

## Charting the trajectory of forgetting: Insights from a working memory period paradigm

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## Abstract

Working memory capacity is commonly measured in terms of its item-span, and much less often in terms of its time-span, or ‘period’. The former measures how many items can be stored in working memory when carrying out episodes of concurrent processing. The latter complements this by determining the duration of processing episodes that can be tolerated while successfully storing a fixed number of items. We investigated the generality of previous evidence that working memory period varies with the distribution of longer and shorter processing episodes within a trial, and that notwithstanding such differences, a global measure of period is a reliable predictor of children’s educational attainment. We describe data from 184 children between 7 and 11 years of age, who completed variants of an operation period task with different distributions of processing episodes together with measures of scholastic attainment. Individual differences in period scores were consistent over two test sessions, and predictive of reading and number skills. We replicated previous effects of the order of longer and shorter processing episodes, but found that they did not generalize fully to other manipulations of order. The results point to the contribution of subtle within-trial sequence configurations for working memory. We make the case for a broader view of what constrains working memory than in current models.

Keywords: working memory; operation period; individual-differences; forgetting rate

## 1. Introduction

Our goal in this paper is to investigate ‘working memory period’, a (complementary) alternative to the widespread use of working memory span as a measure of capacity (Towse, Hitch, Hamilton, Peacock & Hutton, 2005). We introduce the rationale for measuring working memory in this way and examine its characteristics, by replicating and then extending previous analyses of both experimental task manipulations and individual-differences. Cronbach (1957) emphasised the value of such an integrated approach to experimental and differential approaches 60 years ago, and its relevance for working memory theory has been echoed since (Conway et al., 2007a).

The starting point for the present project is the widely held view that working memory is a limited capacity system supporting the maintenance and processing of transient representations (Baddeley, 1986; Baddeley & Hitch, 1974). Within cognitive psychology, the concept of working memory is used to help understand a wide range of phenomena, ranging from the inhibition of saccadic eye movements to such complex activities as playing chess (Crawford, Smith & Berry, 2017; Robbins et al. 1996). Its success has become intertwined with the enduring popularity of a family of tasks designed to measure working memory capacity, such as counting span, reading span and operation span (Case, Kurland & Goldberg, 1982; Daneman & Carpenter, 1980; Turner & Engle, 1989).

These span tasks deliberately have a common structure. On a typical trial participants perform a sequence of processing operations (e.g., counting a display of objects, reading a sentence, carrying out a numerical calculation) and each delivers a memorandum (e.g., an array total, the final word in a sentence, the result of a calculation). At the end of the trial participants attempt to recall the memoranda in the

order they were generated. Trials differ in sequence length, and span estimates the limit on the simultaneous retention of memoranda in working memory. These tasks are referred to generically as ‘complex span’ tasks, and assess how much material can be successfully maintained in working memory when attention is required also for processing operations. They predict an impressive range of cognitive and real world behaviours (see Conway, Jarrold, Kane, Miyake & Towse, 2007b).

Complex span provides tremendous value as a psychometric instrument and has also been deployed in many experiments testing detailed models of working memory processes (e.g. Saito & Miyake, 2004; Towse & Hitch, 1995). Whilst acknowledging this success, one of the potential structural limitations of complex span tasks is that, inherently, they only measure the number of items that can be recalled when working memory is also engaged in episodes of processing. That is, the paradigm is designed to assess the *residual memory* capacity of the working memory system. As Towse, Hitch & Horton (2007) note, complex span maps onto a *suitcase* metaphor for the limit on working memory, such that the key variable is how much material can be packed at any one time. Indeed, because complex span is frequently the *only* task and performance metric, it carries a heavy burden to account for the full range of working memory phenomena researchers may wish to explore (see also Cowan et al., 2008).

Potential alternatives exist to an exclusive focus on evaluating span size. Towse et al., (2005) proposed a measure called ‘working memory period’, designed to assess the endurance or *longevity* of representations when processing also engages working memory. Underlying this paradigm are two notions; first, that some individuals may be able to withstand a longer filled retention interval than others, ie. to recall information over a longer period, and second, the possibility that endurance reflects a separable dimension from storage capacity. This notion maps onto an alternative

*vacuum flask* metaphor for the limit on working memory (Towse et al., 2007) such that a key variable is how long a flask insulates recently activated material from loss to the ambient environment, over and above its volume.

Towse et al., 2005, reported working memory period data from children. Matching complex span, episodes of processing generate accompanying memoranda. However, unlike complex span the number of processing episodes remains fixed as trials progress. Instead, the durations of the processing episodes are increased in steps or levels in order to establish the longest overall duration that can be tolerated for successful serial-order recall of the memoranda. For ease of administration and scoring, working memory period is scored in terms of the number of levels through which participants progress. Towse et al. (2005) studied different versions of the period task (a reading period task – analogous to reading span– and operation period – analogous to operation span). They found that working memory period correlated with complex span and with measures of cognitive ability. Furthermore, the period task showed sensitivity to experimental effects that replicated and extended within-trial order effects found previously in span tasks (these are described in more detail below). Thus, whilst initially developed to permit measurement on a separable dimension, working memory period is not entirely orthogonal to working memory span. Towse, Hitch, Hamilton & Pirrie (2008) also reported that correct recall times (ie. production durations) increased with period difficulty level, consistent with the idea that the fidelity of memory representations was degrading, and so required more time to be prepared for output.

The importance of rapid forgetting fits with a wide range of developmentally-based evidence that faster (general) processing speed is associated with better working memory performance (e.g. the cascade model - Fry & Hale, 1996; Hale,

Myerson, Emery, Lawrence & Dufault, 2007). Potentially, faster processing speed reduces the amount of exposure to information degradation. At a task-specific level, children's complex span covaries with the speed at which accompanying processing operations can be completed (Bayliss, Jarrold, Gunn, & Baddeley., 2003; Case et al., 1982; Towse & Hitch, 1995). Moreover, other work converges on the more specific idea that forgetting rate is a separable working memory parameter. In particular, Bayliss & Jarrold (2015) report that one component of working memory capacity variance can be traced to the rate of forgetting in the Peterson and Peterson task (see also discussion in Jarrold, 2017). Consequently, an endurance-based dimension of working memory, as advocated above, aligns with a range of empirical work. This is broadly consistent with theoretical interpretations such as the task-switching account (Towse & Hitch, 1998) and the Time-Based Resource-Sharing Model (TBRS; Barrouillet & Camos, 2007; Barrouillet, Gavens, Vergauwe, Gaillard & Camos, 2009) that emphasise the importance of rapid forgetting during intervals in which participants undertake processing.

In the present study, we explored the generality of experimental effects in children's operation period that we have previously interpreted in terms of within-trial forgetting, while at the same time seeking further evidence for an association between individual differences in working memory period and scholastic attainment (Towse et al., 2005). We consider these in turn.

