



**Proton pump inhibitor use among adults**  
**Mapping the landscape of PPI use and exploring its effect**  
**on cancer risk and mortality**

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**Thesis for the degree of Philosophiae Doctor**

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**UNIVERSITY OF ICELAND**  
**SCHOOL OF HEALTH SCIENCES**

FACULTY OF MEDICINE

**Notkun prótónupumpuhemla á Íslandi**  
**Kortlagning PPI notkunar og möguleg áhrif hennar á**  
**krabbameinsáhættu og lifun**

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The road to wisdom?

Well, it's plain  
and simple to express.  
Err and err and err again,  
but less and less and less.

- Piet Hein



## Ágrip

**Inngangur og markmið:** Prótónupumpuþemlar (PPI) eru sýrubindandi lyf sem eru almennt notuð við meðferð á ýmsum magasýrusjúkdómum. Notkun PPI lyfja er útbreidd á heimsvísu og sú umtalsverða aukning sem hefur orðið á notkun þeirra hefur verið gagnrýnd í ljósi mögulegrar ofnotkunar og óvissu sem ríkir um mögulegar skaðlegar aukaverkanir. Sýrubindandi virkni PPI hefur verið talin hafa möguleg krabbameinshindrandi áhrif vegna hæfni þeirra til að hindra virkni sérhæfðra sýruseytandi ensíma. Talið er að slík ensím taki þátt í myndun á súru utanfrumuumhverfi krabbameinsfruma.

Markmið okkar var nota lýðgrundaða gagnagrunna á Íslandi til að I) kortleggja notkun PPI lyfja meðal fullorðinna einstaklinga á Íslandi, II) meta áhættu PPI notenda á því að greinast með brjóstakrabbamein, blöðruhálskirtilskrabbamein, eða sortuæxli í húð, og III) meta möguleg verndandi áhrif PPI lyfjanotkunar á lifun sjúklinga með blöðruhálskirtilskrabbamein.

**Aðferðir:** Rannsókn I var lyfjanotkunarrannsókn þar sem við lýstum notkun PPI lyfja á árunum milli 2003 og 2015. Við áætluðum árlegt nýgengi og algengi PPI notkunar, lengd PPI lyfjameðferðar og samhliða notkun lyfja sem geta haft í för með sér blæðingar í meltingarvegi. Rannsókn II var tilfella-viðmiðsrannsókn þar sem tilfelli voru einstaklingar sem greindust með brjóstakrabbamein, blöðruhálskirtilskrabbamein og sortuæxli á milli 2005 og 2014. Hvert og eitt krabbameinstilfelli þöruðum við saman við allt upp að 10 viðmið eftir almanaksári, fæðingarári, og kyni. Við áætluðum PPI notkun þátttakenda, þ.e. hvort leyst hefði verið út að minnsta kosti eina PPI lyfjaávisun, hvort notkun væri  $\geq 1000$  skilgreindum dagskömmtum (DDD) og heildarnotkun og reiknuðum út gagnlíkindahlutföll (ORs) og 95% öryggisbil (CIs) fyrir áhættuna á því greinast. Rannsókn III var hóprannsókn þar sem einstaklingar á aldursbilinu 40 til 85 ára sem greindust með krabbamein í blöðruhálskirtli á milli 2007 og 2012 mynduðu rannsóknarhópinn. Við áætluðum upphaf PPI notkunar (fyrir eða eftir greiningu), heildarnotkun og lagskiptum eftir klínískri stígun. PPI notkun var meðhöndluð sem tímaháð breyta og Cox aðhvarfsgreining var notuð til að reikna út hættuhlutfall (HRs) fyrir dauða af völdum blöðruhálskirtilskrabbameins annars vegar og dauða af öllum orsökum hins vegar með 95% öryggismörkum (CI).

**Niðurstöður:** Niðurstöður úr rannsókn I sýndu að heildarnotkun PPI lyfja á Íslandi fór ört vaxandi á rannsóknartímabilinu. Þótt nýgengi hafi haldist stöðugt jókst algengi PPI notkunar úr 8.5 á hverja 100 einstaklinga árið 2003 yfir í 15.5 á hverja 100 einstaklinga árið 2015. Ennfremur, reyndist algengi hækka með hækkandi aldri og 22% sjúklinga var enn að nota PPI einu ári eftir að meðferð hófst.

Niðurstöður úr rannsókn II bentu ekki til þess að PPI notkun hafi í áhrif á krabbameinsáhættu (ORs 1.03; 95% CI: 0.92-1.16 fyrir brjóstakrabbamein, 1.12; 95% CI: 1.00-1.25 fyrir blöðruhálskirtilskrabbamein og 0.84; 95% CI: 0.69-1.12 fyrir sortuæxli). Sömuleiðis virtist PPI notkun  $\geq 1000$  DDDs ekki hafa áhrif (OR 0.97; 95% CI: 0.78-1.19 fyrir brjóstakrabbamein, 1.20; 95% CI: 0.99-1.47 fyrir blöðruhálskirtilskrabbamein, og 0.59; 95% CI: 0.40-1.13 fyrir sortuæxli). Niðurstöður okkar bentu ekki til þess að tengsl væru á milli heildarnotkunar á PPI og áhættunnar á því að greinast með brjóstakrabbamein, blöðruhálskirtilskrabbamein, eða sortuæxli.

Niðurstöður úr rannsókn III bentu ekki til þess að PPI notkun eftir greiningu hefði áhrif á líkur á dauða af völdum blöðruhálskirtilskrabbameins (HR 0.88; 95% CI: 0.52-1.48) eða dauða af öllum orsökum (HR 1.02; 95% CI: 0.73-1.43). Upphaf PPI notkunar virtist ekki hafa áhrif, en HRs fyrir dauða af völdum blöðruhálskirtilskrabbameins voru 0.45 (95% CI: 0.21-0.98) meðal sjúklinga sem notuðu PPI lyf samfelld bæði fyrir og eftir greiningu og 1.12 (95% CI: 0.61-2.08) á meðal nýrra PPI notanda. HRs fyrir dauða af öllum orsökum voru 0.67 (95% CI: 0.43-1.04) á meðal sjúklinga sem notuðu PPI samfelld og 1.25 (95% CI: 0.82-1.92) á meðal nýrra PPI notanda. Lagskipting eftir heildarnotkun PPI lyfja og klíniskri stigun leiddi ekki ljós tölfræðilega marktækt samband á milli PPI notkunar og lifunar.

**Ályktun:** Niðurstöður verkefnisins benda til þess að PPI notkun hafi aukist umtalsvert á Íslandi yfir síðasta áratuginn; sér í lagi hjá eldri einstaklingum. Þar að auki er stór hluti sjúklinga meðhöndlaður lengur en mælt er með í klínískum leiðbeiningum fyrir lyfin. Niðurstöður okkar benda hvorki til þess að PPI notkun hafi áhrif á áhættu á brjóstakrabbameini, blöðruhálskirtilskrabbameini, eða sortuæxlum, né að hún hafi áhrif á lifun meðal sjúklinga með krabbamein í blöðruhálskirtli.

#### **Lykilorð:**

Prótónupumpuþemlar, lyfjafaraldsfræði, krabbamein, sýruseyting, lýðgrunduð rannsókn

## Abstract

**Background and aims:** Proton pump inhibitors (PPIs) are commonly prescribed drugs that are used to treat acid-related disorders of the gastrointestinal tract. Over the last decade, PPI use has repeatedly been shown to be increasing worldwide, causing concerns due to reports of unsubstantiated long-term use and potential adverse effects. However, PPIs have also been suggested to promote antineoplastic effects in certain cancer settings via inhibition of specialized proton pumps. These proton pumps are involved in pH regulation in eukaryotic cells and believed to act as facilitators for the acidification of the tumor microenvironment (TME).

Our aim was to use the population-based resources available to us in Iceland I) to assess the utilization of PPIs among the adult outpatient population residing in Iceland, II) to explore the potential of PPIs possessing an antineoplastic effect by estimating the risk among PPI users of being diagnosed with a first-time breast cancer, prostate cancer, or malignant melanoma, and III) to assess the potential influence of post-diagnosis PPI use on mortality among prostate cancer patients.

**Materials and methods:** In study I, a drug utilization study, we investigated changes in overall PPI use between 2003 and 2015 among the adult outpatient population in Iceland. We estimated changes in annual incidence and prevalence, duration of PPI treatment, and the concurrent use of ulcerogenic drugs. In study II, a nested case-control study, we identified incident cases of breast cancer, prostate cancer, and malignant melanoma between 2005 and 2014. For each case, up to 10 controls were matched on birth-year, sex, and calendar year using risk-set sampling. Assessing ever use, high use, and cumulative use of PPIs, we calculated odds ratios (ORs) and 95% confidence intervals (CIs) using conditional logistic regression. In study III, a cohort study, we identified patients diagnosed with prostate cancer between 2007 and 2012 among adult residents of Iceland aged between 40 and 85 years. PPI use was modelled in a time-dependent manner. Assessing post-diagnosis use, timing of use, cumulative use and stratifying by clinical stage we estimated the associations with prostate-cancer specific and all-cause mortality using Cox proportional hazard regression models and 95% CIs.



**Results:** In study I, we observed a marked increase in outpatient PPI use over the last decade. Although the annual incidence remained fairly stable between 2003 and 2015, the annual prevalence estimates rose from 8.5 per 100 persons in 2003 to 15.5 per 100 persons in 2015. Furthermore, we found that prevalence increased with age and that 22% of patients were still being treated with PPIs one year after treatment initiation.

In study II, we observed the following adjusted odds ratios (ORs) associated with ever use and high use of PPIs, respectively: 1.03 (95% CI: 0.92-1.16) and 0.97 (95% CI: 0.78-1.19) for breast cancer, 1.12 (95% CI: 1.00-1.25) and 1.20 (95% CI: 0.99-1.47) for prostate cancer, 0.84 (95% CI: 0.69-1.12) and 0.59 (95% CI: 0.40-1.13) for malignant melanoma. In secondary analyses, we did not observe a pattern consistent with a dose-response relationship for these three cancer types.

In study III, we did not observe a statistically significant association between post-diagnosis PPI use and prostate cancer-specific mortality (HR 0.88; 95% CI: 0.52-1.48) or all-cause mortality (HR 1.02; 95% CI: 0.73-1.43). In secondary analyses, stratification by timing of use yielded adjusted HRs of 0.45 (95% CI: 0.21-0.98) among continuous PPI users and 1.12 (95% CI: 0.61-2.08) among new PPI users for prostate cancer-specific mortality. For all-cause mortality, we observed adjusted HRs of 0.67 (95% CI: 0.43-1.04) and 1.25 (95% CI: 0.82-1.92) among continuous users and new users, respectively. Stratification by cumulative dose and clinical stage did not reveal a statistically significant association with post-diagnosis PPI use for the mortality outcomes of interest.

**Conclusions:** In conclusion, our observations indicate that PPI use in Iceland has increased considerably over the last decade; especially among older adults. Additionally, a high proportion of patients were treated for longer periods than clinical guidelines recommend. Furthermore, our findings do not support a chemopreventive role of PPIs in attenuating the risk of being diagnosed with a first-time breast cancer, prostate cancer, or malignant melanoma. Finally, our results do not indicate that post-diagnosis PPI use influences mortality among prostate cancer patients.

**Keywords:**

Proton pump inhibitors, pharmacoepidemiology, population-based, cancer, V-ATPase

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## List of abbreviations

|  |   |
|--|---|
| ADP                                    | Adenosine diphosphate                             |
| ATC                                    | Anatomical Therapeutic Chemical Classification    |
| ATP                                    | Adenosine triphosphate                            |
| BMI                                    | Body Mass Index                                   |
| CI                                     | Confidence Interval                               |
| DDD                                    | Defined Daily Dose                                |
| ER                                     | Estrogen Receptor                                 |
| GERD                                   | Gastroesophageal Reflux Disease                   |
| H <sup>+</sup> -K <sup>+</sup> -ATPase | Hydrogen Potassium ATPase                         |
| HER2                                   | Human Epidermal Growth Factor Receptor 2          |
| HIC                                    | High Income Country                               |
| HR                                     | Hazard Ratio                                      |
| ICD                                    | International Classification of Diseases          |
| LMICs                                  | Low and Middle Income Countries                   |
| NCSF                                   | NOMESCO Classification of Surgical Procedures     |
| NICE                                   | National Institute for Health and Care Excellence |
| NSAID                                  | Non-Steroidal Anti-Inflammatory Drug              |
| OR                                     | Odds Ratio  |
| PPC                                    | Proportion of Patients Covered                    |
| PPI                                    | Proton Pump Inhibitor                             |
| PR                                     | Progesterone Receptor                             |
| PSA                                    | Prostate Specific Antigen                         |
| PUD                                    | Peptic Ulcer Disease                              |
| RCT                                    | Randomized Clinical Trial                         |
| SES                                    | Socioeconomic Status                              |
| TME                                    | Tumor Microenvironment                            |
| TNM                                    | Tumor Node Metastasis                             |
| V-ATPase                               | Vacuolar H <sup>+</sup> -ATPase                   |
| WHO                                    | World Health Organization                         |

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## List of original papers

This thesis is based on the following original publications, which are referred to in the text by their Roman numerals:

- I. Hálf dánarson ÓÖ, Pottegård A, Björnsson ES, Lund SH, Ogmundsdóttir MH, Steingrímsson E, Ogmundsdóttir HM, Zoëga H. *Proton-pump inhibitors among adults: a nationwide drug-utilization study*. Therapeutic Advances in Gastroenterology 2018; 11: 1-11.
- II. Hálf dánarson ÓÖ, Fall K, Ogmundsdóttir MH, Lund SH, Steingrímsson E, Ogmundsdóttir HM, Zoëga H. *Proton pump inhibitor use and risk of breast cancer, prostate cancer, and malignant melanoma: An Icelandic population-based case-control study*. Pharmacoepidemiology and Drug Safety 2018; 1-8.
- III. Hálf dánarson ÓÖ, Pottegård A, Lund SH, Ogmundsdóttir MH, Ogmundsdóttir HM, Zoëga H. *Use of proton pump inhibitors and mortality among Icelandic patients with prostate cancer*. Submitted.

## **Declaration of contribution**

In close collaboration with his supervisor, Óskar Örn Hálfðánarson, planned and designed the research work of papers I – III. Óskar Örn applied for the appropriate ethical and research approvals, conducted the statistical analyses for all three studies, drafted each manuscript, and wrote this thesis under the sound guidance of his doctoral committee.

# 1 Introduction

A rapid increase in proton pump inhibitor (PPI) use has in recent years given rise to concerns about the appropriateness of their use and potential adverse effects that might stem from staying on PPI treatment for extended durations of time (Batuwitage et al., 2007). Given the high usage of these drugs it is important to determine all potentially associated safety considerations for their use; both adverse and beneficial. Due to their inherent function as acid inhibitors, PPIs have been proposed to be able to function as potential antineoplastic agents by promoting pH homeostasis in certain cancer settings (Luciani et al., 2004).

In this PhD project, we explore the landscape of PPI use among the adult population in Iceland, evaluate the risk of incident diagnoses of breast cancer, prostate cancer, and malignant melanoma among PPI users, and finally aim to determine whether PPI use influences mortality among prostate cancer patients.

## 1.1 Pharmacoepidemiology

Pharmacoepidemiology is a rather young discipline that can be seen to bridge the research areas of clinical pharmacology and epidemiology. It is defined as the study of the use of drugs and their consequent effects, both beneficial and adverse, in large numbers of individuals (Strom et al., 2012). Pharmacoepidemiology emerged from the need to address increasing concerns of adverse effects by developing methods to assess the safety profiles of drugs. These concerns were compounded by events such as the “thalidomide disaster” which took place around 1960, where rare birth defects could be traced back to the fetus being exposed to thalidomide via maternal use of the drug during pregnancy (Strom et al., 2012).

Pharmaceutical drugs are extensively evaluated in pre-clinical and clinical phases prior to marketing and randomized controlled trials (RCTs) are considered to be the gold standard when it comes to estimating the effectiveness of the drugs under study (Akobeng, 2005). However, RCTs are not always applicable when estimating the safety of a drug, such as when the adverse effects are rare or take a long time to develop. Furthermore, the small sample sizes often seen in RCTs tend to be relatively homogenous. Thus, they are not always comparable to the general population, making it

hard to predict the overall real-world benefits and risks of a particular drug (Strom et al., 2012; Vandenbroucke, 2004). Furthermore, in addition to pre-marketing clinical trials being limited by duration, extent, and patient characteristics, there are valuable insights to be gained from observational studies that are conducted in the post-marketing phase. In fact, they can sometimes reveal previously undetected beneficial or adverse effects and they can also test the effectiveness of drugs under different conditions and within patient populations that were not adequately represented in pre-marketing clinical trials (Guess, 2005). Pharmacoepidemiological studies are therefore able to provide further insight into the safety profiles and effectiveness of previously marketed drugs due to the possibility of greater follow-up and the tracking of real-world drug use and prescription patterns within large populations (Strom & Tugwell, 1990).

A key aspect of pharmacoepidemiological research is the availability of relevant data on drug use. Since its advent, pharmacoepidemiology has developed hand in hand with technological advancements. The growth of medical databases has led to ever-increasing amounts of data that previously were collected and compiled in a time-consuming manner (Wettermark, 2013). For instance, in the Nordic countries population-based registries have been established in each country that cover the majority of all dispensed drugs within their populations (Furu et al., 2010). However, although the amount of available data has increased considerably over the years, there are still several methodological challenges that arise and need to be considered when data on drug use is utilized in pharmacoepidemiological studies.

### **1.1.1 Nationwide prescription registries in the Nordic countries**

High-quality data sources are essential for registry-based research in the field of pharmacoepidemiology. A crucial component of such data sources is the ability to be able to compile and store data on an individual level, using a personal identification number that is unique to every individual for each record entry. Regional databases that collected individual-level data in specific regions of Sweden (Boethius & Wiman, 1977) and Denmark (Hallas et al., 2017), were important precursors of the national level prescription registries of today. With the ushering in of the computer age, it became possible to efficiently collect individual-level data on every filled prescription from pharmacies within entire countries. Today, all five Nordic countries have established centralized databases that cover each country's entire population that hold information on individual-level data on dispensed drugs dating back

to 1994 in Denmark (Pottegård et al., 2016b) and Finland (Klaukka, 2009), 2003 in Iceland, 2004 in Norway, and 2005 in Sweden (Furu et al., 2010; Wettermark et al., 2007).

There are a lot of similarities between the Nordic prescription registries. First, the data structure is the same and the variables are categorized based on their nature, i.e. whether the data is related to the patient, the prescriber, the drug, or the pharmacy (Furu et al., 2010). Second, the parliaments in the respective countries have given informed consent on behalf of their populations for everyone to be included in the national health registries (Rosén, 2002). Third, the prescription databases are all based on personal identification numbers that are unique to every resident of each country. Personal identification numbers are important because they allow the data to be linked to other national registries that hold data on other variables, potential outcomes and confounding factors, which then facilitates the study of potential effects of drug exposures (Wettermark et al., 2013). Due to the similarities between the Nordic countries, and their shared history and cultural ties, and their frequent collaborations in general, cross-national pooling of pharmacoepidemiological data for research purposes is an intriguing prospect. Such collaborative efforts have several potential benefits, such as allowing for assessment of possible variations between countries, strengthening research competencies, and increasing the sample size of studies. Taken together, the national prescription databases in the Nordic countries cover around 27 million individuals. Thus, the Nordic countries are well placed to collaborate on high-quality pharmacoepidemiological studies with large underlying populations (Wettermark et al., 2013).

However, there are also some challenges that come along with combining cross-national data. While the healthcare systems and access to data on exposures and outcomes are similar between countries (Furu et al., 2010), there can be some administrative and logistic challenges that come with the combining the data in one place. Furthermore, even small differences in the record-linking process and access to clinical variables can cause some difficulties when performing studies where the focus is on the outcome of drug therapy, although descriptive cross-national drug utilization studies might be easier to carry out (Wettermark et al., 2013).

Even though the Nordic prescription registries cover entire national populations, allow for linkage of data with other relevant registries such as cancer registries, cause-of-death registries, population registries, and inpatient registries and contain vast amounts of data on dispensed drugs,

they do not include information on the underlying indications behind each prescription nor the prescribed daily dose, which is a limiting factor for some pharmacoepidemiological studies. Furthermore, the prescription registries do not contain information on the majority of non-prescription over-the-counter (OTC) drugs, which is a potential cause of misclassification bias (Furu et al., 2010).

All in all, the nature of the Nordic prescription registries allows the Nordic countries to collaborate on high-quality population-based, cross-national pharmacoepidemiological studies. Such collaborations may enhance the field of public health by contributing to a deeper understanding of real-world drug use and raising awareness on previously unknown effects of drugs, thus promoting the development of safer and more effective treatment protocols. Today, studies that are based on the Nordic prescription registries have paved the way for new knowledge on drug utilization and effectiveness, and have increased the safety of prescription drug use in the society (Wettermark et al., 2013).

### **1.1.2 ATC/DDD drug classification system**

The Anatomical Therapeutic Chemical Classification System (ATC) with Defined Daily Doses (DDD) was developed in Norway as a modified and extended version of a previous classification system, used by the European Pharmaceutical Market Research Association (EphMRA). This system is recommended by the World Health Organization (WHO) as the international standard for drug utilization studies. The DDD is a measuring unit based on the assumed average maintenance dose per day for a drug that is being used for its main indication in adults. It is a unit that is technical in nature and was originally developed for use in drug utilization studies where it is important to have a clear and stable classification system, as well as a standardized unit of measurement, when presenting and comparing statistics of drug consumption at an international level. The DDD unit should not be assumed to necessarily reflect actual prescribed dosages, since those can drastically vary based on individual characteristics of the patients and other considerations (WHO, 2018). The Nordic prescription registries utilize the Anatomical Therapeutic Chemical/Defined Daily Dose (ATC/DDD) classification to classify the prescription drugs that are recorded in the registries.

The basis of the ATC/DDD system is its five level hierarchical classification of the active substances of the drugs that are being classified. The first level of classification is based on organ or system on which the drugs act, and has

14 distinct anatomical/pharmacological groups. The remaining levels (2nd – 5th) break down the ATC main groups and categorizes them based on their chemical, pharmacological, and therapeutic properties (WHO, 2018).

### **1.1.3 Biases in pharmacoepidemiological studies**

Observational studies in the field of pharmacoepidemiology are subject to some of the same biases as epidemiological studies in general. These include systematic errors such as selection bias and information bias. Selection bias occurs when the selection of study participants is not representative of the total population. This bias makes it impossible to conclude anything meaningful in a larger context from a study suffering from this bias (Guess, 2005). In pharmacoepidemiological studies, selection bias can be avoided by utilizing information on drug exposure from large data sources, such as the Nordic prescription registries, that cover entire populations (Wettermark et al., 2013). Information bias is a result of the inaccurate collection of information relating to the study subjects. These inaccuracies tend to cause misclassification of some of the important variable under study relating to exposures, outcomes, or covariates (Guess, 2005). As an example, studies that require patients to recall previous exposure to specific drugs might be subject to misclassification of the information provided. Also, misclassification in pharmacoepidemiological studies might stem from prescription registries not containing information on OTC drug use where a patient might be misclassified as unexposed due to an OTC purchase of a drug.

Confounding is another common bias in epidemiological studies that is an important source of concern. It is a consequence of an imbalance in the distribution of important patient characteristics between exposed and unexposed subjects (Suissa, 2009). A confounding factor is not situated in the causal pathway between the exposure and outcome of interest but has a strong association with the exposure and is a risk factor for the outcome. If confounding is not adequately dealt with in study design or the analysis phase, then it will bias the effect of the exposure on the outcome under study. The ways in which to control for a confounder in the analysis phase might include stratification, standardization, multivariable regression, and the application of a propensity score, while restriction and matching might be used to control for confounding in study design (Klungel et al., 2004). A prerequisite for confounder adjustment is that the confounding factor is measured. However, it might also be the case that it is poorly measured, or



unmeasured, in which case we refer to such factors as unmeasured, unobserved, or residual confounders (Uddin et al., 2016).

Confounding by indication is another form of confounding bias that is especially relevant when discussing potential biases in pharmacoepidemiological studies. A medical intervention such as drug use is usually supported by an underlying clinical indication based on specific symptoms displayed by a patient. This type of bias is of special concern when the indication itself increases the risk of the outcome that is being studied (Guess, 2005). Examples of studies where confounding by indication was an issue include a study examining the use of calcium channel blockers and the risk of myocardial infarction (Psaty et al., 1995), and studies on PPI use and the risk of gastric cancer (Poulsen et al., 2009); where the underlying indication for drug use is a risk factor for both the exposure and outcome. Another type of bias is protopathic bias, sometimes also referred to as reverse causality or reverse causation. As the name suggests, reverse causation occurs when the outcome precedes and leads to the exposure of interest, i.e. the exposure does not cause the outcome but rather the outcome causes the exposure (Guess, 2005).

Cohort studies are an important observational study design and frequently used in pharmacoepidemiology. One bias that has been frequently seen to arise in pharmacoepidemiological cohort studies is a form of time-related bias called immortal time bias, where an exposure to a drug is determined based on filled prescriptions during follow-up (Suissa, 2007). The concept of immortal time indicates that there is a period during follow-up, often from the moment of cohort entry until an exposure definition has been met, during which a specific end point, i.e. death or another study outcome, is by definition unable to occur (Levesque et al., 2010). In other words, in order to be exposed to a drug, a patient would have been required to survive this time period in order to receive a prescription. Therefore, misclassification of the exposure during the immortal time period, or exclusion of this period altogether, can then give rise to immortal time bias. By failing to appropriately account for the immortal time as an unexposed period of time, the results of an analysis, comparing exposed subjects to those that are unexposed, will ultimately give rise to an effect estimate that will be biased downward. In the context of pharmacoepidemiology, this would provide an artificial association that would, if the bias goes undetected, result in a false conclusion of a drug providing a protective effect in relation to a given outcome. This happens, in essence, because the follow-up time of exposed subjects contains a period of time where they are artificially protected, i.e. they are unable to experience

the study outcome until they become exposed (Suissa, 2007). There are however several ways to circumvent the issue of immortal time bias. One option would be to avoid a time-fixed definition of exposure by conducting a time-dependent cohort analysis, where the immortal time is correctly classified as an unexposed period (Weberpals et al., 2016). Another time-related bias can stem from the failure of accounting for the lengthy period of time that usually pass during carcinogenesis and the latency of potential drug effects (Suissa & Azoulay, 2012). One approach to address this potential bias in the study design by implementing a period of lag-time within which exposure should be disregarded (Pottegård et al., 2017). Another option is to exclude patients who experience the event of interest within the lag-period (Suissa & Azoulay, 2012).

Finally, an alternative study design that could be used to avoid immortal time bias is to use a time matched nested case-control design (Suissa, 2013). Case-control studies themselves can be subject to a form of time-related bias, i.e. time-window bias, that arises if the exposure opportunity times among cases and controls are not comparable (Suissa et al., 2011). However, this bias can be circumvented by ensuring that both cases and controls have a similar exposure opportunity time.

