| 1 | Title: |
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| 2 | A hierarchical watershed model of fluid intelligence in childhood and adolescence |
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Abstract

Fluid intelligence is the capacity to solve novel problems in the absence of task-specific 13 knowledge, and is highly predictive of outcomes like educational attainment and 14 psychopathology. Here, we modelled the neurocognitive architecture of fluid intelligence in 15 two cohorts: CALM (N = 551, aged 5 - 17 years) and NKI-RS (N = 335, aged 6 - 17 years). We 16 used multivariate Structural Equation Modelling to test a preregistered watershed model of 17 18 fluid intelligence. This model predicts that white matter contributes to intermediate cognitive phenotypes, like working memory and processing speed, which, in turn, contribute to fluid 19 intelligence. We found that this model performed well for both samples and explained large 20 amounts of variance in fluid intelligence ($R^2_{CALM} = 51.2\%$, $R^2_{NKI-RS} = 78.3\%$). The relationship 21 between cognitive abilities and white matter differed with age, showing a dip in strength 22 around ages 7 - 12 years. This age-effect may reflect a reorganization of the neurocognitive 23 architecture around pre- and early puberty. Overall, these findings highlight that intelligence 24 25 is part of a complex hierarchical system of partially independent effects.

26

Keywords

27 Working memory, processing speed, fractional anisotropy, watershed model, structural28 equation modeling

Fluid intelligence (g_f) is a core part of human cognition and refers to the capacity to solve 29 novel problems in the absence of task-specific knowledge. It is highly predictive of a number 30 31 of important life span outcomes, including educational attainment (Primi et al. 2010; Roth et 32 al. 2015) and psychopathology (Gale et al. 2010). Despite years of investigation, however, our understanding of the neurocognitive architecture of $g_{\rm f}$ remains limited. Longstanding debates 33 have considered, for instance, how $g_{\rm f}$ relates to more fundamental cognitive functions such 34 35 as working memory and processing speed, and how all of these cognitive functions relate to brain structure and function (Kyllonen and Christal 1990; Fry and Hale 2000; Chuderski 2013; 36 37 Ferrer et al. 2013).

38 Working memory is the ability to hold and manipulate information in the mind short-term. It has been suggested that working memory is a key determinant of $g_{\rm f}$ by limiting mental 39 information processing capacity (Fukuda et al. 2010; Chuderski 2013). Proponents of this 40 working memory account of $g_{\rm f}$ cite high correlations between the two domains ranging from 41 0.5 to 0.9 in meta-analyses (Ackerman et al. 2005; Oberauer et al. 2005). Such high 42 correlations have led some to suggest that g_f and working memory are, in fact, isomorphic 43 (Kyllonen and Christal 1990). However, more recent work has highlighted that this 44 isomorphism only arises under conditions of high time constraints for g_f tasks (Chuderski 45 2013). This suggests that g_f and working memory are, in fact, separable constructs and 46 underlines the importance of processing speed for $g_{\rm f}$. 47

Processing speed, the speed of mental computations, is thought to be rate-limiting to g_f and is therefore sometimes proposed to be a particularly good predictor of g_f (Kail and Salthouse 1994; Salthouse 1996; Ferrer et al. 2013; Kail et al. 2015; Schubert et al. 2017). Proponents of the processing speed account of g_f cite moderate but robust correlations between g_f and

processing speed of 0.2 in meta-analyses (Sheppard and Vernon 2008) as well as longitudinal 52 53 evidence (Finkel et al. 2005; Coyle et al. 2011; Kail et al. 2015). Salthouse (1996) argued in the 54 context of cognitive aging, that processing speed determines high-level cognitive 55 performance because slow processing means that relevant sub-operations cannot be completed in a set amount of time or are not available for successful integration. A 56 complementary explanation of individual differences in $g_{\rm f}$ proposes that processing speed 57 58 may be a direct reflection of fundamental neuroarchitectonic properties of the brain, such as myelination or white matter microstructure (Lu et al. 2011; Chevalier et al. 2015). 59

60 White matter shows protracted development throughout childhood and adolescence, and 61 into the third decade of life (Mills et al. 2016). White matter tracts can be characterised in vivo using diffusion-tensor imaging (DTI), which is sensitive, but not necessarily specific, to 62 white matter microstructural properties such as myelination or axonal density (Jones et al. 63 2013; Wandell 2016). Fractional anisotropy (FA) is the most commonly investigated DTI 64 measure and quantifies the directionality of water diffusion in different white matter tracts 65 66 (Pfefferbaum et al. 2000; Wandell 2016). Working memory, processing speed and $g_{\rm f}$ have 67 each been linked to individual differences in FA (Vestergaard et al. 2011; Kievit, Davis, Griffiths, Correia, CamCAN, et al. 2016; Bathelt et al. 2018). While some studies, using 68 69 Principal Component Analysis, have posited that FA in different tracts can be summarized by 70 sizable single components (Penke et al. 2010; Cox et al. 2016), formal investigations using confirmatory factor analysis have demonstrated that single-factor models of FA generally 71 show poor fit and do not adequately capture individual differences in white matter 72 microstructure (Lövdén et al. 2013; Kievit, Davis, Griffiths, Correia, Cam-CAN, et al. 2016). In 73 a similar vein, there is a growing body of literature showing specific associations between 74 white matter tracts and cognitive abilities, with those connecting frontoparietal regions 75

vsually showing largest contributions to complex cognitive functions like $g_{\rm f}$ (Vestergaard et al. 2011; Kievit et al. 2016; Bathelt et al. 2018).

We here seek to address several critical outstanding issues in the field: First, there is limited 78 systematic evidence on the concurrent relationships between g_f, working memory, 79 processing speed and white matter. This leaves the relative contributions of processing speed 80 and working memory to $g_{\rm f}$ unclear, which, in turn, poses challenges for the design of effective 81 82 cognitive training interventions. Second, studies usually use a single task as a proxy for 83 complex and abstract constructs such as processing speed, working memory, and $g_{\rm f}$. This raises questions about the generalizability of findings (Noack et al. 2014). Third, our 84 understanding of how the relationships between relevant cognitive domains and between 85 brain and cognition change with age remains limited, raising the possibility that brain-86 87 behaviour relationships may change with age (Garrett 1946; Johnson 2000; Tamnes et al. 2017). 88

To address these issues, we here used structural equation modelling (SEM) to model the associations between g_f , working memory, processing speed, and white matter microstructure and age in two large, independent samples: the Centre for Attention, Leaning and Memory sample (CALM, N = 551, aged 5 - 17 years), which consists of children and adolescents referred to a clinic for having problems with attention, learning and memory (Holmes et al. 2018), and the Enhanced Nathan Kline Institute – Rockland Sample (NKI-RS, N =335, aged 6 - 17 years), a community-ascertained sample (Nooner et al. 2012).

