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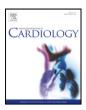
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International Journal of Cardiology xxx (2013) xxx-xxx



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- Time from adenosine di-phosphate receptor antagonist discontinuation to 1
- coronary bypass surgery in patients with acute coronary syndrome: 2
- Meta-analysis and meta-regression 3

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ABSTRACT

Background: Adenosine di-phosphate receptor antagonists (ADPRAs) blunt hemostasis for several days after admin-	2
istration. This effect, aimed at preventing cardiac ischemic complications particularly in patients with acute coronary	2
syndromes (ACS), may increase perioperative bleeding in the case of cardiac surgery. Practice Guidelines recom-	2
mend withholding ADPRAs for at least 5 days prior to surgery, though with a weak base of evidence. The purpose	2
of this study was to systematically review observational and experimental studies of early or late preoperative dis-	3
continuation of ADPRAs prior to coronary artery bypass grafting (CABG) for patients with ACS.	3
Methods: MEDLINE, EMBASE, the Cochrane Library databases up to December 2011; and reference lists.	3
Observational and experimental studies that compared early ADPRA discontinuation with late discontinuation, or	3
no discontinuation, in patients with ACS undergoing CABG.	3
Results: There were 19 studies, including 14,046 participants, 395 deaths and 309 reoperations due to bleeding.	3
ADPRA late discontinuation up to CABG was associated with an increased risk of postoperative mortality (ÔR	3
1.46, 95% confidence interval (CI) 1.10 to 1.93) and reoperations due to bleeding (OR 2.18; 95% CI 1.47 to 2.62).	
Between-study heterogeneity was low. Meta-analysis limited to high quality or prospective studies gave consistent	3
results. In most instances, the 95% prediction intervals for summary risk estimates confirmed the risk across study	3
groups.	4
Conclusions: ADPRA late discontinuation prior to CABG is associated with an increased risk of death and	4
reoperations due to bleeding in patients with ACS. The confidence in the estimates of risk for late discontinuation	4
is moderate to high.	4

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49 1. Introduction

The platelet adenosine di-phosphate receptor antagonists (ADPRAs) 5051are extensively used in patients with coronary artery disease (CAD) [1]. In stable CAD, they are given in combination with aspirin in order to 52prevent stent thrombosis after elective angioplasty. In acute coronary 53 54syndrome (ACS), regardless of stent implantation, they have been shown to reduce the aggregate risk of cardiovascular death, myocardial 55infarction (MI) and stroke by almost 20% by acting on an activated 56thrombotic milieu [2–6]. 57

Treatment prior to angioplasty is recommended by current guidelines 58 in order to prevent early ischemic events [1]. However, since two of these 59 agents (clopidogrel and prasugrel) exert an irreversible antiplatelet effect, 60 and the third (ticagrelor) is a powerful, though reversible, blocker, an 61 increased risk of perioperative bleeding exists in the case of urgent coro- 62 nary artery bypass grafting (CABG). In this setting, preoperative ADPRA 63 administration has been associated with increased rates of transfusion 64 and re-operation to stop bleeding. Regardless of the antiplatelet regimen, 65 reoperation due to bleeding is generally uncommon in patients undergo- 66 ing elective CABG (about 1.5%) [7]. However this risk is increased in the 67 case of urgent bypass surgery (4-15% among ACS patients) possibly due 68 to preoperative antithrombotic treatments [8,9]. Reoperation to stop 69 bleeding is associated with worse clinical outcomes and a 4.5 fold 70 higher mortality [10,11]. On this basis, withholding ADPRAs prior to 71 cardiac surgery might decrease re-exploration rate, chest tube drainage 72

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N. Morici et al. / International Journal of Cardiology xxx (2013) xxx-xxx

and post-operative mortality. Based upon multiple reviews and cohort 73 74 studies, current guidelines recommend discontinuing ADPRAs for at least 5 days, and preferably 7 days, before CABG in order to reduce 7576 bleeding complications [12-16].

However, whereas this recommendation seems to be safe in elective 77 CABG patients, the risk versus benefit of ADPRA discontinuation in ACS 78 79patients remains unclear.

80 Therefore, the aim of this meta-analysis was to determine whether 81 early compared to late discontinuation of ADPRAs affects the postoper-82 ative course of CABG in ACS patients, and to define the optimal timing of 83 ADPRA discontinuation prior to surgery.

84 2. Methods

2.1. Eligibility 85

86 This meta-analysis has been registered with PROSPERO-the NHR International Prospec-87 tive Register of Systematic Reviews (CRD42011001865) [17]. We followed a priori study el-88 igibility criteria for study selection. We included any observational and experimental study 89 that compared any early discontinuation of ADPRA drugs to later discontinuation treatment 90 or no discontinuation for patients with ACS referred to CABG surgery. We excluded studies 91 with fewer than 50% of patients with ACS, those without a comparison group and those pub-92lished in languages other than English

93 Evidence from observational studies was included because it is unlikely that patients 94 were randomized to receive immediate or postponed surgery after ADPRA administration 95 to obtain evidence of the mortality-delay association. Furthermore observational studies 96 may provide important additional information to RCTs with regard to specific populations, 97 administration modes, and outcomes, especially mortality.

We did not define a priori an optimal time for ADPRA discontinuation but accepted 98 99 what the authors claimed as the reference point between early and late discontinuation, 100 within an interval of maximum seven days.

