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Discordant attributes of structural and functional brain connectivity in a two-layer multiplex network

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Several studies have suggested that functional connectivity (FC) is constrained by the underlying structural connectivity (SC) and mutually correlated. However, not many studies have focused on differences in the network organization of SC and FC, and on how these differences may inform us about their mutual interaction. To explore this issue, we adopt a multi-layer framework, with SC and FC, constructed using Magnetic Resonance Imaging (MRI) data from the Human Connectome Project, forming a two-layer multiplex network. In particular, we examine node strength assortativity within and between the SC and FC layer. We find that, in general, SC is organized assortatively, indicating brain regions are on average connected to other brain regions with similar node strengths. On the other hand, FC shows disassortative mixing. This discrepancy is apparent also among individual resting-state networks within SC and FC. In addition, these patterns show lateralization, with disassortative mixing within FC subnetworks mainly driven from the left hemisphere. We discuss our findings in the context of robustness to structural failure, and we suggest that discordant and lateralized patterns of associativity in SC and FC may provide clues to understand laterality of some neurological dysfunctions and recovery.

The relationship between structural connectivity and functional connectivity has attracted much attention in recent years¹⁻⁵. Yet despite numerous empirical^{4,6} and computational studies^{1,2} the nature of their interaction remains only incompletely understood. Several studies have suggested that functional brain networks are constrained by the underlying structural connectivity⁷, and brain-wide comparisons have supported the idea that functional connectivity (FC), measured in the resting state, and structural connectivity (SC) are in general statistically correlated⁸. For example, when there is a strong anatomical connection between two areas of the brain, the corresponding functional connection is likely to be strong as well, but the inverse is not always the case^{5,9-14}. A number of computational models have successfully reproduced some features of empirical functional connectivity, including models based on large-scale dynamics^{15,16} or graph theory metrics¹⁷. While most studies have emphasized the statistical association between SC and FC, important differences and discrepancies remain^{11,14}. This may be expected given that fMRI and dMRI measure different signals and use different statistical approaches for estimating pairwise connections between regions of interest (ROIs). Whereas SC estimates a direct relationship or path between two brain regions, measurements of FC, for instance, estimated by Pearson's moment correlation coefficients, incorporate both direct and indirect relationships between two nodes influenced by other brain areas^{18,19}.

However, not many studies have focused on fundamental topological differences in the network organization of SC and FC, and on how these differences may provide insight into their mutual interaction. To explore this issue, we adopt a multi-layer framework, with SC and FC forming a multiplex network²⁰⁻²². What are the fundamental topological differences that underpin FC and SC, and are these differences biologically meaningful? Are there potential benefits that might arise from topological differences among these two different types of brain

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networks? As previous studies have shown, SC and FC are intricately (and non-trivially) linked – for example, by demonstrating that FC arises from underlying anatomical connections. Thus, if any systematic/consistent topological differences exist between these two networks, are they mere by-products of the generative process or could they represent biologically meaningful features of multi-layer organization that carry benefit or enhance overall functionality?

Specifically, we will examine node strength assortativity within and between the SC and FC layers. Degree assortativity has been extensively studied in the context of network robustness^{23–25}. In isolated networks, assortativity stands for correlation among nodes features (e.g., degree) of directly connected nodes²³. A network is said to be assortative if its connectivity pattern is such that high-degree nodes are frequently attached to other high-degree nodes, and low-degree nodes are preferentially connected to other low-degree nodes. Assortative networks are generally resilient against the random removal of nodes and edges^{23,26,27}. In multiplex networks, correlations among nodes features can be measured both within- and between-layers^{28,29}. The two types of correlations provide different information about the robustness of the interdependent system. In the absence of any correlations between layers, it is well known that an interdependent network undergoes a sudden percolation transition²⁰. An increased within-layer degree assortativity decreases the robustness of the network in terms of the percolation threshold³⁰. On the other hand, positive values of between-layer correlations generally mitigate the abrupt nature of the transition, making the system more robust. Examples include degree-degree correlations³¹, edge overlap^{32–36}, clustering and spatial coordinates^{37–39}.

Robustness is an important feature of brain networks^{40–42}. In many cases, FC network patterns appear to maintain large-scale patterns and functionality even in the face of serious disruptions or disturbance of underlying SC^{43,44}. Can a multi-layer model shed light on the network basis for these observations? Here, we investigate if these findings from theoretical investigations and non-biological networks carry over to human brain networks derived from Magnetic Resonance Imaging (MRI) data from the Human Connectome Project⁴⁵. SC was constructed based on diffusion MRI and tractography and FC was estimated using regularized partial correlation coefficients with the elastic net. The two layers (SC and FC) are coupled by creating links on pairs of corresponding nodes in the two layers, thus creating a multiplex network. We examine assortative mixing by strength both within SC and FC and between the two layers. We divided FC and SC into 7 subnetworks according to a canonical resting-state partition⁴⁶. We find that coupled structural and functional human brain networks exhibit a combination of similarities and differences. In addition, we find heterogeneous node strength correlations across the two layers and within FC and SC subnetworks, as well as interesting contrast between the left and right hemispheres. Our findings may ultimately provide clues to understand why some brain networks are more vulnerable to or more resilient against functional disruption due to brain disorders or injuries.

Methods and Materials

Data and data processing. The dataset was provided by the Human Connectome Project (HCP; <http://www.humanconnectome.org>) from the Washington University–University of Minnesota (WUMinn) consortium⁴⁵, acquired using a modified 3T Siemens Skyra scanner with a 32-channel head coil. Resting-state fMRI data in an eyes-open condition were collected for approximately 14 min (1,200 time points) with TR = 720 ms, TE = 33.1 ms, flip angle = 52, voxel size = 2 mm isotropic, and FOV = 208 × 180 mm² and 72 slices. The data were acquired with opposing phase encoding directions, left-to-right (LR) in one run and right-to-left (RL) in the other run. Scanning parameters of a T₁-weighted structural image were TR = 2,400 ms, TE = 2.14 ms, flip angle = 8, voxel size = 0.7 mm isotropic, FOV = 224 × 224 mm² and 320 slices. Diffusion-weighted images (DWI) were acquired with 270 gradient directions with *b*-values 1000, 2000, 3000 s/mm², two repeats, and in a total of 36 *b*₀ scans: TR = 5520 ms, TE = 89.5 ms, flip angle = 78, FOV = 210 × 180 mm², 111 slices, and voxel size = 1.25 mm isotropic. A T₁-weighted structural image was acquired with TR = 2400 ms, TE = 2.14 ms, flip angle = 8, FOV = 224 × 224 mm², 320 slices, and voxel size = 0.7 mm isotropic. From the minimally preprocessed DWI data, white matter fibers were reconstructed using generalized *q*-sampling imaging⁴⁷ and a modified deterministic streamline tractography⁴⁸ using DSI studio (<http://dsi-studio.labsolver.org>)^{49–52}. Resting-state fMRI Data were realigned and co-registered to the T₁-weighted structural image using FSL⁵³. Linear trends and first order drifts with global effects were removed by regressing out the white matter, ventricle, global mean signals, together with 6 motion parameters using Matlab R2016b (Mathworks Inc., Natick, MA). White matter and ventricle voxels were determined based on the T1 FreeSurfer segmentations⁵⁴. Next, time-series were band-pass filtered with a Butterworth filter (0.01–0.1 Hz)^{55,56}. Motion scrubbing was used to remove scan frames when significant movement was detected in the individual time series^{56,57}. Our study included 484 participants in total from the Q4 release of HCP data.

Structural connectivity (SC) and functional connectivity (FC). Both structural networks and functional networks consisted of 219 cortical nodes using a subdivision parcellation⁵⁸ of the Desikan–Killiany atlas⁵⁹ excluding subcortical areas. For the structural networks, the edge weights were defined by the streamline count between two ROIs derived from diffusion MRI tractography and the edge weights for the functional networks were estimated as regularized partial correlation coefficients for each individual (see Section 2.6 for details). For both SC and FC, networks were constructed for each individual resulting in 484 FC and 484 SC networks and these individual networks were used for subsequent analyses.

