

1 **Distinguishing Emission-Associated Ambient Air PM_{2.5} Concentrations and Meteorological**
2 **Factor-Induced Fluctuations**

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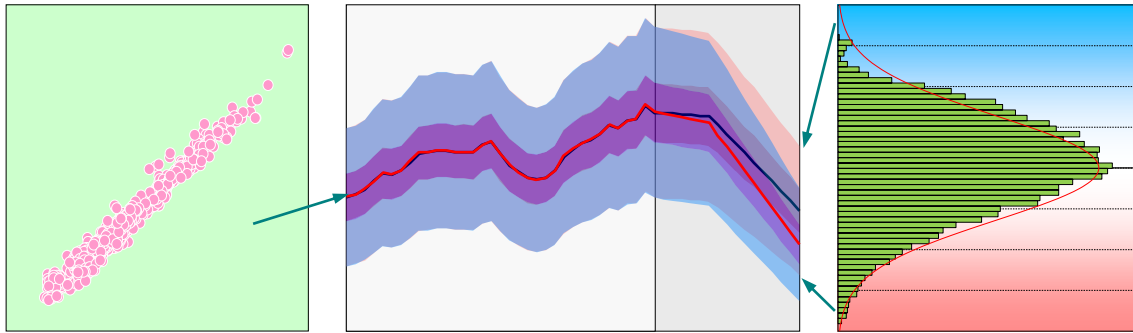
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18 **ABSTRACT**

19 Although PM_{2.5} (particulate matter with aerodynamic diameters of less than 2.5 μm) in the air originates
20 from emissions, its concentrations are often affected by confounding meteorological effects. Therefore,
21 direct comparisons of PM_{2.5} concentrations made across two periods, which are commonly used by
22 environmental protection administrations to measure the effectiveness of mitigation efforts, can be
23 misleading. Here, we developed a two-step method to distinguish the significance of emissions and
24 meteorological factors and assess the effectiveness of emission mitigation efforts. We modeled ambient
25 PM_{2.5} concentrations from 1980 to 2014 based on three conditional scenarios: realistic conditions, fixed
26 emissions, and fixed meteorology. The differences found between the model outputs were analyzed to
27 quantify the relative contributions of emissions and meteorological factors. Emission-related gridded PM_{2.5}
28 concentrations excluding the meteorological effects were predicted using multivariate regression models,
29 whereas meteorological confounding effects on PM_{2.5} fluctuations were characterized by probabilistic
30 functions. By combining the regression models and probabilistic functions, fluctuations in the PM_{2.5}
31 concentrations induced by emissions and meteorological factors were quantified for all model gridcells and
32 regions. The method was then applied to assess the historical and future trends of PM_{2.5} concentrations and
33 potential fluctuations on global, national, and city scales. The proposed method may thus be used to assess
34 the effectiveness of mitigation actions.

35

36 INTRODUCTION

37 PM_{2.5} (particulate matter with aerodynamic diameters of less than 2.5 μm) is a major environmental and
38 health concern^{1,2}. PM_{2.5} in the air originates from the direct emissions of primary aerosols and from the
39 secondary formation of aerosols from various precursors³ and ambient PM_{2.5} concentrations are shaped
40 primarily by the emission rates⁴⁻⁶. In addition to emissions, meteorological conditions are critical to the
41 formation and transport of PM_{2.5} through the air⁷⁻⁹. Interannual climate variability can also affect regional
42 pollution levels¹⁰. Therefore, spatiotemporal variations in PM_{2.5} concentrations in the atmosphere are
43 mainly driven by the combined effects of emissions, chemical reactions, and meteorology¹¹.

44 Although the impacts of emissions and meteorological confounding effects on PM_{2.5} pollution have been
45 studied extensively¹²⁻¹⁴, a lack of understanding of interactions between them has often led to confusion
46 among the public and policymakers. For example, local governments often report on the effectiveness of
47 their mitigation efforts from observed reductions in annual mean PM_{2.5} concentrations ignoring
48 considerable fluctuations in meteorological conditions occurring between years. Such a practice is
49 misleading whenever strong positive or negative meteorological interferences occur. For example, an
50 abnormal increase in PM_{2.5} concentrations occurred following a period of PM_{2.5} decline in northern China
51 in early 2017. The average PM_{2.5} concentration in the first half-year of 2017 (66 μg/m³) was slightly higher
52 than that during the same period in 2016 (64 μg/m³) in Beijing although comprehensive mitigation efforts
53 have been made in recent years. The event has stimulated debate on the effectiveness of recent mitigation
54 actions¹⁵ even though these efforts have already led to a continuous decrease in annual mean PM_{2.5}
55 concentrations in this area in recent several years¹⁶. A recent study has suggested that the abnormal increase
56 during the first six months of 2017 was strongly associated with anomalies in humidity.¹⁷

57 To quantify the contributions of emissions and confounding meteorological factors to ambient PM_{2.5}
58 concentrations, a two-step approach was developed. In brief, global PM_{2.5} concentrations from 1980 to
59 2014 were simulated based on three conditional modeling scenarios: 1. realistic conditions, 2. fixed
60 meteorology (realistic daily emission estimates but fixed meteorological parameters for 2014) and 3. fixed
61 emissions (realistic daily meteorological variables with mean emissions from 1980 to 2014). Based on the
62 results of the simulations, regression models were developed for individual gridcells to predict
63 emission-driven PM_{2.5} trends. Probabilistic functions were established to characterize superimposed

64 meteorology-associated fluctuations. By combining the regression models and probabilistic function, PM_{2.5}
65 concentration trends to be induced by changes in emissions and meteorological factor-associated
66 fluctuations could be distinguished. The effectiveness of emission mitigation measures could thus be
67 evaluated. Moreover, future trends of ambient PM_{2.5} concentrations can be predicted based on projected
68 changes in emissions.

69 **METHODS**

70 **Overall Approach.** Fig. 1 shows the overall scheme of the proposed approach including 1) a simulation
71 based on three scenarios and 2) the development of regression models and probabilistic functions.

