

Research Article

A Note on the Rate of Strong Convergence for Weighted Sums of Arrays of Rowwise Negatively Orthant Dependent Random Variables

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Let $\{X_{ni}; i \geq 1, n \geq 1\}$ be an array of rowwise negatively orthant dependent (NOD) random variables. The authors discuss the rate of strong convergence for weighted sums of arrays of rowwise NOD random variables and solve an open problem posed by Huang and Wang (2012).

1. Introduction

Firstly, let us recall the definitions of negatively associated (NA) random variables and NOD random variables as follows.

Definition 1. A finite collection of random variables $\{X_i; 1 \leq i \leq n\}$ is said to be NA if for every pair of disjoint subsets A_1 and A_2 of $\{1, 2, \dots, n\}$,

$$\text{Cov}(f_1(X_i, i \in A_1), f_2(X_j, j \in A_2)) \leq 0, \quad (1)$$

whenever f_1 and f_2 are nondecreasing functions such that the covariance exists. An infinite collection of random variables $\{X_i; i \geq 1\}$ is NA if every finite subcollection is NA.

An array of random variables $\{X_{ni}; i \geq 1, n \geq 1\}$ is called rowwise NA random variables if for every $n \geq 1$, $\{X_{ni}; i \geq 1\}$ is a sequence of NA random variables.

Definition 2. A finite collection of random variables $\{X_i; 1 \leq i \leq n\}$ is said to be NOD if

$$P(X_1 \leq x_1, X_2 \leq x_2, \dots, X_n \leq x_n) \leq \prod_{j=1}^n P(X_j \leq x_j), \quad (2)$$

$$P(X_1 > x_1, X_2 > x_2, \dots, X_n > x_n) \leq \prod_{j=1}^n P(X_j > x_j),$$

for all $x_1, x_2, \dots, x_n \in \mathbb{R}$. An infinite collection of random variables $\{X_i; i \geq 1\}$ is said to be NOD if every finite subcollection is NOD.

An array of random variables $\{X_{ni}; i \geq 1, n \geq 1\}$ is called rowwise NOD random variables if for every $n \geq 1$, $\{X_{ni}; i \geq 1\}$ is a sequence of NOD random variables.

The concepts of NA and NOD random variables were introduced by Joag-Dev and Proschan [1]. Obviously, independent random variables are NOD, and NA implies NOD from the definition of NA and NOD, but NOD does not imply NA. So, NOD is much weaker than NA. Because of the wide applications of NOD random variables, the notion of NOD random variables has been received more and more

attention recently. Many applications have been found. We can refer to Volodin [2], Asadian et al. [3], Amini et al. [4, 5], Kuczmaszewska [6], Zarei and Jabbari [7], Wu and Zhu [8], Wu [9], Sung [10], Wang et al. [11], Huang and Wang [12], and so forth. Hence, it is very significant to study limit properties of this wider NOD random variables in probability theory and practical applications.

Let $\{X_n; n \geq 1\}$ be a sequence of independent and identically distributed (i.i.d.) random variables and let $\{a_{ni}; i \geq 1, n \geq 1\}$ be an array of real constants. As Bai and Cheng [13] remarked, many useful linear statistics, for example, least-squares estimators, nonparametric regression function estimators, and jackknife estimates, are based on weighted sums of i.i.d. random variables. In this respect, the strong convergence for weighted sums $\sum_{i=1}^n a_{ni}X_i$ has been studied by many authors (see, e.g., Bai and Cheng [13]; Cuzick [14]; Sung [15]; Tang [16]; etc.).

Cai [17] proved the following complete convergence result for weighted sums of NA random variables.

