

#### Spatial patterns of African ungulate aggregation reveal complex but limited risk effects from reintroduced carnivores

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#### 1 Spatial patterns of African ungulate aggregation reveal complex but limited risk effects

#### 2 from reintroduced carnivores

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Abstract. The 'landscape of fear' model, recently advanced in research on the non-lethal effects 16 of carnivores on ungulates, predicts that prev will exhibit detectable antipredator behavior not 17 only during risky times (i.e., predators in close proximity) but also in risky places (i.e., habitat 18 19 where predators kill prey or tend to occur). Aggregation is an important antipredator response in numerous ungulate species, making it a useful metric to evaluate the strength and scope of the 20 landscape of fear in a multi-carnivore, multi-ungulate system. We conducted ungulate surveys 21 over a two-year period in South Africa to test the influence of three broad-scale sources of 22 variation in the landscape on spatial patterns in aggregation: 1) habitat structure, 2) where 23 carnivores tended to occur (i.e., population-level utilization distributions) and 3) where 24 carnivores tended to kill ungulate prey (i.e., probabilistic kill site maps). We analyzed spatial 25 variation in aggregation for six ungulate species exposed to predation from recently reintroduced 26 lion (Panthera leo) and spotted hyena (Crocuta crocuta). Although we did detect larger 27 aggregations of ungulates in 'risky places', these effects existed primarily for smaller-bodied 28 (<150kg) ungulates and were relatively moderate (change of  $\leq 4$  individuals across all habitats). 29 In comparison, ungulate aggregations tended to increase at a slightly lower rate in habitat that 30 was more open. Lion, an ambush (stalking) carnivore, had stronger influence on ungulate 31 aggregation than hyena, an active (coursing) carnivore. In addition, places where lions tended to 32 kill prey more strongly affected ungulate aggregation than places where lions tended to occur, 33 but an opposing pattern existed for hyena. Our study reveals heterogeneity in the landscape of 34 fear and suggests broad-scale risk effects following carnivore reintroduction only moderately 35 influence ungulate aggregation size and vary considerably by predator hunting mode, type of 36 predation risk, and prey species. 37

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38 Key words: antipredator behavior; carnivore predation risk; landscape of fear; predator-prey
39 interaction; risk effects; ungulate aggregation

#### 40 **INTRODUCTION**

Examinations of predator-prey ecology have focused on direct, lethal predation effects on 41 prey (Lima 1998), but recent work has expanded this view by quantifying non-lethal "risk 42 effects" (i.e., altered prey behavior due to risk; Creel and Christianson 2008). These risk effects, 43 which have been well-described in aquatic systems (e.g., Peacor and Werner 2001), are now 44 being assessed in vast carnivore-ungulate systems. From this research, we understand that 45 ungulates employ a variety of strategies to manage their interactions with carnivores. These 46 include migrating to areas with lower carnivore densities (Hebblewhite and Merrill 2007), 47 selecting habitat that provides cover (Valeix et al. 2009a, Thaker et al. 2011), and aggregating 48 into larger groups as a means of additional vigilance, dilution, and defense (Hamilton 1971, 49 Dehn 1990, Hebblewhite and Pletscher 2002). Predation-sensitive migration and habitat 50 selection has been relatively well-documented in a variety of carnivore-ungulate systems (e.g., 51 Hebblewhite and Merrill 2007, Valeix et al. 2009a), although understanding the demographic 52 consequences of these behaviors remains a challenge. Comparatively, the factors shaping 53 ungulate aggregation are more poorly understood, largely due to the complexity of herding 54 behavior in heterogeneous environments (Pays et al. 2007, Fortin et al. 2009). For instance, 55 whereas larger aggregations of ungulates can reduce predation risk via dilution, they can also 56 result in decreased foraging efficiency (Molvar and Bowyer 1994). In addition, larger 57 aggregations offer improved predator detection, but they are also more easily detected by 58 predators (Caro 2005). Given that aggregation is one of the most common antipredator strategies 59 60 employed by ungulates (Caro 2005), a more thorough understanding of the degree to which

spatial variation in risk from multiple carnivores influences ungulate aggregation is crucial toboth predator-prey and behavioral ecology.

Several ungulate species in African systems modulate their group size in response to 63 carnivore predation risk (Valeix et al. 2009b, Thaker et al. 2010, Creel et al. 2014). Whereas 64 evidence suggests that ungulates aggregate when faced with predation risk in temporal 65 dimensions (e.g., Valeix et al. 2009b), it is unclear how spatial aggregation patterns relate to 66 predation risk. The "landscape of fear" model of predator-prev interactions (Laundré et al. 2001) 67 predicts that antipredator processes such as aggregation will be detectable not only during risky 68 times (i.e., when carnivores are in the immediate vicinity) but also in risky places (i.e., where 69 carnivores typically occur or kill prey; Valeix et al. 2009a, Valeix et al. 2009b, Thaker et al. 70 2011, Creel et al. 2014). Although the landscape of fear theory has been presented as a 71 72 framework for understanding broad-scale ecosystem dynamics (Ripple and Beschta 2004), the strength and scope of carnivore-induced risk effects are highly variable. For example, in African 73 ecosystems, prey body size (Hopcraft et al. 2010), climate (Riginos 2015), and forage quality 74 (Anderson et al. 2010) have been found to influence ungulate behavior as much or more than 75 predation risk, while other work suggests interactive and context-specific effects of these factors 76 (Sinclair and Arcese 1995). Furthermore, risk effects depend on predator hunting mode, with 77 ambush predators generally eliciting the strongest prey responses (e.g., Schmitz 2008, Thaker et 78 al. 2010). 79

Here, we aim to clarify the relative contribution of multiple forms of landscape-level predation risk on patterns of aggregation for six species of African ungulates exposed to multicarnivore predation. A key objective of our study is to contextualize the top-down influence of carnivores on ungulate aggregation by comparing its effect size with that of an established

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relationship between ungulate aggregation size and habitat structure (i.e., aggregations tend to be
larger in more open habitat; Evans 1979, Gerard and Loisel 1995, Pays et al. 2007, Thaker et al.
2010, Marino and Baldi 2014). In addition, our study evaluates the hypothesis that carnivore risk
effects on ungulates will vary by hunting mode and by prey species, as has been indicated in
predator-prey systems across a diversity of taxa (Preisser et al. 2007).

To these ends, we conducted surveys over a two-year period in South Africa to quantify average ungulate aggregation size. We then compared spatial variation in ungulate aggregation to two types of risk associated with the primary carnivores in this system: 1) the probability of carnivore occurrence (population-level utilization distributions (UDs)) and 2) the probability of carnivore kill occurrence (probabilistic kill site maps). Given that both aggregation size and predation risk vary by ungulate species (Thaker et al. 2011, Gervasi et al. 2013), we developed this assessment at the level of each ungulate species.

