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## Population Demography in Ecology

## Citation for published version:

Newman, K 2019, Population Demography in Ecology. in Handbook of Environmental and Ecological Statistics. 1 edn, CRC Handbooks of Modern Statistical Methods, Chapman and Hall/CRC.

## Link:

Link to publication record in Edinburgh Research Explorer

## Document Version:

Early version, also known as pre-print

## Published In:

Handbook of Environmental and Ecological Statistics

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## Demography Ken Newman

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### 1.1 Introduction

Demography is often thought of in terms of human populations. The word's origin is a combination of the ancient Greek words demos, meaning "the people", and "graphy", which refers to the "the writing or recording or study of". One definition of demography is "the science of vital and social statistics, as of births, deaths, diseases, marriages, etc, of populations" (Companies, 2005). The focus here is on ecology and ecological populations, and demography will be defined similarly but without the human social science component, e.g., the notion of marriage is not relevant. In particular demography will be defined as the scientific study and characterization of biological populations' structure and dynamics. In the case of a single population, structure can include total abundance at arbitrary points in time. Structure can also refer to how the population might be partitioned into different categories such as sex, age, or spatial location. Dynamics refers both the changes in structure and abundances over time as well as the collection of processes that cause these changes. These processes, sometimes referred to as "vital rates" include reproduction, growth, maturity, movement, and mortality.

Demography is of central importance to scientists and natural resource managers (not that these are mutually exclusive groups) for a variety of reasons. There is of course inherent curiosity, pure scientific interest, about the abundances and dynamics of many species. Why do the numbers of wolves (Canis lupus) on Isle Royale (in Lake Superior) fluctuate the way that they do? What effect will decreased snowpack levels have on the geographic range of American pika (Ochotona princeps) in Yosemite National Park? Answers to such questions require not only estimates of abundances of the species but also understanding of the factors that affect the abundances and dynamics.

For species that are harvested commercially, for sport, or for subsistence, e.g., salmon (Oncorhynchus spp.), red deer (Cervus elaphus), morel mushrooms (Morchella spp.), and black duck (Anas rubripes), there is interest in
the effect of harvest on the population abundances and dynamics. Comparison of alternative harvest regulations is facilitated by predictions of the magnitude and sustainability of harvest levels.. To have some idea as to the effects of setting harvest regulations, e.g., a bag limit of 10 black ducks for a one month hunting season, requires some understanding of how this mortality might interact with other sources of mortality and other processes, like reproduction or movement. Estimates of the degree to which harvest mortality will be compensatory (removes individuals that will die anyway from other factors during that time period) and additive (the number of animals that will be removed over and above those that would have died from other factors) can be useful.

In the case of species declared threatened or endangered by a government agency there is often a legal mandate for actions to be taken, or avoided, by the owners or managers of land or water regions inhabited by the species. Those actions can pertain directly to the population, such as to not take actions that could kill, harm, or harass the species, or pertain to actions which could harm or reduce the species' habitat. More positively, managers wanting to recover the population would like to identify actions that could increase the species abundance by making habitat restoration efforts or eliminating habitat loss. For example, the United States Fish and Wildlife Service (USFWS) has a mandate to develop "Conservation Management Plans" for species listed as threatened or endangered under the US Endangered Species Act. Such plans must include (a) specification of management actions to conserve the species, (b) measurable criteria which would lead to a determination that the species can be "delisted", no longer declared threatened, and (c) estimates of the time and cost to carry out such actions. Demographics models are central to identifying such actions, to predicting the effects of actions, and to prioritizing multiple actions.

Questions about demographics can be coarsely divided into questions about abundances and about processes. How many individuals, or what volume or mass, have there been in the past and how many are there currently? As noted previously questions about population abundances over time can be further refined by partitioning the population into subgroups or subpopulations, e.g., sex, location, age, genotype, etc. Answering the how many question alone can be a challenging problem depending on the magnitude of the abundances, geographic location and range, physical size, mobility, degree of elusiveness, and/or ability to see or to detect individuals. A variety of statistical sampling methods, e.g., mark-recapture, and technological tools and devices, e.g., radio tracking, have been developed to help provide answers to the how many question. Answering the how many question is the primary focus of other chapters in this volume, and further discussion here is limited.

Even if population abundances, however partitioned, were known with certainty, questions about population processes remain. Why were the numbers what they were last year and why are they what they are now? What are the relative effects of each process on abundances at specific points in time? For example, how do adult female fecundity rates of salmon, egg hatching success
rates, and larval to juvenile survival combine to affect the abundance of juveniles? How do environmental conditions, both natural and anthropogenic, affect these processes?

The primary focus of this chapter is on mathematical and statistical approaches to answering such process questions. Answering these questions involves a quantification of the relationship between past abundance and current abundances. In other words, a population dynamics model (PDM) is needed. The more complete and accurate the PDM, the better the understanding of how changes in environmental and anthropogenic factors influence population processes, and how changes in these processes translate into changes in population abundances. Measures of the degree of uncertainty as to the consequences are critical as well. For threatened or endangered species, in particular, PDMs are also central to population viability analysis (PVA, Morris et al., 2002). PVAs use PDMs to make predictions about population trajectories, typically via computer simulation. They are a means of estimating the probability of the species surviving, or not, as a function of environmental conditions and anthropogenic factors, including accidental actions, like oil spills, and deliberate actions, like habitat restoration.

While attention here will be centered largely on general notions about factors influencing population dynamics and some popular quantitative models, it is worth pointing out how answers to initial what, why, and how questions typically lead to deeper investigation and potentially a chain of subsequent what, why, and how questions. Answers at the end of the chain can lead to ideas about management actions to take, assuming some actions are desired, and implementation of a particular action may then be justified by reversing the direction to yield a so-called results chain (Margoluis et al., 2013). For example, a proposed management action is to plant riparian vegetation along a stream where juvenile salmon rear. The results chain is the vegetation grows and provides increasing shade along the stream, the shade causes reductions in water temperatures, which leads to an increase in juvenile survival during the month of May, resulting in an increase in the population abundance. Having such a conceptual understanding in place can then both guide data collection, such as long term biological monitoring programs (Reynolds et al., view), and guide further model development. For example, to assess the effects of planting riparian vegetation, one might collect a time series of measurements of vegetation biomass, hours of shade, stream temperatures, juvenile abundances before and after the month of May (to estimate survival) at both treatment sites and control sites where no planting is done (Before-After-Control-Impact BACI designs, Smith, 2002).

The organization of the remainder of this chapter is the following. Section 1.2 is an overview of components of demography, including the definition of multiple subpopulations and multiple processes. Section 1.3 presents a progression of mathematical models more or less corresponding to the components in section 1.2. Section 1.4 discusses matrix population models (MPMs) and is followed by Section 1.5 on integral projection models (IPMs). Individual based

## TABLE 1.1

Listing of frequently used acronyms and their meaning.

| Acronym | Meaning |
| :--- | :--- |
| PDM | Population Dynamics Model |
| MPM | Matrix Projection Model |
| IPM | Integral Projection Model |
| IBM | Individual-Based Model |
| SSM | State-Space Model |
| PVA | Population Viability Analysis |

models (IBMs) are discussed in Section 1.6 and are contrasted with population based models. Section 1.7 is on state-space models (SSMs). Section 1.8 concludes the chapter with pointers to further literature on MPMs, IPMs, IBMs, and SSMs, comments on topics of demography that were minimally or not at all discussed, and thoughts about the future of biological demography.

There are a number of acronyms used repeatedly and for convenience some of the more frequently used ones are shown in Table 1.1.

### 1.2 Components of demography

At a very simple level, demography can be understood as a time series of the abundance of individuals in the population. To reduce notation the time series indexing is a sequence of integers at discrete evenly spaced points in time, $t=1,2, \ldots, T$. Abundances are denoted $n_{1}, n_{2}, \ldots, n_{T}$. A correspondingly simple mathematical model for the population dynamics relates abundance at time $t$ to abundance at time $t-1$ is the following

$$
\begin{equation*}
n_{t}=\lambda_{t} n_{t-1} \tag{1.1}
\end{equation*}
$$

where $\lambda_{t}$ is the population growth rate, a multiplier which if less than 1 indicates a decline in abundance and if greater than 1 an increase in abundance.

If the population abundance is small enough and easy enough to enumerate without error, not elusive or secretive or difficult to detect, then a succinct and completely accurate characterization of the population and its dynamics is trivial. For example, on July 1, over the period 2011-2015, in a 500 liter tank aquarium populated by angelfish (Pterophyllum spp), which did not receive any introduced fish, there were $n_{2011}=70, n_{2012}=61, n_{2013}=82, n_{2014}=$ 53 , and $n_{2015}=63$ fish. The annual population growth rates were $\lambda_{2012}=$
$0.87(61 / 70), \lambda_{2013}=1.34(82 / 61), \lambda_{2014}=0.65(53 / 82)$, and $\lambda_{2015}=1.19$ (63/53).

Such instances of exact enumeration are relatively rare or uninteresting in isolation. Complexity and more interesting aspects of demographic studies or population dynamics arise in several general ways:

1. Multiple subpopulations of $n_{t}$ : The population abundance at time $t$ can be partitioned or subdivided into two or more groups or subpopulations.
2. Multiple processes underly $\lambda_{t}$ : The mechanisms underlying the population growth rate $\lambda_{t}$ include a multitude of processes, including survival and reproduction.
3. Stochasticity: Environmental and demographic variability add uncertainty to projections of population dynamics.
4. Density dependence: The effect of processes can be affected by population size.
5. Competition and predation: The population dynamics of one population can be affected by other populations, of the same species or different species, can be affected in these two ways.
6. Manipulation of the dynamics: Related to the multiple components of $\lambda_{t}$, humans sometimes want to control or manipulate processes and thus affect the population dynamics.
7. Partial observability (Nichols et al., 1995) of $n_{t}$ : The population abundances cannot be enumerated exactly and at best subsets, or samples, of the population are observed.

Each of these complications are addressed individually below in Sections 1.2.1 - 1.2.7.

### 1.2.1 Multiple subpopulations

Finer characterization of the abundance involve partitioning the total number into numbers in subsets, or subpopulations, where the subsets are distinguished by one or more attributes. Some common distinguishing attributes are listed below.

- sex: female, male, or indeterminate.
- age: $0,1,2,3, \ldots$, max age.
- maturity level: young, immature but developing, mature (capable of reproduction), post-reproduction.
- size or size class: distinguished by weight, height, length, etc.
- spatial location: exact location in space (e.g., in two dimensions, latitude and longitude; in three dimensions, depth or height added).
- genotype: hereditary components that can be characterized at a variety of levels, e.g., allele combinations for a single gene or for two or more genes.
- phenotype: external characters that are a consequence of genotype and environment, e.g., color or color pattern (includes some of the above categories such as size class).

The general term structured population, or in specific cases age-structured or stage-structured populations, is commonly used in ecological literature. When the partitioning is in terms of some other factors, like sex or spatial location, however, the term structured population is less common. Spatially distinction populations are sometimes labeled metapopulations (Levins, 1969). Here the term multiple subpopulations will be used to include any partitioning of a population.

The degree and nature of the partitioning of a population may be a subjective determination, as it depends on who is studying the population and the objectives of analysis. For plentiful non-commercial species, the population partitioning might be coarse for all but scientists studying a particular aspect of the population. The partitioning might be fine for harvested populations distinguished by sex, age, and spatial locations, as well as for closely studied rare populations. Limitations on the available data about the population can limit the degree of partitioning, and restrict the type of population dynamics modeling that can be done.

Arbitrariness of partitioning may also occur when the distinguishing attributes are continuous variables, such as measures of individual size like weight, height, length. For example, if the variable is weight, the number of partitions can vary as can the labeling of the partitions; e.g., small $=<10$ $\mathrm{kg}, 10 \leq$ medium $<20 \mathrm{~kg}$, and large $\geq 20 \mathrm{~kg}$. The partitioning of continuous attributes is an important distinction between MPMs (Section 1.4) and IPMs (Section 1.5).

The finest partitioning of a population is at the individual entity level. For example, each animal is unique in its sex, date of birth (age), size, maturity level, number of offspring, location, and so on, and the values of the individual's characteristics throughout its entire existence, at any given point in time, are the most complete description possible. This is often at least conceptually possible (though not necessarily so for large subterranean vegetative populations), but usually not practically possible. A notable exception is the Soay sheep population on St Kilda Archipelago, off the coast of western Scotland, where near complete individual animal identification and labeling is done. As a mathematical exercise, however, the modeling of individuals in a population can be useful for elucidating population level dynamics and this will be discussed in Section 1.6 on IBMs.

### 1.2.2 Multiple processes

The population growth rate $\lambda_{t}$ in eq'n (1.1) can be a function on multiple processes, including survival, reproduction, immigration, emigration, growth, maturation, and movement between regions. If the population is defined for a specific geographic area, and other individuals located outside that area can move into the area, then immigration is a factor, and conversely, if individuals can move outside the area, then emigration is a factor.

The partitioning of the population affects the complexity of $\lambda_{t}$ with each additional partition adding at least one additional demographic parameter. For example, if the population is partitioned into 3 age classes, ages 0,1 , and $2+$, where $2+$ is age 2 and older, there could be three age-specific survival probabilities, contributions to reproduction, and different propensities to emigrate and immigrate. The reproduction process could be further distinguished on the basis of sex, or maturation class, or size. For some species, the number of females are more critical to reproductive success than the number of males, when a single male can mate with multiple females (are polygynous), e.g., elk (Cervus elaphus), and reproductive output might be viewed entirely as a function of female abundances. Maturation classifications can distinguish sexually immature and sexually mature individuals. Size of the animal can affect reproductive contribution; e.g., larger female fish will have more eggs than smaller fish.

