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Richness, diversity, and factors influencing occupancy of mammal communities across human-modified landscapes in Colombia.

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Highlights

- Agriculture reduces community diversity and evenness, while increases dominance.
- Forests and vicinity to water increase species occupancy; pastures decrease it.
- Forest areas are crucial for pumas, ocelots, raccoons, pacas, and agoutis.
- Wetlands are important for jaguars, the apex predator.
- Key to maintain forests and wetlands and target future crop expansion on pastures.

1 Abstract

2 As human-modified landscapes are increasing in the tropics, it becomes critical to understand 3 how they affect mammal communities to reconcile conservation and development. We 4 combined land cover information and camera-trapping data to explore the effects of 5 agricultural expansion on mammals in the Magdalena river valley of Colombia. We estimated 6 species diversity, evenness, and dominance across two agricultural landscapes, modified by 7 cattle ranching and oil palm cultivation. We further assessed which variables influence species-8 and community-level occupancy using multi-species occupancy models. Results highlight that 9 modified landscapes display lower species richness, diversity and evenness, and higher 10 dominance than more pristine sites. Residual forest cover and distance to water had significant 11 effect on community occupancy (positive and negative respectively). Forests were particularly 12 important for pumas, ocelots, lowland pacas, Central American agoutis, and crab-eating 13 raccoons while wetlands had a positive effect on jaguars, the apex predator in the region. The 14 influence of anthropogenic pressure was not clearly evident, though pastures were not valuable habitats for any mammal species, as they had a negative, yet not robust, effect on species and 15 community occupancy. In light of rapidly expanding agriculture across the tropics, our findings 16 highlight species-specific responses to disturbance that can inform land use planning and 17 18 conservation policies. We stress the conservation value of forest and wetland habitat to 19 mammal occupancy in heterogeneous ecosystems. Moreover, our results demonstrate that oil 20 palm and crop expansion should target existing pastures, which displayed limited conservation 21 value for Neotropical mammals but occupy vast swathes of land across Latin America.

22 Keywords: Panthera onca; camera trap; Bayesian statistics; land-use change; oil palm; pasture.

23 1. Introduction

Agricultural expansion is driving severe habitat loss and degradation, threatening biodiversity 24 worldwide (Foley et al., 2005; Green et al., 2005; Tscharntke et al., 2012; Maxwell et al., 25 26 2016). This is particularly concerning across tropical countries, which are extremely rich in 27 biodiversity but experiencing unprecedented land cover change (Gibbs et al., 2010; Laurance et 28 al., 2014). Therefore, there is an urgent need to understand how species respond to different 29 land cover types across modified landscapes to reconcile biodiversity conservation and 30 agricultural expansion across the tropics (Crooks et al., 2011; Rondinini et al., 2011). Amongst agricultural sector, oil palm cultivation is of particular concern. These plantations 31 32 have a negative effect on biodiversity, including mammals, and are expanding rapidly across the tropics (Fitzherbert et al., 2008; Yue et al., 2015; Wearn et al., 2017; Pardo et al., 2018a). 33 34 Tropical mammals are a conservation priority because they are declining due to land-use 35 change and hunting (Schipper et al., 2008; Visconti et al., 2011; Barlow et al., 2016) with 36 important consequences for ecosystem health and resilience. Mammals make significant contributions to ecosystem functioning including trophic regulation, nutrient cycling, carbon 37 38 storage, seed dispersal and ultimately maintenance of forest structure (Brodie et al., 2009; Jansen et al., 2010; Estes et al., 2011; Sobral et al., 2017). Amongst mammals, large-bodied 39 40 carnivores like jaguars *Panthera onca* are acutely threatened with extinction due to their slow 41 population growth rates and extensive area and dietary requirements (Crooks, 2002; Cardillo et 42 al., 2005; Carbone et al., 2011).

43 Conservation and management have largely moved away from single-species approaches to
44 sustaining ecosystems and communities (Balmford et al. 2005; Santini et al., 2017). However,
45 developing interventions to safeguard forest-dwelling mammals is hindered by their cryptic
46 nature. Statistical approaches, like occupancy modelling, that account and correct for imperfect

detection are, therefore, essential to provide a reliable evidence-base to support environmental
policies (MacKenzie et al. 2002; Brodie et al. 2014; Guillera-Arroita 2017). Multi-species
occupancy models (Dorazio & Royle 2005) go a step further as they enable us to
simultaneously explore habitat and anthropogenic variables that influence both community and
single species distribution, while providing robust inference for species infrequently detected
during ecological surveys (Ahumada et al., 2011; Ahumada et al., 2013; Tobler et al., 2015;
Rich et al. 2016; Wearn et al. 2017; Deere et al., 2017).

Colombia is the 4th largest palm oil producer (Ocampo-Penuela et al., 2018), but it is also a 54 megadiverse country, covering 0.7% of the planet and hosting 10% of known biodiversity 55 56 (Mittermeier et al., 1997). Yet the country's biodiversity is understudied due to a sustained period of conflict that rendered much of the region inaccessible. To our knowledge, this 57 58 research is the first to study terrestrial mammals across agricultural landscapes in Colombia 59 using multi-species occupancy models. We combine high-resolution land cover maps and 60 camera trap data to achieve the following objectives: (1) assess how habitat availability and 61 anthropogenic pressure (pastures, oil palm cultivation, and human settlements) influence community and species occupancy; (2) estimate Shannon diversity and evenness, and Berger-62 63 Parker dominance across two agricultural landscapes in the Magdalena river-valley of 64 Colombia. We hypothesized that mammalian occupancy would respond positively to bottomup resources (i.e. habitat availability) and negatively to top-down anthropogenic pressures, 65 66 depending on species-specific habitat preferences and sensitivity to disturbance. The study sites 67 comprised the dominant land use in Latin America, cattle ranching, (Grau & Aide, 2008) and oil palm cultivation. In face of a rapid land use change in the tropics, this study provides 68 valuable information to inform management, land use planning and policies that reconcile 69 70 agricultural expansion and mammal conservation.

71 **2. Methods**

- 72 2.1 Study sites
- 73 We conducted the study at two sites in the central region of the Magdalena River valley,
- 74 Department of Santander, Colombia (Site-1 (7.3752N -73.8842E to 7.5404N -73.7118E; Site-
- 75 2: 5.3450N -72.8471E to 5.4365N -72.7607E) (Fig. 1). Both sites are situated in the central
- 76 part of the Magdalena river valley, in between the Central and Eastern Andes, in the
- 77 Department of Santander, Colombia. The Euclidean distance between the two study sites is 93
- 78 km. The overall region is part of the tropical forest biome and is rich in wetlands with no
- 79 altitudinal gradient (IDEAM et al., 2007). Mean annual temperature is 27 °C, and annual
- 80 precipitation ranges between 2100-2600 mm (IDEAM et al., 2007). Land tenure is primarily
- 81 *private (different owners)* and there are no national protected areas.
- 82 The region is considered important for several species, including keystone mammals such as
- 83 the jaguar (Payan-Garrido et al., 2013; Boron et al., 2016b) and endangered endemics like the
- 84 brown spider monkey (Ateles hybridus ssp. brunneus) and white-footed tamarin (Sanguinos
- 85 *leucopus*). However, most of the region's historical forest cover has been lost due to the
- 86 expansion of cattle ranching and oil-palm agro-industries, while the remaining natural areas
- 87 facing a high risk of conversion (Etter et al., 2006; Castiblanco et al., 2013; Link et al., 2013).
- 88 We chose these two sites because they are modified, to a large degree, by agriculture but still
- 89 retained top predators like jaguars and pumas *Puma concolor* as well as other declining species
- 90 of conservation importance. The dominant land cover classes across our study sites
- 91 include:pastures (Site 1: 312 km²/35%; Site 2: 244 km²/36%), wetlands (Site-1: 182 km² Site-
- 92 1/20%; Site-2: 233 km²/34% at Site-2), secondary forest (Site-1: 112 km²/12%; Site-2: 129
- 93 km²/19%), oil-palm plantations (Site-1:172 km²/19%, Site-2: 17 km²/2%), water (Site-
- 94 1:190km²/10%;Site-2:Site-2: 43km²/6%)), bare ground (Site-1: 24km²/3%; Site-2: 13