Interdependent relationship between SC and FC. We model the interdependency of SC and FC using a multi-layer network approach^{22,60}. SC and FC form two separate layers that are linked by multiplex coupling, such that a node (ROI) in one layer is connected to the same node in the other layer in a one-to-one correspondence (Fig. 1, left). Building on previous work that has shown significant interactions between SC and FC^{1–5}, the

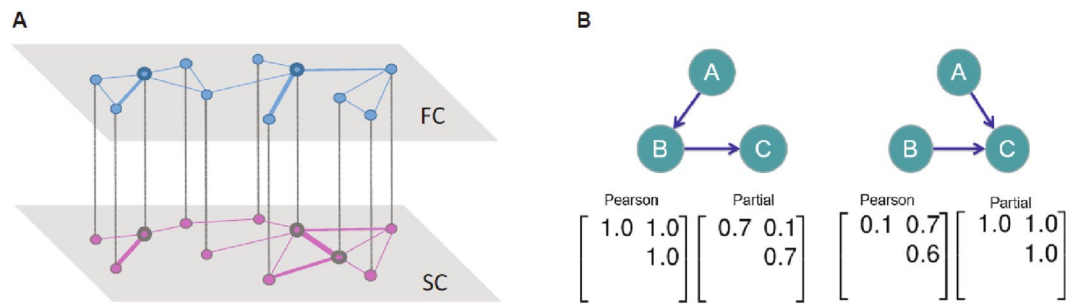


Figure 1. (A) A schematic representation of a brain multiplex network, where the networks of functional connectivity (FC) and structural connectivity (SC) are coupled via one-to-one match between corresponding region of interests; hub nodes in the SC (denoted as bigger circles) may not coincide with high-degree nodes in the FC; nodes linked in SC may not co-activate strongly, resulting in the absence of the edge in FC. (B) Limitations of Pearson's correlation and partial correlation. Depending on the underlying neuronal circuits both approaches can result in undesirable connection weights in the functional graph. (Left) A case where Pearson's correlation coefficient fails to disregard a non-existent connection between node A and C. The visualization stands for the underlying probabilistic graphical model among the variables A, B, and C, with connections standing for dependencies among pairs of variables. The two matrices contain the coefficients of Pearson's correlation and partial correlation coefficients, respectively. Rows and columns of the matrices refer to alphabetically ordered pairs of variable. Given their symmetry, we show only the upper-triangle of the matrices. (Right) A case where the partial correlation coefficient counter-intuitively imposes a high weight on the connection between node A and B due to the dependence on C. (Adapted from Nie 2015⁸¹).

multi-layer approach is designed to consider SC and FC as interdependent networks with a one-to-one correspondence. In this study, we consider a two-layer multiplex network with unweighted dependency links between layers for simplicity with a minimal set of assumptions; however, one could extend the model to represent a more general multi-layer network framework for future studies.

Estimation of functional connectivity from resting-state fMRI. Here, we use ROI-based FC in order to make a direct connection between SC and FC, in particular, we use regularized partial correlation coefficients based on a regression approach⁶¹. In the following sections, we elaborate on why we chose to use regularized partial correlation coefficients. Many different measures have been used to estimate or construct functional brain networks^{62–64}. Among them, the Pearson's moment correlation coefficient has been the most popular choice among brain researchers, and, despite its simplicity, it has provided valuable information regarding the intrinsic functional organization of the brain⁶⁵. However, other measures for estimating pairwise functional connectivity do exist, and they probe different aspects of dynamic interactions. Alternative choices include partial correlation coefficients, coherence measures that estimate linear relationships considering direct/indirect coupling effects in either the time domain or in the frequency domain, or non-linear measures such as mutual information, second-order maximum entropy or generalized synchronization^{65,66}. In addition to the aforementioned ROI-based definition for functional networks, spatial ICA-based functional networks are also widely used^{67–69}. Importantly, network properties of FC may differ depending on which dependency/synchrony measure one chooses as each measure captures different aspects of the functional network⁶⁵. Spatial ICA and seed-based FC have shown to be similar in certain cases⁷⁰ and both have advantages and disadvantages (for more quantitative comparison between the two methods, see⁷¹). Importantly, network properties of FC may differ depending on which dependency/synchrony measure or which type of derived FC one chooses to construct as each method captures different aspects of the functional network^{19,65,66,72,73}.

Full correlation versus partial correlation. The utility of a partial correlation approach to functional brain networks derives from the capacity to remove indirect effects due to remote linear effects propagated from other regions^{74–79}. However, using partial correlation coefficients entails some other issues, such as requiring number of observations larger than the number of ROIs, potential overfitting and less stable estimation^{61,76,80–87}. On the positive side, partial correlation coefficients could estimate connection strengths between two brain regions that are conditionally independent with a small coefficient, reducing or removing indirect connections (often referred to as spurious connections) (Fig. 1, left). Nonetheless, when the underlying structure happens to involve conditional dependence, in other words, 'explaining away' phenomenon in the Bayesian modelling literature (probabilistic graphical models)⁸⁸, estimating partial correlation coefficients can cause Berkson's paradox, inducing a 'spurious' connection, which will not happen when we use full correlation coefficient estimation (Fig. 1B, right)⁸⁹. This can be partially solved by using regularized partial correlation coefficients⁸¹.

Regularized partial correlation coefficients using elastic net. Regularized partial correlation coefficient estimation has been proposed for constructing functional networks based on resting-state fMRI to overcome the limitation of partial correlation coefficient estimation, while measuring direct relationships between two brain areas^{61,76,80–87} and, more commonly, for constructing gene association networks⁶¹. Estimation of regularized partial correlation coefficients has been carried out by applying regularization on Gaussian Graphical Models (GGMs), which can be represented as a graph with edges estimating conditional dependence between nodes^{61,90}.

Of those regularized models, the elastic net has been shown to be a good model to estimate resting-state functional connectivity⁷⁶. Although providing sparser solutions which do not require further statistical thresholding, L_1 -norm regularization can only identify the number of functional connections that is less than or equal to the number of observations (time points) and can detect only a subset of connections when the time series are highly correlated⁹¹. On the other hand, L_2 -norm regularization does not shrink small values of coefficients to zero, and hence we may not achieve the desirable level of sparseness of the network⁹¹. We can overcome these limitations by using the elastic net regression, which uses penalization of both L_1 and L_2 norms, or a linear combination of L_1 and L_2 norm regularization by solving the following problem⁹²:

$$\min_{\beta_0, \beta} \frac{1}{N} \sum_{i=1}^N w_i l(y_i, \beta_0 + \beta^T x_i) + \lambda \left[\frac{1}{2} (1 - \alpha) \|\beta\|_2^2 + \alpha \|\beta\|_1 \right] \quad (1)$$

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} \|\mathbf{y} - \mathbf{X}\beta\|^2 + \lambda [(1 - \alpha) \|\beta\|^2 + \alpha \|\beta\|_1] \quad (2)$$

where l is the negative log-likelihood function and w_i is the contribution of it for an observation x_i . While trying to minimize our objective function (equation (1)), we need to optimize our parameters λ and α . λ controls the overall penalization of the model and α determines how much we would put weight on L_1 -regularization compared to L_2 regularization. For example, if α is 1, our model becomes LASSO or L_1 regularization model, which will give us the sparsest graph. There are several methods to identify optimal parameter values such as grid search, which is slow and unstable because grid density affects the accuracy and depends on heuristic choices for parameter ranges. Alternatively, one could also use a stability selection method, which aims to control the false discovery rate⁸⁶ to determine the proper amount of regularization. Here, we made use of the interval search EPSGO algorithm to tune our parameters λ and α based on 10-fold cross-validation. This algorithm learns a Gaussian process model of the loss function surface in parameter space and samples at points where the expected improvement criterion is maximal^{93–95}. After calculating regularized β s (coefficients of predictors) with the optimized regularization parameters, we obtain partial correlation coefficients from β s^{61,90}. All estimation and optimization are done for each individual network.

$$\hat{\rho}_{ij} = \operatorname{sign}(\hat{\beta}_j^{(i)}) \min \left\{ 1, \sqrt{\frac{\hat{\beta}_j^{(i)} \hat{\beta}_i^{(j)}}{\hat{\beta}_j^{(i)} \hat{\beta}_i^{(j)}}} \right\} \quad \text{if } \operatorname{sign}(\hat{\beta}_j^{(i)}) = \operatorname{sign}(\hat{\beta}_i^{(j)})$$

otherwise 0

(3)

where $\hat{\beta}_j^{(i)}$ is the regularized estimate of β s between brain region i and the rest of the brain regions except the region i , which ensures the partial correlation coefficients are well-defined and in the interval $[-1, 1]$.