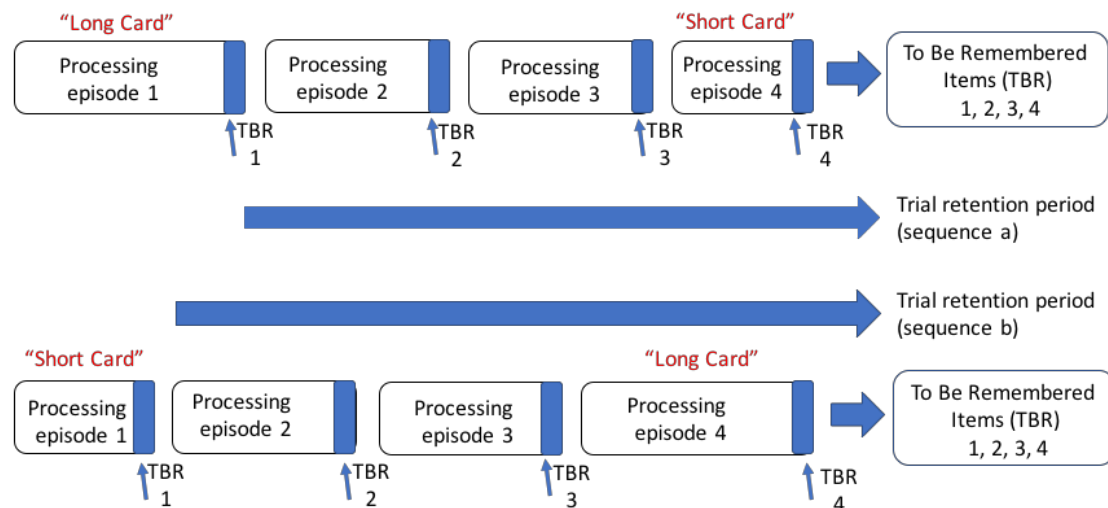
#### Within-trial forgetting

Towse et al. (2005) found that working memory period showed the same 'card order' effects as working memory span (Towse et al., 1998). This terminology reflects the presentation of each processing episode on an image of a card in a computerised

display. The logic of card order effects is as follows. If a trial begins with a short processing episode (card), compared to a long processing episode (card), then the retention demands commence earlier, insofar as retention begins with generation of the first memorandum (e.g. the result of a math operation or the number of array targets counted). Moreover, if in one condition a trial starts with a short processing episode and ends with a long processing episode, and in another condition this order is reversed, then the total ensemble of processing and storage activities is the same, and any differences in recall are attributable to within-trial order effects.

Based on a span paradigm, Towse et al (1998) reported that ‘short-final’ trials ending with a short processing episode led to higher spans than ‘long-final’ trials. This occurred for counting span, operation span and reading span in children (a finding also replicated for counting span by Ransdell & Hecht, 2003), and was taken to support the hypothesis that working memory span is affected by the amount of within-trial forgetting (see also adult data from Maehara & Saito, 2007). Findings were interpreted according to a task-switching model whereby there is no functional opportunity for maintenance during processing episodes which thus serve to postpone the point of recall. This is illustrated in Figure 1, which shows that a short-final trial (completion order a) has a shorter overall retention requirement than a long-final trial (completion order b). The TBRS model of Barrouillet and Camos (2007) differs from the task-switching account because it assumes that ‘attentional refreshing’ during processing episodes can be used to offset forgetting, to an extent that varies inversely with the cognitive load imposed by the processing episodes. However on this account too, short final trials should lead to better performance, benefitting from a combination of a shorter cumulative retention interval and a lower cognitive load.

Figure 1, Schematic representation of 4-item working memory span trial, based on Towse et al., 1998. Blue areas for each processing episode represent the identification or specification of the To Be Remembered (TBR) item. The arrows depict the retention phase, which differs between order (a) and (b).

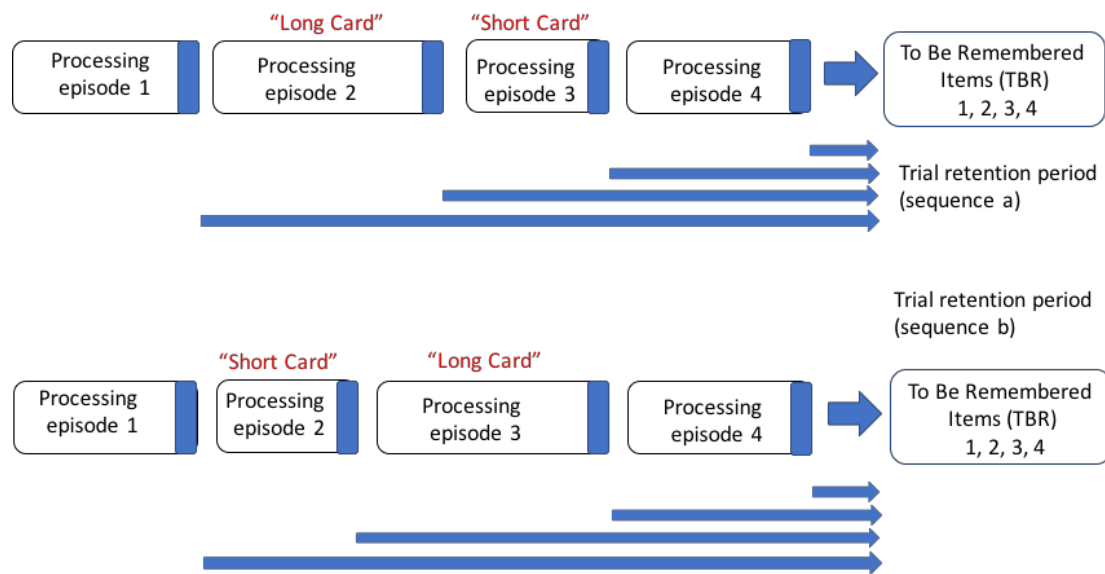


Subsequently, Towse et al (2005; Experiment 3) showed corresponding order effects in the working memory *period* paradigm, with a working memory advantage for short-final trials of four items than long-final trials. Towse et al. found a similar working memory advantage when short and long cards were placed in the central rather than end positions – period level was larger with a *long (second)* and *short (third)* sequence compared to a *short (second)* and *long (third)* sequence. Thus, the order effect was not specific to end-item manipulations. This led to an elaboration of the schematic model for retention effects, depicted in Figure 2. In this case, the durations of processing episodes 2 and 3 were fixed (to be short or long) and those of 1 and 4 were variable and increased progressively through test administration to determine period. A key element to this model is the characterisation of a working memory task as the ensemble of a set of item retention trajectories, not just a unitary composite with a start and stop point. Assessing the adequacy of this characterisation



requires analysis of other novel, sequence permutations. This forms a major element of the present study.