95	km ² /2%), settlements <0.4%, and roads <0.6% (Fig. 1) (Boron et al., 2018). We identified
96	these land cover types and their amounts across the study sites, which we defined by adding a
97	buffer of 9.2 km around the camera stations. This reflects the maximum distance moved by the
98	species with the largest home range in the study region, (jaguars; Nowell & Jackson, 1996).

2.2 Camera trapping

100	We deployed 47 camera stations between April and August 2014 at Site-1 and 26 stations
101	between September and December 2014 at Site-2 using a blocked design. This included dry
102	and wet seasons at both sites (April, May, October, November, December: wet months; June,
103	July, August, September: dry months). The minimum convex polygons connecting camera
104	stations were 154.8 km ² at Site-1 and 85.4 km ² at Site-2. We set-up the cameras following
105	standardized survey techniques for terrestrial mammals (Ahumada et al., 2011, 2013; Rovero et
106	al., 2014), in a grid at intervals of 1.6±0.3 km, and across the main land cover types of the
107	region: forests (N=35), wetlands (N=8), pastures (N=11), and oil palm plantations (N=13). The
108	remaining cameras (N=6) were located in transitional habitat between wetland and oil palm
109	"Edge" (Fig. 1). To optimize detection, we placed the cameras so that their field of view
110	would be facing low resistance travel routes such as wildlife trails (Cusack et al., 2015).
111	Twenty-eight (60%) and 19 (73%) stations were placed on roads/trails at Site-1 and Site-2
112	respectively and we took this into account in our modelling approach. Camera traps are
113	consistently able to detect terrestrial mammals ≥ 0.5 kg, which are what we refer to in this
114	manuscript (Rovero et al., 2010). We deployed Cuddeback Attack (model 1149) and Ambush
115	(model 1170) camera traps and secured them to a tree at a height of 35 cm from the forest
116	floor. Both camera models have an identical trigger speed (0.25 seconds) and a detection zone
117	of 3-4 m due to the high temperatures characteristic of the region.

2.3 Species richness, diversity, and evenness For each study site we produced species accumulation curves using EstimateS (Colwell, 2013).

120 Accumulation curves reach an asymptote once all detectable species have been recorded and 121 were produced using the rarefaction method with 1000 randomizations (Magurran, 2004; Ahumada et al., 2011). We also calculated a) Shannon diversity index ($D_{Shannon} = -\Sigma p_i \ln p_i$ where 122 123 p_i is the proportion of abundance for species i, N_i relative to the total abundance N); b) evenness index ($E_{Shannon} = D_{Shannon}/ln$ (S), where S in the number of species observed) and c) the 124 125 Berger–Parker dominance index ($D = N_i/N$ where N_i is the abundance of the most abundant species relative to the total abundance N) (Shannon, 1948; Berger & Parker, 1970; McCune et 126 127 al., 2002).

We defined species abundances as integer capture rates, and calculated them using the total number of independent capture events of that species divided by the number of trap-nights (TN) and expressed as integer records per 100 trap nights (Carbone et al., 2001; O'Brien et al., 2003). Independent capture events were defined as consecutive photographs of individuals of the same species taken more than 30 min apart (O'Brien et al., 2003). Capture rates may not reflect real abundance, however they still provide more information than just incidence records (Carbone et al., 2001; Sollmann et al., 2013).

135 2.4 Multi-species occupancy modelling

136 We performed statistical analysis at the scale of the camera trap station (e.g. Rovero et al., 137 2014; Rich et al., 2016), and defined a sampling unit as the circular area with a radius of 800m 138 around each camera station, corresponding to half the average distance between neighboring stations (Sollmann et al., 2012). Prior to analysis, we discarded species with fewer than five 139 detections (i.e. eastern cottontail Sylvilagus floridanus and grison Galictis victata). To reduce 140 141 zero inflation in the dataset, we constructed detection histories for each species grouping seven 142 consecutive camera trap nights into one sampling occasion (Alexander et al., 2016; Everatt et 143 al., 2014). This resulted in 12 sampling occasion at Site-1 and 14 at Site-2.

We implemented a hierarchical Bayesian multi-species model formulation (Dorazio & Royle,
2005) to estimate species and community occupancy while accounting for imperfect detection
(MacKenzie et al., 2006) (see Appendix 1 for model code). This class of model is composed of
two components: the state process, describing the ecological system, and the observation
process, defining the sampling protocol. Relative to occupancy, this distinction differentiates
between non-detection and true absence.

We modelled the occurrence (z) of species *i* at site *j* as a realization of a Bernoulli process, $z_{i,j}$ \sim Bern($\psi_{i,j}$), where $\psi_{i,j}$ represents the latent occupancy state. To account for imperfect detection of true occupancy, detection probability was estimated as a function of temporal replicate *k* at site *j*. We denoted detection as a second Bernoulli process, $x_{i,j,k} \sim \text{Bern}(p_{i,j,k} * z_{i,j})$, where $x_{i,j,k}$ represents the observed detection histories, and $p_{i,j,k}$ is the detection probability of species *i* for temporal replicate *k* at site *j*, conditional on species presence ($z_{i,j}=1$) (Zipkin et al., 2010).

We connected occurrence and detection models of observed taxa through an additional 156 157 hierarchical component that modeled coefficients from a community-level distribution 158 governed by hyper-parameters. This protocol assumes similarity in community responses to 159 covariates. Consequently, species-specific parameter estimates are a function of individual 160 detection histories and average community-level responses, which provides more robust estimation precision for infrequently observed species (Pacifici et al. 2014). Hyper-parameters 161 162 provide insights into community level responses to covariates of interest and among species variability (Kery and Royle, 2008). We defined the linear predictor for the state process (ψ) 163 and observation process (*p*) models as: 164

168 $logit(p_{i,j,k}) = v_{(i) \operatorname{Habitat}(j)} + \beta_{1i} \operatorname{Road}_j + \beta_{2i} \operatorname{Season}_j$ 169

¹⁶⁵ $logit(\psi_{i,j}) = \mu_{(i) \operatorname{Site}(j)} + \alpha_{1i} \operatorname{Distance} \operatorname{Settlement}_j + \alpha_{2i} \operatorname{Distance} \operatorname{Water}_j + \alpha_{3i} \% \operatorname{Forest} \operatorname{Cover}_j + \alpha_{4i} \% \operatorname{Pasture} \operatorname{Cover}_j + \alpha_{5i} \% \operatorname{Wetland} \operatorname{Cover}_j + \alpha_{6i} \% \operatorname{Oil} \operatorname{Palm} \operatorname{Cover}_j + \alpha_{7i} \operatorname{Season}_j$ 167