Theorem A. *Let $\{X, X_n; n \geq 1\}$ be a sequence of identically distributed NA random variables, and let $\{a_{ni}; 1 \leq i \leq n, n \geq 1\}$ be an array of real constants satisfying*

$$A_\alpha = \limsup_{n \rightarrow \infty} A_{\alpha,n} < \infty, \quad A_{\alpha,n} = \frac{1}{n} \sum_{i=1}^n |a_{ni}|^\alpha, \quad (3)$$

for some $0 < \alpha \leq 2$. Suppose that $EX = 0$ when $1 < \alpha \leq 2$. If

$$E \{ \exp(h|X|^\gamma) \} < \infty \quad \text{for some } h > 0, \gamma > 0, \quad (4)$$

then, for $b_n = n^{1/\alpha}(\log n)^{1/\gamma}$,

$$\sum_{n=1}^{\infty} \frac{1}{n} P \left(\max_{1 \leq j \leq n} \left| \sum_{i=1}^j a_{ni} X_i \right| > \varepsilon b_n \right) < \infty \quad \forall \varepsilon > 0. \quad (5)$$

Wang et al. [11] extended the above result of Cai [17] to arrays of rowwise NOD random variables as follows.

Theorem B. *Let $\{X_{ni}; i \geq 1, n \geq 1\}$ be an array of rowwise NOD random variables which is stochastically dominated by a random variable X and let $\{a_{ni}; 1 \leq i \leq n, n \geq 1\}$ be an array of real constants. Assume that there exist some δ with $0 < \delta < 1$ and some α with $0 < \alpha < 2$ such that $\sum_{i=1}^n |a_{ni}|^\alpha = O(n^\delta)$ and assume further that $EX_{ni} = 0$ if $1 < \alpha < 2$. If for some $h > 0$ and $\gamma > 0$ such that (4), then*

$$\sum_{n=1}^{\infty} n^{\alpha p - 2} P \left(\max_{1 \leq j \leq n} \left| \sum_{i=1}^j a_{ni} X_{ni} \right| > \varepsilon b_n \right) < \infty \quad \forall \varepsilon > 0, \quad (6)$$

where $p \geq 1/\alpha$ and $b_n = n^{1/\alpha}(\log n)^{1/\gamma}$.

Recently, Huang and Wang [12] partially extended the corresponding theorems of Cai [17] and Wang et al. [11] to NOD random variables under a mild moment condition.

Theorem C. *Let $\{X_n; n \geq 1\}$ be a sequence of NOD random variables which is stochastically dominated by a random*

variable X and let $\{a_{ni}; i \geq 1, n \geq 1\}$ be a triangular array of real constants such that $a_{ni} = 0$ for $i > n$. Let

$$A_\beta = \limsup_{n \rightarrow \infty} A_{\beta,n} < \infty; \quad A_{\beta,n} = n^{-1} \sum_{i=1}^n |a_{ni}|^\beta, \quad (7)$$

where $\beta = \max(\alpha, \gamma)$ for some $0 < \alpha \leq 2, \gamma > 0$, and $\alpha \neq \gamma$. Assume that $EX_n = 0$ for $1 < \alpha \leq 2$ and $E|X|^\beta < \infty$. Then,

$$\sum_{n=1}^{\infty} \frac{1}{n} P \left(\left| \sum_{i=1}^n a_{ni} X_i \right| > \varepsilon b_n \right) < \infty, \quad (8)$$

where $b_n = n^{1/\alpha}(\log n)^{1/\gamma}$.

As Huang and Wang [12] pointed out, Theorem C partially extends only the case of $\alpha > \gamma$ of Theorems A and B. They left an open problem whether the case of $\alpha = \gamma$ of Theorem C holds for NOD random variables.

The main purpose of this paper is to further study strong convergence for weighted sums of NOD random variables and to obtain the rate of strong convergence for weighted sums of arrays of rowwise NOD random variables under a suitable moment condition. We solve the above problem posed by Huang and Wang [12].

We will use the following concept in this paper.