101 2.3. Search strategy

Studies were identified by searching electronic databases and scanning reference lists of 102103 articles. This search was applied to Medline, and adapted for EMBASE and the Cochrane Library [i.e. Cochrane Central Register of Controlled Trials, the Cochrane Database of 104105 Systematic Reviews and Database of Abstracts of Reviews of Effects (DARE)] for studies 106 published in English between January 2001 and December 2011. The strategy was devel-107 oped using the search terms "clopidogrel", "thienopyridine", "ticagrelor", "prasugrel", ADP receptor antagonist", "antiplatelet therapy", "coronary artery bypass", "coronary artery 108 bypass graft", "graft occlusion", and "graft patency". The reference lists of relevant papers, 109 110including other systematic reviews focusing on this topic, and abstracts presented at the European and American Cardiology Congresses (2006-2012) were also searched. 111

112 2.4. Data extraction

The primary outcome was postoperative mortality (<30 days) and the secondary out-113 114 come was re-operation due to bleeding.

The exposure under consideration was ADPRA administration during 2-7 days 115 preceding CABG. The control group was defined as any other antiplatelet treatment 116 117 during 2-7 days preceding CABG, such as aspirin or no treatment.

118 We developed a data extraction sheet, pilot-tested it on five randomly-selected studies, and refined it accordingly. One review author (NM) extracted the following data from stud-119 120 ies included and entered in the data extraction form: patient demographics and baseline characteristics, the ADPRA loading dose, the number of ADPRA-free days before surgery, 121122the perioperative use of antifibrinolytic drugs, primary and secondary endpoints. A second 123author (JAO) checked the extracted data to ensure quality. Disagreements were solved by 124discussion between the two review authors; if no agreement was reached, a third author 125(VR) could decide.

126For all studies, we recorded the number of treated-patients, the number of non treated-127patients and the number of events in each group in order to estimate the odds ratios (OR) 128 and corresponding 95% confidence intervals (CI).

Data concerning transfusion requirements, myocardial infarction, chest-tube drainage, 129130and the duration of hospitalization were not included because of the heterogeneity of definitions, indications and measures applied in the different studies. 131

2.5. Methodological quality 132

133 Methodological quality was independently assessed by two review authors (NM and 134AS). The Newcastle-Ottawa (NOS) scale for cohort and case-control studies was used 135[18,19]. This scale has three groups of items: selection, exposure/outcome and compara-136bility. A study can be awarded a maximum of one star for each numbered item in 'patient 137 selection' (four items) and 'exposure or outcome' (for case-control or cohort studies re-138 spectively) (three items) and a maximum of two stars in the 'comparability of study 139groups' (two items), for a total of nine stars. Since we were interested in mortality of op-140 erated patients, we expected three items would be scored positively across all studies, 141

specifically ascertainment of exposure (secure ADPRA administration), demonstration

that outcome of interest was not present at the start of the study, and assessment of out- 142 come (record linkage). The same three items were verified for studies reporting only the 143 secondary endpoint. In fact, in our meta-analysis, the NOS scale could have ranged be- 144 tween three and nine. For randomized controlled trials (RCTs) we summarized the risk 145 of bias for mortality within study across the following specific domains: sequence gener- 146 ation, allocation concealment, and incomplete outcome data [18]. We decided a priori that 147 only observational studies that met eight or nine of the Newcastle-Ottawa Scale criteria 148 were to be considered of high quality, whereas RCTs were considered of high quality if 149 they satisfied two or more components. The Newcastle-Ottawa quality scoring assess-150ment is reported in Appendix, Table 1. 151

2.6. Statistical analysis

We did an overall quantitative synthesis using all ORs for mortality computed from the 153 frequencies obtained from each study. The results were pooled using the Mantel-Haenszel 154 random effects model described by DerSimonian and Laird [20] and ordered by study year. 155 Random effects model was used to synthesize data rather than the fixed effects model because they incorporate within- and between-study variability. This model was selected a 157 priori as the meta-analysis was expected to include primarily observational studies with in- 158 herently more variability than RCTs. A Mantel-Haenszel estimate was also computed and 159 compared to the DerSimonian and Laird estimate to investigate any influence of small 160 study effects on the pooled OR, since the DerSimonian and Laird methods tend to attribute 161 greater weight to small studies with increasing heterogeneity. The heterogeneity across 162 studies was assessed by the I-squared statistic and corresponding p-value. We explored 163 meta-analytic prediction intervals as means for providing a clear, appropriate and robust fu-164 ture treatment summary reflecting current estimates [21]. The prediction interval estimates 165 the possible treatment effect in a future study, and if it includes the null value of one it is 166 possible that the direction of the treatment effect in a single study may not be the same 167as that from the meta-analysis. 168

Pooled estimates were computed for each stratum of time from ADPRA discontinua-169tion to surgery. Pooled ORs were obtained even for the secondary endpoints throughout 170 the same approach. We also computed the predictive interval for the approximate predic-171 tive distribution of a future trial, based on the extent of heterogeneity, for both outcomes. Sensitivity analysis was performed to account for differences between the studies. Data 173 were synthesized for study design (RCT, prospective and retrospective studies), for 174 study quality (high and low qualities), for time from ADPRA discontinuation to surgery 175(<3, <4, <5, <7 days), and percentage of diabetic patients (>30% and <30%). 176

The extent to which study-level variables explained heterogeneity in predicting mortal-177 ity and re-operation due to bleeding was explored by fitting random effect meta-regression 178 models to account for the time from ADPRA discontinuation to surgery (<3, <4, <5, <7 days). 179 We checked for potential publication and small study effects by the visual inspection of con-180 tour enhanced funnel plot [22,23], and the test proposed by Harbord [24]. 181

