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# Cumulative Evidence for Relationships Between 8q24 Variants and Prostate Cancer 

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Multiple independent cancer susceptibility loci at chromosome $8 q 24$ have been identified by GWAS (Genome-wide association studies). Forty six articles including 60,293 cases and 62,971 controls were collected to conduct a meta-analysis to evaluate the associations between 21 variants in 8 q24 and prostate cancer risk. Of the 21 variants located in 8q2\5 were significantly associated with the risk of prostate cancer. In particular, both homozygous AA and heterozygous CA genotypes of rs16901979, as well as the AA and CA genotypes of rs1447295, were associated with the risk of prostate cancer. Our study showed that variants in the 8 q 24 region are associated with prostate cancer risk in this large-scale research synopsis and meta-analysis. Further studies are needed to explore the role of the $8 q 24$ variants involved in the etiology of prostate cancer.

Keywords: 8q24, genetic variant, prostate cancer, susceptibility, meta-analysis

## INTRODUCTION

Prostate cancer ( PCa ) is the commonest non-cutaneous malignancy in men all over the world. Based on epidemiological and biological data, there is growing evidence that many influencing factors, including geography, ethnicity, genetic factors, and so on(Rebbeck, 2017), are associated with the risk of PCa. PCa exhibits high heritability, however, the exact etiology of PCa is still unknown. Identification of genetic factors regulating the susceptibility and progression of PCa contributes to improvement of preventive measures and therapeutic outcomes.

Multiple risk loci for prostate cancer have been identified by GWAS. In 2007, a two-stage GWAS from 1,854 prostate cancer patients and 1,894 population-screened controls was conducted. In this study, common loci at 8 q 24 were identified to be associated with prostate cancer (Eeles et al., 2008). It was proved that 8 q 24 region was associated with lots of cancers, including breast (Pereira et al., 2016), prostate (Hubbard et al., 2016), bladder (Kiltie, 2010), colon (Ling et al., 2013), lung (Zhang et al., 2012), gliomas (Rice et al., 2013), and so on. These susceptibility loci actually do not affect coding DNA, interestingly, these loci showed strong linkage disequilibrium (LD) as they often tightly linked with many SNPs. However, further study found that there are many enhancers in $8 q 24$ region, and the rs6983267-containing enhancer interacts with the MYC gene by binding with TCF7L2 (TCF4), and alter the sensitivity to WNT signaling (Tuupanen et al., 2009). Another recent study found that the rs378854-containing region can interact with the promoters of both MYC and MYC activator PVT1(Meyer et al., 2011). Based on the above compelling evidence, it was supposed that the 8 q 24 variants played important roles in prostate carcinogenesis.

Here we performed a comprehensive meta-analysis, involving a total of 60,293 cases and 62,971 controls, to evaluate all genetic studies that investigated associations between 15 variants in 8 q 24 and risk of prostate cancer.

## METHODS

## Search Strategy and Selection Criteria

We systematically searched PubMed and Embase to identify genetic association studies published in print or online before January 10th, 2018 in English language using key terms " 8 q 24 " and "polymorphism or variant or genotype" and "prostate carcinoma or prostate tumor or prostate cancer". Two investigators (Yu Tong and Tao Yu) independently assessed the eligibility of each study. All studies included in this meta-analysis must meet all the following inclusion criteria: (i) evaluating the associations of the 8 q 24 variants with prostate cancer risk; (ii) providing sufficient data or multivariate-adjusted risk estimates [e.g., odds ratios (ORs), hazard ratios (HRs), relative risks (RRs), $95 \%$ confidence intervals (CIs) or standard errors (SEs)] to calculate these estimates. The exclusion criteria were as follows: (i) insufficient data; (ii) they were published as letters to editors or conference abstracts; (iii) they were studies about cancer mortality.

## Data Extraction

Guidelines recommended were used to report meta-analyses of observational studies by an investigator (Yu Tong and Tao $\mathrm{Yu})$ to extract data. Extracted data efrom each eligible study included name of first author, study design, publication date, source population, ethnicity, sample size, variants, alleles, and genotype counts, Hardy-Weinberg equilibrium (HWE) among controls. Ethnicity was classified as Caucasian, African, Asian, or others such as Latinos and Hawaiians. In this meta-analysis, 46 eligible publications are available with sufficient data.

