# Out-of-pocket costs, primary care frequent attendance and sample selection: estimates from a longitudinal cohort design

Pymont, C., McNamee, P., & Butterworth, P

Abstract

This paper examines the effect of out-of-pocket costs on subsequent frequent attendance in primary care using data from the Personality and Total Health (PATH) Through Life Project, a representative community cohort study from Canberra, Australia. The analysis sample comprised 1,197 respondents with two or more GP consultations, and uses survey data linked to administrative health service use (Medicare) data which provides data on the number of consultations and out-of-pocket costs. Respondents identified in the highest decile of GP use in a year were defined as Frequent Attenders (FAs). Logistic regression models showed that out-of-pocket costs incurred during respondents' prior two consultations were significantly associated with subsequent FA status. Respondents who incurred higher costs (\$15 to \$35; or >\$35) were less likely to become FAs than those who incurred no or low (<AUS\$15 per consultation) costs, with no difference evident between the no and low-cost groups. Differences were not explained by taking account of previous years FA status or out-ofpocket expenses, survey data on health conditions, socio-economic circumstances including receipt of income support payments, or other established risk factors for FA status. However, a counterfactual model that adjusted for factors associated with the selection into payment levels did not find an influence of payment, with only a non-significant gradient in the expected direction was evident. Hence these findings raise doubts that price drives FA behaviour and suggests that co-payments are unlikely to be an effective intervention strategy.

Keywords: Frequent attendance; out-of-pocket expenses; Primary health care

### Introduction

In 2014 the Australian Health Minister noted that a small group of patients accounted for a large proportion of overall government funded health services and proposed a need to reconsider the way these patients were managed.[1] The most prolific users of primary health care services (frequent attenders (FAs)), commonly defined as the top 10% of attenders in a year, have been found to account for 33% of GP consultations[2] and generate five times as many prescriptions and hospital contacts as other patients.[3] FA's 3-year expenditures have been found to be higher than non-FAs, even after adjustment for patient and health care provider characteristics.[4] Data from Australia's universal health insurance scheme (Medicare) from 2012-13 show the top 12.5% of general practitioner (GP) attenders accounted for 41% of (non-hospital) Medicare expenditure.[5] Despite FA being signalled as a potential point for intervention it is not clear if frequent attendance necessarily reflects overuse of services or whether out-of-pocket costs are a determinant of their behaviour.

Under the Australian Medicare system, a scheduled fee is set for each type of health service or consultation. GPs can opt to accept this scheduled fee from Medicare and not charge their patients directly. Alternatively, GPs can charge patients an amount greater than the scheduled fee. In these cases, the patient can receive a rebate of up to 100% of the scheduled fee but the difference represents an out-of-pocket expense borne by the patient. Recent Medicare data indicated that over 80% of in-scope GP consultations incur no patient out-of-pocket costs.[6] Consultations that incur no out-of-pocket costs are positively associated with chronic disease, having a concession card, and having private health insurance and negatively with larger practice size, having an appointment for the visit, higher household income, and region of residence.[7] While Medicare covers the majority of consultations it does not cover all primary care consultations in Australia, excluding Department of Veterans' Affairs beneficiaries, patients receiving treatment under compensation agreements, and some telephone helpline or extended hours (nurse-led) walk in clinics. (for more information see[8])

Since universal health insurance was first introduced to Australia in 1975 there has been considerable variability in the proportion of GP consultations with no additional cost, suggesting that GPs decision on price charged may be sensitive to policy parameters and incentives. Concerns about unsustainable growth in health expenditure have prompted calls to introduce a price signal to reduce unnecessary and over use of health services.[9, 10] There has been much discussion about a mandatory co-payment[11, 12] and a freeze on the level of scheduled fees[13] which, over time, would increase pressure on GPs to charge above the set rebate.[14]

It is unclear if, and in what context, a mandatory co-payment or cost sharing would change attendance behaviour. In the USA, Medicaid recipients (who are enrolled in a private health plan which covers all or most of the recipient's healthcare needs) were more likely to be FAs than others.[15] Out-of-pocket expenses at the point of use have been shown to influence overall attendance at health care institutions[16-18] particularly when free.[19] There is some evidence of effects of co-payments in vulnerable populations including evidence of adverse health consequences for patients with heart failure and diabetes mellitus[20], and an impact on adherence to cardiovascular disease treatment [21] and attendance at obstetric emergency rooms.[22] Increases in co-payments in the US have been found to be related to decreased utilization of inpatient care, physician visits, brand-name medications, and emergency department visits.[23] The most comprehensive examination of co-payments, the RAND health insurance experiment, which has collected over 40 years of data, found that higher out-of-pocket expenses led to fewer medical visits and hospital admissions, [24, 25] and detrimental health effects for the sickest and poorest patients. In Australia, 14% of adults reported not attending the GP or getting appropriate care due to the cost [26, 27] including 24% of individuals with chronic health problems, [27] consistent with international evidence.[27, 28] While the ramifications of introducing co-payments in Australia are still being debated, the relationship between out-of-pocket expenses and frequent GP attendance remains unclear.[29]

We have previously used administrative Medicare data linked to longitudinal survey data to identify the characteristics of Australian primary-care FAs[2] and found that health related risk factors assessed in the survey explained over 50% of FA status and this increased a further 10% to 17% when the time varying nature of the risk factors was considered.[30] This research added to the literature linking FA status to a range of patient characteristics including socioeconomic status,[31] employment status (particularly unemployed),[32] being an immigrant,[3, 33] insecure attachment,[34] distress,[35] number of medical issues,[36] and with somatising and somatic illness.[37-39] However, there is little research examining the role of out-of-pocket expenses on frequent attendance behaviour.