#### **1.1.4 Studying drug-cancer associations in pharmacoepidemiology**

Cancer is sometimes referred to as a family of diseases, i.e. not one but many diseases, displaying several different faces, all characterized by an uncontrolled and abnormal growth of cells (Mukherjee, 2011). Cancers are complex and heterogenous and their development within the human body contains multiple different stages such as initiation, promotion, and progression of cancerous growth, the invasion of cancer cells from a site of origin into adjacent tissues, and the spread of malignant cells to regional lymph nodes and beyond; forming secondary tumors in other organs (Hanahan & Weinberg, 2011). Thus, carcinogenesis can be a very long process; in some types of cancer it can take up to 20-30 years before they become detectable and clinical symptoms appear (Umar et al., 2012).

Several observational studies have established that exposure to certain pharmaceutical drugs has the potential to either increase or decrease the risk of a cancer related outcome (Drew et al., 2016; Dubach et al., 1991; Jensen et al., 1989). However, the long developmental period of cancer growth provides a challenge in elucidating the real effect of drug exposure on cancer development since it is highly unlikely that drug initiation would have an

immediate impact on the manifestation of cancer, encouraging the use of lag-time in data analysis (Pottegård et al., 2017).

Pharmacoepidemiological studies on drug-cancer associations are an important area of research that has the potential to have a significant public health importance, especially as time passes and more and more high-quality data on exposures and outcomes become available. It is vitally important to identify drugs, often widely used, that either exhibit a potential carcinogenic effect or are associated with a potential beneficial effect. This applies to establishing an association between drug exposure and cancer incidence, as well as on cancer mortality. Furthermore, pharmacoepidemiological studies are valuable when it comes to establishing that a drug does not exhibit a carcinogenic effect, which holds a significant value.

#### **1.1.4.1 Assessment of drug exposure**

For exposure assessment it is important to keep in mind that drug exposure does not usually occur in one single treatment episode. Rather, in many cases drug use is characterized by continuous starts and stops, often over a long period of time. This makes it important to obtain detailed individual-level drug history so that the exposure variable can be handled appropriately, i.e. in a time-dependent manner allowing researchers to account for the changes in exposure status over a long period of time (Pottegård et al., 2017). Therefore, data sources, such as the Nordic prescription registries which can contain information on the use of prescription drugs over a long time-period, provide an excellent source of exposure information for drug-related cancer studies. From these registries, one should be able to obtain information on the initiation of exposure and the duration of use. The prescribed dose for each prescription is an important piece of information to be able to estimate cumulative exposure, a crucial variable for dose-response analyses. Although the prescribed dose is not always available from prescription registries, a comparable variable like the number of dispensed DDDs could be obtained which allows for the approximation of the duration of each prescription.

Drug exposure that affects cancer development, exerts its effect within a given period that can be defined as a 'risk period' for a particular drug. This could be the time from exposure until the manifestation of cancer, i.e. the induction period, or the time from manifestation of cancer until diagnosis, i.e. the latency period. In observational studies, it is customary to refer to the time that passes from the initiation of drug exposure until the ascertainment of outcome as latency, since actual induction and latency periods cannot be accurately identified (Pinheiro et al., 2016). A drug that contributes to

initiation of the first steps of cancerous growth, i.e. an initiator, would elicit its effect prior to carcinogenesis. Considering the long period of cancer growth mentioned previously, assessing the exposure to such a drug would require a significant period of recorded drug use. On the other hand, assessing the effect of a promoter, i.e. a drug whose cancer related function it is to promote the growth of a tumor that is already established within a given tissue, might require a shorter period of recorded drug use. Although it might also require a long period of recorded drug use depending on the nature of its effect. However, whether a drug is categorized as an initiator or a promoter; whether its latency period is thought to be long or short; whether its effect is chemopreventive or carcinogenic, it is extremely difficult, and almost impossible, to accurately determine the exact moment when a drug elicits its effect in this context. Which in turn makes it challenging to define a relevant exposure window for suspected drug related associations with cancer (Pottegård et al., 2017).

Due to the long period of cancer growth and latency of drug effects, the use of lag-time in observational studies on drug-cancer associations has been recommended, as mentioned above. This is because drug exposure that is initiated shortly before a cancer diagnosis, should not realistically be expected to have had a carcinogenic or chemopreventive effect. Additionally, it should be considered that, although a patient may have discontinued drug treatment, there might be a period of time after that discontinuation might be influenced by the drug exposure. Furthermore, the implementation of a lag-time might counteract the potential effect of reverse causation (Pottegård et al., 2017).

#### **1.1.4.2 Ascertainment of cancer outcome**

Individual-level information on cancer incidence for ascertainment of outcome is a requirement for studies on potential drug-cancer associations. Population-based nationwide registries, e.g. the Nordic cancer registries (Pukkala et al., 2018), are generally the preferred choice of data source and is considered the gold standard for obtaining the necessary information on each cancer diagnosis (Pinheiro et al., 2016).

Bearing in mind the heterogenic nature of cancer (Hanahan & Weinberg, 2011), even when tumors residing within the same organ are compared, the clustering of all cancers into one group in an analysis on drug-cancer associations should be avoided. Any associations observed for 'cancer overall' is likely driven by an effect on higher incidence cancers (Pottegård et al., 2017). Therefore, separating different cancer types by International

Classification of Diseases (ICD) codes, or tumor sites, is recommended. It might even be argued that a further separation might be desirable in some organs, e.g. by histological subtype, although that might not always be feasible (Pottegård et al., 2017).

## 1.2 Proton pump inhibitors

Proton pump inhibitors (PPIs) are a class of drugs that inhibit acid secretion by forming a covalent bond with their target; cysteine residues on the gastric hydrogen potassium ATPase ( $H^+K^+$ -ATPase). The  $H^+K^+$ -ATPase, also known as the gastric acid pump, is a proton pump that can be found in the canalicular membrane of parietal cells of the stomach (Sachs et al., 1995). Omeprazole, the first PPI substance, became commercially available in 1989 and since then several other PPI substances have been introduced (Strand et al., 2017). Other PPI substances with a marketing licence in Iceland include pantoprazole, lansoprazole, rabeprazole, and esomeprazole.

PPIs are effective in the treatment of gastroesophageal reflux disease (GERD), peptic ulcer disease (PUD) (Dent et al., 2005; Lundell et al., 2009; Mahon et al., 2005; Mehta et al., 2006) and have also been shown to be useful if included in the treatment of *Helicobacter pylori* infection, which is a risk factor for ulcer bleeding (Yuan et al., 2013). Furthermore, PPIs are considered effective in treating non-steroidal anti-inflammatory drug (NSAID) associated ulcers and can be used for prophylaxis treatment among patients taking NSAIDs and low-dose aspirin (Scheiman, 2013). Clinical guidelines generally recommend treatment durations of 4 to 12 weeks, although the duration of therapy can depend on the severity of symptoms (NICE, 2018; Scarpignato & Blandizzi, 2016).

PPIs are prodrugs, i.e. they remain inactive after intake until they are absorbed in the small intestines and carried to acidic environments, e.g. the acidic secretory canaliculi of a parietal cell, where they undergo two protonations that render them active and able to react with the active form of the  $H^+K^+$ -ATPase, situated in the canalicular membrane of a parietal cell (Sachs et al., 2006). The serum half-life of PPIs is one to two hours, which is relatively short. Therefore, to maximize their effectiveness, clinical instructions recommend pre-prandial intake since canaliculi expression of  $H^+K^+$ -ATPases is activated in response to a meal; a prerequisite for the covalent binding of PPIs and subsequent inhibition of the proton pump activity (Strand et al., 2017).

Their potency and clinical effectiveness as acid inhibitors, accompanied by their excellent safety profile, has led PPIs to becoming the mainstay in treatment of upper gastrointestinal related disorders (Strand et al., 2017). PPIs are therefore widely prescribed, and their use has been increasing over time in a number of populations and patient groups (Książczyńska et al., 2015; Moriarty et al., 2016; Pottegård et al., 2016a; Wallerstedt et al., 2017). Although they are considered to be safe, the increased use and elevated popularity has given rise to concerns related to reported adverse effects that might be associated with their use, i.e. increased risk of bone fractures (Zhou et al., 2016), kidney disease (Lazarus et al., 2016), microscopic colitis (Law et al., 2017), hypomagnesemia (Cheungpasitporn et al., 2015), *Clostridium difficile* infection (Naito et al., 2018), and chronic liver disease (Llorente et al., 2017). Additionally, stepping down from PPI have been shown to be a cause of rebound acid hypersecretion following discontinued use (Lødrup et al., 2013; Waldum et al., 1996). Additionally, inappropriate PPI use and overprescribing, driven by vague indications and unsubstantiated long-term use have been reported as a cause of concern in relation to the use of PPIs (Batuwitage et al., 2007; Grant et al., 2006; Ladd et al., 2014; Naunton et al., 2000), especially among the elderly (Cahir et al., 2010; Moriarty et al., 2016). Furthermore, there have also been reports on the potential drug-drug interactions associated with metabolic inhibition of drugs that are coadministered together with some PPI substances (Hagymási et al., 2011).

There is an ongoing debate over the long-term use of PPIs and whether they might possess a carcinogenic potential and therefore increase cancer risk, especially in digestive organs. However, the question whether PPIs are associated with increased risk of gastrointestinal related cancer is complex due to the nature of symptoms and underlying indications for PPI use. Observational studies examining this issue have not provided a definitive conclusion and many of the studies reporting an association with increased risk might be prone to reverse causation (Kearns et al., 2017), confounding by indication (Poulsen et al., 2009; Rodriguez et al., 2006; Tamim et al., 2008), or time-related biases (Cheung et al., 2017; Suissa & Suissa, 2018).

Although H<sup>+</sup>-K<sup>+</sup>-ATPases are their main original target, PPIs are potent inhibitors of acid secretion in general and have been reported to also inhibit the activity of another type of proton pump that controls the intracellular and extracellular pH of cells and cellular compartments; the vacuolar H<sup>+</sup>-ATPase (V-ATPase) (Ikemura et al., 2017; Moriyama et al., 1993; Sabolic et al., 1994).

## 1.3 Vacuolar H<sup>+</sup>-ATPase

### 1.3.1 Structure and function

The V-ATPase is a complex multisubunit ATP-dependent proton pump that operates through a rotary mechanism and is involved in the regulation of intracellular and extracellular pH (Forgac, 2007). It is a highly conserved membrane-bound enzyme in eukaryotic cells made up of several subunits that are arranged into two domains; the peripheral V1 domain, and the integral V0 membrane domain. The V1 domain is responsible for ATP hydrolysis while the V0 domain has a role to play in the translocation of protons across the membrane (Nishi & Forgac, 2002). Eight subunits, A-H, make up the V1 domain while the V0 domain is composed of the a, d, e, c, c'/Ac45 and c'' subunits (Wilkens et al., 2005). It has been shown that the two domains do reversibly disassociate, resulting in the inhibition of the ATPase activity of the V1 domain, which is an important regulatory mechanism of the activity of the V-ATPase (Cotter et al., 2015b).

The V-ATPase is found in a variety of cellular membranes and of importance to the diversity of their biological functions is the ATP-dependent proton transport from the cytoplasm and across cellular membranes; either into intracellular compartments or the extracellular space (Nishi & Forgac, 2002). Within intracellular membranes, V-ATPases function in various cellular processes such as receptor endocytosis and vesicular trafficking of lysosomal enzymes (Pamarthy et al., 2018). Additionally, the V-ATPase serves a critical role in regulating pH within digestive organelles, such as lysosomes, and securing the acidic pH that is required for the activation of digestive enzymes within these organelles (Forgac, 2007). V-ATPases have also been shown to be expressed in the plasma membrane of various specialized cells where they function to acidify the external environments by facilitating the transport of protons from the cytosol and across the plasma membrane. This function of plasma membrane V-ATPases in specialized cells such as renal intercalated cells, osteoclasts, and clear cells of the epididymis, is critical for maintaining pH homeostasis via acid secretion into the renal tubule, bone resorption, and sperm maturation, respectively (Breton & Brown, 2013; Marshansky et al., 2014). Furthermore, overexpression of the V-ATPase has been observed in the plasma membrane of tumor cells and their presence there is believed to contribute to the acidification of the tumor microenvironment (TME) (Stransky et al., 2016; Webb et al., 2011).

### 1.3.2 Extracellular acidification

Growth promoting metabolic alterations are emerging as one of the hallmarks of cancer (Hanahan & Weinberg, 2011). Driven by increased expression of the glucose transporter GLUT1, which facilitates the transport of glucose through the plasma membrane, there is a marked increase in the uptake of glucose in proliferating tumor cells, compared to non-proliferating normal tissue (Lunt & Vander Heiden, 2011; Pavlova & Thompson, 2016). In the presence of oxygen (O<sub>2</sub>), normal cells convert glucose to pyruvate, via glycolysis, and the pyruvate is then processed further to produce ATP via oxidative phosphorylation in the mitochondria, while surplus pyruvate is converted to lactate in the cytoplasm (DeBerardinis & Chandel, 2016). A characteristic of cancer cells is their ability to be able to shift away from oxidative phosphorylation, even in the presence of O<sub>2</sub>, by reprogramming their metabolism of glucose and increasing the rate of glycolysis; a metabolic switch that is generally termed “aerobic glycolysis” or “the Warburg effect” (Hanahan & Weinberg, 2011; Warburg, 1956). The metabolic switch along with the increased consumption of glucose results in a higher glycolytic rate which leads to the accumulation of lactate and protons within the cytoplasm (Chen et al., 2007; Gladden, 2004). To avoid intracellular acidification, the cells seem to adapt to this accumulation by increasing the activity of membrane-bound proteins that are involved in pH regulation, such as NA<sup>+</sup>/H<sup>+</sup> exchangers (NHE), carbonic anhydrases (CAs), bicarbonate transporters (HCO<sub>3</sub>-transporters), monocarboxylate transporters (MCTs), and V-ATPases (Granja et al., 2017). This creates a reversed pH gradient by facilitating the extrusion of protons across the plasma membrane, or into internal vacuoles, thus promoting alkalization of the cytoplasm and acidification of the extracellular environment (Webb et al., 2011). The increased acidity disrupts pH homeostasis in the TME and creates an environment within the tumor tissue that is believed to enhance invasiveness, metastatic behavior, and drug resistance (Martínez-Zaguilán et al., 1996; Rofstad et al., 2006; Spugnini et al., 2015; Wachsberger et al., 1997). Furthermore, it has recently been reported that exposure to low pH in the extracellular environment, in prostate cancer cells and a model of lung metastasis, resulted in prolonged mobility of cancer cells which eventually leads to establishment of distant metastases (Riemann et al., 2016).

A potential role of V-ATPases in regulating pH in various human cancer cells was initially suggested in 1993 (Martinez-Zaguilan et al., 1993). Invasive breast cancer cells were later reported to exhibit enhanced V-ATPase activity at their plasma membrane, compared to poorly metastatic BC cells



(Sennoune et al., 2004). As of today, several studies have reported that plasmalemmal V-ATPase activity is elevated in a number of cancer types, including breast cancer (Capecci & Forgac, 2013; Cotter et al., 2015a; Hinton et al., 2009), prostate cancer (Michel et al., 2013; Riemann et al., 2016), and melanoma (Nishisho et al., 2011).

### 1.3.3 Inhibition of V-ATPase function in cancer

Given the mounting evidence suggesting that enhanced V-ATPase activity correlates with cancer cell invasion and migration, metastasis, and drug resistance, all major attributes of a malignant phenotype, inhibitors of V-ATPase activity have been increasingly studied as potential therapeutic candidates that could hinder progression to malignancy and multidrug resistance. A number of studies, both *in vitro* and *in vivo*, have shown that V-ATPase inhibition reduces invasion and migration. Knockdown of subunit C of the V-ATPase V1 domain in a mouse xenograft model of breast cancer was shown to inhibit tumor growth and metastatic tendencies (Feng et al., 2013). Also, knockdown of subunit a3 inhibited metastasis in a mouse model of melanoma (Nishisho et al., 2011). Treatment with V-ATPase specific inhibitors, such as archazolid and bafilomycin, have also been shown to reduce BC tumor growth *in vivo* (Schneider et al., 2015) and significantly inhibit the invasive behaviour of highly metastatic BC cells (Sennoune et al., 2004). Furthermore, exposure to bafilomycin A and concanamycin A, another V-ATPase specific inhibitor, has been shown to significantly reduce invasion of prostate cancer cells *in vitro* (Michel et al., 2013). However, since V-ATPase specific inhibition involves a high degree of toxicity for normal cells, due to the ubiquitous expression of V-ATPase, other avenues have increasingly been explored with regard to clinical applications (lessi et al., 2017). As previously mentioned, and depicted in Figure 1, PPIs have been shown to exhibit an affinity for V-ATPases (Moriyama et al., 1993; Sabolic et al., 1994) and one of the first studies to demonstrate their potential efficacy in anticancer therapy reported that pre-treatment with PPIs *in vitro* enhanced the effect of chemotherapeutic agents in cancer cells derived from human melanoma, adenocarcinoma, and lymphoma (Luciani et al., 2004). The same study also reported that oral pre-treatment with omeprazole enhanced cisplatin sensitivity *in vivo*, using a mouse xenograft melanoma model (Luciani et al., 2004) and pre-treatment with lansoprazole was recently reported to increase the efficacy of paclitaxel, a chemotherapeutic agent, in the treatment of human melanoma (Azzarito et al., 2015). Additionally, PPIs have been reported to inhibit proliferation and inducing tumor cell death of melanoma cells *in vitro*, while also reducing tumor growth in mice engrafted with human melanoma cells (De Milito et al., 2010).

The therapeutic benefit of PPIs in tandem with chemotherapy have also been evaluated in studies among human cancer patients. One of those studies reported that pre-treatment with PPIs, among patients with osteosarcomas, increased the effectiveness of neoadjuvant chemotherapy (Ferrari et al., 2013). Furthermore, a phase II clinical study performed among patients with metastatic breast cancer, showed that intermittent treatment

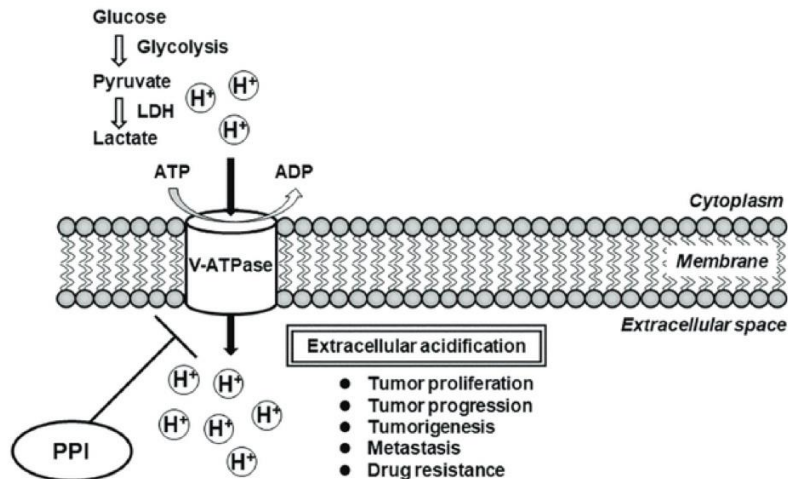


Figure 1. A schematic overview of the potential therapeutic benefits of PPI inhibition of V-ATPase, which has been reported to promote alkalization of the tumor microenvironment. PPI inhibition is thought to prevent tumor progression and drug resistance, which is otherwise induced by extracellular acidification. Figure adjusted from Ikemura et al., 2017.

with high-dose PPIs increased the efficacy of chemotherapy in breast cancer patients with a metastatic disease (Wang et al., 2015). However, a recent Danish observational study reported that PPI use was associated with increased mortality among patients diagnosed with any cancer, as well as certain site specific cancer types such as breast cancer and prostate cancer (Tvingsholm et al., 2018).

## 1.4 Epidemiology of breast cancer, prostate cancer, and malignant melanoma

### 1.4.1 Breast cancer

On a global scale, it is estimated that in 2018 there will be approximately 2.1 million diagnosed cases of incident female breast cancer, making it the most

commonly diagnosed cancer among women (Bray et al., 2018). There is however a considerable diversity in incidence rates when the underlying numbers are examined on a regional level, with the highest rates being observed in Australia/New Zealand, North America, and Europe (excluding Eastern Europe) and the lowest in South Central Asia, Middle Africa, and Eastern Africa (Bray et al., 2018). Generally, although breast cancer incidence rates in many low- and middle-income countries (LMICs) have been steadily increasing, high income countries (HICs) tend to have the highest incidence (Torre et al., 2016). In the case of female breast cancer, this in part reflects a varying degree of access to early detection programs, i.e. screening as well as differences in the prevalence of established risk factors; reproductive and hormonal factors that have been shown to increase breast cancer risk, such as long menstrual history, nulliparity, late age at first birth, recent use of oral contraceptives, and hormone replacement therapy, while breastfeeding has been shown to be a protective factor (Jemal et al., 2011; Torre et al., 2016). Other risk factors include age, family history of the disease and inherited mutations (e.g. *BRCA1*, *BRCA2*), obesity in postmenopausal women, alcohol use, low socio-economic status, and physical inactivity (Barnard et al., 2015; Ginsburg et al., 2017; Torre et al., 2016).

As well as being the most commonly diagnosed cancer, breast cancer is also the leading cause of cancer related death among women worldwide (Bray et al., 2018; Torre et al., 2016). Overall, there is less variation in the age-standardized rates for mortality compared to incidence but unlike that pattern the breast cancer mortality rate has actually been decreasing in many HICs while simultaneously increasing in some LMICs, likely due to restricted access to early detection and treatment accompanied by increasing prevalence of risk factors (Bray et al., 2018; Torre et al., 2016). On a specific population level, in 2016 the age-standardized incidence rate of female breast cancer in the Icelandic population was 85.5 per 100.000 persons and the mortality rate was 15.8 per 100.000 persons (Laufey Tryggvadottir et al., 2018).

Given the fact that breast cancer is a highly heterogenous disease there have been endeavors to characterize individual tumors based on tumor size, lymph node and metastasis status (TNM), and histological grade (Provenzano et al., 2018). Additionally, tumors are also clinically categorized and grouped by their varying expression of the estrogen receptor (ER), progesterone receptor (PR), and the *neu* oncogene (HER2), thus creating three distinct tumor subgroups, i.e. the ER positive group, the HER2

amplified group, and the triple negative/basal-like group, which lacks expression of ER, PR, and HER2 (Koboldt et al., 2012; Slamon et al., 1987). Furthermore, at the turn of the last century, studies reporting on the outcome of hierarchical clustering analyses of gene expression profiling within breast tumors revealed expression patterns that would come to define the intrinsic molecular subtypes, categorized as the luminal A, luminal B, HER2 overexpressing, and basal-like subtypes (Perou et al., 2000; Sørlie et al., 2001, 2003). Clinical practice guidelines recommend that factors such as histological type, grade, TNM staging, ER and PR status, HER2 gene expression, and the molecular intrinsic subtypes should be considered when estimating prognosis and for the purposes of treatment decision making (Rakha et al., 2010; Senkus et al., 2015).

#### **1.4.2 Prostate cancer**

According to global cancer statistics, it is estimated that a total of 1.3 million incident cases of prostate cancer will be diagnosed in 2018 and it is the most frequently diagnosed cancer type in 105 out of 185 countries listed in the GLOBOCAN estimates for 2018, making it the second most commonly diagnosed cancer among men worldwide after lung cancer (Bray et al., 2018). The prevalence of prostate cancer is especially high in HICs, where one in six among those that have reached the age of 79 years are expected to be diagnosed with the disease, compared to one in 47 in LMICs (Global Burden of Disease Cancer Collaboration et al., 2017).

With regard to mortality on a global scale, prostate cancer is the fifth most common cause of cancer related death and counter to the incidence pattern the mortality rate reveals itself to be higher in LMICs compared to HICs (Bray et al., 2018; Pernar et al., 2018). Within the Icelandic population, the age-standardized incidence and mortality rates were 79.5 and 15.0 per 100.000 persons in 2016, respectively (Laufey Tryggvadottir et al., 2018). Prostate cancer risk is heavily influenced by age, with rising incidence estimates generally observed with increasing age (Laufey Tryggvadottir et al., 2018).

There is a considerable degree of variation in the global pattern of prostate cancer incidence. This is in part due to varying levels of prostate-specific antigen (PSA) screening between individual countries (Pernar et al., 2018). PSA screenings were intensively used after they became commercially available in the late 20<sup>th</sup> century accompanied by a rapid increase in incidence rates and a shift in diagnostic patterns, reflected in a higher proportion of patients being diagnosed with localized disease and at an earlier age (Etzioni et al., 2008; Hassanipour-Azgomi et al., 2016;

Seamonds et al., 1986). However, a variation in incidence predating the use of PSA tests hints that the observed differences in the number of new cases between countries cannot be entirely due to PSA screening variability, emphasizing the effect of potential lifestyle-related differences and other factors that might modulate prostate cancer risk, such as age, family history of the disease and ethnicity (Brawley, 2012; Pernar et al., 2018).

Prostate cancer is a biologically heterogeneous disease and pathologically complex. The identification of specific prognostic determinants has therefore proven to be quite cumbersome. However, there are several clinical and pathological characteristics that have been investigated as potential prognostic factors. These factors, measured around the time of diagnosis, include Gleason score, TNM status, and PSA levels as measured at diagnosis (Martin et al., 2011).

### **1.4.3 Malignant melanoma**

Malignant melanoma is one of the deadliest forms of skin cancer. Worldwide, it is estimated that around 287 thousand new cases will be diagnosed and about 60 thousand melanoma related deaths will occur in 2018 (Bray et al., 2018). There is a considerable variation in both incidence and mortality rates when different countries and regions are compared. In 2012, the lowest incidence, in both men and women, of melanoma was observed to be under 0.5 per 100.000 persons in South-Eastern and South-Central Asia while the highest incidence was seen to be 40.3 per 100.000 persons and 30.5 per 100.000 persons in Australia and New Zealand, respectively. Meanwhile, in North-America and Northern- and Western-Europe incidence rates over 10 per 100.000 persons were observed (Ferlay et al., 2015). The regions that are most affected by this cancer are those that inhabit predominantly fair-skinned populations. In 2012, the age-standardized mortality rates ranged from 0.1 per 100.000 persons in South-East Asia to 4.7 per 100.000 persons in New Zealand (Schadendorf et al., 2018). In the Icelandic population, the age-standardized incidence rates in 2012 were 9.1 and 13.3 per 100.000 persons among men and women, respectively. The estimated age-standardized mortality rates in Iceland were under 3 per 100.000 persons among both sexes (Laufey Tryggvadottir et al., 2018).

Established risk factors that are known to enhance the risk of malignant melanoma include ultraviolet radiation and subsequent sunburns following sun exposure or use of indoor sunbeds (Bonniol et al., 2012; Gandini et al., 2011), a personal or family history of the disease as well as the presence of melanocytic birthmarks (Berwick et al., 2009), high socioeconomic status

(SES) (Jiang et al., 2015), and certain phenotypic characteristics such as having fair skin that has a tendency to freckle, light eye color, and light hair color (Berwick et al., 2009).



## **2 Aims**

The overarching aim of this study was to use the population-based resources available in Iceland to assess the use of proton pump inhibitors among the adult outpatient population in Iceland, explore whether PPIs possess a chemopreventive effect on malignant melanoma, breast or prostate cancer, and to assess a potential antineoplastic effect of PPI use on mortality among prostate cancer patients.