Spatial partitioning of the population into metapopulations leads to the addition of a movement process. The probabilities of moving from one region to another region within the overall defined population boundaries can be affected by other categorizations of the population, e.g., age or sex or maturity level.

If size is a distinguishing characteristic, defined ordinally (e.g., small, medium, and large) or continuously (e.g., length in cm ), then growth is a process affecting dynamics. In particular, the individual growth dynamics are quantified in terms of the probability of moving from one size class to another (as in MPMs, section 1.4) or by a conditional probability density function for size $z_{t+1}^{\prime}$ given previous size $z_{t}$ (as in IPMs, section 1.5).

Whether or not processes occur sequentially, partially overlap, or are simultaneous affects the modeling of processes and the temporal partitioning of abundances. In the simplest case where processes occur sequentially and do not overlap, abundances for points in time immediately following the occurrence of a process need to be considered. For example, the life history of coho salmon starts with egg fertilization in freshwater, egg hatching and larval emergence, fry and then smolt stages while rearing in freshwater, migration to the ocean, and a period of time in the ocean followed by migration back to the freshwater, spawning, and death. Within for a single cohort, and letting $t$ be the time of egg fertilization, abundances that could be acknowledged include the number of eggs fertilized, $n_{t_{1}}$, the number of smolts two months later, $n_{t_{2}}$, the number of immature adults in the ocean just prior to return migration,
$n_{t_{3}}$, and the number of adults surviving to spawn, $n_{t_{4}}$. The life cycle of the species is thus guiding the temporal indexing of the abundance vector. When multiple cohorts overlap in time, the abundance vector becomes even more complicated as multiple cohorts will be at different life stages at a single point in time.

### 1.2.3 Stochasticity

Population dynamics are complicated by two types of stochasticity, environmental and demographic. Environmental stochasticity refers to between year (or any time period) variation in underlying vital rates, such as survival, reproduction, or the overall population growth rate. Such between year variation is typically due to variation in environmental conditions, e.g., precipitation and air temperature. For example, the underlying average probability that a juvenile salmon in a given stream will survive from 1 May to 1 June is 0.7 in one year, and 0.6 in the following year.

Demographic stochasticity, on the other hand, reflects inherent betweenindividual variability conditional on a specific vital rate. For example, in 2016, the underlying average May to June survival probability for the juvenile Chinook salmon is 0.7 and there are 1000 salmon present on 1 May. While the expected number surviving to 1 June is 700 , assuming independence between the fish which means the number surviving is a binomial random variable, the observed number surviving will with high probability (0.972) not equal 700 , with $95 \%$ probability the observed numbers will lie between 671 and 728 . That variation around 700 reflects demographic stochasticity.

Unless population numbers are relatively low, as for a severely endangered species, the influence of demographic stochasticity on population dynamics will be minor relative to the influence of environmental stochasticity. As population size increases, the deviation of the observed value from the expected value will be relatively small. For a survival probability of 0.7 , if $n=50$, the $95 \%$ interval is $(28,41)$ a relative range of $37 \%,(41-28) / 35$, but with $n=10,000$, the $95 \%$ interval is $(6910,7090)$ with a relative range of $2.6 \%$. The inherent survival probability for a group of animals is unlikely to be the same, however, even within a year. Kendall and Fox (2002) examined the effect of between individual variation in the survival probability within a year and showed that a binomial distribution based measure of demographic stochasticity can be an overestimate of true demographic stochasticity. A rule of thumb regarding whether or not to account demographic variation when doing PVA, provided by Morris et al. (2002), is that demographic variation can be ignored when just a single population is of interest and there are at least 100 individuals, and for PVA with multiple subpopulations, or life stages, there are at least 20 individuals in the most important subpopulations.

### 1.2.4 Density dependence

Consider a simple univariate population model similar to eq'n (1.1) but where a constant value of $\lambda$ is substituted for the $\lambda_{t} \mathrm{~s}$. If $\lambda>1$, the population will grow exponentially, and if $\lambda<1$ the population will go extinct. While exponential decline is possible, exponential growth cannot continue indefinitely as resources available for survival and reproduction, e.g., food, water, and space, are finite. For example, a population of plants introduced to a burned over acre of land may initially grow at a near exponential rate. As there is a limited amount of space for the plant to take root and grow, the population size will eventually reach an upper bound, and this bound is sometimes referred to a carrying capacity.

Limits on population growth can be viewed as limits on population processes like survival, reproduction, growth, and movement with these vital rates generally decreasing as population abundance increases. In other words these vital rates are abundance dependent; equivalently, scaling abundance by the available area or volume, the vital rates are density dependent.

Conversely as population abundance decreases, survival and reproduction rates typically increase, due to less competition for finite resources. However, there are situations where decreases beyond a certain lower bound can lead to decreases in vital rates. For example, if the numbers of animal population get so low that individuals have difficulty finding mates, reproductive rate will decline. This is known as an Allee effect, and can be a concern for critically endangered species.

### 1.2.5 Competition and predation

Denote a population and species of primary interest as species A, and populations of two different species as species B and C. If species B uses some of the same finite resources as species A, i.e., it is a competitor for those resources, and, similar to intra-specific density dependence, increases in the population size of species B can reduce survival, reproduction, growth, etc, of species A.

If individuals in the population of species C consumes members of species A, i.e., A is prey for C , then the abundance of species C affects the vital rates of species B, most obviously survival. Conversely, if individuals of species C are consumed by species A, namely A is a predator of C , and species C is a primary food source for A , a decline in the abundance of C can lead to decreases in vital rates of A.

### 1.2.6 Manipulation of dynamics

There are various motivations for manipulating the dynamics of a population. One is that the population is being used by humans. For example, portions of the population are being removed by fishing, hunting, or otherwise harvesting. In managed populations there are questions about the number that
can be removed from the population without causing irreparable damage, i.e., a sustainable harvest, or the harvest rate that yields the maximum sustainable harvest. The harvesting of a population can be quite selective and based on particular attributes, i.e., subpopulations. For example, only mature male moose (Alces alces) can be harvested during a summer time period, or only white sturgeon (Acipenser transmontanus) between 100 and 150 cm in length can be caught and kept, or morels (Morchella esculenta) can only be harvested in a designated area of a state park. In terms of characterizing population dynamics such removals are in effect components of the survival process for a specific subpopulation.

For endangered populations, manipulation of dynamics is done by regulating human activities and carrying out actions to increase the quality and quantity of the habitat of the population. In the case of populations harvested by humans, reductions in allowable harvest, including complete harvest bans, are often mandated by regulatory agencies. Such actions translate into effects on survival probabilities for potentially different subpopulations. Habitat alterations can affect survival, reproduction, growth, maturation, and movement. Projections of the effects of such manipulations, and those of factors that are not manipulable, on processes and subsequent population dynamics is central to PVA.

### 1.2.7 The role of data: Partial observability

In most cases abundances of a population or subpopulations are unknown. Some sort of sampling is required and many methods have been developed since the early 1900s to produce estimates of abundance, as well as estimates of process related parameters like survival, reproductive success, and movement. For abundance estimation, the simplest methods are based on classic sample survey designs and related estimators (Thompson, 2012). These can be used when the area occupied by the population can be subdivided into a set of sampling units, thus constructing a sample frame. Largely non-mobile populations, such as plants, are more amenable to such methods; e.g., the population of interest is an invasive non-native grass on a wildlife refuge and the refuge can be subdivided into 100,000 square meter plots, a simple random sample of 50 plots is drawn, and appropriate measures of grass biomass, say, are made on the selected plots. For more mobile populations, abundances can be estimated using mark-recapture methods (Williams et al., 2002), line transect and point transect (Buckland et al., 2001) and presence/absence sampling (occupancy modeling, MacKenzie et al., 2005). See Chapters X, Y, and Z for further discussion of abundance estimation procedures.

The fact that only estimates of abundance are available has two effects on the modeling of population dynamics. One obvious effect is that additional uncertainties about abundances and processes are introduced. Suppose abundances are estimated annually at the same time of the year. Population growth rate, $\lambda_{t}$, can be estimated by $\hat{\lambda}_{t}=\hat{n}_{t} / \hat{n}_{t-1}$, and uncertainty in the estimate
is a function of uncertainty in the abundance estimates. A second, and less obvious, effect is that mathematical model formulation can be affected by the time at which the estimates are made. Rees et al. (2014) give an example of a sequence of processes, beginning with reproduction, then mortality, and then growth. If abundance estimates are made just before reproduction, $\lambda_{t}$ includes a term for the probability of the previous year's reproduction (recruits) living an entire year. If abundance estimates are made just after reproduction, the $\lambda_{t}$ does not reflect the survival of this year's reproduction as the estimates were made before subsequent mortality, and the survival of the previous year's reproduction is entangled with the survival of the previous year's abundance of old entities (non-recruits). Inserting additional sampling or estimation points in the year, say, is one means of disentangling the effects of multiple processes, but partial observability remains a factor.

### 1.3 General mathematical features of PDMs

Demographic models can be classified by the factors given in Section 1.2 such as the number and types of subpopulations, the number and types of processes, type of stochasticity, and degree of density dependence. In this section we present various mathematical and probabilistic formulations of such demographic models. We begin with the simplest demographic models, a single population, a single process, deterministic, and density independent. Viewing the growth rate parameter in eq'n (1.1) as deterministic means that $\lambda_{t}$ is by definition $n_{t} / n_{t-1}$. For convenience this model is shown again.

$$
n_{t}=\lambda_{t} n_{t-1}
$$

If $\lambda_{t}=\lambda, n_{t}=\lambda^{t} n_{0}$. As noted previously, if $\lambda>1$, population will grow exponentially.

### 1.3.1 Multiple subpopulations

Partitioning a single population into two or more populations extends the scalar $n_{t}$ to a vector $\mathbf{n}_{t}$. For example, if a population of deer is distinguished by three life stages, young, immature, and mature, then the abundance vector is

$$
\left[\begin{array}{c}
n_{y, t} \\
n_{i, t} \\
n_{m, t}
\end{array}\right]
$$

The length of the abundance vector over time need not remain fixed. The ef-
reproduction, or to shrink, following an aggregation of age classes (Buckland et al., 2007).

### 1.3.2 Multiple processess

Survival and reproduction.
Simple single population growth models can be modified by partitioning the population growth rate into survival and reproduction processes. One way of doing so is a balance equation with additions and subtractions.

$$
\begin{equation*}
n_{t}=n_{t-1}+R_{t}-M_{t} \tag{1.2}
\end{equation*}
$$

where $R_{t}$ is the number of (surviving) young produced between $t-1$ and $t$ and $M_{t}$ is the number of mortalities from the number $n_{t-1}$. Such additive formulations do not make clear the relationship between the previous abundance, $n_{t-1}$, and reproduction and mortality, however. The effect of previous abundance on reproduction and mortality can be made clearer with a multiplicative model which extends eq'n (1.1). Assume that the survival, a fraction $\phi_{t}$, precedes reproduction rate, $\rho_{t}$, and there is no mortality of the young between the time of reproduction and time $t$.

$$
\begin{equation*}
n_{t}=\left(1+\rho_{t}\right) \phi_{t} n_{t-1} \tag{1.3}
\end{equation*}
$$

Implicit to this formulation is that the survival rates for the young and adult components which made up $n_{t-1}$ was the same, $\phi_{t}$. Given this assumption about survival, the model remains the same even if the abundance denotes the number just prior to reproduction.

