# Rising adiposity curbing decline in the incidence of myocardial infarction: 20-year follow-up of British men and women in the Whitehall II cohort 

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Received 12 November 2010; revised 9 March 2011; accepted 11 April 2011; online publish-ahead-of-print 8 June 2011


#### Abstract

Aims To estimate the contribution of risk factor trends to 20-year declines in myocardial infarction (MI) incidence in British men and women.

Conclusion In men and women, over half of the decline in MI risk could be accounted for by favourable risk factor time trends.

\section*{Methods} and results

Keywords

From 1985 to 2004, 6379 men and 3074 women in the Whitehall II cohort were followed for incident MI and risk factor trends. Over 20 years, the age-sex-adjusted hazard of MI fell by $74 \%$ ( $95 \%$ confidence interval $48-87 \%$ ), corresponding to an average annual decline of $6.5 \%$ (3.2-9.7\%). Thirty-four per cent (20-76\%) of the decline in MI hazard could be statistically explained by declining non-HDL cholesterol levels, followed by increased HDL cholesterol ( $17 \%, 10-32 \%$ ), reduced systolic blood pressure ( $13 \%, 7-24 \%$ ), and reduced cigarette smoking prevalence ( $6 \%$, $2-14 \%$ ). Increased fruit and vegetable consumption made a non-significant contribution of $7 \%(-1-20 \%)$. In combination, these five risk factors explained $56 \%$ ( $34-112 \%$ ). Rising body mass index (BMI) was counterproductive, reducing the scale of the decline by $11 \%(5-23 \%)$ in isolation. The MI decline and the impact of the risk factors appeared similar for men and women. The adverse role of BMI emphasizes the importance of addressing the rising population BMI.

Myocardial infarction • Incidence • Time Trends • Population • Prevention • Risk factors


## Introduction

Coronary heart disease (CHD) incidence has declined appreciably in the UK (by almost two-thirds among men in the last 25 years), ${ }^{1}$ contributing to a substantial fall in CHD mortality. ${ }^{2}$ However, the reasons for the decline in CHD incidence are not well understood. Furthermore, the possible impact of rising adiposity in restraining
the decline remains unknown. Addressing these uncertainties is important, because despite the declines CHD remains the leading cause of death in the UK, ${ }^{2}$ the USA, and other wealthy countries, ${ }^{2}$ while control of the emerging CHD epidemic in the developing world is an increasing priority. ${ }^{3}$

Although several studies have examined the contribution of risk factor trends to changes in incidence or mortality in the UK ${ }^{1,4,5}$

[^0]and in other locations, ${ }^{6-8}$ most have used aggregate data sources and are subject to the limitations of ecological analyses. Few studies have used individual-level data from single populations. ${ }^{1,6}$ One previous UK-based investigation assessed major risk factors only twice and did not include women. ${ }^{1}$ The aim of this present analysis was therefore to estimate the contribution of risk factor trends to recent trends in the incidence of major CHD in the Whitehall II cohort of British men and women followed over a period of 20 years, with clinic visits every 5 years.

## Methods

## The Whitehall II study

The Whitehall II study, described elsewhere, ${ }^{9}$ recruited 10308 men and women, aged $35-55$, from 20 civil service departments in London in 1985-88. At baseline (1985-88), phase 3 (1991-93), phase 5 (1997-99), and phase 7 (2002-04), the participants completed clinical examinations and questionnaires on health and lifestyle. Participants were flagged at the National Health Service Central Registry, which provided information on the date and cause of death. Ethical approval for the Whitehall II study was obtained from the UCL Medical School Committee on the ethics of human research. Informed consent was obtained from the study participants.

## Coronary events

The outcome was a first myocardial infarction (Ml, fatal or non-fatal) between baseline and 2002-04 (end of phase 7) (mean follow-up of 15.4 (SD 4.2) and 9.0 (SD 4.5) years for participants who were censored and those who experienced the outcome, respectively). Participants who developed angina (either before baseline or during follow-up) were retained for analysis. Fatal MI was identified as a record of death with CHD as the underlying cause, including sudden death of presumed cardiac origin (international classification of diseases, ninth revision, codes 410-414). Potential new cases of non-fatal MI were ascertained by questionnaire items on chest pain and the physician's diagnosis of heart attack in all four phases listed above. Only those cases confirmed according to MONICA criteria using electrocardiograms, markers of myocardial necrosis, and chest pain history from the medical records, were included. ${ }^{10}$

