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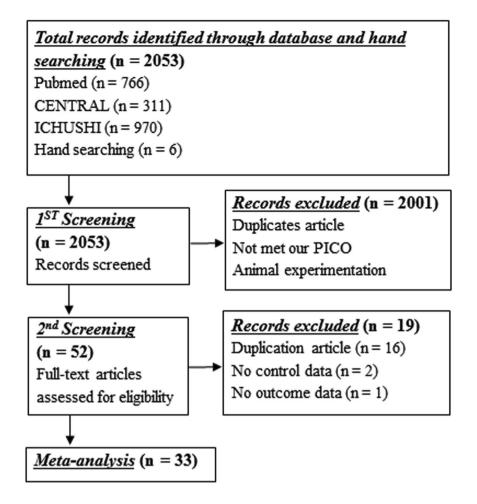
## Meta-analysis of the efficacies of amiodarone and nifekalant in shock-resistant ventricular fibrillation and pulseless ventricular tachycardia

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Amiodarone (AMD) and nifekalant (NIF) are used in the treatment of ventricular fibrillation or tachycardia; however, only few studies have been conducted on their efficacies. Therefore, a meta-analysis was conducted. Relevant sources were identified from PubMed, Cochrane Central Register of Controlled Trials, and Igaku Chuo Zasshi. The outcomes were short-term and long-term survival in patients with shockresistant ventricular fibrillation /pulseless ventricular tachycardia. Thirty-three studies were analysed. The results showed that, compared to the control treatment, AMD did not improve short-term survival (odds ratio (OR): 1.25, 95% confidence interval (CI): 0.91–1.71) or long-term survival (OR: 1.00, 95% CI: 0.63– 1.57). However, compared to the control treatment, NIF significantly improved short-term survival (OR: 3.23, 95% CI: 2.21–4.72) and long-term survival (OR: 1.88, 95% CI: 1.36–2.59). No significant difference was observed in short-term survival (OR: 0.85, 95% CI: 0.63–1.15) or long-term survival (OR: 1.25, 95% CI: 0.67–2.31) between AMD- and NIF-treated patients. The results suggest that NIF is beneficial for shortterm and long-term survival in shock-resistant ventricular fibrillation/pulseless ventricular tachycardia; however, the efficacy of AMD in either outcome is not clear.

The incidence of sudden cardiac arrest (SCA) is high. Moreover, SCA patients have a poor prognosis for survival. Out-of-hospital cardiac arrest (OHCA) that results in transportation in an ambulance to the hospital occurs in approximately 120,000 people per year<sup>1</sup>. In addition, the incidence of ventricular fibrillation (VF) in patients who have suffered OHCA is about 60% in Japan<sup>2</sup>. The current resuscitation guidelines recommend the use of amiodarone (AMD) for the treatment of shock-resistant adult VF/pulseless ventricular tachycardia (pVT) during cardiopulmonary resuscitation (CPR). However, lidocaine or nifekalant (NIF, approved for use only in Japan)

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#### Abbreviation: CENTRAL, the Cochrane Central Register of Controlled Trials; ICHUSHI, Igaku Chuo Zasshi.

#### Figure 1. Literature screening process and results.

may be administered if AMD is unavailable<sup>3-5</sup>. AMD and NIF are type III antiarrhythmic drugs. Randomized controlled trials (RCTs) and observational studies have shown that AMD and NIF cause better rates of survival until hospital admission than placebo or lidocaine do in VF/pVT patients<sup>6-8</sup>. Only a few RCTs and observational or retrospective studies have been conducted on the efficacies of AMD and NIF in the management of SCA. One of the reasons for this is that SCA occurs unexpectedly; therefore, large-scale clinical trials on it are difficult to conduct. In addition, there is a possibility that differences in treatment cannot be detected due to power shortage in a single research. In such a case, meta-analysis is a useful method of analysis because it integrates and evaluates multiple studies.

The aim of this study was to clarify the effects of AMD or NIF on the survival outcome of patients with shock-resistant VF/pVT. We conducted a meta-analysis of studies that compared the efficacies of the following: AMD versus control treatment (lidocaine, placebo, or non-treatment antiarrhythmic drugs), NIF versus control treatment (lidocaine, placebo, or non-treatment antiarrhythmic drugs), and AMD versus NIF.

