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Duration and Time Trends in Hospital Stay for Very Preterm Infants Differ Across European Regions*

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***See also p. 1175.**

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Objectives: To compare duration and changes over time in length of hospital stay for very preterm and extremely preterm infants in 10 European regions.

Design: Two area-based cohort studies from the same regions in 2003 and 2011/2012.

Setting: Ten regions from nine European countries.

Patients: Infants born between 22+0 and 31+6 weeks of gestational age and surviving to discharge (Models of Organising Access to Intensive Care for Very Preterm Births cohort in 2003, $n = 4,011$ and Effective Perinatal Intensive Care in Europe cohort in 2011/2012, $n = 4,336$).

Interventions: Observational study, no intervention.

Measurements and Main Results: Maternal and infant characteristics were abstracted from medical records using a common protocol and length of stay until discharge was adjusted for case-mix using negative binomial regression. Mean length of stay was 63.6

days in 2003 and varied from 52.4 to 76.5 days across regions. In 2011/2012, mean length of stay was 63.1 days, with a narrower regional range (54.0–70.1). Low gestational age, small for gestational age, low 5-minute Apgar score, surfactant administration, any surgery, and severe neonatal morbidities increased length of stay. Infant characteristics explained some of the differences between regions and over time, but large variations remained after adjustment. In 2011/2012, mean adjusted length of stay ranged from less than 54 days in the Northern region of the United Kingdom and Wielkopolska, Poland to over 67 days in the Ile-de-France region of France and the Eastern region of the Netherlands. No systematic decrease in very preterm length of stay was observed over time after adjustment for patient case-mix.

Conclusions: A better understanding of the discharge criteria and care practices that contribute to the wide differences in very preterm length of stay across European regions could inform policies to optimize discharge decisions in terms of infant outcomes and health system costs. (*Pediatr Crit Care Med* 2018; 19:1153–1161)

Key Words: cohort studies; intensive care; length of stay; neonatal

Very preterm infants require care in specialized neonatal units before they can be discharged home. Their length of stay (LOS) in hospital is influenced primarily by their gestational age (GA) at birth and medical conditions leading to longer stays, for example, bronchopulmonary dysplasia (BPD), persistent apnoeic spells, or need for tube feeding (1–5). The family's socioeconomic circumstances as well as the neonatal unit's policies and other health system factors, including the supply of beds or the availability of postdischarge home care, may also influence LOS. Although many interventions in this high-risk population are standardized, guidelines concerning when to discharge a very preterm infant are rare, reflecting the absence of evidence-based, consensual criteria (6).

Prolonged stay in hospital lengthens exposure to risks associated with the hospital environment, including nosocomial infections as well as noise and bright lights which may impact on future development (7). It may also disturb the establishment of interactions between parents and the infant (6). On the other hand, discharging a very preterm infant too early may expose the infant to the risk of life-threatening events and increase the risk of rehospitalization (8, 9). From a health system perspective, longer LOS reduces availability of beds and may limit the admission of other preterm infants to a neonatal ICU (NICU). LOS also determines healthcare costs, although other factors can offset the costs of longer hospital stays (10).

Trends toward shorter hospital stays have been observed for other hospital populations in industrialized countries, including stay in hospital after birth at term (11). For very preterm infants, however, it is possible that improvements in survival may have led to higher LOS as more infants at very low GAs or with more severe health conditions survive to discharge (12, 13). Some studies, however, show that morbidity among very preterm infants is declining, so survivors may be healthier and be discharged sooner (12–14). How these changing population

characteristics affect the overall LOS for very preterm infants is not known.

The aim of this study was to compare LOS in hospital for very preterm infants across two time periods in European regions that have similar standards of living and social systems guaranteeing insurance coverage. We also sought to identify maternal, pregnancy, and newborn characteristics associated with variability in LOS. Data come from two area-based studies conducted in the same European regions in 2003 and 2011/2012 (15, 16).