Figure 2. Schematic representation of a 4-item operation period trial. Long and short processing episodes occur at middle positions, which affects (only) the duration of intermediate sequence items. Based on Towse et al. (2005, Figure 7).



In the present study, we sought to replicate and extend these sequence order manipulations. Therefore we first created two sequence versions that exactly replicated the order effects described above –manipulating either end positions or middle positions (the outer and inner segments of a four-item sequence). Second, we also created two additional and novel sequences: manipulating order in the first half of a trial, and in the second half of a trial. These are all shown in Figure 3. Together, the sequences explore the conditions under which the ensemble of retention intervals matter for recall accuracy of the set.

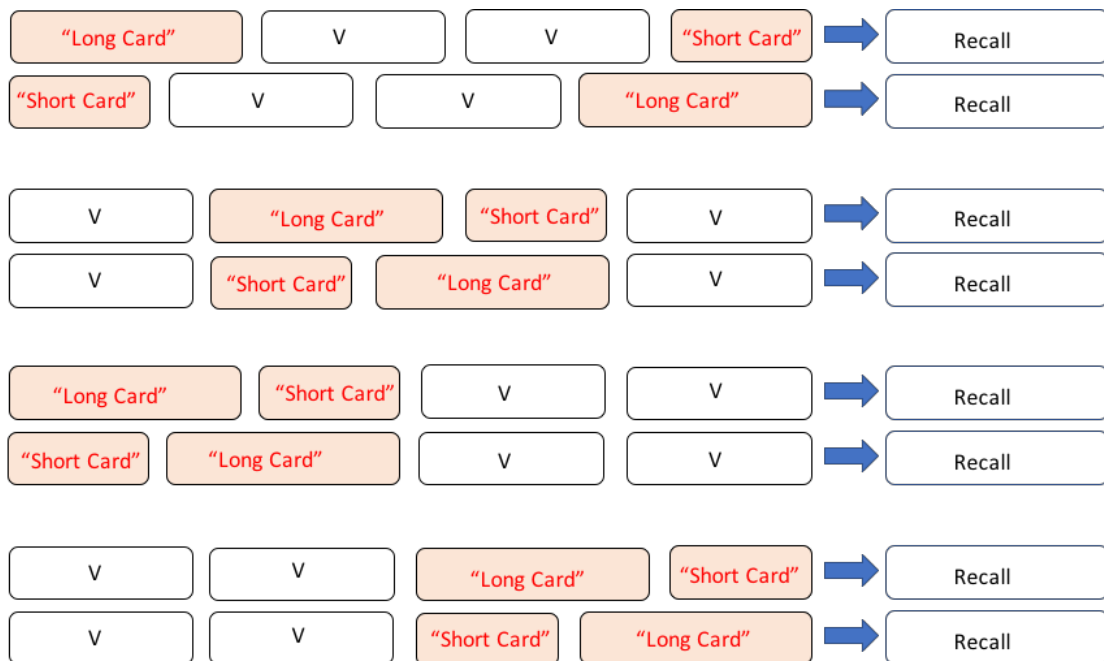
Figure 3. Illustration of the 4 permutations of the card order effect

manipulated in the study “V” cards are variable in that they increase with task level.

Thus, from top to bottom, this represents LVVS / SVVL; VLSV / VSLV; LSVV /

SLVV; VVLS / VVSL. Participants complete each version of the pair in

counterbalanced order.



It is important to point out that our investigation of these additional conditions was exploratory, in that our main aim was to establish whether previous evidence for the importance of item retention trajectories in working memory tasks extends to novel permutations of such trajectories. Our task-switching framework is deliberately simplistic and consequently, it underspecifies complex patterns of performance. It does nevertheless make the general prediction that for each pair of conditions illustrated in Figure 3, working memory performance will be superior when the short processing episode comes after the long one. Our main objective was to see if we do indeed find such a pattern, as this would encourage further elaboration and testing of

more detailed models within the task-switching framework. If, on the other hand, the general prediction is not upheld, this would suggest the framework is too simplistic, most probably in its assumption that participants are passive throughout. If participants are more active, task strategies represent a powerful extra factor to incorporate within performance.

### Individual differences

A key aspect of complex span is its ability to predict individual differences in higher-level cognition. This is applicable to a range of tasks completed by adults (Engle, Kane & Tuholski, 1999). It is also relevant to children at primary school level, where complex span predicts aspects of reading and mathematics skill both concurrently and longitudinally (Hitch, Towse, & Hutton, 2001). In this context, an important outcome from previous studies is that working memory period associates with measures of scholastic attainment. Complementing this, working memory period shows some statistical overlap with its better-known span paradigm. This individual-difference perspective supports the idea that the endurance of working memory – or put another way the rate of forgetting or item loss (Bayliss & Jarrold, 2015) – is a relevant construct that contributes to individual differences and is to some extent distinguishable from capacity measures.

In the present study, we attempted to replicate this link between individual differences in working memory period and scholastic attainment across children of primary school age. In addition, we sought to confirm the reliability of the period measure (Towse et al., 2005, Experiment 1 test-retest reliability estimate was .72). We also investigated whether any relationship between working memory and scholastic attainment could be isolated to particular segments, that is serial positions, within a

working memory period sequence. The fixed list-length architecture of the period paradigm makes this a feasible question to address systematically, unlike orthodox span trials wherein list length varies across trials and participants. Recent evidence distinguishing secondary memory (early list segments) and primary memory (late list segments) components of free recall lists add weight to this question (Roome, Towse & Jarrold, 2014; Unsworth & Engle, 2007) especially since Unsworth and Engle (2007) suggest, at least in adult data, that secondary memory provides the cornerstone of complex span performance.

## 2. Method

We report how we determined our sample size, all data exclusions (if any), all manipulations, and all measures in the study (Simmons, Nelson, & Simonsohn, 2012).

### 2.1 Participants

The analysis sample comprised 184 children from both rural and urban primary schools in the North West of England (8 additional children did not complete working memory assessments and are not described further). Date of birth was missing for one child, but for the remaining sample, mean age was 9 years 6 months (range 7;8 to 11;9) There were three age groups separated by class assignments. No more than one week after initial assessment, all but 10 of the sample were available to undertake a second working memory assessment. Written parental consent was provided for each participating child.

Based on prior work (ie using comparability not formal power estimates) sample sizes were initially proposed in a grant funding application. These were adjusted through reviewer-suggested design modifications (to collapse conditions), and finalized through availability of children in class for whom parental consent was provided.