Definition 3. An array of random variables $\{X_{ni}; i \geq 1, n \geq 1\}$ is said to be stochastically dominated by a random variable X if there exists a positive constant C such that

$$P(|X_{ni}| > t) \leq CP(|X| > t), \quad (9)$$

for all $t \geq 0, i \geq 1$, and $n \geq 1$.

2. Main Results

Now, we will present the main results of this paper; the detailed proofs will be given in the next section.

Theorem 4. *Let $\{X_{ni}; i \geq 1, n \geq 1\}$ be an array of rowwise NOD random variables which is stochastically dominated by a random variable X and let $\{a_{ni}; 1 \leq i \leq n, n \geq 1\}$ be an array of real constants satisfying $\sum_{i=1}^n |a_{ni}|^\alpha = O(n)$ for some $0 < \alpha \leq 2$. Assume further that $EX_{ni} = 0$ for $1 < \alpha \leq 2$ and $E|X|^\alpha \log(1 + |X|) < \infty$. Then,*

$$\sum_{n=1}^{\infty} n^{-1} P \left(\left| \sum_{i=1}^n a_{ni} X_{ni} \right| > \varepsilon b_n \right) < \infty \quad \forall \varepsilon > 0, \quad (10)$$

where $b_n = n^{1/\alpha}(\log n)^{1/\alpha}$.

Similar to the proof of Theorem 4, we can obtain the following result for NOD random variable sequences.

Corollary 5. *Let $\{X_n; n \geq 1\}$ be a sequence of NOD random variables which is stochastically dominated by a random variable X and let $\{a_{ni}; 1 \leq i \leq n, n \geq 1\}$ be an array of real constants satisfying $\sum_{i=1}^n |a_{ni}|^\alpha = O(n)$ for some*

$0 < \alpha \leq 2$. Assume further that $EX_n = 0$ for $1 < \alpha \leq 2$ and $E|X|^\alpha \log(1 + |X|) < \infty$. Then,

$$\sum_{n=1}^{\infty} n^{-1} P\left(\left|\sum_{i=1}^n a_{ni} X_i\right| > \varepsilon b_n\right) < \infty \quad \forall \varepsilon > 0, \quad (11)$$

where $b_n = n^{1/\alpha}(\log n)^{1/\alpha}$.

Remark 6. In Theorem 4 and Corollary 5, we consider the case of $\alpha = \gamma$ for $0 < \alpha \leq 2$ and obtain some strong convergence results for arrays of rowwise NOD random variables and NOD random variable sequences without assumption of identical distribution. The main result settles the open problem posed by Huang and Wang [12]. In addition, it is still an open problem whether

$$\sum_{n=1}^{\infty} n^{-1} P\left(\max_{1 \leq j \leq n} \left|\sum_{i=1}^j a_{ni} X_{ni}\right| > \varepsilon b_n\right) < \infty \quad \forall \varepsilon > 0 \quad (12)$$

holds true under the same moment condition of Theorem 4.

3. Proofs

In order to prove our main results, the following lemmas are needed.

Lemma 7 (see Bozorgnia et al. [18]). Let $\{X_i; 1 \leq i \leq n\}$ be a sequence of NOD random variables, and let $\{f_i; 1 \leq i \leq n\}$ be a sequence of Borel functions all of which are monotone nondecreasing (or all are monotone nonincreasing). Then, $\{f_i(X_i); 1 \leq i \leq n\}$ is a sequence of NOD random variables.