To investigate the neurocognitive architecture of g_f in a principled way, we used a watershed model of individual differences. Based on the metaphor of a watershed, the model predicts a hierarchical many-to-one mapping of partially independent effects such that upstream

99 tributaries (e.g. brain structure) contribute to intermediate cognitive phenomena (cognitive 100 endophenotypes, e.g. working memory and processing speed), which then contribute to 101 downstream, complex cognitive phenomena such as g_f (Cannon and Keller 2006; Kievit, Davis, 102 Griffiths, Correia, CamCAN, et al. 2016). See Figure 1 for a representation of the model.

SEM, as a statistical technique, is uniquely suited to modeling the kinds of complex 103 multivariate brain-behavior associations posited by the watershed model (Kievit et al. 2011; 104 Kline 2015). SEM combines factor analysis and path analysis (a variant of regression analysis). 105 It can model abstract cognitive constructs like $g_{\rm f}$, by estimating latent variables from 106 107 observed task scores (i.e. manifest variables). This feature of SEM allowed us to model $g_{\rm f}$, 108 working memory, and processing speed in two independent samples, and thereby provided a direct test of the generalizability of our findings. Second, SEM can test the simultaneous 109 relations between multiple cognitive and neural variables, allowing us to address the relative 110 contributions of different white matter tracts and different cognitive endophenotypes to $g_{\rm f}$. 111 Finally, using SEM Trees (Brandmaier et al. 2013), a novel, decision-tree-based extension of 112 113 SEM, we investigated whether the associations in the watershed model change with age.

Based on the watershed model we made the following preregistered predictions(http://aspredicted.org/blind.php?x=u5pf6z):

116 1. Working memory, g_f and processing speed are separable constructs.

117 2. Individual differences in *g*_f are predicted by working memory and processing speed.

118 3. White matter microstructure is a multi-dimensional construct.

119 4. There is a hierarchical relationship between white matter microstructure, cognitive 120 endophenotypes (working memory and processing speed) and $g_{\rm f}$, such that white

121 matter contributes to working memory and processing speed, which, in turn 122 contribute to $g_{\rm f}$.

123

5. The contribution of working memory and processing speed to $g_{\rm f}$ changes with age.

124

Materials and Methods

125 Samples

We analysed data from the CALM and NKI-RS sample, as described in detail by (Holmes et al. 126 2018) and (Nooner et al. 2012) respectively. See also Simpson-Kent et al. (2019). We had also 127 preregistered to analyse data from the ABCD cohort (Volkow et al. 2018). The latter cohort 128 contains only data for 9 - and 10 - year olds at present, however, which limits comparability 129 130 to CALM and NKI-RS, and makes it unsuitable for investigations of developmental differences. We therefore opted to not analyse ABCD data here and instead recommend a replication of 131 the analyses presented here in ABCD once longitudinal data is available. The CALM sample 132 consists of children and adolescents referred by health and educational professionals as 133 having difficulties in attention, learning and/or memory. The NKI-RS is a community-134 135 ascertained, lifespan sample, and representative of the general population of Rockland, New York, and the United States as a whole, in terms of ethnicity, socioeconomic status etc. For 136 NKI-RS, we included data for participants under the age of 18 only to match the age range of 137 138 CALM and excluded data that were completed more than half a year after enrolment. The latter criterion was implemented to ensure that age at assessment did not differ 139 substantively between cognitive measures. The final samples included 551 participants from 140 141 CALM (30.85% female, aged 5.17 - 17.92 years, N_{Neuroimaging} = 165) and 335 participants from NKI-RS (43.48% female, aged 6.06 - 17.92 years, N_{Neuroimaging} = 67). See Table 1 for prevalence 142 of relevant disorders and learning difficulties in the samples. 143

| 144 | Table 1. Prevalence of Relevant Disorders and Learning Difficulties in the CALM and NKI-RS |
|-----|--|
| 145 | cohorts |

| Variable | Percentage | Percentage |
|--------------------------|------------|------------|
| | CALM | NKI-RS |
| ADHD | 31.94 | 17.01 |
| Dyslexia | 5.81 | 5.67 |
| Autism | 6.72 | 0.60 |
| Mood disorder | 0.54 | 0.90 |
| Anxiety disorder | 2.36 | 18.21 |
| Medicated ¹ | 10.53 | 17.01 |
| Speech/language problems | 38.11 | 19.40 |

146 *Note.*¹ unspecified medication for NKI-RS, ADHD-medication for CALM

147 Cognitive Tasks

We included cognitive tasks measuring the domains of g_f , working memory or processing speed for CALM and NKI-RS. See Table 2 for the complete list of tasks used, and the Supplementary Methods for task descriptions. Supplementary Figure 1 and 2 show raw scores on all tasks. The tasks modelled here were preregistered for CALM but not NKI-RS.

152 Table 2. Cognitive Tasks Modelled

| | CALM | NKI-RS |
|------------------|--|----------------------------|
| Working memory | AWMA Digit Recall (forward digit span) | WISC-R Forward Digit Span |
| | AWMA Backward Digit Span | WISC-R Backward Digit Span |
| | AWMA Dot Matrix | - |
| | AWMA Mr X | - |
| | - | CNB N-back task |
| <i>B</i> f | WASI-II Matrix Reasoning | WASI-II Matrix Reasoning |
| | - | WASI-II Block Design |
| | - | WASI-II Similarities |
| | - | CNB Verbal Reasoning |
| Processing speed | DKEFS Trail-Making | DKEFS Trail-Making |
| | PhAB Rapid Naming | - |
| | TEA-Ch RBBS | - |
| | - | CNB Motor Speed |
| | - | CNB Sensory Motor Speed |

Note. See the Supplementary Methods for task descriptions. Abbreviations: AWMA -153 Automated Working Memory Assessment (Alloway 2007), CNB - Computerized 154 Neurocognitive Battery (Gur et al. 2001), DKEF - Delis-Kaplan Executive Functioning System 155 (Delis et al. 2004), PhAB – Phonological Assessment Battery (Gallagher and Frederickson 156 1995), TEA-Ch RBBS - Test of Everyday Attention for Children, Red & Blues, Bags & Shoes 157 subscale (Manly et al. 2001), WASI - Wechsler Abbreviated Scale of Intelligence - Second 158 Edition (Wechsler 2011), WISC-R - Wechsler Intelligence Scale for Children – Revised 159 (Kaufman 1975). 160