## 2.2 Procedure

Both working memory assessments were administered across a delay of approximately one week, with task completion order varied (absences meant three children could contribute only partial data). In the first round of data collection, children also completed British Abilities Scale II subtests for word reading and maths attainment

An Apple Macintosh Powerbook 5300c controlled experimental tasks, using a RuntimeRevolution software environment, a form of HyperCard stack programming. Administration of operation period followed the procedure reported in Towse et al. (2005). With the use of laminated instruction cards, the experimenter initially explained that arithmetic sums, of the type shown on the cards, would appear on the computer. The child calculated the answer to a problem and reported this verbally. In addition, eight computer-presented sums (comparable to experimental stimuli) were presented without any concurrent memory task as further practice. Feedback appeared after each response to encourage calculation accuracy.

2.2.1 Operation period. Each trial comprised four self-paced sums followed by an auditory and visual cue to recall – verbally – the four answers in serial order. In each experimental condition two of the processing episodes were fixed and two were variable (see Figure 3). For each condition testing involved sets of three trials at a given stage level with two variable processing episodes determined by independent data on average solution times (see below). Stage level and thus processing duration progressed in successive sets of three trials. Provided that at least two of the three trials at a given level were recalled correctly, a further set was presented at the next level. <sup>1</sup> Testing continued in this way until children either failed to recall correctly two

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<sup>1</sup> An early version of period software was developed and made available here,

of the three trials at a particular level, or they reached the maximum level. An audio-visual message congratulated the child whenever s/he successfully recalled at least two of three lists in a set.

We implemented exactly same task rule as Towse et al., 2005 – item recall was scored as correct when it matched the derived answer supplied by the child to the original problem, even if their arithmetic calculation was erroneous. When children made multiple processing errors on a trial, they received feedback encouraging processing accuracy.

The corpus of arithmetic problems is described in Hamilton, Towse, Hitch & Hutton, (2001), who used empirical data from an independent sample of children to derive estimates of computational duration and accuracy. Problems were selected for each of seven levels of duration and depicted in Towse et al. (2005; Figure 1).

Examples of problems for each duration from level 1 (short and fast) to level 7 (long and slow) are as follows; “ $4+0=$ ”; “ $3+1=$ ”; “ $5-1+0$ ”; “ $4+1-1$ ”; “ $3-1+2$ ”; “ $5-1-1+1$ ”; “ $2+1+2-0-1$ ”. As can be seen, levels differed with respect to the number of arithmetic steps and the size of computational operations (0, 1 or 2), while keeping the answer the same. Towse et al (2005) showed that on average, children took almost twice as long to complete long cards ( $M=6.1s$ , drawn from stage level 6 above) than short cards ( $M=3.2 s$ , drawn from stage level 2).

2.2.2 Scholastic attainment. Children received the Number Skills and Word Reading sub-scales of the British Abilities Scale (II) (BAS), the former in a group setting, the latter individually. Older children began at a higher basal level as per BAS

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<http://www.lancs.ac.uk/staff/towse/wmperiod.html> and a configurable software implementation of the working memory period paradigm, written by James Stone, also available here: <http://www.cognitivetools.uk/cognition/tasks/Verbal-WM/workingMemoryPeriod/>

instructions. Both are graded tests of attainment in core curriculum activities. The Number Skills test covers for example, oral number transcription, written arithmetic computations, and at higher levels, fractions and long division. Word Reading involves reading aloud a series of (regular and irregular) single words.

### 3. Results

Raw data for this study are available at

<https://dx.doi.org/10.17635/lancaster/researchdata/260>

Towse et al. (2005) presented operation period data in terms of the highest stage-level at which recall was accurate (analogous to the largest list length that permits accurate recall in span tasks). That is, from level 1-7, at what point did children fail to recall correctly all the 4 derived solutions within a trial, for the majority of trials? We report the same measure here, likewise incorporating partial credit for recall accuracy at the terminal trial level (Towse & Hitch, 1995)<sup>2</sup>.

#### 3.1 Trial configuration effects

Initially we segregated the different conditions whereby processing length is varied (ie. the *card order effect*) into the four permutations shown in Figure 3. Towse et al. (2005) reported a working memory advantage in the short-late compared to long-late condition (where short-late reduces the retention profile). We aggregated data over the two test sessions, and we similarly calculated a card order effect for each participant, i.e., the difference between working memory period levels (the period score when short cards appears after long, minus the period score when long cards appeared after short).

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<sup>2</sup> The accompanying raw data also document performance based on the number of successful list recalls (see Conway et al., 2005, for an analysis of analogous working memory span scoring algorithms). Statistical conclusions are not affected by choice of scoring measure.

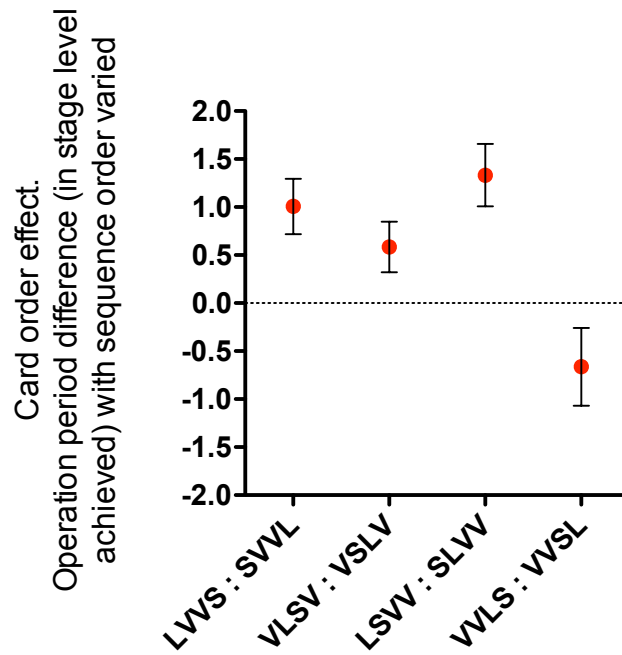
In the current data, the card order effect was systematic, positive and substantial with a short-final (and thus long-first) sequence (ie. LVVS vs SVVL in Figure 3),  $t(43)=3.492$ ,  $p=.001$ ,  $\eta^2= .221$ . This is illustrated in Figure 4. Next, we examined the card order effect established by Towse et al, where short and long late cards are manipulated in central rather than end positions, (ie. VLSV vs. VSLV in Figure 3). We again found a positive and systematic short-late recall advantage,  $t(43)=2.213$ ,  $p=.032$ ,  $\eta^2= .102$ .

We then tested what is to our knowledge a novel order effect: where short and long cards are both placed in the first half of the sequence (ie. LSVV vs. SLVV). As with the configurations assessed above, this also yielded a systematic short-late advantage,  $t(42)=4.098$ ,  $p<.001$ ,  $\eta^2= .286$ . The effect size was again substantial. Of course, the term short-late here is relative, since both short and long cards are positioned in the first half of the sequence.