Lemma 8 (see Asadian et al. [3]). Let $M \geq 2$ and let $\{X_n; n \geq 1\}$ be a sequence of NOD random variables with $EX_n = 0$ and $E|X_n|^M < \infty$ for all $n \geq 1$. Then, there exists a positive constant $C = C(M)$ depending only on M such that, for all $n \geq 1$,

$$E\left(\left|\sum_{i=1}^n X_i\right|^M\right) \leq C \left[\sum_{i=1}^n E|X_i|^M + \left(\sum_{i=1}^n EX_i^2\right)^{M/2} \right]. \quad (13)$$

Lemma 9. Let $\{X_n; n \geq 1\}$ be a sequence of random variables which is stochastically dominated by a random variable X . For any $u > 0$ and $t > 0$, the following two statements hold:

$$E|X_{ni}|^u I(|X_{ni}| \leq t) \leq C_1 (E|X|^u I(|X| \leq t) + t^u P(|X| > t)), \quad (14)$$

$$E|X_{ni}|^u I(|X_{ni}| > t) \leq C_2 E|X|^u I(|X| > t), \quad (15)$$

where C_1 and C_2 are positive constants.

Lemma 10 (see Sung [15]). Let X be a random variable and let $\{a_{ni}; 1 \leq i \leq n, n \geq 1\}$ be an array of real constants satisfying

$\sum_{i=1}^n |a_{ni}|^\alpha = O(n)$ for some $\alpha > 0$. Let $b_n = n^{1/\alpha}(\log n)^{1/\gamma}$ for some $\gamma > 0$. Then,

$$\begin{aligned} & \sum_{n=1}^{\infty} n^{-1} \sum_{i=1}^n P(|a_{ni} X| > b_n) \\ & \leq \begin{cases} CE|X|^\alpha, & \text{for } \alpha > \gamma, \\ CE|X|^\alpha \log(1 + |X|), & \text{for } \alpha = \gamma, \\ CE|X|^\gamma, & \text{for } \alpha < \gamma. \end{cases} \end{aligned} \quad (16)$$

Lemma 11 (see Sung [19]). Let X be a random variable and let $\{a_{ni}; 1 \leq i \leq n, n \geq 1\}$ be an array of real constants satisfying $a_{ni} = 0$ or $|a_{ni}| > 1$ and $\sum_{i=1}^n |a_{ni}|^\alpha = O(n)$ for some $\alpha > 0$. Let $b_n = n^{1/\alpha}(\log n)^{1/\alpha}$. If $q > \alpha$, then

$$\begin{aligned} & \sum_{n=1}^{\infty} n^{-1} b_n^{-q} \sum_{i=1}^n E|a_{ni} X|^q I(|a_{ni} X| \leq b_n) \\ & \leq CE|X|^\alpha \log(1 + |X|). \end{aligned} \quad (17)$$

Throughout this paper, let $I(A)$ be the indicator function of the set A . C denotes a positive constant, which may be different in various places and $a_n = O(b_n)$ stands for $a_n \leq Cb_n$.

Proof of Theorem 4. Without loss of generality, suppose that $\sum_{i=1}^n |a_{ni}|^\alpha \leq Cn$ and $a_{ni} \geq 0$, for all $1 \leq i \leq n, n \geq 1$. For fixed $n \geq 1$, define

$$\begin{aligned} X_i^{(n)} &= -b_n I(a_{ni} X_{ni} < -b_n) + a_{ni} X_{ni} I(|a_{ni} X_{ni}| \leq b_n) \\ & \quad + b_n I(a_{ni} X_{ni} > b_n), \quad i \geq 1, \end{aligned} \quad (18)$$

$$T_n^{(n)} = \sum_{i=1}^n (X_i^{(n)} - EX_i^{(n)}).$$

Denote

$$\begin{aligned} A &= \bigcap_{i=1}^n (a_{ni} X_{ni} = X_i^{(n)}), \\ B &= \bar{A} = \bigcup_{i=1}^n (a_{ni} X_{ni} \neq X_i^{(n)}) \\ &= \bigcup_{i=1}^n (|a_{ni} X_{ni}| > b_n), \end{aligned} \quad (19)$$

$$E_n = \left(\left| \sum_{i=1}^n a_{ni} X_{ni} \right| > \varepsilon b_n \right).$$