#### 96 METHODS

#### 97 *Study Area*

We assessed spatial patterns in ungulate aggregation in the 132 km<sup>2</sup> fenced Main Camp 98 section of Addo Elephant National Park, South Africa (Addo; 33°30'S, 25°45'E; Appendix A). 99 The dominant habitat type in Addo is thicket with grassland sections in the east that are 100 derivative of an agricultural history (Tambling et al. 2012). Unlike typical African savanna 101 ecosystems, Addo lacks pronounced wet and dry seasons (e.g., average wet and dry season 102 rainfall only differs by ~20mm; South African Weather Service, unpublished data). Using 103 digital vegetation maps provided by South African National Parks, we identified six habitat types 104 within Addo, including bushclump, gully thicket, thicket, Zuurkop woodland, long grassland, 105 106 and short grassland (South African National Parks, unpublished data). The two primary

carnivores in the study area, lion (*Panthera leo*) and spotted hyena (*Crocuta crocuta*), were
reintroduced in the Main Camp following a 100-year absence (Hayward and Hayward 2006).
All lions (4M, 2F) and four hyenas (3M, 1F) were unrelated and were reintroduced in late 2003;
an additional four unrelated hyenas (2M, 2F) were reintroduced in the winter of 2004 following a
period of disease testing (see Hayward et al. 2007a, 2009). One leopard (*Panthera pardus*) also
occurred in the study area but was not included in the analysis due negligible predation on the
ungulate species of interest.

114 Ungulate surveys

Between December 2003 and November 2005, we conducted surveys to describe spatial 115 variation in aggregation size for Cape buffalo (Syncerus caffer; hereafter buffalo), eland 116 (Tragelaphus oryx), greater kudu (Tragelaphus strepsiceros; hereafter kudu), red hartebeest 117 (Alcelaphus buselaphus caama), warthog (Phacochoerus africanus) and zebra (Equus quagga). 118 These species occurred across the entire study area and were common prey of lion and hyena in 119 Addo (Hayward 2006, Hayward et al. 2007b). We conducted vehicle-based surveys on a 120 network of 51 transects ( $\bar{x}$  transect length = 913 m, range 567 – 1,029 m; Appendix A). We 121 conducted surveys three hours preceding and following dawn and dusk, using spotlights when 122 necessary. These time periods correspond with peak activity times for lions and hyenas in Addo 123 (Hayward and Hayward 2006). We conducted 30 surveys per year on randomly selected days 124 and varied the start location for surveys to eliminate temporal bias at any given location. We 125 126 recorded aggregation size at the start, middle, and end points of each transect. The same observer conducted each survey alongside a scribe. Detections occurred without the aid of 127 128 binoculars, although counts were confirmed using binoculars. In most cases (i.e., >80%), only 129 one aggregation of a given species was present at a transect point during a given survey. When

multiple, distinct groups of the same species were seen at a given location, we first corrected 130 group size via distance sampling to accommodate variation in visibility by regressing the log of 131 group size against the estimated detection probability in each habitat type (Buckland et al. 2001). 132 We then summed group sizes to obtain species-specific aggregation sizes. We considered groups 133 distinct when individuals were closer to other members of their group than the nearest peripheral 134 members of another group. We counted all individuals (i.e., adults and juveniles) and assumed 135 counts close to the line were accurately recorded (Buckland et al. 2005). Thus, aggregations 136 represent all individuals of the same species at a transect point at a given time, whether one 137 distinct aggregation or an aggregation of two or more groups in close proximity to the transect 138 point. 139

140 *Carnivore occurrence models* 

We developed population-level UDs for lions and hyenas to model predation risk as a 141 function of carnivore occurrence (Hebblewhite and Merrill 2007, Valeix et al. 2009, Thaker et al. 142 2011). We VHF-tracked the entire lion population (N = 6) and 75% (N = 6) of the adult hyena 143 population for two years following release in Addo. Both lions and hyenas were consistently 144 tracked throughout the two-year study period via continuous, 96-hour follows conducted on 145 average once every six weeks; opportunistic observations of carnivore locations and kills were 146 also recorded outside these times (see Hayward et al. 2009). As with ungulate surveys, carnivore 147 location data were primarily collected during times when lions and hyenas were most active (i.e., 148 within 3 three hours of dawn and dusk; Hayward and Hayward 2006, Hayward et al. 2009). In 149 this way, we attained an average of 48 locations (SD = 18) per hyena (N = 285 total locations for 150 6 hyenas) and 103 locations (SD = 23) per lion (N = 615 total locations for 6 lions). Beginning 151 152 in November 2003, we partitioned lion locations into three, four-month periods per year through

October 2005: summer (November - February), autumn/early winter (March - June), and late 153 winter/spring (July – October). We then mapped population-level lion UDs for each period to 154 account for temporal variation in lion ranging patterns following reintroduction (i.e., range 155 expansion; Hayward et al. 2009). Given the lower sample size of hyena locations (N = 285), we 156 developed a pooled UD representing population-level hyena space use over the two-year period 157 (Seaman et al. 1999). We developed UDs in R (R Development Core Team 2011) using a 158 bivariate plug-in matrix (Gitzen et al. 2006). We depicted the interpolated density estimates as 159 UD percentiles at a  $30m^2$  resolution such that the highest percentile (100) represented the very 160 core of the UD, while the lowest percentile (1) represented the periphery. Areas not 161 encapsulated by any portion of the UDs had a value of 0, representing a negligible probability of 162 carnivore occurrence. 163

164 *Kill occurrence models* 

We identified carnivore-killed ungulates via continuous follows and incidental sightings. We distinguished lion kills from those of hyenas based on observations of hunts and kills, signs and direct observations of feeding on fresh carcasses. We saw no indication of intraguild scavenging between lions and hyenas during continuous follows and have no other evidence of its occurrence, as both species were at low densities during our study period.