170 Occupancy and detection probabilities were defined on the logit scale, with site- and habitat-171 specific intercepts on the state and observation processes respectively. We specify intercepts and slopes as a function of individual taxa. Intercepts were specific to site (1 or 2) for 172 173 occupancy and land cover class for detection (i.e. forest, pasture, oil palm, wetland, and edge 174 between oil palm and wetland). We derived covariates from Landsat 8 satellite imagery using 175 Object Oriented Image Analysis (Bock et al., 2005) and eCognition Developer 9 software (see 176 Boron et al. 2018 for more details). Dominant land cover types (specified as % 177 Forest/Pasture/Wetland/Oil Palm Cover) were extracted as proportions across each sampling 178 unit with ArcMap 10.3. We measured the Euclidean distances of each camera station to water 179 and settlements, which averaged 0.60 ± 0.07 km and 4.85 ± 0.29 km respectively. We tested 180 for collinearity amongst covariates using a threshold value of |r| = 0.7 (Dormann et al., 2013) 181 and no covariate was highly correlated to others. As mammals can use roads and trails to 182 facilitate their movement (Cusack et al., 2015) we included a categorical covariate on detection 183 probability (0/1 representing on/off roads/trails respectively). We grouped roads and trails 184 together because the roads we refer to are not paved, ≤ 3 metres wide, not open to the public 185 (inside private lands), not used regularly and thus not very different to trails. We also tested the 186 effect of season (wet vs. dry) on both occupancy and probability of detection. To improve model convergence and place covariates on a comparable scale, we centered and standardized 187 188 (by subtracting the mean and dividing by the standard deviation of all the sites) all continuous 189 predictor covariates prior to analysis.

Parameter posterior distributions were estimated using Markov chain Monte Carlo (MCMC)
simulation and conducted in WinBUGS version 1.4.3, called through R version 3.3.0 using the
package "R2WinBUGS" (Sturtz *et al.* 2005). Hierarchical models were constructed using
uninformative priors for all parameters. Unless stated otherwise, we present parameter
estimates as mean values of the posterior distribution, accompanied by 95% Bayesian Credible

195 Intervals (BCI) to express uncertainty. Parameters were considered significantly influential if 196 their 95% BCI did not overlap zero and weakly influential if their 90% BCI did not overlap 197 zero. We ran three parallel MCMC chains for 75,000 iterations each, discarding the first 25,000 198 iterations during the burn-in process and thinning posterior samples by a value of 10. To assess convergence, trace plots were visually inspected for satisfactory mixing and the Gelman-Ruben 199 200 statistic (Gelman and Hill, 2007) was observed to ensure a value of <1.1 for all parameters. We implemented a posterior predictive check to assess model fit compared to a simulated dataset, 201 202 extracting Bayesian P-values as a numerical overview of the predictive distribution (~0.5 203 indicative of good model fit; 0.05<*P*>0.95 indicative of poor model fit).

204

205 **3. Results**

The total sampling effort resulted in 3069 and 1903 trap nights at Site-1 and Site-2

respectively. Cameras were active for an average of 66 trap nights at Site-1 and 73 at Site-2.

208 On average camera stations detected 5 species at both sites (Site 1: 5, range 0-12; Site 2: 5,

range 0-10). We recorded a total of 17 terrestrial mammal species at each site consisting of

210 different guilds and threat categories (Table 1) and three arboreal species (varied capuchins

211 Cebus versicolor, howler monkey Alouatta seniculus and red-tailed squirrel Sciurus

212 granatensis) that we excluded from analysis. Species accumulation curves indicate that we

213 likely recorded most species at Site-1, whereas a larger sampling effort would have been

required to characterise the mammal community at Site-2, as the curve does not reach an

asymptote (Fig. 2). Overall sites displayed identical Shannon species diversity (Site 2: 2.01 vs.

Site 1: 2.02) and species evenness (Site 2: 0.71 vs. Site 1: 0.71) and similar species dominance

217 (Site 2: 0.32 vs. Site 1: 0.40).

218 Species occupancy varied from 0.54 (ocelot *Leopardus pardalis*) to 0.09 (Spiny rat *Proechymis*

219	chrysaeolus) at Site-1 and from 0.31 (ocelot) to 0.12 (nine-banded armadillo Dasypus
220	novemcinctus) at Site-2 (Table 1). Bayesian p values detailing model adequacy fit for single
221	species (0.08-0.78) and the full model (0.51) indicated good to optimal fit (Table A2A1).
222	Proportion of forest cover and distance to water were found to have significantly positive and
223	negative influences respectively on community-level mammalian occupancy (Fig. 3 and Table
224	2). Pumas (0.985; 95% BCI: 0.185-1.876), ocelots (0.811; 95% BCI: 0.032 - 1.618), Central
225	American agoutis Dasyprocta punctuata (1.008; 95% BCI: 0.222 - 1.884), lowland pacas
226	Cuniculus paca (0.825; 95% BCI: 0.014 - 1.668), and crab-eating racoons Procyon
227	cancrivorus (0.771; 95% BCI: 0.012 - 1.538) (Fig. 3 and Table A2) demonstrated a strong
228	preference for areas dominated by forest cover. A further nine species were positively
229	influenced by forest cover, though to a lesser extent (90% BCI non-overlapping zero) (Fig. 3
230	and Table A2). Consistent negative impacts of pasture cover on occupancy was documented
231	for all species, though this trend was not significant at the community level t (Fig. 3 and Table
232	2). The negative effect of pasture was weakly influential for lowland pacas (-0.644; 90% BCI: -
233	1.3010.025), Central American agouti (-0.693; 90% BCI: -1.3250.092), and jaguarundis
234	(Herpailurus yagouaroundi) (-0.696; 90% BCI: -1.3780.045) (Fig. 3 and Table A2).
235	Wetlands increased jaguar occupancy (0.787; 90% BCI: 0.112 – 1.497); however, oil palm,
236	distance to settlements, and season had unclear effects on community occupancy driven by
237	high species variability (Tables 2 and A2, and Fig. 3).

Probability of detection across the community was not significantly affected by any of the

covariates (land cover types, season, and camera placement on roads) (Table 2), although

240 placing cameras on roads/trails significantly improved detection for jaguars (1.054; 95% BCI:

241 0.152-1.939), ocelots (0.869; 95% BCI: 0.171-1.612), crab-eating foxes *Cerdocyon thous*

242 (1.024; 95% BCI: 0.297 - 1.753) and jaguarundi detection (0.723; 90% BCI: 0.148-1.296)

243 (Table A3). We document species-specific covariate effects on occupancy and detection in

245

246 **4. Discussion**

As agricultural and human-modified landscapes are increasing across the tropics it becomes 247 248 crucial to understand how they affect species and communities to inform evidence-based 249 conservation interventions for species vulnerable to land-use change. Mammals are a key 250 component of tropical forest ecosystems yet their populations continue to decline (Schipper et 251 al., 2008; Visconti et al., 2011). We synthesized land cover information, and camera trapping 252 data using multi-species occupancy models to produce an assessment of Neotropical mammal 253 persistence in landscapes dominated by agriculture. Our results demonstrate that: a) modified 254 landscapes display lower species diversity and evenness, and higher dominance compared to non-modified landscapes in the Neotropics in the literature; b) remaining forest areas and 255 256 distance to water significantly influenced community level occupancy confirming our 257 hypothesis; c) the effect of anthropogenic pressures was not clearly evident, although pastures 258 were not valuable habitats for any mammal species and could therefore be targeted for future 259 crop expansion.

