

# The cooler the better? Indirect effect of spring–summer temperature on fecundity in a capital breeder

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**Abstract.** Female reproductive performance is a central component of ungulate population dynamics, and it can be influenced by individual, social, and environmental factors. Researchers have often assumed direct effects of different predictors on reproduction, yet more complex relationships should be considered when investigating temporal variations in life-history traits within a broader eco-evolutionary context. In this study, we explored direct effects of individual, social, and environmental predictors on female reproductive performance and investigated potential causal chains among variables. We analyzed the variation in fecundity, measured as the probability of being pregnant, in 215 adult female deer *Cervus elaphus* culled on the Italian Alps, with respect to age, body mass, kidney fat, jaw length, lactation status, population size, temperature and precipitation in spring–summer, temperature and snow depth in winter, and the delayed effect of spring–summer temperature. We used random forest and logistic regression models to select variables whose direct effects best explained variation in fecundity. Path analysis was used to test for alternative hypotheses of direct/indirect effects between pre-selected weather (spring–summer temperature) and individual (age, kidney fat index [KFI]) predictors. The most important direct predictors of fecundity were age, kidney fat, and the interaction between kidney fat and spring–summer temperature. Path analysis supported the hypothesis that higher spring–summer temperature had negative, indirect effects on the probability of being pregnant, mediated by decreasing values of KFI. Our study revealed some complex, cause–effect relationships between weather stochasticity, body condition, and reproduction, possibly suggesting a conditional trade-off between opportunity for reproduction and survival, and emphasizing how environmental variations and individual characteristics may interact to shape life-history traits in ungulate populations.

**Key words:** cause–effect relationships; *Cervus*; climate; deer; demography; fecundity; life-history; path analysis; ungulates.

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## INTRODUCTION

Reproduction is a central component of animal population dynamics. Several individual, social, and environmental factors such as age, body

condition, population density, and climatic conditions are known to affect reproductive performance of females (Gaillard et al. 2000). The importance of each factor may vary among species, owing to different life-history strategies

(Coulson et al. 2000, Hamel et al. 2010), and within species, depending on different habitat characteristics (Balbontín and Ferrer 2008).

In many species of birds and mammals, female reproduction is strongly related to age in a non-linear fashion, and different reproductive parameters generally increase from sexual maturity to prime age and then decline later in life (Nussey et al. 2008). The age effect is intimately associated with the long-term variations in body condition that occur over an individual's lifetime, which influence its reproductive performance (Nussey et al. 2011, Flajšman et al. 2017). Body condition may also undergo great short-term variations, in response to internal and/or external pressures. For example, reproductive events may impose high energetic costs on females and negatively affect individual condition, thus lowering the probability of reproducing in the following breeding season. This pattern has been observed in mammals and in other *taxa* (Gustafsson and Sutherland 1988, Yurewicz and Wilbur 2004, Hamel et al. 2010), especially in food-limited populations (Clutton-Brock et al. 1989). Density-dependent availability of food supply (Bonenfant et al. 2009) or climate effects may in fact trigger short-term changes in body condition and in turn variation in reproductive performance. For example, several studies showed evidence for negative responses of reproductive traits to increasing ambient temperature in different *taxa*, possibly mediated by variations in individual conditions (Grazer and Martin 2012). Individual, social and environmental variables may thus influence female reproductive traits in a complex manner, operating through pathways that include both direct and indirect relationships between different factors, possibly reflecting adaptive responses to optimize the trade-offs between reproduction and survival in different environmental conditions (Sand 1996).

A deeper understanding of the adaptive mechanisms underlying variations in female reproductive traits may benefit from explicitly assuming non-independence among predictive variables. Recent studies have supported the occurrence of more complex causal relationships among potential drivers of reproduction: In Norwegian red deer *Cervus elaphus*, for example, body condition during summer is indirectly affected by climate through plant phenology, as higher spring temperatures accelerate plant development

(Mysterud et al. 2008). This in turn suggests that climatic variables may exert an important indirect effect on deer reproductive performance. Indirect consequences of rising temperatures on life-history traits are increasingly suggested to occur also in Alpine ungulates. In highly seasonal, energy-limited mountain temperate environments, fat reserves are mainly deposited before autumn: Warmer spring–summer periods may cause a reduction in food acquisition through several mechanisms (Pettorelli et al. 2007, Mason et al. 2014, 2017) and thus negatively impact on body condition and, possibly, reproduction in capital breeders (Rughetti and Festa-Bianchet 2012). Different hypotheses may be put forward to explain this pattern. A first hypothesis suggests that warm temperatures in spring–summer may accelerate plant development and reduce the availability of high-quality food resources over the summer (Pettorelli et al. 2007). An alternative hypothesis suggests that in mountain-dwelling ungulates such as the chamois *Rupicapra rupicapra*, body condition in autumn is not limited by summer resource availability: Rather, the high spring–summer temperatures may reduce the time spent foraging before autumn thus limiting the ability of individuals to acquire resources (Mason et al. 2014). More recently, for the Alpine ibex *Capra ibex* it has been suggested that warmer temperatures in summer would force animals to thermoregulate by using less productive areas at higher elevations, and thus consume lower quality forage, without compensating their foraging effort (Mason et al. 2017). Despite their potential demographic and evolutionary significance, however, the indirect effects between individual, social, and climatic variables on the reproductive performance of female ungulates in mountainous landscapes have received comparatively little attention.