#### Results

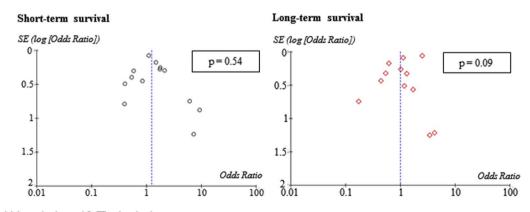
A total of 2053 studies were initially retrieved from the databases. After reviewing titles and abstracts, 2001 studies were excluded from the analysis. Therefore, the full texts of 52 reports were evaluated; however, 19 studies were further excluded. Finally, 33 studies<sup>6-36</sup> were used in the qualitative synthesis and meta-analysis. The literature screening process and results are depicted in Fig. 1. The general characteristics of the included studies are shown in Supplementary Table S1. Seven of the included studies were RCTs<sup>6,7,9,18,22,28,30</sup>, 6 were observational studies<sup>8,11,19,21,25</sup>, and 20 were retrospective studies<sup>10,12–17,19,20,23,24,26,27,29,31–36</sup>.

**Methodological quality of studies.** Assessment of methodological quality revealed that seven studies were RCTs (Supplementary Table S2). Of these, three had a low risk of bias, three had a high risk of bias, and one had an unclear risk of bias. Twenty-six studies were non-RCTs (Supplementary Table S3). Of these, two had a low risk of bias, 12 had a high risk of bias, and 12 had an unclear risk of bias.

a) Meta-analysis

	AMD Control		rol	Odds Ratio			Odds Ratio	
Study or Subgroup		Total	Events	Total	Weight	M-H, Random, 95% CI	Year	M-H, Random, 95% Cl
1.1.1 Short-term surv	ival							
Kudenchuk 1999	108	246	89	258	12.6%	1.49 [1.04, 2.13]	1999	-
Dorian 2002	41	180	20	167	10.0%	2.17 [1.21, 3.88]	2002	
Somberg 2002	14	18	3	11	2.8%	9.33 [1.65, 52.68]	2002	
Perzanowski 2003	13	152	15	102	7.8%	0.54 [0.25, 1.19]	2003	
Skrifvars 2004	38	75	67	105	9.7%	0.58 [0.32, 1.06]	2004	
Pollak 2006	24	36	49	59	6.3%	0.41 [0.15, 1.08]	2006	
Ito 2010	6	11	1	7	1.5%	7.20 [0.64, 81.54]	2010	
Tahara(1) 2010	10	30	29	78	6.9%	0.84 [0.35, 2.05]	2010	
Watanabe 2011	11	16	5	19	3.6%	6.16 [1.42, 26.78]	2011	
Kubo 2011	13	20	14	17	3.3%	0.40 [0.08, 1.87]	2011	
Piccini 2011	141	160	779	966	10.9%	1.78 [1.08, 2.95]	2011	
Amino 2015	151	380	20	73	10.3%	1.75 [1.00, 3.04]	2015	
Kudenchuk 2016	445	974	887	2051	14.4%	1.10 [0.95, 1.29]	2016	t
Subtotal (95% CI)		2298		3913	100.0%	1.25 [0.91, 1.71]		◆
Total events	1015		1978					
Heterogeneity: Tau <sup>2</sup> =	0.18; Ch	i <sup>2</sup> = 39.	50, df = 1	2 (P < 0.	0001); P	= 70%		
Test for overall effect:	Z=1.35	(P = 0.1	8)					
1.1.2 Long-term survi	val							
Kudenchuk 1999	33	246	34	258	10.6%	1.02 [0.61, 1.71]	1999	
Dorian 2002	9	180	5	167	7.0%	1.71 [0.56, 5.20]		
Skrifvars 2004	21	75	44	105	9.9%	0.54 [0.29, 1.02]	2004	
Pollak 2006	13	36	33	59	8.5%	0.45 [0.19, 1.04]		
Rea 2006	36	79	29	74	9.8%	1.30 [0.68, 2.47]	2006	_ <b>_</b>
Tahara(1) 2010	7	30	16	78	7.6%	1.18 [0.43, 3.23]		
Ito 2010	4	11	1	7	2.7%	3.43 [0.30, 39.64]		
Watanabe 2011	3	16	i	19	2.8%	4.15 [0.39, 44.57]		
Kubo 2011	4	20	10	17	5.4%	0.17 [0.04, 0.75]	2011	
Piccini 2011	94	160	675	966	11.5%	0.61 [0.44, 0.87]	2011	
Huang 2015	530	6459		19517	12.2%	2.52 [2.24, 2.83]		•
Kudenchuk 2016	237	970	455	2041	12.0%	1.13 [0.94, 1.35]		+
Subtotal (95% CI)	2.07	8282	400		100.0%	1.00 [0.63, 1.57]	2010	+
Total events	991		1972					
Heterogeneity: Tau <sup>2</sup> =				11 (P < I	0.00001);	I <sup>2</sup> = 92%		
Test for overall effect:	Z = 0.01	(P = 0.9	9)					
								0.01 0.1 i 10 100
								Favours [Control] Favours [AMD]