The aim of this paper is to contribute to the ongoing policy debate in Australia and examine whether the costs patients incur for GP consultations influence their likelihood of becoming a FA. We assess whether the average costs incurred by a patient in two consultations influences subsequent attendance in the following 12-month period after controlling for a range of patient health and social circumstances (e.g., chronic physical conditions, medication use, mental health, and socioeconomic characteristics), and their previous year's health service use and costs incurred. The linkage of administrative data of attendance at primary health care with rich survey data on health (which provides an independent marker of need) allows for a unique investigation of potential drivers of attendance and the relative effect of out-of-pocket costs in relation to need based drivers which have been identified as important in previous research.[e.g. 2, 3, 4, 38] However, it is difficult to demonstrate causality with observational data as individuals are not randomly assigned to treatment groups. Patients who receive reduced or no cost consultations are potentially different from those who are charged more, and this introduces a possible source of bias in estimates of a causal effect of out-of-pocket costs on subsequent attendance. Previous studies using observational data to examine attendance have not controlled for potential selection bias. To strengthen the existing evidence base, our analysis also addresses such potential selection effects. Using a range of health and socioeconomic factors to model selection, we evaluate whether charging patients small, medium or large out-of-pocket costs influences the likelihood of frequent use of GP services compared to the application of no out-of-pocket expenses.

### Method

# Design

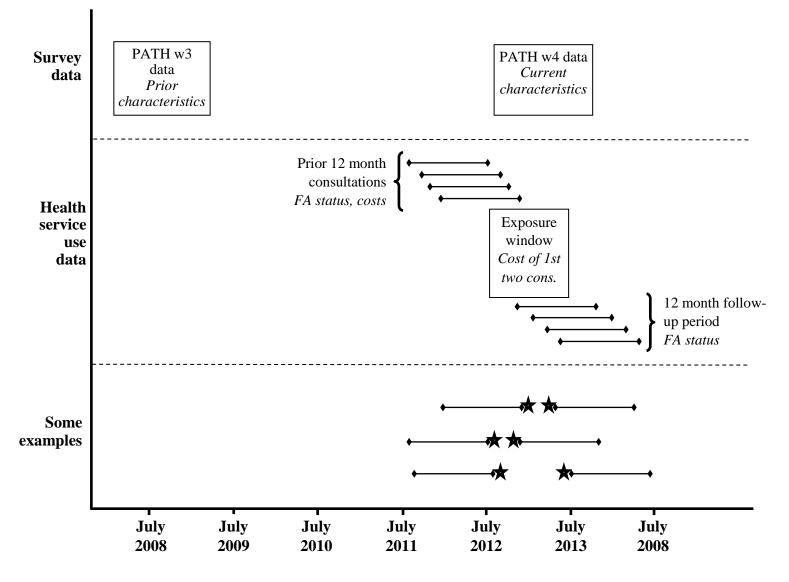
This study draws on data from the Personality and Total Health (PATH) Through Life Project, a longitudinal community study of health and wellbeing. The data, methods, and individual scales and measures are described in detail elsewhere [40] but briefly: the PATH project follows three narrow age-range cohorts, randomly sampled from the electoral rolls for Canberra and Queanbeyan and reassessed on four occasions. This analysis considers data from wave four interviews conducted in 2012/13 with the mid-aged cohort who were then aged between 52 and 58 years. Overall, 2257 respondents remained in-scope for wave four and were invited to participate. In addition, data from the previous wave are used in the counterfactual model to adjust for prior health and financial status. Respondents who remained resident in the local region (n = 1615) were invited to participate in a face-to-face interview, which included physical, cognitive and clinical assessment, and asked to complete a comprehensive survey questionnaire online. Of these, 1570 (97%) participated. The remaining 642 in-scope respondents who had moved from the Canberra region were invited to complete an online survey alone, with 236 (36.8%) participating. Participants were asked to consent to release their administrative health service use (Medicare) data from a four-year period, and 1591 (88%) gave consent. The analysis is further restricted to those respondents identified with at least two GP consultations during a 12-month study period (thereby excluding 399 respondents). Thus, the analysis sample comprises 1192 respondents.

The Human Research Ethics Committee of the Australian National University approved all aspects of the PATH study including data linkage and participants provided written informed consent.