Owing to its wide geographical distribution, the direct role of individual, social, and environmental factors on female reproductive performance has been largely investigated in red deer. The importance of age on hind reproduction, for example, has been supported in several populations (e.g., Albon et al. 1986, Bertouille and de Crombrughe 2002), while the role of other individual drivers appears to be less consistent, varying with latitude and habitats. Body mass and fat reserves had positive effects on hind pregnancy probability in Scottish populations (Albon et al. 1986), but no

relationship was found in central Europe (Borowik et al. 2016). Body size also showed contrasting effects on pregnancy probability in northern and central Europe (Mitchell and Brown 1974, Albon et al. 1986, Bertouille and de Crombrughe 2002). The negative effect of lactation on the probability of being pregnant in the following reproductive season is strong in food-limited environments (Clutton-Brock et al. 1989), but little information is available in rich environments. Similarly, the direct effects of social and environmental variables on hind reproduction show site-dependent variations. Negative density-dependent relationships between pregnancy probability and body mass were found in Scotland (Albon et al. 1983), while no density-dependent effects were found in central Europe (Bonenfant et al. 2002, Borowik et al. 2016). In Scottish deer, summer precipitation, winter temperature, and snow negatively influenced female fecundity (Albon and Clutton-Brock 1988, Langvatn et al. 1996), whereas adult pregnancy ratio in Rocky Mountain elk increased following summers with high precipitation (Proffitt et al. 2014). More recently, no significant effects of winter and summer temperature on deer fecundity were found in Poland (Borowik et al. 2016). Overall, the direct effects of individual, social, and climatic variables on hind reproduction do not show congruent patterns over large geographic scale, and little is known about their potential indirect effects. A deeper understanding of the adaptive responses to environmental conditions should thus account for more complex interactions (cf. Stopher et al. 2014).

In this paper, we first investigate the direct effect of individual, social, and environmental factors on reproductive performance in an Alpine population of red deer, as limited information is available in this environment. We then explore potential causal pathways affecting female fecundity, accounting for the occurrence of direct and indirect relationships between variables (sensu Shipley 2016). In particular, we test hypotheses about the potential effect of weather conditions on fecundity, already suggested for other Alpine ungulates, which indicate that increasing spring–summer temperatures could negatively and indirectly affect reproductive performance through a decline in body condition (Pettorelli et al. 2007, Rugghetti and Festa-Bianchet 2012, Mason et al. 2014).

## METHODS

### *Study area and population*

The study site Valfurva lies in the northwestern part of the Stelvio National Park, within the Province of Sondrio, Central Italian Alps (10°25' N, 46°27' E). Valfurva is the wintering site of a large population management unit for red deer, defined by tracking individuals collared with GPS (global positioning system) devices (cf. Corlatti et al. 2016), and extends over 4975 ha between 1200 and 2400 m a.s.l. (Fig. 1). About 73% of its surface is dominated by spruce *Picea abies*, larch *Larix decidua*, and Stone pine *Pinus cembra* forests, while the remaining 27% consists of open areas with mesic meadows at lower elevations, and Alpine grasslands of *Carex* spp., *Festuca halleri*, and *Sesleria coerulea* above the treeline. The climate is alpine continental, with mean temperatures between 15.7°C in July and –2.8°C in January and yearly precipitation of about 765 mm. Between 2011 and 2015, the winter density of the red deer population in Valfurva was about 27.4 ind./km<sup>2</sup> (±2.5 SD; Corlatti et al. 2016). The large increase in deer density that occurred in the Park over the last two decades, owing to the good environmental conditions within the Park and possibly the absence of hunting pressure inside the protected area, severely impacted on forest regeneration, on agricultural activities, and on the ecosystem biodiversity. In 2011, the National Park Agency therefore started a culling program aimed at reducing population density: Given the initial high density of deer (about 31 ind./km<sup>2</sup> in winter), the effects of culling were apparent only since 2015 (cf. Corlatti et al. 2016). In 4 yr, from 2011 to 2016, within the study site a total of 358 female deer (0.5 yr:  $n = 81$ ; 1.5 yr:  $n = 39$ ; 2+ yr:  $n = 238$ ) were culled by professional hunters, under the supervision of the Park Authority. Culling was conducted between late October and early February of each year (no cull occurred in 2013/14). No restrictions were imposed on adult females in terms of age or lactation status; therefore, the sample of females of 2+ yr of age likely reflected the structure of their population.