## Assessment of risk factors

At each of the three study phases: baseline (1985-88), phase 3 (199193), and phase 5 (1997-99), cigarette smoking status, physical activity levels, elements of diet, and alcohol consumption were ascertained by questionnaire, while fasting lipid levels, systolic blood pressure (SBP), and body mass index (BMI) were obtained from clinical examinations, using consistent techniques on each occasion. ${ }^{11,12}$ At baseline, 9065 participants ( $88 \%$ ) had no HDL cholesterol measurement, but serum apolipoprotein-A1 was available for almost $80 \%$ of participants. ${ }^{11}$ Age- and gender-adjusted linear regression of the available HDL data on apolipoprotein-A1 was used to estimate the relationship between the two variables and then predict the baseline HDL for those participants with no data (see Supplementary material online, Appendix 1). Alcohol consumption in the previous week was measured as units per week, then categorized as none, within recommended limit for gender ( $<21$ units for men, $<14$ units for women), over recommended limit, and very heavy ( $>50$ units for men and $>35$ units for women). Cigarette smoking categories were non-smoker, ex-smoker, and current smoker. Dietary data available were usual milk consumption (categorized as none, whole milk, semi-skimmed, skimmed and other), usual bread consumption (white, wholemeal,
granary or wheatmeal, other brown bread, other), and usual fruit and vegetable consumption (less than three times per week, three to four times per week, five to six times per week, daily, two or more times per day). Physical activity levels were categorized as low ( $<2 \mathrm{~h}$ per week of moderate activity and $<1 \mathrm{~h}$ of vigorous activity), high ( $\geq 2.5 \mathrm{~h}$ per week of moderate activity or $>1 \mathrm{~h}$ of vigorous activity), or medium (levels in between low and high). ${ }^{13}$

## Statistical analyses

Cox regression was used to estimate associations between each risk factor at phase 1 and subsequent MI hazard to justify inclusion in the main analyses prior to computing attributable proportions. All risk factors except type of milk were significantly associated (positively or negatively) with MI hazard. To estimate secular time trends in the risk factors and in MI, the follow-up for each participant was split into three consecutive periods, each of approximately 5 years, separated by the different examination phases: a first period from phase 1 to phase 3; a second from phase 3 to phase 5, and a third from phase 5 to phase 7. The MI incidence during, and risk factor levels at the start of, different periods are then compared, adjusting for age and gender, to assess secular trends over time. In particular, age-adjusted secular time trends among men and women from phases $1-5$ in the risk factors were estimated from the regression of the risk factor on calendar time (of start of period), in this split data set using generalized estimating equations with robust standard errors to take account of dependency between repeated measures for each participant. Poisson models were fitted for percentage change in prevalence of being a current cigarette smoker, having medium or more physical activity levels, consuming alcohol over the recommended limit, usually eating white bread, and usually eating fruit and vegetables twice or more daily; and linear models for time trend in mean BMI, SBP, HDL, and non-HDL cholesterol. Cox regression on this split data set was used to estimate the time trend in the hazard of MI overall and by gender, again using robust standard errors to account for dependency between repeated observations for each participant. Age was used as the underlying time scale, enabling automatic adjustment for age. ${ }^{14}$ There was no evidence of departure from the proportional hazards assumption of the Cox regression tested using Schoenfeld residuals. ${ }^{15}$
The extent to which the secular time trends in each of the risk factors statistically explained the trend in MI hazard was estimated by the expression $\left(\beta_{0}-\beta_{1}\right) / \beta_{0}$, where $\beta_{0}$ is the coefficient of calendar time in the Cox model with calendar time as the single covariate and $\beta_{1}$ is the coefficient of calendar time in a Cox model adjusting additionally for the risk factor(s). ${ }^{16}$ Bias-corrected bootstrap resampling gave an approximate $95 \%$ confidence interval (CI) for this estimate. ${ }^{17}$ Squared terms in the continuous risk factors (BMI, SBP, and HDL and non-HDL cholesterol) were added to the models to test for nonlinearity; squared-terms for HDL and non-HDL cholesterol were significant and so retained in the final models. The above analyses were applied to men and women combined, adjusting for gender. Further exploratory analyses were carried out to estimate risk factor contributions to the decline in $M I$ in men and women separately. Participants missing data on one or more risk factors in a particular phase were excluded from that phase and the associated follow-up, but included in other phases. Participants missing data in all phases were excluded from analyses altogether. The number of participants included in each phase is given in Table 1. A $P$-value of 0.05 was used as the threshold for statistical significance and all tests were two-sided. Stata, version 11.1 (Stata Corp., College Station, TX, USA), was used for all analyses.