161 White Matter Microstructure

162 We modelled mean FA for all ten tracts of the Johns Hopkins University (JHU) white matter

tractography atlas (Hua et al. 2008) averaged over the hemispheres (Figure 2). See

164 Supplementary Methods for details of the MRI acquisition and processing and Supplementary

165 Figure 3 and 4 for raw FA values in all tracts.

166 Analysis Methods and Structural Equation Modelling

167 Covariance matrices and scripts replicating key analyses can be obtained from:

168 https://github.com/df1234/gf_development. Supplementary Figure 5 and 6 show correlation

169 matrices of all tasks and white matter tracts modelled. We modelled raw scores for $g_{\rm f}$ and

working memory tasks, as preregistered. Raw scores on processing speed tasks were 170 transformed. This step was not preregistered, but found necessary to achieve model 171 172 convergence to ensure interpretability of scores. First, we inverted response time scores (using the formula y = 1/x) to obtain more intuitive measures of 'speed' for all but the CNB 173 Motor Speed task, for which raw scores were already a measure of speed. Afterwards, we 174 applied a log-transformation to reaction time tasks to increase normality and aid estimation. 175 176 For the CNB Motor Speed task only, we additionally removed values ± 2 SD of the mean (N = 6) because the presence of these outliers had caused convergence problems. 177

We modelled the associations between cognition and white matter microstructure using SEM in R (R core team 2015) using the package lavaan (Rosseel 2012). All models were fit using maximum likelihood estimation with robust Huber-White standard errors and a scaled test statistic. Missing data was addressed using full information maximum likelihood estimation.

182 We used SEM Trees to investigate whether the associations among cognitive and neural 183 measures differed with age. SEM Trees use decision tree methods to hierarchically split a dataset into subgroups if parameter estimates differ significantly based on a covariate of 184 185 interest - in this case age (Brandmaier et al. 2013). We first ran the watershed model in 186 OpenMx (Boker et al. 2011) and then passed this model object to semtree to compute the 187 SEM Trees. We ran one SEM Tree for each parameter of interest (e.g. the covariance between working memory and processing speed). All other parameters in each semtree 188 object were set to be invariant across groups to ensure that splits were specific to the 189 190 parameter of interest. We used a 10 - fold cross-validation estimation method as recommend by (Brandmaier et al. 2013). For the path from the cingulate to working memory only we 191 used 5 - fold cross-validation because the model did not converge using 10 - fold cross-192

validation. Minimum sample size in age group was set to N = 50 to ensure reliable estimation of standard errors. Note that this choice effectively limited search space for potential splits to ages 6.58 - 12.42 years for CALM and 8.08 - 15.49 years for NKI-RS.

196

Results

To evaluate the hypotheses generated by the watershed model, we built up the watershed 197 model in steps and carried our comprehensive tests of model fit at each step. First, we 198 assessed the overall fit of our models to the data using the chi-square test, root mean square 199 error of approximation (RMSEA), comparative fit index (CFI) and standardized root mean 200 square residual (SRMR). Good absolute fit was defined as RMSEA < 0.05, CFI > 0.97 and SRMR 201 202 < 0.05; acceptable fit as RMSEA = 0.08 - 0.05, CFI = 0. 95 - 0.97, SRMR = 0.05 - 0.10 (Schermelleh-Engel et al. 2003). Second, we assessed specific predictions from our models by 203 comparing them to alternative models. Comparative model fit for nested models was 204 205 assessed using the chi-square difference test. Non-nested models were compared using the Akaike (AIC) weights, which indicates the probability of a model being the data-generating 206 207 model compared to all other models tested (Wagenmakers and Farrell 2004). Lastly, we evaluated the significance and strength of relationships between specific variables in our 208 models by inspecting the Wald test for individual parameters, noting the joint R^2 where 209 210 relevant and reporting standardized parameter estimates. Absolute standardized parameter estimates above 0.10 were defined as small effects, 0.20 as typical and 0.30 as large (Gignac 211 212 and Szodorai 2016).

213 The Measurement Model of Cognition

To examine the neurocognitive architecture of $g_{\rm f}$, we started by modelling the cognitive components of the watershed model: $g_{\rm f}$, working memory and processing speed. Specifically, we fit a three-factor model of cognition (Figure 3) and compared it to alternative measurement models. This approach allowed us to test Hypothesis 1: namely that $g_{\rm f}$, working memory and processing speed form three separable, albeit likely correlated cognitive factors.

The Three-Factor Model (Figure 3) showed excellent *absolute fit* for both the CALM and NKI-RS sample (Table 3), indicating that overall, the data was compatible with a model of g_{f} , working memory and processing speed as three separate factors.

222 The Three-Factor Model also showed very good comparative fit for NKI-RS as well, with a 96.60% probability of being the data-generating model compared to all alternative models 223 tested, as indicated by its AIC weight (Figure 3). The evidence was more mixed for CALM, for 224 which the Three-Factor Model showed a 27.15% probability of being the data-generating 225 model, while Two-Factor Model B (Figure 3, treating working memory and $g_{\rm f}$ as a unitary 226 factor) showed a 72.85% probability of being the data-generating model, highlighting a close 227 228 relationship between $g_{\rm f}$ and working memory for this sample. The Single-Factor Model and Two-Factor Model A (Figure 3, treating speed and g_f as a unitary factor) showed a very low 229 230 (approximately 0%) probability of being the data-generating model, indicating that speed and 231 $g_{\rm f}$ were clearly separable in both samples.

| Single-Factor Model | Two-Factor Model A | Two-Factor Model B | Three-Factor Model | | | |
|---------------------------------|---|--------------------------------|--------------------------------|--|--|--|
| | CALM | | | | | |
| $\chi^2(20) = 70.28, p < .001$ | χ ² (19) = 67.99, <i>p</i> < .001 | $\chi^2(19) = 41.66, p = .002$ | $\chi^2(18) = 41.74, p = .001$ | | | |
| RMSEA = .068 | RMSEA = .068 | RMSEA = .047 | RMSEA = .049 | | | |
| [.051085] | [.052086] | [.027066] | [.030068] | | | |
| CFI = .963 | CFI = .964 | CFI = .983 | CFI = .983 | | | |
| SRMR = .047 | SRMR = .043 | SRMR = .032 | SRMR = .032 | | | |
| AIC = 9697.18 | AIC = 9696.44 | AIC = 9668.58 | AIC = 9670.55 | | | |
| BIC = 9800.66 | BIC = 9804.24 | BIC = 9776.37 | BIC = 9782.66 | | | |
| $AIC_{weight} = 0\%$ | $AIC_{weight} = 0\%$ | $AIC_{weight} = 72.85\%$ | $AIC_{weight} = 27.15\%$ | | | |
| | N | KI-R | | | | |
| $\chi^2(35) = 109.96, p < .001$ | χ ² (34) = 108.15, <i>p</i> < .001 | $\chi^2(34) = 64.85, p = .001$ | $\chi^2(32) = 54.15, p = .009$ | | | |
| RMSEA = .080 | RMSEA = .081 | RMSEA = .052 | RMSEA = .045 | | | |
| [.064097] | [.064098] | [.033071] | [.024065] | | | |
| CFI = .936 | CFI = .936 | CFI = .974 | CFI = .981 | | | |
| SRMR = .045 | SRMR = .044 | SRMR = .035 | SRMR = .030 | | | |
| AIC = 7155.64 | AIC = 7155.74 | AIC = 7109.43 | AIC = 7102.74 | | | |
| BIC = 7270.07 | BIC = 7273.98 | BIC = 7227.67 | BIC = 7228.60 | | | |
| $AIC_{weight} = 0\%$ | $AIC_{weight} = 0\%$ | $AIC_{weight} = 3.40\%$ | $AIC_{weight} = 96.60\%$ | | | |