## Statistical Analysis and Assessment of Cumulative Evidence

For each study, the odds ratio (OR) was used as the metric of choice. Pooled odds ratios were computed by the fixed effects model and the random effects model based on heterogeneity estimates, according to Prof. Michael Borenstein's suggestion (Borenstein et al., 2010). Once an overall gene effect was confirmed, the genetic model-free approach suggested by Minelli et al. (2005) was used to estimate the genetic effects and mode of inheritance. Assessment of protection from bias also considered the magnitude of association. OR less than 1.15 implicated presence of bias, unless the association had been replicated prospectivelywith no evidence of publication bias by several studies, such as GWAS or GWAS meta-analysis from collaborative studies. Heterogeneity between studies was evaluated by Cochran's Q test and calculated $I^{2}$ statistic h . $I^{2}$-values $<25 \%, 25-50 \%$, and $>50 \%$ represent no or little heterogeneity, moderate heterogeneity, and large heterogeneity, respectively. Sensitivity analyses were conducted to examine if exclusion of first published study deviated from HWE in controls influence the significant association. Harbord's test was
performed to evaluate publication bias. Small study bias was calculated by egger's test. All analyses were conducted using Stata, version 14.0 (StataCorp, 2017), with the metan, metabias commands.

## RESULTS

## Eligible Studies

Our initial database search identified 268 potentially relevant studies. Based on a review of titles and abstracts, 85 articles were retained. The full text of these 85 articles was reviewed in detail, and 46 studies were eligible in this meta-analysis. The specific process for identifying eligible studies and inclusion and exclusion criteria are summarized in Figure 1.

## Allelic Associations

Of the 21 variants located in $8 \mathrm{q} 24,15$ were significantly associated with the risk of prostate cancer, including rs16901979, rs1447295, rs6983561, rs7000448, rs6983267, rs13254738, rs7017300, rs7837688, rs1016343, rs7008482, rs4242384, rs620861, rs10086908, DG8S737 Allele-8, and rs10090154. No significant associations were found between rs4242382, rs4645959, rs7837328, rs16901966, rs10505476, rs13281615 and prostate cancer (data not shown).

## rs16901979 C > A

Twenty-four studies were included (Table 1), and a significant association with prostate cancer risk was found ( $p=1.08 \times$ $10^{-12}$, random effect $O R=1.48,95 \% C I: 1.33,1.65 ; Q=141.34$, $p=0.00, I^{2}=83.7 \%$, Figure 2A). A similar pattern was observed for Africans $\left(p=1.26 \times 10^{-26}\right.$, random effect $O R=1.33,95 \%$ CI: 1.26, 1.40; $\left.Q=2.76, p=0.949, I^{2}=0.0 \%\right)$, Asians $(p=8.49$ $\times 10^{-5}$, random effect $O R=1.36,95 \% C I: 1.17,1.59 ; Q=12.31$, $p=0.031, I^{2}=59.4 \%$ ) and Caucasians ( $p=6.48 \times 10^{-6}$, random effect $O R=1.72,95 \% C I: 1.36,2.17 ; Q=50.60, p=0.00$, $\left.I^{2}=84.2 \%\right)$. No publication bias was found in the eligible studies (Harbord's test $p=0.757$, Table 2).

## rs1447295 C>A

Thirty-seven studies were included (Table 1), a significant association was found with the risk of prostate cancer $(p=3.20$ $\times 10^{-14}$, random effect $O R=1.29,95 \% C I: 1.21,1.37 ; Q=160.1$, $p=0.00, I^{2}=77.5 \%$, Figure 2B).Significant association was also found for Asians ( $p=2.08 \times 10^{-11}$, random effect $O R=1.41$, $\left.95 \% C I: 1.27,1.56 ; Q=7.77, p=0.354, I^{2}=9.9 \%\right)$ and Caucasians ( $p=2.52 \times 10^{-23}$, random effect $O R=1.41,95 \%$ $\left.C I: 1.31,1.50 ; Q=50.80, p=0.00, I^{2}=64.6 \%\right)$. However, no significant association was found for Africans ( $p=0.168$, random effect $O R=1.05,95 \% C I: 0.98,1.11 ; Q=9.68, p=0.289$, $I^{2}=17.3 \%$ ), No publication bias was found in the eligible studies (Harbord's test $p=0.587$, Table 2).