Table I Number of participants contributing data in each study phase by age group (participants with complete risk factor data)

|  | Age group (years) |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 35-39 | 40-44 | 45-49 | 50-54 | 55-59 | 60-64 | 65-68 | All |
| Men |  |  |  |  |  |  |  |  |
| Number of participants |  |  |  |  |  |  |  |  |
| Phase |  |  |  |  |  |  |  |  |
| 1 (1985-88) | 1333 | 1354 | 928 | 1048 | 129 | 0 | 0 | 4792 |
| 3 (1991-93) | 100 | 1537 | 1401 | 1015 | 1029 | 143 | 0 | 5225 |
| 5 (1997-99) | 0 | 9 | 783 | 964 | 614 | 653 | 180 | 3203 |
| Number of subsequent incident major CHD events ${ }^{\text {a }}$ |  |  |  |  |  |  |  |  |
| Phase |  |  |  |  |  |  |  |  |
| 1 (1985-88) | 6 | 4 | 9 | 25 | 5 |  |  | 49 |
| 3 (1991-93) | 1 | 21 | 12 | 31 | 36 | 6 |  | 107 |
| 5 (1997-99) |  | 0 | 4 | 9 | 14 | 21 | 4 | 52 |

## Women <br> Number of participants

Phase

| $1(1985-88)$ | 547 | 562 | 532 | 629 | 99 | 0 | 0 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $3(1991-93)$ | 45 | 549 | 582 | 487 | 575 | 76 | 0 |
| $5(1997-99)$ | 0 | 2 | 291 | 336 | 288 | 289 | 79 |

Number of subsequent incident major CHD events ${ }^{\text {a }}$
Phase

| $1(1985-88)$ | 0 | 2 | 3 | 6 | 1 | 12 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| $3(1991-93)$ | 0 | 4 | 1 | 6 | 2 | 22 |
| $5(1997-99)$ |  | 0 | 1 | 1 | 8 | 4 |

${ }^{\text {a }}$ Events occurring between phases 1 and 3 for participants at phase 1, between phases 3 and 5 for participants at phase 3, and between phase 5 and phase 7 (2002-4) for participants at phase 5.

## Results

## Time trends in myocardial infarction incidence

Of the 10308 participants recruited, 1 had no follow-up data, 35 reported an MI before baseline, and 819 had missing data on one or more risk factor at all phases and were excluded from the analysis. The remaining 9453 participants included 6379 (67\%) men and 3074 women. A total of 256 first MI events occurred during 107892 person-years of follow-up; 208 first MI events occurred among men during 74474 person-years of follow-up and 48 first MI events occurred during 33418 person-years of follow-up among women. There was an average annual age-sex-adjusted decline in MI hazard over the course of the follow-up of $6.5 \%$ ( $95 \% \mathrm{Cl} 3.2-9.7$ ) corresponding to a 20 -year fall from baseline (1985-88) to 2004 of $74 \%$ ( $95 \% \mathrm{Cl} 48-87$ ). Men experienced a 20 -year fall of $73 \%$ ( $95 \%$ $\mathrm{Cl} 42-87$ ) and women experienced a 20 -year fall of $82 \% ~(95 \%$ $\mathrm{Cl}-5-97)$. There was no evidence of a gender-time interaction ( $P=0.7$ ).

## Time trends in risk factors

The levels of each risk factor according to age group and study phase for men and women are shown in Appendix 2 in the Supplementary material online. Favourable time trends occurred in several risk factors between 1985-88 and 1997-99: adjusting HDL cholesterol rose, prevalence of consumption of fruit and vegetables twice or more daily rose (comparable statistically significant changes for men and women); cigarette smoking prevalence fell (statistically significant among women only), and prevalence of at least moderate physical activity levels fell (statistically significant among men only) (Table 2). Bread consumption did not change among men or women. There were unfavourable increases in mean BMI and alcohol consumption, adjusting for age.