We used the Grading of Recommendations Assessment, Development, and Evaluation 182 (GRADE) guidelines in order to rate the quality of evidence [25]. Factors that affect the 183 confidence in the estimate of effect include risk for bias (also known as detailed design 184 and study limitations), imprecision, indirectness (directness in the GRADE approach in-185cludes generalizability and applicability), inconsistency of results (heterogeneity), publi-186 cation bias, dose-effect responses, magnitude of effect, and issues of residual plausible 187 confounding. The confidence in the estimate of effect is categorized into 4 levels, ranging 188 from very low to high. The completed evidence summaries and GRADE assessments were 189 discussed by several investigators and reviewed by the methodological and clinical senior 190investigators. Evidence summaries were prepared for each research question by using 191 GRADE Profiler, version 3.6 (McMaster University, Hamilton, Ontario, Canada). 192

All statistical calculations were performed using Review Manager (RevMan), version 1935.0.24 and STATA version 11.1. 194

3. Results

3.1. Search results

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Database searches yielded 602 references, whereas one reference was 197 yielded through other sources ("Prasugrel as an anti-thrombotic therapy 198 in patients with ACS", presented to the Cardiovascular and Renal Drugs 199 Advisory Committee, 3 February 2009) (Fig. 1). After screening the ab- 200 stracts, 62 full-text articles were assessed for eligibility. The final meta- 201 analysis was based on 19 articles [8,26-42], including 13 observational 202 studies [8,28-30,32,34-37,39-42], two randomized clinical trials (RCTs) 203 [27,33], and four post hoc analyses of RCTs [26,31,38,43]. Although the 204 latter derived from RCTs, in this meta-analysis they are classified as obser- 205 vational since the randomization explored a subject different from the 206 optimal timing of discontinuation. The other 43 full-text articles were ex- 207 cluded because they included patients with stable coronary artery disease 208 (21 studies), or due to incomplete data availability in order to estimate 209 the primary and secondary endpoints (19 studies), or due to unknown 210

N. Morici et al. / International Journal of Cardiology xxx (2013) xxx-xxx

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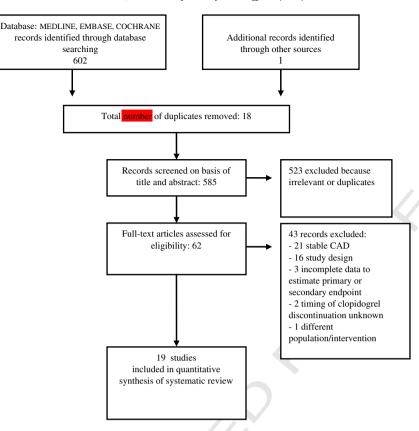


Fig. 1. Selection for studies exploring the association between post-operative outcome (mortality and reoperation due to bleeding) and optimal time to cardiac surgery in patients treated with ADPRA therapy.

timing of ADPRA discontinuation (2 studies) or, finally, different population/intervention (1 study).

213 3.2. Characteristics of studies

We identified 19 studies including 14,046 participants (24.04% females), with a mean age of 64.5 years. All but two included participants who were receiving clopidogrel. Another study included clopidogrel and ticagrelor [26] and the last included clopidogrel and prasugrel [43]. For the purpose of our meta-analysis, the results of the two study populations were disaggregated and each analyzed independently.

Table 1 summarizes the characteristics of the studies included in the 220meta-analysis. Most studies (12 studies) [8,27-29,31,34-39,43] consid-221222ered 5 days prior to surgery as cutoff for early vs late discontinuation, whereas the remainder used 3 days (4 studies) [32,33,40,41], 4 days 223(1 study) [30] and 7 days (1 study) [42]; only the PLATO (Platelet 224Inhibition and Patients Outcomes) –CABG trial used 5-day cut-off for 225clopidogrel administration and 3 day cut-off for ticagrelor administra-226227tion [26]. Seventeen of the 19 studies included information about con-228comitant aspirin treatment [8,26–34,36–38,40–43], the daily doses of which were 80–150 mg in the RCTs, 75–325 mg in the post-hoc analysis 229of RCTs, and 75–325 mg in nine of the observational studies (in the other 23026 the dose was not given). Thirteen studies (including the two RCTs) 231232did not mention the concomitant administration of GPIIb/IIIa inhibitors [27-29,31-37,39,41,42]. Antifibrinolytic therapies such as tranexamic 233 acid or aprotinin were not used in 3 studies [40-42], and not mentioned 234 in 8 [26-29,31,34,36,38]. The administration of ADPRA loading dose was 235not specifically mentioned in 10 studies (two RCTs) [27,29,30,32,33,35, 23636,38,41,42]; in the remainder between 15.9% and 100% of the patients 237received a loading dose. 238

Three studies [27,34,35] did not mention the percentage of diabetic patients; in the others this value varied from 21% to 54%. Five studies (two RCTs) included only patients undergoing first time CABG [27,28,33,40,41], whereas 12 also included patients undergoing 242 redo CABG (with percentages varying from 0.3% to 18.5%) [8,29–32, 243 34,36–39,42], and 2 did not specify [26,35]. A high rate of internal mam-244 mary artery grafting was reported in about 80% of the studies. Two studies 245 [40,42] (all observational) included only off-pump procedures, and 6 246 (two RCTs) only on-pump procedures [27,33,35,37,39,41]; the others in-247 cluded both, with the rate of on-pump procedures ranging from 1% to 248 87%. Ten (one RCT) [8,30,32,36,39–42] considered isolated CABG and 7 249 (one RCT) [26,28,31,33–35,38] did not specify; the percentage of concom-250 itant valve surgery in the remaining papers was between 6.6% and 19.5%. 251