We used the locations of killed ungulates to develop probabilistic spatial models of kill occurrence throughout the study area as a function of landscape characteristics, which were assigned using a lattice comprised of  $250m^2$  cells in ArcMap Version 10. We identified relevant landscape covariates *a priori* for both lion and hyena kill occurrence models. These included the proportion of each habitat type and the length of hard edge within a cell, in addition to the Euclidean distance from kill site to fenceline, waterholes (N = 11), and campsites (N = 2). Given

that lions use cover to ambush prey (Loarie et al. 2013) and frequently kill near edge habitat 176 (Prins and Iason 1989), we defined hard edge as the boundary between open (short and long 177 grassland) and closed habitats (gully thicket, thicket, Zuurkop woodland, and bushclump). We 178 179 included the fenceline, waterhole, and campsite covariates to account for increased carnivore hunting success near fences (Davies-Mostert et al. 2013), variation in ungulate water dependency 180 (Valeix et al. 2009b) and carnivore/ungulate avoidance of anthropogenic infrastructure (Howe et 181 al. 2013). For lions, we evaluated both a pooled model (using all locations of lion-killed 182 ungulates) and species-specific models (e.g., using only locations of lion-killed buffalo to 183 develop a lion-buffalo kill occurrence map) for ungulate species with sufficient sample size of 184 kill sites (N  $\ge$  25; buffalo, kudu, and warthog). However, preliminary model analyses using 185 Akaike's Information Criterion corrected for sample size (AIC<sub>c</sub>) indicated that, for each ungulate 186 species, the pooled model of lion-killed ungulates was better supported by the data than species-187 specific models. Therefore, we used the pooled model in subsequent analyses for all six 188 ungulate species. For hyena, we did not have sufficient sample sizes to evaluate species-specific 189 models, thus we built the hyena kill occurrence model using pooled locations of all hyena-killed 190 ungulates. The smaller sample size of hyena-killed ungulates was unsurprising given that hyenas 191 exhibit non-specific predation tendencies and lower preference for our focal ungulate species 192 than lions (Hayward 2006). 193

Using a case-control design, we compared the landscape covariates of the locations of each recorded lion (N = 172) or hyena (N = 50) kill site to landscape covariates associated with 20 random locations in *k* strata, where each stratum is a matched set consisting of the kill location and 20 random locations (Hosmer and Lemeshow 2000, Manly et al. 2002). We enforced a minimum distance of 500m such that none of the random controls were associated

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with the exact habitat conditions of the kill sites (Gervasi et al. 2013). We used a mixed effects logistic regression procedure and modeled the  $k_{ith}$  stratum as a random effect. We evaluated all linear, non-interacting combinations of the *a priori* covariates described above and ranked models according to AIC<sub>c</sub> weight ( $w_i$ ; Burnham and Anderson 2002). Given that our primary goal was prediction, we averaged parameter estimates over all models with a cumulative AIC<sub>c</sub> $w_i$ of 0.95 (Burnham and Anderson 2002:150). We used the resultant averaged parameter estimates to produce an odds ratio of kill occurrence (Keating and Cherry 2004) that took form:

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$$\Psi(\chi|\chi_R) = \exp \left[\beta_1(\chi_1 - \chi_R) + \beta_2(\chi_2 - \chi_R) + \dots + \beta_n(\chi_n - \chi_R)\right] (1)$$

where  $\Psi_{(x)}$  is the relative probability of kill occurrence in a given 250 m cell,  $\chi_n$  is the covariate value of that cell, and  $\chi_R$  is the mean covariate value of all cells in the study area. Given that predation events are rare throughout the study area, we can interpret the odds ratio as an approximation of the relative probability of kill occurrence (Keating and Cherry 2004), where  $\Psi(\chi|\chi_R) = 1.0$  represents the mean probability of kill occurrence. Values lower or higher than 1.0 represent changes in the odds ratio of kill occurrence (e.g.,  $\Psi(\chi|\chi_R) = 2.0$  indicates a location with twice the relative probability of kill occurrence than the mean; Fig. 1).

To assess the predictive ability of the kill occurrence models, we conducted k-folds cross validation tests (Boyce et al. 2002). We divided the data between training and testing portions maintaining an 80:20 ratio and fit kill occurrence models using the training data and validated the predictions with the testing data. We conducted this test five times and partitioned the predictive output into 10 categories. We compared the training and testing data using Spearmanrank correlation coefficients ( $r_s$ ; Boyce et al. 2002).

220 Ungulate aggregation models

 $Y_{ij} = X\beta_{ij} + Zu_j + e \ (2)$ 

We used the average ungulate aggregation size for each species at each transect point and within each period as the response variable in our ungulate aggregation models. We fit spatiallyexplicit mixed effects linear regression models to predict spatial variation in average ungulate aggregation in relation to carnivore occurrence and kill occurrence. These models took the form:

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where  $Y_{ij}$  is the log-transformed average ungulate aggregation size at the *i*th transect point in period *j*,  $X\beta_{ij}$  is a vector of the predictor covariates at the *i*th transect point in time period *j* (only lion UDs were separated by period), and  $Zu_j$  is the random effect term for each time period in the assessment (3 periods in each year such that  $N_j = 6$ ). For all models, we accounted for spatial autocorrelation among the *i*th transect points by using an exponential covariance structure, as determined by AIC<sub>c</sub> (Zuur et al. 2009). Prior to model fitting, we examined multicollinearity to avoid excessive covariate correlation (0.7 tolerance level).

We took a two-step information theoretic approach to model selection to balance 233 precision and bias in our final models (Burnham and Anderson 2002). In the first step of model 234 selection, we built a global model *a priori* that included carnivore UDs, carnivore kill 235 occurrence, amount of hard edge (linear length in meters), and degree of openness (proportion of 236 open habitat within a 250m<sup>2</sup> cell). We included hard edge and openness covariates due to 237 associations between open habitats, increased visibility, and aggregations of African ungulates 238 (Evans 1979, Gerard and Loisel 1995, Pays et al. 2007, Creel et al. 2014, Marino and Baldi 239 2014). We then evaluated models developed from all possible linear combinations of these six 240 covariates. Interaction terms were not considered due to lack of clear justification and a desire to 241 keep the candidate model set as small as possible (Burnham and Anderson 2002:17). We ranked 242 models based on AIC<sub>c</sub> and assessed models based on AIC<sub>c</sub> $w_i$  (Burnham and Anderson 2002). 243

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Aggregations of a given species might be influenced by the presence of other ungulates 244 (Scheel 1993). Accordingly, the second step of our model selection process tested whether the 245 top model for each ungulate species could be improved by the sequential addition of covariates 246 247 representing aggregations of other ungulate species at the *i*th transect point in the *j*th period. These covariates were not considered in the first step of the model selection process because we 248 could not hypothesize ungulate species associations in an *a priori* fashion and because an all 249 combinations approach would have undermined model parsimony by resulting >2000 candidate 250 models (Burnham and Anderson 2002:174). Thus, we assessed whether the fit of the top-ranking 251 model from the first model selection step could be improved by the iterative addition of 252 covariates representing the other ungulate species, as assessed by a significant likelihood ratio 253 test ( $\alpha = 0.05$  cutoff; Zuur et al. 2009). We concluded this process when the addition of 254 parameters no longer improved model fit and ranked final models using AIC<sub>c</sub> and AIC<sub>c</sub> $w_i$ . We 255 averaged parameter coefficients across all models within  $1/8^{\text{th}}$  AIC<sub>cWi</sub> of the top-ranking model 256 (Burnham and Anderson 2002:171). We calculated the relative importance of covariates  $X_i$  by 257 summing Akaike weights across models where covariate *j* occurred (Burnham and Anderson 258 2002:168). Finally, we plotted model-averaged coefficients for all parameters with a 259 significance of  $\alpha < 0.01$  to examine covariate effect size. 260

261 **RESULTS** 

We recorded 4,964 observations of ungulates, 615 lion locations, and 285 hyena locations during the two-year survey period. Population-level lion UDs varied by time period, with home range expansion occurring over the two-year reintroduction period (Appendix B). Populationlevel hyena UDs were focused around denning sites in the southwestern portion of the Main Camp (Appendix B). Ungulate aggregation distributions were right-skewed; mean aggregation sizes for the six species ranged from 2.8 to 16.7, with warthog forming the smallest aggregations

and buffalo forming the largest (Appendix C). We located 172 lion-killed ungulates and 50 hyena-killed ungulates. For both lion and hyena, kudu was the most commonly killed ungulate (34.3% [N = 59] of all lion kills and 40.0% [N = 20] of all hyena kills) and zebra was the least common (1.7% [N = 3] of all lion kills and 2.0% [N = 1] of all hyena kills).