### *Data collection*

In the literature on ungulate ecology, different terms such as “pregnancy probability,” “fecundity,” and “fertility” are often used interchangeably to indicate reproductive performance. For

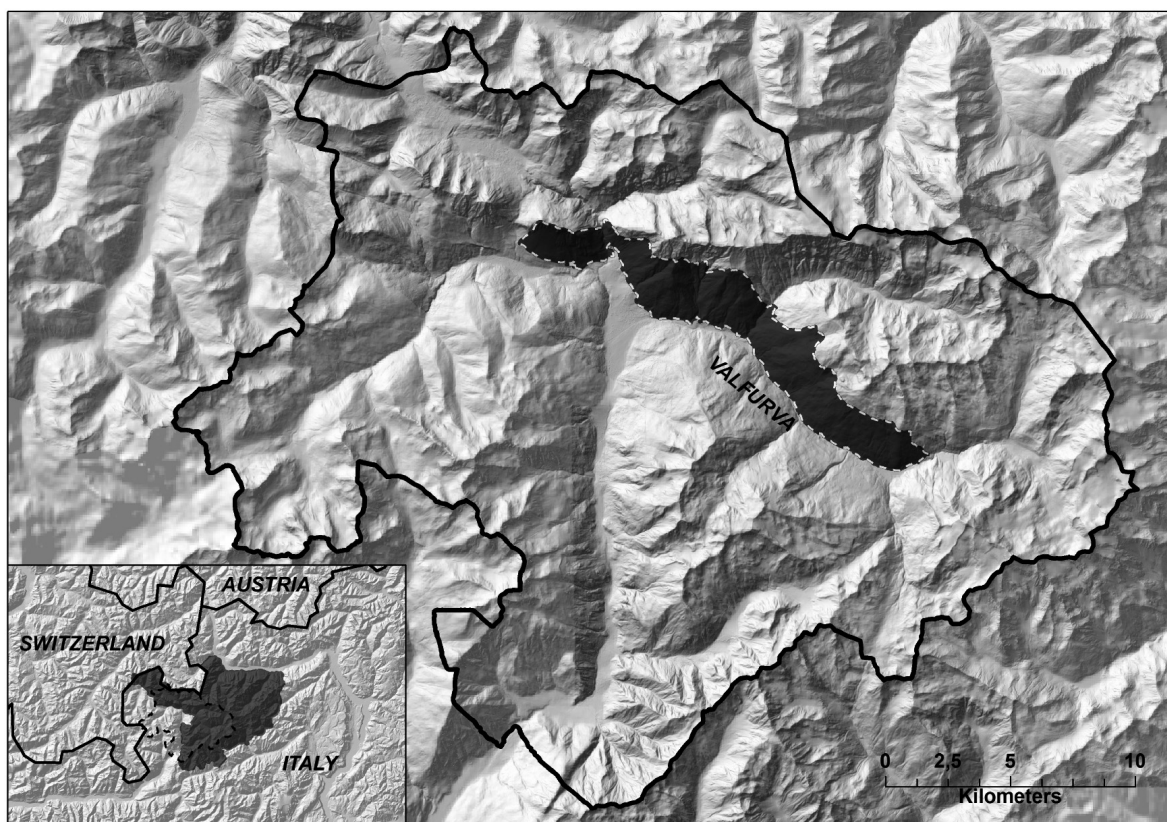


Fig. 1. Location of the study site Valfurva (gray-shaded area on the right), within the deer management unit (dashed line on the left, solid line on the right), in the Stelvio National Park (gray-shaded area on the left), Central Italian Alps.

the sake of clarity, hereafter we refer to fecundity as to the probability of being pregnant. To investigate deer fecundity, we focused on adult females (2+ yr of age) only, as very few yearlings breed. All females were brought to a control center within 2 h from culling. For some individuals, it was not possible to collect all the parameters needed for the analysis, and our sample size reduced to  $n = 215$  adults. Pregnancy status was investigated through the presence/absence of *corpora lutea* (i.e., endocrine structures that develop from ovarian follicles during the luteal phase of the oestrous cycle) by dissecting ovaries: For nearly all individuals, pregnancy was confirmed by the presence of fetuses, which always occurred in conjunction with *corpora lutea*, thus supporting the suitability of this last parameter to assess reproductive status. Potential drivers of hind fecundity included individual and external (social and environmental) variables. Individual

variables comprised age (estimated by counting the tooth cementum annuli), lactation status (evaluated through the presence of milk, or cutting into the udder), jaw length (in mm, measured by means of an electronic calliper), dressed body mass (in kg), and kidney fat index (KFI, measured following Riney 1955). We considered KFI as the most suitable proxy for individual condition in red deer (Riney 1955), as it was the most direct proxy of body fat reserves available to us. External variables included population size in the previous spring (based on mark–resight estimates: Corlatti et al. 2016) and meteorological variables. On the Alps, deer fertilization occurs at the beginning of October, and fat reserves are mainly deposited in spring–summer: Because in this environment warmer temperatures are linked to anticipated vegetation growth (Pettorelli et al. 2007), to predict fecundity in each year  $t$  we used the mean daily temperature (in °C) during the entire

vegetative season, that is, from April, when vegetation growth begins, to September, when the seasonal course of shoot biomass drops, at year  $t$ . Additionally, we also considered the effects of other climatic variables on fecundity: the cumulative precipitation (in mm) in spring–summer (April–September) at year  $t$ ; the mean daily temperature (in °C) and the mean snow depth (in cm) during winter (January–March) at year  $t$ ; the mean daily spring–summer temperature (in °C) at year  $t - 1$ . Data were retrieved from a meteorological station within the study area (cumulative precipitation in spring–summer at year  $t - 1$  was not available for the year 2010; hence, this variable was not included in the analysis).