## rs6983561 A>C

Eleven studies were included (Table 1), a significant association was found with the risk of prostate cancer $(p=0.036$, random effect $O R=1.29,95 \% C I: 1.02,1.64 ; Q=128.51, p=0.00$, $I^{2}=92.2 \%$, Figure 2C). No significant association was found for


FIGURE 1 | Flow diagram of included and excluded studies.

Africans ( $p=0.269$, random effect $O R=1.17,95 \% C I: 0.88,1.56$; $\left.Q=21.67, p=0.000, I^{2}=86.2 \%\right)$ and Caucasians $(p=0.241$, random effect $O R=1.36,95 \% C I: 0.81,2.27 ; Q=105.31$, $\left.p=0.00, I^{2}=95.3 \%\right)$. No publication bias was found in the eligible studies (Harbord's test $p=0.977$, Table 2).

## rs7000448 C>T

Eight studies were included (Table 1), a significant association was found with the risk of prostate cancer ( $p=0.003$, random effect $O R=1.11,95 \% C I: 1.04,1.19 ; Q=9.41, p=0.152$, $I^{2}=36.2 \%$, Figure 2D). Further evaluation by ethnicity showed that significant association was found for Africans ( $p=2.92 \times$ $10^{-5}$, random effect $O R=1.21,95 \% C I: 1.11,1.32 ; Q=1.82$, $p=0.403, I^{2}=0.0 \%$ ) and Caucasians ( $p=0.018$, random effect $\left.O R=1.08,95 \% C I: 1.01,1.14 ; Q=3.18, p=0.37, I^{2}=5.6 \%\right)$. No publication bias was found in the eligible studies (Harbord's test $p=0.868$, Table 2).

## rs6983267 T>G

Twenty-eight were included (Table 1), and a significant association with risk of prostate cancer was found ( $p=0.003$, random effect $O R=1.15,95 \% C I: 1.05,1.25 ; Q=275.92$, $p=0.00, I^{2}=90.2 \%$, Figure 2E). A similar pattern was observed for Asians ( $p=0.003$, random effect $O R=1.13,95 \% C I: 1.04$, 1.22; $Q=4.35, p=0.501, I^{2}=0.0 \%$ ) and Caucasians ( $p=0.001$, random effect $O R=1.21,95 \% C I: 1.08,1.36 ; Q=189.54$, $\left.p=0.00, I^{2}=93.1 \%\right)$. No significant association was found for Africans ( $p=0.269$, random effect $O R=0.98,95 \% C I: 0.68$, 1.42; $Q=69.39, p=0.000, I^{2}=91.4 \%$ ). No publication bias was found in the eligible studies (Harbord's test $p=0.577$, Table 2).

## rs13254738 A>C

Six studies were included (Table 1), a significant association was found with the risk of prostate cancer $(p=0.026$, random effect $O R=1.11,95 \% C I: 1.01,1.22 ; Q=12.44, p=0.029, I^{2}=59.8 \%$, Figure 2F). Significant association was found for Caucasians ( $p=0.08$, random effect $O R=1.06,95 \% C I: 0.99,1.14 ; Q=2.52$, $\left.p=0.47, I^{2}=0.0 \%\right)$. No publication bias was found in the eligible studies (Harbord's test $p=0.599$, Table 2).

## rs7017300 A>C

Four studies were included, a significant association with prostate cancer risk was found ( $p=0.001$, random effect $O R=1.39,95 \%$ CI: 1.15, 1.68; $Q=17.93, p=0.000, I^{2}=83.3 \%$, Figure 2G). No publication bias was found in the eligible studies (Harbord's test $p=0.564$, Table 2).

## rs7837688 G>T

Eight studies were included (Table 1), a significant association was found with the risk of prostate cancer $\left(p=1.66 \times 10^{-10}\right.$, random effect $O R=1.51,95 \% C I: 1.33,1.72 ; Q=35.02$, $p=0.000, I^{2}=80.0 \%$, Figure 2H). Significant association was also found for Caucasians ( $p=3.64 \times 10^{-9}$, random effect $\left.O R=1.53,95 \% C I: 1.33,1.77 ; Q=26.07, p=0.000, I^{2}=80.8 \%\right)$. No publication bias was found in the eligible studies (Harbord's test $p=0.921$, Table 2).

