



## Seabird Trophic Position Across Three Ocean Regions Tracks Ecosystem Differences

Tyler O. Gagné<sup>1\*</sup>, K. David Hyrenbach<sup>2</sup>, Molly E. Hagemann<sup>3</sup>, Oron L. Bass<sup>4</sup>, Stuart L. Pimm<sup>5</sup>, Mark MacDonald<sup>6</sup>, Brian Peck<sup>7</sup> and Kyle S. Van Houtan<sup>1,5</sup>

<sup>1</sup> Conservation and Research Department, Monterey Bay Aquarium, Monterey, CA, United States, <sup>2</sup> Department of Natural Science, Marine Science Programs at Oceanic Institute, Hawaii Pacific University, Honolulu, HI, United States, <sup>3</sup> Vertebrate Zoology Collections, Bernice Pauahi Bishop Museum, Honolulu, HI, United States, <sup>4</sup> South Florida Natural Resources Center, Everglades National Park, Homestead, FL, United States, <sup>5</sup> Nicholas School of the Environment, Duke University, Durham, NC, United States, <sup>6</sup> Department of Marine and Wildlife Resources, Pago Pago, AS, United States, <sup>7</sup> Rose Atoll Marine National Monument, United States Fish and Wildlife Service, Pago Pago, AS, United States

We analyze recently collected feather tissues from two species of seabirds, the sooty tern (*Onychoprion fuscatus*) and brown noddy (*Anous stolidus*), in three ocean regions (North Atlantic, North Pacific, and South Pacific) with different human impacts. The species are similar morphologically and in the trophic levels from which they feed within each location. In contrast, we detect reliable differences in trophic position amongst the regions. Trophic position appears to decline as the intensity of commercial fishing increases, and is at its lowest in the Caribbean. The spatial gradient in trophic position we document in these regions exceeds those detected over specimens from the last 130 years in the Hawaiian Islands. Modeling suggests that climate velocity and human impacts on fish populations strongly align with these differences.

Keywords: trophic ecology, commercial fisheries, ocean memory, global change, machine learning, stable isotopes, food webs

#### INTRODUCTION

The scale and accessibility of marine ecosystems, anthropogenic impact, and industry removal presents management and conservation challenges. Consequently, monitoring and evaluation often rely on proxies to assess marine ecosystems. Such proxies include manually surveyed biological indicators, remotely sensed environmental data, and state-aggregated fisheries dependent metrics like mean trophic level (MTL) (Pauly et al., 1998; Holt and Miller, 2011; Hunsicker et al., 2016). For these to be effective decision-making tools, identifying indicators and metrics that accurately and independently quantify impacts is key (Link et al., 2009).

Food web status has been a widely used metric for gauging ecosystem state, especially in relation to commercial fishery impacts (Pimm, 1982). In particular, fishery-derived MTL is repeatedly used to gauge food web status and trends; though its interpretation has been contested (Pauly et al., 1998). Fishery-dependent MTL assessments have inherent biases that may fail to detect multiple ecosystem signals such as high trophic-level species loss, prey release, and market targeting shifts (Essington et al., 2006; Estes et al., 2011). Ideal biological indicators are those sampled and approximated from species representing broad ecosystem patterns (Lyday et al., 2015; Reed et al., 2016). The techniques must be robust and reproducible at scale (Gagné et al., 2018). Increasingly, metrics independent of fisheries and often involving non-target upper trophic level predators, like

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> \*Correspondence: Tyler O. Gagné tgagne@mbayaq.org

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marine birds and mammals, offer promising approaches for circumventing biases (Holt and Miller, 2011).

To assess the status of pelagic food webs, we use compoundspecific stable isotope analysis of amino acids (CSIA-AA) to compare seabird tissues across three ocean regions. This approach helps resolve previous questions of bias from fisherydata based metrics and from those relying on bulk stable isotope techniques which have come under scrutiny due to isotopic baseline shifts affecting trophic position estimates (Nielsen et al., 2015).