Finally, we examined the effect of manipulating the order effect in the second half of the sequence, which forms the natural complement to the previous configuration (ie. VVLS vs. VVSL). In this instance, there was no systematic short-late advantage,  $t(42)=-1.632$ ,  $p=.110$ ,  $\eta^2= .060$ , and the effect size was the smallest of the comparisons made. The contrast between this configuration and others is also evident in Figure 4.

Figure 4. The size and direction of the card order effect. Operation period difference score in stage level, between a short card following a long card, and a long card following a short card, for each sequence manipulation. Error bars describe one standard error on each side.

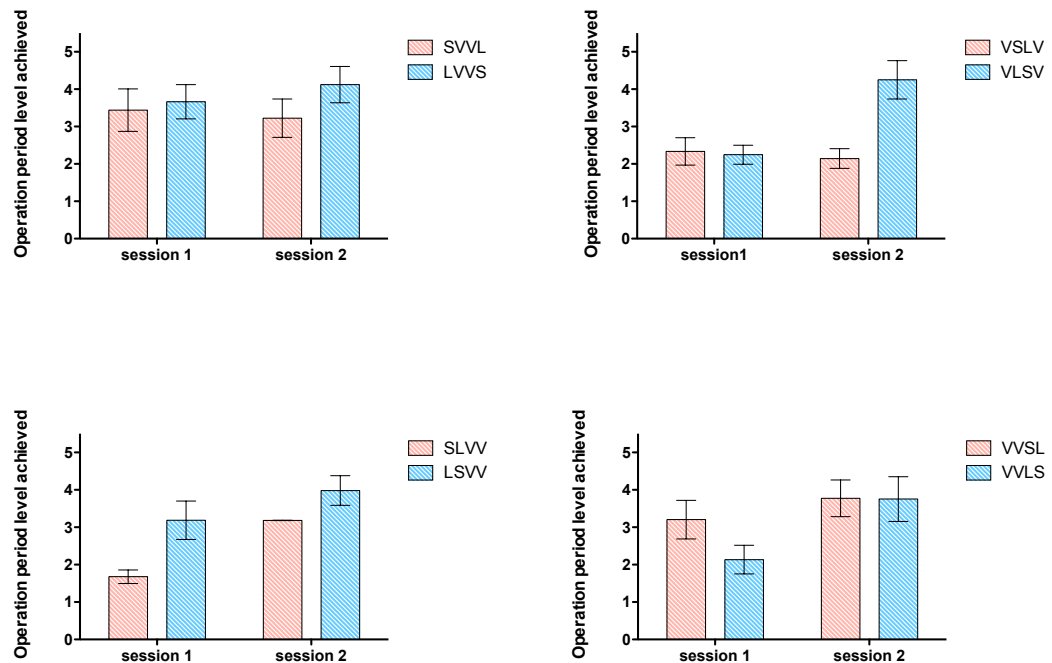




These analyses provide a focused investigation of the card order effect in its various permutations, comparable with Towse et al. (2005). However, we observed that task performance in general shows a robust and sizeable practice or repetition effect, with a 30.2% improvement in operation period on second administration,  $t(173)=5.038, p<.001, \eta^2=.128$ . The above card order analyses collapse across this general improvement in performance, over which card order configurations were counterbalanced. Therefore, we next examined the relationship between card order effects and session.

First we examined the original card order effect (separating SVVL followed by LVVS from LVVS followed by SVVL), finding a significant interaction with session,  $F(1,42)=4.603, p=.038, \eta_p^2=.099$ . As can be seen in Figure 5, the short-final advantage was stronger in the second session. This prompted us to re-analyse data originally reported by Towse et al (2005). A (previously unanalysed) session by card order effect interaction was evident there also. Thus, the pattern reported here is not unique to the current study.

Figure 5. Card order effects and session effects for period level achieved. Error bars describe one standard error on each side.



Second, we examined the card-order effect in central rather than terminal positions (separating VSLV followed by VLSV from VLSV followed by VLSV). This also yielded a significant interaction with session,  $F(1,42)=14.938, p<.001, \eta_p^2=.262$ . Again the short-late advantage was stronger when tested in the second session<sup>3</sup>.

Third, we examined short and long cards manipulated in the first half of the sequence (SLVV and LSVV alongside its counterbalanced pair). We found once more an interaction with session,  $F(1,41)=18.078, p<.001, \eta_p^2=.306$ . In this case, the recall advantage to a short-late card was stronger in the first session comparison<sup>4</sup>.

<sup>3</sup> In this analysis, we also found a significant overall practice effect,  $F(1,42)=10.686, p=.002, \eta_p^2=.203$ , and a difference between the two orders,  $F(1,42)=5.902, p=.019, \eta_p^2=.123$ .

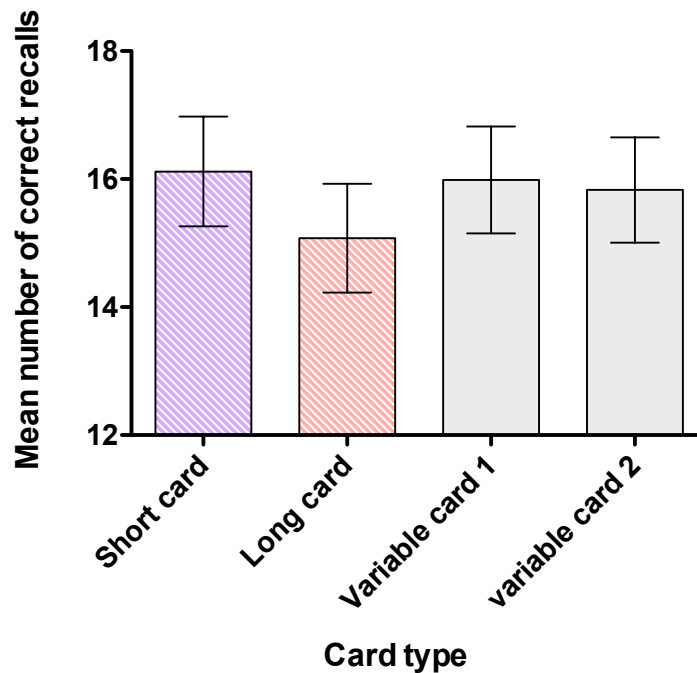
<sup>4</sup> Here, the main effect (ie. benefit) of practice was also significant,  $F(1,41)=15.086, p<.001, \eta_p^2=.269$ , but no systematic difference between order configuration,  $F<1, \eta_p^2=.015$

Finally, we examined the effect of manipulating the card order effect in the second half of the sequence, (VVSL and VVLS, alongside the counterbalanced match). In this case we did not obtain a reliable interaction with session,  $F(1,41)=2.427, p=.127, \eta_p^2=.056$ . The main effect (ie. benefit) of practice was reliable,  $F(1,41)=7.454, p=.009, \eta_p^2=.154$ , while the difference between orders was not,  $F<1, \eta_p^2=.020$ .