3.3. Mortality

Seventeen studies considered the association between preoperative 253 ADPRA administration and post-operative death, including 3869 patients 254 with ADPRA early discontinuation (171 deaths) and 8975 patients 255 with **ADPRA** late discontinuation (223 deaths). Late or no preoperative 256 ADPRA discontinuation was associated with an increased rate of post- 257 operative death (OR 1.46; 95% CI 1.10–1.93) (Fig. 2). The estimated pre- 258 dictive interval was 0.82–2.59, meaning that an adverse effect of early 259 surgery after discontinuation might be a plausible finding in a new 260 study. Between-study heterogeneity was low (I-squared 15.7%). The 261 Mantel-Haenszel OR was 1.47 (95% CI, 1.16-1.86), suggesting an un- 262 likely impact of small studies on the random effects estimate towards 263 more beneficial values. Summary estimates for post-operative mortality 264 across strata of study design, quality, time from ADPRA discontinuation 265 to surgery, and percentage of diabetic patients are presented in Fig. 3. 266 All of the strata were consistent with the overall pooled estimate, al- 267 though the strata estimates were only significant for prospective (OR 268 1.69, 95% CI: 1.01-2.81) and high quality (OR 1.45, 95% CI: 1.06-1.99) 269 studies, and for studies with percentage of diabetic patients >30% (OR 270 1.43, 95% CI: 1.08–1.88). Timing of discontinuation had a uniform impact, 271

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N. Morici et al. / International Journal of Cardiology xxx (2013) xxx-xxx

t1.2 Study characteristics.

Table 1

1st author, year	Period of inclusion	Trial type	No of subjects	Diagnosis	Outcome	Concomitant ASA	Concomitant Antifibrinolytic	REDO-CABG	Concomitan valve surge
Gansera, 2003	2000-2002	Obs	64 Clo 64 Not Clo	ACS ^e	Death; re-op	Unknown	Unknown	Unknown	Unknown
Chu, 2004	1999_2001	Obs	41 Clo 271 Not Clo	ACS	Death; re-op	70%	76.5%	<1%	NO
Fox, 2004	1998_2000	Post-hoc	436 Clo 476 Not Clo	ACS	Re-op	100%	Unknown	4.6%	Unknown
Ascione, 2005	2001-2002	Obs	91 Clo 379 Not Clo	ACS	Death; re-op	47%	Unknown	NO	Unknown
Akowuah, 2005	2002-2003	RCT ^g	25 Clo 24 Not Clo	ACS	Death; re-op	100%	100%	NO	NO
Metha, 2006	2003-2004	Obs	739 Clo 113 Not Clo	ACS	Death	96%	Unknown	13.3%	Unknown
Berger, 2008	2007-2007	Obs	298 Clo 298 Not Clo	ACS	Death; re-op	91.9%	61%	5.0%	NO
Hyung-Jun-Kim, 20	08 1999-2003	Obs	332 Clo 4462 Not Clo	Mixed	Death; re-op	76.7%	10.4%	21.9%	NO
Filsoufi, 2008	1998_2005	Obs	72 Clo 72 Not Clo	Mixed	Death; re-op	89%	100%	4.5%	NO
Song, 2008	2004-2006	Obs	70 Clo 102 Not Clo	ACS	Death; re-op	100%	NO	NO	NO
Tabary, 2008	2003-2006	Obs	154 Clo 136 Not Clo	ACS	Re-op	63%	NO	NO	NO
Blasco, 2009	2000-2003	Obs	194 Clo 1483 Not Clo	Mixed	Death; re-op	100%	Unknown	8%	6.4%
Ebrahimi, 2009	2003-2005	Post-hoc	524 Clo 249 Not Clo	ACS	Death; re-op	97.7%	Unknown	Yes, Unknown%	Unknown
Firanescu, 2009	2006-2007	RCT	80 Clo 38 Not Clo	Mixed	Death; Re-op	100%	100%	NO	Unknown
Vaccarino, 2009	2003-2006	Obs	123 Clo 981 Not Clo	Mixed	Death; re-op	100%	NO	4.7%	NO
Nesher, 2010	2005-2008	Obs	189 Clo 262 Not Clo	ACS	Death; re-op	Unknown	100% Clo; %unknown not Clo	0.3%	NO
Mariscalco, 2011	2005-2010	Obs	225 Clo 225 Not Clo	Mixed	Death; re-op	100%	100%	1%	19.5%
Held, 2011	2008–2008	Post-hoc	Clo:412vs217 Tica:304vs328	ACS	Death	100%	Unknown	Unknown	Unknown
Smith, 2012	2004-2007	Post-hoc	Clo:91vs97 Pra:72vs105	ACS	Death	100%	Unknown	2.9%	NO

*1.42 $^{a}Obs = observational studies; ^{B}re-op = reoperation for bleeding; ^{c}Clo = clopidgrel group; ^{d}Mixed = studies with unknown percentage of unstable coronary syndrome or with percentage t1.43 >50% but <100%; ^{e}ACS: acute coronary syndrome; ^{fh} = studies including 2 different comparisons; ^{g}RCT = randomized controlled trial.$

although the majority of studies considered five days as cut-off (OR 1.45, 95% CI: 1.01–2.08).

The risk of postoperative mortality in patients who received new generation ADPRA therapy (prasugrel and ticagrelor) was not increased in patients with early discontinuation (OR 0.76; 95% CI 0.10–5.57) (Appendix Fig. 1).