232 Table 3. Model Fit of Competing Measurement Models

Note. See Figure 3 for the configuration of different models. Abbreviations: Akaike Information Criterion (AIC),
 Bayesian Information Criterion (BIC), Akaike weight (AIC_{weight})

Overall, these result provide mixed evidence for Hypothesis 1: Even though working memory, 235 236 processing speed and g_f were highly correlated (Table 4), processing speed formed a clearly 237 separable factor from working memory and $g_{\rm f}$ in both samples. Working memory and $g_{\rm f}$, however, were clearly separable only in NKI-RS, but not CALM, suggesting greater similarity 238 between g_f and working memory in the CALM sample. To facilitate comparison across 239 samples and in accordance with our preregistered analysis plan we nonetheless used the 240 three-factor measurement model (Table 4, Supplementary Table 1) in all subsequent 241 242 analyses.

| Sample | Path | Standardized Estimate |
|--------|---------------------------|---|
| CALM | g _f <-> memory | 0.71, <i>z</i> = 28.42, <i>p</i> < .001 |
| | g _f <-> speed | 0.55, <i>z</i> = 12.20, <i>p</i> < .001 |
| | memory <-> speed | 0.79, <i>z</i> = 19.35, <i>p</i> < .001 |
| NKI-RS | g _f <-> memory | 0.91, <i>z</i> = 19.51, <i>p</i> < .001 |
| | g _f <-> speed | 0.81, <i>z</i> = 24.73, <i>p</i> < .001 |
| | memory <-> speed | 0.87, <i>z</i> = 17.43, <i>p</i> < .001 |

243 Table 4. Covariance between Cognitive Measures in the Three-Factor Model

244 *Note.* See Supplementary Table 1 for factor loadings.

The Relationship between Working Memory, Processing Speed and $g_{\rm f}$

We next examined the relationships between working memory, processing speed and g_f in more detail. Specifically, we fit a SEM including regression paths between working memory and g_f , as well as speed and g_f , to test Hypothesis 2 - that working memory and processing speed each predict individual differences in g_f . We found that this model showed good absolute fit for both samples (CALM: $\chi^2(18) = 41.74$, p = .001; RMSEA = .049 [.030 - .068]; CFI = .983; SRMR = .032, NKI-RS: $\chi^2(32) = 54.15$, p = .009; RMSEA = .045 [.024 - .065]; CFI = .981; SRMR = .030), indicating that, overall, the data was compatible with our model.

To further scrutinize the relationship between g_f, working memory and speed, we compared 253 our freely-estimated model to a set of alternative models with different constraints imposed 254 upon the regression paths. First, to test whether working memory and speed each made 255 256 different contributions, we tested an alternative model in which the paths from processing speed and working memory to $g_{\rm f}$ were constrained to be equal. In CALM ($\Delta \chi^2(1) = 15.53$, p < 15.53257 .001), but not NKI-RS ($\Delta \chi^2(1) = 3.25$, p = .072), the freely-estimated model fit better than the 258 259 equality-constrained model, indicating that working memory and speed each made different 260 contributions in CALM but not NKI-RS. Next, we tested whether the freely estimated model fit better than a model in which the path between $g_{\rm f}$ and working memory was constrained 261

to zero. We found that the freely estimated model fit better for both samples (CALM: 262 $\Delta \chi^2(1) = 20.77$, p < .001; NKI-RS: $\Delta \chi^2(1) = 12.97$, p < .001). In line with our hypothesis, this 263 result indicates that working memory makes a significant incremental contribution to $g_{\rm f}$. 264 Finally, we tested a model in which the path between $g_{\rm f}$ and processing speed was 265 constrained to zero. This model showed no difference in fit to the freely estimated model for 266 CALM ($\Delta \chi^2(1) = 0.02$, p = .875) or NKI-RS ($\Delta \chi^2(1) = 0.04$, p = .849). Contrary to our hypothesis, 267 268 this indicates that there was no clear incremental contribution of processing speed to $g_{\rm f}$. Finally, we inspected standardized path estimates of the freely estimated model to assess the 269 270 effect seizes of working memory and processing speed. Parameter estimates showed that 271 working memory showed a greater effect on $g_{\rm f}$ than processing speed, particularly in CALM (Table 5) even though raw correlations between $g_{\rm f}$ and speed were high in both samples 272

273 (Table 4).

274 *Table 5. Regression Path Estimates.*

| Sample | Path | Standardized Estimate |
|--------|--------------------------|--|
| CALM | speed -> g _f | -0.01, <i>z</i> = -0.16, <i>p</i> = .876 |
| | memory -> g _f | 0.72, <i>z</i> = 7.65, <i>p</i> < .001 |
| NKI-RS | speed -> g _f | 0.06, <i>z</i> = 0.21, <i>p</i> = .208 |
| | memory -> g _f | 0.86, <i>z</i> = 1.81, <i>p</i> = .070 |

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Overall these results provide mixed evidence for Hypothesis 2: There was good evidence that working memory and speed made a significant joint contribution to g_f , and that working memory made an incremental contribution to g_f in CALM. Contrary to our hypothesis, and the watershed model, however, processing speed showed no significant incremental contribution to g_f , above and beyond working memory. We explore likely explanations for this finding in the Discussion.