#### b) Funnel plots and Egger's test for the assessment to publication bias



Abbreviation: AMD, Amiodarone.

**Figure 2.** Meta-analysis and publication bias assessment for the effects of AMD on short-term survival and long-term survival compared with control treatment.

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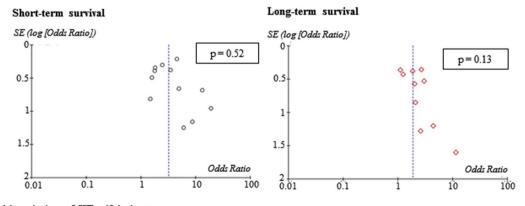
**Meta-analysis and assessment of publication bias, AMD versus control treatment.** Fifteen studies (RCTs, 4; non-RCTs, 11) compared the effects of AMD (8831 patients) with those of a control treatment (23510 patients) during CPR. Furthermore, the risk of bias was low, high, and unclear in 5, 6, and 4 studies, respectively. Compared to the control group, AMD-treated patients did not show improved short-term survival (OR: 1.25, 95% CI: 0.91–1.71) or long-term survival (OR: 1.00, 95% CI: 0.63–1.57). Moreover, both short-term survival ( $I^2 = 70\%$ ) and long-term survival ( $I^2 = 92\%$ ) showed high heterogeneity when the AMD-treated and control groups were compared (Fig. 2a). Funnel plots and Egger's tests indicated no possibility of publication bias in short-term survival (p = 0.54); however, there was possible publication bias in long-term survival (p = 0.09) (Fig. 2b).

**Meta-analysis and assessment of publication bias, NIF versus control treatment.** Thirteen studies (RCTs, 2; non-RCTs, 11) compared the effects of NIF (490 patients) with those of a control treatment (1630