## rs1016343 C>T

Six studies were included (Table 1), a significant association with risk of prostate cancer was found ( $p=8.25 \times 10^{-10}$, random effect $O R=1.37,95 \% C I: 1.24,1.52 ; Q=20.42, p=0.001$, $I^{2}=75.5 \%$, Figure 2I). Significant association was also found for Caucasians ( $p=3.64 \times 10^{-9}$, random effect $O R=1.41,95 \%$

TABLE 1 | Characteristics of the included articles.

| Study, year | Study design | Country/region | Ethnicity | Variant | Cases/controls |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Geraldine Cancel-Tassin, 2015 (Cancel-Tassin et al., 2015) | Population-based case-control study | France | African | rs16901979 | 489/534 |
| Mian Li, 2011 (Li et al., 2011) | Case-control study | China | Asian | rs16901979 | 432/782 |
| Maurice P Zeegers, 2011 (Zeegers et al., 2011) | Cohort Study | Netherlands | Caucasian | rs1447295 | 281/267 |
| Marcelo Chen, 2010 (Chen et al., 2010) | Case-control study | China | Asian | rs16901979 | 331/335 |
|  |  |  |  | rs6983561 | 324/336 |
| Prodipto Pal, 2009 (Pal et al., 2009) | Case-control study | USA | Caucasian | rs16901979 | 596/567 |
|  |  |  |  | rs1447295 |  |
|  |  |  |  | rs6983267 |  |
|  |  |  |  | rs4645959 |  |
|  |  |  |  | rs1016343 |  |
| Marcelo Chen, 2009 (Chen et al., 2009) | Hospital-based case-control study | China | Asian | rs1447295 | 340/337 |
| Andreas Meyer, 2009 (Meyer et al., 2009) | Hospital-based case-control study | Germany | Caucasian | rs1447295 | 486/462 |
|  |  |  |  | rs13281615 | 488/462 |
| Iona Cheng, 2008 (Cheng et al., 2008) | Case-control study | USA | Caucasian | rs16901979 | 417/416 |
|  |  |  | African |  | 89/87 |
|  |  |  |  | rs1447295 | 417/417 |
|  |  |  |  |  | 89/89 |
|  |  |  |  | DG8S737 | 416/417 |
|  |  |  |  |  | 89/89 |
|  |  |  |  | rs6983561 | 417/417 |
|  |  |  |  |  | 88/89 |
|  |  |  |  | rs10090154 | 417/414 |
|  |  |  |  |  | 89/88 |
|  |  |  |  | rs7000448 | 416/417 |
|  |  |  |  |  | 89/89 |
|  |  |  |  | rs6983267 | 417/417 |
|  |  |  |  |  | 89/89 |
|  |  |  |  | rs13254738 | 506/506 |
|  |  |  |  |  | 89/88 |
| Christiane Robbins, 2007 (Robbins et al., 2007) | Case-control study | USA | African | rs16901979 | 490/567 |
|  |  |  |  | rs1447295 |  |
|  |  |  |  | DG8S737 |  |
|  |  |  |  | rs6983267 |  |
|  |  |  |  | rs7008482 |  |
| Miia Suuriniemi, 2007 (Suuriniemi et al., 2007) | Population-based case-control study | USA | Caucasian | rs1447295 | 582/538 |
| Fredrick R. Schumacher, 2007 <br> (Schumacher et al., 2007) | Nested case-control study | Multiple countries | Caucasian | rs1447295 | 5505/6270 |
|  |  |  | African |  | 676/643 |
| Julius Gudmundsson, 2007 (Gudmundsson et al., 2007) | Case-control study | Iceland | Caucasian | rs16901979 | 2663/5509 |
|  |  |  | African |  | 373/372 |
|  |  |  | Caucasian | rs1447295 |  |
|  |  |  | African |  |  |
| Gianluca Severi, 2007 (Severi et al., 2007) | Case-control study | Australia | Caucasian | rs1447295 | 821/732 |
| Dominika Wokołorczyk, 2008 (Wokolorczyk et al., 2008) | Case-control study | Poland | Caucasian | rs6983267 | 1910/1885 |
| S. Lilly Zheng, 2007 (Zheng et al., 2007) | Case-control study | USA | Caucasian | rs16901979 | 1563/576 |