Assortativity mixing within and between layers. Assortativity quantifies the tendency for nodes to connect to other nodes that are similar in some way^{23,25}. For example, modularity Q ^{23,96} is an assortativity-based measure which expresses the actual connection density of nodes within the same community compared to the value of the connection density expected in a suitably defined null model. We can also measure the tendency of ‘similar’ nodes being connected to each other (actual vs. expected) based on some scalar nodal attribute such as degree, or betweenness. One of the most common cases where we define assortative mixing according to a scalar quantity is assortativity mixing by degree; positive degree assortativity implies that high-degree nodes are preferentially connected to high-degree nodes on average and low-degree nodes mainly connect to low-degree nodes on average. Since FC edges can carry either positive or negative weights, we considered a version of assortativity that takes into account node strengths, called node strength assortativity. The strength of a node is defined by the sum of its all weights⁹⁷. In FC a node’s strength is close to zero when its neighbors maintain positive and negative weights that nearly balance out. In this study, high values of the assortativity for the FC layer reflect a connectivity pattern where nodes with high strength are tendentially connected, through positively valued edges, to nodes with high strength. We did not weight the connection strength between two nodes, rather we measured the strength correlation between two nodes. In other words, we calculated Pearson’s correlation coefficient between a pair of nodes based on their strengths:

$$r = \frac{\sum_{ij} (A_{ij} - k_i k_j / 2m) s_i s_j}{\sum_{ij} (k_i \delta_{ij} - k_i k_j / 2m) s_i s_j} \quad (4)$$

where k_i is the degree of node i and $2m = \sum_i k_i$. $A_{ij} = 1$ if a connection between nodes i and j exists, otherwise $A_{ij} = 0$. $s_i = \sum_j w_{ij}$ is the sum of the weights of all connections departing from node i . $\delta_{ij} = 1$ if $i = j$ and $\delta_{ij} = 0$, otherwise. The numerator of equation (4), is the covariance of the pair of s_i and s_j on the edge (i, j) averaged over all pairs of edges. We define the mean μ of the s_i at the end of an edge as $\mu = \frac{\sum_{ij} A_{ij} s_i}{\sum_{ij} A_{ij}} = \frac{\sum_i k_i s_i}{\sum_i k_i} = \frac{1}{2m} \sum_i k_i s_i$, which is the average over edges rather than over all vertices. Then the covariance of s_i and s_j over edges is the following.

$$\begin{aligned}
cov(s_i, s_j) &= \frac{\sum_{ij} A_{ij} (s_i - \mu)(s_j - \mu)}{\sum_{ij} A_{ij}} \\
&= \frac{1}{2m} \sum_{ij} A_{ij} (s_i s_j - \mu s_i - \mu s_j + \mu^2) \\
&= \frac{1}{2m} \sum_{ij} A_{ij} s_i s_j - \mu^2 \\
&= \frac{1}{2m} \sum_{ij} A_{ij} s_i s_j - \frac{1}{(2m)^2} \sum_{ij} k_i k_j s_i s_j \\
&= \frac{1}{2m} \sum_{ij} \left(A_{ij} - \frac{k_i k_j}{2m} \right) s_i s_j
\end{aligned} \tag{5}$$

To normalize this, we divide equation (5) by the following equation (6) where all edges connect two nodes with the equal values of s_i . When we replace s_j with s_i , we have

$$\frac{1}{2m} \sum_{ij} \left(A_{ij} s_i^2 - \frac{k_i k_j}{2m} s_i s_j \right) = \frac{1}{2m} \sum_{ij} \left(k_i \delta_{ij} - \frac{k_i k_j}{2m} \right) s_i s_j \tag{6}$$

Thus, our node strength assortativity is $r = cov(s_i, s_j) / var(s_i)$, which is equation (4)⁹⁸. We measure the coefficient r within the SC and FC layers. Furthermore, we subdivide our networks into Yeo's 7 networks⁴⁶ (See Fig. S1) and investigated strength assortativity within each subnetwork for FC and SC as well as the strength assortativity between FC and SC for each network. In addition, we compared the left and the right hemispheres for each subnetwork.

Statistical analysis. Paired permutation test was used to compare the left and right hemisphere median differences for each subnetwork⁹⁹ by approximating the exact conditional distribution using conditional Monte Carlo procedures (10000 permutations) and corrected by Bonferroni method¹⁰⁰. A two-tailed test was used with alpha level 0.05 and adjusted p -values were reported based on Bonferroni correction. All statistical tests and calculations were performed using Matlab R2016b (Mathworks Inc., Natick, MA) with Brain Connectivity Toolbox¹⁰¹ and R with R packages (<http://www.R-project.org/>)¹⁰².

Results

Assortative within-layer structural connectivity (SC), disassortative within-layer functional connectivity (FC) and between-layer assortativity (SC and FC). We first calculated node strength assortativity in each individual data set across the entire cerebral cortex. Within each layer, FC was characterized by disassortative connectivity (i.e., negative assortativity, Fig. 2C Left), while SC showed assortative connectivity with a somewhat broader variability among individuals (Fig. 2C Right). The coupling between FC and SC assuming multiplexity between the two layers demonstrated a weak but positive assortativity in general (Fig. 2D). Moreover, we have explored the impact of the number of regions on the assortativity by using 68 cortical regions (Desikan-Killiany Atlas)⁵⁹, Lausanne 114 cortical regions and Lausanne 219 cortical regions⁵⁸ (Fig. S2) since degree disassortative networks have shown a decreasing degree disassortativity as the network size increases¹⁰³. The different sizes of the networks affected node strength assortativity coefficients; however, the disassortativity of FC remained the same and the contrasting assortativity properties between FC and SC were retained as well in three different network sizes (Fig. S2). Assortativity estimation based on Pearson's correlation coefficient is shown to be influenced by the distribution of the degrees of the nodes in the network as well, which is the main factor for the decreasing assortativity in larger networks with broader degree distributions¹⁰³. Thus, we examined the potential bias based on the difference in the ranges between SC and FC. In particular, we investigated if the disassortativity of FC was from a broader strength distribution of nodes in FC. In fact, the strength distribution of SC is much broader than that of FC. Hence, the smaller assortativity value (or disassortativity) for FC is not attributable to a wider range of the strength distribution of FC (Fig. S3). In addition, we have used Spearman's rank correlation coefficient (ρ) as recommended¹⁰³ to account for a possible bias due to the strength distribution differences between SC and FC and in different network sizes. Our results showed no qualitative differences when we used Spearman's ρ instead of Pearson's correlation coefficient (Fig. S4).

Within-layer assortativity between canonical resting-state networks. Next, we investigated if these assortativity patterns were distributed homogeneously across the whole brain or whether they exhibited local order within or between subnetworks. The overall disassortative mixing in FC was more prominent within functional subnetworks derived from the canonical Yeo parcellation (Fig. 3B), showing a contrast between diagonal (within a subnetwork) and off-diagonal (between subnetworks) elements. SC demonstrated overall assortative mixing between subnetworks (Fig. 3A), while showing a slightly higher range of assortativity within a subnetwork (diagonal) and some more strongly disassortative mixings between certain subnetworks.

Between-layer versus within-layer assortativity between 7 subnetworks. The overall assortative linkage between SC and FC could be also displayed between SC subnetworks and FC subnetworks assuming one-to-one correspondence (Fig. 4). We found that the magnitude of subnetwork coupling between SC and FC ranged fairly heterogeneous across different networks. For instance, Dorsal Attention and Fronto-Parietal networks exhibited between-layer assortativity that was twice as large as compared to the rest of the subnetworks.