278 3.4. Re-operation due to bleeding

Sixteen studies considered the association between preoperative 279ADPRA administration and the rate of re-operation due to bleeding, 280281 including 2883 patients with ADPRA early discontinuation (116 re-282 operations) and 8685 patients with ADPRA late discontinuation (193 re-operations). The frequency of re-operation in each study varied be-283tween 1.2% and 12.2% (median 3.2%). Later preoperative ADPRAs were 284associated with an increased risk of re-operation due to bleeding (OR 2852.18; 95% CI 1.47-3.25) (Fig. 4). The estimated predictive interval was 286 0.72–6.62, meaning that no effect or an adverse effect of late discontinu-287 ation might be a plausible finding in a new study. Between-study hetero-288 geneity was significant (p = 0.046) although modest (I-squared 41.7%). 289Seven studies were classified as high quality and meta-analysis limited 290to prospective studies gave consistent results. All the strata were consis-291tent with the overall pooled estimate, although the strata estimates 292were only significant for prospective (OR 2.82; 95% CI 1.52-5.24) and 293low quality (2.49; 95% CI 1.53-4.07) studies, and for percentage of diabet-294295 ic patients <30% (OR 3.59; 95% CI 1.63-7.89). Discontinuation gave

similar benefit at all days although the majority of studies considered 296 five days as cut-off (OR 1.77; 95% CI 1.20–2.62) (Appendix Fig. 2). 297

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3.5. Meta-regression analysis

Time from ADPRA discontinuation to surgery studied with meta- 299 regression yielded no significant effect on mortality or re-operation due 300 to bleeding, being the p-values > 0.20. Moreover, none of the ORs 301 changed after adjustment for percentage of diabetic patients (Table 2). 302

Visual inspection of the contour-enhanced funnel plot (Fig. 5) indi- 304 cated that pooled data did not appear to be heavily influenced by pub- 305 lication bias. This means that slight asymmetry of the plot is possible, 306 with few studies insisting in the area of significance and the majority 307 midway in the area of non-significance. The Harbord's test was not sta- 308 tistically significant (p = 0.167).

3.7. Summary findings

The summary findings following the GRADE guidelines are reported 311 in Table 3. Postoperative mortality is increased by one-half whereas the 312 risk of re-operation is approximately two times more likely for patients 313 discontinuing ADPRAs later. This means that, out of 1000 patients 314 discontinued from ADPRAs early, about 45 would die and about 40 315 would require re-operation for bleeding. However out of 1000 patients 316

N. Morici et al. / International Journal of Cardiology xxx (2013) xxx-xxx

Study	OR (95% CI)	Events, Treatment	Events, Control	% Weight
<=3				
Filsoufi - 2008	7.00 (0.37, 133.12)		0/72	0.89
Song - 2008	3.00 (0.12, 72.40)	1/70	0/70	0.76
Held - 2011 (a)	1.44 (0.69, 2.99)	16/304	12/328	10.76
Firanescu - 2009	(Excluded)	0/38	0/80	0.00
Subtotal (I-squared = 0.0% , p = 0.543)	1.63 (0.81, 3.25)	20/484	12/550	12.42
•				
<=4				
Chu - 2004	0.66 (0.09, 5.03)	1/41	10/271	1.83
Subtotal (I-squared = $.\%$, p = .)	0.66 (0.09, 5.03)	1/41	10/271	1.83
<=5				
Akowuah - 2005	4.81 (0.24, 95.25)	2/25	0/24	0.87
Ascione - 2005	7.29 (2.18, 24.37)	7/91	4/379	4.77
Metha - 2006	0.66 (0.28, 1.57)	26/739	6/113	8.34
Berger - 2008	4.00 (0.45, 35.58)	4/298	1/298	1.59
Hyung-Jun Kim - 2008	1.29 (0.66, 2.53)	9/332	94/4462	12.09
Blasco - 2009	1.66 (0.85, 3.24)	10/194	46/1483	12.28
Ebrahimi - 2009	1.90 (0.79, 4.59)	24/524	6/249	8.10
Nesher - 2010	1.39 (0.45, 4.23)	6/189	6/262	5.47
Mariscalco - 2011	1.00 (0.33, 3.05)	6/225	6/225	5.47
Held - 2011 (b)	1.73 (0.97, 3.08)	46/412	14/217	14.93
Smith - 2012 (b)	0.71 (0.26, 1.92)	6/91	9/97	6.68
Smith - 2012 (c)	0.16 (0.01, 2.95)	0/72	4/105	0.91
Gansera - 2003	(Excluded)	0/64	0/64	0.00
Subtotal (I-squared = 35.6% , p = 0.105)	1.45 (1.01, 2.08)	146/3256	196/7978	81.51
<=7				
Vaccarino - 2009	1.60 (0.44, 5.81)	4/88	5/176	4.24
Subtotal (I-squared = $.\%$, p = .)	1.60 (0.44, 5.81)	4/88	5/176	4.24
Overall (I-squared = 15.7% , p = 0.270)	1.46 (1.10, 1.93)	171/3869	223/8975	100.00
.5 1 3				

favours early discontinuation favours late discontinuation

(a): Ticagrelor-assigned patients.
(b): Clopidogrel-assigned patients.
(c): Prasugrel-assigned patients.
ADPRAs: Adenosine diphosphate receptor antagonists.
CABG: Coronary artery bypass graft surgery.
Treatment: late-no discontinuation
Control: early discontinuation

discontinued from ADPRAs later, about 70 (25 more patients) would die
and 87 (47 more patients) would need to undergo re-operation to stop
bleeding.