We consider the implications of these card order effects, and the asymmetric transfer effects across session, in the Discussion. At this point, we simply note that working memory endurance is systematically affected by *several but not all* sequence permutations, with large effect sizes in two cases, a moderate effect in one, and a small-moderate reverse effect in the condition where order permutation did not significantly affect recall. Thus, previous card order phenomena have been replicated and extended, but interestingly, a short-late advantage is not found under all circumstances.

The data also permit a different type of question to be investigated focusing on recall accuracy associated with each type of processing card (i.e. long, short, first variable card, second variable card). This is described in Figure 6. Analysis confirmed that answers to long cards were less well recalled than answers to short cards  $t(183)=6.993, p<.001, \eta^2=.211$  and variable cards (e.g. long vs. first variable card,  $t(183)=6.579, p<.001, \eta^2=.211$ ), while the two variable cards were not significantly distinguishable,  $t(183)=1.328, p=.186, \eta^2=.010$ . We defer interpretation of these data to the Discussion.

Figure 6. Recall performance associated with the card type products. Error bars describe one standard error on each side.



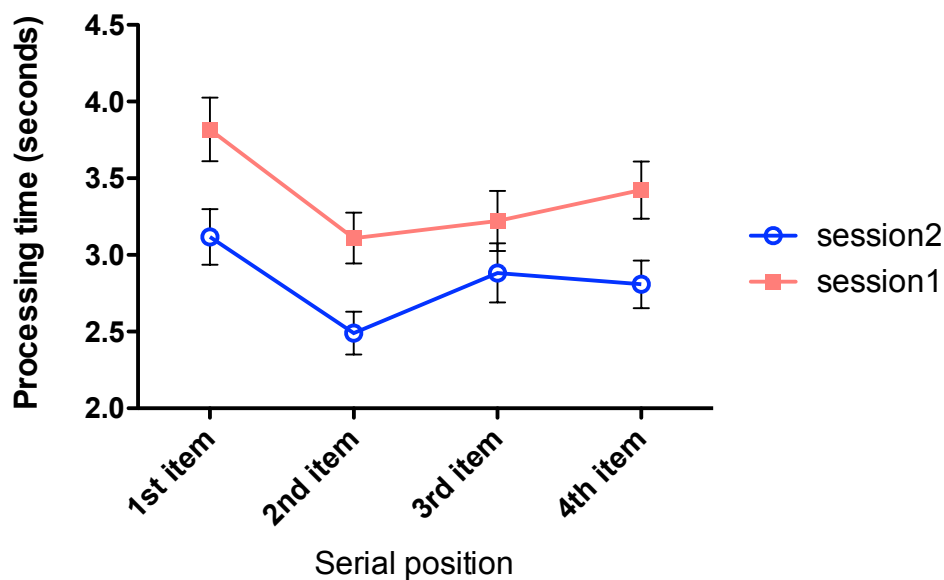
### 3.2 Arithmetic operations: processing time analysis

To analyse the time taken to respond to arithmetic problems, we based performance on the first set of 3 trials only (and not subsequent sets, even when these were potentially available). Even though this produced a sparse dataset, we wanted to derive a measure that was equivalent (all children experienced the first set of trials) and thus comparable (beyond the first set of trials, arithmetic operations changed, so the sampling space differed). Overall performance is described in Figure 7. These data emphasise several features. First, solution times on the second session are consistently quicker than those on the first session (approximately .5s for each answer calculated, or 2s from the initial problem presentation time until the recall cue). This offers one insight into the practice benefits in the memory recall data— on the second session, quicker responses reduced the retention duration for the trial by over 10%

Second, for neither session do the data exhibit a monotonic increase in duration across serial position. This is noteworthy insofar as accounts of working memory that draw upon the notion of resource-sharing or a trade-off between

processing and memory (Case et al, 1982; Barrouillet et al., 2004), should predict that processing times will be longer at later serial positions, because of the increasing burden of handling computations alongside retention of prior answers. Furthermore, even if one accepts that the first serial position may be affected by “start of trial” or “startup” costs. we found no main effect of serial position when examining just serial positions 2 – 4,  $F(2,182)=1.896, p=.153, \eta_p^2=.020$  (there was a strong session effect,  $F(1,183)=22.9, p<.001, \eta_p^2=.109$ , but no reliable interaction,  $F<1, p=.650, \eta_p^2=.005$ ).

Figure 7. Processing times for arithmetic problems as a function of serial position and session. Error bars describe one standard error on each side.



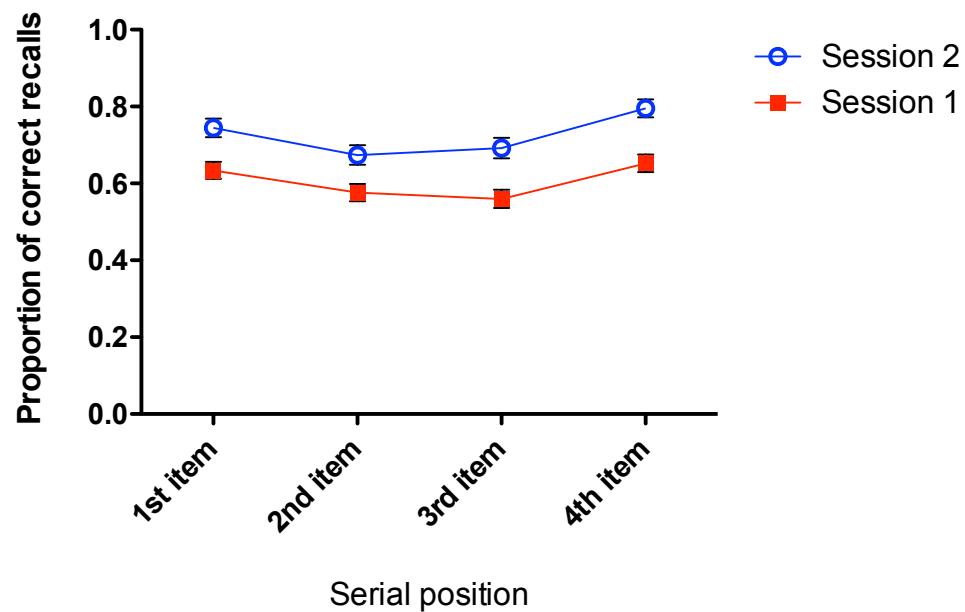
### 3.3 Recall accuracy as a function of serial position

We calculated the accuracy of recall as a function of each serial position on the first three task trials (ie. level 1). These are reported in Figure 8 for the two sessions. A conventional (albeit mild), bow-shaped recall function replicates the pattern of data reported in Towse et al. (2005; see Figure 2). The data also illustrate

the improved second session recall that is visible at each serial position. More detailed analysis of serial position accuracy across period task levels is reported in the appendix.

