



Original article

Sexual mixing in opposite-sex partnerships in Britain and its implications for STI risk: findings from the third National Survey of Sexual Attitudes and Lifestyles (Natsal-3)

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Editorial decision 24 September 2018; Accepted 9 October 2018

Abstract

Background: The extent to which individuals are similar to their sexual partners influences STI-transmission probabilities, yet there is a dearth of empirical data, especially those representative of the population.

Methods: Analyses of data reported by 10 759 sexually active people aged 16–74 y interviewed for a British national probability survey undertaken in 2010–12. Computer-assisted self-interviews asked about partner numbers and characteristics of participants' three most recent partnerships (MRPs). Opposite-sex MRPs were weighted to represent *all* such partnerships in the past year ($N = 16\,451$). Estimates of disassortative age mixing ($\geq \pm 5$ -y difference), ethnic mixing (partner of a different ethnic group) and geographical mixing (partner from a different region/country when they first met) were calculated, stratified by gender, age group and partnership status (casual/steady). Multivariable regression examined how these disassortative mixing measures were associated with STI-risk measures: condom use at first sex together at the partnership-level and, at the participant-level, STI-risk perception and reporting STI diagnoses.

Results: Disassortative age mixing occurred in around one-third of opposite-sex partnerships, with men ≥ 5 y older in most cases, although this proportion varied by participant's gender and age group. Ethnic mixing occurred less frequently (11.3% of men's and 8.6% of women's partnerships) as did geographical mixing (14.1 and 16.3%, respectively). Disassortative mixing was more common among casual vs steady partnerships. Condom use at first sex was less likely in women's partnerships that were age-disassortative [adjusted odds ratio (AOR): 0.79, 95% confidence interval (CI): 0.69–0.95], whereas men reporting disassortative ethnic mixing were more likely to perceive themselves at STI risk (AOR: 1.76, 95% CI: 1.23–2.52) and report STI diagnoses (AOR: 2.37, 95% CI: 1.22–4.59).

Conclusions: Disassortative mixing, although uncommon among opposite-sex partnerships in Britain, is independently associated with STI risk, warranting consideration in STI-prevention efforts.

Key words: sexual behaviour, heterosexual, sexual mixing, assortative mixing, disassortative mixing, sexual partners, sexually transmitted infection, survey, population

Key Messages

- Understanding STI risk requires knowledge not only of an individual's socio-demographic and behavioural characteristics, but also those of their partner(s) and the extent to which these characteristics are similar ('assortative mixing') or not ('disassortative mixing').
- Unlike previous studies that have tended to focus on the most recent partner, thus underestimating casual partnerships, we weighted national probability survey data on individuals' three most recent partners to represent all opposite-sex partnerships in the past year.
- In most opposite-sex partnerships in Britain, individuals were similar to their partners as *disassortative* age mixing ($\geq \pm 5$ -y age difference) occurred in approximately one-third of partnerships, whereas *disassortative* ethnic and geographical mixing were less common (fewer than one in six partnerships).
- Disassortative sexual mixing was associated—albeit weakly—with STI-risk indicators, including after controlling for the number and type of sexual partners.
- STI-prevention efforts may benefit from taking account of sexual mixing patterns as well as individual-level risk factors.

Background

Sexually transmitted infection (STI) prevalence varies within populations according to individuals' sexual behaviour, such as the number and timing of sexual partners, the type(s) of sexual practice engaged in, the extent to which condoms are used, as well as socio-demographic characteristics such as gender, age and ethnicity, and access to and use of health-care services.^{1–3} As such, STI-transmission probabilities vary when sex occurs between individuals from population groups with different STI prevalences, e.g. between commercial sex workers and their clients,^{4,5} which in turn influences the rate at which STIs are spread at a population level.⁶ Understanding the extent to which individuals are similar to their partners is therefore important for improving our understanding of STI transmission, including in parameterizing mathematical models that can inform the design and evaluation of STI-control interventions,⁶ such as England's National Chlamydia Screening Programme.⁷ Where the degree of assortative mixing (i.e. individuals tending to have sexual partners with characteristics and engaging in behaviours similar to themselves⁸) is high, STI-prevention efforts may be more effective if they focus on those at highest risk. In contrast, where mixing is largely disassortative (i.e. individuals tending to have sexual partners with different characteristics and behaviours to themselves⁸), then more generalized approaches may be

more appropriate, e.g. population screening. Between these two extremes lies mixing at random, i.e. people having sex with partners at random such that there is no pattern in how their own characteristics and behaviours relate to those of their partners.

Studies that have sought to empirically investigate sexual mixing using data representative of the population are rare,^{9,10} with analyses of convenience survey data more common, often collected from sexual health clinic attendees,^{11–13} who are known to report greater STI-risk behaviour than observed in the general population.² One exception analysed data from the Health Survey for England (HSE), a national probability survey of all individuals in randomly selected households.¹⁴ In its 2010 round, data on sexual behaviour were collected for the first time, permitting analyses of the extent of sexual mixing among 943 heterosexual cohabiting couples.¹⁵ A positive correlation between individuals and their partners' characteristics, health behaviours, and sexual histories examined (current mental illness), with 12 of the 17 considered more than moderately assortative, suggesting high levels of assortativity in cohabiting partnerships.¹⁵ Whereas the vast majority of the adult population in Britain lives with a partner,¹⁶ non-cohabiting partnerships—who are not captured by the HSE—correspond to a large proportion of all

partnerships.¹⁷ This is an even greater issue for young people, who bear the greatest burden of STIs,^{1,2} the majority of whom do not live with their sexual partners.¹⁶ Furthermore, there is evidence that non-cohabiting partnerships are more likely to be disassortative than cohabiting partnerships,¹⁸ and their often shorter duration¹⁹ enables higher rates of partner change, meaning that they are important for STI transmission within the population.

Using a previously published method,²⁰ this paper aims to build upon earlier analyses¹⁷ as, in addition to describing the extent of age mixing, it describes the extent of ethnic and geographical mixing among all opposite-sex partnerships in the past year in Britain using data from the third National Survey of Sexual Attitudes and Lifestyles (Natsal-3). It also seeks to examine how sexual mixing relates to STI risk, specifically, condom use, STI-risk perception and history of STI diagnosis/es, and whether any associations remain after adjusting for the confounding effects of the number and types of partners participants reported.

Methods

Participants and procedures

Full details of the methods used in Natsal-3 have been reported elsewhere.^{16,21} Briefly, Natsal-3 involved a multi-stage, clustered, stratified probability sample design and interviewed 15 162 men and women aged 16–74 y, resident in households in Britain between September 2010 and August 2012. The response rate was 57.7%.^{16,21} One randomly selected person per household was invited to participate in a face-to-face computer-assisted personal interview (CAPI). More sensitive questions were asked in a computer-assisted self-interview (CASI), including those about participants' sexual partners (defined as 'people who have had sex together—whether just once, or a few times, or as regular partners, or as married partners'; having sex together was defined as 'vaginal, oral and anal sexual intercourse'). Detailed questions about participants' three (where applicable) most recent partners (MRPs) in the past 5 y were also asked in the CASI. This paper focuses on sexual mixing within opposite-sex partnerships; sexual mixing within men's same-sex partnerships is the focus of a separate paper.

Measures

The age difference between a participant and their partner was calculated as the male's age minus the female's age. Age differences –5 years or +5 years are referred hereon as disassortative age mixing, as defined in previous studies.^{17,22} Ethnic group was categorized as White, Asian/Asian British, Black/Black British or other, and disassortative ethnic

mixing was defined as the participant and partner being from different groups. However, because of our sample size and the relatively low prevalence of people of non-White ethnicity in the British population,²³ we do not examine the effect of mixing between particular ethnic groups. Disassortative geographical mixing was defined as the partner living in a different region within the same country or another country when they first met.