A more realistic model has different survival fractions for just born young and the older individuals, $\phi_{y, t}$ and $\phi_{a, t}$, and the model should have subpopulations for young and adult. The subpopulations for young will be those just born if the time period is just after reproduction, denoted $n_{0, t}$, while, if the time period is just before reproduction, the young will be nearly age 1 (if the time interval is one year), denoted $n_{1, t}$.

$$
\begin{align*}
t \text { just after reproduction } & {\left[\begin{array}{l}
n_{a, t}=\phi_{a, t} n_{t-1}+\phi_{y, t} n_{0, t-1} \\
n_{0, t}=\rho_{t}\left(\phi_{a, t} n_{t-1}+\phi_{y, t} n_{0, t-1}\right)
\end{array}\right] }  \tag{1.4}\\
t \text { just before reproduction } & {\left[\begin{array}{l}
n_{a, t}=\phi_{a, t} n_{t-1} \\
n_{1, t}=\phi_{y, t} \rho_{t} n_{t-1}
\end{array}\right] } \tag{1.5}
\end{align*}
$$

Immigration and emigration.
Immigration and emigration can be included in the dynamics. Extending the balance equation model,

$$
\begin{equation*}
n_{t}=n_{t-1}+R_{t}-M_{t}+I_{t}-E_{t} \tag{1.6}
\end{equation*}
$$

where $I_{t}$ is the number immigrating into the population and $E_{t}$ is the number emigrating from the population. The multiplicative model can be extended
but does not necessarily remain multiplicative. The ordering of processes is again important. Assuming that immigrants arrive, and emigrants leave after mortality occurs, but prior to reproduction, the model is

$$
\begin{equation*}
n_{t}=\left(1+\rho_{t}\right)\left(\zeta_{t} \phi_{t} n_{t-1}+I_{t}\right) \tag{1.7}
\end{equation*}
$$

where $\zeta_{t}$ is the fraction of the survivors from $n_{t-1}$ that stay and $I_{t}$ is again the number of immigrants. If the order of processes change, the model changes. For example, suppose that immigrants arrive and emigrants leave after mortality and reproduction:

$$
\begin{equation*}
n_{t}=\zeta_{t} \rho_{t} \phi_{t} n_{t-1}+I_{t} \tag{1.8}
\end{equation*}
$$

## Movement.

When components of the population are distinguished by spatial location, sometimes called metapopulations, the process of movement becomes relevant. Immigration and emigration is of course a movement process but where the individuals are coming from (immigrants) or going to (emigrants) are not distinguished. With multiple locations there is often a movement transition matrix which may or may not be time invariant. For example with three regions labeled $\mathrm{A}, \mathrm{B}$, and C , a time invariant transition matrix, $M$, has the following structure.
$M=\left[\begin{array}{c|cccc} & & t+1 & & \\ t & \rightarrow A & \rightarrow B & \rightarrow C & \\ \hline A & 1-\mu_{A \rightarrow B}-\mu_{A \rightarrow C} & \mu_{A \rightarrow B} & \mu_{A \rightarrow C} & (1.9) \\ B & \mu_{B \rightarrow A} & 1-\mu_{B \rightarrow A}-\mu_{B \rightarrow C} & \mu_{B \rightarrow C} & \\ C & \mu_{C \rightarrow A} & \mu_{C \rightarrow B} & 1-\mu_{C \rightarrow A}-\mu_{C \rightarrow B}\end{array}\right]$
where the $\mu \in[0,1]$ and the columns sum to 1 .
Growth.
When subpopulations are defined in terms of size categories, movement between size classes could be defined in terms of the fractions of a given size category moving from one category to another. The process is analogous to that for movement between spatial regions.

In the case of finest scale partitioning of populations to the individual entity level, growth from one size at time $t, z_{t}$, to another size at time $t+1$, $z_{t+1}$, could be modeled by the addition of an individual growth increment, $x_{t+1}$,

$$
\begin{equation*}
z_{t+1} \mid z_{t}=z_{t}+x_{t+1} \tag{1.10}
\end{equation*}
$$

${ }^{474} x_{t+1}$ could be a function of the size at time $t, z_{t}$. Such fine scale handling of

### 1.3.3 Stochasticity

As discussed previously, environmental stochasticity is often more important to account for than demographic stochasticity so long as the population is not too small. An example of the mathematical distinction between the two forms is presented by modifying the simple population model in equation (1.1). Assume that the underlying annual population growth rate does not vary between years, namely, does not have environmental stochasticity, and denote that rate $\lambda$. However, there is between-individual variation in the growth rate contribution. A demographic stochasticity extension to eq'n (1.1) is

$$
\begin{equation*}
n_{t} \mid n_{t-1} \sim \operatorname{Poisson}\left(n_{t-1} \lambda\right) \tag{1.11}
\end{equation*}
$$

Environmental stochasticity alone is reflected by population growth rates that vary between years but the per individual contribution to the growth rate is the same for all individuals. An example model is

$$
\begin{align*}
\lambda_{t} & \sim[\operatorname{Gamma}(\alpha, \beta)],  \tag{1.12}\\
n_{t} \mid n_{t-1} & =\lambda_{t} n_{t-1}
\end{align*}
$$

where [ ] denotes the nearest integer function. An example of a mathematically convenient alternative formulation for environmental stochasticity, which ignores the discrete nature of many populations is the following:

$$
\begin{equation*}
n_{t} \mid n_{t-1} \sim \text { Lognormal }\left(\log \left(n_{t-1} \lambda\right)-0.5 \sigma_{\epsilon}^{2}, \sigma_{\epsilon}^{2}\right) \tag{1.13}
\end{equation*}
$$

The term $0.5 \sigma_{\epsilon}^{2}$ is a bias correction which ensures that the expected abundance, $E\left[n_{t} \mid n_{t-1}\right]$, is $n_{t-1} \lambda$.

Environmental and demographic variation typically coincide. One way to demonstrate this is with a hierarchical model:

$$
\begin{align*}
\lambda_{t} & \sim \operatorname{Gamma}(\alpha, \beta)  \tag{1.14}\\
n_{t} \mid n_{t-1}, \lambda_{t} & \sim \operatorname{Poisson}\left(n_{t-1} \lambda_{t}\right) \tag{1.15}
\end{align*}
$$

Specific ways of introducing demographic and environmental stochasticity to other processes, e.g., survival, are described later.

Asymptotic results for environmentally stochastic growth rates.
The long term, or asymptotic, behavior of a single population trajectory with environmentally stochastic annual growth rates, such as in eq'n (1.12) is tractable and has similarities with deterministic exponential growth models. Consider the following general single population model:

$$
\begin{equation*}
n_{t}=\lambda_{t} n_{t-1}, \text { where } \lambda_{t} \stackrel{i i d}{\sim} \operatorname{Distribution}\left(\mu, \sigma^{2}\right) \tag{1.16}
\end{equation*}
$$

where $E\left[\lambda_{t}\right]=\mu$ and $V\left[\lambda_{t}\right]=\sigma^{2}$. Given an initial abundance $n_{0}, n_{t}$ can be rewritten as

$$
n_{t}=n_{0} \prod_{i=1}^{t} \lambda_{i}
$$

Taking the natural logarithm of both sides of the equation,

$$
\ln \left(n_{t}\right)=\ln \left(n_{0}\right)+\sum_{i=1}^{t} \ln \left(\lambda_{i}\right)
$$

which can be re-expressed as

$$
\begin{equation*}
\frac{\ln \left(n_{t}\right)-\ln \left(n_{0}\right)}{t}=\frac{1}{t} \sum_{i=1}^{t} \ln \left(\lambda_{i}\right) \tag{1.17}
\end{equation*}
$$

The righthand side of eq'n (1.17) is the mean of a sequence of independent random variables, $\ln \left(\lambda_{i}\right), i=1, \ldots, t$. Adding the assumption that the $E\left(\ln \left(\lambda_{t}\right)^{2}\right)<\infty$, the strong law of large numbers says that the average converges to $E[\ln (\lambda)]$. Further, by the Central Limit Theorem, the asymptotic distribution the mean of the log of the "annual" growth rates is normal. Denoting the sample average $\log$ growth rate by $\overline{\ln (\lambda))}$

$$
\begin{equation*}
\overline{\ln (\lambda))}=\frac{1}{t} \sum_{i=1}^{t} \ln \left(\lambda_{i}\right) \quad \sim \quad \text { Asymptotic Normal }(E[\ln (\lambda)], V(\ln (\lambda)))(1 \tag{1.18}
\end{equation*}
$$

Another way to express this result, using the lefthand side of eq'n (1.17),

$$
\begin{equation*}
\ln \left(n_{t}\right) \sim \text { Asymptotic Normal }\left(\ln \left(n_{0}\right)+t E[\ln (\lambda)], t V(\ln (\lambda))\right) \tag{1.19}
\end{equation*}
$$

or

$$
\begin{equation*}
n_{t} \sim \text { Asymptotic Lognormal }\left(n_{0} \exp (t E[\ln (\lambda)]), t V(\ln (\lambda))\right) \tag{1.20}
\end{equation*}
$$

Thus, the average population abundance at $t$ is identical to a deterministic exponential growth model.

Survival.
Environmental and demographic variation in the survival fractions can be modeled with the same hierarchical structure used for population growth (eq'ns 1.14 and 1.15). One such combination is a logit-normal model for environmental stochasticity and a binomial distribution for demographic stochasticity. Letting $\phi_{c, t}$ be the survival probability for subpopulation $c$ at time $t$

$$
\begin{align*}
\operatorname{logit}\left(\phi_{c, t}\right) & \sim \operatorname{Normal}\left(\beta_{0, \phi, c}, \sigma_{\phi, c}^{2}\right)  \tag{1.21}\\
n_{c, t} & \sim \operatorname{Binomial}\left(n_{c, t-1}, \phi_{c}\right)
\end{align*}
$$

where $\operatorname{logit}(x)=\ln (x /(1-x))$. Beta distributions for the $\phi_{c, t}$ are alternative for environmental stochasticity,

$$
\phi_{c, t} \sim \operatorname{Beta}\left(\alpha_{\phi, c}, \beta_{\phi, c}\right)
$$

but this distribution may be more awkward when it comes to including co-

## Reproduction.

One way to characterize reproduction is in terms of the number of progeny produced by a given individual in the population, perhaps only mature and female individuals. Letting $\rho_{c, t}$ be the expected progeny multiplier for subpopulation $c$ at time $t$, a hierarchical model for both levels of stochasticity is a Lognormal and Poisson combination.

$$
\begin{align*}
\rho_{c, t} & \sim \text { Lognormal }\left(\ln \left(\mu_{c}\right)-0.5 \sigma_{\mu, c}^{2}, \sigma_{\mu, c}^{2}\right)  \tag{1.22}\\
n_{c, 0, t} & \sim \operatorname{Poisson}\left(\rho_{c, t} n_{c, t}\right)
\end{align*}
$$

where $n_{c, 0, t}$ is the number of progeny produced. An alternative distribution for $\rho_{c, t}$ is the Gamma,

$$
\rho_{c, t} \sim \operatorname{Gamma}\left(\alpha_{\rho}, \beta_{\rho}\right)
$$

but again the inclusion of covariates may be more awkward than for the lognormal.

## Immigration and Emigration.

Only the case of multiplicative survival, reproduction, and emigration processes with additive immigration is considered and the only partitioning of the population is of adults and young. Immigrants are generated from an arbitrary non-negative integer valued distribution, here a Poisson will be used. The hierarchical structure for survival and reproduction used previously (1.21 and 1.22 ) is used again and the emigration probability is dealt with the same as the survival probability. The order of processes is survival, reproduction, emigration, and immigration where just born individuals do not emigrate. The notation $t^{\prime}$ denotes time just after survival but before emigration.

$$
\begin{align*}
\text { Survival probability } & \operatorname{logit}\left(\phi_{t}\right) \sim \operatorname{Normal}\left(\beta_{\phi}, \sigma_{\phi}^{2}\right)  \tag{1.23}\\
\text { Survivors } & n_{t^{\prime}} \sim \operatorname{Binomial}\left(n_{t-1}, \phi_{t}\right)  \tag{1.24}\\
\text { Reproductive rate } & \rho_{t} \sim \operatorname{Lognormal}\left(\ln \left(\beta_{\rho}\right)-0.5 \sigma_{\rho}^{2}, \sigma_{\rho}^{2}\right)  \tag{1.25}\\
\text { Reproduction } & n_{c, 0, t} \sim \operatorname{Poisson}\left(\rho_{t} n_{t^{\prime}}\right)  \tag{1.26}\\
\text { Emigration probability } & \operatorname{logit}\left(\zeta_{t}\right) \sim \operatorname{Normal}\left(\beta_{\zeta}, \sigma_{\zeta}^{2}\right)  \tag{1.27}\\
\text { Non-emigrants } & n_{t} \sim \operatorname{Binomial}\left(n_{t^{\prime}}, \zeta_{t}\right)  \tag{1.28}\\
\text { Immigrants } & I \sim \operatorname{Poisson}(\eta) \tag{1.29}
\end{align*}
$$

Movement.
Environnmental stochasticity can be introduced by randomly drawing a vector of movement probabilities for each row of the movement matrix, eq'n (1.9), and then conditional on these probabilities the numbers moving to each location are drawn at random. One hierarchical formulation is a multi-vector
extended logit-multivariate normal (Newman and Brandes, 2010) and multinomial combination. Just one vector of the formulation, the case of being in location $A$ at time $t$, is shown here, and subscripting notation for parameters indicating time are omitted.

$$
\begin{align*}
{\left[\begin{array}{l}
\ln \left(\frac{\mu_{A \rightarrow B}}{1-\mu_{A \rightarrow B}-\mu_{A \rightarrow C}}\right) \\
\ln \left(\frac{\mu_{A \rightarrow C}}{1-\mu_{A \rightarrow B}-\mu_{A \rightarrow C}}\right)
\end{array}\right] } & \sim \operatorname{MVN}\left(\left[\begin{array}{l}
\beta_{0, \mu_{A}, B} \\
\beta_{0, \mu_{A}, C}
\end{array}\right],\left[\begin{array}{cc}
\sigma_{A, A \rightarrow B}^{2} & \sigma_{A, A \rightarrow B, A \rightarrow C_{1}} \\
\sigma_{A, A \rightarrow B, A \rightarrow C} & \sigma_{A, A \rightarrow C}^{2}
\end{array}\right]\right) \\
n_{A \rightarrow B, t}, n_{A \rightarrow C, t} & \sim \operatorname{Multinomial}\left(n_{A, t-1}, \mu_{A \rightarrow B}, \mu_{A \rightarrow C}\right) \tag{1.31}
\end{align*}
$$

An alternative distribution for environmental stochasticity for the $\mu$ 's is a Dirichlet distribution but it does not necessarily lend itself as easily to the inclusion of covariates as the extended logit-MVN.