260 4.1 Species richness and community structure

261 Species detection was commensurate with survey effort. At Site-2 the survey effort (26

stations) was not sufficient to detect all mammal species. At Site 1, where sampling effort was
higher (47 stations), the number of terrestrial mammal species (17) was similar to what
recorded with comparable survey efforts in an oil palm landscape in the Colombian Llanos (16
sp.) (Pardo & Payan, 2015) and in the Volcan Barva Transect in Costa Rica (15 sp.), which is
situated in a highly fragmented landscape (Ahumada et al., 2011), while lower than in pristine

267 areas such as the Peruvian Amazon (Tobler et al., 2008; Tobler et al., 2015), suggesting 268 declining mammal richness with increasing habitat modification. However, richness may not 269 be the most informative metric. The effect of habitat modification on species richness may not 270 be immediately apparent because generalist species spread and/or colonise modified 271 landscapes, as suitable niches become available (Ewers & Disham, 2006; Bogoni et al., 2016). 272 Despite detecting different guilds (e.g. carnivores, herbivores, omnivores, insectivores) we recorded fewer species classified as threatened on the IUCN Red List (Vulnerable, Endangered 273 274 and Critically Endangered) than other studies focussing on Neotropical mammals (e.g. Tobler 275 et al. 2008; Payan, 2009; Ahumada et al. 2011), alluding to the sensitivity of vulnerable species 276 to human modification. The complete absence of browsing herbivores like tapirs Tapirus 277 terrestris and deer Mazama and Odocoileus sp. suggests that historical land transformation and 278 overhunting could have been drivers of mammal decline and local extinction (Rodriguez-279 Mahecha et al., 2006). Deer and tapir have been the preferred quarry of hunters for centuries as 280 their significant body size yields more bushmeat per unit of hunting effort (Redford & 281 Robinson n.d.; Jerozolimski & Peres, 2003).

282 Both sites displayed lower diversity and evenness, and higher dominance than the equivalent 283 values reported by Ahumada et al. (2011) for Neotropical mammal communities (Diversity: 284 2.5-3.0; Evenness: 0.91-0.93; Dominance: 0.09-0.14). Our results discussed so far confirm that while agricultural landscapes with remaining natural habitat cover still hold some potential for 285 286 medium-large mammal conservation (Daily et al., 2003; Cassano et al., 2012; Magioli et al., 287 2016), habitat loss changes mammal communities, decreasing diversity and increasing 288 dominance (Ahumada et al., 2011; Bogoni et al., 2016). Top predators like jaguars and pumas 289 were still present in both study areas, however, their prey community seems impoverished as armadillos, pacas, peccaries (Pecari tajacu), capybaras, and deer were absent or rare across 290 291 both sites. Therefore it is likely that puma survival depends on smaller prey such as widespread

- agoutis, while jaguars rely on aquatic prey like caimans *Caiman crocodilus* and turtles
- 293 *Podocnemis* and *Trachemys* sp. (Da Silveira et al., 2010).

294 4.2 Factors influencing species and community occupancy and detection probability In accordance with our hypotheses, the proportion of forest cover in each sampling unit had a 295 robust effect on mammalian occupancy, confirming that maintaining connectivity and forest 296 cover in agricultural regions is crucial to preserve functional assemblages (Magioli et al., 2016; 297 298 Zimbres et al., 2017; Pardo et al., 2018a;b), and wider biodiversity (Prescott et al., 2016). The 299 effect of forest cover was particularly strong for pumas, ocelots, lowland pacas, raccoons, and Central American agoutis. Pumas' and ocelots' preference for forest is well documented (e.g. 300 301 Paviolo et al., 2009; Davis et al., 2011; Massara et al. 2015), reflecting higher prey occupancy 302 (e.g. agoutis and pacas) coupled with lower human presence and disturbance. Proximity to 303 water was also a strong determinant of community occupancy (distance to water had a 304 significantly negative effect meaning community occupancy increases nearer to water). Water 305 is a fundamental need of all species. Positive associations with water likely reflect the use of 306 riparian forests for movement and dispersal as thanks to legislation these forests tend to be the 307 only ones to remain in modified landscapes like our study sites (Nunez-Regueiro et al., 2015). 308 Contrary to expectation, anthropogenic pressures did not show a significant effect on 309 community or species occupancy. The presence of oil palm, for example, did not have a 310 significant effect on species, but this is likely because it only occupies a small percentage of 311 the study sites. For example, Pardo et al. (2018b) report that Neotropical mammal communities 312 drastically change when oil palm cover reaches 45–75%. Considering the documented negative effect of oil palm on mammals (Fitzherbert et al., 2008; Yue et al., 2015; Wearn et al., 313 314 2017; Pardo & Payan, 2015; Pardo et al., 2018a;b), the expansion of this crop in the Neotropics remains a concern. Recent estimates suggest that 21.1 million hectares of land is potentially 315 316 suitable for oil palm expansion in Colombia (Pirker et al. 2016). Ensuring that this expansion is not met at the expense of tropical forest habitat should represent a regional conservationpriority.

When unavoidable, new plantations should be established on already modified areas like 319 320 pastures, which displayed limited conservation value for most mammal species. We found a negative (although not significant) effect of pasture on community occupancy. This was more 321 322 evident for jaguarundis, lowland pacas, and agoutis. The negative effect of pasture on pacas and agouti could be due to lack of habitat cover and food resources. Our findings add to a 323 324 growing body of literature demonstrating the low conservation value of pastures for a range of 325 taxa (e.g. birds, beetles, and herpetofauna) (Gilroy et al., 2015; Prescott et al., 2016). Pastures account for a large proportion of agricultural land in Latin America and generally have low 326 327 productivity (Grau & Aide, 2008). Thus directing oil palm expansion on pastures, albeit 328 intensifying use of land, would additionally enable to maximise food security, carbon storage, 329 and natural habitat cover (Garcia-Ulloa et al., 2012).

Finally, wetlands were important for jaguars, the apex predators, in agreement with previous
evidence (Soisalo & Cavalcanti, 2006; Quigley et al. 2017). Jaguar conservation in the study
region will ultimately depend on the preservation of wetlands and the aquatic prey they host
(Da Silveira et al., 2010).

334 In interpreting these outputs, it is important to note that covariate influences on rare species 335 occupancy may be underestimated due to the effects of shrinkage on parameter estimates. 336 Given the prevalence of generalist species, community averages are driven by the response of 337 these dominant species, thus, when statistical strength is borrowed for rare species, parameter 338 estimates are drawn towards the community mean. At the same time shrinkage allows us to 339 estimate occupancy for species that we would otherwise not be able to make reliable inferences for due to data limitations. In this respect, slight bias in parameter estimates is a fair trade-off 340 341 (Pacifici et al., 2014).