## Role of risk factors trends in myocardial infarction incidence trends

Four risk factor trends contributed in isolation to the $74 \%$ decline in MI hazard among all participants (Table 3). Percentage contributions of these risk factors in order of size were: declining non-HDL

Table 2 Age-adjusted population-averaged time trends in coronary risk factors among men and women over 12 years from 1985-88 (baseline) to 1997-99 (phase 5)

| Risk factor | Men |  |  | Women |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Change in mean levels per annum ( $95 \% \mathrm{Cl}$ ) | $P$-value | Change over 12 years ( $95 \% \mathrm{Cl}$ ) | Change in mean levels per annum (95\% Cl) | $P$-value | Change over 12 years ( $95 \% \mathrm{Cl}$ ) |
| BMI (kg/m ${ }^{2}$ ) | 0.10 (0.08, 0.11) | $<0.001$ | 1.16 (0.99, 1.33) | 0.07 (0.03, 0.10) | $<0.001$ | 0.78 (0.41, 1.15) |
| Systolic blood pressure ( mmHg ) | $-0.35(-0.42,-0.28)$ | $<0.001$ | $-4.19(-5.02,-3.35)$ | $-0.52(-0.63,-0.41)$ | $<0.001$ | -6.21 (-7.52, -4.90) |
| Non-HDL cholesterol (mmol/L) | $-0.033(-0.038,-0.028)$ | $<0.001$ | $-0.40(-0.46,-0.33)$ | $-0.047(-0.054,-0.039)$ | $<0.001$ | $-0.56(-0.65,-0.47)$ |
| HDL cholesterol (mmol/L) | 0.011 (0.009, 0.012) | $<0.001$ | 0.13 (0.11, 0.15) | 0.006 (0.004, 0.009) | $<0.001$ | 0.08 (0.04, 0.11) |
|  | \% change in prevalence per annum ( $95 \% \mathrm{Cl}$ ) | $P$-value | \% change over 12 years ( $95 \% \mathrm{Cl}$ ) | \% change in prevalence per annum ( $95 \% \mathrm{Cl}$ ) | $P$-value | \% change over 12 years ( $95 \% \mathrm{Cl}$ ) |
| Current smoker | $-0.80(-1.89,0.30)$ | 0.2 | $-9.2(-20.4,3.6)$ | $-3.78(-4.94,-2.62)$ | $<0.001$ | $-37.1(-45.5,-27.2)$ |
| At least moderate physical activity | $-1.06(-1.35,-0.76)$ | $<0.001$ | -12.0 (-15.1, -8.8) | $-0.48(-1.16,0.21)$ | 0.2 | $-5.6(-13.1,2.5)$ |
| Consume alcohol over recommended limit | 6.12 (5.15, 7.10) | $<0.001$ | $104(82.8,128)$ | 7.96 (5.79, 10.17) | $<0.001$ | 151 (96.5, 220) |
| White bread as usual bread type | $-0.26(-0.53,0.01)$ | 0.06 | $-3.1(-6.2,0.1)$ | 0.12 (-0.24, 0.47) | 0.5 | $1.4(-2.8,5.8)$ |
| Consume fruit and vegetables twice or more daily | 7.99 (7.01, 8.98) | $<0.001$ | 151 (125, 180) | 8.73 (7.56, 9.92) | $<0.001$ | 173 (140, 211) |

cholesterol $34 \%$ (bootstrap $95 \% \mathrm{Cl} 20-76$ ), rising HDL cholesterol 17\% (bootstrap $95 \%$ Cl 10-32), declining SBP 13\% (bootstrap 95\% Cl7-24), and declining cigarette smoking $6 \%$ (bootstrap $95 \% \mathrm{Cl} 2-$ 14). Together, they explained a total of $54 \%(95 \% \mathrm{Cl} 34-105)$ of the decline (the upper bound of the Cl indicates that the data are consistent at the $95 \%$ confidence level with the risk factors explaining a greater decline than that observed). The contribution of increased fruit and vegetable consumption did not reach statistical significance ( $7 \%$, bootstrap $95 \% \mathrm{Cl}-1$ to 20 ), the combined contribution with the four other risk factors being $56 \%$ (bootstrap $95 \% \mathrm{Cl} 34-112$ ). Trends in physical activity, alcohol consumption, and bread consumption had no notable impact. The rise in mean BMI was adverse, explaining $-11 \%$ (bootstrap $95 \% \mathrm{Cl}-23$ to -5 ) of the decline in MI hazard in isolation. The proportion of the decline explained by the risk factors combined reduced from 56 to $48 \%$ (bootstrap $95 \% \mathrm{Cl} 27-96$ ) with additional adjustment for the adverse trend in BMI. This suggests that the MI decline could be $8 \%$ greater in the absence of rising BMI .