Both the lion and hyena kill occurrence models revealed substantial spatial variation in 272 kill occurrence across the study area (Fig. 1). The k-folds cross validation tests indicated a good 273 fit of the models to the data, with an average correlation among the training and testing data of  $r_s$ 274 = 0.921 and  $r_s = 0.916$  for the lion and hyena models, respectively. For the lion model, the 275 relative probability of lion kill occurrence was positively associated with length of hard edge and 276 proportion gully thicket and negatively associated with distance to campsites; the highest 277 probability of lion kill occurrence was where gully thicket habitat formed a hard edge with open 278 habitat (Appendix D, Fig. 1). For the hyena model, the relative probability of hyena kill 279 occurrence was negatively associated with distance to waterhole and distance to fenceline and 280 positively associated with length of hard edge (Appendix E). Hyena kill occurrence was greatest 281 in open habitat (Fig. 1). 282

For each ungulate species, the addition of covariates representing aggregations of the 283 other ungulate species failed to improve the top model and thus those covariates are absent in 284 final model sets. Buffalo aggregation size was best described by a univariate model representing 285 degree of openness and the open covariate was the most important covariate in the model set 286 (Table 1, Fig. 2a). No covariate was significantly associated with buffalo aggregation size at the 287  $\alpha < 0.05$  level (Appendix F). For eland, hyena occurrence featured as a covariate in each of the 288 top four models (Table 1) and had the highest relative importance in the model set (Fig. 2b). For 289 290 kudu and hartebeest, the top four models for both species featured the open, lion kill occurrence,

and hyena occurrence covariates and these three covariates had the highest relative importance 291 over the model set (Table 1, Fig. 2c,d). In the case of kudu, lion kill occurrence and proportion 292 open habitat were both highly significant predictors of aggregation size (P < 0.0001, Appendix 293 F). For example, kudu aggregations increased from an average of  $\sim 2$  individuals in completely 294 closed habitat to  $\sim 4.5$  individuals in completely open habitat (Fig. 3a). Similarly, as the odds 295 ratio of lion kill occurrence increased twelve-fold from 0.5 to 6.0, the average aggregation size of 296 kudu increased from  $\sim 2$  individuals to  $\sim 5$  (Fig. 3b). For red hartebeest, these two covariates 297 (proportion open habitat and lion kill occurrence) were also highly significant (P < 0.01, 298 Appendix F) and showed a similar positive association. For example, average hartebeest 299 aggregations increased from an average of  $\sim$ 3 individuals in completely closed habitat to  $\sim$ 5 300 individuals in completely open habitat (Fig. 3c) and from ~3 individuals to ~7 as the odds ratio 301 of lion kill occurrence increased from 0.5 to 6.0 (Fig. 3d). For warthog, the open and lion kill 302 occurrence covariates featured in all of the top four models (Table 1) and both had a relative 303 importance near 1.0 over the model set (Figure 2e). The positive association between warthog 304 aggregation size and lion kill occurrence was highly significant (P = 0.01, Appendix F); as the 305 odds ratio of kill occurrence increased from 0.5 to 6.0, warthog aggregations increased from an 306 average of  $\sim 2$  individuals to  $\sim 3$  (Fig. 3e). Finally, for zebra the open and edge covariates 307 featured in all of the top four models (Table 1) and these two metrics both had a relative 308 importance near 1.0 over the model set (Fig. 2f). Of these, the open covariate had a highly 309 significant (P < 0.01) positive association with zebra aggregation size; zebra aggregations 310 increased from an average of  $\sim$ 3 individuals in completely closed habitat to  $\sim$ 4 in completely 311 open habitat (Fig. 3f). 312

313 DISCUSSION

A recent wave of research has advanced our understanding of non-lethal effects of 314 carnivores on ungulates, but generalizing these risk effects over broad terrestrial scales has 315 proven difficult (e.g., see Kauffman et al. 2010). Over a two-year period, we found that ungulate 316 aggregation size was positively associated with increased predation risk and habitat openness, 317 but the effects of both of these factors were moderate and varied by carnivore hunting mode and 318 prey species (Preisser et al. 2007). The final models predicting ungulate aggregation size were 319 largely multivariate and often contained several important covariates (Fig.2), highlighting the 320 idea that no single influence drives ungulate aggregation (Creel et al. 2014). Given that top 321 predators were absent from Addo for 100 years preceding our study and that reintroduced 322 carnivores existed in low densities, the risk effects we observed might be smaller in magnitude 323 when compared to systems with non-naïve prey (Berger 2001) or high predator densities (Peacor 324 and Werner 2001). Moreover, given the coarse-grain scale of our study (i.e., population-level 325 spatial tendencies of carnivore occurrence and kill sites over two years), we view our 326 conclusions as relevant to broad spatial scales of predation risk rather than finer-scaled responses 327 (e.g., increased vigilance as a predator approaches; Périquet et al. 2010). Therefore, despite the 328 established general principle that prey tend to aggregate in response to heightened predation risk 329 (Caro 2005), application of this antipredator defense strategy is scale-specific and dependent 330 upon habitat structure, the type and duration of predation risk, prey naiveté, and the ungulate 331 species in question. We do not view our study as an explicit test of top-down versus bottom-up 332 effects on ungulate aggregation. However, given the correlation between habitat structure and 333 forage availability and quality, our study suggests that complex top-down effects act in tandem 334 with bottom-up habitat features to shape prey response to predation risk (Jarman 1974, Anderson 335 336 et al. 2010, Hopcraft et al. 2010).