282 The Measurement Model of White Matter

We next examined the measurement model of white matter to test Hypothesis 3, namely 283 284 that white matter microstructure is a multi-dimensional construct. Specifically, we examined 285 absolute model fit of a single factor model to test whether a unidimensional model could adequately capture white matter microstructure. As expected, the single-factor model of 286 white matter microstructure did not fit the data well (CALM: $\chi^2(35) = 124.63$, p < .001; 287 RMSEA = .125 [.103 - .147]; CFI = .933; SRMR = .039; NKI-RS: $\chi^2(35)$ = 132.33, p < .001; RMSEA 288 = .204 [.167 - .242]; CFI = .885; SRMR = .023). This indicates that white matter microstructure 289 290 could not be reduced to a single 'global FA' dimension in our samples, in line with (Lövdén et al. 2013; Kievit, Davis, Griffiths, Correia, CamCAN, et al. 2016) and supporting Hypothesis 3. 291 We therefore modelled each of the ten white matter tracts separately in all subsequent 292 models. 293

294 The Watershed Model: Relationships between Cognition and White Matter

Next, we fit the full watershed model including white matter, working memory, processing speed and $g_{\rm f}$. Following our general analysis procedure, we investigated overall model fit, alternative models and individual path estimates to gain a comprehensive understanding of the relationships in the watershed model and to test Hypothesis 4 - that white matter contributes to working memory capacity and processing speed, which, in turn, contribute to $g_{\rm f}$.

We found largely converging results across samples. The watershed model showed good absolute fit in CALM ($\chi^2(78) = 107.78$, p = .014; RMSEA = .026 [.012 - .038]; CFI = .981; SRMR = .043) and acceptable fit in NKI-RS ($\chi^2(112) = 219.22$, p < .001; RMSEA = .053 [.043 - .064]; CFI = .928; SRMR = .088). White matter explained large amounts of variance in working memory (R^2_{CALM} = 32.3%; R^2_{NKI-RS} = 46.1%) and processing speed (R^2_{CALM} = 38.2%; R^2_{NKI-RS} = 306 54.4%), which, in turn, explained even more variance in g_f (R^2_{CALM} = 51.2%; R^2_{NKI-RS} = 78.3%). In line with Hypothesis 4, this indicates that the watershed model fit the data overall.

Comparing the freely estimated watershed model to alternative, constrained, models 308 showed that white matter contributed significantly to memory and processing speed. 309 Specifically, a model in which paths from white matter to processing speed were constrained 310 to zero fit worse than the freely-estimated model (CALM: $\Delta \chi^2(10) = 50.26$, p < .001; NKI-RS: 311 $\Delta \chi^2(10) = 27.19$, p = .002), as did a model in which paths from white matter to working 312 memory were constrained to zero (CALM: $\Delta \chi^2(10) = 52.26$, p < .001; NKI-RS: $\Delta \chi^2(10) = 25.85$, 313 p = .004). As hypothesised, white matter therefore contributed to both processing speed and 314 working memory. 315

We next inspected that relationship between individual white matter tracts and working memory and speed in more detail. A model in which paths from white matter to working memory and speed were constrained to be equal, fit worse than the freely-estimated watershed model for CALM ($\Delta \chi^2(18) = 47.76$, p < .001) and NKI-RS ($\Delta \chi^2(18) = 30.42$, p = .034), indicating that the role of white matter microstructure in supporting working memory and processing speed differed across tracts. This supports the notion that there is a many-to-one mapping between white matter and cognition - a core tenet of the watershed model.

Investigating individual standardised parameter estimates of the different white matter tracts showed that for CALM, only the anterior thalamic radiation contributed significantly to processing speed, whereas the superior longitudinal fasciculus, forceps major and cingulum were significantly, independently and positively related to working memory (Figure 4). For NKI-RS, the superior longitudinal fasciculus was significantly and positively related to

processing speed and working memory (Figure 5). Two tracts showed an unexpected, 328 strongly negative (< -1), relationship: the forceps minor for CALM and the inferior fronto-329 330 occipital fasciculus for NKI-RS. We found that these negative estimates occurred only when 331 all other brain to cognition pathways were also estimated: When estimated on their own, path estimates were positive (forceps minor to working memory: standardized estimate = 332 0.36, z = 4.05, p < .001; inferior fronto-occipital fasciculus to working memory: standardized 333 334 estimate = 0.14, z = 0.859, p = .390; inferior fronto-occipital fasciculus to processing speed: standardized estimate = 0.26, z = 1.41, p = .158). This sign-flip suggests that the negative 335 336 pathways were potentially due to modelling several, highly-correlated paths at the same time (Jöreskog 1999). Overall, these results further support the watershed prediction that multiple 337 white matter tracts map onto working memory and processing speed. 338

Finally, we probed the watershed model in more detail by testing a set of alternative 339 expressions of the watershed model still compatible with the core tenants of the watershed 340 model - as well as a set of alterative models incompatible with the watershed model. We 341 compared all alternatives (see Figure 6 for graphical representations) to the original 342 343 watershed model by inspecting each models' relative probability of being the data-generating model as indicated by AIC weights (Wagenmakers and Farrell 2004). We found that the 344 original watershed model showed a very high probability (98.58%) of being the data-345 generating model for CALM but only a 0.10% probability for NKI-RS. For NKI-RS, a different 346 expression of the watershed model, such that $g_{\rm f}$ was regressed on working memory, which 347 was regressed on processing speed, which was then regressed on white matter (Alternative 348 A, Figure 6) showed a 95.04% probability of being the data-generating model. This model 349 only showed a 0.37% probability for CALM. Another expression of the watershed model, in 350 which all tasks were modelled separately as manifest, rather than latent, variables 351

(Alternative B, Figure 6), showed no advantage over the watershed model for CALM (0.00% 352 probability) or NKI-RS (0.00% probability). We next tested two alternative models 353 354 incompatible with the tenants of the watershed model. We found that a model in which the 355 hierarchy between cognitive endophenotypes and g_f was inverted (Alternative C, Figure 6) showed comparatively low probability of being the data-generating model for both CALM 356 357 (0.00%) and NKI-RS (2.86%). Similarly, a model in which $g_{\rm f}$ was directly regressed on white 358 matter, working memory and processing speed (Alternative D, Figure 6), showed no clear advantage over the watershed model for CALM (1.05% probability) or NKI-RS (0.00% 359 360 probability). Overall these model comparisons highlight that while the watershed model fit the data for both samples and had large explanatory power (as indicated by R^2 s), the precise 361 configuration of the watershed model may differ somewhat between cohorts. 362

In summary, we found that the watershed model performed well overall for both cohorts. As 363 hypothesised, white matter contributed to working memory and processing speed, which, in 364 turn, contributed to $g_{\rm f}$, and explained large amounts of variance therein. Also as predicted by 365 366 the watershed model, there was a many-to-one mapping between white matter tracts and cognition. The exact configuration of the watershed model, however, may differ slightly 367 between cohorts. These differences may be a function of cohort differences in sample size, 368 average levels of cognitive ability and/or the specific tasks used – a topic we will return to in 369 the Discussion. 370

371 Testing for potential confounds

We carried out a series of supplementary and non-preregistered analyses to examine whether possible confounders influenced our models. These analyses showed that our findings were robust to the inclusion of covariates such as scanner motion or socio-economic

375 status. They were also robust across genders and participants taking or not taking 376 medication. There were no differences in the structure of the model between participants 377 with and without diagnosed disorders for CALM. Potential small differences cannot be ruled 378 out for NKI-RS, likely due to the low number of diagnosed participants of N = 106379 (Supplementary Material).