It is easily seen that, for all $\varepsilon > 0$,

$$\begin{aligned} E_n &= E_n A \cup E_n B \subset \left(\left| \sum_{i=1}^n X_i^{(n)} \right| > \varepsilon b_n \right) \\ & \cup \left(\bigcup_{i=1}^n |a_{ni} X_{ni}| > b_n \right), \end{aligned} \quad (20)$$

which implies that

$$\begin{aligned} P(E_n) &\leq P\left(\left|\sum_{i=1}^n X_i^{(n)}\right| > \varepsilon b_n\right) + P\left(\bigcup_{i=1}^n |a_{ni}X_{ni}| > b_n\right) \\ &\leq P\left(\left|T_n^{(n)}\right| > \varepsilon b_n - \left|\sum_{i=1}^n EX_i^{(n)}\right|\right) + \sum_{i=1}^n P(|a_{ni}X_{ni}| > b_n). \end{aligned} \quad (21)$$

First, we will prove that

$$b_n^{-1} \left| \sum_{i=1}^n EX_i^{(n)} \right| \rightarrow 0, \quad \text{as } n \rightarrow \infty. \quad (22)$$

Actually, for $0 < \alpha \leq 1$, by (14) of Lemma 9, Markov inequality, and $E|X|^\alpha \log(1 + |X|) < \infty$, we have that

$$\begin{aligned} b_n^{-1} \left| \sum_{i=1}^n EX_i^{(n)} \right| &\leq C b_n^{-1} \sum_{i=1}^n |EX_i^{(n)}| \\ &\leq C b_n^{-1} \sum_{i=1}^n E |a_{ni}X_{ni}| I(|a_{ni}X_{ni}| \leq b_n) \\ &\quad + C \sum_{i=1}^n P(|a_{ni}X_{ni}| > b_n) \\ &\leq C b_n^{-1} \sum_{i=1}^n (E |a_{ni}X| I(|a_{ni}X| \leq b_n) \\ &\quad + b_n P(|a_{ni}X| > b_n)) \\ &\quad + C \sum_{i=1}^n P(|a_{ni}X| > b_n) \\ &\leq C b_n^{-1} \sum_{i=1}^n (E |a_{ni}X| I(|a_{ni}X| \leq b_n)) \\ &\quad + C \sum_{i=1}^n P(|a_{ni}X| > b_n) \\ &\leq C b_n^{-\alpha} \sum_{i=1}^n (E |a_{ni}X|^\alpha I(|a_{ni}X| \leq b_n)) \\ &\quad + C b_n^{-\alpha} \sum_{i=1}^n E |a_{ni}X|^\alpha \\ &\leq C b_n^{-\alpha} \sum_{i=1}^n |a_{ni}|^\alpha E|X|^\alpha + C b_n^{-\alpha} \sum_{i=1}^n |a_{ni}|^\alpha E|X|^\alpha \\ &\leq C(\log n)^{-1} E|X|^\alpha \rightarrow 0 \quad \text{as } n \rightarrow \infty. \end{aligned} \quad (23)$$

Next, for $1 < \alpha \leq 2$, by $EX_{ni} = 0$, (15) of Lemmas 9 and 10, Markov inequality, and $E|X|^\alpha \log(1 + |X|) < \infty$, we also have that