### **2.1 Study I – Proton-pump inhibitors among adults**

The aim of our first study was to provide evidence of real-world use of PPIs in Iceland between 2003 and 2015. Specifically, we set out to determine overall use of PPIs by individual PPI substance and specialty of the prescribing physician. Furthermore, our objective was to determine the annual prevalence and incidence of PPI use, treatment duration and the proportion of PPI use that could be associated with gastroprotection.

### **2.2 Study II – Proton pump inhibitor use and risk of breast cancer, prostate cancer, and malignant melanoma**

The aim of our second study was to examine a potential preventive role of PPI use by testing the association between exposure to PPIs and the risk of being diagnosed with a first-time malignant melanoma, breast or prostate cancer among the adult population in Iceland between 2005 and 2014.

### **2.3 Study III – Use of proton pump inhibitors and mortality among Icelandic patients with prostate cancer**

The aim of our third study was to explore whether PPI use decreases prostate-cancer specific and all-cause mortality among patients that were diagnosed with a first-time prostate cancer between 2007 and 2012.





## **3 Materials and methods**

The following chapters give an overview of the data sources, study designs, and methods that we used in this PhD project. A more detailed description of the materials and methods that were used in each of the separate studies can be found within the original publications.

### **3.1 Data sources**

#### **3.1.1 The Icelandic Medicines Registry**

The Icelandic Medicines Registry is an important resource for Icelandic pharmacoepidemiological studies that is maintained by the Directorate of Health ("The Directorate of Health," 2018). It was established in 2005 and contains individual-level information on all filled prescriptions to the outpatient population in Iceland from January 1, 2003 onwards with a completeness in the range between 91% to 98%. Since 2010 the Medicines Registry has also contained information on dispensed prescription drugs within Icelandic nursing homes (Furu et al., 2010).

We used data from the Medicines Registry in all three studies. For each filled PPI prescription by the outpatient population we received information on the name of the drug, ATC code, date of dispensing, number of dispensed DDDs, specialty of the prescribing physician, location of the pharmacy where a prescription for a PPI drug was filled, number of tablets dispensed and tablet strength in milligrams. For study III we also retrieved information on the number of distinct medications, down to the fourth ATC level, that were dispensed to a patient in the 12 months prior to an incident diagnosis of prostate cancer to be used as medication-based comorbidity.

#### **3.1.2 The Icelandic Cancer Registry**

The Icelandic Cancer Registry (Laufey Tryggvadottir et al., 2018) was established in 1954 and is maintained by the Icelandic Cancer Society under the authority of the Directorate of Health. The Cancer Registry is a population-based registry that contains information on every cancer diagnosis in Iceland since 1955 with 99% completeness (Sigurdardottir et al., 2012). Each diagnosis is currently coded according to the 10th revision of the ICD (ICD-10). Previous diagnoses coded based on earlier ICD revisions have been converted to ICD-10 to facilitate reporting and communication of the data (Sigurdardottir et al., 2012).

We used data from the Cancer Registry for studies II and III. For each diagnosis of interest, we obtained information on the date of diagnosis, ICD-10 code, morphological code, clinical stage, number of previous cancer diagnoses, and age at diagnosis. For study III, we additionally retrieved information on Gleason score for prostate cancer patients.

Systematic collection of information regarding TNM pathological staging was initiated at the Cancer Registry in 2011 (Sigurdardottir et al., 2012). Therefore, for the period between 2003 and 2015 there was a considerable amount of missing information on TNM staging for breast cancer and malignant melanoma. For prostate cancer, information on TNM staging has been collected for diagnoses of prostate cancer dating back to 1998 in a collaboration between the Cancer Registry and urologists operating in Iceland.

### **3.1.3 The Icelandic Cause of Death Registry**

The Icelandic Cause of Death Registry is a centralized national registry that is maintained by the Directorate of Health. It contains mortality data for the Icelandic population categorized according to the ICD-10 system (World Health Organization, 2016). This includes data on date of death and the main underlying cause of death for each deceased individual. For study III, we obtained information on the underlying causes of death, enabling us to identify prostate cancer-specific mortality within our cohort of prostate cancer patients.

### **3.1.4 Landspítali – The National University Hospital of Iceland**

Landspítali – The National University Hospital of Iceland is supervised by the Directorate of Health and is the leading hospital in Iceland, providing health care to patients from all health districts in Iceland. For study III we obtained information on chemotherapy, radiotherapy, and relevant surgical operations categorized according to the NOMESCO Classification of Surgical Procedures (NCSP). For surgical operations, data were available from 2003 onwards. For chemotherapy and radiotherapy, complete data were available from 2007 onwards. These variables were obtained from Electronic Health Records that were accessed through the Clinical Data Warehouse at the hospital.

### **3.1.5 Other data sources**

For all of our studies (I-III) the Icelandic Population Register provided us with information on year of people's birth, month of birth, sex, residency, migration

status and date of death (if appropriate). These variables were collected for every resident of Iceland during the relevant study periods, both Icelandic and foreign.

PPIs were not available in Iceland as OTC drugs prior to February 1, 2009. The Icelandic Medicines Agency (“Icelandic Medicines Agency,” 2018) provided us with data on wholesale statistics of PPI drugs which allowed us to determine the annual proportion of OTC PPI use from 2009 onwards.

Table 1 lists basic information about each study that we conducted that is included in this thesis.

## **3.2 Study design and population**

### **3.2.1 Study I – Proton-pump inhibitors among adults**

This was a nationwide population-based drug utilization study covering the entire adult population residing in Iceland from January 1, 2003 through December 31, 2015. Over the study period 313,296 individuals constituted our study population; a dynamic cohort where individuals could enter the cohort once they reached 19 years or immigrated to Iceland and left the cohort if they emigrated from Iceland. We obtained data on outpatient PPI use from the Medicines Registry while the Population Registry provided demographic information on the study population.

### **3.2.2 Study II – Proton pump inhibitor use and risk of breast cancer, prostate cancer, and malignant melanoma**

This was a population-based matched case-control study nested within the adult population of Iceland, which amounted to 220,512 individuals during the study period. Individuals were required to have resided in Iceland from January 1, 2003 to be eligible for inclusion in the study. Those with a previous history of cancer were excluded. Incident cases of breast cancer (N = 1739; ICD-10: C50), prostate cancer (N = 1897; ICD-10: C61), and malignant melanoma (N = 385; ICD-10: C43) that were diagnosed between January 1, 2005 and December 31, 2014 were identified using data from the Cancer Registry. For each case, we selected up to 10 controls from the underlying population that were matched on birth year and sex using risk-set-sampling. The Medicines Registry provided us with data on outpatient PPI use while we obtained demographic information from the Population Register.

Table 1. The studies included in this thesis, the study periods, size of study populations, and data sources that were used.

| Study | Title  | Study period | Study population (N)  | Study design                               | Data sources  |
|-------|--|--------------|---|--|---|
| I     | Proton-pump inhibitors among adults: a nationwide drug-utilization study   | 2003 - 2015  | 313,296   | Population-based cohort study              | Icelandic Medicines Registry, Icelandic Population Registry   |
| II    | Proton pump inhibitor use and risk of breast cancer, prostate cancer, and malignant melanoma: An Icelandic population-based case-control study | 2005 - 2014  | 19,129 BC cases and controls,<br>20,865 PC cases and controls,<br>42,35 MM cases and controls | Population-based nested case-control study | Icelandic Cancer Registry, Icelandic Medicines Registry, Icelandic Population Registry  |
| III   | Use of proton pump inhibitors and mortality among Icelandic patients with prostate cancer  | 2007 - 2015  | 1058  | Population-based cohort study              | Icelandic Cancer Registry, Icelandic Medicines Registry, Icelandic Population Registry, Icelandic Cause Death Register, Landspítali - the National University Hospital of Iceland |

Abbreviations: BC, breast cancer; PC, prostate cancer; MM, malignant melanoma

### **3.2.3 Study III – Use of proton pump inhibitors and mortality among Icelandic patients with prostate cancer**

Using nationwide data from Icelandic health registries we conducted a population-based cohort study. We identified eligible patients that received a first-time diagnosis of prostate cancer (N = 1058; ICD-10: C61) between January 1, 2007 and December 31, 2012 using data from the Cancer Registry. To be eligible for inclusion patients had to be 18 years or older, were required to have resided in Iceland at the start of follow-up and had to have survived the first 12 months following their diagnosis. Outpatient use of PPIs was assessed from January 1, 2003 through December 31, 2015 using data from the Medicines Registry. Data from the cancer registry and the medicines registry were then linked together with the Population Register, the Causes of Death Registry, and data from Landspítali – the National University Hospital of Iceland using unique personal identification numbers. The primary outcomes in study III were prostate cancer-specific mortality (ICD-10: C61) and all-cause mortality. Patients were followed from 12 months after their diagnosis until their date of death, date of emigration or the end of the study period.

### **3.3 Assessment of exposure and ascertainment of outcome**

We obtained data on PPI drug exposure from the Medicines Registry. Every drug we assessed was defined according to the World Health Organization anatomical therapeutic chemical/defined daily doses (ATC/DDD) classification (WHO, 2018). In all studies (I-III), we considered PPI use as the primary exposure and we defined PPIs as those drugs belonging to the ATC-group A02BC (*Proton pump inhibitors*). The four PPI substances that were prescribed to the outpatient population in Iceland during our study periods are listed in Table 2, as well as our classification of higher and lower doses based on tablet strength in milligrams (mg).

#### **3.3.1 Study I – Proton-pump inhibitors among adults**

When estimating prevalence and incidence (per 100 persons) in study I, we defined PPI use as at least one PPI dispensing within the relevant calendar year. We used the National Institute for Health and Care Excellence (NICE) clinical guidelines (NICE, 2018) to define PPI dose strengths (in mg) as either higher or lower dose by defining standard and double doses as higher-dose PPIs and low doses as lower-dose PPIs (Table 2).

Table 2. Proton pump inhibitor substances that were prescribed to the outpatient population in Iceland between 2003 and 2015.

| Drug         | ATC     | DDD (mg) | Available                        | Lower dose (mg) | Higher dose (mg) |
|--------------|---------|----------|----------------------------------|-----------------|------------------|
|              |         |          | package sizes<br>in Iceland (mg) |                 |                  |
| Omeprazole   | A02BC01 | 20       | 10, 20, 40                       | 10              | 20, 40           |
| Lansoprazole | A02BC03 | 30       | 15, 30                           | 15              | 30               |
| Rabeprazole  | A02BC04 | 20       | 10, 20                           | 10              | 20               |
| Esomeprazole | A02BC05 | 30       | 10, 20, 40                       | 10              | 20, 40           |

### 3.3.2 Study II – Proton pump inhibitor use and risk of breast cancer, prostate cancer, and malignant melanoma

In the main analysis of study II, we defined patients as PPI users if they received at least one PPI dispensing prior to index date. We implemented a lag-time where we disregarded filled prescriptions within the 24 months prior to index date. Furthermore, we performed secondary analyses stratified by high use, cumulative dose, and cumulative duration. Additionally, we adjusted observed risk estimates for NSAID use, defined as at least two filled prescriptions prior to index date. We required NSAID users to have received at least two filled prescriptions rather than one since our objective was to approximate longer-term use which has been linked with concurrent use of PPIs. The primary outcome of study II was a registered diagnosis in the Cancer Registry of breast cancer (ICD-10: C50), prostate cancer (ICD-10: C61) or malignant melanoma (ICD-10: C43).

### 3.3.3 Study III – Use of proton pump inhibitors and mortality among Icelandic patients with prostate cancer

In study III, we focused on the use of PPIs after prostate cancer diagnosis. Individuals that received at least two filled PPI prescription following their diagnosis were defined as post-diagnosis users. PPI exposure was modelled in a time-dependent manner where patients were considered to be unexposed until they had received a second PPI prescription, after which they were considered to be exposed for the remainder of follow-up. Exposed person-time was lagged for 12 months after the exposure criteria was met. We then assessed post-diagnosis PPI use in varying ways in several secondary and sensitivity analyses as described in paper 3. The primary outcome of study III was prostate cancer-specific death (ICD-10: C61) but we also assessed death from all causes as a secondary outcome.

### **3.4 Data analysis**

In studies II and III, we analyzed each cancer type separately. R version 3.4.2 (“R: The R Project for Statistical Computing,” 2018) and RStudio (“RStudio – Open source and enterprise-ready professional software for R,” 2018) were used in all analyses for studies I-III.

#### **3.4.1 Study I – Proton-pump inhibitors among adults**

We measured overall use of PPIs as the total amount of dispensed DDDs during the study period stratified by calendar year, PPI substance, and specialty of the prescribing physician. We defined annual prevalence of PPI use as the number of adult individuals who filled at least one PPI prescription within each calendar year per 100 persons in the adult population. Further, we performed a sensitivity analysis after redefining prevalence use as the total number of adults filling at least two PPI prescriptions within a relevant calendar year. We defined annual incidence of PPI use as the number of adults who were dispensed their first PPI drug, after a 24-month period of no PPI dispensing, per 100 persons in the adult population.

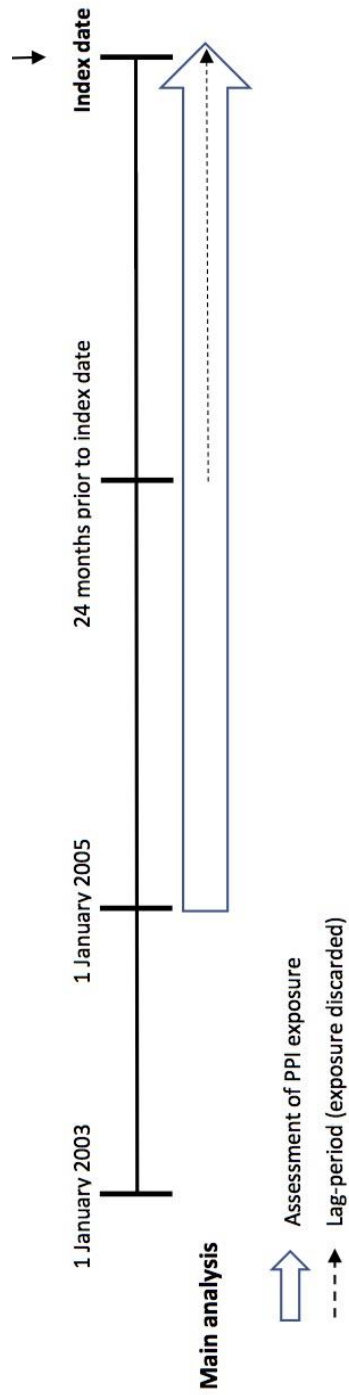
We estimated the duration of each PPI prescription assuming a daily intake of one tablet and added a grace period of 108 days to account for irregular use. We then used the ‘Proportion of Patients Covered’ (PPC) method (Rasmussen et al., 2018) to estimate the duration of PPI treatment for each incident PPI user over a 5-year period from their first PPI prescription. The duration analysis was stratified by age, dose strength, and sex. Additionally, we examined the distribution of dispensed DDDs and tablets over a 5-year period following the start of an initial treatment episode. Lastly, we explored the proportion of PPI use that might be attributable to gastroprotection by measuring concurrent use of drugs that have been shown to increase the risk of gastrointestinal complications.

#### **3.4.2 Study II – Proton pump inhibitor use and risk of breast cancer, prostate cancer, and malignant melanoma**

To estimate the association between PPI use and an incident diagnosis of breast cancer, prostate cancer or malignant melanoma, we compared the risk among cases and controls, matched on birth-year and sex, using conditional logistic regression to calculate the relevant odds ratios (ORs) and 95% confidence intervals (CIs). Figure 2 gives an overview of the main analysis. In subgroup analyses, we stratified the data based on high-use, cumulative dose, cumulative duration of PPI use, and calendar period. All analyses were adjusted for NSAID use prior to index date.



Figure 2. An overview of the main analysis of study II.

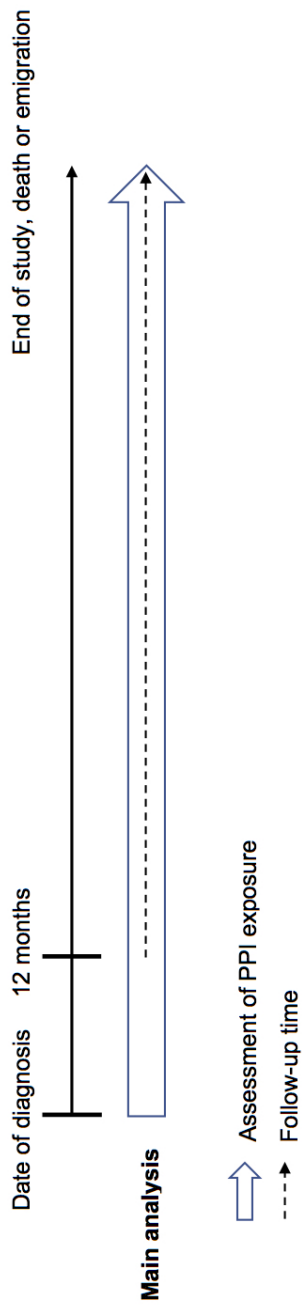


Additionally, we performed several sensitivity analyses to assess the effect of different definitions of a lag-period prior to index date (24 months in the main analysis). Furthermore, we also repeated the main analyses employing a new-user design rather than a prevalent user design. Finally, we performed a post-hoc supplementary analysis considering clinical stage at diagnosis among prostate cancer patients.

### **3.4.3 Study III – Use of proton pump inhibitors and mortality among Icelandic patients with prostate cancer**

We used Cox proportional hazard models to estimate hazard ratios (HRs) with 95% CIs of prostate cancer-specific mortality and all-cause mortality associated with PPI use. All models included adjustment for age at diagnosis, calendar year, clinical stage, Gleason score, radiotherapy, prostate cancer surgery, cancer drug treatment, and medication-based comorbidity. In the main analysis, PPI exposure was considered as a time-dependent covariate. Additionally, we performed secondary analyses by timing of use, clinical stage, and cumulative dose. We performed three sensitivity analyses to assess our original definition of PPI exposure. A graphical overview of the main analysis is given in Figure 3.

Figure 3. An overview of the main analysis of study III.



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### **3.5 Ethical considerations**

We received authorization for all three studies from the National Bioethics Committee and the Data Protection Authority in Iceland. In all instances, personal identification numbers were encrypted by the data manager at the Directorate of Health and we, the researchers, did not at any stage have access to identifiable personal information. Following are the licenses we were granted by the relevant Icelandic authorities.

We obtained ethical approvals from the National Bioethics Committee and the Data Protection Authority October 27, 2015 for studies I and II (reference number VSNb2015080004/03.03). Reprint of the original documents can be found in Appendix A.

For study III we obtained ethical approvals from the National Bioethics Committee and the Data Protection Authority on September 6, 2016 (reference number VSNb2016080001/03.01). Reprint of the original document can be found in Appendix B.

All of our three studies were observational and based on nationwide registry data. Thus, they did not require us to obtain informed consent from the study population.



## **4 Results**

### **4.1 Study I – Proton-pump inhibitors among adults**

In this study, we observed an increase in total PPI use from 3.5 million dispensed DDDs in 2003 to 10.7 million dispensed DDDs in 2015. We found that the majority of all DDDs that were dispensed in this period were prescribed by primary care physicians. Esomeprazole was the most commonly prescribed PPI substance early on in the study period but after 2009 omeprazole became the most commonly prescribed substance. There was an overall increase in annual prevalence of PPI use over time. In 2003 we observed a prevalence of 8.5 per 100 persons while by 2015 it had increased to 15.5 per 100 persons. We did not observe a similar increase when estimating the annual incidence, which we found to be 3.3 per 100 persons in 2005 and 4.1 per 100 persons in 2015. We observed that the prevalence increased with age and was higher among females than among males. After redefining prevalent use for the purposes of a sensitivity analysis we observed that the prevalence estimates decreased somewhat compared with the estimates from the main analysis, i.e. rising from 5.4 per 100 persons in 2003 to 11.0 per 100 persons in 2014 (Figure 4).

Using the PPC method to estimate the duration of PPI treatment among incident PPI users by age and initial dose strength, we found that the duration of treatment tended to be longer among older patients and among patients that started their initial PPI treatment on higher doses. When looking at the proportion of PPI use that might have been attributed to gastroprotection, we observed that the proportion of patients concurrently using PPIs and NSAIDs decreased over time while the opposite was true of concurrent use of PPIs and oral anticoagulants, PPIs and acetylsalicylic acid, and PPIs and platelet inhibitors.

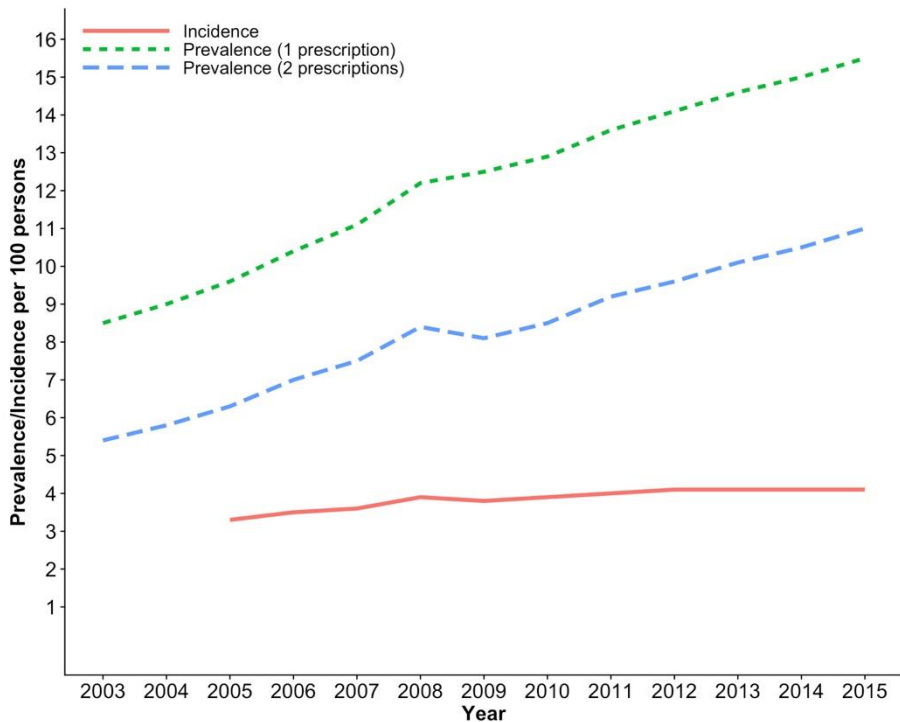


Figure 4. Annual incidence and prevalence (per 100 persons) of proton pump inhibitor use among adults in Iceland. Displaying prevalence estimates from both main and sensitivity analyses.

## 4.2 Study II – Proton pump inhibitor use and risk of breast cancer, prostate cancer, and malignant melanoma

### 4.2.1 Ever use and high use of PPIs

In our analyses of ever use of PPIs, comparing it to non-use, we observed adjusted ORs of 1.03 (95% CI: 0.92-1.16), 1.12 (95% CI: 1.00-1.25), and 0.84 (95% CI: 0.69-1.12) for breast cancer, prostate cancer, and malignant melanoma, respectively. For high use of PPIs, we observed an adjusted OR of 0.97 (95% CI: 0.78-1.19) for breast cancer, 1.20 (95% CI: 0.99-1.47) for prostate cancer, and 0.59 (95% CI: 0.40-1.13) for malignant melanoma.

### 4.2.2 Cumulative use of PPIs

We further explored cumulative use of PPIs by stratifying by cumulative dose in DDDs and cumulative duration in years for all three cancer types. We observed that individual ORs were elevated, with a marginally statistically

significant association for prostate cancer, indicating a potential increase in risk for patients that used over 1096 DDDs (1.26 (95% CI: 1.02-1.55)) and those that were exposed for 1-5 years (1.22 (95% CI: 1.04-1.42)). Additionally, for malignant melanoma we observed ORs that indicated a marginally significant association with decreased risk among patients consuming between 365 and 730 DDDs (0.37 (95% CI: 0.15-0.93)) and those using PPIs for a duration of 1-5 years (0.57 (95% CI: 0.36-0.91)). However, the patterns we observed did not indicate a dose-response relationship for any of the cancer types (see Figure 5 and Figure 6).

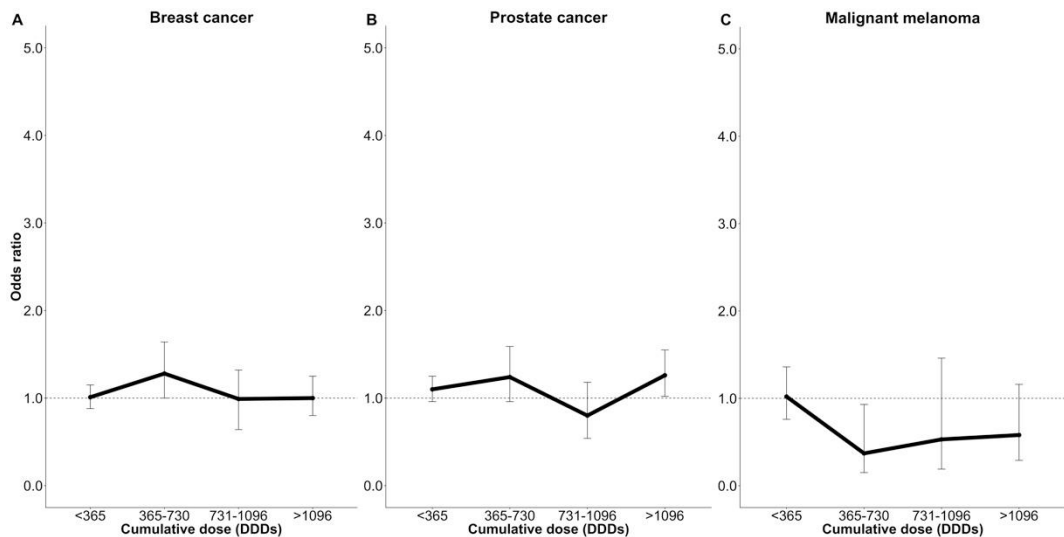


Figure 5. The observed pattern between cumulative PPI dose and risk of breast cancer (A), prostate cancer (B), and malignant melanoma (C).



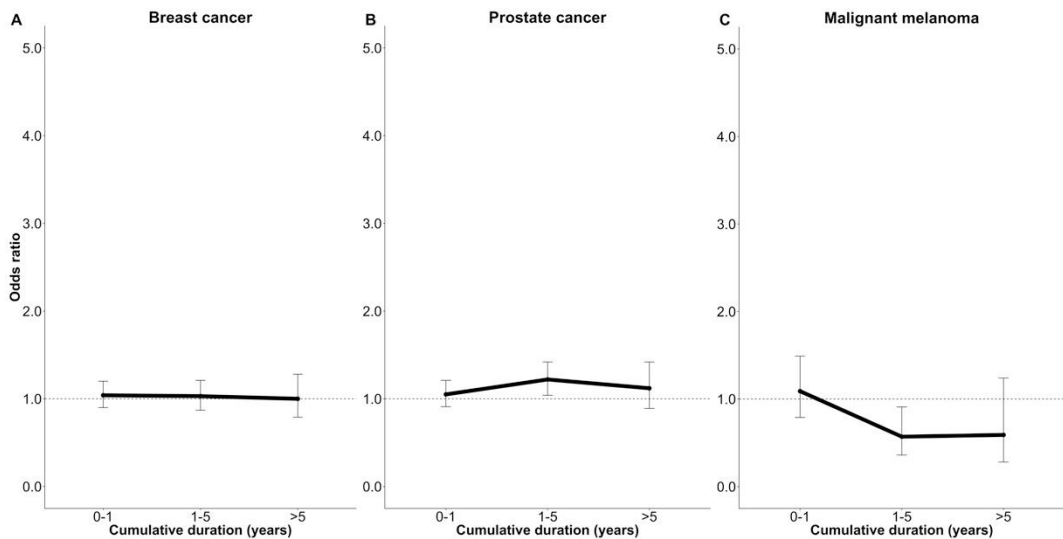


Figure 6. The observed pattern between cumulative duration of PPI use and risk of breast cancer (A), prostate cancer (B), and malignant melanoma (C).