Statistical analysis

We initially restricted analysis for this paper to the 11 340/15 162 (74.8%) of Natsal-3 participants who reported having had at least one opposite-sex partner in the past year to provide a contemporary picture of sexual mixing (4749 men, 6591 women). Because of the paper's use of data collected by the MRP module, participants were also required to have reported data on at least one opposite-sex MRP with whom sex had occurred in the past year. This reduced the sample size from 11 340 to 10 759 (94.9% of eligible participants; 4490 men, 6269 women) who are described in [Supplementary Table 1](#), available as [Supplementary data](#) at *IJE* online.

As partner numbers decrease the more recent the timeframe,¹⁶ focusing on the past year rather than the past 5 y enabled us to increase the proportion of participants with MRP data available for all the partners they reported having in response to a question in the CASI on partner numbers during this time [question 'Het1Yr': 'Altogether, in the last year, how many (women/men) have you had sexual intercourse with?' where 'women' was asked whether the participant was male and vice versa]. 382/4490 (8.5%) eligible men and 334/6269 (5.3%) eligible women reported more than three opposite-sex partners in the past year¹⁶ and so the MRP questions were not asked for all their partners. Using the total number of partners reported by all participants in response to 'Het1Yr' described above ($N=16\,451$), we estimated that the MRP questions were not asked of 19% of all partnerships in the past year (24% of all men's and 14% of all women's partnerships in the past year). These unreported partnerships are typically shorter in duration and less likely to be ongoing compared with the most recent¹⁹ and so, to minimize this potential bias, we weighted participants' ≤ 3 MRPs to represent fourth and higher-order partners if they reported more than three opposite-sex partners using a previously published method.²⁰ Briefly, MRPs that were assumed to have ended in the past year were weighted to represent participants' partnerships that were not captured by the MRP questions, i.e. partnerships that were less recent than the three MRPs. For example, if a participant reported that they had four partners in the year prior to interview, then

data on their three MRPs were obtained, including whether each MRP had ended or was ongoing. If two MRPs were ongoing while the third MRP had ended, then the two MRPs were each weighted as one to represent just those partnerships, whereas the third MRP would be assigned a weight of two so that it represented itself and the fourth partner for which no detailed data were collected. If, however, all three MRPs had ended, then they would be weighted equally to represent the fourth partner, i.e. each having a weight of $4/3 = 1.333$. These partnership-level weights were applied in addition to participant-level weights used to account for unequal selection probabilities and differential non-response.²¹ Doing so results in weighting the 13 824 opposite-sex partners reported in the MRP module to represent the total 16 451 partnerships reported at Het1Yr. Given the study's large sample size, $p < 0.01$ is considered as evidence of an association, whereas $0.01 < p < 0.05$ is considered as weak (or weaker) evidence of an association.

We used Stata (version 14) complex survey analysis functions to incorporate the weighting, stratification of the data²⁴ and geographical clustering of participants.²⁵ After describing the sample characteristics (Supplementary Table 1, available as Supplementary data at *IJE* online), we used boxplots to show the distribution of age differences between partners by gender and age group, which highlight the medians and inter-quartile ranges (IQRs). We then estimated the prevalence of disassortative age, ethnic and geographical mixing, defined above, by gender, age group and partnership type (Supplementary Table 2, available as Supplementary data at *IJE* online, shows how partnership type was categorized). We used the Pearson chi-square statistic (adapted for complex surveys) to determine whether disassortative sexual mixing varied according to these key socio-demographic variables.

We used multivariable logistic regression to examine how each of the three measures of disassortative sexual mixing was associated with three indicators of STI risk at the partnership-level and the participant-level—outcome 1: condom use at first sex in a partnership (yes/no) as a direct measure of risk behaviour within the partnership; outcome 2: participant's perception of their STI risk (at risk or not; asked face to face using showcards); and outcome 3: whether the participant reported STI diagnosis/es in the past year (asked in the CASI). In the participant-level models (outcomes 2 and 3), disassortative sexual mixing was captured as the participant reporting disassortative age, ethnic or geographical mixing in any of their ≤ 3 MRPs, i.e. as a marker of being exposed to disassortative sexual mixing in the past year. Potential confounders in the participant-level models were participant's current age, whether any of their ≤ 3 MRPs were casual, the total

number of opposite-sex partners reported (past year, i.e. Het1Yr) and, for the models looking at disassortative ethnic mixing as a hypothesized explanatory variable, participant's ethnicity. In contrast, in the partnership-level models, participant's age at first sex with the partner was considered as a confounder.¹⁷ Supplementary Table 3, available as Supplementary data at *IJE* online, gives the adjusted odds ratios (AORs) for all variables included in each model. In the partnership-level models, we additionally tested for effect modification by including a term corresponding to the interaction between age and the age mixing indicator. Participant-level models were limited to those 7880 participants aged 16–44 y and partnership-level models were limited to those opposite-sex partners (10 651 unweighted, 11 206 weighted) reported by the 7880 participants aged 16–44 y as STI risk is considerably lower in older people,¹ although, for completeness, modelling results for those aged 16–74 y are presented in Supplementary Table 4, available as Supplementary data at *IJE* online.

Ethical approval

The Natsal-3 study was approved by the Oxfordshire Research Ethics Committee A (reference: 09/H0604/27).

Results

Sample characteristics

The sample was evenly distributed by age for men and women (Supplementary Table 1, available as Supplementary data at *IJE* online). Among those aged 16–74 y, only a minority—approximately 1 in 10—were of non-White ethnicity. The distribution of opposite-sex partner numbers was highly skewed, with men reporting larger numbers of partners on average than women (18.5% of men reported more than one partner vs 12.5% of women). Men were also more likely to report that at least one of their three MRPs was casual (18.3 vs 11.9% of women).

Disassortative age mixing

The median age difference between opposite-sex partners varied slightly by gender: 2 y (IQR: 0, 5 y) in men's partnerships and 1 y (–1, 4 y) in women's partnerships, and also with age, especially among men for whom both the median and IQR increased with age (Figure 1). Around one-third of all men's and women's partnerships involved an age difference of $\geq \pm 5$ y, considered as denoting disassortative age mixing (Table 1). Whereas the proportion of men's partnerships with disassortative age mixing was low (7.9%) at the youngest ages (16–24 y), this proportion

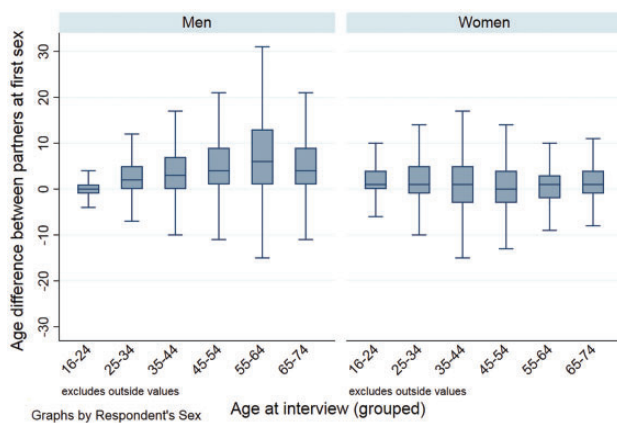


Figure 1. Boxplot of age differences between partners by participant's gender and age group. Denominator: opposite-sex partners in the past year reported by Natsal-3 participants aged 16–74 y who reported 1+ opposite-sex partners in the past year, data on ≥ 1 opposite-sex MRPs in the MRP module, and the age of their MRP(s). Unweighted N=13 604 partners (5 867 men's partners; 7 737 women's partners); Weighted N=16 201 partners (8 816 men's partners; 7385 women's partners).

increased to around a half for those age 35 and above. A higher proportion (19.4%) of women's partnerships than men's at age 16–24 y involved disassortative age mixing. Although this proportion also increased with women's age, it 'plateaued' at a lower level (between 30 and 40%) compared with men. The man was ≥ 5 y older than the woman in the majority of partnerships with disassortative age mixing (84.6% of men's and 71.7% of women's partnerships with disassortative age mixing), although the corresponding percentages for the youngest age group were 96.9% of women vs 25.3% of men. When stratified by partnership type, disassortative age mixing was more common in men's and women's casual partnerships than their steady partnerships, with the gap widening between the corresponding proportions with increasing age, especially for women's partnerships.