(Continued)

TABLE 1 | Continued

| Study, year | Study design | Country/region | Ethnicity |
| :--- | :--- | :--- | :--- |

(Continued)

TABLE 1 | Continued

| Study, year | Study design | Country/region | Ethnicity |
| :--- | :--- | :--- | :--- |
|  |  | Variant |  |
| Jielin Sun, 2008 (Sun et al., 2008) |  |  | rs620861 |

TABLE 1 | Continued

| Study, year | Study design | Country/region | Ethnicity |
| :--- | :--- | :--- | :--- |
| Fang Liu, 2011 (Liu et al., 2011) | Case-control study | China | Asian |

TABLE 1 | Continued

| Study, year | Study design | Country/region | Ethnicity | Variant |
| :--- | :--- | :--- | :--- | :--- |
| Joke Beuten, 2009 (Beuten et al., 2009) | Cohort Study | USA | Caucasian <br> hispanic | rs10505476 |

CI: 1.32, 1.50; $\left.Q=0.76, p=0.859, I^{2}=0.0 \%\right)$. No publication bias was found in the eligible studies (Harbord's test $p=0.922$, Table 2).

## rs7008482 G > T

Four studies were included (Table 1), a significant association was found with the risk of prostate cancer ( $p=0.021$, random effect $O R=0.77,95 \%$ CI: $0.62,0.96 ; Q=6.49, p=0.039$, $I^{2}=69.2 \%$, Figure 2J). No publication bias was found in the eligible studies (Harbord's test $p=0.549$, Table 2).

## rs4242384 A>C

Three studies were included (Table 1), a significant association with prostate cancer risk was found ( $p=0.022$, random effect $O R=1.42,95 \% C I: 1.02,1.92 ; Q=10.71, p=0.005, I^{2}=81.3 \%$,
Figure 2K). No publication bias was found in the eligible studies (Harbord's test $p=0.376$, Table 2).

## rs620861 G > A

Six studies were included (Table 1), a significant association was found with the risk of prostate cancer $\left(p=3.57 \times 10^{-4}\right.$, random effect $O R=0.86,95 \% C I: 0.79,0.94 ; Q=19.28, p=0.002$, $I^{2}=74.1 \%$, Figure 2L). Significant association was also found for Caucasians ( $p=3.64 \times 10^{-9}$, random effect $O R=0.84,95 \% C I$ : 0.77, 0.91; $\left.Q=13.34, p=0.004, I^{2}=77.5 \%\right)$. No publication bias was found in the eligible studies (Harbord's test $p=0.791$, Table 2).

## rs10086908 T>C

Five studies were included (Table 1), a significant association was found with the risk of prostate cancer $\left(p=3.57 \times 10^{-4}\right.$, random effect $O R=0.73,95 \% C I: 0.60,0.88 ; Q=37.54, p=0.000$, $I^{2}=89.3 \%$, Figure 2M). Significant association was also found for Caucasians ( $p=0.036$, random effect $O R=0.70,95 \% C I$ : $\left.0.50,1.00 ; Q=37.13, p=0.004, I^{2}=94.6 \%\right)$. No publication bias was found in the eligible studies (Harbord's test $p=0.339$, Table 2).

## DG8S737 Allele-8 Absent>Present

Five studies were included (Table 1), a significant association with risk of prostate cancer was found ( $p=3.06 \times 10^{-4}$, random effect $O R=1.29,95 \% C I: 1.12,1.47 ; Q=2.32, p=0.803$, $I^{2}=0.0 \%$, Figure 2N). A similar pattern was observed for Caucasians ( $p=0.005$, random effect $O R=1.33,95 \%$ CI: 1.09, 1.62; $\left.Q=1.91, p=0.386, I^{2}=0.0 \%\right)$. No publication bias was found in the eligible studies (Harbord's test $p=0.592$, Table 2).