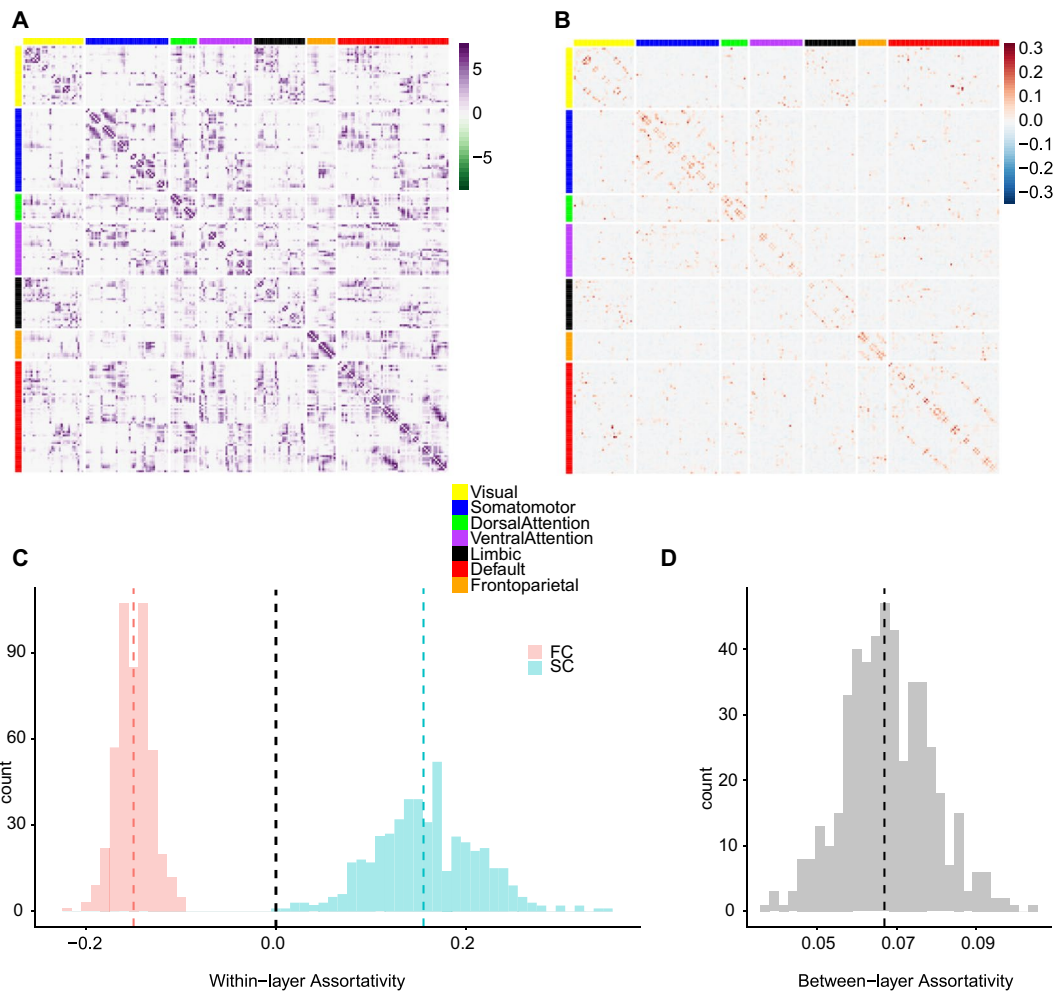


Figure 2. (A) Averaged SC across all subjects, with edges representing log transformed streamline counts. (B) Averaged FC across all subjects with edges representing regularized partial correlation coefficients. Nodes in both panels (A) and (B) are sorted by membership in 7 canonical resting-state networks and (A,B) are only shown for representative examples of SC and FC. For analyses, individual SC and FC were estimated and used in the study. (C) Within-layer assortativity: histograms (tallying numbers of individual subjects) of the strength assortativity within the functional network (FC) and the structural network (SC), respectively. Red: FC, Blue: SC. (D) Between-layer assortativity between FC and SC.

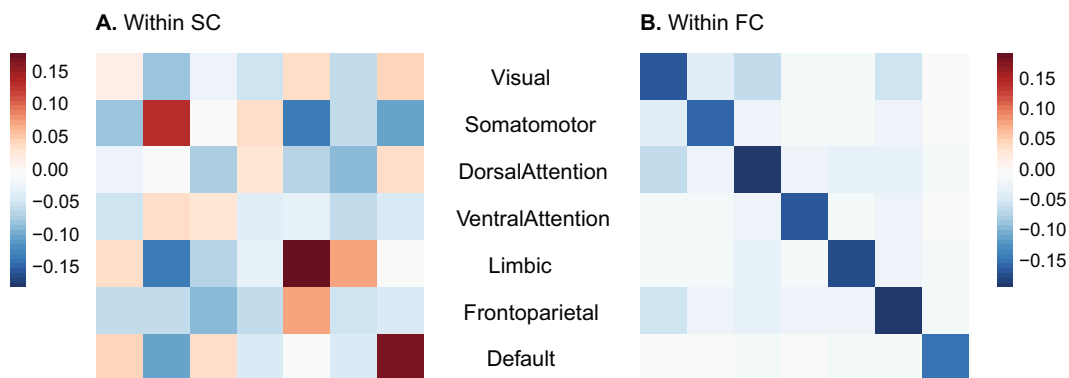


Figure 3. (A) Within-layer assortativity in SC between 7 canonical resting-state networks. (B) Within-layer assortativity in FC between 7 canonical resting-state networks.

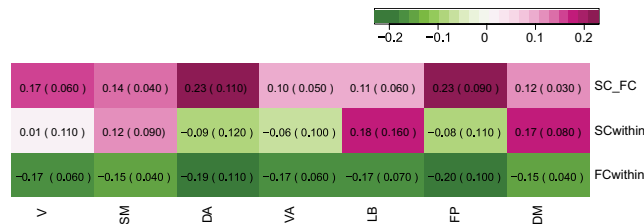


Figure 4. (First row) Between-layer assortativity between SC and FC, (Second row) Within-layer assortativity in SC (Third row) Within-layer assortativity in FC, numbers indicate median between-layer assortativity of the all subjects and numbers in the parentheses are median absolute deviation (MAD), V: Visual, SM: Somatomotor, DA: DorsalAttention, VA: VentralAttention, LB: Limbic, FP: FrontoParietal, and DM: Defaultmode network.

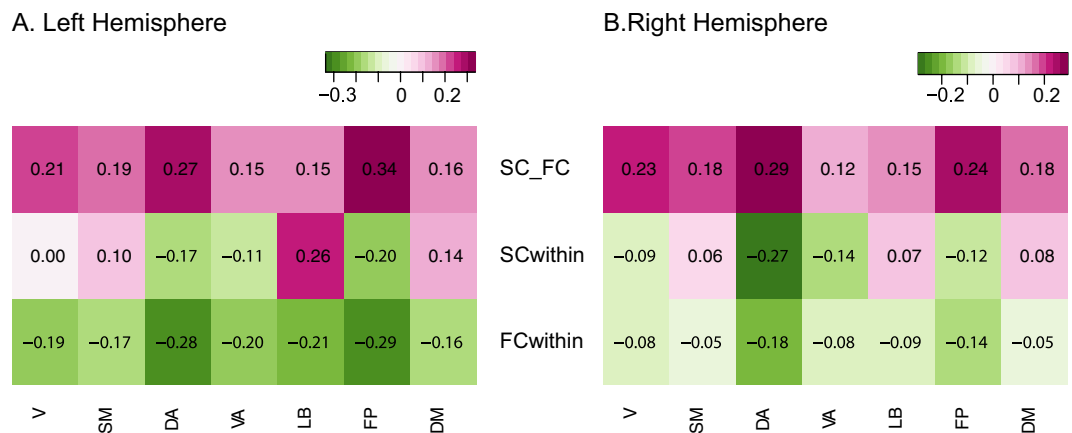


Figure 5. (A) Left Hemisphere, (B) Right Hemisphere. (First row) Between-layer assortativity between SC and FC, (Second row) Within-layer assortativity in SC, (Third row) Within-layer assortativity in FC, V: Visual, SM: Somatomotor, DA: Dorsal Attention, VA: Ventral Attention, LB: Limbic, FP: Fronto Parietal, and DM: Default mode network.

Interestingly, those networks with higher between-layer assortativity (Dorsal Attention and Fronto-Parietal) also showed strong disassortativity within the FC layer and strong disassortativity within the SC layer. We could differentiate within-layer SC assortativity for each subnetwork, showing both assortative and disassortative mixings for subnetworks as opposed to the overall assortative mixing pattern when aggregated in a single network. Within-layer FC assortativity showed disassortativity across all subnetworks although each network exhibited a different assortativity value.

Left versus right hemisphere assortativity differences both within and between layers. When we further investigated the above characteristics by separating the left and the right hemispheres, we found that the characteristic disassortative mixing in within FC layer is mainly driven by the left hemisphere (Figs 5 and 6). In fact, the subnetworks in the right hemisphere showed weak disassortativity within the right FC layer; Dorsal Attention and Fronto-Parietal networks still demonstrated stronger disassortativity compared to other networks but all subnetworks in the right hemisphere showed much weaker disassortativity than those of the left hemisphere (Fig. 6, all p -values $< 10^{-21}$ after Bonferroni adjustment). In contrast, those with higher between-layer assortativity (Dorsal Attention and Fronto-Parietal) in both hemispheres also displayed strong disassortativity within the FC layer and within the SC layer in the similar way when both hemispheres were aggregated (Fig. 5).

We quantified the contrasts between hemispheres in both within and between-layer assortativity using permutation test based on 10000 Monte Carlo resampled approximate distribution (See methods for details). Three common features were observed for all subnetworks (Fig. 6). The assortativity distributions for the left and right hemispheres showed a smaller difference between layers compared to within layer assortativity distributions except Fronto-Parietal and Default networks. Of note, the left and the right hemisphere differences showed opposite patterns between within SC and within FC; within SC, the right hemisphere was characterized with negative and smaller assortativity than those in the left hemisphere except Ventral Attention (not significant) and Fronto-Parietal networks (the trend is reversed) (Fig. 6). In addition, the stark difference between hemispheres demonstrated mainly within FC. Moreover, the pattern is the opposite of within SC difference, which is summarized as box plots in the last panel of Fig. 6; the left hemisphere showed strong disassortative connectivity and the right hemisphere showed weak disassortativity within the FC layer (Fig. 6). However, there are interesting differences among subnetworks. For instance, unlike other subnetworks, the frontoparietal network and Default network showed small but significant strong lateralization in terms of between-layer assortativity.