### Statistical analysis

To investigate the direct relationships between female fecundity and individual, social, or environmental variables, we used both non-parametric and parametric approaches, as different selection criteria can give different results even with the same dataset (Gotelli and Ellison 2013). Prior to analysis, all continuous explanatory variables were standardized by subtracting each sample's mean and then dividing by the sample's standard deviation, to return comparable coefficients and reduce issue of collinearity in presence of interaction terms. To investigate direct and indirect causal relationships among selected variables (age, KFI, spring–summer temperature), we used path analysis on unstandardized data (Wright 1934).

*Direct relationships: non-parametric approach.*—We started exploring the relevance of each variable, with respect to fecundity, with a wrapper algorithm based on a random forest classification method (Breiman 2001), using 99 random forest runs. The algorithm returns a numerical estimate of the singular variable importance, measured as the loss of accuracy of classification caused by the random permutation of variable values across observations (Kursa and Rudnicki 2010). This non-parametric approach allows to identify ecologically important predictors of fecundity and offers some advantages over traditional variable selection procedures, as it is more robust to collinearity (Cutler et al. 2007). The interpretation of high-order interactions in random forest-based approaches, however, is not immediate, and the method is not suitable for hypothesis testing or for identifying ecologically important

subsets of variables in the way model selection does (Cutler et al. 2007).

*Direct relationships: parametric approach.*—We proceeded exploring direct effects of additive and interactive combinations of individual and environmental variables on deer fecundity, by applying a parametric, information–theoretic (IT) model selection approach (Burnham and Anderson 2002) on a set of biologically plausible competing models explicitly tied to underlying mechanisms linking climate and individual/social variables to fecundity. Specifically, binomial linear regression models were fitted with logit link function:

$$\begin{aligned} \text{Fecundity}_i &\sim B(\pi_i, 1) \\ E(\text{Fecundity}_i) &= \pi_i \quad \text{and} \\ \text{var}(\text{Fecundity}_i) &= \pi_i \times (1 - \pi_i) \\ \pi_i &= \frac{e^{\eta_i}}{1 + e^{\eta_i}} \quad \text{where } \eta_i = X1_i \times X2_i + \text{age}_i + \text{age}_i^2 \end{aligned} \quad (1)$$

Fecundity $_i$  represents the pregnancy status (0/1) of individual  $i$  at time  $t$ . Individual age $_i$  at time  $t$  was fitted as a quadratic term (to account for non-linear effect on fecundity) in each model, as preliminary analyses showed that it consistently improved models' fit. The fixed variable  $X1_i$  was represented by either individual body mass $_i$ , KFI $_i$ , jaw length $_i$ , lactation status $_i$ , or by population size $_i$ , all referred to year  $t$ , while the variable  $X2_i$  was represented by mean spring–summer temperature $_i$  at year  $t$ , cumulative spring–summer precipitation $_i$  at year  $t$ , mean winter temperature $_i$  at year  $t$ , average snow cover $_i$  in winter  $t$ , or by mean spring–summer temperature $_i$  at year  $t - 1$ . When assessing the relative importance of variables in the IT framework, it is important to achieve a balance in the number of models that contain each explanatory variable (Burnham and Anderson 2002). Therefore, to explore the effect of every biologically plausible combination of predictors  $X1_i$  and  $X2_i$  a set of 24 models was generated (Appendix S1: Table S1). To account for temporal variation in the value of body mass and KFI, all the models that included these two variables also included individual shooting date $_i$  as a covariate, that is the number of days elapsed from October 30 of each year (i.e., the first day of shooting) to individual culling. A preliminary analysis using AICc showed that fixed-effect

models (Eq. 1) consistently outperformed mixed-effect models fitted with an observation-level random intercept (to account for unexplained heterogeneity among subjects: Harrison 2015). Prior to analysis, a matrix based on Pearson's correlation coefficient ( $r_p$ ) on standardized continuous variables was built to identify potential issues of collinearity. In case of severe collinearity ( $r_p > 0.7$ , Dormann et al. 2013), variables were never fitted in the same model, to avoid bias in parameter estimation. Model parameters were estimated using maximum likelihood. All fitted models ( $n = 24$ ) were subsequently ranked based on their AICc values and retained in the final candidate set if they had  $\Delta\text{AICc} \leq 2$  (Burnham and Anderson 2002). The explained variance of models in the candidate set was measured by studying the Nagelkerke's pseudo- $R^2$ , and the goodness of fit was assessed by calculating the  $P$ -value associated with the Hosmer and Lemeshow test statistic (Hosmer and Lemeshow 2000). Additionally, the predictive accuracy of candidate models was measured using the area under the receiver operating characteristic curve (AUC) which, in this study, refers to the ability of a given model to discriminate between pregnant and non-pregnant females. Bootstrapping is arguably the best alternative for obtaining predictive ability measures, as it provides stable estimates with low bias, especially with small sample sizes (Steyerberg et al. 2001). To estimate the AUC for each model in the candidate set, we thus performed an internal validation using 1000 bootstrap samples following Harrel et al. (1996).