## rs10090154 C>T

Nine studies were included (Table 1), a significant association was found with the risk of prostate cancer $\left(p=2.04 \times 10^{-5}\right.$, random effect $O R=1.33,95 \% C I: 1.17,1.52 ; Q=0.70, p=0.873$, $I^{2}=0.0 \%$, Figure 2O). A similar pattern was observed for Caucasians ( $p=3.63 \times 10^{-5}$, random effect $O R=1.33,95 \%$ $\left.C I: 1.16,1.52 ; Q=0.70, p=0.705, I^{2}=0.0 \%\right)$. No publication bias was found in the eligible studies (Harbord's test $p=0.641$, Table 2).

## GENOTYPE COMPARISON <br> rs16901979 C > A

Of the 24 studies, nine reported genotype information. The effects of genotype for AA vs. CC (OR1) and CA vs. CC (OR2) were calculated. Multivariate meta-analysis was conducted to estimate the pooled risk (Table 2). Individuals with the homozygous AA genotype ( $p=3.86 \times 10^{-9}$, random effect OR1 $\left.=1.71,95 \% C I: 1.43,2.04 ; Q=7.48, p=0.486, I^{2}=0.0 \%\right)$ and heterozygous CA genotype ( $p=3.06 \times 10^{-4}$, random effect $\left.O R 2=1.36,95 \% C I: 1.15,1.61 ; Q=14.29, p=0.074, I^{2}=44.0 \%\right)$ have increased risk of prostate cancer.

## rs1447295 C>A

Of the 38 studies, 19 reported genotype information. The effects of genotype for AA vs. CC (OR1) and CA vs. CC (OR2) were calculated for each study (Table 2). Individuals with the homozygous AA genotype ( $p=0.006$, random effect $O R 1=1.42$,


FIGURE 2 | Forest plots for associations between selected variants in the 8q24 region and prostate cancer risk. Associations of rs16901979 (A), rs1447295 (B) rs6983561 (C), rs7000448 (D), rs6983267 (E), rs13254738 (F), rs7017300 (G), rs7837688 (H), rs1016343 (I), rs7008482 (J), rs4242384 (K), rs620861 (L), rs10086908 (M), DG8S737 Allele-8 (N), and rs10090154 (O) with prostate cancer risk.
TABLE 2 | Details of genetic variants significantly associated with cancer risk in meta-analyses.