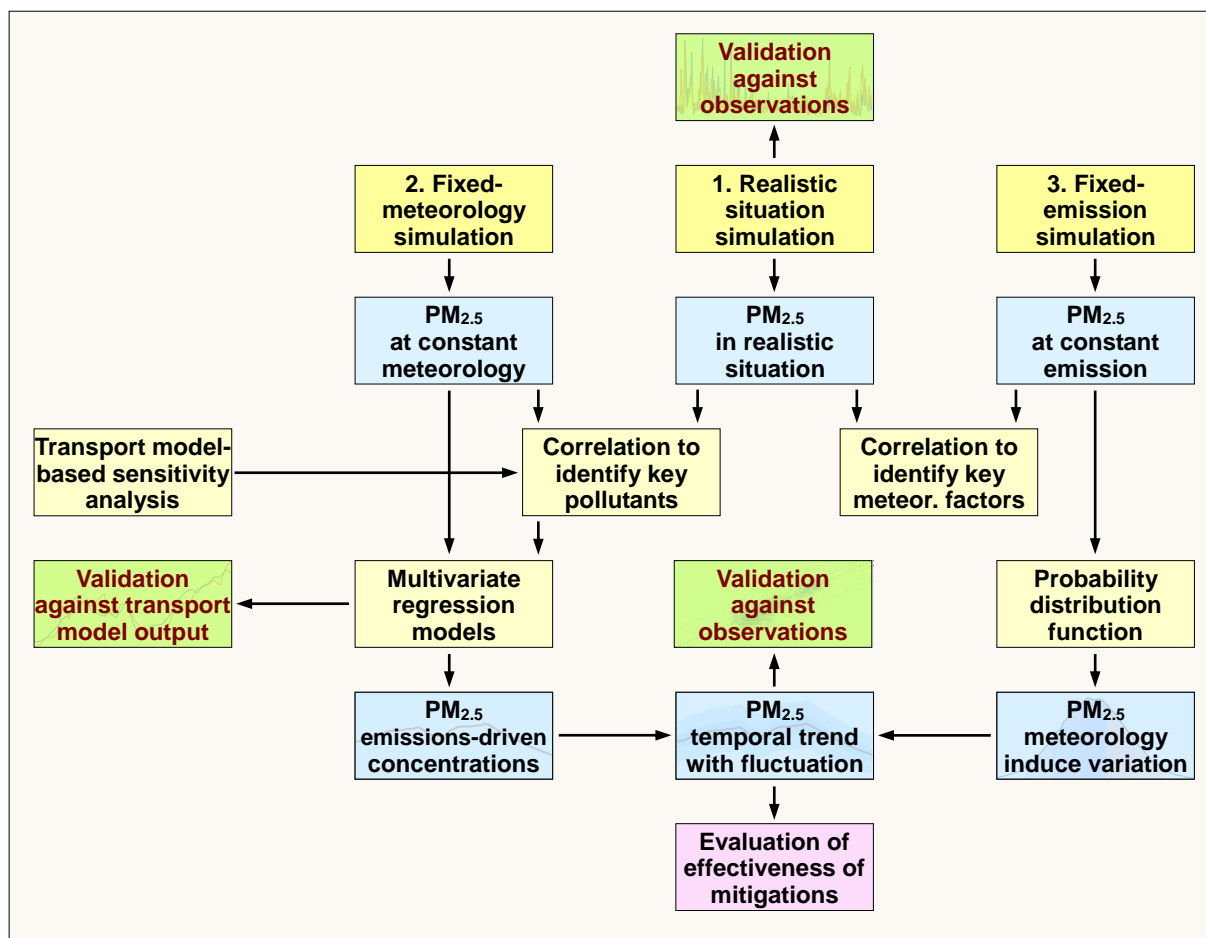


Fig. 1 Flowchart showing the research scheme of this study. Gridded $PM_{2.5}$ concentrations were simulated for the three scenarios from 1980 to 2014. Individual effects of emissions and meteorological factors were measured. Regression models were developed using the second model's scenario simulation output to predict gridded $PM_{2.5}$ concentrations based on emissions. Meteorological confounding effect-induced variations were quantified as probabilistic functions using the third model's scenario simulation output. Using the models, trends in $PM_{2.5}$ concentrations with a variability range were generated, and the effectiveness of mitigation measures were evaluated. The procedures were validated at various stages.

72

73 **Atmospheric Chemical Transport Modeling and Validation.** The MOZART4 (Model for Ozone and

74 Related Chemical Tracers, version 4) was applied to simulate daily $PM_{2.5}$ concentrations from 1980 to 2014

75 on a global scale¹⁸. The model was set with a 1.895° (latitude) \times 1.875° (longitude) horizontal resolution,

76 with 28 vertical layers, and with a 15-minute time step. The species considered include black carbon (BC),

77 organic carbon (OC), unspecified PM_{2.5} (primary PM_{2.5} - BC - 1.3OC), SOA (Secondary Organic Aerosol),
78 sulfate, nitrate, ammonium, dust, and sea salt. Emissions were obtained from the PKU (Peking
79 University)-series for primary aerosols (PM_{2.5}, BC, and OC), SO₂ (sulfur dioxide) and NO_x (nitrogen
80 oxides)¹⁹. Emissions drawn from other inventories were also used in this study, including NH₃ and
81 nonbiogenic NMVOC (Nonmethane Volatile Organic Carbon) data collected from EDGAR (Emissions
82 Database for Global Atmospheric Research) and HTAPv2 (Hemispheric Transport of Air Pollution, version
83 2)^{20,21}, biogenic VOC (Volatile Organic Carbon) data collected from MEGAN (Model of Emissions of
84 Gases and Aerosols from Nature)²², and open-field biomass burning emission data collected from GFED4.1
85 (Global Fire Emissions Database, version 4.1)²³. NCEP/NCAR (National Centers for Environmental
86 Prediction/National Centers for Atmospheric Research) reanalysis products²⁴ were used as offline
87 meteorological inputs. Aerosol optical depths from MODIS (Moderate Resolution Imaging
88 Spectroradiometer)²⁵ were used as a proxy to downscale the model predicted parameters into a fine gridcell
89 of 0.125°×0.125°²⁶. Model performance was evaluated against more than 220 thousand daily monitoring
90 data points collected from around the world (**Fig. S1**), against time series observations for six major cities
91 around the world (**Fig. S2**), and against major components (**Fig. S3**). It can be observed that the majority of
92 data points fall around the 1:1 line without bias and that the deviation of the predicted concentrations from
93 the observations increase as the time scale decreases. For the annual means primarily used in this study,
94 87% of data points are within the two-fold range.