### 4.3 Study III – Use of proton pump inhibitors and mortality among Icelandic patients with prostate cancer

In the main analysis of study III, we observed adjusted HRs of 0.88 (95% CI: 0.52-1.48) and 1.02 (95% CI: 0.73–1.43) for prostate cancer-specific and all-cause mortality, respectively. Thus, our results did not indicate that post-diagnosis PPI use was associated with increased, or decreased, prostate cancer-specific or all-cause mortality. When we stratified our analysis by timing of PPI use we observed adjusted HRs of 0.45 (95% CI: 0.21-0.98) for prostate-cancer specific mortality and 0.67 (95% CI: 0.43-1.04) for all-cause mortality among continuous PPI users, while we observed adjusted HRs of 1.12 (95% CI: 0.61-2.08) for prostate-cancer specific mortality and 1.25 (95% CI: 0.82-1.92) for all-cause mortality among new PPI users. Stratification by clinical stage yielded an adjusted HR of 0.50 (95% CI: 0.22-1.16) among patients with localized disease and 1.00 (95% CI: 0.44-2.27) among patients with non-localized disease. For cumulative PPI use, analyses of prostate cancer-specific mortality yielded adjusted HRs of 0.91 (95% CI: 0.43-1.90) for cumulative use of 1-365 DDDs and 0.86 (95% CI: 0.45-1.61) for >365 DDDs. For all-cause mortality we observed adjusted HRs of 1.19 (95% CI: 0.76-1.87) and 0.91 (95% CI: 0.61-1.37) for cumulative use of 1-365 DDDs and >365 DDDs, respectively.

Figure 7 gives a visual overview of the results from study III for analyses of prostate cancer-specific mortality.

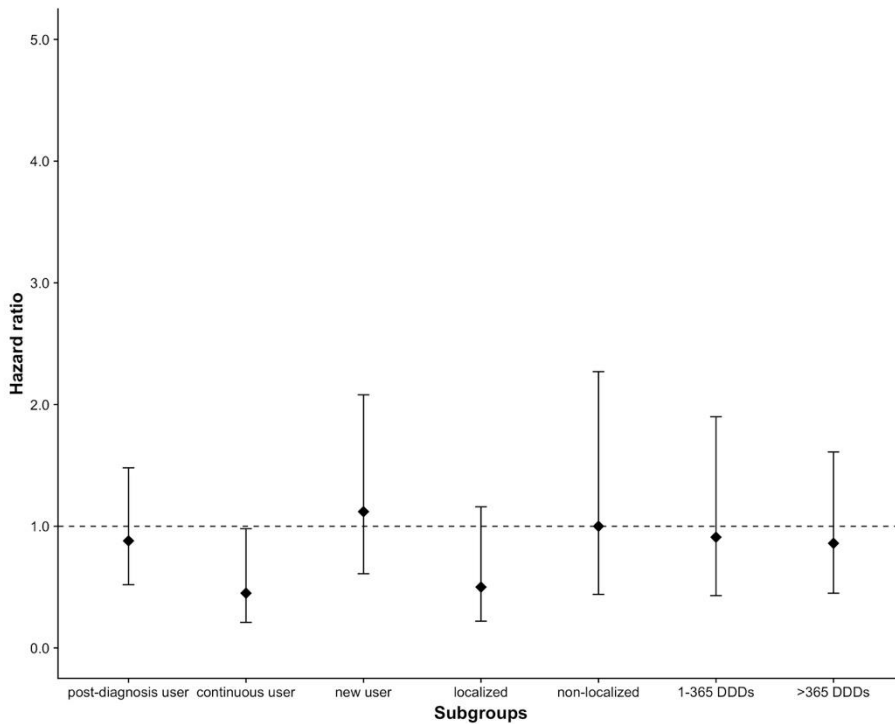


Figure 7. Adjusted hazard ratios and 95% confidence intervals for prostate cancer-specific mortality among PPI user subgroups from the main analysis (post-diagnosis users) and secondary analyses (stratified by timing of use, clinical stage, and cumulative use).



## **5 Discussion**

### **5.1 Main findings**

Overall, we found that PPI use is widespread within the Icelandic population and that it has increased considerably over the past decade. Additionally, patients seem to be treated for longer durations than is generally recommended by clinical guidelines. The findings of our studies do not support a chemopreventive role of PPI use when it comes to the risk of being diagnosed with an incident breast cancer, prostate cancer, or malignant melanoma. Furthermore, our results do not indicate that post-diagnosis PPI use decreases mortality among prostate cancer patients.

### **5.2 General discussion**

#### **5.2.1 Study I – Proton-pump inhibitors among adults**

We made an effort to map the landscape of PPI use within the adult Icelandic population by conducting a nationwide population-based drug utilization study. In line with findings from comparable populations (Haastrup et al., 2014; Pottegård et al., 2016a; Wallerstedt et al., 2017), our results indicate that overall use of PPIs increased considerably during the study period. In 2015, the total use of PPIs within the population had reached 10.7 million dispensed DDDs; an increase of 7.2 million DDDs when compared to the 3.5 million DDDs dispensed in 2003. We found that the rising use over time was driven in large part by a surge in prescriptions from primary care physicians, which accounted for 60% of the overall increase in sold DDDs. There are several possible explanations for this considerable rise in PPI use. First, it might be due to changes in the incidence of underlying clinical indications, e.g. increasing incidence of GERD. Unfortunately, the Medicines Registry does not contain information on the underlying indication for each prescription, which made this difficult to assess. Nonetheless, should this be the case we might expect that to be reflected to some degree in either the prescribing pattern among gastroenterologists or in our estimates of annual incidence, or in both. However, prescriptions issued by gastroenterologists only accounted for 6% of the overall increase in PPI use. Furthermore, we only observed a modest rise in annual incidence during the study period while we found that there was a marked increase in annual prevalence which seems to suggest that the elevated PPI use was driven by rising use among

current users, rather than a surge in the number of new users. Second, it could be that PPIs were increasingly used for prophylactic purposes. However, we found that there were actually fewer PPI users concurrently using ulcerogenic drugs in 2015 (36.2%) than in 2003 (37.6%).

As is discussed in more detail in paper I, our analyses of duration of PPI treatment revealed that 22% of patients were still using PPIs one year after starting their initial treatment and that a higher proportion of older patients stayed on treatment for longer durations compared to younger patients. In general, the observed treatment duration among a considerable proportion of patients were longer than is generally recommended by clinical guidelines (NICE, 2018). However, we were unable to determine whether these prolonged durations of PPI treatment reflected more severe symptoms of appropriate underlying indications. The popularity of PPIs has led to some speculations that their general tolerability and good safety profile might be contributing factors to their potential overuse in some quarters, due to patients receiving prescriptions for PPIs without a clear diagnosis (Heidelbaugh et al., 2012). For example, one US study found that among patients receiving antisecretory treatment for more than 90 days, around 39% did not have a documented upper GI diagnosis (PUD, GERD, dyspepsia, or a combination of the three) (Jacobson et al., 2003). Additionally, the appropriateness of long-term PPI use has been questioned in some cases amid concerns that patients might be receiving repeat PPI prescriptions with automatic renewals, without their symptoms being reevaluated, which is likely to encourage unsubstantiated long-term use (Batuwitage et al., 2007). It has been reported in other studies that many long-term PPI users do not meet with their general practitioner (GP) regularly to discuss their treatment (Krol, Muris, Schattenberg, Grol, & Wensing, 2004). Furthermore, when they do meet, the expected duration of PPI treatment is not necessarily discussed (Haastrup et al., 2014; Krol et al., 2004). As with any drug treatment, it is important that PPI therapy is based on reliable information and appropriate indications.

### **5.2.2 Study II – Proton pump inhibitor use and risk of breast cancer, prostate cancer, and malignant melanoma**

The continuous rise of PPIs since they originally became available approximately 30 years ago has stimulated the conversation around the potential links between long-term PPI use and certain adverse effects, e.g. an increased risk of kidney disease (Lazarus et al., 2016), bone fractures (Zhou et al., 2016), hypomagnesemia (Cheungpasitporn et al., 2015), Clostridium

difficile infection (Naito et al., 2018), microscopic colitis (Law et al., 2017), chronic liver disease (Llorente et al., 2017), as well as changes in the composition of the intestinal microbiota (Marlicz et al., 2014). Furthermore, there have also been some reports that have focused on PPI use in association with cancer outcomes. Specifically, they have focused on cancers related to the digestive tract with conflicting results and some of them are likely to be influenced by confounding by indication (Poulsen et al., 2009; Rodriguez et al., 2006; Tamim et al., 2008), reverse causation (Kearns et al., 2017), and time-related biases (Cheung et al., 2017; Suissa & Suissa, 2018).

The emphasis in most observational studies to date has been on gastrointestinal-related cancers and the potential of PPIs to enhance cancer risk. The focus of our cancer-related studies however, narrowed in on the potential beneficial effects of using PPIs in relation to cancer. Therefore, the underlying hypothesis of studies II and III was that PPI use had a potential preventive role in the context of cancer risk and mortality. We decided to exclude gastrointestinal-related cancers from our studies, due to the high probability of confounding by indication and reverse causation in this context. Rather, we decided to focus our attention on three cancers that were not as likely to be subject to these biases, i.e. breast cancer, prostate cancer, and malignant melanoma; cancer types that have been studied previously both in vitro and in vivo where PPIs were shown to exhibit antineoplastic effects (De Milito et al., 2010; Katara et al., 2016; Marino et al., 2010; Michel et al., 2013; Schneider et al., 2015). Furthermore, breast cancer, in women, and prostate cancer, in men, are commonly diagnosed cancer types; a meaningful consideration given the size of the Icelandic population and the importance of elucidating exposures that could influence disease risk.

The biological plausibility of PPIs having a preventive role in a cancer setting, centers on their function as potent acid inhibitors. As has been discussed previously, although not specifically designed to do so, PPIs are able to bind to proton pumps of the V-ATPase type. The V-ATPase has been shown to play a part in promoting acidification of the tumor microenvironment by facilitating a flow of protons through the plasma membrane. Our underlying hypotheses therefore rest on the assumptions that, in the context of the human body, the PPIs are consumed and then absorbed into the circulation where they are then distributed to cancer sites where the extracellular acidity would have to be acidic enough to attract and activate the PPIs. There they would have to bind the V-ATPases and inhibit the extrusion of protons out into the extracellular environment. However, as presented in

paper II, our findings do not support the hypothesis that PPIs possess a chemopreventive effect in the context of breast cancer, prostate cancer, and malignant melanoma.

As we discussed in paper II, if PPIs actually do possess a chemopreventive effect in these cancers, our results could be explained by a number of reasons. First, it could be that several other proteins participate in regulating pH-levels within the cancer cells. Thus, V-ATPase inhibition in itself might not be enough to cut off the flow of protons into the extracellular environment. Second, the pH-level where these tumors are growing might not be low enough for the PPIs to accumulate at the target sites. PPIs are weak bases that are inactive upon consumption but become active in acidic environments. It has been postulated that PPIs selectively accumulate in the acidic space of the secretory canaliculus of parietal cells of the stomach (Shin & Kim, 2013). The reason for this selective accumulation is that weak bases like the PPIs require a pH < 4.0, which is not found in another region of the body (Shin et al., 2004). Therefore, for the PPIs to accumulate at tumor sites and be activated, the pH would have to be below 4.0. Although the TME around cancer cells has been shown to be acidic compared to the external environment of normal cells, the acidity is only thought to reach pH values around 6.0 (Gatenby & Gillies, 2004), which might not be enough to attract the PPIs to these sites.

Whether the acidity in tumors that are progressing to a metastatic state might reach lower pH values than 6.0 is unclear. Interestingly, highly metastatic cancer cells have been shown to exhibit an increased expression of V-ATPase (Nishisho et al., 2011; Sennoune et al., 2004), which might indicate increased TME acidity in advanced tumors. Although our post-hoc analysis, where we stratified prostate cancer patients by clinical stage, did not return conclusive results, it would be interesting to examine this matter systematically with an increased sample size.

### **5.2.3 Study III – Use of proton pump inhibitors and mortality among Icelandic patients with prostate cancer**

Results from a phase II clinical trial among patients with metastatic breast cancer reported that intermittent high-dose treatment with esomeprazole was associated with increased responsiveness in patients receiving chemotherapy (Wang et al., 2015). Initially, influenced in part by the findings of Wang et al., our aim was to include breast cancer patients in this study. However, the Cancer Registry unfortunately did not contain information on clinical stage among breast cancer patients before 2011. Since clinical stage

is a crucially important prognostic variable, we eventually decided to focus our attention solely on prostate cancer patients in this study. Furthermore, when we originally conceived of this study our aim was to focus on a longer study period, i.e. from 2004 through 2012. However, since the Clinical Data Warehouse at Landspítali hospital did not possess exhaustive information for all patients on chemotherapy and radiotherapy prior to 2007, we decided to adjust the study period accordingly. Furthermore, our initial efforts to conduct secondary analyses stratified by PPI substance as well as PPI pre-treatment among patients receiving chemotherapy were thwarted by the small sample size and eventual low numbers in some of the subgroups.

The association between PPI use and mortality among prostate cancer patients has not been studied extensively in other observational studies. In fact, to our knowledge, the only other study to look into this matter is the study by Tvingsholm et al. (Tvingsholm et al., 2018). Although not solely focused on prognosis among prostate cancer patients, they observed significantly increased prostate cancer-specific mortality among post-diagnosis users of PPIs compared with non-users, in their analyses of selected cancer sites. Motivated by these conflicting results, the previously reported antineoplastic activity of PPIs on prostate cancer cells, and the high incidence of prostate cancer overall, we sought to explore whether PPI use would be associated with mortality among prostate cancer patients. As in study II, our hypothesis was based on the biological plausibility that PPI use might have a beneficial effect; in this case by improving survival among exposed patients. However, our findings did not support this hypothesis.

Our observations of null associations between PPI use and prostate cancer-specific and all-cause mortality is in contrast with the findings of Tvingsholm et al., i.e. that PPI use is associated with increased mortality risk among prostate cancer patients. In their study, Tvingsholm et al. found that the increased mortality they observed seemed to be exclusively associated with new users, while the increased risk was not observed among continuous users (Tvingsholm et al., 2018). Their results seem to suggest that there is some unmeasured confounding at play, since the increased mortality is only observed among patients that start their PPI use after they are diagnosed with prostate cancer. One would think, that if PPI use increased the risk of mortality among post-diagnosis users, that this would also be observed among continuous users, who had been using PPIs for longer durations and consumed a greater cumulative quantity of the drugs. In our study, although we observed lower adjusted HRs for prostate cancer-specific mortality among continuous users of PPIs, compared with new PPI users, our findings did not



indicate that initiating PPI use after diagnosis was associated with excess mortality.

As in study II, the biological rationale for the potential antineoplastic role of PPIs in cancers, via their inhibitory function of acid secretion, depends on a number of factors that have to align for them to be able to have their proposed effect within the human body, i.e. accumulate at the cancer sites and promote the alkalization of the TME. As it stands, the evidence from these observational studies do not suggest that the PPIs are able to elicit these effects. In the context of mortality among cancer patients, the possibility of pre-treatment with PPIs being able to increase the effectiveness of chemotherapy might be best suited to be studied in the controlled surroundings of a RCT or in a well-controlled observational study that is able to account for possible confounding by other diseases likely to increase mortality.

### **5.3 Studies II and III – Potential biases**

In studies II and III, there were several biases we had to take into account. A more detailed overview of these biases, and others, is given in chapter 1.1.1, which focuses on biases in pharmacoepidemiological studies. Here, we discuss the biases we encountered in our two outcome studies.

#### **5.3.1 Immortal time bias**

A simple definition of immortal time is that it refers to a period of follow-up in a cohort during which the outcome of interest is not able to occur (Levesque et al., 2010). Immortal time bias in pharmacoepidemiological research has been shown to be increasingly common, e.g. in a paper by Suissa this bias was identified in 20 observational studies that were studying drug-related effects of commonly used prescription drugs (Suissa, 2007). If unaccounted for, this bias will invariably skew the results of studies on drug effects so that they are likely to suggest a highly protective role of the drug under study relating to a given outcome, e.g. an incident cancer diagnosis. Another manifestation of this bias, in the context of mortality, was observed in a study whose results suggested that Academy Award winners are likely to live longer than their peers that never receive the prestigious award (Redelmeier & Singh, 2001; Sylvestre et al., 2006).

Both in study II and study III, immortal time bias was an issue that we needed to deal with in our study designs. This was especially important, given the underlying hypothesis, because immortal time bias was likely to

skew the resulting estimates from our analyses downward, thereby likely creating a false sense of a protective drug effect. In fact, before we implemented the nested case-control design in study II, we set up a cohort study where the study population consisted of the entire adult population in Iceland. Using that design, we compared those that had ever used PPIs to those that had never used PPIs. Additionally, we performed a secondary analysis where we estimated the effect of cumulative duration of use in a Cox regression analysis (0-3 months, 3-6 months, 6-24 months, 24-60 months, >60 months). Our observations, heavily influenced by immortal time bias and presented in appendix C, highlight the importance of averting this bias. Otherwise, we might have falsely concluded that PPIs possess a chemopreventive effect in all three cancer types.

### **5.3.2 Time-window bias**

Although observational cohort studies have been shown to be susceptible to certain time-related biases, the same cannot be said about studies using the case-control study design. In a paper by Suissa et al (Suissa et al., 2011), the authors investigated the results of a case-control study claiming that statin use drastically reduces lung cancer risk by 45% (Khurana et al., 2007). What Suissa and colleagues found was that the results from the lung cancer study could be explained by a bias referred to as “time-window bias” (Suissa et al., 2011). This bias arises when there is an imbalance in the length of exposure opportunity time between cases and controls, because a patient with a shorter exposure opportunity time is, by definition, not as likely to be exposed to a specific drug, than a patient with a longer exposure opportunity time.

In study II, we ensured that cases and controls would have similar opportunities to become exposed by restricting the underlying study population to those individuals that had resided in the Iceland from January 1, 2003. As presented in Table 1 of paper II, this resulted in a comparable exposure opportunity time between cases and controls, allowing us to avoid time-window bias.

### **5.3.3 Reverse causation (protopathic bias)**

Reverse causation in pharmacoepidemiology refers to a situation where drug use is initiated as a response to initial symptoms caused by a disease that is still undiagnosed when drug use is started. In the context of study II for example, this bias might lead to false conclusions on the association between PPIs and cancer risk, i.e. that PPI use increases cancer risk when in reality the cancer “causes” the PPI use. Reverse causation can be dealt with by

implementing a lag-period within which all drug exposure is disregarded, i.e. within a time period of a given length prior to the cancer diagnosis (Pottegård & Hallas, 2017).

In study II, we implemented various lag-periods in several sensitivity analyses, ranging from 0-24 months using 6-month intervals. As results from these analyses show, reverse causation was not really a problem in study II. This might be due to these three cancers not causing physical symptoms that are likely to lead to the initiation of PPI treatment. To this point, we repeated our main analyses from study II and looked at the effect of implementing different lag-periods when assessing the association between PPI use and gastric cancer risk (ICD-10: C16.0-C16.9). As the results presented in Table 3 show, removing the lag-period yielded a higher risk estimate, indicating that the results are likely influenced by reverse causation. Although we performed this analysis on the association between PPI use and gastric cancer, we do not want to conclude anything from the observed results since these analyses were mainly done to explore the effect of implementing different lag-periods, where the underlying symptoms from an undiagnosed tumor were likely to influence initiation of PPI treatment. These observations should encourage the implementation of various lag-periods in pharmacoepidemiological studies exploring the association between drug use and cancer.

Table 3. Associations between proton pump inhibitor use and gastric cancer, with varying length of lag-time implemented.

| Gastric cancer: Ever use vs. Never use |              |                 |                          |
|--|--------------|-----------------|--------------------------|
| Lag-time (months)                      | No. of cases | No. of controls | Adjusted OR <sup>‡</sup> |
| 0                                      | 154 (64.4)   | 855 (35.8)      | 3.61 (2.70 - 4.83)       |
| 6                                      | 93 (38.9)    | 809 (33.9)      | 1.26 (0.95 - 1.68)       |
| 12                                     | 84 (35.1)    | 766 (32.1)      | 1.16 (0.86 - 1.56)       |
| 18                                     | 80 (33.5)    | 724 (30.3)      | 1.17 (0.87 - 1.58)       |
| 24                                     | 72 (30.1)    | 676 (28.3)      | 1.10 (0.80 - 1.49)       |

Prescribing patterns among patients that are close to death likely reflect worsening physical conditions and PPIs are commonly prescribed to a patient with a life-limiting medical diagnosis (McNeil et al., 2016). In study III, we therefore lagged the exposure by 12 months, following the date that patients met the exposure criteria, to limit the influence of changing prescribing patterns nearing end of life. In a sensitivity analysis where we removed the

lag-period, the observed increase in mortality indicates that without the exposure lag, our results would likely have been influenced by reverse causation (Table 4).

Table 4. Cox proportional hazard regression models for the associations between post-diagnosis PPI use and prostate cancer-specific and all-cause mortality, without lagging the exposure.

| PPI exposure                              | No of deaths | No of person years | Age adjusted HR (95% CI) <sup>b</sup> | Adjusted HR (95% CI) <sup>c</sup> |
|---|--------------|--------------------|---------------------------------------|-----------------------------------|
| <b>Prostate cancer-specific mortality</b> |              |                    |                                       |                                   |
| Non-use                                   | 49           | 3854               | 1.00 (Reference)                      | 1.00 (Reference)                  |
| Post-diagnosis PPI use                    | 49           | 1059               | 3.69 (2.48 - 5.50)                    | 3.95 (2.59 - 6.02)                |
| <b>All-cause mortality</b>                |              |                    |                                       |                                   |
| Non-use                                   | 123          | 3854               | 1.00 (Reference)                      | 1.00 (Reference)                  |
| Post-diagnosis PPI use                    | 80           | 1059               | 2.34 (1.76 - 3.10)                    | 2.29 (1.71 - 3.08)                |

Abbreviations: HR, hazard ratio; CI, confidence interval

<sup>b</sup>Adjusted for age at diagnosis

<sup>c</sup>Adjusted for age at diagnosis, calendar period, clinical stage, Gleason score, medication-based comorbidity, surgery, endocrine and/or chemotherapy, radiotherapy

## 5.4 Strengths and limitations

The most important overall strength of our studies involved nationwide data sources of high quality allowing us, e.g. through the Icelandic Medicines Registry, to assess PPI use among the entire Icelandic adult population and, e.g. through the Icelandic Cancer Registry, to identify all of the cancer diagnoses in Iceland relevant for our studies. Furthermore, the nature of how the data were collected for each data source, independently from one another, allowed our analyses to be carried out without us having to worry about recall bias; a bias that can be problematic if e.g. survey data were used to assess prior drug use. Lastly, the time-varying nature of our analyses in studies II and III enabled us to avoid time-related biases which likely would have skewed our results, leading us to draw false conclusions.

Our studies also had several overall limitations that we were unable to avoid. First, and perhaps most importantly, we lacked information about the underlying reason for PPI use which limited our ability to assess the appropriateness of PPI use in study I and further limited our ability to address potential confounding factors in studies II and III that might have influenced our results. To address the potential of confounding by indication influencing our results in studies II and III, we considered performing analyses using H2RA use as an active comparator, i.e. to compare the observed association for PPIs with the association for another drug used to treat the same clinical condition. However, this was not a feasible option due to the low-level of H2RA use in our underlying study population. In 2014 the total amount of sold H2RA, measured in DDDs, was 387,584 DDDs, out of which 245,375 DDDs were sold OTC. That amounts to 63.3% of the total amount of sold

H2RA drugs during that year. For comparison, there were 10,866,604 DDDs of PPIs sold in 2014, out of which 9.4% were sold OTC.

Second, we were unable to obtain information on individual-level risk factors such as body-mass-index (BMI), SES, alcohol use, smoking, and disease-based comorbidities. The lack of adjustment for these variables might have contributed to some level of unmeasured confounding influencing our results in studies II and III.

Third, information on PPI use before 2003 was unknown which might be a source of misclassification bias. Further, PPIs and NSAIDs are available OTC in pharmacies in Iceland which might have resulted in some misclassification of their use. However, as mentioned in study I the overall OTC use of PPIs, our primary exposure in all three studies, never exceeded 10% of the total use of PPIs after they became available OTC on February 1, 2009. Nevertheless, it cannot be completely ruled out that misclassification of PPI use might have biased the results from studies II and III towards the null.

Fourth, another potential source of misclassification bias might stem from the fact that the Medicines Registry only contains information on outpatient PPI use, leaving us in the dark about their use within hospitals. PPIs have been shown to be used extensively among hospitalized patients, with reports of approximately 50% of inpatients being prescribed PPIs during their stay within the hospital setting. This might have influenced our results to some degree, especially in study III where a higher proportion of patients likely entered the inpatient setting at some point following their diagnosis. Fifth, we lacked information on exact dosing regimens for the PPI prescriptions which forced us to assume a daily intake of either one tablet or one DDD, although it is likely that patients with more severe symptoms might have had a higher daily intake. In study I, this would have further allowed us to evaluate the appropriateness of PPI use. If a high proportion of patients consumed a higher daily dose than one tablet/DDD that would potentially affect our estimates of treatment duration in studies I and II.

Sixth, as in all registry-based studies on drug use we had to assume that patients receiving a dispensing for a drug do actually take them. It remains likely that some patients that receive PPIs only take them occasionally and on-demand. In fact, our observation in study I of lowered prevalence estimates, when prevalent use was redefined by requiring two filled prescriptions rather than one, supports the idea that a number of PPI users can probably be referred to as 'occasional users'. Seventh, the Icelandic Medicines Registry did not contain information on PPI use within nursing

homes until 2010, which might have resulted in PPI use being somewhat underestimated before that time, since the prevalence of PPI use was shown to increase with age.

Eighth, the length of our study periods in studies II and III were limited in part by the information that was available to use. In study II, we were limited by the fact that the Medicines Registry only started in 2003 and to be able to implement a 24-month lag-period we had to limit the start of the study period to the year 2005. In study III, we were limited by the lack of comprehensive information from Landspítali hospital on chemo- and radiotherapy prior to 2007, forcing us to limit the study period to 2007-2012. Additionally, because we only had information on clinical stage prior to 2011 for prostate cancer diagnoses, we were unable to include breast cancer patients in study III, as we initially intended.