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Our results suggest that, like a complex physical landscape, the landscape of fear can be 337 highly heterogeneous. We note two sources of variation in the landscape of fear in Addo: 1) 338 carnivore identity, and 2) the nature of predation risk. In our study, spatial variation in lion kill 339 340 occurrence was a more important factor for predicting ungulate aggregation size than lion occurrence, the latter having generally low importance (Fig. 2). Others have found the presence 341 of lions influences ungulate behavior, either by increasing vigilance (Hunter and Skinner 1998, 342 Périquet et al. 2010), altering habitat selection and space use (Valeix et al. 2009a, Thaker et al. 343 2010, 2011), or limiting access to quality forage (Barnier et al. 2014). Taken in context with our 344 results, this suggests that ungulates employ multiple strategies to mitigate interactions with lions. 345 For example, whereas ungulates might prefer more open habitats and increase vigilance in 346 response to fine-scale risk (e.g, a lion in close proximity; Valeix et al. 2009b), they might 347 aggregate in especially dangerous locations at broader scales (Thaker et al. 2010, this study). 348 Moreover, in the face of consistent, long-term (i.e., >3 years) exposure to predation, some 349 ungulates might form large herds (i.e. > 100 individuals; Tambling et al. 2012) or migrate to 350 areas of lower risk (Hebblewhite and Merrill 2007). 351

In contrast with these lion-related effects, the influence of hyena on ungulate aggregation 352 was weaker and more strongly linked to where hyenas tended to occur rather than where they 353 tended to kill (Fig. 2). This difference appears to be driven in part by the opposing patterns of 354 kill occurrence of lions and hyenas in our study (Fig. 1). Given that covariates for lion and 355 hyena kill occurrence often occurred in top models together (Table 1) but did not always have 356 opposing coefficients (Appendix F), the effect of the opposing patterns in kill occurrence on 357 ungulate aggregation likely arose from differences in hunting behavior rather than as a mere 358 359 artefact of negatively correlated spatial kill patterns between lions and hyenas. Thus, our results

corroborate the disparity of influence from active (coursing) versus ambush (stalking) predators 360 that has been shown in smaller systems (e.g., Schmitz 2008). Creel et al. (2014) found that 361 African ungulate group size was more influenced by hyena presence than lion presence and 362 suggested ambush predators might have weaker effects in carnivore-ungulate systems than in 363 invertebrate systems (e.g., Schmitz 2008). Our study sheds light on this discrepancy by 364 decomposing risk from carnivores with differing hunting modes into two distinct expressions, 365 namely, where they tend to occur and where they tend to kill. With this level of resolution, our 366 results suggest that the active carnivore (hvena) tends to influence ungulate aggregation via 367 occurrence while the ambush carnivore (lion) tends to influence via a tendency to kill in specific 368 locations. However, as in Creel et al. (2014), the influence of these factors varied considerably 369 among our focal ungulate species, highlighting the need for additional research to test this 370 371 generalization in carnivore-ungulate systems

Habitat structure was also associated with ungulate aggregation for several species. Early 372 research posited that ungulates aggregated in open habitat to decrease predation risk via dilution 373 (Jarman 1974). However, more recent theoretical (Gerard and Loisel 1995) and experimental 374 (Pays et al. 2007) studies have shown that open areas facilitate aggregation simply by virtue of 375 improved visibility and conspecific detection. In our study, the degree of habitat openness 376 consistently featured in the top models for most species (Table 1, Fig. 2). However, with respect 377 to lion kill occurrence, open habitat was relatively safe (Fig. 1). Therefore, our work 378 corroborates the notion that predation and openness can act independently to shape group size in 379 ungulates (Marino and Baldi 2014), although these factors might be interactive for some species 380 (Thaker et al. 2010). The edge formed between open and closed habitat is especially dangerous 381 for African ungulates (Prins and Iason 1989). However, the amount of hard edge was generally a 382

less important aspect of habitat structure for ungulate aggregation compared to the level of
openness (Fig. 2), emphasizing the importance of visibility over access to cover, as has been
reported for other African ungulates (Riginos 2015).

Our models for buffalo suggest that neither predation risk nor habitat structure were 386 strongly related to aggregation size (Table 1, Appendix F). This lack of effect corroborates the 387 notions that buffalo "basically ignore lions" (Prins and Iason 1989:290) and that risk-induced 388 behavioral modifications are weaker in large-bodied ungulates (Hopcraft et al. 2012). Rather 389 than aggregating, buffalo might shift habitat selection to avoid risky areas (Valeix et al. 2009a, 390 2009b), a pattern that has also been recorded for giraffe (Giraffa camelopardalis; Creel et al. 391 2014). Recent work in Addo using long-term (i.e., 14 years) data found that, following lion 392 reintroduction, small buffalo herds amalgamated into a large herd that provided protection for 393 juvenile buffalo; however, this change required more than three years to take place (Tambling et 394 al. 2012). Hence, behavioral adaptations in ungulates following reintroduction might need 395 several years to actualize for naïve prey (Berger 2001). Understanding how risk effects vary over 396 time, especially in relation to naïve ungulate prey facing reintroduced carnivores, is an important 397 topic for future research. 398

Species-specific strategies for minimizing predation risk are poorly understood in multicarnivore systems (Thaker et al. 2011, Creel et al. 2014). We noted distinct responses among six ungulate species. The two largest ungulates, buffalo and eland, showed the least response to risk and had the highest model uncertainty (Table 1, Fig. 2, Appendix F). These two species had the largest aggregations on average (Appendix C) and both can successfully defend against lion attacks (Caro 2005), thus their response to predation risk is perhaps more closely tied to a few very large herds capable of self-defense rather than smaller aggregations of individuals

throughout the study area (Tambling et al. 2012). Moreover, the lack of risk effects on buffalo, 406 eland, and zebra (Fig. 2a,b,f) support the conclusion of Sinclair et al. (2003) that the influence of 407 top-down predation rapidly declines as prev body size increases beyond 150kg (zebra, the 408 smallest of these three species, has an average body size of 175kg; Stuart and Stuart 2000). The 409 final models for red hartebeest, kudu and warthog (Table 1, Fig. 2c,d,e; Appendix F), further 410 corroborate stronger influence of predation risk on small-bodied (<150kg) ungulates (Sinclair et 411 al. 2003, Hopcraft et al. 2012). However, whereas the associations between lion predation and 412 aggregation in these three smaller-bodied species were significant (P < 0.01, Appendix F), the 413 magnitude was moderate (Fig. 3b,d,e). For example, the greatest degree of association between 414 spatial aggregation and lion kill occurrence among these three species was in red hartebeest, with 415 aggregations increasing from  $\sim 3$  to  $\sim 7$  individuals on average across the full range of risk (e.g., 416 the odds ratio of lion kill occurrence in the most dangerous locations was >10 times that of the 417 safest places, Fig. 1, Fig. 3d). Nominal effects of predation risk on ungulate group size have 418 similarly been reported in both experimental (Hunter and Skinner 1998) and observational 419 420 studies (Valeix et al. 2009b, for exception see Thaker et al. 2010). However, more pronounced risk effects might be expected in systems with non-naïve prev or higher predator densities 421 (Peacor and Werner 2001). For instance, lion density in our study was 0.045 lions/km<sup>2</sup>, which is 422 lower than systems such as Kruger National Park (0.07 lions/km<sup>2</sup>) or the Serengeti (0.14 423 lions/km<sup>2</sup>; Hayward et al. 2007c). Similarly, prey response to risk might be more pronounced at 424 finer spatial scales (e.g., when a predator is in the immediate vicinity of prey). Risk effects from 425 wide-ranging carnivores vary over spatial scales (Hebblewhite and Merrill 2007), suggesting 426 that, like many ecological processes, comparisons of risk effects are likely scale-dependent 427 (Levin 1992). Accordingly, our inferences apply at a relatively coarse spatiotemporal scale (i.e., 428

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analogous to Johnson's (1980) second or third order) and do not necessarily preclude strongerrisk effects at finer-scales.