Disassortative ethnic mixing

Disassortative ethnic mixing occurred in 11.3% of all men's and 8.6% of all women's partnerships, with these proportions declining with age for both genders (Table 2). For both men and women, disassortative ethnic mixing was more common in casual partnerships than steady partnerships (20.6 vs 7.2%, respectively, of men's partnerships and 14.9 vs 6.9%, respectively, of women's partnerships).

Disassortative geographical mixing

The proportion of partnerships involving someone from a different region or a different country was similar for men and women (14.1 and 16.3%, respectively) (Table 2).

Fluctuations in disassortative geographical mixing by age group were evident but there was no obvious linear trend. Similar proportions of steady and casual partnerships involved disassortative geographical mixing, apart from at the youngest ages, when this type of mixing was twice as likely in casual vs steady partnerships (13.6 vs 6.8% of men's partnerships; 14.0 vs 8.3% of women's partnerships to those aged 16–24 y).

Disassortative sexual mixing and its implications for STI risk

Focusing on those aged 16–44 y, the age group at greatest risk of STIs,¹ condoms were less likely to have been used at first sex in women's partnerships involving disassortative age mixing, including after adjusting for the age of the woman at first sex with her partner [AOR: 0.79, 95% confidence interval (CI): 0.65–0.95] (Table 3). The effect of having an age-disassortative partnership on condom use at first sex was greater the younger the woman was at the start of the partnership. For example, women under 20 y at first sex with their partner had an AOR of 0.56 (95% CI: 0.42–0.75), whereas it was 0.83 (95% CI: 0.64–1.08) for women aged 20 y or older at this time. Furthermore, the AOR for condom use was 0.76 (95% CI: 0.59–0.99) for women's partnerships where specifically the male partner was ≥ 5 y older but there was no difference in the odds where the male was ≥ 5 y younger (both relative to the woman being of a similar age to her partner). Neither ethnic mixing nor geographical mixing was associated with the partnership-level indicator of STI risk for either gender.

At a participant-level, there was some weak evidence that women—but not men—who had at least one MRP involving disassortative age mixing were more likely to perceive themselves to be at risk of STIs (AOR: 1.21, 95% CI: 1.02–1.44). There was stronger evidence that men—but not women—who had at least one MRP involving disassortative ethnic mixing were more likely to perceive themselves to be at risk of STIs (AOR: 1.76, 95% CI: 1.23–2.52). Disassortative geographical mixing was not associated with STI-risk perception among men or women. There was weak evidence of an association for men between ethnic mixing and reporting STI diagnosis/es in the past year (AOR: 2.37, 95% CI: 1.22–4.59). When all three sexual mixing indicators were included in the models, similar effect sizes were observed as for when each sexual mixing indicator was considered in the model by itself (data not shown).

Discussion

In Britain, people tend to be similar to their opposite-sex partners in terms of their age and ethnic group, and to be

Table 1. Age mixing in the population of partnerships by gender, age group and partnership type

	Age group										All ages (16-74)	All ages (16-44)	
	16-24	25-34	35-44	45-54	55-64	65-74	65-74		All ages (16-74)				
%	[95% CI]	%	[95% CI]	%	[95% CI]	%	[95% CI]	%	[95% CI]	%	[95% CI]	%	[95% CI]
Comparing all partnerships by gender and age group													
<i>Men's partnerships</i>													
Denominators (unweighted, weighted partnerships)	1989, 2318	1671, 2101	805, 1593	685, 1451	528, 1034	286, 448			5964, 8945	4384, 5916			
Same age (<5 y younger/older)	92.1 [89.1, 94.4]	63.2 [56.6, 69.3]	53.5 [49.2, 57.8]	47.7 [41.5, 53.9]	40.4 [35.3, 45.8]	50.0 [43.6, 56.5]			63.1 [60.6, 65.5]	71.7 [68.8, 74.5]			
Man ≥5 y younger	5.9 [3.9, 8.9]	6.9 [5.5, 8.5]	6.7 [5.1, 8.6]	4.3 [3.0, 6.2]	3.8 [2.4, 6.0]	4.4 [2.5, 7.8]			5.7 [4.9, 6.6]	6.4 [5.4, 7.7]			
Man ≥5 y older	2.0 [1.2, 3.3]	29.9 [23.8, 36.9]	39.9 [35.6, 44.2]	48.0 [41.6, 54.6]	55.8 [50.3, 61.2]	45.6 [39.1, 52.2]			31.2 [28.8, 33.7]	21.8 [19.1, 24.8]			
Any disassortative age mixing	7.9	36.8	46.6	52.3	59.6	50.0			36.9	28.2			
% of disassortative where man older	25.3	81.3	85.6	91.8	93.6	91.2			84.6	77.3			
<i>Women's partnerships</i>													
Denominators (unweighted, weighted partnerships)	2454, 1823	2622, 1645	1107, 1723	885, 1252	524, 705	240, 317			7832, 7466	6113, 5136			
Same age (<5 y younger/older)	80.6 [77.8, 83.2]	67.2 [64.6, 69.6]	58.4 [54.6, 62.2]	61.9 [58.7, 65.7]	64.2 [59.6, 68.6]	72.7 [66.2, 78.4]			67.5 [65.6, 69.2]	68.9 [66.4, 71.3]			
Man ≥5 y younger	0.6 [0.3, 1.3]	6.4 [5.4, 7.6]	12.0 [8.2, 17.2]	17.4 [14.4, 20.8]	17.2 [13.9, 21.0]	8.4 [5.0, 13.7]			9.2 [8.1, 10.5]	6.2 [5.2, 7.5]			
Man ≥5 y older	18.8 [16.3, 21.6]	26.5 [24.3, 28.9]	29.6 [23.5, 36.5]	20.7 [17.9, 23.8]	18.6 [15.3, 22.5]	18.9 [14.2, 24.8]			23.3 [21.2, 25.6]	24.8 [22.1, 27.9]			
Any disassortative age mixing	19.4	32.9	41.6	38.1	35.8	27.3			32.5	31.0			
% of disassortative where man older	96.9	80.5	71.2	54.3	52.0	69.2			71.7	80.0			
<i>p</i> -value for difference between men's and women's partnerships	<0.001	0.37	0.007	<0.001	<0.001	<0.001			<0.001	0.24			
Comparing men's partnerships by type and age group													
<i>Men's steady partnerships</i>													
Denominators (unweighted, weighted partnerships)	1086, 1178	1204, 1368	653, 1291	561, 1146	443, 836	263, 406			4210, 6225	2916, 3799			
Same age (<5 y younger/older)	93.2 [87.9, 96.2]	67.4 [60.5, 73.6]	57.2 [52.9, 61.4]	53.2 [48.5, 57.8]	46.2 [40.5, 51.9]	54.0 [47.3, 60.6]			63.7 [61.4, 66.0]	71.8 [68.5, 74.8]			
Man ≥5 y younger	4.2 [1.7, 9.8]	6.2 [4.9, 7.9]	6.6 [5.0, 8.8]	4.5 [3.1, 6.6]	3.7 [2.2, 6.0]	4.1 [2.3, 7.3]			5.1 [4.3, 6.2]	5.8 [4.5, 7.3]			
Man ≥5 y older	2.6 [1.3, 5.1]	26.4 [20.1, 33.8]	36.2 [32.1, 40.5]	42.3 [37.8, 46.9]	50.2 [44.3, 56.0]	41.9 [35.4, 48.7]			31.1 [28.9, 33.5]	22.5 [19.6, 25.7]			
Any disassortative age mixing	6.8	32.6	42.8	46.8	53.9	46.0			36.2	28.3			
% of disassortative where man older	38.2	81.0	84.6	90.4	93.1	91.1			85.9	79.5			
<i>Men's casual partnerships</i>													
Denominators (unweighted, weighted partnerships)	903, 1140	467, 733	152, 302	124, 305	85, 197	23, 42			1754, 2719	1467, 2116			
Same age (<5 y younger/older)	91.0 [87.0, 93.9]	55.2 [42.2, 67.5]	37.4 [27.0, 49.1]	26.4 [14.3, 43.5]	15.6 [6.5, 32.8]	9.3 [3.4, 22.8]			61.6 [55.7, 67.2]	71.7 [65.9, 76.9]			
Man ≥5 y younger	7.7 [5.0, 11.7]	8.1 [5.3, 12.1]	6.8 [3.9, 11.7]	3.4 [1.2, 9.4]	4.2 [1.1, 14.3]	7.5 [1.0, 38.8]			7.0 [5.3, 9.0]	7.7 [5.8, 10.1]			
Man ≥5 y older	1.3 [0.7, 2.4]	36.8 [24.8, 50.7]	55.8 [44.0, 67.0]	70.2 [51.8, 83.8]	80.2 [63.3, 90.5]	83.2 [60.2, 94.2]			31.4 [25.8, 37.7]	20.6 [15.6, 26.8]			
Any disassortative age mixing	9.0	44.9	62.6	73.6	84.4	90.7			38.4	28.3			
% of disassortative where man older	14.4	82.0	89.1	95.4	95.0	91.7			81.8	72.8			