| Variants | Cancer risk |  | Initial study influence |  | Deviation from HWE | $p$-value for publication bias | $p$-value for small study bias | Genotype cancer risk |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | OR (95\% CI) | $p$-value | OR (95\% CI) | $p$-value |  |  |  | OR1 (95\% CI) | $p$-value | OR2 (95\% CI) | $p$-value |
| rs16901979 | 1.48 (1.26-1.40) | $1.08 \times 10^{-12}$ | 1.49(1.33-1.66) | $1.67 \times 10^{-12}$ | No | 0.757 | 0.757 | 1.72(1.44-2.05) | $1.97 \times 10^{-9}$ | 1.36(1.15-1.61) | $3.06 \times 10^{-4}$ |
| rs1447295 | 1.29 (1.21-1.37) | $3.20 \times 10^{-14}$ | 1.30 (1.21-1.39) | $9.94 \times 10^{-15}$ | No | 0.559 | 0.664 | 1.42(1.10-1.82) | 0.006 | 1.31(1.18-1.45) | $3.06 \times 10^{-7}$ |
| rs6983561 | 1.29 (1.02-1.64) | 0.036 | 1.29 (1.00-1.66) | 0.048 | No | 0.977 | 0.887 | 0.84(0.62-1.13) | 0.242 | 1.54(1.29-1.83) | $1.84 \times 10^{-6}$ |
| rs7000448 | 1.11(1.04-1.19) | 0.003 | 1.11(1.03-1.20) | 0.004 | No | 0.868 | 0.889 | 0.98(0.80-1.21) | 0.867 | 1.04(0.90-1.20) | 0.64 |
| rs13254738 | 1.11(1.01-1.22) | 0.026 | 1.13(1.04-1.23) | 0.005 | No | 0.599 | 0.601 | 1.19(0.85-1.68) | 0.312 | 1.04(0.94-1.16) | 0.458 |
| rs6983267 | 1.15(1.05-1.25) | 0.003 | 1.14(1.04-1.25) | 0.006 | No | 0.577 | 0.583 | 1.31(0.92-1.86) | 0.134 | 1.05(0.5-1.22) | 0.546 |
| rs7017300 | 1.39(1.15-1.68) | 0.001 | 1.37(1.08-1.75) | 0.009 | No | 0.564 | 0.531 |  |  |  |  |
| rs7837688 | 1.51(1.33-1.72) | $1.66 \times 10^{-10}$ | 1.49(1.30-1.70) | $1.20 \times 10^{-8}$ | No | 0.921 | 0.816 |  |  |  |  |
| rs1016343 | 1.37(1.24-1.52) | $8.25 \times 10^{-10}$ | 1.36(1.20-1.54) | $1.37 \times 10^{-6}$ | No | 0.922 | 0.895 |  |  |  |  |
| rs7008482 | 0.77(0.62-0.96) | 0.021 | 0.86(0.77-0.96) | 0.008 | No | 0.549 | 0.533 |  |  |  |  |
| rs4242384 | 1.42(1.05-1.92) | 0.022 | 1.22(1.01-1.48) | 0.044 | No | 0.376 | 0.340 |  |  |  |  |
| rs620861 | 0.86(0.79-0.94) | $3.57 \times 10^{-4}$ | 0.89(0.81-0.97) | 0.007 | No | 0.791 | 0.795 |  |  |  |  |
| rs10086908 | 0.73(0.60-0.88) | 0.001 | $0.81(0.76-0.86)$ | $1.66 \times 10^{-10}$ | No | 0.339 | 0.428 |  |  |  |  |
| DG8S737-8 allele | 1.29 (1.12-1.47) | $3.06 \times 10^{-4}$ | 1.29 (1.09-1.54) | 0.004 | No | 0.592 | 0.648 | 0.83(0.29-2.38) | 0.733 | 1.25(0.98-1.59) | 0.068 |
| rs10090154 | 1.33 (1.17-1.52) | $2.04 \times 10^{-5}$ | 1.33(1.16-1.52) | $3.63 \times 10^{-5}$ | No | 0.641 | 0.668 | 1.34(0.82-2.19) | 0.245 | 1.40(1.2-1.62) | $1.24 \times 10^{-5}$ |

95\% CI: 1.10, 1.82; $\left.Q=33.56, p=0.010, I^{2}=49.3 \%\right)$ and heterozygous CA genotype ( $p=3.06 \times 10^{-7}$, random effect OR2 $=1.31,95 \% C I: 1.18,1.45 ; Q=38.05, p=0.002, I^{2}=55.3 \%$ ) have increased risk of prostate cancer.

## rs6983561 A>C

Of the 11 studies, five reported genotype information. The genotype effects for CC vs. AA (OR1) and AC vs. AA (OR2) were calculated for each study (Table 2). There was a significantly increased risk of prostate cancer among individuals with heterozygous AC genotype ( $p=1.84 \times 10^{-6}$, random effect $\left.O R 2=1.54,95 \% C I: 1.29,1.83 ; Q=4.10, p=0.393, I^{2}=2.4 \%\right)$. However, no significant association was found among individuals with the homozygous CC genotype.

## rs10090154 C>T

Of the 9 studies, four reported genotype information. The effects of genotype for TT vs. CC (OR1) and CT vs. CC (OR2) were calculated for each study (Table 2). Individuals with heterozygous CT genotype ( $p=1.24 \times 10^{-5}$, random effect $\left.O R 2=1.40,95 \% C I: 1.20,1.62 ; Q=1.58, p=0.663, I^{2}=0.0 \%\right)$ have an increased risk of prostate cancer. However, no significant association was found among individuals with the homozygous TT genotype.

## SENSITIVITY ANALYSIS

Results of sensitivity analysis showed that the obtained results of 8 q 24 variants and risk of prostate cancer were robust statistically and no individual study affected the pooled OR significantly (Table 2).