Figure 8. Proportion of correct recalls as a function of serial position and session.

Error bars describe one standard error on each side.



### 3.4 Individual differences

We have already noted that in absolute terms operation period performance improved from session 1 to session 2. However, performance also showed systematic stability with respect to individual differences. Working memory period correlated across sessions,  $r(172)=.514, p<.001$ , and processing speed – which is itself a common associate / mediator of children’s working memory performance – was also consistent across sessions,  $r(172)=.619, p<.001$ . The reliability estimate of period is strong, although slightly smaller than reported in Towse et al. (2005) which was derived from a smaller and more restricted age sample, without the different card order permutations and a slightly different scoring algorithm (based on aggregating different recall accuracy thresholds). Table 1 reports the relationships between the

principal variables of interest in the study. There were 6 children with missing data on the BAS number skills score, 4 children with missing data on word reading, and 2 children who lacked data on both assessments. We derived a composite measure of number and reading skills by averaging z-scores on each raw score variable (using just the single z-score for the partial data noted above). We also created a composite measure of operation period, by combining z-scores from each assessment where available and a measure of arithmetic (operation) processing speed based on solution times for first set of trials.

Of particular note, working memory period robustly correlated with a composite ability measure from BAS tests, and this relationship remained significant once partialling out children's age. It was also the case that period was correlated with ability specifically for the youngest age group,  $r(59)=.316$ ,  $p=.013$ , for the middle age group,  $r(60)=.537$ ,  $p<.001$ , and for the oldest age group,  $r(57)=.449$ ,  $p<.001$ . Moreover, these relationships also persisted after controlling for age (in months) within each age band, with  $r(57)=.316$ ,  $p=.015$ ,  $r(59)=.498$ ,  $p<.001$ , and  $r(56)=.456$ ,  $p<.001$  respectively. The data support the conclusion that representational endurance, as measured by working memory period, is a stable and meaningful characteristic across the sampled age groups, and specifically at each age band, with a numerical trend towards a stronger relationship amongst older children.

For participants with data on all three measures, operation period correlated specifically with both number skills  $r(170)=.489$ ,  $p<.001$  and word reading,  $r(170)=.449$ ,  $p<.001$ . These two correlations did not significantly differ in strength,  $z=.71$ ,  $p=.239$ . That the numerical trend is for a stronger link with number skills is not at all surprising given that operation period involves simple numerical calculations. Indeed, that these are statistically equivalent reinforces the view that the primary task

demand in this configuration of operation period is the retention element and not the processing content.

To summarise, there is a healthy association between the overall operation period score and the aggregated measure of scholastic skills. Since the operation period task uses a fixed list-length structure, it is also feasible to investigate whether this relationship is carried by all elements of the sequence, or whether for example, early or late serial positions (presumably affected by primacy and recency processes respectively) differentially contribute to the predictive profile. We calculated participant recall accuracy for each serial position, separately for session 1 and 2, and found a strong convergence in that there were consistent associations between position-specific accuracy and BAS scores ( $r_s(182) > .343$ , all  $p_s < .001$ , for positions 1-4). Individual differences in ability scores and working memory endurance were not mediated by specific portions of the recall curve, nor were they reliant solely on the aggregation of positional data.

As noted above, we also replicate a finding consistently obtained across multiple studies of children (see Towse & Hitch, 1995) – individual differences in working memory ability are linked the speed with which the processing elements of the task can be accomplished<sup>5</sup>. Children who quickly answer the arithmetic problems tend to be children with larger working memories – here they are children who can endure a longer retention period and still effectively access the to-be-remembered items.

#### 4. Discussion

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<sup>5</sup> Processing speed is measured with the first set of trials only, so as to avoid the confound of mixing different trial sets that have different length configurations.



In first reporting on the working memory period, paradigm, Towse et al. (2005) argued that “the present research represents an important description of the *potential* of a novel measure of working memory.” Our study confirms and extends that potential, with respect to individual differences and experimental analysis. To justify this claim, we discuss the terminology adopted, key findings and their interpretation.

In this paper, card order effects have been specified by referring to a short-late advantage in sequence order. We should note that this descriptive label is partly one of convenience only, since the data are not available to demonstrate unequivocally whether it is primarily a short-late advantage – foreshortening the delivery of a recall cue – or a long-late disadvantage – delaying the delivery of a recall cue - or indeed an amalgam of both. We believe it would require a different type of experimental design to arbitrate satisfactorily between these possibilities.

We introduced the card order analyses with respect to the schematic model in Figure 1. The key feature captured by that model was that short-final and long-final sequences can be distinguished by when retention commences. This model assumes that recall begins earlier and occupies more time in the long-final (short-first) condition. That model was later elaborated into a second version, Figure 2, (Towse et al., 2005) by making the explicit distinction between overall recall duration for the trial and trajectories for each item in the set. This elaboration helped to understand the finding of a short-late and not just short-final advantage.

We have replicated both the short-final and short-late advantage with working memory period. This confirms the important but often overlooked conclusion that working memory retention and recall is not a single act. Instead, memoranda can exist

at different levels of fidelity. In addition, we have also reported data reinforcing the view that the task-switching model, and its family neighbours, even where it explains the majority of effects, does not explain the totality of data.

In particular whilst there is evidence that a short-late advantage can be obtained when short and long processing episodes appear in the first half of the ensemble, there is no comparable advantage when they appear in the second half of the ensemble. Also, we found that card order effects vary with session repetition, for which task switching does not offer a simple explanation. Accordingly, whilst the data confirm the importance of taking detailed, within-trial temporal perspectives into account, modelling temporal trial effects is likely insufficient on its own.

The impetus for the present work has been to explore predictions from a simple task-switching hypotheses (Towse & Hitch, 1995; Towse et al, 1998; Towse et al., 2005), though this account is not alone in proposing temporal constraints on working memory performance. In particular the TBRS (eg, Barrouillet & Camos, 2007) includes a loss-and-refresh cycle for representational fidelity, based on multiple rehearsals though a trial (micro task-switches; see also Towse et al., 2002). For example, Portrat and Lemaire (2015) model successive decay and refresh trajectories of a single item over an interval, where decay weights more strongly than refresh, and leads overall to loss of activation (see Figure 1, Portrat & Lemaire, 2015). This is in many ways a non-linear version of Figure 1, focusing on a single memorandum.

Consistent with Portrat and Lemaire (2015) the empirical data point to the heuristic value of considering the ensemble activation and time parameters of a working memory trial. Data also emphasise the relevance of how individual representations are retained. Yet, none of current models are complete because some

sequence order manipulations do not impact recall in the predicted way. In this respect, the current patterns of data present challenges for both the task switching account and the TBRS model, because within the loss-and-refresh perspective in the latter model, longer cards should generate additional decay through attentional processes switched away from all working memory representations at that point.