MATERIALS AND METHODS

Study Design and Population

This analysis used data from the Models of Organising Access to Intensive Care for Very Preterm Births (MOSAIC) and Effective Perinatal Intensive Care in Europe (EPICE) studies, which included all stillbirths and live births from 22 + 0 to 31 + 6 weeks of GA. MOSAIC focused on comparing medical practices and the organization of care in European regions (15), whereas the primary aim of EPICE was to investigate how evidence-based medicine is introduced into the clinical care of very preterm infants (16). MOSAIC included 10 regions from nine countries, whereas EPICE included 19 European regions from 11 countries. All of the MOSAIC regions were included in the EPICE study, but several regions in France, Poland, and the United Kingdom had slightly different administrative boundaries. For this study, we included only the common geographic zones in both studies (see study flow chart, **Fig. S1**, Supplemental Digital Content 1, <http://links.lww.com/PCC/A775>; legend: study flow chart showing numbers of infants eligible, excluded, and included). In both studies, recruitment occurred over a 12-month period, except for the French region, where the recruitment period was 6 months.

Investigators abstracted data from medical records in obstetric and neonatal units using a standardized questionnaire with common definitions that was pretested in all regions. The EPICE instrument was developed from the MOSAIC instrument, although modifications were made to achieve the new study objectives. Inclusions were cross-checked against delivery ward registers in maternity units or another external data source. Infants were followed up until discharge home from hospital or into long-term care or death. The methodologies for the studies have been described in more detail previously (15, 16).

Study Population

The full MOSAIC cohort included 7,356 still and live births of which 7,125 occurred in geographic zones included in the EPICE project. Of these births, 5,133 were live births, 4,917 were admitted to a NICU and 4,176 survived to discharge, representing 81.1% of live births and 56.3% of all births. The EPICE cohort included 10,329 births of which 7,009 were in common geographic zones. Of these, 5,265 were live births, 5,079 were admitted to a NICU and 4,504 survived to discharge, 85.5% of live births and 64.2% of all births. Infants

with missing location of discharge ($n = 319$) or missing date of discharge ($n = 14$) were not included in this study, but infants who were discharged into long-term institutional or foster care were included ($n = 116$; 1.4% of the sample). Infants with missing information on location of discharge were 4% of the total in both periods. The final study population comprised 4,011 infants surviving to discharge in 2003 and 4,336 in 2011/2012 (Fig. S1, Supplemental Digital Content 1, <http://links.lww.com/PCC/A775>).

Outcome and Covariables

The primary outcome measure was the length of hospital stay as calculated from the date of birth and the date of final discharge. Covariables were maternal and infant characteristics likely to influence the duration of hospitalization and included maternal and pregnancy characteristics (multiple pregnancy, prenatal corticosteroids, cesarean delivery) and infant characteristics (GA at birth, sex, birthweight, birthweight percentile for GA [< 10 th, 10–24th, ≥ 25 th], and the 5-min Apgar score). These covariables were selected based on the scientific literature assessing risk factors for long LOS, as well as previous studies from these cohorts identifying risk factors for adverse outcomes which could prolong stay in hospital (17–20). GA at birth was based on the best obstetric assessment, according to information on last menstrual period and/or ultrasound measures. Birthweight for GA was computed based on the observed birthweight for GA distribution in each cohort and reported in three groups previously found to be relevant for neonatal outcomes (19) Variables describing the neonatal course included any surfactant, surgery for any reason, neurologic morbidity (intraventricular hemorrhage [IVH] grade III or IV according to the classification Papile et al [21] and cystic periventricular leukomalacia [cPVL] documented on ultrasound or MRI scan), BPD (oxygen dependency or ventilation [including nasal continuous positive airway pressure] at 36 wk of GA), and any congenital malformation. Variables were collected using similar definitions with the exception of congenital malformations, for which more detail was available in the EPICE study, yielding a higher overall prevalence, which we believed to be principally linked to more thorough ascertainment of anomalies. This variable was therefore presented in descriptive tables, but not included in adjustment models.