### Measures

Data on out-of-pocket expenses for each individuals' first two consultations in the 12 months from July 2012 was used to define the exposure groups (average out-of-pocket costs). Analyses categorized the average cost of these two GP consultations as: none (both no-cost); low (less than \$15); medium (\$15 to \$30); and high (greater than \$30). These levels were chosen as \$15 represents a previously proposed co-payment level [41] and \$30 represents the median average payment in our cohort (further, the average ACT patient contribution for GP services in 2012/13 was \$34.40). Attendance during the 12 month period immediately following the second of these visits was used to calculate FA status. Hence, there was a consistent 12-month observation period for all participants but these could have different starting points. The analysis considers a comprehensive list of relevant GP Medicare item numbers (see [2]) representing all face-to-face Medicate services delivered by a GP. A cutpoint was applied to identify the (approximately) 10% of respondents (stratified by gender) with the greatest number of GP consultations consistent with the FA literature [4, 30]. Analysis is stratified by gender as attendance behaviour has previously been found to differ for men and women[30] and it was defined as the cut point which provided as close to 10% of each gender as possible. Similarly lagged FA status was determined by examining the consultations in the 12 months prior to the two consultations used to define the average cost taking the approximate top 10% stratified by gender.

This analysis considers the level of out-of-pocket expenses that participants incurred in these two consultations from July 2012 as the "treatment", and the FA status in the following 12-months as the outcome. For clarity, this is represented in Figure 1. Additionally Medicare codes related to management and referral for chronic conditions (diabetes, heart problems etc.) were included in the above measures as well as coded separately and reported in the demographics of the sample.



*Figure 1.* Timeline of measures. The timeline for which the treatment was defined was dependent on the consultation times of the participant hence some indicative examples are given where  $\bigstar$  indicates the consultations used to define out-of-pock costs for a participant.

Potential covariates in the PATH survey data were selected on the basis of the previous literature.

- Morbidity: Participants were asked if they experienced a range of chronic physical conditions (heart disease, cancer, arthritis, thyroid disorder, epilepsy, cataracts, asthma, diabetes and stroke). An additional question asking if the respondent had experienced a serious illness, injury or assault in the last 6 months was used as a marker of other medical issues not explicitly assessed. Each of the physical conditions and additional question were included separately in the models. While use of self-reported morbidity measures has some limitations (see discussion), it provides a measure of 'need' that is measured in a different manner to the exposure and outcome measures which are based on administrative data.
- Perceived health: Individual items from the SF-12 [42] assessed self-rated health, health-related impairment in daily activities and work, and pain.
- Health anxiety and depression: The 9-item Goldberg Depression scale was used as a measure of severity (the number of symptoms experienced: 0 to 9). Separate items from the Goldberg Anxiety scale were used to assess psychosomatic symptoms and aspects of health anxiety, including reported experience of i) headaches, ii) trembling, iii) sleep issues and iv) general worry about health.
- Socioeconomic characteristics: Labour-force status, educational attainment (higher education, completed high school vs not), receipt of welfare, low household income (less than \$575 per week) and the experience of financial hardship (having pawned or sold something, went without meals, could not heat home or sought help from welfare organisations) were included as measures of individuals' socioeconomic circumstances.
- Medication use: Participants reported on their current medication use, including
  medication for blood pressure, anxiety and depression, sleep problems, memory
  problems, cholesterol, contraception, hormone replacement therapy, pain relief or
  other any problems. All variables were coded to reflect current use/not, aside from use
  of pain relief medication which was coded as a binary variable to reflect more than
  weekly (as compared to less frequent) use.
- Life satisfaction: Life satisfaction was used as an alternative outcome measure, and assessed using the Satisfaction with Life Scale (Diener, Emmons, Larsen, & Griffin, 1985; Pavot & Diener, 1993).

## **Statistical analysis**

A series of longitudinal logistic regression models were initially used to investigate the association between out-of-pocket expenses (in four categories) for the first two consultations of the (financial) year with the likelihood of being subsequently classified as a FA. Model A only controlled for gender, Model B included lagged (prior-year) measures of FA status, average out-of-pocket costs, and lack of GP consultations (considered likely markers of the determinants of subsequent FA status) and Model C added the range of health, socioeconomic and medication-use measures, see Table 2 for the full list.

It is likely that differences in the out-of-pocket expenses paid by patients for their first two consultations (i.e., the "treatment") does not reflect a random process but is based on differences between patients that influence GP charging decisions, such as their level of need (health), their income (reflecting capacity to pay), the predisposition/default policy of the GP practice, or the patient's own active selection of a GP on the basis of billing practice. It may be these factors, rather than cost per se, that cause subsequent FA status. To take account of these limitations of observational data for causal inference, we draw upon the counterfactual framework[43] to implement an augmented inverse propensity weighted estimator (AIPW) to calculate Average Treatment Effects (ATEs) to overcome potential selection bias. [e.g. 44, 45] We used the teffect-procedure in STATA to estimate treatment effects. AIPW is a doubly-robust treatment effect approach that adjusts for factors associated with selection into different payment levels and equates groups via inverse probability weights and regression methods. The model used a multinomial logistic regression model to generate a weight reflecting the inverse probability of being in each payment "treatment" (the treatment model) and applied this to the outcome model which used regression to predict both potential outcomes (i.e., FA or non-FA status) for each respondent. A major advantage of this method is that it is robust against specification error in either the 'treatment model' or the 'outcome model'.[46]

We specify a treatment model including variables that would influence selection into payments levels; prior year FA status, prior (average) out-of-pocket costs, prior number of GP attendances, and previous income support receipt, physical and mental health, and medication use drawn from the wave 3 PATH data. We argue that these factors reflect the respondent differences that could cause differences in out-of-pocket expenses incurred. The outcome model included these terms, in addition to covariates reflecting current physical and mental health, current medication use, and current income support receipt. This model captures both likely long-term selection factors and recent changes that may be driving discontinuity in out-of-pocket costs (e.g., recent onset of chronic illness, improvement in financial circumstances).