*Causal relationships: path analysis.*—Path analysis (Wright 1934) was used to investigate potential

causal relationships between selected climate (spring–summer temperature) and individual (age, KFI) variables and their direct/indirect effects on fecundity. Path analysis requires the creation of diagrams that illustrate the hypothesized relationships among the selected variables (Gotelli and Ellison 2013); thus, we first built three directed acyclic graphs to represent alternative causal models (Fig. 2), including direct and indirect relationships (sensu Shipley 2016) reflecting biologically plausible hypotheses of causal chains with respect to red deer life history. In Model a, temperature had an indirect effect mediated by KFI on fecundity; in Model b, both temperature and KFI directly affected fecundity; in Model c, temperature had both a direct effect and an indirect effect mediated by KFI on fecundity. In all models, age was assumed to have a direct non-linear effect on pregnancy status and on KFI. To verify the consistency of the correlational structures between the hypothesized models and the sample data, we checked the values of the chi-square goodness-of-fit test ( $P$ -values  $> 0.05$  indicate adequate fit) and we calculated the value of the root-mean-square error of approximation (RMSEA: Steiger and Lind 1980) which, compared to other indexes of fit, offers the possibility to calculate confidence intervals (RMSEA values  $\leq 0.6$  and  $P$ -values  $> 0.05$  indicate adequate fit: Tomer and Pugsek 2003). Since we used the diagonally weighted least-squares estimator to fit models with binary response (Rosseeel 2012), AIC values were not available to compare models' fit. We therefore used the scaled chi-square difference test using the Satorra-Bentler method (Satorra and Bentler 2001) to compare competitive structures.

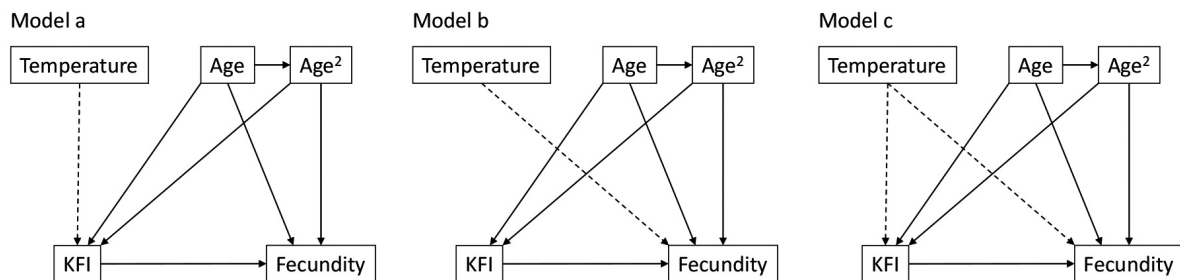


Fig. 2. Graphical representation of the three models built to test for causal relationships among age, temperature, and fecundity in adult female red deer culled within the Stelvio National Park between 2011 and 2016. Dashed arrows indicate the causal links that change among the models.

Table 1. Summary of the final results derived from random forest classification to explain the variation in fecundity in adult female red deer culled within the Stelvio National Park between 2011 and 2016.

Attribute	Mean Z	Median Z	Min. Z	Max. Z	Hits	Decision
Lactation status	1.111	1.096	-2.389	4.176	0.152	Rejected
Body mass	0.918	0.613	-2.188	4.042	0.091	Rejected
<b>KFI</b>	<b>3.491</b>	<b>3.433</b>	<b>-0.646</b>	<b>8.089</b>	<b>0.717</b>	<b>Confirmed</b>
<b>Age</b>	<b>3.586</b>	<b>3.529</b>	<b>-0.274</b>	<b>10.650</b>	<b>0.707</b>	<b>Confirmed</b>
Jaw length	2.124	2.049	-1.692	9.242	0.384	Rejected
Population size	-0.280	-0.250	-2.535	1.797	0.000	Rejected
Spring–summer temperature at $t$	0.126	0.207	-1.249	0.948	0.000	Rejected
Spring–summer precipitation at $t$	-0.376	-0.274	-1.964	1.340	0.000	Rejected
Winter temperature at $t$	-0.867	-0.699	-2.744	0.317	0.000	Rejected
Winter snow at $t$	0.199	0.068	-1.156	1.697	0.000	Rejected
Spring–summer temperature at $t - 1$	-0.195	-0.281	-2.026	2.637	0.000	Rejected

Notes: The table shows the Z score statistics (Mean, Median, Min., and Max.) for each attribute. The Hits column refers to the fraction of random forest runs in which the corresponding attribute was more important than the most important shadow attribute. The final column (Decision) reports whether the attribute was eventually confirmed or rejected (confirmed attributes in bold).

All analyses were conducted using RStudio 1.1.442 (RStudio Team 2016) in R 3.4.4 (R Core Team 2018). We used the Boruta package (Kursa and Rudnicki 2010) to investigate variable importance based on random forest classification. The glm function was used to fit logistic regression models; model selection was performed using the MuMIn package (Bartoń 2015), while the bootstrap internal validation was conducted with the rms package (Harrell 2017). Causal relationships using path analysis were tested with the package lavaan (Rosseel 2012).

## RESULTS

### Direct relationships

The all-relevant feature selection based on the wrapper algorithm around random forest classification showed that age and KFI were the most relevant variables directly related to fecundity in our study population (Table 1; Appendix S2: Fig. S1).