a) Meta-analysis

	NIF		Contr	lo		Odds Ratio		Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	Year	M-H, Random, 95% CI
2.1.1 Short-term sur	vival							
Amino 2003	9	11	4	21	3.4%	19.13 [2.92, 125.32]	2003	
garashi 2003	5	7	2	9	2.5%	8.75 [0.90, 84.80]	2003	
Shimizu 2004	11	27	16	53	9.0%	1.59 [0.61, 4.18]	2004	
Tahara 2006	37	55	24	65	11.6%	3.51 [1.65, 7.48]	2006	
Yoshioka 2006	17	21	8	33	5.8%	13.28 [3.45, 51.19]	2006	
garashi 2006	50	100	169	944	17.0%	4.59 [3.00, 7.02]	2006	
Shiga 2010	23	27	15	28	6.1%	4.98 [1.36, 18.21]	2010	
Tahara(3) 2010	6	14	4	12	4.5%	1.50 [0.30, 7.43]	2010	
Tahara(2) 2010	20	38	104	274	12.7%	1.82 [0.92, 3.59]	2010	
Tahara(1) 2010	61	103	29	78	13.9%	2.45 [1.34, 4.49]	2010	
to 2010	5	10	1	7	2.2%	6.00 [0.52, 69.75]	2010	
Amino 2015	19	47	20	73	11.3%	1.80 [0.83, 3.91]	2015	+
Subtotal (95% CI)		460		1597	100.0%	3.23 [2.21, 4.72]		•
Total events	263		396					
Heterogeneity: Tau <sup>2</sup> =	= 0.17; Ch	P = 20.	39, df = 1	1 (P = 0)	0.04); P=	46%		
Test for overall effect	Z = 6.06	(P < 0.0	0001)					
2.1.2 Long-term sur	vival							
Amino 2003	2	11	0	21	1.1%	11.32 [0.49, 259.07]	2003	
Shimizu 2004	3	27	3	53	3.7%	2.08 [0.39, 11.10]		
Ando 2005	17	30	10	33	9.6%	3.01 [1.07, 8.47]		
garashi 2006	11	100	42	944	21.1%	2.65 [1.32, 5.34]		<b>_</b> _
Tahara 2006	14	55	14	65	14.3%	1.24 [0.53, 2.90]	2006	<b>_</b> _
Tahara(3) 2010	4	14	1	12	1.9%	4.40 [0.42, 46.26]	2010	
Tahara(1) 2010	23	103	16	78	19.9%	1.11 [0.54, 2.29]	2010	
Tahara(2) 2010	12	38	55	274	18.6%	1.84 [0.87, 3.87]	2010	
Shiga 2010	12	27	8	28	8.3%	2.00 [0.65, 6.11]		
to 2010	3	10	1	7	1.6%	2.57 [0.21, 31.71]		
		415		1515	100.0%	1.88 [1.36, 2.59]		•
Subtotal (95% CI)						-		
Subtotal (95% CI) Total events	101		150					
				P = 0.6	9); l² = 09	6		
Total events	= 0.00; Ch	P= 6.5	3, df = 9 (	P = 0.6	9); l² = 09	6		
Total events Heterogeneity: Tau <sup>a</sup> :	= 0.00; Ch	P= 6.5	3, df = 9 (	P = 0.6	9); I² = 09	6		
Total events Heterogeneity: Tau <sup>a</sup> :	= 0.00; Ch	P= 6.5	3, df = 9 (	P = 0.6	9); I² = 09	6	ţ	0.01 0,1 1 10 1

b) Funnel plots and Egger's test for the assessment to publication bias



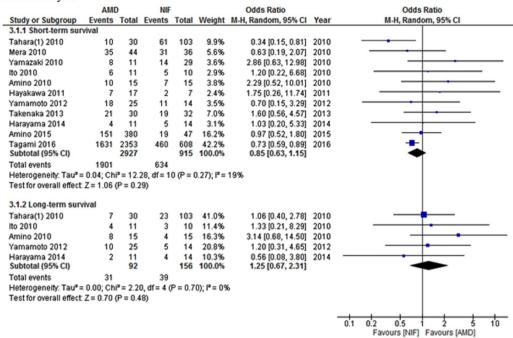


**Figure 3.** Meta-analysis and publication bias assessment for the effects of NIF on short-term survival and long-term survival compared with control treatment.

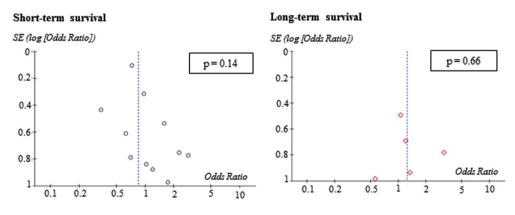
patients) during CPR. The risks of bias of the included studies were as follows: low, 1; high, 8; and unclear, 4. NIF significantly improved short-term survival (OR: 3.23, 95% CI: 2.21–4.72) and long-term survival (OR: 1.88, 95% CI: 1.36–2.59) (Fig. 3a). Furthermore, heterogeneity was low in short-term survival ( $I^2 = 46\%$ ) and long-term survival ( $I^2 = 0\%$ ) when the NIF-treated and control groups were compared. Funnel plots and Egger's tests indicated no possibility of publication bias in short-term survival (p = 0.52) or long-term survival (p = 0.13) (Fig. 3b).