95 **Conditional Scenarios and Relative Contributions.** The simulation was conducted based on three
96 conditional modeling scenarios. The control run was conducted using realistic emission estimates and
97 meteorological fields. For the fixed-meteorological condition scenario, meteorological parameters for 2014
98 (a normal non-El Niño year) were applied to all years with realistic emission estimates data. For the
99 fixed-emission scenario, 35-year-averaged emissions were applied to all years together with realistic
100 meteorological conditions. Deviations in the fixed emissions and fixed meteorological condition
101 simulations from the normal simulation (control run) were normalized to their respective fractions to
102 quantify the overall contributions of emissions (RC_E) and meteorological conditions (RC_M) for a given
103 region (from a gridcell to the globe) and for a given period (from a month to multiple years) of interest.

104 **Sensitivity Analysis.** A sensitivity analysis was conducted to identify major air pollutants governing
105 ambient air PM_{2.5} concentrations through a preliminary simulation for January 2010 (monthly resolution).

106 Modeling was repeatedly performed by reducing or enhancing the emissions of individual pollutants by
107 10%, 25%, 50%, 75%, or 100% each time. The 21 pollutants tested include primary PM_{2.5} (including
108 primary BC, OC and unspecified PM_{2.5}), SO₂, NH₃, NO_x, CO, CH₃SCH₃, C₆H₅(CH₃), BIGALK (lumped
109 alkanes with C > 3), C₂H₄, BIGENE (lumped alkenes with C > 3), C₃H₆, CH₂O, CH₃CHO, CH₃OH,
110 CH₃COCH₂CH₃, C₃H₈, C₂H₅OH, C₂H₆, CH₃COCH₃, C₁₀H₁₆, and C₅H₈. The results of the sensitivity
111 analysis are listed in **Table s1**.

112 **Emission-based Regression Model.** Based on the results of the sensitivity analysis, the four main air
113 pollutants were used for regression model development. Using annual emissions of these pollutants as
114 independent variables and PM_{2.5} concentrations from the fixed-meteorology simulation as a dependent
115 variable for 35 years, multivariate regression models with both dependent and independent variables
116 log-transformed were developed for individual gridcells to predict PM_{2.5} concentrations without
117 meteorological confounding effects. The regression was established for all individual gridcells using data
118 for 35 years. The uncertainty of the regression models based on the fixed-meteorology simulation was
119 characterized by a 90% confidence interval of predicted PM_{2.5} concentrations. Model-predicted PM_{2.5}
120 concentrations were compared against those calculated from the fixed-meteorology simulation (the same
121 data set used for model development). The method cannot be applied to model PM_{2.5} variation on a
122 relatively short time scale such as a daily scale, which can be affected by many occasional extreme
123 emission or meteorological events, as well as the nonlinearity of secondary formation of aerosol.

124 **Meteorology-related Probabilistic Functions.** For each individual gridcell, the frequency distribution of
125 the annual mean PM_{2.5} concentrations for a 35-year period derived from the fixed-emission simulation was
126 used as a meteorology-related probabilistic function to quantify random variations of PM_{2.5} induced by
127 changes in meteorology at each gridcell. The function can also be generated for a region (such as a country)
128 at other time scales (such as a month) of interest. At 84% of all model gridcells, the probabilistic functions
129 calculated follow a normal distribution with a zero mean (K-S test, $p > 0.05$).

130 **Characterization of Emission-Driven Trends with Meteorology-Induced Fluctuations.** This was done
131 by combining emission-based trends with meteorology-induced variations from 1980 to 2030. Using
132 emissions and PM_{2.5} concentrations for 2014 as baselines, the gridcell-specific models were applied to
133 project the trajectory of PM_{2.5} concentrations induced by given emission changes for all gridcells across the

134 globe. When superimposed on predicted $PM_{2.5}$ concentrations derived from regression models, variations
135 induced by fluctuations in meteorological variables presented as UI_{50} (intervals between the 25th and 75th
136 percentiles) and UI_{95} (intervals between the 2.5th and 97.5th percentiles) were derived using the distribution
137 pattern discussed in the previous section. Prior to future projections, combined model simulations were
138 conducted for a period from 1988 (when the first valid observation was available) to 2014 and were
139 validated against 2940 field observations collected from IMPROVE (Interagency Monitoring of Protected
140 Visual Environments) for the United States and from EMEP (The European Monitoring and Evaluation
141 Programme) for European countries at annual scale, and corresponding results are shown in **Fig. S4**.

142 **Other Analysis.** Statistical analysis was conducted using SPSS 23.0²⁷ with a significance level of 0.05.
143 Monte Carlo simulations were performed using MATLAB R2016b²⁸ to generate the frequency distribution
144 functions associated with variation of meteorological parameters for individual gridcells.

145 **Limitations and Uncertainties.** The methodology is affected by limitations and uncertainties. For example,
146 the emission inventories are subject to uncertainty, and meteorological conditions for a single year (2014)
147 are not truly representative. Like other atmospheric chemical transport models¹⁴, MOZART cannot provide
148 model uncertainty information, while Monte Carlo simulation for complex atmospheric chemistry
149 modeling would be unrealistic due to extremely high computation loading. Moreover, many
150 physicochemical processes were not even included^{29,30}. Contribution of SOA to $PM_{2.5}$ formation is often
151 underestimated by the modeling. To date, very limited multiple-year observation data are available on a
152 global scale, which are critical for model validation. Last but not the least, the overall uncertainty of the
153 two-step procedure was unable to be characterized due to the limitations listed above. Nevertheless, there is
154 still room to further improve the method. In addition to updating the inventories, quantifications of the
155 effects of individual pollutants and meteorological factors could help to mitigate such uncertainties.

156 **RESULTS AND DISCUSSION**

157 **Effects of Emissions and Meteorological Factors.** Based on the results of a sensitivity analysis, the
158 relative contributions of various air pollutants to PM_{2.5} concentrations and the responses of PM_{2.5} to these
159 pollutants are shown in **Fig. S5**. As is shown, 97% of the variations in PM_{2.5} concentrations are attributable
160 to the emission of primary PM_{2.5} (56.9±28.6%) followed by the emission of SO₂ (18.9±8.8%), NH₃
161 (12.9±6.6%), and NO_x (8.3±6.8%), respectively. Similar results have recently been reported^{31,32}.
162 Significant ($p < 0.05$) correlations between the emissions of the four pollutants and PM_{2.5} concentrations
163 derived from the fixed-meteorology simulation were found for 70% of land gridcells around the world,
164 denoting the feasibility of predicting emission-driven PM_{2.5} concentrations based on emission densities of
165 these pollutants while excluding confounding meteorological effects. Those land gridcells (30%) not
166 showing significant correlations between pollutant emissions and ambient PM_{2.5} concentrations were
167 mostly identified in desert areas and high-latitude regions with low emissions, such as the Sahara Desert
168 and the Arctic Archipelago.