Finally, one of our original aims was to assess the mortality among patients that received PPIs prior to chemotherapy but this turned out to be infeasible due to the small sample size of patients receiving chemotherapy. Therefore, it is clear that studies II and III would benefit greatly from an increased period of follow-up time. Finally, despite our best efforts, we cannot entirely exclude the possibility that the aforementioned biases in chapter 5.3 might have influenced our results to some degree, although our precautions should have substantially limited their effect on our study results.



## 6 Conclusions and future studies

Our findings suggest that overall PPI use has increased considerably since 2003, driven by a substantial increase of prescriptions in primary care. Our results indicate that the observed increase was mainly due to increased use among current users, especially among the elderly. Furthermore, our observations of extended treatment durations, often on higher doses and well beyond the recommended duration of PPI treatment according to clinical guidelines, should encourage future studies to explore the appropriateness of the extensive PPI use observed in Iceland in this study.

Overall our findings do not support our hypothesis that PPIs possess antineoplastic properties. Specifically, our results do not suggest a chemopreventive role of PPIs in breast cancer, prostate cancer, or malignant melanoma. Future studies on PPI use and cancer risk should focus on clinical stage and whether PPIs influence the risk of being diagnosed with a metastatic disease, given the evidence of increased plasmalemmal V-ATPase expression in metastatic cancer cells. Furthermore, our results do not indicate that post-diagnosis PPI use is associated with decreased mortality among prostate cancer patients. Future observational studies on PPI use and mortality among cancer patients should focus on whether pre-treatment with PPIs among patients receiving chemotherapy influences mortality, possibly by enhancing the chemotherapeutic effect. However, due to the high level of PPI use among cancer patients, likely with various indications, confounding by indication likely needs to be addressed; perhaps by stratifying by underlying clinical indications.





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## **Original publications**



# Paper I





# Proton-pump inhibitors among adults: a nationwide drug-utilization study

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## Abstract

**Background:** The use of proton-pump inhibitors (PPIs) has grown worldwide, and there are concerns about increased unsubstantiated long-term use. The aim of the study was to describe the real-world use of PPIs over the past decade in an entire national population.

**Methods:** This was a nationwide population-based drug-utilization study. Patterns of outpatient PPI use among adults in Iceland between 2003 and 2015 were investigated, including annual incidence and prevalence, duration of use, and dose of tablet used (lower versus higher), as well as the proportion of PPI use attributable to gastroprotection.

**Results:** We observed 1,372,790 prescription fills over the entire study period, of which 95% were for higher-dose PPIs. Annual incidence remained stable across time (3.3–4.1 per 100 persons per year), while the annual prevalence increased from 8.5 per 100 persons to 15.5 per 100 persons. Prevalence increased with patient age and was higher among women than men. Duration of treatment increased with patients' age (36% of users over 80 years remained on treatment after 1 year compared with 13% of users aged 19–39 years), and was longer among those initiating on a higher dose compared with a lower dose. The proportion of PPI users concurrently using nonsteroidal anti-inflammatory drugs decreased over the study period, while the proportion concurrently using acetylsalicylic acid, oral anticoagulants, or platelet inhibitors increased.

**Conclusions:** In this nationwide study, a considerable increase in overall outpatient use of PPIs over a 13-year period was observed, particularly among older adults. Patients were increasingly treated for longer durations than recommended by clinical guidelines and mainly with higher doses.

**Keywords:** incidence, nationwide, pharmacoepidemiology, prevalence, proton-pump inhibitors, treatment duration

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## Introduction

Proton-pump inhibitors (PPIs) are commonly prescribed for several acid-related disorders,<sup>1</sup> such as gastroesophageal reflux disease (GORD) and peptic ulcer disease.<sup>2–5</sup> These drugs are also effective in treating ulcers associated with the use of nonsteroidal anti-inflammatory drugs (NSAIDs) and as prophylactic treatment for patients on NSAIDs and low-dose aspirin.<sup>6–10</sup> Recommended doses and duration of PPI treatment vary by indications. Clinical guidelines rarely recommend PPI treatment for more than

8–12 weeks.<sup>11,12</sup> High-dose treatment is recommended when initiating therapy for GORD and peptic ulcer disease, while low-dose treatment is generally regarded as a maintenance therapy for recovering patients.<sup>12</sup>

PPIs are generally considered safe.<sup>13</sup> However, their use has been associated with increased risks of adverse events, such as bone fractures,<sup>14</sup> kidney disease,<sup>15</sup> microscopic colitis,<sup>16</sup> and hypomagnesemia.<sup>17</sup> Use of PPIs has also been suggested to cause changes in the composition of the intestinal

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microbiota, increasing the risk of *Clostridium difficile* infection<sup>18</sup> and chronic liver disease.<sup>19</sup> Although PPIs have been shown to minimize NSAID-related adverse effects in the stomach, recent evidence suggests that PPIs might cause changes in the composition of the small intestinal microbiota, augmenting unwanted adverse effects of NSAIDs in the small intestines.<sup>20</sup> Furthermore, discontinuation of PPI treatment has been linked to acid hypersecretion<sup>21</sup> and the development of dyspeptic symptoms in healthy volunteers.<sup>22</sup>

PPIs have had undisputed effects on the treatment of symptoms related to excessive acid secretion, but concerns are growing about inappropriate indications and potential overuse, both within hospitals and in the primary-care setting.<sup>23–26</sup> These concerns are compounded by observations of increased long-term use especially in elderly populations,<sup>27–29</sup> where overprescribing has been associated with increased morbidity and mortality.<sup>30</sup>

In light of these concerns, we aimed to provide data on real-world use of PPIs, and changes thereof, across the past decade in an entire national population. Specifically, we aimed to determine patterns of use by patient and prescriber characteristics, including treatment duration contrasting between higher- and lower-dose PPIs. Furthermore, we described the proportion of PPI use attributable to gastroprotection.

### Methods

This was an observational drug-utilization study describing the use of PPIs among the adult Icelandic population (19 years or older) during the period 1 January 2003 through to 31 December 2015.

### Data sources

The Icelandic Medicines Registry (IMR) contains individual information on all dispensed prescription drugs in outpatient care in Iceland since 1 January 2003. We received information from the IMR on PPI dispensing during the study period. As of 2010, the IMR also contained information on dispensed prescription drugs within nursing homes in Iceland.<sup>31,32</sup> Completeness of the IMR ranged from 91% to 98% of all dispensed prescription drugs for the study years. Information on wholesale statistics of PPIs was provided by the Icelandic Medicines Agency.<sup>33</sup>

The Icelandic Population Register provided information about all citizens, Icelandic and foreign, residing in Iceland during the study period, including data on month and year of birth, sex, residency at 1 January 2003, migration status, and date of death (if appropriate).

Using personal identification numbers, unique to every individual residing in Iceland, we linked together the variables from these two registries.

### Study drugs

The drugs of interest were classified according to the World Health Organization anatomical therapeutic chemical/defined daily doses (ATC/DDD) classification.<sup>34</sup> During the study period, four PPI substances were prescribed in Iceland: omeprazole (A02BC01), lansoprazole (A02BC03), rabeprazole (A02BC04), and esomeprazole (A02BC05). We further categorized each PPI type by available tablet strengths in milligrams as higher or lower dose. In the National Institute for Health and Care Excellence (NICE) clinical guidelines, PPI doses (in mg) are defined as standard/full dose, double dose, or low dose.<sup>12</sup> In the current study, standard and double doses were defined as higher-dose PPIs and low doses as lower-dose PPIs (Table 1).

On 1 February 2009, PPIs became available as over-the-counter (OTC) products in Iceland. However, the majority of PPIs during the study period were obtained by prescription rather than OTC, with OTC sales ranging from 1% in 2009 to 10% in 2015 of the total dispensed DDDs in these years (Supplementary Table S1).

Information on the indication for the prescription of PPIs was not available in the IMR. We explored potential reasons for PPI use by assessing the proportion of use attributable to gastroprotection, that is, concurrent use of PPIs with acetylsalicylic acid (ATC codes: B01AC06, N02BA01, B01AC30), NSAIDs (ATC codes: M01, excluding M01AX), oral anticoagulants (ATC codes: B01AA, B01AE, B01AF, B01AX06), and platelet inhibitors (B01AC04, B01AC07, B01AC22, B01AC24, B01AC30).

### Analysis

We presented overall use of PPIs in Iceland as the total number of dispensed DDDs to the adult population stratified by calendar year, PPI substance,

**Table 1.** Proton-pump inhibitors and tablet strengths dispensed to adults in Iceland in 2003–2015.

| PPI          | ATC     | DDD (mg) | Lower dose (mg)* | Higher dose (mg)* |
|--------------|---------|----------|------------------|-------------------|
| Omeprazole   | A02BC01 | 20       | 10               | 20, 40            |
| Lansoprazole | A02BC03 | 30       | 15               | 30                |
| Rabeprazole  | A02BC04 | 20       | 10               | 20                |
| Esomeprazole | A02BC05 | 30       | 10               | 20, 40            |

\*National Institute for Health and Care Excellence clinical guidelines define PPI doses as standard/full dose, double dose or low dose.<sup>12</sup> Here we categorize low PPI doses as lower-dose PPIs while standard and double doses are categorized as higher-dose PPIs. ATC, anatomical therapeutic chemical; DDD, defined daily dose, PPI, proton-pump inhibitor.

and specialty of the prescribing physician (primary care, gastroenterology, and other specialties).

Annual prevalence (per 100 persons) of PPI use was defined as the number of adult individuals who filled at least one prescription in the relevant calendar year (2003–2015) divided by the total adult population residing in Iceland on 1 July of that year. Further we reported the sex- and age-specific prevalence of PPI use in 2015, the last year of the study period (by 1-year age intervals between ages 19–39 years and 80+ years). As a sensitivity analysis, we repeated the analysis of annual prevalence requiring at least two filled PPI prescriptions in the relevant calendar year to be classified as a prevalent user.

Annual incidence (per 100 persons per year) of PPI use was defined as the number of adult individuals who, during the relevant calendar year (2005–2015), filled their first PPI prescription after a period of 24 months during which no PPI prescriptions were filled, divided by the total adult population residing in Iceland on 1 July of that year.

To describe the duration of PPI use we used the ‘proportion of patients covered’ method, which estimates the proportion of subjects that are alive and covered by treatment on a given day after the initiation of an incidence treatment episode. For each patient, we estimated duration of each filled prescription based on days’ supply, assuming one tablet as a daily dose. We allowed for a grace period of 108 days ( $2 \times$  the median number of days between dispensing, that is, the number of days by which 50% of the population had received a subsequent dispensing), to account for irregular prescription fills and added to the duration of each prescription. If a patient did not fill a new

prescription within this time we considered them to have discontinued their PPI treatment. They could then later re-enter the user population upon initiating a new treatment episode. We followed incident PPI users for 5 years, from the date of their first PPI prescription (day 0), and calculated the proportion of patients covered by dividing the number of users that were using the drug at day X (defined by 30-day intervals) by the number of people who were still alive and had not migrated at day X. Furthermore, to assess differences in treatment duration by patient age or by their prescribed PPI dose, we stratified the duration analysis by age (19–39, 40–49, 50–59, 60–69, 70–79, 80+ years), dose strength (higher *versus* lower), and sex. In addition, we explored the distribution in number of dispensed DDDs and tablets in the first 5 years after start of initial treatment episode (0–99, 100–199, 200–299, 300–399, 400–499, 500–599, 600–699, 700–799, 800–899, 900–999,  $\geq 1000$ ).

To assess concurrent use of selected drugs (ATC codes: M01 [excluding M01AX], B01AC06, N02BA01, B01AC30, B01AA, B01AE, B01AF, B01AX06, B01AC04, B01AC07, B01AC22, B01AC24, and B01AC30), we calculated the proportion (%) of prevalent PPI users in each study year who also filled prescriptions for these drugs within 90 days leading up to a PPI prescription fill. To assess the pattern of concurrent use among different age groups we performed a stratified analysis by age (19–39, 40–64, 65+ years).

All analyses were performed using R version 3.4.2<sup>35</sup> and RStudio.<sup>36</sup> The study was approved by the National Bioethics Committee in Iceland (study reference number: VSNb2015080004/03.03). As the study was based on national registry data, we did not obtain informed consent from individuals

in the study population. All personal information was encrypted and de-identified prior to analysis.

## Results

We observed 1,372,790 prescription fills for PPIs over the entire study period. The vast majority (95%) were higher-dose prescriptions. Among 313,296 individuals constituting our source population, a total of 101,909 (33%) filled at least one PPI prescription, including 56,252 women (55%) and 45,657 men (45%). The mean age at first prescription fill was 46 years (interquartile range 30–60). We observed a median of three PPI prescription fills per patient (interquartile range 1–15). The median number of days between prescription fills was 54.

During the study period, there was an increase in total PPI use, measured as the number of dispensed DDDs, from 3.5 million DDDs dispensed in 2003 to 10.7 million DDDs dispensed in 2015 (Figure 1a). Primary-care physicians prescribed the majority (60%) of all dispensed DDDs during the study period, whereas gastroenterologists prescribed 11% and physicians of other specialties prescribed 29%. Prior to 2009, esomeprazole was the most commonly prescribed drug among all specialties. Although esomeprazole remained the PPI of choice among gastroenterologists, omeprazole became the most commonly prescribed PPI thereafter among nongastroenterologists (Figure 1b–d).

Figure 2 shows an increase in annual prevalence of PPI use with calendar time, from 8.5 per 100 persons in 2003 to 15.5 per 100 persons in 2015. Meanwhile, the incidence of PPI use ranged from 3.3 per 100 persons in 2005 to 4.1 per 100 persons in 2015. A more stringent measure of annual prevalence, requiring at least two prescription fills within a relevant year, yielded a prevalence of 5.4 per 100 persons in 2003 to 11.0 per 100 persons in 2015 (Supplementary Figure S1). Prevalence of PPI use was higher among women than men and increased with patient age (Figure 3).

We identified 74,973 incident PPI users in our study population, which we then followed for 5 years to estimate the proportion of users still on treatment over time. Figure 4(a) shows the estimated treatment duration stratified by patient age. The proportion of patients still on PPI treatment after 1 year was highest among those over

80 years of age, (36%) and lowest in those aged 19–39 years (13%). After 5 years, the proportion was highest in those aged 70–79 years (20%) and lowest among the youngest, 19–39 years (7%). The majority of patients filled fewer than 200 DDDs/tablets during the first 5 years after starting PPI treatment (Supplementary Figure S2).

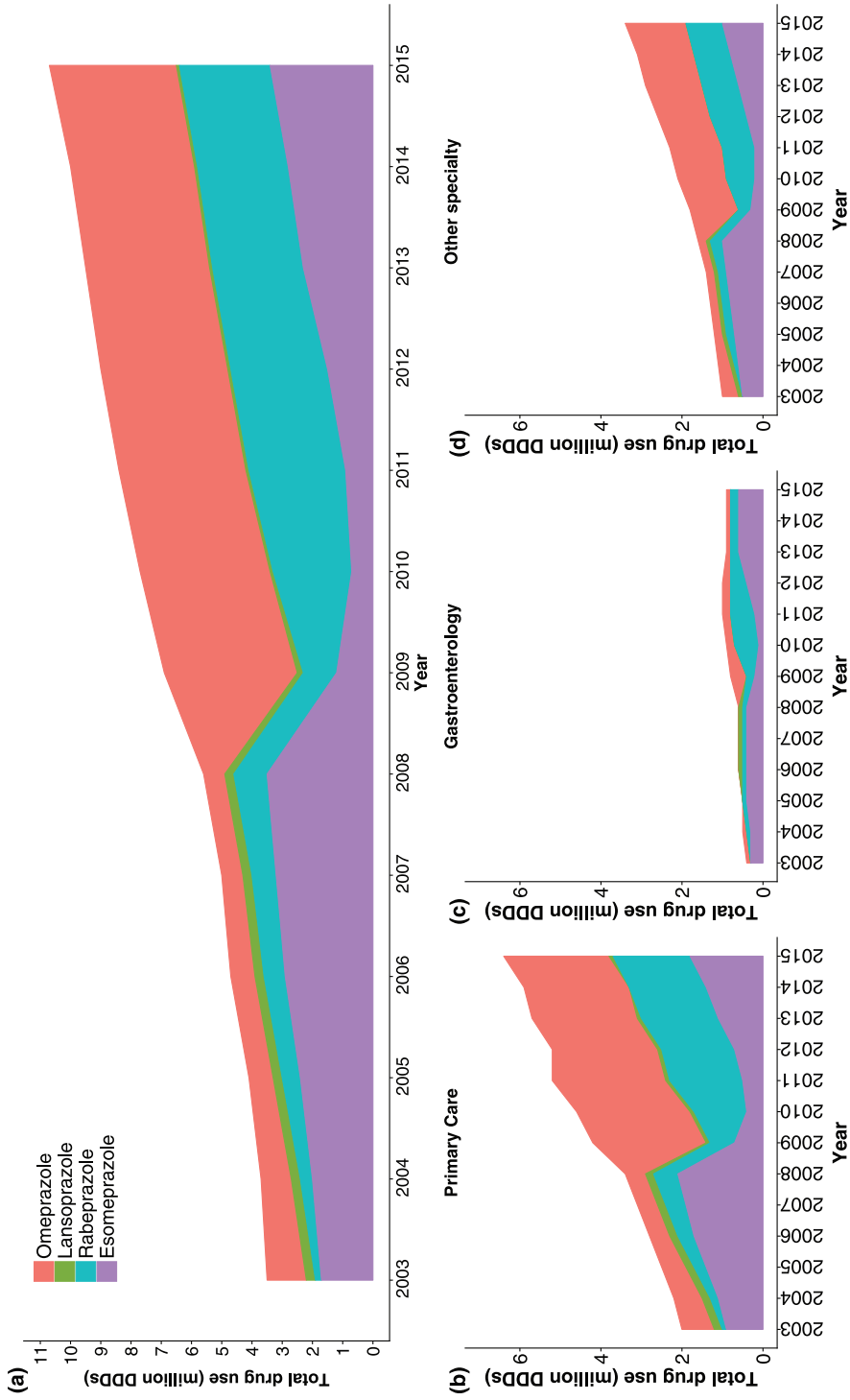
Figure 4(b) shows PPI treatment duration among incident PPI users stratified by strength of PPI dose at treatment initiation. Of the 74,973 incident users, 70,720 (94%) initiated on higher-dose PPIs and 4240 (6%) on lower-dose PPIs. The proportion of patients still treated with the same dose after 1 year was greater among those prescribed higher- (21%) than lower-dose PPIs (9%). The proportion of patients still on the same dose was 13% *versus* 2% after 5 years, respectively on higher- *versus* lower-dose PPIs. Duration of treatment by PPI dose strength was nearly identical for both sexes (Supplementary Figure S3).

We observed a slight decrease in the proportion of PPI users concurrently using drugs that have been shown to be ulcerogenic or increase the risk of bleeding, from 38% in 2003 to 36% in 2015 (Figure 5). The proportion of PPI users concurrently using NSAIDs decreased from 33% in 2003 to 24% in 2015. We observed an increase in concurrent use of oral anticoagulants (3–6%), acetylsalicylic acid (5–8%), and other platelet inhibitors (2–3%). The proportion of PPI users concurrently treated with any of these four drugs was highest among those aged over 65 years (47% in 2003, 47% in 2015) and lowest among the youngest aged 19–39 years (21% in 2003, 17% in 2015) (Supplementary Figure S4).

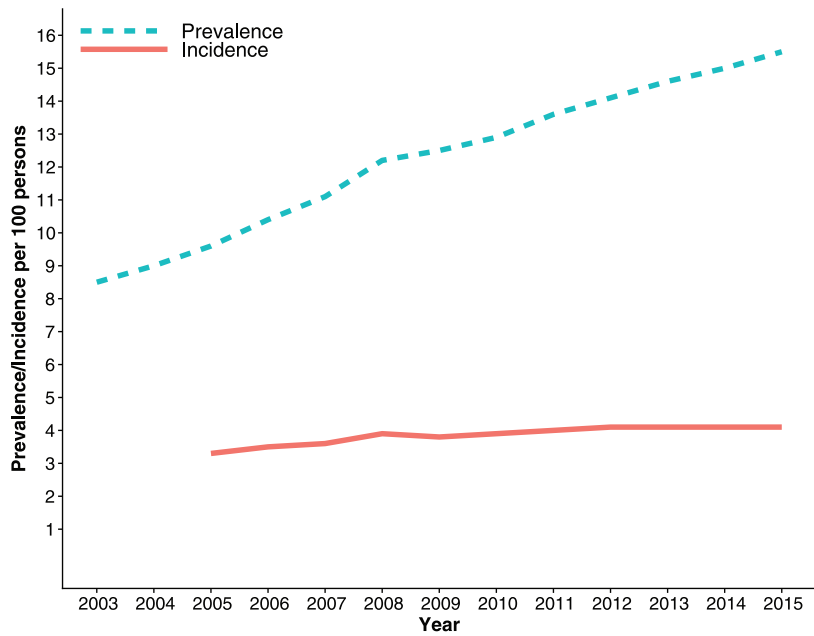
## Discussion

In this study, which covered all PPI dispensing in an entire national population over 13 years, we observed widespread and increasing use of PPIs, especially among the elderly. Primary-care physicians prescribed the vast majority of dispensed PPIs in our study data. While the number of new users remained relatively stable over time, the results suggested that patients were increasingly treated for longer durations than recommended by clinical guidelines and mainly with higher-dose PPIs.

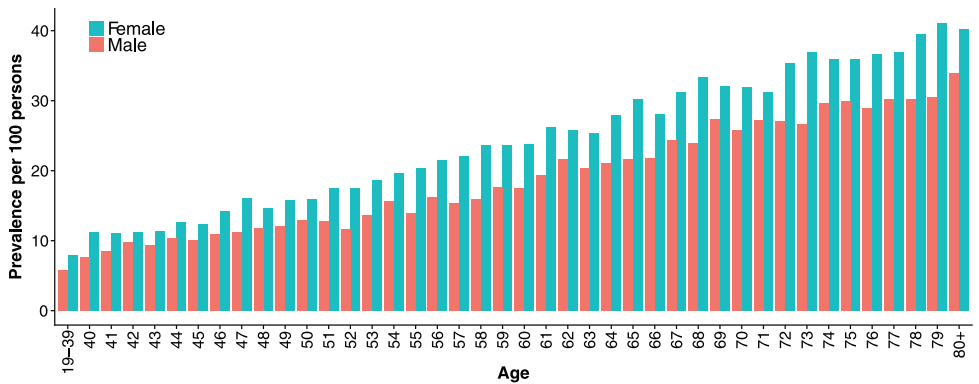
The rising prevalence of PPI use across time observed in our study is in line with recently



**Figure 1.** Proton-pump inhibitor (PPI) use among adults in Iceland during 2003–2015, measured in million dispensed defined daily doses (DDDs). (a) Overall use by PPI type; (b–d) overall use by specialty of prescribing physician and PPI type.



**Figure 2.** Annual prevalence and incidence (per 100 persons) of proton-pump inhibitor use among adults in Iceland.

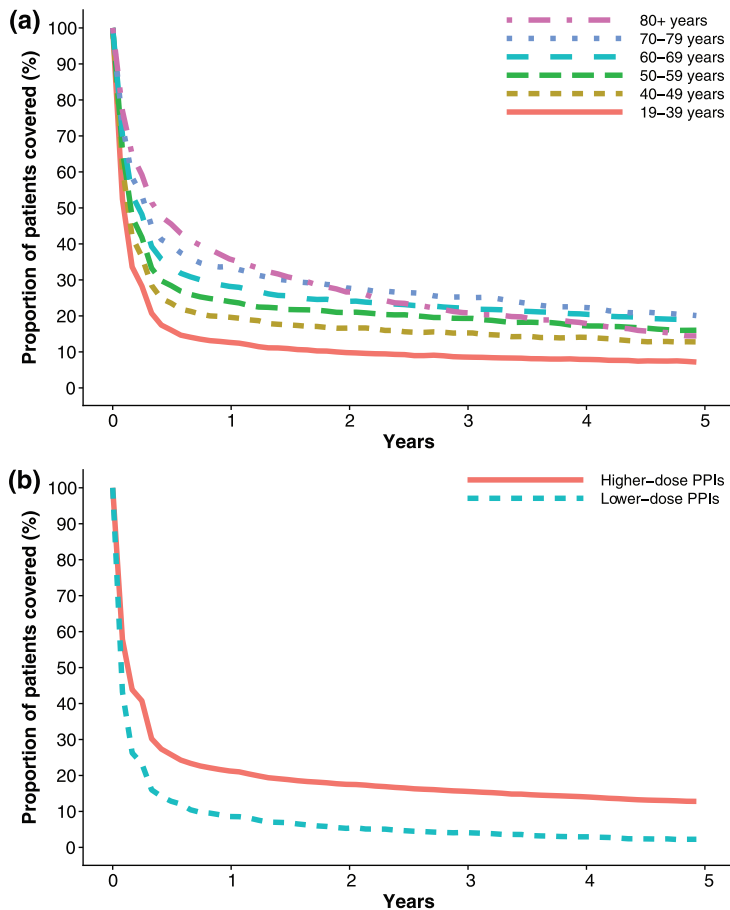


**Figure 3.** Age- and sex-specific prevalence of proton-pump inhibitor use among adults in Iceland in 2015.

published reports in comparable populations.<sup>27,29,37</sup> However, the prevalence in Iceland in 2015 was more than twice that observed among adults in Denmark in 2014 (15.5% versus 7.4%). GORD is the most common indication for PPIs with an estimated prevalence of 9–26% in European populations.<sup>38</sup> Although our use estimates were within this range, we were unable to draw definitive

conclusions on the appropriateness of PPI use in Iceland as we did not have information on the indications for which PPIs were prescribed nor data on the prevalence of GORD or other underlying conditions in the population.