Statistical Analysis

We first compared means and medians in LOS across periods and regions. Then we examined changes in the characteristics of women and infants between 2003 and 2011/2012. In both periods, we analyzed the impact of these population characteristics on LOS. As LOS is not a normally distributed variable, we used a negative binomial model to estimate risk ratios, as recommended (22). We also calculated the adjusted predictions in number of days for each covariable, meaning the LOS that would be observed for infants with this characteristic in a sample with the covariable profile of the overall study population. Adjusted models included all covariables. Region

was entered as a fixed effect, and we adjusted for clustering of patients within region.

We then used our adjusted models to compare LOS across regions taking into consideration patient characteristics in 2003 and in 2011/2012. We computed adjusted risk ratios for each region and time period by running one combined model with interaction terms for period on all covariables and predicted adjusted mean LOS for each region in each period. The estimated mean LOS are those that would be observed if the infants born in each region in each period had the covariable profile observed in the overall study population. Using these adjusted means, we calculated the difference between the periods for each region, and we obtained CIs using the delta method (23).

Finally, we carried out a sensitivity analysis by running this final model using a more homogeneous lower risk sub-population defined as infants born at 29 weeks or over without severe congenital anomalies or severe neonatal morbidity (IVH grades III/IV, cPVL, BPD, necrotizing enterocolitis, retinopathy of prematurity) or surgery during the neonatal hospitalization.

Ethics Approval

Ethics approval was obtained for both studies in each study region from regional and/or hospital ethics committees as required by national legislation. The European studies were also approved by the French Advisory Committee on Use of Health Data in Medical Research and the French National Commission for Data Protection and Liberties.

RESULTS

In 2003, the average LOS among survivors to discharge was 63.6 days (median of 56 d), with a variation across regions from 52.4 days (median of 46 d) in Wielkopolska in Poland to 76.5 days (65 d) in Hesse in Germany (Table 1). In 2011/2012, the overall mean and median were slightly lower (63.1 and 55), but there was no overall trend toward lower LOS; regional differences remained substantial despite a narrowing of the range between regions at the extremes. The small reduction in overall LOS masked contrasting regional trends with sizable decreases in Hesse in Germany and Flanders in Belgium and increases in the Eastern-Central region of the Netherlands.

Table 2 compares the characteristics of very preterm survivors to discharge in the two time periods. The proportion of infants born to older women, as part of a multiple set and to nulliparous women increased. More active perinatal management was evidenced by changes in antenatal steroid use and a slightly higher proportion of cesarean deliveries. The GA distribution changed slightly with more survivors at lower gestations; more infants had low 5-minute Apgar scores. In contrast, the proportions of survivors with severe neurologic morbidity and BPD were significantly lower.

Table S1 (Supplemental Digital Content 2, <http://links.lww.com/PCC/A776>) illustrates how maternal and infant characteristics affected LOS in 2011/2012. GA at birth was strongly related to LOS, with an adjusted mean of 106.7 and 105.2 days for infants born at 23 and 24 weeks, respectively compared with 42.7 days at

TABLE 1. Length of Hospital Stay in the 10 European Regions in 2003 and 2011/2012