Our findings highlight the complexity inherent in broad-scale, multi-carnivore, multi-431 ungulate systems and suggest that understanding the strength and scope of carnivore risk effects, 432 along with associated implications for ungulate demography and ecosystem function, remains a 433 substantial challenge. Moreover, evaluating the numeric prey response to reintroduced predators 434 is also important, both for understanding patterns in ungulate aggregation (Hebblewhite and 435 Pletscher 2002) and predicting the long-term impacts of top predators on trophic interactions 436 (Peterson et al. 2014). Hence, we view our findings of moderate, coarse-grain risk effects on 437 naïve prey over a two-year period as reason to echo the call of Tambling et al. (2012) for robust, 438 long-term monitoring of African carnivore reintroductions in order to place non-lethal behavioral 439 responses within the context of the numeric and functional responses in multi-predator, multi-440 prey systems. Studies that disentangle the effects of 'risky times' (carnivores in close proximity) 441 and 'risky places' (locations of probable carnivore occurrence or kills) at multiple spatial scales 442 (sensu Johnson 1980) would be particularly insightful. Lastly, we encourage researchers to build 443 upon our study by incorporating explicit measures of forage into evaluations of carnivore risk 444 effects on ungulates. 445

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#### 457 LITERATURE CITED

- 458 Anderson, T. M., J. G. C. Hopcraft, S. Eby, M. Ritchie, J. B. Grace, and H. Olff. 2010.
- 459 Landscape-scale analyses suggest both nutrient and antipredator advantages to Serengeti
  460 herbivore hotspots. Ecology 91:1519–1529.
- 461 Barnier, F., M. Valeix, P. Duncan, S. Chamaillé-Jammes, P. Barre, A. J. Loveridge, D. W.
- Macdonald, and H. Fritz. 2014. Diet quality in a wild grazer declines under the threat of
  an ambush predator. Proceedings of the Royal Society B-Biological Sciences
  281:20140446.
- 465 Berger, J., J. E. Swenson, and I. Persson. 2001. Recolonizing carnivores and naïve prey:

466 conservation lessons from Pleistocene extinctions. Science 291:1036-1039.

Boyce M. S., P. R. Vernier, S. E. Nielsen, and F. K. A. Schmiegelow. 2002. Evaluating resource
selection functions. Ecological Modelling 157:281-300.

469 Buckland, S. T., Anderson, D. R., Burnham, K. P., Laake, J. L., Borchers, D. L., and L. N.

- Thomas. 2001. Introduction to distance sampling: estimating abundance of biologicalpopulations. Oxford University Press, Oxford, United Kingdom.
- 472 Buckland, S. T., D. R. Anderson, K. P. Burnham, and J. L. Laake. 2005. Distance sampling.
- 473 Encyclopedia of biostatistics. Second edition. John Wiley & Sons, Inc, Hoboken, New
- 474 Jersey, USA.

475 Burnham, K. P., and D. R. Anderson. 2002. Model selection and multimodel inference: a

- 476 practical information-theoretic approach. Second edition. Springer, NewYork, New York,477 USA.
- 478 Caro, T. 2005. Antipredator defenses in birds and mammals. University of Chicago Press,

479 Chicago, IL, USA.

- Creel, S., and D. Christianson. 2008. Relationships between direct predation and risk effects.
  Trends in Ecology and Evolution 23:194–201.
- 482 Creel, S., J. Schuette, D. Christianson. 2014. Effects of predation risk on group size, vigilance,
- and foraging behavior in an African ungulate community. Behavioral Ecology 24:773784.
- Davies-Mostert H. T., M. G. L. Mills, and D. W. Macdonald. 2013. Hard boundaries influence
  African wild dogs' diet and prey selection. Journal of Applied Ecology 50:1358-1366.
- 487 Dehn, M. M. 1990. Vigilance for predators: detection and dilution effects. Behavioral Ecology
  488 and Sociobiology 26:337-342.
- Evans, P. G. H. 1979. Habitat preferences of ungulates in closed savanna of central Africa.
  Mammal Review 9:19–32.
- Fortin, D., M. E. Fortin, H. L. Beyer, T. Duchesne, S. Courant, and K. Dancose. 2009. Groupsize-mediated habitat selection and group fusion-fission dynamics of bison under
  predation risk. Ecology 90:2480-2490.
- 494 Gerard, J., and P. Loisel. 1995. Spontaneous emergence of a relationship between habitat
- 495 openness and mean group size and its possible evolutionary consequences in large
- 496 herbivores. Journal of Theoretical Biology 176:511–522.

- 497 Gervasi, V., H. Sand, B. Zimmermann, J. Mattisson, P. Wabakken, and J. D. C. Linnell. 2013.
- 498 Decomposing risk: Landscape structure and wolf behavior generate different predation
   499 patterns in two sympatric ungulates. Ecological Applications 23:1722–1734.
- 500 Gitzen, R. A., J. J. Millspaugh, and B. J. Kernohan. 2006. Bandwidth selection for fixed-kernel
- analysis of animal utilization distributions. Journal of Wildlife Management 70:13341344.
- Hamilton, W. D. 1971. Geometry for the selfish herd. Journal of Theoretical Biology 31:295311.
- Hayward, M. W. 2006. Prey preferences of the spotted hyaena (*Crocuta crocuta*) and degree of
  dietary overlap with the lion (*Panthera leo*). Journal of Zoology 270:606-614.
- Hayward, M. W., and G. J. Hayward. 2006. Activity patterns of reintroduced lion *Panthera leo*and spotted hyena *Crocuta crocuta* in Addo Elephant National Park, South Africa.
- 509African Journal of Ecology 45:135-141.
- 510 Hayward, M. W., G. I. H. Kerley, J. Adendorff, L. C. Moolman, J. O'Brien, A. Sholto-Douglas,
- 511 C. Bissett, P. Bean, A. Fogarty, D. Howarth, and R. Slater. 2007a. The reintroduction of
- 512 large carnivores to the Eastern Cape, South Africa: an assessment. Oryx 41:205.
- Hayward, M. W., J. O'Brian, M. Hofmeyr, and G. I. H. Kerley. 2007b. Testing predictions of the
  prey of lion derived from modelled prey preferences. Journal of Wildlife Management
  71:1567-1575.
- Hayward, M. W., J. O'Brian, and G. I. H. Kerley. 2007c. Carrying capacity of large African
  predators: Predictions and tests. Biological Conservation 139:219-229.
- 518 Hayward, M. W., G. J. Hayward, D. J. Druce, and G. I. H. Kerley. 2009. Do fences constrain
- 519 predator movements on an evolutionary scale? Home range, food intake, and movement