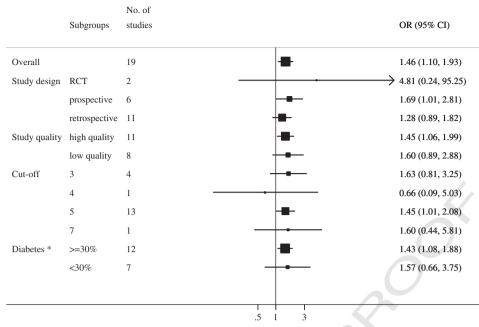
320 4. Discussion

Our findings, derived from 19 randomized and observational studies 321 of ADPRAs in patients with an ACS undergoing urgent CABG, indicate 322 that surgery performed within five days of last drug administration is 323 associated with increased mortality and reoperation due to bleeding. 324 Despite the fact that the majority of the information originates from ob-325servational studies, our analysis provides a moderate to high confidence 326 in the estimates of effect. Moreover, the amount of evidence from obser-327 vational studies increased the precision of the estimates we found in the 328 RCTs, with consistent direction and size of the effects. The prediction in-329 330 terval also suggests that further studies are likely to confirm that longer time of discontinuation is safer than later discontinuation. Unfortunate-331 ly, we could not define the best time period for discontinuation, since 332 the majority of the studies considered five-days as cutoff, and the evidence exploring other cutoff times was sparse and limited. 334

This conclusion is consistent with the pharmacology of currently avail-335 able oral ADPRAs and reinforces present guideline recommendations of discontinuing these agents for at least five days prior to surgery [44,45]. 337 The current recommendation of starting dual antiplatelet therapy with aspirin and an ADPRA in ACS patients immediately upon admission in order to prevent early ischemic events is based on the results of the 340 CURE study [34], but this approach creates a clinical dilemma in those pa-441 tients who will later show at coronary angiography an indication to CABG. 442 From one side, withdrawal of ADPRA before surgery might expose the pa-443 tients to ischemic events in the preoperative period; however, its contin-444 uation up to the time of surgery has been shown to increase postoperative bleeding. Surgeons have been persuaded to operate on aspirin therapy 346

Fig. 2. Risk of postoperative mortality in patients who received ADPRA therapy as compared to those who stopped before CABG, stratified according to the time to discontinuation. The combined OR and 95% CI were calculated using the random-effects models.

N. Morici et al. / International Journal of Cardiology xxx (2013) xxx-xxx



favours early discontinuation favours late discontinuation

*The sum does not add up to the total because of missing values.

ADPRAs: Adenosine diphosphate receptor antagonists.

CABG: Coronary artery bypass graft surgery.

Fig. 3. Subgroup analysis for postoperative mortality in patients who received ADPRA therapy versus who stopped before CABG. The combined OR and 95% CI were calculated using the random-effects models.

since the convincing study by Mangano et al. has shown a 60 percent 347 lower mortality and 50 percent reduction in ischemic complications 348 without excess bleeding among patients receiving aspirin within 48 h 349 after operation [46,47]. However, they are not similarly confident with 350 ADPRAs due to the increased risk of bleeding complications and lack of 351 evidence of a protective effect towards early post-operative ischemic 352 events [48]. Antiplatelet therapy, especially with thienopyridines, may 353 354 significantly contribute to cardiopulmonary bypass-induced platelet dys-355 function, causing an increase in chest drain blood loss, utilization of blood 356 products and incidence of re-explorations. Perioperative blood loss de-357 manding transfusion has been shown to be associated with an eightfold increase of death after surgery [49], whereas re-exploration for bleeding 358 359 after cardiac surgery is an independent predictor of adverse events such as sepsis, renal failure, acute respiratory distress syndrome and prolonged 360 mechanical ventilation [50]. All these consequences can ultimately lead to 361 death of patients as well as to increased hospital costs [51]. 362

The landmark CURE study, which established the benefit of early 363 364 clopidogrel use in addition to aspirin in the secondary prevention of 365 ischemic events among patients with NSTEACS [34], is of little help for investigating the impact of this therapy in ACS patients undergoing ur-366 gent CABG. As clearly reported in the specific subanalysis of the study, 367 the median time from index admission to CABG was 25.5 days, 93% of 368 the patients who proceeded to CABG stopped clopidogrel before 369 CABG, the median time off the study drug before CABG was 17 days 370 (interquartile range, 9 to 33) and the median time after CABG was 371 10 days (interquartile range, 6 to 25) [2,34]. 372

A meta-analysis of 11 observational studies published by Purkayastha et al. [52] involved 4002 CABG patients: pre-operative continuation of clopidogrel was independently associated with a significant increase in major bleeding, reoperation due to bleeding and transfusion requirements. Similar results were reported by Pickard et al. [53]. Both metaanalyses had insufficient power to be conclusive about the risk of death, and suffered from heterogeneity between studies. Several other meta- 379 analyses have been published on this issue over the last years, but they in- 380 cluded both stable CAD and ACS patients, and showed conflicting results 381 [10,48]. A recent meta-analysis [54], specifically focused on ACS patients, 382 showed a trend towards increased mortality (HR 1.44, 95% CI 0.97–2.01, 383 p=0.07) and reoperation rates among patients undergoing CABG with- 384 out clopidogrel discontinuation. 385

Recent data with the newer ADPRA, ticagrelor [26] and prasugrel 386 [43] in patients with ACS show similar or slightly increased rates of 387 bleeding and reoperation, however, with significantly reduced mortali-388 ty as compared to clopidogrel. Also in the case of these agents, current 389 guidelines recommend at least five days of discontinuation prior to 390 surgery. 391