The correlation matrix did not suggest severe issues of collinearity, except between temperature and population size ( $r_p = -0.88$ ) or winter snow ( $r_p = -0.88$ ): These two pairs of variables were thus never included in the same model, to avoid bias in parameter estimation. The model selection procedure retained only one model as candidate to explain variation in fecundity of adult hinds ( $\Delta AICc \leq 2$ , Table 2; see also Appendix S1: Table S2). This model assumes that the effect of spring–summer temperature at year  $t$  on fecundity variation was moderated/mediated by body condition and varied additively with age<sup>2</sup>. The

model fitted the data satisfactorily (Hosmer and Lemeshow GOF test:  $\chi^2 = 6.275$ ,  $df = 8$ ,  $P$ -value = 0.616), explained 16% of the variance (Nagelkerke's pseudo- $R^2$ ), and had acceptable discrimination ability (AUC = 0.71: Hosmer and Lemeshow 2000). The parameter estimates showed that age had a strong non-linear effect on hind fecundity, with relatively lower values for the extreme classes and higher values for the intermediate classes, whereas KFI had a positive effect on pregnancy status (Table 3; Appendix S2: Fig. S2). Spring–summer temperature at year  $t$

Table 2. Selection of models with  $\Delta AICc \leq 4$ , fitted to explain variation in fecundity in adult female red deer culled within the Stelvio National Park between 2011 and 2016.

Model	df	$\Delta AICc$	Weight
<b>Fecundity ~ KFI × Spring–summer temperature at <math>t</math> + Age + Age<sup>2</sup> + Shooting date</b>	7	<b>0.00</b>	<b>0.537</b>
Fecundity ~ KFI × Winter snow at $t$ + Age + Age <sup>2</sup> + Shooting date	7	2.47	0.156
Fecundity ~ KFI × Spring–summer precipitation at $t$ + Age + Age <sup>2</sup> + Shooting date	7	2.89	0.127
Fecundity ~ Body mass × Spring–summer temperature at $t$ + Age + Age <sup>2</sup> + Shooting date	7	3.24	0.106
Fecundity ~ Body mass × Winter snow at $t$ + Age + Age <sup>2</sup> + Shooting date	7	3.95	0.074

Notes: The table reports model formula, degrees of freedom (df), differences in Akaike's information criterion corrected for small sample size ( $\Delta AICc$ ) between each model and the model with the lowest AICc, Akaike's weights (weight). For each model, + and × indicate additive and interactive effects, respectively. Selected models ( $\Delta AICc \leq 2$ ) in bold.

Table 3. Parameter estimates from the model with  $\Delta\text{AICc} \leq 2$ , retained to explain the variation in fecundity in adult female red deer culled within the Stelvio National Park between 2011 and 2016.

Parameters	Estimate	Standard error	95% confidence interval	
			LCL	UCL
<b>(Intercept)</b>	<b>2.488</b>	<b>0.325</b>	<b>1.851</b>	<b>3.125</b>
<b>KFI</b>	<b>0.577</b>	<b>0.268</b>	<b>0.052</b>	<b>1.102</b>
Spring–Summer Temperature at $t$	−0.320	0.249	−0.808	0.168
Age	0.131	0.230	−0.320	0.582
<b>Age<sup>2</sup></b>	<b>−0.415</b>	<b>0.154</b>	<b>−0.717</b>	<b>−0.113</b>
<b>KFI: Spring–Summer Temperature at <math>t</math></b>	<b>0.470</b>	<b>0.236</b>	<b>0.007</b>	<b>0.933</b>

Notes: For each predictor, the table reports the standardized values of beta estimate with the corresponding standard error and 95% confidence interval (lower confidence limit—LCL, upper confidence limit—UCL). Relevant effects (i.e., with 95% CI that do not include zero) in bold.

also played a role in explaining the variance in fecundity, in relation to KFI: With decreasing values of KFI, the probability of being pregnant reduced strongly only with increasing air temperature (Table 3, Fig. 3). As the 2AICc cutoff might arguably be considered overly conservative, in the supporting information we show that consistent results are reached when averaging models using less restrictive cutoffs ( $\Delta\text{AICc} \leq 4$ ; see Appendix S1: Table S3).

### Causal relationships

Path analysis, performed to discriminate alternative causal relationships among the selected meteorological and individual variables, showed that two of the structures hypothesized in our directed acyclic graphs (Model a and Model c) satisfactorily fitted the correlational structure of the sample data ( $\chi^2$  value  $>0.05$ , RMSEA value  $\leq 0.6$ , Table 4). The Satorra-Bentler test did not show a significant difference between the two models ( $\chi^2$  difference = 0.600,  $P$ -value = 0.439). Following the principle of parsimony, this result suggests that the additional path assumed in Model c (Fig. 2) may be considered uninformative (as confirmed by the estimates reported in Appendix S2: Fig. S3). Model a was thus selected as the best model, supporting the hypothesis that temperature had an indirect effect on fecundity, mediated by KFI: The standardized path coefficients (Fig. 4) show that increasing spring–summer temperatures