169 **Fig. 2** presents maps of partial correlation coefficients between emissions and PM_{2.5} concentrations on an
170 annual basis. The four major pollutants in terms of their respective contribution to PM_{2.5} concentrations,
171 including primary PM_{2.5}, SO₂, NO_x, and NH₃, are shown. Primary PM_{2.5}-dominated partial correlations
172 were found for China and India, where coal and biomass fuels used for power generation, industry,
173 residential sectors, and cement production are the most important emission sources^{33,34}. In the United States,
174 PM_{2.5} concentrations are more SO₂ emission-dependent, which is consistent with the large fraction of
175 sulfates in total PM_{2.5} concentrations observed in the country³⁵. For most Western European countries,
176 primary PM_{2.5} and SO₂ made a synthetic contribution to PM_{2.5} mass concentration (such was the case in
177 Germany³⁶), whereas NO_x has a stronger effect on France. The influence of NH₃ mainly occurred in
178 Eastern European countries and Russia (west) because NH₃ exhausted from the agriculture sector (e.g.,
179 fertilizer and domesticated animals) is the leading factor affecting formation of ammonium sulfate and
180 nitrate¹. The significance of the correlation increased as the time scale changed from annual to daily. For
181 example, median correlation p values of SO₂ are 0.14 (0.012-0.46), 0.011 (0.0000038-0.29), and 9.4×10^{-31}
182 (3.1×10^{-105} - 7×10^{-7}) on annual, monthly, and daily scales, respectively.

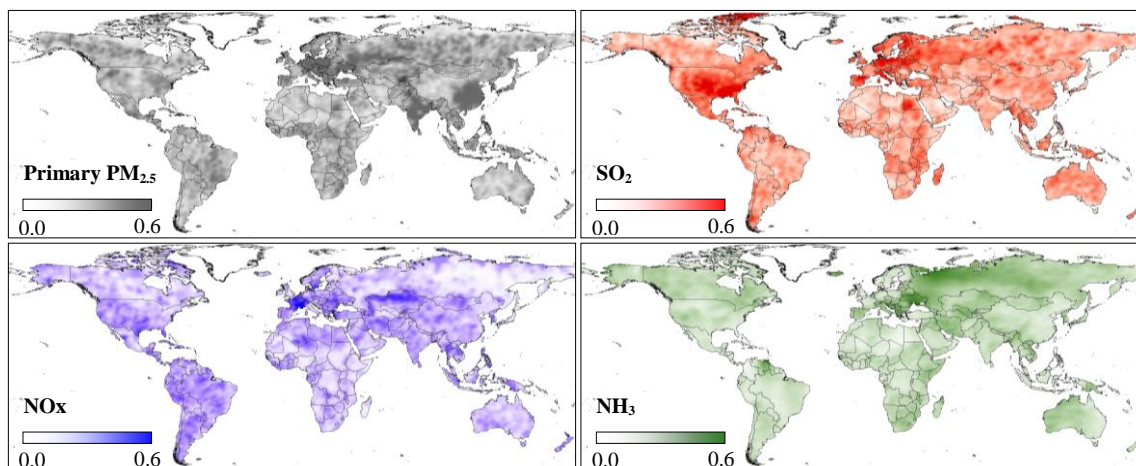


Fig. 2 Geospatial distribution of partial correlation coefficients between the emissions of major air pollutants and $PM_{2.5}$ concentrations. The four pollutants are primary $PM_{2.5}$, SO_2 , NO_x , and NH_3 .

183 Similarly, significant partial correlations ($p < 0.05$) were found between the main meteorological
 184 parameters and $PM_{2.5}$ concentrations derived from the fixed-emission scenario simulation. On average, the
 185 most important parameter is air temperature (T), with a correlation of 0.22 followed by wind speed (WS, r
 186 = -0.16), planetary boundary layer height (PBLH, $r = -0.16$), relative humidity (RH, $r = 0.14$), and surface
 187 pressure (SP, $r = -0.14$). These results correspond with those of a previous study^{7,14,37} The geospatial
 188 distribution of the main meteorological parameters is shown in **Fig. S6**. In cold, high-latitude regions of
 189 North America and Siberia and in warm regions extending from northern Africa to the Arabian Peninsula,
 190 $PM_{2.5}$ concentrations are mostly sensitive to temperature, which is partially associated with
 191 temperature-sensitive SO_2 ³⁸. The effects of WS or PBLH are stronger in regions with relatively high
 192 elevations, where strong winds facilitate dispersion^{7,14,37}, whereas the presence of low PBLH levels predict
 193 a stable atmosphere³⁹. WS and PBLH are also important in many other regions, including Southeast Asia,
 194 Brazil, and the eastern seaboard of Australia, where tropical or subtropical monsoons prevail⁴⁰. In dry
 195 inland regions such as central Eurasia, the formation of secondary $PM_{2.5}$ is more sensitive to RH⁴¹. To
 196 characterize the relationship between emissions and meteorological effects, the relative contributions of

197 emissions (RC_E) and meteorological conditions (RC_M) were measured across all model gridcells based on
198 the results of the three conditional scenario simulations. The mean daily/weekly RC_M values for $PM_{2.5}$
199 ($68\% \pm 5\%/63\% \pm 5\%$) are much higher than the mean daily/weekly RC_E values for $PM_{2.5}$
200 ($32\% \pm 5\%/37\% \pm 5\%$) ($p < 0.05$). Emissions become more significant on a seasonal/annual basis. For
201 example, mean seasonal RC_E is $54\% \pm 7\%$. Changes in emissions on these longer time scales are largely
202 driven by seasonal emission cycles^{23,42} and by long-term socioeconomic patterns⁴³.