In summary most card order manipulations confirm the importance of item retention but they defy a simple, unitary account. One possible interpretation from Figure 4 is that the largest effect occurs when short and long cards are manipulated in the first two positions, and the next largest effect occurs when a short or long card is placed in the first position. Placement of short / long cards in the middle positions generates a smaller effect and placing them in the final two positions leads to the smallest (negative) effect. In other words, a *post-hoc* description of the data is that bigger card order effects occur when they are placed early within the serial position sequence, and smaller effects occur the later in the serial position sequence.

How might these inconsistent sequence permutation effects be explained? In short, we do not have a comprehensive answer. However, one potential clue to understanding all the models' limitations lies in the large main effect of practice on working memory period and interactions whereby card order effects change with practice across session. These interactions suggest that recall benefits not only from processing speed changes but from the deployment of different performance strategies. Different refreshment strategies have been modelled in other work (see Lemaire, Pageot, Plancher & Portrat, 2017) and whilst we cannot identify the specific strategies that account for the practice effects, simply their presence serves as an

important reminder that almost all existing models are dependent on essentially unexplored assumptions about strategies. Strategy changes can be rapid, occurring within a session (Towse, Cowan, Horton & Whytock, 2008) and they can also be slower and longer-term, underlying working memory training (Guye & von Bastian, 2017; Stone, 2016). In terms of the Baddeley and Hitch (1974) working memory framework, these data highlight the need for modelling the *flexibility* with which central executive processes can be deployed to support working memory.

A second clue comes from the observation that answers from long cards were harder to remember than the answers from short cards, with the latter more closely resembling variable length cards (see Figure 6). The equivalence between memory for the answers to short and variable cards may have arisen from many children faltering at a fairly early stage of the period task. This is because in the determination of working memory period, variable cards start off more similar to short cards, and then progressively become more like long cards as the level of difficulty is increased. The greater difficulty of remembering answers to long cards is striking evidence that period is influenced by factors other than within-task retention intervals. We note that long cards involved more arithmetic operations and therefore more interim results, and we suggest that these may interfere with memory for the final result.

Irrespective of these specific interpretations, data underline the value of measuring working memory period. That is, the attempt to measure the endurance of working memory representations, whilst keeping constant the number of independent items held in memory, is illuminating. We note that it would be very difficult to measure different permutations of card order effects within a span paradigm, because list length must vary across trials in order to assess capacity. Likewise, it would be

hard to determine the selective contribution of recall at specific serial positions for predictive power (as we consider in more detail below) when the serial position vary with list length. In summary, the characteristics of working memory period afford novel perspectives into some of the processes that support working memory performance.

It is clear that explaining the retention requirements of representations is important— as within-task retention duration increases, the probability of correct recall declines. Yet whilst this is the case for many card order sequences, it is not true for all of them. The present data implicate a number of other factors that determine the period of working memory, including processing speed and practice effects, strategies, and interference (see also Posner & Konick, 1966). This argues against simple models, but without detracting from the value of measuring working memory period.

The value of working memory period is also demonstrated through the evidence that children's performance correlates robustly with scholastic attainment – indeed despite large practice effects - and this is true also throughout the serial position list. Empirically, an index of representational endurance is shown to be both stable and linked to scholastic ability - there is reliability and predictive validity. Importantly, this offers converging evidence that forgetting rate is a relevant attribute of working memory (see also for example, Jarrold, 2017). Forgetting rate has been often overlooked in studies of complex working memory that focus instead on capacity, but it is increasingly apparent that it affects performance, and there is growing evidence that we can develop tasks to successfully capture this parameter.

Just as we have advocated the value of implementing a period procedure for illuminating working memory issues, we should also point out some of the arguments we are explicitly not making. First, as should be clear from what we have said above, we do not wish to claim that time is necessarily the causal *mechanism* for informational loss. For example, the task structure always delivers four TBR items and has easier trials to start with. Being more likely to be correctly remembered, these items are thereby available to interfere with subsequent trials, clearly providing the opportunity for the build-up of proactive interference across trials. To clarify -our proposal is that endurance is a useful metric for measuring working memory, not a simple temporal causal mechanism of forgetting.

Second, whilst the period task keeps list length constant, we do not claim that volume- or capacity-related issues do not contribute to period task demands. There is a constant volume of things to remember, which is likely non-trivial for some children. These two concepts are intertwined; in just the same way, a working memory capacity metric is not an instantaneous trial format, and thus involves endurance. We refer to different metaphors for memory – suitcases and vacuum flasks (Towse et al., 2007). We regard these as useful perspectives that highlight relevant dimensions. Yet these dimensions cannot be entirely orthogonal and independent, and both neglect the important role of, for example, executive processes in complex task performance. And third, we are not claiming that capacity, endurance or speed of processing represents the sole constraint on working memory performance. There is abundant evidence for other contributory processes that shape the quality of encoding (eg. chunking; Cowan, 2010), maintenance (eg. mapping onto longer-structure representatiuons; Ericsson & Delaney, 1999) and recall (eg. recall reconstruction; Towse, Hitch, Horton & Harvey, 2010)

In conclusion, we reiterate that it is valuable conceptually to show that working memory phenomena do not completely rely on a single paradigm family focusing on volume metrics. That most theoretical and empirical working memory research is informed by some version of a suitcase measurement metaphor, is an important recognition in its own right. The present data shows that this can be usefully augmented by modelling the flask-like properties of memory. Highlighting the endurance of memory representation does not offer a sufficient or complete account of working memory, yet it is we argue feasible, tenable, coherent and informative.

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Table 1. Relationships between variables

	1	2	3	4	5
1.Age (months)					
2.BAS (number)	.597				
3.BAS (word read)	.458	.638			
4.BAS (composite)	.581	.908	.909		
5.Operation period (composite)	.368	.525	.496	.544	.438
6. Processing speed	-.405	-.517	-.470	-.530	-.438

## 7. Appendix

Serial position analysis is reported in the main text. A more detailed analysis is also possible, that describes recall in terms of the period test level administered (ie. the recall difficulty in terms of the endurance required of representations). These data are reported in Figure 9. It may seem paradoxical in these data that accuracy *increases* with difficulty level. However, later test levels are based on successively smaller samples (these are specified on the x-axis labels), and thus comprise performance from the more-able children. The figure also demonstrates that the serial position curve is not constant across all task levels – in particular for session 1 the primacy advantage evident at the start of performance declines. In comparison for session 2 data, the primacy and recency effects are more evident throughout.

Figure 9. Proportion of correct recalls as a function of task level and serial position. Upper panel represents data from session 1; lower panel represents data from session 2.