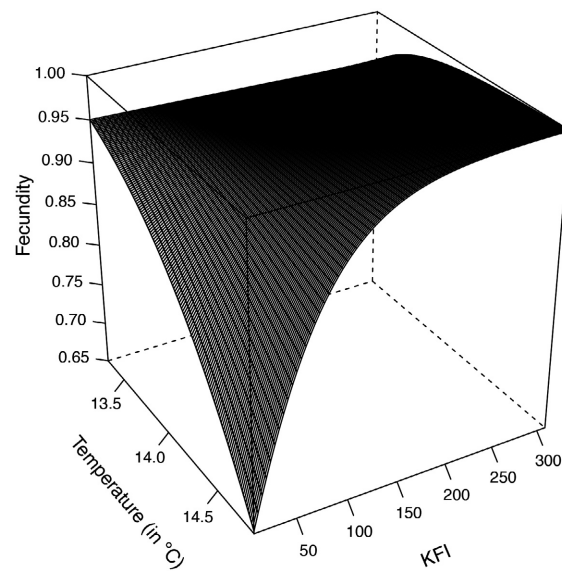


Fig. 3. Relationship between fecundity and kidney fat index (KFI) for increasing values of spring–summer temperature in the female red deer population within the Stelvio National Park between 2011 and 2016, as predicted by the best logistic regression model. With decreasing values of KFI, the probability of being pregnant reduces strongly only when air temperature is increasingly high.

negatively impacted on KFI and, in turn, on pregnancy status (indirect effect: estimate =  $-0.054$ ,  $P$ -value = 0.020). The direct non-linear effect of age on fecundity, however, was stronger than the indirect effect of temperature (Fig. 4).

## DISCUSSION

The parametric and non-parametric variable selection procedures consistently supported the

Table 4. Path models fitted to explain the variation in fecundity in adult female red deer culled within the Stelvio National Park between 2011 and 2016.

Model	$\chi^2$	df	$\chi^2$ $P$ -value	RMSEA	RMSEA $P$ -value
Model a	0.776	2	0.678	0.000	0.793
Model b	21.185	2	0.000	0.212	0.000
Model c	0.102	1	0.750	0.000	0.807

Note: For each model, the table reports values of the chi-square goodness-of-fit test ( $\chi^2$ ), degrees of freedom (df),  $P$ -values for the chi-square test ( $\chi^2$   $P$ -value), RMSEA values (RMSEA), and  $P$ -values for RMSEA (RMSEA  $P$ -value).



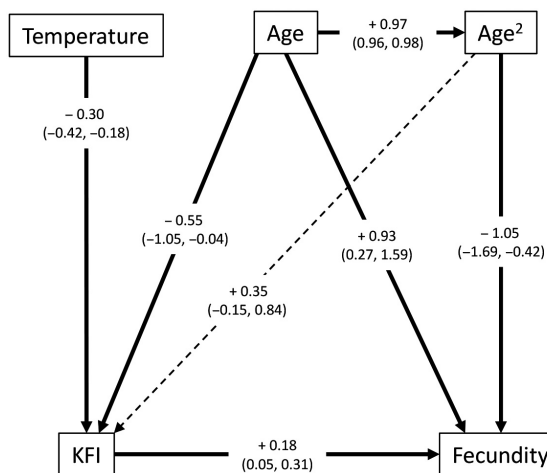


Fig. 4. Pathways of the selected path model (Model a) with standardized coefficients and associated 95% confidence interval. Thicker arrows highlight significant relationships; dashed arrows indicate non-significant relationships.

direct role of age and KFI to explain fecundity variation within our study population. Age had a strong, non-linear effect on fecundity probability, with relatively low values for the extreme age classes and high values for the intermediate age classes, whereas increasing values of KFI had positive effects. In both approaches, the direct, independent effects of climatic variables were negligible. Spring–summer temperature was not retained by the non-parametric analysis since our approach only evaluates the importance of individual effects, but its influence became significant in parametric models when fitted as an interaction with KFI. The result of this interaction suggests that a mediation effect (sensu Baron and Kenny 1986) might occur between these variables: Indeed, air temperature and KFI were assumed to be causally related, as warm spring–summer temperatures might lead to relatively poorer body conditions because of anticipated plant development (Mysterud et al. 2008). The path analysis went beyond the simplistic regression approach, allowing to confirm that increasing spring–summer temperature negatively, indirectly affected hind fecundity through a negative effect on KFI values, thus allowing a deeper understanding of the adaptive response of females to different environmental conditions.

The non-linear relationship between age and adult fecundity has been supported in several

deer populations (e.g., Albon et al. 1986, Stewart et al. 2005, Morano et al. 2013) and in other ungulates (e.g., mountain goats *Oreamnos americanus*, Côté and Festa-Bianchet 2001, Soay sheep *Ovis aries*, Tavecchia et al. 2005). Younger and older individuals are less likely to become pregnant than prime-aged individuals, suggesting an age-dependent cost of investment in reproduction. While the relatively low pregnancy rates in young age classes may be explained by the necessity to reach a threshold body mass to attain primiparity (Gaillard et al. 2000), there is still limited understanding of the causes of decline in fecundity in old females. In red deer, however, the disposable soma and antagonistic pleiotropy theories of senescence have received some support (Nussey et al. 2006). The fecundity probability in relation to age in adult females shows contrasting patterns in different areas: While in our study population we found a clear non-linear relationship between age and fecundity, with relatively low values before 3 yr and after 15 yr of age (cf. Appendix S2: Fig. S2), populations in central Europe showed high values of fecundity already at 2 yr of age (Bertouille and de Crombrugghe 2002, Borowik et al. 2016). Differences in age-dependent fecundity may be explained by variations in local conditions: High population density, for example, may lower the proportion of pregnant young adults (Stewart et al. 2005). The much lower deer density in the study sites of Borowik et al. (2016) compared to our study site (5.3 deer/km<sup>2</sup> vs. 27.4 deer/km<sup>2</sup>, respectively) may explain the different patterns of fecundity in young adult females between the two populations. Kidney fat exerted a positive effect on hind fecundity (cf. Albon et al. 1986), supporting the hypothesis that adult females in good body conditions have higher probability of being pregnant (Gaillard et al. 2000). Similar results were obtained by Morano et al. (2013). The limitations in the use of KFI as a ratio index have been discussed by Serrano et al. (2008), who suggested the use of residuals of the linear regression  $\log(\text{kidney fat}) \sim \log(\text{kidney weight})$  instead, as they have the advantage of being size independent. Other works, however, pointed out that “size independence does not necessarily mean that residual indices predict body fat content better than ratio indices” (Labocha et al. 2014). Given this uncertainty, in this study we