- 520 patterns of large predators reintroduced to Addo Elephant National Park, South Africa.
- 521 Biodiversity and Conservation 18:887-904.
- 522 Hebblewhite, M., and D. H. Pletscher. 2002. Effects of elk group size on predation by wolves.
- 523 Canadian Journal of Zoology 80:800-809.
- Hebblewhite, M., and E. H. Merrill. 2007. Multiscale wolf predation risk for elk: does migration
  reduce risk? Oecologia 152:377-387.
- 526 Hopcraft, J. G. C., H. Olff, and A. R. E. Sinclair. 2010. Herbivores, resources and risks:
- alternating regulation along primary environmental gradients in savannas. Trends in
  Ecology and Evolution 25:119–128.
- 529 Hopcraft, J. G. C., T. M. Anderson, S. Pérez-Vila, E. Mayemba, and H. Olff. 2012. Body size
- and the division of niche space: food and predation differentially shape the distribution of
  Serengeti grazers. Journal of Animal Ecology 81:201-213.
- Hosmer, D. W., and S. Lemeshow. 2000. Applied Logistic Regression. Wiley, New York, New
  York, USA.
- Howe, M., M. M. Okello, and J. M. Davis. 2013. Interspecific variation in the distribution of
- ungulates relative to human infrastructure surrounding Amboseli National Park. AfricanZoology 48:159-166.
- Hunter, L. T. B., and J. D. Skinner. 1998. Vigilance behaviour in African ungulates: the role of
  predation pressure. Behaviour 135:195-211.
- Jarman, P. J. 1974. The social organisation of antelope in relation to their ecology. Behaviour
  48:215–267.
- Johnson, D. H. 1980. The comparison of usage and availability measurements for evaluating
  resource preference. Ecology 61:65–71.

### Moll et al. COLOGICAL SOCIETYOFAMERICA<sup>25</sup>

- 543 Kauffman, M. J., J. F. Brodie, and E. S. Jules. 2010. Are wolves saving Yellowstone's aspen? A
- landscape-level test of a behaviorally mediated trophic cascade. Ecology 91:2742-2755.
- Keating, K. A., and S. Cherry. 2004. Use and interpretation of logistic regression in habitatselection studies. Journal of Wildlife Management 68:774-789.
- Laundré, J. W., L. Hernández, and K. B. Altendorf. 2001. Wolves, elk, and bison: reestablishing
  the "landscape of fear" in Yellowstone National Park, USA. Canadian Journal of Zoology
  79:1401-1409.
- Levin, S. A. 1992. The problem of pattern and scale in ecology: the Robert H. MacArthur Award
  lecture. Ecology 73:1943–1967.
- Lima, S. L. 1998. Nonlethal effects in the ecology of predator-prey interactions. BioScience
  48:25–34.
- Loarie, S. R., C. J. Tambling, and G. P. Asner. 2013. Lion hunting behaviour and vegetation
   structure in an African savanna. Animal Behaviour 85:899–906.
- 556 Manly, B. F. J., L. L. McDonald, D. L. Thomas, T. L. McDonald, and W. P. Erickson. 2002.
- Resource selection by animals: statistical design and analysis for field studies. Second
  edition. Kluwer, Boston, Massachusetts, USA.
- Marino, A., and R. Baldi. 2014. Ecological correlates of group-size variation in a resourcedefense ungulate, the sedentary guanaco. PLoS ONE 9:e89060.
- Molvar, E. M., and R. T. Bowyer. 1994. Costs and benefits of group living in a recently social
- ungulate: the Alaskan moose. Journal of Mammalogy 75:621–630.
- Pays, O., S. Benhamou, R. Helder, and J. F. Gerard. 2007. The dynamics of group formation in
  large mammalian herbivores: an analysis in the European roe deer. Animal Behaviour
- 565 74:1429-1441.

- Peacor, S. D., and E. E. Werner. 2001. The contribution of trait-mediated indirect effects to the 566
- net effects of a predator. Proceedings of the National Academy of Sciences of the United 567 States of America 98:3904–3908. 568
- 569 Périquet, S., M. Valeix, A. J. Loveridge, H. Madzikanda, D. W. Macdonald, and H. Fritz. 2010.
- Individual vigilance of African herbivores while drinking: the role of immediate predation 570 risk and context. Animal Behaviour 79:665-671. 571
- Peterson, R. O., J. A. Vucetich, J. M. Bump, and D. W. Smith. 2014. Trophic Cascades in a 572
- Multicausal World: Isle Royale and Yellowstone. Annual Review of Ecology, Evolution, 573 574 and Systematics 45:325–345.
- Preisser, E. L., J. L. Orrock, and O. J. Schmitz. 2007. Predator hunting mode and habitat domain 575 alter nonconsumptive effects in predator-prey interactions. Ecology 88:2744-2751. 576
- 577 Prins, H. H. T., and G. R. Iason. 1989. Dangerous lions and nonchalant buffalo. Behaviour 108:262-296. 578
- R Development Core Team. 2011. R: A language and environment for statistical computing. R 579 Foundation for Statistical Computing, Berkeley, Califorina, USA. 580
- Riginos, C. 2015. Climate and the landscape of fear in an African savanna. Journal of Animal 581 Ecology 84: 124-133. 582
- Ripple, W. J., and R. L. Beschta. 2004. Wolves and the ecology of fear: can predation risk 583 structure ecosystems? BioScience 54:755-766. 584
- Scheel, D. 1993. Watching for lions in the grass: the usefulness of scanning and its effects during 585 hunts. Animal Behaviour 46:695-704. 586
- Schmitz, O. J. 2008. Effects of predator hunting mode on grassland ecosystem function. Science 587 319:952-954. 588