(Continued)

Table 1. Continued

	Age group										All ages (16-44)					
	16-24	25-34	35-44	45-54	55-64	65-74	All ages (16-74)		All ages (16-44)							
	%	[95% CI]	%	[95% CI]	%	[95% CI]	%	[95% CI]	%	[95% CI]	%	[95% CI]				
<i>p</i> -value for difference between men's steady and casual partnerships	0.19		0.13		0.001		0.003		0.002		0.001		0.35		0.41	
Comparing women's partnerships by type and age group																
<i>Women's steady partnerships</i>																
Denominators (unweighted, weighted partnerships)	1650, 1111		2113, 1281		957, 1386		789, 1141		492, 673		231, 307		6232, 5899		4681, 3748	
Same age (<5 y younger/older)	81.7	[78.9, 84.2]	69.8	[67.4, 72.0]	66.5	[59.0, 73.3]	64.7	[60.8, 68.4]	66.0	[61.2, 70.4]	74.5	[68.0, 80.0]	70.1	[68.0, 72.0]	72.1	[69.4, 74.6]
Man <5 y younger	0.9	[0.4, 2.0]	5.2	[4.3, 6.4]	11.3	[8.3, 15.2]	13.7	[11.2, 16.7]	14.9	[11.7, 18.7]	6.3	[3.6, 10.7]	8.7	[7.7, 9.8]	6.2	[5.1, 7.5]
Man ≥5 y older	17.4	[15.0, 20.2]	25.0	[22.9, 27.2]	22.2	[18.6, 27.6]	21.6	[18.6, 24.9]	19.2	[15.8, 23.2]	19.2	[14.4, 25.2]	21.3	[19.8, 22.9]	21.7	[19.7, 23.9]
Any disassortative age mixing	18.3		30.2		33.5		35.3		34.1		25.5		30.0		27.9	
% of disassortative where man older	95.1		82.8		66.3		61.2		56.3		75.3		71.0		77.8	
<i>Women's casual partnerships</i>																
Denominators (unweighted, weighted partnerships)	804, 712		509, 364		150, 337		96, 112		32, 33		9, 10		1600, 1567		1432, 1387	
Same age (<5 y younger/older)	78.9	[73.3, 83.6]	57.9	[51.3, 64.3]	24.5	[9.6, 49.5]	34.1	[23.2, 46.8]	28.5	[15.3, 46.8]	14.4	[2.7, 50.4]	57.6	[47.1, 67.5]	60.4	[48.2, 71.5]
Man <5 y younger	0.1	[0.2, 1.0]	10.4	[7.7, 14.1]	14.9	[5.6, 33.9]	54.8	[41.5, 67.4]	64.3	[45.4, 79.6]	76.6	[37.6, 94.7]	11.4	[8.6, 15.1]	6.3	[4.5, 8.9]
Man ≥5 y older	21.0	[16.3, 26.6]	31.6	[25.6, 38.3]	60.7	[29.2, 85.2]	11.2	[6.1, 19.7]	7.2	[1.8, 25.4]	8.9	[1.0, 47.9]	31.0	[20.3, 44.1]	33.2	[21.7, 47.2]
Any disassortative age mixing	21.1		42.0		75.6		66.0		71.5		85.5		42.4		39.5	
% of disassortative where man older	99.5		75.2		80.3		17.0		10.1		10.4		73.1		84.1	
<i>p</i> -value for difference between women's steady and casual partnerships	0.093		<0.001		0.013		<0.001		<0.001		<0.001		0.068		0.092	

p-value for overall association

Table 2. Disassortative ethnic and geographic mixing in the population of partnerships by gender, age group and partnership type