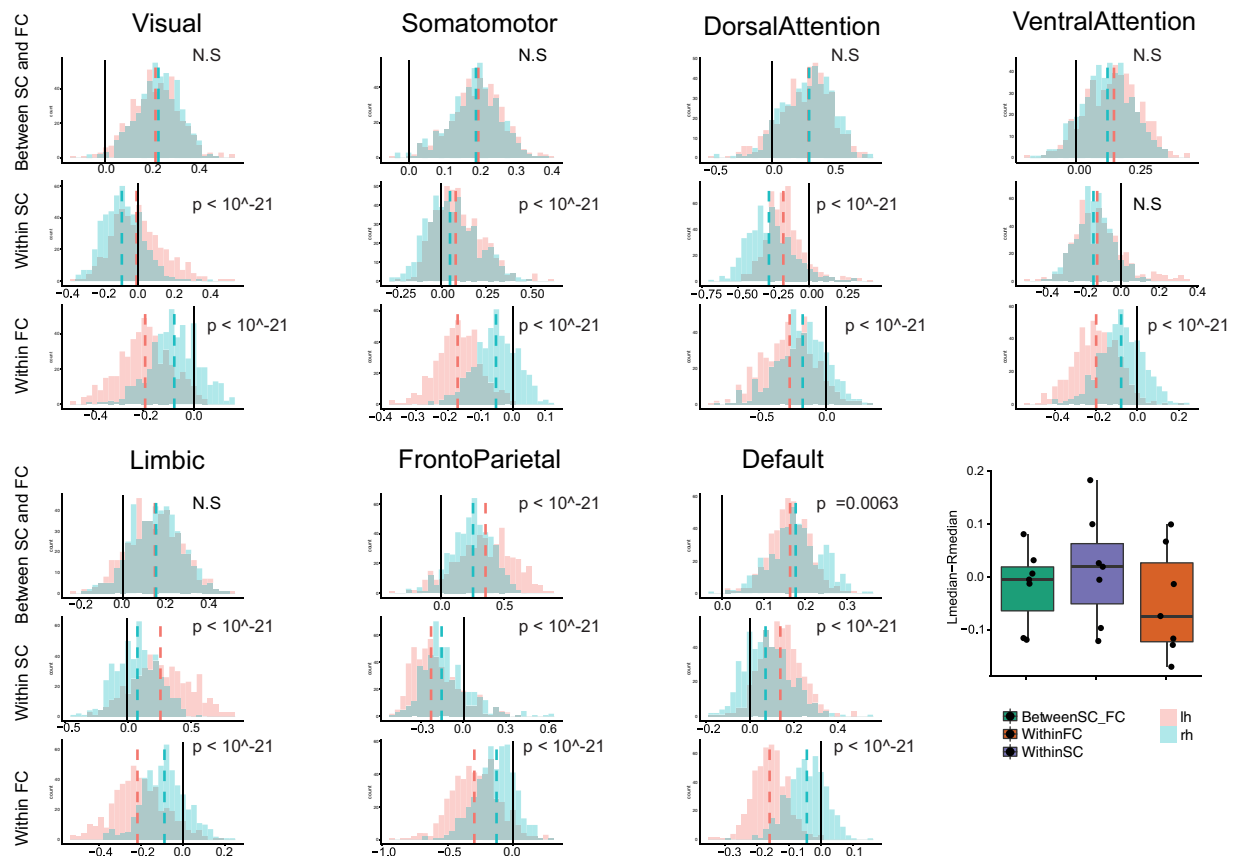


Figure 6. (First row) Between-layer assortativity between SC and FC, (Second row) Within-layer assortativity in SC, (Third row) Within-layer assortativity in FC (Box plots, Dashed lines indicate the medians of two distributions, Black solid line is added to identify zero on the horizontal axis. A summary box plot for the median differences between the left and the right hemispheres. X-axis: Green: between-layer assortativity, Purple: within-FC assortativity, Orange: within-SC assortativity, y-axis: the difference between the median of the left hemisphere distribution of assortativity and the median of the right hemisphere distribution of assortativity in each subnetwork.

Discussion

In this study, we constructed a two-layer multiplex interdependent network from structural and functional connectivity estimated from MRI data to examine if within- and between-layer node strength assortativity differs between SC and FC. In particular, we divided SC and FC into a canonical partition of seven resting state networks (RSNs) to examine within- and between-layer assortativity differences in subnetworks and hemispheric differences as well. We find that, in general, SC is organized in an assortative manner, indicating brain regions are, on average, connected to other brain regions with similar node strengths. On the other hand, FC showed disassortative mixing in node strength. More detailed analysis showed that this discrepancy between SC and FC assortativity was pronounced to a different extent within- and between- RSNs. In SC, brain regions within the same subnetwork are connected with similar node strengths; in contrast, in FC brain regions are more likely connected to brain regions with similar node strengths between subnetworks rather than within its own subnetwork. In addition, these patterns showed lateralization; the overall disassortative mixing within subnetworks in FC was mainly driven from the left hemisphere. The degree of laterality also showed differences among subnetworks.

Assortative SC and Disassortative FC. Degree assortativity, as typically applied in network science is computed as a global network metric and typically ranges between -0.3 and 0.3 ²³. Among biological networks, previous studies have shown disassortativity in several biological networks, including those defined by protein interactions²³. Instead, synaptic networks in *C. elegans*¹⁰⁴ and human structural brain networks estimated with diffusion weighted imaging and tractography^{105,106} are assortative, and this assortativity appears associated with the existence of modules¹⁰⁷. Functional connectivity has been reported to show assortative mixing¹⁰⁸ when edges are computed as standard Pearson correlations. Assortativity in FC networks rises in the course of epileptic seizures¹⁰⁹. To our knowledge, no previous study has examined assortative coupling within a two-layer multiplex SC/FC model. Assortative mixing in networks is known to confer greater robustness against random removal of nodes or edges compared to disassortative networks^{23,26,27}. On the other hand, when it comes to spreading infectious diseases or seizure activity, assortativity makes it easier for these disruptions to spread across the whole network^{23,26,27}. In this study, our goal was to see if the brain's functional and structural networks would show

tendencies that may promote network's robustness or resilience in line with the aforementioned theoretical works. The observed assortative and disassortative organization of SC and FC, respectively, may be relevant to posit a potential complementary interdependent relationship between the human brain's structure and function.

Between-layer assortativity between SC and FC. When more than two networks are coupled, or more generally, in a multi-layer network, the coupling between nodes in different layers affects the robustness of the system³⁷. When nodes in different layers are connected regardless of their degrees, cascading failures of the nodes can destroy the network easily because even if a low-degree node is removed, that node can be connected to a high-degree node in another layer, and its removal could thus fragment the network into disconnected parts^{31,37}. In general, when there is a positive correlation of the degree-degree coupling between layers, the interdependent networks are known to be more robust³¹. In this study, we find that overall SC and FC are coupled in a way that nodes with similar strengths are connected between layers in all subnetworks (Figs 3 and 4). This topological feature of the brain's multi-layer organization may explain the retention of the brain's functionality unless a significant volume of the brain has been affected for example in the course of progressive neurodegenerative disease^{110,111}.

Subnetwork differences of SC and FC in within- and between- layer node strength assortativity. After demonstrating these topological patterns within whole-brain networks, we carried out a more detailed analysis of specific RSNs, or subnetworks, to discern if these effects were predominantly found in specific subdivisions of the cerebral cortex. We found that Visual, Somatomotor, Limbic and Default Mode Networks displayed assortativity patterns within and between layers. In contrast, the Dorsal Attention (DA), Ventral Attention (VA) and Fronto Parietal networks showed disassortativity within both SC and FC layers. Moreover, we also find distinctive differences between Dorsal and Ventral Attention networks, for instance, the between-layer assortativity of the Dorsal Attention network averaged over all participants was approximately twice as high as that of the Ventral Attention network (Fig. 5); SC within-layer assortativity in the right and the left hemispheres showed significant differences in DA but not in VA (Fig. 6, which corroborates previous literature that showed distinctive functional and structural networks between DA and VA^{112–114}).

Possible implications for robustness of brain networks. Although, our study did not measure the brain's 'robustness' *per se*, we may speculate on potential implications of our results in the context of previous theoretical network science literature. For SC, to counter the effects of lesions from injuries or disease processes, having a connected network may be a priority to keep the flow of neural signals and processing as intact as possible, even if signaling paths or delays may increase due to the lesions. Hence, SC may need to be organized with positive assortativity to promote resilience. On the other hand, limiting the extent of shared neuronal information (as expressed in the statistical construct of partial correlations) or controlling the spread of abnormal brain activity such as seizures could be a significant aspect in the architecture of FC, with disassortativity helping FC to maintain its functionality against indiscriminate propagation of perturbations.