203 In addition to annual mean $PM_{2.5}$ concentrations, the number of severely polluted days (NSPD, defined as
204 the number of days with daily $PM_{2.5}$ values of $> 150 \mu\text{g}/\text{m}^3$) is of particular interest not only because the
205 annual mean concentrations are significantly associated with these high values⁴⁴ but also because public
206 responses to extreme conditions are stronger⁴⁵. The occurrence of heavy pollution episodes is often
207 associated with stable meteorological conditions, as emissions do not usually change dramatically on a
208 daily basis⁴⁶. **Fig. S7a** compares temporal variations of the NSPD for Beijing (from the realistic-case
209 simulation) to emissions of major air pollutants for the surrounding area (Beijing-Tianjin-Hebei) for the
210 winter months from 2000 to 2014. Although the NSPD and emissions undergo similar increasing trends,
211 they are not always synchronous on an annual basis due to the influence of meteorological conditions. For
212 example, a sharp increase in the NSPD observed from 2012 to 2013 was not driven by emission increases
213 but by unusual meteorological conditions^{46,47}. During that winter, the seasonal averaged WS dropped from
214 a long-term mean of 2.94 to 2.33 m/s, and the number of days of abnormally high humidity ($RH > 75\%$)
215 and extremely low PBLH ($< 150 \text{ m}$) increased from 3% to 10% and from 6.3% to 8.4%, respectively (**Fig.**
216 **S7b-f**), favoring the growth of secondary aerosols and the accumulation of air pollutants at the ground
217 level^{48,49}.

218 **Emission-based Prediction.** As discussed above, annual mean $PM_{2.5}$ concentrations for the 35-year period
219 derived from the fixed-meteorology simulation are significantly correlated with emissions observed across
220 individual model gridcells. Such a correlation suggests that a set of regression models can be developed to
221 predict $PM_{2.5}$ concentrations based on emissions with meteorological confounding effects excluded. If such
222 models can be validated against the output of the fixed-meteorology simulation, they can be applied to
223 simulate historical $PM_{2.5}$ trends-based exclusively on emissions and to predict emission-driven future $PM_{2.5}$
224 trends. As confounding meteorological effects are eliminated by these models, the proposed method

225 enables us to evaluate the effectiveness of emission-reduction efforts. To do so, the emissions of the four
226 most important air pollutants identified based on a sensitivity analysis, primary PM_{2.5}, SO₂, NH₃, and NO_x,
227 were used as independent variables in developing multivariate regression models, whereas PM_{2.5}
228 concentrations derived from the fix-meteorology simulation were used as a dependent variable. As both
229 emission densities and PM_{2.5} concentrations are log-normally distributed (**Fig. S8**), the multivariate
230 regression models were fitted to all model gridcells using log-transformed data and were applied to
231 calculate annual mean PM_{2.5} concentrations for these gridcells. As the formation of secondary aerosols in
232 the air does not linearly respond to precursor emissions³, several nonlinear equations were tested with no
233 significant improvements observed in the results. Given that the statistical regression models were
234 established to predict annual PM_{2.5}, the nonlinearity of the secondary aerosol formation, which occurred in
235 a short time ranging from seconds to diurnal, was filtered out by the annual means. As such, the following
236 linear model was adopted.

$$237 \log PM_{2.5} = \sum a_i \log E_i + b,$$

238 where $PM_{2.5}$ is annual mean PM_{2.5} concentration, E_i are annual emissions of the i^{th} pollutants, a_i and b are
239 regression coefficients. **Fig. S9** shows the spatial distribution of R^2 values of the regression models,
240 indicating that results for areas characterized by high emission levels and population densities are much
241 better (R^2 values are close to one) than those found for other regions, which is helpful in reducing overall
242 uncertainty. The regression models were validated by plotting the predicted PM_{2.5} concentrations against
243 those derived from the fixed-meteorology scenario simulation shown in **Fig. S10** for China, India, the
244 United States, and Germany. This good agreement suggests that the models could be used to predict annual
245 PM_{2.5} concentrations with reasonable accuracy, while confounding meteorological effects were not taken
246 into account. It should be pointed out that the potential impact of climate change was not taken into
247 consideration.

248 The simplified approach to predicting annual mean ambient PM_{2.5} concentrations at the ground level based
249 on annual total emissions omits the exchanges occurring among gridcells due to transport. Although the
250 association between the emissions and PM_{2.5} concentrations at a given gridcell can be disturbed by the
251 atmospheric transport across gridcells, the influence of the atmospheric transport on the association is
252 weakened by similarities among adjacent model gridcells. Such similarities were demonstrated by the spatial

253 autocorrelation of the regression model parameters. On a global scale, the calculated Moran's
254 autocorrelation indexes are valued at 0.39 (intercepts), 0.50 (slopes for primary $PM_{2.5}$), 0.33 (slopes for
255 SO_2), 0.36(slopes for NO_x), and 0.30 (slopes for NH_3) and are statistically significant ($p < 0.05$). As was
256 expected, such autocorrelation is also significant for the gridded emissions and $PM_{2.5}$ concentrations and
257 Moran's autocorrelation indexes vary from 0.25 to 0.52 for gridded emissions of the four pollutants and are
258 as high as 0.75 for gridded $PM_{2.5}$ concentrations ($p < 0.05$). The most significant autocorrelation of $PM_{2.5}$
259 concentrations is attributed to the dispersion of $PM_{2.5}$ in the air. Due to the autocorrelation of emissions,
260 emissions observed at individual gridcells also shape emissions from the surrounding gridcells.

261 **Meteorology-related Variations.** As discussed above, interannual trends of emission-driven $PM_{2.5}$
262 concentrations excluding meteorological confounding effects can be predicted based on annual emissions
263 from data generated from the fixed-meteorology simulation. Similarly, the outputs of the fixed-emission
264 simulation provide the information on variations in $PM_{2.5}$ concentrations caused by confounding
265 meteorological effects. As the influence of meteorological factors randomly fluctuates based on
266 emission-induced $PM_{2.5}$ concentrations, the following probabilistic function was used to characterize such
267 random effects:

$$268 \quad F(PM_{2.5}) = (2\pi\sigma)^{-0.5}\exp(-PM_{2.5}^2/2\sigma^2),$$

269 where $F(PM_{2.5})$ is a probability function, $PM_{2.5}$ is annual mean $PM_{2.5}$ concentration, and σ is standard
270 deviation associated with change in meteorological conditions under the fixed emission. Based on annual
271 mean $PM_{2.5}$ concentrations calculated from the fixed-emission simulation for the 35 years spanning from
272 1980 to 2014, probabilistic functions were derived for all individual gridcells on an annual scale. For most
273 of the gridcells (84%) the functions are normally distributed ($p > 0.05$). Deviations from the normal
274 distribution are mostly observed in deserts or surrounding areas (**Fig. S11**). For the fixed-meteorology
275 simulation the year 2014 is assumed to be a "normal" year for which most meteorological parameters are
276 approximately equal to the 35-year mean with a standardized deviation of 0.12 ± 0.25 . This assumption is
277 confirmed by calculating the average deviation of annual $PM_{2.5}$ concentrations derived from 2014
278 meteorology trends to average values for 1980 to 2014 based on the fixed-emission simulation. It was
279 found that relative deviations for 95% of all model gridcells are less than 5%, and the overall mean value of
280 deviation for all gridcells is $0.072\%\pm 1.1\%$ (mean and standard deviation), which is not significantly

281 different from zero ($p < 0.05$) as was expected. Therefore, the frequency distribution generated from the
282 fixed-emission simulation represents random variations resulting from confounding meteorological effects.
283 The robustness of the function was also tested using a Jackknife test for a randomly selected gridcell. The
284 test was conducted 35 times by removing calculated 35 $PM_{2.5}$ concentrations from the fixed-emission
285 simulation one by one and by generating probabilistic distributions based on the 34 remaining datasets. The
286 mean and standard deviation of the 35 repeated calculations are $3 \times 10^{-17} \pm 5 \times 10^{-17}$ and 0.04 ± 0.001 ,
287 respectively, indicating a very high degree of robustness.