	Age group																
	16-24	25-34	35-44	45-54	55-64	65-74	All ages (16-74)		All ages (16-44)								
	%	[95% CI]	%	[95% CI]	%	[95% CI]	%	[95% CI]	%	[95% CI]							
Ethnic mixing																	
Comparing all partnerships by gender and age group																	
<i>Men's partnerships</i>																	
Denominators (unweighted, weighted partnerships)	1990, 2319	1667, 2098	804, 1576	683, 1445	527, 1031	286, 448	5957, 8918	4461, 5993									
Disassortative ethnic mixing	13.5	[10.3, 17.4]	16.4	[11.8, 22.4]	8.6	[6.3, 11.6]	8.4	[5.8, 11.9]	7.6	[4.7, 12.1]	3.4	[1.3, 8.3]	<0.001	11.3	[9.7, 13.2]	13.2	[11.0, 15.8]
<i>Women's partnerships</i>																	
Denominators (unweighted, weighted partnerships)	2454, 1823	2614, 1641	1105, 1721	883, 1249	523, 705	239, 316	7818, 7454	6173, 5185									
Disassortative ethnic mixing	13.2	[10.8, 16.1]	10.2	[8.6, 12.1]	6.4	[4.3, 9.5]	7.3	[5.5, 9.6]	3.0	[1.9, 4.8]	3.0	[1.6, 5.9]	<0.001	8.6	[7.5, 9.8]	10.0	[8.5, 11.7]
<i>p</i> -value for difference between men's and women's partnerships	0.92	0.011	0.25	0.53	0.005	0.83								0.006		0.021	
Comparing men's partnerships by type and age group																	
<i>Men's steady partnerships</i>																	
Denominators (unweighted, weighted partnerships)	1086, 1178	1200, 1364	652, 1275	559, 1140	442, 834	263, 406	4202, 6198	2938, 3817									
Disassortative ethnic mixing	8.5	[6.2, 11.7]	9.2	[7.0, 12.0]	8.2	[5.9, 11.2]	6.9	[4.8, 9.7]	3.5	[2.1, 6.0]	2.2	[1.0, 4.8]	0.003	7.2	[6.3, 8.3]	8.7	[7.3, 10.2]
<i>Men's casual partnerships</i>																	
Denominators (unweighted, weighted partnerships)	903, 1140	467, 733	151, 301	124, 305	85, 197	22, 42	1753, 2718	1521, 2174									
Disassortative ethnic mixing	18.6	[13.3, 25.3]	29.9	[19.2, 43.4]	10.3	[5.5, 18.7]	14.1	[6.0, 29.9]	24.6	[12.3, 43.1]	15.3	[2.3, 57.9]	0.09	20.6	[16.3, 25.7]	21.2	[16.4, 27.1]
<i>p</i> -value for difference between men's steady and casual partnerships	0.002	<0.001	0.49	0.11	<0.001	0.031								<0.001		<0.001	
Comparing women's partnerships by type and age group																	
<i>Women's steady partnerships</i>																	
Denominators (unweighted, weighted partnerships)	1649, 1111	2110, 1279	955, 1384	787, 1137	491, 672	230, 306	6222, 5889	4714, 3774									
Disassortative ethnic mixing	11.0	[9.0, 13.3]	7.9	[6.6, 9.5]	6.2	[4.4, 8.7]	6.4	[4.7, 8.7]	2.4	[1.4, 4.1]	2.9	[1.5, 5.8]	<0.001	6.9	[6.1, 7.8]	8.2	[7.1, 9.4]
<i>Women's casual partnerships</i>																	
Denominators (unweighted, weighted partnerships)	804, 712	504, 362	150, 337	96, 112	31, 31	9, 10	1594, 1563	1458, 1410									
Disassortative ethnic mixing	16.8	[12.1, 22.9]	18.1	[13.6, 23.8]	7.3	[2.7, 17.9]	16.1	[8.5, 28.5]	12.5	[4.3, 31.3]	6.4	[0.8, 37.9]	0.15	14.9	[11.3, 19.3]	14.9	[11.0, 19.8]
<i>p</i> -value for difference between women's steady and casual partnerships	0.023	<0.001	0.71	0.008	0.002	0.48								<0.001		<0.001	

(Continued)

Table 2. Continued

		Age group										All ages (16–44)				
		16–24	25–34	35–44	45–54	55–64	65–74	p-value ^a		All ages (16–44)						
%	[95% CI]	%	[95% CI]	%	[95% CI]	%	[95% CI]	%	[95% CI]	%	[95% CI]	%	[95% CI]			
Geographical mixing																
Comparing all partnerships by gender and age group																
<i>Men's partnerships</i>																
Denominators (unweighted, weighted partnerships)	1964, 2283	1654, 2068	803, 1589	683, 1438	528, 1034	284, 443		5916, 8855	4421, 5941							
Disassortative geographical mixing	10.1	[8.0, 12.6]	13.2	[10.7, 16.1]	19.4	[16.3, 22.9]	14.5	[11.2, 18.5]	16.8	[12.6, 22.0]	12.7	[8.1, 19.4]	14.1	[12.8, 15.5]	13.7	[12.2, 15.3]
<i>Women's partnerships</i>																
Denominators (unweighted, weighted partnerships)	2441, 1811	2611, 1639	1105, 1721	884, 1252	522, 704	240, 317		7803, 7444	6157, 5171							
Disassortative geographical mixing	10.5	[8.6, 12.8]	15.4	[13.6, 17.4]	27.4	[11.0, 53.6]	13.9	[11.6, 16.6]	12.6	[9.9, 15.9]	10.9	[7.4, 15.9]	16.3	[11.2, 23.1]	17.7	[10.8, 27.7]
<i>p</i> -value for difference between men's and women's partnerships	0.64	0.89	0.037	0.83	0.16	0.96		0.089	0.46				0.089	0.46	0.31	
Comparing men's partnerships by type and age group																
<i>Men's steady partnerships</i>																
Denominators (unweighted, weighted partnerships)	1085, 1178	1203, 1366	653, 1291	561, 1146	443, 836	262, 404		4207, 6221	2941, 3835							
Disassortative geographical mixing	6.8	[5.1, 9.1]	12.5	[10.0, 15.4]	18.7	[15.4, 22.5]	12.9	[9.9, 16.6]	14.0	[10.5, 18.5]	13.1	[8.2, 20.3]	13.0	[11.7, 14.5]	12.8	[11.2, 14.6]
<i>Men's casual partnerships</i>																
Denominators (unweighted, weighted partnerships)	878, 1104	451, 702	149, 298	122, 292	85, 197	22, 39		1707, 2632	1478, 2104							
Disassortative geographical mixing	13.6	[10.1, 18.0]	14.6	[10.0, 20.7]	22.4	[14.7, 32.7]	20.7	[10.0, 38.0]	28.4	[14.8, 47.6]	8.6	[2.3, 27.2]	16.7	[13.7, 20.1]	15.2	[12.4, 18.5]
<i>p</i> -value for difference between men's steady and casual partnerships	<0.001	0.45	0.44	0.22	0.047	0.52		0.077	0.026				0.077	0.026	0.16	
Comparing women's partnerships by type and age group																
<i>Women's steady partnerships</i>																
Denominators (unweighted, weighted partnerships)	1649, 1109	2106, 1277	956, 1386	789, 1141	490, 671	231, 307		6221, 5890	4711, 3772							
Disassortative geographical mixing	8.3	[6.7, 10.3]	14.2	[12.3, 16.4]	22.8	[11.3, 40.7]	13.3	[10.9, 16.2]	12.3	[9.5, 15.7]	11.1	[7.4, 16.1]	14.6	[11.1, 19.0]	15.6	[10.5, 22.6]
<i>Women's casual partnerships</i>																
Denominators (unweighted, weighted partnerships)	792, 702	505, 362	149, 336	95, 111	31, 32	9, 10		1581, 1552	1446, 1399							
Disassortative geographical mixing	14.0	[10.2, 18.9]	19.7	[15.2, 25.1]	46.6	[13.5, 83.0]	19.7	[11.5, 31.7]	19.0	[8.7, 36.8]	6.4	[0.8, 37.9]	22.8	[12.2, 38.5]	23.3	[11.8, 40.7]
<i>p</i> -value for difference between women's steady and casual partnerships	0.007	0.036	0.021	0.18	0.29	0.59		0.019	0.015				0.019	0.015		

^a *p*-value for overall association.