Inappropriate use of PPIs in the outpatient setting, for example, in the form of inappropriate

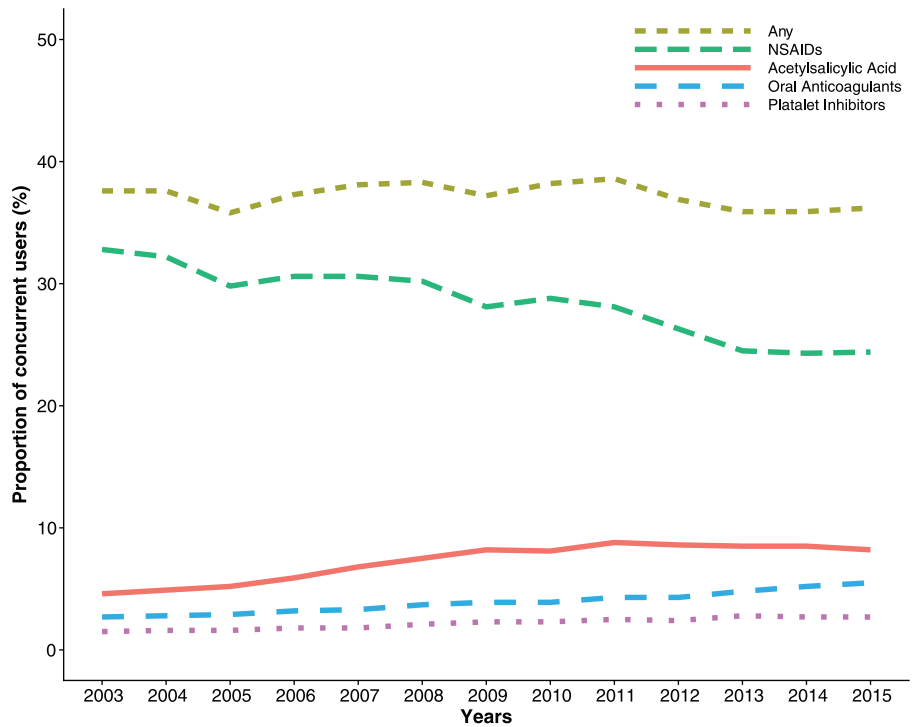


**Figure 4.** Duration of PPI treatment among incident users: (a) by age; (b) by initial dose strength of the proton-pump inhibitors (PPIs), measured as the proportion of patients covered.

indications and automatic renewal of prescriptions without re-evaluation of patients' symptoms, is a looming concern.<sup>25,39</sup> Such concerns were reinforced by Reimer and Bytzer's findings, which showed that only 27% of people receiving long-term treatment had a verified diagnosis justifying the need for long-term treatment.<sup>40</sup> The NICE clinical guidelines recommend long-term PPI therapy for rare conditions like Zollinger–Ellison syndrome or Barrett's esophagus as well as for patients with severe esophagitis, who have not responded to an initial high-dose 8-week treatment, and for patients who have experienced a dilation of an esophageal stricture.<sup>12</sup> In general, the recommended duration of PPI treatment in clinical guidelines rarely exceeds 12 weeks. We

found that 22% remained on treatment 1 year after treatment initiation. The proportion was highest among the oldest age group (36%) and lowest among the youngest (13%). Extended treatment durations among older adults are concerning in light of widespread polypharmacy and increased risk of adverse events with PPI use.<sup>41</sup> In fact, we observed that nearly half of older adults in our data used PPIs concurrently with NSAIDs, acetylsalicylic acid, oral anticoagulants, or platelet inhibitors, reflecting the level of polypharmacy among older adults using PPIs. Given the recent evidence of PPIs potentially facilitating injurious effects of NSAIDs in the small intestines, especially in older people and other high-risk patients,<sup>20</sup> this pattern of high concurrent drug use might be





**Figure 5.** Concurrent use of proton-pump inhibitors with drugs that are ulcerogenic or increase the risk of upper gastrointestinal bleeding. NSAIDs, nonsteroidal anti-inflammatory drugs.

concerning. However, as we were unable to link prescription data with clinical information, we cannot rule out that these patients were appropriately prescribed PPIs as bleeding prophylaxis.

The vast majority of PPI users in our population initiated treatment with higher-dose PPIs and after 1 year 21% remained on that treatment, for example, had not switched to lower-dose PPIs or discontinued treatment. This might indicate that their underlying symptoms are more severe than among those initiating treatment on lower doses and reflect the level of difficulty some users experience when discontinuing treatment due to resurfacing symptoms.<sup>42</sup>

Recently, Helgadottir and colleagues demonstrated that among confirmed GORD patients on long-term PPI treatment, women were more likely than men to be able to lower their dose by half, while still achieving symptom relief.<sup>43</sup> In our study we found no observable difference in treatment durations by patient sex, nor did women

seem more likely to initiate or maintain treatment on lower-dose PPIs. Thus, it is conceivable that women might be able to tolerate lower PPI doses than is mostly used nowadays.

The present study has several limitations. First, as with all register-based drug studies, it is not certain that individuals who filled the PPI prescriptions actually consumed the drugs. To address this, we performed a sensitivity analysis (Supplementary Figure S1) requiring at least two PPI prescription fills within a year to count as a prevalent PPI user, which resulted in lowered prevalence estimates. Actual consumption might thus in reality lie between these two measures of prevalence. Second, the study data did not contain information on clinical characteristics such as indications underlying the PPI prescriptions and/or the severity of symptoms, which prevented us from drawing sound conclusions on the appropriateness of PPI prescribing in our population. Third, information on PPI use within nursing homes was not included in the IMR until 2010,

which presumably resulted in an underestimation of the prevalence of PPI use among the elderly in the first half of the study period. Fourth, information on exact dosing for each prescription was not available in our data preventing us from accurately assessing prescribed doses. Our assessments of PPI doses were based on dispensed tablet strengths and therefore only an approximation of actual doses. Finally, PPIs became available OTC on 1 February 2009. However, the proportion of PPIs sold OTC was relatively low, ranging from 1% to 10% of the total number of DDDs sold annually from 2009 to 2015, and may therefore only have led to a slight underestimation of overall PPI use.

In conclusion, over a 13-year follow-up period we observed a considerable increase of real-world PPI use in a nationwide population setting, particularly among older adults. We found that a number of patients stayed on PPI treatment for longer periods than is recommended by clinical guidelines, mainly on higher doses. In view of these results, further initiatives towards appropriate prescribing of PPIs, especially in terms of the adoption of de-prescribing strategies, are warranted.

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### Conflict of interest statement

The authors declare no conflicts of interest in preparing this article.

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## Paper II





# Proton pump inhibitor use and risk of breast cancer, prostate cancer, and malignant melanoma: An Icelandic population-based case-control study

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## Abstract

**Purpose:** Increased expression of Vacuolar-type H<sup>+</sup> ATPases (V-ATPases), in the plasma membrane of cancer cells has been suggested to contribute to the development of aggressive cancer phenotypes by promoting acidic tumor microenvironments. Accumulating data suggest that proton pump inhibitors (PPIs) may elicit a chemopreventive effect via V-ATPase inhibition in some cancers, but evidence is still limited. Therefore, we aimed to explore a potential preventive role of PPIs in this study.

**Methods:** In this population-based case-control study, we identified incident cases of breast cancer (n = 1739), prostate cancer (n = 1897), and malignant melanoma (n = 385) in Iceland between 2005 and 2014 from the Icelandic Cancer Registry. We assessed varying levels of PPI use through record linkages to the Icelandic Medicines Registry. For each case, we selected up to 10 age-matched, sex-matched, and calendar-matched population controls using risk-set sampling. Using conditional logistic regression, we calculated odds ratios (ORs) and 95% confidence intervals (CIs) controlling for NSAID use.

**Results:** Adjusted ORs associated with ever use of PPIs were 1.03 (95% CI: 0.92-1.16) for breast cancer, 1.12 (95% CI: 1.00-1.25) for prostate cancer, and 0.84 (95% CI: 0.69-1.12) for malignant melanoma. Analyses of high use of PPIs ( $\geq 1000$  DDDs) yielded ORs of 0.97 (95% CI: 0.78-1.19), 1.20 (0.99-1.47), and 0.59 (0.40-1.13) for breast cancer, prostate cancer, and malignant melanoma, respectively. Analyses of cumulative exposure to PPIs did not support a dose-response relationship for any of the three cancer types.

**Conclusions:** Our findings do not support a chemopreventive effect of PPI use on breast cancer, prostate cancer, or malignant melanoma.

## KEYWORDS

breast cancer, melanoma, pharmacoepidemiology, prostate cancer, proton pump inhibitors, V-ATPase

## 1 | INTRODUCTION

Altered energy metabolism of cancer cells, characterized by high-glycolytic rate, has been proposed as one of the hallmarks of cancers.<sup>1</sup>

This high rate of glycolysis generates an excess amount of protons within the intracellular environment of cancer cells.<sup>2</sup> A slightly alkaline intracellular pH is preserved by facilitating the transport of metabolic products out of cancer cells and into the extracellular environment,



via membrane-bound transporters and channels, thus promoting an acidic extracellular environment.<sup>3</sup> Among these, membrane-bound proteins are the vacuolar type H<sup>+</sup> ATPases (V-ATPases), which are complex multisubunit proteins that can be found in a variety of cellular membranes where they facilitate the transport of protons and regulate intracellular and extracellular pH.<sup>4-6</sup> Plasmalemmal expression of V-ATPases has been associated with increased cancer cell survival, enhanced metastatic potential and the development of multidrug resistance through the acidification of the tumor microenvironment.<sup>7-13</sup> Furthermore, inhibition of V-ATPase function, via either V-ATPase specific inhibitors or proton pump inhibitors (PPIs), has been shown to have anticarcinogenic effects in a variety of cell-based and animal-based models, including breast cancer,<sup>14,15</sup> prostate cancer,<sup>16</sup> and melanoma.<sup>17,18</sup>

PPIs are commonly used drugs that are generally well tolerated and routinely prescribed for acid-related disorders of the gastrointestinal tract.<sup>19</sup> They are prodrugs that accumulate and become active in acidic environments where they inhibit acid secretion.<sup>20</sup> Originally developed to inhibit the extrusion of protons through H<sup>+</sup>/K<sup>+</sup> ATPases in the parietal cells of the stomach,<sup>21</sup> PPIs have also been shown to reduce V-ATPase activity.<sup>22,23</sup> Previous observational studies on PPI use and cancer risk have primarily focused on cancers of the digestive organs (ICD-10: C15-C26) and have reported conflicting results. Several studies found that PPI use is not associated with colorectal cancer risk.<sup>24-27</sup> Three studies concluded that PPI use was not associated with increased risk of pancreatic cancer,<sup>28-30</sup> while others have reported the opposite.<sup>31,32</sup> Furthermore, some studies have reported an increased risk of oesophageal and/or gastric cancer associated with PPI use,<sup>33-37</sup> although some of them are likely subject to reverse causality,<sup>31</sup> confounding by indication,<sup>34-36</sup> or time-related biases such as immortal time bias and latency bias.<sup>37,38</sup>

In Iceland, the most commonly diagnosed cancers are prostate cancer among men and breast cancer among women.<sup>39</sup> Both have been studied in relation with the effect of V-ATPase inhibition, yielding promising anticancer effects *in vitro* and *in vivo*.<sup>14-16</sup> Melanoma is less common, but acidic pH has been shown to enhance the invasive potential of melanoma cells,<sup>40,41</sup> suggesting that inhibiting V-ATPase function may have antineoplastic effects.<sup>17,18</sup> Therefore, we aimed to explore a potential preventive role of PPIs by conducting a population-based case-control study using risk-set sampling. Cancer development typically occurs over long periods of time,<sup>42</sup> and it is not inconceivable that an imminent disease may affect intake of medications. Thus, we implemented a lag-time period in our analyses to minimize the risk of reverse causation. To our knowledge, this is the first population-based study to test if PPI use is associated with the risk of breast cancer, prostate cancer, or malignant melanoma.

## 2 | METHODS

### 2.1 | Setting

We conducted a population-based-nested case-control study in Iceland, to assess the association between proton pump inhibitor use and the risk of a first-time diagnosis of breast cancer, prostate cancer,

### KEY POINTS

- Previous studies *in vivo/in vitro* have reported that proton pump inhibitors (PPIs) may have antineoplastic effects
- This is the first epidemiological study to test if PPI use affects the risk of malignant melanoma, breast or prostate cancer
- Our results do not support a clear association between PPI use and malignant melanoma, breast or prostate cancer
- Future well-controlled epidemiological studies should take clinical staging into account, given the available evidence that Vacuolar-type H<sup>+</sup> ATPase (V-ATPase) is highly expressed in the plasma membrane of metastatic cancer cells.

or malignant melanoma among adults (18 years and older). Our study base consisted of all adult residents of Iceland on January 1, 2003, including both prevalent and incident users of PPIs. Using personal identification numbers, unique to every individual residing in Iceland, we linked nationwide data from the Cancer Registry, Medicines Registry, and Population Register.

The Icelandic Cancer Registry contains nationwide information on every cancer diagnosis in Iceland since 1955, categorized according to the 10th revision of the International Statistical Classification of Diseases (ICD-10).<sup>43</sup> The Icelandic Medicines Registry contains individual information on all dispensed prescription drugs in outpatient care in Iceland since January 1, 2003. The drugs of interest were classified according to the World Health Organization (WHO) Anatomical Therapeutic Chemical/defined daily doses (ATC/DDD) classification.<sup>44</sup> As of 2010, the Icelandic Medicines Registry also holds information on dispensed prescription drugs within nursing homes in Iceland.<sup>45,46</sup> The completeness of the Icelandic Medicines Registry is high, ranging from 91% to 98% of all dispensed prescription drugs for the study years.

From the Icelandic Population Register, we obtained information about all citizens, Icelandic, and foreign, residing in Iceland during the study period, including data on: month and year of birth, sex, residency on January 1, 2003, migration status, and date of death (if appropriate).

### 2.2 | Cases

From the Icelandic Cancer Registry, we identified 1739 individuals with a first-time diagnosis of breast cancer (ICD10: C50), 1897 individuals with a first-time diagnosis of prostate cancer (ICD-10: C61), and 385 individuals with a first-time diagnosis of malignant melanoma (ICD-10: C43) between January 1, 2005 and December 31, 2014. The date of diagnosis for each cancer was defined as the index date. We excluded individuals who had previously been diagnosed with any cancer prior to the start of the study period.

## 2.3 | Population controls

We selected controls from the total underlying adult population in Iceland ( $N = 220\,512$ ). Using risk-set sampling, we matched up to 10 controls to each case on birth year, sex, and calendar time. The controls had to be alive and cancer free at the index date. Each case was eligible for sampling as a control before the time of disease onset, and each sampled control was eligible to later become a case. The resulting odds ratios (ORs) should therefore provide estimates of the incidence rates comparable with those expected from a cohort study in the source population.<sup>47</sup> To ensure comparable exposure opportunity time within each risk set between cases and controls (ie, the amount of time prior the index date available for exposure ascertainment), all individuals had to have resided in Iceland from January 1, 2003 to the index date.

## 2.4 | Drug exposure

From the Icelandic Medicine Registry, we obtained information on all dispensed PPIs from 2003 to 2014. Four PPIs were prescribed to patients within the study population during this period: omeprazole (A02BC01), lansoprazole (A02BC03), rabeprazole (A02BC04), and esomeprazole (A02BC05). We defined the exposure as PPI use before the index date for both cases and controls. Individuals with one or more PPI dispensing prior to the assigned index date were considered as "ever-users" of PPIs, while those without any PPI dispensing were classified as "never-users."

Cumulative dose, measured as the total amount of dispensed "defined daily doses" (DDDs) prior to index date, was also estimated for each patient (<365 DDDs, 365-730 DDDs, 731-1096 DDDs, >1096 DDDs). We defined high-level PPI use as dispensed prescriptions for greater than or equal to 1000 DDDs prior to index date. Furthermore, based on a daily intake of one tablet, we estimated the duration of each PPI prescription among ever users, and subsequent dispensings were then added together to estimate the cumulative duration of PPI use prior to the index date (0-1 years, 1-5 years, >5 years).

To minimize the risk of reverse causality biasing our effect estimates, we introduced a lag period where the exposed person time within 24 months leading up to the index date was disregarded.

Since use of nonsteroidal anti-inflammatory drugs (NSAIDs) has been associated with both PPI use<sup>48,49</sup> and cancer,<sup>50-52</sup> we further obtained prescription data for the use of prescription NSAIDs, both aspirin and nonaspirin (ATC codes: M01A [excluding M01AX], B01AC06, N02BA01). To approximate longer-term use of NSAIDs, individuals with at least two NSAID dispensings prior to the index date were considered as NSAID users in our analysis.

## 2.5 | Statistical analysis

We used conditional logistic regression to calculate ORs and 95% confidence intervals (CIs) for the association between PPI use and a first-time diagnosis of the cancers of interest, based on a prevalent user design, analyzing each cancer separately. Analyses were adjusted

for NSAID use prior to index date. We did not add patient sex or birth to the multivariate regression models, as the matching of cases and controls on these variables was successful; only one prostate case could not be matched to any controls and was therefore excluded.

We performed subgroup analyses, assessing the effect of high-PPI use, cumulative dose, cumulative duration of PPI use, and calendar period (2005-2008, 2009-2011, and 2012-2014) on the hypothesized associations. Additionally, we performed several sensitivity analyses by implementing various lag times between 0 and 2 years with 6-month intervals. Furthermore, we repeated the main analysis employing a new-user study design, where we excluded all patients who dispensed a prescription for a PPI drug during 2003 or 2004, the first 2 years of the Icelandic Medicine Registry. Finally, we performed a post hoc supplementary analysis by clinical stage, ie, whether the disease was localized or nonlocalized, among patients diagnosed with prostate cancer between 2005 and 2012. Unfortunately, we did not have information on clinical staging for the years 2013 and 2014 and were thus unable to include them in the analysis. Also, we were unable to perform a similar analysis for breast cancer and malignant melanoma because of large amounts of missing information on clinical stage for these cancers in the years prior to 2012.

All analyses were performed using R<sup>53</sup> and R Studio.<sup>54</sup> The study was approved by the National Bioethics Committee in Iceland (study reference number: VSNb2015080004/03.03). As the study was based on national registry data, we did not obtain informed consent from individuals in the study population. All personal information was encrypted and de-identified prior to analysis.

## 3 | RESULTS

### 3.1 | Baseline characteristics

We identified 1739 cases of breast cancer, 1897 cases of prostate cancer, and 385 cases of malignant melanoma and matched these, respectively, with 17 390, 18 968, and 3850 population controls. The median age at index date was 62 years (Interquartile range [IQR]: 52-72 years) among breast cancer cases, 70 years (IQR: 63-77 years) among prostate cancer cases, and 55 years (IQR: 42-68 years) among melanoma cases (Table 1). Exposure opportunity time was comparable between cases and control for all three cancer types (Table 1).

### 3.2 | Association between PPI use and breast cancer, prostate cancer, or malignant melanoma

We first estimated the ORs for breast cancer, prostate cancer, and malignant melanoma associated with ever use and high use of PPIs, accounting for patient age, sex, calendar time, and NSAID use. These analyses yielded neutral adjusted ORs (Table 2).

We then conducted stratified analyses by cumulative duration of PPI use (0-1, 1-5, >5 years), cumulative dose (<365 DDDs, 365-730 DDDs, 731-1096 DDDs, >1096 DDDs). For breast and prostate cancer, these analyses mainly yielded ORs that were close to unity and similar to those observed for high use (Table 2). For prostate cancer,

**TABLE 1** Baseline characteristics of breast cancer, prostate cancer, and malignant melanoma cases and matched controls

|                                      | Breast Cancer           |                              | Prostate Cancer*        |                              | Melanoma               |                            |
|--------------------------------------|-------------------------|------------------------------|-------------------------|------------------------------|------------------------|----------------------------|
|                                      | Cases (%)<br>(n = 1739) | Controls (%)<br>(n = 17 390) | Cases (%)<br>(n = 1897) | Controls (%)<br>(n = 18 968) | Cases (%)<br>(n = 385) | Controls (%)<br>(n = 3850) |
| Sex                                  |                         |                              |                         |                              |                        |                            |
| Female                               | 1739 (100.0)            | 17 390 (100.0)               | -                       | -                            | 231 (60.0)             | 2310 (60.0)                |
| Male                                 | -                       | -                            | 1897 (100.0)            | 18 968 (100.0)               | 154 (40.0)             | 1540 (40.0)                |
| Age at index date                    |                         |                              |                         |                              |                        |                            |
| 18-29 years                          | 7 (0.4)                 | 70 (0.4)                     | 0 (0.0)                 | 0 (0.0)                      | 34 (8.8)               | 340 (8.8)                  |
| 30-39 years                          | 57 (3.3)                | 570 (3.3)                    | 0 (0.0)                 | 0 (0.0)                      | 46 (12.0)              | 460 (12.0)                 |
| 40-49 years                          | 286 (16.5)              | 2860 (16.5)                  | 28 (1.5)                | 280 (1.5)                    | 74 (19.2)              | 740 (19.2)                 |
| 50-59 years                          | 421 (24.2)              | 4210 (24.2)                  | 232 (12.2)              | 2320 (12.2)                  | 79 (20.5)              | 790 (20.5)                 |
| 60-69 years                          | 468 (26.9)              | 4680 (26.9)                  | 655 (34.5)              | 6550 (34.5)                  | 61 (15.8)              | 610 (15.8)                 |
| 70-79 years                          | 292 (16.8)              | 4680 (16.8)                  | 669 (35.3)              | 6690 (35.3)                  | 54 (14.1)              | 540 (14.1)                 |
| 80+ years                            | 208 (12.0)              | 2080 (12.0)                  | 313 (16.5)              | 3128 (16.5)                  | 37 (9.6)               | 370 (9.6)                  |
| Calendar period (year of index date) |                         |                              |                         |                              |                        |                            |
| 2005-2008                            | 661 (38.0)              | 6610 (38.0)                  | 807 (42.5)              | 8068 (42.5)                  | 158 (41.1)             | 1580 (41.1)                |
| 2009-2011                            | 558 (32.1)              | 5580 (32.1)                  | 523 (27.6)              | 5230 (27.6)                  | 126 (32.7)             | 1260 (32.7)                |
| 2012-2014                            | 520 (29.9)              | 5200 (29.9)                  | 567 (29.9)              | 5670 (29.9)                  | 101 (26.2)             | 1010 (26.2)                |
| Ever use of NSAIDs before index date |                         |                              |                         |                              |                        |                            |
| No                                   | 796 (45.8)              | 8008 (46.0)                  | 836 (44.1)              | 8943 (47.1)                  | 196 (50.9)             | 2105 (54.7)                |
| Yes                                  | 943 (54.2)              | 9382 (54.0)                  | 1061 (55.9)             | 10 025 (52.9)                | 189 (49.1)             | 1745 (45.3)                |
| Exposure opportunity time (days)     |                         |                              |                         |                              |                        |                            |
| Overall—mean                         | 1848                    |                              | 1797                    |                              | 1763                   |                            |
| By case-control status—mean          | 1848                    | 1848                         | 1797                    | 1797                         | 1763                   | 1763                       |

\*One prostate case could only be matched to nine controls but was included in all analyses.

we observed a slightly elevated adjusted OR of 1.26 (95% CI: 1.02-1.55) for cumulative dose of over 1096 DDDs and 1.22 (95% CI: 1.04-1.42) for cumulative use for 1 to 5 years. For malignant melanoma, the effect estimates decreased with increased PPI use but did not indicate a dose-response relationship.

Removing or changing the lag period did not significantly affect the observed associations between PPI use and first-time diagnosis for any of the three cancers of interest (Table S1), suggesting that reverse causality did not have a major impact on the main results. Also, employing a new-user design, rather than a prevalent-user design, did not change the results from the main analysis in any significant way (Table S2). In a supplementary analysis based on clinical stage, we observed similar risk estimates between subgroups (Table S3).

## 4 | DISCUSSION

In this population-based-nested case-control study, we found no clear evidence of a link between PPI use and reduced risks of breast cancer, prostate cancer, or malignant melanoma.

To our knowledge, this is the first epidemiological study to explore the possibility of a chemopreventive effect of PPI use on breast cancer, prostate cancer, and malignant melanoma risk. PPIs are prodrugs that selectively accumulate in acidic spaces where pH is below 4 and become functionally active through protonation.<sup>55</sup>

Previous studies indicating an antitumor effect of PPIs have mainly been conducted using cell-based and animal-based models. Those studies suggest that PPI treatment may inhibit proliferation of cancer cells, induce cytotoxicity, and reduce tumor growth.<sup>17,56,57</sup> The proposed underlying mechanism is that PPIs inhibit V-ATPases residing in the plasma membrane, inducing intracellular acidification and alkalization of the tumor microenvironment, which should hypothetically, have a chemopreventive effect. Although we observed a pattern of reduced risk of malignant melanoma with increased PPI use, the observed ORs did not indicate a dose-response relationship. For prostate cancer, we observed a marginally elevated ORs, but these results are likely a result of unmeasured confounding.