Age-Related Differences in the Neurocognitive Architecture of $g_{\rm f}$

Finally, we tested Hypothesis 5 - that the contribution of working memory and processing speed to g_{f} varied with age. We first inspected cross-sectional differences in g_{f} , working memory and processing speed, and then used SEM trees to investigate potential agedifferences in the relationships between these factors. In additional, non-preregistered, analyses we also used SEM Trees to investigate potential agebetween white matter and cognitive endophenotypes by inspecting paths that were significant in the watershed model (Figure 4 and 5).

388 SEM trees combine SEMs with decision tree methods, separating a dataset into subgroups (in this case age groups) if SEM parameter estimates of interest differ sufficiently (Brandmaier et 389 al. 2013). SEM trees allowed us to investigate age as a potential moderator without imposing 390 391 a-priori categorical age splits. We initially allowed for no more than two age groups. This 392 yielded inconsistent results for CALM and NKI-RS (Supplementary Table 4). To test whether these inconsistencies were an artefact of allowing for only two groups, we repeated our 393 394 analysis and allowed for up to four age groups. This analysis yielded consistent results 395 between CALM and NKI-RS (Table 6). This pattern of results indicates that the initial parameters of our analysis caused us to miss relevant age differences. 396

| Path | Est. | Age | Est. | Age | Est. | Age | Est. |
|--------------------------|--------|-------|--------|-------|-------|-------|-------|
| | Before | Split | Betw. | Split | Betw. | Split | After |
| | | 1 | | 2 | | 3 | |
| | | | CALM | | | | |
| memory <-> speed | 0.85 | 8.46 | 0.97 | 9.46 | 0.74 | - | - |
| memory -> g _f | 0.83 | 9.38 | 0.42 | 10.04 | 1.14 | 10.88 | 0.94 |
| speed -> g _f | 0.04 | 6.88 | -0.19 | 11.21 | 0.17 | - | - |
| SLF -> memory | 0.67 | 7.21 | 0.18 | 11.21 | 0.76 | - | - |
| FMaj —> memory | 0.59 | 7.71 | 0.14 | 9.29 | 0.33 | 11.13 | 0.74 |
| CG –> memory 1 | 0.64 | 6.96 | 0.09 | 11.04 | 0.70 | - | - |
| ATR -> speed | 0.96 | 7.13 | 0.68 | 7.96 | 0.17 | 11.96 | 0.65 |
| | | | NKI-RS | | | | |
| memory <-> speed | 0.90 | 9.82 | 0.48 | 14.72 | 1.11 | - | - |
| memory -> g _f | 1.10 | 8.59 | 0.59 | 12.67 | 1.03 | - | - |
| speed -> g _f | 0.53 | 8.59 | -0.12 | 12.96 | 0.52 | - | - |
| SLF -> memory | 2.15 | 8.30 | 1.47 | 12.15 | 1.93 | - | - |
| SLF -> speed | 3.12 | 8.63 | 1.83 | 15.09 | 2.31 | - | - |

398 Note. The table shows differences in parameter estimates for paths of interest (as shown in 399 Figure 4 and 5) depending on participants' age in years. Our analyses allowed for a maximum 400 of three age splits (and thus four age groups). An absence of a third age split (denoted by '-' 401 in the table), indicates that the SEM tree split only twice, suggesting no further changes in 402 parameter strength after the second split. See Supplementary Figure 7 for a graphical 403 representation of these results.

As shown in Figure 7, g_{f} , working memory and processing speed factor scores increased with 404 age for all three cognitive phenotypes. In line with our hypothesis, SEM trees showed that 405 406 there were pronounced age-related differences in brain-behaviour in childhood and 407 adolescence (Table 6). For both samples and all but one path, there was an initially strong 408 relationship between components of the watershed model, then a dip around ages 7 - 9 409 years for CALM and age 8 for NKI-RS, followed by an increase in path strength around ages 11 - 12 (see Supplementary Figure 7 for a graphical representation of these results). 410 Speculatively, this pattern of results is consistent with an interpretation of a reorganization of 411

412 neurocognitive faculties in late childhood, followed by a consolidation of neurocognitive413 pathways around the onset of adolescence (Johnson 2000, 2011).

414

Discussion

415 We here used multivariate statistical techniques to investigate the neurocognitive architecture of g_f in two large ($N_{CALM} = 551$, $N_{NKI-RS} = 335$) developmental cohorts and, for the 416 first time, investigated how the neurocognitive architecture of $g_{\rm f}$ changes dynamically with 417 age. We tested a preregistered watershed model of $g_{\rm f}$, which predicts a hierarchy of partially 418 independent effects. As might be expected from a multi-cohort study, there were some 419 differences between the community-ascertained cohort (NKI-RS) and the cohort of children 420 421 and adolescents with learning difficulties (CALM) in specific path estimates. Overall however, we found convergent results across these two heterogeneous samples. The watershed model 422 performed well for both CALM and NKI-RS: White matter contributed to working memory 423 and processing speed, which, in turn, contributed to g_f and explained 51% of variance therein 424 for the CALM sample and 78% of variance for NKI-RS. Models were robust across genders, 425 426 participants taking or not taking medication and when controlling for socio-economic status and scanner motion. Investigations of age effects showed that the relationship between 427 cognitive abilities and white matter dipped in strength around ages 7-12 years. Speculatively, 428 429 this age-effect may reflect a reorganization of the neurocognitive architecture during prepuberty and early puberty (Byrne et al. 2017). These findings have implications for 430 understanding and targeting cognitive impairments in populations with learning difficulties. 431