Table 3. Continued

	Male participants					Female participants						
	% of participants	95% CI	Odds ratio	AOR ^a	p-value	Denominator ^c	% of participants	95% CI	Odds ratio	AOR ^a	p-value	Denominator ^c
Participant-level outcome: STI diagnosis in the past year												
Hypothesized explanatory variables												
Disassortative age mixing in any of (max.) 3 MRPs					0.55						0.054	
No, only assortative mixing	1.4	[1.0, 2.0]	1	1	–	2233, 2197	1.5	[1.0–2.1]	1	1	–	3062, 2181
Yes, age mixing in at least one partnership	1.6	[0.9, 2.6]	1.10	1.23	[0.63, 2.40]	906, 1047	2.1	[1.5–3.0]	1.46	1.64	[0.99, 2.72]	1637, 1102
Disassortative ethnic mixing in any of (max.) 3 MRPs					0.011						0.12	
No, only assortative mixing	1.2	[0.8, 1.6]	1	1	–	2762, 2873	1.4	[1.1, 1.9]	1	1	–	4203, 2946
Yes, ethnic mixing in at least one partnership	4.0	[2.3, 6.7]	3.53	2.37	[1.22, 4.59]	377, 372	4.4	[2.5, 7.5]	3.23	1.83	[0.86, 3.89]	496, 336
Disassortative geographical mixing (when first met) in any of (max.) 3 MRPs					0.87						0.062	
No, only assortative mixing	1.5	[1.1, 2.0]	1	1	–	2632, 2678	1.5	[1.2, 2.0]	1	1	–	3969, 2764
Yes, geographical mixing in at least one partnership	1.5	[0.7, 3.1]	1.02	0.94	[0.43, 2.04]	507, 567	2.8	[1.7, 4.5]	1.86	1.68	[0.98, 2.90]	730, 519

^aAOR, adjusted odds ratio; for partnership-level outcome (condom use at first sex with partner), OR-adjusted for participant's age at first sex with that partner; for participant-level outcomes (STI-risk perception and reported STI diagnosis), ORs adjusted for participant's age at interview, whether any of their (max.) 3 MRPs were casual, participant's opposite-sex partner numbers in the past year and participant's ethnicity when the sexual mixing indicator related to ethnic mixing.

^bUnweighted, weighted partnerships in the past year of sexually active participants aged 16–44.

^cUnweighted, weighted participants, defined as those sexually active and aged 16–44.

^dParticipants were asked: 'What do you think about the risks to you, personally, with your present lifestyle of getting a sexually transmitted infection that is not HIV?' Response options were: greatly at risk; quite a lot; not very much; not at all. For the purposes of these (binary) analyses, this was coded as 'not at all' vs 'at some risk'.

from the same geographical region when they meet. Around one-third of all partnerships involved disassortative age mixing, whereas smaller proportions of partnerships—less than one-sixth—entailed disassortative ethnic and/or geographical mixing. Disassortative sexual mixing is more common in men's partnerships than women's, which may in part reflect how a larger proportion of men's reported partnerships are casual, as disassortative sexual mixing was often more prevalent in casual than steady partnerships. We found disassortative sexual mixing to be associated with markers of STI risk independently of key confounders including age and the number and type(s) of partners reported. Of note, the odds of condom use at first sex with a partner were lower for women in age-disassortative partnerships, whereas the odds of men perceiving themselves to be at risk of STIs and/or reported STI diagnoses were higher if they had had recent ethnic-disassortative partnerships. However, we observed no evidence of a consistent relationship between sexual mixing and STI-risk markers, and the effect sizes observed were weaker than those for more established risk factors, such as age and number and type(s) of partner, which we adjusted for in our models (see [Supplementary Table 4](#), available as [Supplementary data](#) at *IJE* online). This may reflect how sexual mixing is more complex conceptually and to measure than factors such as age and partner numbers, in part because it corresponds to not just the individual's characteristics, but also those of their partners and their sexual network more broadly.

A key strength of this study is that we analysed probability survey data, meaning that our findings can be considered as broadly representative of the British general population.²¹ Whereas the response rate was 57.7%, this is consistent with other major social surveys undertaken contemporaneously in Britain^{26,27} and the co-operation rate was 65.8% (of all contacted addresses known to be eligible).²⁸ Nonetheless, we acknowledge that non-response could be a source of bias for our data. We aimed to minimize this bias by weighting the sample so that it was broadly representative of the underlying population with respect to the distribution of the sexes, age and regions as used in the census.²¹ Furthermore, the sampling strategy used for the Natsal studies means that the target population is specifically the population resident in private households in Britain and, as such, excludes individuals living in institutions, whose behaviour could differ from others'. Whereas this sampling strategy is also a potential source of bias, the institutionalized population constitutes a relatively small proportion of the British population.²¹ We also endeavoured to make our sample of partnerships broadly representative of the population of opposite-sex *partnerships* experienced by people living in Britain in the

past year. This involved taking account not just of the current or most recent partner, as in many other studies,^{29–32} but also weighting the three MRPs that were asked about in detail in Natsal-3 to represent less recent partnerships for whom this information was not collected.²⁰ As such, we were able to account for under-represented partnerships, which are more likely to be casual, shorter in duration¹⁹ and involve greater disassortative sexual mixing¹⁸—characteristics that may facilitate STI transmission.

Whereas Natsal-3 collected a wealth of data from participants on their sexual behaviour, it has relatively limited data on their partners' characteristics and specifically only three socio-demographic variables. However, it is worth noting that one study that had data on socio-demographic as well as sexual risk assortativity showed that both are likely to be important determinants of STI transmission.¹⁸ Furthermore, our paper builds upon analyses of Natsal-2 data¹⁷ as, in addition to disassortative age mixing, it investigated the extent of disassortative ethnic and geographical mixing and their significance for STI transmission. However, their relatively low population prevalence meant that we had to use relatively crude binary variables and were limited in the extent to which we could discriminate between categories. For example, limited numbers of partnerships where one or both partners were of an ethnicity other than 'White' meant that we were obliged to use broad ethnic categories, including a 'Black/Black British' group despite epidemiological differences in STI prevalence between Black Caribbean and Black African populations.³³ Similarly, we used a crude measure of disassortative geographical mixing, as this applies to partnerships formed between people from different regions within the same country as well as partnerships involving people from different countries and global regions. This is a particular limitation when considering how disassortative geographical mixing is associated with STI risk, as STI prevalence, and thus transmission probabilities, varies considerably globally.³⁴ With all socio-demographic mixing indicators, we acknowledge that these characteristics must work through the more proximate behavioural or biological determinants to impact on STI transmission. In addition, whereas we included both partnership- and participant-level markers of STI risk, we recognize that our participant-level measures of sexual mixing correspond to any mixing in participants' ≤ 3 MRPs, which may not correspond to the partnership(s) in which STI(s) were acquired or led to the participant perceiving themselves at risk of STIs. Furthermore, we recognize that, by analysing cross-sectional survey data, it is not possible to determine the chronology of events, e.g. whether having an age-disassortative partnership occurred before or after an STI diagnosis, nor assume causality more broadly.

Focusing on partnerships in the past year meant that all partners in this timeframe were captured in our analyses

for a very high proportion of study participants. Whereas those reporting STI diagnosis/es were more likely to be those who reported larger numbers of partners than asked about in the MRP module,¹ we used multivariable models to adjust for the total number of partners in the past year. By taking into account this established driver of STI transmission in our analyses, we have importantly demonstrated an independent, although weak, relationship between disassortative age and ethnic mixing and STI risk.