We focus on two species of colonial terns that are central place foragers, have large foraging ranges (BRNO = 200 km, SOTE = 800 km), and are circumtropically distributed (Hebshi et al., 2008). CSIA-AA determines trophic position robustly by comparing the relative enrichment of the <sup>15</sup>N to <sup>14</sup>N ratio  $(\delta^{15}N)$  in trophic and source amino acids (Nielsen et al., 2015). This approach proposes a fishery-independent metric to document food web status amongst locations. It builds upon our previous research, that applied this method across a chronology of 130-years (Gagné et al., 2018) within one region. To understand potential spatial patterns that the birds may be detecting, we use random forest regression models to investigate the influence of various anthropogenic and climate factors on seabird trophic position. We reaffirm seabirds are reliable indicators of marine systems, especially through the use of CSIA-AA to calculate their trophic position (Lyday et al., 2015), and in recording impacts from commercial fisheries extraction.

#### MATERIALS AND METHODS

#### **Specimen Collection and Preparation**

We sampled brown noddy (*Anous stolidus*) and sooty tern (*Onychoprion fuscatus*) feathers from three locations in three distinct ocean basins: Rose Island, Rose Atoll National Wildlife Refuge, American Samoa (South Pacific); Waimanalo, Oahu, Hawaii (North Pacific); and Bush Key, Dry Tortugas National Park, Florida (North Atlantic). From 2013–2015, we collected senesced, fully emerged flight feathers at nesting colonies and additional samples from dead strandings (USFWS permits MB052060-0, MB180283-1).

Collected feathers were free from debris and other tissues, were stored in heavyweight polyethylene bags (ULINE<sup>TM</sup>, 4 mil) with indicating silica gel desiccant (Fisher<sup>TM</sup> grade 48, 4–10 mesh), and later debrided with compressed air. We homogenized individual feathers (n = 20) and sent samples to the UC Davis Stable Isotope Facility for CSIA-AA. Though our processed specimen data are limited (n = 20, 6 Hawaii, 4 Florida, 10 American Samoa with an even split of species by location), the sample size we collected is characteristic of CSIA-AA studies of marine species (Popp et al., 2007; Chikaraishi et al., 2009; Votier et al., 2010; O'Malley et al., 2012; Ostrom et al., 2017; Peavey et al., 2017) and effective at describing possible population patterns. That said, we discuss future research directions that may improve upon the sampling and analyses presented here.

## **Trophic Position Calculation**

We calculated trophic position, TP, using:

$$TP = \frac{(\delta^{15}N_{Trp} - \delta^{15}N_{Src} - \beta)}{TEF} + 1$$

where  $\delta^{15}N_{Trp}$  is the mean value for six trophic amino acids (alanine, glutamic acid, isoleucine, leucine, proline, and valine),  $\delta^{15}N_{Src}$  is the single source amino acid (phenylalanine), and  $\beta$ (2.42) and *TEF* (trophic enrichment factor = 5.63) are amino acid-specific constants (Nielsen et al., 2015), calculated using established methods (Gagné et al., 2018). Variation of the selected constants does not impact the relative relationships of TP between species or location, only absolute values, and therefore does not influence our aim of detecting differences between regions. For each specimen, we generated 1000 random TP values from the  $\delta^{15}N$  (normal) parameter distributions the analytical lab provided for each amino acid (Figure 1A and Supplementary Table S1). Due to the large (20000 estimates) sample drawn the lab parametrized distributions, tests of difference are highly significant (p < 0.005) with narrow confidence intervals. Therefore, we report and discuss trophic positions effect size and magnitude differences (Sullivan and Feinn, 2012; Greenland et al., 2016).

#### **Model Input Aggregation**

We aggregated covariates tied to environmental and human drivers that may describe spatial differences. As seabird prey items vary between regions, we represented fishing pressure for each region with catch density (catch-per-unit area, or CPUA, measured in tons  $km^{-1}$ ) as per previous studies (Pauly, 2007; Pauly et al., 2013). This metric is calculated by dividing the cumulative total annual catch from the Sea Around Us reconstructed landings by the total area of the EEZ of the location from which the feathers were sampled. We also obtained MTL by EEZ from the SeaAroundUs database. To capture the nature of the potentially relevant climate velocities that may have precipitated the relevant seabird prey composition, we calculated the 10-years change during and immediately prior to the sampling period. We calculated this as the rate of change in remotely sensed sea surface temperature (Supplementary Figures S1A,B, AVHRR POES 0.1°, 14-day composite) during the period of 2006-2016, consistent with previous studies (Hamann et al., 2015; Van Houtan et al., 2015).