## DISCUSSION

To our knowledge, this study is the most comprehensive and largest evaluation of publications on associations between 8 q 24 variants and PCa risk. Preliminary meta-analyses mostly focused on the association between single or less SNPs with prostate cancer. From 46 eligible articles including 60,293 cases and 62,971 controls, we performed meta-analysis to evaluate associations between 15 variants in 8 q 24 region and PCa risk. Our study here provides an update of the previous reports. In addition, more variants were evaluated that have not been analyzed by meta-analyses previously.

Of the 21 variants located in $8 \mathrm{q} 24,15$ were associated with prostate cancer risk significantly. Our primary analysis shows that, the rs16901979 ( $p=1.08 \times 10^{-12}, O R=1.48$ ), rs1447295 $\left(p=4.51 \times 10^{-15}, O R=1.29\right)$, rs6983561 $(p=0.036$, $O R=1.29)$, rs7000448 $(p=0.003, O R=1.11)$, rs6983267 ( $p=0.003, O R=1.15$ ), rs13254738 ( $p=0.026, O R=1.11$ ), rs7017300 ( $p=0.001$, OR $=1.39$ ), rs7837688 $(p=1.66 \times$ $\left.10^{-10}, O R=1.51\right)$, rs1016343 $\left(p=8.25 \times 10^{-10}, O R=1.37\right)$, rs7008482 ( $p=0.021$, OR $=0.77$ ), rs4242384 $(p=0.022$, $O R=1.42)$, rs620861 $\left(p=3.57 \times 10^{-4}, O R=0.86\right)$, rs10086908 $\left(p=3.57 \times 10^{-4}, O R=0.73\right)$, DG8S737 Allele$8\left(p=3.06 \times 10^{-4}, O R=1.29\right), \operatorname{rs} 10090154(p=2.04 \times$
$10^{-5}, O R=1.33$ ) were significantly associated with PCa risk. In particular, both homozygous AA $\left(p=3.86 \times 10^{-9}\right.$, OR1 $=1.71$ ) and heterozygous CA $\left(p=3.06 \times 10^{-4}, O R 2=1.36\right)$ genotypes of rs16901979, as well as the AA $(p=0.005, O R 1=1.41)$ and CA $\left(p=2.14 \times 10^{-8}, O R 2=1.33\right)$ genotypes of rs1447295, were associated with PCa risk. Heterozygous AC genotype ( $p=1.84$ $\times 10^{-7}$, OR2 $=1.54$ ) of rs6983561, CT genotype ( $p=1.24 \times$ $10^{-5}, O R 2=1.40$ ) of rs 10090154 were also found to be associated with the risk of PCa . Our findings were robust in regard to study design and sensitivity analyses according to several gene-variants-association studies and thousands of participants. No evidence of small study bias or publication bias was found.

The 8 q 24 region is dense with SNP (single-nucleotidepolymorphism) associated with risk for prostate, colorectal, breast cancer, et al. There are about five separated different cancer susceptibility loci specific for different cancers within the 8 q 24 "desert" (Huppi et al., 2012). Region 1, including rs16901979, rs13254738 and rs6983561, region 4, including rs7000448 and region 5, including rs1447295 specifically associated with the PCa risk, rs13281615 in region 2 is a breast-specific cancer susceptibility loci, rs10505477 and rs10808556 in a same block in region 3 were confirmed to be associated with colorectal cancer(Ghoussaini et al., 2008). Although the exact biological mechanisms underlying these associations with multiple cancers are confusing, these variants might affect tissue-specific enhancers of one or more genes involved in carcinogenesis. FAM84B, very closest to 8 q 24 , is reported that, during prostate tumorigenesis and follows PCa progression, its

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expression increased (Wong et al., 2017). Another pseudogene of POU5F1P1/POU5F1B, located in $8 q 24.21$ region, was also observed that levels of both the mRNA and protein increased in PCa (Kastler et al., 2010). Therefore, variants in 8 q 24 region themselves or with other variants might be responsible for the associations with prostate cancer.

Our study provides summary evidence that common 15 variants in the 8 q 24 region are associated with PCa risk. To explore the exact mechanisms of 8 q 24 variants involved in parthenogenesis of prostate cancer needs further functional studies.

## AUTHOR CONTRIBUTIONS

Data were extracted by YT and TY. SL, FZ, and JY analyzed the data. YQ and DM wrote the manuscript.

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Conflict of Interest Statement: The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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