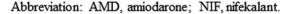
**Meta-analysis and assessment of publication bias, AMD versus NIF.** Eleven studies (RCT, 1; non-RCTs, 10) compared the effects of AMD (2927 patients) with those of NIF (915 patients) during CPR. The risk of bias was low, high, and unclear in 1, 3, and 7 studies, respectively. No significant difference was found in short-term survival (OR: 0.85, 95% CI: 0.63–1.15) or long-term survival (OR: 1.25, 95% CI: 0.67–2.31) (Fig. 4a) when the AMD- and NIF-treated groups were compared. Furthermore, no significant heterogeneity was observed in short-term survival ( $I^2 = 19\%$ ) or long-term survival ( $I^2 = 0\%$ ). Funnel plots and Egger's tests indicated no possibility of publication bias in short-term survival (p = 0.14) or long-term survival (p = 0.66) (Fig. 4b).

a) Meta-analysis



b) Funnel plots and Egger's test for the assessment to publication bias





**Figure 4.** Meta-analysis and publication bias assessment for the effects of AMD on short-term survival and long-term survival compared with NIF.

**Subgroup analyses.** Subgroup analyses were performed to evaluate the effects of the various treatments in the different type of studies (Table 1). Differences in the effects of AMD and the control treatments were observed between the RCT (OR: 1.61, 95% CI: 1.03–2.49) and non-RCT (OR: 1.04, 95% CI: 0.61–1.77) studies with regard to short-term survival; however, this was not the case for long-term survival (RCTs, OR: 1.13, 95% CI: 0.95–1.33; non-RCTs, OR: 0.93, 95% CI: 0.46–1.87). Furthermore, when the NIF-treated and control groups were compared, the results obtained for the RCT and non-RCT studies were not different with regard to short-term survival (RCTs, OR: 15.03, 95% CI: 5.02–44.99; non-RCTs, OR: 2.80, 95% CI: 2.04–3.86)., but different with regard to long-term survival (RCTs, OR: 11.32, 95% CI: 0.49–259.07; non-RCTs, OR: 1.84, 95% CI: 1.33–2.54). Conversely, when the effects of AMD and NIF were compared, no different results were obtained between the RCT and non-RCT studies with regard to short-term survival (RCTs, OR: 1.04, 95% CI: 0.52–10.01; non-RCTs, OR: 0.80, 95% CI: 0.61–1.05) or long-term survival (RCTs, OR: 3.14, 95% CI: 0.68–14.50; non-RCTs, OR: 1.04, 95% CI: 0.53–2.05).

**Sensitivity analysis.** Only studies that had a low risk of bias were analysed. When the AMD-treated and control groups were compared, it was observed that AMD significantly improved short-term survival (OR: 1.50, 95% CI: 1.14-1.98, n = 5) but not long-term survival (OR: 0.95, 95% CI: 0.65-1.39, n = 4). Furthermore, when the NIF-treated and control groups were compared, it was observed that NIF did not improve short-term survival

Comparison	Survival outcome	Study design	Odds ratio (95% CI)	Number of study
	Short-term	RCT	1.61 (1.03-2.49)*	4
AMD vs. Control	Short-term	non-RCT	1.04 (0.61–1.77)	9
AND VS. CONTO	Long-term	RCT	1.13 (0.95–1.33)	3
	Long-term	non-RCT	0.93 (0.46-1.87)	9
	Short-term	RCT	15.03 (5.02-44.99)*	2
NIF vs. Control	Short-term	non-RCT	2.80 (2.04-3.86)*	10
NIF VS. Control	Long-term	RCT	11.32 (0.49–259.07)	1
	Long-term	non-RCT	1.84 (1.33-2.54)*	9
	Short-term	RCT	2.29 (0.52-10.01)	1
AMD vs. NIF	Short-term	non-RCT	0.80 (0.61-1.05)	10
AND VS. INIT	Long-term	RCT	3.14 (0.68–14.50)	1
	Long-term	non-RCT	1.04 (0.53-2.05)	4

**Table 1.** Subgroup analysis in three comparison (RCTs or non-RCTs). Abbreviation: AMD, amiodarone; NIF, nifekalant; RCT, randomized controlled trial; CI, confidence interval. \*p < 0.05.