Placing cameras on roads/established trails did not have an effect on community-level
detection. However, it was important to improve detection of jaguars, ocelots, crab-eating
foxes, and jaguarundis without affecting the detection of other species. Thus placing cameras
on trails remains the preferred option to detect multiple mammal species including carnivores,
which are known to use trails (Cusack et al., 2015).

347 4.3 Conclusions and management implications

Unprotected and increasingly human-modified areas can play a vital for species long-term 348 survival and connectivity and thus it is important to further investigate how their composition 349 350 and structure can affect species persistence and community composition. Furthermore, multi-351 species approaches are particularly valuable for informing conservation strategies as they 352 enable us to move beyond single species to community impact assessments. Data on how 353 mammal assemblages respond differentially to agricultural habitats represents key information 354 to understand disturbance. Thus, similar studies to the one presented in this paper can add 355 empirical evidence to environmental decision making. More specifically this study shows that species diversity tends to be low in human modified landscapes while species dominance 356 357 increases. We can expect similar patterns in agricultural areas around the tropics. These effects 358 may take time to unfold due to extinction debts owed to a legacy of human disturbance 359 (Tilman et al., 1994). There is clear evidence that some species were more associated with 360 natural habitats (e.g. pumas, ocelots, jaguars, pacas) than others. Therefore they are more 361 likely to decline rapidly in modified landscapes with ramifications on their historical ranges 362 and thus Red List categories (Ewers & Didham, 2006; Bogoni et al., 2016).

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Overall, planning for agricultural activity needs to factor in displacement and absence of species vulnerable to land-use change in areas where they naturally occur. As oil palm expansion and agriculture continues across the tropics, including Colombia, it is critical to minimize its negative impact on biodiversity. This study helps refine conservation strategies

368	and policy, having demonstrated that to reconcile agricultural expansion with mammal
369	conservation, it is crucial to maintain natural forest cover, since it had a robust impact on
370	community occupancy and is irreplaceable, as well as wetlands specifically for jaguars, the
371	apex predators. Furthermore, it is important to conserve natural habitat cover across
372	agricultural landscapes because no mammals displayed strong affiliation for oil palm and
373	pasture. These findings can also inform land use planning and highlight that in order to avoid
374	further negative impacts on biodiversity, future oil palm and crop expansion could be directed
375	towards existing pastures, which have low productivity (Grau & Aide, 2008), and, as we
376	attested, hold limited value for terrestrial mammals, and biodiversity in general (Gilroy et al.,
377	2015; Prescott et al., 2016). Under these conditions, oil palm has the potential to expand
378	without posing severe threat to vertebrate species in Colombia (Ocampo Penuela et al., 2018).
379	However, this presuppose the existence of both, stronger regulatory approaches (e.g. land use
380	planning that takes into account High Conservation Value forests and zero deforestation), as
381	well as incentives that could help retain vital natural habitats and thus promote mixed
382	landscapes (Lambin et al., 2014; Boron et al., 2016a).
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688	Figure captions
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Fig. 1 Study map of the two agricultural study sites (Site-1 and Site-2) in the Magdalena river valley of Colombia, including land cover types and camera trap stations (Site-1: N=47; Site-2: N=26).

Fig. 2 Species accumulation curves across two study sites (Site-1 and Site-2) in the Magdalenariver valley of Colombia.

Fig. 3 Caterpillar plots delineating effects of covariates on single species and community-level occupancy across two sites in the Magdalena river valley of Colombia. Mean hyper-parameter values and 95% Bayesian Credible Interval (BCI) values are presented in the grey background to each plot, providing an insight to community level responses to covariates. Species-specific mean posterior summaries are denoted with points, while horizontal lines represent the associated 95% BCI. Significant species-specific associations are in blue (95% BCIs do not overlap zero). Other relevant associations are in light blue (90% BCIs do not overlap zero).

719 **Tables**

Table 1. Terrestrial mammal species recorded across two study sites (Site-1 and Site-2) in the
 Magdalena river valley of Colombia, their IUCN (2015) and Regional Red List (Rodriguez-

722 Mahecha et al. 2006) categories, occupancy (ψ) estimates and 95% Bayesian Credible Intervals

723 (95% BCIs). Site-specific estimates assume covariates are held at their average values DD= Data

724 Deficient, LC= Least Concern, NT= Near Threatened, VU= Vulnerable, EN= Endangered. The

greater grison and eastern cotton tail were excluded from the occupancy modelling due to the

726 low number of records.

Scientific name	Common name	IUCN Red List	Regional Red List	Site-1 ψ (95% BCI)	Site-2 ψ (95% BCI)
Apex carnivores					
Panthera onca	Jaguar	NT	VU	0.47 (0.24-0.77)	0.13 (0.03-0.33)
Puma concolor	Puma	LC	NT	0.32 (0.15-0.55)	0.16 (0.05-0.39)
Meso carnivores					
Cerdocyon thous	Crab-eating fox	LC	/	0.26 (0.08-0.61)	0.13 (0.02-0.36)
Galictis victata	Greater grison	LC	/	Not recorded	Recorded once
Herpailurus yagouaroundi	Jaguarundi	LC	/	0.37 (0.17-0.66)	0.13 (0.02-0.35)
Leopardus pardalis	Ocelot	LC	NT	0.54 (0.33-0.76)	0.31 (0.12-0.66)
Omnivores					
Didelphis marsupialis	Common opossum	LC	/	0.24 (0.09-0.53)	0.14 (0.03-0.38)
Eira barbara	Tayra	LC	/	0.28 (0.12-0.57)	0.28 (0.11-0.66)
Pecari tajacu	Collared peccary	LC	/	0.14 (0.03-0.38)	0.20 (0.05-0.55)
Proechymis chrysaeolus	Spiny rat	DD	/	0.09 (0.02-0.28)	0.16 (0.04-0.43)
Procyon cancrivorus	Crab-eating raccoon	LC	/	0.39 (0.20-0.63)	0.20 (0.08-0.45)
Grazing herbivores					
Sylvilagus floridanus	Eastern cottontail	LC	/	Recorded once	Not recorded
Hydrochoerus isthmius	Lesser capybara	DD	/	0.26 (0.08-0.61)	0.13 (0.02-0.36)
Frugivores	Frugivores				
Cuniculus paca	Lowland paca	LC	/	0.17 (0.05-0.40)	0.17 (0.05-0.41)
Dasyprocta punctata	Central American agouti	LC	/	0.20 (0.08-0.41)	0.19 (0.06-0.43)

727	Insectivores					
728	Dasypus novemcinctus	Nine-banded armadillo	LC	/	0.19 (0.07-0.45)	0.12 (0.02-0.32)
729	Myrmecophaga tridactyla	Giant anteater	VU	VU	0.21 (0.06-0.51)	0.14 (0.03-0.39)
	Tamandua tetradactyla	Lesser anteater	LC	/	0.16 (0.05-0.38)	0.19 (0.06-0.47)
	Total N. of species				17	17

730 Table 2. Hyper-parameter posterior summaries denoting community-level responses to

731 covariates hypothesized to influence occupancy (ψ) and detection (*p*) of 16 mammal species at

two study sites in the Magdalena river valley, Colombia. We present mean predicted posterior

radius summary values and 95% Bayesian Credible Interval (95% BCI). Bold denotes covariates with

right significant effects on community occupancy.