In some cases, most notably in unilateral neglect, lesions may have differential effects on behavior and cognition depending on the laterality of the lesion site. Right hemisphere lesions have been reported to accompany more frequent and severe unilateral spatial neglect up to 80% of the time^{115–118}. Hence, we were interested whether we could observe lateralization of the assortativity, which could be relevant to potential resilience differences in hemispheres. Indeed, the strong disassortativity in the FC layer observed in the whole brain mainly seems to derive from the left hemisphere, although both hemispheres show disassortativity in general within the FC layer. Stronger disassortativity in the left hemisphere may suggest greater robustness to the disruptive effect of the brain injuries – conversely, weaker disassortativity in the right hemisphere suggests greater vulnerability. Increased/decreased robustness in the left hemisphere may also be related to faster/slower recovery post-injury. Indeed, previous studies showed that a right hemisphere stroke is more likely to be followed by more severe spatial neglect and more prolonged recovery^{118–122}, although some studies have shown opposite or inconclusive results on hemispheric differences in functional recovery¹²³ (See the references therein). This pronounced lateralization of the FC layer can also be related to previous studies that showed disrupted laterality in the functional brain network in brain disorders^{124–126}. The Ventral Attention network (as opposed to Dorsal Attention network) showed a large difference in the disassortativity between the left and the right hemispheres; the right hemisphere VA exhibited much weaker disassortativity than that of the left hemisphere, which is consistent with the prevalence of spatial neglect when patients experienced strokes in the right hemisphere in regions associated with the VA^{113,127}. We note that our findings suggest potential hemispheric differences in robustness despite largely symmetric distributions of standard topological measures related to both SC and FC.

Limitations. There are several limitations of our study. First, there are many ways to estimate functional brain networks. Depending which preprocessing steps one chooses to use, the relationship of subnetworks can vary; for instance, negative correlations between some RSNs are observed only when global signal regression is applied¹²⁸. In our study, we adopted an approach to estimate partial correlations in order to allow the assessment of assortativity within a sparse functional network composed of functional links that express specific shared pairwise dependencies. More commonly used full correlation methods yield full networks and are prone to transitivity and spurious dependencies that artifactually boost shared variance. Second, assortativity could in principle be calculated based on other nodal attributes such as node between-ness or page-rank, and node strength can be also defined in different ways depending on how we define weights in FC. Third, as the assortativity is a global measure, estimating it within and between subnetworks might suffer as there are smaller numbers of nodes within each subnetwork than in the network as a whole. A main aim of the paper was to investigate the node strength assortativity in the functional subnetworks and the left and the right hemispheres. However, as brain regions are

heterogeneous even within functionally defined subnetworks, it is worth examining ROI-wise local assortativity. Some initial data on local assortativity results are included in the supplementary information (Figs S1, S6 and S7). These data may offer a starting point for a more detailed analysis and discussion of potential regional contributions to assortativity in the context of network robustness. Future studies could also provide more detailed analyses of various nodal attributes using alternative definitions of weighted assortativity and with different parameters during time series processing.

Conclusion

In this study, we have systematically examined topological discrepancies between SC and FC by estimating node strength assortativity, using a framework of two-layer multiplex interdependent networks. We find that SC is, in general, organized as an assortative network while FC is organized as a disassortative network, with assortative coupling between the layers, which resembles an arrangement that promotes robustness within the interdependent network considered as a multi-layer system in the theoretical framework. Moreover, we find differences in subnetworks for within and between layer assortativity. Finally, we find there is a characteristic lateralization of assortativity expressed in the FC layer. Our study may be a useful starting point to further investigate the robustness of human brain networks, which may ultimately allow predicting individual differences in the response to injury, recovery rate or prognosis.

References

1. Park, H.-J. & Friston, K. Structural and functional brain networks: from connections to cognition. *Science* **342**, 1238411 (2013).
2. Sporns, O. Structure and function of complex brain networks. *Dialogues in clinical neuroscience* **15**, 247 (2013).
3. Damoiseaux, J. S. Effects of aging on functional and structural brain connectivity. *NeuroImage* (2017).
4. Uddin, L. Q., Supekar, K. S., Ryali, S. & Menon, V. Dynamic reconfiguration of structural and functional connectivity across core neurocognitive brain networks with development. *Journal of Neuroscience* **31**, 18578–18589 (2011).
5. Mišić, B. *et al.* Network-level structure-function relationships in human neocortex. *Cerebral Cortex* **26**, 3285–3296 (2016).
6. Betzel, R. F. *et al.* Changes in structural and functional connectivity among resting-state networks across the human lifespan. *NeuroImage* **102**, 345–357 (2014).
7. Sporns, O. The human connectome: origins and challenges. *NeuroImage* **80**, 53–61 (2013).
8. Skudlarski, P. *et al.* Measuring brain connectivity: diffusion tensor imaging validates resting state temporal correlations. *NeuroImage* **43**, 554–561 (2008).
9. Koch, M. A., Norris, D. G. & Hund-Georgiadis, M. An investigation of functional and anatomical connectivity using magnetic resonance imaging. *NeuroImage* **16**, 241–250 (2002).
10. Damoiseaux, J. S. & Greicius, M. D. Greater than the sum of its parts: a review of studies combining structural connectivity and resting-state functional connectivity. *Brain Structure and Function* **213**, 525–533 (2009).
11. Messé, A., Rudrauf, D., Benali, H. & Marrelec, G. Relating structure and function in the human brain: relative contributions of anatomy, stationary dynamics, and non-stationarities. *PLoS computational biology* **10**, e1003530 (2014).
12. de Pasquale, F., Della Penna, S., Sabatini, U., Caravasso Falletta, C. & Peran, P. The anatomical scaffold underlying the functional centrality of known cortical hubs. *Human Brain Mapping* (2017).
13. Honey, C. J., Thivierge, J.-P. & Sporns, O. Can structure predict function in the human brain? *NeuroImage* **52**, 766–776 (2010).
14. Honey, C. *et al.* Predicting human resting-state functional connectivity from structural connectivity. *Proceedings of the National Academy of Sciences* **106**, 2035–2040 (2009).
15. Breakspear, M. Dynamic models of large-scale brain activity. *Nature neuroscience* **20**, 340–352 (2017).
16. Deco, G. *et al.* Resting-state functional connectivity emerges from structurally and dynamically shaped slow linear fluctuations. *Journal of Neuroscience* **33**, 11239–11252 (2013).
17. Goñi, J. *et al.* Resting-brain functional connectivity predicted by analytic measures of network communication. *Proceedings of the National Academy of Sciences* **111**, 833–838 (2014).
18. van den Heuvel, M. P. & Sporns, O. Network hubs in the human brain. *Trends in cognitive sciences* **17**, 683–696 (2013).
19. Zalesky, A., Fornito, A. & Bullmore, E. On the use of correlation as a measure of network connectivity. *NeuroImage* **60**, 2096–2106 (2012).
20. Buldyrev, S. V., Parshani, R., Paul, G., Stanley, H. E. & Havlin, S. Catastrophic cascade of failures in interdependent networks. *Nature* **464**, 1025–1028 (2010).
21. Boccaletti, S. *et al.* The structure and dynamics of multilayer networks. *Physics Reports* **544**, 1–122 (2014).
22. Kivela, M. *et al.* Multilayer networks. *Journal of complex networks* **2**, 203–271 (2014).
23. Newman, M. E. Mixing patterns in networks. *Physical Review E* **67**, 026126 (2003).
24. Callaway, D. S., Newman, M. E. J., Strogatz, S. H. & Watts, D. J. Network robustness and fragility: percolation on random graphs. *Phys. Rev. Lett.* **85**, 5468–5471 (2000).
25. Noldus, R. & Van Mieghem, P. Assortativity in complex networks. *Journal of Complex Networks* **3**, 507–542 (2015).
26. Pechenick, D. A., Payne, J. L. & Moore, J. H. The influence of assortativity on the robustness of signal-integration logic in gene regulatory networks. *Journal of theoretical biology* **296**, 21–32 (2012).
27. Vázquez, A. & Moreno, Y. Resilience to damage of graphs with degree correlations. *Physical Review E* **67**, 015101 (2003).
28. Nicosia, V. & Latora, V. Measuring and modeling correlations in multiplex networks. *Physical Review E* **92**, 032805 (2015).
29. de Arruda, G. F., Cozzo, E., Moreno, Y. & Rodrigues, F. A. On degree–degree correlations in multilayer networks. *Physica D: Nonlinear Phenomena* **323**, 5–11 (2016).
30. Zhou, D., Stanley, H. E., D’Agostino, G. & Scala, A. Assortativity decreases the robustness of interdependent networks. *Physical Review E* **86**, 066103 (2012).
31. Reis, S. D. *et al.* Avoiding catastrophic failure in correlated networks of networks. *Nature Physics* **10**, 762–767 (2014).
32. Cellai, D., López, E., Zhou, J., Gleeson, J. P. & Bianconi, G. Percolation in multiplex networks with overlap. *Physical Review E* **88**, 052811 (2013).
33. Min, B., Lee, S., Lee, K.-M. & Goh, K.-I. Link overlap, viability, and mutual percolation in multiplex networks. *Chaos, Solitons & Fractals* **72**, 49–58 (2015).
34. Radicchi, F. Percolation in real interdependent networks. *Nature Physics* **11**, 597 (2015).
35. Radicchi, F. & Bianconi, G. Redundant interdependencies boost the robustness of multiplex networks. *Physical Review X* **7**, 011013 (2017).
36. Baxter, G. J., Bianconi, G., da Costa, R. A., Dorogovtsev, S. N. & Mendes, J. F. Correlated edge overlaps in multiplex networks. *Physical Review E* **94**, 012303 (2016).
37. Danziger, M. M., Shekhtman, L. M., Bashan, A., Berezin, Y. & Havlin, S. Vulnerability of interdependent networks and networks of networks. In *Interconnected Networks*, 79–99 (Springer, 2016).