How complex and confusing is this issue is demonstrated by an 392 extensive review recently published by Burke et al. [55]. The authors 393 underline how the ACC/AHA 2007 NSTEACS guidelines endorsed the 394 use of clopidogrel upstream of coronary angiography in all patients 395 irrespective of subsequent modality of treatment [56]. The updated 396 ACC/AHA 2009 guidelines have modified this approach by stating that 397 clopidogrel may be administrated "before or at the time of PCI", in pa-398 tients with NSTEACS, but they do not give any new evidence to support 399 this changing recommendation [57]. The same paper, revising data com- 400 ing from the wide population included in the GRACE [58] and CRUSADE 401 [59] registries, underscores that clopidogrel is underused in patients 402 with NSTEACS, and especially in the high risk subgroup, probably because 403 of concerns about CABG-related bleeding. In the CRUSADE study, in spite 404 of the national guidelines, 87% of clopidogrel-treated patients underwent 405 CABG within 5 days after discontinuation of treatment. Both the increas- 406 ing hospitalization costs and the urgent underlying disease make often 407 difficult to wait the suggested period. These facts cast doubt on the rec- 408 ommendation for a fixed safe waiting period following discontinuation 409 of antiplatelet therapy. 410

N. Morici et al. / International Journal of Cardiology xxx (2013) xxx-xxx

Study	OR (95% CI)	Events, Treatment	Events, Control	% Weight
<=3				
Song - 2008	1.00 (0.06, 15.67)	1/70	1/70	1.86
Tabary - 2008	4.42 (0.98, 19.80)	10/154	2/136	5.05
Firanescu - 2009	2.41 (0.12, 48.95)	2/80	0/38	1.58
Filsoufi - 2008	(Excluded)	0/72	0/72	0.00
Subtotal (I-squared = 0.0%, p = 0.640)	> 3.01 (0.90, 10.07)	13/376	3/316	8.49
<=4				
Chu - 2004	7.93 (2.54, 24.81)	6/41	5/271	7.28
Subtotal (I-squared = .%, p = .)	7.93 (2.54, 24.81)	6/41	5/271	7.28
<=5				
Gansera - 2003	→ 11.00 (0.62, 194.90)	5/64	0/64	1.72
Fox - 2004	1.79 (0.85, 3.74)	18/436	11/476	11.20
Akowuah - 2005	0.96 (0.21, 4.30)	3/25	3/24	5.06
Ascione - 2005	4.16 (1.86, 9.30)	11/91	11/379	10.45
Berger - 2008	3.50 (1.17, 10.51)	14/298	4/298	7.60
Hyung-Jun Kim - 2008	1.27 (0.69, 2.34)	11/332	116/4462	12.80
Blasco - 2009	- 2.55 (0.94, 6.93)	5/194	15/1483	8.45
Ebrahimi - 2009	1.90 (0.41, 8.89)	8/524	2/249	4.85
Nesher - 2010	1.07 (0.48, 2.38)	10/189	13/262	10.46
Mariscalco - 2011	0.78 (0.29, 2.05)	7/225	9/225	8.73
Subtotal (I-squared = 35.4% , p = 0.125)	1.77 (1.20, 2.62)	92/2378	184/7922	81.32
<=7				
Vaccarino - 2009	10.00 (1.19, 84.29)	5/88	1/176	2.91
Subtotal (I-squared = .%, p = .)	10.00 (1.19, 84.29)	5/88	1/176	2.91
Overall (I-squared = 41.7% , p = 0.046)	2.18 (1.47, 3.25)	116/2883	193/8685	100.00
.5 1 3				

favours early discontinuation favours late discontinuation

ADPRAs: Adenosine diphosphate receptor antagonists. CABG: Coronary artery bypass graft surgery Treatment: late-no discontinuation Control: early discontinuation

Fig. 4. Risk of re-operation for bleeding in patients who received ADPRA therapy as compared to those who stopped before CABG stratified for the time to discontinuation. The combined OR and 95% CI were calculated using the random-effects models.

The strengths of our review include its focus on ACS patients, a com-411 prehensive search methodology, inclusion of randomized as well as ob-412 413

servational studies, and detailed assessment of the factors that influence

the confidence in the results. It adds data on timing of interventions 414 (such as early vs. late discontinuation), hard outcomes (mortality 415 and re-operation due to bleeding), subgroups of patients (such as 416

t2.1Table 2

Random effects meta-regression analysis of postoperative mortality and re-operation for bleeding in patients who received ADP receptor antagonist as compared to those who t2.2stopped before CABG. t2.3

t2.4		No. of studies	No. of participants	OR (95% CI)	p-value	OR (95% CI) ^a	rp-value ^a
t2.5	Mortality						
t2.6	Cut-off						
t2.7	≤3	4	1034	1 ^b		1 ^b	
t2.8	≤ 4	1	312	0.38 (0.03-5.78)	0.46	0.33 (0.02-5.95)	0.42
t2.9	≤5	13	11,234	0.85 (0.29-2.46)	0.75	0.83 (0.30-2.24)	0.68
t2.10	≤7	1	264	0.94 (0.14-6.51)	0.97	0.72 (0.09-5.49)	0.73
t2.11	Drug			+		*	
t2.12	Ticagrelor/prasugrel	2	809	1 ^b		1 ^b	
t2.13	Other	17	12,035	1.22 (0.46-3.24)	0.67	1.18 (0.46-3.02)	0.70
t2.14				*		*	
t2.15	Re-operation for bleeding						
t2.16	Cut-off						
t2.17	≤3	4	692	1 ^b		1 ^b	
t2.18	≤ 4	1	312	3.74 (0.45-30.92)	0.20	3.62 (0.36-36.46)	0.23
t2.19	≤5	10	10,300	0.84 (0.18-3.87)	0.80	0.86 (0.16-4.62)	0.83
t2.20	≤7	1	264	4.71 (0.25-87.79)	0.27	4.09 (0.17-98.59)	0.33