## Growth.

As for the deterministic case, when growth refers to movement from a categorical size class, a stochastic representation of the advancement is again analogous to the handling of between region movement. In the case of individual entities, assuming that growth is never negative, a stochastic formulation is a Gamma probability density function for the increment in growth:

$$
z_{t} \mid z_{t-1}=z_{t-1}+x_{t}
$$

where

$$
x_{t} \sim \operatorname{Gamma}(\alpha, \beta)
$$

with parametric formulation being $E\left[x_{t}\right]$ being $\alpha \beta$. The parameters $\alpha$ and $\beta$ could be individual specific, say functions of other covariates, or draws from another probability distribution, thus introducing additional stochasticity.

### 1.3.4 Density dependence

A variety of univariate models with density dependent population growth rates, $\lambda\left(n_{t}\right)$, are commonly used. We begin with some well known deterministic, and mostly discrete time, models. These univariate models are readily applied to individual subpopulations of a structured population, but applications where the subpopulations interact, as in predator-prey situations or spatially partitioned partitions, are more involved.

A slight variation on the exponential growth model is the Gompertz model, where the following formulation is taken from Dennis et al. (2006).

$$
\begin{equation*}
\text { Gompertz model : } n_{t}=\lambda n_{t-1}^{1+\alpha} \tag{1.32}
\end{equation*}
$$

where $\alpha<0$ is the parameter inducing density dependence. Substituting $n$ for $n_{t}$ and $n_{t-1}$ the equilibrium abundance for the Gompertz distribution is

$$
n_{\text {Gompertz,e }}=\lambda^{\frac{-1}{\alpha}}
$$

Another well known model is the logistic model, which has discrete and continuous time forms:

$$
\begin{align*}
\text { Logistic (discrete) model } & : \quad n_{t}=n_{t}+\lambda n_{t}\left(1-\frac{n_{t}}{K}\right)  \tag{1.33}\\
\text { Logistic (continuous) model } & : \quad n(t)=\frac{K}{1+\frac{K-n(0)}{n(0)} \lambda^{t}} \tag{1.34}
\end{align*}
$$

In both cases $K$ is the upper bound on the population abundance, i.e., the carrying capacity.

A well-known model that originated with fish populations, but is now applied many other kinds of populations, is the Ricker model.

$$
\begin{equation*}
\text { Ricker model : } \quad n_{t}=\left(\phi_{a}+\phi_{y} b \exp \left(-c n_{t-1}\right)\right) n_{t-1} \tag{1.35}
\end{equation*}
$$

The Ricker model formulation citep[taken from][]gurney1998ecological can be viewed as a balance equation where the product $\phi_{a} n_{t-1}$ is the number of surviving adults and $\phi_{y} f \exp \left(-c n_{t-1}\right)$ is the number of survivng offspring, the recruits. The term $b$ is the fecundity of the average adult (with sex being ignored here; if the sex ratio if $50: 50, b$ could equal half the average female fecundity), $\exp \left(-c n_{t-1}\right)$ is a density dependent reproductive success rate which decreases as abundance increases and reproductive success is the number alive at some point in time prior to time $t$, and $\phi_{y}$ is the survival probability from that time point to $t$. Combining the product $f$ and $\phi_{y}$ into a single term has no effect on the model dynamics. Substituting $n$ for $n_{t}$ and $n_{t-1}$ the equilibrium abundance for the Ricker model is

$$
n_{\text {Ricker }, e}=\frac{-1}{c} \ln \left(\frac{1-\phi_{a}}{\phi_{y} b}\right)
$$

Another well known model from fisheries science is the Beverton Holt model, again taking the formulation from citepgurney1998ecological.

$$
\begin{equation*}
\text { Beverton Holt }: \quad n_{t}=\left(\phi_{a}+\frac{\phi_{y} b}{1+c n_{t-1}}\right) n_{t-1} \tag{1.36}
\end{equation*}
$$

The parameters $\phi_{a}, \phi_{y}$, and $b$ have the same interpretation as for the Ricker model, and the parameter $c$ is again the density dependent parameter. The equilibrium abundance is

$$
n_{\text {BevertonHolt,e }}=\frac{\phi_{y} b\left(1-\phi_{a}\right)}{\left(1-\phi_{a}\right) c}
$$

Figure 1.1 shows population growth trajectories for these five models such that the equilibrium abundances are nearly identical (around 800) and $\phi_{a}$ is 0.7 and $\phi_{y}$ is 0.2 for both the Ricker and Beverton Holt models. The Ricker model can overshoot the equilibrium abundance while the other models converge monotonically.


FIGURE 1.1
Population trajectories with initial abundance $=50$ for Gompertz, Ricker, and Beverton Holt growth models. Equilibrium abundances are around 800 for all three models and $\phi_{a}$ is 0.7 and $\phi_{y}$ is 0.2 for both the Ricker and Beverton Holt models.

The density dependent formulations for population growth $\lambda_{t}$ in the Ricker and Beverton Holt models can be seen to be simply density dependent reproduction rates as opposed to density dependent survival rates by simply removing the number of adults and the adult survival term $\phi_{a}$ from the model. Then the reproductive rate multipliers are:

$$
\begin{array}{r}
\text { Ricker reproduction rate }:\left(\phi_{y} b\right) \exp \left(-c n_{t-1}\right) \\
\text { Beverton Holt reproduction rate }
\end{array}
$$

The difference in the density dependent effects for the two models can be seen in the terms following $\left(\phi_{y} b\right)$. Density dependence in the Gompertz model cannot be disentangled into effects on reproduction and survival as $\lambda$ is not being adjusted. Instead the per capita contribution of each individual to the population growth rate is being diminished abundance increases.

Density dependence at a subprocess level.
Density dependence can be introduced for other vital rates such as survival, movement, and individual animal growth rates. An example formulation of density dependence and environmental stochasticity is a logit-normal distribution with population abundance as a covariate:

$$
\operatorname{logit}\left(\phi_{t}\right) \sim \operatorname{Normal}\left(\beta_{0, \phi}+\beta_{1, \phi} n_{t-1}, \sigma_{\phi}^{2}\right)
$$

where $\beta_{1, \phi}$ is constrained to be negative. Similar adjustments can be made for movement probabilities between spatially distinct subpopulations, an example is given shortly.

Density dependence for multiple (sub)populations.
If the vital rates for one subpopulation do not affect another subpopulation, then the above univariate density dependent models can be applied on a per subpopulation basis for subpopulation specific vital rates. More realistically, subpopulation abundances will impact vital rates for other subpopulations, if the subpopulations occupy the same geographic area this is likely obvious as it is the total resource consumption or usage that matters.

Even when subpopulations are defined by spatially distinct locations, the density in one spatial subpopulation can affect another spatial subpopulation, e.g., one region is becoming crowded and individuals leave to seek less densely populated regions, thus movement probabilities are affected by density dependence. Referring to stochastic movement example given previously, eq'n (1.30), adjustments could be made to the mean parameters in the multivariate normal model. Movement from one location to another is made a function of the relative densities in the originating region and the destination region. In the following formulation, assuming that the slopes ( $\beta_{1}$ 's) are positive valued, the probability of moving from $A$ to either $B$ or $C$ increases as the density in
where $\delta_{x}=n_{x} /\|x\|$ and $\|x\|$ is the area of $x$.
Density dependence, predation and competition
The dynamics of predator-prey interactions and between species competition are implicitly density dependent. For example, as the abundance (density) of a predator increases survival probabilities for the prey population decrease. One of the best known, and simpler, mathematical models for predator-prey dynamics is the Lotka-Volterra model which is a continuous time model defined in terms of differential equations for the instantaneous population growth rate citep[][p. 155]gurney1998ecological:
$\frac{d \operatorname{Prey}}{d t}=(r-\alpha \operatorname{Predator}(t)) \operatorname{Prey}(t) \quad \frac{d \text { Predator }}{d t}=(\epsilon \alpha \operatorname{Prey}(t)-\delta) \operatorname{Predator}(t)$
Following convention, the indexing for continuous time models is in parentheses $n(t)$ in contrast to discrete time indexing with subscripts $n_{t}$.

Referring to the prey equation, in the absence of predators, the instantaneous growth rate of the prey population is $r$ and the population trajectory is exponential growth, $\operatorname{Prey}(t)=\exp (r) \operatorname{Prey}(t-1)$. The addition of predators to the population is causing mortality at rate $\alpha \operatorname{Predator}(t)$, where $\alpha$ is the per capita consumption rate of prey by predators. For predator dynamics, the mortality rate is a constant independent of predator abundance; in the absence of prey, the population trajectory is exponential decline, $\operatorname{Predator}(t)=\exp (-\delta) \operatorname{Predator}(t-1)$. Predator growth is entirely dependent on the presence of prey with instantaneous growth rate $\epsilon \alpha \operatorname{Prey}(t)$, where $\epsilon$ is the per prey item predator offspring rate. If $\operatorname{Prey}(t)$ was a constant, Prey*, the resulting dynamics would be exponential, Predator $(t)=$ $\exp \left(\epsilon \alpha\right.$ Prey $\left.{ }^{*}-\delta\right)$ Predator $(t-1)$. The Lotka-Volterra model is considered a relative simple, and unrealistic model (Gurney and Nisbet, 1998, p. 159), but the resulting dynamics can be relatively complex with oscillating population trajectories. See Gurney and Nisbet (1998) for examples of the trajectories and dynamics as well as extensions, such as a prey model where prey abundance also affects the population growth rate (self-limiting prey).

## Periodicity and chaos.

The periodicity observed in the Lotka-Volterra model, and much more complex dynamics, can occur with discrete time single population models like the Ricker and discrete logistic model. Figure 1.3.4 Gurney and Nisbet (based on Figure 2.6 in 1998) shows what can happen for a variaty of combined recruitment rate values, $\phi_{j} f$, when adult survival, $\phi_{a}$, is zero and the density


FIGURE 1.2
Population trajectories for the Ricker model for a variety of $\phi_{j} f$ combinations with $\phi_{a}=0$ and $c=0.001$.
dependent parameter $c=0.001$. With too little recruitment (a), the population goes extinct, (b) shows a stable equilibrium, (c) and (d) show damped and expanding oscillations, respectively, (e) and (f) have two and four year periodicities, respectively, and (g) is chaotic (no periodicity and apparently random fluctuations).

### 1.3.5 Inclusion of covariates

Similar to density dependent models where abundance was used as a covariate to modify vital rates, environmental and anthropogenic covariates can be included; e.g., bird survival modeled as a function of weather data (North and Morgan, 1979).

Covariates can also be used as a means of assessing the effects of predators, competitors, or prey abundance when the populations' dynamics of predators, competitors, or prey are not modeled. Instead the abundances of these groups are simply treated as fixed input values that affect the vital rates for the population of interest. The legitimacy of such handling of these other populations assumes that these other populations are not receiving any feedback from the population of interest. If the primary population's abundance has relatively little effect on the abundances of the other populations, for example, the primary population is a minor food item for predators, then treating these other
population abundances as covariates may not adequate. Otherwise the more complex models for the joint dynamics, as in the Lotka-Volterra model, would be more appropriate.

### 1.3.6 Simultaneous versus sequential processes

The modeling of multiple processes in previous examples, e.g., eq'ns (1.231.29 ), assumed that the processes occur in a specific sequence. Such a tidy partitioning of processes may be a poor approximation to reality for some populations. In the case of harvested populations, for example, harvest-related mortality is co-occurring with other sources of mortality. One of the simpler treatments of such simultaneous mortality is frequently used in fisheries science. The underlying differential equation in the simplest case of two sources of mortality, say fishing and natural mortality, is the following:

$$
\begin{equation*}
\frac{d n(t)}{d t}=-(F+M) n(t) \tag{1.37}
\end{equation*}
$$

where $F$ and $M$ are instantaneous fishing and natural mortality rates. The solution to the differential equation is

$$
\begin{equation*}
n(t+\Delta)=\exp (-(F+M) \Delta) n(t) \tag{1.38}
\end{equation*}
$$

where $\exp (-(F+M) \Delta)$ is the survival fraction over a time interval of length $\Delta$. The number that die in total and from each source is then

$$
\begin{align*}
\text { Total Mortality } & =(1-\exp (-(F+M) \Delta) n(t)  \tag{1.39}\\
\text { Fishing Mortality } & =\frac{F}{F+M}(1-\exp (-(F+M) \Delta) n(t)  \tag{1.40}\\
\text { Natural Mortality } & =\frac{M}{F+M}(1-\exp (-(F+M) \Delta) n(t) \tag{1.41}
\end{align*}
$$

Formulations and simultaneous and competing sources of mortality can be translated into discrete time population dynamics models in various ways. For example, a hierarchical model for survival which includes environmental and demographic stochasticity and covariates can be constructed.

$$
\begin{align*}
\ln \left(F_{t}\right) & \sim \operatorname{Normal}\left(\beta_{0, F}+\beta_{1, F} x_{1, t}, \sigma_{F}^{2}\right)  \tag{1.42}\\
\ln \left(M_{t}\right) & \sim \operatorname{Normal}\left(\beta_{0, M}+\beta_{1, M} x_{2, t}, \sigma_{M}^{2}\right)  \tag{1.43}\\
n_{t} & \sim \operatorname{Binomial}\left(n_{t-1}, \exp \left[-\left(F_{t}+M_{t}\right)\right]\right) \tag{1.44}
\end{align*}
$$

The mortality due to each source between time $t$ and $t+1$ can be predicted using eq'ns (1.40) and (1.41) with $\Delta=1$.