The watershed model tested here consists of three levels: g_f forms the most down-stream point, with working memory and processing speed as intermediate tributaries, and white

matter microstructural tracts as upstream sources. Previous studies suggested that matter 434 microstructure is best characterised by a single, 'global FA' factor (Penke et al. 2010) while 435 436 others have contended that association patterns among different white matter tracts are 437 more complex (Lövdén et al. 2013; Kievit, Davis, Griffiths, Correia, CamCAN, et al. 2016). Here we found strong evidence for a multifactorial view of white matter tracts – for both samples, 438 a unidimensional model of white matter fit poorly, and for CALM, multiple tracts also showed 439 440 partially independent contributions to distal cognitive outcomes. This is in line with the 441 watershed model. There were some differences between cohorts as to which tracts 442 contributed most to working memory and processing speed: In line with previous research (Kievit, Davis, Griffiths, Correia, CamCAN, et al. 2016; MacPherson et al. 2017; Bathelt et al. 443 444 2018), we found that the anterior thalamic radiation was related to processing speed, as were the forceps major, forceps minor and the cingulum to working memory for CALM. 445 However, these tracts were not significant for NKI-RS. A possible explanation for these 446 differences between samples is the discrepancy in the number of participants with imaging 447 448 data (N = 165 in CALM versus N = 67 in NKI-RS). This discrepancy likely confers differential power to detect weaker pathways. Other, not mutually exclusive, explanations are that the 449 observed differences reflect differences in brain-behaviour mapping between more atypical 450 and typical cohorts (Bathelt et al. 2018), sampling variance across two independent cohorts 451 collected under somewhat different socio-economic conditions (United Kingdom and United 452 453 States of America), or a more uniform age distribution in NKI-RS. While DTI images were processed with the same pipeline across sites, the scanner and MRI acquisition protocol were 454 also different. Although previous work suggests that FA is relatively robust measure in multi-455 site comparisons (Vollmar et al. 2010), we therefore cannot rule out site differences as a 456 potential confound. It will therefore be necessary to replicate these findings in large typical 457

and atypical cohorts collected in the same setting. Of note, however, the superior longitudinal fasciculus was robustly associated with working memory across the two different samples and settings. For NKI-RS, the superior longitudinal fasciculus was also associated with processing speed. The superior longitudinal fasciculus is a large, bilateral association fibre connecting temporal, occipital, parietal and frontal regions (Kamali et al. 2014). It is therefore well-situated for supporting cognitive processes such as *g*_f, which rely on integrative multipledemand systems (Jung and Haier 2007; Fedorenko et al. 2013; Parlatini et al. 2017).

Our findings for the cognitive levels of the watershed model highlighted a close relationship 465 between working memory and $g_{\rm f}$. Previous studies had variably suggested that $g_{\rm f}$ and working 466 467 memory (Kyllonen and Christal 1990; Fukuda et al. 2010), or g_f and processing speed (Kail and Salthouse 1994; Salthouse 1996; Coyle et al. 2011; Ferrer et al. 2013) may be most closely 468 related. We found that all three cognitive factors were highly correlated for both samples. 469 Nonetheless, processing speed formed a cognitive factor clearly separable from working 470 memory and g_{f} . Working memory and g_{f} , in turn, were separable in the community-471 472 ascertained NKI-RS but not in CALM, the cohort of children and adolescents with learning difficulties. This close relationship between $g_{\rm f}$ and working memory was also evident in other 473 models of CALM where processing speed and working memory were used as joint predictors 474 of g_f. Contrary to our hypotheses, processing speed became non-significant after controlling 475 for working memory here. There are several possible, and not mutually exclusive, 476 explanations for this finding and the apparent differences between cohorts. First, a broader 477 478 set of speed tasks (which might be captured by several latent variables for clerical speed, 479 choice reaction time and speed variability) might show higher predictive power than the single latent variable for speed, which could be modelled here. This may be particularly 480

pertinent for heterogeneous cohorts like CALM. Second, and in line with previous work 481 showing that time-constraints increase isomorphism of g_f and working memory (Chuderski 482 483 2013), even standard implementations of g_f tasks may place considerable time-pressure on 484 struggling learners, thereby increasing g_f - working memory covariance in CALM as compared to NKI-RS. Conversely, less subjective or objective time pressure may also confer a 485 differentiation of cognitive domains and the watershed hierarchy in cohorts of older ages 486 487 and/or higher ability levels, such as NKI-RS: There was some evidence that, for this sample, speed formed an intermittent level in the hierarchy between white matter and working 488 489 memory (Alternative A, Figure 6). Future longitudinal research will be necessary to differentiate these different configurations of the watershed model and scrutinize the causal 490 flow of effects. For now, our findings highlight the value of replicating analyses in different 491 cohorts using different tasks: While evidence was mixed for the association between $g_{\rm f}$ and 492 processing speed, the strong associations between $g_{\rm f}$ and working memory across samples 493 indicate a robust and likely generalizable relationship between these two domains, 494 supporting the notion that mental information processing capacity is a key determinant of $g_{\rm f}$ 495 (Kyllonen and Christal 1990; Fukuda et al. 2010). 496

The associations in the watershed model differed between ages in a complex, non-monotonic fashion. Previous research had suggested either a decrease in covariance among cognitive domains with age (age differentiation; Garrett 1946), an increase in covariance with age (age de-differentiation; Blum and Holling 2017), or no changes with age (Tucker-Drob 2009; de Mooij et al. 2018). These investigations have traditionally focussed on relations between cognitive domains, however, not on relationships between brain and cognition - although see de Mooij et al. (2018). Possible linear and non-linear changes in brain-behaviour mapping

with age have remained mostly unexplored (Tamnes et al. 2017). Using structural equation
modelling trees, a novel decision-tree-based technique, we here found evidence of complex
developmental differences consistent across samples and relationships in the watershed
model: Initially strong path estimates showed a pronounced decrease in strength around
ages 7 - 9 years, followed by a renewed increase in the strength, even surpassing initial levels,
around ages 10 - 15.

There are at least two possible explanations for this developmental dip in brain-cognition 510 relationships. First, there may be a true decrease in relationship strength during this time of 511 life. Possibly, other cognitive skills, such as verbal reasoning, temporarily support $g_{\rm f}$, resulting 512 513 in weaker relationships between $g_{\rm f}$ and working memory. Alternatively, the configuration of the watershed model may change temporarily during this time, which could also manifest in 514 an apparently weaker covariance structure. In this case, the true relationship between $g_{\rm f}$, 515 memory, speed and white matter may still be strong, just configured differently from the 516 watershed model. We note that both explanations are compatible with the interactive 517 518 specialization theory (Johnson 2000, 2011), which predicts as remapping of the relationships between brain substrates and cognitive abilities during development. 519

520 On a physiological level, this age effect may be driven by neuroendocrine changes during pre-521 and early puberty. Puberty is driven by a complex and only partially understood set of 522 hormonal events including gonadarche and andrenarche (Sisk and Zehr 2005). Gonadarche 523 begins with the secretion of gonadotropin-releasing hormone from the hypothalamus around 524 ages 10-11 years and closely tracks the overt bodily changes of puberty (Dorn 2006). 525 Andrenarche, beginning with the maturation of the andrenal gland, starts as early as six years 526 of age, and is increasingly recognized as a complimentary driver of puberty and brain

527 development (Byrne et al. 2017). It is possible that the hormonal changes of andrenarche and 528 early gonadarche may lead to a level of neural reorganization, which may initially appear as 529 weaker relationships in the watershed model. The sweeping bodily, social and cognitive 530 changes happening in early adolescence may then drive a consolidation of the neurocognitive 531 architecture of $g_{\rm f}$.