Considering men and women separately, the relative contributions of each of the risk factors to the MI declines within each gender were generally similar to each other and to that in the combined analysis (Table 4). Exceptions were that among women, there was a smaller contribution from HDL cholesterol compared with SBP and the proportion explained by cigarette smoking was not significant. Further, among women, a negative impact of BMI
was not as apparent. In secondary analyses considering fatal events only, results were similar (data not shown).

## Discussion

Over 20 years between 1985 and 2004, there was a substantial decline of $74 \%$ in the age-adjusted hazard of first MI among men and women the Whitehall II cohort. Over half of the MI decline could be explained by a combination of favourable time trends in major risk factors, particularly non-HDL cholesterol, HDL cholesterol, SBP, and cigarette smoking. Rising adiposity had an adverse impact on the declining trend in MI, such that had other risk factor trends not occurred, rising BMI may have led to an increase in MI incidence over the follow-up. The MI decline and the risk factor contributions were broadly similar for men and women.

Multiple repeated measurements of risk factors, using consistent techniques for the measurement of the physical factors on each occasion, are a key strength of this study. We linked risk factor trends to coronary events at an individual level, thus avoiding the limitation of ecological analyses predominantly used to study time trends. Further, this is apparently the first analytical study of MI trends in a cohort following both men and women. We used consistent methods to identify Mls throughout the follow-up period to limit confounding of the estimate of the incidence trend by changes in diagnostic criteria. Silent Mls were not

Table 3 Fall in hazard of a first myocardial infarction per annum among all participants and percentage of this fall explained by risk factor time trends

| Model | Risk factors adjusted for in addition to age and gender | Fall in hazard per annum, \% (95\% CI) | $P$-value | \% of the observed decline in hazard explained by the risk factor(s) $(95 \% \mathrm{CI})^{\text {a }}$ |
| :---: | :---: | :---: | :---: | :---: |
| A | No adjustment | 6.51 (3.22, 9.68) | $<0.001$ |  |
| Effect of adjustment for individual risk factors in isolation |  |  |  |  |
| B | Smoking (current/ex/never) | 6.13 (2.82, 9.33) | <0.001 | $5.9(2.3,13.6)$ |
| C | Physical activity (low/medium/high) | 6.51 (3.20, 9.70) | <0.001 | $0.1(-4.5,5.3)$ |
| D | Alcohol units per week (none/within limit/over limit/heavy) | 6.44 (3.13, 9.65) | $<0.001$ | $1.0(-6.1,8.3)$ |
| E | Usual bread consumption (white/wholemeal/ granary or wheatmeal/other brown bread/ combination) | 6.55 (3.26, 9.72) | $<0.001$ | $-0.6(-3.3,0.3)$ |
| F | Usual fruit and vegetable consumption (less than three times per week/three to four times per week/five to six times per week/ daily/one or more times per day) | 6.07 (2.72, 9.31) | $<0.001$ | 6.8 (-1.1, 19.9) |
| G | BMI ( $\mathrm{kg} / \mathrm{m}^{2}$ ) (continuous) | 7.18 (3.94, 10.32) | <0.001 | -10.8 (-23.2, -4.6) |
| H | Systolic blood pressure ( mmHg ) (continuous) | 5.70 (2.41, 8.87) | 0.001 | 12.8 (7.4, 24.4) |
| I | HDL cholesterol (mmol/L) (continuous) | 5.45 (2.13, 8.67) | 0.001 | 16.6 (9.9, 32.3) |
| J | Non-HDL cholesterol (mmol/L) (continuous) | 4.32 (0.79, 7.72) | 0.02 | 34.4 (20.4, 75.7) |
| Effect of adjustment for combinations of risk factors |  |  |  |  |
| K | Smoking, non-HDL cholesterol, HDL cholesterol, systolic blood pressure | 3.05 (-0.47, 6.44) | 0.09 | 54.4 (34.4, 105) |
| L | Smoking, non-HDL cholesterol, HDL cholesterol, systolic blood pressure, usual fruit and vegetable consumption | $2.92(-0.64,6.36)$ | 0.1 | 55.9 (34.3, 112) |
| M | Smoking, non-HDL cholesterol, HDL cholesterol, systolic blood pressure, usual fruit and vegetable consumption, BMI | 3.44 (-0.15, 6.91) | 0.06 | 47.9 (26.6, 95.5) |