Country: Region	Models of Organising Access to Intensive Care for Very Preterm Births (2003)			Effective Perinatal Intensive Care in Europe (2011/2012)			Mean Difference 2011/2012–2003 (95% CI)
	n	Mean (SD)	Median (IQR)	n	Mean (SD)	Median (IQR)	
Belgium:Flanders	461	68.6 (33.7)	61 (46–81)	641	63.4 (32.5)	54 (44–74)	–5.1 (–8.6 to –1.6)
Denmark:Eastern	286	67.3 (31.3)	60 (48–79)	283	65.5 (42.9)	60 (39–77)	1.9 (–6.9 to –3.1)
France: Ile-de-France	712	67.0 (30.3)	61 (48–80)	656	65.8 (30.4)	60 (46–77)	–1.2 (–3.9 to –1.6)
Germany: Hesse	544	76.5 (46.6)	65 (47–94)	534	70.1 (45.0)	59 (43–84)	–6.4 (–10.8 to –1.9)
Italy: Lazio	361	60.3 (32.8)	53 (38–73)	470	57.9 (33.4)	49 (36–73)	–2.4 (–6.5 to –1.7)
The Netherlands: East Central	300	58.2 (23.1)	51 (41–71)	328	67.0 (36.4)	58.5 (44–80)	8.8 (4.7–12.8)
Poland: Wielkopolska	213	52.4 (29.0)	46 (31–68)	242	54.0 (29.3)	47 (31–70)	1.6 (–3.6 to –6.8)
Portugal: Northern	188	55.2 (48.2)	46 (33–62)	247	60.0 (30.0)	53 (39–71)	4.8 (–0.7 to –10.3)
United Kingdom: Northern	349	57.0 (28.9)	50 (36–73)	353	60.3 (43.2)	50 (37–70)	3.3 (–1.1 to –7.7)
United Kingdom: former Trent	597	57.2 (31.7)	49 (34–72)	582	61.2 (32.7)	51.5 (38–80)	4.0 (0.6 to –7.5)
Total	4,011	63.6 (34.9)	56 (41–77)	4,336	63.1 (36.0)	55 (40–76)	–0.4 (–1.7 to –0.9)

IQR = interquartile range.

31 weeks. Other risk factors and morbidities with a strong independent impact on LOS were small for GA (adjusted mean of 81.0 d for a birthweight percentile < 10th percentile compared with 58.2 d for a birthweight > 25th percentile), any surgery (adjusted mean of 77.5 d vs 59.4 d for no surgery), and BPD (adjusted mean of 74.5 d vs 58.8 d). Other significant predictors, but with a smaller impact, were a low 5-minute Apgar score, surfactant administration, and IVH-cPVL. The impact of these predictors on LOS was very similar in 2003 (Table S2, Supplemental Digital Content 3, <http://links.lww.com/PCC/A777>).

Patient characteristics contributed to the differences in LOS between regions as shown by changes in risk ratios after adjustment for population case-mix (Table 3). In particular, the extreme position in Hesse, Germany was reduced in both periods. On the other hand, the low LOS of the Netherlands in 2003 appeared to be largely explained by its patient case-mix. Nonetheless, although these differences explained some of the variation, the range of the adjusted means was still over 2 weeks: 51.0–71.8 days in MOSAIC and 52.8–68.0 days in EPICE. In general, the adjusted changes over time in mean LOS were less marked than unadjusted changes, with the exception of the Danish and Italian regions. We also observed large regional differences in LOS in sensitivity analyses using a lower risk sample, from 37.1 to 53.6 days in MOSAIC and 33.4–51.3 days in EPICE (Table S3, Supplemental Digital Content 4, <http://links.lww.com/PCC/A778>), with a high correlation between values in the overall and low risk population (Spearman rho of 0.95; $p < 0.001$ in MOSAIC and 0.92; $p < 0.001$ in EPICE).

DISCUSSION

In a large area-based study including more than 8,000 very preterm infants in two time periods, we observed a greater

than 2-week difference in duration of neonatal hospitalization between regions with the lowest and highest mean LOS in both periods. Marked regional differences persisted after adjusting for maternal and infant characteristics. Although overall LOS remained largely stable, there were varying trends across time by region with decreases in Belgium, Denmark, Germany, and Italy and increases in France, the Netherlands, Portugal, and the United Kingdom. In general, regions with higher initial LOS experienced reductions, whereas those with lower initial LOS remained constant or increased which reduced some of the variation between regions over time. These regional differences raise questions about how local and regional factors determine discharge decisions, whether differences in medical practices underlie LOS variation and what consequences these decisions have for health outcomes and costs.

Our study's strengths include its population-based design, large sample size, use of a common standardized protocol to abstract data on perinatal characteristics, and neonatal morbidity from medical records and the existence of data from the same geographical zones in two time periods. We were able to measure the entire length of hospitalization, which is not the same as the stay in the NICU (22). As organization of neonatal care differs between regions and units, it is important to include infants who are transferred to another ward or to another hospital when they do not need further intensive care (24).

