ_2 . In the more general context where $u(\cdot)$ is not separable across attributes, we show that θ may be decomposed in terms of γ_1, γ_2 , and a third index κ that captures the correlation between

the two attributes. Finally, we generalize our decomposition in the context of $p > 2$ attributes.

2 Inequality and the correlation between attributes

It may certainly be advocated that changes in the joint distribution which increase the correlation between attributes for given marginal distributions, ought to modify inequality. Formally, let x_l and x_m denote any two rows of X , and define $y_i \doteq x_i$ for all $i \neq l, m$ and $y_{lj} \doteq \max\{x_{lj}; x_{mj}\}$, $y_{mj} \doteq \min\{x_{lj}; x_{mj}\}$ $j = 1, 2$; following Tsui (1999), we shall state that Y is a correlation increasing transformation (*CIT*) of X . Given that a *CIT* does not affect the marginal distributions, it amounts to a transfer which increases the rank correlation between attributes (i.e. Kendall's *Tau*). The next axiom formalizes the argument $W(Y) \leq W(X)$:

CITAV: For any X and Y , if Y is a *CIT* of X , then $W(Y) \leq W(X)$, (aversion to correlation increasing transformations).

To provide some intuition for the results to follow, suppose x_1 is income and x_2 is health. With $u_{12} < 0$, the marginal utility of income is higher for the unhealthy. Therefore, a redistribution of income from a person in good health to a less healthy one may increase welfare, even if the latter is initially wealthier. The following lemma (Atkinson and Bourguignon, 1982) formalizes this argument.

Lemma *Assume $W(\cdot)$ satisfies the basic axioms. If in addition W has an aversion (preference) for *CITs*, then $u_{12} \leq 0$ ($u_{12} \geq 0$).*

If $W(\cdot)$ satisfies the basic axioms then $u(\cdot)$ is one of three forms (Aczél, 1988; Tsui, 1995):

$$u(x_{i1}, x_{i2}) \doteq x_{i1}^\alpha x_{i2}^\beta \quad \alpha, \beta > 0, \quad \alpha + \beta \leq 1 \quad (3)$$

$$u(x_{i1}, x_{i2}) \doteq -x_{i1}^\alpha x_{i2}^\beta \quad \alpha, \beta < 0 \quad (4)$$

$$u(x_{i1}, x_{i2}) \doteq \alpha \ln x_{i1} + \beta \ln x_{i2} \quad \alpha, \beta > 0 \quad (5)$$

The first of these forms has $u_{12} > 0$, i.e. $W(\cdot)$ is increasing in *CITs*. For the second form $u_{12} < 0$, while for the latter $u_{12} = 0$, i.e. $W(\cdot)$ is insensitive to *CITs*. We begin with this simpler case.

Under (5), $u(x_1, x_2)$ is additively separable in x_1 and x_2 ; accordingly,

$$w^o = \alpha \ln \theta \mu_1 + \beta \ln \theta \mu_2 = \alpha \ln \gamma_1 \mu_1 + \beta \ln \gamma_2 \mu_2 \quad (6)$$

Upon defining $\omega \doteq \alpha/(\alpha + \beta)$, we obtain the elasticity decomposition

$$\ln \theta(X) = \omega \ln \gamma_1(X_1) + (1 - \omega) \ln \gamma_2(X_2) \quad \text{when } u_{12} = 0 \quad (7)$$

so that a 1% increase in *equality* of the distribution X_1 results in a ω % increase in the overall level of equality θ . Notice also here that $\rho_j = \gamma_j = \exp \left[\frac{1}{n} \sum_i \ln x_{ij} - \ln \mu_j \right]$.

For the general case $u_{12} \neq 0$, it is usually not possible to modify the distribution X_1 while holding γ_2 constant. The elasticity decomposition is given in the following proposition.

Proposition 1: *Assume $W(\cdot)$ satisfies the basic axioms. Then*

$$\ln \theta(X) = \omega \ln \gamma_1(X_1) + (1 - \omega) \ln \gamma_2(X_2) + \frac{1}{(\alpha + \beta)} \ln \kappa(X) \quad (8)$$

where γ_j is an AKS index for X_j and κ is given by

$$\kappa \doteq \frac{n \sum_i x_{i1}^\alpha x_{i2}^\beta}{\sum_i x_{i1}^\alpha \sum_i x_{i2}^\beta} \quad (9)$$

Proof: Consider first (3). From the definition of θ , it follows that $w^o = (\theta \mu_1)^\alpha (\theta \mu_2)^\beta = (\gamma_1 \mu_1)^\alpha (\rho_2 \mu_2)^\beta$. Therefore, $(\alpha + \beta) \ln \theta = \alpha \ln \gamma_1 + \beta \ln \rho_2$. In other terms,

$$\ln \theta = \omega \ln \gamma_1 + (1 - \omega) \ln \gamma_2 + (1 - \omega) \ln(\rho_2/\gamma_2) \quad (10)$$

Likewise, $w^o = (\rho_1 \mu_1)^\alpha (\gamma_2 \mu_2)^\beta = \frac{1}{n} \sum_i (\rho_1 \mu_1)^\alpha x_{i2}^\beta$. Hence,

$$\gamma_2 = \frac{1}{\mu_2} \left(\frac{1}{n} \sum_i x_{i2}^\beta \right)^{1/\beta} \quad (11)$$