${ }^{\mathrm{a}} \%$ of the observed fall in hazard rate explained by risk factor $=100 \% \times\left(\beta_{0}-\beta_{1}\right) / \beta_{0}$, where $\beta_{0}$ is the coefficient of calendar time in the Cox regression model which only included calendar time (Model A ) and $\beta_{1}$ is the coefficient of calendar time in the Cox regression model adjusting additionally for the risk factor(s)
included, and the outcome thus corresponds to major CHD events. Risk factor levels were related to MI events up to 5 years ahead, based on the interval between clinic phases, and there is evidence that the benefits of smoking cessation, changes in blood lipids, and blood pressure are realized in this timeframe. ${ }^{18-20}$

There are several limitations. The analyses were necessarily based on participants who re-attended after baseline and provided complete risk factor data at one phase at least. This could introduce survival and response biases which might overestimate the favourable trends observed, due to a healthy participant effect. However, survivor bias is unlikely to be marked as survival in the cohort is high.' Including those participants with missing risk factor data, the 20 -year decline was smaller: $62 \%$ ( $95 \% \mathrm{Cl}$ $34-78$ ) indicating some response bias. As we could arguably expect similar overestimation of the favourable risk factor trends, the percentage explained by each risk factor may still be comparable. HDL cholesterol values at baseline were derived from serum apolipoprotein-A1 for a subgroup of the participants. The likely impact is underestimation of the variance associated with the baseline HDL measurements but without biasing the estimate of the contribution of HDL to the MI decline. Any measurement imprecision of the risk factors, particularly likely for the questionnaire-
derived dietary factors, physical activity, and alcohol consumption, may have led to the underestimation of the contribution to the MI decline. Questions on physical activity at phase 5 were more detailed than those in the earlier phases, giving more opportunity to report activity, which could lead to the underestimation of the physical activity decline and its counterproductive role. The analyses of the risk factor contributions by gender lack precision (Cls for the percentage contributions are wide), particularly for women who experienced few events (48 in total), and should thus be considered exploratory. Diabetes was not considered in this analysis. It is likely that diabetes lies on the causal pathway between several of the risk factors considered here and major CHD risk. Including diabetes in the analysis would therefore be problematic and could lead to the underestimation of the effects of the risk factors. ${ }^{21}$ The limitation of not considering diabetes is that we are unable to ascertain the extent to which the adverse effect of increasing BMI levels operates through an increase in diabetes (particularly type 2 diabetes) incidence. Effort was made to model carefully the relationship between the risk factors and MI incidence, for example, by inclusion of squared terms in the continuous variables in the Cox regression models, where significant. However, if the relationship between the risk factors and MI incidence is not fully captured in the Cox

Table 4 Fall in hazard of a first myocardial infarction per annum and percentage of this fall explained by risk factor time trends among men and women