### 1.3.7 Remarks: Estimability and Data Collection.

It is relatively easy to formulate a population dynamics model where the parameters cannot be estimated given the available data. For example, annual
surveys alone do not allow separate estimation of the survival fraction, $\phi_{t}$, and reproductive rate, $\gamma_{t}$, in the simple univariate model (1.3). Intuitively given estimates of $n_{t}$ and $n_{t-1}$ one can just estimate the combination $\left(1+\gamma_{t}\right) \phi_{t}$.

One way to disentangle such combinations of parameters, in the case of sequential processes, is to have abundance estimates at time points immediately after the end each process. For example, in the model (1.3) abundance should be measured twice a year, once immediately following the survival process, and once after reproduction. The reality of the processes is typically more complicated with such sharp demarcations unlikely, but formulating such models can provide guidance for data collection.

Detailed discussion of issues of estimability, or parameter redundancy, for population dynamics-related models can be found in Catchpole and Morgan (1997, 2001); Cole and Morgan (2010); Cole et al. (2012).

### 1.4 Matrix Projection Models, MPMs

Lewis (1942) and Leslie (1945) independently proposed matrix projection models as a means of modeling the population dynamics of populations with age-specific subpopulations, more commonly referred to as age-structured models. Let $n_{0, t}$ denote the number of young at time $t$ and $n_{a, t}$ be the abundance for ages 1 to $A-1$, and $n_{A+, t}$ be the abundance of age $A$ and older individuals. A deterministic formulation for the dynamics can be written as

$$
\left[\begin{array}{c}
n_{0, t}  \tag{1.45}\\
n_{1, t} \\
n_{2, t} \\
\vdots \\
n_{A, t} \\
n_{A+, t}
\end{array}\right]=\left[\begin{array}{cccccc}
\gamma_{0} & \gamma_{1} & \gamma_{2} & \ldots & \gamma_{A-1} & \gamma_{A} \\
\phi_{1} & 0 & 0 & \ldots & 0 & 0 \\
0 & \phi_{2} & 0 & \ldots & 0 & 0 \\
\vdots & 0 & 0 & \ldots & \phi_{A} & \phi_{A+}
\end{array}\right]\left[\begin{array}{c}
n_{0, t-1} \\
n_{1, t-1} \\
n_{2, t-1} \\
\vdots \\
n_{A, t-1} \\
n_{A+, t-1}
\end{array}\right]
$$

or more compactly as $\mathbf{n}_{t}=L \mathbf{n}_{t-1}$. Lefkovitch (1965) proposed an MPM where the subpopulations were distinguished by life stage rather than age class, e.g., young, immature, and mature, thus a stage-structured model. Of course, partitioning by gender, genotype, and many other subpopulation identifiers is possible.

### 1.4.1 Analysis of MPMs.

Caswell (2001, p. 18) refers to four sets of general questions which can be asked about MPMs for which the answers can elucidate deeper understanding of the dynamics. His questions (paraphrased and italicized) are shown below.

1. What is the asymptotic behavior of the MPM? As time increases, does the
total population grow or decline exponentially? Do the relative proportions of each subpopulation become constant? Does the population approach an upper bound (carrying capacity)? Do the total population and individual subpopulation abundances oscillate (in a damped or undamped manner)? Do the abundances display periodicity? Do the abundances become chaotic?
2. Is the MPM ergodic? In other words, are the asymptotic dynamics independent of the initial conditions, e.g., independent of the actual values of $\mathbf{n}_{0}$ ?
3. What are the transient dynamics? What are the dynamics like in the short term as opposed to the asymptotic or limiting results?
4. How sensitive are the results to the values of the elements of the matrix? The survival probabilities and fecundity rates, for example, are estimates, and will have some degree of estimation error. How much would the population dynamics, including asymptotic and transient dynamics, change if some elements of the matrix were changed "slightly"?

We will not address all these questions further here and refer the interested reader to Caswell (2001). However we will briefly discuss one type of asymptotic behavior, for both deterministic and stochastic MPMs, which is analogous to single population exponential growth models.

### 1.4.2 Limiting behavior of density independent, time invariate MPMs.

Results from linear algebra can be used to describe the asymptotic behavior of a time invariant projection matrix, if the matrix has certain mathematical properties. Three properties are (a) non-negativity (all elements are $\geq 0$ ), (b) irreducible (e.g., every age class can contribute to every other age class at some point in time), (c) primitive (there is some positive integer $k$ such that every element in the matrix raised to the power $\mathrm{k}, L^{k}$, is a positive number). If these conditions are met, in the limit the population dynamics are either exponential growth or decay, i.e., $A^{T} \mathbf{n}_{t}=\lambda \mathbf{n}_{t}$, where $\lambda$ is a scalar value that is multiplied against each component of the vector $\mathbf{n}_{t}$. Further, the relative proportions of each component of $\mathbf{n}_{t}$ will remain constant.

For example, a MPM with three age classes has the following Leslie matrix

$$
L=\left[\begin{array}{lll}
0.0 & 1.2 & 1.4 \\
0.3 & 0.0 & 0.0 \\
0.0 & 0.5 & 0.9
\end{array}\right]
$$

The initial abundance vector is $\mathbf{n}_{0}^{\prime}=(100,50,10)$. The population over 9 iterations is:

| Stage | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Young | 100 | 74 | 84 | 90 | 103 | 116 | 131 | 148 | 167 | 189 |
| Adult | 50 | 30 | 22 | 25 | 27 | 31 | 35 | 39 | 44 | 50 |
| Old | 10 | 34 | 46 | 52 | 59 | 67 | 76 | 86 | 97 | 109 |

The population growth rates, per stage, over time:

| Stage | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Young | 0.74 | 1.13 | 1.08 | 1.14 | 1.12 | 1.13 | 1.13 | 1.13 | 1.13 |
| Adult | 0.60 | 0.74 | 1.13 | 1.08 | 1.14 | 1.12 | 1.13 | 1.13 | 1.13 |
| Old | 3.40 | 1.34 | 1.14 | 1.14 | 1.13 | 1.13 | 1.13 | 1.13 | 1.13 |

Thus after six generations the annual growth rate reaches $13 \%$ and stays there. The fraction of the population in each stage class:

| Stage | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Young | 0.62 | 0.54 | 0.55 | 0.54 | 0.54 | 0.54 | 0.54 | 0.54 | 0.54 | 0.54 |
| Adult | 0.31 | 0.22 | 0.15 | 0.15 | 0.14 | 0.14 | 0.14 | 0.14 | 0.14 | 0.14 |
| Old | 0.06 | 0.25 | 0.30 | 0.31 | 0.31 | 0.31 | 0.31 | 0.31 | 0.31 | 0.31 |

And, after six generations the fractions in the Young, Adult, and Old stages remain $0.54,0.14$, and 0.31 .

The limiting population growth rate and proportions of each category can be determined analytically again using linear algebra, in particular, by carrying out an eigen analysis of $L$. For a $p$ by $p$ matrix $L$, the eigen analysis yields $p$ eigenvalues, $\lambda_{1}, \ldots, \lambda_{p}$, and $p$ corresponding right eigenvectors, $\mathbf{v}_{1}, \ldots, \mathbf{v}_{p}$. An eigenvalue and its corresponding eigenvector have the relationship, $L \mathbf{v}_{i}$ $=\lambda_{i} \mathbf{v}_{i}$. Denote the largest eigenvalue $\lambda_{1}$ and its corresponding eigenvector $\mathbf{v}_{1}$. Then $\lambda_{1}$ is equal to limiting population growth rate, in the example 1.13 (more precisely, 1.12938), and dividing each element of $\mathbf{v}_{\mathbf{1}}$ by its total yields the limiting fractions, here $(0.54,0.14,0.31)$.

### 1.4.3 Stochasticity.

There are various ways to add stochasticity to MPMs, one of which is to randomly draw elements of the matrix from probability distributions, e.g., randomly draw survival probabilities for age $a$ individuals. This has the effect of introducing environmental stochasticity. Under some conditions, in the absence of density dependence for example, the introduction of environmental, or demographic, stochasticity will not appreciably alter the asymptotic dynamics from that of a deterministic MPM, namely that the above eigen analysis results more or less hold, in the limit there is an average growth rate and stable population structure. Caswell (2001, Chap. 14) provides details of these results (with some of earliest work from Cohen, 1976; Tuljapurkar and Orzack, 1980).

Below we closely follow (Caswell, 2001) and somewhat mimic the derivation of the asymptotic distribution of the stochastic univariate model shown in eq'ns (1.16-1.20). Let $A_{1}, A_{2}, \ldots$, be a "stationary, metrically transitive stochastic process with $E\left(\ln ^{+}\|A\|_{1}\right)<\infty$, where $\ln (x+)=\max (0, \ln (x))$. Suppose that the matrices $A_{t}$ are drawn from an ergodic set of matrices," (Caswell, 2001, p. 393). The total population size at time $t$, denoted $N(t)$, is the vector norm of $\mathbf{n}_{t}\left(\sum_{i}\left|n_{t, i}\right|\right)$. Given an initial vector $\mathbf{n}_{0}$ :

$$
\begin{align*}
N(t) & =\left\|\mathbf{n}_{t}\right\|=\left\|\prod_{i=1}^{t} A_{i} \mathbf{n}_{0}\right\|  \tag{1.46}\\
& \Rightarrow \\
\frac{1}{t} \ln (N(t)) & =\frac{1}{t} \ln \left\|\prod_{i=1}^{t} A_{i} \mathbf{n}_{0}\right\| \tag{1.47}
\end{align*}
$$

Furstenberg and Kesten (1960) proved that the limit of (1.47) existed (with probability 1). That limiting value, denoted $\ln \left(\lambda_{s}\right)$, is called the stochastic growth rate.

$$
\begin{equation*}
\lim _{t \rightarrow \infty} \frac{1}{t} \ln (N(t))=\lim _{t \rightarrow \infty} \frac{1}{t} \ln \left\|\prod_{i=1}^{t} A_{i} \mathbf{n}_{0}\right\|=\ln \left(\lambda_{s}\right) \tag{1.48}
\end{equation*}
$$

With further conditions on the matrices, $A_{i}$, including nonnegativity, the asymptotic distribution of the population total is lognormal:

$$
\begin{equation*}
N(t) \sim \text { Asymptotic Lognormal }\left(\exp \left(t \ln \left(\lambda_{s}\right)\right), t \sigma^{2}\right) \tag{1.49}
\end{equation*}
$$

where $\sigma^{2}$ is some constant. Thus, similar to eq'n (1.20), the asymptotic mean of the population total is the same as for a univariate exponential population growth model, and $\lambda_{s}$ is analogous to the largest eigenvalue, $\lambda_{1}$, of a deterministic MPM.

### 1.4.4 Density dependent MPMs.

Density dependence can be introduced into MPMs by simply making some of the elements of the projection matrix density dependent. Consider the example given previously with three age classes, young, adult, and old, and suppose that the fecundity of the old group was a function of the total abundance of adult and old individuals. Then the $(1,3)$ element in the transition matrix is simply written as such a function. The linearity aspect of the MPM is subsequently altered and the analyses carried out for density independent MPMs do not directly apply, e.g., the eigen analysis is no longer directly applicable. Caswell (2001, Chap. 16) provides considerable details on the construction and analysis of density dependent MPMs.

### 1.4.5 Building block approach to matrix construction.

Buckland et al. (2007) present a building block approach to formulating the MPMs and further examples are given in Newman et al. (2014). The essential idea assumes a particular sequence to processes which operate on a vector of population abundances, such as survival, then movement, then reproduction. An example from Newman et al. (2014, eq. 2.11, p. 18)is based on two subpopulations distinguished by size class (small and large) with three processes, survival, then size class changes (only from small to large), and reproduction. The survival probabilities are size specific $\left(\phi_{L}\right.$ and $\left.\phi_{S}\right)$, the probability that a small individual becomes large is $\pi$, and only large individuals can reproduce and they do so with rate $\rho$.

$$
\begin{array}{r}
{\left[\begin{array}{c}
n_{S, t} \\
n_{L, t}
\end{array}\right]=\left[\begin{array}{ll}
1 & \rho \\
0 & 1
\end{array}\right]\left[\begin{array}{cc}
1-\pi & 0 \\
\pi & 1
\end{array}\right]\left[\begin{array}{cc}
\phi_{S} & 0 \\
0 & \phi_{L}
\end{array}\right]\left[\begin{array}{l}
n_{S, t-1} \\
n_{L, t-1}
\end{array}\right]} \\
=\left[\begin{array}{cc}
(1-\pi+\rho \pi) \phi_{S} & \rho \phi_{L} \\
\pi \phi_{L} & \phi_{L}
\end{array}\right]\left[\begin{array}{l}
n_{S, t-1} \\
n_{L, t-1}
\end{array}\right] \tag{1.50}
\end{array}
$$

The matrix in eq'n (1.50) is an example of a Lefkovitch matrix which is arguably more simply constructed by using such a building block approach than by trying to construct the final matrix in a single operation.