Limitations are the absence of some data items relevant for the investigation of discharge decisions, including the family's socioeconomic status, whether infants were discharged with a feeding tube or oxygen and whether postdischarge home care was provided. Although we had a wide range of variables to describe case-mix, we did not have full information on each child's morbidity risk profile, such as sepsis, extended need for

TABLE 2. Maternal and Infant Characteristics in 2003 and 2011/2012

Maternal and Infant Characteristics	Models of Organising Access to Intensive Care for Very Preterm Births (2003) <i>n</i> = 4,011		Effective Perinatal Intensive Care in Europe (2011/2012) <i>n</i> = 4,336		<i>p</i>
	<i>n</i>	%	<i>n</i>	%	
Maternal characteristics					
Maternal age (yr)					
< 25	840	21.0	704	16.3	< 0.001
25–29	1,093	27.3	1,177	27.2	
30–34	1,234	30.8	1,347	31.2	
35 +	834	20.8	1,095	25.3	
Multiple birth	1,144	28.5	1,422	32.8	< 0.001
Nulliparous	2,165	54.7	2,468	57.5	0.01
Antenatal steroids (any)	3,133	80.5	3,838	89.5	< 0.001
Cesarean section	2,543	63.7	2,926	68.0	< 0.001
Foreign born ^a	838	22.1	880	22.0	0.943
Infant characteristics					
Gestational age (completed weeks)					
23	12	0.3	19	0.4	0.007
24	72	1.8	131	3.0	
25	161	4.0	207	4.8	
26	272	6.8	295	6.8	
27	344	8.6	407	9.4	
28	516	12.9	522	12.0	
29	625	15.6	657	15.2	
30	878	21.9	900	20.8	
31	1,131	28.2	1,198	27.6	
Birthweight					
< 750 g	305	7.6	346	8.0	0.533
≥ 750 g	3,701	92.4	3,990	92.0	
Small for gestational age (percentiles)					
< 10th	345	8.6	370	8.5	0.093
10–25th	546	13.7	665	15.3	
> 25th	3,107	77.7	3,301	76.1	
Male	2,153	53.7	2,339	53.9	0.817
5-min Apgar score < 7	469	12.6	610	15.0	0.003
Surfactant	1,919	48.2	2,272	53.7	< 0.001
At least one surgery	393	9.8	451	10.4	0.361
Intraventricular hemorrhage III–IV or cystic periventricular leukomalacia	291	7.3	260	6.1	0.028
Bronchopulmonary dysplasia at 36 wk	606	15.4	590	13.9	0.046
Congenital anomaly	202	5.0	403	9.3	< 0.001

^aIn the United Kingdom, data are collected on ethnicity, rather than country of birth; we used self-defined foreign ethnicity.

TABLE 3. Differences Between Regions in Length of Hospital Stay in 2003 and 2011/2012 and Changes Over Time After Adjustment for the Characteristics of the Patient Populations