$$\begin{aligned} b_n^{-1} \left| \sum_{i=1}^n EX_i^{(n)} \right| &\leq C \sum_{i=1}^n P(|a_{ni}X_{ni}| > b_n) \\ &\quad + C b_n^{-1} \left| \sum_{i=1}^n E a_{ni} X_{ni} I(|a_{ni}X_{ni}| > b_n) \right| \\ &\leq C \sum_{i=1}^n P(|a_{ni}X| > b_n) \\ &\quad + C b_n^{-1} \sum_{i=1}^n E |a_{ni}X_{ni}| I(|a_{ni}X_{ni}| > b_n) \\ &\leq C b_n^{-\alpha} \sum_{i=1}^n E |a_{ni}X|^\alpha \\ &\quad + C b_n^{-1} \sum_{i=1}^n E |a_{ni}X| I(|a_{ni}X| > b_n) \\ &\leq C b_n^{-\alpha} \sum_{i=1}^n E |a_{ni}X|^\alpha \\ &\quad + C b_n^{-\alpha} \sum_{i=1}^n E |a_{ni}X|^\alpha I(|a_{ni}X| > b_n) \\ &\leq C b_n^{-\alpha} \sum_{i=1}^n E |a_{ni}X|^\alpha + C b_n^{-\alpha} \sum_{i=1}^n E |a_{ni}X|^\alpha \\ &\leq C(\log n)^{-1} E|X|^\alpha \rightarrow 0 \quad \text{as } n \rightarrow \infty. \end{aligned} \quad (24)$$

From the above statements, we can get (22) immediately. Hence, for n large enough,

$$P(E_n) \leq P\left(\left|T_n^{(n)}\right| > \frac{\varepsilon b_n}{2}\right). \quad (25)$$

To prove (10), it is sufficient to show that

$$\begin{aligned} I &\triangleq \sum_{n=1}^{\infty} n^{-1} P\left(\left|T_n^{(n)}\right| > \frac{\varepsilon b_n}{2}\right) < \infty, \\ J &\triangleq \sum_{n=1}^{\infty} n^{-1} \sum_{i=1}^n P(|a_{ni}X_{ni}| > b_n) < \infty. \end{aligned} \quad (26)$$

It follows from Lemma 10 and $E|X|^\alpha \log(1 + |X|) < \infty$ that

$$\begin{aligned} J &\triangleq \sum_{n=1}^{\infty} n^{-1} \sum_{i=1}^n P(|a_{ni}X_{ni}| > b_n) \\ &\leq C \sum_{n=1}^{\infty} n^{-1} \sum_{i=1}^n P(|a_{ni}X| > b_n) \\ &\leq E|X|^\alpha \log(1 + |X|) < \infty. \end{aligned} \quad (27)$$

For fixed $n \geq 1$, it is easily seen that $\{X_i^{(n)} - EX_i^{(n)}, i \geq 1, n \geq 1\}$ is still a sequence of NOD random variables with mean zero

by Lemma 7. Hence, it follows from (14) of Lemmas 9 and 8 and Markov inequality (for $M > 2$) that

$$\begin{aligned}
 I &\triangleq \sum_{n=1}^{\infty} n^{-1} P\left(|T_n^{(n)}| > \frac{\varepsilon b_n}{2}\right) \\
 &\leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} E\left(|T_n^{(n)}|^M\right) \\
 &\leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \left[\sum_{i=1}^n E|X_i^{(n)}|^M + \left(\sum_{i=1}^n E|X_i^{(n)}|^2\right)^{M/2} \right] \quad (28) \\
 &\leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \sum_{i=1}^n E|X_i^{(n)}|^M \\
 &\quad + C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \left(\sum_{i=1}^n E|X_i^{(n)}|^2\right)^{M/2} \\
 &\triangleq I_1 + I_2.
 \end{aligned}$$

It follows from Lemma 10, (14) of Lemma 9, and Markov inequality that

$$\begin{aligned}
 I_1 &\triangleq C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \sum_{i=1}^n E|X_i^{(n)}|^M \\
 &\leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \left\{ \sum_{i=1}^n |a_{ni}|^M E|X_{ni}|^M (|a_{ni}X_{ni}| \leq b_n) \right. \\
 &\quad \left. + \sum_{i=1}^n b_n^M P(|a_{ni}X_{ni}| > b_n) \right\} \\
 &\leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \left\{ \sum_{i=1}^n |a_{ni}|^M E|X|^M I(|a_{ni}X| \leq b_n) \right. \\
 &\quad \left. + 2 \sum_{i=1}^n b_n^M P(|a_{ni}X| > b_n) \right\} \quad (29) \\
 &\leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \sum_{i=1}^n |a_{ni}|^M E|X|^M I(|a_{ni}X| \leq b_n) \\
 &\quad + C \sum_{n=1}^{\infty} n^{-1} \sum_{i=1}^n P(|a_{ni}X| > b_n) \\
 &\triangleq I_{11} + I_{12}.
 \end{aligned}$$