A series of sensitivity analyses were conducted to assess the robustness of the current findings. These include repeating the regression models in a subgroup analysis limited to those with greater capacity to pay for services (i.e. excluding those respondents who were identified as not receiving welfare payments or who reported low income), a negative binomial regression considering the number of GP consultations (rather than FA status), and analysis operationalising the period of time consistently for all respondents using with calendar or financial year. Additionally, regression and selection models were rerun with life satisfaction as the outcome to examine if out-of-pocket expenses had an effect on a quality of life.

#### Results

FA status was defined separately by gender: with male FAs (10.1%) defined by seven or more consultations during the year while female FAs (10.2%) were those with 10 or more consultations. Overall FAs (10.18%) were responsible for 37.7% of all consultations. In total 35.8% of consultations incurred no out-of-pocket costs. 17.5% of GP patients were not required to make an out-of-pocket payment for any of their consultations, while 66.6% of patients had no consultations without an out-of-pocket cost. 21.0% of patients identified as FAs made no out-of-pocket payment for any of their consultations. Consistent with our previous findings,[2, 30] FA status was associated with diabetes, number of life events, low income, lower educational achievement, anxiety, depression, cholesterol, and pain medication use as evidenced in the regression models and preliminary analyses. Some demographics are presented in Table 1 while demographics across payment levels are presented in Supplemental Table 1.

# Table 1.

# Demographic summary

		Non-FAs	FAs	Overall	significance
N		1429	162		
Gender	(%Female)	52.8%	53.1%	52.8%	$\chi^2$ =.006, <i>p</i> =.938
Marital status	Married	69.4%	66.4%	69.1%	
	Separated/Divorced/Widowed	19.3%	24.8%	19.9%	
	Never Married	11.3%	8.7%	11.0%	χ <sup>2</sup> =3.30, <i>p</i> =.192
Employment	Employed	84.4%	67.1%	81.7%	
	Unemployed	1.9%	6.2%	2.5%	
	Not in the labour force	13.7%	26.7%	15.8%	$\chi^2$ =32.99, <i>p</i> <.001
Welfare receipt		4.4%	11.8%	5.2%	χ <sup>2</sup> =16.37, <i>p</i> <.001
Number of health	problems, Mean (SD)	1.1 (1.0)	1.7 (1.2)	1.2 (1.1)	t(1558)=-7.37, <i>p</i> <.001
Number of medica	tions, Mean (SD)	1.0 (1.1)	1.9 (1.2)	1.1 (1.1)	t(1558) = -9.44, <i>p</i> <.001
		\$28.31	\$17.01	\$27.00	t(1190) = 8.60, <i>p</i> < .001
Payment across the	e year Mean (SD)	(\$15.96)	(\$13.15)	(\$16.07)	
Percentage provided all no cost consultations		37.8%	21.0%	36.08%	$\chi^2 = 17.81, p < .001$
FA in previous time period (11/12)		6.2%	51.8%	10.8%	χ <sup>2</sup> =315.08, <i>p</i> <.001
Average number o	f chronic problem consultation (SD)	.06 (.31)	.67 (1.2)	.17 (.51)	t(1589)=-15.06, <i>p</i> <.001

\*All measures are for the 12 months following the two consultations which define the treatment (unless indicated with a '11/12') or are derived from the survey

Logistic regression models (see Table 2) showed that small average costs do were not associated with increased risk of FA status relative to no cost. However, respondents with medium or large average out-of-pocket costs had significantly lower likelihood of subsequently being classified as a FA (see Model A). These results held after controlling for prior (lagged) FA status and prior (lagged) access to no cost consultations (Model B: lagged characteristics) and all of the current health and socioeconomic covariates (Model C).

A subgroup analysis was conducted that focused on the effects for those patients with greater capacity to pay for services. This also showed that those with large average payments (but not medium) for their first two consultations were at decreased likelihood of becoming a FA compared to those who incurred no out-of-pocket costs. Similarly, analysis of total number of GP consultations rather than FA status both as the outcome (assessed using a negative binomial regression model) and as a lagged predictor also showed the same pattern of results. As such medium or large payments significantly predicted an increase in the number of subsequent consultations. Additionally all reported results hold whether we consider a consistent period for all participants (rather than a moving one as reported) regardless of whether we use financial year (as reported) or calendar year, with the effects for calendar being slightly weaker due to larger confidence intervals.