Country: Region	Models of Organising Access to Intensive Care for Very Preterm Births (2003)			Effective Perinatal Intensive Care in Europe (2011/2012)			Adjusted Mean Difference 2011/2012-2003 (95% CI)
	Crude Incident Rate Ratio (95% CI)	Adjusted Incident Rate Ratio (95% CI)	Adjusted Mean (95% CI)	Crude Incident Rate Ratio (95% CI)	Adjusted Incident Rate Ratio (95% CI)	Adjusted Mean (95% CI)	
Belgium: Flanders	1.11 (1.11-1.11)	1.12 (1.11-1.13)	68.9 (67.9-70.0)	1.02 (1.02-1.02)	1.10 (1.09-1.10)	67.0 (66.3-67.7)	-1.9 (-2.9 to -0.9)
Denmark: Eastern	1.09 (1.09-1.09)	1.14 (1.13-1.15)	70.1 (69.1-71.2)	1.05 (1.05-1.05)	1.08 (1.06-1.09)	65.6 (64.5-66.6)	-4.5 (-6.3 to -2.8)
France: Ile-de-France	1.09 (1.09-1.09)	1.09 (1.07-1.10)	66.7 (65.3-68.1)	1.06 (1.06-1.06)	1.11 (1.10-1.12)	67.8 (67.0-68.6)	1.1 (0.2-2.1)
Germany: Hesse ^a	1.24 (1.24-1.24)	1.17 (1.16-1.18)	71.8 (70.8-72.7)	1.12 (1.12-1.12)	1.07 (1.06-1.07)	65.1 (64.7-65.6)	-6.6 (-7.7 to -5.6)
Italy: Lazio	0.98 (0.98-0.98)	1.02 (1.01-1.03)	62.7 (61.6-63.8)	0.93 (0.93-0.93)	0.95 (0.94-0.96)	58.1 (57.5-58.8)	-4.6 (-6.0 to -3.1)
The Netherlands: East Central	0.95 (0.95-0.95)	1.04 (1.02-1.05)	63.7 (62.0-65.4)	1.07 (1.07-1.07)	1.12 (1.11-1.12)	68.0 (67.3-68.7)	4.4 (2.8-5.9)
Poland: Wielkopolska	0.85 (0.85-0.85)	0.86 (0.85-0.87)	52.7 (52.0-53.5)	0.87 (0.87-0.87)	0.87 (0.85-0.88)	52.8 (51.8-53.8)	0.1 (-1.1 to -1.1)
Portugal: Northern	0.90 (0.90-0.90)	0.92 (0.91-0.93)	56.3 (55.6-57.1)	0.96 (0.96-0.96)	0.97 (0.96-0.98)	59.0 (58.5-59.5)	2.7 (2.1-3.3)
United Kingdom: Northern	0.93 (0.93-0.93)	0.83 (0.82-0.84)	51.0 (50.7-51.4)	0.97 (0.97-0.97)	0.87 (0.86-0.88)	53.2 (52.5-53.9)	2.2 (1.4-3.1)
United Kingdom: former Trent	0.93 (0.93-0.93)	0.89 (0.88-0.90)	54.7 (54.2-55.1)	0.98 (0.98-0.98)	0.92 (0.91-0.93)	56.1 (55.5-56.6)	1.4 (0.6-2.1)

^aInterpretation: for instance, German infants in 2003 had a length of stay 1.17 times longer than overall mean of the Models of Organising Access to Intensive Care for Very Preterm Births population, adjusted for covariables, which represented an adjusted mean length of 71.8 d; in 2011/2012, this length of stay was of 1.07 times longer than overall mean of the Effective Perinatal Intensive Care in Europe population with an adjusted mean of 65.1 d.

All incident rate ratios and adjusted means have as reference the overall mean of each study population.

total parenteral nutrition, the number of congenital anomalies, and the complexity of surgical interventions. However, sensitivity analyses among lower risk infants with a more homogeneous risk profile found similar differences across regions, suggesting that residual confounding does not explain these findings. Other limitations were the absence of data on readmissions after discharge and on the unit practices that might contribute to variation as well as insufficient sample sizes within regions to analyze differences across units. Finally, our study questions related to infants surviving to discharge and averages for the regions based on all admissions, including in-hospital deaths, would differ.

Patient factors were strong predictors of LOS, as found in previous studies, with GA at birth having a major impact (1, 4, 5). The most immature babies born at 23, 24, and 25 weeks of gestation were hospitalized for over twice as long as those born at 30 to 31 weeks. Other predictors of a longer LOS in our sample were neonatal complications and morbidities, especially BPD and having at least one surgery, and also to a lesser extent, brain lesions, which is concordant with previous knowledge (25, 26). Being small for GA also increased LOS. In contrast, maternal sociodemographic and healthcare characteristics were not significant predictors. In this respect, our results contrast with a recent U.S. study that found that the absence of antenatal steroids prolonged LOS in very low birthweight infants (27). We also observed differences in patient case-mix over time with an increase in the proportion of the highest risk infants at very low GAs, but also a decline in some complications such as severe brain lesions and respiratory morbidity.