From Lemma 10 and $E|X|^\alpha \log(1 + |X|) < \infty$, we can obtain that

$$\begin{aligned}
 I_{12} &\triangleq C \sum_{n=1}^{\infty} n^{-1} \sum_{i=1}^n P(|a_{ni}X| > b_n) \quad (30) \\
 &\leq E|X|^\alpha \log(1 + |X|) < \infty.
 \end{aligned}$$

For fixed $n > 1$, we divide $\{a_{ni}, 1 \leq i \leq n\}$ into three subsets $\{a_{ni} : |a_{ni}| \leq 1/(\log n)^m\}$, $\{a_{ni} : 1/(\log n)^m < |a_{ni}| \leq 1\}$, and $\{a_{ni} : |a_{ni}| > 1\}$, where $m = (1/(M - \alpha))$. Then,

$$\begin{aligned}
 I_{11} &\triangleq C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \sum_{i=1}^n |a_{ni}|^M E|X|^M I(|a_{ni}X| \leq b_n) \\
 &= C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \\
 &\quad \times \sum_{i: |a_{ni}| \leq 1/(\log n)^m} |a_{ni}|^M E|X|^M I(|a_{ni}X| \leq b_n) \\
 &\quad + C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \\
 &\quad \times \sum_{i: 1/(\log n)^m < |a_{ni}| \leq 1} |a_{ni}|^M E|X|^M I(|a_{ni}X| \leq b_n) \\
 &\quad + C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \sum_{i: |a_{ni}| > 1} |a_{ni}|^M E|X|^M I(|a_{ni}X| \leq b_n) \\
 &\triangleq I_{11}^{(1)} + I_{11}^{(2)} + I_{11}^{(3)}. \quad (31)
 \end{aligned}$$

By Lemma 11 and $E|X|^\alpha \log(1 + |X|) < \infty$ again, it follows that

$$\begin{aligned}
 I_{11}^{(3)} &\triangleq C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \sum_{i: |a_{ni}| > 1} |a_{ni}|^M E|X|^M I(|a_{ni}X| \leq b_n) \quad (32) \\
 &\leq E|X|^\alpha \log(1 + |X|) < \infty \quad \text{for } M > 2 \geq \alpha > 0.
 \end{aligned}$$

Noting that $\sum_{i: |a_{ni}| \leq 1/(\log n)^m} |a_{ni}|^\alpha \leq Cn(\log n)^{-m\alpha}$, for $M > \alpha$ and fixed $n > 1$, we have that

$$\begin{aligned}
 I_{11}^{(1)} &\triangleq C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \sum_{i: |a_{ni}| \leq 1/(\log n)^m} |a_{ni}|^M E|X|^M I(|a_{ni}X| \leq b_n) \\
 &\leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-\alpha} \sum_{i: |a_{ni}| \leq 1/(\log n)^m} |a_{ni}|^\alpha E|X|^\alpha I(|a_{ni}X| \leq b_n) \\
 &\leq CE|X|^\alpha \sum_{n=1}^{\infty} n^{-1} b_n^{-\alpha} \sum_{i: |a_{ni}| \leq 1/(\log n)^m} |a_{ni}|^\alpha \\
 &\leq CE|X|^\alpha \sum_{n=1}^{\infty} n^{-1} n^{-1} (\log n)^{-1} n(\log n)^{-m\alpha} < \infty. \quad (33)
 \end{aligned}$$