Table 3 presents the estimates derived from the AIPW treatment effect model. The results showed that, after adjusting for selection into payment levels. There was no difference in the likelihood of being a FA for those who incurred low, medium or high levels of out-of-pocket costs relative to those with no costs. Pairwise comparisons between all payment levels indicated that the treatment model adequately balanced the covariates and, therefore, adjusted for the differences between groups that may reflect selection bias (see supplementary Table 7).

Regression models (Model A, B and C) were repeated predicting life satisfaction instead of FA status. When just considering the effect of payment (Model A) greater out-ofpocket costs was associated with *increased* life satisfaction (small  $\beta$ =1.80, CI.95=[.16,3.45], *p* =.031; medium  $\beta$ =1.76, CI.95=[.59,2.93], *p* =.003; large  $\beta$ =1.81 CI.95=[.74,2.88], *p* =.001]) however this effect was no longer significant after including prior (lagged) FA status and prior (lagged) access to no cost consultations (Model B, all *p*>.1) or other health and socioeconomic covariates (Model C, all *p*>.3). Counterfactual models with overall life satisfaction as the outcome showed that, after controlling for factors likely to select individuals into different payment (treatment) levels, out-of-pocket costs did not significantly predict subsequent life satisfaction. Additional and sensitivity analyses are presented in the supplemental. Table 2.

Odds ratios and 95% Confidence Intervals (CI) from series of logistic regression models examining the association of payment level with subsequent FA status, controlling for previous years status and a range of risk factor covariates.

	Model A		Model B			Model C			
	Initial			Including lagged			Including all covariates		
					eristics				
	Odds	Odds 95% CI Sig		Odds	lds 95% CI S	Sig	Odds	95% CI	Sig
	Ratios			Ratios			Ratio		
Average two payment (no cost)									
Small(Less than \$15)	.66	.36-1.20	.177	.54	.26-1.13	.102	.48	.22-1.24	.112
Medium (\$15 – \$30)	.39	.2562	<.001	.43	.2478	.006	.45	.2492	.024
Large (Over \$30)	.25	.1639	<.001	.32	.1857	<.001	.36	.1973	.004
Gender (female)	.88	.63-1.25	.491	.83	.57-1.21	.334	.66	.47-1.13	.098
FA at lag				8.37	5.58-12.56	<.001	8.14	5.14-12.89	<.001
No cost cons. at lag (none)									
Some				2.54	1.32-4.91	.005	3.01	1.27-6.05	.007
All				1.31	.62-2.79	.481	1.94	.72-4.18	.150
Diabetes							2.06	.98-3.56	.036
Asthma							1.53	.87-2.67	.095
Thyroid							1.24	.61-2.52	.551

Arthritis					.66	.40-1.07	.092
Heart					1.30	.667-2.54	.871
Cancer/Leukaemia					1.06	.54-2.07	.871
Epilepsy					7.62	2.14-27.08	.002
Stroke					1.66	.51-5.32	.395
High blood pressure					1.00	.53-1.89	.995
Other health problem					.42	.2280	.008
Pension					.46	.17-1.23	.145
Low income					1.32	.59-2.95	.501
Refused to disclose income					.91	.27-3.04	.878
Experienced financial problems					1.18	.51-2.76	.698
Employment status (employed)							
Unemployed					3.99	1.33-11.93	.013
Not in labour force					1.61	.89-2.89	.114
Not secondary education							
secondary					1.09	.58-2.06	.788
tertiary					.51	.2989	.018
SF rated health					.61	.21-1.76	.361
SF health limit moderate activities					1.02	.54-1.91	.957
SF health limits climbing stairs					1.08	.69-1.67	.740
SF health means accomplish less					.57	.26-1.28	.176
	I I	I I	I	I		I I	

SF health limits kind of activities			.49	.22-1.10	.084
SF pain interferes			1.07	.63-1.82	.808
Goldberg headaches or neck-aches			1.37	.85-2.22	.193
Goldberg trembling, tingling, dizzy spells etc.			1.25	.75-2.09	.390
Goldberg worried about health			.75	.43-1.30	.306
Goldberg lacking energy			1.12	.71-1.76	.637
Goldberg depression			1.02	.91-1.14	.756
Anxiety and dep medications			1.91	1.15-3.18	.013
Sleep medications			.73	.41-1.29	.277
Blood pressure medications			.81	.41-1.61	.555
Cholesterol medications			1.59	.97-2.60	.066
Pain medications			2.22	1.21-4.08	.101
Other medications			1.21	.78-1.89	.392

Cons.= consultations; values in brackets represent comparison category

Table 3.