(OR: 1.80, 95% CI: 0.83–3.91, n = 1); however, no deductions could be made concerning long-term survival as it was not evaluated. Lastly, when the AMD- and NIF-treated groups were compared, no significant difference was found in the results for short-term survival (OR: 0.97, 95% CI: 0.52–1.80; n = 1). However, long-term survival was not evaluated in the study with low risk of bias.

#### Discussion

AMD and NIF are class III antiarrhythmic drugs; however, their pharmacological profiles are different. The effects of AMD include blocking Na<sup>+</sup> channels,  $Ca^{2+}$  channels, and K<sup>+</sup> channels. It is also a  $\beta$ -adrenergic receptor blocker. In contrast, NIF is a pure K<sup>+</sup> channel blocker. In this study, we evaluated the effects of AMD or NIF on short-term and long-term survival in adult patients with shock-resistant VF/pVT in a meta-analysis. Majority of the studies analysed were non-RCTs (79%). All the studies were conducted within a 20-year period (1994–2015) in various countries including the United States of America, Canada, Hungary, Finland, Taiwan, and Japan. The guidelines used for CPR varied among the studies.

Among the major RCTs studied<sup>6,7</sup>, AMD was reported to improve short-term survival but not long-term survival in out-of-hospital VF patients. AMD is recommended in the 2015 American Heart Association (AHA) Guidelines Update for CPR and Emergency Cardiovascular Care and Japan Resuscitation Council Guidelines 2015 (JRC Guideline 2015) as the first-line drug for treating shock-resistant VF/pVT patients during ACLS<sup>4,5</sup>. However, we could not confirm the effectiveness of AMD for long-term or short-term survival in this meta-analysis. In the subgroup analysis, different results were obtained for the RCT and non-RCT studies. AMD improved short-term survival in the RCTs but not in the non-RCTs. Furthermore, the sensitivity analysis showed that, the improved short-term survival occurred only in the studies with low risk of bias. We considered the following factors as reasons for these results. 1) It was observed that many of the non-RCT reports indicated statistically significant differences in patient background factors, which are considered as confounding factors, between the AMD-treated and control groups. The confounding factors included defibrillation frequency, epinephrine dosage, patient age, and time until drug administration, among others. There was no study that aligned the patient background between AMD treatment group and control group by using such as propensity score. These confounding factors may have influenced the efficacy assessment of AMD and affected survival outcome in the RCTs and non-RCTs. 2) Regarding the ability to improve survival until hospital admission, AMD was found to be superior to placebo treatment, but not to lidocaine<sup>18</sup>. In the present analysis, placebo treatment, lidocaine, and non-treatment antiarrhythmic drugs were termed control group. The results indicate that AMD is more effective than placebo treatment is in treating shock-resistant VF/pVT; however, it appears that AMD and lidocaine have a similar efficacy. 3) In the 2015 AHA and JRC guidelines, the recommended dose of AMD for shock-resistant VF/VT patients is 300 mg as a bolus; however, this usage is not approved in some countries. Various dosages of AMD, such as 125 mg over 5-10 min and 150-300 mg over a few seconds to 1 min, were used until 2013, when the resuscitation dose (300-mg bolus injection) was approved in Japan. AMD has hemodynamic side effects such as bradycardia and blood pressure reduction, which are due to its multichannel blocking action. However, the side effects can be avoided by slowing the infusion rate<sup>37</sup>. Furthermore, it has been reported that, survival rate until hospital admission in Japan is higher in patients who receive 150 mg AMD than in those who receive  $300 \text{ mg AMD}^{21}$ . These indicate that differences in dosage and infusion rate may influence the effects of AMD. 4) Polysorbate 80 is a surfactant that has been reported to lower blood pressure in dogs<sup>38</sup>. The effects of AMD alone<sup>9,18</sup> or polysorbate 80 as a control treatment<sup>6,7</sup> have been studied in RCTs. However, the effects of polysorbate 80 have not been excluded in non-RCTs<sup>10–17,19–21</sup>. It is therefore possible that polysorbate 80 affects the efficacy of AMD. In this meta-analysis, it was not possible to clarify the effects of AMD on short-term survival as reported in the RCTs. 5) From the analysis, fewer non-RCTs had low risk of bias than the RCTs had (non-RCTs, 18%; RCTs, 75%). Therefore, it is possible that some bias (such as that from confounding factors) was introduced into the efficacy evaluation of AMD in the non-RCTs.