- 589 Seaman, D. E., J. J. Millspaugh, B. J. Kernohan, G. C. Brundige, K. J. Raedeke, and R. A.
- Gitzen. 1999. Effects of sample size on kernel home range estimates. Journal of WildlifeManagement 63:739-747.
- Sinclair, A. R. E., and P. Arcese. 1995. Population consequences of predation-sensitive foraging:
  the Serengeti wildebeest. Ecology 76:882–891.
- Sinclair, A. R. E., S. Mduma, and J. S. Brashares. 2003. Patterns of predation in a diverse
  predator-prey system. Nature 425:288–290.
- Stuart, C. and T. Stuart. 2000. Field guide to the larger mammals of Africa. Struik, Cape Town,
  South Africa.
- Tambling, C. J., D. J. Druce, M. W. Hayward, J. G. Castley, J. Adendorff, and G. I. H. Kerley.
- 599 2012. Spatial and temporal changes in group dynamics and range use enable anti-predator
   600 responses in African buffalo. Ecology 93:1297-1304.
- Thaker, M., A. T. Vanak, C. R. Owen, M. B. Ogden, and R. Slotow. 2010. Group dynamics of
- zebra and wildebeest in a woodland savanna: effects of predation risk and habitat density.
  PLoS One 5:e12758.
- Thaker, M., A. T. Vanak, C. R. Owen, M. B. Ogden, S. M. Niemann, and R. Slotow. 2011.
- 605 Minimizing predation risk in a landscape of multiple predators: effects on the spatial 606 distribution of African ungulates. Ecology 92:398-407.
- 607 Valeix, M., A. J. Loveridge, S. Chamaillé-Jammes, Z. Davidson, F. Murindagomo, H. Fritz, and
- D. W. Macdonald. 2009a. Behavioral adjustments of African herbivores to predation risk
- by lions: spatiotemporal variations influence habitat use. Ecology 90:23-30.

- 610 Valeix, M., H. Fritz, A. J. Loveridge, Z. Davidson, J. E. Hunt, F. Murindagomo, and D. W.
- 611 Macdonald. 2009b. Does the risk of encountering lions influence African herbivore
- behaviour at waterholes? Behavioral Ecology and Sociobiology 63:1483-1494.
- E Zuur A., E. N. Ieno, N. Walker, A. A. Saveliev, and G. M. Smith. 2009. Mixed effects models
- and extensions in ecology with R. Springer, New York, New York, USA.

#### 615 SUPPLEMENTAL MATERIAL

- 616 Ecological Archives
- 617 Appendices A-F are available online: URL
- 618 Appendix A. Study area and transect locations.
- 619 Appendix B. Utilization distributions (UDs) of lions and hyenas.
- 620 Appendix C. Ungulate aggregation sizes.
- 621 Appendix D. Lion (*Panthera leo*) kill occurrence model.
- 622 Appendix E. Spotted hyena (*Crocuta crocuta*) kill occurrence model.
- 623 Appendix F. Model-averaged coefficients for ungulate aggregation models.

- Table 1. Models describing aggregation of ungulates exposed to multi-carnivore predation in the 624
- Main Camp section of Addo Elephant National Park, South Africa (2003-2005) after a two-step 625
- information theoretic model selection procedure. Top four models shown for clarity. 626

|            | Model covariates <sup>†</sup>                         | k  | $\Delta AIC_{c}$ | Wi   |
|------------|---|----|------------------|------|
| Duffalo    | Onen  | 6  | 0                | 0.12 |
| Bullalo    | Open  | 0  | 0                | 0.13 |
|            | HyenaKill   | 6  | 1.23             | 0.07 |
|            | Open + LionKill                                       | 7  | 1.43             | 0.06 |
|            | Null (Intercept only)                                 | 5  | 1.45             | 0.06 |
| Eland      | LionKill + HyenaUD                                    | 7  | 0                | 0.07 |
|            | LionUD + HyenaUD                                      | 7  | 0.35             | 0.06 |
|            | HyenaUD   | 6  | 0.42             | 0.05 |
|            | LionKill + LionUD + HyenaUD                           | 8  | 0.48             | 0.05 |
| Kudu       | Open + Edge + LionKill + LionUD + HyenaKill + HyenaUD | 11 | 0                | 0.21 |
|            | Open + LionKill + LionUD + HyenaKill + HyenaUD        | 10 | 0.37             | 0.17 |
|            | Open + Edge + LionKill + HyenaKill + HyenaUD          | 10 | 0.79             | 0.14 |
|            | Open + LionKill + LionUD + HyenaUD                    | 9  | 1.78             | 0.09 |
| Red        | Open + Edge + LionKill + HyenaUD                      | 9  | 0                | 0.36 |
| hartebeest | Open + Edge + LionKill + HyenaKill + HyenaUD          | 10 | 1.28             | 0.19 |
|            | Open + LionKill + HyenaUD                             | 8  | 1.7              | 0.15 |
|            | Open + Edge + LionKill + LionUD + HyenaUD             | 10 | 2.11             | 0.13 |
| Warthog    | Open + LionKill                                       | 7  | 0                | 0.28 |
|            | Open + LionKill + HyenaUD                             | 8  | 1.84             | 0.11 |

|       | Open + Edge + LionKill   | 8 | 1.93 | 0.11 |
|-------|--------------------------|---|------|------|
|       | Open + LionKill + LionUD | 8 | 1.94 | 0.11 |
| Zebra | Open + Edge              | 7 | 0    | 0.2  |
|       | Open + Edge + LionKill   | 8 | 0.76 | 0.14 |
|       | Open + Edge + LionUD     | 8 | 1.49 | 0.1  |
|       | Open + Edge + HyenaKill  | 8 | 1.57 | 0.09 |

627 *Notes*: k = number of model parameters;  $w_i = AIC_c$  weight for the *i*th model.

<sup>†</sup>Covariate definitions: Open = proportion open habitat, Edge = length of hard edge, LionKill =

odds ratio of lion kill occurrence, LionUD = lion UD percentile, HyenaKill = odds ratio of hyena

630 kill occurrence, HyenaUD = hyena UD percentile.

#### 631 FIGURE LEGENDS

- Figure 1. Odds ratios of lion and hyena kill occurrence of six ungulate species (buffalo, eland,
- kudu, red hartebeest, warthog, and zebra) at 250m<sup>2</sup> resolution in the Main Camp section of Addo
- Elephant National Park, South Africa (2003-2005). The odds ratio scale of predation applies to
- both maps, but given the higher prevalence of lion predation, risky areas in the lion map are
- riskier than similarly colored areas in the hyena map. Green and red/purple areas represent the
- 637 lowest and highest odds ratios of kill occurrence, respectively. Open habitat is identified by the
- overlaid dashed polygons; all other habitat was classified as closed.
- 639 Figure 2. Relative importance of covariates in final model sets describing ungulate aggregation
- 640 in Addo Elephant National Park, South Africa (2003-2005). The numbers above each bar are
- relative importance values of a given covariate rounded to two decimal places.

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Figure 3. Statistically significant ( $\alpha \le 0.01$ ) spatially-explicit mixed effects regression trends derived from model-averaged coefficients from models describing ungulate aggregation in Addo Elephant National Park, South Africa (2003-2005). Shaded areas represent 95% confidence intervals.

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651 Covariate definitions: Open = proportion open habitat, Edge = length of hard edge, LionKill =





Figure 3.



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