38. Kleineberg, K.-K., Boguná, M., Serrano, M. Á. & Papadopoulos, F. Hidden geometric correlations in real multiplex networks. *Nature Physics* **12**, 1076 (2016).
39. Kleineberg, K.-K., Buzna, L., Papadopoulos, F., Boguñá, M. & Serrano, M. Á. Geometric correlations mitigate the extreme vulnerability of multiplex networks against targeted attacks. *Physical review letters* **118**, 218301 (2017).
40. Petrosini, L. Neurobiological and psychological aspects of brain recovery (2017).
41. Aerts, H., Fias, W., Caeyenberghs, K. & Marinazzo, D. Brain networks under attack: robustness properties and the impact of lesions. *Brain* **139**, 3063–3083 (2016).
42. Bullmore, E. & Sporns, O. The economy of brain network organization. *Nature Reviews Neuroscience* **13**, 336–349 (2012).
43. Lee, D. *et al.* Analysis of structure–function network decoupling in the brain systems of spastic diplegic cerebral palsy. *Human brain mapping* **38**, 5292–5306 (2017).
44. Odding, E., Roebroek, M. E. & Stam, H. J. The epidemiology of cerebral palsy: incidence, impairments and risk factors. *Disability and rehabilitation* **28**, 183–191 (2006).
45. Van Essen, D. C. *et al.* The wu-minn human connectome project: an overview. *Neuroimage* **80**, 62–79 (2013).
46. Yeo, B. T. *et al.* The organization of the human cerebral cortex estimated by intrinsic functional connectivity. *Journal of neurophysiology* **106**, 1125–1165 (2011).
47. Yeh, F., Wedeen, V. & Tseng, W. Generalized q-sampling imaging. *IEEE transactions on medical imaging* **29**, 1626–1635 (2010).
48. Yeh, F.-C., Verstynen, T. D., Wang, Y., Fernández-Miranda, J. C. & Tseng, W.-Y. I. Deterministic diffusion fiber tracking improved by quantitative anisotropy. *PloS one* **8**, e80713 (2013).
49. de Reus, M. A. & van den Heuvel, M. P. Estimating false positives and negatives in brain networks. *Neuroimage* **70**, 402–409 (2013).
50. de Reus, M. A. & van den Heuvel, M. P. Simulated rich club lesioning in brain networks: a scaffold for communication and integration? *Frontiers in human neuroscience* **8** (2014).
51. van den Heuvel, M. P., Scholtens, L. H., de Reus, M. A. & Kahn, R. S. Associated microscale spine density and macroscale connectivity disruptions in schizophrenia. *Biological psychiatry* **80**, 293–301 (2016).
52. van den Heuvel, M. P., Scholtens, L. H., Barrett, L. F., Hilgetag, C. C. & de Reus, M. A. Bridging cytoarchitectonics and connectomics in human cerebral cortex. *Journal of Neuroscience* **35**, 13943–13948 (2015).
53. Jenkinson, M., Beckmann, C. F., Behrens, T. E., Woolrich, M. W. & Smith, S. M. Fsl. *Neuroimage* **62**, 782–790 (2012).
54. Fischl, B. Freesurfer. *Neuroimage* **62**, 774–781 (2012).
55. van den Heuvel, M. P. *et al.* Proportional thresholding in resting-state fmri functional connectivity networks and consequences for patient-control connectome studies: Issues and recommendations. *Neuroimage* **152**, 437–449 (2017).
56. van den Heuvel, M. P. *et al.* Abnormal rich club organization and functional brain dynamics in schizophrenia. *JAMA psychiatry* **70**, 783–792 (2013).
57. Power, J. D., Barnes, K. A., Snyder, A. Z., Schlaggar, B. L. & Petersen, S. E. Spurious but systematic correlations in functional connectivity mri networks arise from subject motion. *Neuroimage* **59**, 2142–2154 (2012).
58. Cammoun, L. *et al.* Mapping the human connectome at multiple scales with diffusion spectrum mri. *Journal of neuroscience methods* **203**, 386–397 (2012).
59. Desikan, R. S. *et al.* An automated labeling system for subdividing the human cerebral cortex on mri scans into gyral based regions of interest. *Neuroimage* **31**, 968–980 (2006).
60. Boccaletti, S. Complex networks: structure and dynamics. *Phys. Rep.* **424**, 175–308 (2004).
61. Krämer, N., Schäfer, J. & Boulesteix, A.-L. Regularized estimation of large-scale gene association networks using graphical gaussian models. *BMC Bioinformatics* **10**, 384, <https://doi.org/10.1186/1471-2105-10-384> (2009).
62. Bullmore, E. & Sporns, O. Complex brain networks: graph theoretical analysis of structural and functional systems. *Nature Reviews Neuroscience* **10**, 186–198 (2009).
63. Deco, G., Jirsa, V. K. & McIntosh, A. R. Emerging concepts for the dynamical organization of resting-state activity in the brain. *Nature Reviews Neuroscience* **12**, 43–56 (2011).
64. Friston, K. J., Harrison, L. & Penny, W. Dynamic causal modelling. *Neuroimage* **19**, 1273–1302 (2003).
65. Wang, H. E. *et al.* A systematic framework for functional connectivity measures. *Frontiers in neuroscience* **8** (2014).
66. Simpson, S. L., Bowman, F. D. & Laurienti, P. J. Analyzing complex functional brain networks: fusing statistics and network science to understand the brain. *Statistics surveys* **7**, 1 (2013).
67. van de Ven, V. G., Formisano, E., Prvulovic, D., Roeder, C. H. & Linden, D. E. Functional connectivity as revealed by spatial independent component analysis of fmri measurements during rest. *Human brain mapping* **22**, 165–178 (2004).
68. Calhoun, V. D., Adali, T., Pearlson, G. D. & Pekar, J. A method for making group inferences from functional mri data using independent component analysis. *Human brain mapping* **14**, 140–151 (2001).
69. Smith, S. M. *et al.* Functional connectomics from resting-state fmri. *Trends in cognitive sciences* **17**, 666–682 (2013).
70. Van Dijk, K. R. *et al.* Intrinsic functional connectivity as a tool for human connectomics: theory, properties, and optimization. *Journal of neurophysiology* **103**, 297–321 (2010).
71. Joel, S. E., Caffo, B. S., van Zijl, P. & Pekar, J. J. On the relationship between seed-based and ica-based measures of functional connectivity. *Magnetic Resonance in Medicine* **66**, 644–657 (2011).
72. Liang, X. *et al.* Effects of different correlation metrics and preprocessing factors on smallworld brain functional networks: a resting-state functional mri study. *PloS one* **7**, e32766 (2012).
73. Jalili, M. Functional brain networks: does the choice of dependency estimator and binarization method matter? *Scientific reports* **6**, 29780 (2016).
74. van den Heuvel, M., Mandl, R., Luigjes, J. & Pol, H. H. Microstructural organization of the cingulum tract and the level of default mode functional connectivity. *Journal of Neuroscience* **28**, 10844–10851 (2008).
75. Supekar, K. *et al.* Development of functional and structural connectivity within the default mode network in young children. *Neuroimage* **52**, 290–301 (2010).
76. Ryali, S., Chen, T., Supekar, K. & Menon, V. Estimation of functional connectivity in fmri data using stability selection-based sparse partial correlation with elastic net penalty. *Neuroimage* **59**, 3852–3861 (2012).
77. Hampson, M., Peterson, B. S., Skudlarski, P., Gatenby, J. C. & Gore, J. C. Detection of functional connectivity using temporal correlations in mr images. *Human brain mapping* **15**, 247–262 (2002).
78. Marrelec, G. *et al.* Partial correlation for functional brain interactivity investigation in functional mri. *Neuroimage* **32**, 228–237 (2006).
79. Smith, S. M. *et al.* Network modelling methods for fmri. *Neuroimage* **54**, 875–891 (2011).
80. Smith, S. M. The future of fmri connectivity. *Neuroimage* **62**, 1257–1266 (2012).
81. Nie, L., Yang, X., Matthews, P. M., Xu, Z. & Guo, Y. *Minimum Partial Correlation: An Accurate and Parameter-Free Measure of Functional Connectivity in fMRI*, <https://doi.org/10.1007/978-3-319-23344-413>, 125–134 (Springer International Publishing, Cham, 2015).
82. Friedman, J., Hastie, T. & Tibshirani, R. Sparse inverse covariance estimation with the graphical lasso. *Biostatistics* **9**, 432–441 (2008).
83. Huang, S. *et al.* Learning brain connectivity of alzheimer's disease by sparse inverse covariance estimation. *NeuroImage* **50**, 935–949 (2010).