Among small RCT<sup>22,28</sup>, NIF was reported to improve short-term survival, not long-term survival in out-of-hospital VF patients. NIF is recommended in the JRC Guidelines 2015 as the second-line drug for treating shock-resistant VF/pVT patients during ACLS. We were able to confirm the effects of NIF on short-term and long-term survival in this meta-analysis. There were similar results when the RCTs and non-RCTs were compared. In the sensitivity analysis, there was only one study that had a low risk of bias, and no significant effects of NIF on short-term survival were observed. It has been reported that NIF reduces defibrillation threshold. In addition, the time to achieve defibrillation success is shorter with NIF than it is with AMD<sup>30,35</sup>. Furthermore, NIF is a pure potassium channel blocker; therefore, its hemodynamic side effects are probably of a lesser severity than those of AMD are. The above-mentioned facts suggest that NIF is a very useful agent for resuscitation. Moreover, the findings of the present analysis suggest that NIF is effective for both short-term and long-term survival. However, because there was only one study that had a low risk of bias (non-RCTs, 9%; RCT, 0%), it is possible that some bias was introduced into the efficacy evaluation of NIF.

Among small-RCT<sup>30</sup>, AMD and NIF are reported to have no differences in their effects on short-term and long-term survival in out-of-hospital VF patients. In the present analysis, there was no statistically significant difference in short-term or long-term survival between the AMD- and NIF-treated patients or between the RCT and non-RCT studies. The results of an experiment conducted in pigs showed that NIF caused a higher survival rate than AMD does<sup>39</sup>. However, another study using pigs<sup>40</sup> and some clinical trials<sup>19–21,29–36</sup> have indicated that AMD and NIF have a similar efficacy. All the studies evaluated in this meta-analysis were conducted at advanced medical facilities in Japan (special functioning hospitals). In addition, there have been many studies conducted in Japan in which different dosages of the drugs (e.g., AMD, 125 mg over 5 min or 150 mg over 1 min; NIF, 0.15 mg/ kg/min) have been evaluated. The above-mentioned factors may have had some impact on comparing the efficacy of AMD and NIF. The results of this meta-analysis suggest that NIF and AMD have a similar efficacy. However, since the number of studies analysed was few and several of them were of low quality (low risk of bias: RCT, 0%; non-RCTs, 10%), there may have biases in the efficacy evaluation of the drugs.

The possibility of publication bias was considered low in the three comparisons (AMD versus control treatment, NIF versus control treatment, and AMD versus NIF).

In order to clarify the efficacies of AMD and NIF during ACLS, large-scale studies such as multicentre collaborative researches must be conducted and the use of several databases will be required. In addition, it is essential that patient characteristics are similar between groups, especially in non-RCTs, to allow for effective comparisons to be made. This analysis had some limitations. Firstly, several of the reports analysed were non-RCTs and conference articles. In addition, the studies analysed were conducted during different periods and in different countries. Moreover, different CPR guidelines were used in the various studies. Lastly, patients who were concomitantly administered antiarrhythmic drugs were not excluded from the analysis.

#### Conclusion

NIF may be effective for short-term and long-term survival in shock-resistant VF/pVT patients. However, the effects of AMD on either outcome could not be clarified.

#### Methods

We conducted the meta-analysis according to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines<sup>41</sup>. The PRISMA checklist is shown in Supplementary Table S4.

**Eligibility criteria.** We identified studies according to the following PICO criteria: Patients (adult patients who had suffered out/in-hospital cardiopulmonary arrest (CPA) and had VF or pVT were recruited for the studies), Intervention (AMD or NIF was administered during advanced cardiovascular life support (ACLS)), Comparison (as control) (lidocaine, placebo, or a non-treatment antiarrhythmic drug was administered during ACLS), and Outcome (the reporting outcomes were short-term survival (defibrillation success, VF/pVT termination, return to spontaneous circulation, survival until admission to the hospital/intensive care unit, and three-hour survival) and long-term survival (30-day survival, 1-year survival, and survival until discharge from hospital)).