	Mean Sigma		95% BCI
Intercept_ ψ Site-1	-1.02	0.92	-1.71 -0.36
Intercept_ ψ Site-2	-1.56	0.71	-2.40 -0.76
ψ _Distance settlements	-0.11	0.38	-0.43 0.22
ψ_Distance water *	-0.39	0.32	-0.77 -0.03
ψ_Forest *	0.72	0.37	0.09 1.34
ψ_Pasture	-0.42	0.42	-1.14 0.25
ψ_W etland	0.18	0.45	-0.51 0.86
ψ_Oil Palm	0.36	0.44	-0.33 1.07
Ψ_Season	0.23	0.50	-0.45 0.89
Intercept_P_Forest	-1.66	1.07	-2.42 -0.93
Intercept_P_Wetland	-1.36	0.68	-2.17 -0.65
Intercept_P_Edge	-1.19	0.56	-1.88 -0.54
Intercept_P_Pasture	-1.29	0.68	-2.08 -0.60
Intercept_P_Oil Palm	-1.64	1.68	-2.77 -0.59
P_Roads/Trails	0.15	0.75	-0.50 0.73
P_Season	0.09	0.42	-0.29 0.47

⁷³⁵

736

737 Supplementary information

738

739 Appendix 1. Model code used in the analysis.740

741 ### Hierarchical community occupancy model to assess response of neotropical mammals to

```
742
       ### habitat and anthropogenic covariates (written in BUGS language).
743
       ### Based on the Dorazio-Royle Community Model (2005; J. Am. Stat. Assoc., 100, 389-398).
       ### Implemented using code adapted from Zipkin et al. (2010; Biol. Conserv., 143, 479-484).
744
745
       ### Species-specific slopes and intercepts.
       ### Random intercepts on the process/occupancy model to account for sampling at
746
       geographically
747
748
       ### distinct sites (N=2).
749
       ### Random intercepts on the observation model to account for variability in detectability
750
       relative \#\#\# to habitat type (N=5).
751
752
       sink("OM Colombia M1.txt")
753
       cat("
          model{
754
755
756
          ### Hyperparameters for site-specific intercepts
757
          ###=
758
          for (s in 1:2) {
              a1.mean[s] \sim dunif(0,1)
759
              mu.a1[s] \le log(a1.mean[s]) - log(1-a1.mean[s])
760
761
              tau.a1[s] \sim dgamma(0.1,0.1)
              sigma.a1[s] <- 1/sqrt(tau.a1[s])</pre>
762
763
764
          ### Hyperparameters for habitat-specific intercepts
765
          ###==
          for (h in 1:5) {
766
767
              a2.mean[h] \sim dunif(0,1)
              mu.a2[h] \le log(a2.mean[h]) - log(1-a2.mean[h])
768
              tau.a2[h] \sim dgamma(0.1,0.1)
769
770
              sigma.a2[h] <- 1/sqrt(tau.a2[h])
771
          }
772
          ### Hyperparameters for fixed effects on the process/occupancy model
773
774
          ###==
775
          mu.b1 \sim dnorm(0,0.001)
776
          mu.b2 \sim dnorm(0,0.001)
777
          mu.b3 \sim dnorm(0,0.001)
778
          mu.b4 \sim dnorm(0,0.001)
          mu.b5 \sim dnorm(0,0.001)
779
780
          mu.b6 \sim dnorm(0,0.001)
781
          mu.b7 \sim dnorm(0,0.001)
          tau.b1 ~ dgamma(0.1,0.1)
782
          tau.b2 \sim dgamma(0.1,0.1)
783
          tau.b3 \sim dgamma(0.1,0.1)
784
785
          tau.b4 \sim dgamma(0.1,0.1)
786
          tau.b5 \sim dgamma(0.1,0.1)
          tau.b6 \sim dgamma(0.1,0.1)
787
          tau.b7 \sim dgamma(0.1,0.1)
788
789
          sigma.b1 < -1/sqrt(tau.b1)
          sigma.b2 <- 1/sqrt(tau.b2)
790
          sigma.b3 < -1/sqrt(tau.b3)
791
```

```
792
                       sigma.b4 <- 1/sqrt(tau.b4)
793
                       sigma.b5 <- 1/sqrt(tau.b5)
                       sigma.b6 < -1/sqrt(tau.b6)
794
795
                       sigma.b7 < -1/sqrt(tau.b7)
796
797
                       ### Hyperparameters for fixed effects on the observation/detection model
798
                       ###==
799
                       mu.p1 \sim dnorm(0,0.001)
                       mu.p2 \sim dnorm(0,0.001)
800
801
                       tau.p1 ~ dgamma(0.1,0.1)
802
                       tau.p2 ~ dgamma(0.1,0.1)
                       sigma.p1 <- 1/sqrt(tau.p1)</pre>
803
                       sigma.p2 <- 1/sqrt(tau.p2)
804
805
806
                       ### Create priors for species i from the hyperparameters
807
                       ###=
808
                       for (i in 1:(n+nzeroes)) {
809
                       ### Species and site-specific intercepts for occupancy and detection
810
811
                       ###=
812
                                 for(s in 1:2) {
                                            a1[s,i] \sim dnorm(mu.a1[s], tau.a1[s])
813
814
                                            }
815
                                 for(h in 1:5) {
816
817
                                          a2[h,i] \sim dnorm(mu.a2[h], tau.a2[h])
818
                                          }
819
820
                       ### Species-specific occupancy and detection fixed effects drawn from a normal distribution
                       ### governed by community-level hyperparameters
821
822
                       ###=
                                 b1[i] \sim dnorm(mu.b1, tau.b1)
823
                                 b2[i] \sim dnorm(mu.b2, tau.b2)
824
                                 b3[i] \sim dnorm(mu.b3, tau.b3)
825
                                 b4[i] \sim dnorm(mu.b4, tau.b4)
826
                                 b5[i] \sim dnorm(mu.b5, tau.b5)
827
                                 b6[i] \sim dnorm(mu.b6, tau.b6)
828
                                 b7[i] \sim dnorm(mu.b7, tau.b7)
829
830
831
                                 p1[i] ~ dnorm(mu.p1, tau.p1)
                                 p2[i] \sim dnorm(mu.p2, tau.p2)
832
833
834
                       ### Process model of true occurrence
                       ###=
835
836
                                 for (j \text{ in } 1:J) {
                                            logit(psi[j,i]) \le a1[Site[j],i] + b1[i]*dist.sett[j] + b2[i]*dist.water[j] + b2[i]*dis
837
                                                                                  b3[i]*forest[j] + b4[i]*pasture[j] + b5[i]*wetland[j] + b6[i]*OP[j] +
838
839
                                                                                  b7[i]*Season.psi[j]
                                           Z[j,i] \sim dbern(mu.psi[j,i])
840
841
```