84. Lee, H., Lee, D. S., Kang, H., Kim, B.-N. & Chung, M. K. Sparse brain network recovery under compressed sensing. *IEEE Transactions on Medical Imaging* **30**, 1154–1165 (2011).
85. Peng, J., Wang, P., Zhou, N. & Zhu, J. Partial correlation estimation by joint sparse regression models. *Journal of the American Statistical Association* **104**, 735–746 (2009).
86. Meinshausen, N. & Bühlmann, P. Stability selection. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)* **72**, 417–473 (2010).
87. Varoquaux, G., Gramfort, A., Poline, J.-B. & Thirion, B. Brain covariance selection: better individual functional connectivity models using population prior. In *Advances in neural information processing systems*, 2334–2342 (2010).
88. Pearl, J. *Probabilistic Reasoning in Intelligent Systems: Networks of Plausible Inference* (Morgan Kaufmann Publishers Inc., 1988).
89. Berkson, J. Limitations of the application of fourfold table analysis to hospital data. *Biometrics Bulletin* **2**, 47–53 (1946).
90. Whittaker, J. *Graphical models in applied multivariate statistics* (Wiley Publishing, 2009).
91. Zou, H. & Hastie, T. Regularization and variable selection via the elastic net. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)* **67**, 301–320 (2005).
92. Friedman, J., Hastie, T. & Tibshirani, R. Regularization paths for generalized linear models via coordinate descent. *Journal of statistical software* **33**, 1 (2010).
93. Sill, M. *et al.* c060: Extended inference with lasso and elastic-net regularized cox and generalized linear models. *Journal of Statistical Software* **62**, 1–22 (2014).
94. Frohlich, H. & Zell, A. Efficient parameter selection for support vector machines in classification and regression via model-based global optimization. In *Proceedings. 2005 IEEE International Joint Conference on Neural Networks, 2005.*, vol. 3, 1431–1436 (IEEE, 2005).
95. Jones, D. R., Schonlau, M. & Welch, W. J. Efficient global optimization of expensive black-box functions. *Journal of Global optimization* **13**, 455–492 (1998).
96. Newman, M. E. & Girvan, M. Finding and evaluating community structure in networks. *Physical review E* **69**, 026113 (2004).
97. Barrat, A., Barthélemy, M., Pastor-Satorras, R. & Vespignani, A. The architecture of complex weighted networks. *PNAS* **101**, 3747–3752 (2004).
98. Newman, M. *Networks: an introduction* (Oxford university press, 2010).
99. Strasser, H. & Weber, C. On the asymptotic theory of permutation statistics. (1999).
100. Dunn, O. J. Estimation of the medians for dependent variables. *The Annals of Mathematical Statistics* 192–197 (1959).
101. Rubinov, M. & Sporns, O. Complex network measures of brain connectivity: uses and interpretations. *Neuroimage* **52**, 1059–1069 (2010).
102. R Core Team. *R: A Language and Environment for Statistical Computing*. R Foundation for Statistical Computing, Vienna, Austria (2013).
103. Litvak, N. & Van Der Hofstad, R. Uncovering disassortativity in large scale-free networks. *Physical Review E* **87**, 022801 (2013).
104. Chatterjee, N. & Sinha, S. Understanding the mind of a worm: hierarchical network structure underlying nervous system function in *C. elegans*. *Progress in brain research* **168**, 145–153 (2007).
105. Hagmann, P. *et al.* Mapping the structural core of human cerebral cortex. *PLoS biology* **6**, e159 (2008).
106. Van Den Heuvel, M. P. & Sporns, O. Rich-club organization of the human connectome. *Journal of Neuroscience* **31**, 15775–15786 (2011).
107. Avalos-Gaytán, V., Almendral, J. A., Papo, D., Schaeffer, S. E. & Boccaletti, S. Assortative and modular networks are shaped by adaptive synchronization processes. *Physical Review E* **86**, 015101 (2012).
108. Eguiluz, V. M., Chialvo, D. R., Cecchi, G. A., Baliki, M. & Apkarian, A. V. Scale-free brain functional networks. *Physical review letters* **94**, 018102 (2005).
109. Bialonski, S. & Lehnertz, K. Assortative mixing in functional brain networks during epileptic seizures. *Chaos: An Interdisciplinary Journal of Nonlinear Science* **23**, 033139 (2013).
110. Valenzuela, M. J. & Sachdev, P. Brain reserve and dementia: a systematic review. *Psychological medicine* **36**, 441–454 (2006).
111. Yoo, S. W. *et al.* A network flow-based analysis of cognitive reserve in normal ageing and Alzheimer's disease. *Scientific reports* **5** (2015).
112. Fox, M. D., Corbetta, M., Snyder, A. Z., Vincent, J. L. & Raichle, M. E. Spontaneous neuronal activity distinguishes human dorsal and ventral attention systems. *Proceedings of the National Academy of Sciences* **103**, 10046–10051 (2006).
113. He, B. J. *et al.* Breakdown of functional connectivity in frontoparietal networks underlies behavioral deficits in spatial neglect. *Neuron* **53**, 905–918 (2007).
114. Vossel, S., Geng, J. J. & Fink, G. R. Dorsal and ventral attention systems: distinct neural circuits but collaborative roles. *The Neuroscientist* **20**, 150–159 (2014).
115. Heilman, K. M., Valenstein, E. & Watson, R. T. Neglect and related disorders. In *Seminars in Neurology*, vol. 4, 209–219 (© 1984 by Thieme Medical Publishers, Inc., 1984).
116. Hillis, A. E. Neurobiology of unilateral spatial neglect. *The neuroscientist* **12**, 153–163 (2006).
117. Karnath, H.-O., Fruhmann Berger, M., Küker, W. & Rorden, C. The anatomy of spatial neglect based on voxelwise statistical analysis: a study of 140 patients. *Cerebral Cortex* **14**, 1164–1172 (2004).
118. Stone, S., Halligan, P. & Greenwood, R. The incidence of neglect phenomena and related disorders in patients with an acute right or left hemisphere stroke. *Age and Ageing* **22**, 46–52 (1993).
119. Li, K. & Malhotra, P. A. Spatial neglect. *Practical Neurology* **15**, 333–339, <https://doi.org/10.1136/practneurol-2015-001115>. <https://pn.bmj.com/content/15/5/333.full.pdf> (2015).
120. Corbetta, M. & Shulman, G. L. Spatial neglect and attention networks. *Annual review of neuroscience* **34**, 569–599 (2011).
121. Legge, S. D., Saposnik, G., Nilanont, Y. & Hachinski, V. Neglecting the difference: does right or left matter in stroke outcome after thrombolysis? *Stroke* **37**, 2066–2069 (2006).
122. Aszalós, Z., Barsi, P., Vitrai, J. & Nagy, Z. Lateralization as a factor in the prognosis of middle cerebral artery territorial infarct. *European neurology* **48**, 141–145 (2002).
123. Yoo, A. J. *et al.* Predictors of functional outcome vary by the hemisphere of involvement in major ischemic stroke treated with intra-arterial therapy: a retrospective cohort study. *BMC neurology* **10**, 25 (2010).
124. Swanson, N. *et al.* Lateral differences in the default mode network in healthy controls and patients with schizophrenia. *Human brain mapping* **32**, 654–664 (2011).
125. Royer, C. *et al.* Functional and structural brain asymmetries in patients with schizophrenia and bipolar disorders. *Schizophrenia research* **161**, 210–214 (2015).
126. Ocklenburg, S., Güntürkün, O., Hugdahl, K. & Hirnstein, M. Laterality and mental disorders in the postgenomic age—a closer look at schizophrenia and language lateralization. *Neuroscience & Biobehavioral Reviews* **59**, 100–110 (2015).
127. Corbetta, M., Kincade, M. J., Lewis, C., Snyder, A. Z. & Sapir, A. Neural basis and recovery of spatial attention deficits in spatial neglect. *Nature neuroscience* **8** (2005).
128. Murphy, K. & Fox, M. D. Towards a consensus regarding global signal regression for resting state functional connectivity mri. *Neuroimage* **154**, 169–173 (2017).

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Author Contributions

S.L., F.R. and O.S. wrote the main manuscript text. S.L. analysed data and prepared figures. M.V. preprocessed fMRI and DWI data from HCP. All authors reviewed the manuscript.

Additional Information

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