#### **Model Development**

To explore correlative relationships, we implemented the random forest algorithm to model trophic position as a function of the drivers. Random forest models excel in their flexibility to model non-linearity and complex interactions while also maximizing generalizability, minimized overfitting tendency, and easy interpretation (Breiman, 2001). We trained random forest models with conservative hyper-parametrization. Hyperparameters were set at 500 trees, 2 variables tried at each split, a minimum node size of 5 observations, and an out-of-bag sample proportion of 0.4 with replacement (Breiman, 2001). We built centered individual conditional expectation plots to highlight



colors correspond to labeled wing silhouettes, "SOTE" is sooty tern and "BRNO" is brown noddy. **B-G** are conditional repetation plots of predictors from the model for continuous (**B-D**) and categorical model inputs (**E-G**). Relative TP on the y-axis of **B-G** refers to the change in trophic position units observed across the range of values of a predictor. Colored lines (**B-D**) and center quantile (**E-G**) represent the respective mean and median of the partial effect of a covariate on the model prediction of TP. Gray bands and quartiles represent the range and distribution of predictions, akin to a prediction interval. TP declines with increases in catch density (**B**) and rates of SST increase (**D**), and declines as Sea Around Us reported catch landings MTL decreases (**C**). We assessed model variable importance with a leave one out sensitivity analysis. Panel (**H**) shows the proportion of model runs where each covariate appeared at each importance rank. Beyond the splitting power of species and geography, fisheries related factors rank highly.

the range of predictor-response relationships conditional on the range of observations of the predictor values (Goldstein et al., 2015).

#### **Sensitivity Analysis**

We conducted a sensitivity analysis on variable importance rankings using a leave-one-out randomization technique for modeling fitting to measure sample stability on model inferences (**Figure 1H**; Archer and Kimes, 2008). In this process, we built 500 sets of variable importance values, based on 500 models each run from random subsets of the full sample set. As opposed to one set of variable importance rankings, this established a distribution of variable rankings under both sample perturbations and model fitting stochasticity that is inherent to random forest. This analysis provides a means to account for the influence of individual specimens given our relatively limited sample size.

#### RESULTS

The sooty tern and brown noddy share a similar ecomorphology (Gagné et al., 2018). **Figure 1A** shows comparable trophic positions within species that decline amongst locations from American Samoa (mean = 4.00) to Hawaii (3.68) to Florida (3.59). Additionally, unlike bulk stable isotope analysis, CSIA-AA accounts for nitrogen from varied food web bases. Therefore, the observed shifts in trophic position suggest divergence in the

trophic positions of available prey, and food web compression or simplification.

To model potential regional patterns, we use random forest models to investigate the influence of anthropogenic and climate factors on trophic position. We aggregated covariates tied to environmental and anthropogenic drivers that describe spatial differences in fisheries-dependent MTL, fisheries landings, and coupled sea surface temperature rate of change. Our models suggest that commercial fishing and climate align with the food web differences between regions. **Figures 1B–H** shows the model relationships and the estimated variable importance.

The random forest model explained 57.5% of the variance in trophic position, with a root mean square error of 0.17. The three continuous variable inputs show well-defined trends. Trophic position declines with increasing catch (Figure 1B), increasing rates of regional sea surface temperature change (Figure 1D), and as fishery-dependent MTL declines (Figure 1C). The categorical variables (Figures 1E-G, location, basin, and species) all offer strong and comparable splitting power (Figure 1H), though little insight into the changing conditions effecting TP. In the leaveone-out cross validation (Figure 1H), root mean square error on the out-of-test sets averaged 0.18 with a 95% quantile range of 0.098-0.284. This range does highlight some sensitivity to stability of the specimen dataset, which may be improved in future studies with additional specimens. However, due to field logistics and laboratory costs, studies using CSIA-AA for similar purposes often employ a similar number of replicates as we have here (see methods). Importantly, however, our sensitivity

analysis reveals the modeled signal is distinct from the sample variance.