Noting that $\sum_{i:1/(\log n)^m < |a_{ni}| \leq 1} |a_{ni}|^M \leq Cn$ and $m = 1/(M - \alpha)$, for $M > 2$, $0 < \alpha \leq 2$, we have that

$$\begin{aligned}
 I_{11}^{(2)} &\triangleq C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \sum_{i:1/(\log n)^m < |a_{ni}| \leq 1} |a_{ni}|^M E|X|^M I(|a_{ni}X| \leq b_n) \\
 &\leq C \sum_{n=1}^{\infty} b_n^{-M} E|X|^M I(|X| \leq b_n (\log n)^m) \\
 &= C \sum_{n=1}^{\infty} b_n^{-M} \sum_{k=1}^n E|X|^M I\left((k-1)^{1/\alpha} (\log(k-1))^{m+1/\alpha} < |X| \leq k^{1/\alpha} (\log k)^{m+1/\alpha}\right) \\
 &= C \sum_{k=1}^{\infty} E|X|^M I\left((k-1)^{1/\alpha} (\log(k-1))^{m+1/\alpha} < |X| \leq k^{1/\alpha} (\log k)^{m+1/\alpha}\right) \\
 &\quad \times \sum_{n=k}^{\infty} n^{-M/\alpha} (\log n)^{-M/\alpha} \\
 &\leq C \sum_{k=1}^{\infty} E|X|^M I\left((k-1)^{1/\alpha} (\log(k-1))^{m+1/\alpha} < |X| \leq k^{1/\alpha} (\log k)^{m+1/\alpha}\right) \\
 &\quad \times k^{1-M/\alpha} (\log k)^{-M/\alpha} \\
 &\leq C \sum_{k=1}^{\infty} E|X|^\alpha I\left((k-1)^{1/\alpha} (\log(k-1))^{m+1/\alpha} < |X| \leq k^{1/\alpha} (\log k)^{m+1/\alpha}\right) \\
 &\leq CE|X|^\alpha < \infty.
 \end{aligned} \tag{34}$$

Finally, we will prove that

$$I_2 \triangleq C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \left(\sum_{i=1}^n E|X_i^{(n)}|^2 \right)^{M/2} < \infty. \tag{35}$$

Hence, by C_r inequality, Markov inequality, Lemmas 9–11, and $E|X|^\alpha \log(1 + |X|) < \infty$, we have that

$$\begin{aligned}
 I_2 &\triangleq C \sum_{n=1}^{\infty} n^{-1} b_n^{-M} \left(\sum_{i=1}^n E|X_i^{(n)}|^2 \right)^{M/2} \\
 &= C \sum_{n=1}^{\infty} n^{-1} \left(\sum_{i=1}^n b_n^{-2} E|X_i^{(n)}|^2 \right)^{M/2} \\
 &\leq C \sum_{n=1}^{\infty} n^{-1} \left(\sum_{i=1}^n P(|a_{ni}X_{ni}| > b_n) \right)^{M/2}
 \end{aligned}$$

$$\begin{aligned}
 &+ C \sum_{n=1}^{\infty} n^{-1} \left(\sum_{i=1}^n b_n^{-2} E|a_{ni}X_{ni}|^2 (|a_{ni}X_{ni}| \leq b_n) \right)^{M/2} \\
 &\leq C \left(\sum_{n=1}^{\infty} n^{-1} \sum_{i=1}^n P(|a_{ni}X| > b_n) \right)^{M/2} \\
 &\quad + C \left(\sum_{n=1}^{\infty} n^{-1} \sum_{i=1}^n b_n^{-2} E|a_{ni}X|^2 (|a_{ni}X| \leq b_n) \right)^{M/2} \\
 &\leq C(E|X|^\alpha \log(1 + |X|))^{M/2} < \infty.
 \end{aligned} \tag{36}$$

Therefore, the desired result (10) follows from the above statements. This completes the proof of Theorem 4. \square

Conflict of Interests

The authors declare that they have no conflict of interests.

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