The predominance for men—rather than women—to be ≥ 5 y older in the majority of opposite-sex partnerships involving disassortative age mixing is consistent with previous studies in and outside of Britain.^{15,18,35} However, this was not observed for women in casual partnerships in whom the proportion with a male partner who was ≥ 5 y younger increased with age, although the denominator (the number of casual partnerships among women in older age groups) was small in absolute terms and also relative to their male counterparts, which in itself is a notable gender difference. These age mixing patterns and differences in the extent of ethnic and geographical mixing may reflect the availability of eligible partners in an individual's socio-sexual network,³⁶ as well as social norms, e.g. regarding what constitutes an appropriately aged partner, which may be less stringent for casual partnerships.³⁷

Disassortative age mixing has been shown to manifest itself, especially at younger ages, in gender power imbalances,^{29–31} resulting e.g. in condoms being less likely to be used, as we and others have found.^{38,39} However, it was not possible to determine from the quantitative data Natsal-3 collected whether such power imbalances were the reason for non-use of condoms. It is also worth noting that our measure of condom use only corresponds to the first occasion of sex with the partner, so does not capture how well condoms were used on that occasion, nor consistency of use during the partnership. Indeed, evidence suggests that condom use quickly wanes, on average within 3 weeks of first sex together.⁴⁰ Regarding our measure of STI-risk perception, only a single question was used and this may not capture the complexity of this concept, e.g. individuals' risk perception may change over time and people may perceive their risk to be different for different STIs. Furthermore, the placement of this question—after the CASI's detailed questions about sexual behaviour—may have influenced participants' assessment of their risk. Also, as it was asked in the face-to-face section of the interview, albeit using a showcard so that participants only needed to give a response code, it is plausible that responses may still be subject to social-desirability bias.

In addition to differential misreporting by gender,⁴¹ the observed gender differences in the three measures of sexual

mixing reflect how men's and women's partners are not from closed populations.⁴² For example, men are more likely than women to report new sexual partners from outside of the UK,¹⁶ thus partners who are ineligible themselves to participate in Natsal-3; they are also more likely to report paying for sex and so these partners are less likely to be captured by a survey like Natsal.⁴² However, in contrast to analyses of the previous Natsal in which the target population was men and women aged 16–44 y,¹⁷ these latest analyses of Natsal-3 data correspond to people from a much wider age range (16–74 y) meaning that they capture a greater proportion of both male and female partners, as well as the adult population, thus providing a more complete picture of age mixing in the population. Nonetheless, the age mixing patterns we observed are similar to those from our earlier analyses,¹⁷ suggesting that this component of partnerships has remained constant between 2000 and 2010, despite evidence of individual-level behaviour changing over time.¹⁶ Finally, and as in our earlier paper,¹⁷ we were not able to reliably quantify the extent to which the observed amount of disassortative mixing differs from that which would be expected if mixing occurred at random in the population. This relates in part to crudely categorizing ethnicity, differences in how participants may interpret 'region' and also challenges in defining the appropriate population from which to calculate expected numbers. Under mixing at random, most partnerships would involve disassortative age mixing, since, for any individual, most potential partners across the full age range of 16–74 y would be $\geq \pm 5$ y different in age. Our findings therefore indicate that the pattern of mixing in Britain in terms of age is strongly assortative, and likewise for geographical mixing too. Although we also found that only a minority of partnerships involve disassortative ethnic mixing, this is more difficult to interpret. Because 86% of the British population are estimated to be of White ethnicity,²³ this relative homogeneity ensures that the majority of partnerships will be ethnically-assortative whatever the underlying pattern of mixing. In terms of the role of ethnically-disassortative partnerships in STI transmission, we, like others,^{12,43} found that men who had had such partnerships were more likely to report STI diagnoses, but this was not observed for women.

Future research needs to focus on improving understanding the granularities of disassortative mixing, especially given its potential significance in STI transmission. Population-based studies therefore need to be adequately powered, including for ethnic minority groups. It is also vital that studies ask participants about more than just their most recent or cohabiting partner and, whereas it is unlikely to be feasible to ask all participants about all partners, the resulting data should be weighted to take account

of those partnerships that are not asked about, which are more likely to be casual. Stratifying analyses by partnership type is also necessary to avoid masking differences in mixing. Future studies should also seek to obtain data on partners' sexual behaviour, especially their partner's partner numbers given its importance in STI transmission.¹ Ideally, this is obtained directly from the partners—as in HSE-2010¹⁵—as there is evidence that individuals have a poor ability to assess their partners' behaviour, especially that of non-cohabiting partners.⁴⁴

In terms of implications for policy and practice, the evidence that sexual partnerships in Britain are mainly assortative suggests that STI-prevention efforts are best focused on those individuals at highest risk. However, as there was some evidence that disassortative mixing was independently associated with markers of STI risk, it is worth considering, from a public health perspective, whether highlighting disassortative sexual mixing in health-promotion messaging might be helpful e.g. for identifying those most likely to benefit from STI testing. Similarly, asking patients about their 'exposure' to disassortative sexual partnerships may warrant investigation to ascertain its predictive value and utility in clinical triage in contrast, or in addition, to conventional risk-assessment questions such as those on partner numbers and/or condom use. Qualitative research has shown that people are generally willing to answer questions about their partners, at least in the context of a survey,⁴⁵ but further research would be needed to assess the social acceptability in a clinical context of 'labelling', e.g. having an older partner, as an STI risk. Even if this is not the case, these data show that STI-prevention efforts may benefit from taking account of sexual mixing patterns as well as individual-level risk behaviour.

Supplementary data

Supplementary data are available at *IJE* online.

Funding

Natsal-3 was supported by grants from the Medical Research Council (G0701757) and the Wellcome Trust (084840), with contributions from the Economic and Social Research Council and Department of Health. The sponsors of the study had no role in study design, data collection, data analysis, data interpretation or writing of the report. The corresponding author had full access to all the data in the study and had final responsibility for the decision to submit for publication.

Conflict of interest: A.M.J. has been a Governor of the Wellcome Trust since 2011. The other authors declare that they have no conflicts of interest.

References

1. Sonnenberg P, Clifton S, Beddows S *et al.* Prevalence, risk factors, and uptake of interventions for sexually transmitted infections in Britain: findings from the National Surveys of Sexual Attitudes and Lifestyles (Natsal). *Lancet* 2013;**382**:1795–806.
2. Fenton KA, Mercer CH, Johnson AM *et al.* Reported sexually transmitted disease clinic attendance and sexually transmitted infections in Britain: prevalence, risk factors, and proportionate population burden. *J Infect Dis* 2005;**191**:S127–38.
3. Adimora AA, Schoenbach VJ. Social context, sexual networks, and racial disparities in rates of sexually transmitted infections. *J Infect Dis* 2005;**191**:S115–22.
4. Huang ZJ, Wang W, Martin MC *et al.* 'Bridge population': sex workers or their clients?—STI prevalence and risk behaviors of clients of female sex workers in China. *AIDS Care* 2011;**23**: 45–53.
5. Patterson TL, Volkman T, Gallardo M *et al.* Identifying the HIV transmission bridge: which men are having unsafe sex with female sex workers and with their own wives or steady partners? *J Acquir Immune Defic Syndr* 2012;**60**:414–20.
6. Low N, Broutet N, Adu-Sarkodie Y *et al.* Global control of sexually transmitted infections. *Lancet* 2006;**368**:1960–61.
7. Turner KME, Adams EJ, Gay N, Ghani AC, Mercer CH, Edmunds WJ. Developing a realistic sexual network model of chlamydia transmission in Britain. *Theor Biol Med Model* 2006;**3**:3.
8. Morris M (ed). *Network Epidemiology: A Handbook For Survey Design And Data Collection*. New York: Oxford University Press on Demand, 2004.
9. Merli MG, Moody J, Mendelsohn J *et al.* Sexual mixing in Shanghai: are heterosexual contact patterns compatible with an HIV/AIDS epidemic? *Demography* 2015;**52**:919–42.
10. Ford K, Woosung K, Lepkowski J. American adolescents: sexual mixing patterns, bridge partners, and concurrency. *Sex Transm Dis* 2002;**29**:13–19.
11. Garnett G, Hughes J, Anderson R *et al.* Sexual mixing patterns of patients attending sexually transmitted diseases clinics. *Sex Transm Dis* 1996;**23**:248–57.
12. Aral SO, Hughes JP, Stoner B *et al.* Sexual mixing patterns in the spread of gonococcal and chlamydial infections. *Am J Public Health* 1999;**89**:825–33.
13. Gorbach PM, Drumright LN, Holmes KK. Discord, discordance, and concurrency: comparing individual and partnership-level analyses of new partnerships of young adults at risk of sexually transmitted infections. *Sex Transm Dis* 2005;**32**:7–12.
14. <http://content.digital.nhs.uk/healthsurveyengland> (August 2018, date last accessed).
15. Prah P, Copas A, Mercer CH *et al.* Patterns of sexual mixing with respect to social, health and sexual characteristics among heterosexual couples in England: analyses of probability sample survey data. *Epidemiol Infect* 2015;**143**:1500–10.
16. Mercer CH, Tanton C, Prah P *et al.* Changes in sexual attitudes and lifestyles in Britain through the life course and over time: findings from the National Surveys of Sexual Attitudes and Lifestyles (Natsal). *Lancet* 2013;**382**:1781–94.
17. Mercer CH, Copas AJ, Sonnenberg P *et al.* Who has sex with whom? Characteristics of heterosexual partnerships reported in a national probability survey and implications for STI risk. *Int J Epidemiol* 2009;**38**:206–14.