532 On a more general level, these age effects suggest the existence of potential non-linear changes in brain-behaviour mapping during childhood and adolescence and underline the 533 value of modern statistical approaches, such as SEM Trees, for the study of age-related 534 differences. It is worth noting, however, that these findings, which are based on an inherently 535 536 exploratory technique, will need to be replicated in future confirmatory studies with finegrained data on puberty and larger sample sizes. The latter will also allow for detailed 537 investigations of potential gender differences. Moreover, while we were able to investigate 538 *individual differences* in g_f, we could not assess *intra-individual changes* during childhood and 539 adolescence. Although the relatively narrow age range makes large cohort effects unlikely, it 540 541 may still be that there were differences in recruitment and selection that varied across the age range. As such, the cross-sectional nature of our samples limits our ability to make 542 inferences about developmental dynamics. 543

544 Our study illustrates some of the advantages and challenges of preregistered secondary data 545 analyses. We agree with others in the field that secondary data analysis need not be and 546 should not be confounded with purely exploratory research (Mills and Tamnes 2014; Orben 547 and Przybylski 2019; Scott and Kline 2019). Preregistrations, as well as dedicated multivariate 548 methods such as SEM, can help to reduce the scope for analytic flexibility and increase 549 scientific rigour when using rich, secondary datasets. Preregistrations also do not preclude

the use of exploratory methods or the ability to ask exploratory questions, as we did in our 550 analysis of age effects. Preregistrations merely facilitate the distinction between exploratory 551 552 and confirmatory research (Wagenmakers et al. 2012). There are, however, some unique 553 challenges to preregistering secondary data analyses worth noting. First, information on the 554 precise measures collected is not always available prior to data access, which can limit the level of detail in which an analysis can be preregistered. Second, data quality and the level of 555 556 data-processing, the latter being particularly relevant for MRI data, is not always clear a priori (e.g. see Kievit et al. 2018), which can necessitate changes to analyses plans after data 557 558 inspection. Third, convergence issues are fairly common when using complex multivariate methods such as SEM. We found it necessary to transform some of our speed variables, for 559 instance, to achieve model convergence. Such post-hoc modifications, not guided by the 560 palatability of the results, but rather by unforeseen, and sometimes unforeseeable, practical 561 considerations, mean that preregistration can sometimes fall short of full compliance. 562 Nevertheless, we believe that even imperfect preregistrations, alongside shared code, data 563 and the transparent presentation of results, can help the reader distinguish between 564 confirmatory and exploratory results, and adjust their level of confidence in conclusions 565 accordingly. For guidance on maximizing transparency in preregistration of secondary data, 566 see Weston et al. (2018). 567

Finally, the findings from our study have implications understanding and targeting cognitive impairments in populations with learning difficulties. First, the close relationship between working memory and g_f found here and in other studies (Fukuda et al. 2010; Chuderski 2013), indicates that children and adolescents struggling with working memory are likely to also struggle in terms of complex reasoning tasks. Either reducing working memory load,

decreasing time constraints, or training working memory and fluid ability capacity in such 573 574 populations may therefore be promising lines of inquiry for intervention studies. It is worth 575 highlighting, however, that cognitive training studies have so far shown little evidence of (far) 576 transfer: Training abstract reasoning, a common measure of g_{f} , has not resulted in robust increases in working memory (Knoll et al. 2016) and working memory training has not been 577 shown to transfer to reasoning skills or school performance (Dunning et al. 2013; 578 579 Schwaighofer et al. 2015). Similarly, transfer from processing speed to reasoning seems to be limited (Mackey et al. 2011). The results obtained here suggest that interventions may 580 581 increase their chance of success by implementing programs of sufficient complexity to affect the entire neurocognitive architecture of effects (see also Kievit et al. 2016). The level of 582 intensity required to produce sustained benefits may need to be as demanding and 583 consistent as education itself, which shows robust effects in increasing general cognitive 584 abilities over time (Ritchie and Tucker-Drob 2018). This work and work by others (Noack et al. 585 2014) also highlights the value of assessing, modeling, and potentially intervening on, 586 587 multiple tasks, rather than relying on a single task to capture complex cognitive domains such as $g_{\rm f}$. Finally, the age-related differences in the relationships of the watershed model 588 observed using SEM-trees suggest that some interventions may work best at particular 589 590 developmental phases.

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Figure Legends

Figure 1. The Watershed Model. Schematic representation of the watershed model
developed by (Cannon and Keller 2006) and adapted for the present study. Fluid ability is
hypothesized to be the downstream product of working memory and processing speed,
which are, in turn, the product of white matter contributions. Figure adapted from Kievit et
al. (2016).

813 Figure 2. White Matter Tracts Modelled in the Analyses.

814 *Figure 3. Different Measurement Models of Cognition.* Abbreviations: WM: working memory,815 PS: processing speed

Figure 4. The Watershed Model in CALM. See Supplementary Table 2 for regression
estimates. Residual covariances between white matter tracts were allowed but are not
shown for simplicity. Abbreviations: uncinate fasciculus (UF), superior longitudinal fasciculus
(SLF), inferior fronto-occipital fasciculus (IFOF), anterior thalamic radiations (ATR),
cerebrospinal tract (CST), forceps major (FMaj), forceps minor (FMin), dorsal cingulate gyrus
(CG), ventral cingulate gyrus (CH), inferior longitudinal fasciculus (ILF).

Figure 5. The Watershed Model in NKI-RS. See Supplementary Table 3 for regression
estimates. Residual covariances between white matter tracts were allowed but are not
shown for simplicity.

Figure 6. Configuration of Alternative Models. Alternatives A and B are watershedcompatible, while C and D are watershed-incompatible. The best-fitting model for CALM is highlighted in blue; the best-fitting model for NKI-RS is highlighted in green. Regression paths only are shown for simplicity. Square shapes denote manifest variables and oval shapes latent variables.

830 Figure 7. Cognitive Factor Scores by Age.