**Study selection and data extraction.** To identify studies that were relevant to the analysis, we searched the PubMed, Cochrane Central Register of Controlled Trials (CENTRAL), and Igaku Chuo Zasshi databases for articles published until December 2016 using the following keywords: "amiodarone", "nifekalant", "ventricular fibrillation", "pulseless ventricular tachycardia", and "cardiopulmonary arrest". The search formula is shown in Supplementary Table S5. We also manually searched through the reference lists of retrieved articles to trace other relevant studies. The search was not limited to articles published in a particular language. Two authors (S.S. & S.T.) independently screened the titles and abstracts of all selected articles (first screening). We then assessed the full text of each article for eligibility, which constituted the second screening. Articles that did not meet our PICO criteria, such as review articles, paediatric reports, results of animal experimentation, case reports, reports on pharmacokinetics, and duplicate reports, were excluded from the analysis during the screening. Disagreements about which studies to exclude from the analysis were resolved based on discussions between two authors (S.S. & S.T.). The following data were extracted from the included studies: first author's last name, publication year, study design, settings of the study, patients' characteristics, rhythms of arrest, details of sample collection, treatment regimens, and outcomes reported. Inconsistencies in data extraction were resolved through discussion. In addition, authors were contacted for clarifications when needed.

**Evaluation of the methodological quality of included studies.** We assessed the methodological quality of each included RCT by using the risk of bias tool recommended by the Cochrane Collaboration (London,

UK)<sup>42</sup>. In addition, the studies were evaluated using the following quality domains: random sequence generation, allocation concealment, blinding of participants as well as personal and outcome assessors, incomplete outcome data, selective outcome reporting, and other potential threats to validity. Non-RCTs (such as observational and retrospective studies) were evaluated using the Risk of Bias Assessment Tool for Non-randomized Studies<sup>43</sup>. This was followed by assessment using the following quality domains: selection of participants, confounding variables, measurement of exposure, blinding of outcome assessments, incomplete outcome data, and selective outcome reporting. The assessment was performed independently by two authors (S.S. & S.T.). A study was considered as having a low risk of bias if each key domain was found to have a low risk of bias. Alternatively, a study was considered as having high or unclear risk of bias if one or more key domains were found to have high or unclear risk of bias.

**Comparison groups.** The effects of the following were compared: AMD versus control treatment (lidocaine, placebo, or non-treatment antiarrhythmic drugs), NIF versus control treatment (lidocaine, placebo, or non-treatment antiarrhythmic drugs), and AMD versus NIF.

**Meta-analysis.** Meta-analysis was performed using RevMan5.1<sup>®</sup> software (The Cochrane Collaboration). Outcomes were determined using a random-effects model and taking into consideration the variability of CPR practice, the different countries in which the studies were conducted, the different comorbidities of patients, and the different treatment strategies, among other factors. Outcomes were evaluated using odds ratio (OR) with 95% confidence interval (95% CI). Between-study heterogeneity was considered high if  $I^2$  statistic was >50%.

**Assessment of publication bias.** R metafor software (Software Foundation's GNU General Public License) was used to assess the degree of publication bias both graphically and statistically using the funnel plot asymmetry test (Egger's test)<sup>44</sup>. P values < 0.1 were considered statistically significant according to the asymmetry of the funnel plot (indicating the possibility of publication bias).

Subgroup analysis. We performed subgroup analyses to evaluate RCTs and non-RCTs.

**Sensitivity analysis.** We also performed a sensitivity analysis by excluding studies that had a high or unclear risk of bias.

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#### **Author Contributions**

S.S., T.I., Y.Z. and K.I. conceived the study concepts. S.S. and S.T. performed literature search, study selection and data extraction. S.S., S.T. and T.K. performed statistical analyses. T.K., T.N., Y.K., M.C. and K.T. conducted technical support. S.S., T.I., T.K., K.T. and Y.Z. drafted the initial manuscript. Y.K., Y.I., T.S., H.N. and K.I. conducted critical revision of the manuscript for important intellectual content. Y.Z. had primarily responsibility for the final content. All authors reviewed the manuscript.

#### **Additional Information**

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