Estimated proportion of FAs by payment category from augmented inverse propensity weighted estimator (AIPW) counterfactual models taking into account selection into payment categories and covariates reflecting current status

	Full Model with covariates <sup>a</sup>						
	estimated potential-	AIPW coefficient (v none)	Significance				
Payment	outcome means [95% CI]	[95% CI]					
None	.163 [.065260]						
Small	.157 [.094220]	006 [125113]	.920				
Medium	.124 [.090159]	038 [141064]	.461				
Large	.091 [.067115]	072 [172028]	.159				

<sup>a</sup> Lagged FA status, Lagged payment, Lagged GP attendance, wave 3 pension receipt, wave 3 physical and mental health, wave 3 medications in the selection model and covariates current physical and mental health, current medications and lagged payment, and lagged FA status in predictive model; [95% CI] = the 95% confidence interval

#### Discussion

This study explored the association between out-of-pocket costs and subsequent frequent GP attendance over a 12-month period. While the link between patient costs and attendance is a salient policy topic, there is currently not a strong evidence base on which to make policy decisions. The regression models show that the costs patients incurred for two GP consultations were associated with the likelihood of becoming a FA during the following 12 months. While there was no significant difference in the outcomes for those who had no out-of-pocket cost or low costs, the likelihood of becoming a FA was significantly reduced for patients charged a high (average) cost for their initial consultations. Fees around the median charged in the Canberra region did make some difference, but appeared to have a lesser effect than high costs, however regression models are confounded in observational data. Importantly, once we used the extensive data on previous health status, socioeconomic circumstances and prior GP use to model the factors leading to selection into payment levels the effect of payment level was no longer significant. While some evidence of an effect of

out-of-pocket costs remained, as there was a gradient in the expected direction, it was reduced and no longer significant. Hence it appears that those with a history of frequent attendance or those identified with greater need were selected by practitioners to receive low/no-cost consultations or alternatively that need and morbidity is a more important determinate of FA status than costs. Additionally increased patient costs were associated with greater life satisfaction which provides an indication that there is a degree of selection into out-of-pocket costs of individuals in better circumstances.

These results suggest out-of-pocket costs are a relatively minor determinant of frequent health service use. Consistent with previous research, our regression models found that FA status was associated a range of factors including diabetes, number of life events, low income, lower educational achievement, anxiety, depression, cholesterol, and pain medication use. Attendance is also influenced by other factors beyond those related to the individual patient and their circumstances. GP attendance rates can also reflect the nature of the health care system, the intersection between health prevention, primary and tertiary care, broader social and cultural factors that may differ within and between countries, patient to GP communication, and characteristics of the health professional.[47-49] Studies such as this one, are unable to explore all of these issues (in this case due to data limitations) but do make an important contribution through consideration of some components of the relationship.

### Limitations

The current study is, however, limited by an exclusive focus on mid-aged respondents, which may restrict generalizability. Further, generalizability may be limited as the sample is drawn exclusively from residents of the Canberra region in South-Eastern Australia: an area shown to have a relatively low GP consultations and a smaller proportion of consultations with no out-of-pocket costs compared to the Australian average (55.6% in 2015/16 vs 83.7% nationally).[5] Future research should seek to replicate this study in other areas of Australia and with a broader range of patient ages. Further, while our modelling of morbidity is a strength, providing a separate measure of health based need derived from a different data source to the exposure and outcome measures (i.e., linked administrative data) which reduces the likelihood of confounding due to measurement type, there is evidence that self-reported morbidity differs from that reported by a physician. While, studies have shown the predictive utility of self-report morbidity (e.g.[50]) it should be noted, that health condition and levels of educational attainment have been shown to impact the accuracy of

self-reported morbidity [51] and hence this may generate a different pattern of results. The counterfactual model is limited to the variables which were available to account for selection into treatment levels (out-of-pocket costs) and not all factors which may influence this selection were measures in the survey or available in the data (e.g., GP characteristics). However, the comprehensive range of relevant potential selection characteristics that we were able to model did equate the four (cost-level) groups and showed that out-of-pocket costs no longer significantly predicted FA status. It is important to recognise the limitations of this type of research and not perpetuate the myth that frequent attendance is entirely a patient's responsibility. Future research with access to GP level information are important to further address this. One potential reason why calendar year results were slightly weaker than financial year is the Medicare Safetynet which keeps a tally of out-of-pocket expenses of an individual and reimburses more to the patient when they reach a certain cap.[for more information see 8] As the safetynet resets at the beginning of the calendar year this may have resulted in weaker effects when the calendar year was used.

## Strengths

A unique contribution of this study lies in the linkage of rich representative and longitudinal survey data with the objectivity of administrative data. The use of observation data is also taken into account and the lack of random selection controlled for through the use of counterfactual analyses. Together our analyses explicitly model the temporal ordering of exposure and consequence, demonstrating robustness of results with many alternative model specifications, and draw on a range of covariates and methods to control for underlying predisposition and reverse causation. Thus, we provide comprehensive findings on the effects of out-of-pocket patient expenses on FA status, an area currently not well understood, but a pressing policy concern using rich data, removing biases which may influence 'selection' into different levels of health service cost.