| Model |  | Men |  |  | Women |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Fall in hazard per annum, \% (95\% $\mathrm{Cl})$ | $P$-value | \% of decline in hazard explained by risk factor(s), (95\% CI) ${ }^{\text {a }}$ | Fall in hazard per annum, \% (95\% CI) | $P$-value | \% of decline in hazard explained by risk factor(s), (95\% CI) ${ }^{\text {a }}$ |
| A | No adjustment | 6.26 (2.66, 9.73) | 0.001 |  | 8.12 (-0.25, 15.80) | 0.06 |  |
| Effect of adjustment for individual risk factors in isolation |  |  |  |  |  |  |  |
| B | Smoking (current/ex/ never) | 5.96 (2.34, 9.45) | 0.001 | $4.8(1.4,13)$ | 7.51 (-0.96, 15.27) | 0.08 | 7.9 (-3.3, 43.9) |
| C | Physical activity (low/ medium/high) | 6.31 (2.70, 9.79) | 0.001 | $-0.9(-6.9,4.7)$ | 7.78 (-0.75, 15.58) | 0.07 | 4.5 (-3.1, 60.7) |
| D | Alcohol units per week (none/within limit/ over limit/heavy) | 6.25 (2.62, 9.75) | 0.001 | 0.1 (-9.2, 9.3) | $7.84(-0.64,15.60)$ | 0.07 | $3.7(-10.2,32.6)$ |
| E | Usual bread consumption (white/ wholemeal/granary or wheatmeal/other brown bread/ combination) | 6.29 (2.69, 9.75) | 0.001 | $-0.5(-3.8,0.6)$ | 8.16 (-0.26, 15.86) | 0.06 | -0.4 (-9.1, 6.4) |
| F | Usual fruit and vegetable consumption (less than three times per week/three to four times per week/five to six times per week/ daily/one or more times per day) | 5.87 (2.20, 9.40) | 0.002 | 6.3 (-2.2, 23.1) | 7.29 (-1.32, 15.18) | 0.1 | 10.6 (-23.1, 57.3) |
| G | BMI $\left(\mathrm{kg} / \mathrm{m}^{2}\right)$ (continuous) | 7.32 (3.73, 10.77) | $<0.001$ | -17.6 (-41.1, -8.2) | 8.22 (-0.06, 15.81) | 0.05 | -1.2 (-28.0, 14.2) |
| H | Systolic blood pressure (mmHg) (continuous) | 5.58 (1.98, 9.05) | 0.003 | 11.1 (5.7, 25.5) | 6.67 (-1.73, 14.37) | 0.1 | 18.5 (6.8, 69.8) |
| 1 | HDL cholesterol (mmol/ <br> L) (continuous) | 5.10 (1.43, 8.62) | 0.007 | 19.1 (10.2, 39.0) | 7.48 (-0.91, 15.17) | 0.08 | 8.3 (1.0, 44.4) |
| J | Non-HDL cholesterol (mmol/L) (continuous) | 4.19 (0.32, 7.91) | 0.03 | 33.8 (18.2, 87.4) | 5.52 (-3.47, 13.73) | 0.2 | 33.0 (10.8, 214) |

Effect of adjustment for combinations of risk factors

| K | Smoking, non-HDL cholesterol, HDL cholesterol, systolic blood pressure | 2.99 (-0.88, 6.71) | 0.1 | 53.0 (30.7, 123) | 3.66 (-5.35, 11.89) | 0.4 | 56.0 (21.5, 269) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| L | Smoking, non-HDL cholesterol, HDL cholesterol, systolic blood pressure, usual fruit and vegetable consumption | 2.90 (-1.02, 6.67) | 0.1 | $54.4(29.8,126)$ | 3.38 (-5.66, 11.65) | 0.5 | 59.4 (19.2, 221) |
| M | Smoking, non-HDL cholesterol, HDL cholesterol, systolic blood pressure, usual fruit and vegetable consumption, BMI | 3.48 (-0.49, 7.30) | 0.09 | 45.1 (21.7, 119) | 3.76 (-5.27, 12.02) | 0.4 | 54.7 (11.2, 210) |

[^1] included calendar time (Model A ) and $\beta_{1}$ is the coefficient of calendar time in the Cox regression model adjusting additionally for the risk factor(s).
models, this may have led to the underestimation of the association between the risk factors and the MI risk and in turn the underestimation of the percentage of the decline in MI explained by the risk factors.

Trends in non-HDL cholesterol had the greatest single impact on the decline in MI incidence. The favourable time trend in non-HDL cholesterol may reflect the increasing use of lipid-regulating medication or lifestyle (e.g. diet) or some combination of factors. Statin use rose to $11 \%$ of the cohort ( $25 \%$ of those with high LDL cholesterol) by the end of the follow-up in 2004, ${ }^{22}$ suggesting that lipid-regulating medication may have made an appreciable contribution.