If PPIs do indeed possess a chemopreventive effect for these cancer types, our null findings could be explained by a number of factors. For PPIs to have a chemopreventive effect, they would first of all have to be distributed to tumor sites with low pH. Once there, the tumor microenvironment would have to be acidic enough for the PPIs to become functionally active and inhibit the flow of protons through the V-ATPase, from the intracellular environment and into the extracellular environment. And even if this occurs, it might still not be enough to alkalinize the tumor microenvironment. Although V-ATPase expression in the plasma membrane of cancer cells has been associated with the acidification of the tumor microenvironment, there are also other pH-regulating proteins, such as Na<sup>+</sup>/H<sup>+</sup> exchangers (NHE), carbonic anhydrases, HCO<sub>3</sub><sup>-</sup>-transporters, and monocarboxylate

**TABLE 2** Proton pump inhibitor (PPI) use among breast cancer, prostate cancer, and malignant melanoma cases and matched controls in 2005 to 2014

| Subgroups                          | Breast Cancer |               |                          |                                   | Prostate Cancer |               |                          |                                   | Malignant Melanoma |              |                          |                                   |
|------------------------------------|---------------|---------------|--------------------------|-----------------------------------|-----------------|---------------|--------------------------|-----------------------------------|--------------------|--------------|--------------------------|-----------------------------------|
|                                    | Cases (%)     | Controls (%)  | OR (95% CI) <sup>a</sup> | Adjusted OR (95% CI) <sup>b</sup> | Cases (%)       | Controls (%)  | OR (95% CI) <sup>a</sup> | Adjusted OR (95% CI) <sup>b</sup> | Cases (%)          | Controls (%) | OR (95% CI) <sup>a</sup> | Adjusted OR (95% CI) <sup>b</sup> |
| Never use                          | 1208 (69.5)   | 12 189 (70.1) | 1.00 (ref)               | 1.00 (ref)                        | 1342 (70.7)     | 13 899 (73.3) | 1.00 (ref)               | 1.00 (ref)                        | 301 (78.2)         | 2925 (76.0)  | 1.00 (ref)               | 1.00 (ref)                        |
| Ever use                           | 531 (30.5)    | 5201 (29.9)   | 1.03 (0.92-1.15)         | 1.03 (0.92-1.16)                  | 555 (29.3)      | 5069 (26.7)   | 1.14 (1.03-1.27)         | 1.12 (1.00-1.25)                  | 84 (21.8)          | 925 (24.0)   | 0.87 (0.67-1.13)         | 0.84 (0.69-1.12)                  |
| High use (≥ 1000 DDDs)             | 108 (6.2)     | 1110 (6.4)    | 0.97 (0.79-1.19)         | 0.97 (0.78-1.19)                  | 124 (6.5)       | 1028 (5.4)    | 1.23 (1.01-1.49)         | 1.20 (0.99-1.47)                  | 10 (2.6)           | 159 (4.1)    | 0.61 (0.32-1.18)         | 0.59 (0.40-1.13)                  |
| Cumulative dose (DDDs)             |               |               |                          |                                   |                 |               |                          |                                   |                    |              |                          |                                   |
| <365                               | 320 (18.4)    | 3209 (18.4)   | 1.01 (0.88-1.15)         | 1.01 (0.88-1.15)                  | 338 (17.8)      | 3160 (16.7)   | 1.12 (0.98-1.27)         | 1.10 (0.96-1.25)                  | 66 (17.1)          | 597 (15.5)   | 1.06 (0.80-1.41)         | 1.02 (0.76-1.36)                  |
| 365-730                            | 77 (4.4)      | 608 (3.5)     | 1.28 (1.00-1.63)         | 1.28 (1.00-1.64)                  | 73 (3.9)        | 603 (3.2)     | 1.26 (0.98-1.62)         | 1.24 (0.96-1.59)                  | 5 (1.3)            | 121 (3.1)    | 0.39 (0.16-0.97)         | 0.37 (0.15-0.93)                  |
| 731-1096                           | 34 (2.0)      | 374 (2.2)     | 0.92 (0.64-1.32)         | 0.92 (0.64-1.32)                  | 29 (1.5)        | 368 (1.9)     | 0.82 (0.56-1.20)         | 0.80 (0.54-1.18)                  | 4 (1.0)            | 68 (1.8)     | 0.56 (0.20-1.54)         | 0.53 (0.19-1.46)                  |
| >1096                              | 100 (5.7)     | 1010 (5.8)    | 1.00 (0.81-1.25)         | 1.00 (0.80-1.25)                  | 115 (6.1)       | 938 (4.9)     | 1.29 (1.05-1.59)         | 1.26 (1.02-1.55)                  | 9 (2.4)            | 139 (3.6)    | 0.61 (0.31-1.22)         | 0.58 (0.29-1.16)                  |
| Cumulative duration of use (years) |               |               |                          |                                   |                 |               |                          |                                   |                    |              |                          |                                   |
| 0-1                                | 264 (15.2)    | 2554 (14.7)   | 1.04 (0.91-1.20)         | 1.04 (0.90-1.20)                  | 258 (13.6)      | 2525 (13.3)   | 1.06 (0.92-1.23)         | 1.05 (0.91-1.21)                  | 55 (14.3)          | 468 (12.1)   | 1.13 (0.83-1.54)         | 1.09 (0.79-1.49)                  |
| 1-5                                | 184 (10.6)    | 1811 (10.4)   | 1.03 (0.87-1.21)         | 1.03 (0.87-1.21)                  | 210 (11.1)      | 1752 (9.2)    | 1.24 (1.07-1.45)         | 1.22 (1.04-1.42)                  | 21 (5.4)           | 335 (8.7)    | 0.60 (0.38-0.96)         | 0.57 (0.36-0.91)                  |
| >5                                 | 83 (4.8)      | 836 (4.8)     | 1.00 (0.79-1.28)         | 1.00 (0.79-1.28)                  | 87 (4.6)        | 792 (4.2)     | 1.15 (0.91-1.45)         | 1.12 (0.89-1.42)                  | 8 (2.1)            | 122 (3.2)    | 0.62 (0.30-1.30)         | 0.59 (0.28-1.24)                  |
| Calendar period <sup>d</sup>       |               |               |                          |                                   |                 |               |                          |                                   |                    |              |                          |                                   |
| 2005-2008                          | 117 (17.7)    | 1285 (19.4)   | 0.89 (0.72-1.10)         | 0.88 (0.71-1.09)                  | 165 (20.4)      | 1426 (17.7)   | 1.20 (1.00-1.45)         | 1.19 (0.99-1.43)                  | 25 (15.8)          | 242 (15.3)   | 1.04 (0.66-1.64)         | 1.00 (0.63-1.58)                  |
| 2009-2011                          | 202 (36.2)    | 1820 (32.6)   | 1.18 (0.98-1.42)         | 1.16 (0.96-1.40)                  | 164 (31.4)      | 1519 (29.0)   | 1.12 (0.92-1.36)         | 1.11 (0.91-1.35)                  | 28 (22.2)          | 314 (24.9)   | 0.85 (0.54-1.34)         | 0.85 (0.53-1.35)                  |
| 2012-2014                          | 212 (40.8)    | 2096 (40.3)   | 1.02 (0.85-1.23)         | 1.05 (0.87-1.27)                  | 226 (39.9)      | 2124 (37.5)   | 1.11 (0.93-1.32)         | 1.06 (0.88-1.27)                  | 31 (30.7)          | 370 (36.6)   | 0.76 (0.49-1.19)         | 0.71 (0.45-1.12)                  |

Abbreviation: CI: confidence interval; OR: odd ratio.

<sup>a</sup>Matched on birth year, sex, and calendar time.

<sup>b</sup>Matched on birth year, sex, calendar time and controlled for NSAID use.

<sup>c</sup>Year of index date.

transporters (MCTs)<sup>3,58</sup> that participate in the extrusion of protons out into the extracellular environment. Therefore, these membrane-bound transporters might maintain an acidic extracellular pH, lessening the impact of PPI-inhibited V-ATPase function.

V-ATPase expression has been shown to be increased in highly metastatic cancer cells compared with poorly metastatic cells,<sup>59,60</sup> implicating a potentially more important role of V-ATPase in tumor progression and invasiveness rather than cancer initiation. These reports make it plausible that the acidity of the tumor microenvironment during initial tumorigenesis might not be sufficient for the PPIs to accumulate at the primary tumor site. As an attempt to explore this issue, we performed a post hoc analysis among patients diagnosed with prostate cancer, which did not yield conclusive results. Therefore, a systematic analysis taking clinical stage into account for all three cancer types would provide further insight into this matter.

The main strength of our study was that it is nested within a clearly defined population-based cohort and is based on high quality, nationwide data sources. Furthermore, underlying data on exposure and outcome were collected prospectively and independently from each other for the entire Icelandic population, therefore eliminating any potential recall bias. The sampling method, where controls were randomly selected from the underlying population, minimized the risk of selection bias. Additionally, the inherent time varying nature of our study design allowed us to avoid common time related biases, such as immortal time bias and latency bias that have been shown to be an issue in studies of drug-cancer associations.<sup>38,42</sup> Furthermore, the study design ensured the same exposure opportunity time among cases and controls.

Our study has several limitations. First, it lacked important individual level information on common risk factors for PPI use and cancer, such as BMI, smoking, socioeconomic status (SES), and comorbidities. Therefore, residual confounding might explain the slightly elevated risk estimates observed for prostate cancer, eg, among ever users of PPIs and those with 1 to 5 years of cumulative duration of PPI use. Second, individuals already in contact with the healthcare system through prescription use of PPIs may be more likely than those without such prescriptions to receive a cancer diagnosis, yielding elevated risk estimates (detection bias). However, our sensitivity analyses allowing different lag periods to be tested suggest that such mechanisms had limited influence on our findings. Third, in 2009 low-dose PPIs became available over-the-counter (OTC) in Iceland, and OTC use is not recorded in the Medicine Registry. This may have led to some misclassification of PPI use in our study but is unlikely to have impacted the results much as the amount of OTC use was relatively low during the study period, ranging from 1% to 10% of total PPI volume sold annually in 2009 to 2014.<sup>61</sup> Another misclassification of PPI exposure might have modestly biased the study results since we did not have information on PPI use prior to 2003, causing some potential PPI users before 2003 to be considered as never users. Furthermore, we attempted to control for longer-term NSAID therapy prior to index date, but since these drugs are commonly used OTC, misclassification of NSAID exposure is likely to have occurred.

In conclusion, our findings did not support a chemopreventive effect of PPI use against breast cancer, prostate cancer, or malignant melanoma. Future well-controlled epidemiological studies need to

take clinical staging into account, given the available evidence that V-ATPase is highly expressed in the plasma membrane of metastatic cancer cells.

## ETHICS STATEMENT

The study was approved by the National Bioethics Committee in Iceland (study reference number: VSNb2015080004/03.03).

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## CONFLICT OF INTEREST

The authors declare that there is no conflict of interest.

## AUTHORSHIP CONTRIBUTION

Helga Zoega is the guarantor of the article. Helga Ogmundsdóttir, Margret H. Ogmundsdóttir, Eiríkur Steingrímsson, and Helga Zoega conceived the study, and Óskar Ö. Hálfðánarson and Helga Zoega designed it. Óskar Ö. Hálfðánarson, Helga Zoega, Katja Fall, and Sigrún Helga Lund contributed to the data analysis, and all authors contributed to the interpretation of the data. Óskar Ö. Hálfðánarson drafted the manuscript, and all authors participated in the interpretation of the data and revising the content of the manuscript. The final version of the manuscript was revised and approved by all authors.

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#### SUPPORTING INFORMATION

Additional supporting information may be found online in the Supporting Information section at the end of the article.

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## Paper III





# Use of proton pump inhibitors and mortality among Icelandic patients with prostate cancer

## Running title: PPIs and mortality in prostate cancer patients

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32 **AUTHORSHIP STATEMENT**

33 Helga Zoëga is the guarantor of the article. Helga M. Ogmundsdottir, Margret H  
34 Ogmundsdottir and Helga Zoëga conceived the study; and Óskar Ö. Hálfðánarson and  
35 Helga Zoëga designed it. Óskar Ö. Hálfðánarson, Helga Zoëga, Anton Pottegard and Sigrún  
36 Helga Lund contributed to the data analysis and all authors contributed to the interpretation  
37 of the data. Óskar Ö. Hálfðánarson drafted the manuscript and all authors participated in the  
38 interpretation of the data and revising the content of the manuscript. The final version of the  
39 manuscript was revised and approved by all authors.

40

41 There are no competing interests to declare.

42

43  
44 **ABSTRACT**

45 **Aims:** Proton pump inhibitors (PPIs) have both been reported to enhance chemosensitivity  
46 and contribute to increased mortality among cancer patients. Due to conflicting reports, we  
47 aimed to determine whether PPI use is associated with mortality among prostate cancer  
48 patients.

49  
50 **Methods:** In this population-based cohort study, we identified all eligible patients with an  
51 incident diagnosis of prostate cancer in Iceland between 2007-2012 (n = 1058). We used  
52 time-dependent Cox proportional hazard regression models to compute hazard ratios (HRs)  
53 and 95% confidence intervals (CIs) for prostate cancer-specific and all-cause mortality  
54 associated with post-diagnosis use of PPIs, defined as at least  $\geq 2$  filled prescriptions after  
55 diagnosis and lagged by 12 months.

56  
57 **Results:** Among the study cohort, we identified 347 (32.8%) post-diagnosis PPI users and  
58 711 (67.2%) non-users. Out of the 347 patients using PPIs after diagnosis, 59 patients (17.0%)  
59 died due to any cause and 22 patients (6.3%) due to prostate cancer, compared with 144  
60 (20.3%) and 76 (10.7%) among non-users, respectively. Post-diagnosis PPI use was not  
61 statistically significantly associated with prostate cancer-specific mortality (HR 0.88; 95% CI:  
62 0.52-1.48) or all-cause mortality (HR 1.02; 95% CI: 0.73-1.43). Stratification by timing of use  
63 and clinical stage did not reveal any statistically significant associations to the mortality  
64 outcomes of interest. Furthermore, we did not find any evidence of a significant dose-response  
65 relationship.

66  
67 **Conclusions:** Our findings did not indicate an association between post-diagnosis PPI use  
68 and mortality among prostate cancer patients.

69  
70

71 STATEMENT 1

72 **What is already known about this subject?**

- 73 • An acidic tumor microenvironment has been associated with a malignant cancer  
74 phenotype.
- 75 • Although *in vitro* and *in vivo* studies have suggested that proton pump inhibitors have  
76 antineoplastic properties and increase chemosensitivity a recent observational study  
77 reported that PPI use was associated with increased prostate cancer-specific and  
78 overall cancer mortality.

79

80

81

82 STATEMENT 2

83 **What this study adds:**

- 84 • Contrary to a previous report, post-diagnosis PPI use was not associated with  
85 increased mortality among prostate cancer patients.
- 86 • We found no evidence of decreased mortality risk among post-diagnosis PPI users.
- 87 • Timing of use was not statistically significantly associated with the outcome and we  
88 did not observe a dose-response association.

89

90

91

## 92 INTRODUCTION

93 Proton pump inhibitors (PPIs) are commonly used drugs and their use has been increasing  
94 quite rapidly over the last decade.[1] As potent inhibitors of acid secretion, PPIs were  
95 originally developed to inhibit the activity of the H<sup>+</sup>/K<sup>+</sup> ATPase, a type of proton pump  
96 that secretes gastric acid from parietal cells of the stomach.[2] However, they have also been  
97 shown to have an affinity for another proton pump, i.e. the vacuolar H<sup>+</sup>-ATPase (V-  
98 ATPase).[3, 4] The V-ATPase is frequently seen overexpressed in the plasma membrane of  
99 cancer cells where they are believed to promote alkalization of the cytoplasm and  
100 acidification of the tumor microenvironment.[5–10] Increased tumor acidity has been  
101 associated with a malignant cancer phenotype characterized by increased invasiveness,  
102 metastatic potential, and drug resistance.[11–13] Thus, due to the ability of PPIs to inhibit V-  
103 ATPase function their repositioning as potential antineoplastic agents has been suggested.[14]  
104 Studies, *in vitro* and *in vivo*, have reported a potential anticancer activity of PPIs[15–17] and a  
105 phase II trial among breast cancer patients with a metastatic disease reported increased  
106 efficacy of chemotherapy in patients pre-treated with PPIs.[18] Furthermore, a clinical study  
107 among osteosarcoma patients found that pre-treatment with PPIs improved the effectiveness  
108 of chemotherapy.[19] These results highlight a potential avenue for studying whether PPI use  
109 increases the effectiveness of cancer therapy in various cancer types.

110

111 The potential association between PPI use and cancer mortality has not been evaluated  
112 conclusively in epidemiological studies. A study among pancreatic cancer patients found no  
113 association between PPI use and survival.[20] Another study found that PPI use, and use of  
114 histamine receptor-2 antagonist (H2RA), was associated with improved overall survival  
115 among patients with head and neck squamous cell cancer.[21] A recent Danish study

116 reported that PPI use was associated with increased cancer-specific mortality for a number of  
117 cancer types, including prostate cancer.[22]

118

119 Prostate cancer is the second most commonly diagnosed cancer among men and the fifth  
120 most frequent cause of cancer-specific death.[23] Given the conflicting results of the few  
121 epidemiological studies conducted so far, the increasing overall use of PPIs, and the high  
122 incidence of prostate cancer, we aimed to utilize the high-quality nationwide registry data  
123 available in Iceland to examine the association between post-diagnosis PPI use and mortality  
124 among prostate cancer patients.

125

## 126 METHODS

### 127 Data sources

128 This was a population-based cohort study where we used unique personal identification  
129 numbers to link together data from the Icelandic Cancer Registry,[24] the Icelandic  
130 Medicines Registry, the Icelandic Population Register, the Cause of Death Register, and  
131 from electronic health records of Landspítali – The National University Hospital of Iceland.

132

### 133 Study population

134 Eligible patients, identified using the Icelandic Cancer Registry, were all adult Icelandic  
135 residents between 40 – 85 years of age with a verified first-time diagnosis of prostate cancer  
136 (ICD-10: C61) between January 1, 2007 and December 31, 2012.

137

### 138 Follow-up and mortality outcomes

139 The primary outcome in all analyses was prostate cancer-specific mortality. The secondary  
140 outcome was all-cause mortality. Prostate cancer-specific mortality was defined by the  
141 relevant ICD-10 code (C61) as the underlying cause of death. Eligible patients were followed  
142 from 12 months after prostate cancer diagnosis until their death, emigration, or end of the  
143 study period (December 31, 2015). We excluded those patients who died or emigrated from  
144 Iceland within 12 months after diagnosis.

145

### 146 Exposure assessment

147 We obtained information on PPI use from the Icelandic Medicine Registry; a nationwide  
148 prescription registry with a completeness ranging from 91% to 99%. Although PPIs became  
149 available over-the-counter (OTC) in 2009 the majority (>90%) of PPIs between 2009 and  
150 2015 were obtained by prescription.[1] We considered the Anatomical Therapeutic



151 Chemical (ATC)[25] code group A02BC as a PPI dispensing. Four PPI substances were  
152 prescribed within our cohort during the period under study: omeprazole (A02BC01),  
153 lansoprazole (A02BC03), rabeprazole (A02BC04), and esomeprazole (A02BC05). The  
154 information we received for every PPI prescription between 1 January 2003 and 31  
155 December 2015, including date of dispensing, ATC code, and number of dispensed ‘defined  
156 daily doses’ (DDDs).

157

158 The primary exposure was post-diagnosis PPI use, defined as at least two or more filled PPI  
159 prescriptions after prostate cancer diagnosis. In all analyses, we considered the exposed  
160 person-time of post-diagnosis PPI users in a time-dependent manner to avoid time-related  
161 biases such as immortal time bias.[26] In the main analysis, patients were thus initially  
162 considered unexposed until they received a second PPI prescription, after which they were  
163 considered exposed for the remainder of follow-up. Furthermore, the exposed person-time  
164 was lagged by 12 months to account for the possibility of reverse causation and to allow for a  
165 biologically meaningful latency period, since it is unlikely that a short duration of drug use  
166 would influence mortality outcomes in a significant way. Patients that did not receive at least  
167 two PPI dispensing after diagnosis were thus considered as non-users.

168

169 For the purposes of secondary analyses, we explored the timing of PPI use by assessing pre-  
170 diagnosis PPI use. Patients were considered pre-diagnosis users if they received at least two  
171 PPI prescriptions in the 3 years prior to diagnosis. Pre-diagnosis use was modelled as a time-  
172 fixed covariate, i.e. a dichotomous yes/no variable. Thus, patients exposed to PPIs were  
173 either considered to be ‘new PPI users’ or ‘continued PPI users’ based on their exposure  
174 status before and after diagnosis. We defined new users as those patients that only used PPIs  
175 after diagnosis while those who used PPIs prior to and after diagnosis were considered as

176 continuing PPI users. Additionally, we estimated the cumulative dose for each patient based  
177 on the total number of dispensed DDDs during exposed person-time (0 DDDs, 1-365 DDDs,  
178 >365 DDDs).

179

## 180 Covariates

181 We considered a range of demographic and clinical factors for multivariable adjustments.  
182 Patient age at diagnosis and year of diagnosis were modelled as continuous variables. A  
183 medication-based comorbidity score was derived by identifying the number of different  
184 prescription drug groups that were dispensed in the 12 months prior to a cancer  
185 diagnosis[27]. To be categorized in the same group the drugs had to share the same initial  
186 four characters of the ATC classification system. The medication-based comorbidity score  
187 was then modelled as a continuous variable. Clinical stage according to the tumor-node-  
188 metastasis (TNM) system was classified into three categories if information on M was  
189 available: localized (M0), non-localized (M1), and unknown (Mx or information missing). We  
190 adjusted for the following clinical variables: Gleason score was grouped into five distinct  
191 categories (2-5, 6, 7,  $\geq 8$ , unknown). Cancer treatment in the 12 months following diagnosis  
192 was accounted for in the following way: cancer surgery was categorized into three categories  
193 (total excision of prostate, partial excision of prostate, no surgery), cancer drug treatment was  
194 grouped into four categories (chemotherapy, endocrine therapy, combination of  
195 chemotherapy and endocrine therapy, no therapy), and radiotherapy was modelled as a  
196 dichotomous variable (radiotherapy, no radiotherapy).

197

## 198 Data analysis

199 We used a time-dependent Cox proportional hazard regression models, with time since  
200 diagnosis as the underlying time-scale, to estimate crude and multivariable adjusted hazard

201 ratios (HRs) and 95% confidence intervals (CIs) for prostate cancer-specific mortality and all-  
202 cause mortality associated with post-diagnosis PPI use modelled as a time-dependent  
203 covariate where patients were considered unexposed until they had met the exposure criteria,  
204 and then remained exposed throughout follow-up. In multivariable adjusted analyses we  
205 adjusted for the aforementioned covariates, also listed in Table 1.

206

207 In the main analysis, we assessed PPI use following prostate cancer diagnosis; modelled as a  
208 time-dependent covariate as described above. Exposed person-time was then lagged by 12  
209 months following a second dispensing of a post-diagnosis PPI prescription. Furthermore, we  
210 performed three secondary analyses. First, PPI use was stratified by continuing users versus  
211 new users. Second, we stratified by clinical stage (localized versus non-localized). Third, we  
212 stratified PPI use by cumulative dose (0 DDDs, 1-365 DDDs, >365 DDDs).

213

214 We performed three sensitivity analyses to assess the definition of PPI use. In the first one,  
215 post-diagnosis PPI use was defined as at least one filled PPI prescriptions following diagnosis  
216 and the exposure was modelled as a time-dependent covariate as in the main analysis. In the  
217 second sensitivity analysis, we defined post-diagnosis PPI use as at least two filled prescriptions  
218 within 12 months following the diagnosis of prostate cancer. In a third sensitivity analysis, we  
219 defined post-diagnosis PPI use as at least two filled prescriptions and assessed the exposure  
220 continuously throughout follow-up as a time-dependent covariate. Thus, by assuming a daily  
221 intake of one DDD and estimating the duration of each prescription as the number of  
222 dispensed DDDs we allowed patients to move back and forth between periods of non-use and  
223 periods of use.

224

225 All analyses were performed using the survival package[28] in R.[29] This study was  
226 approved by the National Bioethics Committee in Iceland (study reference number:  
227 VSNb2015080004/03.03).

228

## 229 RESULTS

230 We initially identified 1138 prostate cancer patients, but after implementing the exclusion  
231 criteria 1058 were eligible for inclusion in the study (Figure 1). During 4810 person-years of  
232 follow-up, we identified a total of 203 patients (19.2%) that died, thereof 98 patients (9.3%)  
233 that died due to prostate cancer. The median follow-up time was 4.6 years. Among eligible  
234 patients, 347 (32.8%) were identified as post-diagnosis PPI users; thereof 182 (52.4%) were  
235 continuous users and 165 (47.6%) new users. Among the 347 post-diagnosis PPI users we  
236 identified 59 patients (17.0%) that died from any cause and 22 patients (6.3%) that died from  
237 prostate cancer, compared with 144 patients (20.3%) and 76 patients (10.7%) among non-  
238 users, respectively. The median age among post-diagnosis PPI users was 69 years  
239 (interquartile range: 63 – 76) while it was 69 years (interquartile range: 62 – 75) among non-  
240 users. The majority of all patients were diagnosed with a localized disease; 81.6% among  
241 post-diagnosis PPI users and 77.2% among non-users. Compared with non-users, post-  
242 diagnosis PPI users had a higher median of medication-based comorbidity score (Table 1).  
243

244 In the main analysis, we observed adjusted HRs of 0.88 (95% CI: 0.52 – 1.48) for prostate  
245 cancer-specific mortality and 1.02 (95% CI: 0.73 – 1.43) for all-cause mortality among post-  
246 diagnosis PPI users as compared with non-users (Tables 2 and 3). In secondary analyses for  
247 prostate cancer-specific mortality (Table 2), we observed adjusted HRs of 0.45 (95% CI: 0.21  
248 – 0.98) among continuous PPI users and 1.12 (95% CI: 0.61 – 2.08) among new PPI users,  
249 when we stratified by timing of PPI use. Stratifying by clinical stage yielded adjusted HRs of  
250 0.50 (95% CI: 0.22 – 1.16) and 1.00 (95% CI: 0.44 – 2.27) among patients with localized and  
251 non-localized disease, respectively. For cumulative dose, we observed an adjusted HR for  
252 cumulative use of 1-365 DDDs of 0.91 (95% CI: 0.43 – 1.90) and 0.86 (95% CI: 0.45 – 1.61)  
253 for >365 DDDs. For all-cause mortality (Table 3), the adjusted HRs were 0.67 (95% CI:

254 0.43 – 1.04) and 1.25 (0.82 – 1.92) among continuous and new PPI users, respectively.  
255 Analyses stratified by clinical stage yielded an adjusted HR of 0.74 (95% CI: 0.47 – 1.15)  
256 among patients with localized disease and 1.18 (95% CI: 0.58 – 2.34) among patients with  
257 non-localized disease. For cumulative PPI use, we observed adjusted HRs of 1.19 (95% CI:  
258 0.76 – 1.87) and 0.91 (95% CI: 0.61 – 1.37) for patients using 1-365 DDDs and >365 DDDs,  
259 respectively.

260

261 Redefining post-diagnosis use as at least one filled prescription for a PPI drug yielded similar  
262 result as in the main analysis (Table S1). When we redefined the exposure opportunity  
263 window by assessing PPI use only in the 12 months following prostate cancer diagnosis, we  
264 observed HRs that were slightly lower, but mostly in line with those observed in the main  
265 analysis (Table S2). When post-diagnosis PPI use was assessed continuously throughout  
266 follow-up, we observed higher HRs than in the main analysis, but the estimates were not  
267 statistically significant (Table S3).

268

269 **DISCUSSION**

270 In this population-based cohort study among Icelandic prostate cancer patients, we did not  
271 observe a clear association between post-diagnosis PPI use and mortality among prostate  
272 cancer patients.

273

274 To our knowledge, this is only the second observational study to explore the association  
275 between PPI use and mortality among prostate cancer patients. Recently, post-diagnosis use  
276 of PPIs was reported to have led to increased mortality among cancer patients; both among  
277 cancer patients overall and among patients with certain site-specific cancers, including  
278 prostate cancer.[22] PPIs are commonly used among cancer patients,[30] often as a  
279 preventive measure against the risk of gastric damage following chemotherapy, radiotherapy,  
280 and steroid use.[31] Furthermore, PPI use has been shown to be associated with indicators of  
281 worse overall health[32] and among prostate cancer patients PPIs have been suggested to be  
282 related to decreased overall health.[33] However, our results were not consistent with the  
283 findings of Tvingsholm et al., in that we did not observe an increase in mortality among post-  
284 diagnosis PPI users.

285

286 Although the study by Tvingsholm et al., suggests that PPI use is associated with excess  
287 mortality among cancer patients, and that the association might be substance specific,  
288 previous clinical studies have reported that PPIs might enhance the effectiveness of  
289 chemotherapy.[18, 19] However, there have also been reports of unwanted drug interactions  
290 between PPIs and oral anticancer agents suggesting a negative impact of PPIs on  
291 chemotherapeutic efficacy.[31, 34] Unfortunately, we were unable to perform stratified  
292 analyses by chemotherapy or PPI substance in our study due to the small sample size leading  
293 to low numbers in stratified subgroups.

294 The study has several limitations that might have influenced our observations. First, clinical  
295 data on the underlying indications for PPI use was not available, leaving us unable to adjust  
296 for the potential of confounding by indication. Second, we did not have information on  
297 concomitant use of other drugs that might influence our estimates, e.g. statins which have  
298 been reported to be associated with decreased mortality among prostate cancer patients.[35,  
299 36] Third, we lacked information on clinical diagnoses to be able to adjust for disease-based  
300 comorbidities, although we made an attempt to counteract this limitation by using a  
301 medication-based comorbidity score as a proxy for the Charlson comorbidity index. Fourth,  
302 misclassification of PPI use might have resulted from OTC use and from use within the  
303 hospital setting, since we only had information on dispensed PPI drugs to the outpatient  
304 population. Fifth, we were unable to obtain information on the measured level of prostate  
305 specific antigen (PSA) at diagnosis; a variable that is used in clinical staging and could  
306 influence prognosis. Finally, as in all studies of this nature, our assessment of PPI use is based  
307 on dispensed drugs, which we cannot be sure are necessarily consumed. However, we tried to  
308 minimize the influence of this potential bias by the requirement of PPI users having received  
309 at least two filled prescriptions, in the main analysis. The primary strength of our study was  
310 the clearly defined population-based cohort and our utilization of high-quality nationwide  
311 registry data. Furthermore, utilization of registry data removed the risk of recall-bias.