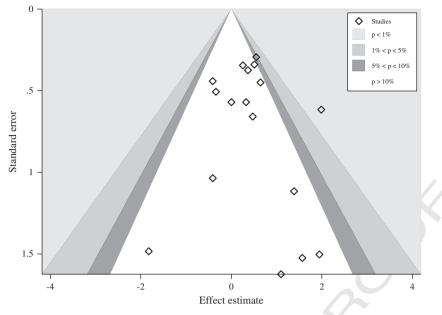
ADP: adenosine diphosphate. t2 21

t2.22CABG: coronary artery bypass graft surgery.

t2.23^a Estimates adjusted for percentage of diabetes.

t2.24 ^b Reference category.

N. Morici et al. / International Journal of Cardiology xxx (2013) xxx-xxx



ADPRAs: Adenosine diphosphate receptor antagonists. CABG: Coronary artery bypass graft surgery.

Fig. 5. Contour enhanced funnel plot of studies comparing patients who received ADPRA therapy versus who stopped before CABG for postoperative mortality.

diabetics) and use of newer agents (prasugrel and ticagrelor) that were
not available from previous meta-analyses. Reporting and publication
bias might be limited: the scientific community is alerted and the outcomes considered relevant.

420 Our review has limitations that deserve attention for both interpreting 421 422 the results and conducting future research. The inclusion of observational studies increases the risk for bias due to the lack of control for con-423 founders and covariates. We could not assess whether different cutoffs 424 in discontinuation yielded different results. The majority of evidence is 425based on a five-day cutoff. The amount of risk might be variable with 426 different length of drug discontinuation, with a skewed distribution of 427the difference between early and late discontinuation. Only few studies 428 explored discontinuation at the extremes. Our meta-regression which 429considered cutoff discontinuation as a continuous outcome has advan-430 431 tages over subgroup analysis: it explores the covariate over an expanded

range of values. The paucity of studies in some time intervals reduces the 432 confidence in the absence of differences between alternative cutoff days. 433 Observational studies can supplement this evidence by contributing data 434 about special populations (e.g. high risk population, as diabetic patients). 435

Risk stratification for both ischemic and bleeding events is 437 recommended by current ACS guidelines, especially for patients 438 presenting without persistent ST-segment elevation [60]. However, 439 recommended risk scores and algorithms are of no help for estimating the 440 probability of urgent CABG whose incidence is extremely variable across 441 centers, and in clinical trials has been shown to range from 10 [31] to 16% 442 [61,62]. Although the extent of ST-segment changes during acute ische-443 mia have been shown to correlate with the presence of three-vessel 444

t3.1 Table 3

8

t3.2 Quality of the evidence according the "Grading of Recommendations Assessment, Development, and Evaluation" (GRADE) approach.

t3.3	ADPRA discontinuation com	pared to ADPRA continuation fo	or patients with ACS under	going CABG			-
t3.4	Bibliography:	•					
t3.5	Outcomes	No. of participants (studies) follow up	Quality of the evidence (GRADE)	Relative effect (95% CI)	Anticipated absolute effects		
					Time frame is short-term mortality		t3.6
					Risk with ADPRA discontinuation	Risk difference with ADPRA continuation (95% CI)	t3.7
$t3.8 \\ Q2 t3.9 \\ t3.10$	Mortality	12489 (19 studies) 30 days	⊕⊕⊕⊕ HIGH ¹	RR 1.56 (1.2 to 2.03)	Study population 45 per 1000	25 more per 1000 (from 9 more to 46 more)	-
					High 100 per 1000	56 more per 1000 (from 20 more to 103 more)	t3.11 t3.12
t3.13 t3.14 t3.15	Re-operation for bleeding	11568 (9 studies) 30 days	$\oplus \oplus \oplus \ominus$ MODERATE ^{1,2} Due to inconsistency	RR 2.18 (1.47 to 3.25)	Study population 40 per 1000	47 more per 1000 (from 19 more to 91 more)	
					High 100 per 1000	118 more per 1000 (From 47 more to 225 more)	t3.16 t3.17 t3.18

*The basis for the **assumed risk** (e.g. the median control group risk across studies) is provided in footnotes. The **corresponding risk** (and its 95% confidence interval) is based on the t3.20 assumed risk in the comparison group and the **relative effect** of the intervention (and its 95% CI).

N. Morici et al. / International Journal of Cardiology xxx (2013) xxx-xxx

disease and left main disease [63], attempts to estimate the probability of 445 early CABG in ACS have shown only moderate discriminative value [61]. 446 The present data provides solid evidence that performing CABG without 447448 allowing at least five days off ADPRAs in ACS patients is associated with a significantly increased risk of reoperation to stop bleeding and mortali-449 ty. In the lack of comparably solid evidence that administration of ADPRAs 450prior to angiography improves outcome in ACS, the present findings 451should suggest caution in recommending ADPRA administration prior 452453to angiography, particularly within the current scenario of very early angiography across the ACS spectrum [64]. Estimating the probability 454455of early CABG requires sound clinical judgment, and selective use of 456short acting GPIIb/IIIa receptor blockers in patients deemed at risk of urgent surgery may provide adequate antiplatelet protection, as well as 457458reduction or perioperative ischemic risk, without paying the price of increased perioperative bleeding [65,66]. 459

Appendix A. Supplementary data 460

461 Supplementary data to this article can be found online at http:// dx.doi.org/10.1016/j.ijcard.2012.12.087. 462

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