The combined contribution of the risk factors to the MI decline in the present study was similar to that found in a national cohort of men over a similar period (46\%), but the individual relative impacts of the risk factors differed between the two cohorts. ${ }^{1}$ The decline in smoking prevalence had greater impact in the national cohort, possibly explained by the already lower prevalence of smokers at a later baseline in the present study ( $23 \%$ among men in Whitehall II compared with $\sim 40 \%$ among men in the national cohort). The trend in and contribution of non-HDL cholesterol was smaller in the national cohort (non-HDL cholesterol fell by $0.4 \mathrm{mmol} / \mathrm{L}$ over 12 years in Whitehall II men, compared with $0.35 \mathrm{mmol} / \mathrm{L}$ over 20 years in the national cohort ${ }^{23}$ ), possibly reflecting greater take-up of effective lipid-lowering medication in the present study. ${ }^{22}$ The differences may reflect the higher socioeconomic status in the present London-based cohort. Indeed, in results stratified by employment grade, the risk factor contributions in the lowest grades corresponded more closely to the national cohort findings (data not shown).

In a comparable analysis of US women, $68 \%$ of the decline in CHD incidence could be explained by combined trends in smoking, diet (decreased saturated fat, increased fibre content), and post-menopausal hormone use. ${ }^{6}$ Dietary trends in isolation accounted for the largest part of the decline (52\%). The greater contribution of diet in the US investigation is likely to reflect the influence of diet on risk factors such as blood pressure and cholesterol not available in that study, but included as explanatory variables in our analysis. Any protective effect of hormone therapy is doubtful in the light of recent evidence from the Women's Health Initiative. ${ }^{24}$ Finally, the WHO MONICA Project suggested that cigarette smoking, SBP and total cholesterol together explained approximately $38 \%$ of the variation in trends in coronary event rates from the mid-1980s to the mid-1990s in men in 27 different populations. ${ }^{7}$ The lower total percentage explained may reflect the ecological analysis, using aggregate data to study variations in trends between populations, rather than studying variation over time in individuals within one population as in the present study.

## Implications

In this cohort of London civil servants, there was a substantial decline in MI over two decades to 2004, more than half of which could be attributed to favourable risk factor trends, highlighting what can be achieved and emphasizing the value of measures to reduce exposure to these risk factors in the population. The risk factor trends were of comparable importance for men and women, suggesting that similar influences have operated
to achieve declines in MI incidence, such that similar prevention strategies may be appropriate for both genders. Further research is needed to determine whether the residual unexplained portion of the decline in MI may be explained by early treatment, underestimated contributions of the major risk factors (reflecting imprecision in the analyses), or the influence of other risk factors. The apparent lack of association of the decline in physical activity with the time trend in MI may reflect the methodological limitations associated with quantifying activity levels or the measured decline in the activity levels was insufficient to influence MI incidence.

While the negative contribution of rising mean BMI over recent decades appears to have been outweighed by the favourable trends in other vascular risk factors, continued increases in BMI may further reduce or even reverse the decline in MI incidence. The extent to which the rise in BMI may have influenced the time trend in MI through an increase in the incidence of diabetes cannot be evaluated from this analysis. The association between type 2 diabetes and CHD risk is well established and previous studies suggest that a concurrent rise in incidence of type 2 diabetes has occurred which may be at least in part explained by rising $\mathrm{BMI},{ }^{25}$ supporting the influence of BMI on the time-trend in CHD operating to some extent through rising diabetes. Sharply rising trends in statin and BP-lowering medication ${ }^{2,23}$ may contribute to continuing favourable Ml incidence trends in the UK and other rich countries but it is unlikely that the healthcare systems in emerging economies will have the necessary resources to provide the level of care needed to compensate for the increasing prevalence of overweight and obesity already taking place. The rising $B M I$ in the UK and in other countries needs therefore urgent attention.

## Supplementary material

Supplementary material is available at European Heart Journal online.

## Funding

S.L.H. is supported by a Medical Research Council Training Fellowship in Health Services Research \& Health of the Public (G0701739). M.J.S. is supported by a grant from the British Heart Foundation. The Whitehall Il study has been supported by grants from the Medical Research Council, British Heart Foundation, Health and Safety Executive, Department of Health, Stroke Association, National Heart Lung and Blood Institute (HL36310), National Institute on Aging (AG13196) and Agency for Health Care Policy Research (HS06516), and the John D and Catherine T MacArthur Foundation. Funding to pay the Open Access publication charges for this article was provided by the Medical Research Council.

## Conflict of interest: none declared.

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[^1]:    ${ }^{\mathrm{a}} \%$ of the observed fall in hazard rate explained by risk factor $=100 \% \times\left(\beta_{0}-\beta_{1}\right) / \beta_{0}$, where $\beta_{0}$ is the coefficient of calendar time in the Cox regression model which only