### 1.4.6 Estimating the elements of projection matrices.

Caswell (2001, p. 22) states that, to fill the elements of the matrix, life tables are used. Life tables contain mortality probabilities, the probability that an individual of age $a$ will die before reaching age $a+1$, and maternity functions, the expected number of offspring that an age $a$ individual will produce in the next year. How the life tables and maternity functions are constructed in the first place may be no trivial task, and is beyond the scope of this section. However, we do note that with wildlife populations, mark-recapture studies where the animals are aged at time of marking, can provide estimates of agespecific survival. For example, a population of ducks were banded, aged at time of banding, and later recovered. Mark-recapture methods are used to calculate $\hat{\phi}_{a}$, the estimated probability that an age $a$ animal lives to age $a+1$. Similarly, the average number of young produced by age $a$ females is used as an estimate of the fecundity parameter for age $a$ females, or when the total population is modeled and a $50: 50$ sex ratio is assumed, the estimated average fecundity is $0.5 \hat{f}_{a}$.

An alternative approach is to integrate inference about vital rates with inference about the entire population dynamics model where stochasticity in the population dynamics is accounted for simultaneously with sampling error in the estimation of vital rate parameters. The SSM framework provides a structure for doing this and is discussed in Section 1.7.

### 1.5 Integral Projection Models, IPMs

The partitioning of a population into discrete subpopulations, namely formulating a structured population, may be arbitrary when natural divisions are lacking. For example, suppose the population cannot be readily aged nor subdivided into distinct life stages and individual size, say weight, is the feature used to subdivide the population. The selection of size classes, small, medium, and large, will necessarily have arbitrary boundaries. The weight classes for the three size classes are $(0,10),[10,20),[20+]$, an animal weighing 9.99 pounds is labeled small and one weighing 10.0 pounds is medium. Those two individuals will be treated differently in terms of population processes, e.g., the survival probability is 0.5 for small individuals and 0.8 for medium individuals, while in reality the survival probabilities for both individuals may be much more similar. Integral Projection Models (IPMS; Easterling et al., 2000), sometimes called integrodifference equation models (see Caswell, 2001, for historical references), are a modeling approach that maintains the continuous nature of a factor that distinguishes population members, while (generally) maintaining the discrete time step characteristic of MPMs.

### 1.5.1 Kernel structure of IPMs.

The core of an IPM, which is analogous to an element in the transition matrix of an MPM, is the kernel denoted $K\left(z_{t+1}^{\prime} \mid z_{t}\right)$. The kernel can be viewed as a conditional probability density function for the "probability" that an animal of size $z$ at time $t$, denoted $z_{t}$, is size $z^{\prime}$ at time $t+1$, denoted $z_{t+1}^{\prime}$. The word probability is put in quotation marks as this is a density not a probability. More accurately $K\left(z_{t+1} \mid z_{t}\right) \Delta$ is an approximate probability for such a movement from size $z_{t}$ to a size in an interval of width $\Delta$ containing $z_{t+1}^{\prime}$, e.g., $z_{t+1}^{\prime} \pm 0.5 \Delta$. The number of individuals in a given size class at time $t+1$ is then the sum of all individuals, of any size class at time $t$, say $n\left(z_{t}^{*}\right)$ where $z_{t}^{*} \in \Omega$ and $\Omega$ is a suitably large range of sizes, that survive, growth, and/or contribute to individuals of size class $z_{t+1}$ at time $t+1$ :

$$
\begin{equation*}
n\left(z_{t+1}^{\prime}\right)=\int_{z_{t} \in \Omega} K\left(z_{t+1}^{\prime} \mid z_{t}\right) n\left(z_{t}\right) d z_{t} \tag{1.51}
\end{equation*}
$$

A simpler version of the kernel is time invariant, $F\left(z^{\prime} \mid z\right)$, where the conditional density for the contribution to size class $z^{\prime}$ at time $t+1$ from size class $z_{t}$ is the same for all times $t$.

The population growth process is the result of other processes, such as survival and reproduction. So the kernel $K$ can be decomposed into survival of the current population and reproduction entering the population. Here, however, individual size is also a factor and survival and reproduction is into a specific size class, $z_{t+1}$, thus growth from size class $z_{t}$ to $z_{t+1}$ is a third
process to account for. The resulting partitioning of the kernel:

$$
\begin{equation*}
K\left(z_{t+1}^{\prime} \mid z_{t}\right)=P\left(z_{t+1}^{\prime} \mid z_{t}\right)+F\left(z_{t+1}^{\prime} \mid z_{t}\right) \tag{1.52}
\end{equation*}
$$

where $P$ is the survival/growth kernel, the combined conditional density for surviving to time $t+1$ and changing to size class $z_{t+1}$, and $F$ is the fecundity kernel, is the conditional density for recruits at time $t+1$ of size $z_{t+1}$ (Merow et al., 2014) .

There are a wide variety of formulations for the survival/growth kernel. One formulation is to treat the two processes as independent, the result being the product of the conditional probability of surviving, $\phi\left(z_{t}\right)$, and the conditional density of moving to size class $z^{\prime}, g\left(z_{t+1}^{\prime} \mid z_{t}\right)$ :

$$
\begin{equation*}
P\left(z_{t+1}^{\prime} \mid z_{t}\right)=\phi\left(z_{t}\right) g\left(z_{t+1}^{\prime} \mid z_{t}\right) \tag{1.53}
\end{equation*}
$$

In principle, though this may be rarely appropiate, a joint density for survival and growth could be used; e.g., movement to a much larger size class could be accompanied by lowered survival probability. The survival probability could be a more complicated function of competing or sequential mortality factors; e.g., there are two mortality processes occurring in sequence, $\phi\left(z_{t}\right)=\phi_{1, z_{t}} \phi_{2, z_{t}}$.

There are also a wide variety of formulations for the fecundity kernel. For example, a female fish has a probability distribution for the number of eggs produced, $f\left(E \mid z_{t}\right)$, there is a probability that the eggs will be fertilized, $p_{E}$, a probability that the fertilized eggs will hatch, $p_{h}$, and then there is density function that hatched larvae will be a particular size, $h\left(z^{\prime}\right)$, where the last three processes are independent of $z_{t}$. Then

$$
\begin{equation*}
F\left(z_{t+1}^{\prime} \mid z_{t}\right)=f\left(E \mid z_{t}\right) p_{E} p_{h} h\left(z^{\prime}\right) \tag{1.54}
\end{equation*}
$$

Merow et al. (2014) note that a common feature of the survival/growth and fecundity kernel formulations is an individual component, e.g., $\phi\left(z_{t}\right)$ in eq'n (1.53) and $\left(f\left(E \mid z_{t}\right) p_{E} p_{h}\right)$ in eq'n (1.54), and a size redistribution component, e.g., $g\left(z_{t+1}^{\prime} \mid z_{t}\right)$ in eq'n (1.53) and $h\left(z^{\prime}\right)$ in eq'n (1.54).

### 1.5.2 Implementation of an IPM.

Equation (1.51) is analogous to the generation of a single component in the state vector of an MPM. With an MPM, the entire state vector at time $t+1$ is $\mathbf{n}_{t+1}=L \mathbf{n}_{t}$, where the $i$ th entry in $\mathbf{n}_{t+1}$ denoted $n_{i, t+1}$ is the following summation:

$$
\begin{equation*}
n_{i, t+1}=\sum_{j=1}^{p} L_{i, j} n_{j, t} \tag{1.55}
\end{equation*}
$$

where $L$ has $p$ columns. Each $L_{i, j}$ in the summation is analogous to a kernel function as it is the per individual contribution from "size" class $i$ at time $t$
to "size" class $j$ from time $t+1$. If the vector $\mathbf{n}_{t}$ is further partitioned into a relative large number of size classes, the summation operation in eq'n (1.55) approaches an integration operation.

Implementation of an IPM is in practice the reverse operation. Referring to the integral in eq'n 1.51 , the interval $\Omega$, which contains the range of size classes that can contribute to size class $z^{\prime}$, is partitioned into $m$ size classes. A finite sum approximation to integration, e.g., the midpoint rule, the trapezoid rule, or Simpson's rule, is used calculate the number of individuals in size class $z^{\prime}$. An example of the midpoint rule: suppose $\Omega$ is an interval $[L, U]$ which is partitioned into $m$ intervals of equal length $(U-L) / h$, and let $z_{i}$ be the midpoint of the $i$ th size class, also known as mesh points (Rees et al., 2014), where

$$
z_{i}=L+(i-0.5) * j, \quad i=1,2, \ldots, m
$$

The integral (1.51) can be approximated by

$$
\begin{equation*}
n\left(z_{t+1}^{\prime}\right) \approx \sum_{i=1}^{m} K\left(z^{\prime} \mid z_{i}\right) h n\left(z_{i, t}\right) \tag{1.56}
\end{equation*}
$$

### 1.5.3 Estimation of kernel components.

Assuming that relevant data on size, survival, reproduction success, etc, are available, there are many standard statistical model fitting procedures, linear regressions, nonlinear regression, generalized linear models including logistic regression, and generalized additive models, that can be utilized to construct the components of the transition kernel, $K\left(z_{t+1}^{\prime} \mid z_{t}\right)$. Likewise, many of the associated model fit diagnostic procedures could, and should be, used to assess the quality of the estimated components of the kernel (Rees et al., 2014).

Referring to the example survival/growth kernel in eq'n (1.53) and the example fecundity kernel in eq'n (1.54), a number of probability and density functions are needed to calculate the transition densities. For individual components that are probabilities, e.g., the conditional probability of survival, sample data on size conditional outcomes can be used to calculate estimates. For example, a mark-recovery study of banded ducks could provide size-specific annual survival probabilities where a smooth fitted survival function, such as a logit model, $\log (\phi /(1-\phi)) \mid z=\beta_{0}+\beta_{1} z$. Whether or not time-specific functions could be fit may depend upon the number of years of data available. Survival probabilities can be a function of size and environmental covariates, e.g., winter temperatures. For size redistribution components, e.g., the conditional density for moving from size class $z$ to $z^{*}$ in eq'n (1.53), longitudinal data is required. Size measurements made over time on multiple individuals are required.

Inference methods for IPMs are continually developing. For example,

Ghosh et al. (2012) use Bayesian hierarchical models where the size distribution is a point pattern on some interval and carry out an integrated analysis that combines the parameter estimation/model fitting stage and the projection stage.

### 1.5.4 Application, use and analysis of IPMs

Plant species were the most common organisms in early applications of IPMs, e.g., Northern Monkhood (Easterling et al., 2000), with the growth transitions between different plant sizes, e.g., stem diameter, and including of processes like flowering strategies. The scope of applications has expanded to include birds (Great tits, Childs et al., 2016), arachnids (soil mites, Brooks et al., 2015), mammals (Soay sheep), diseases (hosts and parasites, Metcalf et al., 2016).

The set of questions asked of MPMs in Section 1.4.1 can be asked of IPMs. For example the asymptotic behavior of IPMs can be examined. Is there a limiting population growth rate, a dominant eigenvalue $\lambda_{1}$ and corresponding stable "size" class distribution? Sensitivity analyses are also possible. See Ellner and Rees (2006) for examples of such analyses. Available software includes the R package IPMpack which does the following: "Construction and analysis of integral projection models and associated measures of population growth, structure, perturbations (sensitivities and elasticities), overall population dynamics, age-specific metrics, etc".

In addition to population dynamics analysis the scope of ecological inference using IPMs includes analysis of evolutionary strategies (Ellner and Rees, 2006). Brooks et al. (2015) separated out the effects of individual body size on developmental rates from the effects of environmental conditions on reproductive rates. Metcalf et al. (2016) examined the feedback between host and parasite in an epidemiological analysis.

### 1.6 Individual Based Models, IBMs

Individual based models (IBMs; DeAngelis and Grimm, 2013), sometimes referred to as agent-based models (but see Roughgarden, 2012, for a definition of agents narrower than individuals), in ecology are computer simulation procedures that, in their most complete form, track the entire life history of multiple individuals simultaneously. Variables that are tracked include emergence into the population (date of birth, germination, hatch date), size at birth, sex, size over time, maturation process including time of sexually mature and reproduction, spatial location and movement over time, senescence, and death. One central feature of many ecological IBMs is the modeling of interactions of individuals with each other, including individuals of the same species, e.g.,
reflecting competition for resources and density dependence, or individuals of different species, e.g., reflecting predator-prey dynamics or more broadly ecological community interactions. The other central feature is the simulation of interactions of individuals with their abiotic environment, e.g., air temperature and precipitation, and their biotic environment excluding like individuals, e.g., vegetative browse and zooplankton.

IBMs are by design forward simulation based approaches to modeling demographics and, as such, the opportunity to insert complexity into dynamic processes is relatively unlimited, constrained in principle by the available computer storage and processing speed. Population level properties can be examined at arbitrary points in time in the simulation process by aggregating the states of individuals in arbitrary ways. For example, the simulation may start with a vector of 1000 individuals where each individual has an associated vector of initial conditions such as age, weight, sex, spatial location, maturity, and whatever attributes are used to distinguish individuals, i.e., attributes that could serve as potential subpopulation categories in an arbitrarily structured MPM. Processes like survival, growth, movement, reproduction are then applied to each individual and at an arbitrary time point $t_{1}$, numbers of individuals in a set of spatial regions further distinguished by sex and age class, say, are tallied to yield abundances of multiple subpopulations. Continuing the simulation to a later time point, $t_{2}\left(>t_{1}\right)$, the aggregation is done again. Repeating the simulation and aggregation $K$ times yields a multivariate time series of subpopulation abundances, $\mathbf{n}_{\mathbf{t}_{1}}, \mathbf{n}_{\mathbf{t}_{2}}, \ldots, \mathbf{n}_{t_{K}}$. Analysis of population level dynamics can then be conducted, studying such things as the effects of region-specific harvest regulations on a population of deer, for example. If the effects of changes to or states of environmental and anthropogenic factors on the population dynamics cannot be readily solved in any analytic sense, IBM output can provide some experiential, albeit simulated, insight.