312

313 In summary, our findings do not indicate that post-diagnosis PPI use influences mortality risk  
314 among prostate cancer patients. However, due to the small size of our cohort and short  
315 follow-up time, the resulting estimates had quite wide CIs, which limits our ability to draw  
316 any definitive conclusions. Future studies should use a larger cohort, longer follow-up time,  
317 and aim to minimize the potential impact of confounding by indication to further elucidate  
318 whether PPI use influences mortality among prostate cancer patients.



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326

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330

331

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## 444 TABLES

445

446 Table 1. Descriptive characteristics of a cohort of Icelandic prostate cancer patients diagnosed  
447 between 1 January 2007 and 31 December 2012 by post-diagnosis PPI user status.  
448

|   |                              | <b>Prostate Cancer</b>    |                         |
|---|------------------------------|---------------------------|-------------------------|
|   |                              | Proton pump inhibitor use |                         |
|   |                              | Non-<br>users             | Post-diagnosis<br>users |
|   |                              | N = 711                   | N = 347                 |
| Pre-diagnosis use (%)                         | Yes                          | 59 (8.3)                  | 182 (52.4)              |
|   | No                           | 652 (91.7)                | 165 (47.6)              |
| Age at diagnosis - years                      | Median (IQR)                 | 69 (62 -<br>75)           | 69 (63 - 76)            |
| Age groups (%)                                | 40-54                        | 47 (6.6)                  | 14 (4.0)                |
|   | 55-69                        | 338 (47.5)                | 169 (48.7)              |
|   | 70-85                        | 326 (45.9)                | 164 (47.3)              |
| Year of diagnosis (%)                         | 2007-2009                    | 349 (49.1)                | 208 (59.9)              |
|   | 2010-2012                    | 362 (50.9)                | 139 (40.1)              |
| Clinical stage                                | Localized                    | 549 (77.2)                | 283 (81.6)              |
|   | Non-localized                | 59 (8.3)                  | 22 (6.3)                |
|   | Unknown                      | 103 (14.5)                | 42 (12.1)               |
| Gleason score                                 | <7                           | 371 (52.2)                | 177 (51.0)              |
|   | 7                            | 195 (27.4)                | 103 (29.7)              |
|   | ≥8                           | 134 (18.8)                | 60 (17.3)               |
|   | Unknown                      | 11 (1.6)                  | 7 (2.0)                 |
| Radiotherapy (%) <sup>a</sup>                 | Yes                          | 196 (27.6)                | 92 (26.5)               |
|   | No                           | 515 (72.4)                | 255 (73.5)              |
| Cancer surgery (%) <sup>a</sup>               | Total excision of prostate   | 173 (24.3)                | 86 (24.8)               |
|   | Partial excision of prostate | 63 (8.9)                  | 44 (12.7)               |
|   | No surgery                   | 475 (66.8)                | 217 (62.5)              |
| Cancer drug treatment (%) <sup>a</sup>        | Yes                          | 62 (8.7)                  | 31 (8.9)                |
| Chemotherapy (%) <sup>a</sup>                 | Yes                          | 8 (1.1)                   | 1 (0.3)                 |
| Endocrine therapy (%) <sup>a</sup>            | Yes                          | 43 (6.0)                  | 24 (6.9)                |
| Chemotherapy & endocrine therapy <sup>a</sup> | Yes                          | 11 (1.5)                  | 6 (1.7)                 |
| Medication-based comorbidity                  | Median (IQR)                 | 5 (3 - 8)                 | 8 (5 - 10)              |

<sup>a</sup>Treatment in first year after diagnosis

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Table 2. Cox proportional hazard regression models for associations between post-diagnosis PPI use and prostate cancer-specific mortality among patients diagnosed with prostate cancer in Iceland between 2007 and 2012

| PPI exposure           | Prostate cancer-specific mortality |                    |                                       |                                   |
|------------------------|------------------------------------|--------------------|---------------------------------------|-----------------------------------|
|                        | No of deaths                       | No of person years | Age adjusted HR (95% CI) <sup>b</sup> | Adjusted HR (95% CI) <sup>c</sup> |
| Non-use                | 76                                 | 3640               | 1.00 (Reference)                      | 1.00 (Reference)                  |
| Post-diagnosis PPI use | 22                                 | 1171               | 0.85 (0.52 - 1.38)                    | 0.88 (0.52 - 1.48)                |
| <b>Timing of use</b>   |                                    |                    |                                       |                                   |
| Continuous PPI use     | 8                                  | 734                | 0.45 (0.22 - 0.93)                    | 0.45 (0.21 - 0.98)                |
| New PPI use            | 14                                 | 437                | 1.39 (0.77 - 2.53)                    | 1.12 (0.61 - 2.08)                |
| <b>Clinical stage</b>  |                                    |                    |                                       |                                   |
| Localized              | 8                                  | 1006               | 0.55 (0.25 - 1.23)                    | 0.50 (0.22 - 1.16)                |
| Non-localized          | 9                                  | 39                 | 0.92 (0.43 - 1.96)                    | 1.00 (0.44 - 2.27)                |
| <b>Cumulative dose</b> |                                    |                    |                                       |                                   |
| 1-365 DDDs             | 9                                  | 390                | 1.04 (0.52 - 2.09)                    | 0.91 (0.43 - 1.90)                |
| >365 DDDs              | 13                                 | 780                | 0.75 (0.41 - 1.37)                    | 0.86 (0.45 - 1.61)                |

Abbreviations: DDD, defined daily doses; HR, hazard ratio; CI, confidence interval

<sup>b</sup>Adjusted for age at diagnosis

<sup>c</sup>Adjusted for age at diagnosis, calendar period, clinical stage, Gleason score, medication-based comorbidity, surgery, endocrine and/or chemotherapy, radiotherapy

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461 Table 3. Cox proportional hazard regression models for associations between post-diagnosis PPI use and all-cause mortality among patients diagnosed with  
 462 prostate cancer in Iceland between 2007 and 2012  
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| PPI exposure           | All-cause mortality |                    |                                       |                                   |
|------------------------|---------------------|--------------------|---------------------------------------|-----------------------------------|
|                        | No of deaths        | No of person years | Age-adjusted HR (95% CI) <sup>b</sup> | Adjusted HR (95% CI) <sup>c</sup> |
| Non-use                | 144                 | 3640               | 1.00 (Reference)                      | 1.00 (Reference)                  |
| Post-diagnosis PPI use | 59                  | 1171               | 1.16 (0.85 - 1.59)                    | 1.02 (0.73 - 1.43)                |
| <b>Timing of use</b>   |                     |                    |                                       |                                   |
| Continuous PPI use     | 28                  | 734                | 0.81 (0.54 - 1.22)                    | 0.67 (0.43 - 1.04)                |
| New PPI use            | 31                  | 437                | 1.57 (1.04 - 2.36)                    | 1.25 (0.82 - 1.92)                |
| <b>Clinical stage</b>  |                     |                    |                                       |                                   |
| Localized              | 33                  | 1006               | 0.99 (0.65 - 1.50)                    | 0.74 (0.47 - 1.15)                |
| Non-localized          | 13                  | 39                 | 1.08 (0.57 - 2.06)                    | 1.18 (0.58 - 2.34)                |
| <b>Cumulative dose</b> |                     |                    |                                       |                                   |
| 1-365 DDDs             | 27                  | 390                | 1.61 (1.06 - 2.44)                    | 1.19 (0.76 - 1.87)                |
| >365 DDDs              | 32                  | 780                | 0.93 (0.63 - 1.38)                    | 0.91 (0.61 - 1.37)                |

Abbreviations: DDD, defined daily doses; HR, hazard ratio; CI, confidence interval

<sup>b</sup>Adjusted for age at diagnosis

<sup>c</sup>Adjusted for age at diagnosis, calendar period, clinical stage, Gleason score, medication-based comorbidity, surgery, endocrine and/or chemotherapy, radiotherapy

465 FIGURES

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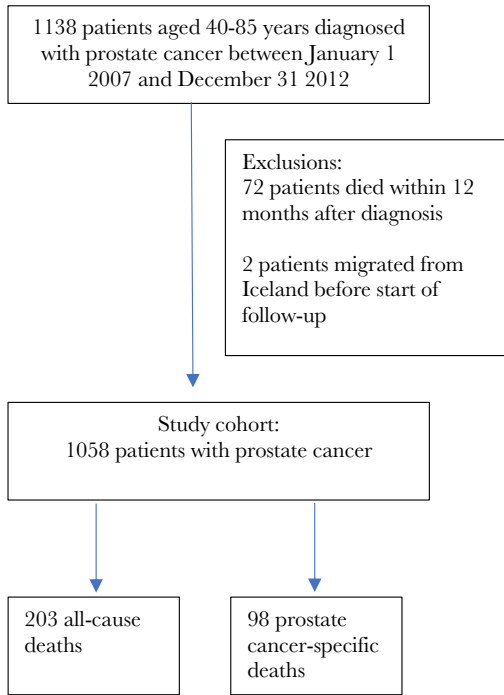
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479 Figure 1. Study flow chart of cohort identification

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483 SUPPLEMENTARY TABLES

484 Table S1. Sensitivity analysis of the association between time-varying post-diagnosis PPI use, prostate cancer-specific mortality, and all-cause mortality among  
 485 prostate cancer patients in Iceland, where post-diagnosis use was defined as at least one dispensed PPI prescription after diagnosis and post-diagnosis users  
 486 were considered unexposed up until the received their first PPI dispensing and then exposed thereafter  
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| PPI exposure                              | No of deaths | No of person years | Age adjusted HR (95% CI) <sup>b</sup> | Adjusted HR (95% CI) <sup>c</sup> |
|---|--------------|--------------------|---------------------------------------|-----------------------------------|
| <b>Prostate cancer-specific mortality</b> |              |                    |                                       |                                   |
| Non-use                                   | 65           | 3203               | 1.00 (Reference)                      | 1.00 (Reference)                  |
| Post-diagnosis PPI use                    | 33           | 1608               | 0.97 (0.63 - 1.50)                    | 1.04 (0.66 - 1.65)                |
| <b>All-cause mortality</b>                |              |                    |                                       |                                   |
| Non-use                                   | 129          | 3203               | 1.00 (Reference)                      | 1.00 (Reference)                  |
| Post-diagnosis PPI use                    | 74           | 1608               | 1.07 (0.79 - 1.43)                    | 0.98 (0.71 - 1.34)                |

Abbreviations: HR, hazard ratio; CI, confidence interval

<sup>b</sup>Adjusted for age at diagnosis

<sup>c</sup>Adjusted for age at diagnosis, calendar period, clinical stage, Gleason score, medication-based comorbidity, surgery, endocrine and/or chemotherapy, radiotherapy

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Table S2. Sensitivity analysis of the association between post-diagnosis PPI use, prostate cancer-specific mortality, and all-cause mortality among prostate cancer patients in Iceland, where post-diagnosis use of PPIs was assessed in the first 12 months following prostate cancer diagnosis.

| PPI exposure                              | No of deaths | No of person years | Age adjusted HR (95% CI) <sup>b</sup> | Adjusted HR (95% CI) <sup>c</sup> |
|---|--------------|--------------------|---------------------------------------|-----------------------------------|
| <b>Prostate cancer-specific mortality</b> |              |                    |                                       |                                   |
| Non-use                                   | 85           | 3930               | 1.00 (Reference)                      | 1.00 (Reference)                  |
| Post-diagnosis PPI use                    | 13           | 881                | 0.64 (0.36 - 1.15)                    | 0.63 (0.34 - 1.18)                |
| <b>All-cause mortality</b>                |              |                    |                                       |                                   |
| Non-use                                   | 163          | 3930               | 1.00 (Reference)                      | 1.00 (Reference)                  |
| Post-diagnosis PPI use                    | 40           | 881                | 1.03 (0.73 - 1.46)                    | 0.84 (0.58 - 1.23)                |

Abbreviations: HR, hazard ratio; CI, confidence interval

<sup>b</sup>Adjusted for age at diagnosis

<sup>c</sup>Adjusted for age at diagnosis, calendar period, clinical stage, Gleason score, medication-based comorbidity, surgery, endocrine and/or chemotherapy, radiotherapy

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513 Table S3. Sensitivity analysis of the association between time-varying post-diagnosis PPI use, prostate cancer-specific mortality, and all-cause mortality among  
 514 prostate cancer patients in Iceland, where post-diagnosis use was assessed continuously throughout follow-up.  
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| PPI exposure                              | No of deaths | No of person years | Age adjusted HR (95% CI) <sup>b</sup> | Adjusted HR (95% CI) <sup>c</sup> |
|---|--------------|--------------------|---------------------------------------|-----------------------------------|
| <b>Prostate cancer-specific mortality</b> |              |                    |                                       |                                   |
| Non-use                                   | 77           | 4012               | 1.00 (Reference)                      | 1.00 (Reference)                  |
| Post-diagnosis PPI use                    | 21           | 798                | 1.33 (0.82 - 2.19)                    | 1.48 (0.88 - 2.47)                |
| <b>All-cause mortality</b>                |              |                    |                                       |                                   |
| Non-use                                   | 158          | 4012               | 1.00 (Reference)                      | 1.00 (Reference)                  |
| Post-diagnosis PPI use                    | 45           | 798                | 1.33 (0.95 - 1.87)                    | 1.30 (0.92 - 1.85)                |

Abbreviations: HR, hazard ratio; CI, confidence interval

<sup>b</sup>Adjusted for age at diagnosis

<sup>c</sup>Adjusted for age at diagnosis, calendar period, clinical stage, Gleason score, medication-based comorbidity, surgery, endocrine and/or chemotherapy, radiotherapy

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# Appendix

A.



VÍSINDASIÐANEFND

Hafnarhúsið, Tryggvagata 17  
101 Reykjavík,

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netfang: [vsni@vsni.is](mailto:vsni@vsni.is) [www.vsn.is](http://www.vsn.is)

Háskóli Íslands, Læknadeild, lýðheilsuvísindi  
Helga Zoega, dósent í lýðheilsuvísindum  
Vatnsmýrarvegur 16  
101 Reykjavík

Reykjavík 27. október 2015  
Tilv.: VSNb2015080004/03.03

Efni: Varðar: 15-115-afg -Faraldsfræðileg rannsókn á áhrifum lyfja úr flokki prótonpumpuhemla (PPI) á krabbamein í mönnum.

Umsókn þinni til Vísindasiðanefndar hefur verið gefið númerið **VSN-15-115**. Við förum vinsamlegast fram á að það númer verði notað í samskiptum vegna þessarar umsóknar.

Á fundi sínum 27.10.2015 fjallaði Vísindasiðanefnd um umsókn þína vegna ofangreindrar rannsóknaráætlunar. Meðrannsakendur þínir eru: Eiríkur Steingrímsson, Óskar Örn Hálfátánarson, Helga M. Ögmundsdóttir og Margrét Helga Ögmundsdóttir.

Eftir að hafa farið vandlega yfir umsókn þína og innsend gögn gerir Vísindasiðanefnd ekki athugasemdir við framkvæmd rannsóknarinnar. Með vísan til 1. mgr. 27. gr. laga nr. 44/2014, heimilar Vísindasiðanefnd aðgang að þeim upplýsingum sem fram koma í kafla B-2 í umsókn til nefndarinnar úr Lyfjagagnagrunni landleknis, Krabbameinsskrá og Þjóðskrá. Með vísan til laga nr. 44/2014, um vísindarannsóknir á heilbrigðisviði, er rannsóknaráætlunin endanlega samþykkt með þeim almenna fyrirvara að lögbindið samþykki skráarhaldara skv. 2. mgr. 27. gr. laga nr. 44/2014 verður að liggja fyrir áður en vinna með heilbrigðisgögn viðkomandi stofnunar/skráarhaldara hefst.

Vísindasiðanefnd áréttar að ábyrgðarmaður rannsóknarinnar ber ábyrgð á að sött sé um viðeigandi leyfi fyrir rannsókninni hjá þeim stofnunum sem við á. Óheimilt er að hefja framkvæmd rannsóknarinnar fyrr en þau liggja fyrir. Afrit leyfa/samstarfsyfirlýsinga þurfa að berast nefndinni. Áréttað er að allar fyrirhugaðar breytingar á þegar samþykktri rannsóknaráætlun þurfa að koma inn til nefndarinnar til umfjöllunar. Jafnframt ber ábyrgðarmanni að sækja um breytingar til þeirra stofnanna, sem veitt hafa leyfi vegna framkvæmdar rannsóknarinnar eða öflunar gagna, um framangreint, ef við á.

Vísindasiðanefnd bendir rannsakendum vinsamlegast á að birta VSN tilvísunarnúmer rannsóknarinnar þar sem vitnað er í leyfi nefndarinnar í birtum greinum um rannsóknina. Minnt er á að tilkynna rannsóknarlok til nefndarinnar.

Með kveðju og ósk um gott rannsóknargengi,  
f.h. Vísindasiðanefndar,

Kristján Eilendsson, lækni, formaður



## B.

Helga Zoëga, prófessor  
Háskóli Íslands  
101 Reykjavík



VÍSINDASÍÐANEFND  
Borgartúni 21 - 4. hæð  
105 Reykjavík.

Sími: 551 7100, Bréfsími: 551 1444  
netfang: vsn@vsn.is www.vsn.is

Reykjavík 6. september 2016  
Tílv.: VSNb201608001/03.01

Efni: Varðar: 16-124 - Prótonpumpuhemlar og möguleg áhrif þeirra á framþróun krabbameina og lífun krabbameinsjúklinga

Umsókn þinni til Vísindasíðanefndar hefur verið gefið númerið **VSN-16-124**. Við förum vinsamlegast fram á að það númer verði notað í samskiptum vegna þessarar umsóknar.

Á fundi sínum 06.09.2016 fjallaði Vísindasíðanefnd um umsókn þína vegna ofangreindrar rannsóknarætlunar. Meðrannsakendur þínir eru: Óskar Örn Hálfðánarson, doktorsnemi, Eiríkur Steingrímsson, prófessor, Helga M. Ögmundsdóttir, prófessor og Sigrún Helga Lund, dósent.

Samkvæmt lið B-1 í umsókn þinni inniheldur rannsóknarúrtakið alla einstaklinga á Íslandi sem greinast með krabbamein á tímabilinu frá 1. janúar 2003 til og með 31. desember 2012.

Þá kemur fram í fylgiskjali 1 í lið B-2 í umsókn að:

„*Unnt verður með fyrirbyggjandi upplýsingar úr Krabbameinsskrá, Lyfjagagnrunni Embættis landlæknis, Dánarheimaskrá og frá Landspítala - Háskólasjúkrahúsi (LSH).*

#### **Krabbameinsskrá**

*Krabbameinsgreiningar frá og með 1. janúar 2003 til 31. desember 2012. Þær breytur sem sótt verður um að fá aðgang að úr skránni eru eftirfarandi: aldur, kyn, dagsetning greiningar, greiningarár, aldur við greiningu, númer æxlis, meingerð æxlis (vefjagerð og tegund æxlis), dánardagur og ICD-10 kóði meins. Með því að byggja á ICD-10 kóðum World Health Organization (WHO) munum við flokka tegundir krabbameins á eftirfarandi hátt: Olli krabbamein (C00-C96), sortuæxli (C43), brjóstakrabbamein (C50) og blóðruhálskirtilskrabbamein (C61).*

#### **Lyfjagagnrunnur**

*Útleystar lyfjaávisanir fyrir prótonpumpuhemlum (PPI) frá 1. janúar 2003 til og með 31. desember 2015. Unnt verður með upplýsingar um aldur (fæðingarár og fæðingarmánuður), kyn, útleystar lyfjaávisanir fyrir PPI lyf, heiti lyfs, fjöldi útleystra DDD á lyfjaávisu, ATC-kóði lyfs, dagsetning lyfjaúttektar, afgreiðslustaður, sérgrein læknis sem skrifar upp á lyfið og skráð lögheimili þann 1. janúar 2003 (flokkað eftir landshlutaskiptingu Hagstofunnar).*

*Óskað er eftir því að Embætti landlæknis bæti einnig við upplýsingum um brotflutning frá Íslandi (dagsetning) og innflutning til Íslands (dagsetning) úr Þjóðskrá Íslands við rannsóknargagnrunninn.*

#### **Dánarheimaskrá**

*Upplýsingar um dánardag og dánarheim einstaklinga í rannsóknarhóp sem láta lífið á rannsóknartímabilinu verða sóttar í Dánarheimaskrá. Flokkun dánarheima verður byggð á alþjóðlegri tölfraediflokkun sjúkdóma og skyldra heilbrigðisvandamála (ICD-10).*

#### **Landspítali - Háskólasjúkrahús**

Við munum nálgast upplýsingar frá Landspítala-Háskólasjúkrahúsi (LSH) um framþróun meinanna hjá einstaklingum í rannsóknarhóp. Þær breytur sem sóttar verða til LSH eru eftirfarandi:

Skurðaðgerð vegna krabbameins (Vöruhús gagna):

- ICD-10 kóði meins (tegund meins)
- NCSF-IS kóði aðgerðar (tegund aðgerðar)
- Dagssetning aðgerðar.

Geislameðferð vegna krabbameins (Geisladeild LSH):

- ICD-10 kóði meins (tegund meins)
- Heildargeislaskammtur
- Daglegur geislaskammtur
- Meðferð hafin, dags.
- Meðferð hætt, dags.

Prótonpumpuheimlar og möguleg áhrif þeirra á framþróun krabbameina og lífun krabbameinssjúklinga.

Lyfjameðferð vegna krabbameins (Vöruhús gagna & Aria MedOnc

lyfjaskráningarkerfið):

- ICD-10 kóði meins (tegund meins)
- Heiti lyfs
- Lyfjaskammtur
- Meðferð hafin, dags.
- Meðferð hætt, dags.

Svörin meðferða vegna krabbameins (Uppfletting í sjúkraskýrslum):

- No response, partial response eða full response."

Samkvæmt umsókn er áætlað að sækja gögn úr Krabbameinsskrá, Lyfjagagnagrunni Embættis landlæknis, Dánameinaskrá, Vöruhúsi gagna LSH, Geisladeild LSH, Aria MedOnc lyfjaskráningarkerfi LSH og úr sjúkraskýrslum.

Áætluð rannsóknarlök eru í september 2018. Samkvæmt umsókn er frambúðarvarðveisla skv. 7. gr. laga nr. 44/2014 ekki fyrirhuguð. Að því virtu ber að eyða gögnum sem aflað er til rannsóknarinnar eða verða til við framkvæmd hennar fyrir árslok 2023.

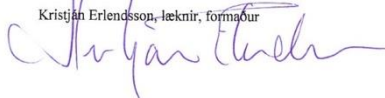
Með vísan til 1. mgr. 27. gr. laga nr. 44/2014 heimilar Vísindasiðanefnd aðgang að framangreindum upplýsingum úr ofangreindum skráum. Með vísan til 1. mgr. 12. gr. laga nr. 44/2014 er rannsóknaráætlun endanlega samþykkt með þeim almenna fyrirvara að lögbundið samþykki skráarhaldara skv. 2. mgr. 27. gr. laganna verður að liggja fyrir áður en vinna með heilbrigðisgögn viðkomandi stofnunar/skráarhaldara hefst.

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Vísindasiðanefnd bendir rannsakendum vinsamlegast á að birta VSN tilvísunarnúmer rannsóknarinnar þar sem vitnað er í leyfi nefndarinnar í birtum greinum um rannsóknina. Minnt er á að tilkynna rannsóknarlök til nefndarinnar.

Með kveðju og ósk um gott rannsóknargengi,  
f.h. Vísindasiðanefndar,

Kristján Erlendsson, læknir, formaður

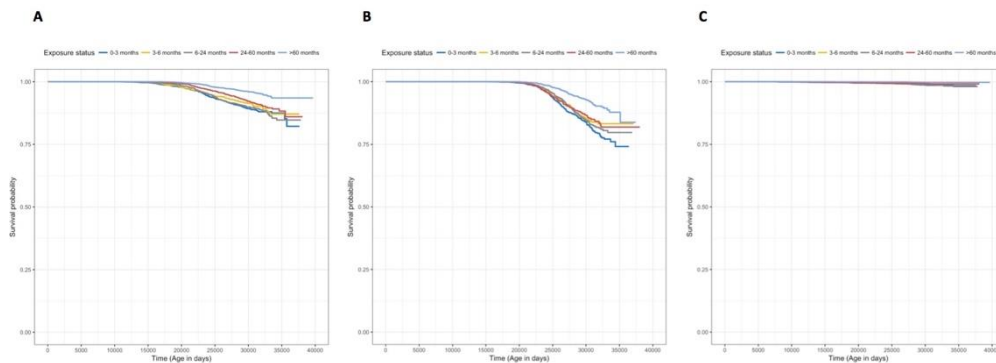


### C.

Using a cohort study design, we conducted analyses on proton pump inhibitor use among patients diagnosed with breast cancer, prostate cancer, and malignant melanoma between 2003 and 2014. We assessed various levels of PPI exposure, i.e. ever use (defined as one filled PPI prescription during follow-up) and cumulative use (in months, assuming the intake of one DDD per day). The exposure was not assessed in a time-dependent manner. We used Cox proportional hazard regression models, with age as the underlying time-scale, to estimate HRs and 95% CIs. The results displayed in the table and the figure are influenced by immortal time bias.

| Subgroups                     | Breast Cancer |                          |             | Prostate Cancer |                          |             | Malignant melanoma |                          |             |
|-------------------------------|---------------|--------------------------|-------------|-----------------|--------------------------|-------------|--------------------|--------------------------|-------------|
|                               | Events        | Adjusted HR <sup>†</sup> | 95% CI      | Events          | Adjusted HR <sup>†</sup> | 95% CI      | Events             | Adjusted HR <sup>†</sup> | 95% CI      |
| Never use of PPI              | 832           | 1.00 (ref)               | 1.00 (ref)  | 984             | 1.00 (ref)               | 1.00 (ref)  | 272                | 1.00 (ref)               | 1.00 (ref)  |
| Ever use of PPI               | 720           | 0.83                     | 0.75 - 0.92 | 779             | 0.85                     | 0.77 - 0.93 | 127                | 0.65                     | 0.52 - 0.80 |
| <b>Cumulative dose (DDDs)</b> |               |                          |             |                 |                          |             |                    |                          |             |
| 0 - 3 months                  | 219           | 1.05                     | 0.90 - 1.22 | 252             | 1.11                     | 0.96 - 1.27 | 52                 | 0.88                     | 0.65 - 1.18 |
| 3 - 6 months                  | 91            | 0.92                     | 0.74 - 1.14 | 96              | 0.89                     | 0.73 - 1.10 | 14                 | 0.55                     | 0.32 - 0.95 |
| 6 - 24 months                 | 184           | 1.00                     | 0.86 - 1.18 | 180             | 0.94                     | 0.80 - 1.10 | 36                 | 0.87                     | 0.61 - 1.24 |
| 24 - 60 months                | 128           | 0.80                     | 0.67 - 0.97 | 144             | 0.86                     | 0.73 - 1.03 | 17                 | 0.53                     | 0.32 - 0.87 |
| >60 months                    | 98            | 0.44                     | 0.35 - 0.54 | 107             | 0.47                     | 0.39 - 0.58 | 8                  | 0.20                     | 0.10 - 0.40 |

<sup>†</sup>Adjusted for age, sex, and prior NSAID use



A, breast cancer; B, prostate cancer; C, malignant melanoma