Since $w^o = \frac{1}{n} \sum_i x_{i1}^\alpha x_{i2}^\beta = \frac{1}{n} \sum_i x_{i1}^\alpha (\rho_2 \mu_2)^\beta$, we obtain

$$\rho_2 = \frac{1}{\mu_2} \left(\frac{\sum_i x_{i1}^\alpha x_{i2}^\beta}{\sum_i x_{i1}^\alpha} \right)^{1/\beta} \quad (12)$$

Together, (11) and (12) give $(1 - \omega) \ln(\rho_2/\gamma_2) = (\ln \kappa)/(\alpha + \beta)$, which completes the proof for (3). For (4), it is readily verified that γ_2 and ρ_2 are also given as in (11) and (12) above. \square

If Y is a *CIT* of X , it is the case that $F_j(Y_j) = F_j(X_j)$. Hence, $\gamma_j(Y) = \gamma_j(X)$. On the basis of (8), we therefore obtain

$$\ln \theta(Y) - \ln \theta(X) = \frac{1}{(\alpha + \beta)} [\ln \kappa(Y) - \ln \kappa(X)] \quad (13)$$

Since κ is a measure of association between the two attributes, it is desirable that this function be rising in correlation increasing transformations. The next proposition formalizes this argument.

Proposition 2: *If Y is a *CIT* of X , then $\kappa(Y) \geq \kappa(X)$.*

Proof: For $j = 1, 2$ let $\underline{x}_j < \bar{x}_j$. If Y is a *CIT* of X , we may take $x_l \doteq (\bar{x}_1, \underline{x}_2)$, $x_m \doteq (\underline{x}_1, \bar{x}_2)$, $y_l \doteq (\bar{x}_1, \bar{x}_2)$, $y_m \doteq (\underline{x}_1, \underline{x}_2)$ and $y_i \doteq x_i$ for all $i \neq l, m$. Since the marginal distributions are unaffected by *CITs*, it follows from (9) that $\kappa(Y) \geq \kappa(X)$ when the difference $\Delta \doteq \bar{x}_1^\alpha \bar{x}_2^\beta + \underline{x}_1^\alpha \underline{x}_2^\beta - \bar{x}_1^\alpha \underline{x}_2^\beta - \underline{x}_1^\alpha \bar{x}_2^\beta$ is positive. Since $\Delta \doteq (\bar{x}_1^\alpha - \underline{x}_1^\alpha)(\bar{x}_2^\beta - \underline{x}_2^\beta)$, we have that $\Delta \geq 0$ when α and β are of the same sign. \square

Returning to (13) we note that when the social welfare function is averse to correlation increasing transformations ($\alpha, \beta < 0$), a *CIT* results in a reduction in θ , that is, an increase in inequality. In other terms, the change in inequality that results from a *CIT* is captured via κ .

3 Some inequalities

Recall that ρ_2 is a measure of equality in X_2 holding the distribution X_1 constant, whereas γ_2 measures the level of equality in X_2 having equalized the distribution X_1 . Consider (4) where $W(\cdot)$ is averse to *CITs*. Then, if the correlation between x_1^α and x_2^β is positive, we expect the level of equality in X_2 to be greater after equalizing the distribution X_1 than before, i.e. that $\gamma_2 \geq \rho_2$. We have:

Proposition 3: *In the context of (3), $\gamma_j \leq \rho_j$ iff the correlation between x_1^α and x_2^β is positive. For (4), $\gamma_j \geq \rho_j$ iff the correlation between x_1^α and x_2^β is positive. For (5), $\gamma_j = \rho_j$.*

Recall that κ is a measure of association between x_1 and x_2 . It is greater than (less than) unity when x_1^α and x_2^β are positively (negatively) correlated. The index κ is a simple transform of the ratio of ρ_j to γ_j :

$$\kappa = \left(\frac{\rho_2}{\gamma_2} \right)^\beta = \left(\frac{\rho_1}{\gamma_1} \right)^\alpha \quad (14)$$

In the context of (3), $\alpha, \beta > 0$ while in (4) $\alpha, \beta < 0$, from which the results of the proposition follow.

4 Generalization

A final point we turn to is the decomposition of inequality in the general context of $p > 2$ attributes. Define $x_i \doteq (x_{i1}, \dots, x_{ip})$. For $u(x_i)$ to satisfy the basic axioms it must take one of the following forms: $u(x_i) \doteq x_{i1}^{\alpha_1} \cdots x_{ip}^{\alpha_p}$, $\alpha_j > 0$ for all j , and $\sum_j \alpha_j \leq 1$; $u(x_i) \doteq -x_{i1}^{\alpha_1} \cdots x_{ip}^{\alpha_p}$, $\alpha_j < 0$ for all j and finally, $u(x_i) \doteq \sum_j \alpha_j \ln x_{ij}$, where $\alpha_j > 0$ for all j . Propositions 1 and 2 generalize to higher dimensions. The decomposition of inequality in the context of p attributes is given by

$$\left(\sum_j \alpha_j \right) \ln \theta = \sum_j \alpha_j \ln \gamma_j + \ln \kappa \quad (15)$$

$$\kappa \doteq \frac{n \sum_i x_{i1}^{\alpha_1} \cdots x_{ip}^{\alpha_p}}{\sum_i x_{i1}^{\alpha_1} \cdots \sum_i x_{ip}^{\alpha_p}} \quad (16)$$

As before, γ_j denotes an AKS equality index for the distribution X_j , whereas now κ is a function of the product moment between the variables $x_1^{\alpha_1}, \dots, x_p^{\alpha_p}$ thus generalizing the earlier bivariate measure of association (9).

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