### 1.6.1 Statistical designs for and analysis of IBMs

The simulation nature of IBMs with multiple attributes and multiple levels to attributes lends itself to using methods from the statistical design of experiments, such as factorial experiments, to construct a set of forward simulations with an arbitrary number of replications, and an arbitrary longitudinal time series length. For example, if three attributes are of particular interest, say sex, spatial location, and age class with corresponding levels of (female, male), (I, II, III, IV) regions, and ages ( $0,1,2,3+$ ), then a factorial design with $2 \times 4 \times 4=32$ "treatment" combinations can be conducted with $r$ replications of each combination. Statistical methods for analyzing data from factorial experiments such as analysis of variance or response surface modeling can then be applied. Aggregated data could also used to construct simple MPMs, like year-specific Leslie matrices, and methods for assessing MPMs, such as calculating annual finite population growth rates for multiple years could be employed (for such an example, see Rose et al., 2013).

The computational burden of IBMs can grow in a number of ways. First, as might have been apparent from the previous $2 \times 4 \times 4$ factorial above, as the number of attributes of interest and the levels of each attribute increases, the number of treatment combinations can grow rapidly. Second, as the level of environmental stochasticity (or demographic) increases, the number of replicates required to provide a desired level of precision for estimates of average population level responses along with the associated stochastic variation increases as well. Third, questions about the effects of the distribution of initial attribute values at time $t_{0}$ as well as questions about the nature of the processes, e.g., density dependent or density independent recruitment success or the chosen value, or distribution of values, for juvenile survival probabilities, can lead to a considerable number of sensitivity analyses.

### 1.6.2 Comparison with population models

In contrast to population models, such as the Ricker model (eq'n 1.35), for which long term population-level behavior such as exponential growth, an asymptotic upper bound, or periodicity can sometimes be determined analytically or by elementary computer simulation, population-level behavior for IBMs is an emergent property that might only be apparent from simulation. The resulting behavior is the result of potentially complex interactions of individuals with each other and with their abiotic environment (DeAngelis and Grimm, 2013), and can demonstrate "the importance of local interactions between individuals in ecological systems" (Judson, 1994).

A succinct way to contrast population-level models and IBMs is top-down versus bottom-up. Population-level models are top-down in that they predict what happens to individuals as function of population level characteristics, e.g., fecundity of the individual decreases as the total population abundance increases (density dependence exists). Conversely, IBMS are are bottom-up in that modeling begins with the characteristics of multiple individuals and manifests characteristics of the population as a whole. An interesting example of the latter is with Anolis lizards in the Caribbean (discussed in Roughgarden, 2012) where an IBM simulated energy gained per unit time after a lizard consumed a prey item as a function of distance from the prey and the optimal foraging distance could then be determined. From that model for the "energy capture" the daily growth rate of the lizard was predicted, with distinction made between growth prior to reproductive stage and during the reproductive stage. Using these results an optimal growth rate, as a function of age, was calculated, which was then used with information on survival probabilities and maternity rates to determine that optimal female body size was 45 mm . As Roughgarden (2012) said " $[t]$ his example illustrates a complete and successful modeling protocol that begins with the properties of an individual and culminates in the an evolutionary prediction of the adult body size for lizards on an island in the absence of congeneric competitors".

### 1.6.3 Applications of IBMs

Some of the earliest applications of IBMs in ecology were in forestry, and such applications remain common. In the IBM JABOWA (Botkin et al., 1972), individual trees were the fundamental entities and the central measure on each tree was its stem diameter (at some height on the tree). Other tree measures such as volume and crown biomass are sometimes functions of diameter. The emergence, growth, and eventual death of a tree is a function of interactions with neighboring trees, their size and proximity and the degree to which their presence led to competition for resources for light and water, for example, and a function of interactions with the abiotic environment, e.g., soil type and chemistry, precipitation, temperature, and light. IBMs have been used for both management purposes, to predict growth and yield of commercially harvested species, and to "explore ecological mechanisms and patterns of structure and functional dynamics in natural forest ecosystem" (Liu and Ashton, 1995).

Applications of IBMs to fish populations have been common as well, where IBMs "track the attributes of individual fish through time and aggregate them to generate insights into population function" (Van Winkle et al., 1993). IBMs simulate how fish of different phenotypes interact with their biotic and abiotic environment. Differences in phenotype can refer to differences in length, weight, sex, and age, the biotic environment can include prey items, such as zooplankton or vegetation, and the abiotic environment can include water temperature, salinity, water clarity. An IBM application to Delta Smelt (Hypomesus transpacificus, Rose et al., 2013) also included bioenergetics considerations, namely the transformation of consumed prey into fish growth.

More generally, the use of IBMs in ecology can be broadly divided into applications for (individual) populations, communities and ecosystems. Single population-level IBMs have been mentioned above, e.g., Anolis lizards and Delta Smelt, but IBMs have used to model predator-prey dynamics (Cuddington and Yodzis, 2002). A community-level application by Weiss et al. (2014) used an IBM to simulate how the dynamics of a community assembly of 90 hypothetical plant types were affected by soil attributes and grazing intensities. The results were then compared to field-based observations of species richness and diversity. Least common are ecosystems level applications; a hypothetical food web system used an IBM to model interactions between three trophic levels, plant, herbivore and carnivore (Schmitz and Booth, 1997).

### 1.6.4 Data needs and structure

There are at least three levels of data or information needs for IBMs. One is the information on the initial attributes to be assigned to the individuals, an attribute vector (Van Winkle et al., 1993), and the biotic and abiotic environment. Due to general case where proximity to other individuals is a factor in the dynamics, information is needed at least about spatial location, thus an initial spatial distribution is needed: should locations be randomly placed as in
a Poisson process, systematically placed, clustered, placed with probabilities proportional to particular habitat conditions? In addition to spatial location other individual attributes, e.g., size, sex, age, need to be assigned. To achieve greater realism, the actual multivariate distribution of such attributes should be mimicked. Attributes of the environment that the individuals populate are also needed, included spatial location. For example, if the population of interest is an herbivore, then the types of plants, abundance and spatial distribution must be specified. Similarly, abiotic features, e.g., soil types, water sources, etc, need to be determined.

The other need is for information about how the individuals interact with each other and with their environment in terms of processes of relevance to the model purpose. For example, how is the probability of survival affected by the availability and proximity to food items? How is movement affected by population density, biotic and abiotic features?

To verify that IBM output, and apparent emergent population level properties, e.g., collective survival, reproduction, and movement rates, are reasonable, field-based observations are needed.

Given these data needs, "IBMs have therefore been criticized as being too 'data hungry'-especially IBMs designed for specific, applied problems (Grimm and Railsback, 2013). The available data may thus constrain and guide the formulation on an IBM, affecting things like the time step resolution, spatial scope, number of attributes followed, and number of interactive processes simulated.

### 1.6.5 Relationship with IPMs

IBMs overlap somewhat with IPMs in that measurements on individuals made over time, i.e., longitudinal data, are central. IPMs use such data to model the relationship between the value of an attribute measured at time $t$ and its value at time $t+1$, say, the probability of transitions from one value to another are then calculated based on the collective pairings of values from multiple individuals. IBMs at times start a sample of $n_{0}$ individuals with randomly chosen initial attribute values, say $x_{0,1}, x_{0,2}, \ldots$, and $x_{0, n}$, and with the relationship between $x_{t, i}$ and $x_{t+1, i}$ for arbitrary individual $i$, typically a stochastic relationship, and then projects the longitudinal trajectory of each of the $n$ individuals forward in time via stochastic simulation or deterministic projection. IBMs can be used to generate data that are then used to evaluate fitting procedures for IPMs and the subsequent performance of IPMs can be evaluated by comparing IPM predictions to the "true" values generated by simulated IBM output (Rees et al., 2014).

### 1.7 State-Space Models, SSMs

State-space models (SSMs) are models for two time series running in parallel. One time series, here referred to as the state process time series, describes the evolution of the true, but generally unknown, state of nature over time. The other time series, here called the observation time series, is a sequence of imperfect or inexact measurements of the state process time series. The state process time series will be denoted $\mathbf{n}_{t}, t=0,1,2, \ldots, T$, and the observation time series is $\mathbf{y}_{t}$ with $t=1,2, \ldots, T$. The state $\mathbf{n}_{0}$ is referred to as the initial state. The discrete integer valued subscripting of the two time series is used here primarily to reduce notation as somewhat arbitrary time points could be used, $t_{1}, t_{2}, \ldots, t_{T}$. The time series indexing for both time series do not necessarily need to coincide, e.g., there could be half the observations if the state is only observed every other time point, although statistical estimation limitations might occur. Also, the dimensions of $\mathbf{n}_{t}$ and $\mathbf{y}_{t}$ need not be the same, although situations where the dimensions differ, in particular the dimension of $\mathbf{n}_{t}$ is greater than the dimension of $\mathbf{y}_{t}$ statistical inference limitations or problems may result. For an ecological application, $\mathbf{n}_{t}$ could be viewed as a vector of true abundances of subpopulations at time $t$ and $\mathbf{y}_{t}$ as estimates of individual components or combinations of components of $\mathbf{n}_{t}$.

Given these two time series, the structure of a SSM is a paired sequence of probability distributions (probability mass functions for integer valued components or probability density functions for continuous valued components) that characterize the evolution of the state process and the relationship between the observation vector and the state vector. The the probability distribution for the state process is typically first order Markov, i.e., the state at time $t$ given the state at time $t-1$ is conditionally independent of all other states. This conditional distribution is sometimes called the state transition "equation". The observation vector at time $t$, given the state vector at time $t$, is conditionally independent all other state vectors and all other observation vectors.

### 1.7.1 Normal dynamic linear models

A classic SSM, originating from Kalman (1960), is the normal dynamic linear model (NDLM). A specific example of an NDLM is the following. form.

$$
\begin{aligned}
\mathbf{n}_{0} & \sim D(\theta) \\
\mathbf{n}_{t} \mid \mathbf{n}_{t-1} & \sim \operatorname{MVN}\left(L \mathbf{n}_{t-1}, \Sigma\right), \quad t=1,2, \ldots, T \\
\mathbf{y}_{t} \mid \mathbf{n}_{t} & \sim \operatorname{MVN}\left(B \mathbf{n}_{t}, \Omega\right), \quad t=1,2, \ldots, T
\end{aligned}
$$

where $D(\theta)$ denotes an arbitrary probability distribution with parameter $\theta$ which may be degenerate, i.e., $\mathbf{n}_{0}$ is a fixed value, $M V N$ is multivariate normal, $L$ and $B$ are matrices (most simply square matrices), and $\Sigma$ and $\Omega$ are
variance-covariance matrices. As denoted here all the matrices are time invariant, but that does not need to be the case. Given the observation time series, $\mathbf{y}_{t}, t=1,2, \ldots, T$, and the values of $\mathbf{n}_{0}, L_{t}, B_{t}, \Sigma_{t}$, and $\Omega_{t}$, an algorithm known as the Kalman filter can be used to calculate the conditional distribution of $\mathbf{n}_{t}$, which is multivariate normal, given $\mathbf{y}_{1}, \ldots, \mathbf{y}_{t}$. The Kalman filter also yields the calculated value of the likelihood (the joint marginal distribution of $\mathbf{y}_{t}$, $t=1,2, \ldots, T)$, which can in principle then be used to estimate unknown parameters of the transition and variance-covariance matrices. In practice there are considerable restrictions on the estimability of the parameters, and correlations between estimates of the parameters of the state process covariance matrix and parameters of the observation covariance matrix (Dennis et al., 2006).

The notation $L$ for the state transition matrix was selected to suggest the notion of a Leslie matrix as for an age-structured MPM, a SSM extensions of MPMs are not uncommon (Sullivan (1992); Newman (1998), and see the gray whale example in section 6.4.2.2 of Newman et al. (2014)). Thus the components of $L$ can include fecundity and survival, for example, or can be considerably more complex, as suggested in the building block approach to MPM formulation discussed previously.