#### DISCUSSION

# Seabirds as an Alternative Trophic Indicator

Does higher seabird trophic position describe a less-impacted ecosystem state? Birds at the Western Atlantic site had a significantly lower trophic position in both species we measured (Figure 1A). Our model indicates lower trophic positions at this site are consistent with chronically high fishing pressure, seasurface warming, and low MTL of fisheries catch. Conversely, our South Pacific site had the highest measured trophic position, paired with low fishing pressure, a recent cooling trend, and high MTL in fisheries catch. Future studies that compare such questions across more extensive geographies and perhaps timelines may be able to decouple model factors, and better reveal the individual effects of climate and fisheries on MTL. Such an approach may provide testing beyond the hypotheses we generated observationally and refine some of the low-resolution signals we detected. Nonetheless, in accord with recent reviews (Estes et al., 2011), we have documented the trophic downgrading of two seabird species across a large spatial gradient and described possible contributing factors.

#### Trophic Position Dissimilarity in Space Appears to Reflect Declines Across Time

The regional differences we describe are consistent with a recent 130-years analysis of trophic position in seabirds from the Hawaiian Archipelago. In that study (Gagné et al., 2018), the ensemble trophic position for eight seabird species declined an average 0.32 units from 1891-2015. (Sooty tern declined by 0.30 and brown noddy by 0.09.) As trophic position may be a reliable metric to track ecosystem changes in one ecosystem over time, it may also reflect food web status across spatially distinct ecosystems. Given this, the contemporary trophic position decline we observe when comparing Florida to American Samoa (0.41) and Hawaii to American Samoa (0.33) is larger than the 130-years decline observed in Hawaii alone (Gagné et al., 2018). Compared more directly, however, the spatial gradient in trophic position between Florida and American Samoa equates to 161 years of change in Hawaii described by Gagné et al. (2018) in their seabird ensemble (-0.258 change in trophic position per century). When we constrain the benchmark to just the sooty tern and brown noddy data from the Hawaii study (-0.157 change per 100 years) the metric jumps to 264 years of change. If we make similar comparisons of our spatial gradient between Hawaii and Florida, those numbers are 34 and 56 years, respectively. Confirming previous studies (e.g., Fitzpatrick and Keegan, 2007; Jackson, 2008; Huettmann, 2012), this suggests historical overfishing and other anthropogenic impacts in Caribbean coastal ecosystems are significantly more advanced than in other tropical regions.

In sum, we reaffirm that seabirds are reliable indicators of marine systems, especially when we use of CSIA-AA to calculate their trophic position. Importantly, our findings align with the widely reported consensus that the Gulf of Mexico is a highly impacted ecosystem (Halpern et al., 2008). To utilize our approach while also improving upon our constrained sample, we suggest continued sampling of seabird feathers from museums and nesting colonies with negligible impacts to wild populations (strandings, fisheries interactions). Such a monitoring framework may better inform sampling and modeling needs for ecosystembased management and decision making. Furthermore, the development of a larger sample with our approach will better elucidate the nature of covariates and their interactions. Active movement tracking and spatially explicit fishery data, combined with the tissue sampling approach here, may further inform these patterns (Cherel et al., 2016). Ultimately, it is becoming clearer that seabird trophic position varies cross space and time and that it reliably informs food web status, and particularly the far-reaching impacts from commercial fisheries extractions.

## DATA AVAILABILITY STATEMENT

Datasets analyzed for this study are included in the data repository at: osf.io/4s9ty.

## ETHICS STATEMENT

This research was performed under USFWS permits #MB052060-0, MB180283-1 to KH. All research followed the Institutional Animal Care and Use Committee criteria under NOAA and HPU.

#### **AUTHOR CONTRIBUTIONS**

KVH and SP designed the study. All authors collected and prepared data. TG and KVH analyzed the data, prepared the figures, and wrote the manuscript. All authors reviewed the manuscript.

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#### SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fmars.2018. 00317/full#supplementary-material

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**Conflict of Interest Statement:** The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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