18. Malagón T, Burchell A, El-Zein M *et al.* Assortativity and mixing by sexual behaviors and sociodemographic characteristics in young adult heterosexual dating partnerships. *Sex Transm Dis* 2017;**44**:329–37.
19. Mercer CH, Jones KG, Johnson AM *et al.* How objectively can we categorise partnership type? A novel classification of population survey data to inform clinical practice and epidemiology. *Sex Transm Infect* 2016;**93**:1–8.
20. Copas A, Mercer CH, Farewell V *et al.* Recent heterosexual partnerships and patterns of condom use: a weighted analysis. *Epidemiology* 2009;**20**:44–51.
21. Erens B, Phelps A, Clifton S *et al.* Methodology of the third British National Survey of Sexual Attitudes and Lifestyles (Natsal-3). *Sex Transm Infect* 2014;**90**:84–89.
22. Badcock PB, Smith AM, Richters J *et al.* Characteristics of heterosexual regular relationships among a representative sample of adults: the Second Australian Study of Health and Relationships. *Sex Health* 2014;**11**:427–38.
23. <https://www.ons.gov.uk/peoplepopulationandcommunity/culturalidentity/ethnicity/articles/ethnicityandnationalidentityinenglandandwales/2012-12-11> (August 2018, date last accessed).
24. StataCorp. 2017. *Stata Statistical Software: Release 15*. College Station, TX: StataCorp LLC.
25. Bottomley C, Kirby MJ, Lindsay SW, Alexander N. Can the buck always be passed to the highest level of clustering? *BMC Med Res Methodol* 2016;**16**:29.
26. Craig R, Mindell J (eds). *Health Survey for England 2010—Volume 1: Respiratory Health*. Leeds: The NHS Information Centre, 2011.
27. Park A, Clery E, Curtice J, Phillips M, Utting D (eds). *British Social Attitudes: The 28th Report*. London: NatCen Social Research, 2012.
28. The American Association for Public Opinion Research. *Standard Definitions: Final Dispositions of Case Codes and Outcome Rates for Surveys*. 2011. http://www.aapor.org/AM/Template.cfm?Section=Standard_Definitions2&Template=/CM/ContentDisplay.cfm&ContentID=3156 (August 2018, date last accessed).
29. Darroch J, Landry D, Oslak S. Age differences between sexual partners in the United States. *Fam Plann Perspect* 1999;**31**:160–67.
30. Juarez F, Martin T. Partnership dynamics and sexual health risks among male adolescents in the Favelas of Recife, Brazil. *Int Fam Plan Perspect* 2006;**32**:62–70.
31. Luke N. Confronting the ‘sugar daddy’ stereotype: age and economic asymmetries and risky sexual behavior in urban Kenya. *Int Fam Plan Perspect* 2005;**31**:6–14.
32. Kraut-Becher J, Aral S. Patterns of age mixing and sexually transmitted infections. *Int J STD AIDS* 2006;**17**:378–83.
33. Furegato M, Chen Y, Mohammed H *et al.* Examining the role of socioeconomic deprivation in ethnic differences in sexually transmitted infection diagnosis rates in England: evidence from surveillance data. *Epidemiol Infect* 2016;**144**:3253–62.
34. Newman L, Rowley J, Vander Hoorn S *et al.* Global estimates of the prevalence and incidence of four curable sexually transmitted infections in 2012 based on systematic review and global reporting. *PLoS ONE* 2015;**10**:e0143304.
35. Gregson S, Nyamukapa CA, Garnett GP *et al.* Sexual mixing patterns and sex-differentials in teenage exposure to HIV infection in rural Zimbabwe. *Lancet* 2002;**359**:1896–903.
36. Matson PA, Chung S, Ellen JM. Perceived neighborhood partner availability, partner selection, and risk for sexually transmitted infections within a cohort of adolescent females. *J Adolesc Health* 2014;**55**:122.
37. Verbakel E, Kalmijn M. Assortative mating among Dutch married and cohabiting same-sex and different-sex couples. *Fam Relat* 2014;**76**:1–12.
38. Harper G, Doll M, Bangi A *et al.* Female adolescents and older male sex partners: HIV associated risk. *J Adolesc Health* 2002;**30**:146–47.
39. Volpe E, Hardie T, Cerulli C *et al.* What’s age got to do with it? Partner age difference, power, intimate partner violence, and sexual risk in urban adolescents. *J Interpers Violence* 2013;**28**:2068–87.
40. Fortenberry JD, Tu W, Harezlak J, Katz BP, Orr DP. Condom use as a function of time in new and established adolescent sexual relationships. *Am J Public Health* 2002;**92**:211–13.
41. Prah P, Copas AJ, Mercer CH *et al.* Consistency in reporting sensitive sexual behaviours in Britain: change in reporting bias in the second and third National Surveys of Sexual Attitudes and Lifestyles (Natsal-2 and Natsal-3). *Sex Transm Infect* 2014;**90**:90–93.
42. Wadsworth J, Johnson AM, Wellings K *et al.* What’s in a mean? An examination of the inconsistency between men and women in reporting sexual partnerships. *J R Stat Soc A* 1996;**159**:111–23.
43. Xiridou M, van Veen M, Prins M *et al.* How patterns of migration can influence the heterosexual transmission of HIV in the Netherlands. *Sex Transm Infect* 2011;**87**:289–91.
44. Drumright LN, Gorbach PM, Holmes KK. Do people really know their sex partners? Concurrency, knowledge of partner behavior, and sexually transmitted infections within partnerships. *Sex Transm Dis* 2004;**31**:437–42.
45. Aicken CRH, Gray M, Clifton S *et al.* Improving questions on sexual partnerships: lessons learned from cognitive interviews for Britain’s third National Survey of Sexual Attitudes and Lifestyles (‘Natsal-3’). *Arch Sex Behav* 2013;**42**:173–85.