The variation in patient characteristics across regions had an impact on LOS, and this could be substantial in some regions. For instance, in the Netherlands, where unadjusted LOS over time increased most, the lower limit for active management of very preterm infants was lowered from 26 to 24 weeks of gestation over this time period (14, 28). In 2003, this policy meant that unadjusted LOS was lower than in the other regions because of the small proportion of extremely preterm infants. This is illustrated by the large increase in LOS once patient characteristics were taken into consideration from 58.2 to 63.7 days. Even after adjusting for case-mix, however, the LOS increased in the Netherlands over time which could reflect other residual changes in the case-mix that were not taken into consideration in our analyses. In Germany and the United Kingdom, adjusted LOS tended to be lower than unadjusted values suggesting care for a higher risk patient population, which may be related to more active management of the highest risk infants (28).

Although case-mix explained some of the variability observed in our study, wide regional differences persisted after adjustment, in line with observations from previous studies on regional and unit differences in LOS (4, 29). Multiple factors could explain these variations, including opinions among professionals about the optimal care for these infants and differences in other practices that might lengthen stay, such as respiratory management or nutritional policies, which are known to vary across units (30). For instance, local mechanical

ventilation and oxygenation strategies, including specific variables used by various units to wean exogenous oxygen or ventilatory support may be a key determinant of LOS. A recent study found that implementing a collaborative quality initiative to promote evidence-based practices for feeding, discharge planning and management of apnea, bradycardia, and oxygen desaturation events made it possible to reduce postmenstrual age at discharge in California NICUs (31).

Health system factors related to financing, demand for beds or organization of postdischarge care could also play a role. In a Swedish cohort where mean postmenstrual age at discharge differed by up to 2 weeks between hospitals, it was lower with breastfeeding and domiciliary care and in hospitals without fixed discharge criteria (32). In our sample, the Hesse region provides an example of the impact of financing as hospital reimbursement was restructured from a system based on per diem charges to diagnosis-related groups, starting in 2003 on a voluntary basis and becoming obligatory in 2004 (33, 34). In this region, case-mix adjusted mean LOS declined steeply between 2003 and 2011/2012 (71.8–65.1 d).

Interventions are increasingly implemented in neonatal units to promote safe, earlier discharge of very preterm infants, including individualized developmental care, more involvement and preparation of parents for discharge, and routine discharge readiness assessment (35–39). Postdischarge organization of care, including hospital-assisted home care, can also reduce LOS in hospital (40). In contrast, gavage feeding of the infant at home has not yet been proven to successfully shorten the stay in hospital (41), while the effectiveness of prescribing monitors for use at home to shorten hospitalization is still debated (42). The usefulness of having models to predict LOS needs to be confirmed (1, 4, 25, 43). These regional differences may also make it possible to carry out research on the determinants as well as the benefits and disadvantages of short versus longer LOS in terms of rehospitalization rates, parental satisfaction, and costs. For instance, it would be possible to select regions with high versus low LOS and collect prospective data on the neonatal stay, including complications and healthcare, how the discharge decision was made for each child and the healthcare provided in the period after discharge. This could also permit use of time-varying models to understand how risk factors affect decisions across the neonatal course. A mixed method approach could complement these quantitative data with structured interviews of clinical decision-makers within the unit and with parents.

CONCLUSIONS

The large differences we found in LOS in hospital after very preterm birth among European regions could not be explained by infant characteristics. Although variation across regions narrowed from 2003 to 2011/2012, substantial differences remained. These raise questions about the diversity of current practices as well as the weak evidence-base for informing clinical and health policy decisions on the optimal hospital stays for these high-risk infants. They also raise questions

about whether some care strategies may make it possible to achieve shorter LOS. Finally, understanding how differences in LOS affect readmissions, infant health, family well-being, and healthcare costs is an area for future research and could provide a basis for best practice guidelines.

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