288 In fact, the probabilistic functions can be derived either on an annual basis or on any temporal scale from a
289 daily to seasonal basis. **Fig. 3** shows typical examples of the probabilistic functions for a typical gridcell
290 (Guangzhou, China) on annual (a), monthly (b), and daily (c) scales. The majority of these functions reflect
291 typical normal distributions, which is more evident on a shorter time scale. On an annual scale, the annual
292 mean $PM_{2.5}$ concentration changes considerably with a coefficient of variation (CV) of 14%. Even without
293 any change in emissions, the annual mean $PM_{2.5}$ concentration presents a 48% chance of increasing or may
294 decrease by more than 10%. This means that while emission-mitigation measures can reduce ambient $PM_{2.5}$
295 concentrations by 10% in a single year for this gridcell, there is a more than 20% chance of the observed
296 annual mean concentration not declining at all or even of increasing. Similarly, the likelihood of the annual
297 mean decreasing by more than 20% is also higher than 20%. Therefore, simply comparing annual mean
298 $PM_{2.5}$ concentrations of two consecutive years without taking meteorological conditions into consideration
299 can be misleading. Upon reducing the time scale from annual to monthly and daily, the variation in
300 probabilistic functions increases. CV values for monthly and daily $PM_{2.5}$ concentrations increase to 29%
301 and 38%, respectively, for the selected gridcell, which are significantly higher values than those found for
302 annual data and which can be explained by the fact that daily and monthly meteorological conditions vary
303 more dramatically than emissions. Therefore, monthly meteorological factor-forced changes are more
304 random than those observed on an annual scale. With constant emissions there is a more than 50%
305 probability of a 20% change occurring in monthly mean $PM_{2.5}$ concentrations. Therefore, it is even riskier
306 to directly compare mean $PM_{2.5}$ concentrations of a given month to those for the same period of a previous
307 year while disregarding random confounding meteorological effects.

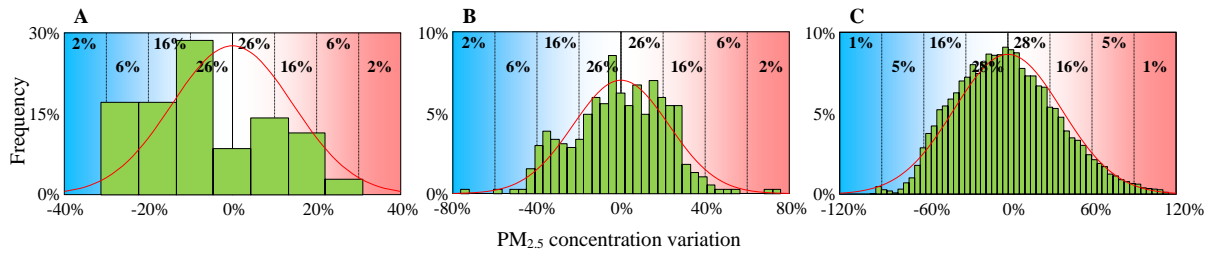


Fig. 3 Probabilistic functions derived from fixed-emission simulations of annual (A), monthly (B), and daily resolutions (C) for a representative gridcell. Bars denote the frequency distribution of the model-calculated $PM_{2.5}$ concentrations normalized by corresponding mean values with a fitted normal distribution curve. The probabilities of individual segments are shown in the background.

308 The random variation observed in the calculated probabilistic function is a direct indicator of the extent of
 309 confounding meteorological effects on individual gridcells. To quantify overall variations on a global scale,
 310 annual mean-based CVs were calculated for all gridcells. Corresponding results are shown in **Fig. S12** as a
 311 cumulative distribution of CVs for all gridcells. The mean and standard deviation of the CV values are
 312 $16 \pm 11\%$ (median is 14.2%) with a maximum value of 109%. On average, confounding meteorological
 313 factors can lead to more than one-sixth of a variation at 28% for all model gridcells. The contribution can
 314 be as high as 100% in extreme cases. As discussed above, short-term variations observed over less than one
 315 year are even larger. When monthly data are used, the mean and standard deviation of the CV values are
 316 $65 \pm 35\%$, showing stronger seasonal variations. The maximum CV of an individual gridcell can reach 200%
 317 on a monthly scale. Again, significant autocorrelations (Moran's index = 0.59, $p < 0.05$) were found for the
 318 probabilistic functions (CVs) on an annual scale, denoting continuity in meteorological effects across
 319 space.

320 The annual change in confounding meteorological effects on globally averaged $PM_{2.5}$ concentrations,
 321 defined as a normalized global average $PM_{2.5}$ anomaly for individual year from the 35-year mean, was
 322 calculated from 1980 to 2014 based on the fixed-emission simulation. The deviations observed reflect the
 323 average influence of annual meteorological conditions on annual mean $PM_{2.5}$ concentrations on a global
 324 scale. It should be noted that the annual deviation observed in 2014 was the smallest, showing that using
 325 meteorological parameters for 2014 as a "normal" year for our fixed-meteorological simulations is the best
 326 choice for the 35 years studied. Such annual changes are often affected by global atmospheric circulation⁵⁰.
 327 It is interesting to observe that the interannual anomalies of meteorological effects are significantly
 328 correlated with Arctic Oscillation (AO), which is shown as solid dots in **Fig. 4** ($r = 0.66$, $p < 0.05$). Some
 329 regional studies also show a similar relationship. For example, it was reported that enhanced dust emissions

330 observed across Saharan regions and the increasing frequency of haze episodes recorded in northern China
331 are associated with the positive phase of AO^{10,51}.