842 843 <i>‡</i>	### Observation model for replicated detections
843 + 844 845 846 847 848	for (k in 1:K[j]) { logit(p[j,k,i]) <- a2[habitat[j],i] + p1[i]*road[j] + p2*season.det[j] mu.p[j,k,i] <- p[j,k,i]*Z[j,i] y[j,k,i] ~ dbern(mu.p[j,k,i])
849 # 850 # 851 # 852 #	 Calculate Pearson's chi-squared residuals to assess goodness of fit Based on Kery and Royle: Applied hierarchical modelling in ecology, pp. 235 Calculate the observed and expected residuals Add small value to prevent division by zero
853 # 854 855 856 857 858	y.sim[j,k,i] ~ dbern(mu.p[j,k,i]) chi2.actual[j,k,i] <- pow(y[j,k,i] - mu.p[j,k,i], 2)/ (mu.p[j,k,i] + 0.0001) chi2.sim[j,k,i] <- pow(y.sim[j,k,i] - mu.p[j,k,i], 2)/ (mu.p[j,k,i] + 0.0001) }
858 859 860 861	chi2.actual.sum[j,i] <- sum(chi2.actual[j,1:n.reps[j],i]) chi2.sim.sum[j,i] <- sum(chi2.sim[j,1:n.reps[j],i]) }
	# Calculate chi-squared discrepency for each species
364 365 366 367 368 369	<pre>fit.sp.actual[i] <- sum(chi2.actual.sum[,i]) fit.sp.sim[i] <- sum(chi2.sim.sum[,i]) c.hat.sp[i] <- fit.sp.actual[i]/fit.sp.sim[i] bpv.sp[i] <- step(fit.sp.sim[i] - fit.sp.actual[i]) }</pre>
#	# Calculate overall chi-squared discrepency measure
f c	fit.actual <- sum(chi2.actual.sum[1:n.sites, 1:n.sp]) fit.sim <- sum(chi2.sim.sum[1:n.sites, 1:n.sp]) c.hat <- fit.actual/fit.sim opv <- step(fit.sim - fit.actual)
#	# Derived quantities # Difference in occupancy between Sites 1 and 2 #====================================
'	for(i in 1:n.sp){ effect1[i] <- a1[2,i] - a1[1,i] }
	# Number of occupied sites #====================================
1	for(i in 1:n.sp) { Nocc.fs[i] $\leq sum(Z[,i])$
	# Number of species occurring at each site #

892 893 894 895 896 897 898	<pre>for(j in 1:n.sites) { Nsite[j] <- sum(Z[j,]) } ",fill=TRUE) sink()</pre>
899	
900	
901	
902	
903	
904	
905	
906	
907	
908	
909	
910	
911	
912	
913	
914	
915	
916	
917	
918	
919 920	Table A1. Bayesian P values detailing model adequacy fit for single species and the full model. Values between 0.025 and 0.975 indicate good fit, 0.5 is deemed optimal fit.

Species	Bayesian P Value
Central American agouti	0.60
Nine-banded armadillo	0.46
Capybara	0.62
Crab-eating fox	0.08
Giant anteater	0.52
Jaguar	0.53
Jaguarundi	0.59
Lesser anteater	0.46
Ocelot	0.39
Opossum	0.53
Lownland paca	0.33
Collared peccary	0.78
Puma	0.42
Crab-eating racoon	0.44
Spiny rat	0.49
Tayra	0.38
Full Model	0.51

Table A2. Posterior means and 95% Bayesian Credible Intervals (95% BCI) detailing the effect of habitat cover (Forest, Pasture, Wetland, Oil Palm), proximity covariates (Dist Settl., Dist Water) and seasonality (Season) on Neotropical mammal occupancy. Covariates are considered significant if their 95% BCIs do not overlap zero and presented here in bold with an asterisk. We additionally present in bold covariates with a weak influence (90% BCI do not overlap zero).

	Forest	Pasture	Dist. settl.	Dist. Water	Wetland	Oil palm	Season
	(95% BCI)						
Capybara	0.469	-0.347	-0.354	-0.596	0.483	-0.004	0.283
	(-0.580 - 1.372)	(-1.360 - 0.629)	(-1.211 - 0.33)	(-1.427 - 0.071)	(-0.464 - 1.581)	(-1.281 - 1.054)	(-0.808 - 1.405)
Central American agouti	1.008*	-0.693	0.259	-0.484	-0.023	0.626	0.395
	(0.222 - 1.884)	(-1.735 - 0.177)	(-0.263 - 0.862)	(-1.114 - 0.083)	(-0.912 - 0.823)	(-0.269 - 1.614)	(-0.513 - 1.351)
Collared peccary	0.820	-0.523	-0.068	-0.354	0.207	0.366	0.260
	(-0.088 - 1.780)	(-1.627 - 0.435)	(-0.816 - 0.702)	(-1.049 - 0.366)	(-0.787 - 1.254)	(-0.612 - 1.37)	(-0.838 - 1.351)
Crab-eating fox	0.656	-0.276	-0.141	-0.213	0.005	0.355	0.432
	(-0.104 - 1.411)	(-1.104 - 0.517)	(-0.643 - 0.351)	(-0.721 - 0.335)	(-0.806 - 0.775)	(-0.466 - 1.188)	(-0.448 - 1.372)
Crab-eating racoon	0.771*	-0.075	-0.306	-0.355	0.207	0.040	0.447
	(0.012 - 1.538)	(-0.906 - 0.774)	(-0.865 - 0.200)	(-0.882 - 0.178)	(-0.626 - 1.043)	(-0.951 - 0.934)	(-0.448 - 1.433)
Giant anteater	0.638	-0.335	0.019	-0.288	0.338	0.341	-0.122
	(-0.275 - 1.546)	(-1.339 - 0.641)	(-0.674 - 0.776)	(-0.948 - 0.423)	(-0.623 - 1.440)	(-0.679 - 1.384)	(-1.483 - 0.972)
Jaguar	0.522	-0.599	0.197	-0.529	0.787	0.269	-0.130
	(-0.364 - 1.360)	(-1.682 - 0.332)	(-0.396 - 0.912)	(-1.234 - 0.098)	(-0.206 - 1.969)	(-0.757 - 1.267)	(-1.322 - 0.879)
Jaguarundi	0.732	-0.696	-0.245	-0.341	0.230	0.768	-0.196
	(-0.116 - 1.591)	(-1.827 - 0.255)	(-0.958 - 0.397)	(-0.984 - 0.316)	(-0.670 - 1.137)	(-0.241 - 2.071)	(-1.458 - 0.83)
Lesser anteater	0.681	-0.186	-0.278	-0.462	-0.191	0.532	0.167
	(-0.155 - 1.515)	(-1.087 - 0.749)	(-0.952 - 0.336)	(-1.123 - 0.157)	(-1.233 - 0.725)	(-0.374 - 1.545)	(-0.872 - 1.177)