# Implications

The findings provided some evidence that out-of-pocket expenses are associated with attendance behaviour in primary care, with those with high out-of-pocket GP costs less likely to be FAs than those with low or no out-of-pocket costs. However, we found that this difference between payment groups was largely explained by health and previous FA status, which were both considered markers of need for the patient. Further, our robust analyses that sought to control for the factors that may determine an individual's out-of-pocket costs (such

as their previous need and attendance history), left no significant difference in FA based on out-of-pocket costs. This aligns with previous research indicating that FA status reflects health and that an individual's FA status varies over time with variability in health and need. As such the use of co-payments is likely to have little to no effect as a major deterrent of frequent attendance or address issues of over use. In fact, based on previous research, there are likely to be adverse unintended consequences of co-payments on some patients, particularly those in the most vulnerable circumstances.[24, 25]

# Conclusion

These current results suggest that co-payments are unlikely to be overly influential on subsequent use of GP services as after controlling for predisposing differences and selection into payment levels, patients who initially paid any amount for GP consultations did not significantly differ in their likelihood of becoming a FA. While regression models did find an effect of larger payments (particularly decreasing the likelihood of becoming a FA with rising cost) such models are confounded. As such efforts to address frequent GP attendance without imposing high out-of-pocket costs and strain on patients who are less able to cope with such increased costs should consider other non-monetary options.

# References

[1] Dutton P. National Press Club Address. 2014.

[2] Pymont C, Butterworth P. Longitudinal cohort study describing persistent frequent attenders in Australian primary healthcare. BMJ Open 2015; 5.

[3] Heywood PL, Blackie GC, Cameron IH, Dowell AC. An assessment of the attributes of frequent attenders to general practice. Family Practice 1998; 48:895-8.

[4] Smits FT, Brouwer HJ, Zwinderman AH, Mohrs J, Smeets HM, Bosmans JE, Schene AH, van Weert HC, ter Riet G. Morbidity and doctor characteristics only partly explain the substantial healthcare expenditures of frequent attenders: a record linkage study between patient data and reimbursements data. BMC Family Practice 2013; 14.

[5] National Health Performance Authority. Healthy Communities: Frequent GP attenders and their use of health services in 2012-2013. 2015.

[6] Commonwealth of Australia. The Annual Medicare Statistics - Financial Year 2007-08 to 2013-14. In: The Department of Health, editor, 2014.

[7] De Abreu Lourenco R, Kenny P, Hass MR, Hall JP. Factors affecting general practitioner charges and Medicare bulk-billing: results of a survey of Australians. Medical Journal of Australia 2015; 202:87-91.

[8] Department of Human Services. Medicare Services. 2017.

[9] Commonwealth of Australia. Budget 2014-15 Health 2014.

[10] Commission of Audit. Towards Responsible Government. 2014.

[11] Del Mar C. Copayments for general practice visits. Medical Journal of Australia 2014; 200:367.

[12] Leraci S. Unnecessary questions. Medical Journal of Australia Insight, 2014.

[13] Commonwealth of Australia. Budget 2016-17. 2016.

[14] Fisher M, Baum F, Kay A, Firle S. Are changes in Australian national primary healthcare policy likely to promote or impede equity of access? A narrative review. Australian Journal of Primary Health 2017; 23:209–15.

[15] Savageau JA, McLoughlin M, Ursan A, Bai Y, Collins M, Cashman SB. Characteristics of Frequent Attenders At a Community Health Center. ScD J Am Board Fam Med 2006; 19:265-75.

[16] Gotsadze G, Bennett S, Ranson K, Gzirishvili D. Health care-Seeking behaviour and out-of-pocket payments in Tbilisi, Georgia. Health Policy and Planning 2005; 20:232-42.

[17] Tamblyn R, Laprise R, Hanley JA. Adverse events associated with prescription drug costsharing among poor and elderly persons. Journal of American Medical Association 2001; 285:421-9.

[18] Trivedi AN, Rakowski W, Ayanian JZ. Effect of cost sharing on screening mammography in medicare health plans. The New Engalnd Journal of Medicine 2008; 358:375-83.

[19] Nolan A, Smith S. The effect of differential eligibility for free GP services on GP utilisation in Ireland. Social Science & Medicine 2012; 74:1644-51.

[20] Gourzoulidis G, Kourlaba G, Stafylas P, Giamouzis G, Parissis J, Maniadakis N. Association between copayment, medication adherence and outcomes in the management of patients with diabetes and heart failure. Health Policy 2017; 121:363-77.

[21] Knott RJ, Petrie DJ, Heeley EL, Chalmers JP, Clarke PM. The effects of reduced copayments on discontinuation and adherence failure to statin medication in Australia. Health Policy 2015; 119:620-7.

[22] Raz I, Novack L, Yitshak-Sade M, Shahar Y, Wiznitzer A, Sergienko R, Warshawsky-Livne L. Effects of changes in copayment for obstetric emergency room visits on the utilization of obstetric emergency rooms. Health Policy 2015; 119:1358-65.

[23] Bisakha S, Blackburn J, Morrisey MA, Kilgore ML, Becker DJ, Caldwell C, Menachemi N. Did copayment changes reduce health service utilization amoung CHIP Enrollees? Evidence from Alabama. Health Service Research 2012; 47:1603-20.

[24] Aron-Dine A, Einav L, Finkelstein A. The RAND Health Insurance Experiment, Three Decades Later. Journal of Economic Perspectives 2013; 27:197-222.

[25] RAND Health. 40 Years of the RAND Health Insurance Experiment. 2017.