reported the results obtained using KFI as a ratio index to ensure comparability with other studies, although preliminary analyses suggested consistent findings when using KFI residuals.

None of the social and environmental variables that we have considered was directly related to temporal changes in fecundity. Spring–summer temperature was found to negatively affect adult deer fecundity only when the interaction with body condition was considered. In absence of randomized or experimentally controlled experiments, however, the use of multiple regression limits the possibility to explore the underlying cause–effect relationships among biological variables. Many studies on deer fecundity used multiple regression to investigate the direct effects of individual, social, and environmental drivers on female reproductive traits, and the occurrence of indirect effects between, for example, climate, density, body condition, and fecundity often remained descriptive (but see Stopher et al. 2014). Nonetheless, the occurrence of indirect effects between different factors is assumed to explain variation in fecundity (Bonenfant et al. 2009) and researchers have the possibility to formulate clear a priori hypotheses for how the individual, social, and environmental factors of interest are related to each other to explain variation in the trait under investigation (Myrsetrud et al. 2008).

Path analysis embraces this philosophical approach, allowing to explicitly consider more complex mechanisms underlying variation in fecundity, through the decomposition of biologically plausible sources of correlations among variables selected a priori. Path analysis supported our hypothesis that in temperate mountain environments, spring–summer temperature may negatively affect deer fecundity by reducing body condition. While several studies already hypothesized this relationship (e.g., Rugghetti and Festa-Bianchet 2012, Mason et al. 2014), to our knowledge this has never been quantitatively investigated in mountain-dwelling ungulates. The observed variation in fecundity likely reflects an adaptive response of females, to optimize the trade-offs between the opportunity for reproduction and survival under different environmental conditions (cf. Sand 1996). It remains unclear, however, which mechanism might account for the observed relationships: Given the predominantly nocturnal behavior of red deer in our study site,

we suggest body condition is most likely limited by summer resource availability (cf. Pettorelli et al. 2007), rather than by the temperature-mediated time constraints on foraging (cf. Mason et al. 2014). Ideally, to disclose which mechanism is at play, individual data on female summer foraging behavior are needed. This information should be integrated with information on forage quality in the area occupied by females over the same period, for example using values of fecal crude proteins, as the use of NDVI in a population that inhabits both forested and open habitats within and between seasons might be problematic (Borowik et al. 2013), at least without marked individuals.

Path analysis represents an appealing approach to investigate cause–effect mechanisms in biology, yet there are limitations in the use of this methodology. Like other frequentist approaches, path analysis is concerned with finding a model that does not reject the null hypothesis (i.e., the hypothesized correlational structure is consistent with the correlational structure of the sample data). If a model is not rejected, however, we cannot be sure it is the true model, as other models may fit the data equally well (Raykov and Marcoulides 2006). Alternative models may include direct or indirect effects of other variables that we did not take into account, and further research is needed to identify missed factors and disclose their mechanisms. In this respect, the association of model selection and path analysis may be useful, as the first allow to discriminate and identify influential variables, while the second may be used to test hypothesized causal relationships, helping to discriminate between potentially opposing mechanisms that can generate similar patterns. Finally, it is worth noting that the direct or indirect effect of a variable in a path analysis should be interpreted as “relative to the other variables that are explicitly invoked in the causal explanation” (Shipley 2016), not with respect to any other variable that might exist (Shipley 2016). The relationship between age or body condition and fecundity, for example, is likely to be mediated by variations in other parameters. Parasites, among the others, may play an important role in shaping fecundity variation, and their effect may be influenced by age, density, and temperature according to the transmission routes of the pathogen (Carlsson et al. 2018). In fact, our study population showed some evidence of negative

consequences of *Toxoplasma gondii* infection on fetal development, and this effect changed with hind age (Formenti et al. 2015). Whether the inclusion of further parameters will enable a better prediction of deer fecundity, and whether increasing temperatures in the future years could have long-term effects on the life history of our study population, however, still remain to be investigated.

Notwithstanding the caveats in the application of path analysis, our study highlights the importance of considering more complex relationships between individual, social, and environmental variables to explain variation in life-history traits (cf. Stopher et al. 2014). In this respect, it appears crucial to formulate sound a priori hypotheses on which factors to include in a causal model, and on the direct and indirect relationships between them. Exploratory data analysis, model selection, and information available in the literature (e.g., to identify the critical periods of the year during which weather conditions may affect the expression of reproductive traits) may help to formulate biologically plausible causal relationships. This, in turn, should allow a better understanding of the mechanisms underlying the adaptive responses of populations living in changing environmental conditions.

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