Lowland paca	0.825 *	-0.644	0.220	-0.307	0.279	0.480	0.284
	(0.014 - 1.668)	(-1.743 - 0.278)	(-0.380 - 0.977)	(-0.917 - 0.340)	(-0.635 - 1.237)	(-0.453 - 1.468)	(-0.695 - 1.268)
Nine-banded	0.629	-0.120	-0.203	-0.405	-0.270	0.130	0.101
armadillo	(-0.238 - 1.467)	(-0.993 - 0.795)	(-0.880 - 0.425)	(-1.051 - 0.216)	(-1.361 - 0.657)	(-0.898 - 1.101)	(-0.982 - 1.109)
Ocelot	0.811 *	-0.532	-0.098	-0.350	0.135	0.370	0.504
	(0.032 - 1.618)	(-1.452 - 0.300)	(-0.624 - 0.427)	(-0.929 - 0.222)	(-0.693 - 0.949)	(-0.49 - 1.243)	(-0.452 - 1.581)
Opossum	0.707	-0.317	-0.105	-0.457	0.144	0.412	0.081
	(-0.149 - 1.555)	(-1.264 - 0.598)	(-0.786 - 0.559)	(-1.140 - 0.171)	(-0.812 - 1.103)	(-0.521 - 1.395)	(-1.040 - 1.107)
Puma	0.985 *	-0.512	-0.182	-0.300	0.183	0.143	0.577
	(0.185 - 1.876)	(-1.450 - 0.349)	(-0.745 - 0.340)	(-0.842 - 0.252)	(-0.691 - 1.062)	(-0.759 - 0.983)	(-0.335 - 1.623)
Spiny rat	0.615	-0.344	-0.319	-0.399	0.225	0.376	0.148
	(-0.291 - 1.475)	(-1.354 - 0.620)	(-1.094 - 0.329)	(-1.079 - 0.263)	(-0.734 - 1.218)	(-0.580 - 1.351)	(-0.960 - 1.173)
Tayra	0.712	-0.498	-0.100	-0.468	0.092	0.534	0.430
	(-0.160 - 1.621)	(-1.505 - 0.415)	(-0.711 - 0.556)	(-1.139 - 0.142)	(-0.848 - 1.027)	(-0.374 - 1.535)	(-0.565 - 1.517)

Table A3. Posterior means and 95% Bayesian Credible Intervals (95% BCI) detailing the effect of camera trap placement (Roads) and seasonality (Season) on Neotropical mammal detection probability. Covariates are considered significant if their 95% BCIs do not overlap zero and presented here in bold with an asterisk. We additionally present in bold covariates with a weak influence (90% BCIs do not overlap zero).

	Roads (95% BCI)	Season (95% BCI)
Capybara	-0.369 (-1.486 - 0.649)	0.022 (-0.784 - 0.778)
Central American agouti	0.177 (-0.741-1.070)	0.637 (-0.008 - 1.374)
Collared peccary	-0.562 (-2.039 - 0.708)	0.017 (-0.853 - 0.810)
Crab-eating fox	1.024* (0.297 - 1.753)	0.266 (-0.244 - 0.795)
Crab-eating racoon	0.286 (-0.518 - 1.083)	0.465 (-0.081 - 1.065)
Giant anteater	-0.195 (-1.578 - 1.096)	-0.114 (-1.020 - 0.677)
Jaguar	1.054 * (0.152 - 1.939)	-0.086 (-0.789 - 0.563)
Jaguarundi	0.723 (-0.159 - 1.634)	-0.031 (-0.811 - 0.698)
Lesser anteater	0.030 (-1.386 -1.346)	0.023 (-0.754 - 0.733)
Lowland paca	-0.042 (-1.117 - 0.942)	0.194 (-0.498 - 0.912)
Nine-banded armadillo	-0.071 (-1.450 - 1.185)	0.082 (-0.700 - 0.844)
Ocelot	0.869 * (0.171 - 1.612)	-0.053 (-0.599 - 0.468)
Opossum	-0.427 (-1.550 - 0.592)	0.080 (-0.711 - 0.858)
Puma	0.350 (-0.468 - 1.181)	-0.018 (-0.664 - 0.608)
Spiny rat	-0.439 (-2.034 - 0.853)	0.085 (-0.744 - 0.891)
Tayra	-0.047 (-1.044 - 0.912)	-0.203 (-0.953 - 0.445)

	Forest	Pasture	Edge	Wetland	Oil palm
	(95% BCI)				
Capybara	0.053	0.214	0.175	0.252	0.045
	(0.009 - 0.218)	(0.052 - 0.535)	(0.041 - 0.402)	(0.109 - 0.487)	(0.001 - 0.582)
Central American agouti	0.555	0.192	0.287	0.156	0.154
	(0.313 - 0.779)	(0.030 - 0.540)	(0.141 - 0.516)	(0.020 - 0.490)	(0.043 -0.435)
Collared peccary	0.038	0.171	0.197	0.178	0.388
	(0.004 - 0.199)	(0.042 - 0.392)	(0.040 - 0.482)	(0.043 - 0.436)	(0.108 - 0.778)
Crab-eating fox	0.308	0.349	0.324	0.300	0.281
	(0.162 - 0.514)	(0.206 - 0.537)	(0.165 - 0.572)	(0.136 - 0.567)	(0.127 - 0.499)
Crab-eating racoon	0.302	0.265	0.224	0.157	0.075
	(0.144 - 0.530)	(0.134 - 0.458)	(0.100 - 0.409)	(0.047 - 0.348)	(0.012 - 0.362)
Giant anteater	0.105	0.227	0.199	0.214	0.048
	(0.024 -0.358)	(0.075 - 0.489)	(0.055 - 0.433)	(0.064 - 0.506)	(0.001 - 0.591)
Jaguar	0.086	0.239	0.285	0.334	0.026
	(0.030 - 0.212)	(0.072 - 0.571)	(0.136 - 0.527)	(0.160 - 0.598)	(0.001 - 0.400)
Jaguarundi	0.077	0.194	0.263	0.210	0.057
	(0.025 - 0.210)	(0.032 - 0.549)	(0.120 - 0.488)	(0.074 - 0.449)	(0.010 - 0.224)
Lesser anteater	0.174	0.170	0.204	0.177	0.099
	(0.050 - 0.491)	(0.052 - 0.365)	(0.044 - 0.503)	(0.027 - 0.514)	(0.019 - 0.343)

Table A4. Average detection probability, including 95% Bayesian Credible Interval, for Neotropical mammals in five reference habitat classes.

Lowland paca	0.180	0.320	0.215	0.159	0.152
	(0.066 - 0.400)	(0.134 - 0.635)	(0.077 - 0.430)	(0.035 - 0.401)	(0.023 - 0.612)
Nine-banded armadillo	0.137	0.186	0.207	0.172	0.318
	(0.041 - 0.366)	(0.054 - 0.407)	(0.043 - 0.503)	(0.025 - 0.505)	(0.046 - 0.864)
Ocelot	0.179	0.166	0.305	0.190	0.285
	(0.094 - 0.313)	(0.048 - 0.363)	(0.146 - 0.566)	(0.085 - 0.358)	(0.156 - 0.456)
Opossum	0.190	0.187	0.194	0.265	0.035
	(0.072 - 0.413)	(0.031 - 0.525)	(0.051 - 0.431)	(0.103 - 0.561)	(0.001 - 0.397)
Puma	0.192	0.174	0.197	0.188	0.488
	(0.091 - 0.357)	(0.027 - 0.496)	(0.079 - 0.381)	(0.054 - 0.438)	(0.193 - 0.790)
Spiny rat	0.105	0.191	0.240	0.176	0.688
	(0.020 - 0.392)	(0.032 - 0.545)	(0.092 - 0.471)	(0.029 - 0.506)	(0.273 - 0.950)
Tayra	0.181	0.205	0.226	0.148	0.133
	(0.072 - 0.398)	(0.077 - 0.406)	(0.085 - 0.450)	(0.031 - 0.375)	(0.032 - 0.409)