### 1.7.2 Non-normal, nonlinear SSMs

The NDLM structure, while it has the advantage of the Kalman algorithm machinery, may often be considered too constricting and unrealistic for population dynamics modeling. More realistic state-space models can on occasion be "shoe-horned" into the NDLM framework. Log transformations to linearize multiplicative relationships can sometimes work depending upon the formulation of the state (and observation) models. For example, Dennis et al. (2006) used a stochastic Gompertz model for the state process distribution.

$$
n_{t} \mid n_{t-1}=\lambda n_{t-1}^{1+\alpha} \exp \epsilon_{t}
$$

where $\alpha \leq 0$ and $\epsilon_{t} \sim \operatorname{Normal}\left(0, \sigma_{\epsilon}^{2}\right)$. A natural $\log$ transform yields a linear normal state distribution.

$$
\ln \left(n_{t}\right) \mid n_{t-1} \quad \sim \quad \operatorname{Normal}\left(\ln (\lambda)+(1+\alpha) \ln \left(n_{t-1}\right), \sigma_{\epsilon}^{2}\right)
$$

Another way to modify an otherwise non-normal, and perhaps nonlinear SSM, into a NDLM approximation is to work with just the first two moments of the state process distribution and then use the mean and covariance structure as the normal mean vector and covariance matrix. Newman (1998) and Newman et al. (2014) give examples of such substitutions. A simplistic univariate example is to suppose that a scalar valued state $n_{t}$ is $\operatorname{Binomial}\left(n_{t-1}, \phi_{t}\right)$, where $\phi_{t}$ is the survival probability, perhaps a function of covariates. The conditional expected value of $n_{t}$ is of course $\phi_{t} n_{t-1} \equiv L_{t} n_{t-1}$, and the conditional variance is $n_{t-1} \phi_{t}\left(1-\phi_{t}\right) \equiv Q_{t}$. Other, perhaps somewhat slight, departures from the

NDLM formulation can be accommodated by Taylor series transformations of the process, using the Extended Kalman Filter (EKF; Einicke and White, 1999). A more recent alternative to the EKF, which has been shown to have at least equal and often far superior performance (Durbin and Koopman, 2012, p. 236) is the Unscented Kalman Filter (Julier and Uhlmann, 2004).

Computer intensive Monte Carlo methods such as Markov chain Monte Carlo (MCMC, Gilks et al., 1996) and Sequential Monte Carlo (SMC Doucet and Gordon, 2001) offer the ultimate flexibility for fitting nonlinear, nonnormal SSMs. With the MC procedures applied to such SSMs, Bayesian inference has been the dominant approach, but not always (see De Valpine, 2003; Ionides et al., 2006, for exceptions). One of the first ecological applications using MC methods was by Meyer and Millar (1999), who used the program BUGS (Bayesian inference Using Gibbs Sampling) to fit an SSM with scalar states and observations. The state was scaled biomass $\left(p_{t}=B_{t} / K\right)$, rather than abundance, where biomass $\left(B_{t}\right)$ was divided by carrying capacity, $K$, thus $0<p_{t} \leq 1$ ), and the observation was a biased measure of scaled biomass, an index $\left(y_{t}\right)$ :

$$
\begin{aligned}
p_{t} \mid p_{t-1} & \sim \operatorname{Lognormal}\left(\ln \left(p_{t-1}+r p_{t-1}\left(1-p_{t}\right)-\frac{c_{t-1}}{K}\right), \sigma_{p}^{2}\right) \\
y_{t} \mid p_{t} & \sim \operatorname{Lognormal}\left(\ln \left(q K p_{t}\right), \sigma_{o}^{2}\right)
\end{aligned}
$$

Thus the SSM was intrinsically nonlinear (no transformation of the state would linearize the mean structure) and non-normal. Environmental stochasticity was implicit to the Millar and Meyer model in the lognormal variation around the median response.

### 1.7.3 Hierarchical and continuous time SSMs

An extension of SSMs is a hierarchical state-space model (HSSM). A general formulation for an HSSM in a Bayesian framework is the following

$$
\begin{array}{rll}
\text { Prior distribution } & : \pi(\eta, \omega) \\
\text { Stochastic variation in parameter } & : & h\left(\Theta_{t}, \eta\right) \\
\text { State process model } & : g_{t}\left(\mathbf{n}_{t} \mid \mathbf{n}_{t-1}, \Theta_{t}\right) \\
\text { Observation model } & : & f_{t}\left(\mathbf{y}_{t} \mid \mathbf{n}_{t}, \Omega\right) \tag{1.60}
\end{array}
$$

Newman and Lindley (2006) used Sequential Monte Carlo to fit a Bayesian HSSM to salmon data which included both environmental and demographic stochasticity. The environmental stochasticity was modeled as above with separate distributions for year-specific survival and maturation probabilities. Demographic stochasticity was incorporated in the state process equations using multinomial distributions to reflect between individual variation in survival and maturation (although given the population size, the influence of demographic stochasticity on the results was likely minimal).

Durbin and Koopman (2012) discuss continuous time SSMs for a couple cases including what is called a continuous time local level SSM. Here $n(t)$ $=n(0)+\sigma_{\epsilon} \omega(t)$, where $\omega_{t}$ arises from a Brownian motion process, which means $\omega(0)=0, \omega(t) \sim \operatorname{Normal}(0, t)$ for $0<t<\infty$, and "jumps" or increments without common endpoints are independent, e.g., $\omega(2)-\omega(1)$ is independent of $\omega(4)-\omega(3)$. For an ecological application of continuous time SSMs see Johnson et al. (2008) who model the location of marine mammals using telemetry data.

### 1.8 Concluding Remarks

### 1.8.1 Omissions and sparse coverage

Demography is vast topic with considerable depth and breadth, as book length treatments of MPMs and IBMs alone indicate. Continuous time demographic models, including models based on deterministic or stochastic differential equations, have been largely ignored here (the Lotka-Volterra predator-prey model being one exception). Williams et al. (2002) provides an introduction to continuous Markov processes, including birth and death processes, and Brownian motion in the context of models for animal populations. Gurney and Nisbet (1998) present several topics in modeling ecological dynamics where discrete time and continuous time models are compared and contrasted. For continuous time SSMs, Durbin and Koopman (2012) is a reference.

Key principles or aspects of ecological theory which have demographic implications that were not mentioned include fitness (of which there are multiple interpretations, with reproductive success an approximate measure), adaptation, mutation. Effective population size, $N_{e}$, of an existing population, here defined as the minimum number of individuals necessary in a hypothetical population that would represent existing populations ability to retain the genetic diversity present, is an important concept for endangered species, and methods for calculating $N_{e}$ were not addressed. Coverage of the demography of multiple populations, communities, and ecosystems was scanty, with some mention made using IBMs, but measures of community structure such as species richness and models for changes in such measures were not mentioned. Demographic modeling of ecosystems has been particularly popular in fisheries (Christensen and Pauly, 1992; Walters et al., 1997, 1999) with Ecopath with Ecosim and Ecospace being the best known software.

### 1.8.2 Recommended literature

For MPMs, the Caswell (2001) book remains an outstanding reference with near encyclopedic coverage of material to 2001. For stochastic MPMs, the Tuljapurkar (1990) book is a classic.

For IPMs, there are two "How To" papers, Rees et al. (2014) and Merow et al. (2014) which provide the basic components of IBMs, ways of estimating the kernel components, and ways of making the projections (using numerical integration methods). The original paper (Easterling et al., 2000) includes detailed discussion of the advantages of IPMs over MPMs, while Ellner and Rees (2006) include detailed examples of stable population analyses often done with MPMs. More sophisticated and integrated IPM fitting and projection approaches are described by Ghosh et al. (2012).

For IBMS, Grimm and Railsback (2013) provide a book length treatment, while DeAngelis and Grimm (2013) is a more recent and considerably briefer overview paper. (Roughgarden, 2012) gives an alternative perspective on the definition of and uses of IBMs, viewing agent-based models as a special case, for example.

For SSMs, Durbin and Koopman (2012) is a book length and extremely thorough treatment of SSMs about two thirds of the book covering linear SSMs, including classical treatment with the Kalman algorithms and extensions, and the remainder of the book on nonlinear, non-normal SSMs including coverage of special cases as well as quite general formulations that are typically fit by Monte Carlo procedures. Specific focus on the use of SSMs for population dynamics modeling is given by Newman et al. (2014).

### 1.8.3 Speculations on future developments

Future developments in biological demography can be placed in three categories: data, model formulations, and model fitting.

## Data.

The volume and complexity of data on individual organisms will continue to grow as the life spans of biological monitoring programs extend, as new monitoring programs are established, and as technology for collecting data advances. Electronic monitoring devices, e.g., radio tag collars, acoustic tags, tags that record the diving depths of marine animals, are providing increasingly fine temporal and spatial resolution information on individual animal movement. Chemical analyses of organisms are providing more information about some aspects of individual life histories, e.g., chemical analyses of bony structures in fish, such as otoliths, can provide information about where the fish were born and migration paths (Secor et al., 1995). Environmental DNA (eDNA) is an emerging tool for indirectly detecting the presence of various species (Thomsen and Willerslev, 2015). Remote sensing data provides increasingly abundant and detailed information about the abiotic environment inhabited by organisms. In short, the "attribute vectors" (Van Winkle et al., 1993) for individuals, for populations, and abiotic and biotic environments will grow in length.

## Model formulations.

Of the four general formulations for demographic modeling, SSMs are unique in making explicit the distinction between process variation (environmental and demographic stochasticity), and observation noise (e.g., sampling errors). Inclusion of both sources of uncertainty in demographic models will be more common. Bolker (2008) has a nice discussion of differences in methods for fitting models with process variation alone, observation noise alone, and both process and observation variation, and argues for the latter. In brief, hierarchical modeling (Kery and Royle, 2016; King et al., 2009), whether in a Bayesian framework (eq'ns 1.57-1.60), or non-Bayesian framework (eq'ns 1.58-1.60) will become more commonplace.

Hierarchical extensions of MPMs which separate process and observation variation within the normal dynamic linear model framework of SSMs date back to the 1990s, e.g., Sullivan (1992) and Newman (1998). Separate accounting of process and observation variation in MPMs (as well as other frameworks) can lead to substantive differences in inferences; An example given by Newman et al. (2014) for modeling the population dynamics of Eastern North Pacific gray whales (Eschrichtius robustus contrasted an MPM with observation error only with a SSM extension. Differences in some of the parameter estimates were considerable, e.g., juvenile survival probability was estimated to be 0.9999 (upper bound) for the observation error only model and 0.8281 for the SSM. NDLM formulations have been used primarily for computational convenience as the Kalman filter provides an efficient means of calculating the likelihood. Given advances in model fitting procedures there is less need to restrict process models to linear formulations, implicit to MPMs, with additive normal (or multiplicative lognormal) distributions. More biologically realistic nonlinear, and non-Gaussian formulations may make applications in the MPM framework less common. However, the MPM structure will remain valuable for formulating approximate deterministic skeletons underlying more realistic models (Buckland et al., 2007).

For IPMs and IBMs, computer simulations can explicitly partition and account for process and observation uncertainty. With IPMs, bootstrapping the process model fitting procedure, namely the estimation of the kernel density components, would provide measures of parameter estimate uncertainty as well as between animal variation. For example, when estimating the growth density, $g\left(z_{t+1}^{\prime} \mid z_{t}\right)$, in the survival/growth kernel (eq'n 1.53), the longitudinal data on sizes would be resampled and a bootstrapped distribution of growth densities would provide a measure of parameter uncertainty. For a given fitted growth density model, the simulated variation of individual sizes around the expected size at time $t+1$ would provide a measure of demographic variation.

For IBMs, computer simulation of between individual variation and parameter uncertainty can be carried out within a designed experiment structure, e.g., factorial designs, to (a) determine the relative import of specific factors
on the model predictions and (b) quantify the degree of uncertainty in model predictions.

## Model fitting.

Extended attribute vectors for increasingly large numbers of individuals, along with increasingly complex demographic model formulations, necessitate increasingly complex model fitting procedures. The pace of development for fitting such models is rapid and the variety of model fitting options available is increasing. Here we focus on some of these options for making inferences for dynamic hierarchical demographic models, like SSMs, with both a state process model and an observation model.

At the heart of the fitting procedures for hierarchical models is mathematical integration and numerical optimization. For some of these models, exact, closed form solutions to the integration and optimization problems do exist. The NDLM is a notable case where the Kalman filter in effect does the integration to yield the conditional distribution for states and the likelihood, which then facilitates maximum likelihood estimation. Numerical approximations to nonlinear, but Gaussian, population dynamics models yield models amenable to such analytic solutions, e.g.,the extended Kalman filter with its first order Taylor series approximation to the state process and observation models.

Inference for more general hierarchical dynamic models requires approximate techniques for integrating over the unobserved state process to yield the likelihood, or the conditional distribution for the states given the observations. Approximate analytic solutions to the integration problems include the Laplace approximation (Tierney and Kadane, 1986) and the Integrated Nested Laplace Approximation (INLA, Rue et al., 2009). The software package, AD Model Builder (ADMB, Fournier et al., 2012), uses Laplace approximations to integrate over the state process distributions to yield the likelihood and then automatic differentiation for calculating maximum likelihood estimates of the parameters.

Computer intensive Monte Carlo (MC) procedures, e.g., MCMC and sequential Monte Carlo, carry out the integration by simulation. In the Bayesian framework, "optimization" per se is not done, as the entire posterior distribution for states and parameters is generated. In the likelihood framework, MC methods can produce estimates of the likelihood function and optimization is then done with that estimate (De Valpine, 2003). Commonly used MCMC software for fitting Bayesian SSMs includes WinBUGS (Lunn et al., 2000) and JAGS (Plummer et al., 2003).

Two recent software programs, NIMBLE and the R package pomp, allow users to choose from a variety of computer intensive model fitting procedures. NIMBLE (de Valpine et al., 2015) extends the BUGS software and allows estimation within Bayesian or likelihood frameworks. The R package pomp, for "partially observed Markov processes", contains a variety of procedures for
fitting state-space models, with including "sequential Monte Carlo, iterated 1454 filtering, particle Markov chain Monte Carlo, approximate Bayesian computation, maximum synthetic likelihood estimation, nonlinear forecasting, and trajectory matching" (King et al., 2016).


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