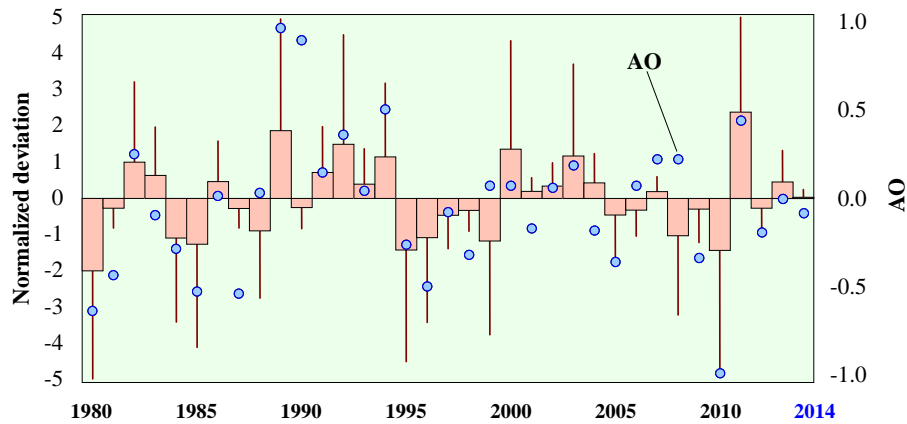


Fig. 4 Normalized global average deviation of PM_{2.5} concentrations from the mean value for the 35 years spanning from 1980 to 2014 (bars). The results are based on a fixed-emission simulation conducted at the global scale. Standard deviations are shown as dark red lines. The blue dots denote Arctic Oscillation.

332 To further illustrate spatial variations in meteorological-induced variation, UI_{95} values are mapped in **Fig.**
333 **S13** in both absolute and relative terms. The global average UI_{95} of annual PM_{2.5} concentrations was
334 measured as 4.9 $\mu\text{g}/\text{m}^3$ (40%). Hot regions of absolute variation exhibit strong meteorological variations. In
335 addition to areas around deserts (e.g., the southern Sahara and the Middle East) where dust forms a major
336 component of PM_{2.5} emissions and where concentrations are subject to synoptic-scale weather patterns⁵¹,
337 strong variations in PM_{2.5} concentrations can be observed in heavily polluted regions such as the North
338 China Plain (NCP) and likely due to interactions between high emissions and highly variable
339 meteorological patterns^{9,46}. On the other hand, relatively large values of relative terms are often observed in
340 regions with low levels of population density and low PM_{2.5} concentrations. For example, very high levels
341 of relative variability were found in high-latitude regions and coastal areas, where background PM_{2.5}
342 concentrations are very low. In most high-emission regions (e.g., eastern China, India, Europe, the United
343 States), although PM_{2.5} variations induced by meteorological conditions are lower, high PM_{2.5} levels can
344 increase absolute variations on a considerable scale. For example, the UI_{95} for northern India and for the
345 NCP are as high as 11.5 $\mu\text{g}/\text{m}^3$ and 20.6 $\mu\text{g}/\text{m}^3$, respectively.

346 **Model Application.** When the regression model predictions and probabilistic functions are combined,
347 annual mean PM_{2.5} concentration trends driven by emissions coupled with meteorological effects can be
348 quantified. The concentration predicted by the regression model provides an estimation of the annual mean

349 PM_{2.5} under given emissions and average meteorological conditions, whereas a range derived from the
350 probabilistic function at a fixed probability (e.g., 95%) shows fluctuation associated with random variations
351 of meteorological parameters. This approach was then applied to simulate global historical temporal trends
352 of PM_{2.5} concentrations from 1980 to 2014 and to project future trends from 2015 to 2030. Emission-driven
353 trends of global annual mean PM_{2.5} concentrations prior to 2015 were calculated from the gridcell
354 regression models based on PKU series emission inventories¹⁹ and from the RCP (Representative
355 Concentration Pathways)2.6 and RCP8.5 emission scenarios model run for after 2014^{52,53} using emissions
356 for 2014 as a baseline. The results are denoted by the solid line shown in **Fig. S14a**. In the figure,
357 meteorological condition-induced variation ranges are shown by the darkly shaded *UI*₅₀ and lightly shaded
358 *UI*₉₅. We further assume that meteorological conditions for 2014 used as a "normal" year can be extended
359 to future years. For the past 35 years, global annual mean PM_{2.5} concentrations decreased slightly from 13.1
360 μg/m³ (10.4~16.1 μg/m³ as *UI*₅₀) to 12.1 μg/m³ (9.8~14.6 μg/m³), and decreasing trends tend to continue in
361 the future at a slightly faster rate, which could be attributed to increasing awareness and to
362 emission-mitigation efforts made by many developing countries, especially China. We found slight
363 differences in projected PM_{2.5} levels between the two emission scenarios on a global scale prior to 2030. It
364 should be noted that the probability functions were developed based on gridded meteorological parameters.
365 When the results are presented on an area with more than one gridcell, such as a country, a city, or even the
366 globe, the calculated *UI* values are simply averaged over gridcells covering the area. This practice is based
367 on the assumption that all meteorological confounding factors do not vary significantly within the region of
368 concern. This applies to a relatively small region such as the NCP, where a somewhat uniform surface
369 pressure with small pressure gradients is often observed, which in turn produces fewer altered wind and
370 temperature fields across the NCP. However, for a larger region such as China or a region with complex
371 terrain, this assumption would lead to an overestimation of *UI* values. Unfortunately, the accuracy of the *UI*
372 estimation is difficult to enhance, as spatial similarities in changes of meteorological parameters are
373 difficult to quantify. To further validate the model-calculated PM_{2.5} concentrations using the regression
374 models, the calculated PM_{2.5} concentrations for before 2014 are compared to those observed from various
375 monitoring stations (gridcells) over various years in **Fig. S14b-c**. Both calculated annual mean
376 concentrations (dots) and *UI* values (bars, b. *UI*₅₀ and c. *UI*₉₅) are shown, indicating a good agreement.