[26] Australian Bureau of Statistics. Health services: patient experences in Australia 2009. Cat. no. 4839.0.55.001. In: ABS, editor. Canberra, 2010.

[27] Schoen C, Osborn R, Squires D, Doty MM. Access, affordability, and insurance complexity are often worse in the United States compared to ten other countries. Health Affairs Web First. published online Nov. 14, 2013, 2013.

[28] Kiil A, Houlberg K. How does copayment for health care services affect demand, health and redistribution? A systematic review of the empirical evidence from 1990 to 2011. European Journal of Health Economics 2014; 15:813-28.

[29] Jones D, Loader N. Should patients pay to see the GP? British Medical Journal 2016; 352.
[30] Pymont C, Butterworth P. Changing circumstances drive changing attendance: A longitudinal cohort study of time varying predictors of frequent attendance in primary health care. Journal of

Psychosomatic Research 2015; 79:498-505.

[31] Scaife B, Gill P, Heywood PL, Neal RD. Socio-economic characteristics of adult frequent attenders in general practice: secondary analysis of data. Family Practice 2000; 17:298-304.

[32] Gili M, Luciano JV, Serrano MJ, Jimenes R, Bauza N, Roca M. Mental Disorders Among Frequent Attenders in primary Care. A comparison with routine attenders. Journal of Nervous and Mental Disease 2011; 199:744-9.

[33] Diaz E, Gimeno-Feliu L-A, Calderon-Larranaga A, Prados-Torrress A. Frequent attenders in general practice and immigrant status in Norway: A nationwide cross-section study. Scandinavian Journal of Primary Health Care 2014; 32:232-40.

[34] Taylor RE, Marshall T, Mann A, Goldberg DP. Insecure attachment and frequent attendance in primary care: a longitudinal cohort study of medically unexplained symptom presentations in ten UK general practices. Psychological Medicine 2012; 42:855-64.

[35] Vedsted P, Fink P, Olesen F, Munk-Jorgensen P. Psychological distress as a predictor of frequent attendance in family practice. Psychosomatics 2001; 42:416-22.

[36] Foster A, Jordan K, Croft P. Is frequent attendance in primary care disease-specific? . Family Practice 2006; 23:444-52.

[37] Bergh H, Baigi A, Marklund B. Consultations for injuries by frequent attenders are found to be medically appropriate from general practitioners' perspective. Scandinavian Journal of Public Health 2005; 33:228-32.

[38] Smits FT, Brouwer HJ, Ter Riet G, van Weert HC. Epidemiology of frequent attenders: a 3 – year historic cohort study comparing attendance, morbidity and prescriptions of one – year and persistent frequent attenders. BMC Public Health 2009; 9.

[39] Vedsted P, Fink P, Sorensen HT, Olesen F. Physical, mental and social factors associated with frequent attendance in Danish general practice. A population-based cross-sectional study. Social Science and Medicine 2004; 59:813-23.

[40] Anstey KJ, Christensen H, Butterworth P, Easteal S, Mackinnon A, Jacomb T, Maxwell K, Rodgers B, Windsor T, Cherbuin N, Jorm AF. Cohort profile: the PATH through life project. International journal of epidemiology 2012; 41:951-60.

[41] National Commission of Audit. Towards responsible government: phase one, February 2014, Recommendation 17. 2014.

[42] Ware JE, Kosinski M, Keller SD. A 12-item short-form health survey. Medical Care 1996; 34:220-33.

[43] Ruben DB. Estimating causal effects of treatments in randomised and nonrandomised studies. Journal of Educational Psychology 1974; 66:688-701.

[44] Molitoris J. The Influence of Grandparental Child Care on Continued Childbearing: Evidence from the Health and Retirement Study. 2016.

[45] Scott JW, Havens JM, Wolf LL, Zogg CK, Rose JA, Salim A, Haider AH. Insurance status is associated with complex presentation amoung emergency general surgery patients. Surgery 2017; 161:320-8.

[46] Glynn AN, Quinn KM. An introduction to the Augmented Inverse Propesnisty Weighted Estimator. Political Analysis 2010; 18:36-56.

[47] Neal RD, Heywood PL, Morley S, Clayden AD, Dowell AC. Frequency of patients' consulting in general practice and workload generated by frequent attenders: comparisons between practices. British Journal of General Practice 1998; 48:895-8.

[48] Salmon P, Peters S, Clifford R, Iredale W, Gask L, Rogers A, Dworick C, Hughes J, R M. Contrasts between general practitioners who consent and decline training to improve the management of medically unexplained symptoms. Journal of General Internal Medicine 2007; 22:565-71.

[49] Reid S, Whooley D, Crayford T, M H. Medically unexplained symptoms; GP's attitudes towards their cause and management. Family Practice 2001; 18:519-23.

[50] Martin L.M., Leff M, Calonge N, Garrett C, D.E. N. Validation of self-reported chronic conditions and health services in a managed care population. American Journal of Preventative Medicine 2000; 18:215-8.

[51] Ferraro K.F., Y. S. Physician-evaluated and self-reported morbidity for predicting disability. American Journal of Public Health 2000; 22:565-71.