377 The method was further applied to various countries to predict annual mean PM_{2.5} concentrations subject to

378 the changes in emissions. Corresponding results are shown in **Fig. 5** for 12 countries. The projected $PM_{2.5}$
379 trend for 1980 to 2030 from the regression model was obtained based on RCP2.6 and RCP8.5 emission
380 scenarios^{52,53}. In general, these trends and *UI* values vary significantly across countries. Relatively high
381 levels of variability observed for some countries are associated with stronger changes in meteorological
382 conditions and especially for monsoon regions (e.g., China and Pakistan) where the strength of prevailing
383 monsoons play an important role in aerosol production and dispersion^{10,54}. The results also show that for
384 developed countries such as the United States, France, and Japan, past declines in $PM_{2.5}$ will remain with
385 slight differences between RCP2.6 and RCP8.5 predictions. Trends for France are an exception, as the
386 RCP2.6 assumes a much stronger decrease in pollutant emissions and hence in $PM_{2.5}$ concentrations.
387 Predicted $PM_{2.5}$ concentration trends vary substantially across developing countries. In China, annual mean
388 $PM_{2.5}$ concentrations tend to decrease continuously, which is consistent with considerable efforts made to
389 curb air pollution in recent years¹⁶. For other developing countries such as India and Indonesia, $PM_{2.5}$
390 concentrations are projected to increase continuously until 2020 if the proposed emission scenarios are not
391 altered. As the RCPs dataset provides emission data at a decadal temporal resolution, tipping points from
392 emission incline to decline cannot be precisely identified. Nevertheless, these trends imply that although
393 severe levels of air pollution have spurred widespread awareness and concern from governments and the
394 public, efficient mitigation is still lacking in most developing countries. Meanwhile, it is very likely that air
395 $PM_{2.5}$ concentrations will increase continuously in coming years in developing countries such as Laos and
396 in Central Africa.

397 **Fig. S15** shows three examples of predicted historical and future trends of $PM_{2.5}$ concentrations for three
398 cities for which recent monitoring data are available, based on the RCP2.6 and RCP8.5 emission
399 scenarios^{52,53} for 1980 to 2030. For the city of New York, $PM_{2.5}$ monitoring data for after 2014 suggest that
400 emission-reduction rates likely range between the two scenarios, which are not remarkably different in the
401 first place. For New Delhi, although the observed values still fall within the UI_{95} range, concentrations
402 reported for the last three years exceed the predicted means. Although unusual meteorological conditions
403 could play a critical role in increasing concentrations, relatively high levels of $PM_{2.5}$ observed for 2014 and
404 2016 may indicate accelerated increases in emission and pollution levels. Numerous studies have reported
405 high levels of air pollution in India in recent years⁵⁵. Beijing is one of the most heavily contaminated cities
406 in northern China. Based on both RCP2.6 and RCP8.5 emission scenarios, we find a slight decline in $PM_{2.5}$

407 concentrations after 2014. However, the measured annual mean $PM_{2.5}$ concentrations from 2014 to 2016
408 are well below the predicted ones and even fall below the lower bound of the 95% uncertainty interval. It is
409 likely that mitigation measures applied in the city were more effective than what was planned in RCP
410 scenarios.

411 In summary, the novel method developed in this study serves as a useful tool for quantifying
412 emission-induced changes in $PM_{2.5}$ concentrations by excluding confounding meteorological effects. The
413 approach involves less computation than an atmospheric chemical transport model; hence it can be used in
414 quantitative environments, for health assessments of $PM_{2.5}$ and to evaluate the effectiveness of mitigation
415 efforts. Importantly, we learned from this study that long-term trends rather than declines occurring over a
416 single year are critical to consider when evaluating the effectiveness of mitigation measures while
417 considering meteorology-induced $PM_{2.5}$ fluctuations.

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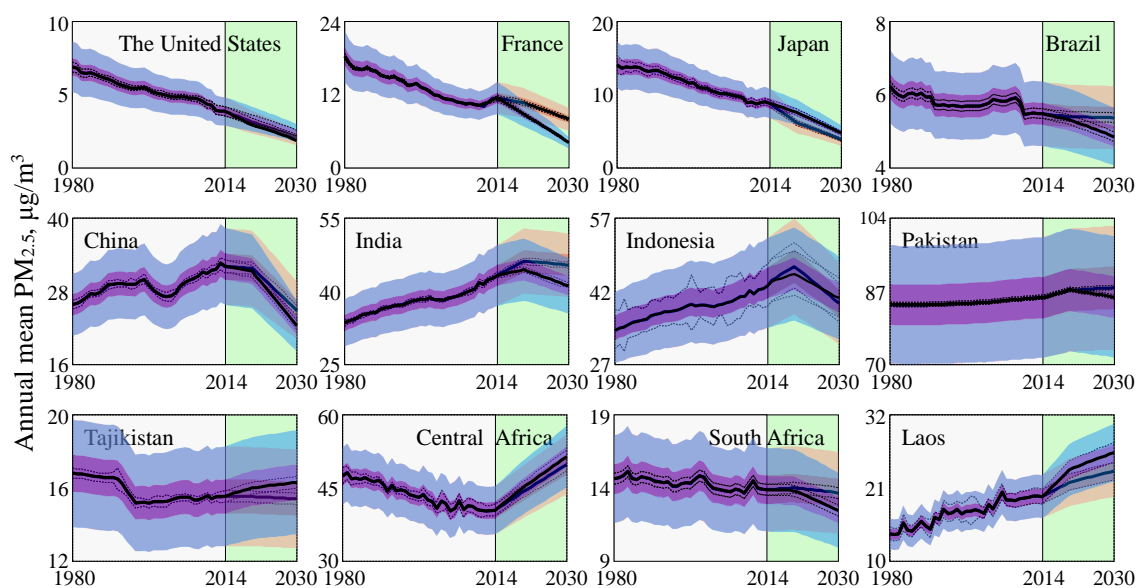


Fig. 5 Temporal trends of $PM_{2.5}$ concentrations for 12 countries for 1980 to 2030 based on the RCP2.6 (blue) and RCP8.5 (orange) emission scenarios. Emission-driven trends are shown as medians (black lines) with a 90% confidence interval (black dash lines). Potential fluctuations induced by meteorological confounding effects are shown as shaded areas as UI_{50} (dark shaded area) and UI_{95} (light shaded area).

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424 **Notes**

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430 **ASSOCIATED CONTENT**

431 Supporting Information. Detailed results of the sensitivity analysis for key pollutants, various model
432 validations, spatial distributions of major meteorological parameters, comparisons drawn between
433 emissions and PM_{2.5} concentrations, frequency distributions of emissions and PM_{2.5} concentrations,
434 spatial distributions of regression model R^2 values, meteorological effect-induced variations,
435 cumulative distributions of CVs, and predicted trends for 3 cities are freely available at
436 <http://pubs.acs.org>.

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